Economic analysis of disease and control of multi-field epidemics in agriculture

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Abstract

Epidemics of plant diseases are estimated to cause significant economic losses in crop production. Fungicide applications are widely used to control crop diseases but incur substantial indirect costs. One essential class of indirect costs arises due to the evolution of fungicide resistance. This indirect cost must be estimated reliably to design economic policy for more sustainable use of fungicides. Such estimation is difficult because the cost depends on economic parameters and the epidemiological/evolutionary properties of crop pathogens. Even a conceptual framework for such estimation is missing. To address this problem, we combined a spatially implicit mathematical model of crop epidemics with an economic analysis at the landscape scale. We investigated how the net economic return from a landscape depends on the proportion of fungicide-treated fields. We discovered a pattern of accelerating (or decelerating) returns, contrary to expected diminishing returns. Next, we calculated the economic cost of the evolution of fungicide resistance as the difference between the optimal net return of the landscape in the absence and presence of resistance. We found that this cost depends strongly on the fungicide price, the degree of resistance, the pathogen's basic reproduction number and the yield loss due to disease. Surprisingly, the cost declines with the fungicide price and exhibits a non-monotonic pattern as a function of the basic reproduction number. Hence, to calculate the cost, we must estimate these parameters robustly, incorporating variations in environmental conditions, crop varieties and the genetic composition of pathogen populations. Appropriate estimation of the

cost of resistance evolution can inform economic policy, design dynamic fungicide pricing, and encourage more sustainable use of fungicides.

1 Introduction

Epidemics of plant diseases are estimated to cause significant economic losses in crop production (Ostfeld et al., 2005; Oerke, 2006; Savary et al., 2019; Global Plant Health Assessment, 2023). To reduce losses, farmers frequently apply fungicides (Global Plant Health Assessment, 2023). They can maximize the net economic return of fungicide applications, e.g., by optimizing the fungicide dose based on the balance between the yield benefit and the cost of fungicide application (Jørgensen et al., 2017; Hillebrandt, 1960). However, this approach is problematic for at least two reasons.

First, fungicide applications not only have direct economic costs, but they also incur considerable indirect costs, such as environmental costs, human health costs, and costs associated with the evolution of fungicide resistance (Pimentel et al., 1980; Pimentel, 2005; Bourguet and Guillemaud, 2016). One opportunity to manage crop diseases in a more sustainable manner is to incorporate indirect costs at least partially into fungicide prices. This opportunity can be realized using economic policy instruments, e.g., by introducing special pesticide taxes or subsidies, as has been done in several countries, including France, Denmark, Norway and Sweden (Finger et al., 2017; Lee et al., 2019). However, it is not easy to design pesticide tax rates or subsidies to reflect indirect costs, as indirect costs are notoriously difficult to estimate (Pimentel et al., 1980; Pimentel, 2005; Bourguet and Guillemaud, 2016).

Second, many crop pathogens of economic relevance disperse over long distances: not only within individual fields but also between fields and across entire regions (Brown and Hovmoller, 2002). Therefore, decisions regarding disease management made at a particular farm can affect epidemic development in other farms across a region; and decisions that are optimal for an individual farm may turn out to be sub-optimal at the scale of a regional cultivated landscape.

An essential class of indirect costs of fungicide applications arises because pathogen populations can adapt to fungicides via the evolution of fungicide resistance, which reduces fungicide efficacy (Fisher et al., 2018). From an economic perspective, these costs represent an externality (Ayres and Kneese, 1969): development of resistance is caused by fungicide applications on specific fields, but resistance can spread between fields and across farms and entire regions. Hence, a much wider community of farmers will eventually suffer the economic costs of fungicide resistance evolution.

These costs are difficult to estimate because they depend not only on the economic parameters of the affected cropping system but also on epidemiological and evolutionary properties of a particular pathosystem. Rough estimates of economic losses in crop production due to insecticide resistance were reported to be in the range of 10-15 % (Pimentel et al., 1980; Pimentel, 2005). These were based on economic data on cotton production in the USA, assuming that insecticide use increases due to resistance and that this increase is the only cost incurred by resistance. However, it is not clear to what extent this assumption is fulfilled, it is also difficult to extrapolate these estimates to other crops and regions, and to extrapolate from insecticide resistance to fungicide resistance. There are no calculations of the economic costs of fungicide resistance in crop pathogens in the literature, and even a conceptual framework to enable such a calculation is lacking. To fill this knowledge gap in this study, we addressed the following questions: (i) "How to calculate the economic cost of fungicide resistance?" and (ii) "How does the economic cost of fungicide resistance depend on the economic and epidemiological/evolutionary parameters of the crop-pathogen system?"

For this purpose, we combined a spatially implicit epidemiological model of crop epidemics with an economic analysis at the landscape scale (bio-economic modeling). Using this approach, we calculated the economic cost of the evolution of fungicide resistance, C_R, as the difference between the optimal net return of the landscape in the absence of resistance and in the presence of resistance. This allowed us to explore how C_R depends on the key epidemiological/evolutionary and economic parameters of the system. Our results provide a conceptual basis for estimation of C_R as a key component of indirect costs of fungicide applications, which can inform economic policy to achieve a more sustainable use of fungicides.

Figure 1: Overview of the multiple field model of plant disease incorporating both epidemiology and economics aspects in the presence of fungicide treatment. Healthy fields (H_u and H_t) are represented in green, while fields infected with wildtype pathogenic fungi (I_{uw} and I_{tw}) are shown in red. Fields infected with resistant pathogenic fungi (I_{ur} and I_{tr}) are represented in orange. Fields treated with fungicides are marked with a light-blue glow around them. (A) Out of a total of N fields, initially, only one is infected, while the rest are healthy. A fraction of θ of all fields are covered with fungicide which comes at a price f. As time progresses, the plants grow and healthy fields become infected, as described in panel (B). The net return ($g(\theta)$) of the total harvest is derived from the yields from healthy and infected fields (1 and y respectively). (B) Figure illustrating the different states within the extended susceptible-infected (SI) epidemic model of plant disease and the corresponding rates of transition between these states.

2 Materials and methods

2.1 Epidemiological model for multiple fields

We study the epidemiological dynamics of a generic fungal pathogen of crop plants at the scale of multiple fields (Fig. 1). We consider a regional cultivated landscape composed of N fields growing the same annual crop (e.g., wheat or maize). The key variables and parameters are given in Table 1. First, we devise a model considering a wild-type pathogen strain with a fungicide treatment. Then, we extend this model to incorporate a fungicide-resistant pathogen strain.

2.1.1 Model with fungicide treatment

Each field can be either healthy (H) or infected (I) with a wildtype pathogen (subscript w) and each field can be treated (subscript t) or untreated (subscript u) with the fungicide. Hence, we have four possible field states: H_u, H_t, I_{uw}, and I_{tw}. H_u and I_{uw} represent the numbers of untreated healthy and infected fields, respectively, while H_t and I_{tw} represent the treated fields. The dynamics of the disease are captured by equations that are similar to the epidemiological susceptible-infected (SI) model:

Table 1: Key variables and parameters of the model

Symbol	Description
Variables	
H_i	Number of healthy fields with treatment status i (u = Untreated, t = Treated)
I_{ij}	Number of fields infected by j type (w = Wildtype, r = Resistant) pathogen with treatment status i (u = Untreated, t = Treated)
Parameters	
N	Total number of fields
θ	Fraction of fungicide-treated fields (or fungicide coverage); 0-1
ϵ_i	The efficacy of fungicide treatment in reducing the transmission rate of the i type (w = Wildtype, r = Resistant) pathogen; 0-1
β_i	The rate of i type (w = Wildtype, r = Resistant) pathogen transmission from infected to the healthy field
μ	Recovery rate of infected fields to become healthy fields
c	The price of fungicide application including both direct and indirect cost
y_H	Yield from a healthy field per season
f	Relative fungicide price with respect to yield from a healthy field
y	Relative yield of a diseased field with respect to the yield from a healthy field
Y	Total yield of the landscape (expressed as a fraction of the monetary equivalent of the maximum yield)
$g(\theta)$	Net return relative to yield from a healthy field, y_H
θ^*	Optimal fraction of treated fields that maximizes the relative net return, $g(\theta)$

$$
\begin{aligned}
\dot{H}_u &= -\beta_w[I_{uw}H_u + I_{tw}H_u] + \mu I_{uw} \\
\dot{H}_t &= -\beta_w(1-\epsilon_w)[I_{uw}H_t + I_{tw}H_t] + \mu I_{tw} \\
\dot{I}_{uw} &= \beta_w[I_{uw}H_u + I_{tw}H_u] - \mu I_{uw} \\
\dot{I}_{tw} &= \beta_w(1-\epsilon_w)[I_{uw}H_t + I_{tw}H_t] - \mu I_{tw}
\end{aligned}
\tag{1}
$$

Given ϵ_w as the fungicide efficacy in transmission reduction, β_w represents the rate at which the wildtype pathogen strain spreads from infected fields to healthy ones. The rate at which fields recover from the disease is set to μ. The fraction of fields treated with fungicide is $\theta = (H_t + I_{tw})/N$. Hence, the number of treated and untreated fields can be calculated as follows:

$$
\begin{aligned}
H_u + I_{uw} &= N(1-\theta) \\
H_t + I_{tw} &= N\theta
\end{aligned}
\tag{2}
$$

We derive the expressions for the fixed points of the dynamics system in Eqs (1), helping us gauge the long-term behavior. Details of the derivation and expressions for the three fixed points are provided in Appendix A.2.1. Of the three fixed points, two are biologically reasonable. Fixed point 1 corresponds to the disease-free equilibrium (no infected fields), and all fields are healthy (either treated or untreated). Fixed point 2 corresponds to the endemic equilibrium, showing the coexistence of healthy and infected fields. Which of these fixed points is realized in the long term depends on the basic reproduction number of the pathogen, R_0 (Anderson and May, 1986; Heffernan et al., 2005). In our case, it can be interpreted as the average number of secondarily infected fields produced by a single infected field.

First, we consider the no-treatment scenario, $\theta = 0$. The basic reproduction number is then given by:

$$
R_0 = \frac{\beta N}{\mu},
\tag{3}
$$

as derived in Appendix A.1. Here, we have two fixed points (see Appendix A.1). The disease-free equilibrium is stable if $R_0 < 1$, i.e. an infected field with pathogenic fungi cannot invade

and will die out if introduced. Conversely, if $R_0 > 1$, the infected field can invade and persist, leading to an endemic equilibrium. The number of secondarily infected fields increases when $R_0 > 1$, decreases when $R_0 < 1$ or remains constant when $R_0 = 1$. Thus, a higher value of R_0 indicates a higher potential for the pathogen to spread.

Including fungicide treatment ($\theta > 0$) as in Eqs. (1), there are two infectious compartments: I_{uw} and I_{tw}. Hence, a reevaluation of the R_0 is necessary. In the limiting case of a high efficacy fungicide ($\epsilon_w = 1$), we derived the invasion properties using the Routh-Hurwitz stability criterion in Appendix A.2.2. In the more general case of arbitrary fungicide efficacy (ϵ_w), R_0 is derived using the next-generation matrix method as in Appendix A.2.3 (Diekmann et al., 1990; van den Driessche, 2017). As a result of this calculation, the *effective* reproduction number can be expressed in terms of R_0, θ, ϵ_w:

$$\mathscr{R}_0 = R_0(1 - \epsilon_w \theta). \tag{4}$$

This derivation quantifies how fungicide treatment affects the pathogen's capacity to invade and spread. According to Eq. (4), the fraction of fungicide-treated fields, θ, and the fungicide efficacy, ϵ_w, have an equal and positive effect in suppressing $\mathscr{R}_0$. Furthermore, higher θ-values are required to prevent the invasion by the pathogen with a higher R_0.

The critical fungicide coverage, for which $\mathscr{R}_0 = 1$ can be calculated as:

$$\theta_{cw} = \frac{1}{\epsilon_w}\left(1 - \frac{1}{R_0}\right). \tag{5}$$

Treating the fields beyond the critical coverage level ($\theta > \theta_{cw}$) prevents an epidemic. This criterion depends on the fungicide efficacy ϵ_w and the pathogen's R_0 without treatment. Thus for a given R_0, when $\theta > \theta_{cw}$, fungicide treatment will suppress $\mathscr{R}_0$ to a value below one if ϵ_w is high enough, i.e. if $\epsilon_w > \epsilon_{cw} = 1 - 1/R_0$. However, at a low fungicide efficacy $\epsilon_w < \epsilon_{cw}$, even treating all fields ($\theta = 1$) will not prevent an epidemic as $\mathscr{R}_0$, will remain higher than one. Similarly, for a fungicide with a given efficacy ϵ_w, increasing θ to values above the critical, θ_{cw}, will drive the pathogen to extinction only for pathogens with $R_0 < R_{0cw}$, where

$$R_{0cw} = 1/(1 - \epsilon_w). \tag{6}$$

The expressions for the critical R_{0cw} and θ_{cw} are helpful already in the epidemiological sense of the problem of fungicide treatment; however, we see that they play a crucial role when performing a bioeconomic analysis below in Sec. 3.1. However, we must regard an evolutionary quantity before considering a comprehensive bioeconomic analysis. In reality, the repeated use of fungicides can lead to the evolution of resistance in pathogen populations (Mikaberidze et al., 2014, 2017; Fisher et al., 2018), which we view next.

2.1.2   Model with fungicide treatment and resistance

We incorporate a resistant pathogen strain that could be fully or partially protected from the fungicide. We assume that either a wildtype (w) or a resistant (r) pathogen strain dominates each infected field (Parnell et al., 2006). Resistance reduces fungicide efficacy. Consequently, fungicide efficacy against the resistant strain is lower than its efficacy against the wildtype strain. However, the transmission rate of the resistant strain can be reduced because of possible fitness costs associated with resistance mutations. As before, each field can be treated or not, and each infected field can be dominated by either a wildtype or a resistant pathogen. Thus, now the epidemiological dynamics are given by:

$$
\begin{aligned}
\dot{H}_u &= -\beta_w[I_{uw} + I_{tw}]H_u - \beta_r[I_{ur} + I_{tr}]H_u + \mu[I_{uw} + I_{ur}] \\
\dot{H}_t &= -\beta_w(1 - \epsilon_w)[I_{uw} + I_{tw}]H_t - \beta_r(1 - \epsilon_r)[I_{ur} + I_{tr}]H_t + \mu[I_{tw} + I_{tr}] \\
\dot{I}_{uw} &= \beta_w[I_{uw} + I_{tw}]H_u - \mu I_{uw} \\
\dot{I}_{tw} &= \beta_w(1 - \epsilon_w)[I_{uw} + I_{tw}]H_t - \mu I_{tw} \\
\dot{I}_{ur} &= \beta_r[I_{ur} + I_{tr}]H_u - \mu I_{ur} \\
\dot{I}_{tr} &= \beta_r(1 - \epsilon_r)[I_{ur} + I_{tr}]H_t - \mu I_{tr}
\end{aligned}
\tag{7}
$$

Here, the fraction of treated fields is denoted by $\theta = (H_t + I_{tw} + I_{tr})/N$, and correspondingly $1 - \theta$ is the fraction of untreated fields.

