Examining the molecular clock hypothesis for the contemporary evolution of the rabies virus

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Abstract

The molecular clock hypothesis assumes that mutations accumulate on an organism's genome

at a constant rate over time, but this assumption does not always hold true. While modelling

approaches exist to accommodate deviations from a strict molecular clock, assumptions about

rate variation may not fully represent the underlying evolutionary processes. There is

considerable variability in rabies virus (RABV) incubation periods, ranging from days to over

a year, during which viral replication may be reduced. This prompts the question of whether

modelling RABV on a per infection generation basis might be more appropriate. We

investigate how variable incubation periods affect root-to-tip divergence under per-unit time

and per-generation models of mutation. Additionally, we assess how well these models

represent root-to-tip divergence in time-stamped RABV sequences. We find that at low

substitution rates (<1 substitution per genome per generation) divergence patterns between

these models are difficult to distinguish, while above this threshold differences become

apparent across a range of sampling rates. Using a Tanzanian RABV dataset, we calculate the

mean substitution rate to be 0.17 substitutions per genome per generation. At RABV's

substitution rate, the per-generation substitution model is unlikely to represent rabies

evolution substantially differently than the molecular clock model when examining

contemporary outbreaks; over enough generations for any divergence to accumulate, extreme

incubation periods average out. However, measuring substitution rates per-generation holds

potential in applications such as inferring transmission trees and predicting lineage

emergence.

Author Summary

Rabies is a neglected disease that kills around 60,000 people each year. After entering the

body, the incubation period of the virus is usually less than one month, but can sometimes

span months to years. While we normally assume a virus accumulates mutations at a constant

rate, it is possible that rabies' occasional long incubation periods mean that mutations

accumulate at varying rates if the virus replicates (and thus mutates) more slowly during the

incubation period. We compared how the rabies virus evolves over time using two simulation

models where mutations either occur per unit time or per infection generation. We also

calculated the mean substitution rate per infection generation, which can be useful for

inferring linkage between related rabies cases. We found that at realistic substitution rates for

the rabies virus, we could not distinguish between the two models. Our calculations show that

in most generations no mutations are expected to occur. Thus, over a time period long enough

to observe genetic divergence, occasional long incubation periods would be "cancelled out"

by shorter than average incubation periods, meaning that the two models are almost

equivalent. However our work suggests that modelling substitution rates per generation may

be useful for epidemiological inference.

Introduction

The molecular clock hypothesis assumes that the genomes of organisms accumulate neutral

mutations at a constant rate over time, either across all lineages (the "strict molecular clock")

or within each individual lineage but with some degree of variation between them (clock

models with this assumption include the relaxed and multirate clock models) (1–3). The

ability to sample viral sequences through time, and the application of the molecular clock

hypothesis to these sequences, has led to massive advances in using viral genetic data to

investigate disease outbreaks (4). The clock rate, measured in substitutions per site per unit

time, can be used to estimate how long ago pathogens diverged (5), and the date of infection

of individual infected hosts (6). Combining the analysis of epidemiological and genetic data

has allowed further insights into the history of outbreaks (7), and the introduction of

geographic data provides estimates as to rates of spread and the frequency and source of

introductions (8,9). However, in order to conduct these phylogenetic analyses, genetic

divergence must increase appreciably over time in the dataset under investigation (10).

Whether or not the viral population is measurably evolving, and thus whether it contains

sufficient temporal signal for phylogenetic analysis, depends mainly on the evolutionary rate,

the sequence length and the length of time sequences are sampled over being sufficiently

high. Various methods exist to assess temporal signal, the most commonly used being root-to-

tip divergence plots (11,12) implemented in tools such as TempEst (13) , but these also

include Bayesian evaluation of temporal signal (BETS) (14) and the date-randomisation test

(15).

