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An Ensemble Penalized Regression Method for Multi-ancestry Polygenic Risk Prediction

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23 **Abstract**

24 Great efforts are being made to develop advanced polygenic risk scores (PRS) to improve the
25 prediction of complex traits and diseases. However, most existing PRS are primarily trained on
26 European ancestry populations, limiting their transferability to non-European populations. In
27 this article, we propose a novel method for generating multi-ancestry Polygenic Risk scores
28 based on ensemble of Penalized Regression models (PROSPER). PROSPER integrates genome-
29 wide association studies (GWAS) summary statistics from diverse populations to develop
30 ancestry-specific PRS with improved predictive power for minority populations. The method
31 uses a combination of \mathcal{L}_1 (lasso) and \mathcal{L}_2 (ridge) penalty functions, a parsimonious specification
32 of the penalty parameters across populations, and an ensemble step to combine PRS generated
33 across different penalty parameters. We evaluate the performance of PROSPER and other
34 existing methods on large-scale simulated and real datasets, including those from 23andMe
35 Inc., the Global Lipids Genetics Consortium, and All of Us. Results show that PROSPER can
36 substantially improve multi-ancestry polygenic prediction compared to alternative methods
37 across a wide variety of genetic architectures. In real data analyses, for example, PROSPER
38 increased out-of-sample prediction R^2 for continuous traits by an average of 70% compared to a
39 state-of-the-art Bayesian method (PRS-CSx) in the African ancestry population. Further,
40 PROSPER is computationally highly scalable for the analysis of large SNP contents and many
41 diverse populations.

42

43

44 **Introduction**

45

46 Tens of thousands of single nucleotide polymorphisms (SNP) have been mapped to human
47 complex traits and diseases through genome-wide association studies (GWAS) ^{1,2}. Though each
48 SNP only explains a small fraction of variation of the underlying phenotype, polygenetic risk
49 scores (PRS), which aggregate the genetic effects of many loci, can have a substantial ability to
50 predict traits and stratify populations by underlying disease risks ³⁻¹². However, as existing
51 GWAS to date have been primarily conducted in European ancestry populations (EUR) ¹³⁻¹⁶,
52 recent studies have consistently shown that the transferability of EUR-derived PRS to non-EUR
53 populations often is less than ideal and in particular poor for African Ancestry populations ¹⁷⁻²².

54

55 Despite growing efforts of conducting genetic research on minority populations ²³⁻²⁶, the gap in
56 sample sizes between EUR and non-EUR populations is likely to persist in the foreseeable
57 future. As the performance of PRS largely depends on the sample size of training GWAS ^{3,27},
58 using single ancestry methods ²⁸⁻³² to generate PRS for a minority population, using data from
59 that population alone may not achieve ideal results. To address this issue, researchers have
60 developed methods for generating powerful PRS by borrowing information across diverse
61 ancestry populations. For example, Weighted PRS ³³ combines single-ancestry PRS generated
62 from each population using weights that optimize performance for a target population.
63 Bayesian methods have also been proposed that generate improved PRS for each population by
64 jointly modeling the effect-size distribution across populations ^{34,35}. Recently, our group

65 proposed a new method named CT-SLEB²², which extends the clumping and thresholding (CT)
66 ³⁶ method to multi-ancestry settings. The method uses an empirical-Bayes (EB) approach to
67 estimate effect sizes by borrowing information across populations and a super learning model
68 to combine PRSs under different tuning parameters. However, the optimality of the methods
69 depends on many factors, including the ability to account for heterogeneous linkage
70 disequilibrium (LD) structure across populations and the adequacy of the models for underlying
71 effect-size distribution^{3,27}. In general, our extensive simulation studies and data analyses
72 suggest that no method is uniformly the most powerful, and exploration of complementary
73 methods will often be needed to derive the optimal PRS in any given setting²².

74

75 In this article, we propose a novel method for generating multi-ancestry Polygenic Risk scores
76 based on an ensemble Penalized Regression (PROSPER) using GWAS summary statistics and
77 validation datasets across diverse populations. The method incorporates \mathcal{L}_1 penalty functions
78 for regularizing SNP effect sizes within each population, an \mathcal{L}_2 penalty function for borrowing
79 information across populations, and a flexible but parsimonious specification of the underlying
80 penalty parameters to reduce computational time. Further, instead of selecting a single optimal
81 set of tuning parameters, the method combines PRS generated across different populations and
82 tuning parameters using a final ensemble regression step. We compare the predictive
83 performance of PROSPER with a wide variety of single- and multi-ancestry methods using
84 simulation datasets from our recent study²² across five populations (EUR, African (AFR),
85 American (AMR), East Asian (EAS), and South Asian (SAS))²². Furthermore, we evaluate these
86 methods using a variety of real datasets from 23andMe Inc. (23andMe), the Global Lipids

87 Genetics Consortium (GLGC)³⁷, All of Us (AoU)³⁸, and the UK Biobank study (UKBB)³⁹. Results
88 from these analyses indicate that PROSPER is a highly promising method for generating the
89 most powerful multi-ancestry PRS across diverse types of complex traits. Computationally,
90 PROSPER is also exceptionally scalable compared to other advanced methods.

91

92 **Results**

93

94 **Method overview**

95

96 PRSOSPER is a method designed to improve prediction performance for PRS across distinct
97 ancestral populations by borrowing information across ancestries (**Figure 1**). It can integrate
98 large EUR GWAS with smaller GWAS from non-EUR populations. Ideally, individual-level tuning
99 data are needed for all populations, because the method needs optimal parameters from
100 single-ancestry analysis as an input; however, even when data is only available for a target
101 population, PRSOSPER can still be performed, and the PRS will be optimized and validated
102 towards the target population. The method can account for population-specific genetic
103 variants, allele frequencies, and LD patterns and use computational techniques for penalized
104 regressions for fast implementation.

105

106 *PROSPER*

107

108 Assuming a continuous trait, we first consider a standard linear regression model for underlying
 109 individual-level data for describing the relationship between trait values and genome-wide
 110 genetic variants across M distinct populations. Let \mathbf{Y}_i denote the $n_i \times 1$ vector of trait values,
 111 \mathbf{X}_i denote the $n_i \times p_i$ genotype matrix, $\boldsymbol{\beta}_i$ denote the $p_i \times 1$ vector of SNP effects, and $\boldsymbol{\epsilon}_i$
 112 denote the $n_i \times 1$ vector of random errors for the i^{th} population. We assume underlying linear
 113 regression models of the form $\mathbf{Y}_i = \mathbf{X}_i \boldsymbol{\beta}_i + \boldsymbol{\epsilon}_i, i = 1, \dots, M$; and intend to solve the linear
 114 regression system by least square with a combination of \mathcal{L}_1 (lasso)⁴⁰ and \mathcal{L}_2 (ridge)⁴¹ penalties
 115 in the form

$$116 \quad \sum_{1 \leq i \leq M} \frac{1}{n_i} (\mathbf{Y}_i - \mathbf{X}_i \boldsymbol{\beta}_i)^T (\mathbf{Y}_i - \mathbf{X}_i \boldsymbol{\beta}_i) + \sum_{1 \leq i \leq M} 2\lambda_i \|\boldsymbol{\beta}_i\|_1 + \sum_{1 \leq i_1 < i_2 \leq M} c_{i_1 i_2} \left\| \boldsymbol{\beta}_{i_1}^{s_{i_1 i_2}} - \boldsymbol{\beta}_{i_2}^{s_{i_1 i_2}} \right\|_2^2$$

117 where $\lambda_i, i = 1, \dots, M$ are the population-specific tuning parameters associated with the lasso
 118 penalty; $\boldsymbol{\beta}_{i_1}^{s_{i_1 i_2}}$ and $\boldsymbol{\beta}_{i_2}^{s_{i_1 i_2}}$ denote the vectors of effect-sizes for SNPs for the i_1 -th and i_2 -th
 119 populations, respectively, restricted to the set of shared SNPs ($s_{i_1 i_2}$) across the pair of the
 120 populations; and $c_{i_1 i_2}, 1 \leq i_1 < i_2 \leq M$ are the tuning parameters associated with the ridge
 121 penalty imposing effect-size similarity across pairs of populations.

122
 123 In the above, the first part, $\sum_{1 \leq i \leq M} 2\lambda_i \|\boldsymbol{\beta}_i\|_1$, uses a lasso penalty. Lasso can produce sparse
 124 solution⁴⁰ and recent PRS studies that have implemented the lasso penalty in the single-
 125 ancestry setting have shown its promising performance^{29, 30}. The second part,

126 $\sum_{1 \leq i_1 < i_2 \leq M} c_{i_1 i_2} \left\| \boldsymbol{\beta}_{i_1}^{s_{i_1 i_2}} - \boldsymbol{\beta}_{i_2}^{s_{i_1 i_2}} \right\|_2^2$, uses a ridge penalty. As it has been widely shown that the

127 causal effect sizes of SNPs tend to be correlated across populations^{42, 43}, we propose to use the
 128 ridge penalty to induce genetic similarity across populations. Compared to the fused lasso⁴⁴,

129 which uses lasso penalty for the differences, we use ridge penalty instead, which allows a small
130 difference in SNP effects across populations rather than truncating them to zero. In addition,
131 the ridge penalty is also computationally more efficient due to its continuous derivative. The
132 solutions for population-specific effect size using the combined lasso and ridge penalties can be
133 sparse.

134

135 The estimate of $\beta_i, i = 1, \dots, M$ in the above individual-level linear regression systems can be
136 obtained by minimizing the above least square objective function. Following the derivation of
137 lassosum²⁹, a single-ancestry method for fitting the lasso model to GWAS summary statistics
138 data, we show that the objective function for individual-level data can be approximated using
139 GWAS summary statistics and LD reference matrices by substituting $\frac{1}{n_i} \mathbf{X}_i^T \mathbf{X}_i$ by \mathbf{R}_i , where \mathbf{R}_i is
140 the estimated LD matrix based on a reference sample from the i -th population, and $\frac{1}{n_i} \mathbf{X}_i^T \mathbf{y}_i$ by
141 \mathbf{r}_i , where \mathbf{r}_i is the GWAS summary statistics in the i -th population. Therefore, the objective
142 function of the summary-level model can be written as

143
$$\sum_{1 \leq i \leq M} (\beta_i^T (\mathbf{R}_i + \delta_i \mathbf{I}) \beta_i - 2 \beta_i^T \mathbf{r}_i + 2 \lambda_i \|\beta_i\|_1) + \sum_{1 \leq i_1 < i_2 \leq M} c_{i_1 i_2} \|\beta_{i_1}^{S_{i_1 i_2}} - \beta_{i_2}^{S_{i_1 i_2}}\|_2^2$$

144 where the additional tuning parameters $\delta_i, i = 1, \dots, M$, are introduced for regularization of
145 the LD matrices across the different populations³⁰. For a fixed set of tuning parameters, the
146 above objective function can be solved using fast coordinate descent algorithms⁴⁵ by iteratively
147 updating each element of $\beta_i, i = 1, \dots, M$ (see the section of **Obtain PROSPER solution in**
148 **Methods**).

149

150 *Reducing tuning parameters*

151

152 For the selection of tuning parameters, we assume we have access to individual-level data

153 across the different populations which are independent of underlying GWAS from which

154 summary statistics are generated. The above setting involves three sets of tuning parameters,

155 $\{\delta_i\}_{i=1}^M$, $\{\lambda_i\}_{i=1}^M$, and $\{c_{i_1 i_2}\}_{1 \leq i_1 < i_2 \leq M}$, totaling to the number of $M + M + \frac{M(M-1)}{2}$. As grid search

156 across many combinations of tuning parameter values can be computationally intensive, we

157 propose to reduce the search range by a series of steps. First, we use `lassosum2`³⁰ to analyze

158 GWAS summary statistics and tuning data from each ancestry population by itself and obtain

159 underlying values of optimal tuning parameters, $(\delta_i^0, \lambda_i^0)$ for $i = 1, \dots, M$; if tuning data is only

160 available for the target population, the $(\delta_i^0, \lambda_i^0)$ for non-target i can be optimized towards the

161 target population. For fitting PROSPER, we fix $\delta_i = \delta_i^0$ for $i = 1, \dots, M$ as these are essentially

162 used to regularize estimates of population-specific LD matrices. We note that the optimal

163 $\{\lambda_i\}_{i=1}^M$ depend on sample sizes of underlying GWAS (**Supplementary Figure 1**), and thus should

164 not be arbitrarily assumed to be equal across all populations. Considering that the optimal

165 tuning parameters associated with the \mathcal{L}_1 penalty function from the single-ancestry analyses

166 should reflect the characteristics of GWAS data, which includes underlying sparsity of effect

167 sizes and sample sizes, we propose to specify the \mathcal{L}_1 -tuning parameters in PROSPER as $\lambda_i =$

168 $\lambda \lambda_i^0$, i.e. they are determined by the corresponding tuning parameters from the ancestry-

169 specific analysis except for the constant multiplicative factor λ . Finally, for further

170 computational simplification, we assume that effect sizes across all pairs of populations have a

171 similar degree of homogeneity and thus set all $\{c_{i_1 i_2}\}_{1 \leq i_1 < i_2 \leq M}$ to be equal to c . By using the

172 above assumptions, the objective function to minimize with respect to $\beta_i, i = 1, \dots, M$,
173 becomes

$$174 \quad \sum_{1 \leq i \leq M} (\beta_i^T (R_i + \delta_i^0 I) \beta_i - 2\beta_i^T r_i + 2\lambda \lambda_i^0 \|\beta_i\|_1) + \sum_{1 \leq i_1 < i_2 \leq M} c \|\beta_{i_1}^{s_{i_1 i_2}} - \beta_{i_2}^{s_{i_1 i_2}}\|_2^2$$

175 where λ and c are the only two tuning parameters needed for lasso penalty and genetic
176 similarity penalty, respectively.

