Title: Mechanistic home range capture–recapture models for the estimation of population density

and landscape connectivity

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Abstract

Spatial capture–recapture models (SCRs) provide an integrative statistical tool for analyzing animal movement and population patterns. Although incorporating home range formation with a theoretical basis of animal movement into SCRs can improve the prediction of animal space use in a heterogeneous landscape, this approach is challenging owing to the sparseness of recapture events.

In this study, we developed an advection–diffusion capture–recapture model (ADCR), which is an extension of SCRs incorporating home range formation with advection–diffusion formalism, providing a new framework to estimate population density and landscape permeability. we tested the unbiasedness of the estimator using simulated capture–recapture data generated by a step selection function. We also compared accuracy of population density estimates and home range shapes with those from an SCR incorporating the least-cost path. In addition, ADCR was applied to real dataset of Asiatic black bear (*Ursus thibetanus*) in Japan to demonstrate the capacity of the ADCR to detect geographical barriers that constrain animal movements.

Population density, permeability, and home range estimates of ADCR were unbiased over randomly determined sets of true parameters. Although the accuracy of density estimates by ADCR was nearly identical to those of existing models, the home range shape could be predicted more accurately by ADCR than by an SCR incorporating the least-cost path. For the application to bear dataset, ADCR could detect the effect of water body as a barrier of movement which is consistent with previous population genetic studies.

ADCR provides unique opportunities to elucidate both individual- and population-level ecological processes from capture–recapture data. By offering a formal link with step selection functions to estimate animal movement, it is suitable for simultaneously modeling with capture–recapture data and animal movement data. This study provides a basis for studies of the interplay between animal movement processes and population patterns.

Introduction

Anthropogenic activity leads to habitat loss and fragmentation, posing a major threat to

biodiversity. In a fragmented landscape, connectivity between habitat patches affects population

patterns and sustainability (Hanski et al. 1994). Functional connectivity in a heterogeneous

landscape is not only of general interest in basic ecology but also a key factor affecting

decision-making in wildlife management and conservation (Goodwin and Fahrig 2002, Royle et

al. 2018, Osada et al. 2019). Circuit theory and least-cost path modeling are two major paradigms

for evaluating landscape connectivity, with different underlying behavioral assumptions and

statistical properties (McRae et al. 2008, Marrotte and Bowman 2017). Permeability, which is a

measure of connectivity based on circuit theory has explicit links with animal movement models,

such as random walks and step selection functions (McRae 2006, Duchesne et al. 2015), and

models of animal gene flow and movement have numerous applications (Dickson et al. 2019).

Empirical determination of connectivity parameters, such as effects of land cover types on

permeability, is important to improve objectivity, and the development of estimators is a major

challenge in ecology (Ovaskainen 2004, Ovaskainen et al. 2008, Beyer et al. 2016, Signer et al.

2017, Osada et al. 2019).

Spatial capture–recapture models (SCRs) are powerful tools with the capacity to

estimate individual movement processes and population abundance in an integrative manner

(McClintock et al. 2022). Although permeability has favorable properties for practical

applications, such as robustness to the resolution of landscape data (Marrotte and Bowman

2017) and a straightforward link with movement models (Diniz et al. 2020), they are not as

well-studied as SCRs incorporating the least-cost path (Royle et al. 2013, Sutherland et al. 2015,

Fuller et al. 2016). In particular, the process of home range formation considering permeability

and site fidelity has been studied extensively by step selection functions in movement ecology

(Thurfjell et al. 2014, Duchesne et al. 2015, Signer et al. 2017, Péron 2019, Potts and Schlägel

2020). The corresponding SCR estimator can provide insight into the relationship between

individual-level displacement and population-level abundance patterns. However, the temporal

sparseness of capture–recapture data limits the application of SCR to complex movement

processes (McClintock et al. 2022).

