

LSD microdosing attenuates the impact of temporal priors in time perception.

Abbreviated title: **LSD-attenuated impact of priors in time perception**

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Abstract

Recent theoretical work embedded within the predictive processing framework has proposed that the neurocognitive and therapeutic effects of psychedelics are driven by the modulation of priors (Carhart-Harris & Friston, 2019). We conducted pre-registered re-analyses of previous research (Yanakieva et al., 2019) to examine whether microdoses of lysergic acid diethylamide (LSD) alleviate the temporal reproduction bias introduced by priors, as predicted by this theoretical framework. In a between-groups design, participants were randomly assigned to one of four groups receiving LSD (5, 10, or 20 μg) or placebo (0 μg) and completed a visual temporal reproduction task spanning subsecond to suprasecond intervals (0.8 to 4 sec). Using mixed-effects modelling, we evaluated the impact of the treatment group, and of the overall history of stimulus intervals (*global* priors) and the local stimulus history (*local* priors), weighted by their respective precision weights (inverse of variance), on temporal reproduction. Our principal finding was that the precision-weighted local priors and their precision weights reduced the under-reproduction bias observed under LSD in the original research. Furthermore, controlling for the precision-weighted local prior eliminated the reduced temporal reproduction bias under LSD, indicating that LSD microdosing mitigated the temporal under-reproduction by reducing the relative weighting of priors. These results suggest that LSD microdosing alters human time perception by decreasing the influence of local temporal priors.

Keywords: temporal prior; time perception; temporal reproduction bias; LSD; microdosing;

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Introduction

53 Over the past decade, there has been a renewed interest in the neurocognitive
54 effects of psychedelics, with recent clinical and basic research demonstrating this
55 trend (Calder & Hasler, 2022; Carhart-Harris & Goodwin, 2017; Johnston et al.,
56 2023; Lewis-Healey et al., 2022; Nichols, 2016). One influential model explaining the
57 effects of psychedelics is the “Relaxed Beliefs Under Psychedelics” (REBUS) model,
58 (Carhart-Harris & Friston, 2019), which proposes that these substances relax the
59 weights on expectations based on prior experiences (priors). Our perception is
60 influenced by our past sensory experiences, and according to the REBUS model, the
61 relaxation of these priors under psychedelics can lead to higher thresholds for
62 susceptibility to illusions and less perceptual bias (Carhart-Harris & Friston, 2019).
63 Here we investigated whether the bias in temporal reproduction may be moderated
64 by LSD’s influence on temporal priors.

65 Information about the external environment processed via sensory channels often
66 varies in quality and is combined with prior predictions for an optimal outcome to
67 guide decision making (Friston, 2009; Körding & Wolpert, 2006; Raviv et al., 2012).
68 When priors and incoming information misalign, prediction errors arise, and the
69 perceptual system strives to reduce the uncertainty. Precise priors are afforded
70 higher weighting and override prediction errors but they are attenuated and updated
71 when higher precision weights are afforded to incoming signals (Clark, 2013).
72 Psychedelics are thought to mitigate the influence of priors, thus yielding increased
73 confidence in bottom-up information, with potential therapeutic benefits for
74 psychiatric disorders hypothesized to be characterized by pathologically over-
75 weighted priors (Cassidy et al., 2018; Powers et al., 2016; Teufel et al., 2015).
76 Recent research has supported this hypothesis, showing that psychedelics can

77 reduce the prior precision weighting, leading to a more flexible and exploratory
78 process of perception and cognition (Leptourgos et al., 2022; Muthukumaraswamy et
79 al., 2013; Rajpal et al., 2022).

80 In addition to their effect on priors, psychedelics have been shown to modulate
81 time perception (Coull, Morgan, et al., 2011; Kenna & Sedman, 1964; Wittmann et
82 al., 2007; Yanakieva et al., 2019). For instance, Yanakieva et al. (2019) found that
83 LSD reduced under-reproduction bias in their participants' perception of stimulus
84 durations compared to the placebo group (Fig. 2 in Yanakieva et al., 2019).
85 However, whether the under-reproduction could be attributed to the influence of
86 temporal priors that was mitigated in the LSD group was not assessed. Accumulating
87 research has shown that interval timing performance can be modelled as a form of
88 Bayesian inference (Karaminis et al., 2016; Sadibolova & Terhune, 2022; Shi et al.,
89 2013; Shi & Burr, 2016), thereby offering the possibility of scrutinizing the predictions
90 of REBUS model through the analysis of the impact of psychedelics on temporal
91 priors. In particular, Bayesian models of time perception posit that the brain forms
92 temporal priors by extracting statistical patterns from the environment such as the
93 mean duration of stimulus intervals (*global prior*, Acerbi et al., 2012; Jazayeri &
94 Shadlen, 2010) and that temporal duration estimates gravitate towards the stimulus
95 duration in previous trials (*local prior*, de Jong et al., 2021; Wiener et al., 2014).
96 These priors are combined with sensory evidence regarding the duration of a
97 stimulus and can introduce temporal biases (Acerbi et al., 2012; Cicchini et al., 2012)
98 that may be mitigated by administration of psychedelics.

