Rapid adaptive evolution of microbial thermal performance curves

Megan H. Liu[1,*], Ze-Yi Han[1], Yaning Yuan[1], Katrina DeWitt[1], Daniel J. Wieczynski[1], Kathryn M. Yammine[2], Andrea Yammine[1], Rebecca Zufall[3], Adam Siepielski[4], Douglas Chalker[5], Masayuki Onishi[1], Fabio A. Machado[6], Jean P. Gibert[1,*]

[1]Department of Biology, Duke University, Durham, NC, 27708, USA
[2]Department of Chemistry, Massachusetts Institute of Technology, Cambridge, MA, 02139, USA
[3]Department of Biology and Biochemistry, University of Houston, Houston, TX 77204, USA
[4]Department of Biological Sciences, University of Arkansas, Fayetteville, AR, USA
[5]Department of Biology, Washington University, Saint Louis, MI, USA
[6]Department of Integrative Biology, Oklahoma State University, Stillwater, OK, USA

To whom correspondence should be addressed: mhliu58@gmail.com, jean.gibert@duke.edu

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ABSTRACT

Microbial respiration alone releases massive amounts of Carbon (C) into the atmosphere each

year, greatly impacting the global C cycle that fuels climate change. Larger microbial population

growth often leads to larger standing biomass, which in turns leads to higher respiration. How

rising temperatures might influence microbial population growth, however, depends on how

microbial thermal performance curves (TPCs) governing this growth may adapt in novel

environments. This thermal adaptation will in turn depend on there being heritable genetic

variation in TPCs for selection to act upon. While intraspecific variation in TPCs is traditionally

viewed as being mostly environmental (E, or plastic) as a single individual can have an entire

TPC, our study uncovers substantial heritable genetic variation (G) and Gene-by-Environment

interactions (GxE) in the TPC of a widely distributed ciliate microbe. G results in predictable

evolutionary responses to temperature-dependent selection that ultimately shape TPC adaptation

in a warming world. Through mathematical modeling and experimental evolution assays we also

show that TPC GxE leads to predictable temperature-dependent shifts in population genetic

makeup that constrains the potential for future adaptation to warming. That is, adaptive

evolution can select for decreased genetic variation which subsequently lowers the evolutionary

potential of microbial TPCs. Our study reveals how temperature-dependent adaptive evolution

shapes microbial population growth, a linchpin of global ecosystem function, amidst accelerating

climate warming.

INTRODUCTION

Microbes play a central role regulating the global carbon (C) cycle that controls climate change (Davidson & Janssens, 2006; Falkowski et al., 2008; Trumbore, 2006). Indeed, soil microbial respiration releases ~94Pg of C into the atmosphere every year (Bond-Lamberty, 2018; Stell et al., 2021; Xu & Shang, 2016) and microalgae provide the bulk of marine C fixation globally (30-50 Pg of C/yr) (Arrigo, 2005; Falkowski, 1994; Litchman et al., 2015). Global warming is expected to alter these microbial processes (Intergovernmental Panel on Climate Change (IPCC), 2023), but anticipating these effects requires a deeper understanding of the biotic and abiotic factors influencing microbial respiration in a warming world (Barton et al., 2020; Rocca et al., 2022; Wieczynski et al., 2023). One such factor is microbial population growth, which influences total standing biomass, and hence, total microbial respiration (Brown et al., 2004; DeLong et al., 2017; Gillooly et al., 2001; Savage et al., 2004).

To understand and anticipate the effects of global warming, we need to characterize the evolutionary processes that shape the microbial Thermal Performance Curves responsible for determining microbial population growth under novel climates. Microbial thermal performance curves are often measured as change in maximum (intrinsic) population growth rates (denoted r, the difference between per capita birth and death rates) across temperatures ('r-TPCs' henceforth, Fig 1). They thus determine how a species population growth rate will change in response to temperature changes. Controlled by temperature-dependent metabolic rates, r-TPCs are typically unimodal: increasing temperatures lead to rising metabolic rates and population growth until an 'optimal' temperature (T_{opt}) is reached (Fig 1a). Beyond T_{opt}, elevated metabolic costs slow down or impede growth (Amarasekare & Savage, 2012; Brown et al., 2004; Rebolledo et al., 2020; Sinclair et al., 2016), Fig 1a). While unimodality is the norm, r-TPC

shape often varies across species (Kellermann et al., 2019) due to differences in microbial traits

(Wieczynski et al., 2021) and genetic expression between closely related species (Jacob &

Legrand, 2021). Ultimately, however, differences in r-TPC shape are attributed to divergent

evolutionary trajectories across species and environmental conditions (Angilletta, 2009;

Kontopoulos et al., 2020; Malusare et al., 2023).

Despite these recent findings, predicting how r-TPCs might adapt to future warming

climates remains an unsolved but central challenge, as r-TPC adaptation underpins a species'

ability to cope with environmental change. Thermal adaptation hinges on the evolution of

intraspecific genetic variation through mutation (Kirkpatrick & Peischl, 2013) and selection

favoring genetic variants that perform better in novel environments (Barrett & Schluter, 2008;

Franks et al., 2007). Characterizing this intraspecific variation in r-TPCs is therefore central to

understanding and predicting how rising temperatures will influence microbial growth in novel

climates (Kling et al., 2023). Most intraspecific variation in microbial r-TPCs is likely coming

from plastic variation in r across temperatures —known as environmental variation (E)—as a

single clonal line or individual can have an entire TPC. However, additive genetic variation (G)

in r-TPCs—upon which selection acts (Frankham, 2005)—is also likely (Kling et al., 2023; Liu

et al., 2020; Singleton et al., 2021), and can ultimately decide how thermal adaptation occurs.

Last, selection may act on plasticity itself, whenever there are fitness consequences associated

with genetic variation in environmental responses (or G × E interactions, (Hoffmann & Sgrò,

2011)). In this case, plasticity can help drive adaptive evolution (Ghalambor et al., 2007; Kling et

al., 2023). Quantifying and characterizing the genetic variation in microbial r-TPCs, as well as

how that variation influences r-TPCs response to selection in novel environments, is thus

paramount.

In this study we ask: 1) What is the extent of heritable G and G × E variation in r-TPCs?

2) What are the evolutionary consequences of G in r-TPCs under warming, and does G influence

r-TPC shape (henceforth 'shape parameters', Fig 1b) in response to selection across

temperatures? Lastly, 3) What are the evolutionary consequences of G × E in r-TPCs under

warming? We address these questions in a globally important ciliate protist species—i.e.,

unicellular Eukaryotes that dominate oceanic biomass (Bar-On & Milo, 2019), hold twice the

biomass of the entire Animal Kingdom (Bar-On et al., 2018), rank third in terrestrial biomass

(Bar-On et al., 2018), and underpin global ecosystem functioning (Geisen et al., 2018; Hu et al.,

2021; Nguyen et al., 2020; Xiong et al., 2020).

RESULTS

r-TPC variation and heritability

Intraspecific variation in r-TPCs can be quantified using classic tools from quantitative

genetics where r-TPCs are interpreted as the reaction norm of *r* across temperatures (Fig 1c-e).

