3 Investigative power of Genomic Informational Field
4 Theory (GIFT) relative to GWAS
5 for genotype-phenotype mapping

Panagiota Kyratzi[1,2], Oswald Matika[3], Amey H Brassington[4], Connie E Clare[5], Juan Xu[6], David A
Barrett[7], Richard D Emes[8], Alan L Archibald[3], Andras Paldi[2], Kevin D Sinclair[4], Jonathan Wattis[9]
and, Cyril Rauch[1]

[1]School of Veterinary Medicine and Science, University of Nottingham, College Road, Sutton Bonington,
LE12 5RD, UK.
[2]École Pratique des Hautes Études, PSL Research University, St-Antoine Research Center, Inserm U938, 34
rue Crozatier, 75012 Paris, France.
[3]Div. Genetics and Genomics, The Roslin Institute and Royal (Dick) School of Veterinary Studies, University
of Edinburgh, Easter Bush, Midlothian EH25 9RG, Scotland, UK
[4]Agriculture and Horticulture Development Board, Middlemarch Business Park Siskin Parkway, East
Coventry CV3 4PE, UK.
[5]School of Biosciences, University of Nottingham, College Road, Sutton Bonington, LE12 5RD, UK.
[6]Shanghai Leadingtac Pharmaceutical Co., Ltd, 781 Cailun Road, China (Shanghai) Pilot Free Trade Zone,
Pudong, Shanghai 201203, China
[7]Centre for Analytical Bioscience, School of Pharmacy, University of Nottingham, Nottingham NG7 2RD,
UK.
[8]Nottingham Trent University, 50 Shakespeare Street, Nottingham NG1 4FQ, UK.
[9]Centre for Mathematical Medicine and Biology, School of Mathematical Sciences, University of
Nottingham, University Park, Nottingham NG7 2RD, UK.

Correspondence: Cyril Rauch (cyril.rauch@nottingham.ac.uk).

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30 ABSTRACT
Identifying associations between phenotype and genotype is the fundamental basis of genetic analyses.
Inspired by frequentist probability and the work of R.A. Fisher, genome-wide association studies (GWAS)
extract information using averages and variances from genotype-phenotype datasets. Averages and
variances are legitimated upon creating distribution density functions obtained through the grouping of
data into categories. However, as data from within a given category cannot be differentiated, the
investigative power of such methodologies is limited. Genomic Informational Field Theory (GIFT) is a
method specifically designed to circumvent this issue. The way GIFT proceeds is opposite to that of GWAS.
Whilst GWAS determines the extent to which genes are involved in phenotype formation (bottom-up
approach), GIFT determines the degree to which the phenotype can select microstates (genes) for its

subsistence (top-down approach). Doing so requires dealing with new genetic concepts, a.k.a. genetic
paths, upon which significance levels for genotype-phenotype associations can be determined. By using
different datasets obtained in *ovis aries* related to bone growth (Dataset-1) and to a series of linked
metabolic and epigenetic pathways (Dataset-2), we demonstrate that removing the informational barrier
linked to categories enhances the investigative and discriminative powers of GIFT, namely that GIFT
extracts more information than GWAS. We conclude by suggesting that GIFT is an adequate tool to study
how phenotypic plasticity and genetic assimilation are linked.

NEW & NOTEWORTHY

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The genetic basis of complex traits remains challenging to investigate using classic GWASs. Given the
success of gene editing technologies this point needs to be addressed urgently since there can only be
useful editing technologies if precise genotype-phenotype mapping information is available initially. GIFT
is a new mapping method designed to increase the investigative power of biological/medical datasets
suggesting, in turn, the need to rethink the conceptual bases of quantitative genetics.

Keywords: Complex traits; GIFT; genotype-phenotype mapping studies; GWAS

INTRODUCTION

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Identifying associations between phenotype and genotype is the fundamental basis of genetic analysis.
The development of high-density genotyping and whole genome sequencing has enabled DNA variants to
be directly identified and Genome-Wide Association Studies (GWASs) have become the method of choice
for mapping genotype to phenotype in large populations of unrelated individuals. GWAS have been
employed in many species, and especially in the study of human disease (1). By 2021 the NHGRI-EBI GWAS
Catalog listed 316,782 associations identified in 5149 publications describing GWAS results (2).
Additionally, extensive collection of data has been initiated through efforts such as the UK Biobank (3),
Generation Scotland (4) and NIH *All of Us* research program (https://allofus.nih.gov/) in the expectation
that large-scale GWAS will elucidate the basis of human health and disease and facilitate precision
medicine.
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While genomic technologies have advanced rapidly, statistical models used to analyze genetic data are
still based on the models developed by Fisher more than 100 years ago (5, 6). GWASs essentially make
use of the Fisher method of partitioning genotypic values by performing a linear regression of phenotype
on marker allelic dosage (7). Regression coefficients estimate the average allele effect size, and the
regression variance is the additive genetic variance due to the locus (8). However, an ongoing debate exists
over whether the present analysis paradigm in quantitative genetics is at its limits for truly understanding
complex traits, namely traits resulting from many genes each with very small effect size (9). As a result,
one may wonder whether alternative statistical model(s) could be invented and used to determine
genotype-phenotype mappings.
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GWASs are fundamentally linked to frequentist probabilities that, defined through relative frequencies,
determines the validity of statistical inferences. In practice, frequentist probabilities are generated
through the grouping of data into bins or categories to generate a bar chart, that is then interpolated to
create a distribution density function (DDF) in the continuum limit. The DDF is, in turn, used to determine
statistical inferences including average, variance, p-value and so on. However, since the DDF approximates

