Featural representation and internal noise underlie the eccentricity effect in contrast sensitivity

Abbreviated title: Computations underlying the eccentricity effect

Shutian Xue[1], Antonio Fernández[1] and Marisa Carrasco[1,2]

Affiliation:

[1]Department of Psychology, New York University, New York, United States

[2]Center for Neural Science, New York University, New York, United States

Corresponding author:

Shutian Xue (shutian.xue@nyu.edu)

Marisa Carrasco (marisa.carrasco@nyu.edu)

Number of pages: 54

Number of figures: 11

Number of tables: 2

Number of multimedia: 0

Number of 3D models: 0

Total number of words: 10,314

Abstract: 250/250

Introduction: 652/650

Discussion: 1,472/1,500

The authors declare no competing financial interests.

Acknowledgements: This research was supported by National Institutes of Health Grants R01-EY027401(M.C.).

Abstract

Human visual performance for basic visual dimensions (e.g., contrast sensitivity

and acuity) peaks at the fovea and decreases with eccentricity. The eccentricity effect is

related to the larger surface area of the visual cortex corresponding to the fovea, but it is

unknown if differential feature tuning contributes to this eccentricity effect. Here, we

investigated two system-level computations underlying the eccentricity effect: featural

representation (tuning) and internal noise. Observers (both sexes) detected a Gabor

embedded in filtered white noise which appeared at the fovea or one of four perifoveal

locations. We used psychophysical reverse correlation to estimate the weights assigned

by the visual system to a range of orientations and spatial frequencies (SFs) in noisy

stimuli, which are conventionally interpreted as perceptual sensitivity to the

corresponding features. We found higher sensitivity to task-relevant orientations and

SFs at the fovea than the perifovea, and no difference in selectivity for either orientation

or SF. Concurrently, we measured response consistency using a double-pass method,

which allowed us to infer the level of internal noise by implementing a noisy observer

model. We found lower internal noise at the fovea than perifovea. Finally, individual

variability in contrast sensitivity correlated with sensitivity to and selectivity for task-

relevant features as well as with internal noise. Moreover, the behavioral eccentricity

effect mainly reflects the foveal advantage in orientation sensitivity compared to other

computations. These findings suggest that the eccentricity effect stems from a better

representation of task-relevant features and lower internal noise at the fovea than at the

perifovea.

Significance

Performance in many visual tasks worsens with eccentricity. Many studies attribute this eccentricity effect to retinal and cortical factors, like higher cone density and a larger cortical surface area representing the foveal than peripheral locations. We investigated whether system-level computations for task-relevant visual features also underlie this eccentricity effect. Measuring contrast sensitivity in visual noise, we showed that the fovea better represents task-relevant orientation and spatial frequency and has lower internal noise than the perifovea, and that individual variability in these two computations correlates with that in performance. These findings reveal that both representations of these basic visual features and internal noise underlie the difference in performance with eccentricity.

Introduction

Visual performance worsens with eccentricity for many visual dimensions, including contrast sensitivity (Cannon, 1985; Baldwin et al., 2012; Jigo et al., 2023) and acuity (Strasburger et al., 2011; Anton-Erxleben and Carrasco, 2013). The behavioral advantage at the fovea relates to higher cone density (Polyak, 1941), smaller retinal receptive fields (Enroth-Cugell and Robson, 1966), and larger V1 area dedicated to processing stimuli (Virsu and Rovamo, 1979; Duncan and Boynton, 2003; Benson et al., 2021; Himmelberg et al., 2021). However, it is unknown whether system-level representations and computations for task-relevant features differ across eccentricities,

and whether they underlie the eccentricity effect in contrast sensitivity. Thus, we

investigated three questions:

(1) Does the representation of orientation and SF differ between fovea and

perifovea? Representation is considered the externalization of the internal neural

processes elicited by perceptual and cognitive processes. Functionally, it serves as a

template for observers to perform a given task (Ahumada, 2002; Gold et al., 2004;

Nagai et al., 2008) and its efficiency depends on resemblance to the signal (Burgess,

Wagner, Jennings and Barlow, 1981; Levi and Klein, 2002) and the ideal observer

template (Abbey and Eckstein, 2009). This representation reflects the system's tuning

properties–the sensitivity to and selectivity for features–and approximate

electrophysiological sensory tuning properties (Neri and Levi, 2006). Representation

modulations reflecting differential system-level computations underlie performance

changes due to expectation (Wyart et al., 2012) and covert spatial (Fernández et al.,

2019, 2022), feature-based (Paltoglou and Neri, 2012) and presaccadic (Li et al., 2016;

Ohl et al., 2017) attention.

Here, we investigated the representation of two fundamental visual features–

orientation and spatial frequency–which are jointly encoded (Blakemore and Campbell,

1969; De Valois et al., 1982a; Moraglia, 1989) and processed by parallel channels

tuned for orientation and SF (De Valois and De Valois, 1988; Graham, 1989). We used

psychophysical reverse correlation (Abbey and Eckstein, 2002; Ahumada, 2002;

Fernández et al., 2019, 2022; Wyart et al., 2012) in conjunction with a Gabor detection

task to simultaneously derive the representation of both orientation and SF.

(2) Does internal noise differ between fovea and perifovea? Our visual system is

perturbed by internal noise, as reflected in contrast (Pelli, 1985; Legge et al., 1987;

Ahumada, 2002), orientation (Heeley et al., 1997; Lu and Dosher, 1998) and motion

(Ling et al., 2009) tasks. Most studies estimate the level of internal noise only at the

fovea (Burgess and Colborne, 1988; Lu and Dosher, 1999) or periphery (Dosher and Lu,

2000; Eckstein et al., 2004; Mareschal et al., 2008b). When eccentricity increases,

internal noise has been reported to increase (motion: Mareschal et al., 2008a) or remain

constant (motion: Falkenberg and Bex, 2007; face: Mäkelä et al., 2001). Whether

internal noise increases with eccentricity for tasks mediated by basic visual features

(e.g., orientation and SF) is unknown.

We investigated whether internal noise limits the processing of basic visual

features–orientation and SF–between fovea and perifovea. First, we inferred the internal

noise level from response consistency, assessed with a double-pass method via

multiple presentations of identical noisy stimuli (Burgess and Colborne, 1988). Second,

we estimated the internal noise level by implementing a noisy observer model, which

predicts trial-wise responses based on the measured response consistency.

(3) Do differential computations mediate behavioral differences between fovea

and perifovea? Individual differences in contrast sensitivity correlate with surface area in

early occipital areas (Himmelberg et al., 2022), and with internal noise in humans

(McAnany and Park, 2018) and macaques (Kiorpes et al., 1999). We tested whether the

eccentricity effect can be accounted for by the location difference in computations by

assessing the correlation between the individual variability in computation and behavior.

In sum, we investigated system-level computations underlying the eccentricity

effect in contrast sensitivity by: (1) comparing the representation of orientation and SF

and the level of internal noise between fovea and perifovea, and (2) testing the

hypothesis that the eccentricity effect in contrast sensitivity reflects discrepancies in

featural representation and internal noise between fovea and perifovea.

102   **Materials and Methods**

Participants

Twelve observers (8 females, age: 23-28, mean=26.25, SD=2.36) participated in

the study. Ten were experienced psychophysical observers and all but one (author, SX)

were naïve as to the purpose of this study. All observers provided written informed

consent and had normal or corrected-to-normal vision. All experimental procedures

agreed with the Helsinki declaration and were approved by the University Committee on

Activities Involving Human Subjects at New York University.

Apparatus

All stimuli were generated and presented using MATLAB (MathWorks, Natick,

MA) and the Psychophysics Toolbox (Kleiner et al., 2007) on a gamma-linearized CRT

monitor (1280×960 screen resolution; 100 Hz; 33 cd/m2 background luminance).

Observers viewed the display at 57 cm with their heads stabilized by a chin rest. An

eye-tracker system (EyeLink 1000) was in front of the observer to track eye position.

Experimental Design

Stimuli

Figure 1 illustrates one trial. A black fixation cross (arm length=0.3º) was

presented at the center of the screen before the stimulus onset. Five black placeholders,

each composed of four corners (line length=0.75°, delimiting a virtual 4°×4° square)

were simultaneously presented at the fovea and four isoeccentric perifoveal locations at

6º eccentricity: two along the horizontal meridian and two along the vertical meridian.

Placeholders indicate the location of the target and four distractors with 100% validity to

minimize spatial uncertainty because it increases with eccentricity (Hess and Hayes,

1994; Michel and Geisler, 2011). Observers were instructed to report whether they

detected a Gabor in the target. In half of the trials, the target consisted of a noise patch

(20% RMS contrast, containing SFs within 1-4 cpd with uniform power). And in the other

half, the target contained a horizontally oriented Gabor embedded in the noise patch.