Under this quantitative genetics framework, purely plastic variation (E) should result in similar r-

TPCs across genotypes (Fig 1c), additive genetic variation (G) should result in parallel r-TPCs

among genotypes (Fig 1d), and G × E interactions should result in non-parallel r-TPCs among

genotypes (Fig 1e). To quantify these different sources of r-TPCs variation, we leveraged a

collection of 22 unique genotypes of the protist *Tetrahymena thermophila* from the Cornell

Tetrahymena Stock Center (see Methods, Appendix S1). We quantified r-TPCs using standard

population growth assays across seven temperatures (13, 19, 22, 25, 30, 32, and 38ºC) replicated

six times each (see Methods). These temperatures span below and above the incubation

temperature of 22ºC and the average summer temperature (~23ºC, (NOAA)) of the species'

native range (US Northeast, Zufall et al., 2013).

All r-TPCs showed strong unimodal temperature effects on r within genotypes (R^2=

0.784, Fig 2a). r-TPCs showed significant variability in shape across genotypes (F= 163.65, p ≤

0.001, D_f = 21, Generalized Eta-Squared (GES, =effect sizes) = 0.832, Fig 2a) and significant G

× E interactions (F=29.7, p ≤ 0.001, D_f = 122, GES = 0.840, Fig 2a). Environmental variation (E)

accounted for 71.7% of all observed variation in r-TPCs; genetic variation (G) explained 6.1% of

all variation; Gene-by-Environment interactions (G × E) explained 11.7% of all variation, and

10.5% was residual variation (Fig 2b). These patterns are in line with what is expected for life

history traits (Hoffmann & Sgrò, 2011). After accounting for experimental error in the form of

inter-treatment and replicate variability (see calculation of broad-sense heritability, Methods), r-

TPCs were strongly heritable ($H^2_{standard}$=0.76, H^2_{cullis}=0.95, H^2_{piepho}=0.91).

Consequences of G: selection and evolvability of r-TPC shape parameters

In the presence of heritable genetic variation (G), r-TPCs shape may evolve under

selection. To understand this phenomenon, we quantified four 'shape' parameters controlling the

rising portion of the TPC, i.e., the 'operational temperature range' (DeLong et al., 2017; Smith et

al., 2021), Fig 1b). We focus on the rising portion because temperatures within this range often

control ecological responses to warming for *T. thermophila* in its native range (Deutsch et al.,

2008; Schoolfield et al., 1981; Schulte et al., 2011). To do so, we fitted a Sharpe-Schoolfield

model (Schoolfield et al., 1981) on r-TPC data (Fig 1a, b, see Methods) and determined the

critical minimal temperature (CT_{min}, Fig 1b), the 'activation energy' (E_a, Fig 1b), the maximum

population growth (r_{peak}, Fig 1b), and the temperature of maximal growth (T_{opt}, Fig 1a, b, see

Methods). We then estimated the intraspecific variation within each parameter, and the form of

selection acting on them with and without taking genetic covariances into account (see Methods).

Genetic covariances among shape parameters can lead to joint parameter evolutionary responses

irrespective of direct selection acting on a given shape parameter, or even preclude parameter

evolution altogether (Hansen & Houle, 2008). To estimate the predicted evolutionary change of

each parameter under different temperature scenarios, we used a modified G-matrix approach

that jointly estimates the genetic variance-covariance among shape parameters and the predicted

parameter change ($\mathbf{\Delta z}$) across temperatures using a modified Price equation (see Methods,

Stinchcombe et al., 2014). This approach allowed us to estimate the selection gradient ($\mathbf{\beta}$) acting

directly on the shape parameters while controlling for effects of environmentally induced trait-

fitness covariances (see Methods), which provides a better estimate of selection on the traits of

interest than the standard multivariate breeder's equation (Lande, 1979; Lande & Arnold, 1983;

Stinchcombe et al., 2014).

Without accounting for genetic associations, selection operated differentially across

shape parameters and was temperature dependent: r_{peak} was under negative directional selection

at low temperatures (<20°C, Fig 3a, Appendix S1), under weakly positive or no selection at

intermediate temperatures (between 20 and 30°C, Fig 3a, Appendix S1), and under strong

positive directional selection in high temperatures. Parameter E_a followed a similar pattern (Fig

3b, Appendix S2). However, CT_{min} was found to be under negative selection at low/intermediate

temperatures (Fig 3c, Appendix S3) but no selection at high temperatures (Fig 3c, Appendix S3).

Lastly, T_{opt} was under no selection at low temperatures but under weak then strong stabilizing

selection at intermediate and high temperatures, respectively (Fig 3d, Appendix S4).

We found clear positive genetic covariances between shape parameters, specifically,

between r_{peak} and E_a, CT_{min} and E_a, and CT_{min} and T_{opt} (Fig 3e). Accounting for these genetic

associations, we again found differences in selection across shape parameters whose magnitude

and direction also shifted with temperature (Fig 3f), resulting in predicted temperature-dependent

shifts in parameter responses (Fig 3g). Specifically, our multivariate selection analysis suggested

that selection would favor higher r_{peak} and E_a at high temperatures (Fig 3f), however, predicted

responses in both cases should result in low values at low temperatures and high values at high

temperatures (Fig 3g), consistent with our univariate analysis. Similarities of the adaptive

landscapes for both shape parameters are mostly given by their strong positive genetic

correlation (Fig 3e), and their response at low temperatures are likely driven by genetic

correlations with CT_{min}. Indeed, we identified mostly negative selection on CT_{min} (except at high

temperatures, where no significant selection was found) and no selection for T_{opt}, also in

accordance with the univariate analysis (Fig 3f). Overall, the evolutionary responses followed

predicted trajectories from the estimates of selection closely (Fig 3f-g), suggesting little effect of

potential antagonistic selection on genetic constraints in r-TPC shape.

Consequences of G × E: differential sorting of standing genetic variation across temperatures

In the presence of G × E—where genotypes express different r-TPCs at different

temperatures—small differences in TPCs can lead to differential growth among genotypes across

temperatures, leading to clonal sorting and swift changes in population genetic makeup, i.e.,

evolution (Fig 4a, b). We tested this form of temperature-mediated rapid evolution induced by r-

TPC GxE through an experimental evolution assay: we competed two fluorescently tagged

genotypes (Fig 4c, AXS and CU4106, see Methods) with different r-TPCs (Fig 4d), hence

differing in relative fitness across temperatures (Fig 4d, inset). We observed significant

temperature-dependent clonal sorting (Fig 4e), which matched theoretical predictions from a

model of genetic evolution (Fig 4f). The model predicts genetic frequencies in a mixture

population using parameters taken from each genotype's r-TPC, as well as patterns of relative

fitness (Fig 4d) between the genotypes from r-TPC data (see Methods). Despite quantitative

discrepancies—notably at 19ºC where the model predicted a polymorphic population, but the

data indicated otherwise (Fig 4e, f)—it correctly predicted observed changes in genetic

frequencies across most temperatures, thus suggesting that temperature-dependent selection

acting on G × E r-TPC variation can drive adaptive evolution in population genetic makeup.

