Variational Bayesian Phylogenies through Matrix Representation of Tree Space

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ABSTRACT

Variational Bayesian methods work well when the distribution of interest can be represented by a Euclidian space using a bijective transformation between phylogenies and their representation. Until now, full phylogenetic inference has been hampered by the lack of such transformation. Here, we study the distance matrix as a representation of a phylogeny through single link clustering. Though such representation does not map to unique trees, restriction to a subspace we call a "cube", makes the representation bijective at the cost of not being able to represent all possible trees. We introduce a variational Bayesian algorithm "cubeVB" based on MCMC and show through well calibrated simulation study that it is possible to recover parameters of interest like tree height and length. Although a cube cannot represent all of tree space, it is a great improvement over a single summary tree, and it opens up exciting new opportunities for scaling up Bayesian phylogenetic inference. An open source implementation of the cubeVB algorithm is available from `https://github.com/rbouckaert/cubevb` as the `cubevb` package to BEAST 2.

1 INTRODUCTION

The workhorse of Bayesian phylogenetics is the Markov chain Monte Carlo (MCMC) algorithm, which is widely used to reconstruct the evolutionary relationships among a set of biological sequences (Huelsenbeck and Ronquist, 2001; Höhna et al., 2016; Suchard et al., 2018; Bouckaert et al., 2019). In practice, MCMC is limited to handling phylogenies of at most a few thousand taxa. With the advent of ever cheaper sequencing techniques, these algorithms are overwhelmed by the amount of available data. Consequently, maximum likelihood methods are employed, which can handle these amounts of data because they are based on optimisation algorithms.

Variation Bayesian (VB) methods (Jordan et al., 1999) are also based on optimisation, so offer the prospect of being more scalable than MCMC. VB methods represent a posterior distribution as a transformation of a multivariate normal distribution on Euclidian space. The form of the transformation is crucial in how accurate the posterior distribution can be represented. Usually, stochastic gradient ascent is used as the optimisation algorithm, which relies on the transformation being a bijection between trees and their transformed representation. This provides a major hurdle for phylogentic inference, since existing tree spaces are either not bijective (like spanning trees between points in hyperbolic space (Matsumoto et al., 2021; Jiang et al., 2022)) or do not allow representation as a Euclidian space (like ultrametric

space (Gavryushkin and Drummond, 2016)).

Prior phylogenetic VB implementations got around these restrictions by being limited to fixed tree topologies (Fourment and Darling, 2019; Zhang, 2020). Others do not handle timed trees (Zhang and Matsen IV, 2018) which are important to many practical and scientific questions.

Here, we explore matrix representations of tree space. The basic idea is to use a symmetric matrix with dimension equal to the number of taxa and apply a hierarchical clustering algorithm like the single link clustering algorithm (Florek et al., 1951) to obtain a tree with internal node heights specified by the values of the matrix. The entries in the matrix form a Euclidian space, however there are many ways to represent the same tree, so the transformation is not bijective. By restricting ourselves to the 1-off-diagonal entries of the matrix (and leave the rest at infinity) the transformation becomes a bijection, but cannot represent all possible trees any more. In this paper, we demonstrate it is possible to capture the most interesting part of posterior tree space using this approach. We implement an algorithm for basic phylogenetic models, infer the parameters of the multivariate normal distribution and show through a well calibrated simulation study that it captures important parameters of interest. We demonstrate that it scales well.

Many post-hoc analyses such as ancestral reconstruction, trait inference or phylogeography are performed on a single summary tree, so these analyses do not take phylogenetic uncertainty in account. Though a cube does not represent all of tree space, in many situations it captures most of the uncertainty of a posterior distribution, and allows taking uncertainty in account in post-hoc analyses. Furthermore, our novel way to represent tree space opens up new avenues to implement online algorithms (Bouckaert et al., 2022), allowing rapid responses to newly available data.

2 METHODS

2.1 Representing trees through a matrix

The main idea is to represent a posterior tree distribution as a distribution over entries of a matrix, and apply the single links clustering algorithm (Algorithm 1) to obtain a tree. The matrix can be interpreted as a distance matrix: it must be symmetric, have positive values and the entries on the diagonal are ignored. So, for a tree with n taxa there are $n(n-1)/2$ entries that are relevant: the entries above the diagonal. We say a matrix entry is *specified* if it is a finite, and *unspecified* if it is positive infinite. We say a matrix is *valid* if the tree associated with it through single link clustering only has finite branch lengths.

