Variability in radiocesium activity concentration in growing hardwood shoots in Fukushima, Japan

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Abstract

The radiocesium contamination caused by the Fukushima Daiichi Nuclear Power Plant accident has made it difficult to use coppice woods as bed logs for mushroom cultivation. Evaluating the variability in the radiocesium activity concentration of logs is necessary in order to predict how many coppice woodlands are available for producing mushroom bed logs. To clarify the variability in radiocesium activity concentrations and to estimate the sample size required to estimate these concentrations with sufficient accuracy, we modeled the log-transformed radiocesium activity concentrations in growing shoots of hardwoods. We designed two models: (1) a model with mean concentrations that varied among stands with a standard deviation that was the same among stands, and (2) a model with varying means and standard deviations. We fit the data pertaining to only *Quercus serrata* to both models and calculated the widely applicable information criterion values. Consequently, we adopted the simpler model (1). Applying the selected model to data for all species, we examined the relationship between the number of measurement individuals and the predictive distribution of the expected concentration. Based on previous recommendations and measurement costs,

we proposed that five individuals would be appropriate for estimating radiocesium activity concentration in a stand.

Introduction

In Japan, coppice woodlands, consisting mainly of oaks, have long been managed for obtaining wood for fuel, charcoal, bed logs for mushroom cultivation, and for other purposes. Woodlands in Japan are a major component of what is referred to locally as "satoyama" landscape, which are areas that comprised of woodlands, grasslands, paddy fields, ponds, and related land use types [2,3]. These areas, which are characterized by periodic human disturbance, have relatively high levels of biodiversity [1]. However, although managed coppice woodlands can maintain biodiversity [4], large areas of them have been abandoned due to socio-economic changes [5].

More than any other areas in Japan, the traditional "satoyama" landscape is most prevalent in the southern Tohoku and northern Kanto regions of Japan. In Fukushima Prefecture, which is located in southern Tohoku, large area of coppice woodlands have been managed to produce the bed logs that are used for mushroom cultivation [6,7]. For this purpose, oak *Quercus serrata* and other hardwood tree species have been cultivated in the coppice woodland. The production of these bed logs previously benefited forestry activities and the regional economy of the area, as well as supporting the biodiversity of the region.

However, the Fukushima Daiichi Nuclear Power Plant accident, which occurred in March 2011, polluted the area with radioactive materials. Most of the fallout consisted of the radiocesium isotopes, ^{134}Cs and ^{137}Cs. Though, as a result of its longer 30-year half-life, only ^{137}Cs remains in the environment. This contamination has had an enormous impact on the communities and industries in the region, including the production of bed logs for mushroom cultivation in the oak coppices.

In order to meet the required safety standards limits for food, $100\,\mathrm{Bq\,kg^{-1}}$, bed logs used for mushroom cultivation must have a radioactivity concentration lower than $50\,\mathrm{Bq\,kg^{-1}}$ [8]. This is because the maximum transfer factor from the bed log to cultivated mushrooms has been estimated as being approximately two [7]. Resuming bed log production is considered to be necessary for the forestry and regional economy

and would also contribute towards maintaining the biodiversity of the region. 29

Recently, studies have been conducted to facilitate the resumption of coppice 30
management. For example, Kanasashi et al. [9] established a strong and negative 31
correlation between the exchangeable potassium present in soil and the radiocesium 32
(^{137}Cs) activity concentrations in the growing shoots of *Q. serrata*. The observation of a 33
constant relationship between radiocesium activity concentration of growing shoots 34
(new branches) and that in woody stem means that the concentration in growing shoots 35
are well suited for use as a proxy for the concentration in woody stems [10,11]. Based 36
on this assumption, coppices situated on the potassium-rich soils may be well suited for 37
the production of bed logs with low radiocesium contamination even in areas that were 38
experienced radioactive fallout. 39

The application of potassium fertilizer to reduce the absorption of radiocesium is a 40
widely used measure in agricultural fields [12]. In woodlands, which are considerably 41
larger than agricultural lands, searching for coppices with low radiocesium activity 42
concentrations could also be a promising measure for sourcing bed logs with 43
radiocesium activity concentrations that are within permissible limits. Indeed, for 44
mushroom cultivation on sawdust media, the relationship between mushroom fruiting 45
body and medium radiocesium activity concentrations have been clarified [13,14]. To 46
accurately predict the proportion of available logs, estimating not only the mean, but 47
also the variances in the radiocesium activity concentrations present in the coppice 48
stems is crucial. It has been shown that a definitive range exists in the radiocesium 49
activity concentrations of growing shoots within and among stands of coppices and 50
plant species [9]. However, this variance is seldom known a priori [15,16]. Moreover, if 51
the variance in the radiocesium activity concentrations varies among stands, then it will 52
be difficult to accurately predict the concentrations from small sample sizes. It is 53
therefore necessary to clarify the distribution of the variance in the radiocesium activity 54
concentration within and among coppice stands. In addition, the variance would be 55
important to determine the required sample size to predict the concentration. 56

