Eye gaze is not inversion-proof: A robust, sex-invariant gaze inversion effect

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Abstract

Humans are adept at distinguishing individual faces, yet inversion dramatically impairs

this ability. This face inversion effect is remarkably robust across observers, but evidence

is mixed as to whether inversion also impairs the perception of facial parts, particularly

the eye region. Some studies have shown that featural processing is preserved or even

enhanced when faces are inverted, whereas others have reported clear inversion-related

impairments in feature-based judgements. These mixed findings may reflect limited

statistical power, unbalanced participant sex ratios, and heterogeneous task designs. To

address these issues, we examined how strongly face inversion affects sensitivity to gaze

direction in a well-powered, sex-balanced sample. A total of 190 participants judged

whether the eyes in briefly presented upright or inverted faces were looking directly at

them or not. Inversion reliably reduced sensitivity to gaze direction, yielding a

medium-to-large effect size. Females showed modestly higher overall sensitivity than

males (a small-to-medium effect), whereas the inversion effect was highly similar for

females and males. These findings show that brief gaze judgements are not immune to

inversion, even in a task that could in principle be based largely on eye-region

information. They provide quantitative constraints that models of gaze perception should

accommodate, including the largely sex-invariant inversion effect.

Keywords:

Gaze inversion effect; Gaze perception; Sex invariance; Sex difference

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dataset analyzed during the current study is available at GitHub:

https://github.com/dicemt/matsuyoshi_gaze_inversion_effect

Abbreviation:

FIE: face inversion effect; GIE: gaze inversion effect

Competing interests:

None of the authors have any potential conflicts of interest.

Authors' contributions:

All authors designed the study and wrote the manuscript. DM and KK collected stimulus

materials. DM collected and analyzed the data.

1   **1. Introduction**

Humans are experts at distinguishing and recognizing faces under various

conditions. People can detect subtle differences and distinguish between thousands of

individuals (Jenkins et al., 2018) even though typical human faces share a strikingly

similar overall structure (i.e., two eyes, a nose, and a mouth). However, this ability

deteriorates markedly when faces are inverted (Yin, 1969). Turning a face upside-down

makes it particularly difficult to recognise, whereas inverting many other types of objects

does not affect recognition performance as much.

This face inversion effect (FIE) is exceptionally robust across observers and tasks. It

has led researchers to speculate that there are *qualitative* differences between upright and

inverted face processing, with a shift from holistic or configural processing for upright

faces to more piecemeal or featural processing for inverted faces (Maurer et al., 2002).

From this qualitative view, perceiving a face as a coherent whole (rather than as a

collection of separate elements) is an idiosyncratic feature that makes face recognition

special (Farah et al., 1998; Moscovitch et al., 1997). A key prediction of this view is that,

because of the shift in processing mode, the processing of individual features should be

comparatively spared when a face is presented upside down. By contrast, more recent

studies have suggested that the differences in processing upright and inverted faces

are *quantitative* (i.e., a matter of degree in processing efficiency) rather than qualitative

(i.e., invoking distinct processing mechanisms) (Gold et al., 2012; Jiang et al., 2006;

Matsuyoshi et al., 2015; Sekuler et al., 2004). Sekuler et al. (2004), for example, used a

response-classification technique to identify the most discriminative stimulus regions that

drove performance and found that observers relied on similar local regions across

orientations but extracted information less efficiently from inverted faces. Gold et al.

(2012) further showed that face perception is no better than would be predicted by

Bayesian integration of information from individual facial features. If observers used a

truly holistic face representation, they should exhibit superoptimal performance for whole

faces (better than the sum of their parts), yet performance was comparable to the Bayesian

prediction derived from features presented in isolation. These findings are difficult to

reconcile with a strict qualitative shift and instead support a quantitative view in which

similar recognition strategies operate for both upright and inverted faces, albeit less

efficiently for the latter. At the same time, the two views are not mutually exclusive, and

dissociable neural pathways may still contribute to relatively holistic and more

feature-based mechanisms (Matsuyoshi et al., 2015).