Algorithm 1 Single link clustering

Input: $n \times n$ matrix M with entries m_{ij}
Let $x_1, \ldots, x_n$ be n taxa
$\forall_{1 \leq i \leq n}$ initialise subtree as single taxon $\{x_i\}$ at height 0
while There is more then 1 subtree **do**
    Let m_{ij} be minimum of M, where x_i and x_j in different subtrees
    Create node x with height m_{ij}
    Add branch between x and root of subtree containing x_i
    Add branch between x and root of subtree containing x_j
Return tree with root at last added node

Some properties of this representation with a short justification:

1. *Note that at least $n-1$ entries must be specified for a matrix to be valid.* This follows from the fact that there are $n-1$ internal nodes in a tree, so the single link clustering algorithm requires at least $n-1$ finite heights to specify a tree.

2. An adjacency graph G over n nodes $v_1, \ldots, v_n$ can be associated with a matrix M having edges e_{ij} between v_i and v_j where ever m_{ij} is specified. *A matrix is valid if and only if the associated adjacency graph is connected.* This follows from the fact the single link clustering algorithm uses entries in the matrix associated with edges in the graph that form a path connecting all nodes. So, it is possible for a matrix to have $n-1$ entries specified, but still not being valid (Fig.1).

3. *To represent a distribution over* all *possible trees, all entries in the matrix must be specified.* This follows from the observation that in the set of all trees, for every pair $i, j, i \neq j$ of taxa x_i and x_j there exists a tree somewhere in the set of all trees such that the pair is connected with the lowest internal node (that is parent to both taxa). Since such a *cherry* can only be formed by the distance in the matrix m_{ij} being the smallest, this implies m_{ij} cannot be infinite, so m_{ij} must be specified.

4. *A matrix with n or more entries specified does not form a bijection to trees.* Since the single link algorithm only uses $n-1$ entries in the matrix to form the tree, if there are n or more values specified in the matrix it is possible to change these values and return the same tree. Therefore, there are multiple ways to represent the same tree with different matrices using hence it cannot be a bijection.

5. *Transforming a tree to a matrix consists of populating the specified entries m_{ij} with height of the most common recent ancestor of x_i and x_j in the tree.* The single link clustering algorithm then retrieves the original tree when applied to the matrix.

If all entries in the matrix on the first row are specified (i.e. entries m_{1j} for $1 < j \leq n$) and the remainder are not specified (Fig.1), all possible caterpillar trees with cherry containing x_1 can be formed. Note that there are $(n-1)!$ topologies that can be formed that way.

If all entries in the matrix that are 1-off-diagonal are specified, the adjacency graph forms a linked chain and $\frac{1}{n+1}\binom{2n}{n}$ (the Catalan number) number of topologies can be represented. We call this configuration a *cube*. Even though the same number of entries are specified, this is less than for the caterpillar trees, but will turn out to be more useful in practice. Note that by reordering the taxa it is possible to get the same kind of adjacency graph containing all nodes in a chain.

We will use a multivariate normal distribution $N(\mu, \Sigma)$ with mean vector μ and covariance matrix Σ of dimension k. Here $n-1 \leq k \leq n(n-1)/2$, where each of the k dimensions is associated with an entry in the matrix. To estimate μ and Σ for a tree set, simply construct the branch length matrix for each tree in the set and take the average for each entry (that is specified) and covariance for each pairs of specified entries.

To get a sample of trees, draw a random sample from this distribution, exponentiate each of the values and populate the matrix with its values. Then run single link clustering to obtain the tree. Algorithm 2 shows this in more detail.

Figure 1. Various matrices where x means the entry is specified, and ∞ the entry is unspecified for $n = 6$, the associated adjacency graphs for $n = 6$ and the number of tree topologies they can represent (in decreasing order).

Algorithm 2 Sample from posterior

Input: Multivariate normal distribution $N(\mu, \Sigma)$ of dimension k
      where μ the vector of mean values and Σ the covariance matrix.
Input: list of k matrix entries
Let $\mathbf{A}\mathbf{A}^T = \Sigma$ be the Cholesky decomposition of Σ
Draw k independent standard normal values $\mathbf{z} = \{z_1, \ldots, z_k\}$
Let $\mathbf{x} = \mu + \mathbf{A}\mathbf{z}$
Fill the matrix M with the k entries of $exp(\mathbf{x})$ and leave the rest unspecified
Return tree by applying single link clustering on M

Figure 2. Cube over 5 taxa consisting of a taxon ordering A, B, C, D, E and heights h_1 for the height between A and B, height h_2 for the heights between B and C, h_3 for C and D and height h_4 for D and E. Dashed lines show how the heights are associated with consecutive pairs of taxa in the taxon ordering.