Two questions arise regarding the measurement of radiocesium activity concentration 57
in growing shoots: (1) Can the variances of the measurement values of the radiocesium 58
activity concentration in growing shoots be regarded as constant or variable among 59
different stands? (2) What number of individual samples are required to measure the 60

stand-level distribution of radiocesium activity concentration? In this study, we ⁠61
evaluated the stand-level variation in radiocesium activity concentrations in the growing ⁠62
shoots of coppice species, and predicted the concentration distribution within a stand. ⁠63
The findings will enable the estimation of the mean radiocesium activity concentration ⁠64
in growing shoots within a specified area, thereby ensuring that the concentration ⁠65
distribution is under a predetermined threshold. ⁠66

Materials and Methods ⁠67

Data collection ⁠68

We established 40 study stands in Miyakoji, Tamura City, Fukushima Prefecture, Japan, ⁠69
and sampled growing shoots of hardwood species in each stand between December 2016 ⁠70
and March 2017. The radiocesium activity concentration is known to be stable during ⁠71
the dormant stage (November to April) [17]. In this study, we assumed that the stands ⁠72
comprised areas of several tens of square meters in size. Permission was obtained from ⁠73
the Fukushima Chuo Forestry Co-operative to collect samples from the study area. The ⁠74
Q. serrata data used in this analysis were essentially the same as those in Kanasashi et ⁠75
al. [9], except that some measurements were updated. The data file is provided in the ⁠76
supporting information (S1 Table). In addition, other hardwood species that were ⁠77
present in the stands with *Q. serrata*, as well as regenerated coppice sprouts and new ⁠78
plantation, were also included in the analysis. A total of 453 individuals of hardwood ⁠79
tree species were collected. Of these, 418 individuals of five species had radiocesium ⁠80
activity concentrations that could be detected by the method described below were ⁠81
included in the analysis (i.e., 35 individuals (7.7%) were excluded) (Table 1). ⁠82

Table 1. Number of individuals of each species that were sampled

Species	No.
Quercus serrata Murray	184
Cerasus jamasakura (Siebold ex Koidz.) H.Ohba	111
Castanea crenata Siebold et Zucc.	86
Quercus acutissima Carruth.	30
Zelkova serrata (Thunb.) Makino	7

The collected samples were processed as described in a previous study [9], and the ⁠83
radioactivity concentrations were measured using a NaI(Tl) scintillation detector (2480 ⁠84

WIZARD^2; PerkinElmer Japan Co., Ltd., Yokohama, Japan) or high-purity germanium detectors (GEM40P4-76 and GEM-FX7025P4-ST; ORTEC, Oak Ridge, TN, USA). All of the detectors were calibrated using the radioactivity standard gamma volume sources, MX035 and MX033U8PP (Japan Radioisotope Association, Tokyo, Japan). For the NaI(Tl) scintillation detector measurements, samples were placed into 20-mL vials and measured from 2,000 to 172,800 s, after which the gamma-ray spectra were analyzed using the Code Fukushima software package [18]. The detection limit is not constant and depends on the radioactivity concentration and measurement time (S1 Table). The 662 keV peaks of Cs-137 were considered to be a distinct peak when the full width at half maximum (FWHM) was greater than 30 keV and the relative counting error was less than 33.3%; peaks meeting these criteria were used for the analysis. For the germanium semiconductor detector, measurements were performed from 1,800 to 86,400 s, and data with relative counting errors of 33.3% or less were used for analysis. The activity concentrations were decay-corrected to the values on December 1, 2016.

Statistical modeling

We assumed that the natural logarithm (hereafter written as logarithm) of the radiocesium activity concentration Y_i (Bq kg^{-1}) in the growing shoots followed a normal distribution [15, 19], namely

$$\log_{\mathrm{e}}(Y_i) \sim \mathrm{Normal}(\mu_{s[i]}, \sigma^2_{s[i]}), \tag{1}$$

where $\mu_{s[i]}$ denotes the mean, and $\sigma_{s[i]}$ denotes the standard deviation for stand $s[i]$ containing the i-th individual. We excluded censored data from the model analysis, as the detection limits of such data were indefinite and varied depending on the measurement time.