The Gabor was generated by modulating a 2-cpd sinewave delimited by a Gaussian

envelope (SD=0.8, width=3º). The phase of the Gabor was random on each trial to

avoid adaptation. The root-mean-square (RMS) contrast of the Gabor was determined

independently at each location for each observer through a titration procedure to yield

70% accuracy. All four distractors were noise patches independently generated at each

location on each trial.

Procedure

Each trial began with a fixation cross at the center of the screen. Then, five

placeholders were presented simultaneously for 400 ms followed by the display of

target and distractors for 100 ms. Target location was blocked and observers were

informed of the target location at the start of the block to eliminate spatial uncertainty.

Placeholders were on the screen until observers made a response. Accuracy was

emphasized and there was no time limit to respond. Auditory feedback was given after

each response: A high-tone beep indicated correct and a low-tone beep indicated

incorrect responses. At the end of each block, accuracy was displayed on the screen.

To infer the amount of internal noise in the system, we measured response

consistency by employing the double-pass method (Burgess and Colborne, 1988). In

each block, the same target patch was presented twice, once in the first half of the trials

and once in the second half in randomized order. Response consistency is the

proportion of trial pairs on which the observer gave the same response, regardless of

the correctness (chance level=50%).

Before the main experiment, each observer completed a titration session using

the same task at each location to determine the RMS contrast of the Gabor to maintain

the overall performance at 70%. Gabor contrast was adjusted using a PEST method

(Pentland, 1980) and averaged over two interleaving staircases each containing 50

trials.

After practice sessions, observers completed on average 4,317 trials (±262 trials)

per location. All observers completed an equal number of trials per location, during 10-

15 one-hour sessions. The Gabor contrast in the main experiment was adjusted

according to the accuracy of the previous block: if the accuracy was higher/lower than

70%, the contrast was reduced/increased proportionally. Contrast sensitivity was

calculated by taking the reciprocal of the contrast threshold of the last ten sessions in

which titrated contrast thresholds were stable.

Eye-tracking

Eye position was monitored online using Eyelink1000. If the observer's gaze

deviated more than 1.5 dva away from the fixation center before the response, the trial

was aborted immediately and repeated at the end of the block.

Reverse correlation

Figure 2 illustrates the procedure of reverse correlation to estimate weights

assigned by the system to different orientations and SFs (Ahumada, 2002; Li et al.,

2016; Fernández et al., 2022). Reverse correlation enables experimenters to discretize

the stimulus feature by varying the fine-grainness of sampling in an efficient and flexible

way. First, we quantified the energy profile ($E_{\theta,f}$) of the noise patch in the noise (N) by

dot-multiplying the noise patch with a pool of Gabor-filters ($g_{\theta,f}$), defined by θ and f, in

quadrature phase and have the same size of the target (**Fig 2A**, *Step 1)*. These Gabor

filters approximate the linear filters in V1 that are tuned to a specific orientation and SF.

$$E_{\theta,f} = \sqrt{\left(N \cdot g_{\theta,f,sin}\right)^2 + \left(N \cdot g_{\theta,f,cos}\right)^2} \qquad [1]$$

where θ stands for orientation, ranging from -90 to 90° in 29 steps relative to and

centered at the Gabor orientation (0°), and f stands for SF, ranging from 1 to 4 cpd in

29 log steps, centered at the Gabor SF (2 cpd). No energy beyond 4 cpd was detected.

Second, we regressed the energy pixels of a certain orientation-SF component

(concatenated across all trials) with the binary behavioral responses using a probit link

function (**Fig 2B**, *Step 2*)

$$p(yes) = \Phi^{-1}\left(\beta_1 E_{\theta,f} + \beta_0\right) \qquad [2]$$

where $\Phi^{-1}(.)$ indicates the probit function (i.e., the inverse cumulative normal function),

β_0 is the intercept and the regression slope β_1 is interpreted as the perceptual sensitivity

to the corresponding orientation and SF. A more positive β_1 indicates a higher sensitivity

while $\beta_1 = 0$ indicates no sensitivity. Before the regression, energy pixels at each

orientation-SF channel across all trials were standardized (i.e., subtracting the mean

across trials and divided by SD) for each Gabor contrast. We report results derived from

Gabor-absent trials, which is a more conservative way to implement reverse-correlation

analysis (e.g., Eckstein et al., 2002, 2004; Shimozaki et al., 2005).

Third, we marginalized the kernel mapping across SF to obtain orientation

sensitivity kernels and across orientation to obtain SF sensitivity kernels (**Fig 2C**, *Step*

3). Before this step, we assessed the separability between orientation and SF,

calculated as the correlation (r^2) between the raw kernel mapping and the reconstructed

mapping (i.e., the outer product of the orientation and SF kernels), to ensure that

marginalizing the 2D mapping would not lose feature information.

Orientation and SF tuning functions

To capture the sensitivity to different orientations, we fitted a scaled Gaussian

function to the orientation sensitivity kernels assuming that sensitivity kernels peaked at

the Gabor orientation (i.e., 0°)

$$\beta(\theta) = \alpha\phi(\theta, 0, \sigma) + b \qquad [3]$$

where $\Phi[\cdot]$ stands for the normal probability function, α stands for the gain, σ stands for

the Gaussian SD, and b stands for the baseline shift.

To capture the sensitivity to different SFs, we fitted a log parabola function to the

SF sensitivity kernels (adapted from Watson & Ahumada, 2005).

$$\beta(f) = \alpha\left[10^{-\frac{log_{10}\left(\frac{f}{f_0}\right)}{\sigma}}\right] + b\ldots\ldots\ldots[4]$$

where α, σ and b indicates SF gain, standard deviation and baseline, respectively.

f_0 stands for the peak of the function. We used the MATLAB *fmincon* function to search

for the parameters (α, σ and b for orientation and α, σ, b and f_0 for SF) using the Least-

Square optimization method. Sensitivity kernels at the fovea and perifovea were fitted

together. The parameter structure was varied by whether a parameter was fixed or free

to vary between fovea and perifovea, yielding 8 candidate models for the orientation

domain and 16 models for the SF domain. The best fitting candidate model was

determined by model comparison with a 10-fold cross-validation method based on

deviance averaged across 10 folds. The optimization procedure was repeated 20 times

using the MATLAB *MultiStart* function with randomized initial parameters each time and

the best fit was selected. Before testing the candidate models, we also compared two

families for orientation (scaled Gaussian vs. difference of Gaussians (DoG)) and for SF

(log parabola with vs. without truncation). In **Figure 7** we will present the model

comparison results.

Tuning characteristics

Finally, we extracted tuning characteristics (i.e., values characterizing the tuning

function, illustrated in **Fig 2D,** *Step 4*) for comparison across locations. See **Table 1** for

a complete list of tuning parameters and characteristics. We calculated the inter-

parameter correlation which is the partial correlation between two parameters across

observers while controlling for location. Compared to estimated parameters (e.g., gain

and width), tuning characteristics (e.g., peak amplitude and bandwidth) have lower inter-

parameter correlations, indicating higher independence among them and thus higher

interpretability (**Figure 3**). Note that the gain and peak amplitude are two

computationally different measures that reflect different tuning properties: Gain indicates

the general amplitude amplification which depends on the baseline, whereas the peak

amplitude indicates the highest sensitivity the system can achieve which is independent

of the baseline.

Unlike many studies that only studied a single location or visual field (Hilz and

Cavonius, 1974; Rovamo et al., 1978; Levi and Klein, 1986), we pooled four polar

angles to represent the perifovea as they reflect the behavioral heterogeneity at

perifovea (Carrasco et al., 2001; Barbot et al., 2021) so that we can generalize the

revealed tuning properties.

Statistical Analysis

After deciding the best fitting function via model comparison to characterize

orientation and SF sensitivity kernels, we used a bootstrapping procedure to derive

median and confidence interval (CI) of the tuning characteristics. In each iteration, we

resampled the trials with replacement, regenerated sensitivity kernels using reverse

correlation, and extracted tuning characteristics at the fovea and perifovea. After

repeating this process for 1,000 times, we calculated and reported the median and 68%

CI (which equals an SE, Wichmann and Hill, 2001) of all measurements and estimates

for each observer. We then reported the mean and ±1 SEM across observers to

represent the group level.

To compare performance (e.g., d-prime, criterion, contrast sensitivity and

response consistency) and the computations (i.e., tuning characteristics and estimated

internal noise) between fovea and perifovea, we conducted within-subjects ANOVAs

and pair-wise two-tailed *t*-tests. When appropriate (e.g., when comparing tuning

functions at each channel), we conducted Bonferroni's correction to adjust for multiple

comparisons.

Noisy observer model

To assess the validity of the sensory representation and to estimate internal

noise, we implemented a *noisy observer model* to predict trial-wise responses based on

the *energy profile* of each trial and the *template* (i.e., the featural representation derived

from Gabor-absent trials via reverse correlation). **Figure 4** illustrates the schematic of

the model. **Table 2** lists all model parameters. Each iteration consisted of four steps.

First, the data at each location were reshuffled and split into two sets (**Fig 4**, *step 1*).