the bar chart (and not the converse), and that it is not possible to differentiate data from within any given
group/category, the DDF is constructed mathematically on the implicit assumption that information is
missing to differentiate data from within any given group/category.
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The notion of 'missing information' can be legitimate and defined experimentally. For example, measuring
the phenotype human height with a ruler with centimetre graduations implies that any height can be
measured to the nearest centimetre. Consequently, one centimetre-width bins/categories need to be
used to generate a frequency table of range of phenotype values upon which the phenotype and genotype
DDFs are defined. In this case, all the resulting statistical inferences are defined with a precision
corresponding to the nearest centimetre. The 'missing information' (i.e., that what cannot be measured
by the ruler) corresponds then to sub-centimetric scales (i.e., distances to the nearest millimetre for this
example). In practice the 'missing information' is therefore linked to the one of 'imprecision' and deciding
to provide more precise statistical inferences implies that the width of categories be reduced, which can
only be achieved by increasing the sample size. It is not by chance that the 'normal distribution' created
by mathematicians and physicists was initially called the 'law of errors', where the notion of error
(misinformation) results from imprecisions in experimental measurements. As a result, GWAS is faced
with a fundamental issue involving the extraction of precise information using a method that,
conceptually, assumes that information is missing or that data is mis-(in)formed.
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In general, the problem concerning the 'missing information' is never mentioned since the DDF in the
continuum limit is never considered as an approximation but as something that has its own reality. Namely
a DDF must exist independently of data measured (i.e., data must fit the DDF and not the converse). The
latter remark leads to an interesting conceptual territory where the notions of average and variance, and
their usage, may be questioned. If one considers the normal distribution (or any other DDFs) is inherent
to life and that data must fit it (them), then the moments of the distribution (e.g., average and variance)
are also essential parameters to describe life, and the variance often interpreted as noise in the data is
then a nuisance. If, on the contrary, data is the important thing, and that the DDF is considered solely as
a tool to interpolate data based on missing information, then average and variance are parameters
derived from a lack of information and are, as a result, poorly informative. The latter point should not
come as a surprise as reducing the huge diversity of populations to a handful of parameters (i.e., average
and variance) is highly reductionist and likely to be poorly descriptive. Thus, while the notions of average
and variance may help representing datasets, they are inventions nonetheless, i.e., thought constructions
akin to the field of frequentist probability. Thus, using average and variance as a starting point to map
genotype-phenotype (GWAS) is a matter of choice. Accordingly, different statistical methods can be
suggested.
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To avoid those conceptual and practical issues a new method called GIFT (Genomic Informational Field
Theory) has been designed and applied to simulated genotype-phenotype data in (10, 11), reviewed in
(12). In short, to associate genotype to phenotype GIFT does not presume that the only important
information concerning the gene effect is found in averages or variances, nor does it presume that DDFs
are central. On the contrary, GIFT starts with the pre-requisite that phenotypic values, or phenotypic
residuals after considering the environment/fixed effects, may be measured with sufficient precision to
be unique in a population. Then, by avoiding grouping data into bins/categories, which would otherwise
create an artificial imprecision, GIFT considers the entire information contained in the data, (i.e., variance
is not a nuisance anymore) making use of the cumulative sum of microstates. Figure 1 provides the
intuition underscoring GIFT as a method.
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The current article extends our previous theoretic studies using simulated data to analyse for the first
time two real datasets:
i.   Dataset-1 is derived from a study concerned with the genetic background of carcass composition in
sheep (*ovis aries*) (13). Using GWAS this study demonstrated a strong association between
chromosome 6 and the carcass composition trait 'bone area at the ischium'. We now apply GIFT to
reanalyse this dataset to benchmark it against GWAS. Since GWAS previously identified a QTL in
chromosome 6, our hypothesis was that GIFT would at least replicate GWAS results and identify
additional putative QTLs.
ii. Dataset-2 comprises biochemical data arising from an ongoing study in sheep which seeks to identify
risk allele variants in genes whose products direct a series of metabolic pathways, collectively referred
to as one carbon (1C) metabolism and associated epigenetic regulators. The gene array was designed
to include all single nucleotide polymorphisms (SNPs) linked to known biochemical enzymes involved
in these pathways. Given that Dataset-2 preselected genes for a targeted analysis of enzymes involved
in these metabolic/epigenetic pathways, it can be considered more specific.
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The present article initially introduces the reader to the way data may be used and analysed differently
using GIFT, contrasting to more conventional methods mostly based on an analysis of averages and
variances. More specifically in Part 1, the null hypothesis defined by GIFT will be established. Using
Dataset-1 the concept of genetic path pertaining to GIFT will be introduced (Part 2) out of which a p-value
for GIFT will be defined (Part 3). Then Dataset-1 (Part 4) and Dataset-2 (Part 5) will be analysed comparing
the informational/investigative power of GIFT relative to GWAS using Manhattan plots prior to performing
enrichment analyses.

MATERIALS AND METHODS

Biological datasets.

The first dataset (Dataset-1) analysed 600 pedigree-recorded Scottish Blackface lambs using CT scans to
determine *in vivo* carcasses composition (13). The trait selected for the present study is the bone areas of
the ischium (BAI) measured in mm^2 from cross-sectional CT scans. The ischium is one of the three bones
that make up the pelvis. It is located beneath the ilium and behind the pubis. The upper portion of the
ischium forms a major part of the concave portion of the pelvis that forms the hip. The BAI crossed a
genome-wide significance threshold on Chromosome 6 (OAR6). The pre-corrected phenotype values were
obtained fitting fixed effects of age of dam, year of birth, the effect of management group (as sheep were
from different farms), sex (males or females) and litter size (singles or twins) and as covariate the day of
birth. Further information can be found in Matika et al. (2016) (13). Supplemental S1 provides the raw
data used (Dataset-1).
The second dataset (Dataset-2) was from previously unpublished data extracted from a large ongoing
programme of research to investigate genome regions (Quantitative trait loci, (QTL)) that determine
metabolic and epigenetic responses to nutritionally induced deficiencies in one carbon metabolism (14,
15). For this study sheep were used as an experimental model. All animal procedures relating to this study
adhered to the Animals (Scientific Procedures) Act, 1986. Associated protocols complied with the ARRIVE
guidelines and were approved by the University of Nottingham Animal Welfare and Ethical Review Body
(AWERB) with Home-Office project licensed authority (30/3376;10th February 2016). Supplemental S2
provides the raw data used (Dataset-2).
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 Dataset-2: Sheep genome resequencing, custom array design and SNP profiling on test subjects.
Twenty-four unrelated Texel ewes were sequenced to a depth of 30x in 2 pools at Edinburgh Genomics.
DNA samples were prepared using Illumina's TruSeq PCR free kits and sequenced on an Illumina HiSeq
2500 Rapid Mode (serial no. D00125), read length of 150PE. Reads were trimmed to remove adapter
sequences and low-quality bases using skewer with commands (-Q 20, -q 3) (16) and mapped to the
reference sheep genome assembly (Oar_v3.1) using bwa mem (options -M -t 4) (17). Following
deduplication using Picard-tools version 1.92, variants were called using GATK pipeline (18) including
realignment around known indels and recalibration of bases, and FreeBayes (--use-best-n-alleles 4 --
pooled-discrete --min-alternate-count 4). Annotation of SNPs was performed using Ensembl variant effect
predictor VEP version ensembl tools release 79 (19). 15,347,831 variants were identified. Of these, ~3
million were novel SNPs and ~12 million were already present in the Ensembl genome database. SNPs
within annotated coding regions (VEP annotated "downstream gene variant" or "intron variant" removed)
and within 3Kb upstream of a gene were retained. SNPs with a minor allele frequency of greater than 0.5
were used to design an Illumina Infinium® iSelect® Custom Array consisting of 4,576 probes. This captured
SNPs in 115 1C metabolism and related genes, and 108 related epigenetic regulators as well as 33 control
SNPs (Supplemental S1).
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Liver samples were next collected post-mortem from 360 male and female Texel lambs (6 to 11 months
of age) representing 11 farms dispersed regionally across the UK. Collections took place at regional
abattoirs and samples immediately snap frozen in liquid N and stored at -80°C until analyses. DNA was
then extracted using AllPrep DNA/RNA Mini kit (Qiagen, Manchester UK). Briefly approximately 20 mg of
liver were mechanically disrupted using a TissueLyser (Qiagen, Manchester, UK) in 600 RLT plus buffer
containing β-mercaptoethanol. Tissue lysates were then used to extract RNA and DNA according to the
manufacturer instructions. The custom designed array was then used to SNP profile DNA from these Texel-
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mercaptoethanol. Tissue lysates were then used to extract RNA and DNA according to the manufacturer
instructions.
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Dataset-2: Metabolic profiling.
For the purposes of the current study the following seven liver metabolites were selected from a larger
pool of 1C metabolites: S-adenosyl methionine (SAM), methylcobalamin (mB12), adenosylcobalamin
(aB12), trimethylglycine (TMG), dimethylglycine (DMG), propionate (PPA) and methylmalonic acid (MMA).
The first four metabolites were selected as representative intermediates of the methionine cycle whilst
the latter two are intermediates in the hepatic synthesis of succinate (15) (Fig.2 & Supplemental S1).
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Hepatic concentrations of four metabolites (i.e., mB12, aB12, TMG and DMG) were determined by
hydrophilic interaction chromatography (HILIC) coupled to electrospray ionization tandem mass
spectrometry (MS/MS) as reported previously (20). For the analysis of SAM (determined separately by
HILIC), the standard was purchased from Sigma-Aldrich (Poole, Dorset, UK). Stock solutions of this
standard were prepared in potassium phosphate extraction buffer (KH_2PO_4 and K_2HPO_4; 40 mmol/L)
containing 0.1% L-ascorbic acid, 0.15% citric acid and 0.1% MCE (adjusted to pH 7 with NaOH), each at a
final concentration of 100 µmol/L. Also, for SAM the mobile phase was modified from that used for the
three other reported metabolites by adjusting the pH of the aqueous ammonium carbonate buffer

solution from 3.5 to 9.1. Mass spectrometer parameters for SAM were as follows: retention time = 7.69
min; Q1mass = 399.1 amu; Q3 mass = 250.1 amu; declustering potential = 56; collision energy = 25;
collision cell exit potential = 16.

