**Exploring the role of cultivar, year and plot age in the incidence of grapevine**

**trunk diseases: insights from 20 years of regional surveys in France**

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**Abstract**

Grapevine trunk diseases cause yield losses and vine mortality in vineyards worldwide.

However, there have been few quantitative studies evaluating grapevine dieback on a large

spatial and temporal scale. Here, we consolidated and standardised databases from the 13 main

wine regions of France, compiling records of foliar symptoms associated with esca and Eutypa

dieback from 2082 plots and 36 cultivars over a 20-year period. This large dataset was used (1)

for quantitative analysis of the prevalence (number of plots with at least one symptomatic plant)

and incidence (percentage of symptomatic plants) of esca and Eutypa dieback; (2) to decipher

the effects of cultivar, year and plot age on both the prevalence and incidence of esca leaf

symptoms by temporal Bayesian modelling. Esca was present on a mean of $74 \pm 2\%$ plots

annually, with an incidence of $3.1 \pm 0.1\%$. Eutypa dieback occurred in $41 \pm 3\%$ of the plots, with

an incidence of $1.4 \pm 0.1\%$. Our modelling approach revealed that the cultivar had a significant

impact on the prevalence of esca, but not on its incidence when prevalence is greater than zero.

Esca prevalence remained stable, whereas esca incidence was higher than the mean value in six

of the years after 2012. We also found a significant non-linear effect of plot age, 10- to 30-year-

old plots significantly more susceptible, depending on the cultivar. This study clearly illustrates

the importance of considering extensive and continuous monitoring to improve our

understanding of the impact and evolution of crop diseases.

## I) Introduction

Perennial plant dieback is characterised by the premature and progressive death of branches, plant weakening and, ultimately, plant death. It has many possible causes, including environmental factors, especially drought (Allen et al. 2010; Cailleret et al. 2019; Hammond et al. 2022), pathogens (Ciesla and Donaubauer, 1994), and their interactions (Jurskis, 2005). Dieback is a well-known phenomenon in forest ecology, observed in various regions of the world (Allen, 2009; Hammond et al. 2022; Hartmann et al. 2018). By contrast, there have been no large-scale quantitative studies assessing dieback in perennial crops.

Dieback of grapevine (*Vitis vinifera* L.) is caused by various abiotic and biotic factors, including diseases affecting mature grapevine wood. Trunk diseases, such as esca, Eutypa dieback, and Botryosphaeriaceae dieback, are the most widespread globally (e.g., Guerin-Dubrana et al. 2019; Mugnai et al. 1999). These diseases cause considerable vine destruction, leading to yield losses, although these are often overestimated (Dewasme et al. 2022) and, often, the death of the vine, resulting in extensive damage to vineyards annually (Bertsch et al. 2009; Gramaje et al. 2018; Mondello et al. 2018). The current epidemic spread of trunk diseases in Europe can be traced back to the early 1990s (Mugnai et al, 1999; Reisenzein et al, 2000; Surico et al, 2000). In the first few years of the 21st century, until at least 2008, the incidence of grapevine trunk diseases, including esca in particular, appears to have increased (Bruez et al. 2013). This increase is a matter of great concern to vine growers, particularly in the wake of the sodium arsenate ban introduced in France in 2001, as this was the only effective treatment for trunk diseases. Several factors may be associated with variation in the incidence of trunk diseases. These factors include cultivar (Murolo and Romanazzi, 2014), year (Dewasme et al. 2022) and plot age (Fussler et al. 2008). Cultivar is a major factor underlying differences in the incidence of esca disease between vineyards and vine-growing regions, as shown by the considerable variability of the proportion of symptomatic plants per cultivar in France during the monitoring of 46 cultivars planted in a common garden vineyard over a period of seven years (Gastou et al. 2024), in Italy during the

monitoring of 67 cultivars for one year (Murolo and Romanazzi, 2014) and in Spain, during the

monitoring of 47 cultivars for three years (Chacón-Vozmediano et al. 2021). Indeed, certain

cultivars, such as Merlot, only rarely display esca leaf symptoms whereas others, such as

Sauvignon Blanc, are frequently affected. Moreover, strong interannual variability has been

observed within vineyards (Bruez et al. 2013; Calzarano et al. 2018). The age of the vines

significantly influences their response to biotic and abiotic stresses, determining their tolerance

or susceptibility. As a result, the incidence and expression of symptoms of stress increase

linearly with plot age (Pandey et al. 2015). The effect of vineyard age on esca incidence is

unclear. A few studies have reported an absence of correlation between plot age and esca

incidence (Bruez et al. 2013; Péros et al. 2008), but others have reported the opposite result

(Kovács et al. 2017), for plots aged from 10 to 21 years. By contrast, other studies have

suggested that the relationship between plot age and symptom expression is quadratic rather than

linear, with disease incidence highest at intermediate ages (Fulchin et al. 2019; Fussler et al.

2008). Incidence appears to be higher in vineyards of between 15 and 25 years of age than in

vineyards of other ages, as shown in a study of 22 cultivars (Fussler et al. 2008) and, more

recently, in a study of 5 cultivars (Fulchin et al. 2019). However, in these studies, age was

considered as a categorical ordered variable (e.g. 'young' if the plot was less than 7 years old and

'old' if more than 11 years old, Romanazzi et al. 2009, or 0-15, 15-25, 25-40, and over 40 years

old, Fussler et al. 2008). No study has ever addressed age as a fully continuous quantitative

variable, probably due to a lack of long-term, large-scale monitoring.

Distinguishing between the effects of year and plot age requires large-scale, long-term

monitoring, to ensure that the variables are not correlated, or at least no more than weakly

correlated. Large-scale monitoring is essential to improve pest surveillance for plants,

particularly over large spatial scales, and to facilitate the implementation of effective measures

for preventing the spread of pathogens and insect pests and for controlling epidemics (Carvajal-

Yepes et al. 2019; Mariette et al. 2023; Parnell et al. 2017). General surveillance encompasses

the collection and analysis of information on plant disease and plays a crucial role in the

detection and effective management of pathogens (Aguayo et al. 2021; ISPM 6 (FAO) 1997;

Parnell et al. 2017). The extensive monitoring of significant pathogens, including native ones, is

also crucial for obtaining spatial indicators of vineyard health, and for tracking temporal trends.

This information enables managers and policymakers to implement sustainable management

practices in vineyards. In France, a National Grapevine Trunk Diseases Survey was performed

between 2003 and 2008, to monitor grapevine trunk disease incidence and mortality and to

assess its significance for viticulture in seven vine-growing regions of France; this survey

included 12 cultivars and 329 vineyard plots (Bruez et al. 2013; Fussler et al. 2008). However,

the collection of more recent information over a longer period would be required to assess

progression of grapevine trunk disease levels over the last few years.

