Main text

The genetic architecture of the load linked to dominant and recessive self-
incompatibility alleles in *Arabidopsis halleri* and *A. lyrata*

Le Veve Audrey[1,2], Genete Mathieu[1], Lepers-Blassiau Christelle[1], Ponitzki Chloé[1], Poux
Céline[1], Vekemans Xavier[1], Durand Eleonore[1], Castric Vincent[1*]

[1]Univ. Lille, CNRS, UMR 8198 – Evo-Eco-Paleo, F-59000 Lille, France

[2]Current address: Department of Botany, Faculty of Science, Charles University, Benátská 2,
CZ-128 01 Prague, Czechia

* Correspondence : vincent.castric@univ-lille.fr

Abstract

The long-term balancing selection acting on mating-types or sex determining genes is expected to lead
to the accumulation of deleterious mutations in the tightly linked chromosomal segments that are
locally "sheltered" from purifying selection. However, the factors determining the extent of this
accumulation are poorly understood. Here, we took advantage of variations in the intensity of
balancing selection along a dominance hierarchy formed by alleles at the sporophytic self-
incompatibility system of the Brassicaceae to compare the pace at which linked deleterious mutations
accumulate among them. We first experimentally measured the phenotypic manifestation of the
linked load at three different levels of the dominance hierarchy. We then sequenced and phased
polymorphisms in the chromosomal regions linked to 126 distinct copies of *S*-alleles in two
populations of *Arabidopsis halleri* and three populations of *A. lyrata*. We find that linkage to the *S*-
locus locally distorts phylogenies over about 10-30kb along the chromosome. The more intense
balancing selection on dominant *S*-alleles results in greater fixation of linked deleterious mutations,
while recessive *S*-alleles accumulate more linked deleterious mutations that are segregating. Hence,
the structure rather than the overall magnitude of the linked genetic load differs between dominant
and recessive *S*-alleles. Our results have consequences for the long-term evolution of new *S*-alleles,
the evolution of dominance modifiers between them, and raise the question of why the non-
recombining regions of some sex and mating type chromosomes expand over evolutionary times while
others, such as that the *S*-locus of the Brassicaceae, remain restricted to small chromosomal regions.

Keywords: Supergene, sheltered load, mating types, balancing selection, genetic dominance, S-locus.

Introduction

The existence of sexes or mating-types leads to one the strongest forms of long-term balancing selection, and is often associated with clusters of polymorphisms around sex/mating-type controlling regions kept together by structural rearrangements. In some cases, such rearrangements can span almost entire chromosomes, *e.g.* sex chromosomes in mammals (Katsura et al., 2012) or mating-type chromosomes in ascomycete fungi (Hartmann et al., 2021), while in others they remain limited to relatively small genomic regions, e.g. chromosomal inversions controlling male reproductive morphs in the ruff (Lamichhaney et al., 2016), mating-type loci in some basidiomycete fungi, segregating indels controlling pin *vs*. thrum floral morphs in Primula (Cocker et al., 2018). The long-term balancing selection acting on these systems is expected to lead to the accumulation of deleterious mutations in the tightly linked chromosomal segments that are "sheltered" from purifying selection by the presence of the balanced polymorphism (Uyenoyama, 1997; Uyenoyama, 2005). These deleterious mutations can have drastic short- and long-term consequences for the evolution of the species, and determining the processes by which they accumulate is crucial to understand how the rearranged regions can either expand along the chromosomes or conversely remain restricted to limited genomic tracts (Jay et al,. 2021; Jay et al,. 2022).

Self-incompatibility (SI) is a genetic mechanism allowing recognition and rejection of self-pollen by hermaphrodite individuals, thereby preventing inbreeding and promoting outcrossing in hermaphroditic plant species (Nettancourt, 2001). In the Brassicaceae family, SI is controlled by a single non-recombining chromosomal region, the *S*-locus (Schopfer et al., 1999; Kusaba et al., 2001). SI is one of the most prominent examples of long-term balancing selection (Uyenoyama, 2003), and as such deleterious mutations are expected to accumulate in very close genetic linkage to the *S*-alleles because of the indirect effects of linked selection (Uyenoyama, 2005). Population genetics models predict that recessive deleterious variants should accumulate within specific *S*-allele lineages (Uyenoyama, 2003; Llaurens et al., 2009a), and should then be reshuffled among them by recombination. However, due to the technical difficulty of phasing polymorphisms, this process has rarely been characterised in detail (Castric and Vekemans, 2004).

A key feature of sporophytic SI systems, also shared by sex chromosomes, is the existence of dominance interactions between *S*-alleles. While most individuals are heterozygous at the *S*-locus and thus carry two different *S*-alleles, only one of them is generally expressed at the phenotypic level. This is especially true for the pollen specificity, where *S*-alleles follow a complex genetic dominance hierarchy (Llaurens et al., 2008; Durand et al., 2014). In *Arabidopsis lyrata* and *A. halleri*, the S-alleles are distributed into four classes of dominance, with increasing dominance from class I to class IV and linear dominance also observed inside some of the classes. Similar patterns of dominance between S-alleles occur in Brassica species, but with only two classes (Hatakeyama et al., 1998; Kakizaki et al., 2003; Yasuda et al., 2017). This system is genetically determined and controlled by small RNAs molecules produced by dominant *S*-alleles that are able to target and repress expression of the recessive *S*-alleles in pollen (Durand et al., 2014; Yasuda et al., 2017). The evolutionary properties of *S*-alleles are expected to vary in a predictable manner along the dominance hierarchy because balancing selection acts more strongly on dominant than on recessive *S*-alleles, as the latter are often masked at the phenotypic level (Billiard et al., 2007). As a result, the dynamics of accumulation of deleterious variation may differ in close linkage with dominant *vs* recessive *S*-alleles. Specifically,

recessive *S*-alleles can form homozygous combinations in natural populations more often than
dominant *S*-alleles (Billiard et al., 2007), such that recombination can occur occasionally between
distinct gene copies of the same recessive *S*-allele, providing the opportunity for the linked recessive
deleterious mutations to be purged from within the *S*-locus itself. In addition, because recessive *S*-
alleles reach higher population frequencies (Billiard et al., 2006; Llaurens et al., 2008), purifying
selection on linked deleterious variants is expected to have higher efficacy among gene copies of
recessive than dominant *S*-alleles. This is expected to result in a higher fixation probability of
deleterious variants linked to the class of dominant *S*-alleles than to the class of recessive *S*-alleles
(Llaurens et al., 2009a). Thus, the level of dominance of S-alleles determines the intensity of purifying
selection acting upon them. This situation closely resembles that of the differential evolution of sex
chromosomes, where Y chromosomes (similar to dominant S-alleles) tend to accumulate more
deleterious variation than X chromosomes (similar to recessive S-alleles, Llaurens et al. 2009a; Goubet
et al. 2012). Empirical support for this simple prediction has been conflicting, though. Based on
phenotypic measurements in *A. halleri,* Llaurens et al. (2009a) observed a decrease of fitness
associated by enforced homozygosity for one of the most dominant *S*-alleles (Ah15) but not for the
most recessive *S*-allele of the allelic series (Ah01). In contrast, Stift et al. (2013) observed no effect of
dominance on the genetic load linked to three dominant *vs*. recessive *S*-alleles in a natural population
of the closely related *A. lyrata*. Hence, the data available so far are inconclusive, but are restricted to
very small numbers of *S*-alleles. They are also based on inherently limited phenotypic measurements,
seriously limiting the power of the comparisons, and preventing proper generalisation of the effect of
the intensity of balancing selection on the accumulation of linked deleterious variation.

In this study, we combined phenotypic, genomic and theoretical approaches to finely dissect the
patterns of accumulation of deleterious variation linked to the *S*-locus supergene in *A. halleri* and *A.*
lyrata, depending on dominance levels of S-allele. We first extended the phenotypic approach of
Llaurens et al. (2009a) to a series of additional *S*-alleles from the same local *A. halleri* population to
evaluate the effect of *S*-allele dominance on the sheltered load. We then used parent-offspring trios
and targeted genome re-sequencing to directly quantify the accumulation of putative deleterious
mutations linked to phased dominant *vs*. recessive *S*-alleles in two *A. halleri* and three *A. lyrata* natural
populations. Finally, we used stochastic simulations to refine the theoretical predictions about the
patterns of accumulation of recessive deleterious mutations linked to dominant *vs.* recessive *S*-alleles.
Overall, our results provide a more nuanced view of the effect of the intensity of balancing selection
on the sheltered load, in which the structure of the sheltered load rather than its magnitude differs
among *S*-alleles from different dominance classes.

Results

The genetic load linked to the S-locus varies among S-alleles, but is not correlated with dominance.

We first expanded the experimental approach of Llaurens et al. (2009a) to phenotypically evaluate the
effect of *S*-allele dominance on the intensity of the sheltered load. The previous study focused on
three *S*-alleles (Ah01, Ah02 and Ah15; Llaurens et al., 2009a). Here we included two *S*-alleles from the
same local population (Nivelle, France): Ah03 and Ah04, and included Ah01 again for comparative
purposes. In the Arabidopsis genus, *S*-alleles have been shown to form a complex dominance hierarchy
(Llaurens et al., 2008; Durand et al., 2014). This hierarchy is largely associated with the phylogeny of
S-alleles (Durand et al., 2014), and at least four phylogenetic classes (I, II, III and IV) have been

described, from the most recessive (class I) to the most dominant of S-alleles (class IV). Dominance interactions also exist among S-alleles within classes, such that these five S-alleles form the following dominance hierarchy (Llaurens et al., 2008; Durand et al., 2014): Ah01<Ah03<Ah02<Ah04<Ah15, from the most recessive (Ah01) to the most dominant (Ah15). To reveal the linked load, we enforced homozygosity at the S-locus using controlled crosses between parental individuals sharing a given S-allele that was masked by different dominant S-alleles (*e.g.*, to obtain Ah_xAh_x homozygotes we deposited pollen from a Ah_xAh_y plant, where $Ah_y>Ah_x$, on pistils of a Ah_xAh_z plant where z≠y, or on a Ah_xAh_x pistil when available; table S1). We obtained 399 offspring from a total of six such crosses. Note that our experimental procedure differs slightly from that of Llaurens et al. (2009a) in that their procedure required a CO_2 treatment to bypass the SI system and obtain selfed offspring, while here we took advantage of the dominance interactions to obtain outcrossed S-locus homozygous individuals that we phenotypically compared to their full-sibs with S-locus heterozygous genotypes. Note also that the S-locus homozygous offspring we obtained contain distinct gene copies of a given S-allele lineage. Hence, they could in principle carry distinct suites of linked deleterious mutations in case these mutations segregate within S-allele lineages.

We first tested whether homozygosity at the S-locus affected survival by measuring for each cross the proportion of homozygotes at the S-locus reaching the reproductive stage for three S-alleles (in two replicate families per S-allele, Table S1). The proportion of Ah01/Ah01 and Ah04/Ah04 homozygotes surviving to the reproductive stage was consistent with mendelian expectations in their respective families. However, we observed a significant decrease of Ah03/Ah03 homozygotes at the reproductive stage compared with Mendelian expectations (Table 1), whereas the observed proportion of the Ah03 S-allele among heterozygous individuals did not depart from expectations (2/3=0.67; Table 1). Thus, the increased mortality is associated with Ah03 homozygosity, rather than with a lower performance of individuals carrying the Ah03 S-allele itself. Overall, a genetic load was thus observed linked to the Ah03 S-alleles, which is at an intermediate level of dominance, but neither to the most dominant (Ah04) nor to the most recessive (Ah01) S-allele. Hence, these observations do not support a positive correlation between S-allele dominance and the magnitude of the sheltered load.

Table 1: Proportion of S-locus homozygous offspring having reached the reproductive stage for three different S-alleles. *The test is performed relative to the expected proportion of homozygous genotypes in the offspring (25% when both parents are heterozygous; 50% when one of the parents is homozygous and the other heterozygous).*

S-allele	Level of dominance	Number of seedlings having reached the reproductive stage	Observed proportion of homozygotes	Ratio of the observed to expected proportion of homozygous genotypes (p-value[a])	Number of heterozygotes with the S-allele (p-values[a])
Ah01	I	35	0.29	1.14 (0.76)	19 (0.40; 0.90)
Ah03	II	27	**0.074**	**0.3 (0.02)**	17 (1; 0.69)
Ah04	III	96	0.479	0.96 (0.39)	50

Next we measured thirteen vegetative and reproductive traits in the resulting families and compared
offspring that were homozygous for their *S*-alleles with their full sibs that were heterozygous (Fig. 1).
We first used permutations to test whether the mean trait value of homozygotes differed from that
in heterozygotes. Overall, with a single exception, we found no effect of homozygosity at the *S*-locus
on variation of the traits measured (Fig. 1; Table S2). The maximum length of flowering stems was the
exception to this general pattern, with longer reproductive stems for *S*-locus homozygous than
heterozygous genotypes, hence in the opposite direction from our expectation of lower fitness in
homozygotes. For this trait, there was significant variation among replicate families for homozygotes
of the recessive allele Ah01 but not of the dominant allele Ah04 (Table S3). We then used generalised
linear models (GLM) to evaluate the effect of dominance (considered as a continuous variable with
fixed effect) on the mean phenotypic value of homozygotes compared to heterozygotes for each trait
(Table S4; treating family of origin, attacks by phytopathogens, phytophagous and oxidative stress as
random effects whenever necessary). We also observed no effect of *S*-allele dominance on the
contrast between *S*-locus homozygotes and heterozygotes for any of these traits. A single of the
thirteen traits was an exception to this general pattern, but again the effect was in the opposite
direction from our expectation, with an earlier rather than delayed appearance of the first leaf for
homozygotes of more dominant *S*-alleles; Table S4). Overall, our phenotypic results confirmed the
presence of a detectable linked load on some phenotypic traits (survival; time to produce the first
leaf), but we could not replicate the observation of Llaurens et al. (2009a) that dominant *S*-alleles carry
a more severe deleterious load than recessive *S*-alleles, even though our samples were obtained from
the same local population.

Figure 1: Effect of homozygosity at the S-locus on 13 phenotypic traits compared to heterozygotes.
For each trait, the phenotypic values in homozygotes (in grey) were normalised relative to the mean phenotypic values in heterozygotes (in black). The distributions were obtained by 10,000 random permutations. SD: Standard Deviation.

S-alleles are associated with specific sets of tightly linked mutations.