We further assumed that μ_s, the mean for stand s, was composed of the overall mean $\bar{\mu}$ and a random effect for the stand $\epsilon_{\mathrm{M}[s]}$,

$$\mu_s = \bar{\mu} + \epsilon_{\mathrm{M}[s]}$$

$$\epsilon_{\mathrm{M}} \sim \mathrm{Normal}(0, \sigma_{\mathrm{M}}^2).$$

where σ_{M} denotes the standard deviation of ϵ_{M}. 109

We constructed two models predicated on distinct methodologies for formulating the 110
standard deviation, denoted as σ_s, as outlined in Eq 1 (Table 2). 111

1. σ_s is constant for all stands, namely $\sigma_s = \bar{\sigma}$. 112

2. σ_s can vary due to a random stand effect, namely $\sigma_s = \bar{\sigma} \exp(\epsilon_{\mathrm{V}[s]})$. 113

The random stand effect ϵ_{V} was assumed to follow a normal distribution, and can be 114
represented as follows: 115

$$\epsilon_{\mathrm{V}} \sim \mathrm{Normal}(0, \sigma_{\mathrm{V}}^2).$$

Table 2. Models to be compared

Model	Structure
Model 1	$\log_{\mathrm{e}}(Y_i) \sim \mathrm{Normal}(\mu_{s[i]}, \sigma_{s[i]}^2), \mu_s \sim \mathrm{Normal}(\bar{\mu}, \sigma_{\mathrm{M}}^2),$ $\sigma_s = \bar{\sigma}$
Model 2	$\log_{\mathrm{e}}(Y_i) \sim \mathrm{Normal}(\mu_{s[i]}, \sigma_{s[i]}^2), \mu_s \sim \mathrm{Normal}(\bar{\mu}, \sigma_{\mathrm{M}}^2),$ $\sigma_s = \bar{\sigma} \exp(\epsilon_{\mathrm{V}[s]}), \epsilon_{\mathrm{V}} \sim \mathrm{Normal}(0, \sigma_{\mathrm{V}}^2)$

Y_i: measured value of the radiocesium activity concentration in the growing shoots (Bq kg^{-1}) of the i-th individual, $\mu_{s[i]}$: mean of $\log_{\mathrm{e}}(Y_i)$ in the stand $s[i]$ containing the i-th individual, $\sigma_{s[i]}$: standard deviation of $\log_{\mathrm{e}}(Y_i)$ in the stand $s[i]$ (fixed in model 1), $\bar{\mu}$: overall mean of $\log_{\mathrm{e}}(Y_i)$ in the study area, $\bar{\sigma}$: overall standard deviation, ϵ_{V}: random stand effect on σ_s, σ_{M}: standard deviation of the random effect ϵ_{M}, σ_{V}: standard deviation of the random effect ϵ_{V}.

The prior distribution of the overall mean was defined as a vague prior as follows: 116

$$\bar{\mu} \sim \mathrm{Normal}(0, 100^2).$$

The priors of the standard deviations ($\bar{\sigma}$, σ_{M}, and σ_{S}) were defined as weakly 117

informative priors [20, 21], as follows:

$$\sigma \sim \mathrm{HalfNormal}(10^2).$$

The parameters in the models were estimated using a Markov chain Monte Carlo (MCMC) approach. The MCMC calculation was conducted using Stan 2.32.2 [22] via the CmdStanR 0.5.3 interface [23] for R [24]. We obtained the MCMC samples from four independent Markov chains, each consisting of 8,000 MCMC iterations after a warmup stage comprising 8,000 iterations. We then evaluated the convergence using the value of $\hat{R}$, or the Gelman-Rubin statistics [25, 26] and a visual census of the degree of chain mixture. The Stan model code is included in the supporting information (S1 File, S2 File).

In addition, we also calculated the widely applicable information criterion (WAIC) [27] values for each model. The WAIC values were calculated under the stipulation of forecasting the radiocesium activity concentration in growing shoots in a new stand. Analogous to the Akaike information criterion (AIC), models characterized by smaller WAIC values are anticipated to exhibit superior mean predictive performance in the context of of model selection. Moreover, WAIC can be applied to complex hierarchical Bayesian models.

We fitted both models to the data containing only *Q. serrata* and selected the optimal model after consideration of WAIC values and model simplicity.

Model for analyzing all data

We then constructed a model to analyze the dataset comprising all five species that was based on the previously selected optimal model. In this model, we incorporated the origin of each individual (i.e., coppice sprout or plantation) as an explanatory variable, as well as an additional random species effect. The model selection affected whether the standard deviation, $\sigma_{s[i]}$ was uniform or varied across different stands.

$$\log_e Y_i \sim \text{Normal}(\mu_{s[i]}, \sigma_{s[i]}^2) \tag{2}$$

$$\mu_s = \bar{\mu} + \beta O + \epsilon_{\text{M}} + \epsilon_{\text{S}} \tag{3}$$

$$\epsilon_{\text{M}} \sim \text{Normal}(0, \sigma_{\text{M}}^2) \tag{4}$$

$$\epsilon_{\text{S}} \sim \text{Normal}(0, \sigma_{\text{S}}^2). \tag{5}$$

where the variable O denotes the origin of the individual (0: sprout, 1: plantation), and β is the coefficient of O. ϵ_{M} is a random stand effect, and σ_{M} is the standard deviation thereof. Similarly, ϵ_{S} denotes a random species effect, and σ_{S} is its standard deviation. Although *Z. serrata* had a small sample size, incorporating the random species effect facilitates estimates of individual species by "borrowing strength" [28].