Hepatic concentrations of PPA and MMA were determined by gas chromatography coupled to mass
spectroscopic-detection (GC-MS). Briefly, for PPA, 750 µL 5-Sulfosalicylic acid (SSA, 0.04 mg/ml) was
added to 150mg frozen liver, homogenised for 2 min and cooled on ice for 10 min. The sample was
centrifuged for 15 min at 14,500 x g and 200 µL liver homogenate transferred to a 2.5 mL screw capped
glass vial. To this, 20 µL internal standard (MBA, 400 µM), 3.5 µl HCl (37%) and 1 mL diethylether were
added, vortexed for 2 min and centrifuged for 10 min at 14,500 x g. 600 µL of the upper layer was
transferred to a screw capped glass vial containing 3.5 µL 1-(tert-butyldimethylsilyl)imidazole (TMDMSIM,
97%), vortexed for 2 min and heated at 60°C for 30 min. GC-MS analysis proceeded after cooling. The
method used a DB-5MS column (J&W Scientific Agilent technology, 30 m x 0.25 mm; 0.25 µm film
thickness). The carrier gas (He) was set at a constant flow rate of 1.3 ml/min. The injection volume was 5
µL for SCAN mode (for qualification) and SIM (selected ion monitoring) mode (for quantification), both
using splitless mode. The injection port and MS selective detector interference temperatures were 260°C
and 250°C respectively. The chromatograph was programmed for an initial temperature of 40°C for 1 min,
increased to 60°C at 70°C min-1, then to 110°C at 15°C min-1, and finally 250°C at 70°C min^{-1}. MS was
tuned regularly and operated in electron impact (EI) ionization mode with the ionization energy of 70eV.
SCAN mode measured at m/z: 30-300 and SIM ions were set at 159 (for MBA) and 131 (for PPA). The same
method was used to produce a calibration curve for PPA using standards at concentrations ranging from
19.5 nmol/g to 5µmol/g. The limit of detection was 19.5 nmol/g. CVs for low, medium and high QCs were
10.4, 6.3 and 6.5% and the inter-assay CV was 4.7%.

For MMA, 250 µL 80% MeOH was added to 50 mg frozen liver, homogenised for 2 min and cooled on ice
for 10 min. The sample was ten centrifuged for 15 min at 14,500 x g and 200 µL liver homogenate
transferred to a 2.5 mL screw capped glass vial. To this, 4 µL internal standard (1 mM 4-chlorobutyric acid
(CBA) in 1 mM HCl) followed by 250 µL 12% BF3-Methanol were added, vortexed for 1 min and heated at
95°C for 15 min. After cooling, 250 µL cold distilled water and 250 µL cold dichloromethane (CH_2Cl_2) were
added to the vial, vortexed for 30s and centrifuged for 10 min at 14,500 x g. The lower dichloromethane
layer was transferred to a screw capped glass auto-sampler vial with insert for GC-MS analysis. The
method used a DB-WAX column (cross-linked polyethylene glycol; J&W Scientific Agilent technology) (30
mm x 0.25 mm; 0.15 µm film thickness). The carrier gas (He) was set at a constant flow rate of 1.0 ml/min.
The injection volume was 1 µL for SCAN mode (for qualification) and SIM mode (for quantification), both
using splitless mode. The injection port and MS selective detector interference temperatures were 260°C
and 280°C respectively. The chromatograph was programmed for an initial temperature of 50°C for 2 min,
increasing to 150°C at 8°C min-1, then to 220°C at 100°C min^{-1} and held for 5 min at the final temperature.
MS was tuned regularly and operated in EI ionization mode with the ionization energy of 70eV. The limit
of detection was 0.75 nmol/g for both MMA and SA and inter-assay CVs were 8.4% for MMA and 11.0%
for SA.

Dataset-2: Determination of GWAS for 1C-metabolites.

Preliminary data analysis indicated the need to log-transform using the natural logarithm (Supplemental
S3) to approximate normality. Transformed data were then pre-corrected for the fixed effects of farm (F)
and sex (S) in ASReml using the following model, $y_{ij} = \mu + F_i + S_j + e_{ij}$, where y_{ij} is the log-transformed
phenotype, that is the log-transformed metabolite concentration studied; μ is the overall mean for the
log-transformed metabolite concentration; F_i is the effect of the ith farm (i =1,..,11); S_j the effect of jth Sex

(Male vs Female) and, e_{ij} is the residual. The genotype dataset was filtered using PLINK (HWE p-value
threshold of 10^{-6}, call rate for genotypes of 10% and a MAF of 5%), the number of independent SNPs was
determined using BCFTOOLS (r^2-threshold=0.1) and the GWAS Manhattan plots, linked to the
determination of p_{GWAS}, were obtained using GEMMA. The same genotype and residual phenotypes as
filtered by GWAS were used by GIFT.

Data representation using GIFT.
Adjusted phenotypic data (i.e., residuals, from Dataset-1 and Dataset-2) were used for this study.
Regarding the representation of GIFT, upon selecting a SNP for all individuals, the different corresponding
genotypes, aa, aA/Aa and AA, were assigned the arbitrary values +1, 0 and -1, respectively. With this
convention any barcode can be represented by a string of numbers from which a GIFT analysis can be
inferred. More specifically, the assignment of values +1, 0 and -1 were done as a function of the base pairs
as follow: AA=TT=+1, GG=CC=-1 and 0 otherwise. As shown schematically in Fig.1, the residuals obtained
were ranked by order of magnitude and the cumulative sum of their corresponding genotypic values
performed to obtain the 'genetic path' for the SNP considered. The genetic path of a SNP is noted $\theta(i)$ in
the text (Fig.1). The null hypothesis for GIFT as well as the notion of significance when GIFT is used will be
introduced and fully explained in the RESULTS section.
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RESULTS
Analyze of the null hypothesis $\theta_0(i)$ for GIFT
While $\theta(i)$ is obtained using phenotypic information (configuration ① in Fig.1 and 'Data representation
using GIFT' in MATERIALS AND METHODS), it is also possible to plot the cumulative sum of microstates
when no phenotypic information is present that is equivalent to 'scrambling' or permuting the string of
microstates in Fig.1A also corresponding to the configuration ② in Fig.1B. Recall that since our focus is
on a given SNP, then the number of microstates, N_+, N_0 and N_-, are identical between the configurations
① and ②. This new cumulative sum noted $\theta_0(i)$ is expected to be a sort of null hypothesis solely
dependent on the bulk microstate frequencies N_+/N, N_0/N and N_-/N, where N_q $q \in \{+,0,-\}$ is the
number of microstates of type q. This is so because there is no further information that could inform on
the positioning of microstates in their list when the scrambled state is considered. However, while $\theta(i)$ is
unique since phenotypic information is used to generate it, $\theta_0(i)$ is not as each time the string of
microstates from Fig.1A is scrambled, a new $\theta_0(i)$ appears. Accordingly, one needs to consider the set of
possible $\theta_0(i)$s generated bounded to the microstate frequencies N_+/N, N_0/N and N_-/N.
Using a selection of theoretic SNPs defined by different microstate frequencies (Table-1). Fig.3A illustrates
the global shape resulting from simulating 1000 $\theta_0(i)$s. The results demonstrate that the global shape of
the $\theta_0(i)$s plotted as a function of the position in the string is ellipsoidal with short and long axes changing
as a function of microstate frequencies involved, and where the different averages of $\theta_0(i)$s represented
by black lines in Fig.3A, are straight lines with slopes linked to the difference, $\Delta N/N = (N_+ - N_-)/N$. The
fact that the averages of $\theta_0(i)$s for a given set of microstates, N_+, N_0 and N_-, is always a straight line
linked to microstate frequencies, N_+/N, N_0/N and N_-/N, can be understood intuitively by the fact that
scrambling or permuting an infinite number of times the string of microstates is equivalent to
determining, for any position i, the presence probability, N_q/N, of each microstate in the string.
Accordingly, for a given set of microstates, N_+, N_0 and N_-, the average of $\theta_0(i)$s, noted $\langle\theta_0(i)\rangle$, is
$\langle\theta_0(i)\rangle = \frac{(N_+ - N_-)}{N} i$. Further theoretic details can be found in (10, 11). Using $\langle\theta_0(i)\rangle$ as a reference for the