In this study, we collected and curated different databases from the 13 main wine regions of

France, to generate a unified database for foliar symptoms of esca and Eutypa dieback covering a

period of 20 years (2003 to 2022). This unified national database covers 2082 plots and 36

cultivars and was used to describe the prevalence (percentage of plots with at least one

symptomatic plant) and incidence (percentage of symptomatic plants per plot), as defined by

Nutter et al. (2006), of esca and Eutypa dieback over the different years, cultivars, regions, and

vineyard ages. Furthermore, the time-series data collected at plot scale were subjected to

modelling by the integrated nested Laplace approximation (INLA) method (Rue et al. 2009,

2017) to explore the effects of cultivar, year and plot age on both the prevalence and incidence of

esca whilst accounting for temporal dependences.

**II) Materials and methods**

**1)  French database of grapevine trunk disease observations**

The database contains 12587 observations of esca leaf symptoms and Eutypa dieback, in the main wine-producing regions in France, stored on the information system of the French Epidemiological Plant Health Surveillance Platform (ESV Platform). The observations were obtained from two sources: (i) regional surveys conducted from 2003 to 2022, and (ii) the historical "National Grapevine Trunk Diseases Survey", which tracked the progression of grapevine trunk diseases throughout France from 2003 to 2008 (Fussler et al. 2008; Grosman and Doublet, 2012; Bruez et al. 2013). The historical and regional surveys were conducted by experts from diverse public or private agronomic institutes or associations in each of the regions (see the Acknowledgements section).

Observations of esca leaf symptoms were obtained in 884 different municipalities (2.4 ± 2.5 plots per municipality, mean ± standard deviation, SD), in 49 provinces, 13 vine-growing regions, and 10 administrative regions. The dates of esca observation coincided with berry ripening, towards the end of August. Champagne was the most closely monitored vine-growing region, with data for 586 plots, whereas only 54 plots in Corse (also known as Corsica) were monitored. Six wine regions were not continuously monitored between 2003 and 2022: Corse, Jura & Savoie, Provence, the Sud Ouest (known as South West), Val de Loire, Vallée du Rhône (known as Rhone valley) and Vendée (Figure 1, Table S1, supplementary materials). Plots were monitored for a mean of 6 ± 4 years (± SD, not always consecutive) and the years with the most observations coincided with the "National Grapevine Trunk Diseases Survey" (2003-2008). Thirty-six different cultivars were monitored, with a mean of 58 ± 69 [range: 1 - 344] plots per cultivar. The most frequently represented cultivars were Chardonnay, Pinot Noir and Meunier (344, 193 and 187 plots, respectively). Three cultivars (Italia, Alphonse Lavallée and Sauvignon Gris) were monitored on only one plot each and were therefore excluded from the analysis. Moreover, only two plots were monitored in the Vendée wine region (Chenin cultivar), and these plots were also, therefore, excluded from the analysis. Finally, mean plot age was 27 ± 13 years,

and plot age ranged from 1 to 101 years (Figure 2). However, the date of plantation was not

recorded for 25% of the plots.

The most prevalent cultivars were Riesling in Alsace Lorraine (36%), Cabernet Sauvignon in the

Bordeaux region (also known as Bordelais, 34%), Pinot Noir in Bourgogne (known as

Burgundy, 38%), Chardonnay in Champagne (46%), Ugni Blanc in Charentes (99%),

Sciaccarellu in Corse (44%), Trousseau in Jura & Savoie (41%), Carignan in Languedoc (34%),

Cabernet Sauvignon in Provence (31%), Malbec in Sud Ouest (21%), Chenin in Val de Loire

(28%), and Grenache in Vallée du Rhône (30%) (Figure S1). Some cultivars were monitored in

more than one region: Fer Servadou, Gamay, Malbec, Muscat de Hambourg, Meunier, and Ugni

Blanc were each monitored in two regions; Cabernet Franc, Carignon, Cinsault, Grenache,

Merlot, Mourvèdre, Muscat Petits Grains and Pinot Noir were each monitored in three regions;

Chardonnay, Sauvignon Blanc and Syrah were each monitored in four regions; and finally,

Cabernet Sauvignon was monitored in five regions (Figure S1).

For 63% of the plots monitored for esca, annotation was also available for Eutypa dieback

symptoms between April and May (Table S1). In total, 7073 observations were available for this

disease, spread over 1310 plots located in 592 municipalities (2.2 ± 2.4 (mean ± SD) plots per

municipality), 39 provinces, 10 wine regions, and 7 administrative regions. All cultivars other

than Meunier, Sciaccarellu, Vermentinu, and Niellucciu were monitored for Eutypa dieback,

with data available only for esca leaf symptoms for these exceptions.

Finally, mortality was recorded, but the methods used to score dead plants were not

homogeneous across regions. We therefore discarded these data from the analysis. Similarly,

information on rootstock, pruning technique, young replanted vines, and apoplexy (total

dehydration of the canopy) was scarce and inconsistently reported, and such data were not,

therefore, considered in this analysis.

**2) Estimations of the prevalence and incidence of foliar symptoms of esca and Eutypa**

**dieback**

The presence of esca or Eutypa dieback symptoms on leaves was recorded on a defined number

of vines in each plot. A mean of $420 \pm 530$ [200 - 1719, 0.025 quantile to 0.975 quantile] vines

were observed per plot (see supplementary materials, Table S2). One plot of Cabernet Sauvignon

was monitored in its entirety as part of an experiment, accounting for the large mean number of

vines observed for this cultivar (Table S2). The definitions of disease prevalence and incidence

proposed by Nutter et al. (2006) were used. Disease prevalence was calculated, for a given year,

as the percentage of plots on which esca, or Eutypa dieback (depending on the disease

considered for the analysis) was observed on at least one plant. Disease incidence was calculated

for each plot in a given year by dividing the number of plants presenting esca foliar symptoms

(as described by Lecomte et al. 2012) or Eutypa dieback foliar symptoms (as described by

Sosnowski et al. 2007) by the number of plants monitored in the plot concerned.

**3) Deciphering the effects of cultivar, year, and plot age on the prevalence and**

**incidence of esca leaf symptoms**

A statistical model was developed to estimate the effect of cultivar, year and plot age on the

prevalence and incidence of esca leaf symptoms at plot scale. We optimised model

identifiability, by considering only observations (for one plot in a given year) from plots no more

than 50 years old (older plots excluded), for which at least five years of observation were

available (not necessarily consecutive), and cultivars for which at least 300 observations were

available. The application of these criteria resulted in a database of 5161 observations (from the

12587 observations initially identified) spread over 12 cultivars, each observed on between 24

and 85 plots (smallest number for Pinot Auxerrois/Riesling, and largest number for Cabernet

Sauvignon). The number of plots monitored per year varied with cultivar, with a mean of five

(Chenin) to 26 (Ugni Blanc) plots (Table 1). Mean plot age ranged from 18 to 32 years (for

Chenin and Pinot Auxerrois, respectively; Table 1). Mean plot age tended to increase from 2003

to 2013, remaining stable thereafter (Figure 3).