The rabies virus (RABV) is a negative-strand RNA virus, with a genome size of

approximately 12 kilobases. While RNA viruses generally have high mutation rates due to a

lack of proofreading by RNA polymerases, RABV has a substitution rate at the lower end of

normal for single-stranded RNA viruses of between 1×10^{-4} and 5×10^{-4} substitutions per site

per year (16–18). This may be due to strong purifying selection (16), or due to peculiarities of

RABV. For example, the RABV genome is longer than average for RNA viruses, and genome

length and evolutionary rate are negatively correlated (19), although this relationship appears

to be weaker in single-stranded RNA viruses (20). A more unusual feature of RABV is that

infections can exhibit extended incubation periods within the host. The median generation

interval (the time between one individual becoming infected and then infecting another) is

estimated to be 17.3 days in domestic dogs (21), with other studies estimating mean serial

intervals of 26.3 days (22) and 45.0 days (23). Symptoms, infectivity, and death from rabies,

however, can occasionally occur years after the initial infection event (24). The length of the

incubation period is influenced by the route of exposure, with bites to the head and neck

leading to more rapid disease progression than bites to lower extremities (25). RABV can

remain in the muscle at the bite site for prolonged lengths of time before invading the host's

motor neurons and progressing through the nervous system, with limited, if any, infection of

other muscle fibres (26). While some replication in the muscle cells has been observed (27),

RABV replication at the inoculation site is not necessary for neural invasion (28). It is

currently unknown precisely how the RABV replication rate in the host muscle cells and

peripheral nervous system compares to the massive replication rate within the cells of the

central nervous system and brain. However, work suggests that RABV replication in muscle

cells may be reduced (29), and RABV replication in cultured rat sensory neurons may be 10-

to 100-fold lower than replication rates in rat and mouse CNS neurons (30). Rabies infections

that involve long incubation periods may, therefore, not lead to more accumulated mutations

than those with shorter incubation periods, as viral mutation is strongly influenced by the

replication process (31).

Changes in mutation rates through time due to long incubation periods may affect how we

analyse RABV sequence data and interpret these analyses. A relaxed molecular clock is

usually required to carry out phylogenetic analyses on rabies datasets, and it is not

uncommon for there to be difficulties in applying these analyses due to "insufficient temporal

signal"; usually referring to either no or a negative relationship between genetic divergence

and time, or this relationship having a lot of noise and a very low R^2 (32–36). RABV shows

variation in substitution rate between lineages (18,37,38) which may be driven in part by

differences in incubation periods. If the variable incubation period of rabies infections does

cause deviation from the molecular clock model (exceeding the variation captured by relaxed

or multirate clock models), this may negatively affect the accuracy of time-scaled

phylogenetic trees and emergence date predictions. Conversely, if mutation does continue at a

consistent rate during the incubation period, attention should be paid to extremely long

incubators which could drive the emergence of new variants, as seen recently in chronic

SARS-CoV-2 infections (39,40).

We hypothesised that reduced replication (and thus mutation) during the incubation period

could cause rabies evolution to be better represented by a per-generation model of mutation

than by the molecular clock model. We aim to clarify the nature of contemporary RABV

evolution using in silico methods, comparing the root-to-tip divergence of sequences

generated from synthetic outbreaks under per-unit time or per-generation mutation models,

and comparing these to RABV genomic data from Tanzania. We also aim to calculate a per-

generation substitution rate for RABV for future use as a parameter in transmission tree

reconstruction algorithms.

Methods

We investigate two contrasting mutational models for RABV – i.e., substitutions occurring on

a per-generation vs. per-unit-time basis – using a simulation approach. We first generated

synthetic RABV outbreaks using a branching process model (21) and then simulated these

two mutation processes over the resulting transmission trees. From the synthetic sequences

generated, we examined root-to-tip divergence and calculated variance explained (R^2) from

linear regressions, and compared these to the root-to-tip divergence of a set of RABV whole

genome sequences from Tanzania. Finally, we developed a method to estimate the per-

generation substitution rate for RABV and tested this on synthetic data before applying it to

the Tanzanian RABV dataset.

