Alpha phase-coding supports feature binding during working memory maintenance

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Abstract

The ability to successfully retain and manipulate information in working memory (WM) requires that objects' individual features are bound into cohesive representations; yet, the mechanisms supporting feature binding remain unclear. Binding (or swap) errors, where memorized features are erroneously associated with the wrong object, can provide a window into the intrinsic limits in capacity of WM that represent a key bottleneck in our cognitive ability. We tested the hypothesis that binding in WM is accomplished via neural phase synchrony and that swap errors result from perturbations in this synchrony. Using magnetoencephalography data collected from human subjects in a task designed to induce swap errors, we showed that swaps are characterized by reduced phase-locked oscillatory activity during memory retention, as predicted by an attractor model of spiking neural networks. Further, we found that this reduction arises from increased phase-coding variability in the alpha-band over a distributed network of sensorimotor areas. Our findings demonstrate that feature binding in WM is accomplished through phase-coding dynamics that emerge from the competition between different memories.

1. Introduction

Working memory (WM) is the ability to retain and manipulate information when it is no longer present in our environment, which introduces the possibility for abstract concepts and plans to influence our behavior [1–4]. The well-described capacity limits of WM [5–7] hence represent a key constraint on higher order cognition such as cognitive control [8]. In real life, WM capacity is not just dependent on the maintenance of individual visual features (e.g., color, orientation, spatial location), but also of the conjunctions or bindings between them [9–12]. Thus, understanding how information is bound in WM is of critical importance to understand the intrinsic limits of this executive function.

One proposed mechanism for feature binding in WM is temporal synchrony between populations of neurons that encode individual features [13,14]. This idea relies on the notion of temporal or phase-coding—that is, the two different neuronal populations encoding the information of distinct features of a conjunction fire at the same phase of an ongoing oscillation, enabling these features to be bound together in WM [15–17]. While the idea of phase-coding is broadly consistent with empirical evidence of synchronized oscillatory dynamics within or between different brain regions during WM [18–23], a biologically-plausible computational implementation of phase-coding has been elusive [24,25] and direct evidence for phase-coding has been lacking. A recent model [26] provides a plausible neural architecture of the cortico-cortical circuits that implement feature binding in WM, in which phase-coding emerges from the lateral inhibition between the different neuronal populations that store information about distinct (competing) conjunctions [27]. In the present study, we sought to test a central prediction of this model: that disruptions in oscillatory phase should be associated with misbinding or 'swap' errors [26], where an inaccurate response to the target item is accurate relative to a non-target item (e.g., if a subject shown a red square and a blue circle mistakenly reports the color of the circle as red) [10,11]. If this prediction were validated, it would provide a key empirical demonstration that phase-coding serves as an organizational principle during WM maintenance.

We tested our hypothesis using magnetoencephalography (MEG) recordings collected from human subjects performing a task designed to induce swap errors. Our results showed a characteristic within-trial phase-locking in the alpha-band during WM maintenance. In parietal-occipital sensors contralateral to the visual stimuli, the consistency over trials of such alpha phase-locking was reduced in swap trials compared to high-performance on-target trials. Importantly, these effects did not generalize to other WM errors, suggesting that such deterioration in phase-locked oscillatory activity is a hallmark of swaps. To understand why phase relationships are compromised in swap trials, we considered a measure of variability in the instantaneous frequency of alpha oscillations and showed that swaps are characterized by increased variability. We further localized these effects in contralateral areas in premotor, motor, parietal, and visual cortices. These results suggest that during WM maintenance feature binding is accomplished via alpha phase-coding, while swaps are produced by unstable phase-locked activity in distributed sensorimotor areas.

2. Results

2.1. Identification of swap errors

We analyzed MEG data previously collected from 26 human subjects who performed a delayed-response task designed to induce swaps [28] (Fig. 1). Subjects were instructed to remember a briefly shown lateralized display of 3 circles (stimulus presentation: 0.2 s). After a brief memory retention interval (delay period: 2 s), subjects reported the location of each of the circles, which were sequentially cued by their color (report period: self-paced) in a random order. Each subject completed a total of 500 trials of this WM task (see section 4.2). We used continuous measures of response errors and employed a maximum likelihood approach to distinguish high-performance (HP) trials (location of all circles reported accurately; Fig. 1a) from swap trials (location of one or more of the non-cued circles mistakenly reported; Fig. 1b), and from low-performance (LP) trials (location of one or more circles reported inaccurately) based on the subjects' responses [29] (https://www.paulbays.com/toolbox/; see section 4.2). Based on previous work [28], we defined HP targets as the trials in which the probability of correctly reporting the target item was greater than 0.95 in all three responses, and swaps as trials with a probability of incorrectly reporting a non-target item greater than 0.70 in at least one response. The remaining trials were defined as LP trials. We observed all three trial types in all subjects: HP trials ranged from 16–439 across subjects (M=92.81, SD=88.06); swaps ranged from 22–196 (M=116.38, SD=47.79), and LP trials ranged from 44–328 (M=198.73, SD=64.03). In a subsequent analysis (see section 2.4), we distinguished between trials that had one (Fig. 1b, upper example) or multiple (Fig. 1b, lower example) swap reports.

Fig. 1 Experimental design. **a–b** Schematic representation of the temporal structure of each trial, with examples of high-performance (HP) response (a) and swap errors (b). In the schematic depiction of the report period, the colored circles represent the cues, while the white circles represent the subject's responses. In b, the examples depict trials in which the location of one (upper) or more (lower) of the non-cued circles are mistakenly reported.

2.2. Less stable phase-locked dynamics in swaps during WM maintenance

Our hypothesis entailed that HP trials are characterized by stable phase-locked activity during WM maintenance, while swap trials are induced by noisy fluctuations in the phase of oscillatory activity related to the maintenance of individual features [26]. To test this prediction, we used the Phase Preservation Index (PPI) [30], which captures for each time point the consistency over trials of the within-trial phase differences with respect to a reference time (in our case, the memory delay onset at t=0.2 s; Fig. 1)—that is, their level of phase clustering in polar space [31]. Thus, PPI provides a measure of the consistency of frequency-specific local phase-locking over trials, as a function of time. We estimated PPI for each MEG sensor over time points in the delay period (0.2–2.2 s; Fig. 1) and frequencies 1–50 Hz separately for each trial type, and then compared PPI between HP trials and swaps. Since PPI is sensitive to the number of trials, we performed the analysis by equating the number of HP and swap trials (see section 4.5).