99 In summary, by integrating the REBUS model into the predictive coding
100 framework, it is possible to gain a more comprehensive understanding of the effects
101 of psychedelics on time perception. Insofar as the pattern of findings by Yanakieva et

102 al. (2019) conforms to the predictions of the REBUS model of the diminished impact
103 of priors under psychedelics, we re-analyzed their data to investigate the possibility
104 that their observations may be accounted for by global and local temporal priors
105 interacting with the drug treatment. We predicted that the reduced impact of priors
106 under psychedelics would remedy the temporal reproduction bias (i.e., the tendency
107 for reproduced intervals to shift toward priors and away from objective interval
108 durations).

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Materials and Methods

112

Participants

114 This study involves re-analyses of previous data (Yanakieva et al., 2019); our
115 analyses were preregistered on the Open Science Framework (<https://osf.io/hfkjr>).
116 The participants were 48 English-speaking adults, aged 55-75 (Mean age = 62.92,
117 21 female [44%], 27 male [56%]). Participants were randomly assigned to one of four
118 groups ($n=12$) that received either LSD (5, 10, or 20 μg) or a placebo (0 μg). The
119 removal of 2 multivariate outliers (one each from the 5 and 20 μg dose groups [see
120 Yanakieva et al., 2019]) reduced the final sample size to 46.

121

Experimental design

123 The data collection took place in an inpatient unit, adhering to standardized pre-
124 screening and medical protocols (Yanakieva et al., 2019) as part of a larger pre-
125 clinical trial on the safety and efficacy of microdose LSD (Family et al., 2020). The
126 study used a randomized, double-blind, placebo-controlled design. The LSD solution

127 was prepared in distilled water at a pharmacy on-site, and placebo groups were
128 administered only distilled water, rendering the LSD and placebos indistinguishable
129 to researchers and participants. The manufacturer of the drug product was Onyx
130 Scientific Limited UK, to cGMP standards (Yanakieva et al., 2020).

131 Participants completed a temporal reproduction task once post-dose on the 4th
132 day of dosing with the specific time of completion varying across participants
133 (Yanakieva et al., 2019). Each trial of the task consisted of a fixed 750 ms cue
134 “*memorize*”, a blank jittered interval (425–650 ms), and a target stimulus interval
135 (blue circle [80 × 80 pixels; ~2 cm in diameter], on a 1280 × 800 pixel-monitor) of
136 varying duration (800, 1200, 1600, 2000, 2400, 2800, 3200, 3600, or 4000 ms). The
137 stimulus was followed by a fixed 500 ms blank interval and a response cue
138 “*reproduce*”, at which point participants responded by holding down the space bar to
139 reproduce the stimulus interval. A blue circle co-appeared with this response and
140 remained on the monitor until the spacebar was released. Trials were separated by a
141 fixed 500 ms blank interstimulus interval. Participants completed one practice block,
142 followed by four experimental blocks of 27 trials (108 trials total).

143

144 **Statistical analysis**

145 We first removed bivariate outlier responses across stimulus intervals for each
146 participant ($M=8.37\%$, $SD=2.95\%$, range: 1.85-14.81%) identified with the median
147 absolute deviation (MAD) method implemented in the robust correlation toolbox
148 (Pernet et al., 2013) in MATLAB (MathWorks, Natick, USA).

149 Data were analyzed with mixed effects modelling using the *lme4* package (Bates
150 et al., 2015) in R (R Core Team, 2021). Our experimental design is characterized by
151 a hierarchical structure (Figure S1 in supplemental materials) with *participants*

152 (categorical; 46 levels) nested in the *dose* (continuous) and the dose nested in
153 *cohorts* (categorical; 4 levels). As per Yanakieva et al. (2019), we created another
154 dichotomous variable *drug* comprised of placebo and LSD levels which was included
155 in our models instead of the *dose*. The models included additional continuous
156 variables *stimulus interval* and *prior*. We further distinguished between different
157 types of priors. The global prior was calculated at each trial (starting from trial 3) as
158 the arithmetic mean of all preceding intervals. There were three local priors: the last
159 preceding stimulus interval (*n-1 prior*), and the arithmetic means of the last preceding
160 two intervals (*n-2 prior*) and three intervals (*n-3 prior*). The continuous variables
161 (reproduced intervals, stimulus intervals, and priors) were mean-centered and the
162 categorical variables (*drug* and *cohort*) were dummy coded.