<u>Rabies outbreak simulation</u>

We simulated RABV mutation on branching-process simulations of rabies outbreaks.

Outbreaks were simulated 100 times over a spatially explicit representation of Mara Region

in northern Tanzania. In Serengeti District, where contact tracing data were available, the

model was initialised with the three cases that occurred in the mean generation interval (g=27

days, based on contact tracing data) prior to 2017 (simulations were run over a dog

population representing that in Mara region between 2017 and 2024). In the rest of Mara

region, where there were no data to guide initialisation, we seeded with (0.01Dg)/365 cases,

where D is the initial dog population in that area. If R_e=1 (endemic transmission), this results

in roughly 1% of the population becoming rabid over a year; contact tracing data suggest that

incidence typically does not exceed that level (41). This led to a total of 273 initial cases in

the region. Each case was assigned a number of offspring cases drawn from a negative

binomial distribution (41) with mean (R_0)=1.05 and dispersion parameter=1.33. The R_0 value

was chosen to result in a median number of cases each month that was roughly constant over

time (over the 100 simulations), mimicking endemic disease. Movement of rabid dogs from

their home locations to and between transmission locations followed a random walk with step

lengths drawn from a Weibull distribution (shape=0.41, scale=0.13). We simulated occasional

long-distance transport of dogs to a random location prior to their first transmission in 2% of

cases (21). At each of a rabid dog's transmission locations, another dog was randomly

selected within the local 1km2 grid cell. If this dog was susceptible (i.e., not vaccinated or

already incubating infection from a prior transmission event), rabies was transmitted. A

generation interval was drawn for each new infection from a lognormal distribution

(meanlog=2.96, sdlog=0.82), describing the time delay before it also became rabid and made

its assigned transmissions. The step-length and generation-interval distributions were fitted

using contact tracing data from Serengeti District, Tanzania (21). Branching process

simulations were continued until 7 years had passed or rabies went extinct. Each synthetic

case was assigned an individual ID, and for every case (except initial seed cases) we recorded

the ID of the associated progenitor case. Dates of infection and transmission were recorded

for each case.

We isolated complete transmission trees descending from each of the 273 initial cases from

within one randomly selected synthetic outbreak. Transmission trees that contained over 100

cases (9 out of 273 trees in total, that ranged in size from 533 - 19,382 cases) were then used

to generate synthetic sequence data. Across these trees, we see a mean generation interval of

26.6 days, and 2.5 and 97.5 percentiles of 3.90 and 94.11 days (Supplementary Figure S1).

For each of the 9 trees the index case was assigned an initial 12kb genome sequence. Under

the per-unit time mutation model, we determined the expected number of mutations by

multiplying the substitution rate, the genome length and the length of the generation interval,

for each case along the resulting transmission tree (because we assume mutations are neutral,

the individual-level mutation rate is the same as the population-level substitution rate). The

realised number of mutations was then drawn from a Poisson distribution, with this mean. We

then randomly chose positions to change and new nucleotides to change them to. The

resulting synthetic sequence data is referred to as the "time-based sequence data". The

generation-based model of mutation works as above, with the exception that the expected

number of substitutions in a generation is constant and produces the synthetic "generation-

based sequence data".

<u>Divergence rate analysis</u>

To investigate patterns of temporal divergence under the mutation models described above,

we generated synthetic data with values of substitution rates ranging from 0.05 to 3

substitutions per genome per generation (or the per unit time substitution rate equivalent) and

4 population sampling regimes (from 1% of cases to 20%, informed by a previous study that

estimated that routine surveillance for rabies rarely confirms more than 10% of circulating

cases (42)). We calculated the genetic divergence as the number of nucleotide differences

from the index case to each sampled case. For each of the nine transmission trees, we then

compared genetic divergence with time under each scenario (substitution rate and sampling

regime combination), using linear regression through the origin.

In order to compare our synthetic patterns of divergence over time to real rabies data, a root-

to-tip divergence plot was also generated for a dataset of real RABV sequences (data from

(43); Figure 1A) using TempEst (v1.5.3 (13)), with the best-fit root located (Figure 1B).