PPI was higher in HP targets than swaps in the alpha-band. We found a cluster in the observed data, extending over frequencies between 8–14 Hz and latencies around ~0.5 s after delay onset (Fig. 2a). PPI differences between trial types were most pronounced over 25 parietal-occipital sensors, contralateral to stimulus presentation (Fig. 2b), with medium effect size (Cohen's d in the range 0.170–0.576 across sensors in the observed cluster). The cluster-based permutation test revealed that PPI was significantly higher in HP targets compared to swaps (p_{perm}=0.0160), with phase preservation reduced for swaps in the alpha-band, showing a steeper decrease than HP targets (Fig. 2c), as predicted by model simulations of synchronization-based feature binding [26].

In order to test the specificity of the observed alpha-band effects, we calculated the z-scores of PPI over frequencies for each time point in the delay, separately for each trial type. We found positive z-scores at low frequencies in the delta/theta range (up to ~6 Hz) and negative z-scores at higher frequencies in the beta/gamma range (above ~15 Hz), while the middle frequencies in the alpha-band (~10 Hz) were characterized by positive z-scores that were higher than those at adjacent frequencies, both low and high (Fig. 2d). These results reveal that while there is a general tendency for PPI to decrease faster at higher frequencies than lower frequencies, alpha deviates from this trend.

Fig. 2 PPI differences between HP targets and swaps. **a** Time-frequency distribution of the sum of PPI differences across MEG sensors (*t*-values). The black contour highlights the statistically significant cluster. The marginal plots on the right and on top represent respectively the time-collapsed frequency distribution and frequency-collapsed time distribution of the differences between trial types. **b** Topography plot with superimposed effect sizes of PPI differences between HP targets and swaps, for each MEG sensor of the observed cluster (*t*-values are shown on the smaller topography plot on the right). **c** Time course of PPI in HP targets (left) and swaps (right) at different frequencies, highlighted by the black horizontal lines in the next panel. **d** Z-scores along frequencies for each time point in the delay, in HP targets (left) and swaps (right). **e** Time course of PPI corrected at 10 Hz for HP targets (red) and swaps (blue), together with the difference between the two (gray). The red and blue shadings (top) represent the standard error of the mean, while the gray shading (difference; bottom) represents 95% confidence intervals (CIs). CIs were estimated using the bias-corrected and accelerated method on a bootstrap distribution of across subjects differences, obtained by resampling with replacement 10,000 times. **f** Time course of PPI corrected at 10 Hz for HP targets (red) and LP trials (green), together with the difference between the two (gray). The red and green shadings (top) represent the standard error of the mean, while the gray shading (difference; bottom) represents 95% CIs (estimated as in panel e). **g** Time distribution of the sum of 10 Hz PPI differences between HP targets and swaps, across source points in source-space (*t*-values; top), and time course of the average PPI across source points of the positive cluster for each condition (bottom). The shadings represent the standard error of the mean. **h** Glass brain with superimposed effect sizes of the observed differences between HP targets and swaps. **i** Time course of PPI corrected at 10 Hz for HP targets (red) and swaps (blue) across source points of the cluster in the observed data, together with the difference between the two (gray; shadings are as in panel e).

To confirm that the observed PPI effects were due to within-trial phase synchronization in the alpha-band, we performed the following control analysis. We repeated the PPI estimation (including the procedure to equate the number of trials between HP targets and swaps – see section 4.5), but this time after randomly shuffling the signals from each MEG sensor over time, which disrupted any phase relationships present in the data. We then computed 'PPI corrected' estimates as the difference between PPI estimates from the main analysis and those obtained from the control analysis. The results obtained from shuffled data (i) showed that the rate of PPI reduction over time increases monotonically with frequency, including for the alpha-band, and (ii) confirmed the presence of a plateau/asymptote for PPI at all frequencies (~0.13 on average across subjects; Supplementary Fig. 2a–b). The PPI corrected results showed, instead, that phase synchronization takes place specifically at alpha frequencies in our WM task (Supplementary Fig. 2c–d), and that HP targets are characterized by more stable dynamics relative to swaps during the delay period (Fig. 2e). In a separate control analysis, we compared PPI estimates between HP targets and LP trials (see section 4.2) and did not find any statistically significant differences (the lowest *p*-value from the permutation test among the identified clusters was $p_{perm}=0.187$; Fig. 2f), suggesting that PPI differences in alpha are specific to swaps and are not a property of impaired behavioral performance.

To identify the underlying cortical sources exhibiting differences in local phase-locked activity between
HP targets and swaps, we used a source reconstruction technique to localize MEG sources of activity, and repeated
the PPI analysis in source-space (see section 4.7). PPI estimation was here restricted to the frequency of interest
of 10 Hz, which was selected as the peak of the time-collapsed frequency distribution of observed PPI differences
in sensor-space (see Fig. 2a). In source-space, we observed a positive cluster (higher PPI in HP targets than swaps)
in the data (Fig. 2g), which included 17 source points localized in the cingulate cortex and cortical areas
contralateral to stimulus presentation (Fig. 2h). Effect size peaked in source points localized in the ventral anterior
cingulate cortex–ACC and dorsal posterior cingulate cortex–PCC, ipsilateral primary motor cortex–M1, and
contralateral areas in primary somatosensory cortex–S1 and precuneus/posterior parietal cortex–PPC (see Table
1). The cluster-based permutation test identified that there was a significant difference in 10 Hz PPI between HP
targets and swaps (p_{perm}=0.0320). Like in the sensor-space analysis, the PPI corrected results showed that alpha
phase-locked activity is characterized by more stable dynamics in HP targets than swaps during the delay period
(Fig. 2i).

Table 1. PPI differences between HP targets and swaps: peak effect sizes

Anatomical region	MNI coordinates (x, y, z)	Cohen's d
Ventral ACC	-17.5, -20, 42.5	0.694
Dorsal PCC	7.5, -44.5, 55	0.702
M1	20, -32.5, 67	0.732
S1	-30, -32.5, 42.5	0.766
Precuneus/PPC	-17.5, -69.5, 2.5	0.616

Note. *ACC* and *PCC* stand for anterior and posterior cingulate cortex, respectively. *M1* stands for primary motor
cortex. *S1* stands for primary somatosensory cortex. *PPC* stands for posterior parietal cortex.