The first set was used to derive the template ($T_{\theta,f}$) using reverse correlation (as

described above) and the second set was used to estimate free parameters by

providing trial-wise energy profiles ($E_{\theta,f}$) and responses.

Second, we calculated the *internal variable* (V_i) for each trial in the test set by

taking the sum of the dot product between the derived template and the energy profile

of each trial (**Fig 4**, *step 2*)

$$V_i = \sum_{\theta}\sum_{f} T_{\theta,f} E_{\theta,f} \dots\dots\dots\dots[5]$$

where V_i stands for the internal variable of each trial. The internal variable forms the

basis of trial-wise decision making, following the decision variable assumption of Pelli,

(1985). Assuming that internal variables of all Gabor-present (V_{S+}) and -absent trials

(V_{S-}) are normally distributed, means (μ_{S+} and μ_{S-}) and variances (σ_{S+}^2 and σ_{S-}^2) were

extracted. Note that the integration of template and energy profile is a convolution in the

spatial domain.

Third (**Fig 4**, *step 3*) we derived noisy internal variable ($V_{i,noisy}$) by adding two

types of internal noise to V_i : (1) *induced noise,* assumed to be drawn from a normal

distribution whose width is proportional to the external variability (Burgess and Colborne,

1988; Pelli and Blakemore, 1990; Lu and Dosher, 2008; Diependaele et al., 2012) and

(2) *constant noise*, whose variance is independent of external variability. Thus, noisy

internal variables for Gabor-present and -absent trials, respectively, are characterized

by a normal distribution

$$V_{S+,noisy} \sim \mathcal{N}\left(\mu_{S+}, \sqrt{\sigma_{S+}^2 + (N_{IE}\sigma_{S+})^2 + \sigma_{constant}^2}\right) \ldots\ldots\ldots\ldots[6]$$

$$V_{S-,noisy} \sim \mathcal{N}\left(\mu_{S-}, \sqrt{\sigma_{S-}^2 + (N_{IE}\sigma_{S-})^2 + \sigma_{constant}^2}\right) \ldots\ldots\ldots\ldots[7]$$

where $\mathcal{N}(.)$ stands for Normal distribution, N_{IE} is the internal-external noise ratio and

$\sigma_{constant}^2$ is the variance of the constant noise.

Finally (**Fig 4**, *step 4*), the response of this trial in the test set was determined by

comparing $V_{i,S+,noisy}$ (or $V_{i,S-,noisy}$), to a threshold (k). The response is present if $V_i \geq k$

and absent if $V_i < k$.

The model was fitted to each observer's responses by minimizing the negative

log-likelihood of correct responses consistent pairs (Lu and Dosher, 2008; Fernández et

al., 2022),

$$L_{Pc} = n_{Pc} ln(Pc) + (m_{Pc} - n_{Pc}) ln(1 - Pc) + f(m_{Pc}, n_{Pc}) \dots [8]$$

$$L_{Pa} = n_{Pa} ln(Pa) + (m_{Pa} - n_{Pa}) ln(1 - Pa) + f(m_{Pa}, n_{Pa}) \dots [9]$$

where n_{Pc} is the total number of correct trials, m_{Pc} is the total number of trials, Pc is the

predicted proportion of correctness, n_{Pa} is the total number of consistent trial pairs, m_{Pa}

is the total number of trial pairs, and Pa is the predicted response consistency.

$f(m_P, n_P)$ is a function to calculate the factorial term in log space (Lu and Dosher, 2008)

$$f(m_P, n_P) = \sum_{m=1}^{m_P} ln(m) - \sum_{n=1}^{n_P} ln(n) - \sum_{o=1}^{m_P - n_P} ln(o) \dots [10]$$

The expression of predicted response consistency is derived below. Take Gabor-

present trials as an example, the response consistency is the proportion of Gabor-

present trials (S^+) when either the responses of two passes (R_A and R_B) are both

present ($R_A = R_B = 1$) or absent ($R_A = R_B = 0$)

$$
\begin{aligned}
Pa_{S+} \quad &= p(R_A = R_B = 1|S^+) + p(R_A = R_B = 0|S^+) \\
&= \int_{x=-\infty}^{\infty} p(V_{S+} = x|S^+)[p(R_A = R_B = 1|V_{S+} = x, S^+) + p(R_A = R_B = 0|V_{S+} = x, S^+)] \\
&= \int_{x=-\infty}^{\infty} p(V_{S+} = x|S^+)[p(R = 1|V_{S+} = x, S^+)^2 + p(R = 0|V_{S+} = x, S^+)^2]
\end{aligned}
$$

318   [11]

where

$$p(V_{S+} = x|S^+) = \Phi\left(x, \mu_{S+}, \sqrt{\sigma_{S+}^2 + (N_{IE}\sigma_{S+})^2 + \sigma_{constant}^2}\right) \dots [12]$$

$$p(R = 1|V_{S+} = x, S^+) = \Phi\left(k, \mu_{S+}, \sqrt{\sigma_{S+}^2 + (N_{IE}\sigma_{S+})^2 + \sigma_{constant}^2}\right) \dots [13]$$

$$p(R = 0|V_{S+} = x, S^+) = 1 - p(R = 1|V_{S+} = x, S^+) \dots [14]$$

x is any possible values V_{S+} could take. Therefore, the predicted response consistency

for Gabor-present and -absent trials are

$Pa_{S+} =$

$\int_{x=-\infty}^{\infty} p(V_{S+}|S^+) \left[\Phi\left(k, \mu_{S+}, \sqrt{\sigma_{S+}^2 + (N_{IE}\sigma_{S+})^2 + \sigma_{constant}^2}\right) \right]^2 +$

$1 - \Phi k, \mu S+, \sigma S+2 + NIE\sigma S+2 + \sigma constant 22$..[15]

$Pa_{S-} =$

$\int_{x=-\infty}^{\infty} p(V_{S-}|S^-) \left[\Phi\left(k, \mu_{S-}, \sqrt{\sigma_{S-}^2 + (N_{IE}\sigma_{S-})^2 + \sigma_{constant}^2}\right) \right]^2 +$

$1 - \Phi k, \mu S-, \sigma S-2 + NIE\sigma S-2 + \sigma constant 22$..[16]

and the predicted response consistency of all trials is $Pa = \frac{Pa_{S+} + Pa_{S-}}{2}$

Finally, to evaluate the validity of the feature representation, we compared the

predictive power of four model variations defined by whether the model contains any

randomness (adapted from Beard and Ahumada, 1999):

(1) *Core* model: neither the template nor the trial-wise energy profiles were

random;

(2) *Random template* model: pixels of the template were randomized on each

trial;

(3) *Random energy* model: pixels in each energy profile were randomized;

(4) *Random template and energy* model: pixels in each energy profile and in the

template were randomized.

The rationale is that a model without any randomness (i.e., to fully utilize the

information of the template and energy profiles) should outperform other models. We

calculated the Bayesian information criterion (BIC) score for each model variation at

each location using $BIC = -2ln(\hat{L}) + kln(n)$, where $\hat{L}$ is the likelihood given the optimal

parameter combinations of the candidate model, k is the number of parameters in the

candidate model and n is the number of data points. We reported differences in BIC

scores $\Delta BIC_i = BIC_i - BIC_{i,min}$. The best model is when $\Delta BIC_i \equiv \Delta_{min} \equiv 0$.

Correlations between behavior and computations

To bridge behavior (i.e., contrast sensitivity) and computations (i.e., tuning

characteristics and internal noise), we assessed the correlations between contrast

sensitivity with tuning characteristics and with estimated internal noise across observers.

To ensure that the correlations are not driven by possible location differences (e.g.,

higher orientation gain at fovea), we conducted a partial correlation analysis while

controlling for the location effect.

Then, to further pinpoint the computation that underlies the eccentricity effect, we

calculated the correlations between the eccentricity effect index (EEI):

$$EEI = \frac{fovea - perifovea}{fovea + perifovea} \quad [17]$$

363   for contrast sensitivity, with tuning characteristics and with estimated internal noise. A

364   positive index means that the assessed value at the fovea is higher than the perifovea

365   and vice versa. A more positive or negative index indicates a larger difference between

366   fovea and perifovea. To visualize the observer effect, we plot individual data with group

367   averages across locations subtracted.

369   **Results**

***d*-prime and criterion are equated between the fovea and perifovea**

Figure 5A- shows that there was no difference between fovea and perifovea in

d-prime (t(11) = -1.02, p > 0.1) or criterion (t(11) = 1.17, p > 0.1), indicating matched

performance across locations. This matched performance ensures that any change in

the perceptual sensitivity is attributed to the trial-to-trial energy fluctuations rather than

to task difficulty or the probability of detecting the target.