These rabies cases occurred between 2001 and 2017 and were primarily from the Serengeti

district and Pemba Island, with the remaining sequences from elsewhere in Tanzania (Figure

1A inset). Sequence acquisition and tree building methods are detailed in (43).

Fig 1. Phylogeny of real rabies virus whole genome sequences from Tanzania and root-to-tip divergence. (A) The time-scaled tree (43) used to generate the root-to-tip divergence plot and to calculate the per-generation substitution rate. The inset map shows the approximate locations that the samples were collected from, and the lineages present in each location. Map point size represents the number of sequences in this dataset from district centroid locations. Base map data is from Natural Earth (naturalearthdata.com), via the *maps* R package. **(B)** The corresponding root-to-tip divergence plot. Point colours represent RABV lineage.

Calculating the per-generation substitution rate

We updated a method of calculating the per-generation substitution rate previously used in

eukaryotes (44) by using Bayesian posterior estimates of the clock rate and the generation

interval. We assessed this method's accuracy using the synthetic outbreak sequence data,

before applying it to the aforementioned set of RABV whole genome sequences.

To estimate the mean per-generation substitution rate, we analysed sequence data with

BEAST, and multiplied the posterior rate estimate for each MCMC sample (excluding the

burn-in period) by the generation-interval lengths sampled from the posterior of a simple

Bayesian analysis and then multiplied again by the genome length. The mean and 95%

credible interval of the estimate of the per-generation substitution rate for the RABV dataset

was calculated by taking the mean and the 2.5% and 97.5% percentiles of the resulting

multiplied posteriors.

To evaluate the accuracy of this method in estimating the mean per-generation substitution

rate, we also applied it to synthetic sequence data generated from outbreaks using the per-

generation mutation model as described above, under different substitution rates (11 values

ranging from 0.05 substitutions per generation to 3 substitutions per generation) and case

sampling rates (1%, 5%, 10% or 20% of cases sampled) across the 9 transmission trees that

contained at least 100 cases. Subsampled synthetic datasets containing more than 2000

sequences were not analysed as this number exceeds the total whole-genome RABV

sequences currently available on the RABV-GLUE database (45), and is unrealistic in the

context of examining individual rabies outbreaks. BEAST log files were generated from these

sequences using BEASTGen version 1.0.2 and BEAST version 1.10.4 (46). We chose to use a

JC substitution model with a strict clock, no site heterogeneity due to our per-generation

mutation model used in the simulations having equal probability of any site or base being

chosen and assumed constant population size. We used a tracelog frequency of 1000 and a

sufficiently long chain length for the effective sample size (ESS) of each parameter to exceed

200 when analysed using Tracer (47), and a 10% burn-in period. We applied the substitution

rate calculation method to these phylogenetic trees, and assessed the accuracy of the resulting

mean per-generation substitution rates by comparing them to the parameter values used to

generate the synthetic sequences, using the natural log of the ratio:

$$Deviation = \ln\left(\frac{M_e}{M_a}\right)$$

where M_e is the mean estimated per-generation substitution rate and $M\square$ is the actual

substitution rate, where a deviation of zero means perfect accuracy.

The same method was applied to the dataset of 153 RABV sequences sampled from across

Tanzania (data from (43); Figure 1A). The mean per-generation substitution rate was

calculated, and distributions were fitted from the multiplied generation interval and clock rate

posteriors (generation interval posteriors based on values from (21) for the Tanzanian dataset,

extracted directly from the lognormal distribution used in simulations, and clock rate

posteriors taken from the BEAST log file of the time-scaled tree from Lushasi et al. (43)) and

genome length as described above. We compared different distributions (Gamma and

Lognormal) for estimating substitution rates and selected the best fitting distribution by AIC.

We also calculated the probabilities of between 0 and 10 SNP differences occurring across 1,

5 or 10 infection generations. For this calculation we simulated mutations arising at a Poisson

rate with lambda drawn from the fitted substitution rate distribution. The means and 95%

confidence intervals were calculated from the 10,000 simulations.