Incorporating mechanistic models of home range formation into SCR can enable

inferences about permeability as well as population abundance from capture–recapture data in a

computationally tractable manner. Mechanistic home range analysis (Moorcroft and Lewis 2006,

Moorcroft and Barnett 2008) is an analytical approach that ties the process of animal movements

to the resultant home range shape at an equilibrium state. The approach has contributed to the

development of advection–diffusion equations to obtain flexible home ranges and

likelihood-based modeling to estimate parameters empirically. The advection–diffusion

equations describe deterministic drift to a focal point (e.g., an activity center) as advection terms

and random walks as diffusion terms. Circuit theory is analogous to a diffusion process with

heterogeneous diffusivity (Hanks and Hooten 2013, Yamaura et al. 2022), and the

advection–diffusion equations with heterogeneous diffusion coefficients in space are applicable

to home ranges formed by animal movement affected by barriers in a landscape. Although

closed-form analytical solutions of advection–diffusion equations rarely exist, computationally

efficient ways to obtain an equilibrium solution have been developed (Hirsch 2007). By

integrating mechanistic home range analysis into SCR, the simultaneous estimation of

permeability and population density is possible, without substantial computational costs.

Hereafter, we refer to this new SCR approach as the advection–diffusion capture–recapture

model (ADCR).

Although ADCRs and SCRs with the least-cost path (Royle et al., 2013; Sutherland et

al., 2015; Royle et al., 2018) differ with respect to connectivity assumptions and cannot be

directly compared, evaluations of the performance of estimators are helpful for users. Applying

statistical models to simulated data generated from a mechanistic animal movement model is a

straightforward way to test the unbiasedness of estimators and robustness to deviation from

assumptions (Dupont et al. 2022, Theng et al. 2022). Also, an application to real dataset is

important to illustrate how statistical models can yield realistic estimates even in situations

where there is unknown heterogeneity in the data generating process.

The aim of this study was to develop a formal framework for ADCR and to compare

the accuracy of the estimated abundance and home range shape with those obtained by SCRs

with the least-cost path. The identification of the optimal model is outside of the scope of this

paper. Rather, our aim is to establish the merit of the ADCR under animal movement modeled

by a step selection function, which is common in movement ecology. We also compared density

estimates with those of basic SCR (Borchers and Efford 2008) as a null model of heterogeneous

connectivity. We predict that all three methods, ADCR, an SCR with the least-cost path, and

basic SCR, will accurately estimate population densities because SCR is robust to the violation

of home range assumptions (Efford 2019). We also predict ADCR will have merit in estimating

environmental signals on permeability correctly hence predicting home range shapes under a

situation that animals movement is driven by exploratory behavior. As an example of model

application, we applied ADCR to estimate factors affecting peameability for Asiatic black bear

(*Ursus thibetanus*) in Japan.

Material and Methods

Model Framework

ADCR is an extension of SCR in which animal home ranges are modeled by equilibrium solutions of the advection–diffusion equation, an Eulerian representation of stochastic animal movement in a heterogeneous connectivity landscape. By integrating the advection–diffusion equation with SCR, it is possible to simultaneously estimate population density while inferring the effect of landscape connectivity on the skewed shape of the home range, using capture–recapture data. Here, we describe the model structure of ADCR.

Detection model and population process

In stationary SCR modeling, a detection model that describes the encounter of an individual to a detector reflects the home range behavior of the animal. The home range of individual is defined by its utilization distribution, which is a bivariate probability distribution of an individual location in a space. In general, the model of utilization distribution has latent variables of an activity center $\boldsymbol{\mu} = (\mu_x, \mu_y)$, where an individual is attracted. For example, a basic SCR (Borchers and Efford 2008, Royle et al. 2009) often assumes a half-normal detection function, which implies that the utilization distribution is a stationary bivariate normal distribution centered to the activity center. To introduce the ADCR, which covers a more flexible class of utility distributions under heterogeneous permeability landscapes, we adopted a detection model that describes an explicit link between the encounter frequency of an individual to a detector and the utilization distribution $p^*(\mathbf{s} \mid \boldsymbol{\mu})$, which defines the probability density at location $\mathbf{s} = (x, y)$ of an individual with an activity center $\boldsymbol{\mu}$. Note that the integration of $p^*(\mathbf{s} \mid \boldsymbol{\mu})$ over $\mathbf{s}$ equals 1 because of the nature of probability distribution.