null hypothesis, Fig.3B show the sur-imposition of the differences, $\Delta\theta_0(i) = \theta_0(i) - \langle\theta_0(i)\rangle$, obtained
from simulations using SNPs from Table-1.
Finally, to assess the impact of the sample size (population size) on the null hypothesis the initial size
(N=565, Table-1) was divided (N=280) and multiplied (N=1130) by a factor ~2 while keeping constant the
microstate frequencies N_+/N, N_0/N and N_-/N from Table-1. The simulations Fig.3A show that the
appearance of ellipsoids is affected when the sample size changes, becoming thinner as the population
size increases. Plotting the standard deviation, $\sigma(i/N)$, as a function of the position once normalized by
the sample size, $\sigma(i/N) = \sqrt{\left[\langle(\theta_0(i/N))^2\rangle - \langle\theta_0(i/N)\rangle^2\right]/N}$, resulting from the different simulations in
Fig.3C demonstrates that the standard deviation from GIFT is quadratic, and independent of the sample
size, as expected from a random allocation of different microstates in the string of positions.
At first sight and with this primary analysis one could suggest that any genetic path departing from the
cloud of genetic paths formed by the set of $\theta_0(i)$s upon the permutation of microstates (grey surface in
Fig.3A or black surface in Fig.3B) would likely result in an association between the genotype and the
phenotype. While true this assumption needs to be handed out carefully as it is not exhaustive. Indeed,
some genetic paths may be highly structured and of relatively small amplitude. Examples of genetic path
using real data from Dataset-1 will demonstrate this point.
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Examples of genetic path using the bone area of the ischium (BAI) as phenotype (Dataset-1)
The resulting average, $\langle\theta_0(i)\rangle$, and variance, $\sigma(i)$, can be used to inform the null hypothesis of a particular
SNP from 'real' datasets. However, since there are as many different sets of $\theta_0(i)$s as number of SNPs,
each SNP will return its own $\langle\theta_0(i)\rangle$ (null hypothesis) upon scrambling. A comparison between SNPs using
GIFT/genetic paths requires then to concentrate on the differences, $\Delta\theta(i) = \theta(i) - \langle\theta_0(i)\rangle$. In the
remaining text one shall rewrite $\langle\theta_0(i)\rangle$ as $\theta_0(i)$ to simplify notations.
Concentrating now on 'real' dataset, the genetic paths were obtained further to ranking BAI residual
values (Dataset 1) using an incremental rank from small to large values. As an example, Fig.4 shows the
two genetic paths $\theta(i)$ and $\theta_0(i)$ for six SNPs, renamed SNP1-6 (see Table-2 for accurate genetic
information) enabling us to appreciate the qualitative difference between the genetic paths. While the
null hypothesis, i.e., $\theta_0(i)$, resulting from the scrambling of phenotypic values many times always returns
a straight line with a different slope for each SNP as seen above, the $\theta(i)$s have different shape. To
represent the set of $\theta(i)$s in relation to the different microstates involved, each datapoint of the $\theta(i)$s is
colour coded as in Fig.1C.
Since $\theta_0(i)$ is linked to the difference between the genetic microstate frequencies of homozygotes, $\Delta N = $
$N_+ - N_-$, in Fig.4 we represent by the angle α such difference. Since $\tan(\alpha) = +N_+/N - N_-/N$ where N
is the total number of positions $(i = 1, 2, ..., N)$ $\theta_0(i)$ can be rewritten as, $\theta_0(i) = \tan(\alpha)i$. As any analysis
must concentrate on the difference, $\Delta\theta(i) = \theta(i) - \theta_0(i)$, such as to cancel the apparent variability in
the null hypothesis across SNPs, we represent the plots of the different $\Delta\theta(i)$s obtained in the right panel
of Figs.4A-4F.
Figs.4A-4B display two distinct genetic paths that are globally similar. While they have different number
of microstates of each type (see Table-2) the $\Delta\theta(i)$s of SNP1 and SNP2 are characterized by their small
amplitudes and the fact that they are erratic crossing several times the axis of position corresponding to
the null hypothesis. In those cases, using the information contained in the phenotypic residuals, namely
ranking the phenotypic residuals from small to large values, does not permit to fully differentiate $\theta(i)$
from $\theta_0(i)$. On the other hand, the right panel in Figs.4C-4D for SNP3 and SNP4 demonstrates, in a more