We note, however, that it is difficult to draw conclusions about the precise timing of effects due to the
intrinsic limitations of the cluster-based permutation testing, which only provides an approximation of the effect
extent [32,33]. Further, PPI may not be best suited for determining the timings of the observed effects. Because it is
equal to 1 at a reference time point and decays quickly over time [26,30], PPI estimates close to said reference are
effectively inflated. This can be appreciated by examining the time course of PPI at different frequencies (e.g., in
HP targets; Fig. 2c): estimates close to the reference are close to 1 in value and tend to decrease over time to an
asymptotic value. Thus, PPI may be most sensitive to differences that are in the early delay period of the task.

2.3. Increased phase-coding variability in swaps

The attractor networks model [26] suggests that phase-locked dynamics allow maintaining color-location
conjunctions in WM, but these phase-locked dynamics can be broken by abrupt noisy fluctuations, which lead to

misbinding of memorized features and swaps. We confirmed the neurophysiological prediction of this model, in that trials containing swaps have a lower PPI compared to HP targets. However, PPI does not directly address the origin of instabilities in the phase of alpha oscillations (Fig. 3a). A degradation of alpha phase preservation could occur either as a result of fluctuations in alpha phase that continues throughout the delay period or as a result of transient fluctuations in alpha phase. To disambiguate these two scenarios, we calculated frequency sliding (FS), which captures the instantaneous temporal fluctuations in oscillation peak frequency (Fig. 3b)—that is, time-varying changes in the instantaneous frequency of the oscillator [34]. In order to demonstrate the feasibility of using FS to capture the noisy fluctuations that induce swaps, we estimated FS from the signals obtained from the two attractor networks model [26]. In Fig. 3c–d, we depict an example trial where a noisy fluctuation at around t=2.7 s reverses the internetwork correlation, ultimately producing a swap (Fig. 3e). Our results demonstrate that FS is able to capture the timing of these fluctuations (Fig. 3f). In the first network, FS varies suddenly at around t=2.7 s, accurately capturing the time when the internetwork correlation reverses. This effect is also visible on the difference in FS between the two networks (see black trace in Fig. 3f).

Unlike PPI, FS is a reference-independent measure; thus, it overcomes the limitation of PPI discussed in section 2.2 to provide a better estimate of the timing of observed effects. This is a useful feature also because previous studies have recognized that, in addition to WM maintenance, encoding and recall are also noisy processes [11]; consequently, swaps may result from a failure in correctly binding visual features during any of these three phases of WM [35–38]. We selected a frequency of interest at the peak of the time-collapsed frequency distribution of PPI differences (10 Hz; see Fig. 2a), and derived FS at single-trial level. We then estimated a measure of 'alpha-FS variability (FSV) over time' in a sliding window during the time between the stimulus presentation and the delay period (0–2.2 s) by taking the median between trials of the standard deviation of FS over time points, separately for HP targets and swaps (see section 4.6). We also estimated the median standard deviation of FS over trials on the sliding window to derive a measure of 'alpha-FSV over trials', separately for HP targets and swaps (see section 4.6). Differences between HP targets and swaps in both measures of alpha-FSV would indicate that swaps are induced by noisy fluctuations that are more sustained over time (scenario 1 in Fig. 3g), while if the source of variability is predominantly over trials, this would suggest that noisy fluctuations in swaps happen quickly and more suddenly (scenario 2 in Fig. 3g and as predicted by the attractor model, Fig. 3f).

Fig. 3 Instantaneous phase fluctuations, frequency sliding (FS), and frequency sliding variability (FSV). **a** Example from a swap trial of the narrowband filtered-signal in the alpha-band (in blue) and its instantaneous phase

(in black) during the delay period. The signal was obtained from the first subject, from one of the contralateral
parietal sensors where we observed significant PPI difference. **b** Single-trial FS estimated from the signal
represented in panel a. In panels a–b, the gray shadings highlight a noisy fluctuation in the alpha phase and how
this can be captured by FS. **c–d** Raster plots of a sample simulation [26] from the two ring attractor networks (left),
together with the delay-period zooms showing clear bump oscillatory activity (right). **e** In the simulated example,
a noisy fluctuation reverses these correlations suddenly producing a swap trial, as shown by the internetwork bump
correlation plot. **f** Frequency sliding (FS) estimated from the simulated field signals of the first (dark orange) and
second network (light orange) at peak-frequency, from the model depicted in panels c–d. The difference in FS
between the two networks is depicted in red (bottom). Vertical arrows mark the time of a large, abrupt noisy
fluctuation, leading to a swap error during the readout. **g** Toy model illustrating the estimation of alpha-FSV over
time points or over trials, and alternative scenarios for how a reduction in alpha phase preservation in swaps is
produced by noisy fluctuations in instantaneous alpha frequency with different temporal profiles.

We did not find any statistically significant differences between HP targets and swaps in the alpha-FSV
over time (the lowest *p*-value from the permutation test among the identified clusters was p_{perm}=0.368). Alpha-
FSV over trials, however, was significantly higher in swaps compared to HP targets (p_{perm}=0.0180) in a cluster
extending into the second half of the delay period (approximately between 1.182–2.045 s; Fig. 4a). The differences
in alpha-FSV over trials were most prominent in 20 contralateral, parietal-occipital sensors (effect size ranging
between *d*=0.088 and *d*=0.321 across cluster sensors; Fig. 4b). This and the previous PPI analysis represent
convergent findings, as the groups of electrodes showing the biggest effects largely overlapped between the two
analyses (compare Fig. 2b and Fig. 4b).

Fig. 4 Differences in alpha-FSV over trials between HP targets and swaps. **a** Time distribution of the sum of alpha-FSV differences across MEG sensors (*t*-values; top), and time course of the average alpha-FSV across sensors of the cluster found in the data for each condition (bottom). The shadings represent the standard error of the mean. **b** Topography plot with superimposed effect sizes of the differences between HP targets and swaps, for each MEG sensor of the observed cluster (*t*-values are shown on the smaller topography plot on the right). **c** Time distribution of the sum of alpha-FSV differences across source points in source-space (*t*-values; top), and time course of the average alpha-FSV across source points of the observed cluster for each condition (bottom). The shadings represent the standard error of the mean. **d** Glass brain with superimposed effect sizes of the differences between HP targets and swaps.