An SCR survey is conducted using an array of detectors in space. Researchers may either iterate the survey on more than one occasion or treat the entire survey period as a single occasion. When the SCR data are collected by count detectors, the encounter frequency of the

*i*th individual to the *j*th detector on the *k*th occasion is written as

$$\lambda_{ijk} = \exp(g_0)p^*(\mathbf{s}_j \mid \boldsymbol{\mu}_i), \quad \text{(eqn. 1)}$$

where g_0 is a parameter of log detectability, $\boldsymbol{\mu}_i$ is the activity center of the *i*th individual, and $\mathbf{s}_j$ is

the location of the *j*th detector. The corresponding observed value, Y_{ijk}, follows a probability

distribution depending on λ_{ijk}, denoted by $p(Y_{ijk} \mid \boldsymbol{\mu}_i, \boldsymbol{\theta})$, where $\boldsymbol{\theta}$ is a vector of parameters that

affects λ_{ijk}, including g_0 and the parameters controlling p^*. The number of detections follows a

Poisson distribution $p(Y_{ijk} \mid \boldsymbol{\mu}_i, \boldsymbol{\theta}) = \text{Poisson}(Y_{ijk} \mid \lambda_{ijk})$ if the conditional independence among

detection events given λ_{ijk} is assumed. Although I implemented an observation model for count

detectors in this study, it can be easily extended to a model of varying effort with various

detector types (e.g., binary proximity and multi-catch trap) as well as the basic SCR (Efford et

al. 2013).

The population process of SCR describes the spatial pattern of activity centers

following a point process in a region $\mathcal{S}$ (Borchers and Efford 2008, Royle et al. 2009). $\mathcal{S}$ is often

defined as a rectangular region that includes an array of detectors. In this study, we applied a

homogeneous Poisson process, the simplest form of the point process, assuming that $\boldsymbol{\mu}$ follows a

uniform distribution over $\mathcal{S}$, $\boldsymbol{\mu} \sim \text{Uniform}(\mathcal{S})$. Let ρ denote the log population density, and the

number of activity centers in an arbitrary region with area A follows a Poisson distribution with

expectation $\exp(\rho)A$. As is the case with the basic SCRs (Borchers and Efford 2008), the model

can be extended to an inhomogeneous Poisson process that can accommodate a situation in

which ρ varies in space depending on the landscape covariates.

Home range formation

In ADCR, the utilization distribution is derived from a fine-scale movement behavior

that forms a stationary home range in a heterogeneous permeability landscape. The

Ornstein–Uhlenbeck process (Smouse et al. 2010) is a basic movement model describing the

home range behavior of an animal that is subject to drift toward the activity center and a random

walk. According to the circuit theory, permeability corresponds to the step length of a random

walk (McRae et al. 2008), and we considered a situation in which permeability depends on

landscape covariates. The Ornstein–Uhlenbeck process converges to a probability distribution of

individual location at the equilibrium state (Smouse et al. 2010), and we applied it to utilization

distribution in ADCR.

To derive the probability distribution of an individual location, we introduced an

Eulerian approach that focuses on the flow of probability density in a space driven by drift and

random walk. Under the Ornstein–Uhlenbeck process, the probability distribution of an

individual location follows the Fokker–Planck equation, a partial differential equation that

describes the time evolution of the probability density in space (Méndez et al. 2013). The

Fokker–Planck equation of probability density p is

$$\frac{\partial p}{\partial t} = -c\left((\mu_x - x)\frac{\partial p}{\partial x} + (\mu_y - y)\frac{\partial p}{\partial y}\right) + D(x, y)\left(\frac{\partial^2 p}{\partial x^2} + \frac{\partial^2 p}{\partial y^2}\right), \qquad \text{(eqn. 2)}$$

where $D(x, y) > 0$ is the spatially-varying diffusion coefficient and c is the intensity of drift. The

first term on the right-hand side is advection corresponding to drift toward the activity center, and

the second term is diffusion. The diffusion coefficient varying in space is interpreted as

permeability (or as conductance in terms of circuit theory) (McRae 2006, McRae et al. 2008). To

model the effect of landscape variables on permeability, we applied the log-linear mode $\ln(D(x,$

$y)) = \beta_0 + \boldsymbol{\beta}\mathbf{z}(x, y)$, where β_0 is the intercept, $\mathbf{z}(x, y)$ is a vector of the landscape covariates at

location (x, y), and $\boldsymbol{\beta}$ is the corresponding regression coefficients that determine the dependence

of permeability to landscape covariates. The intensity of drift, c, takes a positive value, and $\ln(c)$

is denoted by the parameter α_0. To ensure that the total mass of probability does not change with

time, the reflecting boundary condition is given for the space. Eqn. 2 describes the time evolution

of the probability density of an individual location, and we applied it to project the animal

movement paths for the data simulation (see **Simulation Study**).