Interestingly, the results of this experiment were also consistent with our estimated

predicted responses to selection (cf Fig 3 and Fig 4): lower temperatures led to higher

frequencies for the CU4106 genotype, which shows lower E_a and CT_{min}, compared to AXS (Fig

4d), while at higher temperatures, there was selection in favor of AXS, so that the ensuing

population should have an average r-TPC with higher E_a (Fig 4d) as well. Therefore, such

temperature-dependent selection on r-TPC G × E variation could lead to rapid r-TPC evolution

through clonal sorting which could be predictable in nature (but see Nosil et al., 2018 for a

counterpoint). Naturally, our lab-based study by necessity simplifies the complexities of how

organisms contend with nature. Nonetheless, our experimental test of our mathematical

predictions provides an important proof of principle in predicting evolutionary responses to a

warming climate.

DISCUSSION

Our study reveals genetic variation in r-TPCs (Fig 2). While >70% of all TPC variation is

environmental (E), TPCs remain highly heritable ($H^2 > 0.7$, with $G + G \times E \sim 18\%$) once controlled

for inter-treatment and replicate variability, thus allowing selection to shape r-TPCs in new

climates. We also show that different r-TPC shape parameters (Fig 1b) are under different

selection regimes (Fig 3), and this selection is temperature-dependent, in some cases flipping

from negative to positive with temperature (Fig 3). These temperature-dependent selection

regimes should result in lower CT_{min}, r_{peak}, and E_a at cold temperatures while warmer

temperatures should favor higher r_{peak} and E_a with no discernible effect on T_{opt} (Fig 3). Lastly,

we show that $G \times E$ interactions are prevalent in these r-TPCs (Fig 1), which in turn can lead to

rapid—but predictable—shifts in population genetic makeup across temperatures (Fig 4), and

suggests that plasticity will help drive thermal adaptation (Ghalambor et al., 2007).

While the evolution of microbial TPCs in deep evolutionary time is likely the product of

adaptation to local habitats (Kontopoulos et al., 2020; Phillips et al., 2014), how r-TPCs will

adapt to rising temperatures is an open question. The 'Colder-is-Better' (CIB) hypothesis posits

that rising temperatures reduce growth, leading to the evolution of lower r_{peak} in a warming world

(J. Kingsolver & Huey, 2008). Conversely, the Warmer-Is-Better (WIB) hypothesis posits that

growth increases in warmer temperatures, leading to TPCs with higher r_{peak} (Frazier et al., 2006;

Pawar et al., 2015). Lastly, the Generalist-Specialist-Tradeoff (GST) hypothesis posits that

species either exhibit rapid growth within a narrow temperature range (i.e., temperature

specialists), or slower growth over a broader temperature range (i.e., temperature generalists) so

that higher r_{peak} should also result in higher CT_{min} and lower CT_{max} (Seebacher et al., 2015).

There is evidence supporting all three hypotheses (DeLong et al., 2018; Kontopoulos et al.,

2020), but most of it comes from inter-species comparisons that overlook intra-specific variation

and genetic associations between shape parameters, and therefore cannot readily make

predictions about r-TPC evolutionary trajectories for any given species. Without accounting for

genetic associations, our results would suggest, based on selection alone, support for WIB, with

clear directional selection for higher r_{peak} and E_a under warming climates (Fig 3a, b).

Accounting for genetic associations, however, showed support for multiple hypotheses

simultaneously, suggesting a more complex and nuanced evolutionary r-TPC response than

currently predicted by theory. Indeed, our analyses supported WIB, as warming should favor r-

TPCs with high r_{peak} and high E_a (Fig 3f, g), while countering GST, as there was no selection in

favor of higher CT_{min} with higher r_{peak} (Fig 3f, g). Lastly, T_{opt} was predicted to respond the least

to temperature (Fig 3g), as it showed only a small predicted decrease in colder temperatures

(likely though indirect selection)—which is arguably in support of CIB—except >30ºC, in

which case T_{opt} showed no clear pattern of evolutionary response to temperature —arguably

against WIB. Thus, like many studies, our ability to make predictions is context specific, and

identifying those contexts (e.g., temperature) remains an important avenue for future work. Our

results clearly emphasize the importance of intraspecific variation and genetic associations to

predict r-TPC evolution, and both support and generalize existing TPC evolution hypotheses.

We also uncovered an interesting mechanism of rapid thermal adaptation whenever there

is G × E in r-TPCs. While adaptation requires sufficient genetic variation upon which to act, the

adaptive sorting of standing genetic variation often leads to a reduction in genetic variation,

which can slow down the adaptive process or even impede future adaptation altogether (Pauls et

al., 2013). Our results show that, in the presence of extensive G × E variation in r-TPCs, rapid

shifts in genetic makeup are possible (Fig 4), resulting in predictable and rapid local adaptation

to novel conditions—i.e., phenotypic plasticity can be adaptive in novel climates. However, in

our system, rising temperatures led to an initial increase in additive genetic variance t (as both

genotypes become prevalent, Fig 4e, f), which should facilitate adaptation, then a reduction in

additive genetic variation (as genotype CU4106 becomes less prevalent, Fig 4e, f), which in turn

could impede adaptive evolution in the future. It is unclear why at 19ºC AXS is absent from the

experimental populations despite the model predicting that its presence, and otherwise good

agreement between model and data at other temperatures. This temperature-dependent effect on

population genetic makeup may have important but poorly understood consequences for the

conservation of genetically depauperate species under warming whenever large amounts of G $\times$

E are expected (Pauls et al., 2013).

Using a phylogenetic approach, a recent study found that TPC adaptation in deep time

likely occurred gradually across six different *Tetrahymena* species (Montagnes et al., 2022).

While seemingly in contrast with our finding that r-TPCs can evolve rapidly through

temperature-dependent selection on r-TPC G $\times$ E variation, we argue that the mechanism of r-

TPC evolution uncovered here is likely only at play as a form of rapid evolutionary response to

fast-changing environmental conditions, and may or may not result in longer term indefinite r-

TPC change. Indeed, selection often "erases its traces" (Haller & Hendry, 2014) with selection

most often being strongest in novel conditions and diminishing as populations adapt (Caruso et

al., 2017). Montagnes *et al.* (2021) also found poor support for WIB (which they referred to as

thermodynamic-constraint) in contrast with our findings, which show some level of support for

WIB. However, their study did uncover mixed support for CIB (which they call biochemical-

adaptation, or hotter is not better) and poor support for GST, as did we, suggesting that some of

the evolutionary responses uncovered here may be constrained by evolutionary history in deep

time.

Overall, TPCs control the fate of populations (Seebacher & Little, 2021; Sinclair et al.,

2016), ecological interactions (Bideault et al., 2019; Enquist et al., 2015), food web structure and

dynamics (Barbour & Gibert, 2021; Gibert et al., 2022), and ecosystem processes (Antiqueira et

al., 2018; Gibert et al., 2015). Yet, TPC evolution in a rapidly warming world remains a

conspicuous unknown. Here, we shed light on how r-TPC intraspecific variation drives

temperature-dependent evolution in a microbial r-TPC, as well as its consequence for rapid shifts

in population genetic makeup that either facilitates or precludes future thermal adaptation. In

doing so, we emphasize the importance of temperature in mediating rapid microbial evolutionary

change as we grapple with understanding and predicting how organisms in the planet may

respond to an increasingly warm world.