The evaluation of the fixed points can capture the endpoints of the dynamics of the above system. Expressions for the fixed points are presented in Appendices (A.3.1),(A.3.2) for the case of a high efficacy fungicide ($\epsilon_w = 1$) either with absolute resistance ($\epsilon_r = 0$) or partial resistance ($0 < \epsilon_r < 1$), respectively. The general case of partial efficacy fungicide ($0 < \epsilon_w < 1$) and partial resistance ($0 < \epsilon_r < 1$) is investigated in Appendix (A.3.3). The fixed points under consideration can thus be categorized into four distinct, biologically significant solutions. The first solution corresponds to disease-free equilibrium, where no infected fields exist. The second solution corresponds to the wildtype endemic equilibrium, where only fields infected with the wildtype strain exist. The third solution corresponds to the endemic resistant equilibrium, where solely fields infected with the resistant strain exist. Lastly, the fourth solution corresponds to the coexistence between the resistant and wildtype strains.

Using the next-generation matrix approach, we derive the following expression for the effective reproduction number in the general case of partially effective fungicide and partial resistance (Appendix (A.3.4)):

$$\mathcal{R}_0 = \max\left(\frac{N\beta_w}{\mu}(1 - \epsilon_w\theta), \frac{N\beta_r}{\mu}(1 - \epsilon_r\theta) \right). \tag{8}$$

Interestingly, the effective reproduction number depends on the capacity to spread the wildtype strain or the resistant strain. The spread capacity is governed by the relative transmission rates (β_w, β_r) and fungicide efficacy (ϵ_w, ϵ_r). Assuming equal transmission rates for both the wildtype and resistant strains ($\beta_w = \beta_r$, i.e., no fitness costs of resistance) but unequal fungicide efficacy for the two strains ($\epsilon_w > \epsilon_r$), we can safely assume that the resistant strain will outperform the wildtype strain over the long run. As a result, the effective reproduction number takes the following form:

$$\mathcal{R}_0 = R_0(1 - \epsilon_r\theta). \tag{9}$$

The above expression is similar to Eq. (4), but it has ϵ_r instead of ϵ_w as a factor reducing the basic reproduction number, reflecting the fact that here the fungicide suppresses the transmission of the resistant strain. Based on Eq. (9), we determine the critical fungicide coverage, above which the fungicide treatment leads to the extinction of resistance:

$$\theta_{cr} = \frac{1}{\epsilon_r}\left(1 - \frac{1}{R_0} \right). \tag{10}$$

This expression is similar to the critical fraction θ_{cw} given by Eq. (5). It has a similar biological significance, but now the expression is scaled by the fungicide efficacy against the resistant strain, ϵ_r. This difference reflects that the critical fraction here is computed when controlling the resistant rather than the wildtype strain. If $\theta_{cr} > 1$, then the fungicide treatment cannot drive the pathogen population (consisting of the resistant strain) to extinction even for complete fungicide coverage $\theta = 1$. This is the case when the fungicide efficacy against the resistant strain is lower than the critical value, $\epsilon_r < \epsilon_{cr}$, where

$$\epsilon_{cr} = 1 - \frac{1}{R_0}. \tag{11}$$

For a fungicide with a given efficacy ϵ_r against the resistant strain, increasing the fraction of treated fields θ to values above the critical, θ_{cr}, can drive the pathogen to extinction for a sufficiently low basic reproduction number, $R_0 < R_{0cr}$, where

$$R_{0cr} = 1/\left(1 - \epsilon_r\right). \tag{12}$$

In Eq. (1), our model assumes density-dependent transmission or pseudo mass-action (Keeling and Rohani, 2008): disease transmission increases with the density of fields and no explicit spatial structure. The approximation works well when the primary mode of long-distance dispersal of plant pathogens is via networks of transportation/trade (e.g., via movement of infected plants or other materials). For plant pathogens that disperse via air-borne spores, a spatially-explicit model would need to consider a distance-dependent transmission (Medlock and Kot, 2003; Mikaberidze et al., 2016). Our model provides a valuable reference scenario but is expected to overestimate the number of diseased fields at endemic equilibrium.

Although our model considers multiple growing seasons, we simplified it as in (Parnell et al., 2006) by neglecting the cyclic nature of multi-seasonal dynamics due to, e.g. the periodic absence of host plants. Environmental stochasticity associated with such multi-seasonal dynamics may play a role in the emergence of novel, better-adapted pathogens (Willocquet et al., 2020). We also assumed that the yield loss in diseased fields remains constant. However, in reality, it can vary depending on the crop variety, environmental conditions and the genetic composition of the pathogen population. To consider this variation, one would need to incorporate a within-field epidemic model into the landscape scale model.

For simplicity, we assume that environmental conditions and cropping practices are the same across different fields and that all fields have the same size. Hence, the model parameters do not vary across different fields. The model assumes recovered fields are susceptible to re-infection (i.e., no acquired immunity). Neglecting demographic stochasticity (Mikaberidze et al., 2017), our model belongs to the class of spatially implicit metapopulation models and is similar to the model developed by Parnell et al. (2006). However, in contrast to (Parnell et al., 2006), our model does not have an explicit spore compartment, assuming dispersal occurs much faster than other relevant epidemiological processes. While the assumptions come with the above caveats, the upshot is that they allow us to formulate a simple landscape scale model of a generic fungal disease of crop plants. We can capture essential features of disease dynamics and allow for analytical treatment. To describe specific crop-pathogen systems with a higher degree of biological realism, our model will need to relax some of the assumptions at the expense of increased model complexity.

Another benefit of having a simple landscape model of crop infections allows us to extend in a different but necessary direction when considering sustainable fungicide applications - the economic dimension. Thus having incorporated fungicide resistance into the model, we now analyze the economic aspects of fungicide treatment. We aim to determine the optimal fungicide treatment coverage that maximizes the net economic return.

2.2 Economic Analysis: Net return and optimal fungicide coverage

We define the net return as the total income, which is the sales of the harvested yield minus the cost of fungicide treatment across all N fields in the landscape. Mathematically, the net return is:

$$G = y_H \sum_{k=1}^{tot} \left(H(t_k) + I(t_k)y - \theta N f \right), \tag{13}$$

where t_k is the time point corresponding to the end of season k, $y = y_I/y_H$ is the relative yield of a diseased field, $f = c/y_H$ is the relative fungicide price. Here, c, the cost of fungicide application per field includes the cost of spraying material and application cost.

The net return does not change over time when the system reaches a stable equilibrium. Besides being analytically tractable, we consider this case as it corresponds to the limit of sustainable disease management. The outcomes will likely remain the same across a range of times, even before the equilibrium is strictly reached. However, they are expected to be different during the initial stages of dynamics as they are affected by initial conditions.

At equilibrium, we characterize the net return as follows. We compute it for a single season and divide it by the total number of fields N to obtain the average across fields. We also divide it by yield per season from a healthy field y_H, to obtain the relative net return (which we will call "net return" for brevity) $g(\theta) = (H^* + I^* y)/N - \theta f$, where $H^* = H_u^* + H_t^*$ and $I^* = I_{uw}^* + I_{tw}^*$ represent the values of the state variables (numbers of healthy and diseased fields) at a stable fixed point. The yield from healthy and diseased fields boosts the net return (the first term), while the cost of fungicide treatment reduces the net return. Since the total number of fields remains constant, $H^* = N - I^*$, we can re-write the expression for the net return in the form that supports further analyses:

$$g(\theta) = 1 - (1 - y)\left(\frac{I^*}{N} + \alpha\,\theta \right). \tag{14}$$

Here, the net return without disease and no treatment equal to one (the first term in Eq. (14)). The second term in Eq. (14) quantifies the reduction of the net return due to disease, which is

proportional to the fraction of diseased fields I^*/N. The third term in Eq. (14) quantifies the reduction of the net return due to the cost of fungicide treatment. This term is proportional to the cost ratio parameter:

$$\alpha = \frac{f}{1-y}, \tag{15}$$

where f is the relative fungicide price, and $1-y$ is the yield loss in a diseased field (expressed as its monetary equivalent). In fact, the proportion of diseased fields $I^*/N = (I^*_{uw} + I^*_{tw})/N$ also depends on θ albeit in a complex manner according to Eqs. (A4-A11) (Appendix A.3) and the associated linear stability conditions.

Thus given the importance of θ, we determine the optimal fungicide coverage, or simply optimal coverage, θ^*: the proportion of treated fields that results in the highest net return. The optimal coverage θ^* can be determined by calculating the value of $g(\theta)$ across a fine grid of θ-values. The θ-value that maximizes $g(\theta)$ is the optimal coverage θ^*. In cases where multiple values of θ yield the same highest value of $g(\theta)$, we define the lowest among these as the optimal coverage θ^*.

3 Results

3.1 Bio-economic analysis of fungicide use

Optimal coverage without resistance

What fraction of fields should be treated with a fungicide to maximize the net return of a landscape? We address this question by analyzing the expression for the net return in Eq. (14). How the net return across a cultivated landscape depends on the intensity of treatment (expressed as the fraction of treated fields, θ) differs fundamentally from how net return depends on the intensity of treatment in an individual field (expressed as the fungicide dose, D). This difference is illustrated in Fig. 2.

When controlling disease in an individual field, the yield, Y (expressed as its monetary equivalent), typically exhibits a saturating increase as we increase the fungicide dose, D (see Fig. 2A according to Eq. (B2) in Appendix B.1). The total fungicide price, F, increases linearly when increasing the fungicide dose with the slope given by the fungicide price per unit dose (dark red lines in Fig. 2A according to Eq. (B3) in Appendix B.1). The net return, $g(D)$, is then given by the difference between the yield, Y, and the total fungicide price, F (Eq. (B1) in Appendix B.1; Fig. 2A). $g(D)$ reaches a maximum for an intermediate (optimal) fungicide dose. For doses higher than optimal, the cost associated with increasing the dose is no longer compensated by an associated increase in yield. These relationships follow "the law of diminishing returns" that has achieved an almost universal status in the economics of crop production (Kho, 2000) and other areas of economics (Shephard, 1970; Brue, 1993).

In contrast, when controlling disease in a cultivated landscape, the total yield of the landscape (Y, expressed as its monetary equivalent, Eq. (B8)) exhibits an accelerating increase as we increase the fungicide coverage, θ (Fig. 2C). Nevertheless, the total price of the fungicide treatment increases linearly with θ (Fig. 2C), similarly to the case of an individual field above. The net return, $g(\theta)$, is computed as the difference between the monetary yield, Y, and the total fungicide price, F, as for the individual field case. However, here, the outcome is very different. The net return does not exhibit a maximum at any intermediate θ value. Thus the net return can be maximized only at the extreme values of θ: $\theta = 0$ treating no fields or $\theta = 1$ treating all fields. When the fungicide is sufficiently cheap (Fig. 2C) $\theta = 1$ maximizes the net return (Fig. 2D). However, when the fungicide is expensive, then the no-treatment strategy ($\theta = 0$) maximizes the net return (Fig. 2D).

Thus, the law of diminishing returns does not hold when controlling disease in cultivated landscapes. This finding has important implications for the economics of crop disease management, optimal decision making and associated policy-making, which is often done at scale and not for individual fields.

Our analysis reveals that three parameters determine the optimal fungicide coverage, θ^*: (i) the fungicide price, f, (ii) the yield loss in a diseased field, $1-y$ and (iii) the critical fraction of treated fields, θ_{cw} (in Eq. (5)). The first two parameters affect θ^* only through their ratio,

Figure 2: How does the net return depend on the intensity of treatment? (A) In an individual field, the yield Y exhibits a saturating increase as we increase the fungicide dose, D [solid blue; Eq. (B2) in Appendix B.1]. Y is expressed as a fraction of the monetary equivalent of the maximum yield. The total fungicide price, F, is proportional to the dose, D, with the slope given by the price per unit: $q = 0.05$ (solid red); $q = 0.4$ (dashed red). (B) Black curves represent the net return computed as the blue curve in (A) minus one of the red lines in (A), with $q = 0.05$ (solid) and $q = 0.4$ (dashed). The net return exhibits a maximum for an intermediate dose [Eq. (B4)]. Grey vertical lines in (A) and (B) show doses that maximize net return. Other parameter values for (A) and (B): $y_m = 0.55$, $D_{50} = 0.2$, $y_0 = 0.4$. (C) In a cultivated landscape, the total yield of the landscape, Y, [Eq. (B8)] exhibits an accelerating increase as we increase the fraction of treated fields, θ (blue). The total fungicide price, F, is proportional to θ, with the slope given by the relative fungicide price: $f = 0.1$ (solid red), $f = 0.7$ (dashed red). (D) As a result, the net return is maximum when treating either all fields ($\theta = 1$, solid, $f = 0.1$) or when not treating any fields ($\theta = 0$, dashed, $f = 0.7$); according to Eq. (B7) in Appendix B.1. Other parameter values for (C) and (D): $\beta = 0.025$, $\mu = 5$, $N = 1000$ (hence, $R_0 = \beta N/\mu = 5$), $y = 0.5$.

the cost ratio parameter $\alpha = f/(1-y)$ (defined in Eq. (15) above). Thus below, we study the effects of α and θ_{cw}.

When α is sufficiently low, treating all fields ($\theta^* = 1$, if $\theta_{cw} > 1$; Fig. 3A) or most fields ($\theta^* = \theta_{cw}$, if $\theta_{cw} < 1$ Fig. 3A) maximizes the net return (here, θ_{cw} is given by Eq. (5)). When α is sufficiently high, not treating any fields maximizes the net return ($\theta^* = 0$; Fig. 3). Thus, the net return can only be maximized at extreme values of the fungicide coverage, either $\theta^* = 0$ or $\theta^* = 1$ (or $\theta^* = \theta_{cw}$ in case $\theta_{cw} < 1$, whereby θ_{cw} is still close to 1).