Analogous to our PPI analysis, we repeated the FSV analysis in source-space (see section 4.7). The source-space results confirmed that alpha-FSV over trials was higher in swaps than HP targets. A negative cluster was found extending approximately over latencies 1.148–2.200 s (Fig. 4c) and 92 source points, localized for the most part in contralateral cortical areas (Fig. 4d). Effect size peaked in source points localized in premotor cortex–PMC, motor area M1, angular gyrus–AG, visual area V2, and anterior cingulate–ACC (see Table 2). The cluster-based permutation test indicated that there was a significant difference in alpha-FSV over trials between swaps and HP

249 targets (p_{perm}=0.0260). In line with the sensor-space results, we did not find any statistically significant differences

250 in alpha-FSV over time (lowest p_{perm}=0.833 among clusters). Together these results suggest that, in areas

251 contralateral to stimulus presentation, correct feature binding in WM is supported by stable phase-coding in the

252 alpha-band, while swap errors are the result of its perturbations. In swaps, noisy fluctuations in the alpha

253 oscillations peak frequency happen quickly and abruptly rather than in a more sustained fashion over time.

Table 2. Differences in alpha-FSV over trials between HP targets and swaps: peak effect sizes

Anatomical region	MNI coordinates (x, y, z)	Cohen's d
PMC	-30, 5, 55	0.430
PMC	-42.5, 5, 42	0.386
M1	-42.5, -20, 55	0.361
AG	-42.5, -57, 55	0.305
V2	-5, -57, 5	0.291
ACC	-5, -20, 42	0.290

255 *Note*. *PMC* stands for premotor cortex. *M1* stands for primary motor cortex. *AG* stands for angular gyrus. *ACC*

256 stands for anterior cingulate cortex. *V2* stands for secondary visual cortex.

2.4. Distinct possible sources of swaps

The term 'swap' is agnostic as to whether features are symmetrically exchanged between two memorized

items, or asymmetrically exchanged [39]. That is, in a scenario where a subject is presented with a red item at location

1 and a green item at location 2, and, when asked to recall the location of the red item, erroneously reports location

2, it is unclear whether they would necessarily report location 1 when asked to recall the location of the green

item. If so, this would imply a symmetric exchange of features. Here, we defined the trials with only one swap

(i.e., a single non-target report) as 'non-symmetric' and the trials with multiple swaps as 'symmetric' (see Fig. 1b

for examples). The nature of our task, which involved a full report of all memorized items, allows us to ask (i)

whether swaps are necessarily symmetric, and (ii) if both symmetric and non-symmetric swaps exist, whether

phase-coding variability differs between these two types of swaps, potentially implying a different underlying

mechanism. Our analysis focused on FSV for two reasons: first, previous work found that the precise timing of

phase asynchrony may influence how swaps errors arise [26] and second, splitting trials into symmetric and non-

symmetric swaps left an insufficient number of trials for PPI analysis.

The number of trials showing symmetric swaps was in the range 7–81 across subjects (*M*=46.50,

SD=18.51), while for non-symmetric swaps was in the range 15–123 (*M*=69.88, *SD*=31.70), demonstrating that

swaps are not necessarily the result of symmetric feature exchanges. For the alpha-FSV over trials, we observed a

negative cluster (alpha-FSV higher in non-symmetric than symmetric swaps), extending over latencies between

stimulus presentation and the beginning of the delay period (~0–0.740 s; Fig. 5a). This cluster included 63 sensors, mostly in frontal and sensorimotor sites, over both hemispheres (Cohen's d in the range 0.071–0.617 across sensors; Fig. 5b). The cluster-based permutation test revealed a significant difference between swap types in alpha-FSV over trials (p_{perm}=0.0060). We did not observe significant differences later in the delay for alpha-FSV higher in symmetric than non-symmetric swaps. Further, we did not find any statistically significant difference between swap types in the alpha-FSV over time (lowest p_{perm}=0.443 among clusters). When we repeated the analysis in source-space, we observed a trend towards increased alpha-FSV over trials for non-symmetric compared to symmetric swaps, in areas of the lateral and medial prefrontal cortex, temporal pole–TP, and angular gyrus–AG, in particular in the contralateral hemisphere, which is in line with the sensor-space results; however, these effects did not reach statistical significance with a rigorous statistical threshold (positive cluster extending approximately over latencies 0–1.844 s; p_{perm}=0.0919; see Fig. 5c–d).

Fig. 5 Differences in alpha-FSV over trials between symmetric and non-symmetric swaps. **a** Time distribution of the sum of alpha-FSV differences across MEG sensors (*t*-values; top), and time course of the average alpha-FSV across sensors of the observed cluster for each condition (bottom). The shadings represent the standard error of the mean. **b** Topography plot with superimposed effect sizes of alpha-FSV differences between symmetric and non-symmetric swaps, for each MEG sensor of the observed cluster (*t*-values are shown on the smaller topography

292   plot on the right). **c** Time distribution of the sum of alpha-FSV differences across source points in source-space (t-

293   values; top), and time course of the average alpha-FSV across source points of the cluster with the lowest p-value

294   from the permutation test (p_{perm}=0.0919) for each condition (bottom). The shadings represent the standard error of

295   the mean. **d** Glass brain with superimposed effect sizes of differences between symmetric and non-symmetric

296   swaps.

