Polygenic architecture of flowering time and its relationship

with local environments in the grass *Brachypodium distach-*

yon.

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Abstract

Synchronizing the timing of reproduction with the environment is crucial in the wild.

Among the multiple mechanisms annual plants evolved to sense their environment,

the requirement of cold-mediated vernalization is a major process that prevents

individuals from flowering during winter. In many annual plants including crops, both

a long and short vernalization requirement can be observed within species, resulting

in so-called early-(spring) and late (winter)-flowering genotypes. Here, using the grass

model *Brachypodium distachyon*, we explored the link between flowering time-related

trait (vernalization requirement and flowering time), environmental variation, and

diversity at flowering-time genes by combining measurements under greenhouse and

outdoor conditions. These experiments confirmed that *B. distachyon* natural

accessions display large differences regarding vernalization requirements and

ultimately flowering time. We underline significant, albeit quantitative effects of

current environmental conditions on flowering time-related traits. While disentangling

the confounding effects of population structure on flowering time-related traits

remains challenging, population genomics analyses indicate that well-characterized

flowering-time genes may contribute significantly to flowering time variation and

display signs of polygenic selection. Flowering-time genes, however, do not colocalize

with GWAs peaks obtained with outdoor measurements, suggesting that additional

genetic factors contribute to flowering time variation in the wild. Altogether, our study

fosters our understanding of the polygenic architecture of flowering time in a natural

grass system and opens new avenues of research to investigate the gene-by-

environment interaction at play for this trait.

Key words: Flowering time, adaptation, vernalization, *B. distachyon*, grasses, polygenic

selection.

Article Summary

Synchronizing the timing of reproduction with the environment is crucial in the wild.

We used here the model *Brachypodium distachyon* to expand our knowledge on the

adaptive potential and polygenic architecture of flowering time in wild grasses. While

genetic factors play an important role in flowering time variation in our species, we

showed that flowering time correlates with environmental variables and is therefore a

locally adapted trait.

Introduction

The induction of reproduction is a critical fitness-related trait in the wild (Gaudinier and Blackman 2020), as a failure to produce offspring during the growing season may lead to the extinction of the individual's genotype. In many annual plant species adapted to temperate climates, plantlets establish themselves in the fall and overwinter before flowering and producing seeds in more favorable spring conditions (Chouard 1960; Blackman 2017). Vernalization, the prolonged exposure to cold necessary to render plants competent to flower (Chouard 1960), is hence a key component of plant reproduction as it prevents individuals from flowering prior to winter. The adaptive potential and the genetic architecture of flowering time has been studied in an unrivaled manner in *Arabidopsis thaliana* (for review Andrés and Coupland 2012; Blümel et al. 2015; Takou et al. 2019) due to the broad geographical distribution of the species and large genomic resources developed by the community (but see Hall et al. 2006; Monnahan and Kelly 2017; Yan et al. 2021 for works on other *Brassicaceae* and *Mimulus guttatus*). Vernalization is yet controlled by different genes in different plant groups and likely evolved independently multiple times during flowering plant diversification (Ream et al. 2012; Bouché et al., 2017; Raissig and Woods 2022). Moreover, specific crop flowering-time genes (e.g. *ID1* in maize) are for instance lacking homologs in *A. thaliana* (Blümel, Dally, and Jung 2015). As a wild monocot, the model for the temperate grasses *Brachypodium distachyon* constitute a prime system to study the evolution of flowering-time genes. In this context, we made use of the diversity panel developed for this species (Gordon et al. 2017; Gordon et al. 2020; Skalska et al. 2020;

Stritt et al. 2022, Minadakis et al. 2023) to expand our knowledge on the adaptive

potential and polygenic architecture of flowering time in grasses.

Initially established as a model for bioenergy crops (International Brachypodium

Initiative 2010), the grass species *B. distachyon* has more recently become a prime

model for developmental biology (Woods et al. 2017[a]; Woods et al. 2019; Nunes et al.

2020; Hasterok et al. 2022; Raissig and Woods 2022; Zhang et al. 2022; Slawinska et al.

2023), evolutionary genomics (Bourgeois et al. 2018; Gordon et al. 2020; Gordon,

Contreras-moreira, et al. 2017; Stritt et al. 2018; 2020; Eichten et al. 2016) and molecular

ecology (Del'Acqua et al. 2014; Wilson et al. 2019; Skalska et al. 2020; Stritt et al. 2022;

Minadakis et al. 2023). In addition to a chromosome-level genome assembly of 272

Mb (International Brachypodium Initiative 2010), a diversity panel composed of 332

accessions spanning from Spain to Iraq has been sequenced (Gordon et al. 2017;

Gordon et al. 2020; Skalska et al. 2020; Stritt et al. 2022, Minadakis et al. 2023), opening

new avenues of research in this system (Minadakis et al. 2023). We previously showed

that *B. distachyon* accessions cluster into three main genetic lineages (A, B and C),

which further divide into five main genetic clades: the ancestral C clade in Italy and

Balkans, the B_West clade in Spain and France, the B_East clade spanning from Turkey

to Caucasus and Iraq, the A_Italia clade in Italy as well as the A_East clade in Turkey

and Greece (Stritt et al. 2022, Minadakis et al. 2023). These natural accessions are found

in diverse habitats (Bourgeois et al. 2018; Minadakis et al. 2023) making *B. distachyon*

an ideal model to investigate how genetic and environmental factors interact to shape

traits.

B. distachyon accessions display large phenotypic variation with regard to flowering

time (e.g. Ream et al. 2014; Gordon et al. 2017; Woods et al., 2019; Sharma et al. 2017),

with some accessions requiring little to no vernalization to flower rapidly (early flow-

ering accessions) in certain photoperiods, while other accessions require a few weeks

to several months of vernalization in order to flower (late flowering accessions) (Ream

et al., 2014; Gordon et al., 2017; Woods et al., 2019). These flowering differences have

been described as potentially adaptive and responsible for population distribution ac-

cording to climate variation (Gordon et al. 2017; Woods et al. 2019; Skalska et al. 2020).

However, the extent to which variation in vernalization requirement and ultimately

flowering time correlates with local environmental conditions has yet to be formally

tested in this species. For instance, whether late-flowering genotypes, which require

long vernalization treatments, have been selected to complete their life cycle at a

slower rate to overcome harsher or longer winter, as observed in Swedish populations

of *A. thaliana* (Ågren et al. 2017)*,* remains an open question.

In a Mediterranean species like *B. distachyon*, not all accessions might be exposed

to harsh winter and short cold periods might be sufficient to allow permissive flowering

in the wild. In this study, we hence decomposed flowering time into four flowering

time-related traits: the minimum threshold duration (MTD) of vernalization necessary

to be permissive for flowering, the number of day to flower when MTD is received, the

saturating threshold duration of vernalization above which no further reduction in

flowering time is gained and the fastest time for a given accession to flower from the

day plants after vernalization saturation. We combined measurements under

greenhouse and outdoor conditions and asked i) Do flowering time-related traits cor-

relate with environmental variables and show signs of local adaptation? ii) What is the

respective contribution of flowering time genes to flowering time-related trait varia-

tion? iii) Are known flowering-time genes contributing to flowering time variation in

the wild?

Materials and Methods

Biological materials and genomic resources

The *Brachypodium distachyon* diversity panel is composed of 332 natural accessions

for which whole-genome sequencing data are publicly available (Gordon et al. 2020;

Gordon, Contreras-moreira, et al. 2017; Skalska et al. 2020; Stritt et al. 2022). For the

flowering time experiment, we selected a subset of 61 accessions representing all five

genetic clades as first described by Stritt et al. 2022 and occurring along latitudinal

gradiants. Maps were drawn QGIS (version 3.16).

We also made use of the raw vcf produced by Minadakis et al. (2023) for the entire

diversity panel. We used vcftools (Danecek et al. 2011) to apply the following filtering

criteria: --max-alleles 2 --max-missing-count 200 --minQ 20. We further filtered het-

erozygous SNPs as those have been shown to result from duplicated sequenced and

be mostly artifactual in selfing species (Stritt et al. 2022). All the analyses were per-

formed   on   version3   of   the   *B.   distachyon*   genome   (https://phytozome-

next.jgi.doe.gov).

