Examining the molecular clock hypothesis for the contemporary evolution of the rabies virus

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https://github.com/RowanDurrant/Rabies-Mutation

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Abstract

The molecular clock is a method for measuring the rate of virus evolution but the key

assumption that mutations accumulate on the genome at a constant rate over time does not

always hold true. While modelling approaches exist to accommodate deviations from a strict

molecular clock, assumptions about rate variation may not fully represent the underlying

evolutionary processes and can affect the accuracy of clock calibration and divergence time

estimates. There is considerable variability in rabies virus (RABV) incubation periods,

ranging from days to over a year, during which viral replication may be reduced. This

prompts the question of whether modelling RABV on a per infection generation basis might

be more appropriate. We investigate how variable incubation periods affect root-to-tip

divergence under per-unit time and per-generation models of mutation. Additionally, we

assess how well these models represent root-to-tip divergence in time-stamped RABV

sequences. We find that at low mutation rates (<1 substitution per genome per generation)

divergence patterns between these models are difficult to distinguish, while above this

threshold differences become apparent across a range of sampling rates. Using a Tanzanian

dataset, we calculate the mean substitution rate to be 0.17 substitutions per genome per

generation. At RABV's substitution rate, the per-generation substitution model is unlikely to

represent rabies evolution substantially differently than the molecular clock model when

examining contemporary outbreaks; over enough generations for any divergence to

accumulate, extreme incubation periods average out. However, measuring substitution rates

per-generation holds potential in applications such as inferring transmission trees and

predicting lineage emergence.

Introduction

The molecular clock is a method of measuring the rate of evolution of organisms, based on

the assumption that their genomes accumulate neutral mutations at a constant rate over time,

either across all lineages (the "strict molecular clock") or within each individual lineage but

with variation between them (the "relaxed molecular clock") (1,2). The ability to sample viral

sequences through time, and the application of the molecular clock hypothesis to these

sequences, has led to massive advances in using viral genetic data to investigate disease

outbreaks (3). The clock rate, measured in substitutions per site per unit time, can be used to

estimate how long ago pathogens diverged (4), and the date of infection of individual infected

hosts (5). Combining the analysis of epidemiological and genetic data has allowed further

insights into the history of outbreaks (6), and the introduction of geographic data provides

predictions as to rates of spread and the frequency and source of introductions (7,8).

However, in order to conduct these phylogenetic analyses, genetic divergence must

measurably increase over time in the dataset under investigation (9).

The rabies virus (RABV) is a negative-strand RNA virus, with a genome size of

approximately 12 kilobases. While RNA viruses generally have high mutation rates due to a

lack of proofreading by RNA polymerases, RABV has a relatively low clock rate of between

1×10^{-4} and 5×10^{-4} substitutions per site per year (10–12). This may be due to strong

purifying selection (10), or to an unusual feature of RABV: that infections can exhibit

extended incubation periods within the host. The median generation interval (the time

between one individual becoming infected and then infecting another) is estimated to be 17.3

days in domestic dogs (13), with other studies estimating mean serial intervals of 26.3 days

(14) and 45.0 days (15). Symptoms, infectivity, and death from rabies, however, can

occasionally occur years after the initial infection event (16). The length of this incubation

period is influenced by the route of exposure, with bites to the head and neck leading to more

rapid disease progression than bites to lower extremities (17). RABV can remain in the

muscle at the bite site for significant lengths of time before invading the host's motor neurons

and progressing through the nervous system, with limited, if any, infection of other muscle

fibres (18). While some replication in the muscle cells has been observed (19), RABV

replication at the inoculation site is not necessary for neural invasion (20). It is currently

unknown precisely how the RABV replication rate in the host muscle cells and peripheral

nervous system compares to the massive replication rate within the cells of the central

nervous system and brain. However, work suggests that RABV replication in muscle cells

may be reduced (21), and RABV replication in cultured rat sensory neurons may be 10- to

100-fold lower than replication rates in rat and mouse CNS neurons (22). Rabies infections

that involve long incubation periods may, therefore, not lead to more accumulated mutations

than those with shorter incubation periods, as viral mutation is strongly influenced by the

replication process (24).