The model of the sheltered load assumes that distinct *S*-allele lineages carry specific sets of linked deleterious mutations, but to our knowledge this prediction was never tested directly. We combined a parent-offspring trio approach with sequencing of the *S*-locus flanking regions to phase the mutations segregating in the *S*-locus flanking regions with their respective *S*-alleles. Briefly, we used a previously developed sequence capture protocol specifically targeting the nucleotide sequences over 75 kb on each side of the *S*-locus along with a series of 100 control regions from throughout the genome (Le Veve et al., 2023), and we analysed nucleotide sequence polymorphism (including only invariant and biallelic SNPs), based on the *A. lyrata* reference genome (Hu et al., 2011). We define a haplotype as a unique combination of mutations along the phased chromosome, and a *S*-allele lineage as the collection of gene copies of a given functional *S*-allele (different functional *S*-alleles are distinguished based on their strong sequence divergence at the *S*-locus pollen and pistil genes). Different gene copies within an *S*-allele lineage can thus be associated with distinct linked haplotypes in the flanking regions. The *S*-alleles were identified based on short reads sequences according to a previously published method (Genete et al., 2020). We analysed two closely related *A. halleri* populations from Europe (Nivelle and Mortagne) and three allogamous *A. lyrata* populations from North America (IND, PIN and TSS; Foxe et al., 2010). Overall, we were able to reconstruct 34 haplotypes

linked to a total of 12 distinct *S*-allele lineages in Nivelle, 38 haplotypes linked to 11 distinct *S*-allele
lineages in Mortagne and 16, 22 and 16 haplotypes associated with 6, 7 and 5 distinct *S*-allele lineages
in populations IND, PIN and TSS, respectively (Table S5). Nine of the *S*-alleles were shared between
the two *A. halleri* populations (Ah01, Ah03, Ah04, Ah05, Ah12, Ah20, Ah24, Ah25 and Ah59). In the
populations of *A. lyrata*, four *S*-alleles were shared between PIN and TSS (Ah01*, Ah03*, Ah18* and
Ah63*), five *S*-alleles were shared between PIN and IND (Ah01*, Ah03*, Ah46* and Ah63*), four *S*-
alleles were shared between IND and TSS (Ah01*, Ah03*, Ah31* and Ah63*), and three were shared
across all three (Ah01*, Ah03* and Ah63*). Note that for convenience, we used *A. halleri* notations
(with the addition of a *) to refer to the trans-specifically shared *A. lyrata S*-alleles. Altogether, we
were able to obtain the phased flanking sequences of 126 *S*-locus haplotypes, comprising a total of
4,854 variable sites. This provides considerable power to evaluate the local accumulation of linked
mutations across *S*-alleles of different levels of dominance and to examine their patterns of
conservation between populations and between species.

Mutations in the *S*-locus flanking regions can be exchanged between *S*-alleles by recombination, and
between local populations by migration (Charlesworth, 2006). The relative time scale of these two
processes (recombination *vs.* migration) determines the distribution of the linked mutations. To
capture the chromosomal extent of this effect of linkage to *S*-alleles, we developed a new phylogenetic
method comparing the likelihood of two contrasted topologies of interest in overlapping windows
along the chromosome: (1) the topology clustering haplotypes by the populations where they came
from vs. (2) the topology clustering them by the *S*-allele to which they are linked (Fig 2A). This allowed
us to evaluate the progressive shift from a predominant topology by *S*-alleles close to the *S*-locus to a
topology by populations further along the chromosome and in unlinked control regions (Fig. 2B). The
difference in log likelihood between the two topologies decreased significantly with distance to the *S*-
locus (Pearson coefficient = -0.015 and -0.010 for *A. halleri* and *A. lyrata* respectively; *p*-values <2^{e-16}).
In *A. halleri*, the topology grouping haplotypes by populations became more likely than the topology
grouping them by *S*-alleles at a distance of around 30kb from the *S*-locus, but even at a distance of
50kb the phylogenetic structure was still different from that in regions unlinked to the *S*-locus used as
controls for the genomic background (Le Veve et al., 2023; Fig 2B). In *A. lyrata,* the shift was even
more rapid (within 10-15kb), although we note that the phylogenetic structure of the control regions
was less resolved (Fig 2B). To evaluate these patterns more directly we first examined the data using
a Major Component Analysis (MCA, a modified version of PCA adapted to binary data, Fig S1 and S2)
and using simple phylogenetic reconstructions (Fig S3, S4, S5 and S6). We confirmed that haplotypes
linked to a given *S*-allele tended to cluster together in the most tightly linked region, and that this
grouping by *S*-alleles was progressively lost in favour of a grouping by population of origin in the most
distant regions. Following Kamau et al. (2007), we compared the fixation index F_{ST} among local
populations and among *S*-alleles in *A. lyrata* and *A. halleri*. In both *A. halleri* and *A. lyrata*, F_{ST} values
among *S*-alleles were high in regions close to the *S*-locus and quickly decreased to reach the
background level (Fig. S7) as the distance from the *S*-locus increased. In parallel, the differentiation
among populations followed roughly the opposite pattern, i.e. it was initially low in regions close to
the *S*-locus (as expected under strong balancing selection) and increased up to background level within
the first few kilobases (Fig. S7). In line with our phylogenetic analysis, differentiation between
populations started to exceed differentiation between *S*-alleles much closer to the *S*-locus in the *A.
lyrata* than in the *A. halleri* populations (Fig S3, S4, S5 and S6). Finally, we explored the fine-scale
patterns of association within populations between individual *S*-alleles and SNP in the linked and the

control regions (Fig S8). As expected, the vast majority of significant associations were found for the most closely linked SNPs. With a single exception, all *S*-alleles were associated with unique SNPs in the 50kb region around the *S*-locus, albeit with substantial heterogeneity among *S*-alleles in the patterns and extent of associations that they show (Fig S8). Overall, our results indicate that due to limited recombination, the *S*-alleles carry a specific set of polymorphic sites in the linked region. This association fades away for more distant sites over a few kilobases, where population structure becomes predominant, as in the rest of the genome. Hence, different *S*-alleles are associated with specific sets of tightly linked mutations, but only within 10-30kb.

Figure 2: Linkage to the S-locus locally distorts the phylogenetic relationships. A. The two topologies of interest cluster haplotypes either by the S-allele to which they are linked (top) or by the populations where they came (bottom). Different S-alleles are represented by symbols of different colours, different populations of origin are represented by symbols of different shapes. B. Difference in log likelihood of the two topologies of interest. Dots correspond to the difference in log likelihood for overlapping series of 50 SNPs around the S-locus for A. halleri (top panel) and A. lyrata (bottom panel). Positive values correspond to chromosomal positions where the topology by S-alleles explains the phylogeny of haplotypes better than the topology by populations. The right panels show the difference in log likelihood in the control regions. 2.5 and 97.5 percentiles of the distribution in the control regions are indicated by dashed lines.

No overall evidence that dominant S-alleles accumulate more linked deleterious mutations.

Llaurens et al. (2009a) predicted that recessive deleterious mutations should fix more readily when linked to dominant *S*-alleles than when linked to recessive *S*-alleles. To test this prediction, we investigated the correlation between the level of dominance of the *S*-alleles and their total number of 0-fold degenerate mutations (S_{0f}) or the ratio of 0-fold to 4-fold mutations (S_{0f}/S_{4f}) for the phased

haplotypes, assuming that the vast majority of 0-fold degenerate mutations are deleterious. Based on
the results presented above and the results of our previous study (Le Veve et al., 2023), for the rest of
our analyses we focused on the phased haplotypes over the first 25 kb on either side of the *S*-locus.
We found no overall effect of dominance on S_{0f} (p-values= 0.54 and 0.07 for *A. halleri* and *A. lyrata*
respectively; Fig. 3; Table S6) or S_{0f}/S_{4f} (p-values= 0.54 and 0.07 for *A. halleri* and *A. lyrata* respectively;
Table S6). Extending the analysis to all non-synonymous mutations, or to deleterious mutations
predicted by SIFT4G and by SNPeff led to identical conclusions (Table S6). Overall, our genomic results
did not confirm the prediction that dominant *S*-alleles accumulate a larger number of putatively
deleterious mutations in their linked regions. We note that the particular *S*-allele whose sheltered load
was quantified in Llaurens et al. (2009a) (Ah15, red arrow on Fig 3A) appears to be one of the *S*-alleles
associated with the highest number of 0-fold degenerate mutations among all *S*-alleles of the most
dominant class (class IV).

Figure 3: No overall effect of S-allele dominance on the total number of 0-fold degenerate mutations
(S_{0f}) in the linked genomic regions within 25kb. *Each dot represents the mean number of mutations*
observed among haplotypes linked to one S-allele in one population. The correlations evaluated by a
GLM are represented by lines, with confidence intervals represented in grey. The dominance was
considered as a continuous variable. A: A. halleri. Black dots correspond to the Nivelle population, grey
dots to the Mortagne population. The red arrow points to the copy of Ah15, corresponding to the S-
allele whose sheltered load was phenotypically characterised by Llaurens et al. (2009a) in Nivelle. B: A.
lyrata. Red dots correspond to the IND population, black dots to the PIN population and blue dots to
the TSS population.

The structure of the linked genetic load differs between dominant and recessive S-alleles.

Theory predicts that dominant *S*-alleles should fix linked recessive deleterious mutations with a higher
probability than recessive *S*-alleles (Llaurens et al., 2009a), but in natural populations we observed no
difference in the total number of putatively deleterious linked to dominant *vs*. recessive *S*-alleles. To
clarify this discrepancy, we took advantage of our sequencing of multiple copies of *S*-alleles to consider
separately the fixed and the segregating mutations linked to each of the *S*-allele lineages. For each

population, we included only mutations that were segregating, and excluded those that were locally fixed. In agreement with the prediction of Llaurens et al. (2009a), we observed that lineages of dominant S-alleles do indeed tend to fix deleterious mutations more readily (Fig. 4; Table S6). This conclusion held true when extending the analysis to all non-synonymous mutations and to the lowly and the moderately deleterious mutations predicted by SNPeff (Table S6). This was also true using SIFT4G to identify deleterious mutations, with the only exception of a non-significant correlation for *A. halleri,* which might be due to the low number of nucleotide sites included in the SIFT4G database, resulting in low power to detect differences (Table S6). The fact that dominant S-alleles tend to fix deleterious mutations more readily but do not accumulate a larger total number of deleterious mutations is explained by the fact that the structure of the genetic load differs between dominant and recessive S-alleles: the dominant S-alleles tend to have more fixed deleterious mutations, but the recessive S-alleles compensate by accumulating a larger number of segregating mutations, resulting in similar numbers of deleterious mutations overall in most of the populations.

Figure 4: The number of 0-fold degenerate mutations fixed in the 25kb regions flanking the S-locus increases with dominance of the S-allele associated. *Each dot represents the value obtained for haplotypes linked to one S-allele in one population. The correlations evaluated by a GLM are represented by lines, with confidence intervals represented in grey. The dominance was considered as a continuous variable. A: A. halleri. Black dots correspond to the Nivelle population, grey dots to the Mortagne population. B: A. lyrata. Red dots correspond to the IND population, black dots to the PIN population, blue dots to the TSS population.*

Motivated by these empirical observations, we built upon the model by Llaurens et al. (2009a), who showed that linked deleterious mutations (especially fully recessive ones) are expected to fix within dominant S-allele lineages more readily than within recessive S-allele lineages. Here, we adapted the model to focus not only on fixed deleterious mutations, but also on those that are segregating within allelic lineages. Our stochastic simulations confirmed that, at equilibrium, dominant S-alleles tend to accumulate a larger number of recessive deleterious mutations that are fixed among gene copies within S-allele lineages (Fig. 5A). In contrast, the number of segregating linked mutations was higher for recessive than for dominant S-alleles (Fig. 5B). These two effects eventually compensate each

other, such that in the end the mean number of linked deleterious mutations per copy of S-allele was
not expected to change between dominant and recessive *S*-alleles (Fig. 5C). These predictions are in
line with our genomic observations and suggest that the dominance level of *S*-alleles modifies the
structure of the genetic load they shelter: dominant *S*-alleles accumulate more fixed deleterious
mutations, but recessive *S*-alleles accumulate more segregating mutations, resulting in an equivalent
load overall.

Figure 5: Stochastic simulations confirm the contrasted architecture of the load of deleterious mutations linked to dominant vs. recessive S-alleles. *Number of fixed (A), segregating (B) and the mean number of deleterious mutations (C) linked to S-alleles at four different levels of dominance (I<II<III<IV). The means (bold lines) were estimated per S-allele dominance classes over 100 replicate simulations after discarding an initial burn-in of 100,000 generations. h=0. s=0.01.*

Discussion

The genetic load linked to the S-locus is detectable and is manifested on different phenotypes

Our results contribute to a growing body of evidence confirming that the accumulation of deleterious
mutations linked to strongly balanced allelic lines can be substantial, and that their effect can be
detected at the phenotypic level (Lane and Lawrence, 1995; Stone, 2004; Llaurens et al., 2009a; Mena-
Ali et al., 2009; Stift et al., 2013; Vieira et al., 2021). An interesting observation is that the phenotypes
on which the load was revealed varied among these studies. Here, the effect of homozygosity at the
S-locus was apparent on juvenile survival and on the length of the longest flowering stem, but we
detected no effect on any other morphological measurements, including leaf and rosette traits. In the
same population of *A. halleri,* Llaurens et al. (2009a) detected an effect on juvenile survival and on
leaf size. A study in North American outcrossing populations of *A. lyrate* (Stift et al., 2013) detected an
effect on juvenile survival, but not on any other traits that they measured. In the horsenettle *Solanum*
carolinense, the load was associated with reduced seed viability, flower number and germination
(Stone, 2004; Mena-Ali et al., 2009). Hence, the most consistent pattern seems to be a decrease of
overall juvenile survival, possibly because it is a highly integrative measurement of fitness, whereas
other morphological or life history traits can be associated with more specific components of overall
fitness.

A unique genetic load associated with each allele in each population.

The model of the sheltered load posits that each *S*-allele should be associated with a specific set of
linked mutations (Llaurens et al., 2009a). In line with this prediction, the magnitude of the *S*-linked
load varied among *S*-alleles, as the load linked to some *S*-alleles was phenotypically detectable, while
for others it was not. This variation of the genetic load is expected, since deleterious mutations
associated with the different alleles are likely to hit different linked genes, and affect different
phenotypic traits with different effects on fitness. Also in line with the model of the sheltered load,
our phasing of a large number of variants linked to *S*-haplotypes in several natural populations
revealed that the same suite of linked mutations was consistently associated among different copies
of a given *S*-allele when sampled from within the same population, in particular for the dominant *S*-
allele lineages under more intense balancing selection. As expected for outcrossing populations with
short-scale linkage disequilibrium, this association was lost when examining sites at increasing genetic
distances from the *S*-locus along the chromosome (see also Le Veve et al., 2023). Finally, the
association with linked sites was further lost when comparing gene copies of *S*-alleles sampled from
different local populations, suggesting that recombination within populations decouples alleles from
their linked sites faster than migration can homogenise the genetic composition among these natural
populations. We note that the patterns of association and phylogenetic structure differed among
populations, possibly due to their contrasted demographic histories. Indeed, the *A. lyrata* populations
colonised North America from ancestral European populations about 20-30.000 years ago (Clauss et
al., 2006; Ross-Ibarra et al., 2008), and are less diverse overall than the *A. halleri* populations we
studied, who colonised the north of France during the last century from ancestral German populations
(Pauwels et al., 2005). The progressive decoupling between alleles and their linked sites leads to the
simple prediction that *S*-locus homozygous genotypes formed by crossing individuals carrying identical
alleles from distinct populations should not reveal as much load as when they are formed by crossing
individuals within populations. Hence, the *S*-locus region could contribute to overall hybrid vigour.
Testing this prediction will be an interesting next step.