Similar to mechanistic home range analysis (Moorcroft and Lewis 2006), the home

range in this system is defined as the equilibrium probability density distribution $p^*(\mathbf{s} \mid \boldsymbol{\mu})$,

which is the solution to the following system:

197
$$-\left((\mu_x - x)\frac{\partial p}{\partial x} + (\mu_y - y)\frac{\partial p}{\partial y}\right) + \frac{D(x,y)}{c}\left(\frac{\partial^2 p}{\partial x^2} + \frac{\partial^2 p}{\partial y^2}\right) = 0. \quad \text{(eqn.3)}$$

198   To make sure that $p^*(\mathbf{s} \mid \boldsymbol{\mu})$ is the probability distribution in the region $\mathcal{S}$, we imposed a

199   constraint of $\int_{\mathcal{S}} p \, d\mathbf{s} = 1$. Note that the parameters α_0 and β_0 are redundant under the equilibrium

200   assumption because $\ln(D(x, y)/c) = \beta_0 - \alpha_0 + \boldsymbol{\beta}\mathbf{z}(x, y)$. To ensure the identifiability of the model, it

201   was reparametrized as $\gamma_0 = \beta_0 - \alpha_0$. γ_0 is interpreted as the logarithm of the relative intensity of

202   diffusion and drift, which determines the average home range size.

203   Although eqns. 2 and 3 imply that the landscape covariates are spatially continuous,

they are actually given by a discrete dataset in raster or vector format. Accordingly, these

equations can be numerically solved using a variety of discretization techniques, including the

finite difference method and finite element method (Hirsch 2007). In this study, the finite

difference method was applied because it is suitable for common situations in which landscape

factors affecting the diffusion coefficient are available as a raster dataset. In addition, grid cells of

raster data were used to discretize space for simplicity. Technical details for discretization and

numerical derivation of $p^*(\mathbf{s} \mid \boldsymbol{\mu})$ are shown in **Appendix S1**. Importantly, our home range

specification is a generalization of basic SCRs that assume a bivariate normal home range

(Borchers and Efford 2008, Royle et al. 2009), given that the solution of eqn. 3 when $D(x, y)$ is

constant over space (i.e., $\boldsymbol{\beta} = \mathbf{0}$) is a bivariate normal distribution whose variance is $\exp(\gamma_0)$

(Méndez et al. 2013). An example of derived home ranges with different βs in a hypothetical

landscape with a binary covariate for permeability is shown in Fig. 1, which shows that the

larger the effect of a landscape covariate, the more the home range shape skews toward the

habitat patch with high permeability.

Model estimation

We applied the maximum likelihood method to estimate the permeability parameters

and population density simultaneously from the capture–recapture data. Although conditional

likelihood maximization can also be used for parameter estimation (Borchers and Efford 2008),

full likelihood maximization was applied in this study because it has the capacity to extend to

more complex situations in which population density varies across space. In this study, we

applied a quasi-Newton algorithm (Nocedal and Wright 2006) for maximizing log likelihood

function. Technical details in the model estimation was shown in **Appendix S2**.

Simulation study

Animal movement is temporally auto-correlated, and hence so are the locations of detections. To

test the capacity of ADCR to estimate parameters of connectivity and population density

correctly under a realistic movement process, we used a dataset generated via Monte Carlo

simulations of animal movement. The landscape is a rectangular grid system of 50×50 (hence,

the area is 2500), and animal location was described as a discrete coordinate of grid cells.

Arrays of count detectors were set to the range, $x = [9, 41]$ and $y = [9, 41]$, with x and y axis

spacing of 4 and 81 detectors.