2.2 Cube space

An alternative way to represent a cube configuration in a matrix is to consider an order of the taxa and a vector that represents the height of the internal nodes between two consecutive taxa in the ordering. Fig.2 illustrates this: the height h_1 corresponds to a matrix entry for A and B, the height h_2 for B and C etc., this way forming the cube configuration from the previous subsection. But we can consider this an ordering A, B, C, D, E and gaps h_1 to h_4 between consecutive nodes in the ordering.

One nice property of cube space is that it only contains $n - 1$ heights, and forms a bijection under the single link clustering transformation. However, it is obviously limited in that it cannot represent all of tree space, our expectation is that it can represent a large part of posterior space. Consider three taxa, A, B, C ordered alphabetically, then there are three possible topologies: $((A, B), C)$, $(A, (B, C))$ and $((A, C), B)$. The first two are represented in the cube, but the last one is missing, so at least two thirds of the posterior can be represented. In practice, posteriors will be skewed towards a dominant topology, so the first topology may have 60% support, the second 30% and leaving only 10% for the topology that is missed out. In this three taxon example, this means that 90% of the posterior can be captured. This is 30% better than using a summary tree that would only cover the first topology.

The method for learning a multivariate normal distributions over cube heights is the same as for matrices.

2.3 Variational Bayesian approach

The application of cube space we are most interested in is Bayesian inference through variational Bayesian methods. To this end, we have to infer the mean vector μ and covariance matrix Σ from an alignment and phylogenetic model. For each of the parameters of the model, we need a transformation so that the transformed parameter is normally distributed. We employ the following transformations:

- for scale parameters (like κ of the HKY model, birth rate of the Yule model) we use a log-transform.

- for location parameters we use no transform.

- for multi dimensional parameters that are constrained to sum to 1, e.g. frequencies of substitution models, we use a stick breaking transform (Betancourt, 2012).

- for the tree we use the matrix transform. Note that internal node heights are log-transformed.

Given a phylogenetic model that specifies a tree prior, site model and branch rate model, the parameters of the model (including the tree) and associated transformations, Algorithm 3 outlines how we can infer μ and Σ. The strategy is to get a reasonably good value for μ first, and infer Σ afterwards. To this end the first step is finding a state with high probability, ideally the maximum aposteriori probability (MAP) state, since it corresponds to the mode of the multivariate normal distribution, and thus should get us reasonably close to where μ should be.

Algorithm 3 Cube-VB

Input: posterior distribution over a set of parameters
Input: set of parameters (including tree) and their transforms
0. Initialise tree by UPGMA
1. Find high posterior state of other parameters and node heights
2. Randomly initialise ordering of taxa that is compatible with tree from step 1
3. Sample trees restricted to cube space, using MCMC
4. Estimate μ and Σ based on transformed MCMC samples

The tree is initialised using Unweighted Pair Group Method with Arithmetic Mean (UPGMA) RR (1958), a hierarchical clustering algorithm that produces a rooted time tree. The topology of the tree and timing of the tree can potentially be adjusted in the next step.

In our implementation, the tree and all other parameters are then optimised using simulated annealing (Van Laarhoven et al., 1987). This is a general purpose optimisation algorithm that, like MCMC, randomly explores the state space, but it differs from MCMC in the acceptance criterion of a proposed state. Higher posterior proposals are always accepted, and lower posterior proposals are accepted with a probability proportional to the change in posterior weighted by a temperature that decreases over time. We found that using a temperature schedule that decreases fairly fast to the end temperature, but repeatedly resets to the start temperature instead of a long monotonically decreasing temperature allowed escape from local minima more efficiently. For the proposals used in simulated annealing we can use standard MCMC operators implemented in BEAST 2.

To complete the cube-transform for the tree, we need a taxon ordering. Different orderings cover different parts of tree space, so finding a suitable ordering affects the suitability of the representation. This is a research question that we leave open here, but note that in the implementation using a random order compatible with the tree turned out to be sufficiently good to accurately infer parameters of interest such as tree height and tree length.

In order to estimate the covariance matrix Σ, we will generate a representative sample using MCMC. Instead of using the standard tree operators in BEAST 2, we replace the topology changing operators with a pair of operators that ensure

the proposed tree can be represented by the cube restricted by the taxon ordering from Step 2 of Algorithm 3. The first operators is a variant of the narrow exchange operator (a nearest neighbour interchange operator that also proposes a new height of the node that is moved) that ensure it remains within the cube. The second operators is a subtree-slide operator (moves internal node up or down, potentially causing topology changes) and ensures that the target location is compatible with the cube ordering.