To gain a quantitative understanding of what factors determine the optimal coverage θ^*, consider the regimes $\theta_{cw} > 1$ and $\theta_{cw} < 1$ in more detail. When $\theta_{cw} > 1$, it is impossible to wipe out the epidemic even if we treat all fields, i.e. $\theta = 1$. This is the case for pathogens with a sufficiently high R_0, whereby, $R_0 > R_{0cw}$ (Fig. 3A), where R_{0cw} is given by Eq. (6). If these conditions are not met and $\theta_{cw} < 1$, one can wipe out the epidemic by treating a high enough proportion of fields, i.e., $\theta > \theta_{cw}$. This is the case for pathogens with a sufficiently low R_0, whereby, $R_0 < R_{0cw}$ (Fig. 3B). In both cases, the optimal θ-value is determined by how the cost ratio α in Eq. (15) relates to a combination of the epidemiological parameters.

For $\theta_{cw} > 1$, not treating any fields ($\theta = 0$) maximizes the net return when α is sufficiently high, i.e., larger than a threshold value A (such that $\alpha > A$). Treating all fields ($\theta = 1$) maximizes the net return when α is sufficiently low, i.e., $\alpha < A$ (Fig. 3A). Here, A represents a

combination of epidemiological parameters,

$$A = \frac{\epsilon_w}{R_0(1 - \epsilon_w)}. \tag{16}$$

For $\theta_{cw} < 1$, we can maximize the net return by not treating any fields ($\theta = 0$) when α is sufficiently high ($\alpha > \epsilon_w$). If, on the contrary, $\alpha < \epsilon_w$, then we can maximize the net return by treating a fraction of fields that corresponds to the critical value ($\theta = \theta_{cw}$, which constitutes a high fraction of fields; Fig. 3B). We explore how the net return depends on the fungicide coverage in more detail in Appendix B.2.1.

Figure 3: Optimal fraction of fungicide-treated fields and net return with variation in fungicide price and basic reproduction number for the landscape model without fungicide resistance. (A) The net return $g(\theta)$ is plotted against the fraction of treated fields θ for $R_0 > R_{0cw}$ and $\theta_{cw} > 1$ (i.e., $\beta = 0.02$ and $R_0 = \beta N/\mu = 4$), which indicates that it is impossible to eradicate the pathogen even if all fields are treated. The four curves represent different relative fungicide prices (f): 0.05 (solid blue), 0.25 (dashed red), 0.4 (dotted orange), and 0.65 (dash-dotted megenta). (B) The relative net return $g(\theta)$ is plotted against the fraction of treated fields θ for $R_0 < R_{0cw}$ and $\theta_{cw} = 0.714 < 1$ (shown by a vertical grey line, i.e., $\beta = 0.01$ and $R_0 = \beta N/\mu = 2$), which means that it is possible to eradicate the pathogen by treating a high enough fraction of fields (vertical grey line). The four curves represent different values of f: 0.1 (solid blue), 0.33 (dashed red), 0.4 (dotted orange), and 0.65 (dash-dotted megenta). (C) The heatmap shows the optimal fraction of treated fields θ^* for varying fungicide price f and basic reproduction number R_0. In the white region, treating all fields is optimal; in the black region not treating any fields is optimal; in the grey region, treating only a fraction of fields is optimal. Each curve in (A) and (B) corresponds to a point on the green dashed line in the heatmap (C). These points are highlighted by circles of the same color as lines in (A) and (B). Specifically, $R_0 = 4$ and $R_0 = 2$ correspond to (A) and (B) respectively. Other parameter values are constant for all panels: $N = 1000$, $\mu = 5$, $\epsilon_w = 0.7$, and $y = 0.5$.

Figure 3C shows that the optimal coverage, θ^*, depends strongly on both the relative fungicide price, f, and the pathogen's basic reproduction number, R_0. This dependency exhibits three regimes for low, intermediate and high values of R_0. For low $R_0 < 1$, the optimal coverage does not depend on fungicide price, because epidemics will not sustain even without treatment and hence not treating any field is optimal for any f. When $R_0 > 1$, the optimal coverage does depend on f, showing a threshold pattern. For low f-values below the threshold, it is optimal to treat at least some of the fields ($\theta^* > 0$). For high f-values above the threshold, not treating any fields yields a maximum net return ($\theta^* = 0$). For intermediate R_0-values, the threshold in f does not depend on R_0 (hence, the boundary between white/gray and black areas is a vertical line in Fig. 3C). For high R_0-values, the threshold in f becomes lower when R_0 is increased. This threshold pattern, rather than a gradual increase in the optimal coverage with increasing f, occurs because as we have shown above the law of diminishing returns no longer holds when we increase the treatment coverage in multiple fields across a landscape. The identification of the threshold pattern can inform economic policy for achieving more sustainable fungicide use, as we elaborate in Discussion.

Thus, (i) the law of diminishing returns does not hold with respect to treatment coverage across cultivated landscapes. The net return is maximized at extreme θ-values, and (ii) which of the extrema are optimal is determined by how the cost-ratio α relates to a combination of epidemiological parameters of the system.

Optimal coverage with resistance

Now we focus on the fraction of fields to be treated with fungicide to maximize the net return in a landscape model when resistance exists. We consider a high-efficacy fungicide ($\epsilon_w = 1$), no fitness cost of resistance ($\beta_r = \beta_w$), and focus on partial resistance ($0 < \epsilon_r < 1$) (absolute resistance and fitness cost of resistance in Appendix B.2.2). The fungicide can still suppress the partially resistant strain with efficacy ϵ_r, smaller than the efficacy against the wildtype strain, ϵ_w. As before, the optimal fungicide coverage depends on the cost-ratio α (in Eq. (15)).

How the net return changes with θ here is similar to the case without resistance. The dependency is influenced by whether the critical fraction of treated fields, θ_{cr}, in Eq. (10) is higher or lower than 1. When $\theta_{cr} > 1$ ($\epsilon_r < \epsilon_{rc}$, where ϵ_{cr} is given by Eq. (11)), fungicide treatment is ineffectual in clearing the resistance. In this case, the optimal coverage, θ^*, depends on how the cost-ratio α relates to the combination of epidemiological parameters,

$$B = \frac{\epsilon_r}{1 - \epsilon_r} / R_0. \tag{17}$$

The expressions for B is similar to A in Eq. (16), but with ϵ_r instead of ϵ_w. When a fungicide is sufficiently cheap, $\alpha < B$, treating all fields maximizes the net return ($\theta^* = 1$; Fig. 4A). When a fungicide is too expensive, $\alpha > B$, then not buying it is optimal ($\theta^* = 0$; Fig. 4A).

When $\theta_{cr} < 1$ ($\epsilon_r > \epsilon_{cr}$), treating a sufficiently high fraction of fields ($\theta > \theta_{cr}$) will drive the dominant resistant strain to extinction. Here, the optimal θ depends on how the cost-ratio α relates to the efficacy of the fungicide against the resistant strain ϵ_r. If $\alpha < \epsilon_r$, then treating a significant fraction of fields maximizes the net return ($\theta^* = \theta_{cr}$; Fig. 4B). If $\alpha > \epsilon_r$, then not treating any field maximizes the net return ($\theta = 0$; Fig. 4B). Figure 4C provides a broader overview of the parameter space, considering scenarios with continuously varying fungicide prices, f, and fungicide efficacy against the resistant strain, ϵ_r.

Thus, when resistance is absolute, it is impossible to manage it apart from stopping or drastically reducing fungicide treatment (see Appendix B.2.2). Nevertheless, when resistance is partial, the net return can still be maximized by treating all fields or a significant fraction of fields (if the fungicide is sufficiently cheap) or by not treating any fields (if the fungicide is sufficiently expensive).

3.2 Calculation of the economic cost of fungicide resistance

The maximum (or optimal) net return of a landscape becomes lower in the presence of resistance (compare Fig. 3A and Fig. 4A). Thus, the evolution of fungicide resistance can incur a substantial economic cost. Here, we propose a novel way to calculate the economic cost of fungicide resistance across a cultivated landscape using the expression for the net return in Eq. (14). Firstly, we compute the net return in the absence of and presence of resistance according to Eq. (14). Secondly, we determine the optimal fraction of treated fields θ^* to maximize the net return in the absence and presence of resistance. Finally, we calculate the economic cost of resistance, C_R, as the difference between the optimum net return in the absence and presence of resistance:

$$C_R = g_0(\theta^*) - g_1(\theta^*), \tag{18}$$

where $g_0(\theta^*)$ is the optimum net return in the absence of resistance, and $g_1(\theta^*)$ is the optimum net return in the presence of resistance. In this calculation, we assume that, both with and without resistance, the optimum proportion of treated fields is known to farmers across the landscape and that they use these optimum values.

The maximal net return, $g_{\max}$, the optimal fraction of treated fields, θ^*, and the economic cost of resistance, C_R, are plotted versus the relative fungicide price, f. Figure 5A compares the optimal net return in the absence and presence of resistance. The difference between these curves constitutes the economic cost of resistance evolution plotted in Fig. 5C. Herein

Figure 4: Optimal fraction of fungicide-treated fields and net return with variation in fungicide price and its efficiency for the landscape model with fungicide resistance. (A) Net return $g(\theta)$ is plotted against the fraction of treated fields θ for a high degree of partial resistance ($\epsilon_r = 0.6 < \epsilon_{cr} = 0.75$), which indicates that it is impossible to drive the resistant strain to extinction even if all fields are treated. The four curves represent different relative fungicide prices (f): 0.05 (solid blue), 0.15 (dashed red), 0.25 (dotted orange), 0.5 (dash-dotted magenta). (B) Relative net return $g(\theta)$ is plotted against the fraction of treated fields θ for a low degree of partial resistance ($\epsilon_r = 0.8 > \epsilon_{cr}$) which means that it is possible to drive the resistant strain to extinction by treating a high enough fraction of fields ($\theta = \theta_{cr}$, vertical grey line). The four curves represent different values of f: 0.05 (solid blue), 0.25 (dashed red), 0.5 (dotted orange), 0.65 (dash-dotted magenta). (C) Heatmap of the optimal fraction of treated fields θ^* for varying fungicide price f and fungicide efficacy against resistant strain ϵ_r. In the white region, it is optimal to treat all fields with fungicide; in the black region, it is optimal to not treat any fields; in the grey region it is optimal to treat only a fraction of fields. Each curve in (A) and (B) corresponds to a point on the green dashed line in (C). These points are highlighted by circles of the same color as lines in (A) and (B). Specifically, points on green dashed lines for $\epsilon_r = 0.6$ corresponds to (A) and for $\epsilon_r = 0.8$ corresponds to (B). Other parameter values are constant for all panels: $\beta = 0.02$, $R_0 = \beta N/\mu = 4$; $N = 1000$, $\mu = 5$, $\epsilon_w = 1$, $y = 0.5$.

we elucidate the dependence of the cost of resistance on the fungicide price f, the degree of resistance $1 - \epsilon_r$ and the basic reproduction number R_0

Dependence on fungicide price. Figure 5C shows a typical dependence of C_R on f when the basic reproduction number of the pathogen exceeds its critical values ($R_0 > 1/(1 - \epsilon_w)$ in Eq. (6) and $R_0 > 1/(1 - \epsilon_r)$ in Eq. (12)). We explore three ranges of fungicide prices (f) – cheap, intermediate and expensive – corresponding to qualitatively different patterns in the C_R dependence. For cheap fungicides, C_R remains constant and in general, treating all fields maximizes the net return (Fig. 5B). The optimal net return drops linearly with increasing f at the same rate. Hence, the difference between the optimal net return with and without resistance remains constant (the economic cost of resistance in this parameter range - Eq. (C1) in Appendix C). For intermediate f-values, treating all fields remains optimal in the absence of resistance, but in the presence of resistance, the optimum switches to not treating any fields. The difference between the optimal net return without resistance and with resistance (that constitutes C_R) declines linearly with f (the economic cost of resistance in this range - Eq. (C2) in Appendix C). For expensive fungicides, high f-values, not treating any fields becomes optimal in the absence or presence of resistance. Here, the optimal net return is determined by the yield of untreated fields, independent of the fungicide price f, irrespective of resistance Fig. 5C.

When R_0 does not exceed the critical value for the wildtype strain (i.e., $R_0 < R_{0cw}$ [Eq.(6)], C_R increases, reaches a maximum and then decreases, reaching zero as a function of f (region of the plot corresponding to $1 < R_0 < 5$ in Fig. 6B). This pattern arises due to an intricate interplay between the optimal net return with and without resistance and the associated optimal fungicide coverage (we describe this in more detail in Appendix C).

Dependence on the degree of resistance. The dependence of the cost of resistance on the degree of resistance is apparent only for cheap fungicides (low f-values). Herein, C_R remains

Figure 5: Calculation of the economic cost of fungicide resistance, C_R, versus relative fungicide price, f, according to Eq. (18). (A) Optimal net return, $g_{\max}$, without resistance (solid, blue curve) and with resistance (dashed, red curve). (B) Optimal fraction of treated fields, θ^*, without resistance (solid, blue curve) and with resistance (dashed, red curve). (C) Economic cost of resistance, C_R, computed as the optimal net return without resistance minus the optimal net return with resistance. Parameter values: $\beta_r = \beta_w = 0.025$, $N = 1000$, $\mu = 5$ (hence, $R_0 = \beta N/\mu = 5$); $\epsilon_w = 0.8$, $\epsilon_r = 0.5$.

constant at low values of ϵ_r and then decreases linearly with increasing ϵ_r. Thus higher degrees of resistance (lower ϵ_r-values) correspond to higher economic costs of resistance (compare horizontal parts of contour lines in Fig. 6A). This is because more resistance leads to higher numbers of fields becoming diseased despite treatment, which in turn results in lower overall yields of the landscape in the presence of resistance compared to the absence of resistance. Unexpectedly, the range of f-values over which C_R remains constant is larger for lower degrees of resistance. This is because the transition from treating all fields as optimal to treating no fields as optimal occurs at higher f-values for lower degrees of resistance. Since the overall yields are higher for lower degrees of resistance, it takes a more expensive fungicide to tip the cost-benefit balance from treating all fields to treating no fields. For expensive fungicides (high f-values), the degree of resistance no longer affects the economic cost of resistance. If resistance is present in this range, then any treatment is useless.