177

178 *Ensemble*

179

180 Using an ensemble method to combine PRS has been shown to be promising in CT-type
181 methods as opposed to picking an optimal threshold^{22, 36}. In general, a specific form of the
182 penalty function, or equivalently a model for prior distribution in the Bayesian framework, may
183 not be able to adequately capture the complex nature of the underlying distribution of the
184 SNPs across diverse populations. We conjecture that when effect size distribution is likely to be
185 mis-specified, an ensemble method, which combines PRS across different values of tuning
186 parameters instead of choosing one optimal set, is likely to improve prediction. Therefore, as a
187 last step, we obtain the final PROSPER model using an ensemble method, super learning⁴⁶⁻⁴⁸,
188 implemented in the *SuperLearner* R package, to combine PRS generated from various tuning
189 parameter settings and optimized using tuning data from the target population. The super
190 learner we use here was based on three supervised learning algorithms, including lasso⁴⁰, ridge
191⁴¹, and linear regression (see **Methods**).

192

193 Results

194

195 *Methods comparison on simulated data*

196

197 We conducted simulation analyses on continuous traits under various genetic architectures²²

198 to evaluate the performance of different methods that can be categorized into five groups:

199 single-ancestry methods trained from target GWAS data (single-ancestry method), single-

200 ancestry methods trained from EUR GWAS data (EUR PRS based method), simple multi-ancestry

201 methods by weighting single-ancestry PRS (weighted PRS), recently published multi-ancestry

202 methods (existing multi-ancestry methods), and our proposed method, PROSPER. Single-

203 ancestry methods include CT³⁶, LDpred2³¹, and lassosum2³⁰. Existing multi-ancestry methods

204 include PRS-CSx³⁴ and CT-SLEB²². The performance of the methods is evaluated by R^2

205 measured on validation samples independent of training and tuning datasets. Analyses in this

206 and the following sections are restricted to a total of 2,586,434 SNPs, which are included in

207 either HapMap 3 (HM3)⁴⁹ or the Multi-Ethnic Genotyping Arrays (MEGA) chips array⁵⁰. LD

208 reference samples for all five ancestries, EUR, AFR, AMR, EAS, and SAS, in this and the following

209 sections, are from 1000 Genomes Project (Phase 3)⁵¹ (1000G).

210

211 The results (**Figure 2, Supplementary Figure 2-6, Supplementary Table 1.1-1.5**) show that

212 multi-ancestry methods generally exhibit superior performance compared to single-ancestry

213 methods. Weighted PRS generated from methods modeling LD (LDpred2 and lassosum2) can

214 lead to a noticeable improvement in performance (green bars in **Figure 2**). Notably, PROSPER

215 shows robust performance uniformly across different scenarios. When the sample size of the
216 target non-EUR population is small ($N_{target} = 15K$) (**Figure 2a**), PROSPER has comparable
217 performance with other multi-ancestry methods under a high degree of polygenicity ($p_{causal} =$
218 0.01). However, under the same sample size setting and lower ($p_{causal} = 0.01$ and 5×10^{-4}),
219 PRS-CSx and CT-SLEB outperform PROSPER, with the margin of improvement increasing as the
220 strength of negative selection decreases (strong negative selection in **Figure 2a**, mild strong
221 negative selection in **Supplementary Figure 2a**, and no negative selection in **Supplementary**
222 **Figure 3a**). When the sample size of the target population is large ($N_{target} = 80K$) (**Figure 2b**,
223 and **Supplementary Figure 2-5 b**), PROSPER almost uniformly outperforms all other methods,
224 particularly for the AFR population.

225

226 We further compare the computational efficiency of PROSPER in comparison to PRS-CSx, the
227 state-of-the-art Bayesian method available for generating multi-ancestry PRS. We train PRS
228 models for the two methods using simulated data for chromosome 22 using a single core with
229 AMD EPYC 7702 64-Core Processors running at 2.0 GHz. We observe (**Supplementary Table 2**)
230 that PROSPER is 37 times faster than PRS-CSx (3.0 vs. 111.1 minutes) in a two-ancestry analysis
231 including AFR and EUR; and 88 times faster (6.8 vs. 595.8 minutes) in the analysis of all five
232 ancestries. The memory usage for PRS-CSx is about 2.8 times smaller than PROSPER (0.78 vs.
233 2.24 Gb in two-ancestry analysis, and 0.84 vs. 2.35 Gb in five-ancestry analysis).

234

235 *23andMe data analysis*

236

237 We applied various methods to GWAS summary statistics available from the 23andMe, Inc. to
238 predict two continuous traits, heart metabolic disease burden and height; as well as five binary
239 traits, any cardiovascular disease (any CVD), depression, migraine diagnosis, morning person,
240 and sing back musical note (SBMN). The datasets are available for all five ancestries, African
241 American (AA), Latino, EAS, EUR, and SAS. The methods are tuned and validated on a set of
242 independent individuals of the corresponding ancestry from the 23andMe participant cohort
243 (see the section of **Real data analysis** in **Methods** for data description, and **Supplementary**
244 **Table 3-4** for sample sizes used in training, tuning and validation).

245
246 From the analysis of two continuous traits (**Figure 3** and **Supplementary Table 5.1**), we observe
247 that lassosum2 and its related methods (EUR lassosum2 and weighted lassosum2) generally
248 perform better than CT and LDpred2, and their related methods. On the basis of the advantage
249 of lassosum2, PROSPER further improves the performance, and for most of the settings,
250 outperforms all alternative methods, including PRS-CSx and CT-SLEB. PROSPER demonstrates
251 particularly remarkable improvement for both traits in AA and Latino (26.9 % relative
252 improvement in R^2 over the second-best method on average, yellow cells in **Supplementary**
253 **Table 5.2**) (first two panels in **Figure 3a-b**). For EAS and SAS, PROSPER is slightly better than
254 other methods, except for heart metabolic disease burden of SAS (the last panel in **Figure 3a**),
255 which has the smallest sample size (~20K).

256
257 The results from the analysis of the binary traits (**Figure 4** and **Supplementary Table 5.1**) show
258 that PROSPER generally exhibits better performance (7.8% and 12.3% relative improvement in

259 logit-scale variance (see **Supplementary Notes**) over CT-SLEB and PRS-CSx, respectively,
260 averaged across populations and traits) (blue and red cells, respectively, in **Supplementary**
261 **Table 5.2**). A similar trend is observed for the analyses of AA and Latino, where PROSPER
262 usually has the best performance (first two panels in **Figure 4a-e**). In general, no single method
263 can uniformly outperform others. Weighted lassosum2 has outstanding performance for
264 depression (**Figure 4b**), while PROSPER is superior for morning person (**Figure 4d**). PRS-CSx
265 shows a slight improvement in the analysis of migraine diagnosis for EAS populations (last
266 second panel in **Figure 4c**), and CT-SLEB performs the best in the analysis of any CVD for SAS
267 population (last panel in **Figure 4a**).

268

269 *GLGC and AoU data analysis*

270

271 Considering the uncommonly huge sample sizes from 23andMe, we further applied alternative
272 methods for the analysis of two other real datasets, GLGC and AoU. The GWAS summary
273 statistics from GLGC for four blood lipid traits, high-density lipoprotein (HDL), low-density
274 lipoprotein (LDL), log-transformed triglycerides (logTG), and total cholesterol (TC), are publicly
275 downloadable and available for all five ancestries, African/Admixed African, Hispanic, EAS, EUR,
276 and SAS (see **Methods** for data description, and **Supplementary Table 3** for sample sizes).
277 Further, we generated GWAS summary statistics data from the AoU study for two
278 anthropometric traits, body mass index (BMI) and height, for individuals from three ancestries,
279 AFR, EUR, and Latino/Admixed American (see **Methods** for data description, and
280 **Supplementary Table 3** for sample sizes). Both the blood lipid traits and anthropometric traits

281 have corresponding phenotype data available in the UKBB, which we use to perform tuning and
282 validation (see the section of **Real data analysis** in **Methods** for the ancestry composition, and
283 **Supplementary Table 4** for sample sizes). Given the limited sample sizes of genetically inferred
284 AMR ancestry individuals in UKBB, we do not report the performance of PRS on AMR
285 individuals in UKBB.

286

287 Results from analysis of four blood lipid traits (**Figure 5** and **Supplementary Table 6.1**) from
288 GLGC and UKBB show that PRS generated by lasso-type methods substantially outperform
289 alternative methods. In particular, we observe that the weighted lassosum2 always
290 outperforms the other two weighted methods. Furthermore, our proposed method, PROSPER,
291 shows improvement over weighted lassosum2 in both AFR and SAS (13.1% and 12.3% relative
292 improvement in R^2 , respectively, averaged across traits) (green and orange cells, respectively, in
293 **Supplementary Table 6.2**), but not in EAS. Notably, PROSPER outperforms PRS-CSx and CT-SLEB
294 in most scenarios (34.2% and 37.7% relative improvement in R^2 , respectively, averaged across
295 traits and ancestries) (blue and red cells, respectively, in **Supplementary Table 6.2**), with the
296 improvement being particularly remarkable for the AFR population in which PRS development
297 tends to be the most challenging.

298

299 The results from AoU and UKBB (**Figure 6** and **Supplementary Table 7.1**) show that PROSPER
300 generates the most predictive PRS for the two analyzed anthropometric traits for the AFR
301 population. It appears that Bayesian and penalized regression methods that explicitly model LD
302 tend to outperform corresponding CT-type methods (CT, EUR CT, and weighted CT) which

303 excluded correlated SNPs. Among weighted methods, both LDpred2 and lassosum2 show major
304 improvement over the corresponding CT method. Further, for both traits, PROSPER shows
305 remarkable improvement over the best of the weighted methods and the two other advanced
306 methods, PRS-CSx and CT-SLEB (91.3% and 76.5% relative improvement in R^2 , respectively,
307 averaged across the two traits) (blue and red cells, respectively, in **Supplementary Table 7.2**).

308

309 **Discussion**

310

311 In this article, we propose PROSPER as a powerful method that can jointly model GWAS
312 summary statistics from multiple ancestries by an ensemble of penalized regression models to
313 improve the performance of PRS across diverse populations. We show that PROSPER is a
314 uniquely promising method for generating powerful PRS in multi-ancestry settings through
315 extensive simulation studies, analysis of real datasets across a diverse type of complex traits,
316 and considering the most recent developments of alternative methods. Computationally, the
317 method is an order of magnitude faster compared to PRS-CSx³⁴, an advanced Bayesian method,
318 and comparable to CT-SLEB²², which derives the underlying PRS in closed forms. We have
319 packaged the algorithm into a command line tool based on the R programming language
320 (<https://github.com/Jingning-Zhang/PROSPER>).

321

322 We compare PROSPER with a number of alternative simple and advanced methods using both
323 simulated and real datasets. The simulation results show that PROSPER generally outperforms
324 other existing multi-ancestry methods when the target sample size is large (**Figure 2b**).

325 However, when the sample size of the target population is small (**Figure 2a**), no method
326 performed uniformly the best. In this setting, when the degree of polygenicity is the lowest
327 ($p_{causal} = 5 \times 10^{-4}$), CT-SLEB outperforms other methods by a noticeable margin, and
328 PROSPER performs slightly worse than PRS-CSx. Simulations also show that in the scenario of a
329 highly polygenic trait ($p_{causal} = 0.01$), irrespective of sample size, both weighted lasso2
330 and PROSPER tend to enjoy superiority compared to all other methods. In terms of
331 computational time and memory usage, PROSPER is an order of magnitude faster than PRS-CSx in a
332 five-ancestry analysis. The memory usage for PRS-CSx is smaller than PROSPER, but both are
333 acceptable (**Supplementary Table 2**).

334

335 We observe that for the analysis of both continuous and binary traits using 23andMe Inc. data,
336 PROSPER demonstrates a substantial advantage over all other methods for the AA and Latino
337 populations, which have the largest sample sizes among all minority groups. The result is
338 consistent with the superior performance of PROSPER observed in simulation settings when the
339 sample size of the target population is large. However, it is worth noting that even for the two
340 other populations, EAS and SAS, which have much smaller sample sizes, PROSPER still performs
341 the best in half of the settings (the last two panels in **Figure 3a-b** and **Figure 4a-e**). For the
342 prediction of blood lipid traits, methods built upon the lasso penalty (lasso2, weighted
343 lasso2, PROSPER) perform substantially better than all other alternative methods.
344 Intuitively, this might result from the robustness of the heavy-tail lasso penalty function in
345 dealing with large-effect loci that tend to be present for molecular traits, such as lipid levels
346 (**Supplementary Table 8**), and sometimes for complex traits as well. For the analysis of two

347 anthropometric traits using training data from AoU, we observe that methods that explicitly
348 model and account for LD differences (e.g. lassosum2, LDpred2, and their corresponding
349 weighted methods) generally achieve higher predictive accuracy than CT-based methods which
350 discard correlated SNPs. In addition, we observe major improvement in PRS performance using
351 PROSPER over weighted lassosum2 and all other existing multi-ancestry methods. The result is
352 consistent with what we have observed in simulation settings under extreme polygenic
353 architectures as expected for complex traits like height and BMI. In conclusion, our results show
354 that PROSPER is a promising method for handling complex traits of diverse genetic
355 architectures.

356

357 PROSPER, while showing promising results in our simulations and real data analyses, does have
358 several limitations. Specifically, when the sample size for the training sample for a target
359 population is small, particularly for traits with low polygenicity, the method may not perform as
360 well as some of the other existing methods (**Figure 2a**). Additionally, the use of a super learning
361 step in PROSPER can lead to poorer performance compared to weighted lassosum2 when the
362 sample size for the tuning dataset is not adequately large. In the analysis of lipid traits for EAS,
363 for example, we observe lower predictive accuracy of PROSPER than weighted lassosum2 (the
364 middle panel in **Figure 5b** and **d**). This can be attributed to overfitting in the tuning sample, as
365 the number of tuning samples of EAS origin in the UKBB is only ~1000, while the number of
366 PRSs combined in the super learning step is close to 500.

