The successor representation subserves hierarchical abstraction for goal-directed behavior

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ABSTRACT

Humans have the ability to craft abstract, temporally extended and hierarchically organized plans. For instance, when considering how to make spaghetti for dinner, we typically concern ourselves with useful "subgoals" in the task, such as cutting onions, boiling pasta, and cooking a sauce, rather than particulars such as how many cuts to make to the onion, or exactly which muscles to contract. A core question is how such decomposition of a more abstract task into logical subtasks happens in the first place. Previous research has shown that humans are sensitive to a form of higher-order statistical learning named "community structure". Community structure is a common feature of abstract tasks characterized by a logical ordering of subtasks. This structure can be captured by a model where humans learn predictions of upcoming events multiple steps into the future, discounting predictions of events further away in time. One such model is the "successor representation", which has been argued to be useful for hierarchical abstraction. As of yet, no study has convincingly shown that this hierarchical abstraction can be put to use for goal-directed behavior. Here, we investigate whether participants utilize learned community structure to craft hierarchically informed action plans for goal-directed behavior. Participants were asked to search for paintings in a virtual museum, where the paintings were grouped together in "wings" representing community structure in the museum. We find that participants' choices accord with the hierarchical structure of the museum and that their response times are best predicted by a successor representation. The degree to which the response times reflect the community structure of the museum correlates with several measures of performance, including the ability to craft temporally abstract action plans. These results suggest that successor representation learning subserves hierarchical abstractions relevant for goal-directed behavior.

AUTHOR SUMMARY

Humans have the ability to achieve a diverse range of goals in a highly complex world. Classic theories of decision making focus on simple tasks involving single goals. In the current study, we test a recent theoretical proposal that aims to address the flexibility of human decision making. By learning to predict the upcoming events, humans can acquire a 'model' of the world which they can then leverage to plan their behavior. However, given the complexity of the world, planning directly over all possible events can be overwhelming. We show that, by leveraging this predictive model, humans group similar events together into simpler "hierarchical" representations, which makes planning over these hierarchical representations markedly more efficient. Interestingly, humans seem to learn and remember both the complex predictive model and the simplified hierarchical model, using them for distinct purposes.

INTRODUCTION

Classically, research on decision making in humans and animals has focused on comparing choices with differently valued outcomes occurring in the immediate future (1,2). This problem can be formalized in the framework of reinforcement learning (RL; 3) and has led to a rich understanding of the relationship between learning and dopaminergic responses in the midbrain (4). However, naturalistic decision making often involves long-range dependencies between multiple choices and temporally distal outcomes (5). Recently, the interest in biological mechanisms subserving such temporally extended decision making has surged (5–8). Temporally extended behaviors are problematic for standard RL models, which are highly sensitive to the combinatorial explosion of possible states and actions available in real world sequential decisions (9,10). One approach to alleviate this combinatorial explosion is to generalize knowledge over states or actions, which is called hierarchical reinforcement learning (6,11). This involves the learning of effective hierarchical representations that allow for meaningful generalization, but the neurocomputational mechanisms underlying this remain to be determined (12).

Previous accounts of hierarchical representation learning have mainly relied on graph-theoretic analysis of the structure of the task (13–16). However, this requires a fully accurate representation of this task structure to begin with, which in reality often has to be learned during the task itself (17). The successor representation (SR) (18) is a reinforcement learning algorithm specifically concerned with learning task structure based on a biologically plausible learning rule (19–21). Notably, it has been suggested that the successor representation can be leveraged to discover hierarchical representations (22,23). The successor representation learns a multi-step prediction of the task structure, which allows for fast and flexible behavior in the face of changing goals (24). Representing the task structure in this way balances fast computation commonly associated with so-called model-free reinforcement learning against the flexibility commonly associated with so-called model-based algorithms which require more computational time to adjust behavior flexibly (25). Interestingly, it has been shown that the behaviors of both humans and rats exhibit some of the limitations in decision making specifically associated with the successor representation (26,27), which can capture some human and rat learning effects better than typical model-free or model-based RL algorithms do (28). Additionally, the successor representation can explain how people segment events based on the graph topology of the events, capturing higher-order structure in the environment (29–31). However, to our knowledge no one has yet investigated whether the successor representation is leveraged for the purpose of extracting hierarchies relevant for goal-directed behavior.

In the current study, we investigate whether the successor representation provides a basis for hierarchical abstraction that is actually used for planning goal-directed behavior. We designed a task that afforded a simplified choice policy based on community relationships observable in the lower-order task structure. Participants played a computer game wherein they acted as a tour guide in a museum, leading visitors to rooms with specific paintings that they requested to view. A new request was made every time a painting was located, which encouraged the participants to learn an internal model of the museum in order to find the paintings as efficiently as possible. Unbeknownst to the participants, the museum consisted of three different wings, with each wing containing five highly interconnected rooms. The participants made binary choices to navigate between the rooms of the museum, where each choice was probabilistically mapped to two of the four possible neighboring rooms. These action-outcome mappings afforded a simple policy to move between wings, as consistently selecting one key yielded on average shorter paths to one of the two remaining wings, whereas the other key yielded shorter paths to the other wing.

This task design allowed us to test two key predictions. First, we predicted that participants would
learn an internal model of upcoming rooms while they navigated through the museum (23). We
predicted that aspects of this internal model would be reflected in response times, with more
expected rooms yielding faster response times (29), and expectations and prediction errors about
reward also influencing response times (32–36). Second, we predicted that participants would
extract and utilize hierarchical information about the museum, which would be reflected in their
choice behaviors. In particular, we expected that their choices would be sensitive to the higher-order
structure of the museum wings rather than only to the lower-order structure of the museum rooms,
which would enable faster planning using less cognitive effort (6,37).

Following a series of preregistered analyses, we tested the nature of the learned internal model
separately for the response time and the choice data. We compared several candidate models: An
optimal "model-based" model that perfectly represented the lower-order structure of the rooms of
the museum, an "explicit hierarchical" model that perfectly represented the higher-order structure
of the wings of the museum, and a "successor representation" model that learned multistep
predictive representations of the rooms sensitive to higher-order community structure of the wings.
We found that the successor representation model accounted for the response times the best,
whereas the explicit hierarchical model accounted for the choice behavior the best. In an exploratory
follow-up analysis, we examined the "discount factor" parameter of the successor representation,
which arbitrates the degree to which the community structure is reflected in the learned predictions.
We observed that individual differences in the fit of the discount factor correlated with 1) effective
hierarchical abstraction as reflected in choice behavior, 2) accumulated reward compensation at the
task, and 3) performance on a secondary task where the participants were asked to reconstruct the
rooms of the museum from memory. In all cases, discount factors that reflected the higher-order
structure of the wings more strongly predicted better performance, including a more accurate
reconstruction of the lower-order structure of the rooms of the museum in the secondary task. These
results support the hypothesis that the successor representation subserves effective hierarchical
abstraction for goal-directed behavior.

RESULTS

Experimental design

We initially ran a pilot study and preregistered a sample and analysis plan for the current study
(https://osf.io/n2jcz/). Based on the preregistration, we collected a total number of 141 participants
on the online platform Prolific (38). We excluded 21 participants based on preregistered exclusion
criteria that ensured that the participants sufficiently attended to the task (see methods). The
participants played an online game in which they would explore and navigate a simulated museum
consisting of multiple rooms. Because some of the rooms were clustered into local groups or "wings"
of the museum, the subjects could learn a useful hierarchical representation of the task. The
participants were told to imagine that they were a tour guide in this museum who was required to
guide visitors to specific paintings that they requested to see. In total there were 15 rooms in the
museum, each containing a unique painting. Because the requested paintings occurred in different
rooms, the participants needed to plan their route in order to reach the paintings in as few steps as
possible. The 15 rooms were organized according to a "ring-of-cliques" layout as shown in Fig. 1A,
which was originally used by (30). In this graph, each room is connected to four other rooms via four
different hallways. This layout of the museum was never shown to the participants, nor did they
receive any other hints about the layout. During an initial "training phase" (see Methods), the
participants were required to learn the layout of the museum by engaging with the task. Then, during

a subsequent "testing phase", which constitutes the main focus of our analyses, the participants were
required to navigate the museum based on this learned knowledge.

Fig 1 Task design. (**A**) A graph showing the relationships between the rooms in the museum task. The museum is
characterized by three different communities or "wings", each consisting of 5 rooms. Individual rooms are numbered 1-
15. (**B**) "Balanced" action-outcome mappings for each room. Each set of outcomes is randomly assigned to keys <z> and
<m> (labelled "i" and "ii"). The same key mapped to the same set of outcomes across the three different wings. Hence, if
the participant intended to move out of the current wing into a specific different wing, they needed only to select
repeatedly the key leading to the desired wing (labeled as "intention"). Green arrows indicate transitions to the intended
wing, orange arrows indicate transitions ending in nodes with green arrows, and red arrows indicate transitions ending
in nodes without any green arrows. Note that the probabilistic mappings ensure that, even when the subjects select a
response appropriate to their intended direction, the actual transition obtained might or might not be consistent with
their intention (see S1 Appendix for more details). (**C**) Illustration of "preference" in action-outcome contingencies (see
Methods). Blue arrows indicate transitions that are present for both actions (as shown in B). Red arrows indicate

transitions that are not possible for either action. One room does not have any preferred or removed transitions. This
room is labeled with "W" for "wide". (**D**) Illustration of an example miniblock. The first image is the goal cue (i). In this
example, the participant begins the miniblock in a room with a cactus painting (center), and the goal cue indicates that
they must look for the "lamp" painting (top). The participant initiates the miniblock by pressing <space> (ii) (This "start
state", which the participants could anticipate based on the goal-cue (i) step, was excluded from the analyses).
Subsequently, the participants were required to select either <z> or <m> to move between rooms (iii, iv) and <space> to
indicate that they reached the goal (v), after which they received a reward (points later translated into a financial bonus)
(vi). As illustrated, pressing <space> in a non-goal room results in a small penalty ("wrong goal") (vii). Similarly, pressing
<z> or <m> in the goal room also leads to a small penalty ("goal miss") (iix). The miniblock always ends with either a "goal
miss" (iix) or a "reward" (vi) screen, after which the next goal cue (i) is presented.

The experimental task was divided into "miniblocks". At the beginning of each miniblock, the
participant was cued with an image of a specific painting that they would be required to find in that
miniblock. The participant could navigate from room to room in the museum by pressing either the
<z> or the <m> key of their keyboard in order to select the direction that they wished to move next.
However, each key mapped to two possible rooms with equal probability (Fig. 1B). The mappings
were engineered to provide the participant agency at the level of the wings of the museum: Each key
consistently moved the participant out of the current wing in only one direction, allowing them to
choose a direction to "rotate" through the wings of the museum. On reaching the goal painting for
that miniblock, the participant was required to press the <space> key. They were told that if they did
so correctly, then they would receive a small reward that would accumulate throughout the
experiment as a bonus payment. However, if they failed to press <space> in the room with the goal
painting, or if they pressed <space> in a room that was not the current goal, then they would instead
incur a small financial punishment. These requirements ensured that the participants would actively
remember the identity of the goal painting for that miniblock, in order to successfully find it and
receive the reward. The miniblock ended once the room with the goal painting was reached,
irrespective of whether the subject pressed <space> or not, and the next miniblock started with a
new goal cue. The starting location for the subsequent miniblock was always the goal location of the
last miniblock, so the entire experiment consisted of a single continuous walk through the museum.
The new goal painting was always a room located in a wing different from the current starting
location. Fig. 1D provides an overview of the timeline of a single miniblock.

Cognitive models
We fit multiple cognitive computational models (see below) to the response times and the choices,
which we predicted would be sensitive to lower-order task structure and simplified hierarchical
abstraction, respectively. We preregistered the exact structure of the models, including the definition
of all theoretically meaningful and all "nuisance" regressors (https://osf.io/n2jcz/; with exceptions
marked and justified in S2 Appendix). We simulated task performance for each of the proposed
models to confirm that they make different predictions about behavior, and conducted a simulation
study to validate the results for both our model comparison and parameter estimation techniques
(see S1 Appendix). These simulations lend confidence that our main results reported here are
meaningfully interpretable (39).