163 Mixed-effects models were fitted using restricted maximum likelihood (REML)
164 and the nonlinear “*nlopt_in_neldermead*” optimizer (Wächter & Biegler, 2006) with
165 tighter tolerance values ($1e^{-12}$). Model fit improvement was determined by a change
166 in the log-likelihood (increases with goodness of fit), and log-likelihood derived
167 Bayesian Information Criterion (BIC; Schwarz, 1978) and Akaike Information criterion
168 (AIC; Akaike, 1974) (both decrease with superior goodness of fit). We refrained from
169 using the χ^2 -distributed log-likelihood ratio test for model comparison as it’s been
170 reported to produce anti-conservative *p*-values (Pinheiro & Bates, 2000). Instead, we
171 applied Kenward and Roger’s approximation of degrees of freedom for *F*-test *p*-
172 values and we computed the *p*-values for fixed effects parameters with Kenward and
173 Roger’s method, as implemented in the “*pbkrtest*” and “*afex*” R packages (Halekoh &
174 Højsgaard, 2014; Kuznetsova et al., 2017; Singmann et al., 2018). Furthermore, we
175 employed the type III sums of squares method, which is suitable for unbalanced
176 designs (Keppel & Wickens, 2004), due to sample size differences between the drug

177 and placebo groups. To generate reliable estimates of uncertainty, we used the
178 “*bootMer*” function to compute the bootstrapped 95% confidence intervals with 100
179 iterations. To further investigate the interaction effects, we conducted additional post
180 hoc tests such as t-tests.

181 In order to determine the random-effects structure for the mixed-effects model,
182 we began by generating null ('empty') models (Quené & van den Bergh, 2004)
183 comprising distinct random-effects structures (Barr et al., 2013), as outlined in
184 Section S2 of the Supplementary Materials, which also includes the model
185 diagnostics. The final random-effects structure of the model with the lowest BIC and
186 AIC included correlated by-subject within-dose within-cohort intercepts and slopes
187 for stimulus intervals. It was applied in all subsequent mixed-effects model analyses
188 with fixed-effects parameters (stimulus intervals, drug, priors, and interactions
189 thereof). Different priors were evaluated in separate models.

190

191 ***Unregistered prior precision-weights***

192 In our registered analyses, we conceptualized priors as stimulus history, but did
193 not take into account the crucial role of their precision weighting before their
194 integration with sensory evidence (likelihood) (Petzschner et al., 2015; Sadibolova &
195 Terhune, 2022; Shi et al., 2013; Shi & Burr, 2016). To rectify this omission, we
196 replicated our registered analyses using prior precision weights, with higher values
197 reflecting lower uncertainty for prior evidence, and precision-weighted priors. We
198 calculated the precision weights as the inverse of variance for the prior normalized
199 by the sum of inverse variances of the prior and likelihood (Petzschner et al., 2015).
200 The initial prior precision at each trial was the inverse of the variance of all preceding
201 trials (*global* priors), the preceding two trials ($n-2$ priors), and the preceding three

202 trials ($n-3$ priors). We computed the likelihood variance for each trial by measuring
203 the variability in reproduced intervals after regressing out the influence of priors,
204 following the method introduced by Aston and colleagues (2021). For global priors,
205 we regressed the preceding responses on the stimulus intervals and divided the
206 variance of the residuals by the squared slope to remove the central tendency bias
207 (Aston et al., 2021) whereas for the local priors, we regressed the reproduced
208 intervals on each prior itself. As with the prior precision estimates, the initial
209 likelihood precision was computed as the inverse of the likelihood variance.

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212 Results

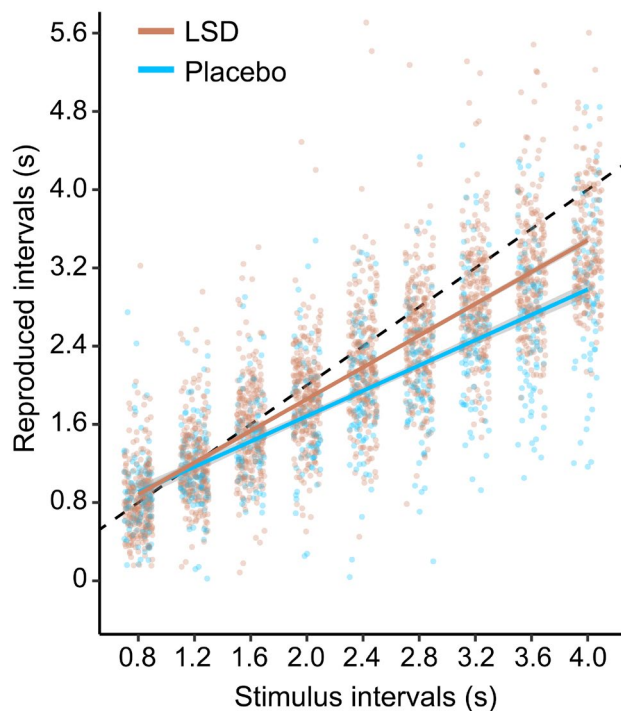
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214 LSD attenuates performance bias in temporal reproduction