Higher contrast sensitivity and higher response consistency at the fovea

Figure 5C and 5D show the contrast sensitivity and response consistency

measured using a double-pass method (Burgess and Colborne, 1988). Contrast

sensitivity was higher at fovea than perifovea (t(11) = 10.07, p < 0.001), confirming the

established eccentricity effect, which was present for all observers. There was also a

higher response consistency at the fovea than perifovea ($t(11) = 6.06$, $p < 0.001$; true

for all but one observer for whom there was no difference), suggesting lower internal

noise in the system at the fovea (Burgess and Colborne, 1988). This response

consistency measurement is the key to directly estimating the level of internal noise via

the implementation of a noisy observer model.

Compare 2D kernel mapping

The sensitivity kernel mapping at the fovea and perifovea was derived using

reverse correlation from Gabor-absent trials only (**Figure 6A-B**). Each pixel indicates

the perceptual sensitivity, i.e., how influential those two channels are in the perception

of the stimuli at a certain orientation and SF channel. **Fig 6C** shows that the most

pronounced positive differences between the fovea and perifovea were found near the

Gabor features (represented by the red horizontal and vertical lines), suggesting that

the fovea is more sensitive to the task-relevant features. **Fig 6D** shows the ideal

template which is the energy profile of the signal Gabor. We confirmed that orientation

and SF dimensions are separable for most observers (**Fig 6E**). The higher the

correlation, the more separable the two dimensions are and the less information is

expected to be lost by examining the two dimensions separately. High separability

justifies marginalization along each dimension to derive orientation and SF tuning

functions (**Figure 7**).

Comparing orientation and SF tuning functions

Figure 7 shows the marginalized kernels and the fitting functions for orientation

(**Fig 7A,** red dots) and SF (**Fig 7B,** blue dots). First, we conducted a within-subjects

two-way ANOVA (fovea vs. perifovea x 29 channels) on the marginalized orientation

and SF kernels separately. There was a main effect of the channel for both orientation

($F(28,308) = 54.00$, $p < 0.001$) and SF ($F(28,308) = 16.93$, $p < 0.001$). Location had a

main effect for both orientation ($F(1,11) = 68.37$, $p < 0.001$) and SF ($F(1,11) = 133.32$, p

< 0.001); kernels were higher at fovea than perifovea across channels. Bonferroni-

corrected post-hoc one-tailed *t*-tests (indicated by grey bands in **Fig 7A-B**, $p < 0.05$)

revealed foveal advantage over perifovea near the task-relevant orientation and SFs

and at higher SFs.

Then, to identify the tuning mechanisms, we captured the shape and magnitude

of the marginalized kernels by fitting a scaled Gaussian function to the orientation

kernels (**Fig 7A,** red curves) and a log parabola function to the SF kernels (**Fig 7B,** blue

curves). Model comparisons revealed that orientation kernels were best fitted by the

candidate model with all parameters free to vary (**Fig 7C**) and SF kernels by the

candidate model with only baseline (*b*) fixed (**Fig 7D)**. The two best candidate models fit

the data well ($R^2 = 83.5\%$ for orientation and $R^2 = 88\%$ for SF on average), allowing us

to extract tuning characteristics from the fitting function as a proxy of tuning properties

(**Fig 2D** and **Table 1**).

Comparing tuning characteristics between the fovea and perifovea showed that

for orientation (**Fig 8, top row**), there was a higher gain (t(11) = 5.62, p < 0.001) and

higher peak amplitude (t(11) = 7.79, p < 0.001) at fovea than perifovea, but no

difference in bandwidth or baseline (p > 0.1). The differences in orientation gain and

peak were highly consistent across individual observers. The results indicate that

observers were more sensitive to orientation in general, especially to the task-relevant

and neighboring orientations at fovea than perifovea.

For SF, there was a higher gain (**Fig 8E**, t(11) = 6.55, p < 0.001) and a higher

peak amplitude (**Fig 8F**, t(11) = 6.39, p < 0.001) at fovea than perifovea, indicating that

observers were more sensitive to SFs in general, especially to the task-relevant SFs.

There was a marginal difference in SF bandwidth in octave (**Fig 8G**, t(11) = -1.83, p =

0.094), 9 out of 12 observers showed narrower SF bandwidth at the fovea. Given the

presence of one outlier at each fovea and perifovea, we calculated both mean (fovea:

0.95, perifovea: 1.15) and median (fovea: 0.95, perifovea: 1.09), and the statistical

results do not change. Finally, tuning functions peaked at a higher SF at fovea than

perifovea (**Fig 8H,** t(11) = 2.57, p = 0.026); this was the case for 10 out of 12 observers.

Moreover, the peak of SF tuning function at the fovea did not differ from the signal SF (2

cpd, indicated by the gray horizontal line in **Fig 8H**, p > 0.1) whereas the peak at the

perifovea was lower than the signal SF (p < 0.001).

Estimate internal noise via computational modeling

We could infer that fovea has lower internal noise because we observed a higher

response consistency at the fovea than perifovea (**Fig 5**). To obtain the numerical

estimation of internal noise, we implemented a noisy observer model to trial-wise

responses by modeling the internal response and induced and constant internal noise

(**Fig 4**). First, the predicted accuracy (**Fig 9A**) and response consistency (**Fig 9B**) match

the data well; most dots lie near the diagonal (p > 0.1), indicating that the model has

good predictive power. Second, response consistency negatively correlated with

induced noise (**Fig 9C**, partial r = -0.80, p < 0.001) and constant internal noise (**Fig 9D**,

partial r = -0.70, p < 0.001) as expected (Burgess and Colborne, 1988). Third, most

observers showed lower induced noise (10 out of 12, t(11) = -2.50, p = 0.029, **Fig 9E**)

and constant noise (9 out of 12, p > 0.1, **Fig 9F**) at the fovea than perifovea. In sum, the

noisy observer model successfully predicted trial-wise responses and found constant

internal noise lower at the fovea than perifovea.

To assess the validity of the template (i.e., the representation for orientation and

SF derived from partial data using reverse correlation), we compared the predictive

power of four model variations (**Figs 9G-H**). The Core model outperformed all three

random models (p = 0.019/0.019/0.016 (fovea) and p = 0.056/0.062/0.068 (perifovea),

respectively). The three random models did not differ in their predictive power (p > 0.1).

The results indicate that featural representation is crucial to predicting trial-wise

responses.

Linking tuning characteristics with contrast sensitivity

So far we have shown that feature representation differs between fovea and

perifovea in sensitivity to and selectivity for orientation and SF. To bridge behavior and

computations, we assessed the partial correlation between contrast sensitivity with

tuning characteristics and with internal noise across observers (**Figure 10**). Contrast

sensitivity correlates (1) positively with orientation gain (**Fig 10A;** partial r = 0.68, p <

0.001) and with orientation peak amplitude (**Fig 10B;** partial r = 0.75, p < 0.001),

indicating that higher contrast sensitivity is associated with higher sensitivity to task-

relevant orientations, (2) negatively with orientation bandwidth (**Fig 10C**; partial r = -0.41,

p = 0.049) and with SF bandwidth (**Fig 10F**; partial r = -0.50, p = 0.016), indicating that

higher contrast sensitivity is associated with higher selectivity for task-relevant features,

(3) negatively with orientation baseline (**Fig 10D**, partial r = -0.42, p = 0.047), indicating

that high contrast sensitivity is associated with the inhibition on orientations that are

distant to the signal orientation, which is task-irrelevant (reflected by low or negative

values in the inset of **Fig 10D**).

Then, we assessed the correlation between contrast sensitivity and internal noise.

There was a negative correlation between contrast sensitivity and both induced noise

(**Fig 10G**; partial r = -0.50, p = 0.014) and constant noise (**Fig 10H**; partial r = -0.68, p <

0.001), indicating that better detection was associated with lower internal noise.

Figure 11 shows the association between the eccentricity-effect index EEI for

contrast sensitivity and computations. We found that an increase in the EEI of contrast

sensitivity is associated with a stronger foveal advantage in the orientation peak

amplitude (**Fig 11A,** r = 0.76, p = 0.004) and with narrower orientation bandwidth at the

fovea (**Fig 11B,** r = -0.79, p = 0.002). No other correlations between EEI for contrast

sensitivity and tuning characteristics were significant. These results indicate that the

sensitivity to and selectivity for orientation play important roles in the underlying

perceptual computations that give rise to the eccentricity effect. However, the EEI of

internal noise does not correlate with the behavioral eccentricity effect (**Fig 11C and D,**

p > 0.1), suggesting that differences in internal noise between fovea and perifovea do

not predict the extent of the eccentricity effect in contrast sensitivity.

Discussion

We investigated whether performance differences between the fovea and

perifovea, which result in an eccentricity effect (Rijsdijk et al., 1980; Carrasco et al.,

1995, 1998; Carrasco and Frieder, 1997; Pointer and Hess, 1989; Baldwin et al., 2012),

can be accounted for by distinct system-level computations –featural representation and

internal noise. Using reverse correlation and the double-pass method, we established a

link between performance and system-level computations and found that the behavioral

advantage at the fovea stems from a better representation of task-relevant orientations

and SFs, as well as lower internal noise.