Null Model
In order to provide a benchmark for the other cognitive models, we included "null models" that
capture behavioral predictions of no theoretical interest. For choice behavior, we estimated
response-coded logistic regressions where key <z> is coded as "0" and key <m> is coded as "1". The
null model included only an intercept term and no regressors, accounting for a bias to select either
<z> or <m> preferentially. For the response times, we estimated a linear model for the mean
parameter of a log-normal distribution. The null model included six nuisance regressors and an

intercept term. These nuisance regressors were also included in all of the other response time
models. Firstly, we hypothesized that participants might get frustrated on longer miniblocks, causing
them to speed up their response times. Hence, a nuisance regressor was included that captured the
(log-transformed) number of "steps" for the current miniblock, indicating the number of button
presses and transitions that were made that miniblock. Secondly, we hypothesized that participants
might better remember rooms that were more recently encountered, causing them to speed up.
Hence, a nuisance regressor was included that captured the (log-transformed) number of trials
previously that the current room was last visited. Thirdly, we hypothesized that participants might
be slower to execute responses on different keys compared to repetitions of the same key. Hence, a
binary nuisance regressor was included that indicated whether the current key press was different
from the previous key press.

The remaining three nuisance regressors were binary, each capturing an irregularity in the action-
outcome contingencies. The reasoning behind these irregularities is further elaborated in the
Methods section (section "action-outcome mapping") and in S1 Appendix. Here, we discuss how these
irregularities might influence response times, and how the three nuisance regressors aim to capture
this influence. Firstly, four transitions per community were possible regardless of which choice the
participant made (Fig. 1C, blue, "preferred"). Following a particular transition will elicit a prediction
error, and larger prediction errors are hypothesized to yield slower responses for choices made in
the subsequent state. Note that this is the basis for our prediction that participants should slow down
following transitions between two different communities. Consider a participant who would behave
completely randomly in the Museum task. In each room, they have a 50% chance of selecting either
action, which yields a 50% chance of following either of two transitions as shown in Fig. 1B. A
preferred transition could be elicited by both actions with a 50% chance, and would thus be
experienced more often. Therefore, it is likely these preferred transitions yield lower prediction
errors (i.e. they are predicted more strongly), yielding faster responses for the choices immediately
following the transition. To capture this, we added a binary regressor that was 1 following a preferred
transition, and 0 otherwise.

Secondly, some transitions were not possible in our design (Fig. 1C, red). Therefore, rooms with an
outgoing red arrow as shown in Fig. 1C have only three unique possible outcomes in total (each action
has two outcomes, but one outcome is shared between the two actions as a preferred transition).
Notice that there is one room that does not have any removed or preferred transitions (labeled with
"W" for "wide" in Fig. 1C). Because this is the only room that has 4 possible unique outcomes,
participants might find it harder to decide on the most appropriate action in this room (as they have
to consider more possible outcomes) predicting slower responses in this room. For this reason, we
included a binary regressor that was 1 when participants occupied this wide room, and 0 otherwise.

Thirdly, transitions out of this wide room might be harder for participants to predict, given that each
action has two unique outcomes as opposed to one unique outcome (with the other outcome shared
between the two actions as a preferred transition in all other rooms). Outcomes that are harder to
predict may yield larger prediction errors, and thus predict increased response times for the choices
that immediately follow these transitions. For this reason, we included a binary regressor that was 1
following transitions out of this wide room, and 0 otherwise.

Model-based
Among the cognitive models, we first considered a canonical model-based reinforcement learning
agent. This agent was equipped with perfect knowledge of the action-outcome contingencies and
could compute the exact (temporally discounted) values of each room with respect to the current

goal. Parameter fitting included the discount factor, which is known to influence community
structure (29). In particular, moderately high values for the discount factor emphasize differences
between the wings, increasing value inside the goal wing relative to values outside the goal wing. We
hypothesized that participants might base their choices on the difference in expected values (EVs)
between choices <m> over <z>, making the choice with the relatively higher value (corresponding to
shorter paths to the goal) more likely to be selected. Hence, to model choice behavior, the model-
based model included the relative difference of EV for <m> over <z> as a regressor. To model
response times, the model-based model included three regressors besides the six nuisance
regressors. Firstly, we hypothesized that participants might slow their responding when they are
closer to the goal, in order to not accidentally trigger a "goal miss" (Fig. 1D). Hence, the model-based
response time model included a regressor for the expected value of the current step. Secondly, we
hypothesized participants might slow their responses for unexpected progress toward or away from
the goal. Hence, the model-based response time model included a regressor for reward prediction
error (RPE), defined as the difference between the EV of the current step and the EV of the previous
step. Thirdly, we hypothesized participants might slow their responses when the value difference
between the two choices is less salient. Hence, the model-based response time model included a
regressor capturing the (inverse) "conflict", which was evaluated as the absolute difference between
the EV of the two available choices.

Explicit hierarchical structure
Secondly, we considered a cognitive model that tested graph-theoretic predictions of hierarchical
abstraction. This agent accessed a perfect representation of the mappings between the lower-level
rooms and the higher-level wings, but not of the exact locations of the rooms within the wings
(whereas the model-based agent represented the lower-level room locations, without any explicit
indication of the mappings with respect to the wings). We hypothesized that the participants might
repeatedly select the choice that directly "rotated" towards the current goal wing. Hence, the explicit
hierarchical model of choice behavior included a binary regressor indicating whether or not <m> was
the choice rotating toward the goal wing. This rotation corresponds to the hierarchically informed
policy where only one key allows for a transition to one of the two other wings. To model response
times, the explicit hierarchical model included three regressors besides the six nuisance regressors.
Firstly, we hypothesized that participants might slow their responding when they are closer to the
goal. Hence, the explicit hierarchical response time model included a binary regressor indicating
whether or not the current room was in the same wing as the goal. This is analogous to the EV
regressor in the model-based model. Secondly, we hypothesized that participants might slow their
responding following a transition between two wings compared to transition within one wing. This
effect has been observed in previous experiments and has been defined as a signature of higher-order
predictive learning (29,40,41). Hence, the explicit hierarchical response time model included a binary
regressor indicating whether or not the previous transition mapped between two wings. Thirdly, we
hypothesized that the participants might slow their responding in rooms that allow for a transition
between two wings, because here one choice will allow for the transition between wings whereas the
other choice excludes this transition. Participants might deliberately seek out or avoid these
transitions and thus take more time to consider an optimal choice. Hence, the explicit hierarchical
response time model included a binary regressor indicating whether or not the current room allowed
for a transition between two wings. This is analogous to the measure of (inverse) conflict defined in
the model-based model.

 Successor representation
Finally, we considered an agent that learned a successor representation, which encodes for each state
the expected future (discounted) states that the agent will occupy. The discount factor parameter
determines the temporal horizon of the expected future state occupations. Moderately high discount
factors emphasize the higher-order wing structure of the museum, whereas lower discount factors
emphasize the lower-order room locations. Importantly, the successor representation differs from
optimal model-based decision making in two key respects. Firstly, the successor representation holds
a multi-step predictive model of future state occupations, while the model-based model holds a one-
step predictive model. This explains why the successor representation can predict slowing between
communities. It also implies generalization of value across different rooms in the same wing. Optimal
model-based decision making requires more intensive computation to deal with changing goals
(18,25) and would predict perfect adherence to the values of individual rooms in each wing. Secondly,
the successor representations' predictive model of states is dependent on the policy being followed.
Because different goal locations imply different optimal choices in our museum task, the internal
model of the task states will be in constant flux across successive task miniblocks, even after a
training phase in which the participants learned a good baseline model of the environment (42). This
differs from the model-based model, which is provided a priori with an accurate model of the
environment that can be used to compute an optimal policy for any possible goal location. After each
transition, the successor representation yields a "state prediction error" that corresponds to whether
the newly observed state was anticipated or not. This prediction error is used to update the
predictive model according to a temporal-difference learning rule (see Methods). The predictive
representation can be used to compute expected values analogous to those computed by the model-
based model. To model choice behavior, the successor representation model included a similar
regressor as for the model-based model, reflecting the relative difference in expected values between
the two choices. To model response times, the successor representation model included similar
regressors reflecting the EV, the RPE, and the conflict. We additionally hypothesized participants
might slow their responding when encountering more surprising rooms. Hence, in contrast to the
model-based model, the successor representation model included a regressor for state prediction
error. Dependent on the discount factor, state predictions are expected to be larger for transitions
between two wings compared to transitions within one wing. Therefore, this state prediction error
regressor is analogous to the binary regressor for transitions between wings in the explicit
hierarchical model.

Choice behavior

All analyses with respect to choice behavior and response switching were preregistered
(https://osf.io/n2jcz/) and thus of confirmatory nature. Random-effects group-level Bayesian model
selection of the computational models of choice behavior indicate a significant group difference in
model prevalence ($BOR < 10^{-26}$) with a strong preference for the explicit graph-theoretic hierarchical
abstraction model ($pxp = 1$). This holds when considering binary model comparisons between the
explicit model and every other model (all $BOR < 10^{-5}$, see Methods for multiple comparison
correction. All $pxp = 1$). Posterior model probabilities for every participant are shown in Fig. 2A.

We tested the group-level significance of the relevant rotation-choice parameter of the explicit
hierarchical model by fitting a hierarchical Bayesian logistic regression with separate participant-
level effects (Fig. 2B). We find that participants are indeed significantly more likely to choose the
action that leads into the rotational direction of the goal wing ($M = 1.172$, $HDI_{95\%} = [0.992, 1.350]$, ER_+
$= \infty$). Fig. 2C shows the difference in proportion of <m> choices when <m> is the correct rotational
direction versus when <z> is the correct rotational direction. A clear effect can be seen in the data

(blue) that is also well captured by the explicit graph-theoretical abstraction model (red). It can also
be seen that the effect of rotational direction appears smaller for <z> than for <m>. Indeed, the
intercept in the hierarchical Bayesian logistic regression is significantly above 0 (M = 0.311, $HDI_{95\%}$
= [0.220, 0.401], ER_+ = ∞) indicating a bias in selecting <m>. This can be explained by the fact that
<m> is placed on the right side of the QWERTY-keyboard and participants might have a bias to
interpret rightward directions as progressive (i.e., reading in most western cultures happens from
left to right), or the fact that the majority of participants are likely to be right-handed (which we did
not assess).

Response switches
Given that participants preferentially selected the key corresponding to the rotational direction of
the goal wing, we examined whether participants also dynamically updated their responses during
task execution. We tested this by asking whether they were more likely to switch their response
immediately after moving out of the goal wing, indicating an intention to reverse their direction of
movement (i.e., go "backwards"). Similarly, we asked whether participants were less likely to switch
their response immediately after moving into the goal wing, indicating an intention to continue in the
same direction. We modeled this with a hierarchical Bayesian logistic regression, the coefficients of
which are shown in Fig. 2D. The model confirms significant effects for goal wing entry indicating
decreased switching (M = -0.722, $HDI_{95\%}$ = [-0.940, -0.509], ER_- = ∞), goal wing leave indicating
increased switching (M = 0.983, $HDI_{95\%}$ = [0.581, 1.349], ER_+ = ∞), and a nuisance regressor for action
repetitions (M = -0.724, $HDI_{95\%}$ = [-0.916, -0.532], ER_- = ∞), which indicated that participant switched
less often when they had already repeated an action multiple times (i.e., action commitment). Fig. 2E
illustrates the proportion of response switches relative to "baseline" for all trials with an entry into
the goal wing ("entry"), vs all trials with an exit from the goal wing ("leave"). The baseline trials
include all trials that are not classified as entry or leave. The empirical results (blue) indicated that
participants were less likely to switch when they entered the goal wing, and were more likely to
switch immediately after leaving the goal wing. Posterior predictions from the hierarchical Bayesian
logistic regression model match these data well (red).