Flowering time-related trait measurements

We performed an experiment in controled conditions from October 2021 until May 2022 in order to test the flowering phenology of the five genetic clades of *Brachypodium distachyon.* Twelve accessions per genetic clade were selected, in addition to the reference accession Bd21. All accessions were treated with five vernalization periods spanning from 2 to 10 weeks, with three replicates per treatment and per accession. Seeds were stratified for at least two weeks before the experiment, and then sowed in pots that were placed in greenhouse conditions (16 h day at 20°C and 8 h dark at 18°C with a light intensity of 200 µMol/m2/s). We distributed the replicates randomly across trays to minimize bias due to position effects. Three weeks after germination the plants were transfered to a cooling chamber (constant temperature at 4°C, 8h light 80 µMol/m2/s, and 16 h dark) for 2, 4, 6, 8, and 10 weeks. At the end of the vernalization treatment, plants were moved back to the greenhouse.

Flowering time was measured as the number of days after return to greenhouse to the first day of spike emergence which is consistent with stage 50 of the Zadoks scale that was used in (Ream et al. 2014). Measurements were taken every two days until the end of the experiment in May. During the experiments, trays were permuted on the table to further limit position effects.

We then extracted from this experiment four flowering time-related traits: i) the minimum threshold duration of vernalization necessary to be permissive for flowering, ii) the number of day to flower when the minimum threshold is received, iii) the saturating threshold duration of vernalization above which no further reduction in

flowering time is gained and, iv) the fastest time for a given accession to flower from

the day plants were moved to the vernalization chamber, i.e. the number of day to

flower when vernalization saturation is received. This latter trait is classically used as a

measure of flowering time in *B. distachyon* studies. Results were plotted in R version

4.0.2 (R Core Team, 2018) with the package ggplot2 (Wickham 2016).

Extraction of current and past bioclimatic variables

Raster maps for current monthly solar radiation and altitude were retrieved from

worldclim (https://www.worldclim.org). In addition, raster maps for monthly Global

Aridity Index (GAI) were obtained from https://cgiarcsi.community/data/global-arid-

ity-and-pet-database/. Bioclimatic variables were then extracted using the R packages

raster (v.3.5-2, Hijmans & van Etten, 2012) and rgdal (v.1.5-27, Keitt et al., 2010) for

each of the 332 accessions. For solar radiation and GAI, data were also average over

spring months (April to June).

For paleo-bioclimatic variables, we used the niche suitability projections for the last

glacial maximum (LGM) computed by Minadakis et al (2022). For each genetic clade,

we extracted the coordinates of a set of 200 random points per clade in highly suitable

habitats (>0.85) with the raster package function rasterToPoints. We retrieved raster

maps for LGM from https://www.worldclim.org/data/v1.4/paleo1.4.html and extracted

the 19 paleo-bioclimatic variables for the corresponding sites as described above.

Redundancy analysis (RDA)

We extracted the 19 classical bioclimatic variables as well as elevation, aridity and solar radiation in spring for the 56 accessions for which flowering time-related trait were measured. We performed a principal component analysis (PCA) with the R base function prcomp using the resulting 22 variables.

Association between flowering time-related traits, bioclimatic variables and genetic cluster of origin were tested with redundancy analysis (RDA) using the R package Vegan 2.6-4 (O'Hara et al. 2022) following (Capblancq and Forester 2023). RDA computes axes that are linear combinations of the explanatory variables. In other words, this method seeks, in successive order, a series of linear combinations of the explanatory variables that best explain the variation of the response matrix (Borcard, Gillet and Legendre 2018). It is therefore especially suited when explanatory variables are correlated.

To identify which variable influence flowering time-related traits, we opted for a forward selection approach (Capblancq and Forester 2023) and first ran empty models where the respective flowering time-related trait were treated independently as a response variable and explained by a fixed intercept alone. We then ran additional models where each flowering time-related traits were treated independently as a response variable while the cluster of origin, the 22 environmental variables mentioned above or both the cluster of origin and the 22 environmental variables were successively entered as explanatory variables. These full models were compared to their respective

empty models with the Vegan ordiR2step function (Capblancq and Forester 2023) and

the following parameters: permutations = 1000, R2scope = T, Pin = 0.01.

We also tested for associations between flowering time and SNPs in known flower-

ing-time genes using RDA. To do so, we first extracted and converted the SNPs located

in flowering-time genes as a dataframe with vcftools -- bed --extract-FORMAT-info

GT, where the bed file contained the position of flowering-time genes. We first fitted

simple models where each flowering time-related traits were treated independently as

a response variable and all SNPs individually as explanatory variables. here again, we

compared those models with their respective empty models with the Vegan

ordiR2step function. For data visualization, the SNPs were pruned to remove SNPs in

perfect linkage disequilibrium within each gene with plink (Purcell et al. 2007) with the

following paramters --indep-pairwise 50 5 1. We kept one focal SNP per gene display-

ing the strongest association with one of the flowering time-related traits for further

analysis. In a second step, we fitted models where all focal SNPs were treated as ex-

planatory variables together. We eventually fitted full model where the cluster of origin

together with all focal SNPs and bioclimatic variable associated with the given flower-

ing time-related trait were treated as explanatory variables. All the plots were pro-

duced with the R package ggplot2.

F_{ST} calculation

Single SNP F_{ST} between accessions of the A and B lineages were calculated with

vcftools (Danecek et al. 2011). To account for shifts in the observed F_{ST} values caused

by the population structure of *B. distachyon*, the expected F_{ST} distribution under neu-

tral evolution was also estimated using forward simulations run in SLiM version 3.4

(Haller and Messer 2019). The population structure, effective population sizes and time

of divergences between lineages of *B. distachyon* during its evolution were modeled

in SLiM based on the results of Minadakis et al. (2023). No migration between the

different populations in the simulations was allowed, as a lack of interbreeding be-

tween the distinct *B. distachyon* clades was reported (Stritt et al. 2022). The simulation,

was run 100 times and single SNP F_{ST} were calculated for each simulation to generate

the expected F_{ST} distribution under neutrality.

Linkage disequilibrium analyses

To plot linkage disequilibrium (LD) decay, we first thinned the vcf with vcftools --thin

20000 to keep one SNP every 20 kilobases (kb). Intrachromosomal LD (r^2) was calcu-

lated with vcftools --geno-r2. We repeated this step by further filtering the vcf for a

minimum allele frequency of 0.05 with vcftools -- maf 0.05. For both outputs, we vis-

ualized LD decay by plotting r^2 and its 95 CI as a function of the physical distance

between SNPs with the R package ggplot2. LD between focal SNPs in flowering-time

genes was calculated separately with vcftools –geno-r2 and added to the LD decay

plot for comparison. LD plots for focal SNPs were produced with the R package gaston

(Pedry and Dandine-Roulland 2020).

Inter-chromosomal LD (r^2) was calculated with vcftools --interchrom-geno-r2 for a

subset of 100,000 loci selected randomly in the genome. We then re-sampled 50,000

times different combination of loci and calculated the mean LD each time. This allowed

us to further calculate the CI around the mean and compare our real data to this dis-

tribution.

Scans of selection

The $X^T X$ analysis was performed with BayPass v2.3 using the five genetic clades or three

genetic lineages as populations (Gautier 2015). We generated the input file by using

vcftools --count to calculate the allele frequency of each SNP present in our vcf (no

filtering on minor minimum allele frequency). We then ran Baypass on our actual da-

taset with the following parameters: -pilotlength 500 -npilot 15 -burnin 2500 -nthreads

6. To calibrate the $X^T X$ and define a threshold of significance for differentiated SNPs,

we then used the simulate.baypass function from baypass_utils (Gautier et al. 2015) to

generate a pseudo-observed dataset (POD) of 100,000 loci based on the co-variance

matrix computed with our real dataset. We then ran Baypass on the POD with the

parameters described above. We used the 0.99 quantile of the $X^T X$ calculated for the

POD as a threshold of significance for the real dataset. Integrated Haplotype Scores

(IHS) were also computed for accessions of the A and B lineages separately with the R

package Rehh 3.1 (Gautier et al. 2017)

Functional effect and age estimates of variants

The functional effect of variants was annotated using SnpEff version 5.0e (Cingolani et

al. 2012) using default parameters and the provided database for *Brachypodium*

distachyon.