Changes in mutation rates through time due to long incubation periods may affect how we

analyse RABV sequence data and interpret these analyses. A relaxed molecular clock is

usually required to carry out phylogenetic analyses on rabies datasets, and it is not

uncommon for there to be difficulties in applying these analyses due to "insufficient temporal

signal"; usually referring to either no or a negative relationship between genetic divergence

and time, or this relationship having a very low R^2 (25–29). RABV clock rates can differ

significantly between hosts (30), geographic locations (31) and circulating lineages (12),

which may be driven in part by differences in incubation periods. If the variable incubation

period of rabies infections does cause deviation from the molecular clock model (exceeding

the variation captured by relaxed or modified clock models), this may negatively affect the

accuracy of time-scaled phylogenetic trees and emergence date predictions. Conversely, if

mutation does continue at a consistent rate during the incubation period, attention should be

paid to extremely long incubators which could drive the emergence of new variants, as seen

recently in chronic SARS-CoV-2 infections (32,33).

We hypothesised that reduced replication (and thus mutation) during the incubation period

could cause rabies evolution to be better represented by a per-generation model of mutation

than by the molecular clock model. We aim to clarify the nature of contemporary RABV

evolution using in silico methods, comparing the root-to-tip divergence of sequences

generated from synthetic outbreaks under per-unit time or per-generation mutation models,

and comparing these to RABV genomic data from Tanzania. We also aim to calculate a per-

generation substitution rate for RABV for future use as a parameter in transmission tree

reconstruction algorithms.

Methods

We investigate two contrasting mutational models for RABV – i.e., substitutions occurring on

a per-generation vs. per-unit-time basis – using a simulation approach. We first generated

synthetic RABV outbreaks using a branching process model (13) and then simulated these

two mutation processes over the resulting transmission trees. From the resulting synthetic

sequences, we examined root-to-tip divergence and calculated variance explained (R^2) from

linear regressions, and compared these to the root-to-tip divergence of a set of RABV whole

genome sequences from Tanzania. Finally, we developed a method to estimate the per-

generation substitution rate for RABV and tested this on synthetic data before applying it to

the real RABV dataset.

<u>Rabies outbreak simulation</u>

We simulated RABV mutation on branching-process simulations of rabies outbreaks.

Outbreaks were simulated 100 times over a spatially explicit representation of Mara Region

in northern Tanzania. In Serengeti District, where contact tracing data were available, the

model was initialised with the three cases that occurred in the mean generation interval (g=27

days, based on contact tracing data) prior to 2017 (simulations were run over a dog

population representing that in Mara region between 2017 and 2024). In the rest of Mara

region, where there were no data to guide initialisation, we seeded with ($0.01Dg$)/365 cases,

where D is the initial dog population in that area. If R_e=1 (endemic transmission), this results

in roughly 1% of the population becoming rabid over a year; contact tracing data suggest that

incidence typically does not exceed that level (34). This led to a total of 273 initial cases in

the region. Each case was assigned a number of offspring cases drawn from a negative

binomial distribution (34) with mean (R_0)=1.05 and dispersion parameter=1.33. The R_0 value

was chosen to result in a median number of cases each month that was roughly constant over

time (over the 100 simulations), mimicking endemic disease. Movement of rabid dogs from

their home locations to and between transmission locations followed a random walk with step

lengths drawn from a Weibull distribution (shape=0.41, scale=0.13). We simulated occasional

long-distance transport of dogs to a random location prior to their first transmission in 2% of

cases (13). At each of a rabid dog's transmission locations, another dog was randomly

selected within the local 1km2 grid cell. If this dog was susceptible (i.e., not vaccinated or

already incubating infection from a prior transmission event), rabies was transmitted. A

generation interval was drawn for each new infection from a lognormal distribution

(meanlog=2.96, sdlog=0.82), describing the time delay before it also became rabid and made

its assigned transmissions. The step-length and generation-interval distributions were fitted

using contact tracing data from Serengeti District, Tanzania (13). Branching process

simulations were continued until 7 years had passed or rabies went extinct. Each synthetic

case was assigned an individual ID, and for every case (except initial seed cases) we recorded

the ID of the associated progenitor case. Dates of infection and transmission were recorded

for each case.