Fig 2 Choice data. (**A**) Horizontal stacked bar charts for every participant, illustrating the posterior model probability derived from random effects Bayesian model comparison between four cognitive models, "Null" (orange), explicit hierarchical "Exp" (pink), model-based "MB" (yellow), and successor representation "SR" (blue). Participants were sorted by posterior model probability for the "Exp" model for interpretability. (**B**) Regression coefficient of the "rotation" variable in the explicit hierarchical logistic regression. Density plot shows the full posterior distribution of the population mean, with orange lines indicating the 95% highest density interval. Dots represent posterior means of individual participant (random) effects. (**C**) Proportion of <m> choices over the course of the testing phase, grouped for every participant by trials where <m> was the optimal rotation vs where <z> was the optimal rotation. Empirical data are shown in blue. Posterior predictions from the explicit hierarchical logistic regression are shown in orange. (**D**) Regression coefficients for the response switch analysis, describing the probability of switching response key upon "entry" of the goal wing, "leave" of the goal wing, and the number of "repetitions" of pressing the current key. Colors similar to B. (**E**) Proportion of trials on which the participant switched their response key, grouped by trials where the participant just entered the goal wing or just left the goal wing. Switch proportion of all other ("baseline") trials is subtracted, so values below 0 reflect a decreased tendency

to switch, and values above 0 an increased tendency. Color coding as in C. (**F**) Given a particular goal outside the current
wing, one action corresponds to the correct "rotation" to follow for most of the rooms in the current wing. However, one
room in each community actually leads to better outcomes when the participants would follow the opposite action
(antirotation, here colored in orange). In this figure, possible outcomes of the rotational action are colored blue, and
possible outcomes of the antirotational action are colored orange. As can be observed, the orange room has a preferred
transition in the direction of the goal wing (colored with an orange to blue gradient, indicating both choices can lead to this
outcome). Interestingly, when following the rotation action (blue transitions), one possible transition leads in the correct
direction whereas the other possible transition leads to a room for which the transition in the direction of the goal wing
was removed (X in the figure). By contrast, when following the antirotation (orange transitions), the other possible outcome
leads to a state that still has a transition in the direction of the goal wing (blue circle). For this reason, model-based and
(converged) successor representation models predict participants would pick the antirotational action in this orange room,
and the rotational action in the blue room. (**G**) Proportion of rotational actions chosen in the orange (antirotation) and blue
(rotation) rooms ("Node type", x-axis, colored as in F), considered separately when the correct rotation would be <z> (left
panel) or <m> (right panel). Black dots with error bars correspond to mean with 95% confidence interval. (**H**) Regression
coefficients for the (anti)rotation selection analysis, describing the probability of selecting the correct rotation when it is
<m> (over <z>, left) and when occupying the blue room (as opposed to the orange room, see F; middle). The interaction
term is displayed on the right. Colors similar to B.

Participants mix abstraction with detail

For each wing, given a goal in each other wing, there was always one room where the optimal choice
according to a model-based (and converged successor representation) model would deviate from the
explicit hierarchical model. This is because both the rotation and the antirotation can lead to the same
desired outcome, but the rotation has a chance of transitioning to a room without connections
towards the goal wing, whereas the antirotation does not (Fig. 2F). We set up an exploratory
hierarchical Bayesian logistic regression to test whether participants are indeed more likely to follow
the rotation when it is the optimal decision (blue, Fig. 2F) compared to when it is not (orange, Fig.
2F). Fig. 2G shows that in both of these rooms, regardless of which direction the rotation follows (<z>
or <m>), participants select the correct rotation on most of the trials, but they do so less often when
the antirotation offers a better alternative. A test of the intercept of the logistic regression confirms
that participants are indeed overall biased towards selecting the rotation, as predicted by the explicit
hierarchical model ($M = 0.580$, $HDI_{95\%} = [0.489, 0.671]$, $ER_+ = \infty$). In accordance with our earlier
observations (Fig. 2C), participants are more likely to do so when the correct rotation is <m>, since
they have a bias towards selecting this response ($M = 0.621$, $HDI_{95\%} = [0.435, 0.809]$, $ER_+ = \infty$; Fig. 2H,
"rotation"). Crucially, when participants occupy the room where the rotation is the optimal direction
(blue, Fig. 2F), they are more likely to select the rotation, compared to when the antirotation is in fact
optimal (orange, Fig. 2F) ($M = 0.097$, $HDI_{95\%} = [0.048, 0.145]$, $ER_+ = 9999.000$; Fig. 2H, "node type").
This effect is numerically, although not significantly, attenuated when the correct rotation is <m> (M
$= -0.057$, $HDI_{95\%} = [-0.156, 0.040]$, $ER_- = 7.061$; Fig. 2H, "interaction"). Overall, this confirms that
participants' choice behavior is mostly explained by an explicit hierarchical model in agreement with
our model comparison, as evidenced by the large influence of the rotation even when this action is
not in fact optimal. However, small influences of more precise (model-based or successor-like)
knowledge is evident at the group level, indicating that the best model might allow for a mixture of
hierarchical abstraction and fine-grained knowledge (43).

Response times

All analyses with respect to response times described here were preregistered and thus of
confirmatory nature (https://osf.io/n2jcz/). Random-effects group-level Bayesian model selection
of the computational models of the response time data indicate a significant group difference in
model prevalence ($BOR < 10^{-10}$) with a strong preference for the successor representation model (*pxp*
$= 1$). This holds when considering binary model comparisons between the successor representation

and every other model (all *BOR* < 0.003, see Methods for multiple comparison correction. All *pxp* >
0.998). Posterior model probabilities for every participant are shown in Fig. 3A.

Fig 3 Response time data. (**A**) Horizontal stacked bar charts for every participant, illustrating the posterior model
probability derived from random effects Bayesian model comparison between four cognitive models, "Null" (orange),
explicit hierarchical "Exp" (pink), model-based "MB" (yellow), and successor representation "SR" (blue). Participant data
were sorted by posterior model probability for the "SR" model for interpretability. (**B**) Regression coefficients of the
successor representation model. Density plots show the full posterior distributions of the population means, with orange
lines indicating the 95% HDI. Dots represent posterior means of individual participant (random) effects. SPE refers to the
state prediction error, EV to expected value, and RPE to reward prediction error. See text for definition of remaining
terms.

We tested whether the estimates of the nuisance regressors included in the successor representation
model were indeed significantly related to response times (Fig. 3B). We do not find any evidence that
participants speed up or slow down based on the (log-transformed) number of trials on the current
miniblock ("logsteps") (M = -0.003, $HDI_{95\%}$ = [-0.014, 0.007], ER_- = 2.510). We do find a significant
effect of response slowing for rooms based on (log-transformed) recency ("logrecency"), indicating
rooms that have not been encountered for longer yield slower responses (M = 0.032, $HDI_{95\%}$ = [0.026,
0.037], ER_+ = ∞). We also find significant slowing for trials where the participant switched their
response key ("response switch") (M = 0.036, $HDI_{95\%}$ = [0.014, 0.059], ER_+ = 1378.310).

Further, we find trending effects for an influence of the action-outcome mapping (see Methods) on
response times (Fig. 3B). In particular, we find a trend for participants to be faster in rooms that have
more possible outcomes ("wide node", which have no "preferred" transitions) (M = -0.008, $HDI_{95\%}$ =
[-0.016, 0.001], ER_- = 28.390) and the transitions following these rooms ("wide trans") (M = -0.008,
$HDI_{95\%}$ = [-0.017, 0.001], ER_- = 26.615). Meanwhile, we find a trend for participants to be slower
following more expected transitions, which occur when both actions can transition to the same room
(Fig. 1C) ("preferred") (M = 0.006, $HDI_{95\%}$ = [-0.001, 0.012], ER_+ = 20.158). Parenthetically, these
trends run opposite to our initial expectations, although they are not statistically reliable according

to our definition, and reverse direction in a control model considering a state-based successor
representation (see S2 Appendix).

With respect to the regressors of theoretical interest (Fig 3B), we find significant slowing for
transitions yielding higher state prediction error ("SPE"), indicating slowing after transitions
between wings compared to transitions within wings (M = 0.015, $HDI_{95\%}$ = [0.007, 0.023], ER_+ =
6152.846). We also find significant slowing for higher expected values ("EV"), indicating that
participants respond slower for rooms closer to the current goal (M = 0.033, $HDI_{95\%}$ = [0.025, 0.042],
ER_+ = ∞). We also find significant slowing for transitions with more negative reward prediction
errors ("RPE"), consistent with post-error slowing when participants move away from the goal, or
conversely, speeding when participants move toward the goal (yielding positive reward prediction
errors) (M = -0.023, $HDI_{95\%}$ = [-0.029, -0.016], ER_- = ∞). Finally, we find no conclusive evidence that
the regressor for conflict, conceptualized as the difference in expected value between the two
different actions, influences response times ("conflict") (M = 0.002, $HDI_{95\%}$ = [-0.004, 0.007], ER_+ =
2.514).

Posterior predictive checks
We further explored the successor representation model by examining specific transition types (Fig.
4A). As illustrated in Fig. 4B, which shows the regressor values assigned by the successor
representation model to these trial types (extracted from the maximum a-posteriori fit for each
participant), the three transition types that cross between wings ("between", "outof", and "into")
yielded increased state prediction errors ("SPE"), as expected. Further, transitions that end in the
goal wing ("inside" and "into") yield larger expected values ("EV"), transitions out of the goal wing
("outof") yield larger negative reward prediction errors ("RPE"), and transitions into the goal wing
("into") yield larger positive reward prediction errors. As a validation of our task design, we also
replicated previous accounts of response time slowing for between-wing transitions (29,40,41) by
conducting a paired t-test of "outside" (green) and "between" (blue) transitions, which
approximately controls for effects related to value (EV, RPE, and conflict). We find significant slowing
for between-wing transitions ($M_{between>outside}$ = 34.797ms, $SD_{between>outside}$ = 85.459, t(118) = 4.442, p <
10^{-4}), comparable to these previous reports.

Fig. 4C plots the mean response times for participants separately for each trial type (left), and
matching mean response times for the full posterior predictive distribution (right). It can be seen the
model predicts response times in a very similar range as the real data, and mostly preserves the order
of different trial types. However, there appears to be a trade-off whereby transitions "inside" the goal
wing appear slightly overestimated, and transitions "into" the goal wing appear slightly
underestimated. Additional exploratory modeling reported in S2 Appendix suggests this can be
accounted for by modeling the absolute (unsigned) reward prediction error, as opposed to the signed
reward prediction error.

Even though the successor representation model accounts well for the response time effects
observed between the different transition types illustrated in Fig. 4A, Fig. 4B indicates substantial
variance in the regressor values between trials of the same transition type, including overlap
between the trials of different transition types. To investigate whether this variance in regressor
values is also reflected in response times, we ordered all trials by regressor value (separately for each
regressor) and binned them by intervals of 0.3. This yields a response time distribution for each bin;
Fig. 4D shows the mean (orange) and the 10th and 90th percentile (blue) for each bin. This reveals a
parametric relationship between response times and the regressors derived from the successor
representation, where for example the response times gradually increase in association with

increasing expected value. This fine-grained analysis provides additional support that participants
learn a detailed, predictive representation of the room locations, beyond just a simplified (explicit
hierarchical) representation of the relationships between the three different wings.