<u>Data and code availability</u>

All code is available at https://github.com/RowanDurrant/Rabies-Mutation. Analyses were

conducted using the R programming language (48). The beta regression curve and prediction

interval in Figure 2C was generated using the 'betareg' R package (49). RABV lineages were

assigned using MADDOG (45).

Results

<u>Root-to-tip divergence analysis</u>

At higher per-generation substitution rates (1 substitution per genome per generation and

above), distinct differences can be seen between root-to-tip divergence plots from the two

models of mutation (Figure 2A). The synthetic data generated from the per-generation

mutation model shows "stray" clusters or ridges of points both above and below the main

funnel of points, illustrated in the example in Figure 2A. Divergence plots from synthetic data

generated from the time-based model of mutation have less variance and do not exhibit this

pattern. At lower substitution rates (below 1 substitution per generation), no such pattern is

clearly distinguishable (Figure 2B). When the cases represented by the high-divergence

points from the per-generation model in Fig. 2A are visualised in a transmission tree, they are

mainly confined to a single chain (Supplementary Figure S1).

Fig. 2: Temporal genetic divergence varies under two models of mutation. (A) Root-to-tip divergence plots for synthetic sequences produced using time-based and generation-based mutation models, equivalent to 2 substitutions per genome per generation and **(B)** equivalent to 0.2 substitutions per genome per generation. Note that the y-axis scales differ by an order of magnitude between A and B. These data are from running mutation models over the same single transmission tree and have a case sampling rate of 5% (i.e., 621 cases sampled of 12,434 total). **(C)** The R^2 values obtained from regression through the origin of root-to-tip divergence of synthetic data from the time-based and generation-based models. Point colour indicates the mutation model used to generate the data. Lines represent beta regressions with logit links fit to data points, and shading represents the 95% prediction interval. The X axis is

log scaled. 5% of cases were sampled here; sampling rate had little effect on R^2

(Supplementary Figure S2).

Root-to-tip divergence plots derived from synthetic transmission trees using the time-based

mutation model had, on average, higher R^2 values than those from synthetic transmission

trees using the per-generation mutation model, although this is more difficult to distinguish

below a substitution rate of 0.5 substitutions per genome per generation (Figure 2C). As the

substitution rate increases, the R^2 values across both mutation models increase. The case

sampling rate appears to have little effect on R^2 (Supplementary Figure S2).

The root-to-tip divergence plot of the Tanzanian RABV dataset more closely resembles those

of lower substitution rate simulations, where it is difficult to determine any difference

between the models of mutation (Figure 1). While most lineages surround the regression line,

some (for example, Cosmopolitan AF1b_B1) group below the line, but without forming a

distinguishable "ridge".

<u>Substitution rate calculation</u>

The accuracy of our method used to calculate per-generation substitution rate remains similar

at all but the lowest values of substitution rate (Figure 3), with a tendency to underestimate

the substitution rate (meaning that the estimated substitution rate is below the substitution

rate parameter used to generate the synthetic data; mean natural log of the ratio of -0.18 and

root-mean-square of 0.54, where values of 0 indicate perfect estimates). Accuracy appears to

be more influenced by the number of sequences used in the BEAST analysis than by the case

sampling rate itself; the mean natural log of the ratio falls to -0.36 when fewer than 50

sequences are used (root-square-mean of 0.95).

Fig. 3: Accuracy of per-generation substitution rate predictions for different numbers of sequences, substitution rates and sampling rates. Facets indicate case sampling rate. The dotted line represents perfect accuracy. X axis and colour scale are log transformed.