The behavioural FIE itself is very strong and has been confirmed by rigorous

replications. However, it is less clear whether inversion impairs the perception of

individual facial features. In particular, whether sensitivity to information conveyed by

the eyes, such as gaze direction, varies with face orientation remains unclear. Several

studies have reported that inversion has little to no effect on the ability to distinguish

faces that differ only in the shape of individual features (Freire et al., 2000; Le Grand et

al., 2001). For example, upright faces with different hairstyles and costumes tend to be

perceived as different individuals even though their internal features are identical (Sinha

& Poggio, 1996), and composite-face paradigms (Murphy et al., 2017) have shown that

inversion can markedly improve the ability to identify individuals based on their internal

features (Hole, 1994; Young et al., 1987). Xu and Tanaka (2013) also reported that

inversion had little or no effect on discriminating featural and configural differences in

the eye region (while severely disrupting the perception of changes in the lower region of

the face). These results suggest that featural processing, at least for some aspects of eye

or gaze information, is preserved or even enhanced when faces are upside-down. By

contrast, numerous studies have reported the opposite pattern, namely that featural

processing becomes more difficult when faces are inverted. One prominent example is

the Thatcher illusion (Thompson, 1980), in which observers fail to notice local feature

changes in an inverted face, even though the changes are obvious in an upright face.

Other studies have suggested that isolated features presented without a full face context

can still produce inversion effects (Leder et al., 2001; Rakover & Teucher, 1997). Taken

together, the literature provides evidence for both relative preservation and clear costs of

inversion in processing individual features, making it difficult to reach consensus on how

inversion affects feature-level information.

Methodological differences among studies may contribute to these discrepancies.

Sample size, participant demographics, study design, and task settings all vary

substantially across the literature. In particular, sample size and participant sex

composition are critical factors in face processing research. Insufficient statistical power

due to small samples can increase the likelihood of false positives and undermine the

reliability of findings (Button et al., 2013; Simmons et al., 2011). Moreover, females

often outperform males in face processing tasks (Bayliss et al., 2005; Goodman et al.,

2012; Heisz et al., 2013; Matsuyoshi et al., 2014; Matsuyoshi & Watanabe, 2021;

McClure, 2000), yet studies have not always monitored or balanced the sex ratio unless

the sex differences were the primary focus. Previous inconsistencies may therefore partly

reflect unaccounted variance in participant demographics. Given the recurring reports of

a female advantage and sex as a potential moderator across diverse face processing tasks

(Cellerino et al., 2004; Lewin & Herlitz, 2002; Matsuyoshi et al., 2014), balancing the

sample for sex and explicitly testing for sex effects may help to obtain more precise

estimates of the inversion cost.

In the present study, we examined the gaze inversion effect, namely the impact of

face inversion on sensitivity to small deviations from direct gaze. Our primary goal was

to obtain a precise estimate of the extent to which face inversion affects sensitivity to

gaze direction under tightly controlled psychophysical conditions in a large, sex-balanced

sample. Observers judged whether the eyes in briefly presented faces were looking

directly at them, with faces shown upright or inverted. The judgement is driven primarily

by eye-region cues, although observers may also use broader facial context. It therefore

provides an informative test of whether inversion costs are confined to clearly global,

configural aspects of faces or also extend to eye-region–based judgements. A second,

more exploratory goal was to examine whether the inversion effect on gaze

discrimination differs between females and males, building on prior reports suggesting

that sex may moderate performance on tasks involving facial cues. This design allowed

us to quantify the magnitude of inversion costs on gaze discrimination and to assess

whether these costs generalise across sexes, thereby providing quantitative constraints

that theories of gaze perception should accommodate.

2. Materials and Methods

2.1 Participants

We assumed a medium-sized effect of sex on gaze direction discrimination based on

studies showing sex differences in face or gaze perception (Bayliss et al., 2005; Goodman

et al., 2012; Matsuyoshi et al., 2014; Matsuyoshi & Watanabe, 2021). To achieve a power

($1 - \beta$) of 0.95 and a medium effect size (Cohen's $f = 0.25$) while assuming a high

correlation between the measures ($r = 0.7$ between upright and inverted gaze processing), a

sample of 180 (90 for each group) is required to detect a significant difference ($\alpha = 0.05$)

between participant sexes (Cohen, 1988; Faul et al., 2007). In addition, we conservatively

assumed the effect size of inversion as small-to-medium (Cohen's $f = 0.15$) because of the

mixed results of previous studies. A sample of 90 (45 for each group) is required to detect a

significant difference ($\alpha = 0.05$, $1 - \beta = 0.95$, $r = 0.7$) between upright and inverted faces

with a small-to-medium effect size, and a sample of 90 (45 for each group) is required to

detect a significant interaction ($\alpha = 0.05$, $1 - \beta = 0.95$, $r = 0.7$) between the face orientation

and sex with a small-to-medium effect size. Therefore, we decided to recruit a sample of

more than 200 participants to confirm small-to-medium-sized effects between conditions

and to overcome the potential loss of analyzable data, often caused by participants'

inappropriate responses or technical errors.