To probe further how the successor representation might explain response times at a fine-grained
(i.e. within-wing) level, we partitioned the trials as shown in Fig. 5A, indicating whether the current
room lies at a boundary allowing for a transition to a different community and specifying whether
this transition would correspond to entering the goal community (orange), leaving the goal
community (yellow), or switching between non-goal communities (blue). These trials can be
compared against trials within the goal community (pink) or outside of it (green). To ensure the
variance in response times on these trials corresponds to knowledge of the within-community
structure of the museum, we excluded all trials that followed a between-community transition. We
conducted three paired t-tests to investigate sensitivity to the within-community structure of the
museum (corrected for multiple comparisons; see Methods). With respect to rooms that are outside
the goal community (Fig. 5B), we found a significant increase in response times in rooms at the
boundary that allows for a transition between two non-goal communities (blue), compared to non-
boundary rooms outside the goal community (green) ($M_{non\text{-}goal>outside}$ = 6.921ms, $SD_{non\text{-}goal>outside}$ =
32.292, $t(119)$ = 2.348, p = 0.041). We also find a significant increase in response times for rooms at
the boundary that allows for a transition "toward" the goal community (orange) compared to rooms
at the boundary that allow for a transition between non-goal communities (blue) ($M_{toward>non\text{-}goal}$ =
10.714ms, $SD_{toward>non\text{-}goal}$ = 42.671, $t(119)$ = 2.751, p = 0.021). Fig. 5C shows posterior predictions
from the successor representation model. This reveals the model predicts a similar increase in
response times for boundary-rooms that allow for transitions into the goal community (orange), but
not for boundary-rooms that allow for transitions away from the goal community (blue; compared
to non-boundary rooms outside the goal community, green). Inspection of the boundary and non-
boundary rooms in the goal community (Fig. 5D) did not reveal statistically significant differences in

response times ($M_{outof>within}$ = 6.445ms, $SD_{outof>within}$ = 79.123, $t(119)$ = 0.892, p = 0.374), nor does the model predict any such differences (Fig. 5E).

Fig 4 Posterior predictive checks of the response time successor representation model between communities. (**A**) Layout of the museum with transitions colored according to type. The goal painting is colored black and labeled "G". Green transitions stay within wings that do not contain the goal painting ("outside"). Blue transitions move between non-goal wings ("between"). Orange transitions move from a non-goal wing into the goal wing ("into"). Pink transitions move from a goal wing into a non-goal wing ("out of"). Yellow transitions stay within the goal wing ("inside"). (**B**) Distributions of regressor values assigned to trials following the different transition types as defined in A. Regressor values were computed based on the discount factor value at the maximum a-posteriori of the joint likelihood of the participant-level successor representation model fit. SPE: state prediction error, EV: expected value, RPE: reward prediction error. (**C**) Mean (log-transformed) response times (left) and mean posterior predictions (right) for each participant for the different

534 transition types as defined in A. Population means with their 95% confidence interval are shown in black. (**D**) Mean
535 (orange) and 10th and 90th percentile (blue) of the response time distribution for trials binned by regressor values (as
536 assigned in B) in intervals of 0.3. Binned separately for the different regressors (SPE, EV, RPE, conflict). Conditional
537 response time distributions are repeatedly plotted side by side for real data (left) and the full posterior predictive
538 distribution (right).

Fig 5 Posterior predictive check of the response time successor representation model within communities. (**A**)
Layout of the museum with rooms colored according to type. The goal painting is colored black and labeled "G". Green
rooms lie within wings that do not contain the goal painting ("outside"). Blue rooms lie at the boundary between two non-
goal wings ("away"). Orange rooms lie at the boundary allowing for a transition into the goal wing ("toward"). Pink rooms
lie within the goal wing ("within"). Yellow rooms lie at the boundary allowing for a transition out of the goal wing ("out
of") (**B**) Mean response times (standardized within participant) in different rooms outside the goal wing, as labeled in A.
(**C**) Same as B, but drawn from posterior predictions of the fitted successor representation for each participant. (**D**) Same
as B, but for rooms inside the goal wing. (**E**) Same as D, but drawn from posterior predictions of the fitted successor
representation for each participant.

Control analyses and convergence

We conducted several control analyses to check the robustness of the above results, and replicated
all results in all cases (see S2 Appendix). Additionally, we explored an alternative formulation of the
successor representation that only tracked state-state predictions (23), rather than state-action
conjunctions (24,26). We found that these two models both perform significantly better than the
other models (null, explicit-hierarchical, and model-based), but do not significantly differ from each
other (see S2 Appendix). Exploring the state-action successor representation further, S1 Appendix
shows that the successor representation is gradually learned over the course of the training phase,
and has saturated by the time the testing phase starts. Additionally, S1 Appendix shows that the
variance explained by the local surprise (as indexed by the "log-recency" nuisance regressor) is
independent from the variance explained by the global surprise (as indexed by the successor
representation prediction error).

Free sort post test

We investigated whether the participants learned an accessible representation of the structure of the
museum by asking the participants after the testing phase to reconstruct the layout of the museum
by arranging the paintings on a grid. The Euclidean distances between the positions of all possible
painting pairs (105 per participant) were then computed and standardized within participants. Fig.
6A shows the individual distances and average distances for all painting pairs, classified according to

four different categories. A hierarchical Bayesian linear regression (preregistered;
) indicated that participants on average placed paintings that were part of the
same community closer together ($M = -0.578$, $HDI_{95\%} = [-0.699, -0.460]$, $ER_- = \infty$) (Fig. 6B), as reflected
in low distances attributed to "community-connected" paintings (Fig. 6A). However, this effect was
attenuated for paintings that were not directly connected (i.e. between the two paintings that lie at
the "boundary" of each community, Fig. 6B) ($M = 0.183$, $HDI_{95\%} = [0.071, 0.295]$, $ER_+ = 1051.632$),
consistent with higher distances reported for these relationships ("community-boundary", Fig. 6A).
The participants also placed paintings that were directly connected but not part of the same
community (i.e. between-community transitions or "bridges", Fig 6B) closer together ($M = -0.489$,
$HDI_{95\%} = [-0.633, -0.349]$, $ER_- = \infty$), as shown by the lower distances for "bridge" relationships (Fig.
6A). This pattern of results is suggestive of fully accurate reconstruction of the graph, since
coefficients associated with all lower-order relationships indicate significant effects. To test for
specific bias introduced by the community structure (as preregistered), we checked whether
participants placed paintings that are connected and part of the same community closer together
than paintings that are connected but not part of the same community ("bridges"). Although this
analysis yielded an effect indicative of bias, the result was not statistically reliable ($M = -0.089$, $HDI_{95\%}$
$= [-0.191, 0.013]$, $ER_- = 12.387$; Fig 6C, "community direct").

We ran an additional exploratory test to isolate the effect of community structure from the effect of
direct connections (44). A hierarchical Bayesian linear regression that included regressors only for
pairs of paintings with the shortest possible distance of 1 intermediate room between them indicated
that the participants placed paintings with a path distance of 1 closer together if they were part of
the same community (i.e. the two paintings at the boundaries of one community) compared to when
they were part of different communities ($M = -0.238$, $HDI_{95\%} = [-0.347, -0.131]$, $ER_- = \infty$) (Fig. 6C,
"community-indirect").

Fig 6 Free sort post-test results. (**A**) Euclidean distances (normalized) for different pairs of paintings, grouped by
whether they were part of the same community and directly connected (community-connected) or not (community-
boundary), or whether they were directly connected but not part of the same community (bridge), or not connected and
not part of the same community (unrelated). Orange dots with error bars represent mean with 95% standard error. (**B**)
Regression coefficients of the free sort regression. Density plot shows the full posterior distribution of the population

mean, with orange lines indicating the 95% highest density interval. Dots represent individual participant posterior
means. Note that the "community" effect captures both "community-connected" and "community-boundary" effects as
shown in A, and that the "boundary" effect captures their difference. (**C**) Effects that quantify bias induced by community
structure. The "community-direct" effect only investigated directly connected rooms, and asked whether paintings of the
same community (community-connected in A) were placed closer together than those of different communities (bridge in
A). The "community-indirect" effect is analogous, but investigating pairs of rooms that had a minimal distance of one
intermediate room. This corresponds to "community-boundary" paintings in A, and a subset of "unrelated" paintings.

Successor representation modularity is associated with better hierarchical structure learning

The degree that a successor representation accounts for community structure depends on the
discount factor parameter for that model. We quantified this by computing the ratio of the state
prediction error yielded by a transition between two wings, over the state prediction error yielded
by a transition within a wing (controlling for "preferred" transitions; see Methods). We called this
measure the "modularity". A modularity of 1 indicates similar state prediction errors for transitions
between wings and within wings, whereas modularities greater than 1 indicate larger state
prediction errors for transitions between wings. For example, a modularity of 2 would indicate the
state prediction error for a transition between wings is twice as large as for a transition within one
wing. This modularity measure depends on the value of the discount factor, but also on the
experienced sequence of states. Fig. 7 (A,B) plot the modularity across a range of possible discount
factors, after a sequence of states as experienced by two representative participants in our dataset.

Fig 7 Modularity of the successor representation model of response times. (**A,B**) Relationship between possible
discount factor values and "modularity" of the successor matrix after a sequence of states as experienced by two

representative participants. Modularity is measured as the ratio of state prediction error for between-wing transitions over within-wing transitions (corrected for "preferred" transitions; see Methods). Notice that A shows a "peak" (orange line) just before 1, whereas B linearly increases with a maximum at 1. These two patterns were ubiquitous in our data set. Not all discount factors lead to modular successor representations (horizontal grey line at modularity = 1). For both these participants, only discount factors above about 0.25 show an effect of community structure. (**C**) Posterior means of the recovered discount factor parameters for all participants (dots) and kernel density estimate over these. (**D**) Modularity of the successor matrix for all participants (blue dots) under a null model and under the fitted successor representation model. (**E**) State prediction errors for all state-action transitions (even those not possible in the actual experiment), based on the estimated successor matrix for an example participant in our dataset (derived from the full posterior distribution; mean estimated discount factor of 0.978). Hotter colors indicate increased prediction error. Both axes index state-action conjunctions, with states labeled 1-15 as in Fig. 1A, and actions labeled as (i) or (ii) as in Fig. 1B. The community structure is visible as 'squares' of decreased prediction error for states (rooms) that are part of the same community (wing). (**F**) Posterior model probabilities of the successor representation response time model (x-axis) and the modularity measure (y-axis) for all participants (blue dots) with line of best fit (orange). (**G**) Similar to F with total accumulated reward bonus on the y-axis.

Fig. 7C shows the recovered posterior mean discount factors for the response time successor representation model. Although the distribution peaks around a value of 0.875, the density is relatively dispersed, decreasing as it approaches 0. Peaks close to (but not exactly at 1) correspond to successor representations that most strongly capture the community structure of the museum environment (e.g. Fig. 7A). We computed the expectation of our modularity measure and tested whether it was significantly larger for the posterior distributions obtained for each participant as compared to a null model (a uniform prior over discount factors) (Fig. 7D). A paired t-test confirmed a significant increase in modularity ($M_{fit>null}$ = 0.132, $SD_{fit>null}$ = 0.182, $t(119)$ = 7.977, $p < 10^{-11}$). Note that the null model also predicts some degree of modularity, since it reflects a uniform distribution over the discount factor, including values that reflect higher modularity. An example of a modular successor representation as estimated for one of our participants is shown in Fig. 7E.

A preregistered (https://osf.io/n2jcz/) analysis indicated that across participants, higher estimated modularities were predicted by a higher posterior model probability for the successor representation model of response times (simple linear regression: β = 0.196, $F(1,118)$ = 4.468, R^2 = 0.036, p = 0.037 ; Fig. 7F). Note that this is a test between two measures both derived from the same fit of the successor representation model of response times. It thus provides insight mainly into the patterns of participant behavior that are uniquely explained by the successor representation model over the competing models (null, model-based, and explicit hierarchical), but not into whether successor representation learning can be observed in independent measures of task performance. By contrast, an additional preregistered regression analysis asked whether the successor representation fit significantly predicted the performance of participants as expressed in terms of their total accumulated bonus payments. However, this analysis did not show evidence for any such relationship (β = -0.073, $F(1,118)$ = 0.007, $R^2 < 0.001$, p = 0.935; Fig. 7G).