Featural representations differ between fovea and perifovea

Orientation sensitivity. We observed a higher orientation gain and peak amplitude

at the fovea than perifovea indicating stronger sensitivity to task-relevant orientation at

the fovea. A possible neural correlate is that the fovea is represented by more retinal

ganglion cells (Wässle, Grünert, Röhrenbeck and Boycott, 1989) and higher neuron

count with V1 neuron with density approximately uniform across the cortex (Hubel and

Wiesel, 1977; Rockel et al., 1980).

Orientation selectivity. We found no difference in orientation bandwidth (fovea:

19.49°±1.65°, perifovea: 19.95°±2.35°), indicating a similar level of orientation selectivity.

The similar orientation bandwidth we found may seem surprising because the receptive

field (RF) size of V1 neurons increases with eccentricity (Hubel and Wiesel, 1962) and

V1 neurons (simple and complex together) with smaller RF have higher orientation

selectivity (Watkins and Berkley, 1974). However, whereas complex V1 neurons'

orientation selectivity decreases with RF size, simple V1 neurons' orientation selectivity

is either less (Watkins and Berkley, 1974) or not (Schiller et al., 1976b) correlated with

RF size, and featural representation estimated by reverse correlation mainly

approximates activity of simple V1 neurons (Neri and Levi, 2006). Indeed, a similar

orientation bandwidth has been reported at 10° eccentricity in other reverse correlation

studies (Li et al., 2016; Fernández et al., 2019). These findings are consistent with

monkey studies reporting that on average simple V1 neurons at fovea and parafovea do

not differ in orientation bandwidth (De Valois et al., 1982b; Xu et al., 2007). We note,

however, that although orientation selectivity did not differ in our detection task, it may

differ in a discrimination task.

SF sensitivity. We found (1) higher gain and peak amplitude at the fovea than

perifovea, indicating higher sensitivity to the signal SF at the fovea, and (2) that SF

tuning function peaked near the signal SF (i.e., 2 cpd) at the fovea (1.86±0.09 cpd) but

at a lower SF at the perifovea (1.53±0.10 cpd), indicating a better task-relevant SF

representation at the fovea. These findings are consistent with a classification imaging

study in which observers assigned more weight to lower SFs at the parafovea (1.25°

and 2.5° eccentricity) than at the fovea (Levi and Klein, 2002). Similarly, in previous

reverse correlation studies the SF tuning function peaked at a lower SF than the signal

SF at 10° eccentricity (Li et al., 2016; Fernández et al., 2019). Together, these studies

indicate that the eccentricity effect in featural representation varies from fovea to

parafovea and perifovea.

The finding that lower SFs are represented as more influential with increasing

eccentricity is consistent with the fact that V1 simple cells prefer lower SF as

eccentricity increases in macaques (De Valois et al., 1982b; Schiller et al., 1976a) and

cats (Movshon et al., 1978a), (2) more cortex is devoted to higher SFs in the fovea than

parafovea (Xu et al., 2007), and (3) the contrast sensitivity function in humans peaks at

a higher SF at the fovea than perifovea (Hilz and Cavonius, 1974; Robson and Graham,

1981; Wright and Johnston, 1983; Jigo et al., 2023). Preference for higher SF at the

fovea can be explained by the fact that V1 simple cells at the fovea have smaller RF

(Hubel and Wiesel, 1962), which yields a preference for higher SF (Movshon et al.,

1978a, 1978b). Similarly, according to fMRI human studies, V1 voxels' sensitivity to the

signal SF decreases with eccentricity (Henriksson et al., 2008; Aghajari et al., 2020).

Importantly, the mismatch between the SF tuning functions at the perifovea and

the signal SF indicates that the task demand (at least for this detection task) cannot

override the preference of lower SFs at the perifovea. This finding questions the

assumptions of models, like the linear amplifier model (Pelli, 1981; Lu and Dosher, 2008)

and perceptual template model (Lu and Dosher, 1998, 2008, 2023), which match the

template to the signal without taking visual field location into account. Future models

(e.g., Lago, Abbey and Eckstein, 2021) should take into account the differential tuning

features across eccentricity.

SF selectivity. We found that SF bandwidth does not differ at the fovea (median:

0.95±0.08 octave) and perifovea (median: 1.09±0.10 octave), indicating similar

selectivity for task-relevant SFs at these locations. This estimation is similar to the

typical 1 octave estimation of neurons (Blakemore and Campbell, 1969; Jakobsson and

Lennerstrand, 1985) and voxels (Aghajari et al., 2020) in human V1. There are mixed

findings regarding SF selectivity across eccentricity. On the one hand, similar SF

bandwidth across eccentricity has been reported for V1 neurons in monkeys (De Valois

et al., 1982a; Foster et al., 1985) and V1 voxels in humans (Broderick et al., 2022). On

the other hand, another fMRI study reported that the SF bandwidth of V1 voxels

increases with eccentricity (Aghajari et al., 2020). This discrepancy might relate to the

task design: Whereas Broderick et al (2022) used stimuli that span a broad range of

orientations and SFs, Aghajari et al. (2020) used stimuli defined by narrow SFs.

Internal noise differs between fovea and perifovea

Our findings indicate that lower internal noise is another important factor

underlying the performance advantage at the fovea. First, we inferred the level of

internal noise from the response consistency using the double-pass method measured

at the fovea (73%±1%) and perifovea (69%±1%). These values are similar to previous

reports at the fovea (Burgess and Colborne, 1988; Murray et al., 2002; Ratcliff et al.,

2018; Vilidaite and Baker, 2017).

Then, we directly estimated the level of internal noise using a noisy observer

model for induced (fovea: 0.14±0.07; perifovea: 0.41±0.12) and constant (fovea:

0.05±0.01; perifovea: 0.06±0.01) noise. Our estimation of the induced noise is lower

than previous reports at the fovea (0.65-1.3; Burgess and Colborne, 1988; Levi and

Klein, 2003; Gold et al., 2004). This difference can be explained by: (1) internal noise

estimated using static stimuli varying in basic features or yes-no detection tasks, like the

current study, tend to fall on the lower end of the distribution (0.2-0.5; Neri, 2010); (2)

the induced noise estimated here is inevitably lower than the estimated noise in

previous studies, which measured both types of internal noise with a common

parameter.

We found internal noise lower at the fovea than perifovea for the basic features

of orientation and SF. For a fixed-size stimulus, more neurons at the fovea may average

out the internal noise (Kara et al, 2000). This finding is consistent with varying internal

noise with eccentricity for motion (Mareschal et al., 2008a) and depth (Falkenberg and

Bex, 2007; Wardle et al., 2012). However, there is similar internal noise across

eccentricities in a face identification task (Mäkelä et al., 2001). Internal noise levels

across eccentricity may differ among tasks because they rely on distinct cortical areas.

Differential computations mediate behavioral differences between fovea and

perifovea

This is the first study to link performance and two system-level computations –

featural representation and internal noise– in humans. First, a correlation analysis

revealed that higher contrast sensitivity at fovea than perifovea can be attributed to (1) a

better representation of task-relevant features and (2) less internal noise in the system.

Correlations have been established between neural structure with performance (Duncan

and Boynton, 2003; Song et al., 2015; Himmelberg et al., 2022) and with computations

(Neri and Levi, 2006). Thus, our findings on the correlation between performance and

computations can inform the interaction among behavior, computation and neural

structure. Second, we tested the hypothesis that the eccentricity effect is reflected in the

differential computations between fovea and perifovea. Our analyses revealed that the

greater the advantage in sensitivity to and selectivity for task-relevant orientations at the

fovea than perifovea, the stronger the observed eccentricity effect. This result allows us

to pinpoint orientation sensitivity and selectivity as dominant computations underlying

the behavioral eccentricity effect in this detection task.

To conclude, we found (1) observers are more sensitive to orientations and SFs,

similarly selective for orientations and SFs at the fovea, and shift their peak sensitivity to

lower-than-signal SFs at the perifovea, (2) internal noise is lower at fovea than perifovea,

(3) a link between individual differences in contrast sensitivity with both tuning

characteristics and internal noise. Together, these findings provide compelling evidence

that the foveal advantage in visual tasks can be attributed to two system-level

computations: a better featural representation and lower internal noise than at

parafoveal locations.

References

Abbey, C. K., Eckstein, M. P. (2002). Optimal shifted estimates of human-observer

templates in two-alternative forced-choice experiments. *IEEE Transactions on Medical*

Imaging, *21*(5), 429–440.

Abbey, C. K., & Eckstein, M. P. (2009). Frequency tuning of perceptual templates

changes with noise magnitude. Journal of the Optical Society of America. A, Optics,

Image Science, and Vision, 26(11), B72–B83.

Aghajari, S., Vinke, L. N., Ling, S. (2020). Population spatial frequency tuning in human

early visual cortex. *Journal of Neurophysiology*, *123*(2), 773–785.