For the parameter estimation, the model was fitted to the whole dataset (S1 Table), i.e. the data for all five species (40 stands and 418 individuals), using Stan 2.32.2. The code for the Stan model is given in the supporting information S3 File. Parameters used in the models are summarized in Table 3.

Table 3. List of parameters

Name	Description
$\bar{\mu}$	Overall mean of log-transformed radiocesium activity concentration
μ_s	Mean of log-transformed radiocesium activity concentration at stand s
β	Coefficient of the trunk origin (0: sprout, 1: plantation)
$\epsilon_{\text{M}[s]}$	Random stand effect on the mean of log-transformed radiocesium activity concentration at stand s
$\epsilon_{\text{V}[s]}$	Random stand effect on the standard deviation of log-transformed radiocesium activity concentration at stand s
ϵ_{S}	Random species effect on the mean of log-transformed radiocesium activity concentration
$\bar{\sigma}$	Overall standard deviation of log-transformed radiocesium activity concentration
σ_s	Standard deviation of log-transformed radiocesium activity concentration at stand s
σ_{M}	Standard deviation of the random stand effect ϵ_{M}
σ_{V}	Standard deviation of the random effect ϵ_{V}
σ_{S}	Standard deviation of the random species effect ϵ_{S}

Prediction interval of the measurements of radiocesium activity concentration for a new stand

Using the selected model with all data, we conducted a simple simulation to predict the radiocesium activity concentration in growing shoots in a new stand. In this simulation, we calculated the relationship between the sample size and the predicted value of the radiocesium activity concentration of growing shoots in the study area.

In addition, we estimated the relationship between the sample size and the standard deviation. Suppose a situation in which the radiocesium activity concentration is measured in n individuals from a new stand.

Let Y_{new} be the radiocesium activity concentration of the growing shoots in the new stand, and $y_{\text{new}} = \log_{\text{e}}(Y_{\text{new}})$. Namely,

$$y_{\text{new}} \sim \text{Normal}(\mu_s, \sigma_s^2). \tag{6}$$

The mean concentration μ_s follows a normal distribution with an overall mean of $\bar{\mu}$ and a variance of σ_{M}^2,

$$\mu_s \sim \text{Normal}(\bar{\mu}, \sigma_{\text{M}}^2). \tag{7}$$

As described above (Table 2), the standard deviation σ_s depends on the model. If model 1 is selected, then σ_s is a constant:

$$\sigma_s = \bar{\sigma}$$

Otherwise, if model 2 is selected, then σ_s varies between stands:

$$\sigma_s = \bar{\sigma} \exp(\epsilon_{\text{V}}[s]).$$

From Eqs 6 and 7,

$$y_{\text{new}} \sim \text{Normal}(\bar{\mu}, \sigma_s^2 + \sigma_{\text{M}}^2).$$

Given the measurements $Y_1, Y_2, \ldots, Y_n$ (Bq kg^{-1}), we denote their logarithm values as $y_i = \log_{\text{e}}(Y_i)$, and y_{n+1} should be distributed as follows [29]:

$$y_{n+1} \sim \text{Normal}(\mu_n, \sigma_s^2 + \sigma_n^2) \tag{8}$$

where the expected value of μ_n after n individuals were measured is given by

$$\mu_n = \frac{\frac{1}{\sigma_{\text{M}}^2}\bar{\mu} + \frac{1}{\sigma_{\text{s}}^2}\sum_{i=1}^{n} y_i}{\frac{1}{\sigma_{\text{M}}^2} + \frac{n}{\sigma_{\text{s}}^2}},$$

and the expected value of $\bar{\mu}_n$ is

$$\bar{\mu}_n = \frac{\frac{1}{\sigma_{\text{M}}^2}\bar{\mu} + \frac{n}{\sigma_{\text{s}}^2}\mu_s}{\frac{1}{\sigma_{\text{M}}^2} + \frac{n}{\sigma_{\text{s}}^2}}, \tag{9}$$

In addition, σ_n can be calculated as follows [29],

$$\frac{1}{\sigma_n^2} = \frac{1}{\sigma_{\text{M}}^2} + \frac{n}{\sigma_s^2}.$$

We calculated the mean μ_n for n = 1, ..., 10, under the assumption that the true value of the stand radiocesium activity concentration is 10 Bq kg^{-1}. We also calculated the standard deviation in Eq 8, $\sqrt{\sigma_s^2 + \sigma_n^2}$.