noticeable way, a paraboloid shape for the $\Delta\theta(i)$s resulting from a segregation of microstates upon
ordering the phenotypic residuals. The segregation of microstates +1 and -1 in opposite direction is
reminiscent of Fisher theoretic works (Fig.1). As it turns out Figs.4C-4D show some similarities with Fig.1C
based on a simulation inspired by Fisher's seminal works. Importantly the ΔN-values of SNP1 and SNP4
while of opposite sign are similar in absolute value, are as those of SNP2 and SNP3, suggesting, in turn,
the ΔN-values do not impact on the ability to differentiate $\theta(i)$ from $\theta_0(i)$. Namely that a segregation of
microstates can be inferred also with relatively large and opposed ΔN-values.
Envisaging the migration of microstates +1 and -1 in opposite direction as initially postulated by Fisher as
the sole framework to associate genotype and phenotype is not always valid. This is demonstrated by
SNP5 and SNP6 and the appearance of structured genetic paths displaying clear sigmoidal shapes for the
$\Delta\theta(i)$s as shown in Figs.4E-4F. Theoretically this phenomenon can be understood and explained by the
presence of non-linear phenotypic fields, see (11) also reviewed in (12), in turn breaking the symmetry
postulated by Fisher assuming the sole presence of linear phenotypic fields. This type of sigmoidal shapes
is of interest since they inform on potential regulation mechanisms involving very probably 'regulatory
variants' (21). Indeed, the right panels in Figs.4E-4F can be envisioned as representing the genetic
organization of two distinct subpopulations of phenotypic residual values, one above the dashed line and
the other one underneath it. Taken separately those two subpopulations draw curves like Figs.4C-4D or
Fig.1C. In this context it is tempting to suggest that sigmoid genetic paths reveal a type of genotype-
phenotype association that is inherently 'scale-dependent', namely function of the magnitude of
phenotypic residuals. Because traditional GWAS concentrates on averages and variances, these sigmoid
paths would be remarkably difficult to characterize with traditional methods. This is so because there is
no clear antisymmetric segregation of microstates. As an example, using SNPs1-6 (from Fig.4) we have
plotted, in Fig.5, the average values of phenotypic residuals for each microstate, and in Table-2 we provide
the resulting gene/size effects and the dominances associated with those. Fig.5 and Table-2 demonstrate
that sigmoid genetic paths (SNP5 and SNP6) are much less detectable with traditional methods while
paraboloid genetic paths (SNP3 and SNP4) are. Note that the numerical determination of '$-\mathrm{Log}_{10}(\mathrm{p_{GIFT}})$'
in Table-2, that is the significance for GIFT, is explained in the next part below.
To conclude, based on Fisher's theoretic works, the traditional GWAS method has been optimized to map
SNPs that, using GIFT, would draw paraboloid genetic paths (see Fig.1C). The potential novelty using GIFT
resides in its ability to provide new information and detect relatively regular/structured sigmoid genetic
paths that would otherwise not be detected by traditional methods.
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$\mathrm{p_{GIFT}}$: p-value for GIFT
GIFT and GWAS extract information on genotype-phenotype associations in totally different ways. While
GIFT concentrates on the significance of curves drawn using $\Delta\theta(i) = \theta(i) - \theta_0(i)$, GWAS focuses solely
on the significance of difference of averages. However, to compare GIFT to GWAS it is essential to
determine a p-value for GIFT that is exhaustive enough such as to also capture the information that GWAS
provides. To this end a p-value was derived that concentrates on the maximal amplitudes difference of
genetic paths (see Figs.6A-6B).
The p-value for GIFT can be understood as follows. Since the number of possible paths is linked to the
number of configuration possible resulting from lodging N_+, N_0 and N_- microstates into a list composed
of $N = N_+ + N_0 + N_-$ components, the number of possible paths is, $N_{path}^0 = \frac{N!}{N_+!N_0!N_-!}$. Let us now divide
the genetic paths into regions, Δi_1, Δi_2 and Δi_3 as shown in Figs.6A-6B. As the number of microstates of
each sort can be determined in each region using an adequate algorithm, then the total number of

possible genetic paths in this first, second and third regions are, respectively, $N_1 = \frac{\Delta i_1!}{(n_+)_1!(n_0)_1!(n_-)_1!}$, $N_2 =$

$\frac{\Delta i_2!}{(n_+)_2!(n_0)_2!(n_-)_2!}$ and, $N_3 = \frac{\Delta i_3!}{(n_+)_3!(n_0)_3!(n_-)_3!}$, where $(n_q)_p$ is the number of microstate of type q in the pth

region, $q \in \{+,0,-\}$ and $p \in \{1,2,3\}$. Consequently, the probability of a genetic path in this context is,

$\hat{p}_{GIFT} = N_1 N_2 N_3 / N^0_{path}$. Using the null hypothesis simulations shown in Fig.3 based on the theoretic SNPs

given in Table-1, $\hat{p}_{GIFT}$ may be determined for each genetic path simulated. Its statistic plotted in Fig.6C

for each SNP demonstrates very little variations across SNPs or when the sample size changes by a factor

two. Based on this observation confidence intervals were determined for all SNPs by averaging the $\hat{p}_{GIFT}$

values obtained. The upper and lower red dashed lines represent the 99% and 95% confidence intervals.

To consider the false discovery rate (FDR) and adjust p-values to remove type-I errors, $\hat{p}_{GIFT}$-values in

Fig.6C were corrected using the Benjamini-Hochberg procedure leading to a new set of adjusted, i.e.,

reduced, p-values, noted p_{GIFT} (see Fig.6D), that may be used to determine the true significance of DNA

variants (SNPs). Returning to Table-2 the numerical value of p_{GIFT} was determined for the genetic paths

shown in Fig.4 demonstrating that GIFT can extract information when sigmoid genetic paths are involved

while traditional GWAS is unable to do so.

Armed with p_{GIFT} an analysis of datasets can now be performed.

Comparison between GWAS and GIFT considering the bone area of the ischium (BAI) as phenotype
(Dataset-1)

The first dataset (Dataset-1) analysed 567 pedigree-recorded Scottish Blackface lambs concentrating on

the bone areas of the ischium measured in mm^2 from cross-sectional CT scans (13). After adjusting

phenotypic values, the work demonstrated a clear involvement of chromosome 6 as shown in Fig.7A. The

genome-wide significant thresholds applied for GWAS in Fig.7A correspond to Bonferroni corrections at

1% (upper red dashed line) and 5% (lower dashed red line) determined by using independent SNPs only.

Formally a 1% (resp. 5%) Bonferroni correction is given by, $-Log_{10}(0.01/N_{ind-SNPs})$ (resp.

$-Log_{10}(0.05/N_{ind-SNPs})$ where $N_{ind-SNPs} = 10433$ is the number of independent SNPs. Using its own

thresholds (Fig.6D) GIFT was applied using the same set of phenotypic residuals. Figs.7A-7B demonstrate

the results obtained by GWAS and GIFT using Manhattan plots.

The significance threshold by GIFT was defined by a null hypothesis using theoretic SNPs. To demonstrate

that the theoretic results obtained from Fig.6D are transferrable to 'real' SNPs (Fig.7B), namely that the

significant SNPs obtained in Fig.7B have null hypotheses with similar properties like those shown in Fig.6D,

each significant SNP (Fig.7B) had its genetic path randomly permutated a thousand times to determine

the distribution of $-Log_{10}(p_{GIFT})$-values corresponding to their null hypothesis. Results show that the

null hypotheses are remarkably similar across SNPs and that the threshold determined using theoretic

SNPs (Fig.6D) holds when 'real' SNPs are used (Supplemental S5).

Overall, Fig.7A and Fig.7B demonstrate that there is an agreement between GWAS and GIFT that

chromosome 6 is involved. However, differences exist that are shown through the involvement of several

chromosomes when GIFT is used. Considering the thresholds involved, for GWAS the phenotype studied

may be considered as a sort of 'single gene trait' while for GIFT, the phenotype looks very much like a

'complex trait' involving more chromosomes than chromosome 6. Detailed information of all significant

SNPs by GWAS or GIFT is given in Supplemental S6.