The age of each single SNP was computed with GEVA (Albers and McVean 2020) to

estimate the average SNP age for each annotated gene (https://phytozome-

next.jgi.doe.gov) in the derived A and B genetic lineages as well as in the five genetic

clades. All private SNPs to the combined A and B lineages were polarized using the

ancestral C lineage using custom R scripts. GEVA was run on the five main scaffolds

(corresponding to the five chromosomes) using the genetic map produced by (Huo et

al. 2011) and the polarized SNP dataset.

Genome-wide association analysis

We further used the flowering time measurements performed outdoor in Zurich in

2017 (Stritt et al. 2022). We extracted the 19 classical bioclimatic variables as well as

altitude, aridity and solar radiation for the 332 accessions as well as a site in Zurich (lat

= 43.73693085, lon = 3.69295907) and performed a principal component analysis

(PCA) as described above. Linear-mixed model analyses were performed as described

above for the greenhouse experiment.

To identify loci associated with flowering time variation, we performed a genome-

wide association analysis with GEMMA (Zhou and Stephens 2012). Due to population

structure, we corrected the analysis by a centered relatedness matrix, first produced

with the option -gk 1. Association tests were performed using the option -maf 0.05 to

exclude SNPs with minor allele frequency with values less than 0.05. Regions were con-

sidered significantly associated if displaying at least four markers above FDR threshold

in 8Kb windows (overlap of 4Kb). Upset plots were drawn with the R package UpSetR

(Gehlenborg 2019).

Overlaps with Environmental association analyses

We made use of the environmental association analyses performed by Minadakis et al.

(2023) to assess whether flowering-time genes and the candidate gene identified by

the GWAs were associated with current environmental variables. Upset plots were

drawn in R with the R package UpSetR (Gehlenborg 2019).

Gene duplication

We   checked   for   potential   gene   duplication   with   detettore

(https://github.com/cstritt/detettore), a program developed to study structural varia-

tion based on short-read sequences (Stritt 2021). We also calculated the proportion of

heterozygous sites over the 22 flowering-time genes and 332 accessions, using the

raw vcf produced by Minadakis et al. (2023) not filtered for heterozygous SNPs but

with the following criteria: --max-alleles 2 --max-missing-count 200 --minQ

Results

Flowering time measurements under greenhouse conditions

We selected 61 accessions (Figure 1A) from the *B. distachyon* diversity panel (Minada-kis et al. 2023) for our flowering time experiment. Those accessions were chosen to represent all five genetic clades and occur, when possible, along latitudinal gradients. Briefly, we submitted plants to five vernalization treatments (2, 4, 6, 8 or 10 weeks at 4°C) with three replicates each and measured how long plants took to flower after the return to warm conditions (Figure S1, Table S1 for the raw data). For five out of the 61 accessions (Veg12 from A_East; Ren4, Lb1, Lb23 and Cm7 from A_Italia), none of the replicates flowered by the end of the experiment (Figure 1a, Figure S1) despite normal growth. All subsequent analyses were hence performed on the 56 accessions for which flowering time data were collected.

However, *B. distachyon* accessions span a large range of habitats (Minadakis et al. 2023). While some accessions of the diversity panel can experience up to -11 °C as a minimum temperature during the coldest months (Bio6), the mean temperature during the coldest quarters (Bio11) is above 4 °C for a large number of accessions (Table S2, Figure S2). As such, not all accessions may experience long vernalization at 4 °C in the wild. We thus extracted four traits link to vernalization requirements and *in fine* flowering time (Table S3) for further analyses: i) the minimum threshold duration of vernalization necessary to be permissive for flowering (hereafter MTD), ii) the number of day to flower when MTD is received (hereafter day to flower after MTD), iii) the saturating threshold duration of vernalization above which no further reduction in

flowering time is gained (hereafter vernalization saturation) and, iv) the fastest time for

a given accession to flower from the day plants were moved to the vernalization cham-

ber, i.e. the number of day to flower when vernalization saturation is received (hereaf-

ter day to flower after vernalization saturation). Note that all plants germinated within

5 days and germination time was therefore not included for further analyses.

Flowering-related traits show significant albeit moderate correlation among each

other (Figure 1b) and, with the exception of day to flower after MTD, a partitioning of

the phenotypes per genetic clades (Figure 1c and Table S4 for P-values). Accessions

from the C, B_East and B_West clades show a significantly reduced MTD compared to

accessions from the A_East and A_Italia clades. Accessions from the B_East and B_West

Fig. 1: Geographical origins of the samples used for greenhouse experiment and flowering time-related trait variation. (a) Map displaying the location of a given accession as well as their genetic clade of origin (C in yellow, A_East in magenta; A_Italy in turquoise; B_East in green and B_West in dark blue). Accessions that did not flower by the end of the experiment are depicted in black. (b) Correlogram of the four flowering time-related traits among accessions (Pearson correlation assuming linear relashionship). DFT: day to flower; MTD: minimum threshold duration of vernalization (c) Schematic phylogeny and distribution of the four flowering time-related traits.

also tend to show a shorter saturating threshold duration of vernalization compared

to accessions from the three other clades (Figure 1c and Table S4). Eventually, acces-

sions from the C, B_East and B_West clades flower significantly faster when vernaliza-

tion is saturated (Figure 1b and Table S4). Altogether, these results reflect that acces-

sions from the A lineage display overall a longer life cycle (late-flowering genotypes)

than the ones from the B and C lineages (early-flowering genotypes) due to longer

vernalization time requirements but longer time to flower after the return to warm

conditions (Figure S1).

Association between flowering time-related traits and bioclimatic variables.

To test whether flowering time-related traits correlate with environmental variation,

we extracted the 19 classical worldclim bioclimatic variables (Bio1 to Bio19) as well as

solar radiation in spring, global aridity index in spring and altitude for each locality and

performed a PCA with the resulting 22 variables (Table S3). The first two axes of the

PCA explained together about 58% of the variation among our samples, indicating that

our set of selected accessions occur in different environmental conditions.

We have previously shown that the five *B. distachyon* occur in different ecological

niches (Minadakis et al. 2023) leading to a confounding effect of population structure

and environmental variation. As mentioned above, we also observed in the current

study some partitioning of the phenotypes per genetic clades for the flowering time-

related traits we investigated. Eventually, most bioclimatic variables display correlation

among each other in our system (Minadakis et al. 2023, Figure S3). We therefore used

a redundancy analysis (RDA) to test for associations between flowering time-related

traits and bioclimatics variables, as this approach typically accounts for confounding

factors by maximizing the genetic variance explained by a set of environmental pre-

dictors.

Specifically, to model the relationships among flowering time traits, bioclimatic var-

iable and population structure, we fitted models where flowering time-related traits

were independently entered as response variable while the cluster of origin

(model = cluster alone), the 22 environmental variables (model = all bioclimatic varia-

bles) or both the cluster of origin and the 22 bioclimatic variables (model = cluster +

all bioclimatic variables) were successively entered as explanatory variables. We com-

pared those full-models to empty models where the respective flowering time-related

trait was explained by a fixed intercept alone (Capblancq and Forester 2023).

Consistent with the phenotype distributions (Figure 1b), the genetic cluster of origin

alone contribute to MTD (R^2 = 0.25), day to flower after vernalization saturation (R^2 =

0.44) and vernalization saturation (R^2 = 0.1) but not to day to flower after MTD (Table1).