Dependence on basic reproduction number. The dependence of the cost of resistance evolution on R_0 is counter-intuitive (Fig. 6B). When the fungicide is very expensive (the rightmost region in Fig. 6B), optimal fungicide coverage goes to zero both with and without resistance, and resistance incurs no economic cost. However, in a more interesting case of low- to intermediately-priced fungicide, C_R exhibits complex, non-monotonic patterns versus R_0 and f (Fig. 6B).

Figure 6: How does the economic cost of fungicide resistance, C_R, depend on fungicide price (f), degree of resistance $(1 - \epsilon_r)$, basic reproduction number (R_0), and yield loss due to disease $(1 - y)$? Each panel illustrates how C_R depends on f along the x-axis and one of the other variables along the y-axis: (A) the effect of the degree of resistance, $(1 - \epsilon_r)$, with $R_0 = 5$ and $y = 0.5$; (B) the effect of the basic reproduction number, R_0, with $\epsilon_r = 0.5$ and $y = 0.5$ ($R_{0cw} = 5$, $R_{0cr} = 2$); (C) the effect of the yield loss due to disease, $1 - y$, with $\epsilon_r = 0.5$ and $R_0 = 5$. Values of ϵ_r, R_0 and y used to plot Fig. 5 are highlighted with an arrow on the y-axis of panels (A), (B) and (C), respectively. Note that each panel uses a different grayscale mapping for C_R as shown in the grayscale bars. Parameter values that remain the same across all panels are: $N = 1000$, $\mu = 5$, and $\epsilon_w = 0.8$

For cheap fungicides (low f-values), the economic cost of resistance first remains at zero, then increases with R_0, reaches a maximum and then decreases (compare different contour lines in Fig. 6B for small f-values). When $R_0 < 1$, no epidemic occurs, and hence the relative net return remains at its highest possible value $g_{\max} = 1$ both in the absence and presence of resistance, resulting in zero C_R. When $1 < R_0 < R_{0cr}$ [Eq. (12)], an epidemic can occur both in the absence and presence of resistance, but in both cases treating an optimal, intermediate proportion of fields with a fungicide drives the pathogen to extinction ($\theta^* = \theta_{cw}$ without resistance; $\theta^* = \theta_{cr}$ with resistance). Hence, the optimal net return is given by $g_{\max} = 1 - f\theta_{cr}$ with resistance and $g_{\max} = 1 - f\theta_{cw}$ without resistance. The difference between the two gives us $C_R = f(\theta_{cr} - \theta_{cw})$ proportional to $1 - 1/R_0$ (see Eq. (5) and Eq. (10)), which increases with R_0. In other words, both with and without resistance, a higher fungicide coverage is required to extinguish the pathogen with higher R_0 values, incurring a higher cost. This cost grows faster with R_0 in the presence of resistance than in the absence, simply because the fungicide has a lower efficacy. When $R_{0cr} < R_0 < R_{0cw}$, C_R continues to increase with R_0. Applying a fungicide to an optimal, intermediate proportion of fields in this regime still drives the pathogen to extinction without resistance. Nevertheless, treating all fields in the presence of resistance becomes optimal ($\theta^* = 1$), but even this does not drive the pathogen to extinction. Here, $g_{\max}$ without resistance declines with R_0 for the same reason as above: a higher fungicide coverage is needed to extinguish pathogens with higher R_0-values, leading to a higher price of treatment. $g_{\max}$ with resistance also declines with R_0 due to an increased proportion of diseased fields at higher R_0-values. Optimal net return declines faster as a function R_0 with resistance than without resistance, simply because the fungicide suppresses the resistant strain with less efficacy. As a result, the cost of resistance continues to increase versus R_0 in this regime until it reaches its maximum at $R_0 = R_{0cw}$.

Finally, and unexpectedly, when $R_0 > R_{0cw}$, C_R drops with increasing R_0. In this case, fungicide treatment cannot extinguish the pathogen population, irrespective of the resistance state. To clarify, consider a simplified scenario of a cheap fungicide ($f \to 0$). In this case, treating all fields ($\theta^* = 1$) becomes optimal both with and without resistance (Fig. 5B). Both with and without resistance, the optimal net return declines with R_0 as the fraction of diseased fields increases with R_0. Without resistance, this fraction (at endemic equilibrium) is given by $i_w = 1 - 1/[R_0(1 - \epsilon_w)]$. With resistance, this fraction (also at endemic equilibrium, but now with the resistant strain) is given by $i_r = 1 - 1/[R_0(1 - \epsilon_r)]$. C_R is proportional to the difference $i_r - i_w$. Both i_r and i_w increase in a saturating manner with R_0 (eventually tending to one at high R_0-values) but do so at different, nonlinear rates in such a way that the difference

between the two diminishes at higher R_0-values, and so does the economic cost of resistance, C_R. In other words, resistance incurs a less economic cost at higher R_0-values because even without resistance, it becomes challenging to suppress the disease due to the high invasiveness of the pathogen (as quantified by the basic reproduction number).

Dependence on yield loss. The economic cost of fungicide resistance C_R is influenced by the yield loss in diseased fields, $1 - y$ (relative to the yield of healthy fields; Fig. 6C). For cheap fungicides (low f-values), higher yield losses are associated with higher economic cost of resistance. However, the transition to a linear decrease of C_R versus f occurs at lower f-values. In the range of intermediate and high f-values, the economic cost of resistance can be higher for lower yield losses (e.g., compare different contour lines in Fig. 6B).

4 Discussion

Farmers can maximize the net return of fungicide application by optimizing the amount of fungicide applied based on the balance between the yield benefit and the economic costs of fungicide application (Jørgensen et al., 2017; Hillebrandt, 1960). However, this analysis neglects the indirect costs of fungicide application as in resistance evolution, rendering decisions based solely on this analysis unsustainable. Designing economic policy instruments incorporating indirect costs into fungicide prices, such as pesticide taxes or subsidies encouraging a reduction in pesticide application, is one way of mitigating the issue (Finger et al., 2017; Lee et al., 2019). However, this approach requires a robust estimation of the indirect costs, and a conceptual framework for such an estimation is absent from the existing literature. We build such a framework based on a simple but generic epidemiological model of crop epidemics across multiple fields combined with an economic analysis (Fig. 1). Using this framework, we demonstrated how to estimate the economic cost of the evolution of fungicide resistance in crop pathogens with long-distance dispersal.

Deriving the economic cost of fungicide resistance evolution is a multifaceted issue. However, after making some well-motivated assumptions, we derived analytical expressions for it (C_R). This derivation allows us to investigate how C_R depends on the system's critical epidemiological and economic parameters across a range of values.We found that this cost is strongly influenced by fungicide price, degree of resistance, pathogen's basic reproduction number, and yield loss due to disease. While intuitively, C_R increases with the degree of resistance and yield loss due to disease, it remains constant or decreases with the fungicide price. Unexpectedly, C_R shows a non-monotonic pattern as a function of the basic reproduction number, R_0 (exhibiting a maximum at a critical value $R_0 = R_{0cw}$).

We might consider the range of low R_0-values ($1 < R_0 < R_{0cw}$) to be unrealistic for most fungal pathogens of crop plants. Firstly, according to empirical knowledge accumulated in plant pathology, fungal crop pathogens are challenging to eradicate (Lucas, 2020). Secondly, existing estimates of the basic reproduction number, e.g., for the major fungal pathogens of wheat *Zymoseptoria tritici* and *Puccinia striiformis* tend to be relatively high: in the range $4 - 10$ for *Z. tritici* (Mikaberidze et al., 2017) and $20 - 30$ for *P. striiformis* (Mikaberidze et al., 2016). The two arguments indicate that R_0-values for major fungal crop pathogens will likely exceed the eradication thresholds. Accordingly, we consider more relevant the range of higher R_0-values, where it exceeds the eradication thresholds ($R_0 > R_{0cw}$, $R_0 > R_{0cr}$). Here we used R_0 as determined on the landscape level. However, most experimentally determined R_0 for fungal pathogens are not evaluated at the landscape but at an individual field level. Given this discrepancy, we expect lower R_0 estimates at the landscape scale, a hypothesis that needs empirical validation.

Along the way to calculating the economic cost of fungicide resistance, we also found that the law of diminishing returns does not hold for the economic return from fungicide applications over a cultivated landscape (Fig. 2D). Optimal fungicide dosing for individual fields still exhibits the law of diminishing returns for the net economic return, where the net return is curved upwards and exhibits a maximum with an increase in the fungicide dose (Fig. 2B). Hence, the econonically optimum dose decreases gradually with increasing the fungicide price. Instead, we find that for multiple fields across the landscape, the optimal proportion of treated field also decreases with the ratio of fungicide price and yield loss (α) but does so via a sudden jump (see Fig. 5B). Thus, adjusting the magnitude of fungicide treatment to maximize the net return is more difficult at the landscape scale owing to coordinating fungicide applications across multiple

farms and handling the discontinuous nature of how θ^* depends on α. At the same time, this presents an opportunity for an economic policy to minimize pesticide use because a limited increase in α induced by a policy instrument may lead to a disproportionately high, sudden drop in the optimal fungicide coverage, θ^*. However, achieving this optimum requires cooperation between farmers across a regional landscape.

The three epidemiological parameters (degree of resistance, R_0 and yield loss due to infection) can vary significantly across different crop diseases (Lucas et al., 2015; Mikaberidze et al., 2016; de Vallavieille-Pope et al., 2000; Savary et al., 2012). Also, fungicide price varies across countries and products and fluctuates over time. Even for a single disease, the three epidemiological parameters can vary substantially depending on environmental conditions, crop varieties and genetic composition of pathogen populations. A strong dependence of C_R on the three parameters highlights the need for a robust empirical estimation of these three factors. Even after estimation, designing economic policy instruments that work across different regions and crops would be demanding. These instruments are more likely to succeed if they are adjusted for specific regions and crops and over time according to changes in cropping systems and pathogen populations. Implementing economic and ecological interventions in a realistic system will further channel feedback into the parameter values. Thus, like adaptive therapy in medicine, the directives must be adapted to specific conditions and responsive in real-time. To achieve this, economic policy aiming for more sustainable use of fungicides needs to account for multidisciplinary research combining biology and economics, boosted by theoretical approaches.

Our study assumes a binary choice in fungicide application. In reality, farmers can adjust the fungicide dose, use mixtures or alternations of different fungicides, and combine them with disease-resistant crop varieties and cultural control measures (such as crop rotations (Bargués-Ribera and Gokhale, 2020)). These scenarios can be incorporated into our modelling framework to explore how different control measures can protect each other from pathogen adaptation (Carolan et al., 2017) and mitigate the economic cost of pathogen adaptation.

We calculated the economic cost of fungicide resistance as a benchmark corresponding to an optimistic scenario when the farmers across the landscape have agreed to adjust their fungicide application to reach the optimum proportion of treated fields that gives the maximum net return for the entire landscape. An economically optimum decision for an individual farm does not necessarily coincide with the landscape-scale optimum and could represent a social dilemma. Addressing this would need to incorporate farmers' dynamic decision-making into the modeling framework. This would allow us to adjust the fungicide coverage and explore how close to the landscape-scale optimum one can get using different decision-making approaches. Including a socioeconomic point of view brings us closer to identifying policy interventions most likely to maximize crop protection's economic return at a regional scale.

The methodological novelty of our study lies in extending a spatially implicit metapopulation model of plant disease epidemics (Parnell et al., 2006) to incorporate an economic analysis of fungicide use. Landscape perspective is recognized as crucial in plant disease epidemiology (Plantegenest et al., 2007), and more recently, spatially explicit landscape-scale models addressed a range of epidemiological/evolutionary questions focusing mainly on pathogen adaptation to disease resistance in plants (Fabre et al., 2012; Papaïx et al., 2014; Papaïx et al., 2015; Djidjou-Demasse et al., 2017; Rimbaud et al., 2021). While some studies recognized the importance of the economic aspects of plant disease epidemics, they did not incorporate an explicit analysis. Qualitative conceptual frameworks that integrate plant pathogen adaptation to control measures with socioeconomic aspects are only starting to appear (Geffersa et al., 2023) and remain a mathematical challenge (Cunniffe et al., 2014).

Extending the marginal theory of pesticide use (Hillebrandt, 1960; Jørgensen et al., 2017) from the scale of individual farms to a broader scale of cultivated landscapes, we incorporate an economic view in our epidemiological study. In other words, we extended the concept of economic yield to the landscape scale. Earlier, Te Beest et al. (2013) extended the marginal theory of pesticide use by incorporating the stochasticity of plant disease epidemics in their calculation of the optimal fungicide dose that maximizes net return from a fungicide application programme at the scale of an individual farm. This approach has been subsequently developed to incorporate the evolution of fungicide resistance (van den Bosch et al., 2018; Bosch et al., 2020). Nevertheless, focusing on individual farm management, they neglected the externalities of the evolution of resistance. Here, we addressed this issue by extending the marginal theory of pesticide use (Hillebrandt, 1960; Jørgensen et al., 2017) from individual farms to cultivated landscapes in a mathematical approach that incorporates the epidemic development and the

evolution of fungicide resistance and links these processes with an economic analysis. Given that our framework builds on fundamental equations of epidemiology and economic cost assessment, we envision to apply our framework to other crop threats such as insect-pests and weeds.

5   Conclusion

We have developed a bioeconomic framework combining a plant disease epidemiological model with an economic analysis at the landscape scale. Using this approach, we found that the law of diminishing returns does not hold for economic returns from fungicide applications across cultivated landscapes. This is surprising given the almost universal status of this law. The breakdown of diminishing returns suggests that adjusting the amount of fungicide applied to maximize the net return is challenging at the landscape scale, as it needs to be coordinated across multiple farms, and the discontinuous nature of how the optimal coverage depends on fungicide price requires a drastic change in fungicide coverage.

Using the bioeconomic modeling framework, we calculated the economic cost of the evolution of fungicide resistance, C_R, which constitutes our study's primary outcome. We found that C_R depends strongly on the fungicide price, the basic reproduction number of the pathogen, its degree of fungicide resistance and the consequential yield loss. While intuitively, C_R increases with the degree of resistance and yield loss due to disease, surprisingly, it declines with the fungicide price and exhibits a complex non-monotonic pattern as a function of the basic reproduction number. Hence, to estimate C_R, robust estimations of these parameters are necessary, incorporating environmental variation, crop varieties and genetic composition of the pathogen population. Estimating the economic cost of resistance would then inform economic policy instruments, to encourage more sustainable fungicide use. We envision this study providing a conceptual and methodological basis for a fruitful integration of economics and eco-evolutionary dynamics necessary for achieving more sustainable agriculture.