Concentrating on Chromosome 6 to address the overlap of information provided by GIFT and GWAS, a

Venn-diagram including highly significant SNPs only, namely SNPs beyond the upper red dashed-line in

Figs.7A-7B, was plotted. The Venn-diagram (Fig.7C) reveals that most SNPs deemed significant by GWAS

were also deemed significant by GIFT. Curiously, only one SNP seemed highly significant by GWAS but

irrelevant for GIFT. As p_{GIFT} was designed to collect exhaustive information from GWAS, the SNP was

identified (OAR6_40311379) and its genetic path, i.e., its $\Delta\theta(i)$, plotted (Fig.7D-left) together with its

GWAS-representations (Fig.7D-right). The genetic path, being erratic of relatively small amplitude and
crossing several times the axis of positions, did not display any obvious 'parabolic or sigmoidal'
associations at first sight, in turn justifying its small p_{GIFT}-value. The GWAS-representation of
OAR6_40311379 however, demonstrated the absence of microstate '-1' as well as a near overlap of
microstates '0' and '+1' further demonstrated by the similarities between their boxplots, suggesting the
occurrence of a false-positive. To confirm this a comparison of phenotypic means for the microstates '0'
and '+1' was performed returning a t-test value of 1.1485 (p-value of 0.2512), confirming the presence of
a false-positive.
In order to assess the overlap of information between GWAS and GIFT we plotted in Fig.7E the first 100
more significant SNPs detected by GIFT and GWAS. Results confirm an overlap of SNPs associated with
the phenotypic residuals for large values of p_{GIFT} and p_{GWAS} (see purple dots in Q_2 in Fig.7E).
Interestingly, two SNPs considered as significant by GWAS (two blue dots in Q_2) were not by GIFT. That is
because the p_{GIFT}-values for these dots were less than other SNPs detected by GIFT. As already stated
above many SNPs from other chromosomes were considered significant by GIFT that were not by GWAS
(see red dots in Q_4). Finally, the quadrant Q_1 in Fig.7E confirms that OAR6_40311379, i.e., the false
positive detected by GWAS, is a standalone SNP among the 100 SNPs for which $p_{GWAS} > p_{GIFT}$. Finally,
the biotype of significant SNPs on Chromosome 6 for GIFT and GWAS are also presented in Fig.7F.
The primary conclusion provided by Figs.7A-F is that, when compared to GWAS, GIFT returns substantially
more genetic information.
However, a central question concerns the genetic pertinence of the significant SNPs obtained by GIFT. As
GIFT has been designed with the aim to increase the investigative power of biological datasets, we may
assume that the significant SNPs obtained by GIFT once translated into gene names should underline some
level of non-random gene-gene interactions. The latter point is particularly relevant since GIFT is expected
to detect regulatory variants (c.f. sigmoidal genetic paths). To assess this point we performed an
enrichment analysis based on gene names using the String database, which helps determine known and
predicted protein-protein interactions. In order to apply String the significant SNPs obtained using GWAS
and GIFT were mapped to the reference sheep genome assembly from ensembl (Oar_v3.1) to obtain the
gene names. Using those gene names String analyses were performed for GWAS and GIFT using a
minimum required interaction score of 0.4. Fig.7G and Fig.7H show the networks obtained. With
enrichment p-values for GWAS and GIFT of 0.176 and 0.00008, respectively, these results confirm that the
set of genes determined by GIFT have more interactions among themselves than what would be expected
for a random set of genes of the same size and degree distribution drawn from the genome. Namely that
GIFT increases the investigative power of biological datasets.
At present, we do not know how the whole information provided by GIFT may inform on the putative
biology of the phenotype studied (BAI). As it turns out, a full validation of the information provided by
GIFT on Dataset-1 would require an in-depth mutational/deletion/insertion/gene-editing analyses in live
animals, extending beyond the scope of this present article.
To demonstrate the relevance of the information provided by GIFT we decided to challenge GIFT using a
different dataset (Dataset-2) concentrating on a complex trait related to 1C-metabolism.
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Comparison between GWAS and GIFT considering 1C-metabolites as phenotype (Dataset-2)
Dataset-2 concerns biochemical data which seeks to identify risk allele variants in genes whose products
direct a specific series of metabolic pathways, known as one carbon (1C) metabolism (Fig.2). The
significance of 1C metabolism is that it is a complex trait involving a series of interlinking metabolic
pathways that provide 1C units (methyl groups) for the synthesis and methylation of biological molecules.
After 1% and 5% Bonferroni corrections for GWAS and the Benjamini-Hochberg procedure applied to GIFT,
the Manhattan plots were obtained (Fig.8A). Note that the number of independent SNPs in this case is
624 (out of 3923 SNPs from the gene array). Fig.8A demonstrates clearly that the informational power of

GWAS is less than that of GIFT. Finally, in Fig.8B we provide the biotypes of the most significant SNPs
shown by the upper red dashed lines obtained using GIFT. Detailed genetic information of the most
significant SNPs obtained using GIFT is provided in Supplemental S8.
Since the gene array was synthesized using SNPs from known genes involved in 1C metabolism, the
relevance of String analyses (i.e., enrichment p-values) would be minimal and of little interest.
Besides validating that GIFT may extract more information from genotype-phenotype datasets, it is worth
underlying the biological importance and novelty of results obtained. One carbon metabolism in sheep is
comparable to that in humans. The significance of 1C metabolism is that it is a complex trait involving a
series of interlinking metabolic pathways that provide 1C units (methyl groups) for the synthesis and
methylation of chromatin among other molecules (15). S-adenosylmethionine (SAM) is a potent methyl
donor within these cycles and serves as the principal substrate for methylation of DNA, associated
proteins, and RNA. It was previously demonstrated in sheep, cattle, rodent and human studies that
disrupting these cycles during early pregnancy, by either dietary means (i.e., reducing dietary vitamin B12,
folate, choline and/or methionine), or through exposure to environmental chemicals such as cigarette
smoking, can lead to epigenetic dysregulation and impaired foetal development with long-term
consequences for offspring cardiometabolic health (22–25). It was also advocated that interindividual and
ethnic variability in epigenetic gene regulation arises because of single-nucleotide polymorphisms (SNPs)
within 1C genes, associated epigenetic regulators, and differentially methylated target DNA sequences
(15). However, information concerning the nature and extent of interactions between parental genotype,
diet and EC exposure was, until now, limited to just a few 1C genes in humans (15). Consequently, data
obtained by the current study provide new evidence concerning significant genetic variants in 1C-
metabolism and directly associated metabolic genes and epigenetic regulators that rely on SAM as the
methyl donor, potentially applicable to the human species.

DISCUSSION

While statistical association methods should not favor any biases when analyzing datasets, the way they
are built mathematically is often indicative of a particular way of thinking. For example, with GWAS the
phenotype is decomposed onto more fundamental sub-distributions characterized by the distribution of
microstates (see Fig.1A). This approach underlines a sort of bottom-up approach that, within a
reductionist framework, defines genes as biological agents controlling the phenotype aligned with the
'Neo-Darwinian synthesis'. However, nothing prevents considering the opposite as far as statistical
association methods are involved, and GIFT uses this degree of freedom. By using the full range of
phenotypic information, GIFT transforms a random or disordered string of microstates (the straight line
in the asymptotic limit seen in Fig.1C or Fig.3A) into an 'ordered' configuration of microstates (see Fig1C
or Figs.4C-4F), in turn providing the signature of a genotype-phenotype association. Accordingly, since the
phenotypic information controls the configuration of microstates it is a top-down approach, which turns
out to be remarkably sensitive. GIFT has been estimated to be ~1000 more sensitive than GWAS (11).