METHODS

Tetrahymena thermophila genotypes

We quantified intraspecific variation in r-TPCs in the protist *Tetrahymena thermophila*—

a freshwater species that is distributed across the eastern United States (Zufall et al., 2013) and

part of a genus of cosmopolitan distribution and importance (Lynn & Doerder, 2012). We used

22 unique *T. thermophila* genotypes: 19 from the Cornell Tetrahymena Stock Center and 3

strains from the Chalker lab (Washington University, Appendix S5). The genotypes vary in

geographic origin and have specific genetic differences (Appendix S5). Our goal in using these

genotypes was simply to have a source of genetic variation, not to select any particular genotypes

on the basis of their functional significance. Because most of these genotypes are derived from

laboratory cultures, this assemblage of genotypes likely harbors—collectively—less genetic

variation than would be found in natural isolates from across the species' distribution (Zufall et

al., 2013). Upon reception, we transferred the cultures from axenic Proteose Peptone growth

medium to Timothy Hay growth medium inoculated with a bacterial community from Duke

Forest Gate 9 pond/Wilbur pond (Lat 36.013914, Long -78.979720, fully described elsewhere

(Rocca et al., 2022) and a wheat kernel as a Carbon source (Altermatt et al., 2015). We

maintained these stock cultures in Percival (Perry, IA) AL-22 growth chambers under light (12hr

day-night cycle) and temperature controlled conditions (22°C) in 250mL borosilicate jars filled

with 150mL of liquid medium. Because *T. thermophila* is a natural bacterivore, these

experimental conditions are more realistic than purely axenic ones.

Quantifying TPCs and TPC shape parameters

We quantified the r-TPCs of all genotypes in 3cm diameter Petri dish microcosms with

3mL of growth medium through growth assays at seven temperatures (13, 19, 22, 25, 30, 32,

38°C), each replicated six times, totaling 1056 microcosms. We chose these temperatures

because they capture the range of temperatures at which *T. thermophila* is known to grow well

(Zufall et al., 2013). Microcosms were initialized by pipetting three individual cells from stock

cultures under a scope (Leica stereomicroscope model M205C) and allowing them to grow for

24hrs, after which we censused the microcosms through whole-population counts under the

scope, and calculated intrinsic growth rate, *r*, as log(final density/initial density)/time (Voronov,

2005; Wieczynski et al., 2021), with time = 1 day. We fitted a Sharpe-Schoolfield model

("nls.multstart" v1.3.0 package in R, (Padfield, 2023)) to obtain r-TPC shape parameters CT_{min}

(minimum temperature at which the population can grow), CT_{max} (maximum temperature at

which the population can grow), r_{peak} (maximum *growth rate*), T_{opt} (temperature at which r_{peak} is

achieved), E_a (thermal sensitivity of the rising portion of the TPC) and E_d (thermal sensitivity of

the declining portion of the r-TPC, Fig 1b). Only parameters E_a, r_{peak}, CT_{min} and T_{opt} could be

unequivocally estimated with our data, as fits were less well constrained at higher temperatures,

and so we focused all subsequent analyses on those (Fig 1b). These parameters also control the

rising portion of the r-TPC in the so-called 'operational temperature range' (DeLong et al., 2017;

Smith et al., 2021), which also is the temperature range within which ecological responses to

temperature are expected for organisms in their native geographic ranges.

Quantifying sources of phenotypic variation in r-TPCs and r-TPC heritability

To address whether and how r-TPCs may adapt to changing temperatures, we quantified

standing genetic variation in r-TPC shape and how much of this variation was heritable. To do

so, we estimated how much of the total observed r-TPC variation was explained by

environmental variation (E), genetic variation (G) or Gene-by-Environment interactions (G × E,

Fig 1c-e). The shape of r-TPCs are often assumed to be controlled by plastic physiological

responses as a single genotype can often express an entire r-TPC ((J. G. Kingsolver et al., 2004),

Fig 1c). We therefore expected to find a large amount of E. However, different genotypes may

still express different TPCs (Singleton et al 2021, Fig 1d), and those could do so differentially

across temperatures, leading to both G and G × E ((Ørsted et al., 2019), Fig 1e).

We estimated E, G and G × E using the function *gxeVarComps*() in R package

statgenGxE v1.0.5. To do so, the procedure fits two models: first, it fits a fixed effects linear

model with *r* as the response variable, and temperature, genotype, and the interaction between

temperature and genotype as predictors to calculate effect sizes, significance levels, and Best

Linear Unbiased Estimators (BLUEs, (Baksalary & Puntanen, 1990; Henderson, 1975)). BLUEs

are subsequently used to calculate r-TPC broad-sense heritability (see below). Second, it re-fits

the model with all terms as random effects to calculate the variance component of each term (i.e.,

G, E and G × E) and Best Linear Unbiased Predictors (BLUPs, (Henderson, 1975)), which are

also subsequently used in the calculation of heritability. Using function *H2cal*() in R package inti

v0.6.2 we calculated the broad-sense heritability (H^2) in three ways: 1) standard heritability,

where $H^2=G/P$, $P = G+(G×E/m)+(ResidVar/(m*r))$, *m* is the number of temperature treatments

and *r* the number of replicates (which accounts for inter-treatment and replicate variability , in

ways that G/P does not, (Baksalary & Puntanen, 1990; Henderson, 1975)), 2) *Cullis* heritability,

where $H^2= 1-vBLUE/2G$, and vBLUE is the mean variance of the difference of two BLUEs

(Cullis et al., 2006), and, 3) *Piepho* heritability, where $H^2= G/(G+vBLUP)$, and vBLUP is the

mean variance of a difference of two BLUPs (Piepho & Möhring, 2007).

Consequences of G: selection on r-TPC shape parameters and evolutionary potential

To understand how TPCs might evolve in different temperatures, we assessed: 1)

selection direction/form and magnitude on TPC shape parameters, 2) the impact of temperature

on such selection, and, 3) potential evolution of shape parameters under these selection regimes.

We measured selection in two different ways: one that neglects genetic correlations between

shape parameters but can be used to estimate non-linear selective effects (e.g., stabilizing

selection), and one that accounts for genetic correlations but neglects non-linear terms but

enables predictions of evolutionary potential. The first approach answered the first two questions

while the second approach complemented the answer to the second question and addressed the

third. In doing so, we are explicitly quantifying how selection on r-TPC shape parameters

changes across temperatures, or what MacColl (2011) called "eco-evo landscapes" (MacColl,

2011).