Based on the short MCMC run, we can transform the MCMC samples and directly estimate μ and Σ from the set of transformed values. There are many tools for visualising and otherwise post-processing a posterior based on an MCMC sample, such as Tracer (Rambaut et al., 2018), DensiTree (Bouckaert, 2010), etcetera. Once we obtain μ and Σ, Algorithm 2 can be used to create such a sample.

3 RESULTS

In this section, we validate the matrix and cube representation of tree distributions and show that variational Bayesian methods on a cube passes well calibrated simulation studies and get and impression of its performance. Finally, we consider using a method for inferring summary trees.

3.1 Matrices and cubes can accurately represent tree distributions

To validate whether we can accurately represent a tree posterior obtained from a full MCMC run, we construct the distance matrix for each tree by taking the logarithm of the pairwise taxon path distance, then estimate the mean and covariance matrix for all relevant entries. Next, we sample a new tree distribution using Algorithm 2 repeatedly to obtain a new tree sample. We compare clade support and clade height distribution with the original tree posterior using the `CladeSetComparator` from the Babel package[1] in BEAST 2.

3.1.1 Apes

Fig.3 shows a small illustrative example of a tree distribution over six apes. There is uncertainty at the root with 83% support for siamang being outgroup (blue trees in panels b & d), 13% orangatan being outgroup (red trees) and 4% siamang and orangatan being ougroup (green trees). A cube can only represent two out of these three cases. By representing the first two, it covers 96% of the posterior distribution, which for many practical cases may be sufficient.

Representing the third case where both siamang and orangatan are ougroup is possible with a matrix representation, as shown Fig.3d.

Panels a & c show comparison of the individual clades. Red dots indicate clade support, running from zero probability (bottom left corner) to 1 (top right corner). Blue dots are placed at the mean of the height of a clade. Both x and y axes are scaled so that the largest tree height (from among both data sets) is at the top right corner. Blue crosses indicate 95% highest probability density (HPD) intervals of clade height. For both red and blue dots, bigger dots mean more clade support. The x-axis represent clade values from one set, the y-axis that of another. If both sets are the same, all dots will be on the x=y line (black diagonal line). The two blue lines indicate there is 20% difference: points within these lines have less than 20% difference, points outside have more, and can be considered as evidence there is substantial difference between the two posteriors.

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[1]Available from `https://github.com/rbouckaert/Babel`.

Figure 3. Comparing the ape dataset with a cube representation (a & b) and matrix representation (c & d). Plates a & c show clade support in red and clade height in blue (see text for details). Plates b & d show the original tree set (using different colours for different topologies) on the left and the cube (b) and matrix representation (d) on the right hand side of the plate. A cube accurately represents most clades, but does not cover all clades, where a matrix can represent all clades but is less accurate in clade support.

Figure 4. Comparing the AARS dataset with a cube representation. a) clade support as in Fig.3 and b) DensiTree of original and cube tree distribution. Clade support and heights show good correspondance as does the DensiTree.

What is striking in both panels a and c in Fig.3 is that the the mean of the clade heights and 95% HPD intervals are very close. However, clade support is slightly distorted: for the cube, the trees with both siamang and urangatan as outgroup is completely missing. Since this only covers 4% of the posterior, this may not be such a big issue in general. The matrix has this clade represented, but instead of the original 83%, 13%, 4% distribution has 59%, 21%, 20% distribution, heavily over representing the siamang and urangatan as outgroup. This is visible in Fig.3c by the red dots being far from the diagonal. It appears the correlation matrix does not accurately capture possible non-linear correlations between the entries in the matrix. A potential remedy is to use normalising flows (Rezende and Mohamed, 2015) to capture such non-linearities.

3.1.2 AARS

A tree distribution with more uncertainty is from aminoacyl-tRNA synthetase (AARS) data for 204 Class I protozyme sequences (Carter Jr et al., 2022). Based on the conditional clade distribution (Larget, 2013), it has an entropy of 124.1, so more than half of the internal nodes are uncertain. The corresponding cube has an entropy of just 96.3, which shows that a cube representation can have a large entropy, thus representing a highly uncertain tree distribution.