We isolated complete transmission trees descending from each of the 273 initial cases from

within one randomly selected synthetic outbreak. Transmission trees that contained over 100

cases (9 out of 273 trees in total, that ranged in size from 533 - 19,382 cases) were then used

to generate synthetic sequence data. For each of the 9 trees the index case was assigned an

initial 12kb genome sequence. Under the per-unit time mutation model, we determined the

expected number of mutations by multiplying the clock rate, the genome length and the

length of the generation interval, for each case along the resulting transmission tree. The

realised number of mutations was then chosen from a Poisson distribution, with this mean.

We then randomly chose positions to change and new nucleotides to change them to. The

resulting synthetic sequence data is referred to as the "time-based sequence data". The

generation-based model of mutation works as above, with the exception that the expected

number of substitutions in a generation is constant and produces the synthetic "generation-

based sequence data".

Divergence rate analysis

To investigate patterns of temporal divergence under the mutation models described above,

we generated synthetic data with values of substitution rates ranging from 0.05 to 3

substitutions per genome per generation (or the per unit time substitution rate equivalent) and

4 population sampling regimes (from 1% of cases to 20%, informed by a previous study that

estimated that routine surveillance for rabies rarely confirms more than 10% of circulating

cases (35)). We calculated the genetic divergence as the number of nucleotide differences

from the index case to each sampled case. For each of the nine transmission trees, we then

compared genetic divergence with time under each scenario (substitution rate and sampling

regime combination), using linear regression through the origin.

In order to compare our synthetic patterns of divergence over time to real rabies data, a root-

to-tip divergence plot was also generated for a dataset of real RABV sequences (data from

(36); Figure 1A) using TempEst (v1.5.3 (37)), with the best-fit root located (Figure 1B).

These rabies cases occurred between 2001 and 2017 and were primarily from the Serengeti

district and Pemba Island, with the remaining sequences from elsewhere in Tanzania (Figure

1A inset). Sequence acquisition and tree building methods are detailed in (36).

Fig 1. Phylogeny of real rabies virus whole genome sequences from Tanzania and root-

to-tip divergence. (A) The time-scaled tree (36) used to generate the root-to-tip divergence

plot and to calculate the per-generation substitution rate. The inset map shows the

approximate locations that the samples were collected from, and the lineages present in each

location. Map point size represents the number of sequences in this dataset from district

centroid locations. **(B)** The corresponding root-to-tip divergence plot. Point colours represent

RABV lineage.

<u>Calculating the per-generation substitution rate</u>

We developed a method of calculating the per-generation substitution rate using the clock

rate, the genome length, and the mean generation interval. We assessed this method's

accuracy using the synthetic outbreak sequence data, before applying it to the aforementioned

set of RABV whole genome sequences.

To estimate the mean per-generation substitution rate, we analysed sequence data with

BEAST, and multiplied clock rate values from the BEAST posterior by the generation-

interval lengths sampled from the posterior of a simple Bayesian analysis and then multiplied

again by the genome length. The mean and 95% confidence interval of the prediction of the

per-generation substitution rate for the RABV dataset was calculated by taking the mean and

the 2.5% and 97.5% percentiles of the resulting multiplied posteriors.

To evaluate the accuracy of this method in predicting the mean per-generation substitution

rate, we also applied it to synthetic sequence data generated from outbreaks using the per-

generation mutation model as described above, under different substitution rates (11 values

ranging from 0.05 substitutions per generation to 3 substitutions per generation) and case

sampling rates (1%, 5%, 10% or 20% of cases sampled) across the 9 transmission trees that

contained at least 100 cases. A subsample of synthetic sequences was selected and saved to a

FASTA file. BEAST log files were generated from these sequences using BEASTGen version

1.0.2 and BEAST version 1.10.4 (38). We chose to use a JC substitution model with a strict

clock, no site heterogeneity due to our per-generation mutation model used in the simulations

having equal probability of any site or base being chosen and assumed constant population

size. We used a chain length of 10 million and tracelog frequency of 1000 and used a 25%

burn-in period. We applied the substitution rate calculation method to these phylogenetic

trees, and assessed the accuracy of the resulting mean per-generation substitution rates by

comparing them to the parameter values used to generate the synthetic sequences, using the

natural log of the ratio:

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$$Deviation = \ln\left(\frac{M_p}{M_a}\right)$$

where M_p is the mean predicted per-generation substitution rate and M_a is the actual

substitution rate, where a deviation of zero means perfect accuracy.