In the first approach, we quantified the relationship between parameter values and

absolute fitness—estimated here as, r—which is often assumed to be a suitable proxy for

absolute fitness (Lande, 1976, 1979). This is because r reflects the birth and death rates per

individual, two important fitness components (Lande, 1982; Partridge & Harvey, 1988). We

considered three temperature ranges: low (<20°C), medium (between 20°C and 30°C), and high

(>30°C). For each temperature range, we considered the relationship between r and the observed

parameter value across genotypes (i.e., the adaptive landscape) (Lande, 1976, 1979). A positive

relationship would be evidence of positive directional selection, a negative relationship, evidence

of negative directional selection, and the absence of a relationship would occur whenever there is

no directional selection. Stabilizing selection would result in a hump-shaped concave-down

relationship, with intermediary values having higher fitness than extreme values, and disruptive

selection with a concave-up relationship where extreme values have higher fitness (Lande &

Arnold, 1983). We analyzed these data using polynomial regression in R v4.3.1 with r as the

response variable, both linear and quadratic terms for the shape parameter and temperature as

explanatory variables with additive effects, an interaction between the linear and quadratic

effects of the shape parameter with temperature (See Appendices S1-S4). We doubled our

quadratic regression coefficients (Stinchcombe et al., 2008).

To predict possible r-TPC shape evolution across temperatures and evaluate the effects of

the genetic association among shape parameters in this evolutionary response, we used the

modified **G**-matrix approach proposed by Stinchcombe *et al.* (2014). This approach fuses the

multivariate breeder's (Lande, 1979) and Price's (Price, 1972) equations to estimate the direct

(i.e., acting on the focal trait) and indirect effects of selection (i.e., acting on another trait the

focal trait is genetically linked to) on an evolving population. The potential evolutionary

response of each r-TPC shape parameter to natural selection, $\mathbf{\Delta z}$, is derived from the covariance

of the parameters with a measure of fitness. In practical terms, we estimated the additive genetic

variance-covariance matrix, $\mathbf{G}$, of all shape parameters, and the mean-standardized r as a

measure of fitness, as done before (Lande 1976, 1972, 1989). We then created a $\mathbf{G_w}$-matrix,

which includes $\mathbf{G}$ and a vector of predicted trait change, $\mathbf{\Delta z}$, as the last column and row. $\mathbf{\Delta z}$ is

defined element-wise as $\Delta z_i = cov_a(w_i, z_i)$, with cov_a being the additive genetic covariance, w_i the

relative fitness of the i-th shape parameter, and z_i the i-th shape parameter. Because the Δz_i

calculated in this way is equivalent to the genetic selection differentials, we can then calculate

the selection gradient as $\mathbf{\beta} = \mathbf{G_w}^{-1} \mathbf{\Delta z}$. The sign of Δz_i indicates the direction of the response.

Alignment between Δz_i and β_i indicates direct responses to selection (Hansen and Houle 2008),

that is, shape parameters that directly respond to selection as imposed by temperature, while

misalignment would be indicative of indirect selection through correlated responses with other

shape parameters. This approach correctly estimates $\mathbf{\Delta z}$ even in the absence of information on all

pleiotropic effects in the system, while also estimating selection on each trait individually (Heath

& Stinchcombe, 2014; Stinchcombe et al., 2014). Furthermore, $\mathbf{\beta}$ estimated in this way controls

for the potential effects of environmentally induced trait-fitness covariances.

To calculate $\mathbf{G}$, for each temperature, we compiled all r-TPC shape parameters and r for

each genetic variant and calculated the between-genotype covariance matrix, $\mathbf{L}$. For lines derived

from the same population, $\mathbf{L}$ is proportional to the additive genetic covariance matrix of the

original population, or $\mathbf{G_r}$. As inbreeding tends to increase between-genotype genetic variation

due to drift, $\mathbf{L}$ is inflated in relation to $\mathbf{G}$ by a factor equal to two times the inbreeding

coefficient, F (Falconer & Mackay, 1996). Genetic correlation among traits were obtained by

estimating the correlation version of $\mathbf{G}$ as $\mathbf{G_{cor}} = \mathbf{SGS}$ where $\mathbf{S}$ is a diagonal matrix containing the

inverse of each trait's standard deviation. We estimated uncertainty using a Bayesian posterior

sample of covariance matrices as implemented in the evolqg v3.0 package (Melo et al. 2015).

This method derives the covariance matrices analytically from a multivariate normal likelihood

function and an inverse Wishart prior (Murphy 2012). We then took 1000 posterior samples for

$\mathbf{G_w}$ and derived statistics for $\mathbf{\Delta z}$, $\boldsymbol{\beta}$ and $\mathbf{G}$. To evaluate if $\mathbf{\Delta z}$, $\boldsymbol{\beta}$ and $\mathbf{G_{cor}}$ estimates were

statistically meaningful, we inspected the 95% maximum density interval to evaluate if the

posterior distribution significantly overlapped with the expected value due to lack of signal

(value=0).

Experimental test of consequences of G × E for population genetic makeup across temperatures

If r-TPCs show G × E in shape, slight differences in intrinsic growth rates can lead to

temperature-mediated clonal sorting (Fig 4a, b). We assessed how differences in r-TPCs across

genotypes could influence thermal adaptation through sorting of standing variance using the

coefficient of selection, *s*, which equals 1- w, where w is the relative fitness of each genotype

(measured as the quotient of the *r* of each focal genotype, Fig 4d). The coefficient of selection

measures the magnitude of selection acting against a given genotype relative to the focal

genotype; the larger the value of *s*, the stronger the selection against the given genotype.

Differences in *s* across genotypes and temperatures would indicate large potential for

temperature-mediated clonal sorting.

We then tested whether temperature-mediated clonal sorting occurs by setting up an

experimental evolution assay that competed two fluorescently marked strains that showed

significantly different in r-TPCs and relative fitness (AXS and CU4106, Fig 4d, Fig4d inset)

across six temperature treatments (19, 22, 25, 30, 32, 38ºC) each replicated seven times, with

additional single-strain controls per temperature. Microcosms were initialized at equal densities

(5 individuals per genotype) in 3cm diameter Petri dishes. We added $CdCl_2$ on days 2 and 7 to

induce fluorescence and counted on days 3 and 8. Fluorescently tagged genotypes may lose their

ability to fluoresce over time—however, they carry a Paromomycin resistance gene so that

fluorescing individuals can be selected for through Paromomycin exposure. We treated the

microcosms with 100µg/mL of Paromomycin prior to censusing, then used a Novocyte 2000R

flow cytometer (Agilent, Santa Clara, CA) to count individual cells and estimate relative

frequencies based on fluorescence (See Appendix S6 and Appendix S7). Because antibiotics can

influence the protists and the bacterial communities they feed on in multiple ways, we replicated

the entire experiment in Paromomycin-free conditions, but this did not qualitatively alter our

results (see Appendix Fig S8).

To confirm fluorescence of the two strains (Fig 4c), cells were mounted on glass slides

with mounting medium (Winey et al., 2012). Images were taken with a Leica Thunder Cell

Culture inverted microscope equipped with an HC PL APO 63X/1.40 N.A. oil-immersion

objective lens. Fluorescent signals were captured using a 510-nm excitation laser and a 535/15-

nm emission filter for YFP (expressed in genotype AXS), and a 395-nm excitation laser and a

Leica DFT51011 quad-band filter set for autofluorescence (exhibited by both genotypes). All

images were captured at a single Z plane using the same exposure settings; the resulting images

were processed in ImageJ (see Fig 4c and Appendix Figure S9).