Although there was strong evidence in favor of a successor representation model of response times across most individuals (see Fig. 3A), there were also large individual differences in the modularity of these successor representations. Note that the modularity is directly derived from the state prediction error metric. Hence, the degree of modularity of the successor representation for a specific participant could be completely driven by the relative slowing on between-community transitions (compared to within-community transitions). Additionally, the estimated modularity of the successor representation could be informed by response time variance within communities, such as the slowing for rooms at the boundaries of a community as reported in Fig. 5. This implies that modularity is in fact a more informative measure than response time slowing. We conducted a series of exploratory analyses that predicted independent task performance metrics based on the

modularity of the response time-fitted successor representation models. To control for pure between-community response time slowing, we constructed hierarchical regression models where we first entered between-community response time slowing as the sole independent variable, and compare this against a second regression model including both response time slowing and modularity as independent variables. Between-community response time slowing was defined as the difference in mean response times between "outside" (green) and "between" (blue) transitions as indicated in Fig. 4A. The modularity of the successor representation is defined as the ratio of the state prediction error elicited by between-wing transitions over within-wing transitions (corrected for "preferred" transitions; see Methods). This tests whether the modularity of the successor representation provides information above and beyond the between-community response time slowing. In total, this exploratory analysis yielded 15 regression models (all individual panels of Fig. 8). When reporting the second-step model, we report the F and R^2 statistics of the model itself, the slope of the newly added modularity regressor, and the ΔF statistic corresponding to the model comparison between the first- and second-step regression models. The reported p-value corresponds to this ΔF statistic. We corrected p-values for multiple comparisons across all first-step regression models and all comparisons with the second-step regression models (15 tests in total; see Methods for multiple comparison correction).

Fig 8 Modularity correlations with task performance. (**A**) Response time slowing (x-axis) for transitions between communities (blue in Fig. 4A) compared to transitions within communities (green in Fig. 4A), shows a positive relationship with Modularity (y-axis), defined as the ratio of the state prediction error elicited by between-wing transitions over within-

wing transitions (corrected for "preferred" transitions; see Methods). (**B**) Between-community response time slowing (top)
and modularity (bottom) show a positive trend and significant positive relationship with acquired total reward bonus
respectively. (**C**) Between-community response time slowing (top) and modularity (bottom) show a positive relationship
with the explicit hierarchical "rotation" choice policy. (**D**) Between-community response time slowing (top) and modularity
(bottom) show a positive relationship with action-switching immediately upon leaving the goal wing ("Leave"). (**E**)
Between-community response time slowing (top) and modularity (bottom) do not show a relationship with action-
repetition when entering the goal wing ("Entry"). (**F**) Between-community response time slowing (top) and modularity
(bottom) show a negative relationship with distance between paintings of the same community ("Within") (i.e., the
paintings are placed closer together). (**G**) Between-community response time slowing (top) and modularity (bottom) show
a positive relationship with distance between paintings at the boundaries of the same community ("Boundary"). (**H**)
Between-community response time slowing (top) and modularity (bottom) show a negative relationship with distance
between paintings connected across communities ("Bridge"). Significance of regression models indicated as *: $p < 0.05$, **:
$p < 0.01$, ***: $p < 0.001$. For RT slowing, this significance corresponds to a simple linear regression. For SR modularity, the
significance corresponds to a model comparison of a multiple regression model (including both RT slowing and SR
modularity as independent variables) against a simple linear regression (including only RT slowing as independent
variable).

Our exploratory analysis confirmed a significant positive relationship between response time
slowing and modularity ($\beta = 0.001$, $F(1,117) = 7.472$, $R^2 = 0.060$, $p = 0.044$; Fig. 8A). Contrary to the
posterior model probability of the successor representation (Fig. 7G), between-community response
time slowing significantly predicted total accumulated reward bonus ($\beta = 0.005$, $F(1,118) = 6.485$, R^2
$= 0.052$, $p = 0.049$; Fig. 8B, top). Including the modularity of the successor representation as a second
independent variable leads to a significant improvement in model fit ($\beta = 3.433$, $F(2,117) = 12.435$,
$R^2 = 0.175$, $\Delta F(1,117) = 17.478$, $p < 0.001$; Fig. 8B, bottom). Note that the total bonus payment
depended solely on the participant reaching a new goal in the minimum number of steps possible.
The time taken to generate these choices (the response time) was not related to reward, as
participants could perform the task entirely at their own pace. Instead, this relationship is suggestive
of a link between successor representation learning and task performance, where participants who
are more sensitive to the community structure of the museum generate more efficient paths to the
goal.

In accordance with the positive relationship between response time slowing and reward, we
observed a significant positive relationship between response time slowing and the posterior means
for the explicit hierarchical "rotation" regressor of the choice data ($\beta = 0.003$, $F(1,117) = 14.281$, $R^2 =$
0.109, $p = 0.002$; Fig. 8C, top), indicating that participants whose response times were more affected
by the community structure also reflected this community structure more strongly in their choice
behavior. Including the modularity of the successor representation as a second independent variable
leads to a significant improvement in model fit ($\beta = 0.768$, $F(2,116) = 10.694$, $R^2 = 0.156$, $\Delta F(1,116)$
$= 6.443$, $p = 0.037$; Fig. 8C, bottom). In addition, the data shows a positive relationship between
response time slowing and the degree of response-switching when participants left the goal wing (β
$= 0.003$, $F(1,117) = 14.642$, $R^2 = 0.111$, $p = 0.002$; Fig. 8D, top). Including the modularity of the
successor representation as a second independent variable leads to a significant improvement in
model fit ($\beta = 0.995$, $F(2,116) = 11.810$, $R^2 = 0.169$, $\Delta F(1,116) = 8.090$, $p = 0.032$; Fig. 8D, bottom). By
contrast, no relationship is observed between response time slowing and response switching upon
goal entry of the goal wing ($\beta = 0.000$, $F(1,117) = 0.061$, $R^2 = 0.001$, $p = 1$; Fig. 8E, top). Including the
modularity of the successor representation does not yield any improvement in model fit ($\beta = 0.005$,
$F(2,116) = 0.031$, $R^2 = 0.001$, $\Delta F(1,116) = 0.000$, $p = 0.979$; Fig. 8E, bottom).

Although response times and choices reflect independent dimensions of the museum task, they are
derived from the same behavioral unit (i.e. "the choice"). By contrast, the post-test free sorting task
is completely independent from the behavior used to estimate modularity. Nevertheless, the data

show a significant relationship between response time slowing and how close participants would sort paintings from adjacent rooms in the same wing (β = -0.002, $F(1,116)$ = 18.926, R^2 = 0.140, $p <$ 0.001; Fig. 8F, top). Including the modularity of the successor representation as a second independent variable leads to a significant improvement in model fit (β = -0.991, $F(2,115)$ = 26.401, R^2 = 0.315, $\Delta F(1,115)$ = 29.264, $p < 10^{-5}$; Fig. 8F, bottom). Interestingly, the data also showed a relationship between response time slowing and how far participants placed rooms that lie at the boundaries of the same wing, and are thus not connected (β = 0.001, $F(1,116)$ = 16.069, R^2 = 0.122, $p < 0.001$; Fig. 8G, top). Including the modularity of the successor representation as a second independent variable leads to a significant improvement in model fit (β = 0.377, $F(2,115)$ = 26.851, R^2 = 0.318, $\Delta F(1,115)$ = 33.177, $p < 10^{-5}$; Fig. 8F, bottom). Higher values of both between-community slowing and successor representation modularity thus correspond to participants who correctly placed these rooms further apart, which in fact provides evidence against community structure and in favor of them accurately learning the true adjacencies. Similarly, the data showed a relationship between response time slowing and how closely participants would place two rooms that are connected between different wings (β = -0.002, $F(1,116)$ = 17.548, R^2 = 0.131, $p < 0.001$; Fig. 8H, top). Including the modularity of the successor representation as a second independent variable leads to a significant improvement in model fit (β = -0.872, $F(2,115)$ = 28.219, R^2 = 0.329, $\Delta F(1,115)$ = 33.912, $p < 10^{-6}$; Fig. 8H, bottom). Higher values of both between-community slowing and successor representation modularity thus correspond to participants who correctly placed these rooms closer together, which is also in line with more accurate learning of true adjacencies. Collectively, these results indicate that participants who learned community structure using successor representations (as evidenced in response times), were better able to reconstruct the museum layout later from memory, while doing so with *less* bias towards reconstructing the underlying community structure.

DISCUSSION

Human behavior in the real world is characterized by long-range dependencies between action plans and outcomes (5,6). Much of this information is high-dimensional and complex, but can be compressed to yield more efficient, generalizable representations (9,10,45–48). Although how humans learn new behaviors according to such principles has remained elusive, previous work has shown that the computational burden involved in naturalistic behavior can be alleviated by setting "subgoals" that constrain planning over low-level actions (11,13–15,37,49). It has been suggested that predictive representations such as the successor representation can be leveraged to identify such abstractions (17,23). Predictive learning is ubiquitous in humans (50–53) and individual differences in predictive learning have previously been linked to individual differences in goal-directed decision making (54). Specific evidence exists that humans can learn predictive representations that are sensitive to higher-order aspects of tasks (29,30,40,41). As far as we are aware, this study is the first to show that individual differences in higher-order predictive learning ability predict individual differences in hierarchical abstraction during goal-directed decision making.

In order to link predictive representation learning to hierarchically-informed decision making, we adapted a task that has been previously leveraged to show higher-order influences on predictive learning (30). The museum task requires participants to navigate through a virtual museum, where the connections between rooms were characterized by community structure. Crucially, a navigational strategy based on a learned successor representation should differ from a strategy based on abstract, hierarchically structured representations, the latter of which are argued to be less cognitively demanding and faster to execute, even if otherwise sub-optimal (6,37,55–58). This allows

for investigating the emergence of hierarchical abstraction on a single task, and establishes a clear
link with earlier work on predictive learning (23,29,59). We find participants follow such a
hierarchically abstracted strategy, consistently selecting the action that is guaranteed to reach the
desired community (subgoal) and dynamically switching their behavior immediately following
undesired transitions out of this community.

It has been shown participants are generally slower to respond to more surprising events (29,60,61).
Crucially, events that are part of separate communities are experienced as more surprising, even
though these events are not in fact less predictable (29,40,41). Notably, we replicated this effect in
our novel goal-directed decision making task, and observed it to be significantly related to the
modularity associated with the successor representation model.

Although previous studies on predictive learning did not investigate goal-directed decision making,
recent work has shown that response times are influenced by the values of the different options in a
value-based decision making task (32–36). Note that these studies investigate decisions between
stimuli leading to immediate outcomes instead of sequential decisions that require planning.
Interestingly, it has been observed that learned values can invigorate the responses of rodents
engaged in temporally extended behaviors, including by increasing their speed (62–64). In particular,
it has been argued that gradually increasing vigor when approaching a goal (i.e., a "goal gradient")
(65,66) may be a core feature of flexible, goal-directed behavior (67–70). We also observed a gradient
in the response times in our task, indicating the use of a learned cognitive map to flexibly compute
values for changing goals. However, the participants slowed rather than sped up when they
approached the goal. We posit that this occurred because the task encouraged participants to slow
their responses when nearing the goal location in order to identify the target painting and obtain the
financial reward. By contrast, unlike higher positive values (which predicted slower responding),
higher positive *reward prediction errors* predicted faster responding. However, our data suggest that
instead of a relationship between response time and signed reward prediction errors (i.e., that
distinguish between positive and negative outcomes), participants might instead slow down for
larger absolute reward prediction errors (i.e., that collapse across the valence of the outcome).
Exploratory modeling indeed suggested this relationship held more strongly (see S2 Appendix). This
suggests that the effect of reward prediction error on response times is most parsimoniously
explained as surprise-based slowing with respect to task progress.

Crucially, the museum task required participants to revise their value predictions in a flexible
manner, as the goal locations were constantly changing. Tasks that involve changing goal locations
cannot be simulated using "model-free" algorithms that have been previously leveraged to
investigate the relationship between response times and value predictions (32–36). Although this
process is consistent with a canonical "model-based" reinforcement learning algorithm, model-based
algorithms do not make specific predictions about the observed influence of community structure on
response times. By contrast, a computationally simpler but less flexible successor representation
algorithm (18,25) can do all of this. Here, successor representation learning specifically accounted
for the hierarchical abstraction observed in our choice data (23). Further, the degree to which
individual participants learned the community structure through the successor representation
correlated with how effectively they applied a hierarchically abstract choice policy, although we note
that this result is derived from an exploratory (i.e. not pre-registered) analysis. This pattern of results
agrees with a theoretical model wherein participants learn a predictive representation of the task by
engaging with it (as revealed by the successor representation model), and apply a secondary
computational operation to this predictive representation in order to find an effective hierarchical
representation useful for choice behavior (as revealed by the explicit hierarchical model).