215 Our first expectation was for the mixed effects model to replicate the difference in
216 reproduced supra-second intervals between the placebo and LSD treatment groups
217 (Yanakieva et al., 2019). Our analyses replicated the reduced under-reproduction of
218 stimulus intervals longer than 2 s (Yanakieva et al., 2019) in the LSD relative to the
219 placebo condition (Figure 1). In particular, the null model including the by-subject
220 within-dose within-cohort intercepts and slopes for intervals (AIC=4020, BIC=4050)
221 improved with the inclusion of Stimulus interval as a fixed-effects parameter
222 (*Baseline model*), $F(1,44.99) = 829.49$, $p < .001$, $AIC=3885$, $BIC=3922$. The fit was
223 further improved with the inclusion of Drug and Stimulus interval*Drug fixed-effects
224 predictors (*Replication model*), $F(2,42.66) = 4.91$, $p = .012$, $AIC=3880$, $BIC=3929$
225 (Table 1). The Stimulus interval, $\beta = .64$ (95% CI [.54, .73]), $SE = .05$, $t = 13.44$, $p < .001$,
226 Drug, $\beta = .24$ (95% CI [.08, .39]), $SE = .10$, $t = 2.37$, $p = .022$, and Stimulus interval x Drug

227 interaction, $\beta=.17$ (95% CI [.07, .27]), $SE=.06$, $t=3.04$, $p=.004$, predictors were all
228 statistically significant. These results suggest a .64s increase in reproduced intervals
229 for each 1s stimulus increment irrespective of the drug condition, reflecting a
230 shallower slope than would be expected with veridical performance. However,
231 relative to the placebo condition, the reproduced intervals under LSD were longer by
232 .24s on average and they rose by .17s per 1s actual increment across stimulus
233 intervals. Cumulatively, these results demonstrate that LSD induced a decrease in
234 the temporal under-reproduction bias, resulting in a shift towards a more accurate
235 and veridical performance.

236



237

238 *Figure 1.* Temporal reproduction of stimulus intervals in LSD and placebo groups.

239 The black dashed line represents veridical performance. Markers represent
240 individual reproduced intervals. Shaded error bars are standard error smoothed with
241 the linear *geom_smooth* function in R.

242

243

244 **The influence of priors on temporal reproduction**

245 In the next subsections (also see Tables **1** and **S2**), we describe how the
246 inclusion of global and local priors improved the model already including the stimulus
247 and drug predictors. Motivated by the proposal that psychedelics modulate cognition
248 and perception by attenuating prior precision weighting (Carhart-Harris & Friston,
249 2019), we predicted that temporal priors would predict reproduced intervals
250 (*Baseline model* improvement) and that this effect would differ between LSD and
251 placebo conditions (*Replication model* improvement).

252

253 Table 1
 254 Comparison of baseline and replication mixed-effects models of temporal
 255 reproduction performance in placebo and LSD conditions.

	Baseline model		Replication model		AIC	BIC	LL
	improvement		improvement				
	<i>F</i>	χ^2	<i>F</i>	χ^2			
Baseline model					3885.9	3922.1	-1937.0
Replication model					3880.4	3928.8	-1932.2
Global prior unweighted	3.96	23.63	3.54	14.16	3874.3	3946.8	-1925.1
Global prior precision-weighted	5.38	30.93	5.64	21.46	3867.0	3939.5	-1921.5
Global prior precision weights	4.54	25.65	4.32	16.18	3872.3	3944.7	-1924.1
n-1 prior unweighted	12.46	74.63	16.44	65.17	3823.3	3895.8	-1899.6
n-2 prior unweighted	10.13	60.73	12.90	51.26	3837.2	3909.7	-1906.6
n-2 prior precision-weighted	4.99	29.38	4.97	19.91	3868.5	3941.0	-1922.3
n-2 prior precision-weights	4.50	26.58	4.28	17.11	3871.3	3943.8	-1923.7
n-3 prior unweighted	9.03	54.17	11.23	44.70	3843.7	3916.2	-1909.9
n-3 prior precision-weighted	3.93	22.58	3.28	13.11	3875.3	3947.8	-1925.7
n-3 prior precision-weights	4.89	28.94	4.90	19.47	3869.0	3941.4	-1922.5

Notes. The baseline model included the intercept and Stimulus interval as fixed-effects parameters. The replication model additionally included Drug and Stimulus interval*Drug interaction as fixed-effects parameters. Rows represent the mixed effects model with a prior parameter added to the two different models. All *F* statistics are significant, $p \leq .01$. AIC = Akaike information criterion, BIC = Bayesian information criterion, LL = Log-likelihood. Bolded values reflect models exhibiting superior fit to both the baseline and replication model. The color gradient for each, the AIC, BIC and LL values indicates the model fit quality (yellow = worse, green = better) by comparing all models in the respective table column.