Simulations were iterated 100 times, and true parameter values and landscape structure

were randomly determined in each iteration. For each iteration, a heterogeneous landscape was

generated using the R package rflsgen (Justeau-Allaire et al. 2022). Random patches were

generated in the landscape by rflsgen with parameters shown in Table S1. For covariate values,

0.5 and −0.5 were assigned to patch and non-patch habitats, respectively. True parameter values

were sampled from uniform distributions for each iteration. The ranges of true parameters were

set as log population density $\rho = (-3, -1)$, log intensity of drift $\alpha_0 = (-1, 0)$ and coefficient of

landscape variable $\beta = (-2, 2)$. g_0 and β_0 were fixed to −3 and 1, respectively. At the beginning

of an iteration, the number of individuals in the landscape was determined by a Poisson

distribution with expected value of $\exp(\rho) \times 2500$ and activity centers were drawn from a

discrete uniform distribution.

The animal movement simulation was run for $t = 5000$ after 100 burn-in steps for

randomizing initial locations of individuals. For each time step, we calculated the probability

vector of individual locations for all cells given the location at the previous step using the

advection–diffusion equation with the explicit method (eqn. 4 in Appendix S1), and the location

was sampled from a categorical distribution. When there was a detector in a cell in which an

individual was located, the number of detection events was simulated by a Poisson distribution

with expectation $\exp(g_0)$. Then, single-occasion capture–recapture data were obtained by

censoring undetected individuals. An example map of simulated movement and observations of

an individual is shown in Fig. 2.

We applied ADCR, SCR with the least-cost path, and basic SCR to the simulated

dataset to evaluate the accuracy of estimations. For the three estimators, the deviation of

estimates from true values was determined and root mean square error (RMSE) and mean bias

error (MBE) were calculated. To show how well ADCR and SCR with the least-cost path

reproduce home range shapes, the agreement between true (p^*_{true}) and predicted (p^*_{pred})

probability distributions of individual locations was calculated by an overlapping coefficient

defined by the volume of overlapping region between two probability distributions (Inman and

Bradley 1989). The overlapping coefficient ranges from 0 (no overlap) to 1 (perfect overlap). To

evaluate the average performance over locations for activity center μ, we calculated the

overlapping coefficients averaged over μ within the range of trap array x = [9, 41] and y = [9,

41] for each model and iteration.

All models were estimated using maximum likelihood methods with R 4.1.3 (R Core

Team 2022). The movement simulation and log likelihood function of ADCR were written

largely in R language, with some subroutines in C++. The source code is available online

(https://github.com/kfukasawa37/adcrtest2). The log likelihood values of SCR with the least

cost path were derived based on the R script by Sutherland et al. (2015), with some

modifications by Keita Fukasawa to deal with count detectors and the negative effect of

landscape on the cost function. The basic SCR was estimated using the R package secr (Efford,

2022).

Case study

We used a capture-recapture dataset of Asiatic black bear population in eastern

Toyama prefecture, Japan, collected over three years (2013-2015) by authors. The data was

obtained using video-recording camera traps at 86 locations with a field design for efficient

photographing of chest marks (Higashide et al. 2013) and photograph-based individual

recognitions (Higashide et al. 2012). Years were treated as occasions. The dataset was constituted

with the locations and year of detections for 109 individuals. For the details of the

capture-recapture survey, see **Appendix S3**.

In Japan, human-bear conflicts are increasing (Sakurai and Jacobson 2011).

Understanding population status (Horino and Miura 2000) and space use (Ohnishi et al. 2019)

of bears are crucial for developing plans to manage bear population and their threat to human

well-being. Population and landscape genetic studies suggested that rivers (Saitoh et al. 2001,

Ohnishi et al. 2007) and agricultural lands (Takahata et al. 2014, Ohnishi et al. 2019). To test the

effect of rivers and agricultural lands as barriers or corridors for black bears, we compiled

0.5km grid data of area ratio covered by water surface and agricultural lands (Fig. S1). These

areas were derived from the national vegetation map created by Biodiversity Center of Japan,

Ministry of Environment (https://www.biodic.go.jp/index_e.html, accessed 24 May 2024)

released in 2010s. We estimated ADCR including the two landscape covariates of connectivity

and compared the coefficients and estimated population densities.

Results

For all iterations, estimations of ADCR converged successfully. All parameters of ADCR,

including the population density and effect of landscape on connectivity, were estimated in an

unbiased way. The true and estimated parameters were on the line y = x, on average (Fig. 3).

RMSE and MBE of g_0 were 0.0258 and 0.0109, respectively. The true and predicted home

ranges overlapped fairly well. The overlapping coefficients marginalized over activity centers

were (0.950, 0.9996) over 100 iterations, and the mean was 0.988 (Fig. 4).