The model consisted of two hierarchical structured components: one describing the prevalence of

esca at plot scale, and the other the incidence of esca conditional on its occurrence (i.e. given that

the plot prevalence is 1). For the prevalence component, the response variable was the presence

(denoted 1) or absence (denoted 0) of leaf symptoms in a given plot in a given year. A Bernoulli

distribution with a logit link was used for this component. For the incidence component, the

response variable was the number of vines with esca leaf symptoms divided by the total number

of vines observed. For this component, we aimed to establish a binomial model conditional on

there being at least one success, defined as the presence of at least one vine with esca symptoms.

Thus, a zero-inflated binomial distribution parameterised to exclude zero with a logit link was

used for this component in R-INLA. We describe below the "full" model including all

explanatory variables. Intermediate models with fewer explanatory variables were also estimated

and compared to the full model on the basis of two information criteria: the deviance information

criterion and the widely applicable information criterion (DIC and WAIC) (Table 2). Both the

prevalence and incidence components included cultivar and plot identity as independent random

effects. Both components also included the year as an autoregressive process of order 1 (AR1).

These effects account for interannual dependence arising from underlying meteorological or

environmental factors and agricultural practices affecting esca symptoms. Finally, the incidence

component included plot age as an AR1 process specific to each cultivar.

More specifically, let $y_i$ denote the esca prevalence of observation $i$ ($i = 1, …, 5161$), where the

disease is either observed (value $y_i = 1$) or not observed (value $y_i = 0$). Let $n_i$ denote the number

of plants with esca symptoms among the $n_i^{tot}$ monitored for this same observation $i$. We will

consider the following explanatory variables associated with observation $i$. First, the variable

*plot(i)* denotes the identity of the plot (554 levels for prevalence and 546 levels for incidence) on

which observation $i$ was performed; *cultivar(i)* denotes the cultivar (12 levels, see Table 1) on

the plot corresponding to observation $i$. The effects of these two variables were modelled as

independent random effects. The variable *year(i)* for prevalence denotes the year of monitoring

(indexed by $t = 2003, \dots, 2022$). Finally, *age(i)* denotes plot age, a variable considered only for

the incidence component (indexed by $z = 1, \dots, 49$). The effects of *year(i)* for both prevalence

and incidence and *age(i)* for incidence only were modelled as first-order autoregressive

dynamics (i.e. AR1 process), the parameters of age being specific to each cultivar.

The first component of the model can be written as follows:

$$y_i \sim Bernoulli(\pi_i), \ i = 1, \ \dots, \ 5161$$

$$logit(\pi_i) = \beta_0 + \mu_0(n_i^{tot}) + \alpha_0[plot(i)] + \gamma_0[cultivar(i)] + \delta_0[year(i)]$$

$$\alpha_0[plot(i)] \sim Normal\big(0, \ \sigma_{plot0}^2\big)$$

$$\gamma_0[cultivar(i)] \sim Normal(0, \sigma_{cultivar0}^2);$$

$$\delta_0[2003] \sim Normal\big(0, \sigma_{year0}^2(1 - \rho_{year0}^2)^{-1}\big)$$

$$\delta_0[t] = \rho_{year0}\big[\delta_0[t-1]\big] + \varepsilon_0[t], \ t = 2004, \dots, 2022$$

$$\varepsilon_0[t] \sim Normal\big(0, \sigma_{year0}^2\big)$$

225  ZIB[+] is a modified version of the zero-inflated binomial distribution supported from 1 to $n_i^{tot}$

226  (exclusion of the 0 value), $S$ being the subset of index $i$ associated with positive results for esca

227  prevalence (4476 observations). The second component of the model can therefore be written as

228  follows:

$$n_i|y_i = 1 \sim ZIB^+(n_i^{tot}, \ \omega_i), \ i \in S$$

$$logit(\omega_i) = \beta_1 + \alpha_1[plot(i)] + \gamma_1[cultivar(i)] + \delta_1[year(i)] + \tau_{cultivar(i)}[age(i)]$$

$$\alpha_1[plot(i)] \sim Normal\left(0, \sigma_{plot1}^2\right)$$

$$\gamma_1[cultivar(i)] \sim Normal\left(0, \sigma_{cultivar1}^2\right)$$

$$\delta_1[2003] \sim Normal\left(0, \sigma_{year1}^2(1 - \rho_{year1}^2)^{-1}\right)$$

$$\delta_1[t] = \rho_{year1}\left[\delta_1[t-1]\right] + \varepsilon_1[t], \ t = 2004, \dots, 2022$$

$$\varepsilon_1[t] \sim Normal\left(0, \sigma_{year1}^2\right)$$

$$\tau_{cultivar1}[1] \sim Normal\left(0, \sigma_{age1,cultivar}^2(1 - \rho_{age1,cultivar}^2)^{-1}\right), \ cultivar = 1, \ \dots, \ 12$$

$$\tau_{cultivar1}[z] = \rho_{age1,cultivar}\left[\tau_{cultivar}[z-1]\right] + \varepsilon_{age1,cultivar}[z], \ z = 2, \dots, 49$$

$$\varepsilon_{age1,cultivar}[z] \sim Normal\left(0, \sigma_{age1,cultivar}^2\right)$$

In the model, $\beta_0$ and $\beta_1$ are the intercepts of prevalence and incidence, respectively. The

parameter $\mu_0$ takes into account the different numbers of plants monitored per plot (note that $n_i^{tot}$

was standardised to ensure robust estimation). The autoregressive parameters $\rho_{year0}$, $\rho_{year1}$,

$\rho_{age1,cultivar}$ (all of which have absolute values < 1) correspond to the correlation coefficients for

the prevalence ($\rho_{year0}$ for year) or incidence ($\rho_{year1}$ for year and $\rho_{age1,cultivar}$ for plot age for a given

cultivar) of esca between two consecutive time points. They indicate the extent to which the

temporal trend in a given time step $t$ depends on the trend from the previous time step $t$-1.

Typically, estimates close to 0 indicate the absence of a significant temporal trend. By contrast,

the closer the estimates are to 1, the more similar the temporal trends between two consecutive

time steps (Zuur et al. 2017). The parameters $\sigma_{year0}^2$, $\sigma_{year1}^2$, $\sigma_{age1,cultivar}^2$ define the precision of

the AR1 process, with lower values corresponding to smoother temporal trends. The model was

fitted in a Bayesian framework by the integrated nested Laplace approximation (INLA) method

(Rue et al. 2009, 2017). Bayesian inference requires the specification of prior distributions for

the model parameters and hyperparameters, and we used the default internal vague priors

recommended in R-INLA. Specifically, the different $\sigma_{plot0}^2$, $\sigma_{cultivar0}^2$, $\sigma_{plot1}^2$, $\sigma_{cultivar1}^2$ values for

the independent random effects and the $\sigma^2_{year0}$, $\sigma^2_{year1}$, $\sigma^2_{age1,cultivar}$ values of the AR1 process are

assigned a log-gamma distribution with parameters 1 and 0.00005. For the cultivar and year

effects, we then calculated the probability of direction (*pd*), which varies between 0.5 and 1 and

indicates the probability that a parameter, as defined by its posterior distribution, is either strictly

positive or negative. This *pd* is the proportion of the posterior distribution that has the same sign

as the median of the posterior distribution (Makowski et al. 2019), and can be used to identify

trends in the distribution of the parameter. Finally, we compared fitted and observed values to

evaluate model fit. For the probability of presence, we used ROC curve analyses (Hoo et al.