367

368 PROSPER and a number of other recent methods have been developed for modeling summary
369 statistics data across discrete populations typically defined by self-reported ancestry
370 information. However, there is an emerging need to consider the underlying continuum of
371 genetic diversity across populations in both the development and implementational of PRS in
372 diverse populations in the future ⁵². Towards this goal, a recent method called GAUDI ⁵³ has
373 been proposed based on the fused lasso penalty for developing PRS in admixed population
374 using individual-level data. While GAUDI shares similarities with PROSPER in terms of the use of
375 the lasso-penalty function, the two methods are distinct in terms of the specification of tuning
376 parameters and use of the ensemble step. Future studies are merited to extend PROSPER for
377 handling data with continuous genetic ancestry information.

378

379 To conclude, we have proposed PROSPER, a statistically powerful and computationally scalable
380 method for generating multi-ancestry PRS using GWAS summary statistics and additional tuning
381 and validation datasets across diverse populations. While no method is uniformly powerful in
382 all settings, we show that PROSPER is the most robust among a large variety of recent methods
383 proposed across a wide variety of settings. As individual-level data from GWAS of diverse
384 populations becomes increasingly available, PROSPER and other methods will require additional
385 considerations for incorporating continuous genetic ancestry information, both global and local,
386 into the underlying modeling framework.

387

388

389 **Author Contribution Statement**

390 J.Zhang and NC conceived the project. J.Zhang, J.Zhan, JJ, and HZ carried out all data analyses
391 with supervision from NC; HZ created all simulated data and ran GWAS on simulated training
392 data with the supervision from NC; J.Zhan, JOC, YJ run GWAS for training data from 23andMe
393 Inc. with the supervision from BLK; RZ ran GWAS on AoU training data with the supervision
394 from NC and HZ; J.Zhang and CM developed the PROSPER software; J.Zhang and NC drafted the
395 manuscript, and HZ, JJ provided comments. All co-authors reviewed and approved the final
396 version of the manuscript. The following members of the 23andMe Research Team contributed
397 to this study: Stella Aslibekyan, Adam Auton, Elizabeth Babalola, Robert K. Bell, Jessica
398 Bielenberg, Katarzyna Bryc, Emily Bullis, Daniella Coker, Gabriel Cuellar Partida, Devika Dhamija,
399 Sayantan Das, Sarah L. Elson, Nicholas Eriksson, Teresa Filshtein, Alison Fitch, Kipper Fletez-
400 Brant, Pierre Fontanillas, Will Freyman, Julie M. Granka, Karl Heilbron, Alejandro Hernandez,
401 Barry Hicks, David A. Hinds, Ethan M. Jewett, Yunxuan Jiang, Katelyn Kukar, Alan Kwong, Keng-
402 Han Lin, Bianca A. Llamas, Maya Lowe, Jey C. McCreight, Matthew H. McIntyre, Steven J.
403 Micheletti, Meghan E. Moreno, Priyanka Nandakumar, Dominique T. Nguyen, Elizabeth S.
404 Noblin, Jared O'Connell, Aaron A. Petrakovitz, G. David Poznik, Alexandra Reynoso, Morgan
405 Schumacher, Anjali J. Shastri, Janie F. Shelton, Jingchunzi Shi, Suyash Shringarpure, Qiaojuan
406 Jane Su, Susana A. Tat, Christophe Toukam Tchakouté, Vinh Tran, Joyce Y. Tung, Xin Wang, Wei
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408

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427

428 **Code and Data availability**

429 PROSPER command line tool: <https://github.com/Jingning-Zhang/PROSPER>

430 CT: <https://www.cog-genomics.org/plink/1.9/>

431 Lassosum2 and LDpred2: <https://github.com/privefl/bigsnpr>

432 PRS-CSx: <https://github.com/getian107/PRScsx>

433 CT-SLEB: <https://github.com/andrewhaoyu/CTSLEB>

434 PLINK: <https://www.cog-genomics.org/plink/1.9/>; <https://www.cog-genomics.org/plink/2.0/>

435 Simulated genotype data for 600K subjects from five ancestries:

436 <https://dataverse.harvard.edu/dataset.xhtml?persistentId=doi:10.7910/DVN/COXHAP>

437 The full GWAS summary statistics for the 23andMe discovery data set could be made available

438 through 23andMe to qualified researchers under an agreement with 23andMe that protects

439 the privacy of the 23andMe participants. Please visit

440 <https://research.23andme.com/collaborate/#dataset-access/> for more information and to

441 apply to access the data. Participants provided informed consent and participated in the

442 research online, under a protocol approved by the external AAHRPP-accredited IRB, Ethical &

443 Independent Review Services.

444 GWAS summary level statistics for five ancestries from GLGC:

445 http://csg.sph.umich.edu/willer/public/glgc-lipids2021/results/ancestry_specific/

446 GWAS summary level statistics for three ancestries from AoU are available upon request.

447 Codes for simulation and data analyses in this paper: <https://github.com/Jingning->

448 [Zhang/PROSPER_analysis](#)

449 The full GWAS summary statistics for the 23andMe discovery data set could be made available

450 through 23andMe to qualified researchers under an agreement with 23andMe that protects

451 the privacy of the 23andMe participants. Please visit

452 <https://research.23andme.com/collaborate/#dataset-access/> for more information and to

453 apply to access the data. Participants provided informed consent and volunteered to

454 participate in the research online, under a protocol approved by the external AAHRPP-
455 accredited IRB, Ethical & Independent (E&I) Review Services. As of 2022, E&I Review Services is
456 part of Salus IRB (<https://www.versiticlinicaltrials.org/salusirb>)

457

458 References

- 459 1. Buniello, A. *et al.* The NHGRI-EBI GWAS Catalog of published genome-wide association
460 studies, targeted arrays and summary statistics 2019. *Nucleic Acids Res.* **47**, D1005-D1012
461 (2019).
- 462 2. Visscher, P. M. *et al.* 10 years of GWAS discovery: biology, function, and translation. *The*
463 *American Journal of Human Genetics* **101**, 5-22 (2017).
- 464 3. Chatterjee, N. *et al.* Projecting the performance of risk prediction based on polygenic
465 analyses of genome-wide association studies. *Nat. Genet.* **45**, 400-405 (2013).
- 466 4. Chatterjee, N., Shi, J. & García-Closas, M. Developing and evaluating polygenic risk
467 prediction models for stratified disease prevention. *Nature Reviews Genetics* **17**, 392
468 (2016).
- 469 5. Sugrue, L. P. & Desikan, R. S. What are polygenic scores and why are they important?
470 *JAMA* **321**, 1820-1821 (2019).
- 471 6. Aragam, K. G. & Natarajan, P. Polygenic scores to assess atherosclerotic cardiovascular
472 disease risk: clinical perspectives and basic implications. *Circ. Res.* **126**, 1159-1177 (2020).
- 473 7. Ma, Y. & Zhou, X. Genetic prediction of complex traits with polygenic scores: a statistical
474 review. *Trends in Genetics* **37**, 995-1011 (2021).
- 475 8. Lambert, S. A., Abraham, G. & Inouye, M. Towards clinical utility of polygenic risk scores.
476 *Hum. Mol. Genet.* **28**, R133-R142 (2019).
- 477 9. Wray, N. R. *et al.* From basic science to clinical application of polygenic risk scores: a
478 primer. *JAMA psychiatry* **78**, 101-109 (2021).
- 479 10. Mavaddat, N. *et al.* Polygenic risk scores for prediction of breast cancer and breast
480 cancer subtypes. *The American Journal of Human Genetics* **104**, 21-34 (2019).

- 481 11. Dikilitas, O. *et al.* Predictive utility of polygenic risk scores for coronary heart disease in
482 three major racial and ethnic groups. *The American Journal of Human Genetics* **106**, 707-
483 716 (2020).
- 484 12. Li, R., Chen, Y., Ritchie, M. D. & Moore, J. H. Electronic health records and polygenic risk
485 scores for predicting disease risk. *Nature Reviews Genetics* **21**, 493-502 (2020).
- 486 13. Fatumo, S. *et al.* A roadmap to increase diversity in genomic studies. *Nat. Med.* **28**, 243-
487 250 (2022).
- 488 14. Popejoy, A. B. & Fullerton, S. M. Genomics is failing on diversity. *Nature* **538**, 161-164
489 (2016).
- 490 15. Peterson, R. E. *et al.* Genome-wide association studies in ancestrally diverse
491 populations: opportunities, methods, pitfalls, and recommendations. *Cell* **179**, 589-603
492 (2019).
- 493 16. Sirugo, G., Williams, S. M. & Tishkoff, S. A. The missing diversity in human genetic
494 studies. *Cell* **177**, 26-31 (2019).
- 495 17. Martin, A. R. *et al.* Clinical use of current polygenic risk scores may exacerbate health
496 disparities. *Nat. Genet.* **51**, 584-591 (2019).
- 497 18. Cavazos, T. B. & Witte, J. S. Inclusion of variants discovered from diverse populations
498 improves polygenic risk score transferability. *Human Genetics and Genomics Advances* **2**,
499 100017 (2021).
- 500 19. Tanigawa, Y. *et al.* Significant sparse polygenic risk scores across 813 traits in UK
501 Biobank. *PLoS Genetics* **18**, e1010105 (2022).
- 502 20. Duncan, L. *et al.* Analysis of polygenic risk score usage and performance in diverse
503 human populations. *Nature communications* **10**, 1-9 (2019).
- 504 21. Duncan, L. *et al.* Analysis of polygenic risk score usage and performance in diverse
505 human populations. *Nature communications* **10**, 1-9 (2019).
- 506 22. Zhang, H. *et al.* Novel Methods for Multi-ancestry Polygenic Prediction and their
507 Evaluations in 3.7 Million Individuals of Diverse Ancestry. *bioRxiv* (2022).
- 508 23. Wojcik, G. L. *et al.* Genetic analyses of diverse populations improves discovery for
509 complex traits. *Nature* **570**, 514-518 (2019).
- 510 24. Mahajan, A. *et al.* Multi-ancestry genetic study of type 2 diabetes highlights the power of
511 diverse populations for discovery and translation. *Nat. Genet.* **54**, 560-572 (2022).

- 512 25. Bentley, A. R. *et al.* Multi-ancestry genome-wide gene–smoking interaction study of
513 387,272 individuals identifies new loci associated with serum lipids. *Nat. Genet.* **51**, 636-
514 648 (2019).
- 515 26. Partanen, J. J. *et al.* Leveraging global multi-ancestry meta-analysis in the study of
516 Idiopathic Pulmonary Fibrosis genetics. *Cell Genomics* **2**, 100181 (2022).
- 517 27. Dudbridge, F. Power and predictive accuracy of polygenic risk scores. *PLoS genetics* **9**,
518 e1003348 (2013).
- 519 28. Vilhjálmsón, B. J. *et al.* Modeling linkage disequilibrium increases accuracy of polygenic
520 risk scores. *The American journal of human genetics* **97**, 576-592 (2015).
- 521 29. Mak, T. S. H., Porsch, R. M., Choi, S. W., Zhou, X. & Sham, P. C. Polygenic scores via
522 penalized regression on summary statistics. *Genet. Epidemiol.* **41**, 469-480 (2017).
- 523 30. Privé, F., Arbel, J., Aschard, H. & Vilhjálmsón, B. J. Identifying and correcting for
524 misspecifications in GWAS summary statistics and polygenic scores. *Human Genetics and*
525 *Genomics Advances* **3**, 100136 (2022).
- 526 31. Privé, F., Arbel, J. & Vilhjálmsón, B. J. LDpred2: better, faster, stronger. *Bioinformatics*
527 **36**, 5424-5431 (2020).
- 528 32. Ge, T., Chen, C., Ni, Y., Feng, Y. A. & Smoller, J. W. Polygenic prediction via Bayesian
529 regression and continuous shrinkage priors. *Nature communications* **10**, 1-10 (2019).
- 530 33. Márquez-Luna, C., Loh, P., South Asian Type 2 Diabetes (SAT2D) Consortium, SIGMA
531 Type 2 Diabetes Consortium & Price, A. L. Multiethnic polygenic risk scores improve risk
532 prediction in diverse populations. *Genet. Epidemiol.* **41**, 811-823 (2017).
- 533 34. Ruan, Y. *et al.* Improving polygenic prediction in ancestrally diverse populations
534 . *Nat. Genet.* **54**, 573-580 (2022).
- 535 35. Cai, M. *et al.* A unified framework for cross-population trait prediction by leveraging the
536 genetic correlation of polygenic traits. *The American Journal of Human Genetics* **108**, 632-
537 655 (2021).
- 538 36. Privé, F., Vilhjálmsón, B. J., Aschard, H. & Blum, M. G. Making the most of clumping and
539 thresholding for polygenic scores. *The American Journal of Human Genetics* **105**, 1213-1221
540 (2019).
- 541 37. Graham, S. E. *et al.* The power of genetic diversity in genome-wide association studies of
542 lipids. *Nature* **600**, 675-679 (2021).
- 543 38. All of Us Research Program Investigators. The “All of Us” research program. *N. Engl. J.*
544 *Med.* **381**, 668-676 (2019).