We found that a large part of the variance in flowering time-related traits is explained

by bioclimatic variables (precipitation levels and temperature) linked to warm months

but only marginally by elevation or variables linked to cold month (e.g. bio19; Table 1)

when the models did not include the cluster of origin. Eventually, environmental (bio-

climatic variables) still explained a significant part of the variance in flowering time-

Table1: RDA output for each flowering time-related trait and outdoor experiment

Trait	Model	Variable	Cumulative R2adj	AIC	F	P-value
MTD	cluster alone	cluster	0.25	248.18	11.35	0.002
	All bioclim variables	bio18	0.18	256.40	28.29	0.002
		bio16	0.22	250.64	7.81	0.006
		bio13	0.28	243.22	9.47	0.004
		all variables	0.45			
	Cluster + all bioclim variables	cluster	0.25	248.18	11.35	0.002
		bio18	0.42	218.35	34.53	0.002
		elevation	0.46	210.31	9.88	0.002
		all variables	0.58			
Day to flower after MTD	cluster alone	cluster	-	-	-	NS
	All bioclim variables	aridity_spring	0.26	15.07	45.31	0.002
		bio9	0.30	10.32	6.77	0.006
		bio5	0.33	5.08	7.21	0.01
		bio18	0.37	-0.46	7.46	0.006
		all variables	0.52			
Vernalization saturation	cluster alone	cluster	0.10	126.14	4.31	0.004
	All bioclim variables	aridity_spring	0.22	105.17	35.61	0.002
		elevation	0.27	98.59	8.66	0.004
		all variables	0.42			
	Cluster + all bioclim variables	aridity_spring	0.22	105.17	35.61	0.002
		elevation	0.27	98.59	8.66	0.006
		all variables	0.44			
Day to flower after veranlization saturation	cluster alone	cluster	0.44	108.19	25.53	0.002
	All bioclim variables	aridity_spring	0.31	132.17	56.02	0.002
		bio16	0.39	116.82	18.17	0.002
		all variables	0.54			
	Cluster + all bioclim variables	cluster	0.44	108.19	25.53	0.002
		aridity_spring	0.60	68.74	46.84	0.002
		bio13	0.65	51.19	19.98	0.002
		bio16	0.68	42.35	10.60	0.002
		srad_spring	0.69	37.56	6.47	0.004
		all variables	0.78			
Day to flower (outdoor experiment)	cluster alone	cluster	0.38	2347.00	93.32	0.002
	All bioclim variables	bio18	0.42	2309.20	428.44	0.002
		bio12	0.45	2276.00	36.18	0.002
		bio19	0.51	2214.60	66.43	0.002
		all variables	0.60			
	Cluster + all bioclim variables	bio18	0.42	2309.20	428.44	0.002
		cluster	0.48	2243.40	19.46	0.002
		bio12	0.53	2188.60	58.93	0.002
		bio19	0.56	2146.50	45.21	0.002
		all variables	0.66			

bio18 : Precipitation of Warmest Quarter ; bio9: Mean Temperature of Driest Quarter; bio5: Max Temperature of Warmest Month; bio13: Precipitation of Wettest Month; bio16: Precipitation of Wettest quarter; bio12 : Annual Precipitation; bio19 : Precipitation of Coldest Quarter. We only display predictors that increase the cumulative R^2 by more than 0.03 while "all variable" display the cumulative R^2 explained by all significant predictors together.

related traits (Table 1) even in the full model including also the genetic clade of origin

(model clutser + all bioclimatic variables). This indicates that flowering time-related

traits, when measured under greenhouse conditions, are significantly shaped by the

current environmental conditions we tested in *B. distachyon* and are locally adapted

trait.

Note that variable selection optimizes the variance explained by explanatory varia-

bles but does not necessarily identify the ecological drivers of phenotypic variation. In

our system, where bioclimatic variables are highly correlated with each other and

sometimes with the genetic clade of origin (Figure S2), while one variable may maxim-

ize the variance, another one maybe the mechanistic driver of variation (Capblancq

and Forester 2023). As such, it remains difficult to disentangle the effect of the cluster

of origin from the effect of the environment for MTD and day to flower after vernali-

zation saturation (Table 1) and to estimate the actual part of the variance explained by

the environment for those two traits.

It is yet clear that accessions from more arid environments tend to flower faster

once vernalization saturation has been reached. Indeed, kendall rank correlations be-

tween flowering time and aridity in spring or annual precipitation levels (bio12) for

instance show a significant (P-value = 7.143e-06 and 4.982e-11 respectively) but only

partial correlations (Tau = -0.25 and 0.36 respectively), implying that while local envi-

ronment are driving flowering time variation and flowering time-related trait at large

(Table 1), it only partly explains the early/late-flowering partitioning of the phenotypes

we observe among genetic lineages (Figure 1c).

Contribution of flowering time genes to flowering time-related trait variation

The genetic basis of flowering time has been extensively characterized at the molecular level in *B. distachyon* (e.g. Raissig and Woods, 2022; Woods et al., 2023). In a second step, we also aimed to assess the contribution of genetic factors to flowering time-related traits. Due to our relatively small sample size (56 accessions with phenotypes), we opted for a targeted approach rather than a classical GWAs and selected 22 flowering-time genes (Table S5) molecularly characterized and described as impacting flowering time in our study system (Higgins et al. 2010; Wu et al. 2013; Woods et al. 2014; Sharma et al. 2017; Woods et al. 2017[a,b]; Lomax et al. 2018; Qin et al. 2019; Woods et al. 2019; Cao et al. 2020; Kennedy and Geuten 2020; Woods et al. 2020).

Using the SNP calling performed by Minadakis et al. (2023), we extracted 502 SNPs across these 22 flowering-time genes and 56 accessions and performed stepwise RDA. We first fitted models where flowering time-related traits were independently entered as response variable and each of the 502 SNPs were entered independently as explanatory variables. Here again, those models were compared to empty models where the respective flowering time-related trait was explained by a fixed intercept alone. 302 SNPs across 20 flowering-time genes were significantly associated with at least one of the four flowering time-related traits (Table S6). Due to linkage disequilibrium over short distances (Figure S4a), many SNPs gave precisely the same signal of association within a given gene (Table S6). We thus pruned our dataset. Overall, the resulting 143 SNPs show a strong association with day to flower after vernalization saturation and MTD but very mild association with the two other traits (Figure 2a). Individual SNP

explained as much as 42% of the variance in day to flower after vernalization saturation

(Table S6 and Figure 2b). SNPs in *VERNALIZATION1* (*VRN1*), *FLOWERING LOCUS T-*

LIKE 1 and *10* (*FTL1* and *FTL10*) and *POLD3* show the strongest association with day to

flower after vernalization saturation and MTD (Figure 2a and b). Interestingly, *VRN1*

and *VERNALIZATION2* (*VRN2*), two genes known to play a role in vernalization require-

ment, show a stronger association with MTD than with vernalization saturation (Figure

2a).

Fig. 2: Association between SNPs in flowering-time genes and flowering time-related traits. (a) RDA output for the 143 pruned SNPs. (b). For each gene, the boxplots contrast the effect of the reference (Bd21) and alternative alleles on day to flower after vernalization saturation for the SNP showing the strongest association with the trait. The color code is the same as in Figures 1.

Because SNPs in flowering-time genes show a marginal association with vernaliza-

tion saturation and day to flower after MTD, we focused on day to flower after vernal-

ization saturation and MTD for the rest of the analysis. To further disentangle the os-

tensible confounding effect of population structure (Figure 2b), SNPs in flowering-time

genes and environmental variables on day to flower after vernalization saturation and

MTD, we fitted additional models that included the phenotype as a response variable

and i) the cluster of origin alone (model = cluster alone), ii) all the 20 SNPs (1 SNP per

gene, figure S4b) showing the strongest association with either day to flower after

vernalization saturation or MTD (model = All SNPs) and, iii) the cluster of origin, all the

20 SNPs and the environmental variables associated with each respective trait (as

shown in Table1) as explanatory variables (model = cluster + all SNPs + envt.). We

compared these models successively with the empty model as previously described.