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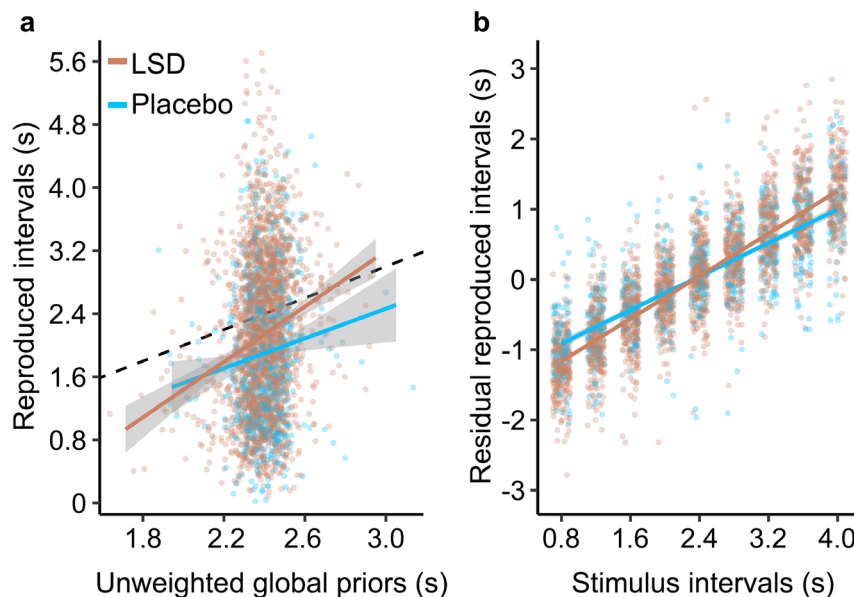
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258 ***Global priors (unweighted)***

259 We first considered the role of global priors calculated as the mean of previous
260 stimulus intervals at each trial. The global prior distribution is conventionally centered
261 at the mean interval range, and research has demonstrated its association with
262 central tendency bias in time perception, resulting in an overestimation of short
263 intervals and underestimation of long intervals (Acerbi et al., 2012; Cicchini et al.,
264 2012). The impact of including global priors in the *Baseline model* and *Replication*
265 *model* is inconclusive (Table 1), as the AIC and Kenward-Roger's test p-values
266 indicate improvement whereas the BIC values greater by more than 10 points
267 suggest deterioration (Burnham & Anderson, 2004). Further results (see Table S1)
268 did not support a differential role of global priors in temporal reproduction of long
269 intervals under LSD (3-way interaction), $\beta=.01$ (95% CI [-.46, .39]), $SE=.21$, $t=.03$,
270 $p=.98$. These observations seemingly contradict our prediction that the influence of
271 global prior on central tendency bias is reduced in LSD condition. Indeed, rather than
272 diminishing the influence of priors, LSD appears to have produced an increased
273 reliance on them that generalized across stimulus intervals. In the placebo condition,
274 the reproduced intervals follow the pattern of the global prior (albeit with overall
275 under-reproduction), whereas in the LSD condition, the slope is steeper (Figure 2a),
276 Drug * Global priors, $\beta=.50$ (95% CI [.02, .94]), $SE=.22$, $t=2.22$, $p=.030$.

277 To test whether this effect contributed to reduced under-reproduction of longer
278 stimulus intervals under LSD (Figure 1; Yanakieva et al., 2019), we fitted a model
279 including Stimulus interval, Drug, and Stimulus interval * Drug fixed-effects to
280 responses after removing trends for global priors for each participant (Figure 2b).
281 Although the resulting residual reproduced intervals continued to show a steeper
282 increase with increasing stimulus intervals under LSD, $\beta=.16$ (95% CI [.063, .272]),

283 SE=.05, $t=3.02$, $p=.004$, the difference in under-reproduction bias across drug
284 groups was no longer significant for three out of five previously significant long
285 intervals (2.4, 2.8 and 4s, Holm-Bonferroni corrected $p>.05$). The mean reproduced
286 intervals did not differ across group, $\beta= -.01$ (95% CI [-.039, .034]), SE=.02, $t=-.25$,
287 $p=.80$. These analyses demonstrate that global priors partly shaped the pattern of
288 the original observations (Yanakieva et al., 2019) although not via their reduced
289 impact under psychedelics (Carhart-Harris & Friston, 2019).
290



291
292 **Figure 2.** The effects of global priors on temporal reproduction performance in
293 placebo and LSD conditions. **a.** Reproduced intervals (individual trial responses) as
294 a function of unweighted global priors in each drug condition. The black dashed line
295 represents the responses identical to global prior values. **b.** Residual reproduced
296 intervals after removing the linear trend for global prior for each participant (see
297 Figure 1 for raw reproduced intervals). Both panels: Markers are individual
298 datapoints (trials). Shaded error bars are standard error smoothed with the linear
299 *geom_smooth* function in R.