As predicted, population densities estimated by SCR with the least cost path and basic

SCR were robust to true connectivity settings in our simulation experiments. RMSEs and bias of

log population density estimated by SCR with the least cost path and basic SCR were very

similar to those of ADCR and were not sensitive to the true effect of landscape on connectivity

(Table 1). Effects of landscape on the cost distance estimated by SCR with the least cost path

were linear, on average, but showed a slightly weaker relationship with the true parameters of

connectivity (Pearson's $r = -0.988$, Fig. S2) than that of ADCR (Pearson's $r = 0.995$, Fig. 3).

The overlapping coefficients of home range predicted by SCR with the least cost path were

lower than those for ADCR over the range of true connectivity parameters, and the discrepancy

between predicted and true home range was large when the true effect of landscape on

connectivity was far from 0 (Fig. 4). SCR with the least cost path tended to predict a more

blurred and isotropic home range shape than the true home range shape (Fig. S3).

Standardized coefficients of agricultural land and water surface on the connectivity of

black bears were 0.361 (SE: 0.113; 95% CI: 0.175, 0.547) and -6.41 (SE: 1.26; 95% CI -8.88,

-3.94), respectively. A predicted home range showed a skewed shape, avoiding water surface

(Fig. S4). Estimated log population density was -1.22 (SE: 0.113; 95% CI: -1.41, -1.03).

Discussion

Individual movement is an important driver of spatial patterns of populations, and incorporating

movement processes in spatial capture–recapture analysis will facilitate studies of the interplay

between individual- and population-level ecological processes (Royle et al. 2018, McClintock et

al. 2022). Predicting connectivity is a focus of research in landscape ecology and crucial for

conservation planning (Dickson et al. 2019). The integration of a least cost path modeling with

spatial capture recapture analysis (Royle et al. 2013, Sutherland et al. 2015) was a great

methodological advancement for the simultaneous estimation of population density and

ecological distance, but accurate estimation of movement parameters exclusively from

capture–recapture data is still difficult (Dupont et al. 2022). In this study, we developed a formal

framework for a SCR model including mechanistic home range formation with an explicit

theoretical link to animal movement models. Through individual-based simulations of animal

movement and detection, we showed ADCR accurately estimated the population density and

effect of landscape. For the application of models to black bear dataset, ADCR could detect

clear environmental signals on the connectivity. These results suggest that ADCR is an effective

approach for estimating population density and random walk-based landscape connectivity.

Random walk-based circuit theory and least cost path are contrasting frameworks for

modeling landscape connectivity focusing on exploration under imperfect knowledge and

optimization with perfect knowledge of the landscape, respectively (Panzacchi et al. 2016).

Generally, animals have to search their foods and mates for reproduction under highly variable

and uncertain environment, and random walk is a central theoretical framework for analyzing

animal home range behavior (Smouse et al. 2010). The ability of ADCR to estimate the effects

of landscape variable on animal movement correctly (Fig. 3) is result from the formal link

between microscopic animal movement and the macroscopic utility distribution, corresponding

to Lagrangian and Eulerian descriptions of a system. Estimated effects of landscape variables on

diffusion coefficients by ADCR can be directly applied to predict relative conductance for

calculating resistance distance. In this study, ADCR provides an important tool for objective

reasoning in landscape connectivity studies.

The effect of a skewed home range on the accuracy of density estimation by SCR has

been studied extensively in ecology (Ivan et al. 2013, Royle et al. 2013, Sutherland et al. 2015,

Efford, 2019, Theng et al. 2022). The results of this study were consistent with those of recent

studies (Efford 2019, Theng et al. 2022) showing that density estimates are generally robust to

deviations in assumptions on the distribution of home range shapes. Although some studies

(Royle et al. 2013, Sutherland et al. 2015) have reported that density estimates by basic SCR

can be biased under the specific spatial patterns of landscape connectivity, our results indicated

that the bias in density averaged over a randomly generated landscape was small. Efford (2019)

emphasized the importance of sampling design in space, such as the use of a two-dimensional

detector array and avoidance of detector array alignment along landscape features, for the

correct estimation of density. Obviously, such sampling designs are also favorable for the

precise estimation of landscape effects on home range shape by ADCR.