The same method was applied to the dataset of 153 RABV sequences sampled from across

Tanzania (data from (36); Figure 1A). The mean per-generation substitution rate was

calculated, and distributions were fitted from the multiplied generation interval and clock rate

posteriors (generation interval posteriors based on values from (13) for the Tanzanian dataset,

extracted directly from the lognormal distribution used in simulations, and clock rate

posteriors taken from the BEAST log file of the time-scaled tree from Lushasi et al. (36)) and

genome length as described above. We compared different distributions (Gamma and

Lognormal) for predicting substitution rates and selected the best fitting distribution by AIC.

We also calculated the probabilities of between 0 and 10 SNP differences occurring across 1,

5 or 10 infection generations. For this calculation we simulated mutations arising at a Poisson

rate with lambda drawn from the fitted substitution rate distribution. The means and 95%

confidence intervals were calculated from the 10,000 simulations.

<u>Data and code availability</u>

All code is available at https://github.com/RowanDurrant/Rabies-Mutation. Analyses were

conducted using the R programming language (39). The beta regression curve and prediction

interval in Figure 2C was generated using the 'betareg' R package (40). RABV lineages were

assigned using MADDOG (41).

Results

<u>Root-to-tip divergence analysis</u>

At higher per-generation substitution rates (1 substitution per genome per generation and

above), distinct differences can be seen between root-to-tip divergence plots from the two

models of mutation (Figure 2A). The synthetic data generated from the per-generation

mutation model shows "stray" clusters or ridges of points both above and below the main

funnel of points, illustrated in the example shown in Figure 2A. Divergence plots from

synthetic data generated from the time-based model of mutation have less variance and do not

exhibit this pattern. At lower substitution rates (below 1 substitution per generation), no such

pattern is clearly distinguishable (Figure 2B). When the cases represented by the high-

divergence points from the per-generation model in Fig. 2A are visualised in a transmission

tree, they are mainly confined to a single chain (Supplementary Figure S1).

Fig. 2: Temporal genetic divergence varies under two models of mutation. (A) Root-to-tip divergence plots for synthetic sequences produced using time-based and generation-based mutation models, equivalent to 2 substitutions per genome per generation and **(B)** equivalent to 0.2 substitutions per genome per generation. Note that the y-axis scales differ by an order of magnitude between A and B. These data are from running mutation models over the same single transmission tree and have a case sampling rate of 5% (i.e., 621 cases sampled of 12,434 total). **(C)** The R^2 values obtained from regression through the origin of root-to-tip divergence of synthetic data from the time-based and generation-based models. Point colour indicates the mutation model used to generate the data. Lines represent beta regressions with logit links fit to data points, and shading represents the 95% prediction interval. The X axis is

log scaled. 5% of cases were sampled here; sampling rate had little effect on R2

(Supplementary Figure S2).

Root-to-tip divergence plots derived from synthetic transmission trees using the time-based

mutation model had, on average, higher R^2 values than those from synthetic transmission

trees using the per-generation mutation model, although this is more difficult to distinguish

below a substitution rate of 0.5 substitutions per genome per generation (Figure 2C). As the

substitution rate increases, the R^2 values across both mutation models increase. The case

sampling rate appears to have little effect on R^2 (Supplementary Figure S2).

The root-to-tip divergence plot of the Tanzanian RABV dataset more closely resembles those

of lower substitution rate simulations, where it is difficult to determine any difference

between the models of mutation (Figure 1). While most lineages surround the regression line,

some (for example Cosmopolitan AF1b_B1) group below the line, but without forming a

distinguishable "ridge".

<u>Substitution rate calculation</u>

The accuracy of our method used to calculate per-generation substitution rate remains similar

at all but the lowest values of substitution rate (Figure 3), with a tendency to underestimate

the substitution rate (meaning that the predicted substitution rate is below the substitution rate

parameter used to generate the synthetic data; mean natural log of the ratio of -0.13 and root-

mean-square of 0.54, where values of 0 indicate perfect prediction). Accuracy appears to be

more influenced by the number of sequences used in the BEAST analysis than by the case

sampling rate itself; the mean natural log of the ratio falls to -0.32 when fewer than 50

sequences are used (root-square-mean of 0.91). The reverse is also observed, as when more

than 2000 sequences are used the mean natural log of the ratio rises to an overestimation of

0.38 (root-mean-square of 0.85). This is likely due to the chain length of 10,000,000 being

insufficient for such a large number of sequences, and so should be taken into consideration if

using this method on very large datasets.