Last, observed changes in genetic frequencies were compared to predicted frequencies by

a modified version of a classic model of genetic evolution in discrete time parameterized with

the r-TPC of each strain in the experimental evolution assay (Fig 4f). The model tracks the

frequency of each strain f_i, and assumes that their absolute fitness, W_i, is a function of their

reproductive output (r), following early work by Lande (Lande, 1976, 1979). The frequency of

each strain in the population is then determined by the classic recursive equation, $f_i(t + 1) =$

$f_i(t)W_i/\overline{W}$ [Eq 1.], where $\overline{W}$ is the average fitness in the population ($\sum_{i=1}^{n} f_i r_i$), such that $W_i/\overline{W}$ is

the relative fitness of the i-th strain. Alternatively, we can calculate the fitness of each strain

relative to that of a focal strain (we call it S_{focal}, Fig 4d inset). Then the relative fitness equation

becomes $W_i/\overline{W}_{S_{focal}}$. In our experiment, we only have two strains (AXS and CU4106), so either

strain can be the focal strain, but this choice does not alter the evolutionary dynamics. Our focal

strain was CU4016. We used each strain's r-TPC (Fig 4d) to calculate the relative fitness of each

strain as W_{AXS}/W_{CU4016} for AXS and W_{CU4016}/W_{CU4016} for CU4106 (Fig 4d, inset) –thus

replacing each strain's fitness, W_i, with their r-TPC (Lande 1976). We then replaced these

empirically parameterized measures of relative fitness in our recursive equation (Eq 1.) and

numerically solved the recursive equation over time to make testable predictions of how genetic

frequencies should change across temperatures. Despite this model not accounting for density,

frequency, and other forms of selection and ecological processes, is does a good job at

qualitatively reproducing the observed evolutionary dynamics (Fig 4e, f), therefore supporting

the hypothesis that observed shifts in genetic frequencies are due to GxE in r-TPCs and

temperature-mediated selection.

Fig. 1: **a.** General shape of the r-TPC. **b.** r-TPC shape parameters. Blue-colored shape
parameters indicate those measured in this study, namely, r_{peak}, E_a, CT_{min} and T_{opt}. **c.**
Environmental variation (E) in r-TPCs results from the expression of a different *r* across
temperatures. In the classic reaction norm approach, this would be observed as different *r* values
across temperatures, but equal across genotypes within temperatures (middle panel). **d.** Genetic
variation (G) in r-TPCs occurs whenever different genotypes express TPCs with different
heights, such that in the classic reaction norm approach observed *r* values vary across but in such
a way that the slope of the effect of the genotype on *r* is additive (middle panel). **e.** Gene-by-
Environment interactions (G × E) occur whenever TPCs vary in both heights and slopes across
genotypes, such that, in the classic reaction norm approach, *r* varies across genotypes and
temperatures in a multiplicative fashion.

Fig. 2: a. Observed r-TPCs for all 22 genotypes. Dots represent observed *r* values, bold lines represent Sharpe-Schoolfield model fits, and dash lines represent average TPCs across all experimental genotypes. **b.** All 22 TPCs are superimposed and dashed lines represent the average r-TPC. Inset: amount of variation due to residual, G × E, G, and E effects.

Fig. 3: a. Estimated adaptive landscape across temperatures (i.e., change in fitness with a change in the underlying shape parameter) for r_{peak}. Color indicates temperature (blue: 10–20ºC, yellow: 20–30ºC, red: 30–40ºC) **b.** As in **a**, but for E_a. **c.** As in **a**, but for CT_{min}. **d.** As in **a**, but for T_{opt}. **e.** Observed genetic associations between shape parameters. Each dot is a genotype color-coded as in Fig. 2. In gray, 95% confidence ellipses. ρ represents correlation coefficients. **f.** Estimated multivariate selection coefficient (**β,** 95% maximum density intervals) for all shape parameters across temperatures, or eco-evo landscapes (MacColl, 2011). **g.** Predicted evolutionary change (**Δz,** 95% maximum density intervals) for all shape parameters, across temperatures.

Fig. 4: **a.** G × E variation in r-TPCs leads to differential growth of each genotype across temperatures. **b.** Differential growth across temperatures leads to clonal sorting and rapid shifts in genetic frequencies across temperatures. **c.** First column: Differential Interference Contrast (DIC) microscopy for two genotypes of the protist *Tetrahymena thermophila* (CU4106 and AXS). Second column: fluorescence microscopy image overlayed on DIC. Only AXS fluoresces (green) due to the expression of Yellow Fluorescent Protein (YFP). Third column: as in the second column, but for autofluorescence (A. Flu, in pink), which both genotypes exhibit. Fourth column: Overlayed DIC, YFP and A.Flu images showing how the different strains fluoresce once all sources of fluorescence are accounted for. **d.** r-TPC for genotypes CU4106 and AXS. Inset: Measures of relative fitness for both CU4106 and AXS. This predicts an increase in AXS frequency relative to CU4106 at intermediate temperatures relative to low or high temperatures. **e.** Observed genetic frequencies across temperatures. **f.** Predicted genetic frequencies across temperatures.

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Online Supplementary Materials: Temperature-dependent selection shapes microbial thermal performance curves and population genetic makeup

Megan H. Liu[1], Ze-Yi Han[1], Yaning Yuan[1], Katrina DeWitt[1], Daniel J. Wieczynski[1], Kathryn M. Yammine[2], Andrea Yammine[1], Rebecca Zufall[3], Adam Siepielski[4], Douglas Chalker[5], Masayuki Onishi[1], Fabio A. Machado[6], Jean P. Gibert[1]

[1]Department of Biology, Duke University, Durham, NC, 27708, USA
[2]Department of Chemistry, Massachusetts Institute of Technology, Cambridge, MA, 02139, USA
[3]Department of Biology and Biochemistry, University of Houston, Houston, TX 77204, USA
[4]Department of Biological Sciences, University of Arkansas, Fayetteville, AR, USA
[5]Department of Biology, Washington University, Saint Louis, MI, USA
[6]Department of Integrative Biology, Oklahoma State University, Stillwater, OK, USA

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Appendix S1. r_peak Stats Table

	Estimate	SD	P-value
Linear term	1.81	0.16	**≤0.001**
Quadratic term	-0.28	0.12	0.24
Low Temp	-6.25	0.26	**≤0.001**
Med Temp	-3.20	0.26	**≤0.001**
Linear*LowTemp	-2.13	0.22	**≤0.001**
Linear*MedTemp	-1.27	0.22	**≤0.001**
Quadratic*LowTemp	0.14	0.17	0.70
Quadratic*MedTemp	-0.08	0.17	0.82

Appendix S2. E_a Stats Table

	Estimate	SD	P-value
Linear term	1.61	0.15	**≤0.001**
Quadratic term	-0.12	0.13	0.62
Low Temp	-6.07	0.27	**≤0.001**
Med Temp	-3.06	0.27	**≤0.001**
Linear:LowTemp	-2.03	0.21	**≤0.001**
Linear:MedTemp	-1.20	0.21	**≤0.001**
Quadratic:LowTemp	-0.22	0.18	0.53
Quadratic:MedTemp	-0.34	0.18	0.33