For both traits, single SNPs (e.g. Bd1_5866489 in *VRN1* for day to flower after ver-

nalization saturation and Bd2_5415945 in FTL1 for MTD) gave separately identical or

similar association as the cluster of origin alone (Table 2). This results in a reduced or

null cluster effect in the full model compared the one of single SNPs (Table 2). Note

that the variance explained by a given SNP (e.g. Bd1_5866489 in *VRN1*) differ from the

one obtained above (Figure 2b). RDA do not allow missing value and the models with

20 SNPs (Table2) are therefore based on a slightly different sample size than the model

performed with single SNP (Figure 2b) which results in different R^2 estimates.

Table 2: Output of the RDA between flowering time-related trait, SNPs in flowering time genes and bioclimatic variables

Trait	Model	Variable	Cumulative R2adj	AIC	F	P-value
Flowering time after vernalization saturation	cluster alone	cluster	0.56	10.72	26.44	0.002
	All SNPs	Bd1_5866489	0.56	7.58	103.39	0.002
		Bd3_8185926	0.58	4.51	5.04	0.01
		all variables	0.68			
	Cluster + SNPs + aridity_spring + bio13	Bd1_5866489	0.56	7.58	103.39	0.002
		aridity_spring	0.65	-10.08	21.41	0.002
		bio13	0.69	-18.38	10.44	0.002
		Bd2_5415945	0.72	-25.66	9.22	0.002
		Bd4_41691478	0.75	-35.32	11.62	0.004
			0.87			
MTD	cluster alone	cluster	0.32	132.53	10.57	0.002
	All SNPs	Bd2_5415945	0.37	123.86	48.04	0.002
		all variables	0.57			
	Cluster + SNPs + elevation + bio18	Bd2_5415945	0.37	123.86	48.04	0.002
		bio18	0.46	112.20	14.32	0.002
		cluster	0.52	105.56	3.66	0.008
		Bd3_8185926	0.57	98.44	8.71	0.006
		all variables	0.61			

As mentioned above, RDA optimize the variance explained but do not allow to iden-

tify the mechanistic driver of variation. Considering the allele distribution of SNPs in

flowering time-genes across genetic cluster (Figure 2a), disentangling and quantifying

the effect of population structure from the one of genetic factors or single gene re-

mains therefore challenging for most SNPs in flowering time-genes.

Differentiation of flowering time genes

Most SNPs in flowering-time genes nonetheless appear to be highly differentiated

among genetic clades. This could result from two processes. On the one hand, *B. dis-*

tachyon populations underwent bottlenecks during the last glaciation (Minadakis et al.

2023), which may have led to reduced genetic diversity and highly differentiated alleles

among genetic lineages/clades genome-wide (neutral scenario). On the other hand,

flowering-time genes might be under positive selection and therefore highly differen-

tiated among genetic clades (selection scenario).

To first assess the extent to which our SNPs of interest are more differentiated com-

pared to genome-wide levels, we computed single SNP F_{ST}. To limit the number of

comparison and considering the partitioning of the reference and alternative alleles

among accessions in the flowering-time genes (Figure 2b), we only computed F_{ST} be-

tween accessions of the A and B lineages. We found that the large majority of SNPs

significantly associated with flowering time are above the 3rd quartile of the genome-

wide distribution or even belong to the top 5% outliers (Figure 3a). Hence, with the

exception of SNPs in *EARLY FLOWERING 3* (*ELF3*), *FD*, and *FD1*, all flowering-time

genes harbor SNPs that tend to be more differentiated than the rest of the genome.

To gain power for our analyses, we further made use of the *B. distachyon* diversity

panel composed of 332 natural accessions (Minadakis et al. 2023, Figure 4a) and found

very similar levels of allele differentiation at flowering-time genes (Figure S4c). These

results are in line with a forward simulation we performed under a neutral scenario

using the demographic estimates computed by Minadakis et al. (2023). The F_{ST} distri-

bution calculated between the simulated A and B lineages, as the distribution obtained

with the real data, is indeed largely shifted towards low values (Figure 3b). These latter

results demonstrate that bottlenecks did not lead to elevated genetic differentiation

at the genome-wide level. As such, the high F_{ST} value we observed suggest that most

flowering-time genes are under selection.

(a)

(b)

(c)

Fig. 3: Flowering-time gene differentiation (a) Top: Single SNP FST computed for the 56 accessions of the A and B lineages (dashed line = 5% outliers); bottom: single SNP XTX computed at the genetic clade level (dashed line = 1% POD threshold). (b) Distribution of FST calculated between the A and B lineage with real and forward-simulated data under a neutral scenario. (c) Upset plot displaying the overlap between gene sets identified with the GEAs performed by Minadakis et al. (2023) for 23 bioclimatic variables and flowering-time genes.

Genome-wide scans of positive selection

To properly test for positive selection while accounting for the structure and demo-

graphic history of our populations, we computed X^TX statistics, a measure comparable

to single SNP F_{ST} that accounts for the neutral covariance structure across populations.

In brief, we computed X^TX with our actual SNP dataset over the entire diversity panel

using the five genetic clades as focal populations. We then simulated a pseudo-ob-

served dataset (POD) of 100,000 SNPs under the demographic model inferred from

the covariance matrix of the actual SNP dataset. X^TX statistics were then computed for

the POD to determine the probability of neutrality for each SNP. The threshold of sig-

nificance was thus set to 11.2, a value slightly lower than the 1% outlier threshold

(13.6). Many flowering-time genes display SNPs more differentiated than expected un-

der a neutral scenario (Figure 3b) suggesting that those genes have evolved under

positive selection. Nine flowering-time genes *PHYTOCHROME C* (*PHYC*), *VRN1*, CON-

STANS 1 and 2 (*CO1* and *CO2*), *POLD3*, *FTL1*, *FTL10*, *FTL9* and *VRN2* display extremely

differentiated SNPs ($X^TX > 15$) both in the subset of accessions (Figure 3a) and in the

entire diversity panel (Figure S4c). We also tested for extended haplotypes and foot-

print of selective sweeps with the integrated Haplotype Score (iHS). However, none of

those highly flowering-time genes are located in regions displaying significantly longer

haplotypes (Figure S5).

We ran SnpEff to test which SNPs in flowering-time genes are more likely to have a

functional impact. Only few SNPs in *CO2*, *ES43*, and *Photoperiod-H1* (*PPD1*) were cat-

egorized as variants with high impact while the large majority of SNPs in other genes

were categorized as variants with moderate and low effect (Figure S6), rendering the

identification of the potential targets of selection challenging for most of our genes of

interest.

In order to characterize which environmental factors might have shaped diversity at

flowering-time genes, we eventually made use of genotype-environment association

(GEA) analyses performed by Minadakis et al. (2023) with bioclimatic variables associated to precipitation levels, temperature or elevation. We found that only one of the 22 flowering-time genes (Bradi2g59119, *ODDSOC2*-like) showed an overlap with the gene sets significantly associated with current bioclimatic variables (Figure 3c).

Yet, the GEAs performed by Minadakis et al. corrected for population structure and can hence result in false negative since the genetic clades occupy different ecological niches. We indeed found significant associations between SNPs at flowering-time gene and bioclimatic variables such as Bio14 (Precipitation of driest quarter) and aridity levels in spring, when not accounting for population structure (Kendall correlation; Figure S7a), indicating that the confounding effect of population structure and adaptation at a regional scale may mask the significant effect of the environment on flowering-time genes diversity.

Potential adaptation to past climatic conditions

We estimated the age of flowering-time alleles and found that those arose relatively recently between 9,000 to 38,000 years ago (Table S7) as most alleles in our system (Minadakis et al. 2023). We speculated that variation in flowering time could also reflect adaptation to recent past conditions, potentially to the last glacial maximum (LGM) 22,000 years ago. We therefore tested whether the delay in flowering time we observed in accessions from the A_East and A_Italia clades could result from an adaptation to past colder climates. To do so, we used the niche suitability projections under LGM conditions computed by Minadakis et al (2023). We selected a set of 200 random

points per clade in highly suitable habitats and extracted the 19 LGM bioclimatic vari-

ables for the corresponding sites. The PCA performed with these 19 LGM bioclimatic

variables do not allow us to separate the five genetic clades. Altogether, accessions

from the A_East and A_Italia clades neither occurred in colder nor in wetter environ-

ments than the B_East, B_West and C accessions (Figure S8) which suggests that the

extended vernalization requirement and delay in flowering we observed for the A_East

and A_Italia accessions were not selected during LGM.