2017, PRROC R package, Grau et al. 2015). For disease incidence, we calculated Pearson's R

coefficient for the correlation between the mean of the fitted values and observed incidence

(function *cor.test* package *stats*).

**III) Results**

**1) Prevalence and incidence of leaf esca and Eutypa dieback symptoms in France**
**between 2003 and 2022, by cultivar and wine region**

**1.1) Variability of the prevalence and incidence of esca leaf symptoms**

Only 10% of the 2075 plots monitored over the 2003-2022 period contained no symptomatic

vines (1861 of 2075 plots contained at least one vine presenting symptoms). The annual

prevalence of esca ranged from 54% in 2010 to 84% in 2017, with a mean of 74 ± 2% (mean ±

SEM, standard error of the mean) plots containing at least one symptomatic plant (Figure 4a).

During the monitoring period, for all years considered together, the prevalence of esca in the

regions ranged from 100% of the plots containing at least one vine with esca (Corse and Jura &

Savoie, 54 and 85 plots, respectively) to 78% (Champagne, 581 plots) (data not shown).

Including plots without esca (prevalence = 0), the mean annual incidence of esca over the 2003-

2022 period was 3.1 ± 0.1% (mean ± SEM). The lowest mean annual incidence of esca was

observed in 2003 (1.6 ± 0.1%, mean ± SEM), whereas the highest incidence was that observed in

2012, 2014, and 2017, when 4% of plants were symptomatic (Figure 4a). The annual incidence

of esca increased between 2003 and 2008 (from 1.6% to 3.9%) and then fluctuated between

2.2% and 4% until 2022. If we excluded plots with no vines presenting esca symptoms in a given

year (prevalence = 0), the mean annual incidence increased to 4.2 ± 0.1% (mean ± SEM, Figure

S2).

We observed considerable variability in the incidence of esca foliar symptoms between cultivars.

The mean incidence by cultivar ranged from 0.6% to 10.6% (Figure 5a). The Trousseau cultivar

had the highest incidence, with a mean of 10.6 ± 0.6% of the vines symptomatic in the 30 plots

monitored, followed by Savagnin, with a mean incidence of 8.5 ± 0.6% in 30 plots, and Ugni

Blanc with a mean incidence of 6.4 ± 0.3% in 72 plots. Incidence was lowest for Meunier, for

which a mean of 0.6 ± 0.03% symptomatic vines per plot was obtained for 187 plots, followed

by Pinot Noir, with a mean incidence of 0.6 ± 0.04% in 193 plots, and Syrah, with a mean

incidence of 0.6 ± 0.1% in 24 plots.

The incidence of esca also varied between wine regions, the highest incidence being recorded in

Jura & Savoie (mean of 8.5 ± 0.4% of vines with esca leaf symptoms in 85 plots; Figure 5a(i)).

Esca incidence was lowest in the Champagne wine region, with a mean of 0.5 ± 0.01% affected

vines and the largest number of plots monitored (581; Figure 5a(i)). The incidence of esca by

year, region and cultivar was not correlated with the number of plots monitored (Pearson's r, $p >$

0.13)

**1.2) Variability of the prevalence and incidence of Eutypa dieback**

In total, 1296 plots were monitored for Eutypa dieback over the 2003-2022 period, and 31%

were Eutypa-free (895 of the 1296 plots contained vines displaying Eutypa dieback). The annual

percentage of plots with Eutypa dieback (i.e, the prevalence) ranged from 20% in 2016 to 61% in

2005, with a mean value of 41 ± 3% (mean ± SEM) (Figure 4b). During the monitoring period,

the regional prevalence of Eutypa dieback ranged from 96% in Charentes (72 plots) to 40% in

the Sud Ouest region (127 plots).

If plots without Eutypa dieback (prevalence = 0) were included in the calculation, then the mean

observed incidence of Eutypa dieback was 1.4 ± 0.1% (mean ± SEM) (Figure 4b). The incidence

of Eutypa dieback tended to decrease from 2003 to 2022 (Figure 4b). Specifically, we observed a

decrease from 2003 to 2008 and then a steady decrease since 2013 (except for 2020). The

incidence of Eutypa dieback was highest in 2003, and mean incidence values of up to 3% were

obtained. In addition, 2017 and 2022 were the years in which the incidence of Eutypa dieback

was lowest, with a mean value of 0.5%. Finally, if plots in which none of the vines were affected

by Eutypa dieback (prevalence = 0) in a given year were excluded, the mean incidence increased

to 3.3 ± 0.1% (mean ± SEM, Figure S2).

The variability of Eutypa dieback incidence between cultivars was high, with values ranging

from 0% to 9.5% (Figure 5b). Ugni Blanc had the highest incidence, with a mean of 9.5 ± 0.5%

symptomatic vines per plot for the 72 plots monitored, followed by Chenin, with a mean of 2.3 ±

0.2% symptomatic vines per plot in 60 plots and Cabernet Sauvignon with a mean of 1.5 ± 0.1%

in 143 plots. Chasselas had the lowest percentage of affected vines, with a mean of 0% per plot

for the 25 plots monitored, followed by Savagnin with a mean value of < 0.1% in 30 plots, and

Riesling, with a mean of 0.1 ± 0.01% in 43 plots. The region with the highest Eutypa dieback

incidence was Charentes, with a mean of 9.5 ± 0.5% vines affected. Languedoc and Alsace

Lorraine were the regions with the lowest incidence of Eutypa dieback, with a mean of 0.3 ±

0.02% affected vines (Figure 5b(ii)). The incidence of Eutypa dieback by year, region and

cultivar was not correlated with the number of plots monitored (Pearson's r, $p > 0.30$). Finally,

there was a weak but significant correlation between the incidences of esca and Eutypa dieback

assessed by plot/year observation (Pearson's r = 0.08, $p < 0.001$).

**2) Deciphering the effects of cultivar, year, and plot age on the prevalence and incidence of esca leaf symptoms**

The two model selection criteria considered (DIC and WAIC) supported the full model (Table 2), suggesting that all the explanatory variables considered had a significant impact on esca dynamics. We therefore consider only the full model hereafter. The model fit was satisfactory for both the prevalence and incidence components. Specifically, the area under the ROC curve (AUC) for esca prevalence was 0.93. Pearson's r for the correlation between adjusted and observed values of esca incidence was 0.83 ($p < 0.001$).

The model included the number of vines monitored as a covariable to account for the sampling effort in each plot used to determine incidence. As expected, the mean effect was positive (0.33), confirming that increasing the number of plants inspected in a plot increases the likelihood of at least one plant being symptomatic. The associated probability of direction was 0.95 ($\mu_0$, Table 3). The identity of the plot had a significant impact on both the prevalence and incidence of esca symptoms. For prevalence, 13% of the individual plot effects were negative (i.e. associated with a q-97.5% < 0), < 1% were positive (q-2.5% > 0) and the remaining 87% were associated with individual effects including zero (the mean value) in their 95% credible intervals. For incidence, 25% of the plots were negative, 30% were positive and the remaining 45% were associated with individual effects including zero in the 95% credible interval. These values correspond to the percentage of plots characterised by an esca prevalence or incidence lower than, higher than or not significantly different from the overall mean, respectively.