- 545 39. Allen, N. E., Sudlow, C., Peakman, T., Collins, R. & Uk biobank. UK biobank data: come
546 and get it. *Science translational medicine* **6**, 224ed4 (2014).
- 547 40. Tibshirani, R. Regression shrinkage and selection via the lasso. *Journal of the Royal*
548 *Statistical Society: Series B (Methodological)* **58**, 267-288 (1996).
- 549 41. Hoerl, A. E. & Kennard, R. W. Ridge regression: Biased estimation for nonorthogonal
550 problems. *Technometrics* **12**, 55-67 (1970).
- 551 42. Brown, B. C., Ye, C. J., Price, A. L., Zaitlen, N. & Asian Genetic Epidemiology Network Type
552 2 Diabetes Consortium. Transethnic genetic-correlation estimates from summary statistics.
553 *The American Journal of Human Genetics* **99**, 76-88 (2016).
- 554 43. Mishra, A. *et al.* Stroke genetics informs drug discovery and risk prediction across
555 ancestries. *Nature* **611**, 115-123 (2022).
- 556 44. Tibshirani, R., Saunders, M., Rosset, S., Zhu, J. & Knight, K. Sparsity and smoothness via
557 the fused lasso. *Journal of the Royal Statistical Society: Series B (Statistical Methodology)* **67**,
558 91-108 (2005).
- 559 45. Friedman, J., Hastie, T. & Tibshirani, R. Regularization paths for generalized linear
560 models via coordinate descent. *Journal of statistical software* **33**, 1 (2010).
- 561 46. Van der Laan, M. J., Polley, E. C. & Hubbard, A. E. Super learner. *Statistical applications in*
562 *genetics and molecular biology* **6** (2007).
- 563 47. Polley, E. C. & Van Der Laan, M. J. Super learner in prediction. (2010).
- 564 48. Van der Laan, M. J. & Rose, S. in *Targeted learning: causal inference for observational and*
565 *experimental data* (Springer, 2011).
- 566 49. International HapMap 3 Consortium. Integrating common and rare genetic variation in
567 diverse human populations. *Nature* **467**, 52 (2010).
- 568 50. Bien, S. A. *et al.* Strategies for enriching variant coverage in candidate disease loci on a
569 multiethnic genotyping array. *PloS one* **11**, e0167758 (2016).
- 570 51. 1000 Genomes Project Consortium. A global reference for human genetic variation.
571 *Nature* **526**, 68-74 (2015).
- 572 52. Ding, Y. *et al.* Polygenic scoring accuracy varies across the genetic ancestry continuum
573 in all human populations. *bioRxiv*, 2022.09. 28.509988 (2022).
- 574 53. Sun, Q. *et al.* Improving polygenic risk prediction in admixed populations by explicitly
575 modeling ancestral-specific effects via GAUDI. *bioRxiv* (2022).

576 Online Methods

577

578 **Data preparation and formatting in PROSPER.** We match SNPs and their alleles in GWAS
579 summary statistics and genotypes of individuals for tuning and validation purposes to that in
580 1000G reference data (phase 3)⁵¹. To simplify computing huge-dimensional LD matrix, we use
581 existing LD block information from EUR²⁹ to divide the whole genome, and assume the blocks
582 to be independent. We use PLINK1.9⁵⁴ with flag --r bin4 to compute the LD matrix within each
583 block in each ancestry for common SNPs (MAF>0.01) either in HM3⁴⁹ or the MEGA⁵⁰. For SNPs
584 not common in all populations, we only model them in the populations where they are
585 common; if an SNP is population-specific that is only common in one population, we model it
586 only using the lasso penalty without the genetic similarity penalty. The parameter path of the
587 tuning parameter λ for the scale factor in lasso penalty is set to a sequence evenly spaced on a
588 logarithmic scale from $\lambda^{\max} = \min_{1 \leq i \leq m} \left(\frac{\max_{1 \leq k \leq p} (|r_{ik}|)}{\lambda_i^0} \right)$ to $\lambda^{\min} = 0.001 \times \lambda^{\max}$ which is set to
589 guarantee non-zero solutions, where r_{ik} is the GWAS summary statistics for the k -th SNP in the
590 i -th population, and λ_i^0 is the underlying values of optimal tuning parameter λ for the i -th
591 population. The parameter path for the tuning parameter c for the genetic similarity penalty is
592 set to a sequence of that evenly spaced on a quad-root scale from $c^{\min} = 0.5$ to $c^{\max} = 100$.
593 For all analyses excluding 23andMe, the length of sequences of both parameters are set to be
594 10, while for the analysis of 23andMe, it is set to be 5 to reduce the computation workload
595 caused by the confidential requirements of the 23andMe dataset.

596

597 **Obtain PROSPER solution.** For M populations, the objective function to minimize for p_i -
 598 dimensional vector of SNP effect, $\boldsymbol{\beta}_i, i = 1, \dots, M$, is

$$599 \quad L(\boldsymbol{\beta}_1, \dots, \boldsymbol{\beta}_m) = \sum_{1 \leq i \leq M} (\boldsymbol{\beta}_i^T (\mathbf{R}_i + \delta_i \mathbf{I}) \boldsymbol{\beta}_i - 2 \boldsymbol{\beta}_i^T \mathbf{r}_i + 2 \lambda_i \|\boldsymbol{\beta}_i\|_1) \\
 600 \quad + \sum_{1 \leq i_1 < i_2 \leq M} c_{i_1 i_2} \left\| \boldsymbol{\beta}_{i_1}^{s_{i_1 i_2}} - \boldsymbol{\beta}_{i_2}^{s_{i_1 i_2}} \right\|_2^2$$

601 where \mathbf{R}_i is an estimate of p_i -by- p_i LD matrix based on a reference sample from the i -th
 602 population, \mathbf{r}_i is the p_i -dimensional vector of GWAS summary statistics in the i -th population,
 603 $\boldsymbol{\beta}_{i_1}^{s_{i_1 i_2}}$ and $\boldsymbol{\beta}_{i_2}^{s_{i_1 i_2}}$ denote the effect vectors for the SNPs shared across i_1 -th and i_2 -th
 604 populations (the set of SNPs is denoted by $s_{i_1 i_2}$); δ_i, λ_i and $c_{i_1 i_2}$ are tuning parameters as
 605 defined in above sections.

606 This optimization can be solved using coordinate descent algorithms by iteratively updating
 607 each element in the vectors. We take derivative for SNP k in i -th population, $k = 1, \dots, p_i, i =$
 608 $1, \dots, M$

$$609 \quad \frac{\partial L(\boldsymbol{\beta}_1, \dots, \boldsymbol{\beta}_m)}{\partial \beta_{ik}} \\
 610 \quad = 2 \left(1 + \delta_i + \sum_{i' \neq i, 1 \leq i' \leq M} c_{ii'} \right) \beta_{ik} + 2 \lambda_i \frac{\partial |\beta_{ik}|}{\partial \beta_{ik}} \\
 611 \quad - 2 \left(r_{ik} - \sum_{k' \neq k, 1 \leq k' \leq p} R_{i, k' k} \beta_{ik'} + \sum_{1 \leq i' \leq M, \text{s.t. } k \in S_{i, i'}} c_{ii'} \beta_{i' k} \right)$$

612 where β_{ik} denotes the SNP k in $\boldsymbol{\beta}_i$, r_{ik} denotes the SNP k SNP in \mathbf{r}_i , and $R_{i, k' k}$ denotes LD
 613 between the SNP k and the SNP k' in \mathbf{R}_i .

614 By solving $\frac{\partial L(\beta_1, \dots, \beta_m)}{\partial \beta_{ik}} = 0$ after the (t) -th iteration, we can get the updating rule for the $(t +$

615 $1)$ -th iteration

$$616 \quad \beta_{ik}^{(t+1)} = \frac{\text{sign}(u_{ik}) \cdot \max\{0, |u_{ik}| - \lambda_i\}}{1 + \delta_i + \sum_{1 \leq i' \leq M, s.t. k \in S_{i,i'}} c_{ii'}}$$

617 where

$$618 \quad u_{ik} = r_{ik} - \sum_{k' \neq k, 1 \leq k' \leq p} R_{i,k'k} \beta_{ik'}^{(t)} + \sum_{1 \leq i' \leq M, s.t. k \in S_{i,i'}} c_{ii'} \beta_{i'k}^{(t)}$$

619

620 **Super learning.** After getting PRSs for all populations under all tuning parameter settings, we
621 further apply super learning to combine them to be trained on the tuning samples to get the
622 final PROSPER model and tested on the validation samples. We use the function “*SuperLearner*”
623 implemented in the R package with the same name, and include three linear prediction
624 algorithms: lasso, ridge, and linear regression for continuous outcomes; and two prediction
625 algorithms: lasso and linear regression for binary outcomes. We did not include ridge for binary
626 outcomes due to the unavailability of ridge for binary outcomes in the function. For the
627 included algorithms which have parameters: (1) in lasso, we use 100 values in lambda path
628 calculated in the default setting in glmnet package; (2) in ridge, we use a lambda path of
629 sequence from 1 to 20 incrementing by 0.1. We use Area under the ROC curve (AUC) as the
630 objective function for binary outcomes and thus use the flag “method = method. AUC” in the
631 function.

632

633 **Existing PRS methods.** We compare five groups of PRS methods. The first group is: single-
634 ancestry method, which contains commonly known single-ancestry methods, including CT,
635 LDpred2, and lassosum2, that are trained from the GWAS data from the target population. The
636 second group is: EUR PRS based method, which is the three above single-ancestry methods
637 trained from EUR GWAS data. The third group is: weighted PRS, which uses the weights
638 estimated from a linear regression to combine the PRSs estimated from the corresponding
639 single-ancestry method from all populations. The fourth group is: existing multi-ancestry
640 methods, which includes two recently published and well-performed multi-ancestry methods,
641 PRS-CSx and CT-SLEB. The last group is our proposed PROSPER. For all algorithms that have
642 tuning parameters or weights, the optimal ones are determined based on predictive R^2 or AUC
643 on tuning samples and finally evaluated on validation samples.

644 **CT** is implemented in our analysis by using r^2 -cutoff of 0.1 in the clumping step and then
645 thresholding by treating p-value-cutoff as a tuning parameter and being chosen from
646 $5 \times 10^{-8}, 1 \times 10^{-7}, 5 \times 10^{-7}, 1 \times 10^{-6}, \dots, 5 \times 10^{-1}, 1.0$.

647 **LDpred2** is a PRS method that uses a spike-and-slab prior on GWAS summary statistics and
648 modeling LD across SNPs. We implement LDpred2 by the function “*snp_ldpred2_grid*” in the R
649 package “bigsnpr”. The two tuning parameters in the algorithm include: the proportion of
650 causal SNPs, which is chosen from a sequence of length 17 that are evenly spaced on a
651 logarithmic scale from 10^{-4} to 1; per-SNP heritability, which is chosen from 0.7, 1, or 1.4 times
652 the total heritability estimated by LD score regression divided by the number of causal SNPs.
653 We fix the additional “sparse” option (for truncating small effects to zero) to FALSE.

654 **lassosum2** is a PRS method that uses lasso regression on GWAS summary statistics for a single
655 ancestry. We implement lasso2 by the function “*snp_lassosum2*” in the R package
656 “*bigsnp*”. The two tuning parameters in the algorithm include: tuning parameter for the lasso
657 penalty, which is chosen from a sequence of length 20 that are evenly spaced on a logarithmic
658 scale from $0.01 \times \max_{1 \leq k \leq p} (|r_k|)$ to $\max_{1 \leq k \leq p} (|r_k|)$; and regularization parameter for LD matrix, which
659 is chosen from a sequence of length 10 that are evenly spaced on a cube-root scale from 0.01
660 to 100.

661 **EUR PRS** are the PRSs trained from EUR GWAS using the above single-ancestry methods, CT,
662 LDpred2, and lasso2, that are then applied to individuals of the target population. There is
663 no need to perform tuning for them because the models have been tuned in EUR tuning
664 samples. When computing scores for EUR PRS based method, we exclude SNPs that are not
665 presented in the validation samples from the target population.

666 **Weighted PRS** linearly combines the corresponding single-ancestry method trained from all
667 populations. The weights in the linear combination are estimated by a simple linear regression
668 in the tuning samples from the target population.

669 **PRS-CSx** is a Bayesian multi-ancestry PRS method that jointly models GWAS summary statistics
670 and LD structures across multiple populations using a continuous shrinkage prior. It has a
671 further step to linearly combine the posterior effect-sizes estimates for EUR and the target
672 population using weights in a simple linear regression in the tuning samples from the target
673 population. We implement PRS-CSx using their python-based command line tool “*PRS-CSx*”. The
674 parameter phi was chosen from the default candidate values, 1, 10^{-2} , 10^{-4} and 10^{-6} . Due to
675 the package restriction, the models are fitted with only HM3 SNPs.

676 **CT-SLEB** is a multi-ancestry PRS method that starts from clumping and thresholding, then uses
677 Empirical-Bayes (EB) method to estimate the coefficients of PRS, and finally combines PRS by a
678 super learning model. The three tuning parameters in the algorithm include: r^2 -cutoff and base
679 size of the clumping window size used in the clumping step, which are chosen from (0.01, 0.05,
680 0.1, 0.2, 0.5) and (50kb, 100kb), respectively; and p-value cutoffs for EUR and the target
681 population, which are chosen from 5×10^{-8} , 5×10^{-7} , 5×10^{-6} , ..., 5×10^{-1} and 1.0.

682

683 **Computational time and memory usage.** The computational time and memory usage of
684 PROSPER and PRS-CSx are compared based on the analysis using simulated data on
685 chromosome 22. The analysis starts from inputting all required data into the algorithms, such as
686 summary statistics and LD reference data, and ends with outputting the final PRS coefficients
687 from the algorithms. PROSPER requires an input of optimal parameters in single-ancestry
688 analysis, so we also include the step of running the single-ancestry analysis, lassosum. The
689 analyses are performed using a single core with AMD EPYC 7702 64-Core Processors running at
690 2.0 GHz. The reported results are averaged over 10 replicates. The sample size for training
691 GWAS summary statistics is 15,000 for non-EUR populations and 100,000 for EUR population.
692 The sample size for the tuning dataset is 10,000 for each population.

693

694 **Real data analysis.** Training GWAS summary statistics are from 23andMe, GLGC, and AoU.
695 Tuning and validation individual-level data are from 23andMe and UKBB. Detailed descriptions
696 of those datasets are listed below.

697 **23andMe Data.** We analyzed two continuous traits, heart metabolic disease burden and height;
698 and five binary traits, any CVD, depression, migraine diagnosis, morning person and SBMN,
699 using GWAS summary statistics obtained from 23andMe Inc.. The information of individuals
700 included in our analyses from the 23andMe participant cohort has consent and answered
701 surveys online according to the human subject protocol reviewed and approved by Ethical &
702 Independent Review Services, a private institutional review board
703 (<http://www.eandireview.com>) as described in a previous paper ²². Data on the seven traits are
704 available for all five populations: AA, EAS, EUR, Latino, and SAS. The LD reference panels used
705 for the five populations, respectively, are unrelated individuals from 1000G of AFR, EAS, EUR,
706 AMR, and SAS origins. The tuning and validation are performed on a set of independent
707 individuals of the corresponding ancestry from 23andMe participant cohort. Please see
708 **Supplementary Table 3** for training sample sizes and **Supplementary Table 4** for tuning and
709 validation sample sizes. The details of the data, including genotyping, quality control,
710 imputation, removing related individuals, ancestry determination, and the preprocessing of
711 GWAS, are also described in the previous paper ²². For continuous traits, we evaluate PRS
712 performance by the predictive R^2 of the PRS for residualized trait values obtained from
713 regressing the traits on covariates. For binary traits, we evaluated PRS performance by the AUC
714 by using the `roc.binary` function in the R package RISCA version 1.0 ⁵⁵. To compare the PRS
715 performance for two different methods, we used the relative increase of logit-scale variance.
716 The logit-scale variance of binary traits is converted from AUC by the formula $\sigma^2 =$
717 $2\phi^{-1}(AUC)$, where ϕ is the cumulative distribution function of the standard normal
718 distribution.