Several algorithms have been proposed that implement this secondary computation by taking an
adjacency matrix as (part of) their input (13,15,16,23). This work generally assumes that the
adjacency matrix has been accurately learned by the participants without considering how the
specifics of the primary learning process could influence the abstraction process. Instead, we show
that individual differences in learning a successor representation correlate directly with the efficacy
of hierarchical abstraction. This finding is contrary to algorithms that rely on spectral graph theory
wherein the relevant eigenvectors are constant over the entire range of possible discount factors
(23), thus predicting no relationship between the estimated discount factors and the effectiveness of
the hierarchical abstraction. Further, if abstractions rely on graph-theoretic measures of
betweenness-centrality as described elsewhere (16), then the abstractions could be instantly
computed from the successor representation as the "simple randomized shortest paths"
betweenness, where the "fundamental matrix" corresponds directly to the successor representation
(71, algorithm 1). However, the quality of this abstraction increases with smaller discount factors,
which is the opposite of what we find. Additionally, subgoal identification through betweenness-
centrality does not explicitly assign lower-level states to higher-level state abstractions, leaving open
the question why people would apply a simplified general "rotational" policy to all states in one
community.

By contrast, our findings are more consistent with a Bayesian model of hierarchical abstraction as
proposed by (15). This model considers several factors for identifying efficient hierarchical
representations, including dense connectivity within state abstractions, and sparse connectivity
between them. When relying on a successor representation as input, as opposed to a ground-truth
adjacency matrix, this density metric is by definition maximized when discount factors are most
sharply tuned to capture the modularity of the environment. Similarly, (13, equation 7) propose an
algorithm for computing the "entropic centrality" of different nodes. This computation can be
performed directly by using the successor representation—which directly corresponds to the
"fundamental tensor"—and substituting the degree of each node by the summed entries in the
successor representation corresponding to that node. They observe that normative hierarchical
abstractions often minimized differences in entropic centrality within state abstractions, and
maximized this between state abstractions. This difference should be most sharply tuned in a
successor representation when the discount factor emphasizes the modular structure most strongly,
in line with our reported results.

Why would participants need to develop high-level (hierarchical) representations when the low-
level successor representations already get the job done? Notably, both (13) and (15) explicitly
assign lower-order states to higher-order state abstractions, and run separate planning algorithms
at both levels. In each of these models, the lower-level planner should in principle be able to find
policies with more fine-grained detail than following the simplified "rotational" policy. However, if
we assume running the lower-level planner is computationally (72) (or cognitively, (6,56)) costly,
agents might learn to omit running this planner, specifically when they can also learn that always
executing the same action for every state in one state abstraction reliably yields good results (as in
the Museum task). Although it has previously been argued that policies can be directly computed by
multiplying the successor representation with a reward vector (24), which should not be
computationally expensive, in more complex domains or for longer timescales, generative sampling
may still be required to reliably asses the outcomes of the possible actions on a single trial based on
a learned successor representation (73–75). Note that we only observe evidence for value
computations based on a successor representation in the response times by fitting the model across
many trials and many participants. Sampling from the successor representation would likely incur

computational or cognitive costs (e.g. 76), which could be alleviated by hierarchical abstraction. We
speculate that participants follow the rotational policy in the Museum task mainly for this reason,
despite the fact that their response times reveal sensitivity to relatively detailed representations of
the underlying task structure. That participants' choice behavior can access such details is also
illustrated by the significant decrease in the proportion of rotation choices when the rotation is in
fact suboptimal (Fig. 2F-H).

Note that the particular successor representation model implemented here is relatively simple, as
more advanced methods that yield better generalization in the face of changing goals have been
suggested (77,78). For example, (79) proposes a stable successor-like "default" representation that
(typically) corresponds to a random policy. This representation can instantaneously adapt to "skew"
with respect to the appropriate policy for a newly presented goal, unlike the typical successor
representation which is updated through experience and thus skews towards recently pursued (but
no longer relevant) goals. Some evidence for more advanced forms of successor-based generalization
has been observed in human choice behavior (80). Other research has shown that graph-topological
features such as community structure can be inferred based on prior learning (81). The latter
research has also shown that this inference can lead to efficient navigation between task states by
selecting the boundary nodes when navigating between communities. However, this is not evidence
for a simplified hierarchical representation of the task, because any non-hierarchical planning
mechanism also makes this same prediction. More generally, many prior studies investigating
learning and generalization only analyze choice behavior. Instead, the current study illustrates how
examination of response times can reveal signatures of predictive learning that may underlie more
advanced forms of generalization in goal-directed learning.

In order to capture sensitivity to the probabilistic action-outcome contingencies in the task design,
we modeled participant behavior using successor representations that predicted state-action
conjunctions (24,26). By contrast, the original formulation of the successor representation, later
invoked as a model of hippocampal function, only predicts relationships between task states (18,23).
The state-state formulation is well-suited for simulating decision making in spatial contexts where
the outcomes of specific actions are fixed (e.g., the action "go left" always results in moving left).
Interestingly, the prediction of state-action-outcome conjunctions has been invoked as an
explanation for several effects related to cognitive control and flexible decision making in the medial
prefrontal cortex (82–86). Additionally, recent evidence suggests that distinct dopaminergic
populations may encode the prediction error associated with executing specific actions (although not
their actual identity) (87–90). The museum task was not specifically designed to disambiguate
between predictive representations of states or state-action conjunctions, and although we observed
a numerical preference for a state-action predictive model, we did not find any statistically reliable
differences between these models (see S2 Appendix). When inspecting posterior predictions of
response times for various rooms within a single community (Fig. 5), we found that the state-action
successor representation does not predict slowing in the boundary room that allows for a transition
between two non-goal communities (blue). Because participants are prone to select the action that
allows for rotation towards the goal community, they typically select the action allowing for a
transition between communities when in the boundary room that allows for a transition toward the
goal community (orange). Conversely, participants are not very likely to select the action allowing
for a transition between communities when occupying the boundary room that allows for a transition
between two non-goal communities (blue). This implies a smaller state prediction error for the latter
rooms for the state-action successor representation as compared to the state-state successor
representation, because outcomes that transition between communities are excluded in the former

but not the latter case. As a consequence, specifically the state-action model incorrectly predicts less
slowing for rooms at the boundary of two non-goal communities. Additionally, the boundary room
that allows for a transition towards the goal community is assigned a substantially higher expected
value than the other two types of rooms, further contributing to the model-predicted slower response
times.

Besides various implementations of the successor representation, our data could be explained by
alternative sequence learning mechanisms, for example "free energy minimization" (29) or (cloned)
hidden Markov models (91), in so far as these can implement a mechanism that is isomorphic to the
discount factor of the successor representation. Note that the prediction of action (sequences) can
form an alternative basis for hierarchical reinforcement learning (57,92–94), but this implies
inflexible "open-loop" control. The museum task contained probabilistic action-outcome mappings
and required participants to search for a new goal location on every miniblock, preventing against
this strategy. The interaction between action sequences and state based hierarchical learning could
be an interesting direction for future research.

Note also that the successor representation yields a vector error signal that indicates the degree to
which predictions of multiple possible upcoming events need to be adjusted. This computation
agrees with the distributed nature of the prediction error signals observed in the medial prefrontal
cortex (82,84,86,95) and in the midbrain dopamine system (20,96). However, modeling of surprise-
based slowing of response times requires that this error vector be converted to a scalar signal.
Toward this end, previous work has modeled response times using a successor representation by
normalizing the successor matrix to resemble a transition probability matrix. The scalar surprise
signal was then obtained by looking at the transition probability from the original to the outcome
state (97). In line with the intuition that the successor representation estimates the "expected future
discounted state occupations", this work specified the target for the update as a one-hot vector of the
outcome state (44,97). However, it has been noted that the correct target for the update is a one-hot
vector of the original state summed with the discounted successor vector of the outcome state
(18,24). Hence, the surprise signal should reflect this latter component, which is vectorial. For this
reason, we modeled the response times with a formal distance metric (the angular distance) in order
to yield a scalar surprise signal.

Not only were the individuals who showed greater sensitivity to community structure in their
response times (as measured through the discount factor of the successor representation) more
sensitive to hierarchical abstractions in their choice behavior, they also reconstructed the lower-
order structure of the museum more accurately from memory. This included more accurate
separation of the "boundary" rooms within the same community, which counterintuitively indicates
less bias with respect to community structure. We speculate that the individual differences in the
estimation of the discount factor reflect how engaged the participants were with the task. In
particular, we suggest that individuals who paid more attention to the task learned more accurate
successor representations of the environment, which simultaneously provided more information
about both low-level (room) and high-level (wing) task structure. Interestingly, (29) proposed that
the learning of community structure is mediated by "mental errors", which should be more prevalent
for participants with lower task engagement. Our data instead suggest that the learning of
community structure improves with higher task engagement.

In addition to the behavioral evidence for successor representation learning (26,28), functional
magnetic resonance imaging of humans performing incidental learning tasks has revealed successor-
like representations in the hippocampus (98–101). During decision making behavior, the medial

frontal cortex and specifically the anterior cingulate cortex show sensitivity to hierarchical features
of tasks (102–104). Our results can be interpreted within this context, where the hippocampus might
be involved in general statistical learning (105), whereas the anterior cingulate cortex represents
variables relevant for decision making, specifically models that can be leveraged for hierarchical
reinforcement learning (6,106). These hierarchical models may be extracted from successor-like
hippocampal representations as a form of "schema" learning or meta-learning (107–109). Future
work can leverage neuroimaging techniques to search for task representations and prediction errors
associated with successor learning and hierarchical reasoning in the museum task, and whether
these are represented in distinct neural regions, potentially following a rostro-caudal abstraction
gradient associated with frontal cortical function (110).

METHODS

Museum tour guide game

A preregistration of the museum task and our main analyses can be found at https://osf.io/n2jcz/,
including all of the relevant code and data. The task was programmed with JsPsych version 7.1.0
(111). Data for 141 participants were collected on the Prolific platform (38). All participants gave
their informed consent prior to participation and the study was conducted according to the
guidelines of the General Ethical Protocol of the Faculty of Psychology and Educational Sciences
(Ghent University) and the ethical standards prescribed in the 1964 Declaration of Helsinki. The data
of 21 of participants were excluded based on preregistered criteria: 3 of these participants chose the
same action on over 90% of the trials, and 18 of these participants selected the wrong goal or missed
the correct goal on over half the miniblocks, both of which are indicative of task disengagement. Data
of the free sorting task was corrupted for 1 participant, meaning this participant was not included in
the analysis of the free sort distances (but still included in all other analyses).

Training and testing phase

We wanted participants to learn the structure of the environment from experience, including the
three communities or "wings" of the museum. The learning process can be captured with a
computational model as described below. Importantly, this means participants were not instructed
on any aspect of the layout of the rooms of the museum, so as not to bias their learning and decision
making behavior prior to task experience. Instead, we designed a training phase where participants
would see every goal painting and possible distance between start and goal location exactly the same
number of times. This ensures an unbiased predictive representation of the museum can be learned
from experience. We designed a separate testing phase where we could then investigate the final
decision making behavior of the participants.

In the training phase, the participants were given exactly 75 goal paintings to search for (i.e. 75
"miniblocks"). They were provided an opportunity for a short rest every 25 miniblocks. By contrast,
the test phase did not limit the number of miniblocks to perform. Instead, the participants were given
a "budget" of 1000 room transitions where each room they visited deducted 1 "step" from this budget.
Importantly, efficient policies enabled participants to visit more goals in a limited number of steps
and thus accrue more reward (see S1 Appendix). The participants were given the opportunity for a
short rest approximately midway through the test phase, namely after completing the miniblock
when their budget dropped below 500 steps. The testing phase terminated after the miniblock on
which their budget dropped below 0 steps. The reward for finding each goal painting in the training
phase was £0.10. This increased to £0.15 in each testing phase. The penalty for indicating the wrong
goal or missing the correct goal was always £0.02.