Different properties of the linked load according to S-allele dominance

The question of whether variations in the intensity of balancing selection, mediated by *S*-allele
dominance, could explain variation of the linked load has received conflicting support in the literature.
In line with Stift et al. (2013), but in contradiction with Llaurens et al. (2009a), we observed no overall
effect of dominance on the magnitude of the load. Several technical and biological reasons could
explain the contrasted results obtained in these different studies. First, phenotypic quantification of
the linked load is experimentally demanding, such that these studies relied on the comparison of a
limited number of alleles (three *S*-alleles in each of the studies) and therefore each of them had
inherently low power. Second, the experimental procedures to reveal the load varied slightly. Llaurens
et al. (2009a) used CO_2 treatment to by-pass the SI system and obtain homozygous progenies from
crosses that would otherwise have been incompatible, whereas we used the "natural" masking by
dominant *S*-alleles to enable the obtention of recessive homozygous genotypes. Our approach is
experimentally simpler and avoids the possible contamination by offspring obtained by selfing, which
may confound the effect of the sheltered load with that of genome-wide inbreeding depression (see
Stift et al., 2013 for a detailed discussion of this caveat). Third, a limitation of our approach is that it is
restricted to *S*-alleles that are recessive or intermediate along the dominance hierarchy, and is thus
not applicable to quantify the load associated with the most dominant *S*-alleles under more intense
balancing selection. It is therefore possible that the *S*-alleles we examined did not exhibit sufficiently
contrasted levels of dominance, in particular if only the most dominant ones are generating a
substantial load, as suggested for fully linked recessive deleterious mutations (Llaurens et al., 2009a).
In addition, since the homozygous *S*-allele genotypes we created correspond to different gene copies
from the population, they may carry distinct sets of linked variants, especially for the more recessive
S-alleles. The variation we observed in the phenotypic magnitude of the load among families confirms
that linked deleterious variants are unlikely to be fixed within all allele lineages. Finally, we note that
our genomic analysis of the genetic load shows that the dominant allele Ah15 previously associated
with reduced fitness in homozygotes (Llaurens et al., 2009a), is indeed unusual in terms of the number
of mutations it carries. In fact, it is one of the most "loaded" alleles among all the dominant *S*-alleles
present in this population, possibly explaining why Llaurens et al. (2009a) observed a significant effect
despite the inherently limited experimental power of their analysis.

A possible caveat of the population genetics approach we used is that simply counting up the number
of putatively deleterious linked mutations is a very crude estimate of the genetic load. We note that
our conclusions are robust to differences in the way we define deleterious mutations: as variants at
0-fold degenerate sites, at non-synonymous sites, or using methods to quantify the severity of
mutations such as SNPeff of SIFT4G. An obvious limitation is that none of these approaches allow for
evaluation of recessivity - a concept critical to ideas concerning the sheltered load. Our stochastic
simulations could be improved in several ways. First, we examined the accumulation of linked
deleterious mutations that were assumed to be fully recessive. This choice was guided by the
observation by Llaurens et al. (2009a) that fully recessive mutations accumulate more substantially,
but it remains possible that the dynamics of deleterious mutations that are only partially recessive
may involve a complex interaction with dominance of the S-alleles to which they are linked. Second,
allowing for partial recombination between S-alleles and their linked deleterious mutations in the
simulations would also be necessary to predict the length of the chromosomal haplotypes associated
with dominant vs. recessive *S*-alleles. In spite of these limitations, our stochastic simulations and
genomic analyses concur to the conclusion that the variation of the intensity of balancing selection
among *S*-alleles affect the genetic architecture of the linked load: a larger proportion of putatively

deleterious mutations are fixed among gene copies of the dominant as compared to the recessive *S*-
alleles, while gene copies of the recessive *S*-alleles tend to accumulate more segregating deleterious
variation. While these two processes eventually compensate one another, they may have distinct
consequences for the evolution of *S*-alleles. Uyenoyama (2003) showed that the existence of a
sheltered load should influence the evolutionary dynamics of new *S*-alleles through self-compatible
intermediates. Specifically, antagonistic interactions are expected between ancestral and derived
functional specificities because they would initially share their linked deleterious mutations, slowing
down the establishment of new *S*-alleles. Our observation that partially different sets of linked
mutations are associated with *S*-alleles from the different populations raises the question of whether
the (short) time scale at which recombination decouples *S*-alleles from their sets of linked mutation is
sufficiently fast to impede such antagonistic interactions to take place. In other words, the effect of
the load on the diversification dynamics should be most important if the two mutational steps
required for the emergence of new *S*-alleles under this model take place within local populations,
rather than involving a metapopulation-scale process. As shown by Stetsenko et al. (2023), this is
expected to occur under very low dispersal only. In addition, the observation that the architecture of
the sheltered load differs between dominant and recessive *S*-alleles suggests that their diversification
dynamics may also differ. Specifically, the self-compatible intermediates required for the formation of
new *S*-alleles (Gervais et al., 2011; Bergero and Charlesworth, 2009) are expected to be capable of
selfing as well as forming homozygous genotypes that would otherwise be prevented. While the
consequences of selfing may be equivalent for all alleles (because the overall number of mutations to
which they are linked are equivalent), the consequences of the formation of homozygotes allowed by
the crossing of separate individuals sharing a given *S*-allele are expected to be more severe for
dominant *S*-alleles. The segregation of distinct deleterious variants linked to different gene copies of
recessive *S*-alleles implies that linked recessive deleterious mutations are likely to remain masked
when two distinct gene copies of a given recessive *S*-allele are brought together. Hence, our results
lead to the prediction that in natural populations self-compatible mutants may segregate more readily
for the more recessive than for the more dominant *S*-alleles, and more generally for allelic lineages
under lower intensity of balancing selection. Considering that self-compatible mutants are a necessary
intermediate stage in the formation of new *S*-alleles, one may predict that the diversification dynamics
should be more efficient for lineages of recessive than dominant *S*-alleles. This prediction is in line
with the observation in Arabidopsis that the most dominant S-alleles exhibit the deepest phylogenetic
divergence among them (Durand et al., 2014). Detailed quantification of the presence of self-
compatible variants in natural populations will now be necessary to test this hypothesis. At this stage,
however, a proper model of allelic diversification taking into account dominance interactions among
S-alleles is still missing.

Variations of the genetic load among balanced allelic lines is a general phenomenon. The classical case
of Y or W sex chromosomes are indeed examples where one balanced line accumulates a greater
genetic load than the other (X or Z, respectively), eventually leading to substantial genetic
degeneration (Wright et al., 2016; Ponnikas et al., 2018). Another example is the supergene
controlling variation in male plumage phenotypes of the ruff, where the genetic load on the derived
"Satellite" haplotype is higher than on the ancestral "Independent" haplotype (Lamichhaney et al.,
2016; Hill et al., 2022). Similarly, in the butterfly *Heliconius numata*, the inverted haplotypes conferring
mimetic wing patterns tend to accumulate a greater load than the non-inverted haplotypes (Rosser et

al., 2022). Interestingly, in all these cases, the haplotypes with the greatest load also act genetically in
a dominant manner, establishing a clear parallel with our observations.

It is clear from our results that *S*-allele dominance affects the linked load, but in turn the differences
in structure of the linked load may affect the conditions under which dominance can evolve. The
Brassicaceae *S*-locus is a unique system, where dominance is controlled by "dominance modifiers"
(Durand et al., 2014; Durand et al. 2020). The presence of deleterious mutations linked to *S*-alleles has
been shown to affect the evolution of dominance modifiers, favouring evolution towards greater
dominance than towards greater recessivity (Llaurens et al., 2009b). This asymmetry arises from the
fact that *S*-alleles that become recessive (e.g. following acquisition of a recessivity modifier such as a
small RNA target) will start forming homozygous genotypes, leading to expression of their linked load,
while *S*-alleles that become dominant will not. Our observation that many deleterious mutations
linked to recessive *S*-alleles are indeed segregating, suggests that expression of the load will be less
severe for recessive than for dominant *S*-alleles, hence decreasing this predicted asymmetry. It will
now be essential to modify models for the evolution of dominance to allow for such differential load
among *S*-alleles.

Material and Methods

Source plant material

We worked on natural accessions from two closely related species, *A. halleri* and *A. lyrata*, represented by two population samples named Mortagne (50°47'N, 3°47'E, France, *n*=60) and Nivelle (50°47'N, 3°47'E, France, *n*=61) for *A. halleri*, and three highly outcrossing population samples from the North American Great Lakes, named IND (Indiana Dunes National Lakeshore in Michigan, *n*=9), PIN (Pinery Provincial Park in Ontario, *n*=11) and TSS (Tobermory Provincial Park in Ontario, *n*=8; Foxe et al., 2010) for *A. lyrata* (Fig. S9). The *A. lyrata* populations colonised North America from ancestral European populations about 20-30.000 years ago (Clauss and Mitchell-Olds, 2006; Ross-Ibarra et al., 2008) and the *A. halleri* populations are peripheral and likely colonised the north of France during the last century from ancestral German populations (Pauwels et al., 2005).

We performed 92, 91, 40, 43 and 21 controlled crosses between randomly chosen individuals within the Nivelle, Mortagne, IND, PIN and TSS populations, respectively. We successfully obtained seeds from 60, 66, 21, 21 and 10 of these crosses, respectively. Because we were not interested in estimating population frequencies of *S*-alleles, we instead tried to maximise the number of reconstructed haplotypes and avoid over representing the most recessive *S*-allele (Ah01) that tends to segregate at very high frequencies in natural populations (Llaurens et al., 2008). To do this, we performed PCR with *S*-allele-specific primers (Llaurens et al., 2008; Goubet et al., 2012) to screen the parents of the crosses and we removed from the experiment offspring with two parents carrying allele Ah01. For *A. halleri*, we selected 19 individuals from the Nivelle population and 19 individuals from the Mortagne population, based on their genotype at the *S*-locus (Fig. S9; Table S7). We also selected one offspring of 9, 11, 5, 6 and 5 pairs of selected individuals in the Nivelle, Mortagne, IND, PIN and TSS populations respectively for the phasing of *S*-haplotypes (Table S8; Fig. S9). To increase sample size for the phenotypic measurements, we included offspring from five additional crosses from the Nivelle population (Table S8).

Library preparation, capture and sequencing

We used a previously developed sequence capture approach to specifically sequence genomic regions of interest (Le Veve et al., 2023). Briefly, indexed genomic libraries were constructed for each individual and libraries were pooled in equimolar proportions. Fragments matching a series of regions of interest (including in particular the 75kb upstream and downstream of the non-recombining *S*-locus region as well as a series of 100 unlinked 25kb regions used as genomic controls; Le Veve et al., 2023), were then enriched using synthetic 120bp RNA probes and sequenced by Illumina MiSeq (a total of 159 million paired-end reads).

For six individuals (Table S7, S8), we completed the sequencing with genome-wide resequencing (WGS) in order to distinguish the homozygous and heterozygous genotypes at the *S*-locus based on read depth (Genete et al., 2020), which is not possible using data from the capture protocol. The prepared libraries were sequenced by Illumina NovaSeq (2x 150pb, paired-end) from the GenoScreen platform (Lille, France).

Determination of the S-locus genotypes and dominance of S-alleles

We used a dedicated pipeline for genotyping the *S*-locus based on short reads sequencing (Genete et al., 2020) obtained from each individual (Table S7 and S8). The level of dominance of *S*-alleles found in our study was determined based on either previous assessment of dominance in *A. lyrata* and *A. helleri* (Schierup et al., 2001; Mable et al., 2003; Bechsgaard,et al., 2004; Llaurens et al., 2008; Goubet et al., 2012) or indirectly inferred based on the observed association between the phylogeny of *S*-alleles and levels of dominance (Prigoda et al., 2005).

Read mapping and variant calling in A. halleri and A. lyrata populations

Raw reads were mapped on the complete *A. lyrata* reference genome V1.0.23 (Hu et al., 2011) using Bowtie2 v2.4.1 (Langmead and Salzberg, 2012), as described in Le Veve et al. (2023). File formats were then converted to BAM using samtools v1.3.1 (Li et al., 2009) and duplicated reads were removed with the MarkDuplicates program of picard-tools v1.119 (http://broadinstitute.github.io/picard). These steps were performed by the custom Python script sequencing_genome_vcf.py available at https://github.com/leveveaudrey/analysis-of-polymorphism-S-locus.

We obtained an average of 620 million properly mapped paired-end 300bp reads per population sample. For consistency, we conserved only reads which mapped to the *S*-locus flanking or control regions, even for samples sequenced by WGS, using the *targetintercept* option of bedtool v2.25.0 (Quinlan and Hall, 2010). We called all SNPs within the chromosomal segment comprising 50 kb upstream from the first base of the gene *Ubox* in 3' and 50 kb downstream from the last base of the gene *ARK3* in 5' of the *S*-locus using the Genome Analysis Toolkit v. 3.8 (GATK; DePristo et al., 2011) with the option GVCF and a quality score threshold of 60 using vcftool v0.1.15 (Danecek et al., 2011). This region contains 20 annotated protein-coding genes. In this study we excluded the genes inside the *S*-locus itself (*SCR*, *SRK*). For each sample independently, we computed the distribution of coverage depth across control regions using samtools depth (Li et al., 2009). We excluded sites with either less than 15 reads aligned or coverage depth above the 97.5 % percentile, as the latter are likely to correspond to repeated sequences (e.g. transposable elements or paralogs). Finally, we removed SNPs fixed in each population using the script 1_fix_pos_vcf.py (https://github.com/leveveaudrey/dominance_and_sheltered_load), thus retaining only nucleotide sites that were variable in the population.

Quantifying the sheltered load of deleterious mutations

We examined deleterious mutations based on the accumulation of either 1) mutations on 0-fold degenerate sites, 2) all non-synonymous mutations, 3) mutations predicted to be deleterious based on the SIFT4G database (Vaser et al,. 2016) or 4) mutations predicted to be lowly, moderately and highly deleterious by SNPeff (Cingolani et al,. 2012). The 0-fold and 4-fold degenerate sites were identified and extracted from the reference genome and the gene annotation using the script NewAnnotateRef.py (Williamson et al., 2014). None of the tools used to predict deleterious mutations are able to determine dominance levels of the mutation. Thus, all the deleterious mutations were considered as recessive. Details of the number of deleterious for each type is presented in Table S6.