Cultivar had a significant impact on the prevalence of esca symptoms but a much weaker effect on esca incidence. Indeed, cultivar was significant if WAIC was used (64436 without the cultivar effect vs. 64418 for the full model) but not with the DIC (37835 without the cultivar effect vs.

37842 for the full model, Table 2). The individual effects of each of the 12 cultivars considered

are shown in Figure 6. Chardonnay, Merlot and Pinot Noir all had a prevalence of esca below the

overall mean (mean = 0, q-97.5% < 0) whereas Gewurztraminer and Cabernet Sauvignon had a

prevalence above the overall mean (q-2.5% > 0) (Figure 6a). No such differences were observed

for the incidence of esca. The 12 estimated 95% credible intervals included zero, meaning that

esca incidence did not differ between cultivars, provided that esca was observed (prevalence > 0;

Figure 6b). However, the probability of direction (*pd*) showed that cultivars such as

Gewurztraminer, Riesling and Melon had a strong tendency to have a high incidence whereas

Merlot and Chardonnay had a low incidence (*pd* > 0.90, Figure 6a, b).

Unlike the effect of cultivar, the effect of year did not affect the prevalence of esca symptoms,

but their incidence. However, the 95% credible interval [0.25, 0.89] of the hyperparameters $\rho_{year0}$

of the AR1 process modelling the effect of year on esca prevalence suggests that esca prevalence

is strongly correlated across successive years. No such significant correlation was detected for

esca incidence (the 95% CI of $\rho_{year1}$ was [-0.20, 0.71]) (Table 3; Figure 6c, d). An analysis by

year showed that the prevalence or incidence (if esca was recorded in the plot) of esca was

significantly lower than the mean value (=0) in the first year (2003) and that esca incidence was

lower in 2022. The prevalence and incidence of esca tended to be below the mean value in 2004,

as was the incidence in 2005 (probability of direction, *pd* > 0.95). In 2017, the prevalence and

incidence of esca were significantly higher than the mean. Moreover, esca incidence has

remained above the mean value since 2012, including the years 2013, 2014, 2016, 2017, and

2021. However, *pd* values showed that between 2010 and 2012, prevalence tended to be high (pd

> 0.92, as shown in Figure 6c, d).

The hyperparameters $\rho_{age1,cultivar}$ of the AR1 process, modelling the effect of plot age on esca

incidence for each cultivar, were strictly positive, suggesting a positive correlation between plot

age and incidence (from 0.43 [0.09, 0.71] for Gewurztraminer to 0.95 [0.87, 0.98] for Cabernet

Sauvignon, mean [0.025 quantile, 0.975 quantile], Table 3). A pattern of the effect of age on esca

incidence common to all cultivars is emerging. Indeed, in most cultivars, esca symptoms peak at

an intermediate age. With the exception of Merlot, all cultivars had an incidence of esca above

the mean value for plots aged between 10 and 40 years (in red, Figure 7). Conversely, the

youngest plots (under 10 years old) sown with these cultivars had an esca incidence below the

mean value (in blue, Figure 7). Finally, Cabernet Franc, Cabernet Sauvignon, Gewurztraminer,

Merlot and Pinot Noir had incidences significantly below the mean value after the age of 30

years. Overall, when age had a significant positive effect on esca incidence (when esca

prevalence $> 0$), esca occurred at ages between 11 and 45 years, with a peak of susceptibility

between 15 and 39 years, depending on the cultivar (Figure 8). Even though the cultivars with

the lowest incidence had a narrower age range for peak incidence (Cabernet Franc, Cabernet

Sauvignon, Chardonnay, Pinot Noir), a pattern remained in the relationship between cultivar

susceptibility (from left to right, according to modelling, Figures 6 and 8) and the age ranges for

susceptibility to esca and peak incidence.

**388    IV) Discussion**

We compiled and homogenised regional and historical databases to obtain a national

quantification of the prevalence and incidence of foliar symptoms of esca and Eutypa dieback in

France between 2003 and 2022. Despite considerable variability between years and cultivars, the

incidence and prevalence of esca leaf symptoms appear to be stable when looking at the raw

data, whereas the incidence of Eutypa dieback tended to decrease over the 2003-2022 period. We

used a hierarchical Bayesian model to decipher the responses of two components — the

prevalence (percentage of symptomatic plots, i.e. plots with at least one symptomatic plant) and

the incidence (percentage of symptomatic plants per plot, when prevalence $> 0$) of esca disease

— to the effects of year, cultivar, and plot age. This framework is compatible with the objectives

of surveillance programmes in applied plant pathology, which often aim to estimate disease

occurrence at national, regional, and local scales. The framework developed here takes into

account disease occurrence at two nested spatial scales (a set of plots within each agricultural

region and a set of plants within each plot). In our case study, focusing on the prevalence of esca

leaf symptoms is particularly relevant when comparing a wide range of cultivars with different

levels of susceptibility, including some for which no leaf symptoms were reported on a large

number of plots. Once the presence of leaf symptoms is confirmed in a plot, analyses of the

variation of incidence are then required to identify the factors driving leaf symptom expression,

as esca is characterised by fluctuating expression between years and plots. This approach showed

that since the incidence of esca leaf symptoms was significantly higher than the mean value in

six of the years since 2003, all occurring after 2012. We also detected a significant effect of

cultivar on the prevalence of esca, but not on its incidence (when prevalence $> 0$). Finally, we

found a significant effect of plot age on the incidence of esca leaf symptoms, with, in particular,

a peak at intermediate ages between 15 and 39 years, depending on the cultivar.

This study highlights the strong annual variability of esca incidence. Analysis of the whole

database showed an increase in the annual mean incidence of esca from 2003 (1.6%) to 2008

(3.9%). This trend was already reported by Bruez et al. (2013). However, our longer-term

monitoring revealed that this trend ceased after 2008. More specifically, the upward trend seems

to have continued after 2008 for certain cultivars, such as Ugni Blanc and Riesling in particular,

and their associated regions (Figure S3 and S4), whereas the opposite was observed for

Chardonnay and Meunier (Figure S4). A statistical analysis of the 12 cultivars with the largest

numbers of observations revealed that esca prevalence did not differ from the mean value in

most years (except 2003 and 2017), whereas esca incidence values significantly higher than the

mean were obtained for several years after 2012. When esca is observed in a plot, foliar

symptoms may be expressed over a number of years. The differences between years may reflect

climatic variation between years. Indeed, the lowest esca incidence was recorded in 2003, a year

characterised by a very intense drought and heatwave (Chuine et al. 2004). The lowest mean

incidences seem to coincide with hot years (e.g., 2015, 2019, 2022,

https://climate.copernicus.eu/europe-continued-swelter-july, site visited on February 22$^{nd}$, 2024).