Appendix S3. CT_min Stats Table

	Estimate	SD	P-value
Linear term	0.44	0.18	**0.02**
Quadratic term	-0.12	0.13	0.66
Low Temp	-6.15	0.32	**≤0.001**
Med Temp	-3.21	0.32	**≤0.001**
Linear*LowTemp	-1.19	0.26	**≤0.001**
Linear*MedTemp	-0.88	0.26	**≤0.001**
Quadratic*LowTemp	-0.04	0.19	0.90
Quadratic*MedTemp	-0.06	0.19	0.89

Appendix S4. T_opt Stats Table

	Estimate	SD	P-value
Linear term	-0.32	0.20	0.11
Quadratic term	-2.10	0.21	**≤0.001**
Low Temp	-7.43	0.40	**≤0.001**
Med Temp	-3.95	0.40	**≤0.001**
Linear*LowTemp	0.01	0.28	0.98
Linear*MedTemp	-0.23	0.28	0.42
Quadratic*LowTemp	2.52	0.30	**≤0.001**
Quadratic*MedTemp	1.42	0.30	**0.02**

Appendix S5. Genotype Table

Name	Original Provenance	Provenance	Stock ID	Mutations
20395-1	Lake Warren in Alstead, NH, lat. 43 07.310, long. -72 17.840	Cornell Tetrahymena Stock Center	SD01557	NA
SB3539-I		Cornell Tetrahymena Stock Center	SD00660	*chx1[C3]-1/chx1[C3]-1 (CHX1[C3]; cy-s, I)*, C3 strain - functional heterokaryon carrying cycloheximide resistance in micronucleus.
B*VII		Cornell Tetrahymena Stock Center	SD00023	B strain star line. Lacks a genetically functional micronucleus.
B2192 III	Frankel lab (Leslie Jenkins)	Cornell Tetrahymena Stock Center	SD01754	Derived from a cross of B2086 II x B2086 VIa. Isogenic with B2192 IVB
CU428.2		Cornell Tetrahymena Stock Center	SD00178	*mpr1-1/mpr1-1* (*MPR1*; mp-s, VII)
DMCK72H	Chalker Lab, Washington University in St. Louis			
19877	SG69-4 in Guys Mills, PA (lat. 41 38.023, long. -79 53.514, elevation 1660 ft)	Cornell Tetrahymena Stock Center	SD01555	Cech's self-splicing intron is present. Cytochrome oxidase I haplotype = WPA1
19617-1	FS136E in PA (latitude 41.46, longitude -78.88)	Cornell Tetrahymena Stock Center	SD03089	

IMB6 (GFP)	Chalker Lab, Washington University in St. Louis			
A*V		Cornell Tetrahymena Stock Center	SD00014	Lacks a genetically functional micronucleus.
20441-1	Gregg Lake in Antrim, NH (lat. 43 02.605, long. -71 59.383	Cornell Tetrahymena Stock Center	SD01560	
CU427-4		Cornell Tetrahymena Stock Center	SD00715	chx1-1/chx1-1 (CHX1; cy-s, VI)
SB1518		Cornell Tetrahymena Stock Center	SD01537	*gal1-1/gal1-1; tyr-14/tyr-14*
IA388		Cornell Tetrahymena Stock Center	SD01454	*elo1-1/elo1-1* (*elo1*; II)
21157-1	FS343S in PA (latitude 41.45, longitude -78.88)	Cornell Tetrahymena Stock Center	SD03114	
CU438-1		Cornell Tetrahymena Stock Center	SD00189	*pmr1-1/pmr1-1*
CU304		Cornell Tetrahymena Stock Center	SD00051	CHX1/CHX1; chx2-1/chx2-1; mpr1-1/mpr1-1
CU4106		Cornell Tetrahymena Stock Center	SD01010	*mpr1-1/mpr1-1*
AXS	Chalker Lab, Washington University in St. Louis			
C*III		Cornell Tetrahymena Stock Center	SD00024	C strain star line. Lacks a genetically functional micronucleus.

Appendix S6. Flow Cytometry Frequency Calculations

Each microcosm was censused with a Novocyte 2000R flow cytometer and analyzed using NovoExpress software v15.0. The flow cytometer detects particles (e.g., cells, debris, bacteria) based on how they scatter light and fluoresce [1]. Light scattering properties can be used to quantify cell size (FSC-H), and we detected fluorescence in the Phycoerythrin (PE-H, yellow) and Fluorescein isothiocyanate (FITC-H) channels. We gated the data in NovoExpress to select for the largest particles which in our case were all *Tetrahymena thermophila* cells. The data are plotted in Fig 8.1. We used a PE-H versus FITC-H plot to parse the different fluorescent signals between the two experimental strains: CU4106 (which autofluoresces exclusively) and AXS (which autofluoresces and expresses Yellow Fluorescent Protein, or YFP). Control microcosms, which contained exclusively one of either strain for each temperature treatment, were used to determine the exact expected fluorescence range for each individual cell. An "AutoF" and "YFP" gate were created based on these controls (Figure 1). These control gating filters were then applied over each experimental microcosm, allowing us to identify cell strain based on their fluorescence pattern.

Appendix S6 Figure 1. This plot shows an AXS control sample, where 65% of AXS individuals fluoresced in the autofluorescence gating range.

We used two CU4106 control microcosms and nine AXS control microcosms per temperature. Additional AXS controls were necessary to increase detection precision while gating. Non-YFP tagged cells (CU-4106) fluoresced more weakly in the FITC-H channel than YFP-tagged cells (AXS) and generally fluoresced more strongly in the PE-H channel than non-YPF tagged cells fluorescence. CU4106 controls were detected exclusively in the autofluorescent gate. However, AXS controls were detected in both the autofluorescent ("AutoF") and YFP gates, meaning that we could expect AXS cells to show up in the AutoF gate under experimental conditions, thus making it harder to parse YFP-tagged from non-YFP tagged cells.

To resolve that, we thus used the control microcosms to adjust the relative frequencies of each strain in each experimental microcosm, calculated as follows. For each single-strain control microcosm, we calculated the proportion of cells detected in each of the two gates across temperatures (Table X). We then used the proportion of AXS cells across control replicates to adjust the observed number of AXS cells (i.e., cells showing up on the YFP gate) to ensure our estimate was as accurate as possible. We then subtracted this adjusted count from the total number of individuals in each microcosm to generate our final adjusted experimental count of CU4106 individuals. See Appendix Table 10 below for pertinent data.