3. Discussion

Previous models have proposed alternative neural processes for how the binding of information between properties or features that belong to an object is accomplished in WM [10,11]. An influential class of theories and models proposed that feature binding is accomplished via the phase synchronization of signals of neurons (phase-coding) that store the different feature values corresponding to the memorized item [13,14,18,24–26]. Here, we tested the central neurophysiological prediction of these models. We found that the correct feature binding is supported by stable phase-locked oscillatory activity at alpha frequencies, while swaps (binding errors in memory) are characterized by reduced phase preservation during WM maintenance, with sources localized in the hemisphere contralateral to stimulus presentation. This reduction did not generalize to other types of errors, suggesting that alpha phase inconsistencies are a hallmark of swaps. An independent analysis found that these phase inconsistencies arose from increased phase-coding variability. We localized the increase in phase-coding variability for swaps in contralateral premotor, motor, parietal, and visual cortical areas. Thus, we provide convergent evidence that, in swaps, the sources of reduced neural phase synchrony and higher phase-coding variability are in contralateral areas. Together our results suggest that feature binding in WM is accomplished through alpha phase-coding dynamics, which supports both detailed computational models of feature integration in WM [26] and more general theories of time-based binding for cognitive flexibility [40]. More broadly, this work represents a demonstration of how mechanistic predictions generated by biophysically-realistic models can be tested using human neuroimaging.

Modeling studies suggest that phase-coding is supported by the competition between different memories, and may result from lateral inhibition between neural populations selective for different features. In particular, Barbosa et al. [26] proposed a model where the maintenance of conjunctions is accomplished through phase-coding between two coupled one-dimensional attractor networks. In the model, oscillatory dynamics emerge within each network through the local interplay of fast recurrent excitation and slower feedback inhibition. In turn, different populations active in each network (representing different colors, or different locations) oscillate out-of-phase based on the competition between the activity bumps, accomplished by gamma-aminobutyric acid (GABA)-mediated lateral inhibition between them. Finally, feature conjunctions are established via weak non-specific excitatory connections between the two attractor networks, whereby pairs of bumps oscillating in each network synchronize selectively. Several empirical observations are consistent with this mechanistic account. Strong feedback inhibition is a prominent feature of cortical networks [41], and it strongly controls their activity fluctuations [42]. Also, brain activity is known to be dominated by an alternation in the firing of different neuronal populations [43], and the 'gating by inhibition hypothesis' has proposed that alpha activity in sensory regions implements a mechanism of pulsed inhibition that silences neuronal firing [44,45], mediating the observed alternation in neuronal firing. Our results suggest that such alternating, phase-coding mechanism supports feature binding during WM maintenance: stable oscillatory target dynamics result in high-performance reports of cued items, while noisy

fluctuations produce abrupt alpha phase instabilities, inducing shifts to non-target dynamics that ultimately result
in swaps.

We provide converging evidence that sources of reduced phase synchrony in swaps and increased phase-
coding variability are localized in areas contralateral to the visual stimuli, with results showing overlapping effects
in both sensor-space topographies and source-space reconstructions. However, the time windows of these effects
appear disjoint between the two analyses. This apparent inconsistency can be explained by the differences between
the two measures, Phase Preservation Index–PPI and frequency sliding variability–FSV over trials, as well as the
intrinsic limitations of the cluster-based permutation approach. Cluster-based permutation testing does not allow
to precisely determine the exact timing of the observed effects, but it only provides an approximation of the extent
of effects [32,33]. In terms of the measures, PPI estimates exhibit a rapid asymptotic decay over time from the
reference point [26,30], which, in our case, is the onset of the memory delay. Therefore, they can reliably detect only
differences that occur early in the delay. On the other hand, FSV does not suffer from this problem. Alpha-FSV
over trials does seem to differentiate between swaps and HP targets not only late in the delay but also early,
especially in the sensor-space analysis (see Fig. 4a, section 3.2). However, this was not significantly detected by
the cluster analysis. Taking into account these methodological limitations, our data is consistent with the view that
quick, rapid phase changes occurring at any time during the delay are responsible for producing swap errors [26].

Our findings that alpha phase-coding supports feature binding in WM agree broadly with previous studies
indicating that mid-frequency alpha and beta oscillations play a central role in the control over WM [46–49].
Nonhuman primate WM studies using invasive recordings in the monkeys' prefrontal cortex, for example, have
found an inverse relationship between gamma (~40–120 Hz) and beta power (~20–35 Hz), mediated by oscillatory
bursts [50–52]. During WM maintenance these anti-correlated beta and gamma dynamics have been found to underlie
a reduction in spiking variability [53], which supports the notion that mnemonic representations are maintained
through mechanisms of phase-coding [18,23]. While conceptually consistent with our observations, these findings
differ from our own both in the regions (prefrontal cortex vs. sensorimotor areas) and frequency band (beta vs.
alpha). A potential reconciliation comes from a recent fMRI study that used longitudinal training to show that
item-selective mnemonic representations become detectable in the prefrontal cortex over long-term learning [54].
Hence, one possibility is that our observed phase-coding mechanisms supporting the storage of mnemonic
representations, may spread from posterior alpha to frontal beta over the course of training or, more in general,
when task-specific categories, associations, and rules are learned.

The fact that phase synchrony disruptions were localized to sensory and motor cortices is consistent with
the 'sensorimotor recruitment' model of WM, which states that visual information is maintained in the same
stimulus-selective regions that are responsible for perceiving that information [2,55–58], as well as with the 'distributed
system view' of WM [59,60], which states that WM storage is distributed across multiple brain regions, among which
both visual and parietal cortices play an important role in storing visual information during WM maintenance. The
model proposed by Barbosa and colleagues relies on two ring-attractor networks, one for each feature space (color

and spatial location), thus it explicitly simulates the independent storage of individual features constituting the item, in line with the increasing evidence that different features are stored in independent brain systems (different cortical areas) [11,61]. In our task, storage for color and spatial location may occur in distinct neural populations. For example, the colors of the items may be stored in color sensitive areas in visual cortex, while their spatial locations may be integrated into salience maps by premotor/motor and parietal areas, by storing them in systems that represent them not only as perceptual information but also as targets for future motor responses during the report period. Our data is consistent with the view that sudden phase instabilities in either system produce abrupt changes in the phase relationship between the two systems, changing the correct conjunctions between features and producing binding errors.

The presence of binding errors has been examined also in cued-recall tasks where two non-spatial features were used as the cue and report features [62–64]; however, spatial location has been recognized as having a privileged role in feature binding and in the generation of swaps compared to non-spatial features [64,65]. Previous studies have shown that the likelihood of swaps depends on the feature employed as the memory cue, and the frequency of swaps is higher when spatial location is the report feature compared to when it is the cue feature [66–68]. In the present study, we employed color as the cue feature and spatial location as the report feature. While the number of swaps would be lower when using location as cue feature and color as report feature, we expect that our results would be similar in terms of modulations of alpha phase dynamics. Future work should explicitly evaluate this scenario to confirm this hypothesis. Also, the mechanistic conditions for this asymmetry should be addressed in computational models, as they currently only consider symmetric feature networks [26]. We see at least two possibilities to break this symmetry: having stronger within-region noise for color networks or asymmetric readouts.