Ahumada, A. J. (2002). Classification image weights and internal noise level

estimation. In *Journal of Vision* (Vol. 2, Issue 1, p. 8). https://doi.org/10.1167/2.1.8

Anton-Erxleben, K., Carrasco, M. (2013). Attentional enhancement of spatial

resolution: linking behavioural and neurophysiological evidence. *Nature Reviews.*

Neuroscience, *14*(3), 188–200.

Baldwin, A. S., Meese, T. S., Baker, D. H. (2012). The attenuation surface for

contrast sensitivity has the form of a witch's hat within the central visual field.

Journal of Vision, *12*(11). https://doi.org/10.1167/12.11.23

Barbot, A., Xue, S., Carrasco, M. (2021). Asymmetries in visual acuity around the

visual field. *Journal of Vision*, *21*(1), 2.

Benson, N. C., Kupers, E. R., Barbot, A., Carrasco, M., Winawer, J. (2021). Cortical

magnification in human visual cortex parallels task performance around the visual

field. *eLife, 10*. https://doi.org/10.7554/eLife.67685

Blakemore, C., Campbell, F. W. (1969). On the existence of neurones in the human

visual system selectively sensitive to the orientation and size of retinal images. *The*

Journal of Physiology, *203*(1), 237–260.

Broderick, W. F., Simoncelli, E. P., Winawer, J. (2022). Mapping spatial frequency

preferences across human primary visual cortex. *Journal of Vision*, *22*(4), 3.

Burgess, A. E., Colborne, B. (1988). Visual signal detection. IV. Observer

inconsistency. *Journal of the Optical Society of America. A, Optics and Image*

Science, *5*(4), 617–627.

Cannon, M. W., Jr. (1985). Perceived contrast in the fovea and periphery. *Journal*

of the Optical Society of America. A, Optics and Image Science, *2*(10), 1760–1768.

Carrasco, M., Evert, D. L., Chang, I., Katz, S. M. (1995). The eccentricity effect:

target eccentricity affects performance on conjunction searches. *Perception*

Psychophysics, *57*(8), 1241–1261.

Carrasco, M., Frieder, K. S. (1997). Cortical Magnification Neutralizes the

Eccentricity Effect in Visual Search. In *Vision Research* (Vol. 37, Issue 1, pp. 63–

82). https://doi.org/10.1016/s0042-6989(96)00102-2

Carrasco, M., McLean, T. L., Katz, S. M., Frieder, K. S. (1998). Feature

asymmetries in visual search: effects of display duration, target eccentricity,

orientation and spatial frequency. *Vision Research*, *38*(3), 347–374.

Carrasco, M., Talgar, C. P., Cameron, E. L. (2001). Characterizing visual

performance fields: effects of transient covert attention, spatial frequency,

eccentricity, task and set size. *Spatial Vision*, *15*(1), 61–75.

De Valois, R. L., Albrecht, D. G., Thorell, L. G. (1982a). Spatial frequency

selectivity of cells in macaque visual cortex. *Vision Research*, *22*(5), 545–559.

De Valois, R. L., De Valois, K. K. (1988). *Spatial Vision.* Oxford University Press.

De Valois, R. L., Yund, E. W., Hepler, N. (1982b). The orientation and direction

selectivity of cells in macaque visual cortex. *Vision Research*, *22*(5), 531–544.

Diependaele, K., Brysbaert, M., Neri, P. (2012). How Noisy is Lexical Decision?

Frontiers in Psychology, *3*, 348.

Dosher, B. A., Lu, Z. L. (2000). Mechanisms of perceptual attention in precuing of

location. *Vision Research*, *40*(10-12), 1269–1292.

Duncan, R. O., Boynton, G. M. (2003). Cortical magnification within human primary

visual cortex correlates with acuity thresholds. *Neuron*, *38*(4), 659–671.

Eckstein, M. P., Pham, B. T., Shimozaki, S. S. (2004). The footprints of visual

attention during search with 100% valid and 100% invalid cues. *Vision Research*,

44(12), 1193–1207.

Eckstein, M. P., Shimozaki, S. S., Abbey, C. K. (2002). The footprints of visual

attention in the Posner cueing paradigm revealed by classification images. *Journal*

of Vision, *2*(1), 25–45.

Enroth-Cugell, C., Robson, J. G. (1966). The contrast sensitivity of retinal ganglion

cells of the cat. *The Journal of Physiology*, *187*(3), 517–552.

Falkenberg, H. K., Bex, P. J. (2007). Sources of motion-sensitivity loss in glaucoma.

Investigative Ophthalmology Visual Science, *48*(6), 2913–2921.

Fernández, A., Li, H.-H., Carrasco, M. (2019). How exogenous spatial attention

affects visual representation. *Journal of Vision*, *19*(10), 100b.

Fernández, A., Okun, S., Carrasco, M. (2022). Differential Effects of Endogenous

and Exogenous Attention on Sensory Tuning. *The Journal of Neuroscience: The*

Official Journal of the Society for Neuroscience, *42*(7), 1316–1327.

Foster, K. H., Gaska, J. P., Nagler, M., Pollen, D. A. (1985). Spatial and temporal

frequency selectivity of neurones in visual cortical areas V1 and V2 of the macaque

monkey. In *The Journal of Physiology* (Vol. 365, Issue 1, pp. 331–363).

https://doi.org/10.1113/jphysiol.1985.sp015776

Gold, J. M., Sekuler, A. B., Bennett, P. J. (2004). Characterizing perceptual

learning with external noise. *Cognitive Science*, *28*(2), 167–207.

Hayes, R. D., Merigan, W. H. (2007). Mechanisms of Sensitivity Loss due to Visual

Cortex Lesions in Humans and Macaques. *Cerebral Cortex* , *17*(5), 1117–1128.

Heeley, D. W., Buchanan-Smith, H. M., Cromwell, J. A., Wright, J. S. (1997). The

oblique effect in orientation acuity. *Vision Research*, *37*(2), 235–242.

Henriksson, L., Nurminen, L., Hyvärinen, A., Vanni, S. (2008). Spatial frequency

tuning in human retinotopic visual areas. *Journal of Vision*, *8*(10), 5.1–13.

Hess, R. F., & Hayes, A. (1994). The coding of spatial position by the human visual

system: effects of spatial scale and retinal eccentricity. Vision Research, 34(5),

625–643.

Hilz, R., Cavonius, C. R. (1974). Functional organization of the peripheral retina:

sensitivity to periodic stimuli. *Vision Research*, *14*(12), 1333–1337.

Himmelberg, M. M., Kurzawski, J. W., Benson, N. C., Pelli, D. G., Carrasco, M.,

Winawer, J. (2021). Cross-dataset reproducibility of human retinotopic maps.

NeuroImage, *244*, 118609.

Himmelberg, M. M., Winawer, J., Carrasco, M. (2022). Linking individual

differences in human primary visual cortex to contrast sensitivity around the visual

field. *Nature Communications*, *13*(1), 3309.

Hubel, D. H., Wiesel, T. N. (1962). Receptive fields, binocular interaction and

functional architecture in the cat's visual cortex. *The Journal of Physiology*, *160*,

106–154.

Hubel, D. H., Wiesel, T. N. (1977). Ferrier lecture. Functional architecture of

macaque monkey visual cortex. *Proceedings of the Royal Society of London.*

Series B, Containing Papers of a Biological Character. Royal Society , *198*(1130),

1–59.

Jakobsson, P., Lennerstrand, G. (1985). Binocular interaction in the VEP to grating

stimulation. II. Spatial frequency effects. *Acta Ophthalmologica*, *63*(3), 290–296.

Jigo, M., Tavdy, D., Himmelberg, M., Carrasco, M. (2023). Cortical magnification

eliminates differences in contrast sensitivity across but not around the visual field.

eLife, *12*. https://doi.org/10.7554/eLife.84205

Kiorpes, L., Tang, C., Movshon, J. A. (1999). Factors limiting contrast sensitivity in

experimentally amblyopic macaque monkeys. *Vision Research*, *39*(25), 4152–4160.

Kleiner, M., Brainard, D., Pelli, D. (2007). What's new in Psychtoolbox-3.

Perception, *36*(14), 1–16.

Klein, S. A., Levi, D. M. (2009). Stochastic model for detection of signals in noise.

JOSA A, *26*(11), B110–B126.

Lago, M. A., Abbey, C. K., & Eckstein, M. P. (2021). Foveated Model Observers for

Visual Search in 3D Medical Images. *IEEE Transactions on Medical Imaging,* 40(3),

1021–1031.

Legge, G. E., Kersten, D., Burgess, A. E. (1987). Contrast discrimination in noise.

Journal of the Optical Society of America. A, Optics and Image Science, *4*(2), 391–

404.

Levi, D. M., Klein, S. A. (1986). Sampling in spatial vision. *Nature*, *320*(6060), 360–

362.

Levi, D. M., Klein, S. A. (2002). Classification images for detection and position

discrimination in the fovea and parafovea. *Journal of Vision*, *2*(1), 46–65.