719 **GLGC Data.** We analyzed four blood lipid traits, LDL, HDL, logTG and TC, using GWAS summary
720 statistics computed without UKBB samples that are publicly available from GLGC
721 (<http://csg.sph.umich.edu/willer/public/glgc-lipids2021/>). Detailed information about the
722 design of the study, genotyping, quality control, and GWAS is described in Graham, S. E. *et al.*
723 (2021)³⁷. Data on the four traits are available for all five populations: admixed African or
724 African, EAS, EUR, Hispanic, and SAS. The LD reference panels used for the five populations,
725 respectively, are unrelated individuals from 1000G of AFR, EAS, EUR, AMR, and SAS origins. The
726 tuning and validation are performed on UKBB individuals (as described below) from the same
727 reference ancestry label as the LD reference panel. Please see **Supplementary Table 3** for
728 sample sizes and the number of SNPs included in the analysis. The cleaning and preprocessing
729 of the GWAS data are described in a previous paper²².

730 **AoU Data.** We analyzed two anthropometric traits, BMI and height, using GWAS summary
731 statistics trained from AoU. The information of individuals included in our analyses has been
732 collected according to All of Us Research Program Operational Protocol
733 (https://allofus.nih.gov/sites/default/files/aou_operational_protocol_v1.7_mar_2018.pdf).
734 Details of the data and GWAS summary statistics are previously described²². Data for the two
735 traits are available for three ancestries: AFR, Latino/Admixed American, and EUR. The LD
736 reference panel used for the three populations, respectively, are 1000G unrelated individuals of
737 AFR, AMR, and EUR origins. The tuning and validation are performed using UKBB individuals (as
738 described below) from the same reference ancestry label as the LD reference panel. Please see
739 **Supplementary Table 3** for sample sizes and the number of SNPs included in the analysis. The
740 cleaning and preprocessing of the GWAS data are described in a previous paper²².

741 **UKBB data.** We used UKBB data only for tuning and validation purposes. The four blood lipid
742 traits and two anthropometric traits mentioned above have direct measurements in UKBB. The
743 ancestry label of UKBB individuals is determined by genetically predicted ancestry, which are
744 described in a previous paper ²². Tuning and validation are based on R^2 of the PRS regressed on
745 the residuals of the phenotypes adjusted by sex, age and PC1-10. Please see **Supplementary**
746 **Table 4** for sample sizes. We note that for PRS we tested in UKBB validation samples, we use
747 the ancestry labels in UKBB (AFR, AMR, EAS, EUR or SAS), instead of ancestry labels in the
748 GWAS training data, to report the R^2 in the figures, result, and discussion sections of this paper.

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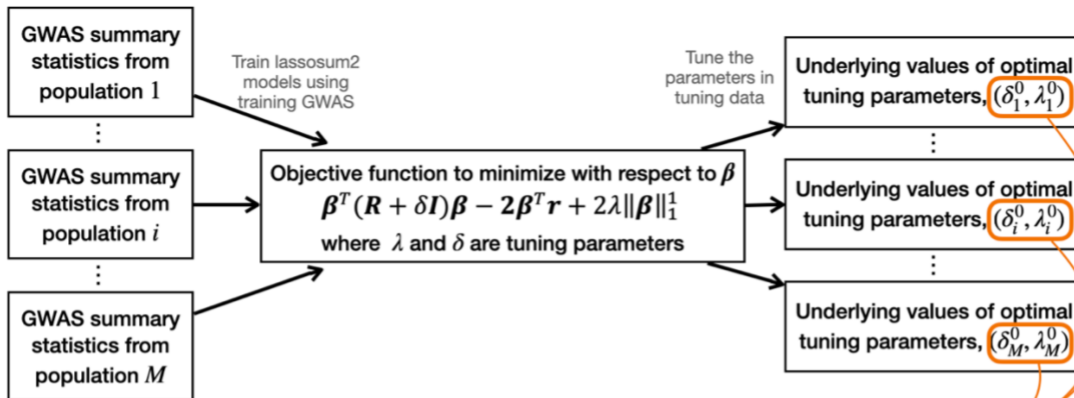
751 References

- 752 54. Purcell, S. *et al.* PLINK: a tool set for whole-genome association and population-based
753 linkage analyses. *The American journal of human genetics* **81**, 559-575 (2007).
- 754 55. Chatton, A. *et al.* G-computation, propensity score-based methods, and targeted
755 maximum likelihood estimator for causal inference with different covariates sets: a
756 comparative simulation study. *Scientific reports* **10**, 1-13 (2020).

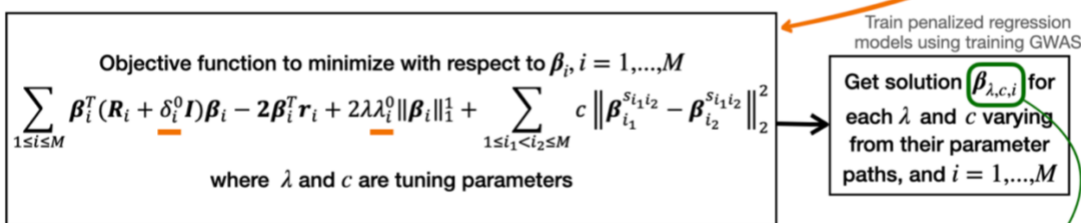
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759 **Figure 1: Detailed flowchart of PROSPER.** The analysis of M populations in PROSPER involves
 760 three key steps: 1. Separate single-ancestry analysis for all populations $i = 1, \dots, M$; 2. Joint
 761 analysis across populations using penalized regression; 3. Ensemble regression. In step 1, the
 762 training GWAS data is used to train lassosum2 models, and the tuning data is used to obtain the
 763 optimal tuning parameters in a single-ancestry analysis. In step 2, the training GWAS and the
 764 optimal tuning parameter values from step 1 are used to train the joint cross-population
 765 penalized regression model, and obtain solution $\beta_{\lambda,c,i}$ for each λ and c . In step 3, the tuning
 766 data is used to train the super learning model for the ensemble of PRSs computed from the
 767 solutions in step 2, $PRS_{\lambda,c,i} = X\beta_{\lambda,c,i}$. The final PRS is computed as $PRS = X(\sum w_{\lambda,c,i}\beta_{\lambda,c,i})$,
 768 where $w_{\lambda,c,i}$ are the weights from the super learning model. Refer to the “Method Overview”
 769 section in the main text for a full explanation of all notations in the flowchart.

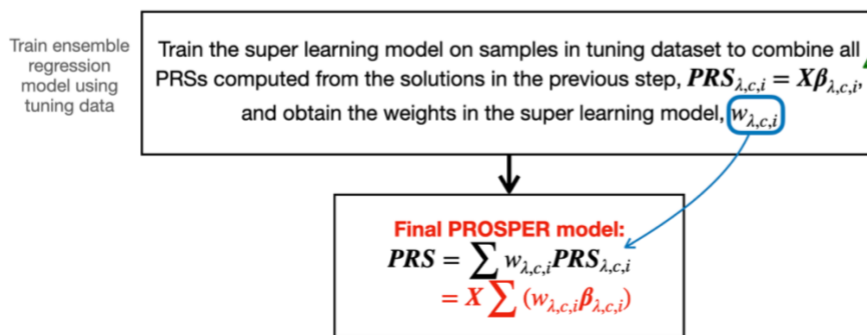
Step 1: Separate single-ancestry analysis for all populations



Step 2: Joint analysis across populations using penalized regression



Step 3: Ensemble regression

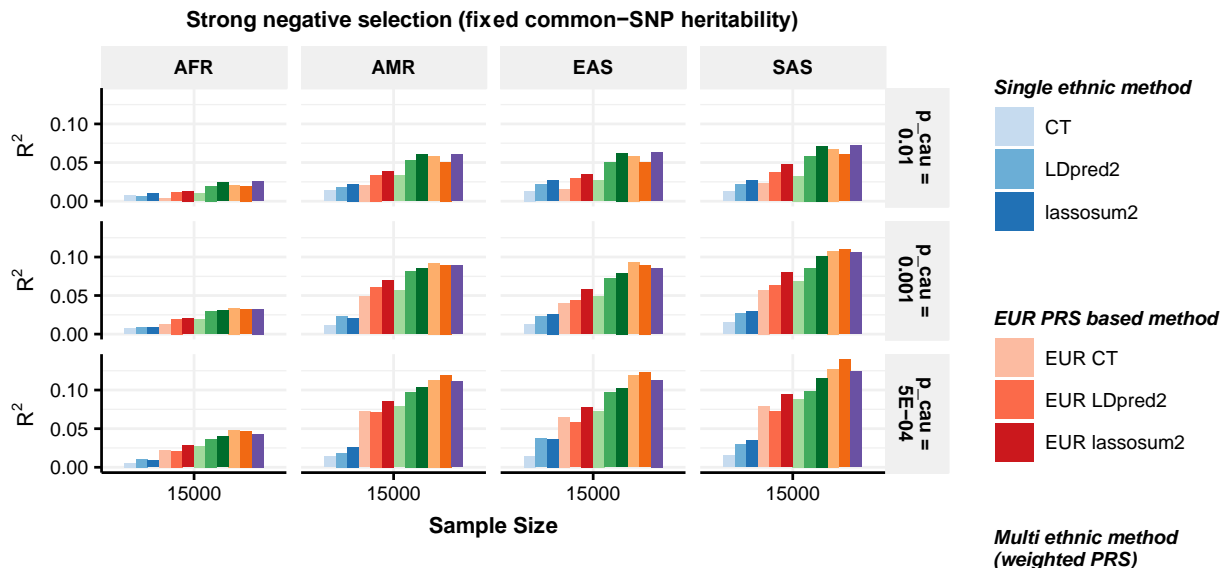


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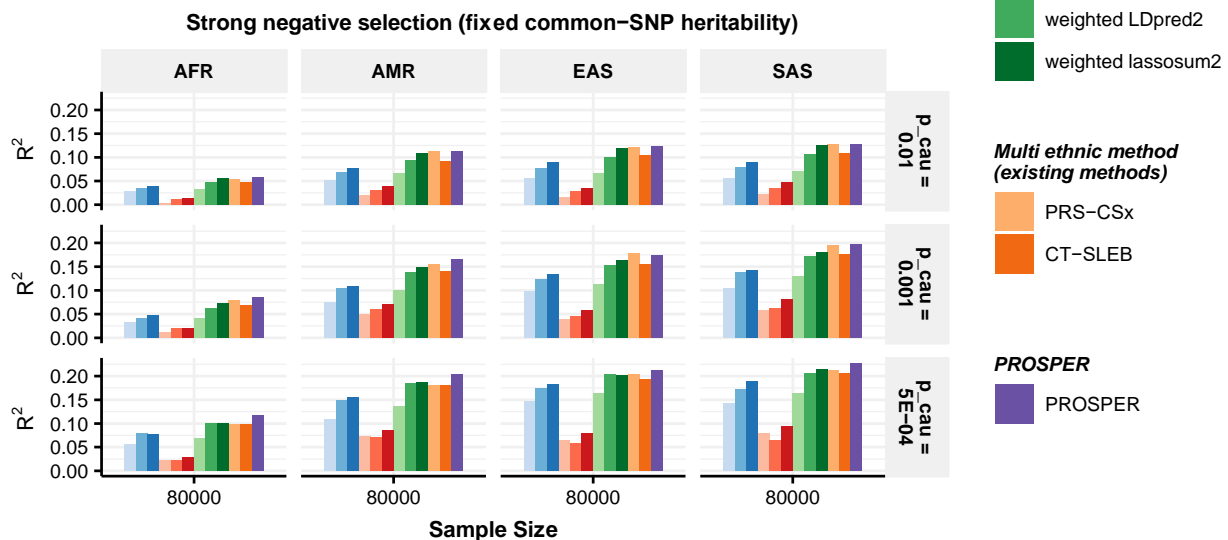
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772 **Figure 2: Performance comparison of alternative methods on simulated data generated with**
 773 **different sample sizes and genetic architectures under strong negative selection and fixed**
 774 **common-SNP heritability.** Data are simulated for continuous phenotype under a strong
 775 negative selection model and three different degrees of polygenicity (top panel: $p_{causal} = 0.01$,
 776 middle panel: $p_{causal} = 0.001$, and bottom panel: $p_{causal} = 5 \times 10^{-4}$). Common SNP
 777 heritability is fixed at 0.4 across all populations, and the correlations in effect sizes for share
 778 SNPs between all pairs of populations is fixed at 0.8. The sample sizes for GWAS training data
 779 are assumed to be (a) 15,000, and (b) 80,000 for the four non-EUR target populations; and is
 780 fixed at 100,000 for the EUR population. PRS generated from all methods are tuned in 10,000
 781 samples, and then tested in 10,000 independent samples in each target population. The PRS-
 782 CSx package is restricted to SNPs from HM3, whereas other alternative methods use SNPs from
 783 either HM3 or MEGA.