The origin of swaps in WM is still a topic of current debate. Some studies have proposed that swaps derive from an informed guessing strategy, rather than being true binding errors in memory [69,70]. Other studies, however, have argued against this account by providing evidence that swaps are better accounted for by cue-dimension variability [68] or by demonstrating that it is possible to reconstruct a mnemonic representation of the swapped location but not of the cued location, suggesting that swaps involve the active maintenance of a non-target memory item [37]. Our work adds to this literature by leveraging the combined spatial and temporal resolution of MEG to demonstrate increased variability in alpha phase-coding for swaps during WM maintenance in distributed sensorimotor areas. This supports the notion that swaps are not mere response guesses but true binding errors in memory that occur prior to the initiation of a response.

Beyond demonstrating that swaps are associated with disruptions in phase-coding dynamics, we found that swaps can arise through both a symmetric and asymmetric feature exchange. When we compared these two types of swaps, we found higher alpha phase-coding variability in non-symmetric swaps (single non-target report) than symmetric swaps (at least two non-target reports). These effects were observed predominantly at latencies between the stimulus presentation and the beginning of the delay period, over medial, frontal, and sensorimotor sensors. These results suggest that non-symmetric swaps may also capture situations in which a wrong association

between visual features is encoded during the initial perception and encoding of the items' features into WM, even

before memory retention [35,36], and/or in the early moments of WM maintenance. Based on the topology of the

differences, these swaps may derive from phase-coding variability in a different network of brain regions,

suggesting that they may thus derive from a different process. A caveat of this analysis is, however, that the number

of trials was limited for the two swap types. Future studies are required to demonstrate whether distinct sources of

swaps can be characterized on the basis of modulations in phase-coding variability.

4. Methods

4.1. Subjects

This study was performed in compliance with the Declaration of Helsinki on "Medical Research Involving Human Subjects", and was approved by the IRB ethics committee at the New York University (NYU) Abu Dhabi. Healthy human subjects from the community of NYU Abu Dhabi were recruited to participate in the study. Subjects participated as paid volunteers (50 AED/h) and gave written informed consent prior to the experimental sessions. All subjects had normal or corrected-to-normal vision. In a preliminary behavioral session, subjects performed the WM task presented in Fig. 1. Only subjects who made 'symmetric swaps' (see section 4.2) in at least 5% of the trials were invited for a following MEG session. Twenty-six subjects participated in the MEG session (age range 19–37 years, M=24.33, SD=4.57; 4 female; all right-handed). Previous papers reporting Phase Preservation Index have ranged from 6 to 12 subjects and typically do not report effect size [71–74]; we based our sample size on an estimated effect size of 0.5 and power of 0.8. Both behavioral and MEG sessions were characterized by the same experimental design and procedures (see below).

4.2. Stimuli, design, and procedures

Visual stimuli consisted of 3 different items (colored circles), each of which measured 0.55 degrees of visual angle (DVA) in diameter, presented at an eccentricity of 4.5 DVA (Fig. 1). Stimulus items were presented on a gray background for 0.2 s, followed by a retention interval (delay period) during which the screen was blank (2 s). A white fixation point (0.2 DVA) was presented at the center of the screen during both stimulus presentation and delay period. In a pre-fixation period (1 s before the beginning of each trial), the fixation point changed color to red when the subjects broke fixation. Trials in which the subjects broke fixation were removed from successive analyses.

Subjects were asked to remember the 3 stimulus items, and to hold in mind their colors and locations during the delay period. After memory delay, subjects were asked to report the location of each of the items (report period), which were sequentially cued by presenting their color at the center of the screen. Subjects had no control over the order in which the items were cued during the report period. During the report, a white circle (same size and eccentricity of the items) was also presented on the screen. Subjects were asked to respond by adjusting the position of this white circle by using a rotating dial, and pressing a button on a response box to confirm their response. The report period was self-paced and was followed by a feedback period (1 s), showing the cued items ringed by white circles and the subjects' responses ringed by black circles. After the feedback, a self-paced inter-trial-interval (1 s) was used to allow the subjects to control their blinking. Subjects had to press a button to move onto the next trial.

Each subject performed 500 trials of the WM task in one session, divided into 10 blocks (50 trials each). During stimulus presentation, the items were always presented on one single hemifield (5 blocks on the left, 5 blocks on the right; randomly ordered), never across hemifields. In each trial, the spatial location of each stimulus item was randomly sampled from a space that depended on the hemifield used for that block, obtained by excluding locations that were 10 DVA away from the vertical meridian, and using a minimum gap of 15 DVA. This was done to prevent physical overlap between items and to avoid swaps purely induced by misperception of spatial location. The color of each item was selected randomly from a color wheel of 180 color segments, using a minimum gap of 15 color segments between items to avoid swaps that were purely induced by colors misperception.

We used a maximum likelihood approach to assign to each response a likelihood of being a binding error (non-target) or a target response. The approach was based on a mixture model that comprised the probability of correctly reporting the target item, the probability of incorrectly reporting a non-target item, and the probability of responding randomly (guess) [29] (https://www.paulbays.com/toolbox/). This procedure allowed categorizing each trial on the basis of the subjects' responses: trials with a probability of correctly reporting the target item greater than 0.95 in all three responses were considered high-performance (HP) targets, while those with a probability of incorrectly reporting a non-target item greater than 0.70 in at least one response were considered swaps. The remaining trials were considered low-performance (LP) reports. After MEG data preprocessing (see section 4.3), the number of trials across subjects was in the range 16–439 for HP targets (M=92.81, SD=88.06), 22–196 for swaps (M=116.38, SD=47.79), and 44–328 for LP trials (M=198.73, SD=64.03). For a specific analysis on swaps (see section 2.4), these were further divided into symmetric (at least two non-target reports) and non-symmetric swaps (single non-target report). The number of trials for the two swap types was in the range 7–81 (M=46.50, SD=18.51) and 15–123 (M=69.88, SD=31.70) across subjects, respectively.