Levi, D. M., & Klein, S. A. (2003). Noise provides some new signals about the

spatial vision of amblyopes. *The Journal of Neuroscience: The Official Journal of*

the Society for Neuroscience, 23(7), 2522–2526.

Levi, D. M., Klein, S. A., Chen, I. (2008). What limits performance in the amblyopic

visual system: seeing signals in noise with an amblyopic brain. *Journal of Vision*,

8(4), 1.1–23.

Li, H.-H., Barbot, A., Carrasco, M. (2016). Saccade Preparation Reshapes Sensory

Tuning. *Current Biology: CB*, *26*(12), 1564–1570.

Ling, S., Liu, T., Carrasco, M. (2009). How spatial and feature-based attention

affect the gain and tuning of population responses. *Vision Research*, *49*(10), 1194–

1204.

Luzardo, F., Yeshurun, Y. (2021). Inter-individual variations in internal noise predict

the effects of spatial attention. *Cognition*, *217*(104888), 104888.

Lu, Z. L., Dosher, B. A. (1998). External noise distinguishes attention mechanisms.

Vision Research, *38*(9), 1183–1198.

Lu, Z. L., Dosher, B. A. (1999). Characterizing human perceptual inefficiencies with

equivalent internal noise. *Journal of the Optical Society of America. A, Optics,*

Image Science, and Vision, *16*(3), 764–778.

Lu, Z. L., Dosher, B. A. (2008). Characterizing observers using external noise and

observer models: assessing internal representations with external noise.

Psychological Review, *115*(1), 44–82.

Lu, Z. L., Dosher, B. A. (2023). Hierarchical Bayesian perceptual template modeling

of mechanisms of spatial attention in central and peripheral cuing. *Journal of Vision*,

23(2), 12.

Mäkelä, P., Näsänen, R., Rovamo, J., Melmoth, D. (2001). Identification of facial

images in peripheral vision. *Vision Research*, *41*(5), 599–610.

Mareschal, I., Bex, P. J., Dakin, S. C. (2008a). Local motion processing limits fine

direction discrimination in the periphery. *Vision Research*, *48*(16), 1719–1725.

Mareschal, I., Morgan, M. J., Solomon, J. A. (2008b). Contextual effects on

decision templates for parafoveal orientation identification. *Vision Research*, *48*(27),

2689–2695.

McAnany, J. J., Park, J. C. (2018). Reduced Contrast Sensitivity is Associated With

Elevated Equivalent Intrinsic Noise in Type 2 Diabetics Who Have Mild or No

Retinopathy. *Investigative Ophthalmology Visual Science*, *59*(6), 2652–2658.

Michel, M., & Geisler, W. S. (2011). Intrinsic position uncertainty explains detection

and localization performance in peripheral vision. *Journal of Vision*, 11(1), 18.

Moraglia, G. (1989). Visual search: spatial frequency and orientation. *Perceptual*

and Motor Skills, *69*(2), 675–689.

Movshon, J. A., Thompson, I. D., Tolhurst, D. J. (1978a). Spatial and temporal

contrast sensitivity of neurones in areas 17 and 18 of the cat's visual cortex. In *The*

Journal of Physiology (Vol. 283, Issue 1, pp. 101–120).

https://doi.org/10.1113/jphysiol.1978.sp012490

Movshon, J. A., Thompson, I. D., Tolhurst, D. J. (1978b). Spatial summation in the

receptive fields of simple cells in the cat's striate cortex. *The Journal of Physiology*,

283, 53–77.

Murray, R. F., Bennett, P. J., Sekuler, A. B. (2002). Optimal methods for calculating

classification images: weighted sums. *Journal of Vision*, *2*(1), 79–104.

Nagai, M., Bennett, P. J., Sekuler, A. B. (2008). Exploration of vertical bias in

perceptual completion of illusory contours: Threshold measures and response

classification. *Journal of Vision*, *8*(7), 25.1–17.

Neri, P. (2010). How inherently noisy is human sensory processing? *Psychonomic*

Bulletin Review, *17*(6), 802–808.

Neri, P., Levi, D. M. (2006). Receptive versus perceptive fields from the reverse-

correlation viewpoint. *Vision Research*, *46*(16), 2465–2474.

Ohl, S., Kuper, C., Rolfs, M. (2017). Selective enhancement of orientation tuning

before saccades. *Journal of Vision*, *17*(13), 2.

Paltoglou, A. E., Neri, P. (2012). Attentional control of sensory tuning in human

visual perception. *Journal of Neurophysiology*, *107*(5), 1260–1274.

Pelli, D. G. (1981). *Effects of visual noise*.

http://citeseerx.ist.psu.edu/viewdoc/download?doi=10.1.1.706.822&rep=rep1&type

=pdf

Pelli, D. G. (1985). Uncertainty explains many aspects of visual contrast detection

and discrimination. *Journal of the Optical Society of America. A, Optics and Image*

Science, *2*(9), 1508–1532.

Pelli, D. G., Blakemore, C. (1990). The quantum efficiency of vision. *Vision: Coding*

and Efficiency, 3–24.

Pentland, A. (1980). Maximum likelihood estimation: the best PEST. *Perception*

Psychophysics, *28*(4), 377–379.

Pointer, J. S., Hess, R. F. (1989). The contrast sensitivity gradient across the

human visual field: with emphasis on the low spatial frequency range. *Vision*

Research, *29*(9), 1133–1151.

Polyak, S. L. (1941). *The retina* (Vol. 607). Univ. Chicago Press The retina.

Ratcliff, R., Voskuilen, C., McKoon, G. (2018). Internal and external sources of

variability in perceptual decision-making. *Psychological Review*, *125*(1), 33–46.

Rijsdijk, J. P., Kroon, J. N., van der Wildt, G. J. (1980). Contrast sensitivity as a

function of position on the retina. *Vision Research*, *20*(3), 235–241.

Robson, J. G., Graham, N. (1981). Probability summation and regional variation in

contrast sensitivity across the visual field. *Vision Research*, *21*(3), 409–418.

Rockel, A. J., Hiorns, R. W., & Powell, T. P. (1980). The basic uniformity in

structure of the neocortex. *Brain: A Journal of Neurology*, 103(2), 221–244.

Rovamo, J., Virsu, V., Näsänen, R. (1978). Cortical magnification factor predicts

the photopic contrast sensitivity of peripheral vision. *Nature*, *271*(5640), 54–56.

Schiller, P. H., Finlay, B. L., Volman, S. F. (1976a). Quantitative studies of single-

cell properties in monkey striate cortex. III. Spatial frequency. *Journal of*

Neurophysiology, *39*(6), 1334–1351.

Schiller, P. H., Finlay, B. L., Volman, S. F. (1976b). Quantitative studies of single-

cell properties in monkey striate cortex. II. Orientation specificity and ocular

dominance. *Journal of Neurophysiology*, *39*(6), 1320–1333.

Shimozaki, S. S., Eckstein, M. P., Abbey, C. K. (2005). Spatial profiles of local and

nonlocal effects upon contrast detection/discrimination from classification images.

Journal of Vision, *5*(1), 45–57.

Song, C., Schwarzkopf, D. S., Kanai, R., Rees, G. (2015). Neural population tuning

links visual cortical anatomy to human visual perception. *Neuron*, *85*(3), 641–656.

Strasburger, H., Rentschler, I., Jüttner, M. (2011). Peripheral vision and pattern

recognition: a review. *Journal of Vision*, *11*(5), 13.

Van Surdam Graham, N. (1989). *Visual Pattern Analyzers*. Oxford University Press.

Vilidaite, G., Baker, D. H. (2017). Individual differences in internal noise are

consistent across two measurement techniques. *Vision Research*, *141*, 30–39.

Virsu, V., Rovamo, J. (1979). Visual resolution, contrast sensitivity, and the cortical

magnification factor. *Experimental Brain Research. Experimentelle Hirnforschung.*

Experimentation Cerebrale, *37*(3), 475–494.

Wardle, S. G., Bex, P. J., Cass, J., Alais, D. (2012). Stereoacuity in the periphery is

limited by internal noise. *Journal of Vision*, *12*(6), 12.

Watkins, D. W., Berkley, M. A. (1974). The orientation selectivity of single neurons

in cat striate cortex. *Experimental Brain Research. Experimentelle Hirnforschung.*

Experimentation Cerebrale, *19*(4), 433–446.

Watson, A. B., & Ahumada, A. J., Jr. (2005). A standard model for foveal detection

of spatial contrast. Journal of Vision, 5(9), 717–740.

Wichmann, F. A., & Hill, N. J. (2001). The psychometric function: I. Fitting, sampling,

and goodness of fit. Perception & Psychophysics, 63(8), 1293–1313.

Wright, M. J., Johnston, A. (1983). Spatiotemporal contrast sensitivity and visual

field locus. *Vision Research*, *23*(10), 983–989.