Fig.4a shows the clade set comparison for the AARS data, which shows that most of the clades that are supported in the cube have reasonably similar support as in the original tree set. Although there is a clade with 38.9% difference in clade support, and there are 5 other clades with more than 25% clade support difference, the average clade support difference is just 4.9% for clades with over 1% support. Note the set of red dots at the x-axis, which represent the set of clades that are not represented by the cube. Furthermore, clade heights also show very close correspondences: the mean height estimate differs just 2.2%. In summary, the clades match remarkably well given the high entropy of the tree set.

Fig.4b shows a DensiTree of the original tree distribution on the left hand side and the cube reprentation on the right hand side. Visually, the images correspond

Figure 5. Comparing the Pama-Nyungan dataset with a cube representation. a) clade support as in Fig.3 and b) DensiTree of original and cube tree distribution.

quite well, confirming the information we have from the clade comparison panel. Note that the cube represents a large number of clades judging from the diversity of trees in the DensiTree as well as the large number of points in the clade comparison plot. Even though a considerable number of clades are missing, the majority of well supported clades are represented by the cube.

3.1.3 Pama-Nynguan languages

Fig.5 shows another tree distribution, here for 306 Pama-Nynguan languages (Bouckaert et al., 2018) with a large entropy of 91.7 for the original tree distribution and 78.6 for the cube. Fig.5 shows the clade comparison, showing just 3 clades having more than 25% clade support, and two of them are 2 taxon clades. Again, we see good correspondence between original distribution obtained through MCMC and its cube representation.

3.1.4 Cichlids

A more troublesome example is a tree distribution over 121 Cichlid sequences sourced from (Matschiner et al., 2017). The entropy of the original dataset is 38.4 but that of the cube representation is higher at 49.8. Different orderings were tried based on the ordering heuristics in DensiTree, but all of them resulted in a cube representation with significant number of clades having substantially different support.

Fig.6 shows the clade comparison and DensiTree. There are 23 clades that have more than 25% difference in support and the maximum difference in clade support is 80% for a clade in the original set that is missing from the cube representation. The mean height of clades differ only 0.57%, so correspond quite well. In this case, a cube representation appears to be less satisfactory than in the other examples.

3.1.5 Indo-European languages

To demonstrate the cube representation works well for internal nodes, even though tip dates are sampled, we considered a recently published tree distribution over 161 languages (Heggarty et al., 2023). The tree set has an entropy of 42.3,the cube has 36.6.

Figure 6. Comparing the cichlids dataset with a cube representation. a) clade support as in Fig.3 and b) DensiTree of original and cube tree distribution. Significant number of clades have large difference in support in the cube representation, though heights of those clades that are captured show good correspondance.

Figure 7. Comparing the Indo-European dataset with a cube representation. a) clade support as in Fig.3 and b) DensiTree of original and cube tree distribution. Even with dated tips the cube representation remains accurate.

Sequence length:	100	250	500
Mean and 95% HPD Entropy	10.6 (5.9-14.6)	6.4 (3.2-10.4)	3.8 (1.3-6.3)
Tree height	95	94	93
Tree treeLength	94	95	96
BirthRate	92	94	94
Kappa	94	90	93

Table 1. Coverage (in %) of 100 runs over 50 taxa with Yule, HKY strict clock for three different sequence lengths. Longer sequences result in lower entropy, but coverage of the parameters of interest (last four lines in table) are all in acceptable range.

Maximum difference in clade support is 28.5% but only 2 clades have difference of 25% and on average the clade support difference is 4.1% for clades with more than 1% support. Mean height difference of clades (not counting tip heights) is less than 0.79% on average. So, even though the original distribution has dated tips, and the cube does not, clade support and clade heights remain quite closely represented when compared to the original.

3.2 Variational Bayesian methods on a cube passes well calibrated simulation study

To verify that the variational Bayesian algorithm on cube works, we performed a well calibrated simulation study (Mendes et al., 2023; Talts et al., 2018) using a tree with 50 taxa sampled from the Yule tree prior with birth rate narrowly distributed as normal (mean=6, sigma=0.1). This results in tree heights with mean 0.58 and 95% HPD of 0.34 to 0.79. Sequence data is generated under a HKY model with kappa lognormally distributed (mean=1,sigma=1.25), and a strict clock with rate fixed at 1 is used to complete the model. So, the parameter space being sampled consists of the kappa and birth rate parameters and a tree. We use sequence lenghts of 100, 250 and 500 resulting in observed entropy in the inferred tree posteriors of 10.6, 6.4 and 3.8 respectively (Table 1). We run 100 instances for each sequence length, and expect the 95%HPD of parameters of interest to contain the true value from 91 to 99 of the cases for the test to pass. Table 1 shows the results and the simulation study indeed passes.