4.3. MEG data acquisition and preprocessing

The locations of the marker coils on each subject, as well as fiducial locations (including nasion–NAS, left and right Pre-Auricular points–LPA/RPA, left and right PreFrontal–LPF/RPF), were recorded before MEG data acquisition. The shape of the subject's head was also recorded using a Polhemus dual source handheld FastSCAN-II system (Colchester, VT, United States of America), collecting between 2,500 and 4,500 points per subject. MEG data were acquired continuously using a 208-channel axial gradiometer system (Kanazawa Institute of Technology, Kanazawa, Japan), with sampling frequency of 1,000 Hz, and applying an online low-pass filter (cutoff frequency 200 Hz). During recording, the subject was lying in the supine position, while performing the WM task. The visual stimuli were generated using MATLAB (The MathWorks Inc., Natick, USA) and projected onto a screen positioned at 85 cm from the face of the subject.

We applied a noise reduction procedure to the continuous MEG data using the rotated spectral Principal Component Analysis (rsPCA), as implemented in the MEG laboratory software MEG160 (Yokogawa Electric

Corporation, Tokyo, Japan), with block width of 5000 and 30 shifts, by using eight magnetometer reference

channels that were located away from the subject's head. We used the FastICA algorithm implemented in MNE-

Python [75] to decompose the MEG data by independent component analysis (ICA), and we removed the components

identified as eye blinks or heartbeat artifacts from the data. We defined trials using the time epoch from 1 s before

stimulus onset (pre-fixation period) to 2.7 s after stimulus onset. We then applied temporal demean and low-pass

filtering   (cutoff   frequency   140   Hz)   procedures   to   the   data   using   FieldTrip   [76]

(http://www.ru.nl/neuroimaging/fieldtrip/). Noisy trials and channels were automatically identified (high variance)

and rejected by visual inspection. We rejected trials showing residual ocular or muscular artifacts by visual

inspection. We interpolated the rejected channels using the average of all their neighbors. As a final step, we

switched the MEG channels between the right and left side in all trials in which the stimulus was presented on the

left hemifield. This procedure was performed to compound together the trials from both left and right stimulus

presentations, as if the stimulus presentation occurred always on the right hemifield.

4.4. Statistics

We performed all statistical comparisons with a cluster-based permutation approach [32], using a two-tailed

dependent t-test (df=25, $p<0.05$, alpha level distributed over both tails by multiplying the p-values with a factor of

2 prior to thresholding), 1,000 permutations, and $p_{perm}<0.05$ for the permutation test. We used the same settings in

every comparison, unless specified otherwise, while the space for clustering in each analysis is specified in the

sections below.

In each comparison between trial types, we estimated the effect sizes of the differences using Cohen's d

[77], after averaging the estimates separately per trial type, either over the time points and frequencies identified by

the significant cluster (Phase Preservation Index–PPI analysis; see section 4.5), or simply over time points

(frequency sliding variability–FSV analyses; see section 4.6–4.7).

4.5. Phase Preservation Index

497 We used the Phase Preservation Index–PPI [30] to estimate the consistency over trials of time-frequency

498 phase differences with respect to a reference phase, separately for each trial type (HP, LP, and swaps). Specifically,

499 PPI measures the intertrial consistency in phase differences as a function of time (t), with respect to a reference

500 time point (t_{ref}), at a specific frequency of interest f_0:

$$PPI(f_0, t) = \frac{1}{N} \left| \sum_{k=1}^{N} e^{i\phi^k(f_0, t_{ref}) - i\phi^k(f_0, t)} \right| \qquad (1)$$

For each trial ($k=1,..., N$) we derived the instantaneous phase $\Phi^k(f_0, t)$ using a Morlet wavelet transform (central

frequency parameter ω_0=6; zero-padding to solve edge effects problem) [78–80], in each MEG sensor. The frequency

of interest f_0 varied in the range 1–50 Hz (1 Hz steps). We used the delay onset as the reference time point (t_{ref}=0.2

s) for PPI estimation. Since PPI is sensitive to the number of trials, this was balanced between trial types (HP

targets and swaps) at the level of each subject. We performed this procedure by i) considering 100 combinations

of trials subsets in the type with more trials, to match the number in the type with less, ii) estimating PPI for each

subset, and finally iii) computing the median of PPI estimates across subsets. We estimated PPI only once in the

type with fewer trials. We then compared PPI between HP targets and swaps using a cluster-based permutation

approach (see section 4.4), over all time points in the delay period (0.2–2.2 s), frequencies (1–50 Hz), and MEG

sensors. Here, we defined neighboring sensors based on their 3D geometrical distance (max distance of 25 mm

for neighbors definition).

4.6. Frequency sliding (instantaneous frequency)

We employed the frequency sliding–FS [34] to capture the single-trial temporal fluctuations in oscillation

peak frequency. FS estimation relies on five core steps: (i) extract the phase angle time series at a certain frequency

(Fig. 6a); (ii) unwrap the phase angle time series (Fig. 6a); (iii) compute the first order derivative over time

(measure in rad); (iv) scale the measure by multiplying by the sampling frequency and dividing by 2π (measure

converted into Hz; Fig. 6b); (v) apply a median filter to attenuate neurophysiological noise spikes (Fig. 6b). We

employed a Morlet wavelet transform to extract the phase angle time series (step i), using ω_0=6 and zero-padding

[78–80]. We used an order of 10 and a maximum window size of 0.1 s for the median filter settings (step v), which

reassigned each time point to be the median of a distribution made from surrounding points [34].

Fig. 6 Estimation of frequency sliding–FS (instantaneous frequency). **a** Example of phase angle time series at 10

Hz, before (top, in black) and after (bottom, in blue) the unwrapping procedure. **b** Frequency sliding pre-filtered

(black) and after median filtering (red) for the phase angle time series shown in panel a.

To derive measures of FS variability (FSV) over time and over trials, we first derived single-trial FS

estimates at 10 Hz, separately for each trial type; then, we estimated the median between trials of the standard

deviation of FS over time points (alpha-FSV over time) and the median between time points of the standard

deviation of FS over trials (alpha-FSV over trials) on a sliding window of 0.4 s (i.e., four cycles of the 10 Hz

oscillation), with maximum overlap in time. We finally compared the two measures of alpha-FSV between trial

types using a cluster-based permutation approach (see section 4.4), over time points during the stimulus

presentation and in the delay period (0–2.2 s) and sensors, where neighboring sensors were defined based on their

3D geometrical distance (max distance of 25 mm for neighbors definition).