Phasing S-haplotypes

For each of the 9, 11, 5, 6 and 5 trios analysed in the Nivelle, Mortagne, IND, PIN and TSS populations respectively, we phased mutations in the flanking regions, resulting in 130 phased haplotypes (Fig. S9). Briefly, we used sites that were heterozygous in the offspring to resolve parental haplotypes by assuming no recombination between parent and offspring, thus attributing the allelic state that was shared between a parent and its offspring to their shared *S*-allele, and the allelic state that was not shared to the other (untransmitted) haplotype of the parent. Twelve of the parents had been used in more than one cross, and in these cases we phased their haplotypes only once (Table S8). We implemented the phasing procedure in the script 3_phase_S_allele.py available at https://github.com/leveveaudrey/dominance_and_sheltered_load.

Study of the structure of S-haplotypes

We first developed a new method to evaluate the distortion of the phylogenetic patterns caused by linkage to *S*-alleles. To do this, we used phyml v.3.3 (Guindon et al., 2010) to calculate the likelihood

of two contrasted topologies of interest: (1) the topology clustering haplotypes by the populations
where they came from vs. (2) the topology clustering them by the *S*-allele to which they are linked (Fig
2A). We used sliding windows of sequences with 50 SNPs to obtain the variation of the difference in
log-likelihood between these two topologies along the chromosome. We then compared these values
to their distribution throughout the genome obtained by random draws of sequences with 50 SNPs
from the control regions. Second, we visualised the relationships among the phased haplotypes using
maximum likelihood phylogenies based on the Tamura-Nei model (Tamura and Nei, 1993), with 1,000
replicates in MEGA X (Kumar et al., 2018). Third, we followed Kamau et al.'s (2007) approach and
examined the variation of F_{ST} among populations within each species (Nivelle and Mortagne for *A.
halleri* and IND, PIN and TSS for *A. lyrata*) along the flanking region in non-overlapping windows of
5kb. We also examined the variation of F_{ST} along the flanking region obtained by grouping haplotypes
by their linked *S*-allele rather than by population of origin. Then, we compared these F_{ST} values
computed in the *S*-locus flanking regions with their genomic distribution as determined from the 100
control regions. The F_{ST} values were estimated with the DNAsp 6 software (Rozas et al., 2017). Fourth,
we performed a major component analysis (MCA) based on SNPs in the first 5kb, SNPs between 5 and
25kb and SNPs between 25 and 50kb around the *S*-locus, using the R packages 'ggplot2' (version 3.4.0),
'factoextra' (version 1.0.7) and 'FactoMiner' (version 2.7). We compared the patterns obtained by
these MCAs with those obtained from identical numbers of SNP (+/- 1%) from the control regions.
Finally, we analysed genetic association in each population independently between each of the locally
segregating variants and the *S*-alleles considered as phenotypes, using STRAT V1.1 (Pritchard et al.,
2000) combined with Structure V2.3 (Pritchard et al., 2010). We examined the distribution of the top
0.1% most significant associations detected specifically for each *S*-allele in each population.

Estimation of the number of fixed and segregating deleterious mutations within S-allele lineages

For each variable position considered in the phased haplotypes, we estimated the number of
mutations on 0-fold (S_{0f}) and 4-fold degenerate sites (S_{4f}) compared with the reference genome. We
distinguished SNPs that were fixed from those that were segregating within each of the allelic lines.
We used GLM with a Poisson distribution to test whether the number of fixed and segregating
mutations were associated with *S*-allele dominance, considering populations as random effects. The
dominance of the S-allele was considered a continuous variable. We reiterated the GLM analysis with
the number of non-synonymous (S_{NS}), synonymous (S_S), lowly and moderately deleterious mutations
predicted by SNPeff and deleterious mutations predicted by SIFT4G mutations.

Estimation of the phenotypic impact of homozygosity at the S-locus for three S-alleles

To determine if the genetic sheltered load putatively linked to the *S*-locus has a detectable phenotypic
impact, we performed 45 crosses (Table S1) between offspring of the Nivelle individuals that we chose
so that they shared one *S*-allele (Fig. S9). Based on the dominance hierarchy in pollen (Durand et al.,
2014; Table S1), these crosses should correspond to compatible partners. The general principle of the
experiment was to take advantage of the dominance hierarchy to mask recessive *S*-alleles and
generate full sibs that were either homozygous (because they inherited the *S*-allele that was shared
by their two parents) or heterozygous at the *S*-locus, and thus isolate the effect of homozygosity at
the *S*-locus. Note that all offspring in our experiments were thus "naturally" outcrossed, whereas
Llaurens et al. (2009a) based their comparisons on outcrossed progenies obtained by enforced
incompatible crosses and Stift et al.[19] based their comparisons on enforced selfed progenies. These
crosses generated 399 seeds overall, with homozygous genotypes expected for the *S*-alleles Ah01,
Ah03 and Ah04 forming the following dominance relationship: Ah01<Ah03<Ah04.

Seedlings were grown in a greenhouse between 14.5 and 23.1°C and a photoperiod of 16 hr day/8 hr
night. Offspring from the six families were placed on tables, and their position randomised every three
days. After three months of growing, all the germinated plants were vernalised under a temperature
between 6 and 8°C and a natural photoperiod for two months (January-February). Then, all surviving

plants began reproduction in a greenhouse under temperature between 10.6 and 25.3°C and a natural
photoperiod. The genotypes at the *S*-locus were determined in surviving plants by a PCR approach,
using *S*-allele-specific primers for the pistil-expressed *SRK* gene. We assessed the reproductive success
of offspring from the different crosses on the basis of fourteen phenotypic traits (detailed below) and
computed the mean difference for the trait between homozygotes and heterozygotes within each
family. We also tested for departures from mendelian proportions of each *S*-locus genotypic category
in the family after the apparition of the first stem. Significant departures were interpreted as reflecting
differences in survival between homozygous and heterozygous *S*-locus genotypes. We performed
10,000 replicate simulations of mendelian segregation based on the *S*-locus genotype of the parents.
We used GLM to test whether the phenotypic impact of homozygosity at the *S*-locus increased with
dominance of the *S*-alleles, considered as a continuous variable. The models used for GLM depended
on the type of trait analysed (poisson for the counts like the number of leaves, flowers by stems or
days; gaussian for continue traits like the lengths, widths and areas).

We measured the following fourteen phenotypic traits: the time (days) to the first leaf measured by
visual control every day during seven weeks after sowing the seeds, the number of leaves, the area of
the rosette (cm²), the mean length and width of leaves (cm), the standard deviation of length and
width of leaves (cm) and the mean area of leaves (cm^2) measured by ImageJ (Schneider et al., 2012)
based on photographs taken seven weeks (+/- five days) after the first leaf. At reproduction, we
measured the time to the first flower bud for the end of vernalisation (day), scored by visual control
every three days during nine weeks, the number of flower buds per flower stem produced during four
week after the appearance of the first bud, the number of flower stems, the length of the highest
flower stem produced four weeks after the appearance of the first bud (cm), and finally the total
duration of buds production (days), scored by visual control every three days during eleven weeks
after the appearance of the first bud. The last trait we measured was the proportion of homozygotes
per family that survived until reproduction assuming mendelian proportions in the seeds. During the
whole experiment, the presence of phytophagous insects, pathogens and stress markers were scored
as binary variables. The presence of phytophagous insects and pathogen attacks were detected by the
occurrence of gaps in leaves. Oxidative stress was scored qualitatively based on the occurrence of
purple leaves. We also controlled the effect of the family on the phenotypic trait. These effects were
controlled by redistributing 10,000 times the values observed in groups of the same size observed for
each effect (for example, presence or absence of pathogen attack) and comparing the difference for
the trait observed with the distribution of the differences obtained in the permutations. We
considered the impact of the effect on the trait if the observed difference between groups was higher
than the 95% percentile of the distribution obtained randomly (Table S9). When the test was
significant, the effect was implemented as a random effect in the GLM. We used the same method to
control for the family effect, which was included as a random effect in GLM if necessary (Table S10).
The general experimental procedure is summarised in Fig. S9 and all data analyses were done in R ver.
3.1.2 (R Development Core Team 2014).

Simulations

Finally, we refined the model of Llaurens et al. (2009a), in several ways. We simulated a panmictic
population of *N* diploid individuals with non-overlapping generations. Each individual was defined by
its genotype in a non-recombining genomic region. This region contains the *S*- locus, and a D locus
where deleterious mutations accumulated. For the *S*-locus, we used a simple model of sporophytic SI,
with four dominance classes, as observed in *A. halleri* (Genete et al., 2020; only three classes were
considered before in Llaurens et al., 2009a), and fourteen *S*-alleles (eight alleles in the class IV, three
in the class III, two in the class II and one allele in the class I). This distribution mirrors that of the
Nivelle population (Table S7), with the exception that a class II allele has been added because its
presence has been reported in previous studies (Llaurens et al., 2008). Alleles within classes were
assumed to be codominant with each other, and dominant over all alleles of the more recessive
classes, with the following linear hierarchy between classes : classI<classII<classIII<classIV). We also

assumed that no new *S*-allele could appear by mutation during the simulations. The population size
was 10,000 diploid individuals, so as to be large enough to avoid *S*-allele loss by drift during the
simulations (previously it was 1,000). The "D locus" comprised one hundred fully linked biallelic
positions (versus a single one in Llaurens et al., 2009a). Fully recessive deleterious mutations were
recurrently introduced (at a rate 10^{-4}), and reverse mutations were possible (at a rate 10^{-5}). We ignored
partially recessive deleterious mutations because these mutations were predicted to be effectively
eliminated by natural selection in Llaurens et al. (2009a). The survival probability *p* of a zygote
depended on its genotype at the D locus: $p = (1 - s)^n$ with *s* the selection coefficient and *n* the number
of positions homozygous for the mutated allele. We explored different values of the selection
coefficient (0.1, 0.05, 0.03, 0.01 and 0.005). Under strong selection (s=0.1, 0.05 and 0.03), the
combined effect of multiple mutations led to low-fitness individuals, eventually causing population
extinction. Under weak selection, (s=0.005), we observed near fixation of the deleterious mutations
under the influence of asymmetrical mutation. Hence, we focused on the intermediate value of the
selection coefficient (s=0.01), where deleterious mutations segregated stably in the simulations.

We first ran simulations without deleterious mutations until a deterministic equilibrium for *S*-allele
frequencies was reached, which was considered to be attained when allelic frequencies changed by
less than 10^{-3} between generations. Recessive deleterious mutations were then allowed to accumulate
at the positions within the D locus. Each simulation was performed with 100 independent replicates
of 100,000 generations, and the frequency of the deleterious alleles was recorded every 1,000
generations. At the end of the simulation runs, we estimated the number of deleterious mutations
found in each haplotype associated with each *S*-allele to determine the expected patterns of
association between the sheltered load and dominance at the *S*-locus.

The code of the program of simulations developed in Llaurens et al. (2009a) and used in our study is
available in Github (https://github.com/leveveaudrey/model_ssi_Llaurens).

Data Availability

All sequence data are available in the NCBI Short Read Archive (SRA;
https://www.ncbi.nlm.nih.gov/sra) with accession codes: PRJNA744343, PRJNA755829.

All scripts developed are available in Github
(https://github.com/leveveaudrey/dominance_and_sheltered_load
https://github.com/leveveaudrey/analysis-of-polymorphism-S-locus).

The code of the program of simulations developed in Llaurens et al. (2009a) and used in our study is
available in Github (https://github.com/leveveaudrey/model_ssi_Llaurens).

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Author contributions

728 AL, ED, VC and XV developed and designed the experiments for the study. AL, CBL and CP performed
729 the experiments. AL and MG analysed and interpreted the data. AL, ED, VC and XV wrote the
730 manuscript. All authors edited the manuscript.

Declaration of interests

732 The authors declare no competing interests.

References

Bechsgaard J, Bataillon T, and Schierup M.H (2004) Uneven segregation of sporophytic self-
incompatibility alleles in *Arabidopsis lyrata*. Journal of Evolutionary Biology 17, 554–561.
https://doi.org/10.1111/j.1420-9101.2004.00699.x.

Bergero R, and Charlesworth D (2009) The evolution of restricted recombination in sex chromosomes.
Trends in Ecology & Evolution 24, 94–102. https://doi.org/10.1016/j.tree.2008.09.010.

Billiard S, Castric V, and Vekemans X (2007) A general model to explore complex dominance patterns
in plant sporophytic self-incompatibility systems. Genetics 175, 1351–1369.
https://doi.org/10.1534/genetics.105.055095

Castric V, and Vekemans X (2004) Plant self-incompatibility in natural populations: a critical
assessment of recent theoretical and empirical advances. Molecular Ecology 13, 2873–
2889.https://doi.org/10.1111/j.1365-294X.2004.02267.x.

Charlesworth D (2006) Balancing selection and its effects on sequences in nearby genome regions.
PLOS Genetics 2, e64. https://doi.org/10.1371/journal.pgen.0020064.

Cingolani, P., Platts, A., Wang, L.L., Coon, M., Nguyen, T., Wang, L., Land, S.J., Lu, X., Ruden, D.M., 2012.
A program for annotating and predicting the effects of single nucleotide polymorphisms, SnpEff. Fly
(Austin) 6, 80–92. https://doi.org/10.4161/fly.19695

Clauss M.J, and Mitchell-Olds T (2006) Population genetic structure of *Arabidopsis lyrata* in Europe.
Molecular Ecology 15, 2753–2766. https://doi.org/10.1111/j.1365-294X.2006.02973.x.

Cocker J.M, Wright J, Li J, Swarbreck D, Dyer S, Caccamo M, and Gilmartin P.M (2018) *Primula vulgaris*
(primrose) genome assembly, annotation and gene expression, with comparative genomics on the
heterostyly supergene. Sci Rep 8, 17942. https://doi.org/10.1038/s41598-018-36304-4.

Danecek P, Auton A, Abecasis G, Albers C.A, Banks E, DePristo M.A, Handsaker R.E, Lunter G, Marth
G.T, Sherry S.T, et al (2011) The variant call format and VCFtools. Bioinformatics 27, 2156–2158.
https://doi.org/10.1093/bioinformatics/btr330.

DePristo M.A, Banks E, Poplin R.E, Garimella K.V, Maguire J.R, Hartl C, Philippakis A.A, del Angel G,
Rivas M.A, Hanna M, et al (2011) A framework for variation discovery and genotyping using next-
generation DNA sequencing data. Nat Genet 43, 491–498. https://doi.org/10.1038/ng.806.

Durand E, Méheust R, Soucaze M, Goubet P.M, Gallina S, Poux C, Fobis-Loisy I, Guillon E, Gaude T,
Sarazin A, et al (2014) Dominance hierarchy arising from the evolution of a complex small RNA
regulatory network. Science 346, 1200–1205. https://doi.org/10.1111/eva.12933.

Durand E, Chantreau M, Le Veve A, et al (2020) Evolution of self-incompatibility in the Brassicaceae: Lessons from a textbook example of natural selection. Evolutionary Applications. https://doi.org/10.1111/eva.12933.

Foxe J.P, Stift M, Tedder A, Haudry A, Wright S.I, and Mable B.K (2010) Reconstructing origins of loss of self-incompatibility and selfing in North american *Arabidopsis lyrata*: a population genetic context. Evolution 64, 3495–3510. https://doi.org/10.1111/j.1558-5646.2010.01094.x.