Long-range linkage disequilibrium among flowering-time genes

A striking pattern eventually emerges from the former analyses. The fact that the flow-

ering-time genes gave nearly identical signals with regard to their association with

flowering time suggests, as displayed in Figure 2b, that SNPs at flowering-time genes

harbor similar allele frequencies. These similar allele frequencies indicate that some

flowering-time genes might be co-selected (polygenic selection) and remain in strong

linkage disequilibrium despite the large physical distances that separate them (Zan

and Carlborg 2019; Gupta et al. 2023).

To formerly test this hypothesis, we computed LD among the 20 focal SNPs used

for the RDA analyses. LD computed among pairs of focal SNPs in the 56 selected ac-

cessions as well as in the entire diversity panel show a similar pattern: many flowering-

time genes display high LD despite the large distances that separate them on a given

chromosome (Figure 4b). This is especially true for flowering-time genes located on

chromosome 1 and 2 (Bd1 and Bd2), which are found in much higher LD than expected

considering the 95% confidence interval (CI) of the genome-wide LD-decay (Figure 4c).

Strikingly, we also found very high LD among flowering time-genes located on dif-

ferent chromosomes both in the selected accessions and entire diversity panel (Figure

4b). Located on chromosome 1, *VRN1* for instance, is in strong LD with eight other

Fig.4 Linkage disequilibrium (LD) in the flowering time pathway a) Map displaying the locality of the 332 sequenced accessions composing the diversity panel b) LD among flowering-time genes in the selected accessions and diversity panel. For the diversity panel, the black lines delimit the focal genes used for the interchromosomal LD analysis. c) LD decay computed over the entire diversity panel. The orange and grey lines display the 95% CI of the LD decay calculated with or without filtering for a minimum allele frequency (maf) respectively. Dots display LD between pairs of flowering-time genes located on the same chromosome d) Distribution of LD among loci located on different chromosomes. The violin plots display the means of LD calculated among eight, ten, eleven and 13 loci. Numbers in square brackets indicate the confidence interval around the mean.

genes located on chromosome 2 (*POLD3*, LD = 0.84; *FTL9*, LD = 0.45, *FTL1*, LD = 0.76;

FTL10, LD = 0.83), chromosome 3 (*VRN2*, LD = 0.65; *CO2*, LD = 0.55, *BdFT6*, LD = 0.54)

and chromosome 4 (*BdFT4*, LD = 0.71). To test to what extent this long-range LD

among the 14 focal flowering-time genes deviates from genome-wide patterns, we

computed LD for a subset of 50,000 random pairs of genic loci located on different

chromosomes across the genome. This allowed us to estimate the average inter-chro-

mosomal LD among loci (Figure 4d). For each combination, the inter-chromosomal LD

we observed among eight loci (for focal flowering-time genes located on chromo-

somes 1), ten loci (for focal flowering-time genes located on chromosomes 2), eleven

loci (for focal flowering-time genes located on chromosomes) and 13 loci (for focal

flowering-time genes located on chromosomes 4) was largely outside the CI calculated

with random genic loci (Figure 4d).

Note that *FRUITFULL-like* (*FUL2*) and *PPD1* are the only genes displaying putative

footprint of insertion polymorphisms in the diversity panel (Table S5). Furthermore, in

a selfing species like *B. distachyon*, heterozygous SNPs are hallmarks of gene duplica-

tion (Jaegle et al. 2023; Stritt et al. 2022). Using a vcf file not filtered for heterozygous

SNPs, we found that only 0.6% of the SNPs located in flowering-time genes are heter-

ozygous across the 332 accessions. These results imply that flowering-time genes are

occurring in single copy and that structural rearrangements or duplications did not

alter their position in non-reference genomes. Taken together, these results indicate

that the long-range LD among flowering-time genes as well as their differentiation

levels deviate from genome-wide patterns and are difficult to explain by the

demographic history of the population alone. We hence conclude that 14 flowering-

time genes (Figure 4d) are co-evolving and undergo polygenic selection.

Genome-wide association for flowering time variation measured in outdoor
conditions

Greenhouse conditions are far from any natural optimum and may obscure association

with relevant environmental cues, while taking an environment closer to the natural

conditions encountered by of one or several clades acts as a useful pivot to contrast

clades. Hence, we eventually made use of flowering time data we collected for a subset

of 131 accessions (Stritt et al. 2022) in Zürich, Switzerland, to assess flowering time

variation in semi-natural conditions. While *B. distachyon* do not occur at such latitudes

(Minadakis et al. 2023), a PCA performed with the 22 bioclimatic variables on our 332

natural accessions and including a site in Zurich shows that the environmental

conditions in Zürich are not different from the ones encountered by ABR9 for instance

(Figure 5a, Table S2 for bioclimatic data). While other biotic (vegetation type and

density) and abiotic (soil characteristic) factors are certainly preventing the species

from occurring in northern latitude, Zürich harbours climatic conditions similar to ones

encountered by the species at the northern margin of its natural distribution, making

it a valid experimental site to study flowering time in *B. distachyon*.

In brief, we planted seeds for 131 accessions (Figure 5b) outdoors in November

2017, in the Botanical Garden of Zurich, Switzerland, with six replicates per accession

and recorded flowering time in spring. 31 of the accessions used in the greenhouse

experiment are used in the outdoor experiment, including four (Ren4, Lb1, Lb23 and

Cm7) of the five accessions that did not flower by the end of the experiment. All plants

flowered within 20 days in April (Figure 5c). Flowering time for plants from the B_East

and C clades were not significantly different (Kruskal-Wallis test, P-value = 0.96) but

those accessions flowered significantly faster than accessions from the other clades

(Kruskal-Wallis test, all P-values < 0.01). Plants from the A_Italia and B_West flowered

significantly later than plants from the other clades (Kruskal-Wallis test, all P-values <

0.02) but were not significantly different from each other (P-value = 0.06). As such, the

data collected outdoor contrasts with the ones collected in the greenhouse.

Fig. 5: Genome-wide association analysis with flowering time (outdoor experiment) a) PCA performed with 19 classical wordlclim variables combined to solar radiation in spring, global aridity in spring and altitude for the 332 accessions of the diversity panel and a locality in Zurich b) Geographical origin of the 131 accessions used for the experiment c) Flowering time per cluster (displaying replicates) d) Manhattan plot displaying the association between SNPs and flowering time. The red and dark dots display SNPs located into AFT- or other flowering-time genes respectively. The dashed line corresponds to FDR threshold of significance e) XTX values for the top GWAs SNP on chromosome 2 (Bd2).

As for the greenhouse experiment, we used RDA to fit models where flowering-time

was independently entered as a response variable while the cluster of origin

(model = cluster alone), the 22 environmental variables (model = all bioclimatic varia-

bles) or both the cluster of origin and the 22 bioclimatic variables (model = cluster +

all bioclimatic variables) were successively entered as explanatory variables. We com-

pared those full-models to empty models where flowering-time was explained by a

fixed intercept alone (Capblancq and Forester 2023). We found that a large part of the

variance in flowering-time is explained by bioclimatic variables linked to precipitation

levels in warm months whether the models did or did not include the cluster of origin

(e.g. bio18 and bio12; Table 1).

A genome-wide association analysis (GWAs) performed with GEMMA (Zhou and

Stephens 2012) on the 131 accessions and 2,266,225 filtered SNPs (Figure S9 for marker

density) identified one significant peak which does not overlap with any of the 22

flowering-time genes we studied above (Figure 5d). We also made use of the GEAs

performed by Minadakis et al. (2023) and found that the gene underlying the peak

(Bradi2g11490, a carbohydrate-binding-like gene) does not overlap with the gene sets

significantly associated with current bioclimatic variables. Yet, similarly to what we

found for the AFT-genes, we found significant associations (Figure S7b) between SNPs

at the GWAs peak and bioclimatic variables, especially with Bio14

(Precipitation_of_Driest_Month), when not taking into account population structure.