Two hundred and thirteen young adults participated. All had normal or

corrected-to-normal vision, and none reported a history of neurological or developmental

disorders. The study was conducted in accordance with the Declaration of Helsinki and

was approved by the Waseda University Ethics Committee. Each participant gave written

informed consent before the experiment. Twenty-three participants were excluded

because their proportion of 'direct' responses on 0° trials was below 60%. This criterion

also served as a basic post-hoc screen to ensure participants were engaged with the task

and could perform it at a minimal level of proficiency. After these exclusions, data from

190 participants (self-reported 92 females, 98 males; mean age: 21.3 years; range: 18-29

years) remained for analysis.

2.2 Stimuli

Face images were obtained from three male and three female models, looking either

directly at the camera (0°) or at 10°, 20°, or 30° to the left or right, with their faces square

to the camera, and were photographed by the authors (Matsuyoshi et al., 2014). Images

were grey-scaled and subtended approximately 12° × 17° of the visual angle. The mean

luminance was normalized across images. A mask stimulus was generated by

grid-scrambling the average face stimuli.

2.3 Procedure

Participants viewed the stimuli on a 21-inch CRT monitor at a distance of 57 cm

with their heads resting on a chinrest and judged whether the eyes were looking directly

at them (Matsuyoshi et al., 2014). All stimuli were presented against a black background.

Each trial began with an upright or inverted face stimulus (20 ms) followed by a brief

blank screen (20, 40, or 60 ms), a visual mask (100 ms), and finally, a fixation cross (until

response) (Fig. 1). Face images were either direct gaze (0°) or left- or right-averted gaze

(10°, 20°, or 30°). The vertical positions of the eyes were aligned between upright and

inverted faces and located along the vertical centre of the display. We used a variable

interstimulus interval (ISI) between the face and the mask to control task difficulty. Each

left or right angle was presented 12 times, except for 0° stimuli, which were presented 24

times, for each of the three ISIs for both face orientations, resulting in 576 trials in total

(6 gaze angles [left or right 10-30°] × 3 ISIs × 2 face orientations × 12 trials + 1 gaze

angle [0°] × 3 ISIs × 2 face orientations × 24 trials). The two face orientations were

presented separately in four alternating blocks of 144 trials each, with the order

counterbalanced across participants (i.e., either upright-inverted-upright-inverted or

inverted-upright-inverted-upright). Gaze directions and ISIs were randomised across

trials. Participants performed the main experiment after 30 practice trials.

2.4 Analysis

Responses from left- and right-averted gaze stimuli were pooled for each absolute

angle. In addition, responses from different ISIs were also pooled. The gaze threshold

was defined as the angle at which a 50% response was obtained, as estimated by a local

linear fit of a nonparametric psychometric function (Żychaluk & Foster, 2009) to each

observer's 'direct' responses as a function of gaze angle. This threshold represents the

boundary of the 'cone of direct gaze' (i.e., the range of perceived direct gaze) (Gamer &

Heiko, 2007; Jun et al., 2013), or the point of subjective equality for 'direct' vs. 'averted'

judgments, rather than a discrimination threshold between two non-zero averted gaze

angles as often used in typical psychophysical paradigms. A higher threshold indicated

lower sensitivity to gaze direction. The significance level was set at $\alpha = 0.05$.