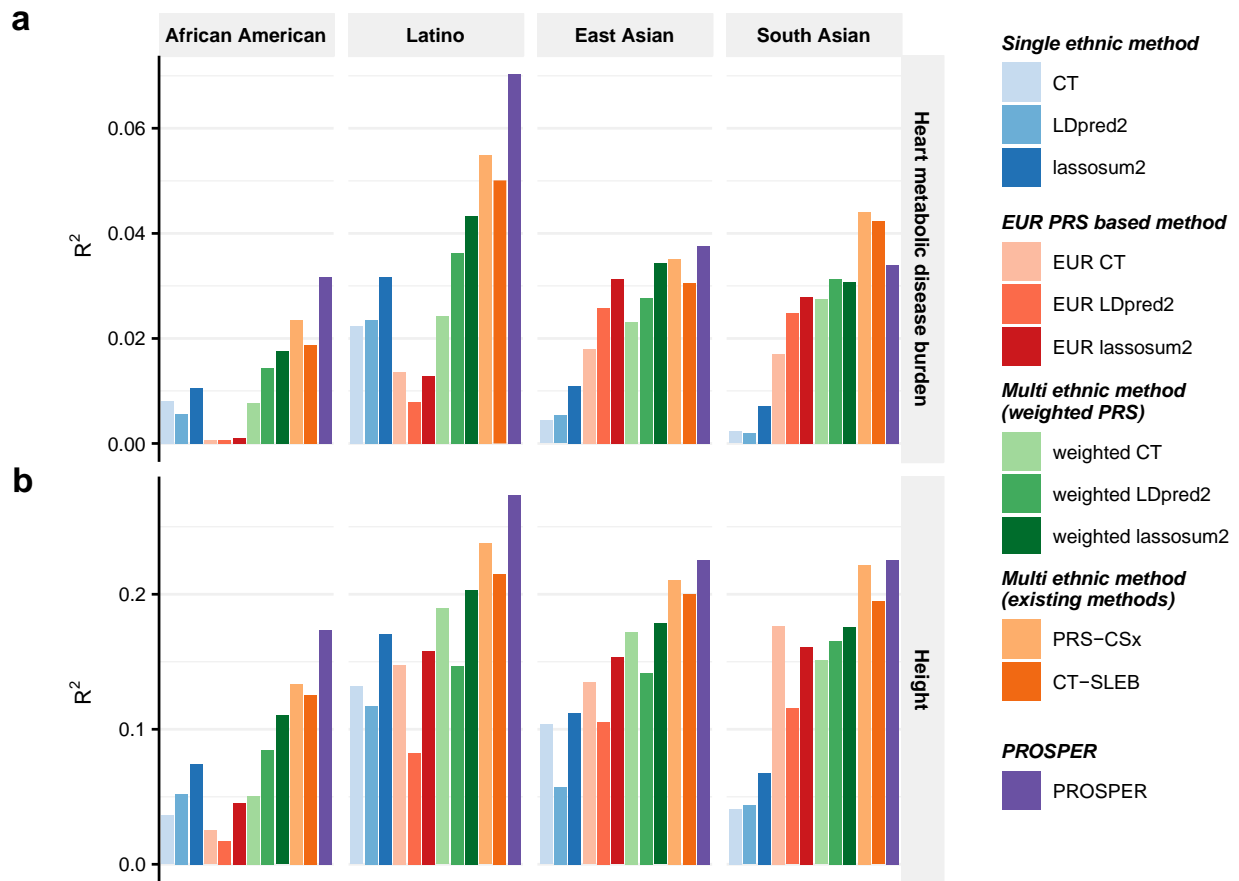
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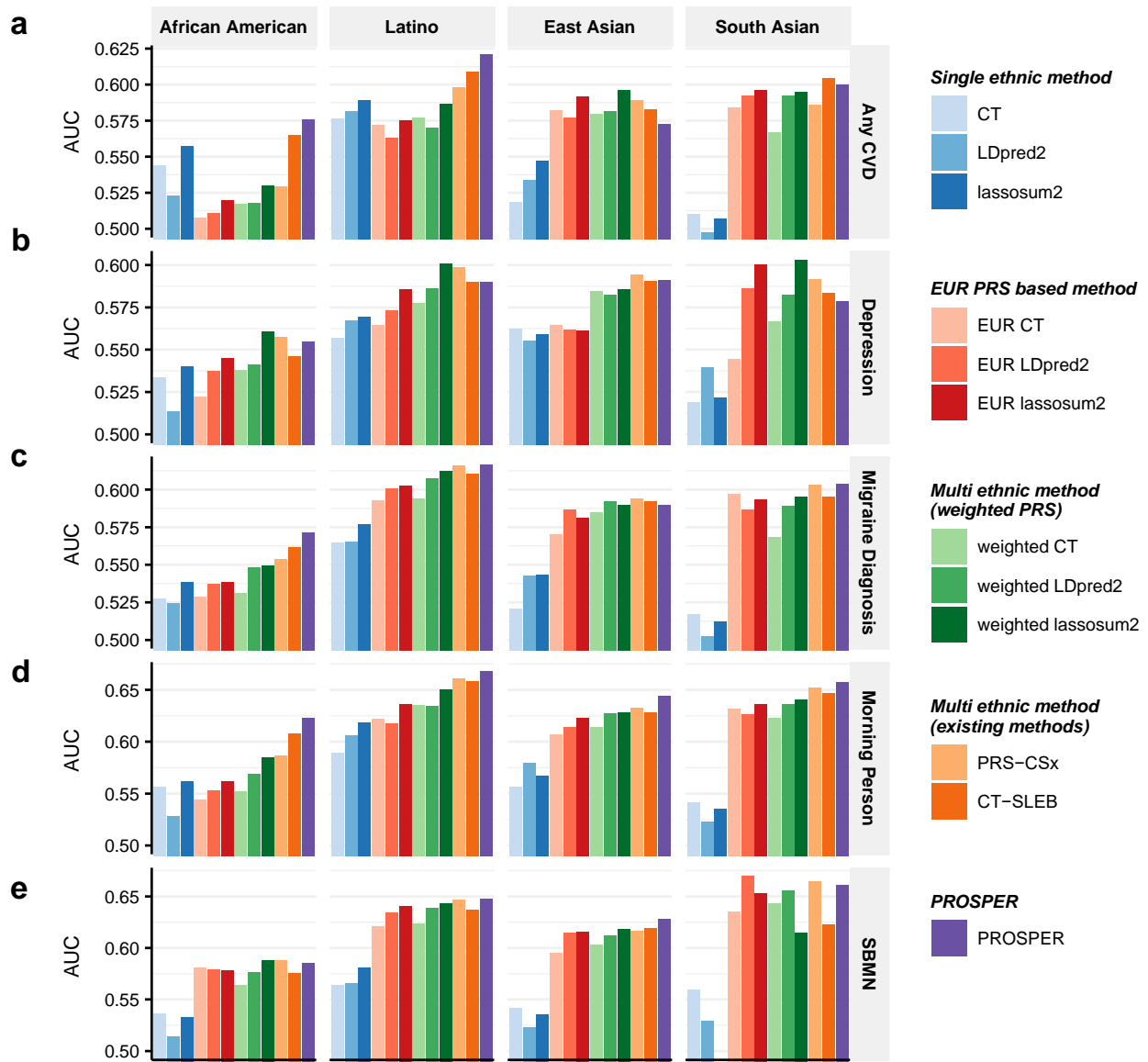


785 **Figure 3: Performance comparison of alternative methods for prediction of two continuous**
 786 **traits in 23andMe.** We analyzed two continuous traits, (a) heart metabolic disease burden and
 787 (b) height. PRS are trained using 23andMe data that available for five populations: African
 788 American, Latino, EAS, EUR, and SAS, and then tuned in an independent set of individuals from
 789 23andMe of the corresponding ancestry. Performance is reported based on adjusted R^2
 790 accounting for sex, age and PC1-5 in a held-out validation sample of individuals from 23andMe
 791 of the corresponding ancestry. The ratio of sample sizes for training, tuning and validation is
 792 roughly about 7:2:1, and detailed numbers are in **Supplementary Table 3-4**. The PRS-CSx
 793 package is restricted to SNPs from HM3, whereas other alternative methods use SNPs from
 794 either HM3 or MEGA.
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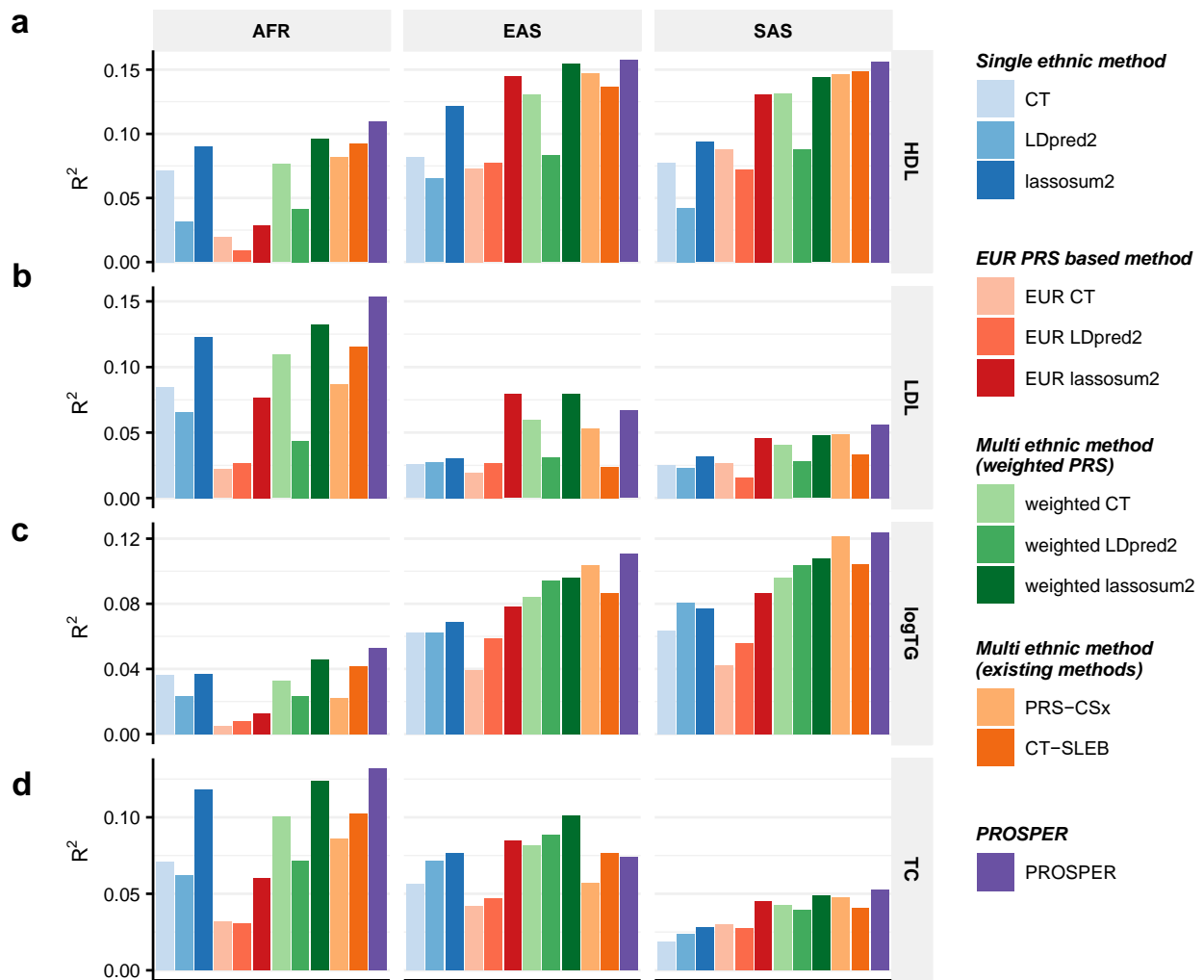


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798 **Figure 4: Performance comparison of alternative methods for prediction of five binary traits**
 799 **in 23andMe.** We analyzed five binary traits, (a) any CVD, (b) depression, (c) migraine diagnosis,
 800 (d) morning person and (e) SBMN. PRS are trained using 23andMe data that available for five
 801 populations: African American, Latino, EAS, EUR, and SAS, and then tuned in an independent
 802 set of individuals from 23andMe of the corresponding ancestry. Performance is reported based
 803 on adjusted AUC accounting for sex, age, PC1-5 in a held-out validation sample of individuals
 804 from 23andMe of the corresponding ancestry. The ratio of sample sizes for training, tuning and
 805 validation is roughly about 7:2:1, and detailed numbers are in **Supplementary Table 3-4.** The
 806 PRS-CSx package is restricted to SNPs from HM3, whereas other alternative methods use SNPs
 807 from either HM3 or MEGA.