Genete M, Castric V, and Vekemans X (2020) Genotyping and *de novo* discovery of allelic variants at the Brassicaceae self-incompatibility locus from short read sequencing data. Mol Biol Evol. https://doi.org/10.1093/molbev/msz258.

Gervais C.E, Castric V, Ressayre A, and Billiard S (2011) Origin and diversification dynamics of self-Incompatibility haplotypes. Genetics 188, 625–636. https://doi.org/10.1534/genetics.111.127399.

Goubet P.M, Bergès H, Bellec A, Prat E, Helmstetter N, Mangenot S, Gallina S, Holl A.-C, Fobis-Loisy I, Vekemans X, et al (2012) Contrasted patterns of molecular evolution in dominant and recessive self-incompatibility haplotypes in Arabidopsis. PLoS Genetics 8, e1002495. https://doi.org/10.1371/journal.pgen.1002495.

Guindon S, Dufayard J-F, Lefort V, et al (2010) New algorithms and methods to estimate maximum-likelihood phylogenies: assessing the performance of PhyML 3.0. Syst Biol 59, 307–321. https://doi.org/10.1093/sysbio/syq010.

Hartmann F.E, Duhamel M, Carpentier F, et al (2021) Recombination suppression and evolutionary strata around mating-type loci in fungi: documenting patterns and understanding evolutionary and mechanistic causes. New Phytologist 229:2470–2491. https://doi.org/10.1111/nph.17039.

Hatakeyama K, Watanabe M, Takasaki T, Ojima K, Hinata K. (1998). Dominance relationships between S -alleles in self-incompatible Brassica campestris L. Heredity 80, 241–247. https://doi.org/10.1046/j.1365-2540.1998.00295.x

Hill J, Enbody E, Bi H, Lamichhaney S, Schwochow D, Younis S, Widemo F, and Andersson L (2022) Low mutation load in a supergene underpinning alternative male mating strategies in ruff. https://doi.org/10.1101/2022.04.27.489720.

Hu T.T, Pattyn P, Bakker E.G, Cao J, Cheng J.-F, Clark R.M, Fahlgren N, Fawcett J.A, Grimwood J, Gundlach H, et al (2011) The *Arabidopsis lyrata* genome sequence and the basis of rapid genome size change. Nat Genet 43, 476–481. https://doi.org/10.1038/ng.807.

Jay P, Chouteau M, Whibley A, Bastide H, Parrinello H, Llaurens V, and Joron M (2021) Mutation load at a mimicry supergene sheds new light on the evolution of inversion polymorphisms. Nat Genet 53, 288–293. https://doi.org/10.1038/s41588-020-00771-1.

Jay P, Tezenas E, Véber A, and Giraud T (2022) Sheltering of deleterious mutations explains the stepwise extension of recombination suppression on sex chromosomes and other supergenes. PLOS Biology 20, e3001698. https://doi.org/10.1371/journal.pbio.3001698

Kakizaki T, Takada Y, Ito A, Suzuki G, Shiba H, Takayama S, Isogai A, Watanabe M. (2003). Linear
Dominance Relationship among Four Class-II S Haplotypes in Pollen is Determined by the Expression
of SP11 in Brassica Self-Incompatibility. Plant and Cell Physiology 44, 70–75.
https://doi.org/10.1093/pcp/pcg009

Kamau E, Charlesworth B, Charlesworth D (2007) Linkage Disequilibrium and Recombination Rate
Estimates in the Self-Incompatibility Region of Arabidopsis lyrata. Genetics 176, 2357–2369.
https://doi.org/10.1534/genetics.107.072231

Katsura Y, Iwase M, Satta Y (2012) Evolution of genomic structures on mammalian sex chromosomes.
Current Genomics 13:115–123. https://doi.org/10.2174/138920212799860625.

Kumar S, Stecher G, Li M, Knyaz C, and Tamura K (2018) MEGA X: Molecular evolutionary genetics
analysis across computing platforms. Molecular Biology and Evolution 35:1547-1549.
https://doi.org/10.1093/molbev/msy096.

Kusaba M, Dwyer K, Hendershot J, Vrebalov J, Nasrallah J.B, and Nasrallah M.E (2001) Self-
incompatibility in the genus Arabidopsis: characterization of the S locus in the outcrossing *A. lyrata*
and its autogamous relative *A. thaliana*. Plant Cell 13, 627–643.

Lamichhaney S, Fan G, Widemo F, Gunnarsson U, Thalmann D.S, Hoeppner M.P, Kerje S, Gustafson U,
Shi C, Zhang H, Chen W, Liang X, Huang L, Wang J, Liang E, Wu Q, Lee S.M.-Y, Xu X, Höglund J, Liu X,
and Andersson L (2016) Structural genomic changes underlie alternative reproductive strategies in the
ruff (*Philomachus pugnax*). Nat Genet 48, 84–88. https://doi.org/10.1038/ng.3430.

Lane M.D, and Lawrence M.J (1995) The population genetics of the self-incompatibility polymorphism
in *Papaver rhoeas*. X. An association between incompatibility genotype and seed dormancy. Heredity
75, 92–97. https://doi.org/10.1038/hdy.1995.108.

Langmead B, and Salzberg S.L (2012) Fast gapped-read alignment with Bowtie 2. Nat Methods 9, 357–
359. https://doi.org/10.1038/nmeth.1923.

Le Veve A, Burghgraeve N, Genete M, et al (2023) Long-term balancing selection and the genetic load
linked to the self-incompatibility locus in *Arabidopsis halleri* and *A. lyrata*. Molecular Biology and
Evolution msad120. https://doi.org/10.1093/molbev/msad120.

Li H, Handsaker B, Wysoker A, Fennell T, Ruan J, Homer N, Marth G, Abecasis G, Durbin R, and 1000
Genome Project Data Processing Subgroup (2009) The Sequence Alignment/Map format and
SAMtools. Bioinformatics 25, 2078–2079. https://doi.org/10.1093/bioinformatics/btp352.

Llaurens V, Billiard S, Leducq J.-B, Castric V, Klein E.K, and Vekemans X (2008) Does frequency-
dependent selection with complex dominance interactions accurately predict allelic frequencies at the
self-incompatibility locus in *Arabidopsis halleri* ? Evolution 62, 2545–2557.
https://doi.org/10.1111/j.1558-5646.2008.00469.x.

Llaurens V, Gonthier L, and Billiard S (2009a) The sheltered genetic load linked to the S locus in plants:
new insights from theoretical and empirical approaches in sporophytic self-incompatibility. Genetics
183, 1105–1118. https://doi.org/10.1534/genetics.109.102707.

Llaurens V, Billiard S, Castric V, and Vekemans X (2009b) Evolution of dominance in sporophytic self-
incompatibility systems: I. Genetic load and coevolution of levels of dominance in pollen and pistil.
Evolution 63, 2427–2437. https://doi.org/10.1111/j.1558-5646.2009.00709.x.

Mable B.K, Schierup M.H, and Charlesworth D (2003) Estimating the number, frequency, and
dominance of S -alleles in a natural population of *Arabidopsis lyrata* (Brassicaceae) with sporophytic
control of self-incompatibility. Heredity 90, 422–431. https://doi.org/10.1038/sj.hdy.6800261.

Mattila T.M, Laenen B, Horvath R, Hämälä T, Savolainen O, and Slotte T. (2019) Impact of demography
on linked selection in two outcrossing Brassicaceae species. Ecology and Evolution 9, 9532–9545.
https://doi.org/10.1002/ece3.5463

Mena-Ali J. I, Keser L. H, Stephenson A. G (2009) The effect of sheltered load on reproduction in
Solanum carolinense, a species with variable self-incompatibility. Sex. Plant Reprod. 22 63–67.
https://doi.org/10.1007/s00497-008-0092-x.

Nettancourt D (2001) Incompatibility and incongruity in wild and cultivated plants. (2nd ed. Springer-
Verlag, Berlin Heidelberg).

Pauwels M, Saumitou-Laprade P, Holl A.C, Petit D, and Bonnin I (2005) Multiple origin of metallicolous
populations of the pseudometallophyte *Arabidopsis halleri* (Brassicaceae) in central Europe: the
cpDNA testimony. Molecular Ecology 14, 4403–4414. https://doi.org/10.1111/j.1365-
294X.2005.02739.x.

Ponnikas S, Sigeman H, Abbott J.K, and Hansson B (2018) Why do sex chromosomes stop recombining?
Trends Genet 34, 492–503. https://doi.org/10.1016/j.tig.2018.04.001.

Prigoda N.L, Nassuth A, and Mable B.K (2005) Phenotypic and genotypic expression of self-
incompatibility haplotypes in *Arabidopsis lyrata* suggests unique origin of alleles in different
dominance classes. Molecular Biology and Evolution 22, 1609–1620.
https://doi.org/10.1093/molbev/msi153.

Pritchard J. K, Stephens M, Rosenberg N. A, and Donnelly P (2000) Association mapping in structured
populations. The American Journal of Human Genetics 67:170–181. https://doi.org/10.1086/302959.

Pritchard J. K, Wen X, & Falush D (2010) Documentation for structure software: Version 2.3. University
of Chicago, Chicago, IL, 1-37.

Quinlan A.R, and Hall I.M (2010) BEDTools: a flexible suite of utilities for comparing genomic features.
Bioinformatics 26, 841–842. https://doi.org/10.1093/bioinformatics/btq033.

Ross-Ibarra J, Wright S.I, Foxe J.P, Kawabe A, DeRose-Wilson L, Gos G, Charlesworth D, and Gaut B.S
(2008) Patterns of polymorphism and demographic history in natural populations of *Arabidopsis
lyrata*. PLOS ONE 3, e2411. https://doi.org/10.1371/journal.pone.0002411.

Rosser N, Edelman N.B, Queste L.M, et al (2022) Complex basis of hybrid female sterility and Haldane's
rule in Heliconius butterflies: Z-linkage and epistasis. Molecular Ecology 31:959–977.
https://doi.org/10.1111/mec.16272.

Rozas J, Ferrer-Mata A, Sánchez-DelBarrio J.C, Guirao-Rico S, Librado P, Ramos-Onsins S.E, and
Sánchez-Gracia A (2017) DnaSP 6: DNA sequence polymorphism analysis of large data sets. Molecular
Biology and Evolution 34, 3299–3302. https://doi.org/10.1093/molbev/msx248.

Schierup M.H, Mikkelsen A.M, and Hein J (2001) Recombination, balancing selection and phylogenies
in *MHC* and self-incompatibility genes. Genetics 159, 1833–1844.
https://doi.org/10.1093/genetics/159.4.1833

Schneider C.A, Rasband W.S, and Eliceiri K.W (2012) NIH Image to ImageJ: 25 years of image analysis.
Nat Methods 9, 671–675. https://doi.org/10.1038/nmeth.2089.

Schopfer C.R, Nasrallah M.E, and Nasrallah J.B (1999) The male determinant of self-incompatibility in
Brassica. Science 286, 1697–1700. https://doi.org/10.1126/science.286.5445.1697.

Stetsenko R, Brom T, Castric V, and Billiard S (2023) Balancing selection and the crossing of fitness
valleys in structured populations: diversification in the gametophytic self-incompatibility system.
Evolution 77:907–920. https://doi.org/10.1093/evolut/qpac065

Stift M, Hunter B.D, Shaw B, Adam A, Hoebe P.N, and Mable B.K (2013) Inbreeding depression in self-
incompatible North-American *Arabidopsis lyrata*: disentangling genomic and S-locus-specific genetic
load. Heredity 110, 19–28. https://doi.org/10.1038/hdy.2012.49.

Stone J.L (2004) Sheltered load associated with S-alleles in *Solanum carolinense*. Heredity 92, 335–
342. https://doi.org/10.1038/sj.hdy.6800425.

Tamura K, and Nei M (1993) Estimation of the number of nucleotide substitutions in the control region
of mitochondrial DNA in humans and chimpanzees. Molecular Biology and Evolution 10, 512–526.
https://doi.org/10.1093/oxfordjournals.molbev.a040023.

Uyenoyama M.K (1997) Genealogical structure among alleles regulating self-incompatibility in natural
populations of flowering plants. Genetics 147, 1389–1400.
https://doi.org/10.1093/genetics/147.3.1389.

Uyenoyama M.K (2003) Genealogy-dependent variation in viability among self-incompatibility
genotypes. Theoretical Population Biology 63, 281–293. https://doi.org/10.1016/S0040-
5809(03)00020-0.

Uyenoyama M.K (2005) Evolution under tight linkage to mating type. New Phytol. 165: 63–70.
https://doi.org/10.1111/j.1469-8137.2004.01246.x.

Vaser R, Adusumalli S, Leng S.N, Sikic M, Ng P.C. (2016). SIFT missense predictions for genomes. Nat
Protoc 11, 1–9. https://doi.org/10.1038/nprot.2015.123

Vieira J, Pimenta J, Gomes A, Laia J, Rocha S, Heitzler P, and Vieira C.P (2021) The identification of the
Rosa S-locus and implications on the evolution of the Rosaceae gametophytic self-incompatibility
systems. Sci Rep 11, 3710. https://doi.org/10.1038/s41598-021-83243-8.

Williamson R.J, Josephs E.B, Platts A.E, Hazzouri K.M, Haudry A, Blanchette M, and Wright S.I (2014)
Evidence for widespread positive and negative selection in coding and conserved noncoding regions
of *Capsella grandiflora*. PLOS Genetics 10, e1004622. https://doi.org/10.1371/journal.pgen.1004622.

Wright A.E, Dean R, Zimmer F, and Mank J.E (2016) How to make a sex chromosome. Nat Commun 7,
12087. https://doi.org/10.1038/ncomms12087.

Yasuda S, Wada Y, Kakizaki T, Tarutani Y, Miura-Uno E, Murase K, Fujii S, Hioki T, Shimoda T, Takada Y,
Shiba H, Takasaki-Yasuda T, Suzuki G, Watanabe M, Takayama S. (2017). A complex dominance
hierarchy is controlled by polymorphism of small RNAs and their targets. Nature Plants 3.
https://doi.org/10.1038/nplants.2016.206

 Figures legend

 Figure 1: Effect of homozygosity at the S-locus on 13 phenotypic traits compared to heterozygotes.
 For each trait, the phenotypic values in homozygotes (in grey) were normalised relative to the mean
 phenotypic values in heterozygotes (in black). The distributions were obtained by 10,000 random
 permutations. SD: Standard Deviation.

 Figure 2: Linkage to the S-locus locally distorts the phylogenetic relationships. *A. The two topologies*
 of interest cluster haplotypes either by the S-allele to which they are linked (top) or by the populations
 where they came (bottom). Different S-alleles are represented by symbols of different colours, different
 populations of origin are represented by symbols of different shapes. B. Difference in log likelihood of
 the two topologies of interest. Dots correspond to the difference in log likelihood for overlapping series
 of 50 SNPs around the S-locus for A. halleri (top panel) and A. lyrata (bottom panel). Positive values
 correspond to chromosomal positions where the topology by S-alleles explains the phylogeny of
 haplotypes better than the topology by populations. The right panels show the difference in log
 likelihood in the control regions. 2.5 and 97.5 percentiles of the distribution in the control regions are
 indicated by dashed lines.