Changes in temperature and rainfall can affect pathogen growth; it has been suggested that

higher maximum temperatures between June and July are associated with a lower incidence of

esca leaf symptoms, whereas the incidence of esca leaf symptoms seems to be positively

correlated with precipitation levels, as high rainfall levels during July tend to intensify leaf

symptom development (Calzarano et al. 2018; Serra et al. 2018). Indeed, water availability has

been shown to affect esca leaf symptom development as, under control conditions, intense

drought inhibits esca leaf symptom onset, suggesting an antagonistic interaction between drought

and esca pathogenesis and a key role for plant water status and climatic conditions (Bortolami et

al. 2021). Vascular diseases with similar characteristics to esca are rare. One example is provided

by a trunk disease of kiwifruits in which there is a significant relationship between temperature

and leaf symptoms (Di Marco and Osti, 2008).

An analysis of the whole database suggested that, unlike esca, Eutypa dieback decreased steadily

in incidence during the 2003-2022 period. These patterns were observed across the most

observed cultivars, years, and wine regions (Figure S5 and S6). The difference between the

development of esca and Eutypa dieback symptoms over time may partly reflect the nature of

these diseases. Esca is generally caused by a complex community of fungi, whereas Eutypa

dieback is caused by a single fungus, *Eutypa lata*. Climate can affect the expression of Eutypa

dieback (Sosnowski et al. 2007), but its gradual decrease in frequency over the years suggests

that climate has a relatively limited impact. The observed gradual decrease in Eutypa dieback

frequency may be due to the use of effective control methods, such as appropriate pruning,

wound protection and the removal of infected wood from vines (Lecomte et al. 2006), or to a

decrease in pathogen aggressiveness (Molyneux et al. 2002). In any case, Eutypa dieback has

been relegated to the status of a secondary disease, attracting less attention from field

technicians, as indicated by the decrease in the number of plots monitored for this disease over

the years (Figure S5 and S6; except for Ugni Blanc).

Our findings reveal considerably variability in the incidence of esca between the 36 cultivars

monitored in the whole database. Cultivar differences in esca incidence have already been

documented in France (Bruez et al. 2013; Gastou et al. 2024), Italy (Murolo and Romanazzi,

2014) and Spain (Chacón-Vozmediano et al. 2021). In these studies, Cabernet Sauvignon and

Sauvignon Blanc are generally identified as much more susceptible than Merlot and Chardonnay

(Murolo and Romanazzi, 2014; Chacón-Vozmediano et al. 2021; Gastou et al. 2024). However,

our statistical analysis revealed a lack of significant difference between the cultivars if esca

incidence was calculated only for plots with at least one symptomatic plant. Interestingly, the

prevalence of esca was significantly higher than the mean for two cultivars and lower than the

mean for three cultivars. Specifically, vineyards planted with Cabernet Sauvignon and

Gewurztraminer have a higher than average probability of having plants displaying esca

symptoms; conversely, this probability is below average for Chardonnay, Merlot, and Pinot Noir.

The cultivars with the highest prevalence of esca also had the highest estimated incidence values

(Figure 6). Only Cabernet Sauvignon displayed contrasting levels of esca prevalence and

incidence, as most of the plots had at least one symptomatic plant but an intermediate incidence.

Several confounding factors could potentially cause the cultivar effect highlighted in our

analysis. Typically, some viticulture practices (e.g. the use of certain rootstocks or types of

pruning) may be specific to a particular cultivar and wine region. Such factors are known to

affect susceptibility to esca (Lecomte et al. 2018). It was not possible to investigate these

confounding factors here, but the observed differences in esca incidence between the cultivars

studied here were correlated with the range of susceptibility of the same cultivars quantified over

several years in a common garden vineyard (Gastou et al. 2024).

Our results revealed a major impact of plot age (estimated as the mean age of the plants in the

plot) on esca incidence, which was lower in younger vineyards (less than 10 years old) than in

older ones. This age effect has also been observed in several field monitoring programmes

(Fulchin et al. 2019; Fussler et al. 2008; Pollastro et al. 2000). The underlying mechanisms

remain to be explored. We could hypothesize that certain traits known to differ with plant age,

such as plant metabolism, wood properties, and pathogen or endophyte communities, are

involved. As suggested by Fischer and Peighami-Ashnaei (2019), older plants might be more

susceptible to disease due to the wounds they have accumulated over time, as cultivated vines

are pruned annually. Older plants may therefore have undergone more infection cycles than

younger plants. In addition, older plants tend to accumulate more endophytes in their tissues than

younger plants (Dissanayake et al. 2018). They also have different fungal communities (Bruez et

al. 2016) and may undergo microbiome modifications during the course of their lifetime

(Bettenfeld et al. 2020; Fournier et al. 2022). Indeed, many endophytic fungi are pathogens and

can induce plant dieback (e.g., Úrbez-Torres et al. 2009).

Interestingly, the effect of age on the incidence of esca differed between cultivars. There was a

trend towards a relationship between cultivar susceptibility (i.e. esca incidence) and the age

range during which the cultivar was most susceptible (i.e. highest incidence), the least

susceptible cultivars had a narrower age range for susceptibility. In four of the 12 cultivars

studied, esca incidence decreased significantly once a certain age was reached, this age

depending on the cultivar but generally being greater than 30 years. This pattern may reflect a

higher rate of plant mortality in older plots, with the most tolerant plants remaining, the others

generally being replaced by new replanted vines. Indeed, in our database, we observed a slight

gradual ageing of the plots from 2003 to 2014 and an increase in the diversity of plot age

(monitoring of new younger plots to replace the older ones) starting in 2015 (Figure 3). If we

wish to study the effect of vine age on the incidence of esca more precisely, we need to be able

to collect the age of the vines rather than the age of the plot (based on the year of first

plantation). Alternatively, ontogenic resistance (Ficke et al. 2002) could be involved, leading to a

lower susceptibility to the development of esca leaf symptoms in the oldest vines, but little is

known about this process in wood diseases. Secondary metabolism may change qualitatively or

quantitatively with plant ageing (Haffner et al. 1991), potentially accounting for the persistence

of long-lived perennial plants (Cui et al. 2024). With regard to other diseases, a higher incidence

of Botryosphaeriaceae (Carlucci et al. 2013; Gubler et al. 2005) and *Phaeoacremonium sp.*

(Carlucci et al. 2013, 2015) has been observed in older grapevines and olive trees. Similarly, an

increase in branch canker and dieback incidence with the age of the tree has been observed in

avocado orchards (Valencia et al. 2022). By contrast, for kiwifruit (Di Marco and Osti, 2008), no

correlation was found between plot age (based on the year of first plantation) and the percentage

of symptomatic plants, although the age range considered was rather small. However, there have

been no studies of a possible non-linear effect of age on wood diseases in perennial crops other

than vines. Clearly, exploring the underlying mechanisms of wood disease resistance or

tolerance according to plant age is a promising research avenue that could improve our

understanding of the pathogenesis of such diseases.

By constituting and unifying large spatial and temporal surveillance databases, as in this study, it

is possible to visualise and statistically quantify the long-term progression of wood diseases.