A slightly more challenging case is where we sample trees from a coalescent tree prior with constant population size sampled from a lognormal (mean=0.1, sigma=0.1). The tree heights have a mean of 2.3 with 95% HPD interval of 0.6 to 5.0. We use a site model with gamma rate heterogeneity where the shape parameter is sampled from an exponential with mean 1, and use a GTR model with rate CT fixed to 1 and rate AG sampled from a log normal with mean 1 in real space and sigma=1. The other four transversions rates are sampled from a log normal with mean in real space of 0.5 and sigma=1. Stationary frequencies for the GTR model were sampled from a Dirichlet with $\alpha = 4$ for all four dimensions. For this setup, getting accurate estimates of the uncertainty required running the MCMC a bit longer than for the case using HKY. Table 2 shows the coverage we get for all paramaters of interest and they ar all in an acceptable range, although population size coverages were on the low side.

Sequence length:	100	250	500
Mean and 95% HPD Entropy	13.1 (5.9-15.7)	5.7 (3.0-9.1)	4.7 (1.5-6.5)
Tree.height	91	90	97
Tree.treeLength	90	91	97
rateAC	92	95	98
rateAG	100	98	95
rateAT	92	95	94
rateCG	93	96	94
rateGT	92	99	95
freqParameter.1	94	92	96
freqParameter.2	96	92	96
freqParameter.3	95	95	97
freqParameter.4	95	96	98
gammaShape	93	100	97
popSize	90	90	90

Table 2. As Table 1 but for coalescent with constant population, GTR with estimated frequencies using gamma rate heterogeneity and strict clock for three different sequence lengths. Coverage of the parameters of interest (last thirteen lines in table) are all in acceptable range.

Number of taxa	50	100	250	500	750	1000
MAP finding	1.17 (0.03)	1.48 (0.03)	2.36 (0.08)	3.91 (0.08)	5.30 (0.19)	6.43 (0.40)
MCMC	1.15 (0.03)	2.77 (0.05)	12.08 (0.24)	41.84 (0.72)	94.74 (3.29)	172.25 (8.24)
Decomposition	0.45 (0.01)	0.60 (0.02)	0.93 (0.03)	1.69 (0.04)	2.88 (0.20)	4.47 (0.10)
Sampling output	0.23 (0.02)	0.36 (0.02)	0.68 (0.02)	1.40 (0.03)	2.18 (0.05)	3.27 (0.15)
Total runtime	3.00 (0.05)	5.20 (0.08)	16.05 (0.27)	48.83 (0.79)	105.10 (3.33)	186.42 (8.60)

Table 3. Run times in seconds averaged over 25 instances for various taxon set sizes (in columns). Numbers in brackets are standard deviations.

This demonstrates that the variational Bayesian approach using cubes can work well estimating a tree, scale parameters as well as constrained sum parameters (i.e. frequencies) for a number of popular phylogenetic models. Many model implementations are available in BEAST 2 that only require these types of parameters, in particular nucleotide and amino acid substitution models.

3.3 Variational Bayesian method on a cube performs well

To see how well the variational Bayesian algorithm scales with number of taxa, we ran a simulation study with 50, 100, 250, 500, 750 and 1000 taxa. For each taxon set size, we ran 25 instances of the Yule/HKY/strict clock model used for the well calibrated simulation study. For each instance, alignments of 250 sites were generated. We checked that coverage of 95% HPD intervals for tree height, tree length, kappa and birth rate parameters was between 22 and 25 to verify that the inferred posterior indeed matches the true distribution. Experiments were performed on a 2021 MacBook pro laptop in OS X using BEAST v2.7.5 and BEAGLE v4.0.0 (Suchard and Rambaut, 2009).

Table 3 shows run times in seconds, demonstrating it is feasible to perform large

| | Sum of clade probability | | Log product of clade probability | |
Data set	MCC	Matrix	MCC	Matrix
Apes	4.8 =	4.8	-0.2 =	-0.2
IE	136.9 <	138.9	-35.0 <	-29.2
AARS	133.6 <	142.9	-128.4 <	-90.8
Cichlids	91.7 <	93.7	-42.3 <	-34.7*
NZ votes	26.7 <	34.1	-297.1 <	-217.6*
Pama-Nyungan	255.6 <	262.4	-80.5 <	-59.3
Species	27.5 =	27.5	-2.6 =	-2.6

Table 4. Sum and log product of posterior clade support over all clades in the summary tree for MCC tree and matrix representation based tree. Higher is better. The entries marked with asterisk ignore zero support clades (1 clade for cichlids, 3 for NZ votes). The matrix approach is never worse and sometimes considerably better.

analyses in a short span of time. Sampling a representative posterior trace and tree log took the least time taking up no more than 2% of total runtime. MCMC for accurately estimating the covariance matrix took most of the time, from 40% for smaller taxon sets to over 90% for the larger tree sets. MAP finding and Cholesky decomposition in comparison took less time and are taking up relatively less time with growing taxon set sizes.