Data Availability Statement

All codes and datasets used in this study are available on the project's GitHub repository (https://github.com/prateek1verma/EpiEcoFarmers). Any necessary dependencies, including specific versions of software packages and libraries, instructions for using the codes, are also documented in the repository's README file.

References

Anderson RM, May RM (1986) The invasion, persistence and spread of infectious diseases within animal and plant communities. Philos Trans R Soc Lond B Biol Sci 314(1167):533–570, doi:10.1098/rstb.1986.0131

Ayres RU, Kneese AV (1969) Production, consumption, and externalities. Am Econ Rev 59(3):282–297, doi:10.1007/978-3-642-27922-5_24

Bargués-Ribera M, Gokhale CS (2020) Eco-evolutionary agriculture: host-pathogen dynamics in crop rotations. Plos Comput Biol 16(1):e1007546, doi:10.1371/journal.pcbi.1007546

van den Bosch F, Lopez-Ruiz F, Oliver R, Paveley N, Helps J, van den Berg F (2018) Identifying when it is financially beneficial to increase or decrease fungicide dose as resistance develops. Plant Pathol 67(3):549–560, doi:10.1111/ppa.12787

Bosch F, Blake J, Gosling P, Helps JC, Paveley N (2020) Identifying when it is financially beneficial to increase or decrease fungicide dose as resistance develops: an evaluation from long-term field experiments. Plant Pathol 69(4):631–641, doi:10.1111/ppa.13155

Bourguet D, Guillemaud T (2016) The hidden and external costs of pesticide use, Springer International Publishing, pp 35–120. doi:10.1007/978-3-319-26777-7_2

Brown JKM, Hovmoller MS (2002) Aerial dispersal of pathogens on the global and continental scales and its impact on plant disease. Science 297(5581):537–41, doi:10.1126/science.1072678

Brue SL (1993) Retrospectives: the law of diminishing returns. J Econ Perspect 7(3):185–192, doi:10.1257/jep.7.3.185

Carolan K, Helps J, van den Berg F, Bain R, Paveley N, van den Bosch F (2017) Extending the durability of cultivar resistance by limiting epidemic growth rates. Proc Royal Soc B 284(1863), doi:10.1098/rspb.2017.0828

Cunniffe NJ, Koskella B, Metcalf CJE, Parnell S, Gottwald TR, Gilligan CA (2014) Thirteen challenges in modelling plant diseases. Epidemics 10:6–10, doi:10.1016/j.epidem.2014.06.002

Diekmann O, Heesterbeek JAP, Metz JAJ (1990) On the definition and the computation of the basic reproduction ratio R0 in models for infectious diseases in heterogeneous populations. J Math Biol 28(4):365–382, doi:10.1007/bf00178324

Djidjou-Demasse R, Moury B, Fabre F (2017) Mosaics often outperform pyramids: insights from a model comparing strategies for the deployment of plant resistance genes against viruses in agricultural landscapes. New Phytol 216(1):239–253, doi:10.1111/nph.14701

van den Driessche P (2017) Reproduction numbers of infectious disease models. Infect Dis Model 2(3):288–303, doi:10.1016/j.idm.2017.06.002

van den Driessche P, Watmough J (2002) Reproduction numbers and sub-threshold endemic equilibria for compartmental models of disease transmission. Math Biosci 180(1-2):29–48, doi:10.1016/s0025-5564(02)00108-6

Fabre F, Rousseau E, Mailleret L, Moury B (2012) Durable strategies to deploy plant resistance in agricultural landscapes. New Phytol 193:1064–1075, doi:10.1111/j.1469-8137.2011.04019.x

Finger R, Möhring N, Dalhaus T, Böcker T (2017) Revisiting pesticide taxation schemes. Ecol Econ 134:263–266, doi:10.1016/j.ecolecon.2016.12.001

Fisher MC, Hawkins NJ, Sanglard D, Gurr SJ (2018) Worldwide emergence of resistance to antifungal drugs challenges human health and food security. Science 742:739–742, doi:10.1126/science.aap7999

Geffersa AG, Burdon JJ, Macfadyen S, Thrall PH, Sprague SJ, Barrett LG (2023) The socio-economic challenges of managing pathogen evolution in agriculture. Philos Trans R Soc Lond, B, Biol Sci 378(1873), doi:10.1098/rstb.2022.0012

Global Plant Health Assessment C (2023) A global assessment of the state of plant health. Plant Dis doi:10.1094/pdis-01-23-0166-fe

Heffernan JM, Smith RJ, Wahl LM (2005) Perspectives on the basic reproductive ratio. J R Soc Interface 2(4):281–293, doi:10.1098/rsif.2005.0042

Hillebrandt PM (1960) The economic theory of the use of pesticides part i. J Agr Econ 13(4):464–472, doi:10.1111/j.1477-9552.1960.tb01667.x

Jørgensen LN, van den Bosch F, Oliver RP, Heick TM, Paveley ND (2017) Targeting fungicide inputs according to need. Annu Rev Phytopathol 55(1):181–203, doi:10.1146/annurev-phyto-080516-035357

Keeling MJ, Rohani P (2008) Modeling infectious diseases in humans and animals. Princeton University Press, doi:10.1515/9781400841035

Kho RM (2000) On crop production and the balance of available resources. Agric Ecosyst Environ 80(1):71–85, doi:10.1016/s0167-8809(00)00135-3

Lee R, den Uyl R, Runhaar H (2019) Assessment of policy instruments for pesticide use reduction in europe; learning from a systematic literature review. Crop Prot 126:104929, doi:10.1016/j.cropro.2019.104929

Lucas JA (2020) Plant pathology and plant pathogens. John Wiley & Sons

Lucas JA, Hawkins NJ, Fraaije BA (2015) The evolution of fungicide resistance, vol 90. Elsevier, doi:10.1016/bs.aambs.2014.09.001

Medlock J, Kot M (2003) Spreading disease: integro-differential equations old and new. Math Biosci 184(2):201–222, doi:10.1016/S0025-5564(03)00041-5

Mikaberidze A, McDonald BA, Bonhoeffer S (2014) Can high risk fungicides be used in mixtures without selecting for fungicide resistance? Phytopathology 104:324–331, doi:10.1094/PHYTO-07-13-0204-R

Mikaberidze A, Mundt CC, Bonhoeffer S (2016) Invasiveness of plant pathogens depends on the spatial scale of host distribution. Ecol Appl 26:1238–1248, doi:10.1890/15-0807

Mikaberidze A, Paveley ND, Bonhoeffer S, van den Bosch F (2017) Emergence of fungicide resistance: the role of fungicide dose. Phytopathology 107:1–16, doi:10.1094/PHYTO-08-16-0297-R

Nutter F, Teng P, Royer M (1993) Terms and concepts for yield, crop loss and disease thresholds. Plant Dis 77:211, doi:10.1094/pd-77-211

Oerke EC (2006) Crop losses to pests. J Agric Sci 144(01):31, doi:10.1017/S0021859605005708

Ostfeld RS, Glass GE, Keesing F (2005) Spatial epidemiology: an emerging (or re-emerging) discipline. Trends Ecol Evol 20(6):328–336, doi:10.1016/j.tree.2005.03.009

Papaïx J, Touzeau S, Monod H, Lannou C (2014) Can epidemic control be achieved by altering landscape connectivity in agricultural systems? Ecol Model 284:35–47, doi:10.1016/j.ecolmodel.2014.04.014

Papaïx J, Burdon JJ, Zhan J, Thrall PH (2015) Crop pathogen emergence and evolution in agro-ecological landscapes. Evol Appl 8(4):385–402, doi:10.1111/eva.12251

Parnell S, van den Bosch F, Gilligan C (2006) Large-scale fungicide spray heterogeneity and the regional spread of resistant pathogen strains. Phytopathology 96(5):549–555, doi:10.1094/PHYTO-96-0549

Pimentel D (2005) Environmental and economic costs of the application of pesticides primarily in the united states. Environ Dev Sustain 7(2):229–252, doi:10.1007/s10668-005-7314-2

Pimentel D, Andow D, Dyson-Hudson R, Gallahan D, Jacobson S, Irish M, Kroop S, Moss A, Schreiner I, Shepard M, Thompson T, Vinzant B (1980) Environmental and social costs of pesticides: a preliminary assessment. Oikos 34(2):126, doi:10.2307/3544173

Plantegenest M, Le May C, Fabre F (2007) Landscape epidemiology of plant diseases. J R Soc Interface 4(16):963–972, doi:10.1098/rsif.2007.1114

Rimbaud L, Fabre F, Papaïx J, Moury B, Lannou C, Barrett LG, Thrall PH (2021) Models of plant resistance deployment. Annu Rev Phytopathol 59(1):125–152, doi:10.1146/annurev-phyto-020620-122134

Savary S, Ficke A, Aubertot JN, Hollier C (2012) Crop losses due to diseases and their implications for global food production losses and food security. Food Secur 4:519–537, doi:10.1007/s12571-012-0200-5

Savary S, Willocquet L, Pethybridge SJ, Esker P, McRoberts N, Nelson A (2019) The global burden of pathogens and pests on major food crops. Nat Ecol Evol 3(3):430–439, doi:10.1038/s41559-018-0793-y

Shephard RW (1970) Proof of the law of diminishing returns. Zeitschrift Für Nationalökonomie Zeitschrift Für Nationalökonomie Zeitschrift Für Nationalökonomie 30(1/2):7–34, doi:10.1007/bf01289990

Te Beest DE, Paveley ND, Shaw MW, Van Den Bosch F (2013) Accounting for the economic risk caused by variation in disease severity in fungicide dose decisions, exemplified for mycosphaerella graminicola on winter wheat. Phytopathology 103(7):666–672, doi:10.1094/phyto-05-12-0119-r

de Vallavieille-Pope C, Giosue S, Munk L, Newton AC, Niks RE, Østergård H, Pons-Kühnemann J, Rossi V, Sache I (2000) Assessment of epidemiological parameters and their use in epidemiological and forecasting models of cereal airborne diseases. Agronomie 20(7):715–727, doi:10.1051/agro:2000171

Willocquet L, Savary S, McDonald BA, Mikaberidze A (2020) A polyetic modelling framework for plant disease emergence. Plant Pathol 69(9):1630–1643, doi:10.1111/ppa.13249

Zadoks JC, Schein RD (1979) Epidemiology and plant disease management. Oxford University Press Inc., doi:10.2307/2402664

A Appendix. Linear stability analysis and computation of fixed points and basic reproduction number

A.1 Model without fungicide treatment

We will first calculate the basic reproduction number for the case when the fungicide treatment is not allowed i.e. $\theta = 0$. It left us with only two variables H_u and I_{uw}, since $H_t = 0$ and $I_{tw} = 0$. In the simple case of no fungicide treatment i.e. $\theta = 0$, the two fixed points for the set of ODEs Eq.(1) are $(H_u^*, I_{uw}^*) = (\frac{\mu}{\beta}, N - \frac{\mu}{\beta})$ and $(N, 0)$. Depending on the rate of infection (β) and recovery (μ), the linear stability analysis of the system produces either of the two scenarios:

- $(N, 0)$ is unstable and $(\frac{\mu}{\beta}, N - \frac{\mu}{\beta})$ is stable when $\beta N > \mu$
- $(N, 0)$ is stable and $(\frac{\mu}{\beta}, N - \frac{\mu}{\beta})$ is unstable when $\beta N < \mu$

The Jacobian of the dynamical system for no fungicide treatment is given by $J(I_{uw}) = \beta(N - 2I_{uw}) - \mu$. Assuming, $H_u(0) \approx N$ and $I_{uw}(0) \approx 0$, we have the following expression for the basic reproduction number using the Routh-Hurwitz criterion:

$$R_0 = \beta N / \mu \tag{A1}$$

At the beginning, when almost all the fields are healthy and $R_0 > 1$; there is an exponential growth in the number of untreated infected fields given by $I_{uw}(t) = I_{uw}(0) \exp[(R_0 - 1)\mu t]$.

A.2 Model with fungicide treatment

A.2.1 Computation of fixed points and their stability for the treatment scenario

When we include fungicide treatment ($\theta > 0$), the epidemiological dynamics of untreated and treated infected fields (I_{uw} and I_{tw}) is given by:

$$\dot{I}_{uw} = \beta[N(1-\theta) - I_{uw}][I_{uw} + I_{tw}] - \mu I_{uw} \tag{A2}$$

$$\dot{I}_{tw} = \beta(1-\epsilon_w)[N\theta - I_{tw}][I_{uw} + I_{tw}] - \mu I_{tw} \tag{A3}$$

Here, we have used equations (2) to replace H_u and H_t from equation (A2)-(A3). In this case, there are three fixed points: FP1, FP2 and FP3. Fixed point 1 (FP1) corresponds to the disease-free equilibrium

$$I_{uw}^{\text{FP1}} = 0 \tag{A4}$$

$$I_{tw}^{\text{FP1}} = 0 \tag{A5}$$

$$H_u^{\text{FP1}} = N(1-\theta) \tag{A6}$$

$$H_t^{\text{FP1}} = N\theta. \tag{A7}$$

Fixed point 2 (FP2) corresponds to the endemic equilibrium

$$I_{uw}^{\text{FP2}} = -\frac{2N\beta_w\epsilon_w\theta - N\beta_w\epsilon_w - N\beta_w + \epsilon_w\mu + \sqrt{\Lambda_1}}{2\beta_w\epsilon_w} \tag{A8}$$

$$I_{tw}^{\text{FP2}} = -\frac{N\beta_w\epsilon_w - N\beta_w - \epsilon_w\mu - 2\left(N\beta_w\epsilon_w^2 - N\beta_w\epsilon_w\right)\theta + \sqrt{\Lambda_1}}{2\left(\beta_w\epsilon_w^2 - \beta_w\epsilon_w\right)} \tag{A9}$$

$$H_u^{\text{FP2}} = N(1-\theta) - I_{uw}^{\text{FP2}} \tag{A10}$$

$$H_t^{\text{FP2}} = N\theta - I_t^{\text{FP2}}, \tag{A11}$$

where

$$\Lambda_1 = \Lambda_0 + 2N\beta_w\epsilon_w\mu(1-\epsilon_w)(1-2\theta) \tag{A12}$$

and

$$\Lambda_0 = N^2\beta_w^2\epsilon_w^2 - 2N^2\beta_w^2\epsilon_w + N^2\beta_w^2 + \epsilon_w^2\mu^2, \tag{A13}$$

Fixed point 3 (FP3) also corresponds to non-trivial values of the four state variables

$$I_{uw}^{\text{FP3}} = -\frac{2 N\beta_w\epsilon_w\theta - N\beta_w\epsilon_w - N\beta_w + \epsilon_w\mu - \sqrt{\Lambda_1}}{2\,\beta_w\epsilon_w} \tag{A14}$$

$$I_{tw}^{\text{FP3}} = -\frac{N\beta_w\epsilon_w - N\beta_w - \epsilon_w\mu - 2\left(N\beta_w\epsilon_w^2 - N\beta_w\epsilon_w\right)\theta - \sqrt{\Lambda_1}}{2\left(\beta_w\epsilon_w^2 - \beta_w\epsilon_w\right)} \tag{A15}$$

$$H_u^{\text{FP3}} = N(1-\theta) - I_{uw}^{\text{FP3}} \tag{A16}$$

$$H_t^{\text{FP3}} = N\theta - I_{tw}^{\text{FP3}}. \tag{A17}$$

However, H_u^{FP3} and H_t^{FP3} exhibit negative values across biologically plausible parameter regimes and therefore do not correspond to a biologically meaningful solution. It is important to note that the names of the fixed points provided in the table A1 may not be strictly correct, as they depend on the stability conditions of the eigenvalues expressions derived using Matlab code in the following Github folder. The stability conditions and the expressions of the fixed points can also be found in Github folder.