300

301

302 ***Local priors (unweighted)***

303 Our next analyses examined the roles of local priors in the impact of LSD on
304 temporal reproduction. Incorporating local priors improved the model fit of both the
305 *Baseline model* and *Replication model* (Table 1). Reproduced intervals increased
306 with longer preceding ($n-1$) stimuli, $\beta=.07$ (95% CI [.04, .10]), $SE=.01$, $t=4.63$,
307 $p<.001$, and this effect varied across Stimulus intervals, Prior * Stimulus interval
308 interaction: $\beta=.03$ (95% CI [-.01, .06]), $SE=.01$, $t=1.99$, $p=.04$. However, as with
309 global priors, we did not find significant evidence for a Prior x Stimulus interval x
310 Drug interaction, $\beta= -.01$ (95% CI [-.05, .02]), $SE=.02$, $t=.52$, $p=.60$, and we did not
311 observe an interaction between the Prior x Drug group, $\beta= -.01$ (95% CI [-.05, .02]),
312 $SE=.02$, $t=.80$, $p=.42$. Analyses using unweighted $n-2$ and $n-3$ priors exhibited similar
313 patterns of results (see Supplementary Table S1). To summarize, our observations
314 align with previous research (de Jong et al., 2021; Wiener et al., 2014) and suggest
315 that reproduced intervals are reliably influenced by local priors not weighted by their
316 precision, and that LSD does not moderate these effects.

317

318 ***Precision-weighted priors and precision weights***

319 The foregoing approach can be expanded upon by considering the level of
320 (un)certainty or precision of prior evidence, which determines how influential the
321 priors are in shaping the responses (Petzschner et al., 2015; Sadibolova & Terhune,
322 2022). Therefore, we repeated the analyses but substituted priors with prior precision
323 weights accounting for how impactful the priors are, and with priors weighted by their
324 precision. We performed these analyses with all predictors *except* the $n-1$ priors, due
325 to the inability to calculate their precision on trial-by-trial basis.

326 The impact of including precision weights and precision-weighted *global* priors in
327 the *Baseline model* and *Replication model* remains inconclusive (Table 1), as the
328 AIC and Kenward-Roger's test p-values continue to indicate improvement whereas
329 the BIC values greater by more than 10 points suggest deterioration (Burnham &
330 Anderson, 2004). Further results (Supplementary Table S1) indicate that the
331 variation in reproduced time intervals between drug groups attributed to unweighted
332 global priors (Figure 2) lost statistical significance after accounting for global prior
333 precision weighting. Altogether, these results do not furnish compelling
334 substantiation for the global prior accounting for Yanakieva et al.'s (2019) findings.

335 Reproduced intervals decreased with increasing n-2 prior weights, $\beta = -.15$ (95%
336 CI [-.26, -.05]), $SE = .07$, $t = -2.19$, $p = .03$, and n-3 prior weights, $\beta = -.36$ (95% CI [-.55, -
337 .17]), $SE = .09$, $t = -3.89$, $p < .01$, indicating a greater local history bias for higher
338 precision weights, as would be expected. The reduction in reproduced intervals
339 driven by the prior precision weights occurred at a slower pace in the LSD condition
340 compared to the placebo condition, $\beta = .19$ (95% CI [.07, .33]), $SE = .08$, $t = 2.38$, $p = .02$
341 (n-2 precision weights * Drug), and $\beta = .34$ (95% CI [.11, .55]), $SE = .11$, $t = 3.07$, $p < .01$
342 (n-3 precision weights * Drug). For n-2 precision weights, this effect was more
343 pronounced for long stimulus intervals (n-2 precision weights * Drug * Stimulus),
344 $\beta = .19$ (95% CI [.03, .35]), $SE = .08$, $t = 2.50$, $p = .01$. Accordingly, the decrease in
345 reproduced intervals (under-reproduction bias) as a function of n-2 priors adjusted by
346 these precision weights was reduced for longer stimulus intervals in the LSD relative
347 to the placebo condition, $\beta = .07$ (95% CI [.01, .13]), $SE = .03$, $t = 2.30$, $p = .02$ (Table S1).
348 In order to determine whether this effect accounts for the observation of reduced
349 under-reproduction of long stimulus intervals in the LSD condition (Fig 2; Yanakieva
350 et al., 2019), we fitted the model with Stimulus interval, Drug, and Stimulus interval *

351 Drug fixed-effects to residual reproduced intervals after removing the linear trend for
352 precision-weighted n-2 priors for each participant. Our results suggest that the LSD-
353 mediated reduced bias in the original study by Yanakieva and colleagues (2019) was
354 indeed no longer significant after accounting for the influence of the precision-
355 weighted n-2 priors, $\beta=-.01$ (95% CI [-.05, .03]), $SE=.06$, $t=-0.29$, $p=.77$, and the
356 across-condition differences for individual long stimulus intervals were not
357 statistically significant (Holm-Bonferroni corrected $ps>.05$). Taken together, our
358 results suggest that reduced precision-weighting of local temporal priors underlies
359 altered time perception in LSD (Yanakieva et al., 2019).