 Figure 3: No overall effect of S-allele dominance on the total number of 0-fold degenerate mutations
 (S_{0f}) in the linked genomic regions within 25kb. *Each dot represents the mean number of mutations*
 observed among haplotypes linked to one S-allele in one population. The correlations evaluated by a
 GLM are represented by lines, with confidence intervals represented in grey. The dominance was
 considered a continuous variable. A: A. halleri. Black dots correspond to the Nivelle population, grey
 dots to the Mortagne population. The red arrow points to the copy of Ah15, corresponding to the S-
 allele whose sheltered load was phenotypically characterised by Llaurens et al. (2009a) in Nivelle. B: A.
 lyrata. Red dots correspond to the IND population, black dots to the PIN population and blue dots to
 the TSS population.

 Figure 4: The number of 0-fold degenerate mutations fixed in the 25kb regions flanking the S-locus
 increases with dominance of the S-allele associated. *Each dot represents the value obtained for*
 haplotypes linked to one S-allele in one population. The correlations evaluated by a GLM are
 represented by lines, with confidence intervals represented in grey. The dominance was considered as
 a continuous variable. A: A. halleri. Black dots correspond to the Nivelle population, grey dots to the
 Mortagne population. B: A. lyrata. Red dots correspond to the IND population, black dots to the PIN
 population, blue dots to the TSS population.

 Figure 5: Stochastic simulations confirm the contrasted architecture of the load of deleterious
 mutations linked to dominant vs. recessive S-alleles. *Number of fixed (A), segregating (B) and total*
 (C) deleterious mutations linked to S-alleles at four different levels of dominance (I<II<III<IV). The
 means (bold lines) were estimated per S-allele dominance classes over 100 replicate simulations after
 discarding an initial burn-in of 100,000 generations. h=0. s=0.01.

955 ***Table 1: Proportion of S-locus homozygous offspring having reached the reproductive stage for three***
956 ***different S-alleles.*** *The test is performed relative to the expected proportion of homozygous genotypes*
957 *in the offspring (25% when both parents are heterozygous; 50% when one of the parents is*
958 *homozygous and the other heterozygous).*

S-allele	Level of dominance	Number of seedlings having reached the reproductive stage	Observed proportion of homozygotes	Ratio of the observed to expected proportion of homozygous genotypes (p-value[a])	Number of heterozygotes with the S-allele (p-values[a])
Ah01	I	35	0.29	1.14 (0.76)	19 (0.40; 0.90)
Ah03	II	27	**0.074**	**0.3 (0.02)**	17 (1; 0.69)
Ah04	III	96	0.479	0.96 (0.39)	50

959 [a] *p-values were obtained after 10,000 random permutations, respectively. Values departing from*
960 *Mendelian expectations are figured in bold. For Ah04, the maternal parent was Ah04/Ah04, so all*
961 *heterozygous offspring carried the S-allele.*

Supplementary data

Figures legend

Figure S1: Analysis of Major Components obtained for haplotypes of A. halleri of the Nivelle (black dots) and Mortagne (grey dots) populations) based on SNPs in the first 5kb, between 5 and 25kb, and between 25kb and 50kb away from the S-locus. The S-allele to which the SNPs are linked are represented by different symbols. The right panels show the analysis on control regions, each time matching the number of SNPs with that of the corresponding linked regions (left panels).

Figure S2: Analysis of major components (AMC) obtained for haplotypes of A. lyrata (of the PIN: grey dots, IND: red dots and TSS: blue dots) populations based on the SNPs in the first 5kb, between 5 and 25kb, and between 25kb and 50kb away from the S-locus. The S-allele to which the SNPs are linked are represented by different symbols. The right panels show the analysis on control regions, each time matching the number of SNPs with that of the corresponding linked regions (left panels).

Figure S3: Phylogenetic tree obtained by Maximum Likelihood for haplotypes of A. halleri (populations Nivelle and Mortagne) across the first 25kb flanking the S-locus. The Tamura-Nei model was used and the percentage of trees in which the associated haplotypes clustered together is shown next to the branches. The tree is drawn to scale, with branch lengths measured in the number of substitutions per site. The black braces indicate haplotypes clustering by populations. The overall pattern shows a structure by S-alleles, with exceptions highlighted in yellow.

Figure S4: Phylogenetic tree obtained by maximum likelihood for haplotypes of A. halleri (populations Nivelle and Mortagne) based on the nucleotide positions between 25kb and 50kb away from the S-locus. The Tamura-Nei model was used and the percentage of trees in which the associated haplotypes clustered together is shown next to the branches. The tree is drawn to scale, with branch lengths measured in the number of substitutions per site. The black braces indicate haplotypes clustering by populations. The overall pattern shows a structure by S-alleles, with exceptions highlighted in yellow.

Figure S5: Phylogenetic tree obtained by maximum likelihood for haplotypes of A. lyrata (populations PIN, IND, TSS) across the first 5kb flanking the S-locus. The Tamura-Nei model was used and the percentage of trees in which the associated haplotypes clustered together is shown next to the branches. The tree is drawn to scale, with branch lengths measured in the number of substitutions per site. The black braces indicate haplotypes clustering by populations. The overall pattern shows a structure by S-alleles, with exceptions highlighted in yellow.

Figure S6: Phylogenetic tree obtained by maximum likelihood for haplotypes of A. lyrata (populations PIN, IND, TSS) based on the nucleotide positions between 5kb and 10kb away from the S-locus. The Tamura-Nei model was used and the percentage of trees in which the associated haplotypes clustered together is shown next to the branches. The tree is drawn to scale, with branch lengths measured in the number of substitutions per site. The black braces indicate haplotypes clustering by populations. The overall pattern shows a structure by S-alleles, with exceptions highlighted in yellow.

Figure S7: The genetic structure of SNPs in the S-locus flanking regions in A. halleri and A. lyrata.
Left: Mean F_{ST} (lines) and F_{ST} by pair (point) analysed among S-alleles (grey) or among populations
(black) in 5 kb windows in the S-locus flanking regions. Right: Distribution (count) of F_{ST} in the control
regions analysed among S-alleles (grey bars) or among populations (black bars). The 95% percentiles
of the distributions are represented by dotted lines and the medians by solid lines.

Figure S8: Patterns of genetic associations between S-alleles and SNPs across the genome. *Each dot*
corresponds to a SNP showing statistically significant association (top 0.1%) with a given S-allele. The
left panel confirms that SNPs physically linked to the S-locus (in red) are considerably more likely to
show statistical associations with S-alleles. The chromosomes were signified by alternance of black and
grey. The middle panel shows a zoom on the 50kb regions flanking the S-locus and shows that the
statistical association extends over long distances, at least for some S-alleles. Each region of 25 kb was
delimited by vertical lines. The right panel shows that the observed number of associated SNPs in the
linked regions far exceeds that in regions of identical size from control regions. The histogram shows
the distribution across control regions of the mean number of significantly associated SNPs per S-allele.
The vertical lines correspond to the mean number of significantly associated SNPs in the first 25kb
(solid lines) and the 25-50kb interval (dashed lines) away from S-locus.).

Figure S9: Experimental protocol. *A) We randomly crossed A. lyrata individuals from the PIN, TSS and*
IND populations in North America (left) and A. halleri from the Nivelle (middle) and Mortagne (right)
populations. Individuals were sequenced by a capture protocol. Numbers between parentheses
represent the number of individuals per dataset. B) One offspring from each cross was sequenced along
with its two parents for trio haplotyping. Offspring from the Nivelle population (black circle) were
conserved for the study of the impact of homozygosity at the S-locus on fitness (G1 population).
C)Individuals sequenced in A and B were used to reconstruct haplotypes linked to each copy of S-allele,
assuming no recombination between the S-locus and its flanking regions between parents and
offspring. D) We used the dominance hierarchy between S-alleles expressed in pollen to cross G1
individuals from the Nivelle population and obtained six G2 families constituted of heterozygous and
homozygous individuals for the alleles Ah01 , Ah03 and Ah04 . E) Description of the traits measured,
and the methods used to estimate the impact of homozygosity at the S-locus in homozygotes. Traits 1-
8 are related to biomass and traits 10-14 are related to reproductive success.

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1049 *Figure S1: Analysis of Major Components obtained for haplotypes of A. halleri of the Nivelle (black*

1050 *dots) and Mortagne (grey dots) populations) based on SNPs in the first 5kb, between 5 and 25kb,*

1051 *and between 25kb and 50kb away from the S-locus.*

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1053 *Figure S2: Analysis of major components (AMC) obtained for haplotypes of A. lyrata (of the PIN: grey*
1054 *dots, IND: red dots and TSS: blue dots) populations based on the SNPs in the first 5kb, between 5*
1055 *and 25kb, and between 25kb and 50kb away from the S-locus.*

1056

1057 *Figure S3: Phylogenetic tree obtained by Maximum Likelihood for haplotypes of A. halleri*
1058 *(populations Nivelle and Mortagne) across the first 25kb flanking the S-locus.*

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Figure S4: Phylogenetic tree obtained by maximum likelihood for haplotypes of A. halleri (populations Nivelle and Mortagne) based on the nucleotide positions between 25kb and 50kb away from the S-locus.

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Figure S5: Phylogenetic tree obtained by maximum likelihood for haplotypes of A. lyrata (populations PIN, IND, TSS) across the first 5kb flanking the S-locus.

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1068 *Figure S6: Phylogenetic tree obtained by maximum likelihood for haplotypes of A. lyrata*
1069 *(populations PIN, IND, TSS) based on the nucleotide positions between 5kb and 10kb away from the*
1070 *S-locus.*

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1072   *Figure S7: The genetic structure of SNPs in the S-locus flanking regions in A. halleri and A. lyrata.*

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1074   *Figure S8: Patterns of genetic associations between S-alleles and SNPs across the genome.*

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1076   *Figure S9: Experimental protocol.*

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Supplemental items

Table S1: Crosses performed to obtain homozygotes for three S-alleles.

Table S2: Trait variation in S-locus homozygous individuals.

Table S3: Trait variation in homozygous at the S-locus for the S-alleles Ah01 and Ah04 between families.

Table S4: Effect of dominance on variation of phenotypic traits in S-locus homozygous individuals.

Table S5: Number of phased haplotypes linked to S-alleles in each sample.

Table S6: Effect of dominance on the accumulation of genetic load in S-flanking regions.

Table S7: S-locus genotypes of individuals sequenced using the capture protocol.

Table S8: S-locus genotypes of the offspring selected for haplotype phasing and the crosses for the study of phenotypic traits.

Table S9: Effect of phytopathogen, phytophagous attacks and oxidative stress on the phenotypic traits.

Table S10: Difference on the phenotypic traits variation between the two families for each allele tested.

1098 ***Table S1: Crosses performed to obtain homozygotes for three S-alleles.*** *For each S-allele, we*
1099 *obtained homozygous and heterozygous plants within two separate families (listed as separate*
1100 *rows).*

Genotype pollen donor [a]	Genotype stigma recipient	Homozygote studied	Dominance level	Number of fruits obtained	Number of seeds
Ah20/Ah01	Ah12/Ah01	Ah01	1	5	48
Ah20/Ah01	Ah12/Ah01	Ah01	1	2	23
Ah02/Ah03	Ah03/Ah01	Ah03	2	6	69
Ah24/Ah03	Ah03/Ah01	Ah03	2	3	27
Ah20/Ah04	Ah04/Ah04	Ah04	3	21	167
Ah20/Ah04	Ah04/Ah04	Ah04	3	8	65

1101 *[a] The S-alleles in bold represent the dominant S-allele expressed on pollen of each donor genotype.*
1102 *Each line corresponds to one pair of individuals used as parents.*

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1106 ***Table S2: Trait variation in S-locus homozygous individuals.***

Trait (unit)	Mean heterozygotes (SD)	Mean homozygotes (SD)	P value [a]
Time first leaf (day)	15.7 (1.04)	14.79 (1.19)	0.14
Number of leaves	13.22 (0.65)	13.46 (0.96)	0.35
Rosette area (cm²)	34.99 (5.11)	35.5 (7.65)	0.46
Mean leaf length (cm)	1.86 (0.15)	1.86 (0.21)	0.49
St dev leaf length (cm)	0.28 (0.02)	0.29 (0.04)	0.31
Mean leaf width (cm)	1.49 (0.11)	1.46 (0.14)	0.36
St dev leaf width (cm)	0.27 (0.02)	0.25 (0.03)	0.24
Mean leaf area (cm²)	3.16 (0.43)	3.11 (0.6)	0.45
Time first flower (day)	36.28 (1.14)	34.98 (1.25)	0.07
Number of flowering stems	12.9 (0.91)	13.3 (1.32)	0.31
Maximum flowering stem length (cm)	**56.19 (2.26)**	**64.55 (3.17)**	**$<1e^{-04}$**
Number of flowers by flowering stem	58.87 (5.96)	66.56 (10.71)	0.09
Flowering duration (day)	44.15 (1.7)	43.81 (1.88)	0.39

1107   *[a] The P value represents the proportions of the distribution with the expected difference of mean*
1108   *between heterozygotes and homozygotes at the S-locus after 10,000 random resamples equal to or*
1109   *less than the value observed. Significant values are represented in bold.*

1115 **Table S3: Trait variation in homozygous at the S-locus for the S-alleles Ah01 and Ah04 between**
1116 **families.**

Allele	Trait (unit)	Mean first familly (SD)	Mean second familly (SD)	p value
Ah01[b]	**Time first leaf (day)**	**17.11 (2.26)**	**32 (NA)**	**<1.00^{e-04}**
	Number of leaves	14 (2.83)	9 (NA)	0.10
	Rosette area (cm²)	27.65 (18.22)	14.21 (NA)	0.50
	Mean leaf length (cm)	1.64 (0.68)	1.52 (NA)	0.49
	St dev leaf length (cm)	0.28 (0.11)	0.17 (NA)	0.30
	Mean leaf width (cm)	1.38 (0.43)	1.38 (NA)	0.49
	St dev leaf width (cm)	0.24 (0.07)	0.16 (NA)	0.30
	Mean leaf area (cm²)	2.52 (1.59)	2.1 (NA)	0.50
	Time first flower (day)	35.86 (3.13)	36 (NA)	0.25
	Number of flowering stems	12.29 (3.9)	7 (NA)	0.25
	Maximum flowering stem length (cm)	**43.43 (6.91)**	**62.4 (NA)**	**<1.00^{e-04}**
	Number of flowers by flowering stem	31.73 (24.16)	52.25 (NA)	0.12
	Flowering duration (day)	**44.71 (6.73)**	**52 (NA)**	**<1.00^{e-04}**
Ah04	Time first leaf (day)	13.15 (2.68)	14.88 (4.36)	0.11
	Number of leaves	13.43 (2.99)	14.04 (3.9)	0.24
	Rosette area (cm²)	36.63 (27.74)	45.86 (33.48)	0.08
	Mean leaf length (cm)	1.93 (0.81)	2.13 (0.89)	0.14
	St dev leaf length (cm)	0.29 (0.15)	0.3 (0.12)	0.33
	Mean leaf width (cm)	1.54 (0.57)	1.44 (0.51)	0.20
	St dev leaf width (cm)	**0.28 (0.13)**	**0.22 (0.09)**	**0.02**
	Mean leaf area (cm²)	3.4 (2.45)	3.46 (2.27)	0.45
	Time first flower (day)	35.57 (4.73)	33.2 (3.66)	0.06
	Number of flowering stems	**13.21 (4.25)**	**16.4 (5.25)**	**0.02**
	Maximum flowering stem length (cm)	64.98 (9.65)	62.41 (9.34)	0.13
	Number of flowers by flowering stem	**72.38 (33.07)**	**45.25 (15.45)**	**1.00^{e-04}**
	Flowering duration (day)	43.68 (6.96)	41.56 (7.38)	0.23

[a] *The P value represents the proportions of the distribution with the expected difference of mean*
between each family after 10,000 random resamples equal to or less than the observed value.
Significant values are represented in bold. The Ah03 S-allele is excluded because only two
homozygotes survived. [b] *the standard deviation for the second family was not estimated because*
only one homozygote survived.