<div align="center">**3. Results**</div>

3.1 Psychometric functions

We performed a mixed ANOVA (with Greenhouse–Geisser correction applied

where sphericity was violated) on the proportion of 'direct' responses, with angle (0°, 10°,

20°, 30°) and orientation (upright, inverted) as within-subject factors and participant sex

(male, female) as a between-subject factor (Fig. 2a). The main effects of the angle ($F_{(1.99,}$

$_{373.89)} = 4796.912$, $p = 9.854 \times 10^{-270}$, $\eta_p^2 = 0.964$, $\omega_G^2 = 0.881$, Cohen's $f = 2.737$), face

orientation ($F_{(1, 188)} = 158.223$, $p = 1.007 \times 10^{-26}$, $\eta_p^2 = 0.457$, $\omega_G^2 = 0.064$, Cohen's $f =$

0.091), and participant sex ($F_{(1, 188)} = 9.758$, $p = 0.002$, $\eta_p^2 = 0.049$, $\omega_G^2 = 0.014$,

Cohen's $f = 0.044$) were significant. This pattern indicates a decrease in 'direct' responses

with increasing angle, higher 'direct' responses in the inverted compared to the upright

face condition, and higher 'direct' responses in male compared to female participants,

respectively.

We found significant interactions between the angle and orientation ($F_{(2.14, 402.85)} =$

276.829, $p = 9.499 \times 10^{-80}$, $\eta_p^2 = 0.596$, $\omega_G^2 = 0.117$, Cohen's $f = 0.127$) and between

the angle and participant sex ($F_{(1.99, 373.89)} = 6.847$, $p = 0.001$, $\eta_p^2 = 0.035$, $\omega_G^2 = 0.009$,

Cohen's $f = 0.035$), while no significant interactions were found between the orientation

and sex ($F_{(1, 188)} = 1.622$, $p = 0.204$, $\eta_p^2 = 0.009$, $\omega_G^2 = 0.000$, Cohen's $f = 0.009$) and

between the angle, orientation, and sex ($F_{(2.14, 402.85)} = 0.962$, $p = 0.388$, $\eta_p^2 = 0.005$, ω_G^2

$= -1.840 \times 10^{-5}$, Cohen's $f = 0.007$). Follow-up tests revealed higher 'direct' responses in

the upright compared with the inverted face condition at an angle of 0° ($F_{(1, 188)} = 68.290$,

$p = 2.486 \times 10^{-14}$, $\eta_p^2 = 0.267$, $\omega_G^2 = 0.094$, Cohen's $f = 0.325$), and higher 'direct'

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responses in the inverted compared to the upright face condition at angles between 10 and

30° ($F_{(1, 188)}$ values > 65.813, p values < 6.274×10^{-14}, η_p^2 values > 0.259, ω_G^2 values >

0.084, Cohen's f values > 0.298). In addition, male participants showed higher 'direct'

responses at angles between 10 and 30° than female participants ($F_{(1, 188)}$ values > 8.673,

p values < 0.004, η_p^2 values > 0.044, ω_G^2 values > 0.031, Cohen's f values > 0.165). By

contrast, male and female participants showed similar 'direct' responses at an angle of 0°

($F_{(1, 188)} = 0.215$, $p = 0.643$, $\eta_p^2 = 0.001$, $\omega_G^2 = -0.003$, Cohen's $f = 0.027$).

3.2 Gaze thresholds

A 2 (upright, inverted) × 2 (male, female) mixed ANOVA on the 50%

gaze-perception threshold (Fig. 2b) showed a significant main effect of face orientation

($F_{(1, 188)} = 109.666$, $p = 1.650 \times 10^{-20}$, $\eta_p^2 = 0.368$, $\omega_G^2 = 0.111$, Cohen's $f = 0.352$) with

a larger threshold in the inverted compared to the upright face condition

(medium-to-large effect size), and a significant main effect of sex ($F_{(1, 188)} = 7.425$, $p =$

0.007, $\eta_p^2 = 0.038$, $\omega_G^2 = 0.026$, Cohen's $f = 0.165$) with a larger threshold for male

compared to female participants (small-to-medium effect size). However, there was no

significant interaction between the face orientation and participant sex ($F_{(1, 188)} = 2.073$, p

$= 0.152$, $\eta_p^2 = 0.011$, $\omega_G^2 = 0.001$, Cohen's $f = 0.046$), showing that both sexes showed

comparable inversion effects (females, $t_{(91)} = 6.639$, $p = 2.244 \times 10^{-9}$, Cohen's $d = 0.692$

[95% CI: 0.394, 0.989]; males, $t_{(97)} = 8.174$, $p = 1.142 \times 10^{-12}$, Cohen's $d = 0.826$ [95%