360

361

362

Discussion

363 In this re-analysis of existing data (Yanakieva et al., 2019), we examined whether
364 LSD alters time perception by modulating the impact of temporal priors on temporal
365 reproduction. We predicted that the REBUS model's theorized reduced impact of
366 priors under psychedelics (Carhart-Harris & Friston, 2019) would remedy the
367 temporal reproduction bias (i.e. the tendency for reproduced intervals to shift toward
368 priors and away from objective interval durations). We found that the impact of *global*
369 priors (unweighted by their precision) on temporal reproduction was more
370 pronounced under LSD, contrary to the REBUS model predictions. However, the
371 difference in reproduced intervals across drug groups was not fully eliminated when
372 the influence of the global priors was controlled for, suggesting that they only partly
373 explained group variation in Yanakieva et al.'s data. Moreover, the impact of the
374 global priors on temporal reproduction was similar in both drug groups once they
375 were weighted by their precision. By comparison, *local* prior precision and precision-

376 weighted priors were associated with tempered under-reproduction of longer
377 stimulus durations in the LSD group. Reproduced intervals decreased with
378 increasing local prior precision weights, indicating a greater local history bias for
379 higher prior precision, and the reduction in reproduced intervals driven by precision
380 weights occurred at a slower pace in the LSD group compared to the placebo group.
381 Further analyses showed that these effects accounted for the original observation
382 (Yanakieva et al., 2019), given that the reproduced long intervals did not significantly
383 differ across drug groups once the impact of the precision-weighted local prior was
384 controlled for. These results suggest that altered temporal reproduction under LSD
385 (Yanakieva et al., 2019) may be explained by local temporal priors in line with the
386 proposal that psychedelics reduce the impact of priors (Carhart-Harris & Friston,
387 2019; Safron et al., 2020).

388 The REBUS model (Carhart-Harris & Friston, 2019) supports these findings by
389 suggesting that psychedelics decrease the confidence of priors and reduce their
390 constraining effect on the processing of incoming information (prediction errors).
391 According to this model, the relaxation of priors is most evident at high levels of the
392 processing hierarchy, such as those associated with the default-mode network,
393 which are linked to self-hood, identity, and ego (Carhart-Harris et al., 2012).
394 However, the model also proposes that a wide range of functional levels will be
395 impacted, including priors at intermediate levels of the processing hierarchy, albeit
396 potentially with less conspicuous psychological effects (Carhart-Harris & Friston,
397 2019). Accordingly, our findings suggest that temporal reproduction under LSD is
398 less influenced by temporal priors and therefore exhibits less bias under LSD. Our
399 observations broadly align with previous research on the LSD-mediated reduced
400 influence of priors on perception. For instance, LSD was found to reduce the brain's

401 ability to detect and respond to unexpected or deviant auditory stimuli, as indicated
402 by the diminished amplitude of the mismatch negativity component in response to
403 auditory stimuli (Timmermann et al., 2018). This was attributed to LSD reducing the
404 precision of the brain's internal models of expected input. Additionally, the reduced
405 Kanizsa illusion, which relies on top-down predictions from higher visual areas,
406 provides further evidence of the impact of psychedelics on priors (Kometer et al.,
407 2011). Altogether, these findings provide valuable insights into the effects of
408 psychedelics on sensory processing and prior predictions in shaping perception.

409 It has been repeatedly demonstrated that LSD and germane psychedelics induce
410 pronounced alterations in perception of time, yet the neurocognitive and
411 neurochemical mechanisms underlying these effects have not been fully understood
412 (Altman, 1977; Aronson, 1959; Kenna & Sedman, 1964; Passie et al., 2008;
413 Wittmann et al., 2007). Recent predictive processing theories, such as the REBUS
414 model, offer a new perspective, proposing that psychedelics exert their influence on
415 cognition by increasing the excitability of deep-layer pyramidal neurons that express
416 5-HT receptors (Carhart-Harris & Friston, 2019; Nichols, 2016; Safron et al., 2020).
417 According to the model, the overly excitable neurons fail to synchronize, thereby
418 reducing the influence of top-down priors and increasing the likelihood of the system
419 being updated by unsuppressed ascending prediction errors. In this way, the REBUS
420 model suggests that psychedelics afford a greater latitude for belief updating
421 throughout the processing hierarchy (Carhart-Harris & Friston, 2019; Nichols, 2016;
422 Safron et al., 2020) although the evidence for this neurocognitive mechanism and its
423 neurochemical instantiation in temporal perception is currently underspecified.
424 Future research will benefit from investigating the neural mechanisms underpinning

425 our observations to shed further light into how they align with these theoretical
426 accounts.