There are three main reasons as to why GIFT is more sensitive. The first is that GIFT determines the
significance of curves composed of an entire population of datapoints. As curves provide a greater level
of significance than considering differences between microstate/phenotypic averages/variances as
advocated by GWAS, hence GIFT is statistically more powerful. The second reason is that the null
hypothesis for GIFT, namely $\theta_0(i)$, is contained in the definition of $\Delta\theta(i)$ and is therefore specific to the
genome position, or SNP, studied. With GIFT there are as many null hypotheses as SNPs. This contrasts
with GWAS defining a null-hypothesis valid for all SNPs at the population level when the average of
microstate distributions overlap. Consequently, the discriminative power of GIFT is amplified. The third
reason is that GIFT is simpler than GWAS. Indeed, based on R.A. Fisher's seminal work, GWAS is based on
a complex theory that seeks to determine genotype-phenotype associations on one hand (aim 1), and the

heritability of phenotypes/traits studied on the other (aim 2). To achieve those two aims, the GWAS
approach relies on frequentist probability to determine the validity of statistical inferences giving the
notions of average and variance fundamental meanings related to aim 1 and 2, respectively. However,
because average and variance are antinomic it is nearly impossible to have a clear picture of associations
(size effects) since the noise (variance/heredity) blurs the average(s). On the other hand, by concentrating
on genetic paths (curves) GIFT determines a global association. This does not mean that GIFT rules out the
notions of size effect, dominance, and heritability, on the contrary, it encapsulates them under the generic
notion of phenotypic field, i.e., size effect, dominance and heritability can be rederived from the
phenotypic field. The term 'field' in the acronym GIFT is used to explain the disorder-order transition in
the string of microstates using an analogy related to physics field theory, see (11, 12) for more details.

Finally, it is important to reframe GIFT within current debates in the field of biology. With GIFT it is the
(information on the) phenotype that selects which SNP is required for its subsistence and it is interesting
to note that, at the conceptual level and as a top-down approach, GIFT has some familiarity with the
notion of phenotypic plasticity. Phenotypic plasticity refers to the ability of phenotypes to respond to a
change in the environment favoring a divergence from the ancestor phenotype. As the phenotype relies
on traits (modules), the responsiveness to any new input(s) must involve a re-organization of the
phenotype architecture by allowing phenotypic sub-components (modular traits) to adapt the changes
(26). Namely that genetic accommodation linked to a standing pool of genetic variations characterizing
any trait is central to phenotypic plasticity that, through persistence, may genetically assimilate the new
architecture (selection) (26, 27). In this context the top-down method GIFT, which is essentially a
phenotype-genotype (and not genotype-phenotype) association method, can pull out any standing genes
awaiting to be used by phenotypes.

To conclude, we provide evidence that GIFT enhances the investigative power of biological datasets.
Additionally, we provide evidence also for the need to rethink the conceptual bases of genotype-
phenotype association methods, such as use more information from the whole biodiversity of data.

DATA AVAILABILITY

561 Data including supplementary materials are available using the link:

SUPPLEMENTAL MATERIAL

S1 provides the raw data for Dataset-1; S2 provides the raw data for Dataset-2. S3 provides the
statistical summary for the phenotypic adjustment prior to running GWAS on Dataset-2; S4 provides the
code to obtain Fig.3 and Fig.6C. S5 represents the permutation analysis of significant SNPs obtained by
GIFT from dataset-1. S6 represents the list of significant SNPs obtained by GWAS and GIFT when applied
on Dataset-1; S7 provides the code to obtain Fig.4, Fig5 and Fig.7. S8 provides the list of significant SNPs
by GIFT for Dataset-2. S9 provides the code to obtain Fig.8.

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DISCLOSURES

Authors declare no conflict of interest, financial or otherwise.

AUTHOR CONTRIBUTIONS

CR conceptualized GIFT; CR & JW formalized GIFT; PK coded GIFT simulations; Dataset-2 was designed and obtained by KDS, OM, AHB, CEC, JX, DAB, RDE; Dataset-1 and Dataset-2 were analysed by PK, CR, JW, KDS, OM, AP; paper was written by CR and KDS, and proofread by CR and KDS.

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FIGURE LEGENDS

Figure 1: (A) For diploid organisms and for a binary (bi-allelic, A or a) genetic marker, any microstate
(genotype) can only take three values that we shall write as '+1', '0' and '-1' corresponding to genotypes
aa, Aa and AA, respectively. The genotypes are color-coded to facilitate the representation of GIFT (+1: aa
(red), 0: aA/Aa (black) and -1: AA (blue)). GWASs rely on probability density functions formed through the
grouping of data into bins/categories. The phenotype distribution density function (A-top left) is then
decomposed onto the distribution density function of genetic microstates (A-top right) for every single
nucleotide polymophism (SNP). Using an analysis of averages and variances such decomposition
determines whether the SNP studied is associated with the phenotype by comparing the average and
variances of distributions. Repeating the same operation for every SNP in the genome permits to map
genotype to phenotype. However, as more precise inferences can only come with, and are only legitimized
by, a reduction in the width of categories, larger sample sizes are needed. To overcome this issue one way
to proceed is to deconstruct density functions and wonder what would happen if one were able to reduce
the width of categories, that is increasing the precision in the measurement of the phenotype or
equivalently getting access to the whole information of datasets, without changing the sample sizes (**A**
from top-to-bottom). The mathematical object that emerges is then a coloured barcode that is a list of
microstates that can be analysed precisely by GIFT. **(B)** Such barcode can be obtained simply at the
practical level through field studies. Assume a flock of sheep has been genotyped and that their phenotype
has been measured sufficiently precisely such as to exclude the possibility that any two phenotypic values
are identical. In the figure the magnitude of the phenotypic value for each sheep is characterised by the
(unique) 'size' of the sheep. The barcode is obtained by ranking animals as a function of the magnitude of
their phenotypic values (configuration ①in Fig.1B). The null hypothesis is obtained via the random
ranking of sheep that is equivalent to a lack of information on phenotypic values (configuration ② in
Fig.1B). As GWAS works on phenotypic residual values after adjusting for fixed/environmental effects a
similar barcode can be generated considering the magnitude of residual phenotypic values. **(C)** GIFT

proceeds by plotting the cumulative sum of microstates as a function of the position in the list generating
a curve called genetic path that is represented by $\theta(i)$ in Fig.1C and is unique to the SNP considered. While
the curve $\theta(i)$ does not provide any significant information on its own, one may generate, for the same
SNP, a curve (genetic path) corresponding to a sort of null hypothesis when ranking the phenotype does
not bring any informational value. This is possible by scrambling (permutating) the string of microstates
an infinite number of times. It is then possible to show that, in the asymptotic limit, the null hypothesis
returns a straight line, noted $\theta_0(i)$ (Fig.1C) out of which inferences may be suggested regarding potential
association between the genotype and the phenotype by comparing $\theta_0(i)$ to $\theta(i)$. Note, the simulation
shown in **(A)** adhering to Fisher seminal model is based on a constant sample size of 1000 involving an
arbitrary normally distributed phenotype of mean and variance 68 and 4 units, respectively. Each
microstate is normally distributed with a gene effect identical to the standard deviation of the phenotype
but without dominance. The frequency of the genotypes aa (red), Aa/aA (Grey) and AA (blue) are 64%,
32% and 4%, respectively and within Hardy-Weinberg ratio.

Figure 2: Linked methionine and propionate metabolism adapted from Clare et al. (2019) where all
metabolites studied for this study are in red. The methionine cycle facilitates the re-methylation of
homocysteine (Hcy) to methionine (Met) and ultimately S-adenosylmethionine (**SAM**) with methyl (CH$_3$)
groups donated either from folate (5-mTHF) or betaine (trimethylglycine; **TMG**), thus leading to the
formation of dimethylglycine (**DMG**). Methylcobalamin (**mB12**) serves as a cofactor for the reduction of
the inactive form of methionine synthase to its active state (MTR), which then transfers a methyl group
from 5-mTHF to Hcy. The linked metabolism of propionate (**PPA**) to succinate (an intermediary metabolite
in the tricarboxylic cycle) requires adenosylcobalamin (**aB12**), which serves as a cofactor for
methylmalonyl-CoA-mutase (MUT) leading to the generation of succinyl-CoA and methylmalonic acid
(**MMA**) in this pathway. Other intermediary metabolites and enzymes listed: glycine (Gly), sarcosine (Sar),
S-adenosylhomocysteine (SAH), tetrahydrofolate (THF), serine (Ser), cystathionine (Cth), cysteine (Cys),
alpha-ketobutyrate (α-KB), methylmalonic acid (MMA); Betaine homocysteine methyltransferase (BHMT),
Methionine adenosyl-transferase (MAT), Glycine methyl-transferase (GNMT), Adenosyl-homocysteinase
(AHCY), Cystathionine beta-synthase (CBS), cystathionine gamma-lyase (Cth).