Fig. 3: Accuracy of per-generation substitution rate predictions for different numbers of

sequences, substitution rates and sampling rates. Facets indicate case sampling rate. The

dotted line represents perfect accuracy. X axis and colour scale are log transformed.

The Tanzanian RABV dataset from which we estimated the per-generation substitution rate

contains 153 sequences in total, and the accompanying time-scaled phylogenetic tree has a

root-to-tip height of approximately 65 years, although the sequences spanned just 16 years as

they were sampled from 2001 to 2017 (with 46.7% from years 2011-2012). These sequences

were largely complete; 98% of sequences were >95% complete (>11,327 kb in length). The

mean per generation substitution rate of RABV in this dataset was estimated to be 0.171

(95% confidence interval: 0.127 - 0.219). The best fitting distribution by AIC to the output of

the multiplied Bayesian posteriors was a Gamma distribution with the parameters shape =

51.69 and rate = 301.8 (table 1).

Table 1: Parameters and distributions used to estimate the per generation substitution

rate. The distributions fitted to the multiplied Bayesian output are shown, along with their

parameters and AIC.

Distribution	Parameters	Δ AIC
Gamma	shape = 51.69, rate = 301.8	0
Lognormal	meanlog = -1.774, sdlog = 0.1402	1277.4

318   Using the calculated per generation per genome substitution rates, we calculated the

319   probability of different numbers of substitutions occurring over 1, 5 and 10 generations,

320   drawing the per-generation substitution rate (λ) from the above distribution (Figure 4). Over

321   many generations it is still quite likely for zero substitutions to occur; after 10 generations,

322   the probability of zero substitutions having occurred is 0.19.

Fig. 4: Probability distributions of the mean per-generation substitution rate and

substitutions occurring over generations. **(A)** predicted probability distribution of the per

genome per generation substitution rate from Tanzanian RABV sequences, with underlying

histogram of multiplied Bayesian posteriors of clock rate and generation interval. **(B)**

probability distribution of SNPs occurring over 1, 5 and 10 generations. The λ value for a

Poisson rate of SNP occurrence is drawn from the SNPs per generation distribution fitted in

Figure 4A. Black bars represent the 95% confidence intervals (which are very tight).

Discussion

It can be difficult to get sufficient temporal signal for RABV sequence datasets, potentially

due to its variable incubation periods. We hypothesised that a per-generation model of

mutation may be more representative of RABV evolution than a purely time-based model.

We found that substantial differences in root-to-tip divergence patterns between synthetic

outbreaks using generation-based and time-based models of mutation could be observed only

at high underlying substitution rates. The substitution rate for the Tanzanian RABV

sequences examined (~0.17 substitutions per genome per generation) was in the range where

divergence patterns in the two models were extremely similar. We can thus assume that the

two models will give extremely similar results on the relevant time scale. As we observed

increasing divergence over time with reasonable R^2 values within this substitution rate range,

it implies that variable incubation periods alone may not fully account for the challenge in

obtaining temporal signal. Other factors such as sampling biases (i.e., insufficient sequence

diversity captured over a short timescale) may be responsible. This is an important

consideration for analysing RABV sequences from new outbreaks, or from endemic areas

where sampling is opportunistic.

The observation of little difference between root-to-tip divergence plots derived from the two

mutation models at substitution rates below 1 substitution per genome per generation is likely

because of averaging; multiple generations of infection are expected to have occurred per

substitution that arises on the viral genome. Over the many generations needed before

significant levels of viral genetic diversity are reached, the influence of any unusually long

incubation periods will be damped by the opposite influence of unusually short incubation

periods, eventually becoming indistinguishable from clock-like behaviour. On the other hand,

at higher substitution rates ridges form on the root-to-tip divergence plots under the per-

generation model of mutation but not under the per unit time model. While not affecting the

overall clock rate, these ridges reduce the overall R^2, and may be better analysed using a

separate local clock (42). The cases in these ridges almost all descend from a common

ancestor (Supplementary Figure S1), suggesting that a single unusually long or short

incubation period can affect which phylogenetic analyses we perform. Studies examining the

number of substitutions occurring between successive sequenced cases, and whether this

increases when the secondary case's incubation period is unusually long, could clarify the

exact relationship between substitutions, generations, and time. More detailed data will be

required to investigate this further.