Table A1: Fixed points for the treatment scenario

Fixed Point	Nature	I_{uw}^*	I_{tw}^*	H_u^*	H_t^*
FP1	Disease-free	0	0	$\neq 0$	$\neq 0$
FP2	Wildtype endemic	$\neq 0$	$\neq 0$	$\neq 0$	$\neq 0$
FP3	Biologically infeasible	$\neq 0$	$\neq 0$	$\neq 0$	$\neq 0$

A.2.2 Computation of basic reproduction number in the limiting case of a high-efficacy fungicide

Now, let us analyze the special case when the fungicide efficacy $\epsilon_w = 1$ i.e. the treated fields are completely immune to the pathogen's infection. The above equations (A2)-(A3) then becomes,

$$\dot{I}_{uw} = \beta_w[N(1-\theta) - I_{uw}][I_{uw} + I_{tw}] - \mu I_{uw} \tag{A18}$$

$$\dot{I}_{tw} = -\mu I_{tw}. \tag{A19}$$

There are two fixed points of the system of equations ((A18)-(A19)) $(I_{uw}^*, I_{tw}^*) = (0,0)$ and $\left(N(1-\theta) - \frac{\mu}{\beta_w}, 0\right)$. The Jacobian matrix of the above equation is $J = \begin{bmatrix} a_{11} & a_{12} \\ a_{21} & a_{22} \end{bmatrix}$ where the elements are given by:

$$\begin{cases} a_{11} = \beta_w[N(1-\theta) - 2I_{uw}^* - I_{tw}^*] - \mu \\ a_{12} = \beta_w[N(1-\theta) - I_{uw}^*] \\ a_{21} = 0 \\ a_{22} = -\mu \end{cases} \tag{A20}$$

The stability of the fixed point can be evaluated from the trace $\text{tr}(J)$ and the determinant $\det(J)$ of the Jacobian:

$$\text{tr}(J) = a_{11} + a_{22} \tag{A21}$$

$$\det(J) = a_{11}a_{22} - a_{12}a_{21} \tag{A22}$$

Let J_1 and J_2 be the Jacobian matrix of the of the two fixed points $(0,0)$ and $\left(N(1-\theta) - \frac{\mu}{\beta_w}, 0\right)$ respectively. The fixed point $(I_{uw}^*, I_{tw}^*) = (0,0)$ is stable if $\text{tr}(J_1) < 0$ and $\det(J_1) > 0$ which happens for $N\beta_w(1-\theta) - \mu > 0$ or $R_0(1-\theta) < 1$. Similarly, the fixed point $\left(N(1-\theta) - \frac{\mu}{\beta_w}, 0\right)$ is stable when $\text{tr}(J_2) < 0$ and $\det(J_2) > 0$ which happens for $N\beta_w(1-\theta) - \mu > 0$ or $R_0(1-\theta) > 1$. Hence, the infected field can invade the population of healthy fields if $R_0(1-\theta) > 1$ for 100% fungicide efficacy $\epsilon_w = 1$. Thus, $\mathscr{R}_0 = R_0(1-\theta)$ is the effective reproduction number of the system under-treated scenario. Let us now derive the expression for $\mathscr{R}_0$ for a general value of efficacy ϵ_w using next-generation matrix method.

A.2.3 Computation of basic reproduction number for the general case of a partial-efficacy fungicide

In the above section, we derive the condition for the invasion of healthy fields by infected using stability criteria for the the Jacobian matrix. An alternative approach would be to directly derive the reproductive number $\mathscr{R}_0$ for our system by using the next generation matrix (Diekmann et al., 1990; van den Driessche and Watmough, 2002). The proof and detailed discussion of this method is given in Diekmann et al. (1990); van den Driessche and Watmough (2002). Here we will follow the prescription outlined in van den Driessche (2017) to evaluate $\mathscr{R}_0$.

Consider the equation, $\frac{dx_i}{dt} = \mathscr{F}_i(x) - \mathscr{V}_i(x)$ for $i = 1, 2, ..., m$ where $x = [x_1, x_2, ..., x_n]$ is the number of individuals in various compartments and $m < n$ is the total number of infected compartment. Now if $F = \frac{\partial \mathscr{F}_i(x_0)}{\partial x_j}$ and $V = \frac{\partial \mathscr{V}_i(x_0)}{\partial x_j}$ for $1 \le i, j \le m$ and x_0 is the infection free state then $\mathscr{R}_0$ is the maximum real part of the eigenvalues of FV^{-1}. When there is no infected field $x_0 = [H_u, H_t, I_{uw}, I_{tw}] = [N(1-\theta), N\theta, 0, 0]$ then

$$F = \begin{bmatrix} \beta_w N(1-\theta) & \beta_w N(1-\theta) \\ \beta_w(1-\epsilon_w)N\theta & \beta_w(1-\epsilon_w)N\theta \end{bmatrix}; \quad V = \begin{bmatrix} \mu & 0 \\ 0 & \mu \end{bmatrix}; \tag{A23}$$

Hence FV^{-1} is given by

$$FV^{-1} = \begin{bmatrix} N\beta_w(1-\theta)/\mu & N\beta_w(1-\theta)/\mu \\ N\beta_w(1-\epsilon_w)\theta/\mu & N\beta_w(1-\epsilon_w)\theta/\mu \end{bmatrix} \tag{A24}$$

The $\mathscr{R}_0$ is the root corresponding to the maximum real part of the characteristic equation,

$$z^2 - \left(\frac{N\beta_w(1-\epsilon_w\theta)}{\mu} \right) z \tag{A25}$$

Hence, $\mathscr{R}_0 = \frac{N\beta_w(1-\epsilon_w\theta)}{\mu} = R_0(1-\epsilon_w\theta)$ where $R_0 = N\beta_w/\mu$. The expression for the basic reproduction number for the general case of a partial-efficacy fungicide has also been derived via Matlab code provided in the Github repository.

A.3 Model with fungicide treatment and resistance

A.3.1 Special case 1: Computation of fixed points and their stability for a high-efficacy fungicide and absolute resistance

The equations corresponding to the epidemiological dynamics for the fungicide resistance case, using Eq. (7) and assuming $\epsilon_w = 1$ and $\epsilon_r = 0$, are given by:

$$\dot{I}_{uw} = \beta_w(I_{uw} + I_{tw})(N(1-\theta) - I_{uw} - I_{ur}) - \mu I_{uw} \tag{A26}$$

$$\dot{I}_{tw} = -\mu I_{tw} \tag{A27}$$

$$\dot{I}_{ur} = \beta_r(I_{ur} + I_{tr})(N(1-\theta) - I_{uw} - I_{ur}) - \mu I_{ur} \tag{A28}$$

$$\dot{I}_{tr} = \beta_r(I_{ur} + I_{tr})(N\theta - I_{tw} - I_{tr}) - \mu I_{tr} \tag{A29}$$

Table A2: Fixed points of the ODE system with fungicide resistance in the limiting case of $\epsilon_w = 1$ and $\epsilon_r = 0$ (Special Case 1)

Fixed Point	Nature	I_{uw}^*	I_{tw}^*	I_{ur}^*	I_{tr}^*
FP1	Disease-free	0	0	0	0
FP2	Wildtype endemic	$\neq 0$	0	0	0
FP3	Resistant strain endemic	0	0	$\neq 0$	$\neq 0$
FP4	Coexistence	$\neq 0$	0	$\neq 0$	$\neq 0$

It is important to note that the names of the fixed points provided in the table A3 may not be strictly correct, as they depend on the stability conditions. Using Eq (A26)-(A29), we can compute four Jacobian matrices for four fixed point conditions. We know that a fixed point is

stable if all the eigenvalues of the corresponding Jacobian matrix are negative. Hence, we can derive the conditions on the parameters of the systems for all the fixed point to remain stable. The expressions for the stability conditions and the fixed points can be found in the following Github folder while the corresponding Matlab code can be found in the following Github folder.

A.3.2 Special case 2: Computation of fixed points and their stability for a high-efficacy fungicide and partial resistance

When the fungicide has the highest efficacy ($\epsilon_w = 1$), but resistance is partial ($0 < \epsilon_r < 1$), there are five fixed points described below:

Table A3: Fixed points of the ODE system with fungicide's partial resistance in the limiting case of $\epsilon_w = 1$ (Special Case 2)

Fixed Point	Nature	I_{uw}^*	I_{tw}^*	I_{ur}^*	I_{tr}^*
FP1	Disease-free equilibrium	0	0	0	0
FP2	Resistant strain endemic (unstable)	0	0	$\neq 0$	$\neq 0$
FP3	Wild type strain endemic	$\neq 0$	0	0	0
FP4	Coexistence	$\neq 0$	0	$\neq 0$	$\neq 0$
FP5	Resistant strain endemic	0	0	$\neq 0$	$\neq 0$

It is important to note that the names of the fixed points provided in the table A3 may not be strictly correct, as they depend on the stability conditions. The non-zero expressions in the table A3 for I_{uw}^*, I_{tw}^*, I_{ur}^* & I_{tr}^* and the stability conditions of the fixed point are too large to be displayed in the above table hence we provide them in the following Github folder while the corresponding Matlab code can be found in the this Github folder.

A.3.3 General Case: Computation of fixed points and their stability with the fungicide resistance

When ϵ_w and ϵ_r have general values then there are six fixed points for the epidemiological dynamics for Eq. (7). These fixed points are given in the table A4.

Table A4: Fixed Points of the ODE System with fungicide resistance (General Case)

Fixed Point	Nature	I_{uw}^*	I_{tw}^*	I_{ur}^*	I_{tr}^*
FP1	Disease-free equilibrium	0	0	0	0
FP2	Coexistence	$\neq 0$	$\neq 0$	$\neq 0$	$\neq 0$
FP3	Wildtype endemic	$\neq 0$	$\neq 0$	0	0
FP4	Resistant endemic	0	0	$\neq 0$	$\neq 0$
FP5	Resistant endemic	0	0	$\neq 0$	$\neq 0$
FP6	Wildtype endemic	$\neq 0$	$\neq 0$	0	0

It is important to note that the names of the fixed points provided in the table A4 may not be strictly correct, as they depend on the stability conditions. The non-zero expressions in the table A4 for I_{uw}^*, I_{tw}^*, I_{ur}^* & I_{tr}^* and the stability conditions of the fixed point are too large to be displayed in the table, hence we provide them in the following Github folder while the corresponding Matlab code can be found in this Github folder.

A.3.4 Computation of basic reproduction number for the general case with the fungicide resistance

We follow the previous formulation of computing $\mathscr{R}_0$ through the next-generation matrix method for the case of fungicide resistance. The epidemiological dynamics of the pest infection in the field are given in the equations (7). Using equation epidemiological equations (7) for the pest infection, matrix F and V can be evaluated as:

$$F = \begin{bmatrix} \beta_w N(1-\theta) & \beta_w N(1-\theta) & 0 & 0 \\ \beta_w(1-\epsilon_w)N\theta & \beta_w(1-\epsilon_w)N\theta & 0 & 0 \\ 0 & 0 & \beta_r N(1-\theta) & \beta_r N(1-\theta) \\ 0 & 0 & \beta_r(1-\epsilon_r)N\theta & \beta_r(1-\epsilon_r)N\theta \end{bmatrix} \tag{A30}$$

$$V = \begin{bmatrix} \mu & 0 & 0 & 0 \\ 0 & \mu & 0 & 0 \\ 0 & 0 & \mu & 0 \\ 0 & 0 & 0 & \mu \end{bmatrix} \tag{A31}$$

Here, we assume that initially there are no infected fields $x_0 = [H_u, H_t, I_{uw}, I_{tw}, I_{ur}, I_{tr}] = [N(1-\theta), N\theta, 0, 0, 0, 0]$. $\mathcal{R}_0$ is the maximum real part of the eigenvalues of FV^{-1}. FV^{-1} has two real valued eigenvalues $\lambda_1 = N\beta_w(1-\epsilon_w\theta)/\mu$ and $\lambda_2 = N\beta_r(1-\epsilon_r\theta)/\mu$. Hence, $\mathcal{R}_0 = \max(\lambda_1, \lambda_2)$.

B Appendix. Maximizing economic return of fungicide treatments

B.1 Optimal intensity of fungicide treatments in individual fields versus cultivated landscapes

B.1.1 Optimal net return in an individual field

When managing disease in an individual field, growers can adjust the intensity of control by optimizing the fungicide dose. To determine the optimal dose, we consider the expression for the net return that is similar to Eq. (14) in the main text, but here we formulate it for the case of an individual field

$$g = Y - F, \tag{B1}$$

where

$$Y = Y_0 + Y_m \frac{D}{D + D_{50}} \tag{B2}$$

is the yield (expressed as the fraction of the maximum yield converted into its monetary equivalent) when a dose D of the fungicide has been applied. The first term in Eq. (B2) is the yield without fungicide application, Y_0. The second term in Eq. (B2) is the yield benefit of the fungicide application, where the functional form used is a simplification of the Hill function, which describes well typical empirical dose-response relationships for fungal diseases in cereal crops Mikaberidze et al. (2017). The parameter Y_m represents the maximum effect at large doses, and D_{50} corresponds to the dose for which half of the maximum effect is achieved. Hence, $Y_0 + Y_m$ corresponds to the attainable yield Zadoks and Schein (1979); Nutter et al. (1993).