CI: 0.533, 1.117]; females − males, $t_{(187.75)} = 1.444$, $p = 0.150$, Cohen's $d = 0.210$ [95%

CI: −0.076, 0.495]), whereas female participants had a higher sensitivity than male

participants for discriminating gaze direction irrespective of face orientation (upright,

$t_{(170.18)} = 2.309$, $p = 0.022$, Cohen's $d = 0.335$ [95% CI: 0.048, 0.621]; inverted, $t_{(177.05)} =$

2.6104, $p = 0.010$, Cohen's $d = 0.379$ [95% CI: 0.091, 0.666]).

There was a significant correlation between the upright and inverted gaze thresholds

(Fig. 2c; overall, $r = 0.631$, $p = 1.556 \times 10^{-22}$; males, $r = 0.693$, $p = 2.531 \times 10^{-15}$;

females, $r = 0.461$, $p = 3.844 \times 10^{-6}$), indicating that there are partly shared mechanisms

in upright and inverted gaze processing. The correlation was significantly greater in male

than in female participants ($z = 2.415$, $p = 0.015$). Note, however, that the difference was

not robust to the exclusion of two outliers (one male and one female participant with the

largest inverted gaze thresholds), after which it was no longer significant ($z = 1.883$, $p =$

0.060).

4. Discussion

Using a well-powered cohort with an a priori determined sample size, the present

study showed that inverting the face reliably impaired sensitivity to small deviations from

direct gaze. At the same time, we observed higher overall gaze sensitivity in females than

in males, while the magnitude of the inversion cost was closely comparable across sexes.

We recognise that gaze perception involves distinctive signalling properties, such as the

high contrast between the iris and sclera, which may set it apart from other facial

information (Kano, 2023; Kobayashi & Kohshima, 1997). It would therefore be

inappropriate to assume that the present results generalise automatically to all facial

features or that they fully resolve the broader debate between qualitative and quantitative

views of inversion effects. Rather, they provide a useful boundary condition for these

theories. Specifically, they are difficult to reconcile with a strict qualitative view that

treats feature-level processing as largely preserved under inversion. The fact that

inversion produces a clear decrement in performance for a task that can in principle be

supported by relatively local eye information suggests that this assumption does not hold

for at least one highly salient facial cue (eye gaze). Our data do not establish that all

feature processing is governed by purely quantitative differences, but they do show that

the qualitative prediction of preserved featural processing cannot be universal.

4.1 Qualitative and quantitative views of the inversion effects

Several studies have suggested that inversion preserves featural processing while

severely impairing holistic or configural processing, such as the distances between

features (Freire et al., 2000; Le Grand et al., 2001; Rossion, 2008; Tanaka & Sengco,

1997; Young et al., 1987). On this qualitative view, inversion is assumed to prompt a shift

toward feature-based encoding, so that performance on tasks driven by individual

features should be relatively similar for upright and inverted faces (Farah et al., 1995). In

particular, the eyes have sometimes been regarded as the least susceptible feature to

inversion. For example, Xu and Tanaka (2013) reported that inversion did not affect the

perception of changes in the eye region, whereas it degraded sensitivity to changes in the

mouth region (N = 22, 40% female). Although their finding appears inconsistent with

ours, their task presented faces for 500 ms, which is substantially longer than the 20-ms

exposure used here. Such long encoding durations may have attenuated or masked any

disadvantage for inverted relative to upright faces, and behavioural performance in their

study was near ceiling (about 90%), leaving limited room for inversion costs to emerge.

By contrast, quantitative views posit that similar mechanisms are engaged for

upright and inverted faces but operate less efficiently when faces are inverted (Gold et al.,

2012; Jiang et al., 2006; Sekuler et al., 2004). Jenkins and Langton (2003) reported that

inversion impaired sensitivity to gaze direction, although their study used a

between-participants design, a small sample size (N = 6 per group), and an unknown sex

ratio. Furthermore, Rakover and Teucher (1997) (N = 16, 73% female) and Leder et al.

(2001) (N = 20, 90% female) likewise found it more difficult to process individual

features when they were inverted. However, the limited and often unbalanced samples in

these studies make it difficult to draw firm conclusions about the size and robustness of

the effect. Our findings help clarify the extent to which inversion affects gaze perception

in a large, sex-balanced sample, using tightly controlled psychophysical methods.