In the training phase, combinations of start and goal locations were exactly balanced, preventing
against the development of model-free preferences for specific rooms (see S1 Appendix). In the
testing phase, the start and goal locations, and the distances between them, were no longer balanced
on the assumption that the participants learned an unbiased representation of the museum layout in
the training phase. In the testing phase, the goal locations were sampled with a uniform probability
from the 10 rooms outside the starting wing. The previous goal location was always the next starting
location, just as in the training phase.

Action-outcome mapping
Participants were requested through Prolific to perform the task using a QWERTY-layout keyboard.
On each trial, they were required to select between key <z> and key <m> on their personal keyboard.
For each state, each key mapped at random with equal probability to two possible outcome states
(Fig. 1B). Which key mapped to which set of outcomes was determined randomly for each
participant, but consistent across wings, meaning there were two possible action-outcome mappings.
A primary feature of this action-outcome mapping was that each key consistently mapped the
participant out of the current wing in only one direction (labeled as "intention" in Fig. 1B).

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Fig 9 Transition types with respect to preferred transitions. Arrows colored by transition type with respect to the
preferred and removed transitions as shown in Fig. 1C. Here, preferred transitions are coded as orange ("preferred")
(blue arrows in Fig. 1C). These can be the outcomes of selecting both <z> and <m> in their outgoing states. These
transitions were included as nuisance regressors in our models, since we expect them to yield faster response times as
they are more predictable. It can be seen that only the top node does not have any outgoing preferred transitions, and
when referencing Fig. 1B, it can be seen this is the only node with 4 possible immediate outcomes (2 distinct outcomes for
each action). Hence, this node is referred to as "wide node" and transitions into this node are labeled as such (yellow).
Transitions out of this node are labeled as "wide trans" (pink). Both these transitions can be expected to yield slower
response times, since the more diffuse nature of transitions out of this node makes them harder to anticipate. Therefore,
both these transition types are included as nuisance regressors in our models. It can be seen only one within-wing
transition is not preferred and not associated with the wide node (blue, "within"). This transition can be contrasted with
the "between" wing transition (green) to yield an unbiased estimate of higher-order surprise-based slowing.

This consistent "rotational" mapping ensured that the participants could exercise direct control over
movement between the museum wings, but not between the individual rooms. In fact, the action-
outcome contingencies were specifically designed to be policy independent, such that irrespective of
whether the participant behaved completely randomly or they consistently selected the same key
over the entire course of the experiment, the expected number of times each room was visited was

equated for all of the rooms (see S1 Appendix), as implemented previously in tasks relying on random
walks (29,30,40,41). In order to maintain the possibility to dissociate hierarchically organized
behavior from detailed model based behavior, this required constraining the action-outcome
contingencies by replacing 4 specific directed (action-dependent) transitions with a transition that
was also available for the other action (Fig. 1C). Extensive justification for these constraints, which
were controlled for statistically, are provided in S1 Appendix. Fig. 9 illustrates further how these
constraints affect the transition types within a wing. In our analyses, we included binary regressors
coded as 1 for the "preferred" transitions (orange), which are common to both <z> and <m> choices
from the outgoing room, and for the "wide node" transitions (yellow), which end in the only room
that does not have preferred outgoing transitions (and is thus the only room that has 4 possible direct
outcomes instead of just 3). Finally, we included a binary regressor coded as 1 for "wide trans" (pink),
which are less predictable outgoing transitions from this wide node.

Computational modeling

As preregistered, miniblocks that contained a "goal miss" or "wrong goal" trial (see Fig. 1D) were
excluded from the model fits. Computational models for the response times and choices were
estimated with Stan (112) using Hamiltonian Monte Carlo (HMC). All models had no divergences and
converged appropriately to a Gelman-Rubin $\hat{R}$-statistic lower than 1.01. For each model for each
participant we computed the approximate leave-one-out cross-validation score using Pareto
Smoothed Importance Sampling (PSIS-LOO) (113). PSIS-LOO has been shown to penalize complex
models more appropriately than AIC and WAIC (114). PSIS-LOO yields an estimation for the
"expected log pointwise predictive density" $\widehat{elpd}_{loo}^k$ for each model k for every participant. We use
this to compute the relative evidence for each model for every participant as Akaike-like exponential
weights $w_k = \frac{\exp(\widehat{elpd}_{loo}^k)}{\sum_{k=1}^{K} \exp(\widehat{elpd}_{loo}^k)}$ and correct them for estimation uncertainty, regularizing them further
away from 0 and 1, using the Bayesian bootstrap (115). This is easily calculated from HMC chains
sampled with Stan and referred to as *Pseudo-BMA+* weighting. Given enough data, Pseudo-BMA+
weighting should converge to 1 for the model that most closely resembles the true data generating
process. It is therefore considered an $\mathcal{M}$-closed model comparison procedure. It is important to
validate that the winning model is not only better than the other considered models, but also
describes the data adequately in the first place (116). In order to show this, we used posterior
predictive simulations of the winning model and compared these to the real data.

Participant-level model evidences were submitted to random-effects group-level Bayesian model
selection (GroupBMS) (117). This allowed us to compute a Bayesian omnibus risk (*BOR*) value, which
indicates whether any model among a set of models is more or less prevalent in the population than
others. It also allowed us to compute a protected exceedance probability (*pxp*) for each model, which
is an estimate of how likely that model is to be the most prevalent. Lastly, it allowed us to get a
participant-level estimate of the posterior model probability, indicating how likely the data from each
participant are to be generated by each of the considered computational models.

Response time model specification

1080 Response times were modeled as a regression on the mean parameter of a shifted log-normal
1081 distribution. We fitted 4 chains with 1000 warmup and 2000 post-warmup samples each, for a total
1082 of 8000 post-warmup samples. We used a treedepth of 13 and a warmup-acceptance of 0.999. The
1083 structure of the model is

$$\left(y_{ij} - ndt_{im}\right) \sim lognormal(\alpha_{im} + \boldsymbol{\beta}_{im}\boldsymbol{x}_{im}, \sigma_{im})$$

1085	$$\alpha_{im} \sim \mathcal{N}(6, 1.5)$$
1086	$$\boldsymbol{\beta}_{im} \sim \mathcal{N}(0, 0.1)$$
1087	$$\sigma_{im} \sim half\,\mathcal{N}(0,1)$$
1088	$$\gamma_{im} \sim Beta(1,1)$$
1089	$$ndt_{im} \sim \mathcal{U}(0, u_i)$$

Where *i* indexes the relevant participant, *j* the relevant trial, and *m* the relevant computational model.
y represents the response time and α represents and intercept. $\boldsymbol{\beta}$ represents a vector of regression
slopes and $\boldsymbol{x}$ a vector of independent variables. *ndt* represents the non-decision time with u_i
representing the minimum response time for participant *i*, and σ represents a variance parameter of
the log-normal distribution. Note that the discount factor γ is only part of the model-based and the
successor representation model, where it is used respectively to discount expected value directly, or
the expected future state occupations, which are entered as independent variables into the model
(see below) hence allowing for the estimation of the discount factor.

Model-based model
In order to compute the expected value for the model-based agent, we precomputed a value iteration
(3) until convergence for each of the 15 possible rooms as the goal room. The discount factor is part
of the value iteration algorithm. Since we estimate the discount factor, the value iteration was
performed within the Stan program, again for every new sample Stan would take. We assumed a
completely accurate transition model of the environment and a reward vector **r** where every entry
was a small cost value (-0.08) except for the relevant goal room, which was set to 1. This
parameterization of the cost value is irrelevant for our modeling, as all values below 0 yield exactly
collinear regressors (118).

In the regression model, we considered specifically the EV of the action that was selected on that trial.
The reward prediction error (RPE) was then computed as the EV of the current room minus the EV
of the previous room. The conflict was computed as the absolute difference in EV between the two
available actions and referred to as EV_{diff}. Since both RPE and EV_{diff} were derived from EV, there was
moderate correlation between these regressors. To control for this, we successively orthogonalized
them using the Gram-Schmidt process before fitting the model. The EV regressor was left intact. The
RPE regressor was orthogonalized with respect to the EV regressor $\boldsymbol{RPE_\perp} = \boldsymbol{RPE} - \frac{<EV,RPE>}{<EV,EV>} \boldsymbol{EV}$
where $<\boldsymbol{v_1, v_2}>$ denotes the dot product of two vectors. The EV_{diff} regressor was then orthogonalized
with respect to both the EV and the orthogonalized RPE regressors $\boldsymbol{EV_{diff\perp}} = \boldsymbol{EV_{diff}} -$
$\frac{<EV,EV_{diff}>}{<EV,EV>} \boldsymbol{EV} - \frac{<RPE_\perp,EV_{diff}>}{<RPE_\perp,RPE_\perp>} \boldsymbol{RPE_\perp}$. Regressors were standardized after orthogonalization. This
way, each regressor only captured unique variance associated with it. Significant parameter
estimates could then be directly interpreted as evidence that participants are sensitive to that
independent variable. Note that shared variance is attributed to the regressor that comes "earlier" in
the order of orthogonalization. Since these independent variables are dependent on the discount
factor and the value iteration, they were recomputed for every new sample Stan would take.
Orthogonalization and standardization thus happened within the Stan program.

Successor representation model
The successor representation agent learns a matrix **M** of expected future (discounted) state visits. If
there are *A* different actions and *S* different states in the environment, $\mathbf{M} \in \mathbb{R}^{|SA| \times |SA|}$. Each entry
$\mathbf{M}_{sa,sa\prime} = \mathbb{E}[\sum_{t=0}^{\infty} \gamma^t \mathbb{I}_{sa=sa\prime} | s_0 = s, a_0 = a]$ where *sa* indicates the previous state-action pair, and *sa'*

indicated the next state-action pair. This can be learned following a SARSA update rule. Starting **M**
out as an identity matrix for each state-action pair, updates follow:

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$$\mathbf{M}[sa_t,:] \leftarrow \mathbf{M}[sa_t,:] + \lambda(\mathbb{1}_{sa_t} + \gamma\mathbf{M}[sa_{t+1},:] - \mathbf{M}[sa_t,:])$$

where $\mathbb{1}_i$ indicates a one-hot vector with a 1 at index i. The SR introduces two parameters to estimate:
The discount factor γ and the learning rate λ. For all our model fitting, we kept the learning rate λ
fixed at 0.1 (44,97).

The state prediction error (SPE) regressor for trial t was computed as the angular distance
$\frac{1}{\pi}\arccos(\frac{v_1 \cdot v_2}{||v_1||||v_2||})$ between $\mathbf{M}[sa_t,:]$ and $\mathbf{M}[sa_{t-1},:]$. The angular distance is a formal distance metric
computed from the cosine similarity of two vectors. An angular distance metric is useful in this case,
as it is not sensitive to the magnitude of the successor representation vectors, which grow larger for
larger discount factors, but only to the relative difference in which states are predicted more or less.
We computed the EV as $\mathbf{M}[sa_t,:]\mathbf{r}$ to model proximity to the goal, where $\mathbf{r}$ is defined similarly as for
the model-based agent. RPE was the difference between EV of the current trial and the previous trial.
Conflict was the difference between EV for the two different available actions. Orthogonalization
similar to the model-based agent was applied to the EV, RPE, and conflict regressors, after which all
regressors were standardized. Note that the discount factor was part of the Stan program, and thus
M and the regressors derived from it are different for each new sample of γ that Stan would take.
Orthogonalization and standardization thus happened within the Stan program for each new sample
Stan would take.