Total calculation time fits a quadratic function ($1.70E - 4x^2 + 1.43E - 2x$) with $R^2 = 0.9997$, so the algorithm seems to scale quadratically. The chain length used are a linear function of the number of taxa, and tree likelihood calculation time (which dominates posterior calculation time here) both scale linear in the number of taxa. Changing the former could change how the algorithm scales. The most striking feature of the table though is that all times can comfortably be expressed in seconds using just three digits before the decimal point.

3.4 Summary tree

Observe that by setting $\mathbf{z} = \mathbf{0}$ in Algorithm 2 we can obtain a tree that is at the mode of the distribution. This tree should be representative of the complete distribution so can be used as a summary tree. Note that it can have a topology that differs from all trees in the tree set to be summarised, unlike popular methods like the maximum clade credibility (MCC) tree implemented in TreeAnnotator. Table 4 shows some the sum of posterior clade support in the summary tree for MMC and matrix representation based summary trees. This is proportional to the number of internal nodes minus the Robinson-Foulds distance averaged over the tree set. The table shows that matrix based summary tree never produced worse trees and sometimes does considerably better. The log product of clade probability is used in MCC trees to find the best fitting topology. Tab 4 shows even according to that criterion the matrix approach can do better. Further evaluation with less crude measures of summary tree quality are required to establish when matrix based summary trees outperform MCC trees.

4 DISCUSSION

Matrix and cube representations can accurately represent tree distributions: For tree distributions with low to medium entropy (Lewis et al., 2016) (according to the conditional clade distribution (Larget, 2013)), we demonstrated that a substantial part of the posterior can be represented using a single cube. For those cases it is possible to do fast inference of rooted time trees using a variational Bayesian approach outlined in Algorithm 3. For tree distributions with high entropy, it may not be possible to capture a substantial portion of the posterior by a single cube. Using more than one cube might elevate this somewhat, but since there can be overlap between the trees represented by the various cubes, and this overlap is hard to characterise[2], this may not be a fruitful avenue. A more productive way may be to use more entries in the distance matrix. This breaks the bijective property of the transformation, but facilitates a greater coverage of the full tree space. Normalising flows (Rezende and Mohamed, 2015) may provide a way to deal with the non-linear dependencies between the various matrix entries that may arise.

Optimisation based variational Bayesian: Variational Bayesian algorithms are usually optimisation based, where the evidence lower bound (ELBO) is maximised. The ELBO is proportional to the posterior plus a term that measures the distance between the true and approximate distribution. The benefit of maximisation over MCMC is that there are no restrictions on optimisation proposals unlike with MCMC where one needs to be careful to use appropriate Hastings ratios in order to ensure there is no bias in sampling from the posterior. One way to optimise the ELBO is by the hugely popular stochastic gradient ascent algorithm, which randomly draws a new state, calculates the gradient of the ELBO with respect to a random subset of parameters and moves parameters in the state in the direction of the gradient. This process is iterated untill convergence is reached. However, there are some issues with gradient ascent: a gradient is required, which can be computationally expensive. Furthermore, it has domain specific parameters, such as learning rate, learning rate scheme, and parameter subset size. Also, random subsets may need to be tuned for efficient exploitation of subsets. Finally, the convergence criterion has large impact on the number of iterations. Since points are drawn stochastically, the convergence criterion needs to deal with noise. Instead of dealing with all these issues, here we introduced an MCMC based variational Bayesian algorithm that benefits from all MCMC related optimisations implemented in BEAST 2 (such as adaptive operator tuning). It serves as a proof of concept demonstrating that cube spaces are suitable for variational inference. Of course, this still leave opportunities for ELBO maximisation implementations of cube based variational inference.

Variational Bayesian parameters: Notwithstanding the issues of ELBO optimisation based VB, the MCMC based variational Bayesian approach we introduced in Algorithm 3 has a number of parameters that need to be specified, in particular the start and end temperatures and number of repeats for simulated annealing for finding a (close to) MAP state. We chose a start temperature of 0.1 and end temperature of 0.01, which worked quite well for our experiments, so this seems a reasonable option. Putting in more than 25 repeats barely led to improvement of the posterior of the MAP state, but halving it turned out to be detrimental (well calibrated simulation studies did not pass).