The second term, F, in Eq. (B1) is the total fungicide price. We consider it to be proportional to the fungicide dose D

$$F = qD \tag{B3}$$

where q is the price per unit fungicide. To compute the net return in Eq. (B1), we measure F in the same units as Y: as the fraction of the maximum yield converted into its monetary equivalent. We substitute Eq. (B2) and Eq. (B3) into Eq. (B1) and obtain the following expression of the net return as a function of the fungicide dose

$$g = Y_0 + Y_m \frac{D}{D + D_{50}} - qD. \tag{B4}$$

Net return, g, in Eq. (B4) exhibits a maximum at the fungicide dose

$$D^* = -D_{50} + \sqrt{D_{50}y_m/q}. \tag{B5}$$

If the fungicide becomes sufficiently expensive ($q > Y_m/D_{50}$), the expression for the optimum dose in Eq. (B5) becomes negative, meaning that the net return is maximized at when no fungicide is applied. For less expensive fungicides, the net return does exhibit a maximum at positive doses. The associated optimum dose is larger for cheaper fungicides, and larger for fungicides with a higher maximum effect, Y_m.

The yield that corresponds to the maximum net return is then given by

$$Y_{\text{econ}} = Y_0 + Y_m - \sqrt{Y_m D_{50} q}, \tag{B6}$$

and is defined as the economic yield Zadoks and Schein (1979); Nutter et al. (1993). The economic yield decreases from the attainable yield, $Y_0 + Y_m$ (achieved if the fungicide is very cheap, i.e. $q \approx 0$), by a value proportional to the square root of the fungicide price per unit q. We use the expression for the net return in Eq. (B4) above and its decomposition into Y and F to illustrate the law of diminishing returns in the context of an individual field in Fig. 2A and B of the main text.

B.1.2 Optimal net return in a cultivated landscape

As in the section above, we decompose the expression for the relative net return into the total yield, $Y(\theta)$, and the total fungicide price, $F(\theta)$, but now in the context of a cultivated landscape and hence both of these are functions of θ:

$$g(\theta) = Y(\theta) - F(\theta), \tag{B7}$$

where

$$Y(\theta) = 1 - (1 - y)I^*(\theta)/N. \tag{B8}$$

Here, $I^*(\theta)$ is number of diseased fields at a stable fixed point that can be computed from the model Eq. (7) in the main text.

$$F(\theta) = f\theta, \tag{B9}$$

where $f = c/y_H$ is the relative fungicide price. We use the expression for the net return in Eq. (B7) above and its decomposition into $Y(\theta)$ and $F(\theta)$ to illustrate the absence of diminishing returns when controlling disease across multiple fields in a landscape in Fig. 2C and D of the main text.

B.2 More detailed analysis of the optimal fungicide coverage across landscapes

We investigate the scenario when the system has reached a stable fixed point and analyze the expression for the associated net gain given by Eq. (14) in the main text

$$g(\theta) = (H^* + I^* y)/N - \theta f, \tag{B10}$$

B.2.1 Optimal fungicide coverage without resistance

The case of high R_0 when fungicide treatment cannot drive the pathogen population to extinction ($\theta_{cw} > 1$). In this regime, the fixed point that corresponds to the endemic equilibrium (FP2 in Appendix A.2.1) is stable and is therefore realized (over sufficiently large times) across the entire range of θ-values. The associated equilibrium values of the state variables are given by Eqs. (A8)-(A11) in Appendix A.2.1. We calculate the net return by substituting these values into Eq. (B10). The dependence $g(\theta)$ is illustrated in Figure 3A. The four curves in Fig. 3A differ only in α-values and represent four qualitatively different scenarios. The four curves start from the same value at $\theta = 0$, because in the absence of fungicide application, the net return $g(\theta)$ in Eq. (B10) does not depend on the fungicide price f and therefore is also independent of α. Which of the four regimes is realized depends on how α relates to the three constants A_1, A_2 and A_3 that are independent of θ, but do depend on the pathogen's basic reproductive number in the absence of treatment $R_0 = \beta N/\mu$ and on the fungicide efficacy ϵ_w:

$$A_1 = \frac{1}{R_0 \left(1 - \epsilon_w\right)/\epsilon_w + 1}. \tag{B11}$$

$$A_2 = \frac{\epsilon_w}{1 - \epsilon_w}/R_0 \tag{B12}$$

$$A_3 = \frac{1}{R_0(1 - \epsilon_w)/\epsilon_w - 1}. \tag{B13}$$

When $\alpha < A_1$, the net return grows monotonically as we treat more fields and thereby increase θ (see Fig. 3A). When $A_2 < \alpha < A_1$, $g(\theta)$ exhibits a minimum at an intermediate value of θ (see Fig. 3A). If in addition $\alpha < A_3$ (i.e., $A_2 < \alpha < A_3$), then $g(\theta = 1) > g(\theta = 0)$, meaning that treating all fields gives a higher net gain than no treating any fields at all; moreover $\theta = 1$ gives the highest net gain among all θ-values under this scenario (see Fig. 3A). Alternatively, if $A_3 < \alpha < A_1$ the dependency $g(\theta)$ still exhibits a minimum at an intermediate value of θ, but now $g(\theta = 0) > g(\theta = 1)$, meaning that not applying the fungicide in any of the fields gives a higher net return than applying the fungicides on all fields (see Fig. 3A). In this case, $\theta = 0$ gives the highest net return among all θ-values. Finally, when $\alpha > A_1$, the net return $g(\theta)$ decreases monotonically with increasing θ and not treating any fields at all (i.e., $\theta = 0$) maximizes the net gain (see Fig. 3A). To summarize, when $\theta_{cw} > 1$, to maximize the net return, one should either treat all fields ($\theta^* = 1$ if $\alpha < A_2$), or treat no fields ($\theta^* = 0$ if $\alpha > A_2$).

The case of low R_0 when fungicide treatment can drive the pathogen population to extinction ($\theta_{cw} < 1$). Next, consider the regime when the critical fraction of treated fields in (5) is lower than one ($\theta_{cw} < 1$). In this regime, two different fixed points can become stable and be realized (over sufficiently large times) depending on the values of θ. At $\theta < \theta_{cw}$, the fixed point corresponding to the endemic equilibrium is stable (FP2), while at $\theta > \theta_{cw}$, the fixed point corresponding to the disease-free equilibrium is stable. Also here, $g(\theta)$ exhibits four qualitatively different regimes depending on how the parameter α relates to the three constants A_1, A_2^* and A_3^*. Here, A_1 is given by Eq. (B11), A_2^* and A_3^* are given by

$$A_2^* = \epsilon_w \tag{B14}$$

$$A_3^* = \frac{\epsilon_w/(2 - \epsilon_w)}{1 - R_0(1 - \epsilon_w)/(2 - \epsilon_w)}. \tag{B15}$$

A_2^* and A_3^* do not depend on θ, but do depend on R_0 and ϵ_w.

The net return versus the fungicide coverage $g(\theta)$ exhibits different patterns in the two different ranges of θ-values. At $\theta < \theta_{cw}$, the behavior of $g(\theta)$ is the same as in the scenario $\theta_{cw} > 1$ described above: it exhibits the four qualitatively different regimes depending on where the α-value lies with respect to A_1, A_2^* and A_3^*. The four regimes are illustrated by the four curves in Fig. 3B. The four regimes are essentially the same as the four regimes in the case $\theta_{cw} > 1$ illustrated in Fig. 3A, but here the range of θ-values in which these regimes are realized is restricted to $0 < \theta < \theta_{cw}$. Beyond this range, when $\theta > \theta_{cw}$, $g(\theta)$ decreases monotonically as we increase θ in all four cases. This is because in this range, the disease-free equilibrium is realized and fungicide treatment does not provide any benefit in terms of disease control, but still incurs a cost. The optimal fungicide coverage, θ^* is determined by where α lies with respect to A_2^*. For $\alpha < A_2^*$, it is best to treat the fraction $\theta = \theta_{cw}$ of all fields, while for $\alpha > A_2^*$ it is best not to treat any fields at all, i.e., to have $\theta^* = 0$.

B.2.2   Optimal fungicide coverage with resistance

What fraction of fields should be treated with a fungicide to maximize the net return if resistance is present?

High-efficacy fungicide and absolute resistance For simplicity, we first consider the case of a fungicide with the highest efficacy ($\epsilon_w = 1$) and absolute resistance ($\epsilon_r = 0$). As we did above in the scenario without resistance, we investigate the scenario when the system has reached a stable fixed point and analyze the expression for the associated net return given by Eq. (B10) above (Eq. (14) in the main text) But here, $H^* = H_u^* + H_t^*$ and $I^* = I_{uw}^* + I_{tw}^* + I_{ur}^* + I_{tr}^*$. These represent the values of the state variables at a stable fixed point. The system has four fixed points described in Sec. A.3.1 together with their stability conditions. These enter into the expression Eq. (B10) for the net return and determine its dependency on θ.

When fungicide resistance confers no fitness cost to the pathogen ($\gamma = 0$), the fixed point that corresponds to the endemic equilibrium of the resistant strain (FP3 in Sec. A.3.1) is stable (and therefore realized) across the entire range of θ-values, i.e. $0 < \theta < 1$. We substitute the expression of the state variables corresponding to this fixed point into the relative net gain in Eq. (A.3.1) above to find out how $g(\theta)$ depends on θ in this regime: $g(\theta) = y + (1 - y)/\left[R_0(1 - \gamma)\right] - r\theta$, where $R_0 = \beta N / \mu$ is the basic reproduction number of the wildtype strain in the absence of fungicide treatment [Eq. (A1)]. In this case, net gain, $g(\theta)$, is a decreasing linear function of θ, where the slope is given by the fungicide price f. This is because in the entire range of θ-values, the resistant strain outcompetes the sensitive strain leading to a complete loss of fungicide efficacy. Note that for more expensive fungicides, the net return drops faster as a function of θ (Fig. B1A). In this regime, not treating any fields with the fungicide maximizes net return; hence the optimal θ is zero.

When fungicide resistance does confer a fitness cost to the pathogen ($\gamma > 0$), there are three fixed points that can be stable within the permissible range of θ-values: (i) the endemic equilibrium of the wild type strain ("wt wins", FP2 in Sec. A.3.1) at small θ-values; (ii) the endemic coexistence of the wild type and the resistant strains ("wt and res coexist", FP4 in Sec. A.3.1) at intermediate θ-values, and (iii) the endemic equilibrium of the resistant strain ("res wins", FP3 in Sec. A.3.1) at large θ-values. The three domains of stability are demarcated by the grey vertical lines in Fig. B1B. The left vertical line represents the transition from FP2 ("wt wins") to FP4 ("wt and res coexist") when θ is increased and is given by

$$\theta_{\text{FP2}\to\text{FP4}} = \gamma / \left[R_0(1 - \gamma)\right], \tag{B16}$$

The right vertical line represents the transition from FP4 ("wt and res coexist") to FP3 ("res wins") and is given by $\theta = \gamma$.

In this regime, there can be two qualitatively different scenarios of how net gain depends on the proportion of treated fields. In the first scenario, the fungicide is sufficiently cheap: it costs less than the yield loss due to disease, $f < 1 - y$. In this case, the net return increases over small θ-values, reaches a maximum at $\theta = \theta_{FP2->FP4}$ and then decreases in a monotonic linear manner as θ increases further with the slope given by the relative fungicide price f (see in Fig. B1B). In this scenario, the optimal proportion of treated fields that maximizes the net gain is given by $\theta_{FP2->FP4}$ in Eq. (B16) (grey vertical line in Fig. B1B on the left). It is proportional to the fitness cost of resistance γ and inversely proportional to the basic reproductive number R_0. Thus, for sufficiently cheap fungicides, there is a non-zero optimal proportion of treated fields which tends to assume rather small values, as the fitness cost is not expected to exceed 0.3. In the second scenario, the fungicide is more expensive than the yield loss due to disease, $f > 1 - y$. In this case, net return decreases monotonically across the entire range of θ-values (see Fig. B1B for $f = 0.5$ and $f = 0.75$). This means that the net return is maximum when no fields are treated ($\theta = 0$). Thus, in the case of absolute resistance and high-efficacy fungicides, the net return is maximized either at zero or small θ-values, and the optimal fraction of treated fields does not depend on the cost ratio α. Hence, the economically sound way to deal with fungicide resistance in this case is to strongly reduce or stop the fungicide treatment.

High-efficacy fungicide and partial resistance When resistance is partial ($0 < \epsilon_r < 1$) the situation is quite different. Here, the qualitative pattern of how $g(\theta)$ depends on θ is determined by the cost ratio $\alpha = f/(1 - y)$ [Eq. (15)]. We first focus on a simpler case of no fitness cost ($\gamma = 0$), and then consider the case when a fitness cost is present ($\gamma > 0$). How the net return depends on θ here is qualitatively similar to the case without resistance discussed above. Since here the resistance is only partial, the fungicide can still suppress the resistant pathogen strain to an extent given by the efficacy parameter ϵ_r. If ϵ_r is high enough, then treating a sufficiently high proportion of fields can drive the resistant strain to extinction. This

Figure B1: What is the optimal fraction of treated fields in the presence of fungicide resistance (absolute resistance)? Relative net gain $g/(y_H N)$ is plotted versus the fraction of treated fields θ according to Eq. (B10), assuming a high-efficacy fungicide ($\epsilon_w = 1$) and absolute resistance ($\epsilon_r = 0$). Panel A: resistance incurs no fitness cost ($\gamma = 0$). Panel B: resistance does incur a fitness cost ($\gamma = 0.3$). The three curves in each panel represent different values of the relative fungicide price f: 0.1 (solid blue), 0.25 (dashed red), 0.5 (dotted orange), 0.75 (dash-dotted black). Other parameter values are the same for all curves in both panels: $\beta = 0.02$, $R_0 = \beta N/\mu = 4$; $N = 1000$, $\mu = 5$, $y = 0.6$.

critical proportion of treated fields is given by

$$\theta_{cr} = \frac{1}{\epsilon_r}\left(1 - \frac{1}{R_0}\right).$$ (B17)

The expression in Eq. (B17) is similar to the critical proportion θ_c given by Eq. (5), but here the entire expression is scaled by the fungicide efficacy against the resistant strain, ϵ_r. This difference reflects the fact that the critical proportion here is computed when controlling the resistant rather than the wildtype strain. The associated critical efficacy against the resistant strain above which $\theta_{cr} < 1$ is given by

$$\epsilon_{cr} = 1 - \frac{1}{R_0}$$ (B18)

The regime when $\theta_{cr} > 1$ is realized when the fungicide efficacy against the resistant strain is below the critical value, $\epsilon_r < \epsilon_{cr}$. In this regime, the only fixed point that is stable and realized across the entire permissible range of θ-values is the endemic equilibrium of the resistant strain (FP5 "res wins" in Sec. A.3.2). The other regime, when $\theta_{cr} < 1$ is realized when the fungicide efficacy against the resistant strain is above the critical value, $\epsilon_r > \epsilon_{cr}$. In this regime, two fixed points can be stable within the permissible range of θ-values: the endemic equilibrium of the resistant strain (FP5 "res wins" in Sec. A.3.2) at low θ-values and the disease-free equilibrium (FP1 "disease-free" in Sec. A.3.2) at high θ-values. The value of θ at which the domain of stability changes from FP5 to FP1 as we increase θ is given by θ_{cr} in Eq. (10) above.