Using the above relationships, we inferred how many individuals are required to adequately evaluate the expected value and error of the radiocesium activity concentration in a new stand.

Results 179

Radiocesium activity concentration measurements 180

The radiocesium activity concentration in the growing shoots ranged from 4.91 to $\qquad$ 181

$1.13 \times 10^4 \, \text{Bq} \, \text{kg}^{-1}$, and the mean and median were 4.65×10^2 and $2.02 \times 10^2 \, \text{Bq} \, \text{kg}^{-1}$, $\quad$ 182

respectively. The stand-level mean of the log-transformed values of the radiocesium $\quad$ 183

activity concentrations ranged from 1.73 to 8.78, and their mean and median was 5.08 $\quad$ 184

and 5.29, respectively (Fig 1.) The standard deviation in the stands ranged from 0.33 to $\quad$ 185

1.09 and the mean was 0.70. It should be noted that 7.7% of the total individuals $\quad$ 186

surveyed were excluded from the analysis for the average value, as their samples could $\quad$ 187

not be quantified due to low radioactivity levels; for the standard deviation, the value $\quad$ 188

might therefore be larger if the excluded samples were included. $\quad$ 189

Fig 1. Stand-level mean (left) and standard deviation (right) of log-transformed values of radiocesium activity concentration measurements in the growing shoots obtained for each species. Values larger than the third quartile + 1.5 × interquartile range (IQR) or smaller than the first quartile - 1.5 × IQR were treated as outliers. *Zelkova serrata* is not shown because the species occurred in only two stands.

Model selection 190

The results of the MCMC calculations for *Quercus serrata* showed that $\hat{R}$ values were $\quad$ 191

smaller than 1.1 for all of the parameters, and that the Markov chains appeared to be $\quad$ 192

well mixed. We therefore considered that the Markov chains converged successfully to $\quad$ 193

stable distributions. $\quad$ 194

Table 4 shows the result of the model selection. Since there was only a small $\quad$ 195

difference in the WAIC values between the two models (499 and 498), and because the $\quad$ 196

values might vary slightly due to the sampling error, we adopted the simpler model, $\quad$ 197

Model 1. Consequently, the standard deviation could be perceived as uniform across the $\quad$ 198

stands, thereby dismissing the variations intrinsic to individual stands in predicting the $\quad$ 199

radiocesium activity concentrations. $\quad$ 200

Table 4. Comparison of models (posterior means of the parameters and WAIC)

Model	$\bar{\mu}$	$\bar{\sigma}$	σ_{M}	σ_{V}	WAIC
Model 1	5.62	0.74	1.09		499
Model 2	5.64	0.70	1.08	0.26	498

Parameter estimates for the whole dataset

Since the results of the model selection allowed us to treat the standard deviation σ_s as being uniform among the stands, we let $\sigma_s = \bar{\sigma}$ and estimated the parameters of the model (Eq 2–5) using MCMC.

Table 5 shows a summary of the posterior distributions for the parameters. The posterior mean of $\bar{\sigma}$, which is the standard deviation of radiocesium activity concentration in the growing shoots, was 0.74 on a logarithmic scale, and the standard deviation was 0.03. The value of 0.74 on the logarithmic scale translates to a geometric standard deviation of 2.09. Using this value, we can expect that approximately 95% of the measurement values are within an 18-fold $(\exp(2 \times 1.96 \times 0.74) = 18)$ on the usual linear scale, excluding the effects of the species and origin. The posterior mean of the coefficient β, which represents the origin of the individuals, was estimated as -0.39, with the entire 95% confidence interval exhibiting negative values. The negative value indicates that the individuals from plantations would have a lower radiocesium activity concentration than the individuals produced from sprouts, and that the log-transformed mean concentration would decrease from 4.92 to 4.53. The posterior mean of σ_{M}, which is the standard deviation of the random stand effect on the log-transformed mean concentration, was 1.33, and larger than that of σ_{S}, 0.62, which is the standard deviation of the random species effect on the log-transformed mean concentration.

Table 5. Summary of the posterior distribution of parameters for all species in the model. SD: standard deviation.