Although prior evidence hinted that inversion can impair gaze direction sensitivity

(Jenkins & Langton, 2003), the strength and generality of this effect had not been firmly

established across paradigms and larger cohorts, leading to ambiguity in the literature.

The present study was designed to address this gap by quantifying the gaze inversion cost

with greater precision in a well-powered cohort with a balanced sex ratio.

It is also important to recognise that the questions of 'qualitative versus quantitative'

and 'configural versus featural' are conceptually distinct, even though they are

intertwined in empirical work (Gold et al., 2012; Maurer et al., 2002; Sekuler et al., 2004).

The qualitative view typically links upright faces with configural processing and inverted

faces with featural processing. Quantitative accounts, by contrast, do not assume such an

orientation-dependent processing style; instead, they propose that inversion generally

reduces processing efficiency. Our design targeted a key implication of a *strict* qualitative

account: if featural processing is preserved or even enhanced for inverted faces, then

performance on a task based largely on information from the eye region should show

little or no impairment under inversion. In contrast to this prediction, we observed a clear

impairment in gaze discrimination when faces were inverted. This suggests that, at least

for gaze judgements of the sort studied here, inversion costs are not confined to clearly

global, configural aspects of faces but extend to tasks that can be supported largely by

eye-region cues.

4.2 The effects of task settings on inversion effects

Although we demonstrated an inversion-induced impairment in gaze sensitivity,

many other studies (Freire et al., 2000; Hole, 1994; Le Grand et al., 2001; Young et al.,

1987), including large-sample investigations (N = 242–282) (Rezlescu et al., 2017; Susilo

et al., 2013), have reported preserved or even improved performance when comparing

individual features in inverted faces. These apparent discrepancies are unlikely to result

from low statistical power or an imbalanced sex ratio, since both these large-sample

studies and the present one used sizeable cohorts, and we (48% female) and Susilo et al.

(2013) (52% female) achieved approximately balanced sex ratios. We suggest instead that

task settings may be a major source of divergence. Featural processing is sometimes

preserved (or enhanced) despite the diminished sensitivity, perhaps not because

parts-based processing is dominant for inverted faces (Farah et al., 1995; Tanaka & Farah,

1993), but because the tasks require discrimination between exemplars from the same

facial parts (e.g., comparing the eyes of different people). Thus, even though perceptual

sensitivity per se decreases, the ability to compare two signals may be relatively unlikely

to deteriorate because sensitivity to both signals, but not to either, decreases. Moreover,

the apparent dominance of feature-based processing for inverted faces may result from

greater difficulty decomposing facial parts in upright faces, rather than from enhanced

featural processing for inverted faces. Processing an upright face likely involves tight

integration of internal features, making it difficult to parse the face into isolated features

(Maurer et al., 2002); however, an inverted face is less likely to undergo this integration.

Under such circumstances, explicit behavioural performance may depend less on whether

inversion alters the underlying processing style and more on how the specific task elicits

comparisons between stimuli, the distribution of attention across features, and the overall

processing load. In this context, our finding that inversion impairs gaze sensitivity in a

task that emphasises brief, single-stimulus judgements complements prior work using

comparison tasks and suggests that different paradigms may reveal different facets of

how inversion affects feature-level information.

Note that the effect size of the GIE was smaller than that of the typical FIE. Despite

a huge effect of the FIE in a face recognition task ($d = 1.298$) (Matsuyoshi et al., 2015),

we found a medium-to-large effect of the GIE in the present study ($d = 0.761$, across all

participants). One reason for the tremendous difference in effect size might be that a

whole face accumulates the impairments in the processing of multiple facial features

(Gold et al., 2012). Although susceptibility to inversion may differ across parts (Rakover

& Teucher, 1997; Xu & Tanaka, 2013), these differences may aggregate to produce a

large effect size for the FIE. Alternatively, the large effect size difference may simply

reflect differences in processing load between the tasks. The present task encouraged

observers to focus on the eyes. By contrast, typical FIE tasks do not explicitly encourage

observers to focus on a facial feature but rather on a whole face. Encoding a whole face

may be more challenging than encoding a single facial feature.