Appendix S7. Table of Frequency Calculations

	temp	rep	AB +/-	proportion	CU4106 count	adjusted CU4106 count	AXS count	adjusted AXS count	total
1	19	Sample1	NoAB	0.51421189	5	5.000000	0	0.000000	5
2	19	Sample2	AB	0.01271186	71	71.000000	0	0.000000	71
3	19	Sample2	NoAB	0.51421189	20	18.110553	2	3.889447	22
4	19	Sample3	NoAB	0.51421189	37	34.165829	3	5.834171	40
5	19	Sample4	AB	0.01271186	53	53.000000	0	0.000000	53
6	19	Sample4	NoAB	0.51421189	17	17.000000	0	0.000000	17
7	19	Sample6	AB	0.01271186	49	49.000000	0	0.000000	49
8	19	Sample6	NoAB	0.51421189	9	8.055276	1	1.944724	10
9	19	Sample7	AB	0.01271186	76	76.000000	0	0.000000	76
10	19	Sample7	NoAB	0.51421189	32	32.000000	0	0.000000	32
11	22	Sample1	AB	0.58564815	73	72.292490	1	1.707510	74
12	22	Sample1	NoAB	0.40142857	105	99.035587	4	9.964413	109
13	22	Sample2	AB	0.58564815	77	74.877470	3	5.122530	80
14	22	Sample2	NoAB	0.40142857	34	11.633452	15	37.366548	49
15	22	Sample3	AB	0.58564815	12	9.877470	3	5.122530	15
16	22	Sample3	NoAB	0.40142857	37	32.526690	3	7.473310	40
17	22	Sample4	AB	0.58564815	22	2.897233	27	46.102767	49
18	22	Sample6	AB	0.58564815	13	12.292490	1	1.707510	14
19	22	Sample6	NoAB	0.40142857	13	11.508897	1	2.491103	14
20	22	Sample7	AB	0.58564815	7	6.292490	1	1.707510	8
21	22	Sample7	NoAB	0.40142857	21	21.000000	0	0.000000	21
22	25	Sample1	AB	0.27131609	162	33.084581	48	176.915419	210
23	25	Sample1	NoAB	0.55903491	254	129.370065	158	282.629935	412
24	25	Sample2	AB	0.27131609	358	113.597851	91	335.402149	449

25	25	Sample2	NoAB	0.55903491	179	73.301194	134	239.698806	313
26	25	Sample3	AB	0.27131609	309	169.341629	52	191.658371	361
27	25	Sample3	NoAB	0.55903491	192	104.443526	111	198.556474	303
28	25	Sample4	AB	0.27131609	246	165.427863	30	110.572137	276
29	25	Sample4	NoAB	0.55903491	261	212.094582	62	110.905418	323
30	25	Sample5	AB	0.27131609	263	220.028194	16	58.971806	279
31	25	Sample5	NoAB	0.55903491	180	169.745638	13	23.254362	193
32	25	Sample6	AB	0.27131609	597	438.541464	59	217.458536	656
33	25	Sample6	NoAB	0.55903491	193	127.529844	83	148.470156	276
34	25	Sample7	AB	0.27131609	291	145.970153	54	199.029847	345
35	25	Sample7	NoAB	0.55903491	184	137.460974	59	105.539026	243
36	30	Sample1	AB	0.13819840	1445	902.470286	87	629.529714	1532
37	30	Sample1	NoAB	0.44750154	263	235.838158	22	49.161842	285
38	30	Sample2	AB	0.13819840	732	588.572604	23	166.427396	755
39	30	Sample2	NoAB	0.44750154	244	202.022608	34	75.977392	278
40	30	Sample3	AB	0.13819840	1566	1191.841576	60	434.158424	1626
41	30	Sample3	NoAB	0.44750154	460	399.503170	49	109.496830	509
42	30	Sample4	AB	0.13819840	848	93.447179	121	875.552821	969
43	30	Sample4	NoAB	0.44750154	341	249.637440	74	165.362560	415
44	30	Sample5	AB	0.13819840	328	128.448841	32	231.551159	360
45	30	Sample5	NoAB	0.44750154	372	307.799282	52	116.200718	424
46	30	Sample6	NoAB	0.44750154	354	287.330024	54	120.669976	408
47	30	Sample7	AB	0.13819840	1908	1459.009892	72	520.990108	1980
48	30	Sample7	NoAB	0.44750154	233	209.542045	19	42.457955	252
49	32	Sample1	AB	0.21434368	536	169.459535	100	466.540465	636
50	32	Sample1	NoAB	0.28314509	161	90.110769	28	98.889231	189
51	32	Sample2	AB	0.21434368	580	492.030288	24	111.969712	604

52	32	Sample2	NoAB	0.28314509	215	171.960110	17	60.039890	232
53	32	Sample3	AB	0.21434368	578	515.688121	17	79.311879	595
54	32	Sample3	NoAB	0.28314509	252	219.087143	13	45.912857	265
55	32	Sample4	AB	0.21434368	2101	1873.744911	62	289.255089	2163
56	32	Sample4	NoAB	0.28314509	121	77.960110	17	60.039890	138
57	32	Sample5	AB	0.21434368	550	201.786558	95	443.213442	645
58	32	Sample5	NoAB	0.28314509	109	60.896593	19	67.103407	128
59	32	Sample6	AB	0.21434368	473	418.018930	15	69.981070	488
60	32	Sample6	NoAB	0.28314509	259	183.047253	30	105.952747	289
61	32	Sample7	AB	0.21434368	2796	1890.645050	247	1152.354950	3043
62	32	Sample7	NoAB	0.28314509	201	157.960110	17	60.039890	218
63	38	Sample2	AB	0.07118688	1135	273.862817	66	927.137183	1201
64	38	Sample2	NoAB	0.07573333	795	380.056343	34	448.943657	829
65	38	Sample3	AB	0.07118688	1051	385.575813	51	716.424187	1102
66	38	Sample3	NoAB	0.07573333	1912	264.429596	135	1782.570404	2047
67	38	Sample4	AB	0.07118688	5506	3209.634179	176	2472.365821	5682
68	38	Sample4	NoAB	0.07573333	1146	718.852117	35	462.147883	1181
69	38	Sample5	AB	0.07118688	537	484.809868	4	56.190132	541
70	38	Sample5	NoAB	0.07573333	710	295.056343	34	448.943657	744
71	38	Sample6	AB	0.07118688	1323	1179.477136	11	154.522864	1334
72	38	Sample6	NoAB	0.07573333	1307	1111.732397	16	211.267603	1323
73	38	Sample7	AB	0.07118688	1160	886.001805	21	294.998195	1181
74	38	Sample7	NoAB	0.07573333	959	324.380289	52	686.619711	1011

Appendix S8. Alternative Experimental Conditions With and Without Antibiotics

Fig Appendix S8: a) r-TPC for genotypes CU4106 and AXS. Inset: Measures of relative fitness for both CU4106 and AXS. This predicts an increase in AXS frequency relative to CU4106 at intermediate temperatures relative to low or high temperatures. b) Predicted genetic frequencies across temperatures. c) Observed genetic frequencies across temperatures with Paromomycin. d) As in c) but without Paromomycin.

Appendix S9. Differential Fluorescence Across Strains

1st column: Differential Interference Contrast (DIC) microscopy for CU4016 and AXS protist Tetrahymena thermophila (CU4106 and AXS). Subsequent columns display raw fluorescence microscopy images. Photos are unimposed and uncorrected for relative fluorescence levels.

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