Wyart, V., Nobre, A. C., Summerfield, C. (2012). Dissociable prior influences of

signal probability and relevance on visual contrast sensitivity. *Proceedings of the*

National Academy of Sciences of the United States of America, *109*(9), 3593–3598.

Xu, X., Anderson, T. J., Casagrande, V. A. (2007). How do functional maps in

primary visual cortex vary with eccentricity? *The Journal of Comparative Neurology*,

501(5), 741–755.

Figure 1. Trial sequence and stimuli. In half of the trials, the target patch (indicated by

the white placeholder) consisted of a horizontal Gabor (SF=2 cpd) embedded in filtered-

white noise; in the other half, the stimulus was just independently generated noise.

Observers performed a yes/no detection task by reporting whether they observed a

Gabor in the target. Auditory feedback was provided after the response was given.

Figure 2. *Reverse correlation. (**A**) Step 1. We computed energy profiles by dot-*

*multiplying a pool of Gabor filters in quadrature phase with the target patch. (**B**) Step 2.*

We regressed the trial-wise energy fluctuations with behavioral responses and used the

*regression slope (β_1) to index perceptual sensitivity. (**C**) Step 3. The derived β_1 of all*

orientation-SF channels in Step 2 yielded a 2-D kernel mapping. Then, we took the

*mean (marginals) of the 2-D mapping to derive orientation and SF kernels. (**D**) Step 4.*

We fitted tuning functions to the kernels and characterized the tuning function by

extracting tuning parameters (i.e., gain) and characteristics (i.e., peak amplitude,

bandwidth and baseline for orientation and SF, and peak SF, see Table 1).

For both orientation and SF		
Name	Computation	Interpretation
Gain (α)	The scaling factor estimated by fitting a tuning function to the kernels	Overall sensitivity to the feature
Peak amplitude (A)	The height of the peak of the tuning function	Sensitivity to the task-relevant feature
Bandwidth (B)	Orientation: full width of function at half-height SF: difference in log transform of SFs at half-height	Selectivity for the task-relevant feature
Baseline (b)	The height of the rightmost kernel	The baseline sensitivity
For SF only		
Peak SF (f_{peak})	The SF corresponding to the peak of the tuning function	The preferred SF of the system

Table 1. A list of tuning characteristics extracted from the orientation and SF tuning

functions.

Figure 3. _Correlations among estimated parameters and tuning characteristics derived_

from orientation (left panel) and spatial frequency (right panel) tuning functions across

observers. The white boxes indicate estimated parameters and the black boxes indicate

tuning characteristics. The symbols at the connecting lines indicate the strength of inter-

parameter correlations. The pair being compared is indicated by the color of the box.

(Note that the estimated baseline is equivalent to the extracted baseline). ***p < 0.001,

_**p < 0.01, *p < 0.05, n.s., p > 0.1._

Figure 4. *Schematic of the noisy observer model.* **Step 1a**. *A representation of*

orientations and SFs, which is assumed to be the template adopted by the observer to

complete the detection task at one location, is derived from the training set (75% of

trials per location). **Step 1b**. *Energy profiles of the test set (25% trials) are calculated.*

Step 2. *The internal variable for each trial is calculated by dot multiplying the template*

and the energy profile of each trial, separately for Gabor-present and -absent trials.

Step 3. *For each trial, a 'noisy internal variable' was created by adding two sources of*

internal noise to the internal variable. **Step 4**. *A decision is made by comparing the*

noisy internal variable to a threshold.

Model expressions		Model fitting	
$T_{\theta,f}$	Template estimated by applying reverse correlation to the *training* set	n_{Pc}	The total number of correct trials
$E_{i,\theta,f}$	Energy profile of the i-th trial in the test set	m_{Pc}	The total number of trials
V_i	The internal variable of the i-th trial	Pc	The predicted proportion of correctness
μ_{S+}, μ_{S-}	The mean of the internal variables for the Gabor-present and -absent trials, respectively	n_{Pa}	The total number of consistent trial pairs
σ_{S+}^2, σ_{S-}^2	The variance of the internal variables for the Gabor-present and -absent trials, respectively	m_{Pa}	The total number of trial pairs
$V_{S+,noisy}$, $V_{S-,noisy}$	The theoretical expression of noisy internal variables for the Gabor-present and -absent trials, respectively	Pa	The predicted response consistency
Free parameters			
N_{IE}	The ratio of the SD of induced noise to the external variability (i.e., σ_{S+}^2 and σ_{S-}^2)	$\sigma_{constant}^2$	The variance of the constant noise
		k	The threshold

Table 2. Parameters in the noisy observer model.

Figure 5. *(A) d-prime, (B) criterion, (C) contrast sensitivity (reciprocal of contrast*

threshold) and (D) response consistency at the fovea and perifovea. Bars indicate group

averages and error bars are ±1 SEM. Gray markers plotted between two bars represent

individual data (the median of 1K bootstrapping) and each observer is indicated by a

unique marker type. Error bars representing ±1 SEM. The error bars intersecting the

black horizontal line indicate ±1 SEM of the difference. ***p < 0.001, n.s., p > 0.1.

Figure 6. 2-D sensitivity kernel mapping and separability derived from Gabor-absent

trials. Pixels in A-C represent group-averaged sensitivity kernels derived from data at

the fovea (**A**) and perifovea (**B**), indicating the regression slope representing the

correlation between the feature energy and the behavioral responses. The mapping's

horizontal and vertical red lines depict the Gabor SF and orientation, respectively. (**C**)

Difference between fovea and perifovea. The red border indicates the orientation-SF

channel at which the fovea is higher than the perifovea (p < 0.05, Bonferroni corrected

one-way t-test). (**D**) The ideal template which is the energy profile of the signal derived

from step 1 of **Fig 2**. (**E**) Separability (i.e., the correlation between reconstructed

sensitivity kernels and original kernels). A correlation close to 1 implies high separability.

Figure 7. *Marginalized kernels and the sensory tuning function for orientation (**A**) and*

*SF (**B**) with model comparison (**C-D**). (**A and B**) Red (fovea) and blue (perifovea) dots*

are group-averaged marginalized kernels and error bars represent ±1 SEM. Curves with

corresponding colors are group-averaged sensory tuning functions and bands represent

±1 SEM. Asterisks indicate channels at which the fovea is higher than the perifovea

(Bonferroni-corrected post-hoc one-tailed t-test, p < 0.05). Black dots below asterisks

are the averaged difference between fovea and perifovea (indicated by the y-axis on the

*right) and bands in black are ±1 SEM. (**C and D**) Comparison among candidate models*

of the winning family (gray bars, scaled Gaussian for orientation and log parabola for

SF), and with the best-fitting model of other families (white bar with gray outline, DoG

for orientation and truncated log parabola for SF). Deviance is the difference between

data and prediction using 10-fold cross-validation averaged across 10 folds. Empty

*squares: free to vary; filled-in squares: shared. (**E and F**) The frequency of each model*

being the best-fitting model when fitted to individual data.

Figure 8. *Comparison of orientation (**A-D**) and SF (**E-H**) tuning characteristics between*

fovea (red dots) and perifovea (blue dots). Dots represent group averages with ±1 SEM.

Gray dots indicate individual data. ***p < 0.001, marginal, p > 0.05 and p < 0.1, n.s., p >

0.1.

Figure 9. *Evaluate noisy observer model. (**A-B**) Comparison of the predicted accuracy*

*(**A**) and response consistency (**B**) with data. (**C-D**) Correlation between the response*

*consistency with the induced noise (**C**) and constant noise (**D**). (**E-F**) Comparison of the*

*induced noise (**E**) and constant noise (**F**) between fovea and perifovea. Gray markers*

*indicate individual data. (**G and H**) compare the predictive power (ΔBIC) of four model*

variations: the Core mode (white bar), the Random template model (black bar), the

Random energy model (dark gray bar) and the Random template and energy model

*(light gray bar) for fovea (**G**) and perifovea (**H**). Bars represent group averages and*

*error bars are ±1 SEM. ***p < 0.001.*p < 0.05, n.s., p > 0.1.*

Figure 10. Correlation between contrast sensitivity and orientation tuning characteristics (**A-D**), SF tuning bandwidth (**E and F**), and estimated internal noise (**G and H**). Values in the figure were zero-meaned to reveal the observer effect. Markers are the median of individual data at the fovea (red) and perifovea (blue). Each individual data is represented by a unique marker. The black diagonal line is the linear regression line.

Figure 11. *Correlation between the eccentricity effect index (EEI) of contrast sensitivity*

*and (**A**) orientation gain (peak amplitude yielded the same result), (**B**) orientation*

*bandwidth, (**C**) induced noise, and (**D**) constant noise. A positive EEI indicates fovea*

being higher than perifovea and vice versa. Greater positive/negative EEI indicates a

larger fovea-perifovea difference. Markers are the median of EEI for each individual

observer. The diagonal black line is the linear regression line. **p < 0.01, *p < 0.05, n.s.,

p > 0.1