Consistent with selection by the environment, the top GWAs SNP constitutes also an

$X^T X$ outlier (Figure 5e).

Discussion

Flowering time-related traits and adaptation to local climate

Flowering time has been shown to play an important role in local adaptation across many plant systems including crops (Izawa 2007; Anderson et al. 2012; Anderson, et al. 2013; Ågren et al. 2017; Navarro et al. 2017; Wadgymar et al. 2017; Takou et al. 2019; Qian et al. 2020; Yan et al. 2021). In *B. distachyon*, flowering time has been mainly studied from a molecular viewpoint, clarifying the role of specific genes in sensing environmental variation such as day-length and temperature (Ream et al. 2012; Ruelens et al. 2013; Ream et al. 2014; Woods et al. 2014; Woods et al. 2017[a,b]; Lomax et al. 2018; Cao et al. 2020; Kennedy and Geuten 2020; Woods et al. 2020; Bouché et al. 2022; Raissig and Woods 2022). While those studies underline the complex archi-tecture of flowering-time and revealed large variation regarding vernalization require-ment among natural accessions for instance, they provide little information with regard to adaptation.

In fact, the wide range of vernalization requirement observed in *B. distachyon* suggests that local adaptation of flowering time could take place through distinct de-velopmental thresholds (i.e the accumulated amount of 'developmental time' required for a given developmental stage to allow a transition to the next developmental stage see Donohue et al. 2014) among accessions. Indeed, in a widespread Mediterranean species like *B. distachyon*, not all accessions are exposed to long and harsh winter (Fig-ure S2). The expression of traits such as the minimum duration of vernalization neces-sary to permit flowering might be contributing *in fine* to flowering time variation in

the wild and be especially relevant to investigate in the context of local adaption. We

indeed observed a significant albeit partial correlations among the four flowering time-

related traits we investigated (Figure 1b) indicating that these traits constitute different

components of flowering time.

Estimating the effect of the environment remains a challenge in such structured

populations. Yet, we show a significant association between flowering time-related

traits and bioclimatic variables in addition to the cluster effect, indicating that flower-

ing time-related traits are locally adapted trait in our system. Because bioclimatic var-

iables are largely correlated among each other's, pinpointing the selective constraints

acting on flowering time-related traits is equally difficult (Capblancq and Forester

2023). Our RDA analyses, however, point at bioclimatic variables which are biologically

meaningful as linked to aridity levels in warm months. As such, precipitation levels and

hence water limitation seem especially important for flowering time-related traits as

already shown for broader fitness effects in our species (Des Marais et al. 2017). In our

system, MTD and day to flower after MTD have never been investigated. It is therefore

very interesting to see that those two traits also display signs of local adaptation. Our

results therefore advocate for a complex adaptation of flowering time and suggest

that vernalization saturation might not be necessary in the wild to synchronize flower-

ing time with the environment.

Early- and late-flowering genotypes are unlikely driven by adaptation

The greenhouse experiment further confirms that *B. distachyon* natural accessions

vary greatly regarding their vernalization saturation requirement and ultimately flow-

ering time and can be broadly grouped into early- and late-flowering genotypes (Fig-

ure S1 and Figure 1, but see Ream et al., 2014; Gordon, et al. 2017 for a finer classifi-

cation scheme). Despite within clade variation regarding flowering time after vernali-

zation saturation (Figure 1c), accessions from the A lineage (A_East and A_Italia) display

a large delay in flowering compared to accessions from the B (B_East and B_West) and

C lineage, which typically result from longer vernalization saturation requirements. This

pattern has been suggested to be a sign of adaptation at a regional scale in Turkey

(Skalska et al. 2020) and could be interpreted as a footprint of diversifying selection at

the genetic clade level. Based on this hypothesis, accessions from the A_East and A_Ita-

lia clades may display a delay in flowering as globally adapted to colder and less arid

environmental conditions, as shown for instance in *A. thaliana* (Ågren and Schemske

2012). Yet, our results do not fully support this scenario, as we only find partial corre-

lations between the partitioning of the phenotypes and the bioclimatic variables asso-

ciated with flowering time (Figure 1 and Figure S3). Our niche modelling projections

for the last glacial maximum (Minadakis et al. 2023), do not further support an adap-

tation to the last glacial maximum conditions (Figure S8). As such, the partitioning of

the early vs. late flowering accessions remains difficult to explain with an "adaptive"

scenario. In summary, we conclude that local environmental conditions are partly

driving flowering time variation within genetic clades but not the early- and late-flow-

ering partitioning of the phenotypes we observed among clades in the greenhouse.

We have not tested different combinations of vernalization temperature and length.

In *A. thaliana* for instance, natural ecotypes have characteristic vernalization tempera-

ture profiles and vernalization temperature ranges from 0 to 14 °C (Duncan et al. 2015).

While *B. distachyon* does not occur in a wide range of latitude as *A. thaliana*, it is clear

that interactions between vernalization temperature and length are likely to play a role

as well. As already shown by Stritt et al. (2022) with a smaller subset of accessions,

variation in flowering time is for instance largely attenuated in outdoor conditions and

although significant differences are observed among genetic clades, all plants flowered

within 20 days when grown outdoors in Zürich (Figure 5c), including the four acces-

sions that did not flower with the greenhouse experiment. Hence, greenhouse experi-

ments, as largely performed in *B. distachyon* (Ream et al. 2014; Gordon, et al. 2017;

Sharma, et al. 2017) may only weakly capture how accessions behave in the wild.

Flowering-time genes and flowering time-related traits

We also selected 22 flowering-time genes known to play a major role in vernalization-

(e.g., *VRN1*, *POLD3*) or photoperiod-sensing (e.g. *PPD1*, *PHYC*) to characterize the

magnitude of their allelic effects on flowering time-related traits. We found that SNPs

in those 22 genes display relatively large association with the number of days to flower

after vernalization saturation and to some extent with MTD or vernalization saturation

but only poorly with the number of days to flower after MTD (Figure 2a). Interestingly,

genes known to play a role in cold-mediated vernalization such as *VRN1* or *POLD3*

(Raissig and Woods 2022) show a larger association with MTD than with vernalization

saturation, suggesting that those genes are important for cold perception regardless

of its length.

Here again, our RDA analyses show that disentangling the effect of the cluster of

origin from the one of single SNPs is challenging. We indeed observed a strong parti-

tioning of the reference and alternative alleles in flowering-time genes among genetic

clades and lineages (Figure 2b). The high genetic differentiation of flowering-time

genes between the A and B lineages (Figure 2b) could result from the recent bottle-

necks experienced by *B. distachyon* in the recent past (Stritt et al. 2018; Minadakis et

al. 2023) and produce spurious association with flowering-time related traits. Three

lines of evidence rule out this demographic scenario.

First, our 22 candidate genes have been shown to play an important role in photo-

period sensing, vernalization and *in fine* in flowering variation among diverse *B. dis-*

tachyon accessions not only by sequence homolgy but also via molecular, biochemical,

and genetic methods (Wu et al. 2013; Ream et al. 2014, Woods et al. 2014, Woods et

al. 2017[a], Woods et al. 2017[b], Lomax et al. 2018; Qin et al. 2019, Woods et al. 2019,

Woods et al. 2020, Raissig and Woods 2022, Bouché et al. 2022; Alvarez et al. 2023;

Woods et al. 2023). It is therefore legitimate to assume that they might play a role in

regulating flowering time-related traits in genetically diverse accessions as well.