Parameter	Mean	SD	2.5% quantile	97.5% quantile
$\bar{\mu}$	4.92	0.41	4.06	5.65
β	-0.39	0.11	-0.61	-0.17
$\bar{\sigma}$	0.74	0.03	0.69	0.79
σ_{M}	1.32	0.17	1.04	1.71
σ_{S}	0.62	0.44	0.19	1.76

Fig 2 shows the posterior distributions of the random species effect. *Q. serrata* had a

relatively high value while that of *Zelkova serrata* was lower, corresponding the ₂₂₁

log-transformed mean (Fig. 1). Values in the figure are shown using a logarithmic scale; ₂₂₂

consequently, -1 on this scale is equivalent to a 0.37-fold decrease compared to the ₂₂₃

overall mean radiocesium activity concentration of the growing shoots on a linear scale. ₂₂₄

In the same manner, -0.5 on this scale is equivalent to a 0.61-fold decrease, while 0.25 ₂₂₅

and 0.5 are equivalent to increase of 1.6- and 2.7-fold, respectively. ₂₂₆

Fig 2. Posterior distribution of the random species effect (on a logarithmic scale). Dots denote medians, thick lines denote 90% credible intervals, and thin lines denote 95% credible intervals.

Fig 3 shows the relationship between the number of measurement individuals (n) ₂₂₇

and the expected value of the log-transformed radiocesium activity concentration ₂₂₈

($\log_e(\mu_n)$), assuming a true mean of $10\,\mathrm{Bq\,kg^{-1}}$. This figure omits the random species ₂₂₉

effect and assumes the origin of the stems to be sprouts. It shows that the expected ₂₃₀

value progressively approximates the true value as the number of measurements ₂₃₁

increases. Further, this relationship holds for other true values. ₂₃₂

Fig 3. Relationship between the number of measurement individuals and the expected value of the log-transformed radiocesium activity concentration when the true mean is $10\,\mathrm{Bq\,kg^{-1}}$ ($\log_e 10 = 2.30$). The curve denotes the mean, and the dark region denotes the 95% prediction interval of the expected value.

As a result, the standard deviation of the predicted radiocesium activity ₂₃₃

concentration on a logarithmic scale within a stand is given as ₂₃₄

$$\sqrt{\sigma_s^2 + \sigma_n^2} = \sqrt{\sigma_s^2 + \frac{1}{\frac{1}{\sigma_M^2} + \frac{n}{\sigma_s^2}}}$$
$$= \sqrt{\bar{\sigma}^2 + \frac{1}{\frac{1}{\sigma_M^2} + \frac{n}{\bar{\sigma}^2}}}. \tag{10}$$

Discussion ₂₃₅

In this study, we employed a Bayesian approach using MCMC methods to estimate ₂₃₆

parameters in the mixed-effect models. Using WAIC values facilitated the selection of a ₂₃₇

model for predicting radioactivity concentrations in the developing shoots of coppice ₂₃₈

hardwoods. In addition, the Bayesian approach facilitated straightforward management ₂₃₉

of uncertainties [29].

Variation in radioactivity concentrations within a stand

Within the context of food safety, the public has a strong interest in radiocesium contamination of agricultural and forest products [7]. Local government bodies measure the radioactivity concentrations of foods and items utilized for food production, such as mushroom bed logs in the forestry sector, to prevent the distribution of foods and items that have radioactivity concentrations that exceed standard limits. Consequently, understanding the frequency distribution characteristics of the radioactivity concentration in forest products obtained under certain conditions is very important for the effective management and monitoring of these resources. In this study, we investigated the variability in radiocesium activity concentrations in growing shoots of hardwood species, which has been utilized to establish indices of radiocesium absorption [10, 11], with a particular focus on the hardwood bed logs that are commonly used in mushroom cultivation. The radiocesium activity concentration in the forest products has been demonstrated to follow a log-normal distribution [15, 19]. After the two largest nuclear power plant accidents to date, Chornobyl and Fukushima, radioactivity concentrations in many agricultural and forest products have been measured, and transfer factors (TF) and aggregated transfer factors (Tag) have been reported [30–32]. However, few studies have examined the variability — characterized by variance or standard deviation — in radioactivity concentrations of vegetation samples after log-transformation to attain normality. Khomutinin et al. [19, 33] analyzed large systematic grid sampling data, with variations in the dimensions of sampling units and increments (distances between the grid nodes). In their research on vegetation, they found that sampling independence in the radiocesium activity concentration could be achieved by sampling over distance of 8–10 m in stands with uniform soil contamination [15]. Their findings were reflected in the sampling guidelines recommended and published by the International Atomic Energy Agency (IAEA) [16]. Khomutinin et al. [19] also reported that the log-transformed standard deviation of radioactivity concentration by ^{137}Cs sampled from 13 sites in Ukraine ranged from 0.22 to 0.62, and the mean and standard deviation was 0.38±0.11 . Their sampling sites covered a variety of different land use types including cultivated lands.