For the case study of Asiatic black bears, the results from ADCR supported the

previous population genetic studies demonstrated rivers as physical barriers (Saitoh et al. 2001,

Ohnishi et al. 2007). Limited movement can lead genetic isolation of neighboring populations

on both sides of the river hence drive formation of macro-scale genetic structure. We also

detected the effect of farmlands as corridors of bear movement. As Takahata et al. (2014)

discussed, farmlands in the mountenious region with low population density tend to be fallowed

or abandoned (Osawa et al. 2016, Mameno and Kubo 2022), offering open habitat for black

bears. Farmlands within our camera trapping array are dispersed in mountanious areas (Fig. S1),

and they do not seem to serve as barriers against bears. For successful detection of the

biologically plausible signals on connectivity, the random walk-based home range model of

ADCR might be suitable for black bears which are omnivorous and oppotunistic foragers

(Landers et al. 1979, Hashimoto 2002, Bastille-Rousseau et al. 2011).

Joint statistical modeling of animal movement trajectories obtained by GPS loggers

and spatial capture–recapture data is an important approach for understanding detailed animal

movement processes, while ensuring the generality of results (Dupont et al. 2022). Both ADCR

and the step selection function are based on biased random walk models (Duchesne et al. 2015),

and ADCR can be applied to joint modeling, retaining consistency between the animal

movement model and resultant home range. Although the static home range assumption of

ADCR has an obvious advantage in the stability of estimates under a severe data deficiency,

intercepts of movement speed and site fidelity are aggregated into a single parameter and not

identifiable. Joint modeling of the step selection function and ADCR can overcome this

limitation by estimating these parameters separately.

The broad application of SCR in ecology can largely be attributed to its flexibility to

incorporate heterogeneity in ecological processes underlying spatial capture recapture data, such

as environment-dependent abundance and sex-biased detection (Efford, 2022). Future advances

in ADCR should focus on similar functionality. In particular, incorporating heterogeneity in

home range size (Efford et al., 2016) as well as heterogeneity in connectivity is a major

challenge. Although we focused only on heterogeneous connectivity for simplicity, the

simultaneous estimation of environmental dependence for both home range size and

connectivity can be achieved by incorporating γ_0 varying with the location of the activity center.

Also, the generalization of ADCR for dealing with temporally open populations driven by

birth-death processes (Gardner et al. 2010, Efford & Schofield 2022) would be helpful for

analyzing long-term capture–recapture data. The development of software tools that enable

users to benefit from the flexible nature of ADCR is an urgent task for expanding the application

of the method.

Similar to step selection analysis linked with advection–diffusion equations (Potts and

Schlägel 2020), an important assumption of ADCR is that sites of animal detection do not show

substantial temporal autocorrelation. In general, this assumption would be valid for

capture–recapture data in which the number of detections per individual is not large. However,

our approach can be biased if the detectability of individuals is quite high and animal steps are

highly correlated due to non-uniform turning angles. Modeling correlated movement with a

PDE formalism is still a challenging issue (Patlak 1953, Wang and Potts 2017), and future

research on this topic has the potential to improve home range representation under more

realistic situations.

In landscape ecological analyses, effects of choice of spatial resolution on the results

of study can often be problematic (Turner et al. 1989, Qi and Wu 1996). We found that ADCR

was robust to decreased spatial resolution of a landscape variable by supplementary simulatios

(Appendix S4). This robustness is a favorable characteristic of connectivity measures in circuit

theory, which shares the same theoretical foundation of random walk descriptions of animal

movement (Hanks and Hooten 2013).

Spatial capture–recapture is a popular method for abundance analyses in ecology and

conservation (Iijima 2020), and large amounts of data have accumulated. Moreover, recent

technical advances in individual identification from environmental DNA (Dugal et al. 2022) and

machine observations, such as camera traps (Schneider et al. 2022), will expand the

applicability of capture–recapture approaches to broad taxa. Such capture–recapture data have

the potential to advance our understanding of the process of animal home range formation.

ADCR will be a useful tool for integrated analyses of capture recapture data incorporating

individual- and population-level ecological processes.