Second, FST analyses performed with our real- as well as forward-simulated data

show that the population size reduction experienced by *B. distachyon* did not lead to

genome-wide highly differentiated alleles and SNPs at flowering-time gene constitute

clear FST outliers. The X^TX analysis we performed with the five genetic clades as focal

populations also shows that most of the flowering-time genes harbor SNPs above the

threshold of significance, *POLD3* and *CO1* presenting more extreme outliers. This foot-

print of positive selection (Gautier 2015) is yet not accompanied by extended haplo-

types around our candidate genes (except for *CO2*) which implies that the initial selec-

tive sweeps may have eroded with time and that selection did not occur in a recent

past. Flowering-time gene are yet not colocalizing with regions we previously identi-

fied with GEA analyses (Figure 3c). Thus, although we find evidence of positive selec-

tion at single flowering-time genes, the selective constraint at play remains yet to be

identified. It is yet essential to keep in mind that GEA performed in species with strong

population structure can lead to high rate of false negatives as they typically include

structure correction (Lottheros 2023; Booker et al. 2023). We did find significant asso-

ciations between bioclimatic variables and SNPs at flowering-time genes when not

correcting for population structure (Figure S6). Common garden experiments will thus

be key to test the effect of genotype-by-environment interactions in our system.

Third, we highlighted a striking pattern of long-range linkage disequilibrium (LD),

both intra- and inter-chromosomal, among 14 flowering-time genes associated with

flowering time variation under greenhouse conditions (Figure 4). Polygenic selection

is expected to result in LD between regions under selection (Gupta et al. 2023; Yeaman

et al. 2016; 2018) but only few cases have been reported so far in plants and animals

(Hohenlohe et al. 2012; Yeaman et al. 2016; Gupta et al. 2023; Park 2019). In *A. thali-*

ana, Zan and Carlborg (2019) also identified long-range LD among four clusters of

flowering-time genes. We demonstrated that the levels of long-range LD (intra- and

inter-chromosomal) we observed among 14 flowering-time genes (Figure 4) are not

observed among random genic loci in the genome. As flowering-time genes do not

display signs of duplication or insertion polymorphisms, we imply that structural rear-

rangements did not bring flowering-time gene physically together in non-reference

accessions. We therefore believe that we present clear evidence for polygenic selection

on key genes involved in the flowering pathway. Expression analyses further support

the functional connection among these loci in *B. distachyon* and grasses in general.

For example, *VRN1* and *FTL1* are expressed in a positive feedback loop, which over-

comes the flowering repression of *VRN2* (Ream et al. 2014; Woods et al. 2016; Woods

et al. 2017[b]). Additionally, there is an intricate connection between many of these flow-

ering-time genes both transcriptionally and at the protein level. For example, among

the pairwise interactions tested between *PHYC, PHYB, ELF3, PPD1, VRN2, CO1,* and

CO2 more than 80% showed positive interactions in yeast two hybrid assays some of

which have been verified *in planta* (Shaw et al. 2020; Alvarez et al. 2023). Thus, many

flowering-time genes in LD can interact at multiple levels.

Altogether, this set of analyses show that the association between flowering-time

genes and flowering time-related traits, especially days to flower after vernalization

saturation, is a result of selection and unlikely due to population structure.

Interestingly, while *pold3* and *elf3* loss of function mutants flower faster than the

wildtype (Bouché et al. 2022;. Woods et al. 2020), their contribution to flowering time

after vernalization saturation for instance (39% and 15% respectively) differ drastically.

This discrepancy might be due to the fact that mutants used to characterize gene func-

tions usually harbor mutations with deleterious or large effect size (loss of function).

In contrast, our natural flowering-time gene variants are mostly predicted as having

low to moderate suggesting that they modulate flowering time-related traits quanti-

tatively. As such, one should remain cautious while extrapolating the effect of flower-

ing-time variants from mutant-based experiments.

GWAs detect additional candidate genes for flowering time

None of the flowering-time genes co-localize with the GWAs candidate identified

with the outdoor experiment (Figure 5d). It is clear that the partitioning of the refer-

ence and alternative variants at flowering-time genes (Figure 2a) largely dimmish our

detection power as we applied a correction for population structure. Yet, while we

should extend our common garden experiment to a larger number of sites, our results

are in line with a previous studies in *A. thaliana* which, by using natural accessions and

RILs, found that flowering time variation scored in the field experiment poorly corre-

lated with the flowering time variation obtained under greenhouse conditions (Wilczek

et al. 2010; Brachi et al. 2010; Malmberg et al. 2005; Weinig et al. 2002). As a conse-

quence, a limited overlap is observed between the genomic regions detected in field

experiments and those detected under greenhouse conditions (Brachi et al. 2010).

Flowering time has a complex polygenic architecture (Buckler 2009; Brachi et al. 2010;

Navarro et al. 2017; Zan and Carlborg 2019; Gaudinier and Blackman 2020) and the

use of EMS-induced mutants suggested that many additional genes might play a role

in shaping this trait in *B. distachyon* (Raissig and Woods 2022), some of which are not

described or only play a minor role in flowering time in other plant models (Woods et

al. 2020). In addition, alleles may affect phenotypes only in specific populations (Yan

et al. 2021; Gloss et al. 2022; Zan and Carlborg 2019) or seasons (Weinig et al. 2002;

Gould and Stinchcombe 2017). Taking into account the polygenic architecture, gene-

by-environment association and phenotypic plasticity (Gaudinier and Blackman 2020;

Yan et al. 2021) will therefore be essential to better capture the adaptive potential of

flowering time and flowering-time genes in natural populations of *B. distachyon.* Alt-

hough flowering-time genes are undoubtedly essential in the perception of environ-

mental cues and overwintering (for review Raissig and Woods 2022), our results show

that the effect of their variants in the wild are more difficult to predict and that addi-

tional genes may be at play.

Conclusion and perspectives

Our results suggest that i) flowering time-related traits are a locally adapted trait in

B. distachyon but ii) part of the variation in those traits can be explained by the

environment and by SNPs in key flowering-time genes and, iii) key flowering-time

genes are co-evolving but their effect in the wild remains to be clarified. In the face of

global warming, plant phenology has recently advanced significantly (e.g. Anderson et

al. 2012). Investigating the polygenetic architecture of flowering time therefore

remains a timely question. Polygenic selection, epistatic and pleiotropic effects might

limit the evolution of traits (Yan et al. 2021; Yeaman 2022) and future experiments in

B. distachyon should focus on disentangling these effects. Mimicking the combination

of the various clues that trigger flowering is yet impractical under greenhouse

conditions and common garden experiments will thus be essential to place flowering

time in a natural context in this system. We, however, know relatively little about the

basic ecology of *B. distachyon*. Fundamental questions now need to be addressed to

more broadly characterize the process of local adaptation in this species. These

include: when are plants emerging in the wild?; how plastic are flowering time and

phenotypes in general?; and what is the contribution of seed banks to the effective

population size and hence selection strength? Answering these will undoubtedly

represent remarkable progress in our understanding of *B. distachyon's* ecology.

Author contribution

NM conceived the study, performed the greenhouse experiment, contributed to the

analysis and the writing of the manuscript. MT collected samples, contributed to the

analysis. WX, RH and LK contributed to the analysis. DPW contributed with the writing

of the manuscript. ACR conceived the study, contributed to the analysis and with the

writing of the manuscript.

Data availability

Seeds will be distributed through GRIN (https://www.ars-grin.gov). Accession numbers for sequencing data can be found in (Gordon et al. 2017; Gordon et al. 2020; Skalska et al. 2020; Stritt et al. 2022; Minadakis et al. 2023). The raw data of the flowering time measurements are available in the supplementary materials. Bioclimatic variable can be downloaded at https://www.worldclim.org/data/bioclim.html and extracted using the R code "download_worldclim2.1.R" provided by Minadakis et al. 2023 (https://zenodo.org/records/8354318). RDA analyses were adapted from the detailed scripts provided by Capblancq and Forester (2023; https://github.com/Capblancq/RDA-landscape-genomics/tree/main)

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Supplementary Materials

The supplementary materials consist in nine figures (File Supplementary figures) and seven tables.

Table S1: Raw results of the flowering time experiment.

Table S2: Bioclimatic variable extracted for the 332 accessions of the diversity panel

Table S3: Flowering time-related traits extracted from the greenhouse experiment

Table S4: Differences in flowering time-related traits among the five genetic clades.

Table S5: Flowering time gene position and ID.

Table S6: RDA output for the four flowering time-related traits again the 502 SNPs.

Table S7: Flowering-time gene age estimates.

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