The Tanzanian RABV dataset from which we estimated the per-generation substitution rate contains 153 sequences in total, and the accompanying time-scaled phylogenetic tree has a root-to-tip height of approximately 65 years, although the sequences spanned just 16 years as they were sampled from 2001 to 2017 (with 46.7% from years 2011-2012). These sequences were largely complete; 98% of sequences were >95% complete (>11,327 kb in length). The

mean per generation substitution rate of RABV in this dataset was estimated to be 0.171

(95% credible interval: 0.127 - 0.219). The best fitting distribution by AIC to the output of

the multiplied Bayesian posteriors was a Gamma distribution with the parameters shape =

51.69 and rate = 301.8.

Using the calculated per generation per genome substitution rates, we calculated the

probability of different numbers of substitutions occurring over 1, 5 and 10 generations,

drawing the per-generation substitution rate (λ) from the above distribution (Figure 4). Over

many generations it is still quite likely for zero substitutions to occur; after 10 generations,

the probability of zero substitutions having occurred is 0.19.

Fig. 4: Probability distributions of the mean per-generation substitution rate and

substitutions occurring over generations. **(A)** estimated probability distribution of the per

genome per generation substitution rate from Tanzanian RABV sequences, with underlying

histogram of multiplied Bayesian posteriors of clock rate and generation interval. **(B)**

probability distribution of SNPs occurring over 1, 5 and 10 generations. The λ value for a

Poisson rate of SNP occurrence is drawn from the SNPs per generation distribution fitted in

Figure 4A. Black bars represent the 95% confidence intervals (which are very tight).

Discussion

It can be difficult to get sufficient temporal signal for RABV sequence datasets, which we

hypothesised could be due in part to its variable incubation periods. We hypothesised that a

per-generation model of mutation may be more representative of RABV evolution than a

purely time-based model. We found that substantial differences in root-to-tip divergence

patterns between synthetic outbreaks using generation-based and time-based models of

mutation could be observed only at high underlying substitution rates. The substitution rate

for the Tanzanian RABV sequences examined (~0.17 substitutions per genome per

generation) was in the range where divergence patterns in the two models were extremely

similar. We can thus assume that the two models will give extremely similar results on the

relevant time scale. As we observed increasing divergence over time with reasonable R^2

values within this substitution rate range, it implies that variable incubation periods alone do

not fully account for the challenge in obtaining temporal signal. Therefore, other factors such

as insufficiently long sampling windows for the substitution rate are likely to be responsible

(50). This is an important consideration for analysing RABV sequences from new outbreaks,

or from endemic areas where sampling is opportunistic. As RABV has a substitution rate

lower than many other RNA viruses, longer sampling windows are required to achieve a

sufficient temporal signal.

The observation of little difference between root-to-tip divergence plots derived from the two

mutation models at substitution rates below 1 substitution per genome per generation is likely

because of averaging; multiple generations of infection are expected to have occurred per

substitution that arises on the viral genome. Over the many generations needed before

significant levels of viral genetic diversity are reached, the influence of any unusually long

incubation periods will be damped by the opposite influence of unusually short incubation

periods, eventually becoming indistinguishable from clock-like behaviour. On the other hand,

at higher substitution rates ridges form on the root-to-tip divergence plots under the per-

generation model of mutation but not under the per unit time model. While not affecting the

overall clock rate, these ridges reduce the overall R^2, and may be better analysed using a

separate local clock (51). The cases in these ridges almost all descend from a common

ancestor (Supplementary Figure S1), suggesting that a single unusually long or short

incubation period can affect which phylogenetic analyses we perform. Ridges caused by these

incubation periods can be distinguished from ridges caused by rate variation between lineages

as they will be parallel to the main cluster of points in the plot, whereas points belonging to

lineages with a different substitution rate will have a different slope. Studies examining the

number of substitutions occurring between successive sequenced cases, and whether this

increases when the secondary case's incubation period is unusually long, could clarify the

exact relationship between substitutions, generations, and time. More detailed data will be

required to investigate this further.