Table S4: Effect of dominance on variation of phenotypic traits in S-locus homozygous individuals.

Trait (unit)	Random effect implemented in GLM model	Distribution of trait tested	Linear effect	p value
Time first leaf (days)	*Family*	*Poisson*	*-0.15*	*5.5^{e-04}*
Leaves (counts)	Oxydative stress	Poisson	0.006	0.91
Rosette area (cm²)	Oxydative stress + family	Gaussian	8.25	0.19
Mean leave length (cm)	Oxydative stress + family	Gaussian	0.17	0.28
St dev leave length (cm)	None	Gaussian	0.02	0.54
Mean leave width (cm)	Oxydative stress + phytophagous attack	Gaussian	0.05	0.83
St dev leave width (cm²)	Oxydative stress + phytophagous attack	Gaussian	0.01	0.91
Mean leave area (cm²)	Oxydative stress	Gaussian	0.39	0.36
Time first flower (days)	Phytopathogen + phytophagous attack + family	Poisson	-0.0002	1
Flowering stems (counts)	Family	Poisson	0.15	0.08
Maximum flowering stem length (cm)	Phytopathogen + family	Gaussian	7.6	0.06
Flowers by stem (counts)	Family	Poisson	0.15	0.44
Flowering duration (days)	Phytopathogen + family	Poisson	-0.05	0.1

1123 *Dominance at the S-locus was implemented as a fixed effect. The dominance was considered a*
1124 *continuous variable.*

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1128   *Table S5: Number of phased haplotypes linked to S-alleles in each sample.*

Species	Population	Number of phased haplotypes	Number of *S*-alleles	Number of *S*-alleles with more than one copy	Mean number of gene copies per *S*-allele [a]
A. halleri	Nivelle	34	12	7	4.1
	Mortagne	38	11	9	4
A. lyrata	TSS	16	5	3	4.7
	IND	16	6	4	3.5
	PIN	22	7	3	6

1129   [a] *based on S-alleles with more than one copy.*

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1131 *Table S6: Effect of dominance on the accumulation of genetic load in S-flanking regions.*

Type of mutations	Species	Number of variable sites	Genetic load estimated	Linear effect	p value
0fold	*A. halleri*	369	**Number of fixed mutations by S-allele**	**0.08**	**2.57^{e-04}**
			Number of segregating mutations by S-allele	**-0.49**	**2.21^{e-03}**
			Mean total number of mutation by S-allele (S_{of})	0.01	0.54
	A. lyrata	243	**Number of fixed mutations by S-allele**	**0.08**	**4.50^{e-03}**
			Number of segregating mutations by S-allele	-0.44	0.09
			Mean total number of mutation by S-allele (S_{of})	-0.07	0.07
4fold	*A. halleri*	292	**Number of fixed mutations by S-allele**	**0.23**	**8.95^{e-04}**
			Number of segregating mutations by S-allele	**-0.48**	**2.02^{e-10}**
			Mean total number of mutation by S-allele (S_{4f})	1.53^{e-03}	0.95
	A. lyrata	189	**Number of fixed mutations by S-allele**	**0.23**	**8.95^{e-04}**
			Number of segregating mutations by S-allele	**-0.46**	**0.04**
			Mean total number of mutation by S-allele (S_{4f})	-0.04	0.32
S_{of}/S_{4f}	*A. halleri*		Ratio	0.02	0.42
	A. lyrata		Ratio	-0.11	0.11
Non-synonymous	*A. halleri*	376	**Number of fixed mutations by S-allele**	**0.10**	**6.36^{e-05}**
			Number of segregating mutations by S-allele	**-0.09**	**2.98^{e-03}**
			Mean total number of mutation by S-allele (S_{NS})	0.03	0.07
	A. lyrata	250	**Number of fixed mutations by S-allele**	**0.24**	**1.74^{e-08}**
			Number of segregating mutations by S-allele	-0.01	0.89
			Mean total number of mutation by S-allele (S_{NS})	**-0.11**	**2.06^{e-16}**
Synonymous	*A. halleri*	546	**Number of fixed mutations by S-allele**	**0.10**	**4.36^{e-07}**
			Number of segregating mutations by S-allele	**-0.12**	**3.74^{e-07}**
			Mean total number of mutation by S-allele (S_S)	**0.03**	**0.02**
	A. lyrata	369	**Number of fixed mutations by S-allele**	**0.27**	**2.3^{e-11}**
			Number of segregating mutations by S-allele	**-0.17**	**3.42^{e-03}**
			Mean total number of mutation by S-allele (S_S)	**-0.11**	**3.06^{e-25}**
S_{NS}/S_S	*A. halleri*		Ratio	-5.82^{e-03}	0.81
	A. lyrata		Ratio	0.02	0.60
Deleterious SIFT4G	*A. halleri*	58	Number of fixed mutations by S-allele	0.28	0.06
			Number of segregating mutations by S-allele	0.01	0.47
			Mean total number of mutation by S-allele (S_D)	0.02	0.69
	A. lyrata	33	**Number of fixed mutations by S-allele**	**0.66**	**3.36^{e-06}**
			Number of segregating mutations by S-allele	**-0.05**	**0.01**

			Mean total number of mutation by S-allele (S_D)	-0.10	5.48^{e-02}
Lowly deleterious SNPeff		548	**Number of fixed mutations by S-allele**	**0.09**	**1.34^{e-06}**
	A. halleri		**Number of segregating mutations by S-allele**	**-0.12**	**1.46^{e-06}**
			Mean total number of mutation by S-allele (S_{LD})	**0.03**	**0.02**
	A. lyrata	373	**Number of fixed mutations by S-allele**	**0.17**	**1.34^{e-13}**
			Number of segregating mutations by S-allele	**-0.17**	**0.47^{E-03}**
			Mean total number of mutation by S-allele (S_{LD})	**-0.11**	**3.51^{e-25}**
Moderately deleterious SNPeff		376	**Number of fixed mutations by S-allele**	**0.11**	**1.31^{e-05}**
	A. halleri		**Number of segregating mutations by S-allele**	**-0.10**	**6.39^{e-06}**
			Mean total number of mutation by S-allele (S_{MD})	0.03	0.07
	A. lyrata	247	**Number of fixed mutations by S-allele**	**0.13**	**4.45^{e-06}**
			Number of segregating mutations by S-allele	-0.02	0.74
			Mean total number of mutation by S-allele (S_{MD})	**-0.11**	**2.21^{e-16}**

1132 *Dominance at the S-locus was implemented as a fixed effect. Populations were implemented as random effects. The distribution of the genetic load*

1133 *estimated followed a Poisson distribution, except the ratios that followed normal distributions.*

1134   **Table S7: S-locus genotypes of individuals sequenced using the capture protocol.**

Individual	Allele 1ᵃ	Allele 2ᵃ	Population	Species	SRA
Mor_19_13	Ah04 (III)	Ah25 (III)	Mortagne	A. halleri	SAMN20088356
Mor_19_14	Ah04 (III)	Ah01 (I)	Mortagne	A. halleri	SAMN20844087
Mor_19_19	Ah12 (IV)	Ah36 (IV)	Mortagne	A. halleri	SAMN20088358
Mor_19_2	Ah03 (II)	Ah24 (IV)	Mortagne	A. halleri	SAMN20088359
Mor_19_23	Ah20 (IV)	Ah25 (III)	Mortagne	A. halleri	SAMN20844088
Mor_19_24	Ah12 (IV)	Ah24 (IV)	Mortagne	A. halleri	SAMN20088362
Mor_19_3	Ah03 (II)	Ah01 (I)	Mortagne	A. halleri	SAMN20844098
Mor_19_37	Ah59 (IV)	Ah20 (IV)	Mortagne	A. halleri	SAMN20088357
Mor_19_38	Ah36 (IV)	Ah01 (I)	Mortagne	A. halleri	SAMN20088371
Mor_19_4	Ah25 (III)	Ah12 (IV)	Mortagne	A. halleri	SAMN20088360
Mor_19_41	Ah25 (III)	Ah25 (III)	Mortagne	A. halleri	SAMN20088372
Mor_19_42	Ah12 (IV)	Ah04 (III)	Mortagne	A. halleri	SAMN20088373
Mor_19_45	Ah36 (IV)	Ah02 (III)	Mortagne	A. halleri	SAMN20088374
Mor_19_46	Ah03 (II)	Ah12 (IV)	Mortagne	A. halleri	SAMN20088375
Mor_19_51	Ah20 (IV)	Ah05 (IV)	Mortagne	A. halleri	SAMN20088361
Mor_19_53	Ah36 (IV)	Ah12 (IV)	Mortagne	A. halleri	SAMN20088365
Mor_19_54	Ah24 (IV)	Ah12 (IV)	Mortagne	A. halleri	SAMN20088366
Mor_19_55	Ah20 (IV)	Ah05 (IV)	Mortagne	A. halleri	SAMN20088377
Mor_19_56	Ah12 (IV)	Ah03 (II)	Mortagne	A. halleri	SAMN20088378
Niv_19_18	Ah15 (IV)	Ah03 (II)	Nivelle	A. halleri	SAMN20088349
Niv_19_19	Ah20 (IV)	Ah24 (IV)	Nivelle	A. halleri	SAMN20088350
Niv_19_22	Ah24 (IV)	Ah01 (I)	Nivelle	A. halleri	SAMN20088351
Niv_19_23	Ah20 (IV)	Ah12 (IV)	Nivelle	A. halleri	SAMN20088352
Niv_19_3	Ah04 (III)	Ah59 (IV)	Nivelle	A. halleri	SAMN20088336
Niv_19_31	Ah04 (III)	Ah24 (IV)	Nivelle	A. halleri	SAMN20088337

Niv_19_4	Ah01 (I)	Ah24 (IV)	Nivelle	A. halleri	SAMN20088382
Niv_19_42	Ah22 (IV)	Ah12 (IV)	Nivelle	A. halleri	SAMN20088341
Niv_19_45	Ah12 (IV)	Ah04 (III)	Nivelle	A. halleri	SAMN20088344
Niv_19_5	Ah25 (III)	Ah67 (IV)	Nivelle	A. halleri	SAMN20088340
Niv_19_52	Ah20 (IV)	Ah22 (IV)	Nivelle	A. halleri	SAMN20088331
Niv_19_53	Ah05 (IV)	Ah20 (IV)	Nivelle	A. halleri	SAMN20088339
Niv_19_54	Ah01 (I)	Ah12 (IV)	Nivelle	A. halleri	SAMN20844097
Niv_19_58	Ah02 (III)	Ah59 (IV)	Nivelle	A. halleri	SAMN20088343
Niv_19_59	Ah01 (I)	Ah67 (IV)	Nivelle	A. halleri	SAMN20844096
Niv_19_60	Ah04 (III)	Ah20 (IV)	Nivelle	A. halleri	SAMN20088332
Niv_19_7	Ah04 (III)	Ah67 (IV)	Nivelle	A. halleri	SAMN20088345
Niv_19_8	Ah04 (III)	Ah12 (IV)	Nivelle	A. halleri	SAMN20844090
Niv_19_9	Ah04 (III)	Ah67 (IV)	Nivelle	A. halleri	SAMN20088338
Pin_15_1	Ah03 (II)	Ah01 (I)	PIN	A. lyrata	SAMN20088324
Pin_16_1	Ah01 (I)	Ah46 (IV)	PIN	A. lyrata	SAMN20088318
Pin_4_24	Ah03 (II)	Ah29 (III)	PIN	A. lyrata	SAMN20088314
Pin_16_3	Ah18 (IV)	Ah29 (III)	PIN	A. lyrata	SAMN20088323
Pin_4_54	Ah03 (II)	Ah03 (II)	PIN	A. lyrata	SAMN20088319
Pin_5_1	Ah01 (I)	Ah01 (I)	PIN	A. lyrata	SAMN20844102
Pin_5_12	Ah29 (III)	Ah01 (I)	PIN	A. lyrata	SAMN20088325
Pin_5_2	Ah03 (II)	Ah01 (I)	PIN	A. lyrata	SAMN20088311
Pin_8_15	Ah01 (I)	Ah01 (I)	PIN	A. lyrata	SAMN20088328
Pin_8_2	Ah29 (III)	Ah63 (III)	PIN	A. lyrata	SAMN20088326
Pin_9_1	Ah42 (IV)	Ah29 (III)	PIN	A. lyrata	SAMN20088312
Tss_14_3	Ah01 (I)	Ah31 (IV)	TSS	A. lyrata	SAMN20088315
Tss_21_10	Ah01 (I)	Ah03 (II)	TSS	A. lyrata	SAMN20088329
Tss_22_24	Ah01 (I)	Ah31 (IV)	TSS	A. lyrata	SAMN20088306

Tss_22_7	Ah01 (I)	Ah31 (IV)	TSS	A. lyrata	SAMN20088317
Tss_3_10	Ah31 (IV)	Ah01 (I)	TSS	A. lyrata	SAMN20088330
Tss_23_2	Ah03 (II)	Ah63 (III)	TSS	A. lyrata	SAMN20088313
Tss_3_23	Ah18 (IV)	Ah01 (I)	TSS	A. lyrata	SAMN20088307
Tss_5_1	Ah01 (I)	Ah31 (IV)	TSS	A. lyrata	SAMN20088316
Ind_1_1	Ah01 (I)	Ah01 (I)	IND	A. lyrata	SAMN20088320
Ind_10_3	Ah31 (IV)	Ah24 (IV)	IND	A. lyrata	SAMN20088327
Ind_15_1	Ah03 (II)	Ah24 (IV)	IND	A. lyrata	SAMN20088321
Ind_15_2	Ah63 (III)	Ah03 (II)	IND	A. lyrata	SAMN20088309
Ind_9_3	Ah01 (I)	Ah46 (IV)	IND	A. lyrata	SAMN20088322
Ind_18_1	Ah01 (I)	Ah01 (I)	IND	A. lyrata	SAMN20844103
Ind_6_1	Ah01 (I)	Ah63 (III)	IND	A. lyrata	SAMN20088308
Ind_8_1	Ah03 (II)	Ah63 (III)	IND	A. lyrata	SAMN20088305

1135   [a] *The individuals were mainly sequenced by capture approach. The genotypes of homozygotes were*
1136   *confirmed after a whole genome sequencing. The class of dominance is specified in parenthesis.*

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Table S8: S-locus genotypes of the offspring selected for haplotype phasing and the crosses for the

study of phenotypic traits.