Author contributions

415 KF conceived the ideas, designed the methodology, conducted the simulation study, conducted

416 field survey, and applied the model to actual dataset. DH conducted field survey, and complied

417 the dataset. KF led the writing and all authors contributed to the revision of the manuscript.

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Conflict of interest

425 The authors have no conflict of interest to be declared.

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Figures and tables

Fig. 1 Example of simulated utility distributions under various dependencies of

permeability on a hypothetical landscape variable. Black dots indicate cells with

high permeability; other cells have low permeability. The red X indicates the

activity center.

Fig. 2 Example of simulated movement and the detection of an individual. The

transparent solid line is the movement path. Open rectangles are traps, and blue

circles indicate detection counts. The background tile map indicates zones with

high (green) and low (pale red) permeability.

Fig. 3 True vs. estimated parameters by ADCR. Dots and error bars indicate maximum

likelihood estimates and 95% confidence intervals. The solid black line is $y = x$.

Fig. 4 Overlap between true and predicted home ranges by ADCR and SCR with the

least-cost path.

Fig. 5 Estimated effect of landscape on connectivity by ADCR for black bear dataset.

Error bars are 95% CIs.

Table 1 RMSE and MBE of log density for the three estimators for the baseline scenario. The values binned into three classes of the true landscape

effect, β, were also shown.

	RMSE			MBE		
	ADCR	SCR with the least-cost path	basic SCR	ADCR	SCR with the least-cost path	basic SCR
Total	0.0788	0.0789	0.0822	-0.00606	0.00123	0.0125
$\beta < -0.5$	0.0752	0.0762	0.0795	-0.0205	-0.0102	-0.00296
$|\beta| <= 0.5$	0.0677	0.0676	0.0730	0.0130	0.0139	0.0333
$\beta > 0.5$	0.0890	0.0883	0.0906	-0.00499	0.00380	0.0133

Supporting information

610		
611	**Appendix S1**	Technical details in the finite difference method.
612	**Appendix S2**	Technical details in model estimation with maximum likelihood methods.
613	**Appendix S3**	Detailed descriptions on capture-recapture survey data used for case study.
614	**Appendix S4**	The effect of spatial resolution on parameter estimates.
615	**Appendix S5**	Supplementary figures.
616		**Fig. S1**
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Appendix S1 Technical details in the finite difference method.

Appendix S2 Technical details in model estimation with maximum likelihood methods.

Appendix S3 Detailed descriptions on capture-recapture survey data used for case study.

Appendix S4 The effect of spatial resolution on parameter estimates.

Appendix S5 Supplementary figures.

Fig. S1 (a) Study area, and (b) locations of camera traps (red
rectangle), agricultural lands (orange) and water surface
(blue), and 0.5km grid cells.

Fig. S2 Relationship between the true effect of landscape on
permeability and estimated effect of landscape on cost by
SCR with the least-cost path. Error bars are 95% CIs.

Fig. S3 Example of true and predicted home ranges by ADCR and
SCR with the least-cost path. Red x symbols mark the
home range center. Black dots are zones with high
permeability.

Fig. S4 An example of predicted home ranges by ADCR for the
black bear. Red x symbols mark the home range center.
Black dots are cells contains water surface.

Fig. S5 Comparison between the estimates of baseline simulations
and 1/2 resolution scenarios for (a) effects of landscape
and (b) log population densities by ADCR. Dots and error
bars are maximum likelihood estimates and 95% CI,
respectively. Glay lines are principal components of
parameter estimates.

Appendix S6 Supplementary table.

Table S1 Target parameter ranges for the neutral landscape generator,
rflsgen. The parameter names were based on
Justeau-Allaire et al. (2022). Parameters not shown were
set to default values.

Fig. 1 Example of simulated utility distributions under various dependencies of permeability on a hypothetical landscape variable. Black dots indicate cells with high permeability; other cells have low permeability. The red X indicates the activity center.

$\beta = 0$
(circular home range)

Large β
(strong dependency on the landscape variable)

Fig. 2 Example of simulated movement and the detection of an individual. The transparent solid line is the movement path. Open rectangles are traps, and blue circles indicate detection counts. The background tile map indicates zones with high (green) and low (pale red) permeability.

Fig. 3 True vs. estimated parameters by ADCR. Dots and error bars indicate maximum likelihood estimates and 95% confidence intervals. The solid black line is y = x.

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