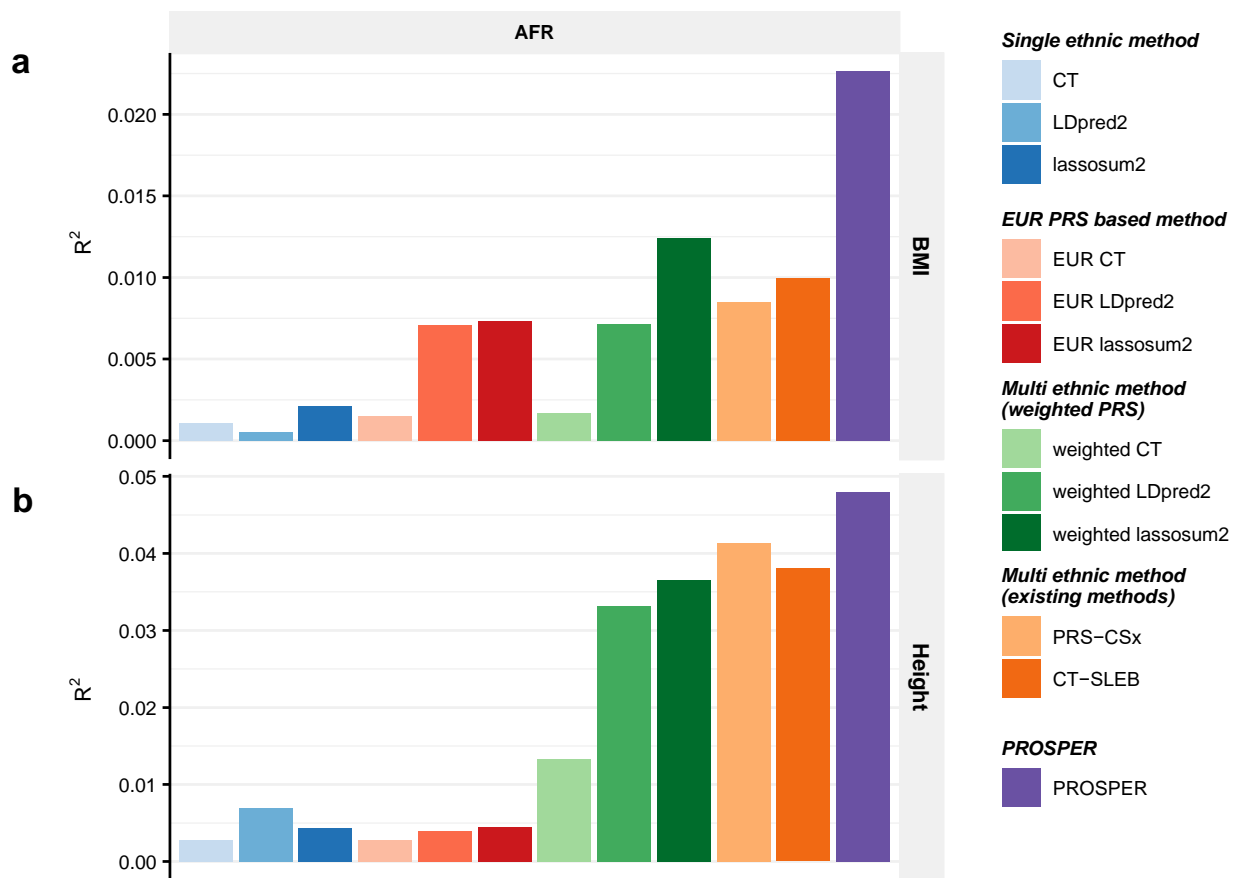


809 **Figure 5: Performance comparison of alternative methods for prediction of four blood lipid traits (GLGC-training and UKBB-tuning/validation).**
 810 We analyzed four blood lipid traits, (a) HDL,
 811 (b) LDL, (c) logTG and (d) TC. PRS are trained using GLGC data that available for five populations:
 812 admixed African or African, East Asian, European, Hispanic, and South, and then tuned in
 813 individuals from UKBB of the corresponding ancestry: AFR, EAS, EUR, AMR, and SAS (see the
 814 section of **Real data analysis** in **Methods** for ancestry composition). Performance is reported
 815 based on adjusted R^2 accounting for sex, age, PC1-10 in a held-out validation sample of
 816 individuals from UKBB of the corresponding ancestry. Sample sizes for training, tuning and
 817 validation data are in **Supplementary Table 3-4**. Results for AMR are not included due to the
 818 small sample size of genetically inferred AMR ancestry individuals in UKBB. The PRS-CSx
 819 package is restricted to SNPs from HM3, whereas other alternative methods use SNPs from
 820 either HM3 or MEGA.
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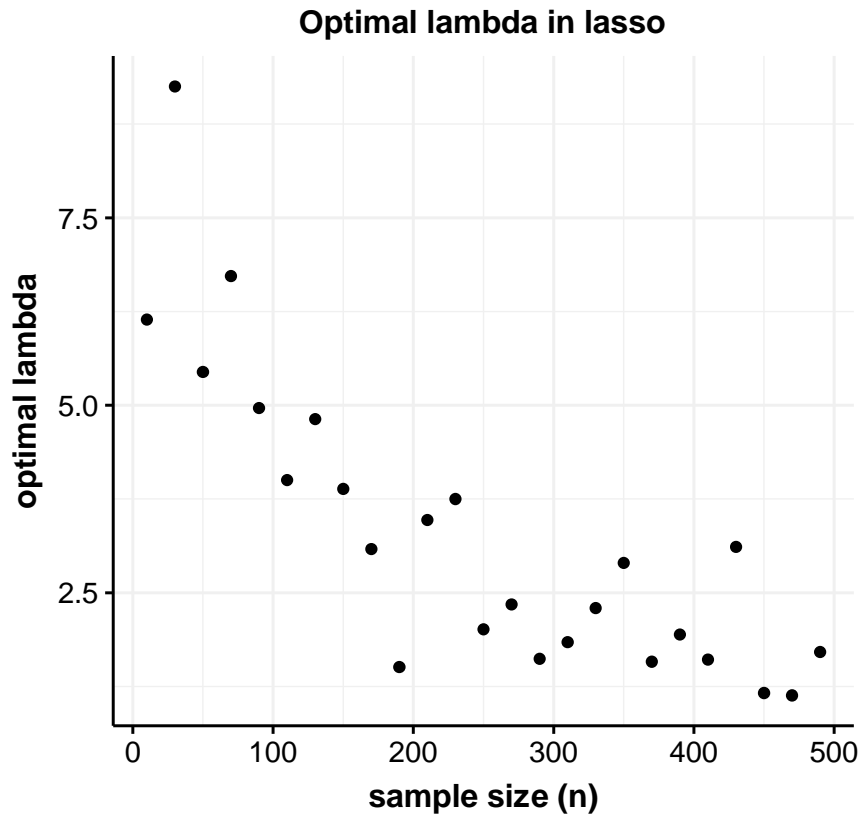
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825 **Figure 6: Performance comparison of alternative methods for prediction of two**
 826 **anthropometric traits (AoU-training and UKBB-tuning/validation).** We analyzed two
 827 anthropometric traits, (a) BMI and (b) height. PRS are trained using AoU data that are available
 828 for three populations: African, Latino/Admixed American, and European and then tuned in
 829 individuals from UKBB of the corresponding ancestry: AFR, AMR, and EUR (see the section of
 830 **Real data analysis in Methods** for ancestry composition). Performance is reported based on
 831 adjusted R^2 accounting for sex, age, PC1-10 in a held-out validation sample of individuals from
 832 UKBB of the corresponding ancestry. Sample sizes for training, tuning and validation data are in
 833 **Supplementary Table 3-4**. Results for AMR are not included due to the small sample size of
 834 genetically inferred AMR ancestry individuals in UKBB. The number of SNPs analyzed in AoU
 835 analyses is much smaller than other analyses because the GWAS from AoU is on array data only
 836 (see **Supplementary Table 3** for the number of SNPs). The PRS-CSx package is restricted to SNPs
 837 from HM3, whereas other alternative methods use SNPs from either HM3 or MEGA.
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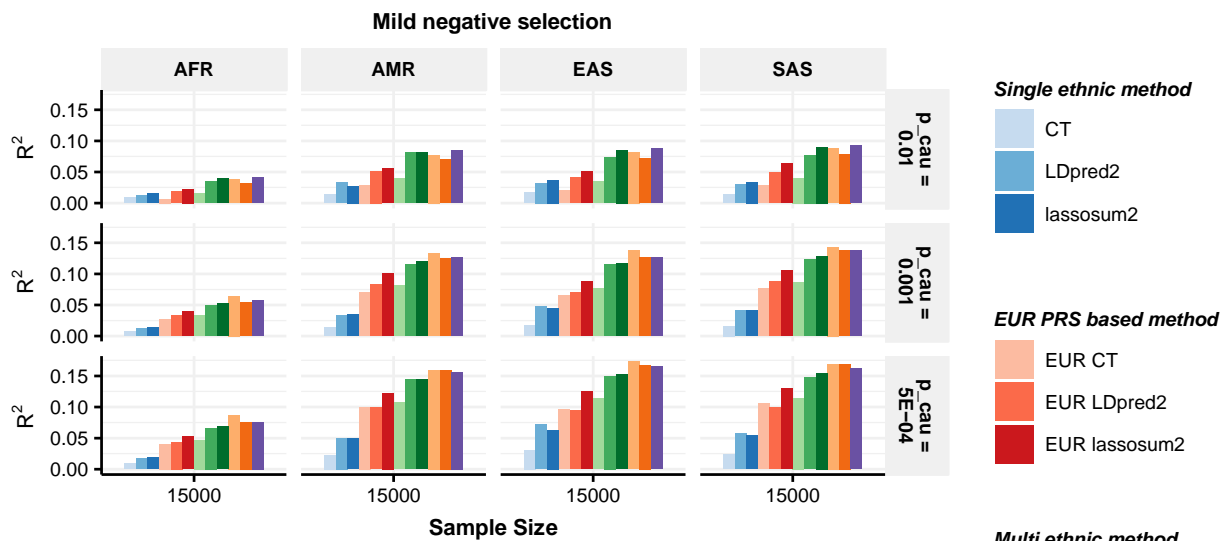
842 **Supplementary Figure 1: Optimal tuning parameter lambda in lasso.** The simulation is
843 performed for design matrix with 1000 predictors ($p = 1000$), and 5% of them are randomly
844 selected to be causal. Correlation structure of those predictors is AR1 with $\rho = 0.4$. The total
845 heritability is simulated to be 0.2.



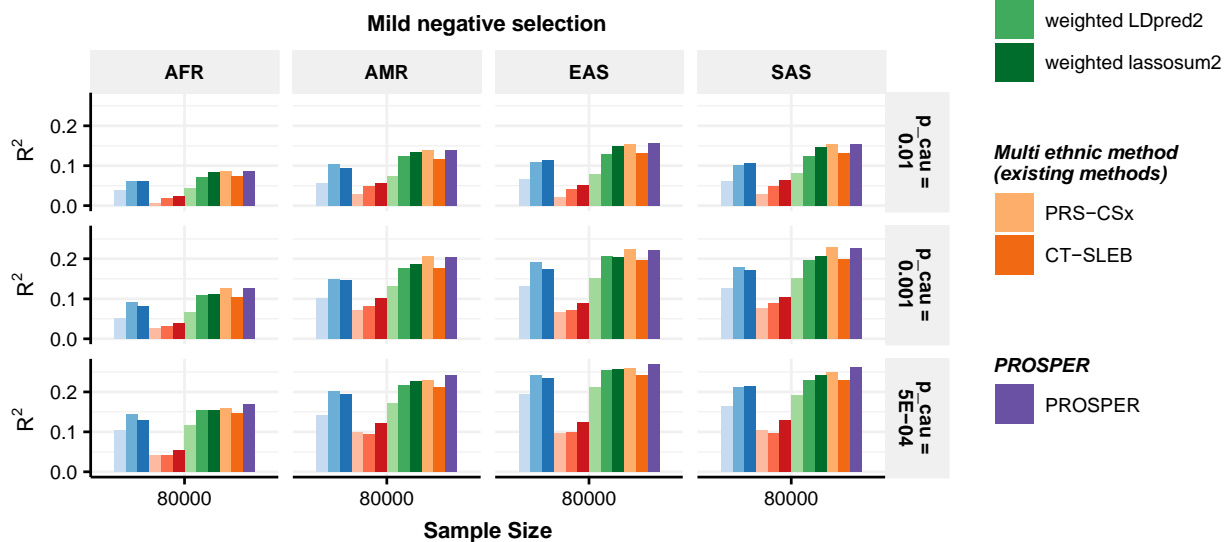
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848 **Supplementary Figure 2: Performance of alternative methods on simulated data generated**
 849 **with different sample sizes and different genetic architectures.** Data are simulated for
 850 continuous phenotype under a mild negative selection model and three different degrees of
 851 polygenicity (top panel: $p_{causal} = 0.01$, middle panel: $p_{causal} = 0.001$, and bottom panel:
 852 $p_{causal} = 5 \times 10^{-4}$). Common SNP heritability is fixed at 0.4 across all populations, and the
 853 correlations in effect sizes for share SNPs between all pairs of populations is fixed at 0.8. The
 854 sample sizes for GWAS training data are assumed to be (a) 15,000, and (b) 80,000 for the four
 855 non-EUR target populations; and is fixed at 100,000 for the EUR population. PRS generated
 856 from all methods are tuned in 10,000 samples, and then tested in 10,000 independent samples
 857 in each target population. The PRS-CSx package is restricted to SNPs from HM3, whereas other
 858 alternative methods use SNPs from either HM3 or MEGA.

a



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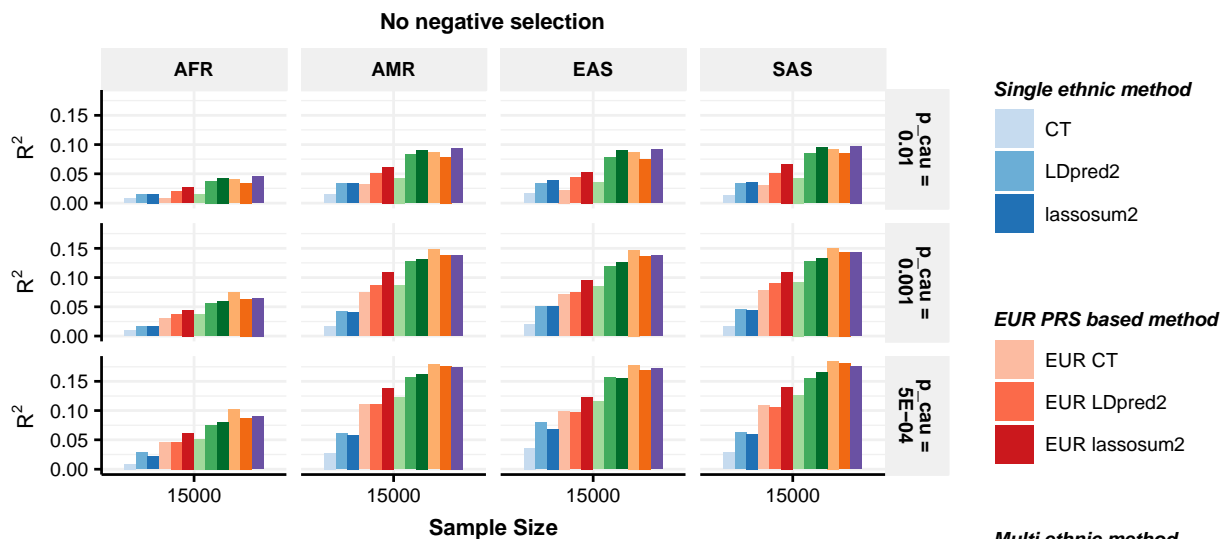