Population-level estimate of parameters
Assuming there is a clear winner in the model comparison procedure (i.e *BOR* < 0.05), we wanted to
test whether a particular independent variable of that model m had a significant effect on the
response time in the population. It is reasonable to pool information across participants and fit a
hierarchical model for the winning model m if there is strong agreement on the optimal model in the
sample (119). A hierarchical model fits all participants simultaneously and combines information
across participants to acquire more accurate parameter estimates. However, the successor
representation model is too computationally demanding to fit simultaneously across all participants.
For this reason, we opted to first fit the models separately to each participant. We then combined the
single-participant estimates of the regression slopes into a population distribution and tested if this
was significantly different from 0. Note that this is not identical to fully hierarchically estimating the
model: Different parameters cannot constrain each other in this approach (e.g. allowing for
population-level influences on the individual regression slopes might affect the estimation of the
discount factors, but this cannot be captured by our current approach). Separately fitting models to
each participant and then combining their estimates to test for an effect is also called "two-stage
regression" (120), and is also common in e.g. functional magnetic resonance imaging (fMRI) analysis
(121).

Instead of testing point-estimates against 0 as typical in maximum-likelihood two-stage regression,
we wanted to maintain participant-level information of the width of the posterior distribution,
acknowledging the amount of uncertainty for each participant. We implemented a model analogous
to the one suggested by (122), where we summarized the participant-level posterior distributions
for each regression slope with a Gaussian parametric density estimation of the HMC samples. While
this can omit information about co-dependencies between parameters (123), it is not expected to
yield false positive results, since these (unmodeled) co-dependencies should widen the marginal

posterior distributions, effectively reducing our power. We computed posterior means $\widehat{\boldsymbol{\beta}_\iota}$, and
posterior standard deviations $\boldsymbol{\varepsilon}_i$ for each participant i, which are O-dimensional vectors with an
entry o for each independent variable. We wanted to estimate the population-level parameters $\boldsymbol{\mu}$ and
$\boldsymbol{\Sigma}$, where $\boldsymbol{\mu}$ is an O-dimensional vector of population means μ_o for each independent variable, and $\boldsymbol{\Sigma}$
is an $O \times O$ covariance matrix. We fitted 4 chains with 2000 warmup samples and 20000 post-
warmup samples each. We used a treedepth of 13 and a warmup-acceptance of 0.999. The structure
of the model is

$$\boldsymbol{\mu} \sim \mathcal{N}(0, 0.1)$$
$$\boldsymbol{\Sigma} \sim \text{diag}(\boldsymbol{\tau}) \cdot \boldsymbol{\Omega} \cdot diag(\boldsymbol{\tau})$$
$$\boldsymbol{\tau} \sim half\,\mathcal{N}(0, 0.1)$$
$$\boldsymbol{\Omega} \sim LKJ(1)$$
$$\boldsymbol{\beta}_i \sim \mathcal{N}(\boldsymbol{\mu}, \boldsymbol{\Sigma})$$
$$\widehat{\boldsymbol{\beta}}_i \sim \mathcal{N}(\boldsymbol{\beta}_i, diag(\boldsymbol{\varepsilon}_i))$$

Where LKJ stands for the Lewandowski-Kurowicka-Joe distribution, a distribution over correlation
matrix $\boldsymbol{\Omega}$, and $\boldsymbol{\tau}$ is a vector of population standard deviations for the O independent variables. The
model gives us regularized estimates of participant level regression slopes $\boldsymbol{\beta}_i$ and, of main interest,
population level effects $\boldsymbol{\mu}$. The latter were tested for significance as described in the section on
Statistical analyses.

Choice model specification
Choices were modeled based on a response-coded logistic regression, where key <z> was coded as
"0" and key <m> is coded as "1". We only considered trials outside the goal wing, since this is where
the explicit hierarchical model makes interpretable and differentiable predictions. We fitted 4 HMC
chains with 1000 warmup and 2000 post-warmup samples each, for a total of 8000 post-warmup
samples. We used a treedepth of 13 and a warmup-acceptance of 0.999. The structure of the model
is

$$y_{ij} \sim bernoulli\big(\alpha_{im} + \beta_{im} x_{ijm}\big)$$
$$\alpha_{im} \sim \mathcal{N}(0,2)$$
$$\beta_{im} \sim \mathcal{N}(0,2)$$
$$\gamma_{im} \sim Beta(1,1)$$

Where i indexes the participant, j the relevant trial, and m the relevant computational model. Note
that the discount factor γ is only part of the model-based and the successor representation model.
We only analyzed choices outside the goal wing, since the hierarchical structure of the task only
informs choices leading to the goal wing, and not inside that wing. We were specifically interested in
whether participants use that hierarchical structure to guide their choices.

For the logistic regression, the null-model only contains an intercept, so in actuality there was no β
parameter for this model. It only models bias to select either <z> or <m>. The three cognitive models
all contained one regressor (besides an intercept). For the model-based agent, we computed the
expected value of the action <m> minus the value of action <z>, and standardized this within-
participant. For the explicit hierarchical model, we computed an effect-coded regressor which was -
0.5 when <z> was the current "correct rotation" and 0.5 when <m> was the current correct rotation.
For the successor representation agent, we computed the expected value difference of <m> minus
<z>, and standardized this within-participant.

 Population-level estimate of parameters
In case one model clearly wins the model comparison (i.e. *BOR* < 0.05), we tested whether the
relevant regression slopes differ from 0 in the population. We tested this hypothesis with a
hierarchical logistic regression implemented in *brms* (124). The prior for the population-level effect
was left at the default, being a student-t distribution with 3 degrees of freedom, mean 0, and scale
2.5. The model included separate participant-level intercepts and slopes. We fitted 4 HMC chains each
with 1000 warmup and 5000 post-warmup samples each, for a total of 20000 post-warmup samples.

Posterior predictive checks
For the posterior predictions of choice and response switches (Fig. 2) and of the response times (Fig.
4,5), we draw one posterior prediction per trial for every sample accepted by HMC. In every case, this
leads to 8000 posterior predictions per trial. For the response times, the posterior predictions were
computed from the participant-level model estimation, again drawing one posterior prediction per
trial for every accepted sample. For each trial, we took the mean of the 8000 posterior predictions as
the mean posterior prediction for that trial. Separately, we assigned each trial regressor values based
on the successor representation model, based on the discount factor of the maximum a-posteriori
(MAP) estimate of the model. This means uncertainty about the discount factor is not taken into
account in our posterior predictive check. This was necessary to overcome limitations in RAM.

Statistical Analyses
Where we mention the use of multiple comparison correction, used in the binary model comparisons
and in the exploratory individual difference analysis, we multiplied the reported *BOR-* or *p*-values
according to the Holm-Bonferroni procedure before reporting them (125). For the significance
testing of population-level parameters in hierarchical Bayesian linear models fitted for the winning
choice and response time computational models, we report the posterior mean (*M*), the 95% highest
density interval ($HDI_{95\%}$), and the evidence ratio (*ER*) of every parameter (126). $HDI_{95\%}$ corresponds
to the smallest continuous interval that includes 95% of the total posterior density. If 0 is not included
in this interval, this is suggestive of a meaningful effect of that independent variable with respect to
the dependent variable. Additionally, we report the evidence ratio in the favored direction (for
positive effects ER_+, for negative effects ER_-). In the case of ER_+, this is the ratio of the posterior density
above 0 over the density below 0. ER_- is the inverse of this. If 95% of the posterior density lies
exclusively above or below 0, the corresponding *ER* will be higher than 19. *ER* will be infinitely large
if all HMC samples lie either exclusively above or exclusively below 0. We consider an effect to be
reliable when both the $HDI_{95\%}$ excludes 0, and the relevant *ER* is larger than 19. The computation of
the *ER* and the definition of a cutoff value provide a 1:1 correspondence with frequentist p-values
(127).

Response switches
Response switching behavior was modeled with *brms* and the same prior structure as the population
Choice model. The outcome variable was binary coded, with response repetitions coded as "0" and
response switches indicating the participant responded different on trial *t* than on trial *t-1* coded as
"1", only considering trials that are part of the same miniblock. Entry and leave regressors were
dummy coded with "1" indicating a trial immediately after the participant entered or left the goal
wing. We also included a nuisance regressor that counted the (log-transformed) number of times an
action has been repeated successively, which was standardized within-participant.

Rotation and antirotation
We modeled the probability of selecting the correct rotation with *brms* and the same prior structure
as the population Choice model. The outcome variable was binary coded, with selection of the correct
rotation coded as "1" and selection of the antirotation coded as "0". A regressor corresponding to
whether the correct rotation was currently <z> or <m> was added, coding <z> as -0.5 and <m> as 0.5.
A second regressor was added, considering whether the current room implied the rotation was in
fact optimal (blue, Fig. 2F; coded as 0.5) or whether the antirotation was in fact optimal (orange, Fig.
2F; coded as -0.5). We also included their interaction, to investigate whether the latter effect is
especially present for either <z> or <m> rotations.

Free sort
We modeled the Euclidean distances between the 105 pairs of paintings sorted by the participants
using *brms* and the same prior structure as the population Choice model. First we standardized the
Euclidean distances within participant, to correct for individual differences in participants who
would use the available grid up to a different width (i.e. some place most paintings near the center,
some place paintings all the way out in the corners). We dummy-coded three possible relationships
between pairs of paintings. We defined a "community" regressor which was 1 when two paintings
were part of the same community, and 0 otherwise. We also defined a "bridge" regressor which was
1 when two paintings were connected between two communities (i.e. not part of the same
community) and 0 otherwise. Additionally, we defined a "boundary" regressor which was 1 for
paintings that were part of the same community but not connected (i.e. the two boundary paintings
within the same wing) and 0 otherwise. Note that for these paintings the "community" regressor is
also 1, meaning this boundary regressor models whether participants place the boundary paintings
further apart than other same-community paintings. Paintings that were neither part of the same
community, nor had any direct connection between them, were captured by the intercept of the
model. In order to inspect bias induced by the community structure of the environment, we tested
whether participants placed connected paintings from the same wing (community) closer together
than connected paintings from different wings (bridge), by subtracting the "bridge" samples from the
"community" samples and testing whether the new chain was significantly below 0. We ran an
additional hierarchical Bayesian linear regression considering only pairs of paintings that had a
shortest path length of 1 intermediate room between them. This allowed us to test whether the
community-boundary paintings were placed closer together than other paintings with the same
distance between them. This model consisted of only the "community" regressor and an intercept.

Modularity measure
The successor representation can represent the community structure present in the task more or
less strongly dependent on the value of the discount factor parameter. Our work differs from
previous modeling work on community structure (29) in three important ways. (i) We do not
standardize the successor matrix to a transition probability matrix after each new experience (see
discussion), (ii) we model sequences of state-action conjunctions instead of states, and (iii) some
states can succeed specific states more often than others ("preferred" transitions, Fig. 9). We
designed a measure analogous to the one reported by (29), but addresses the additional complexities
of our task design. Our measure is based on the notion that, by definition of modularity, the successor
representation prediction error should be higher for transitions between wings than for transitions
within wings. However, since some transitions are "preferred" (see section "Action-outcome
mapping"), these are expected to yield lower prediction errors, thus inflating our measure of
community structure (see S1 Appendix). We thus compared specific non-preferred within-wing
transitions ("within", blue; Fig. 9) to between-wing transitions which are always non-preferred

("between", green; Fig. 9). These transitions are topologically identical except for their distinction as
between- or within-wing, and thus provide an unbiased estimate of the sensitivity to community
structure. We defined the "modularity" measurement as the ratio between the mean angular distance
of the successor representation vectors (the successor representation state prediction error) for all
possible between-wing transitions and these non-preferred within-wing transitions. Note that this
yields a "difference of distances" measurement analogous to representational similarity analysis
techniques (128).

We computed the modularity measurement for each participant using the successor representation
matrix after the testing phase. Note that we obtained a full posterior distribution over the discount
factor γ, and thus over successor matrices. We reduced this to a point-estimate for the modularity by
computing the expectation of modularity over the discount factor. To test for evidence of modularity
at the population level, we first computed the expected modularity under the uniform prior
distribution of the discount factor $\pi(\gamma) \sim Beta(1,1)$. Because each participant made different choices
and experienced a different sequence of states, there are variations in the relationship between the
discount factor and the modularity (see Fig 7A,B), so this provides a proper "null hypothesis" for the
modularity of each specific participant. We then computed the expected modularity under the
recovered posterior distributions $p(\gamma|y)$ for each participant. We compared these prior and
posterior expectations using a paired-samples t-test.

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