We calculated RABV's mean per-generation substitution rate to be approximately 0.17

substitutions per genome per transmission generation. This estimate is lower than those for

other RNA viruses, such as SARS (2 substitutions per genome per human passage (52)),

SARS-CoV-2 (0.52 substitutions per genome per 5.8-day generation interval (53)) and Ebola

virus (0.875 substitutions per genome per 14-day generation interval (54)). RNA viruses that

undergo periods of reduced replication or complete latency often show reduced substitution

rates, with one extreme example being HTLV-1/2 (55,56). However, we would not expect this

to affect the per-generation rate. The low per-generation substitution rate seen in rabies is

therefore likely due to mutation being constrained by other factors, such as strong purifying

selection (16), and likely contributes to the difficulties in obtaining sufficient temporal signal

for phylogenetic analyses. Previous studies suggest that for viruses in this substitution rate

range, sampling windows of up to 30 years may be required to overcome the phylodynamic

threshold (15); for comparison, SARS-CoV-2 achieved sufficient temporal signal within two

months of the start of the pandemic (50).

We can predict from the estimated per-generation substitution rate that identical sequences

are likely to have less than 5 intermediate generations between them (probability of fewer

than five generations occurring before a mutation occurs > 0.49 by repeated sampling of a

Poisson distribution with a lambda of 0.17), but still have a non-negligible probability of

being more distantly related. While the low substitution rate means that comparing the

number of SNPs between sequences alone may not be an effective method of determining

infector-infectee relationships, it could be used in conjunction with temporal and location

data to make more accurate predictions of transmission events by ruling out relationships

between more distantly related transmission chains co-circulating in the same area, as in (57).

Notably, our Poisson distribution of the number of substitutions occurring in one generation

is visually very similar to the genetic signature distribution reported in Cori *et al*. ((57), Fig

S1), despite different methods and RABV datasets being used in their calculations. It is likely,

however, that our estimate of the per-generation substitution rate is lower than the mean

number of SNPs expected between sequences from a primary and secondary case, due to the

time-based substitution rate being affected by purifying selection (58). Further analysis

comparing the estimated per-generation substitution rate to realised SNP distances between

primary-secondary case pairs could quantify this difference.

While the Jukes-Cantor model was the most appropriate to use on our synthetic data due to

the simplicity of the mutation models, phylogenetic analyses on real RABV genomes usually

use a more complex model, such as the GTR + G substitution model used to generate the

Tanzanian tree shown in this study (43). This, along with the simplicity of our mutation

model as well as sampling biases in the real dataset, may affect how comparable synthetic

root-to-tip divergence plots are to the real data.

While the molecular clock has proven critical for gaining insights into the history and

dynamics of disease outbreaks, the epidemiological characteristics of a virus should be

considered when choosing how to measure viral evolution. In this study, we determine that

the per-generation model is not likely to produce substantially different results from the

molecular clock model when analysing contemporary RABV evolution. We also estimate the

mean per-generation substitution rate of RABV for future use in transmission tree

reconstruction and efforts to estimate outbreak sizes and lineage emergence rates. Given that

many different lineages circulating simultaneously is seemingly a common occurrence in

areas with endemic rabies, it is important to investigate whether these lineages vary in

evolutionary rate and generation interval length, and ascertain the potential effects on

phylogenetic analyses.

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Supporting information captions

Supp. Fig. S1: histogram of generation intervals from the simulated outbreaks. Vertical

dashed lines represent the median (blue) and mean (red) generation interval.

Supp. Fig. S2: points in the offshoot ridge predominantly occur in one transmission tree.

(A) root-to-tip divergence plot (2 SNPs/genome/generation, 5% of cases sequenced) with

offshoot ridge points highlighted in red. Offshoot ridge points are defined in this plot as

having a divergence rate above 8×10^{-6} substitutions/day and occurring after day 750. **(B)**

transmission tree of the simulated outbreak with offshoot ridge cases highlighted in red.

Graph edge length is not proportional to time or divergence.

Supp. Fig. S3: Sampling rate does not impact the R^2 of root-to-tip divergence plots from

synthetic data. Plot is faceted by the proportion of the total number of cases in the outbreak

sequenced, point colour represents mutation model.

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