In each of these regimes, we computed the equilibrium values of the state variables that correspond to stable fixed points and substitute them into the expression for the net gain Eq. (B10) to investigate its dependency on the proportion of treated fields θ. This dependency is illustrated in Fig. 4, where we distinguished the regimes when $\theta_{cr} > 1$ (panel A) and $\theta_{cr} < 1$ (panel B). The four curves in Fig. 4A and 4B differ only in α-values ($\alpha = r/(1 - y)$ Eq. (15)) and represent four qualitatively different scenarios.

Consider the case when the degree of partial resistance is high ($\epsilon_r < \epsilon_{cr}$, i.e. $\theta_{cr} > 1$; illustrated in Fig. 4A). Which of the four regimes are realized depends on how the parameter α relates to the three constants B_1, B_2 and B_3 that are independent of θ, but do depend on the basic reproductive number R_0 and the fungicide efficacy against the resistant strain ϵ_r:

$$B_1 = \frac{1}{R_0\left(1 - \epsilon_r\right)/\epsilon_r + 1},$$ (B19)

$$B_2 = \frac{\epsilon_r}{1 - \epsilon_r}/R_0. \tag{B20}$$

$$B_3 = \frac{1}{R_0\left(1 - \epsilon_r\right)/\epsilon_r - 1}, \tag{B21}$$

The expressions for B_1, B_2 and B_3 above are the same as the expressions for A_1, A_2 and A_3 in (B11)-(B13) above, but with ϵ_r instead of ϵ_w.

Consider when the degree of partial resistance is low ($\epsilon_r > \epsilon_{cr}$, i.e. $\theta_{cr} > 1$; illustrated in Fig. 4B). As before, which of the four regimes is realized here depends on how the parameter α relates to the three constants B_1, B_2^* and B_3^* that are independent of θ, but do depend on the basic reproductive number R_0 and the fungicide efficacy against the resistant strain ϵ_r. The constant B_1 is given by Eq. (B19) and the expressions for B_2^* and B_3^* read:

$$B_2^* = \epsilon_r, \tag{B22}$$

$$B_3^* = \frac{\epsilon_r/(2 - \epsilon_r)}{1 - R_0(1 - \epsilon_r)/(2 - \epsilon_r)}. \tag{B23}$$

When the fitness cost is present ($\gamma > 0$), and $\theta_{cr} > 1$, three fixed points can be stable within the permissible range of θ-values: the endemic equilibrium of the wildtype strain (FP2 "wt wins" in Sec. A.3.2) at low θ-values, the stable co-existence of the wild type and the resistance strains (FP3 "wt and res coexist" in Sec. A.3.2) and the endemic equilibrium of the resistant strain (FP5 "res wins" in Sec. A.3.2) over intermediate θ-values. In the different scenario, when $\theta_{cr} < 1$, four fixed points can be stable within the permissible range of θ-values: the three fixed points described above plus the disease-free equilibrium (FP1 "disease-free" in Sec. A.3.2) at high θ-values.

Based on the stability conditions of FP2, we calculated the θ-value at which the domain of stability switches from FP2 to FP3 as we increase θ:

$$\theta_{\mathrm{FP2 \to FP3}} = \gamma/\left[R_0(1 - \epsilon_r)(1 - \gamma)\right]. \tag{B24}$$

Eq. (B24) represents a generalization of Eq. (B16) for the case of partial resistance.

Figure B2: What is the optimal fraction of treated fields in the presence of fungicide resistance (partial resistance)? Relative net gain $g/(y_H N)$ is plotted versus the fraction of treated fields θ according to Eq. (B10), considering a fitness cost of resistance ($\gamma = 0.1$), a high-efficacy fungicide ($\epsilon_w = 1$) and partial resistance ($0 < \epsilon_r < 1$). Panel A: a low degree of partial resistance $\epsilon_r = 0.6 < \epsilon_{cr} = 0.75$, with values of α: 0.1 (solid blue), 0.5 (dashed red), 1.1 (dotted orange), 1.5 (dash-dotted black). Panel B: a high degree of partial resistance $\epsilon_r = 0.8 > \epsilon_{cr}$, with values of α: 0.1 (solid blue), 0.85 (dashed red), 1.1 (dotted orange), 1.5 (dash-dotted black) Other parameter values are the same for all curves in both panels: $\beta = 0.02$, $R_0 = \beta N/\mu = 4$; $N = 1000$, $\mu = 5$, $y = 0.5$.

As we did above, we computed the equilibrium values of the state variables that correspond to stable fixed points and substituted them into the expression for the net return Eq. (B10) to investigate its dependency on the proportion of treated fields θ. This dependency is illustrated in Fig. B2, where we distinguished the regimes when $\theta_{cr} > 1$ (panel A) and $\theta_{cr} < 1$ (panel B).

The four curves in Fig. B2A differ only in α-values ($\alpha = f/(1-y)$ Eq. (15)) and represent the three qualitative scenarios that correspond to different optimal θ-values. In the first regime, treating all fields maximizes the net gain and hence $\theta = 1$ is optimal (see Fig. B2A for $\alpha = 0.1$). This is realized when $\alpha < 1$ and $\alpha < C_1$ or when $\alpha > 1$ and $\alpha < C_2$, where

$$C_1 = \frac{\epsilon_r(1-\gamma)}{R_0\left(1-\gamma\right)\left(1-\epsilon_r\right) - \gamma}, \tag{B25}$$

$$C_2 = \frac{\epsilon_r + \gamma(1-\epsilon_r)}{R_0\left(1-\gamma\right)\left(1-\epsilon_r\right)}. \tag{B26}$$

In the second regime, an intermediate value $\theta = \theta_{FP2 \to FP3}$ (given by Eq. (B24), left vertical line in Fig. B2A for for $\alpha = 0.5$) is optimal, which is realized when $\alpha < 1$ and $\alpha > C_1$ (see Fig. B2A). In the third regime, not treating any fields maximizes the net gain and hence $\theta = 0$ is optimal (see Fig. B2A for $\alpha = 1.5$ and $\alpha = 1.1$). This is realized when $\alpha > 1$ and $\alpha > C_2$.

The scenario $\theta_{cr} < 1$ is illustrated in Fig. B2B. Here, there are also three qualitative regimes corresponding to the different optimal values of θ. But in this regime, $\theta = 1$ is never optimal, because the disease-free equilibrium is stable at high θ-values and in this range (to the right from the right-most vertical line in Fig. B2B) the net gain given by $g(\theta) = 1 - f\theta$ decreases linearly with θ. The θ-value at the point of transition from the domain of stability of FP5 ("res wins") to FP1 ("disease-free equilibrium") is given by Eq. (10).

Under this scenario, in the first regime, treating the proportion $\theta = \theta_{cr}$ of fields maximizes the net gain. This is realized when $\alpha < 1$ and $\alpha < C_1^*$ or when $\alpha > 1$ and $\alpha < C_2^*$, where

$$C_1^* = \epsilon_r, \tag{B27}$$

$$C_2^* = \frac{\epsilon_r(1-\gamma)(R_0 - 1)}{R_0(1-\gamma) - 1}. \tag{B28}$$

In the second regime, a lower intermediate proportion of treated fields $\theta = \theta_{FP2 \to FP3}$ maximizes net gain, which occurs when $\alpha < 1$ and $\alpha > C_1^*$. In the third regime, treating no fields maximizes the net gain ($\theta = 0$ optimal), which is realized when $\alpha > 1$ and $\alpha > C_2^*$.

C   Appendix. Economic cost of fungicide resistance

C.1   Calculation of the economic cost of fungicide resistance

Here, we provide a more detailed mathematical description of the the economic cost of resistance, C_R. We present analytical expressions for C_R as a function of the relative fungicide price, f, in different parameter regimes. First, consider a typical dependence of C_R on f shown in Fig. 5C in the main text. For low f-values, C_R remains constant and is given by

$$C_R = \frac{(\epsilon_w - \epsilon_r)(1-y)}{(1-\epsilon_w)(1-\epsilon_r)}\frac{1}{R_0}. \tag{C1}$$

For intermediate f-values, C_R decreases linearly with f according to

$$C_R = \frac{\epsilon_w(1-y)}{R_0(1-\epsilon_w)} - f. \tag{C2}$$

The expressions in Eq. (C1) and Eq. (C2) are valid when the pathogen's basic reproduction number, R_0, exceeds both of its critical values R_{0cw} [Eq. (6)] and R_{0cr} [Eq. (12)]. This means that when all fields are treated, $\theta = 1$, the pathogen's basic reproduction number exceeds one both for the wildtype and the resistant strain. But when the basic reproduction number is in the range $R_{0cr} < R_0 < R_{0cw}$, the wildtype strain can be driven to extinction at an intermediate treatment coverage $\theta = \theta_{cw} < 1$ (where θ_{cw} is given by Eq. (5)), but not the resistant strain. For

this reason, at small f-values the net return without resistance is maximized at this intermediate treatment coverage $\theta = \theta_c$, while the in the presence of resistance treating all fields ($\theta^* = 1$) remains the optimum. In this regime, the economic cost of fungicide resistance is given by

$$C_R = \frac{1 - R_0(1 - \epsilon_w)}{\epsilon_w R_0} f - \frac{(1 - R_0(1 - \epsilon_r))(1 - y)}{R_0(1 - \epsilon_r)}. \tag{C3}$$

Here, C_R increases linearly as a function of the fungicide price, f.

When the pathogen's basic reproduction number is high enough ($R_0 > R_{0cw}$ and $R_0 > R_{0cr}$) and the fungicide is cheap enough ($f < f_1$), the economic cost of fungicide resistance is inversely proportional to the pathogen's basic reproduction number. This can be understood on the basis of the expression for the net return in Eq. (14). The reduction in the net return due to disease is proportional to the fraction of diseased fields I^*/N. This fraction is given by $1 - 1/[R_0(1 - \epsilon_w)]$ in the absence of resistance when endemic equilibrium is reached for the wildtype strain. When the resistance is present and reaches an endemic equilibrium, the fraction of diseased fields is $1 - 1/[R_0(1 - \epsilon_r)]$. In order to compute the economic cost of fungicide resistance, we need to subtract the net return with resistance from the net return without resistance. For small f-values, when treating all fields maximizes the net return both with and without resistance, the second term in the expression for the net return [Eq. (14)] is the only term that depends on the presence of resistance. Hence, the economic cost of resistance is proportional to the difference between the fraction of diseased fields without resistance and the fraction of diseased fields with resistance. This difference is inversely proportional to the basic reproduction number. The fraction of diseased fields is higher for pathogens with a higher basic reproduction number, R_0, both in the absence and in the presence of resistance. But this increase is higher in the presence of resistance than in the absence of resistance. For this reason, resistance incurs an economic cost. However, because of the nature of the inverse proportionality, at higher R_0-values the difference between the proportion of diseased fields with resistance and without resistance diminishes, and hence the economic cost of resistance is reduced.

C.2 Dependence of the economic cost of resistance on fungicide price and basic reproduction number

In Sec. 3.2 of the main text, we focused on the scenario when the pathogen's basic reproduction number exceeds both of its critical values ($R_0 > R_{0cw}$ in Eq. (6) and $R_0 > R_{0cr}$ in Eq. (12)), because we consider this scenario to be more relevant to fungal crop pathogens. Here, we describe in more detail two alternative scenarios: (i) when R_0 exceeds its critical value for the resistant strain, but not for the wildtype strain (i.e., $R_{0cr} \leq R_0 < R_{0cw}$, and (ii) when R_0 does not exceed its critical value either for the wildtype pathogen strain or for the resistant pathogen strain ($R_0 < R_{0cw}$ and $R_0 < R_{0cr}$).

In both scenarios (i) and (ii), C_R increases, reaches a maximum and then decreases reaching zero as a function of f (region of the plot corresponding to $1 < R_0 < 5$ in Fig. 6B), but this pattern arises due to somewhat different reasons in each of the two scenarios. In scenario (i), the maximum net return without resistance is reached at an intermediate, critical fungicide coverage $\theta^* = \theta_{cw} = (1 - 1/R_0)/\epsilon_w$ (Eq. (5)), while the maximum net return in the presence of resistance is reached when treating all fields $\theta^* = 1$. As a result, when the fungicide price is increased, the net return without resistance drops at a higher rate compared to the net return in the presence of resistance. This leads to a linear increase in the economic cost of resistance versus fungicide price f. But when the fungicide becomes more expensive, the optimal fungicide coverage switches to zero in the presence of resistance. As a result, the net return with resistance remains constant versus f, while the net return without resistance continues to decline as we increase f. In this range of intermediate fungicide prices, the economic cost of resistance declines linearly with f until it reaches zero. Then, the range of high fungicide prices is reached, where no fungicide treatment is economically viable ($\theta^* = 0$ both with and without resistance) and hence the economic cost of resistance remains at zero (lower right quadrant of Fig. 6B).

In scenario (ii), in the range of cheap fungicides, optimal fungicide coverage, θ^* has an intermediate value both in the absence and presence of resistance. The θ^*-value is higher with resistance than without resistance because a higher number of diseased fields due to resistance justifies a more intense fungicide treatment. For this reason, the net return with resistance declines faster versus the fungicide price than the net return without resistance. Consequently,

C_R increases linearly at low f-values. As we further increase f, C_R reaches a maximum, decreases linearly until it reaches zero and remains there as a function of f (for the same reasons as in the scenarios described above).