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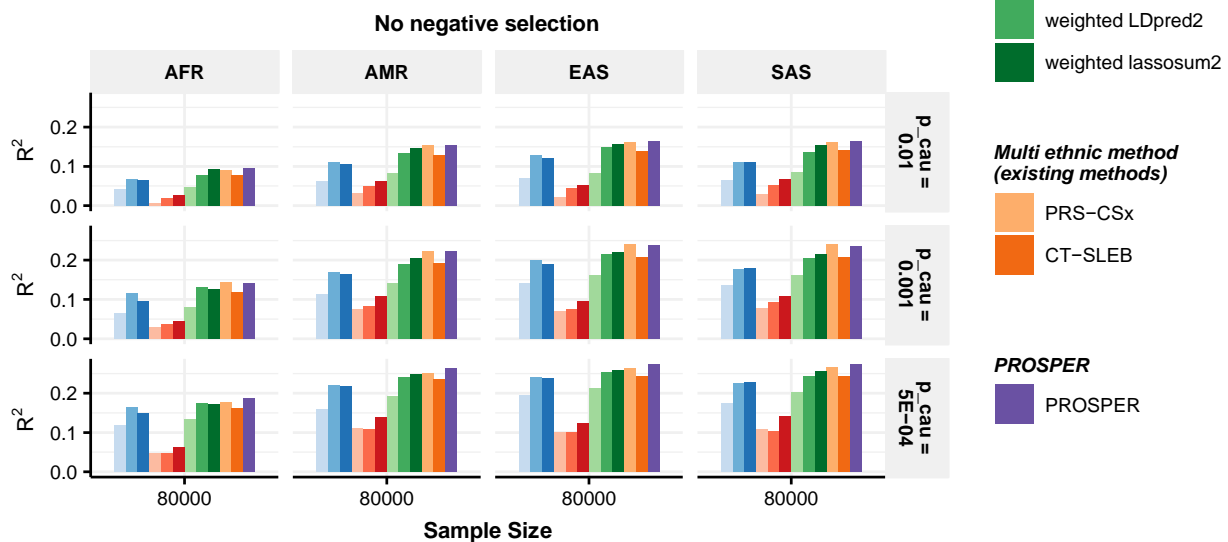
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861 **Supplementary Figure 3: Performance of alternative methods on simulated data generated**
 862 **with different sample sizes and different genetic architectures.** Data are simulated for
 863 continuous phenotype under a no negative selection model and three different degrees of
 864 polygenicity (top panel: $p_{causal} = 0.01$, middle panel: $p_{causal} = 0.001$, and bottom panel:
 865 $p_{causal} = 5 \times 10^{-4}$). Common SNP heritability is fixed at 0.4 across all populations, and the
 866 correlations in effect sizes for share SNPs between all pairs of populations is fixed at 0.8. The
 867 sample sizes for GWAS training data are assumed to be (a) 15,000, and (b) 80,000 for the four
 868 non-EUR target populations; and is fixed at 100,000 for the EUR population. PRS generated
 869 from all methods are tuned in 10,000 samples, and then tested in 10,000 independent samples
 870 in each target population. The PRS-CSx package is restricted to SNPs from HM3, whereas other
 871 alternative methods use SNPs from either HM3 or MEGA.

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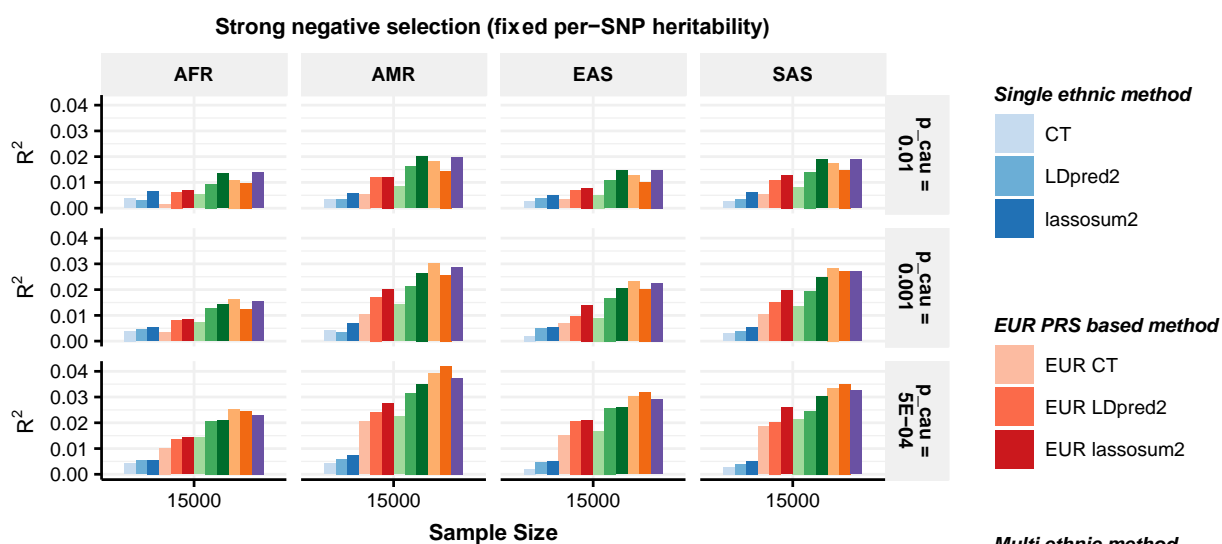


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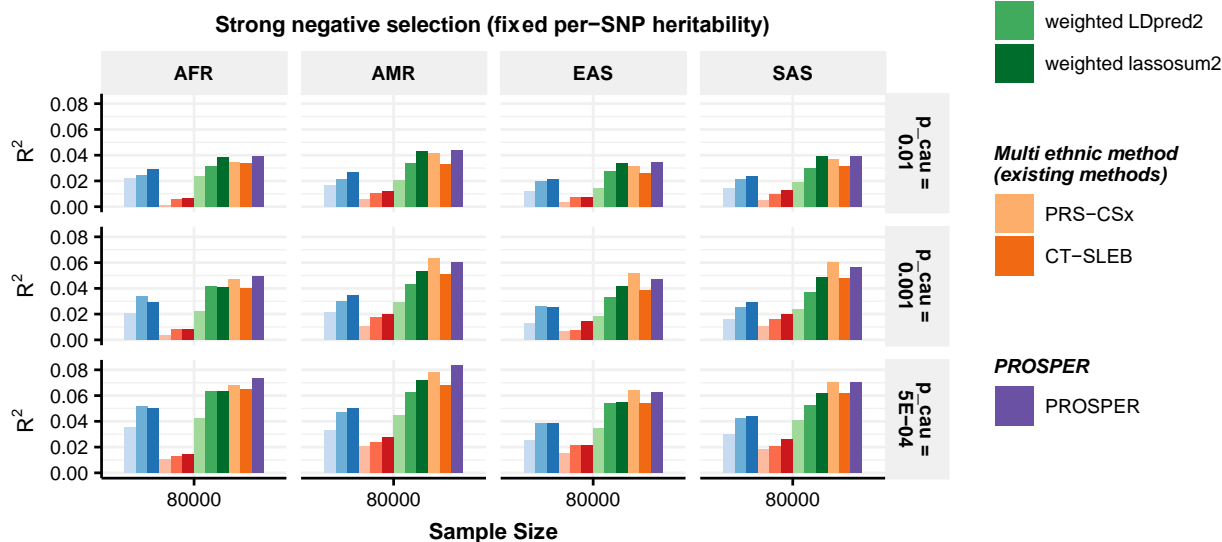
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874 **Supplementary Figure 4: Performance of alternative methods on simulated data generated**
 875 **with different sample sizes and different genetic architectures.** Data are simulated for
 876 continuous phenotype under a strong negative selection model and three different degrees of
 877 polygenicity (top panel: $p_{causal} = 0.01$, middle panel: $p_{causal} = 0.001$, and bottom panel:
 878 $p_{causal} = 5 \times 10^{-4}$). Per-SNP heritability is assumed to be the same across all populations
 879 and thus leads to the common SNP heritability value of 0.32, 0.21, 0.16, 0.19 and 0.17 for AFR,
 880 AMR, EAS, EUR and SAS, respectively. The correlations in effect sizes for share SNPs between all pairs
 881 of populations is fixed at 0.8. The sample sizes for GWAS training data are assumed to be (a)
 882 15,000, and (b) 80,000 for the four non-EUR target populations; and is fixed at 100,000 for the
 883 EUR population. PRS generated from all methods are tuned in 10,000 samples, and then tested
 884 in 10,000 independent samples in each target population. The PRS-CSx package is restricted to
 885 SNPs from HM3, whereas other alternative methods use SNPs from either HM3 or MEGA.

a

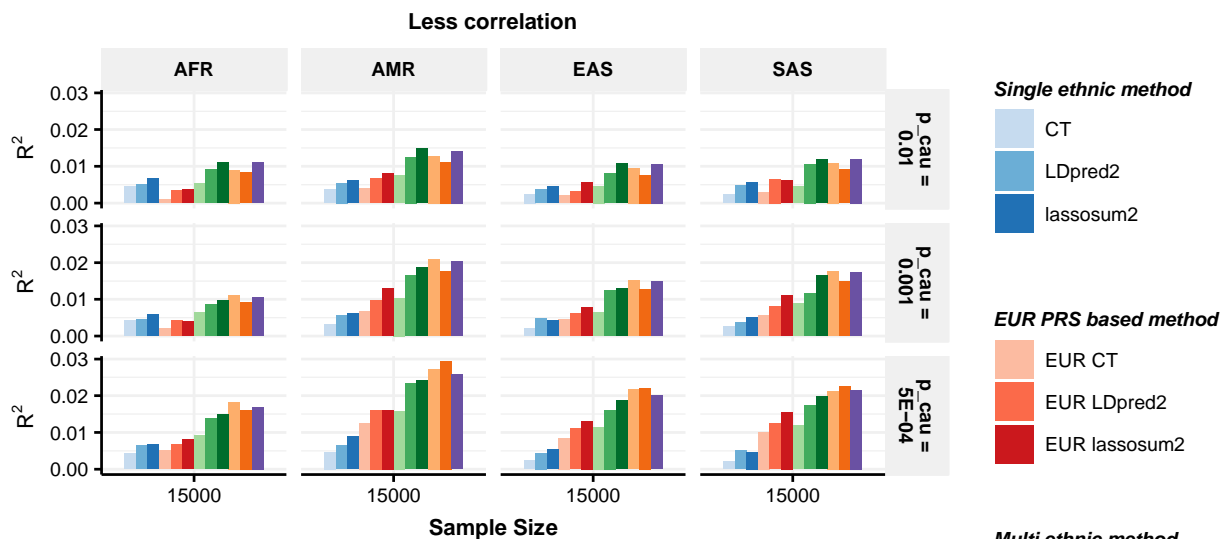


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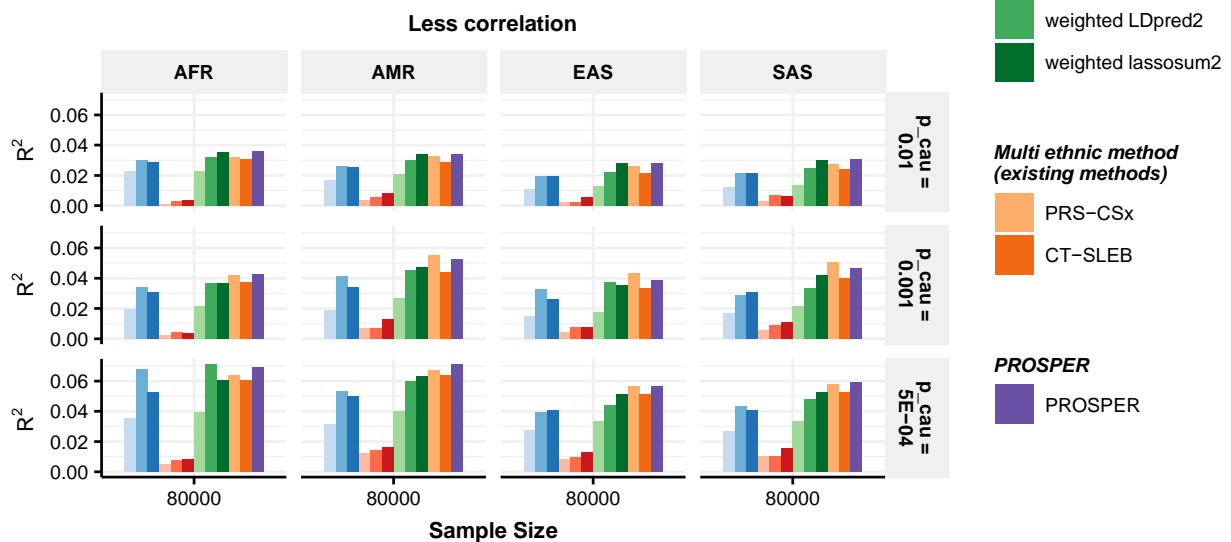


887 **Supplementary Figure 5: Performance of alternative methods on simulated data generated**
 888 **with different sample sizes and different genetic architectures.** Data are simulated for
 889 continuous phenotype under a strong negative selection model and three different degrees of
 890 polygenicity (top panel: $p_{causal} = 0.01$, middle panel: $p_{causal} = 0.001$, and bottom panel:
 891 $p_{causal} = 5 \times 10^{-4}$). Per-SNP heritability is assumed to be the same across all populations, and
 892 the correlations in effect sizes for share SNPs between all pairs of populations is fixed at 0.6.
 893 The sample sizes for GWAS training data are assumed to be (a) 15,000, and (b) 80,000 for the
 894 four non-EUR target populations; and is fixed at 100,000 for the EUR population. PRS generated
 895 from all methods are tuned in 10,000 samples, and then tested in 10,000 independent samples
 896 in each target population. The PRS-CSx package is restricted to SNPs from HM3, whereas other
 897 alternative methods use SNPs from either HM3 or MEGA.

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