Compared to the study in Ukraine, in this study, we estimated the standard ²⁷¹
deviation within 40 individual survey stands, focusing on the growing shoots (N=5) of ²⁷²
five hardwood species of forests in Fukushima. The range in the radiocesium activity ²⁷³
concentrations showed a marked discrepancy between soil and growing shoots; while the ²⁷⁴
range in the soils was relatively stable with a 10-fold difference, that in the growing ²⁷⁵
shoots showed fluctuations that ranged up to 1000-fold [9]. Based on the WAIC values ²⁷⁶
(Table 4), we could select the simpler model with a constant standard deviation, instead ²⁷⁷
of the model with a varying standard deviation among stands. Thus, we could treat the ²⁷⁸
variability of radiocesium contamination of forest products as uniform irrespective of ²⁷⁹
stands, as shown in crops and grasses [19]. The standard deviation of log-transformed ²⁸⁰
concentration within each stand ($\bar{\sigma}$) was estimated to be 0.74 as the posterior mean, ²⁸¹
and its standard deviation was estimated to be 0.03 (Table 5). ²⁸²

However, the within-stand standard deviation observed in this study and the study ²⁸³
in Ukraine (0.38) [19] showed a two-fold difference. In Japan, woodlands are typically ²⁸⁴
situated on steep undulating slopes, and the topographical characteristics of an area can ²⁸⁵
change markedly within a relatively small area. We consider that these variations in the ²⁸⁶
terrain may affect the soil properties. Furthermore, in terms of sampling methods, soil ²⁸⁷
type, land usage, plant species, etc., numerous differences exist in the dataset used in ²⁸⁸
the present study. Consequently, the factor that caused the two-fold difference between ²⁸⁹
the present study and the preceding study in Ukraine was not specified. The standard ²⁹⁰
deviations in the radioactivity concentration in vegetation samples are fundamental ²⁹¹
data that are essential for establishing radiation protection standards. Further studies ²⁹²
are therefore required to clarify the principal factors that influence representative values ²⁹³
and to comprehensively understand the prevalence of associated uncertainties. ²⁹⁴

In this study in Fukushima, we used sampling areas spanning several tens of square ²⁹⁵
meters. To achieve statistical independence in sampling, Barnekow et al. [16] proposed ²⁹⁶
that the center-to-center distance of sampling points for vegetation samples should not ²⁹⁷
be less than 8–10 m. We consider that our sampling protocol satisfied this criterion in ²⁹⁸
most cases. In contrast, in the survey of Khomutinin et al. [19], sampling grids ranged ²⁹⁹
from 1 to 10 m in grid size, with 7 of the 13 surveyed sites having a grid size of 2.5 m or ³⁰⁰
less. We considered the possibility that differences in sampling distance (increments) ³⁰¹
may have affected the differences observed in standard deviations, but the standard ³⁰²

deviation in Khomutinin et al. [19] did not appear to be affected by the difference in grid size. All of the study sites in Khomutinin et al. [19] were located in crop fields and former agricultural fields and did not include forests; this is a notable difference between the datasets used in their study and this study. It is possible that forests show a greater variability of radiocesium activity concentrations in vegetation samples. For example, our Fukushima survey was conducted in mountainous areas with slopes of 6 to 30 degrees (11 to 73% gradient) [9], wheres all. 13 study sites in their Ukraine survey were located on flat terrain (confirmed on Google Maps from latitude and longitude). These differences may have accounted for the differences observed in the variability of radiocesium activity concentrations between these studies.

Factors affecting the radioactivity concentration

The species random effect was larger in *Q. serrata*, which is most commonly used for bed logs; the posterior mean was 0.44. The random effect for *Q. acutissima*, another major hardwood species that is used for mushroom logs, was lower than that for *Q. serrata* (Fig 2); the posterior mean was -0.14. The difference in the posterior mean of the random effect was 0.57, which corresponds to a 1.8-fold difference when converted to a linear scale. The lowest value was observed in *Z. serrata*; the posterior mean was -0.55 or 0.37-fold of that of *Q. serrata* when converted to a linear scale. The difference between the maximum and minimum mean values for the five tree species was 2.7-fold. Ipatyev et al. [34] reported a minimum 10-fold difference in Tag values ($m^2\,kg^{-1}$) for needles and leaves of five tree species in Belarus, with a range from 1.6×10^{-3} for birch to 3.8×10^{-2} for pine. However, Calmon et al. [35] examined the Tag of trees after the Chornobyl nuclear accident and found that the influence of tree species exerted only a minor influence on these values, affecting them by a factor of 2 to 3. The present study also suggested that the interspecific difference in radiocesium absorption from the soil was small.