Identity[a]	Allele 1	Allele 2	Pollen donor[c]	Stigma recipient[c]	SRA
d32	Ah03	Ah12	Mor_19_2	Mor_19_19	SAMN20844104
d33	Ah01	Ah12	Mor_19_3	Mor_19_4	SAMN20844105
d38[d]	Ah04	Ah04	Mor_19_13	Mor_19_14	SAMN20844106
d42	Ah20	Ah12	Mor_19_23	Mor_19_24	SAMN20844107
d72	Ah20	Ah36	Mor_19_37	Mor_19_38	SAMN20844108
d74	Ah25	Ah12	Mor_19_41	Mor_19_42	SAMN20844109
d76	Ah03	Ah36	Mor_19_46	Mor_19_45	SAMN20844110
d89	Ah03	Ah05	Mor_19_56	Mor_19_55	SAMN20844111
d217	Ah12	Ah59	Mor_19_19	Mor_19_37	SAMN20844112
d250	Ah12	Ah20	Mor_19_24	Mor_19_51	SAMN20844113
d265	Ah36	Ah24	Mor_19_53	Mor_19_54	SAMN20844114
d3.1	Ah24	Ah25	Niv_19_4	Niv_19_5	SAMN20844115
d10.1[f]	Ah03	Ah24	Niv_19_18	Niv_19_19	SAMN20844116
d12.1[f]	Ah01	Ah20	Niv_19_22	Niv_19_23	SAMN20844117
d30.1	Ah02	Ah01	Niv_19_58	Niv_19_59	SAMN20844118
d48.1	Ah20	Ah25	Niv_19_52	Niv_19_5	SAMN20844119
d50.1[e]	Ah12	Ah04	Niv_19_8	Niv_19_60	SAMN20844120
d51.1	Ah12	Ah04	Niv_19_54	Niv_19_7	SAMN20844121
d191.1[e,f]	Ah04	Ah04	Niv_19_3	Niv_19_31	SAMN20844122
d206.1	Ah22	Ah25	Niv_19_42	Niv_19_5	SAMN20844123
d205.1[f]	Ah20	Ah04	Niv_19_53	Niv_19_9	SAMN20844124
d208.1[f]	**Ah03**	**Ah02**	**Niv_19_58**	**Niv_19_18**	**SAMN20844125**
d17.1[f]	**Ah03**	**Ah01**	**Niv_19_33**	**Niv_19_32[b]**	**SAMN20844126**
d24.1[f]	**Ah20**	**Ah01**	**Niv_19_47[b]**	**Niv_19_46[b]**	**SAMN20844127**

d29.1[f]	**Ah12**	**Ah01**	**Niv_19_57**[b]	**Niv_19_56**[b]	**SAMN20844128**
d50.2[f]	**Ah04**	**Ah04**	**Niv_19_8**	**Niv_19_60**	**SAMN20844129**
d266.1	Ah04	Ah12	Niv_19_7	Niv_19_45	SAMN20844130
d122	Ah01	Ah63	Pin_16_1	Pin_8_2	SAMN20844131
d127	Ah01	Ah03	Pin_5_1	Pin_4_54	SAMN20844132
d173	Ah03	Ah18	Pin_5_2	Pin_16_3	SAMN20844133
d176	Ah03	Ah42	Pin_15_1	Pin_9_1	SAMN20844134
d177	Ah01	Ah29	Pin_16_1	Pin_5_12	SAMN20844135
d239	Ah29	Ah01	Pin_4_24	Pin_8_15	SAMN20844136
d113	Ah18	Ah31	Tss_3_23	Tss_22_7	SAMN20844137
d273	Ah03	Ah01	Tss_23_2	Tss_22_24	SAMN20844138
d275	Ah63	Ah31	Tss_23_2	Tss_14_3	SAMN20844139
d284	Ah01	Ah31	Tss_21_10	Tss_5_1	SAMN20844140
d285	Ah01	Ah31	Tss_21_10	Tss_3_10	SAMN20844141
d118	Ah03	Ah01	Ind_9_3	Ind_8_1	SAMN20844142
d163	Ah03	Ah01	Ind_8_1	Ind_1_1	SAMN20844143
d166	Ah03	Ah01	Ind_6_1	Ind_15_1	SAMN20844144
d170	Ah01	Ah63	Ind_9_3	Ind_15_2	SAMN20844145
d232	Ah01	Ah31	Ind_18_1	Ind_10_3	SAMN20844146

[a] *Offspring figured in bold were not used to reconstitute haplotype of parents.* [b] *Parents were not sequenced.* [c] *Italics indicate parental haplotypes that were suppressed from the dataset because they had already been phased through another offspring .* [d] *The homozygous genotype of d38 was confirmed by genome-wide sequencing.* [e] *The homozygous genotypes of d191 and d50.1 were confirmed by the proportion of homozygous S-alleles in their offspring found by PCR after cross.* [f] *The individuals used for crosses in the phenotypic experiments.*

Table S9: Effect of phytopathogen, phytophagous attacks and oxidative stress on the phenotypic traits.

Effect tested		Mean with attack (sd)	Mean without attack (sd)	p value [a]
		15.37 (4.84)	15.38 (5.24)	0.49
		13.03 (2.9)	13.46 (3.71)	0.23
		31.93 (27.19)	37.08 (27.14)	0.13
		1.78 (0.78)	1.91 (0.8)	0.16
		0.29 (0.14)	0.28 (0.13)	0.38
		1.42 (0.56)	1.51 (0.53)	0.17
		0.25 (0.11)	0.27 (0.12)	0.23
		2.94 (2.33)	3.25 (2.15)	0.2
		37.43 (4.56)	**31.34 (5.55)**	**2e-03**
		12.05 (4.4)	11.84 (4.62)	0.05
	Maximum flowering stem	56.78 (12.04)	52.55 (11.82)	0.1
		56.56 (30.99)	56.22 (34.55)	0.11
Phytopathogen		**41.44 (8.13)**	**39.51 (7.32)**	**9e-03**
		14.48 (4.4)	15.56 (5.2)	0.16
		12.59 (3.76)	13.44 (3.35)	0.13
		32.15 (32.24)	35.77 (26.16)	0.27
		1.68 (0.83)	1.9 (0.78)	0.1
		0.29 (0.1)	0.28 (0.14)	0.36
		1.29 (0.55)	**1.51 (0.53)**	**0.03**
		0.22 (0.1)	**0.27 (0.12)**	**0.03**
		2.59 (2.3)	3.25 (2.19)	0.08
		34.78 (6.93)	**33.36 (4.92)**	**0.04**
Phytophagous	Number of flowering stems	10.78 (4.76)	12.15 (4.55)	0.34

	Maximum flowering stem	**53.53 (9.81)**	**54.24 (12.2)**	**0.05**
		47.46 (25.45)	58.12 (34.46)	0.17
		37.56 (8.1)	40.76 (7.76)	0.45
		14.92 (3.93)	15.46 (5.28)	0.34
		10.04 (3.88)	**13.93 (2.96)**	**$1e^{-04}$**
		11.79 (16.54)	**39.64 (26.56)**	**$1e^{-04}$**
		1.05 (0.49)	**2.02 (0.74)**	**$1e^{-04}$**
		0.25 (0.1)	0.29 (0.14)	0.09
		0.88 (0.31)	**1.59 (0.5)**	**$1e^{-04}$**
		0.23 (0.12)	**0.27 (0.12)**	**0.04**
		1.05 (0.98)	**3.54 (2.17)**	**$1e^{-04}$**
		31.77 (5.6)	33.95 (5.25)	0.06
		10.38 (4.85)	12.21 (4.54)	0.44
	Maximum flowering stem	50.3 (13.1)	54.85 (11.71)	0.11
		53.97 (55.98)	56.8 (28.05)	0.21
Oxidative stress	Flowering duration (day)	37.35 (6.5)	40.78 (7.94)	0.08

[a] *The P value represents the proportions of the distribution with the expected difference of mean between individuals with and without attack after 10,000 random resamples equal to or less than the value observed. Significant values are represented in bold.*

Table S10: Difference on the phenotypic trait variation between the two families for each allele

tested.

Allele	Family (Male/Female)	Trait (unit)	Mean (sd)	p value[a]
Ah01	*1 (d29.1/d24.1)*	**Time first leaf (day)**	**20.8 (5.43)**	**2e⁻⁰³**
		Number of leaves	11.9 (3.75)	0.09
		Rosette area (cm²)	24.9 (25.58)	0.11
		Mean leaf length (cm)	1.53 (0.79)	0.08
		St dev leaf length (cm)	0.21 (0.09)	0.03
		Mean leaf width (cm)	1.4 (0.58)	0.33
		St dev leaf width (cm)	0.24 (0.14)	0.32
		Mean leaf area (cm²)	2.47 (1.93)	0.17
		Time first flower (day)	28.8 (5.29)	0.44
		Number of flowering stems	**8 (4.07)**	**0.02**
		Maximum flowering stem length (cm)	**40.65 (6.76)**	**0.02**
		Number of flowers by flowering stem	**35.85 (10.72)**	**0.04**
		Flowering duration (day)	**40.35 (12.45)**	**8e⁻⁰³**
Ah01	*2 (d12.1/d29.1)*	**Time first leaf (day)**	**18.76 (5.57)**	**1e⁻⁰⁴**
		Number of leafs	13.45 (3.22)	0.42
		Rosette area (cm²)	31.99 (22.61)	0.25
		Mean leaf length (cm)	1.74 (0.71)	0.19
		St dev leaf length (cm)	0.29 (0.1)	0.42
		Mean leaf width (cm)	1.44 (0.51)	0.34
		St dev leaf width (cm)	0.29 (0.11)	0.11
		Mean leaf area (cm²)	2.84 (1.89)	0.21
		Time first flower (day)	**29.86 (5.43)**	**0.04**
		Number of flowering stems	9.59 (3.54)	0.32
		Maximum flowering stem length (cm)	**34.85 (9.35)**	**1e⁻⁰⁴**
		Number of flowers by flowering stem	**30.27 (25.45)**	**1e⁻⁰⁴**

		Flowering duration (day)	*31.62 (7.17)*	*0.06*
Ah03	*3 (d208.1/d17.1)*	*Time first leaf (day)*	*13.33 (5.27)*	*0.22*
		Number of leaves	*12.05 (4.25)*	*0.17*
		Rosette area (cm²)	*31.23 (25.7)*	*0.35*
		Mean leaf length (cm)	*1.67 (0.78)*	*0.32*
		St dev leaf length (cm)	*0.28 (0.15)*	*0.42*
		Mean leaf width (cm)	*1.39 (0.62)*	*0.5*
		St dev leaf width (cm)	*0.24 (0.13)*	*0.42*
		Mean leaf area (cm²)	*2.87 (2.24)*	*0.45*
		Time first flower (day)	***34.81 (7.87)***	***0.02***
		Number of flowering stems	***7.67 (3)***	***$3e^{-04}$***
		Maximum flowering stem length (cm)	***44.4 (13.87)***	***0.03***
		Number of flowers by flowering stem	*61.55 (46.06)*	*0.06*
		Flowering duration (day)	***39 (6.6)***	***$5e^{-03}$***
Ah03	*4 (d10.1/d17.1)*	Time first leaf (day)	18 (6.2)	0.12
		Number of leaves	12.33 (4.46)	0.22
		Rosette area (cm²)	**17.71 (11.23)**	**0.04**
		Mean leaf length (cm)	1.52 (0.53)	0.14
		St dev leaf length (cm)	0.23 (0.07)	0.14
		Mean leaf width (cm)	1.23 (0.35)	0.13
		St dev leaf width (cm)	0.23 (0.08)	0.23
		Mean leaf area (cm²)	2.01 (1.34)	0.1
		Time first flower (day)	33.33 (4.23)	0.09
		Number of flowering stems	11.67 (4.23)	0.22

		Maximum flowering stem length (cm)	57.25 (3.69)	0.34
		Number of flowers by flowering stem	75.05 (29.69)	0.12
		Flowering duration (day)	45.5 (8.67)	0.33
Ah04	*5 (d205.1/d191.1)*	***Time first leaf (day)***	***13.42 (3.28)***	***1e^{-04}***
		Number of leaves	*13.45 (2.97)*	*0.32*
		Rosette area (cm²)	*36.29 (27.39)*	*0.31*
		Mean leaf length (cm)	*1.92 (0.8)*	*0.22*
		St dev leaf length (cm)	*0.29 (0.15)*	*0.34*
		Mean leaf width (cm)	*1.54 (0.56)*	*0.11*
		St dev leaf width (cm)	*0.27 (0.13)*	*0.15*
		Mean leaf area (cm²)	*3.36 (2.41)*	*0.14*
		Time first flower (day)	*35.1 (4.66)*	*0.33*
		Number of flowering stems	*12.87 (4.16)*	*0.32*
		Maximum flowering stem length (cm)	***62.96 (9.48)***	***1e^{-04}***
		Number of flowers by flowering stem	***70.13 (32.5)***	***1e^{-04}***
		Flowering duration (day)	*42.44 (7.05)*	*0.3*
Ah04	*6 (d205.1/d50.2)*	*Time first leaf (day)*	*14.31 (4.36)*	*0.30*
		Number of leaves	*13.5 (3.9)*	*0.13*
		Rosette area (cm²)	***44.1 (33.48)***	***0.02***
		Mean leaf length (cm)	***2.04 (0.89)***	***0.04***
		St dev leaf length (cm)	*0.29 (0.12)*	*0.21*
		Mean leaf width (cm)	*1.38 (0.51)*	*0.36*
		St dev leaf width (cm)	***0.21 (0.09)***	***0.03***
		Mean leaf area (cm²)	*3.33 (2.27)*	*0.21*
		Time first flower (day)	***31.92 (3.66)***	***2e^{-03}***
		Number of flowering stems	***15.77 (5.25)***	***1e^{-04}***
		Maximum flowering stem length (cm)	*60.01 (9.34)*	*0.07*

Number of flowers by flowering stem	43.51 (15.45)	$5e^{-04}$
Flowering duration (day)	39.96 (7.38)	0.04

1161 *[a] The P value represents the proportions of the distribution with the expected difference of mean*
1162 *between each family and the other individuals after 10,000 random resamples equal to or less than*
1163 *the value observed. Significant values are represented in bold.*

1164