Figure 3: **(A-left panel)** Simulations of genetic paths corresponding to null hypotheses using GIFT as a
method. The data used for the simulation are given in Table-1. **(A-right panel)** Simulations of genetic paths
corresponding to null hypotheses when the sample size is divided or multiplied by a factor two. **(B)**
Representation of $\Delta\theta_0(i) = \theta_0(i) - \langle\theta_0(i)\rangle$ for the microstates data as given in Table-1. **(C)** Plots of the
standard deviation normalised by the square root of the sample size and where the position is also
normalised by the sample size. The code for the simulations is given in Supplemental S4.

Figure 4: A sample of genetic paths selected from Dataset-1. The details of the different SNPs displayed
are given in Table 2.

Figure 5: Analysis of averages (GWAS) for SNP1-6 (see Fig.4 and Table-2). Values for the size/gene effects
(a) and dominances (d) are given in Table 2.

Figure 6: To provide a p-value extracting genotype-phenotype associations in an exhaustive manner for
both GWAS and GIFT a method concentrating on the largest and smallest extreme values of the genetic
path was focused upon. This method can be applied to paraboloid (GWAS or GIFT-like) **(A)** and sigmoid
(GIFT-like) **(B)** genetic paths. The overall idea consists in determining how many paths N_1, N_2 and N_3 can
be generated from the respective interval of positions Δi_1, Δi_2 and Δi_3 given that the constraints for the

extrema are Φ_1 and Φ_2. Then a p-value ($\hat{p}_{GIFT}$) can be determined as seen in the text. **(C)** Using
simulations (K=1000 replicates) a statistic of $\hat{p}_{GIFT}$ for the null hypothesis can be generated using theoretic
SNPs (Table-1). Simulations demonstrate that $\hat{p}_{GIFT}$ is relatively independent of the microstate's
frequencies upon which a 99% (upper dashed line) and 95% (lower dashed line) interval confidences can
be generated. **(D)** $\hat{p}_{GIFT}$-values were adjusted to consider FDR using Benjamini-Hochberg procedure
leading to a new set of p_{GIFT}-values. The code for the simulations is given in Supplemental S4.

Figure 7: Manhattan plots based on p-values obtained by GWAS **(A)** and GIFT **(B)** demonstrating significant
differences between the methods concerning potential genotype-phenotype associations. Note that the
presence of a chromosome '0' results from the fact that some SNPs identified by (Matika et al., 2016)
were not allocated to specific chromosomes/genomic positions due to lack of information at the time. A
fathom chromosome (chromosome zero) was created to allocate those SNPs. **(C)** Venn-diagram
representing the most significant SNPs by GWAS and GIFT. One SNP (OAR6_40311379) demonstrated a
large p-value for GWAS and a small p-value for GIFT. A representation of its genetic path **(D-left)** did not
underscore any 'parabolic' or 'sigmoidal' associations. As it turned out this SNP was a false-positive by
GWAS since the difference between the phenotypic means was not significative **(D-right)**. **(E)** The 100
most significant SNPs by GWAS and GIFT were extracted, and their p-values plotted against each other.
The dashed lines represent the threshold applied for GWAS (blue dashed line) and GIFT (red dashed line).
The SNP OAR6_40311379 pointed by the black arrow is the single one standing out in Q_1 confirming its
false-positive status. **(F)** Biotypes of the most significant SNPs by GIFT and GWAS. **(G)** String analysis
performed to determine gene networks using significant SNPs by GWAS. **(H)** String analysis performed to
determine gene networks using significant SNPs by GIFT, note that the dashed square underlines mTOR
and FOXO3 determined by GWAS. The code for obtaining Figs.7B, 7C, 7D, 7F is given in Supplemental S7.

Figure 8: (A) Comparison of the information extracted by GWAS and GIFT using Manhattan plots for the
metabolites presented in red in Fig.2. We recall the acronyms, S-adenosyl methionine (SAM),
methylcobalamin (mB12), adenosylcobalamin (aB12), trimethylglycine (TMG), dimethylglycine (DMG),
propionate (PPA) and methylmalonic acid (MMA). It should be noted that due to inherent difficulty linked
to the measure of metabolite the sample sizes were not similar across metabolites, that is the values for
N differ between the Manhattan plots (SAM: N=344; mB12: N=183; aB12: N=338; DMG: N=338; TMG:
N=340; MMA: N=348; PPA: N=345). **(B)** Biotypes corresponding to the most significant SNPs for each
metabolite determined by GIFT (a detailed list of information concerning those SNPs in given in
supplemental S8). The code for the Manhattan plots and the determination of biotypes is given in
Supplemental S9.

Table 1: Theoretics SNPs used to capture the null hypothesis associated with GIFT upon 1000 simulations of microstates permutation*.

SNP NAME	N_+	N_0	N_-	N
SNP1	25	25	515	565
SNP2	25	125	415	565
SNP3	25	225	315	565
SNP4	25	325	215	565
SNP5	25	425	115	565
SNP6	25	525	15	565

(*): The difference between consecutive SNPs in the table is linked to the transfer of 100 microstates from the microstates '-1' to the microstate '0' leaving the number of microstates '+' invariant. By permutating the microstates '+' and '-' in the table similar plots as those obtained in Fig.3A could have been obtained, the only difference would have been the slopes of the average $\langle\theta_0(i)\rangle$ changing sign.

Table 2: Determination of gene/size effect (a) and dominance (d) for SNP1-6 from dataset-1. The level of significance for GIFT and GWAS is colour coded: red=not-significant, green=significant.

CHR	NAME	POSITION	-$\text{Log}_{10}(p_{GIFT})$	-$\text{Log}_{10}(p_{GWAS})$	N_+	N_0	N_-	a*	d**
9	OAR9_58767921 (**SNP1**)	56039025	2.7895	0.2735	391	160	16	N/A	N/A
3	s02120 (**SNP2**)	213625709	2.8893	0.0018	198	291	78	N/A	N/A
6	OAR6_40855809 (**SNP3**)	36655091	28.5105	9.8639	229	262	76	96.85	-13.01
6	OAR6_38315830 (**SNP4**)	34256151	20.7541	3.7366	24	222	321	-70.02	-0.05
23	OAR23_35510473 (**SNP5**)	33556377	19.7239	0.2301	254	260	53	N/A	N/A
25	OAR25_30372586 (**SNP6**)	29046746	18.5806	1.0692	90	266	211	N/A	N/A

(a*): The gene/size effect is calculated considering the mid-distance between the average values of phenotypic residuals of microstates '-1' and '+1'. (d**): The dominance is calculated considering the difference between the gene/size effect (a) and the position of the average value of phenotypic residuals for the microstate '0'.

Fig.1

Fig.2

Fig.3

Fig.4

Fig.5

Fig.6

Fig.7

Fig.8

Fig.8 (continued)