427 Notably, the effects of LSD have been linked not only to its psychedelic
428 properties via the activation of serotonin 5-HT receptors, but also to its
429 psychostimulant properties via dopamine D1 and D2 receptors and dissociative
430 properties via NMDA glutamate transmission (Marona-Lewicka & Nichols, 2007;
431 Nichols, 2004; Passie et al., 2008). The dopamine system, in particular, has been
432 widely implicated in altered temporal perception (for reviews, see Agostino & Cheng,
433 2016; Coull, Cheng, et al., 2011; Marinho et al., 2018). Interestingly, whereas the
434 serotonin 5-HT agonism has been discussed in the context of psychedelic relaxation
435 (down-weighting) of priors for its therapeutic effects in psychopharmacological
436 models of psychotic hallucinations (Corlett et al., 2019; Haarsma et al., 2021; Rajpal
437 et al., 2022), there is evidence for elevated striatal dopamine being associated with
438 an over-reliance on priors (Cassidy et al., 2018). Furthermore, the LSD has been
439 shown to decrease dopamine firing activity through 5-HT, D2 and TAAR1 receptors
440 (De Gregorio et al., 2016; Marona-Lewicka & Nichols, 2007). Accordingly, an
441 alternative interpretation of the present results is that an LSD-mediated reduction in
442 striatal dopamine levels yielded the observed attenuation of precision-weighting of
443 local temporal priors underlying the observed altered timing performance (Yanakieva
444 et al., 2019). Further work is required to discriminate between these competing
445 neurochemical interpretations. Additionally, the relationship between dopamine and
446 serotonin systems in the effects of LSD on temporal perception warrants further
447 investigation, as both systems may have opposing effects on temporal priors
448 (Carhart-Harris & Friston, 2019; Cassidy et al., 2018).

449 Alternatively, our findings could be interpreted as resulting from increased
450 information flow due to reduced thalamic gating under the cortico-striatal
451 thalamocortical (CSTC) model (Preller et al., 2019). This model proposes that LSD,
452 acting as a serotonergic agonist, diminishes the striatal influence on the thalamus,
453 thus opening the thalamic filter. This interpretation is consistent with predictive
454 processing accounts, which frame thalamic gating and increased "bottom-up"
455 information flow to cortical areas in terms of the heightened precision of ascending
456 prediction errors (likelihood) (Clark, 2013, 2016). This process reflects the gain on
457 incoming information at the expense of priors, which in the case of temporal priors
458 and their associated biases, leads to a reduction in reproduction biases under LSD,
459 as observed in this study. Recent accounts of the CSTC model have further
460 expanded its scope by including the claustrum, which is densely populated by 5-HT
461 receptors and has been implicated in the temporal integration of cortical and
462 thalamic oscillations for interval timing and working memory (Doss et al., 2022; Yin et
463 al., 2016). Thus, future research would additionally benefit from investigating the
464 roles of the thalamus and claustrum in mediating LSD's effects on temporal priors in
465 temporal perception.

466 Moving from the discussion of neurocognitive and neurochemical bases of the
467 observed effects, it is worth exploring the significance of LSD micro-dosing. A typical
468 full dose of LSD ranges from 75-150 µg and produces a diverse array of
469 phenomenological effects such as hallucinations, ego dissolution and altered
470 perception of space and time (Passie et al., 2008). By comparison, the
471 phenomenological effects of LSD micro-dosing are minimal (Yanakieva et al., 2019).
472 Therefore, unlike explicit psychoactive effects of LSD (e.g., mystical experiences)
473 and putative therapeutic effects (e.g., antidepressant effects), which are plausibly

474 shaped in part by participants' expectancies (Burke & Blumberger, 2021; Butler et
475 al., 2022; Olson et al., 2020), it is unlikely that participants expect to exhibit reduced
476 temporal under-reproduction under LSD. This strongly suggests that the observed
477 results are not attributable to explicit expectations, as has been suggested for other
478 effects of microdose psychedelics (Kaertner et al., 2021). Future research could
479 expand upon the present work by more stringently manipulating temporal priors
480 under different LSD doses and investigating the role of specific neurochemical
481 systems.

482 To summarize, our study has shed light on the intricate ways in which LSD
483 impacts temporal reproduction by modulating local and global temporal priors.
484 Specifically, we have demonstrated that even small doses of LSD can interact with
485 local temporal priors, resulting in alterations in temporal reproduction. Our findings
486 are underpinned by existing theoretical models based on the predictive coding
487 framework, which offer a potential explanation for the observed effects. The results
488 of this study pave the way for future research to further explore these underlying
489 mechanisms. Overall, our findings align with the REBUS model of the effects of
490 psychedelics on cognition (Carhart-Harris & Friston, 2019), suggesting that low
491 doses of psychedelics have the potential to alleviate temporal biases imposed by
492 temporal priors. These exciting results represent an important step in advancing our
493 understanding of the effects of psychedelics on temporal perception.

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740 *Figure 1.* Temporal reproduction of stimulus intervals in LSD and placebo groups.
741 The black dashed line represents veridical performance. Markers represent
742 individual reproduced intervals. Shaded error bars are standard error smoothed with
743 the linear *geom_smooth* function in R.

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745 *Figure 2.* The effects of global priors on temporal reproduction performance in
746 placebo and LSD conditions. **a.** Reproduced intervals (individual trial responses) as
747 a function of unweighted global priors in each drug condition. The black dashed line
748 represents the responses identical to global prior values. **b.** Residual reproduced
749 intervals after removing the linear trend for global prior for each participant (see
750 **Figure 1** for raw reproduced intervals). Both panels: Markers are individual
751 datapoints (trials). Shaded error bars are standard error smoothed with the linear
752 *geom_smooth* function in R.

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