4.7. Source reconstruction

The solution of the MEG forward problem requires models of the head volume conduction, the sensors,

and the cortical sources. To derive these models, we estimated a transformation matrix (rotation and translation)

through an automatic procedure that aligned the five fiducials (NAS, LPA, RPA, LPF, and RPF) between subject-

specific Polhemus head shape and markers. We then used the transformation matrix to align the sensors' locations

to the head shape. We used the boundary element method (BEM) template provided in FieldTrip [76] for the head

volume conduction. We coregistered the template head model with the head shape, and successively refined the

transformation using a method that relies on fitting spheres to the head shapes (which allows deriving translation

and global scaling). We computed the head volume conduction based on the refined template head models, by

creating a single-shell model on the basis of the brain compartment, and taking its inner shell. As a final step, we

manually realigned the head model and head shape, and the registration was also checked manually. The source

model was constructed using template brain coordinates (12.5 mm resolution; 10 mm outward shift; 1,499

equivalent current dipoles), and then transformed into head coordinates. We computed the subject-specific lead

field matrix considering a forward operator with unconstrained orientation, where each source point was modeled

as three orthogonal equivalent current dipoles.

Once the lead field matrix was computed, we solved the MEG inverse problem using the linearly

constrained minimum variance (LCMV) beamformer [81]. We obtained an estimate of sensor-space covariance

matrix from the 1 s pre-fixation period. We extracted source-reconstructed time series from all source points using

a method based on singular value decomposition [82], which allows to derive single-trial time series for each source

point that best explain the variability in dipole orientations and signal strengths across trials for that point. After

source reconstruction, we derived measures of PPI, single-trial FS and FSV at 10 Hz for each source point,

following the same procedures used in the sensor-space analyses (see sections 4.5–4.6). We then compared PPI

and the two measures of FSV between HP targets and swaps using a cluster-based permutation approach (see

section 4.4), over time points either in the delay period only (0.2–2.2 s; for PPI) or in both stimulus presentation

and delay period (0–2.2 s; for alpha-FSV), and source points in source-space. We defined neighboring points based

on their distance in MNI space (max distance of 15 mm for neighbors definition).

4.8. Computational model

We used the attractor model of spiking neural networks described by Barbosa et al. [26] to generate simulated

field potentials from two networks implementing phase-encoded feature binding in working memory, in order to

test the ability of FS to detect synchronization changes corresponding to swap trials (Fig. 3). In brief, we built two

identical networks of excitatory and inhibitory integrate-and-fire spiking neurons, such that within each network

neurons were disposed on a ring and the connection strength of NMDA and AMPA receptors (for excitatory

connections) or $GABA_A$ receptors (for inhibitory connections) between any two neurons was defined by a

Gaussian function of their arc distance on the ring (standard deviation, or footprint of the connectivity, σ). Strong

recurrent excitation and larger footprint for inhibitory than excitatory connections facilitate tuned persistent

activity (a bump) after brief stimulation of a subset of neurons. In turn, when two bumps are sustained in one ring

network, broader inhibitory footprints impose out-of-phase synchronization of the two bumps (Fig. 3c–d).

Connections between excitatory neurons in one ring and neurons in the other ring were instead non-specific,

allowing for the arbitrary synchronization of bump pairs across the two networks. More specific details of the

network implementation can be found in Barbosa et al. [26], or in the code repository

(https://github.com/comptelab/binding). We computed the extracellularly recorded field potential for each of the

network by applying a Gaussian filter (σ=5 ms) centered on each spike from neurons in one half of the network

(simulating biases in representing the visual space), to obtain a time series for each network in each trial (Fig. 3c–

d).

4.9. Data availability

579   Analysis scripts, behavioral data, and preprocessed MEG data are available on the Open Science

580 Framework (https://osf.io/s5mcy/; https://osf.io/37hb6/). The computational model can be obtained from

581 https://github.com/comptelab/binding. Supplemental data can be requested via email to the lead author.

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Acknowledgements

This work was supported by the Swiss National Science Foundation (Grants P2FRP3_195083 and P500PB_214404 to M.F.P.), the National Institutes of Health (Grant R01MH063901 to M.D.), and the NYU Research Enhancement Fund and the NYUAD Center for Brain and Health, funded by Tamkeen under NYUAD Research Institute (grant CG012 to K.K.S.). A.C. and J.B. acknowledge support from the Spanish Ministry of Science, Competitiveness and Universities co-funded by the European Regional Development Fund (Refs: BFU 2015-65318-R, RTI2018-094190-B-I00). J.B. was supported by the Bial Foundation (ref: 356/19). This research was conducted using resources provided by the High Performance Computing Cluster and the Core Technology Platform at NYUAD. We would like to thank Vahan Babushkin for assistance with piloting and technical support.

Author Contributions

M.F.P. and K.K.S. conceptualized the study. K.K.S. and A.T. developed the experimental paradigm. A.T. collected the data. M.F.P. and A.S. preprocessed the data. M.F.P., A.S., A.T., J.B., A.C., and K.K.S. developed the analytical framework. M.F.P. and K.K.S. wrote the manuscript. A.S., J.B., A.C., and M.D. provided feedback and edited the manuscript.

Competing interests

The authors declare that the study has been conducted in the absence of any conflict of interest.

 Supplementary material:

 # Alpha phase-coding supports feature binding during

 # working memory maintenance

Supplementary Fig. 2 PPI estimates between HP targets and swaps. **a** Time course of PPI obtained from shuffling the signals over time in HP targets (left) and swaps (right) at different frequencies, highlighted by the black horizontal lines in the next panel. **b** PPI (shuffled signals over time): z-scores along frequencies for each time point in the delay, in HP targets (left) and swaps (right). **c** Time course of PPI corrected in HP targets (left) and swaps (right) at different frequencies, highlighted by the black horizontal lines in the next panel. **d** PPI corrected: z-scores along frequencies for each time point in the delay, in HP targets (left) and swaps (right).