Furthermore, the data suggest a disparity in radiocesium contamination levels contingent on the origin of the trunk, with stems that originated from sprouts exhibiting a greater degree of contamination by radiocesium than stems that originated from plantations. In this study, the ages of the measured stems ranged from one year to a maximum of seven years. Therefore, the differences in radiocesium absorption among

species and the origin of the stems may differ from those of stems aged around 20 years old, which are best suited for use as bed logs. Long-term monitoring is therefore considered necessary in order to asses whether the observed relationships remain consistent into the future.

The stand-level mean radiocesium concentrations observed in this study ranged from 1.7–8.8 in log-transformed values (Fig. 1). However, since 7.7% of the total number of individuals were excluded due to their low radioactivity levels, it is plausible the true distribution lies more towards the left tail.

Kanasashi et al. [9] pointed out that the amount of exchangeable potassium in the surface soil largely influenced the radiocesium concentrations in the growing shoots. In comparison, although species and origin differences also affect radiocesium concentrations in the growing shoots, their impact is markedly diminished compaired to the growing environment. These findings are consistent with a review [35], where post-Chornobyl nuclear accident analysis of Tag values for trees showed a pronounced influence of soil type, quantified as a level of 100, whereas the effect of tree species was 2. Given this large variability in the mean values, it is important to investigate the radiocesium activity concentrations across stand to determine if the forest owners can resume the bed log production for mushrooms.

Required sample size for appropriate estimation of radiocesium activity concentration

The objective of this study, evaluating the radiocesium activity concentrations of growing shoots in the coppices, is to protect communities from radiation exposure. This is particularly pertinent given the regulations that have been imposed on forest products used in food production, such as mushroom bed logs [7]. Assuming a normal distribution for the log-transformed radioactivity concentration estimates, the two parameters, i.e., the mean and standard deviation, are required to facilitate the accurate estimation of the proportion of the bed logs produced in a stand, which would adhere to the stipulated radioactivity regulation thresholds. Importantly, these parameter estimates have uncertainties. In this study, we showed that the within-stand standard deviation of the log-transformed radioactivity concentration of growing shoots of coppice woods could be regarded as uniform among stands and that the mean value was 0.74

(Table 5). The sample mean value of the log-transformed radioactivity concentration of growing shoots has been reported to vary markedly among stands [9]. Fig. 3 shows the relationship between the number of measurement individuals (n) in a stand and the expected value of the log-transformed radiocesium activity concentration for the case where true mean is $10\,\mathrm{Bq\,kg^{-1}}$ (2.30 in the log-transformed value). When the number of measurements is limited, the expected value approaches toward the overall mean of 4.92; however, with an increase in the data points, the curve progressively approximates the true value, and the slope becomes gentler. Thus, the expected value would be 2.92 in the case where $n = 1$, 2.46 in the case where $n = 5$, and 2.38 in the case where $n = 10$ (Eq 9, Fig. 3). Therefore, as the number of measurement individuals increases, the standard deviation of the predicted radiocesium activity concentration (log-transformed) within a stand diminishes; the standard deviation would be 0.98 in the case where $n = 1$, 0.80 in the case where $n = 5$, and 0.77 in the case where $n = 10$ (Eq 10). While the rate of decrease in the values slows down, the measurement cost increases in proportion to the number of individuals measured. Considering the tendency of decreasing standard deviation and the delicate balance between precision and cost, we propose that sampling five individuals is sufficient to estimate the radiocesium activity concentration in a coppice stand. This aligns with the recent guidelines issued by the International Atomic Energy Agency (IAEA) for soil and vegetation sampling [36]. However, it should be noted that the standard deviation of radioactivity concentrations in trees in Fukushima is approximately twice as large as observed in the previous study in Ukraine [19].

Conclusion

In predicting radiocesium activity concentration in the growing shoots of hardwood species in forests exposed to radioactive fallout, the variance of the log-transformed measurement values within a stand can be assumed to be invariant irrespective of the stands. In this study, we found the mean standard deviation to be 0.74, a value that remained constant despite variations in the mean value of radiocesium activity concentration across stands and species. Furthermore, the findings showed that measurements from approximately five individuals would be sufficient for estimating the mean radiocesium activity concentration in a stand. The variance and mean of

vegetation samples are fundamental parameters used in radiation protection measures, 395
such as setting standard limits for radiation protection and determination surveys based 396
on reference values. Numerous vegetation surveys have been conducted in forests, 397
agricultural lands, and grasslands after the Fukushima and Chornobyl nuclear accidents. 398
The statistical description and understanding of radioactive contamination in terrestrial 399
plant ecosystems through integrated analysis of these data is an important topic that 400
required further work. 401

Supporting information

S1 Table. Data file of the radioactivity concentration in the growing shoots of five hardwood species in Miyakoji. The significant digits of the measurements are three.

S1 File. Stan model code file for model 1.

S2 File. Stan model code file for model 2.

S3 File. Stan model code file for the model with all species.

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