We calculated RABV's mean per-generation substitution rate to be approximately 0.17

substitutions per genome per transmission generation. This estimate is lower than predictions

for other RNA viruses, such as SARS (2 substitutions per genome per human passage (43)),

SARS-CoV-2 (0.52 substitutions per genome per 5.8-day generation interval (44)) and Ebola

virus (0.875 substitutions per genome per 14-day generation interval (45)). RNA viruses that

undergo periods of reduced replication or complete latency often show reduced substitution

rates, with one extreme example being HTLV-1/2 (46,47). However, we would not expect this

to affect the per-generation rate. The low per-generation substitution rate seen in rabies is

therefore likely due to mutation being constrained by other factors, such as strong purifying

selection (10).

We can predict from the estimated per-generation substitution rate that identical sequences

are likely to have less than 5 intermediate generations between them (probability of fewer

than five generations occurring before a mutation occurs > 0.49 by repeated sampling of a

Poisson distribution with a lambda of 0.17), but still have a non-negligible probability of

being more distantly related. While the low substitution rate means that comparing the

number of SNPs between sequences alone may not be an effective method of determining

infector-infectee relationships, it could be used in conjunction with temporal and location

data to make more accurate predictions of transmission events by ruling out relationships

between more distantly related transmission chains co-circulating in the same area, as in (48).

Notably, our Poisson distribution of the number of substitutions occurring in one generation

is visually very similar to the genetic signature distribution reported in Cori *et al*. ((48), Fig

S1), despite different methods and RABV datasets being used in their calculations. The mean

per-generation substitution rate produced through the methods described could also

potentially be used in the future to estimate the size of rabies outbreaks from genetic data and

calculate predicted rates of RABV lineage emergence based on the classification guidelines

of MADDOG (41).

While the Jukes-Cantor model was the most appropriate to use on our synthetic data due to

the simplicity of the mutation models, phylogenetic analyses on real RABV genome data

usually use a more complex model, such as the GTR + G substitution model used to generate

the Tanzanian tree shown in this study (36). This, along with the simplicity of our mutation

model as well as sampling biases in the real dataset, may affect how comparable synthetic

root-to-tip divergence plots are to the real data. We also observed that when a very large

number of sequences (> 2000) were used to predict the per-generation substitution rate, the

prediction accuracy decreased. Upon inspecting the BEAST log files using Tracer (49), it

became apparent that our one-size-fits-all approach of using a BEAST chain length of

10,000,000 was insufficient for such large amounts of data. While it is unlikely that this

number of sequences will be used for this purpose in the future, the prediction accuracy is

likely to be increased simply by using an appropriate chain length.

While the molecular clock has proven critical for gaining insights into the history and

dynamics of disease outbreaks, the epidemiological characteristics of a virus should be

considered when choosing how to measure viral evolution. In this study, we determine that

the per-generation model is not likely to produce substantially different results from the

molecular clock model when analysing contemporary RABV evolution. We also estimate the

mean per-generation substitution rate of RABV for future use in transmission tree

reconstruction and efforts to estimate outbreak sizes and lineage emergence rates. Given that

many different lineages circulating simultaneously is seemingly a common occurrence in

areas with endemic rabies, it is important to investigate whether these lineages vary in

evolutionary rate and generation interval length, and ascertain the potential effects on

phylogenetic analyses.

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Supplementary Information

Supp. Fig. S1: points in the offshoot ridge predominantly occur in one transmission tree.

(A) root-to-tip divergence plot (2 SNPs/genome/generation, 5% sampling rate) with offshoot

ridge points highlighted in red. Offshoot ridge points are defined in this plot as having a

divergence rate above 8×10^{-6} substitutions/day and occurring after day 750. **(B)** transmission

tree of the simulated outbreak with offshoot ridge cases highlighted in red. Graph edge length

is not proportional to time or divergence.

Supp. Fig. S2: Sampling rate does not impact the R^2 of root-to-tip divergence plots from synthetic data. Plot is faceted by sampling rate as a proportion of the total number of cases in the outbreak, point colour represents mutation model.