However, some biases associated with the selection of the plots monitored must be taken into

account. In particular, in some wine regions, plots planted with particular cultivars were selected

on the basis of their susceptibility to disease rather than as a representative sample of the

cultivars in the region. The use of different regional strategies may have decreased the

representativeness of the data. This bias would result in a low or high estimated incidence in the

region, distorted by the selection of the cultivars monitored. Thus, in such surveillance databases,

incidence at regional levels should be interpreted with caution, with particular attention paid to

the regional sampling strategy. These potential biases do not call into question the relevance and

urgent need for the combination and unification of different regional datasets to develop a

526    national epidemiological surveillance programme for complex diseases, such as esca, and to

527    make it possible to perform additional statistical analyses in both space and time. Pursuing

528    efforts to develop surveillance initiatives and improve both the quantity and quality of data is a

529    crucial challenge in the context of global changes. Moreover, climate projections may have far-

530    reaching long-term effects that can only be understood in the light of such long-term surveys.

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**Competing interests**

The authors declare that they have no competing interests.

**Data availability statement**

We agree to archive an anonymized version of the trunk disease dataset in the public repository

'INRAE dataverse' if the manuscript is accepted.

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**Figures**

Figure 1: Map of the French wine-growing regions showing the locations of the 884

municipalities in which plots were monitored for esca leaf symptoms (black dots). These

municipalities are located in 49 provinces, 13 vine-growing regions, and 10 administrative

regions. Eutypa dieback was also monitored in 592 of these municipalities in 39 provinces, 10

wine regions, and 7 administrative regions. Note that Bourgogne, Vallée du Rhône, Corse, Sud

Ouest and Bordelais regions are internationally known as Burgundy, Rhone Valley, Corsica,

South West, and Bordeaux, respectively.

Figure 2: Number of plots or observations (plot-year pairs) monitored in France for grapevine

trunk diseases between 2003 and 2022 and recorded in the database, by year (a), wine region (b),

plot age (c), and *Vitis vinifera* cultivar (d). An observation corresponds to one plot monitored in

a given year. The red dotted line in (c) represents the mean age.

772    Figure 3: Distribution of plot age (range: 1 to 49 years) for each year for the 12 *V. vinifera*

773    cultivars used in the modelling approach.

Figure 4: The mean prevalence and incidence of grapevine trunk diseases per year in France over
the 2003-2022 period. The prevalence (shown as curves in the top panel) was determined as the
number of plots with at least one symptomatic vine recorded per year. The incidence
(represented by bar charts in the bottom panel) was determined as the percentage of symptomatic
vines observed per plot each year. The numbers displayed correspond to the number of plots

 monitored for the incidence of (a) esca (blue dots, lines and bars) and (b) Eutypa dieback

 (mustard dots, lines and bars). The error bars indicate the standard error of the mean.

786 Figure 5: The mean incidence of grapevine trunk diseases by cultivar and wine region (inset, i :

787 esca; ii : Eutypa dieback) in France over the 2003-2022 period. Numbers represent the number of

788 plots monitored for the incidence of (a) esca (blue bars) and (b) Eutypa dieback (mustard bars).

789 The error bars represent the standard error of the mean.

Figure 6: Effects of cultivar (upper panels, means in descending order) and year (lower panels) on esca prevalence (left panels) or esca incidence if esca was observed in the plot (right panels). In the upper panels, black points represent the posterior median (50th percentiles, middle) and vertical lines the lower (2.5th percentiles) and upper (97.5th percentiles) limits of the credible interval for the independent random effects associated with each cultivar. The segments between the upper panels (a and b) show the relationship between the estimated prevalence and incidence of the different cultivars. The lower panels display the year effects for esca prevalence and incidence, as fitted with an autoregressive model of order 1. Estimates in blue correspond to years associated with significantly negative effects (97.5th percentile < 0) relative to the mean (0), whereas those in red correspond to years with significantly positive effects (2.5th percentile > 0). The values shown correspond to the probability of direction for each cultivar (see the materials and methods section).

Figure 7: Impact of plot age (in years), by cultivar, on standardised estimates of the esca

incidence for the plots in which esca was observed. The posterior median (50th percentiles,

middle), lower (2.5th percentiles) and upper (97.5th percentiles) limits of the credible interval

are shown. Estimates in blue correspond to ages associated with significantly negative effects

(97.5th percentile < 0) relative to the mean (0) whereas those in red correspond to ages

associated with significantly positive effects (2.5th percentile > 0). Estimates are shown in colour

if the findings for at least two consecutive years are significant.

Figure 8: Range of ages during which vines are susceptible to esca (black points) and peak (red

triangle) susceptibility for cultivars with significant and positive results between the first and the

last plot age. The cultivars are ordered by estimated mean incidence and only cultivars with an

age significantly and positively above the mean of 0 are shown (Figure 6b, Merlot is therefore

not represented). The wide horizontal grey band represents the standard deviation from the mean,

which is represented by the horizontal dark grey bar, for all cultivars combined.

Table 1: Summary statistics for plot age and number of plots (total and per year) monitored for

the 12 *Vitis vinifera* cultivars retained for statistical modelling. Maximum plot age was fixed at

49 years.

| Cultivar | Plot age (mean ± SD [min, max] years) | Total number of plots | Mean (± SD) number of plots per year |
|---|---|---|---|
| Cabernet Sauvignon | 23 ± 9 [2, 49] | 85 | 14 ± 13 |
| Sauvignon Blanc | 20 ± 10 [3, 49] | 62 | 8 ± 10 |
| Chardonnay | 30 ± 9 [7, 49] | 61 | 6 ± 10 |
| Cabernet Franc | 23 ± 13 [2, 49] | 53 | 9 ± 12 |
| Merlot | 24 ± 10 [7, 49] | 52 | 13 ± 11 |
| Pinot Noir | 29 ± 9 [7, 49] | 47 | 6 ± 8 |
| Ugni Blanc | 26 ± 12 [4, 49] | 41 | 20 ± 28 |
| Melon | 21 ± 10 [2, 49] | 39 | 6 ± 4 |
| Chenin | 18 ± 8 [1, 40] | 38 | 5 ± 5 |
| Gewurztraminer | 28 ± 9 [2, 49] | 28 | 9 ± 10 |
| Pinot Auxerrois | 32 ± 8 [11, 49] | 24 | 8 ± 8 |
| Riesling | 28 ± 9 [10, 49] | 24 | 8 ± 10 |

830    Table 2: Deviance information criterion (DIC) and widely applicable information criterion

831    (WAIC) for the eight model structures considered. Number of vines monitored per plot: $n_{vines}$.