4.3 Sex differences in face and gaze processing

Consistent with numerous studies reporting female advantages in face processing

(Bayliss et al., 2005; Goodman et al., 2012; Heisz et al., 2013; Matsuyoshi et al., 2014;

Matsuyoshi & Watanabe, 2021; McClure, 2000), we found that females outperformed

males in gaze direction sensitivity. The effect size for this female advantage in upright

gaze sensitivity (Cohen's $d = 0.3351$) was notably smaller than that reported for upright

face recognition in a previous study from our group (Cohen's $d = 0.643$, N = 180, 45%

female) (Matsuyoshi & Watanabe, 2021). This may indicate that tasks that rely more on

relatively local feature information, such as brief gaze judgements, are likely to manifest

weaker effects than whole-face processing, perhaps because this task draws primarily on

a single feature, so inversion-related costs may not compound across multiple features.

The observed female superiority in gaze sensitivity may reflect that females use

more efficient scanning strategies or have a lower threshold for detecting subtle social

signals. While the exact origin of this difference remains debated, prior work suggests

that both biological factors (Bölte et al., 2023; Kiesow et al., 2020; Kundakovic &

Tickerhoof, 2024) and psychological factors, such as internalized social norms or

stereotypes (Briton & Hall, 1995; Crandall et al., 2002; Gavrilets & Richerson, 2017;

Rutland et al., 2005), may contribute to sex and gender differences in social cognition. In

the present study, females showed higher baseline sensitivity, perhaps reflecting some

combination of these factors, but the relative cost of inversion was similar for females

and males. This pattern suggests that the computations supporting gaze discrimination are

similarly affected by inversion across sexes, even though baseline performance levels

differ.

A notable limitation of our study is that we treated sex/gender in a binary fashion,

and findings solely depended on self-reported sex/gender. Outcomes can vary across

individuals with different birth-assigned sexes and gender identities (Bölte et al., 2023;

Joel, 2021). Definitions of sex/gender are continually evolving, and sex/gender is now

understood as a spectrum rather than as simple binary categories (Ainsworth, 2015;

Kundakovic & Tickerhoof, 2024; Rosenthal, 2021). Future research that more fully

incorporates variation across the sex/gender spectrum will be necessary for understanding

how biological and sociocultural factors jointly shape face and gaze processing.

4.4 Conclusion

In conclusion, our study provides robust evidence for an inversion-induced

impairment in gaze direction sensitivity (GIE). While our findings do not resolve the

broader debate over the nature of the FIE, they directly challenge a strict interpretation of

the qualitative view, which posits that the processing of individual facial features is

preserved under inversion. Thus, they nonetheless provide a critical boundary condition

for theories assuming preserved parts-based processing. Our results demonstrate that, at

least for the crucial feature of eye gaze, sensitivity is significantly diminished when the

face is inverted, even in a task that can in principle be supported by information within

the eye region. Since we did not incorporate tasks other than discriminating gaze

direction, further research is required to clarify the conditions under which processing of

facial feature information is preserved or impaired by inversion. Nevertheless, our

findings establish the GIE as a reliable phenomenon for brief gaze judgements and

highlight the value of using well-powered, sex-balanced samples when investigating how

inversion shapes both face perception and individual differences in gaze processing.

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Figure Legends

Figure 1. Experimental Paradigm.

Schematic presentation of the experimental task. Upright and inverted faces were presented in different blocks (4 [2 upright + 2 inverted] blocks × 144 trials).

Figure 2. Experimental Results.

(a) Gaze angle psychometric functions. Mean percentage of 'direct' responses (y-axis), plotted as a function of gaze angle (x-axis) × participant sex (female = red, male = blue) × face orientation (upright face = filled circle with a solid line, inverted face = open circle with dashed line). Shaded areas represent 95% confidence intervals of the mean. (b) Mean gaze threshold (angle of 50% reported as direct, which was estimated from the psychometric function) as a function of participant sex (female = red, male = blue) × face orientation (upright face = filled, inverted face = open). Error bars represent 95% confidence intervals of the mean. (c) Scatter plot and density curves of individual upright and inverted gaze thresholds (female = red, male = blue). The grey dashed line indicates identical gaze thresholds between upright and inverted faces (i.e., no gaze inversion effect).

Figures

Figure 1

(a)

(b)

(c)

Figure 2