Another parameter to set is the length of MCMC chain for estimating the covari-

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[2]Lars Berling personal correspondence.

ance matrix. The heuristic implemented in the `cubevb` package for BEAST 2 is to count the number of parameters and multiply with a user specified value. In our experiments with the HKY model, a multiplier of 1000 was sufficient. The more samples, the more accurate the matrix estimate but the longer it takes for the algorithm to finish. In that sense, it nicely fits the class of online algorithms (Bouckaert et al., 2022) that at any time can provide an answer, but the answer becomes more accurate the longer the algorithm runs. Detecting when longer runs do not increase accuracy of the posterior any more is an open question where automated convergence techniques employed for standard MCMC (Berling et al., 2023) could be exploited. These are all based on running two analyses in parallel and measuring the difference in the distributions that are obtained.

Why MCMC based variational Bayesian works: Since plain MCMC is used in Algorithm 3 one may wonder why this is more efficient than just running full MCMC and not being restricted to a cube space. The reason for this stems from the assumption that the true distribution can be represented as a transformation of a multivariate normal distribution. Consider the κ parameter for the HKY model and a log-transformation. In practice, we see that κ is close to log-normally distributed, so representing the posterior as a normal with log-transform leads to little distortion. Now notice that we only have to estimate the mean and variance of the log-normal under this assumption, instead of estimating a full unparameterised distribution, which is what we would be doing with full MCMC. So, even though a small MCMC sample will not be sufficient to accurately represent the distribution of κ and statistics like 95% HPDs, there will be sufficient information to accurately estimate two parameters of a log-normal, from which statistics like 95% HPDs then can accurately be inferred.

Variational Bayesian for matrix representation: The variational Bayesian algorithm can be adapted to suitable matrix representations by replacing Step 2 by some method of identifying which entries in the matrix need to be specified and which entries can be left unspecified. Step 3 would need adjustment in the operators to ensure all possible tree topologies that can be represented by the matrix are available to be proposed. It is an open problem what the best strategy is for identifying which and how many entries should be specified. Since a cube can capture a substantial part of a posterior, identifying a cube could be a first step followed by filling in entries where clades represented have

5 CONCLUSION

Cube space is introduced as a representation of a significant part of tree space. We demonstrated for simple phylogenetic models it allows fast inference of the most relevant part of the tree posterior, thus capturing most of the uncertainty in the posterior distribution.

There are many venues to extend and develop this work, in particular into the area of more sophisticated phylogenetic models. Here we mention just a few of the more commonly used models and potential issues implementing them in VB with cubes:

- Finding the optimal ordering for a cube is an obvious open problem. It is probably related to finding a good ordering for drawing trees in DensiTree (Bouckaert, 2010) for which various heuristics have been developed. It is not

clear how to transfer these heuristics to the VB setting and new heuristics may need to be developed.

- Flexible nonparametric tree priors take epochs in accounts like the Bayesian skyline plot (Drummond et al., 2005) and birth death skyline plot (Stadler et al., 2013) and are widely used but may be challenging to parameterise in a variational Bayesian setting due to their discrete components. Models that integrate out most parameters and have only a single continuous parameter, like BICEPS and multi Yule (Bouckaert, 2022), can be good alternatives and should be straightforward to implement in out cube VB algorithm.

- Relaxed clocks (Drummond et al., 2006; Douglas et al., 2021) need to be adapted for cube space. Due to discontinuities between different topologies, simply associating a (log)normal branch rate distribution with each taxon and each dimension of the cube may not be sufficient. Normalising flows (Rezende and Mohamed, 2015) may be required to capture relaxed branch rate distributions.

- Dated tips put restrictions on the height of internal nodes. When tip dates are uncertain and need te be sampled from a distribution, this adds extra dimensions to the multivariate normal representation of the posterior.

On the other hand, extensions into empirical models of amino acids, models that integrate out internal states like ancestral reconstruction (Lemey et al., 2009) and phylogeography (Bouckaert, 2016) should be straightforward to adapt.

In summery, there are many exciting new opportunities waiting to explore cube and matrix space for fast phylogenetic inference.

SOFTWARE AND DATA AVAILABILITY

The software is open source and available as `cubevb` package to BEAST 2.7 from `https://github.com/rbouckaert/cubevb`. Tree sets and data used in the well calibrated simulation study, and performance study is available from `https://github.com/rbouckaert/cubevb/releases/tag/v1.0.0`.

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