| Model ID | Prevalence model | Incidence model | DIC | WAIC |
|---|---|---|---|---|
| M8 "Full" | plot + cultivar + year + $n_{vines}$ | plot + cultivar + year + age x cultivar | 37842 | 64418 |
| M7 | plot + cultivar + year + $n_{vines}$ | plot + cultivar + year + age | 40272 | 80397 |
| M6 | plot + cultivar + year + $n_{vines}$ | plot + cultivar + year | 39771 | 111086 |
| M5 | plot + cultivar + year + $n_{vines}$ | plot + cultivar | 44832 | 142323 |
| M4 | plot + cultivar + year + $n_{vines}$ | plot | 44871 | 142351 |
| M3 | plot + cultivar + year + $n_{vines}$ | 1 | 132737 | 121358 |
| M2 | plot + cultivar + year | 1 | 132736 | 121357 |
| M1 | plot + cultivar | 1 | 132813 | 121436 |
| M0 | plot | 1 | 132913 | 121497 |

Table 3: Summary statistics for the parameters of the full model of the prevalence and incidence

of esca leaf symptoms. For both model components (prevalence and incidence for plots with a

prevalence of 1), the table summarises the posterior mean, standard deviation (SD), 0.025, 0.5,

0.975 quantiles (quant.) and mode for intercepts ($\beta_0$: intercept for esca prevalence; $\beta_1$: intercept

for esca incidence), the number of vines monitored ($\mu_0$), the variance of the independent random

effects associated with cultivars and plots ($\sigma^2_{cultivar0}$, $\sigma^2_{cultivar1}$, $\sigma^2_{plot0}$ and $\sigma^2_{plot1}$), the correlation

coefficients and variance of the AR1 process associated with year ($\rho_{year0}$ and $\rho_{year1}$) and the

correlation coefficients and variance of the AR1 process associated with age for each cultivar

($\rho_{age1,cultivar}$ and $\sigma^2_{age1,cultivar}$ to be detailed).

| | Mean | SD | 0.025quant. | 0.5quant. | 0.975quant. | Mode |
|---|---|---|---|---|---|---|
| Esca prevalence | | | | | | |
| $\beta_0$ | 2.75 | 0.47 | 1.83 | 2.75 | 3.69 | 2.74 |
| $\mu_0$ | 0.33 | 0.20 | -0.05 | 0.33 | 0.73 | 0.33 |
| $\sigma^2_{cultivar0}$ | 0.90 | 0.40 | 0.35 | 0.83 | 1.88 | 0.70 |
| $\sigma^2_{plot0}$ | 0.36 | 0.04 | 0.29 | 0.36 | 0.44 | 0.36 |
| $\rho_{year0}$ | 0.64 | 0.16 | 0.25 | 0.67 | 0.89 | 0.72 |
| $\sigma^2_{year0}$ | 5.11 | 2.40 | 1.84 | 4.65 | 11.08 | 3.82 |
| Esca incidence (when prevalence > 0) | | | | | | |
| $\beta_1$ | -3.91 | 0.20 | -4.41 | -3.90 | -3.46 | -3.88 |
| $\sigma^2_{cultivar1}$ | 3.53 | 2.42 | 0.94 | 2.89 | 9.95 | 2.01 |
| $\sigma^2_{plot1}$ | 1.25 | 0.09 | 1.09 | 1.25 | 1.43 | 1.25 |
| $\rho_{year1}$ | 0.30 | 0.24 | -0.20 | 0.32 | 0.71 | 0.36 |
| $\sigma^2_{year1}$ | 12.58 | 4.35 | 5.87 | 11.98 | 22.79 | 10.84 |
| $\rho_{age1,Cabernet\ Franc}$ | 0.80 | 0.09 | 0.59 | 0.82 | 0.93 | 0.85 |

| | | | | | | |
|---|---|---|---|---|---|---|
| $\sigma^2_{age1,\text{Cabernet Franc}}$ | 2.81 | 1.32 | 1.03 | 2.54 | 6.12 | 2.09 |
| $\rho_{age1,\text{Cabernet Sauvignon}}$ | 0.95 | 0.03 | 0.87 | 0.95 | 0.98 | 0.96 |
| $\sigma^2_{age1,\text{Cabernet Sauvignon}}$ | 0.72 | 0.38 | 0.24 | 0.64 | 1.68 | 0.50 |
| $\rho_{age1,\text{Chardonnay}}$ | 0.80 | 0.10 | 0.55 | 0.82 | 0.93 | 0.85 |
| $\sigma^2_{age1,\text{Chardonnay}}$ | 2.62 | 1.36 | 0.89 | 2.32 | 6.09 | 1.83 |
| $\rho_{age1,\text{Chenin}}$ | 0.82 | 0.09 | 0.60 | 0.83 | 0.94 | 0.87 |
| $\sigma^2_{age1,\text{Chenin}}$ | 2.59 | 1.27 | 0.87 | 2.35 | 5.77 | 1.89 |
| $\rho_{age1,\text{Gewurztraminer}}$ | 0.43 | 0.16 | 0.09 | 0.44 | 0.71 | 0.46 |
| $\sigma^2_{age1,\text{Gewurztraminer}}$ | 1.33 | 0.39 | 0.72 | 1.28 | 2.23 | 1.19 |
| $\rho_{age1,\text{Melon}}$ | 0.77 | 0.13 | 0.45 | 0.80 | 0.94 | 0.85 |
| $\sigma^2_{age1,\text{Melon}}$ | 6.04 | 2.92 | 2.14 | 5.46 | 13.39 | 4.42 |
| $\rho_{age1,\text{Merlot}}$ | 0.74 | 0.13 | 0.44 | 0.76 | 0.93 | 0.81 |
| $\sigma^2_{age1,\text{Merlot}}$ | 1.62 | 0.80 | 0.52 | 1.47 | 3.59 | 1.18 |
| $\rho_{age1,\text{Pinot Auxerrois}}$ | 0.66 | 0.14 | 0.33 | 0.69 | 0.88 | 0.73 |
| $\sigma^2_{age1,\text{Pinot Auxerrois}}$ | 4.83 | 2.20 | 1.82 | 4.41 | 10.30 | 3.67 |
| $\rho_{age1,\text{Pinot Noir}}$ | 0.91 | 0.05 | 0.80 | 0.92 | 0.97 | 0.94 |
| $\sigma^2_{age1,\text{Pinot Noir}}$ | 0.97 | 0.50 | 0.30 | 0.87 | 2.22 | 0.68 |
| $\rho_{age1,\text{Riesling}}$ | 0.80 | 0.10 | 0.55 | 0.82 | 0.94 | 0.86 |
| $\sigma^2_{age1,\text{Riesling}}$ | 6.85 | 3.45 | 2.27 | 6.15 | 15.53 | 4.90 |
| $\rho_{age1,\text{Sauvignon Blanc}}$ | 0.82 | 0.08 | 0.64 | 0.84 | 0.94 | 0.86 |
| $\sigma^2_{age1,\text{Sauvignon Blanc}}$ | 2.17 | 0.99 | 0.74 | 2.01 | 4.54 | 1.67 |
| $\rho_{age1,\text{Ugni Blanc}}$ | 0.90 | 0.06 | 0.76 | 0.91 | 0.97 | 0.93 |
| $\sigma^2_{age1,\text{Ugni Blanc}}$ | 5.37 | 2.89 | 1.61 | 4.77 | 12.66 | 3.67 |