**Complex genetic architecture underlying the plasticity of tobacco leaf width provides insight into across-environment genomic prediction**

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**Abstract**

Phenotypic plasticity is the property of a given genotype to produce multiple phenotypes in response to changing environmental conditions. Understanding the genetic basis of phenotypic plasticity and establishing a predictive model is highly relevant for future agriculture under changing climate. Here, we investigated the genetic basis of leaf width plasticity in a tobacco MAGIC population across four different environments. Environmental changes not only resulted in differences in leaf width mean but also in leaf width plasticity. A total of 45 QTL were identified, including 14 QTL for leaf width mean and 43 for leaf width plasticity, with 12 overlap. Changes in the environment affected the magnitude of several QTL, thereby influencing phenotypic plasticity. We identified a QTL, qLW14, associated with leaf width plasticity and leaf width mean at Zhucheng, but had no significant impact at Guiyang, indicating that changes in the environment contributed to variations in leaf width plasticity. By integrating genetic diversity, environmental variation, and their interactions into a unified model, we were able to build a model for cross-environment predictions, and improved prediction accuracy by 7.2%. Overall, this study reveals complex genetic basis involving multiple alleles, and genotype interactions underlying variations of leaf width mean and plasticity. These findings contribute to assessing the role of plasticity in responding to climate or other environmental changes.

## Introduction

Phenotypic plasticity refers to the phenomenon where individuals of the same genotype exhibit different phenotypes under different environmental conditions (Bradshaw, 1965; Pigliucci, 2005), or the ability of organisms to produce different phenotypes in response to varying environmental factors (Kliman, 2016; Nicotra et al., 2010). For organisms, phenotypic plasticity is a trait and possibly an important strategy to quickly adjust phenotypes according to environments, which potentially have effects on evolution (Bradshaw, 1965; Kusmec, de Leon, & Schnable, 2018; Sommer, 2020). With rapid climate changes, as sessile organisms, plants have evolved unique mechanisms to respond to environmental changes, including physiological, metabolic, growth, and developmental adjustments. Revealing the genetic basis of plasticity is crucial for enhancing our understanding of genetic variation of these complex traits in heterogeneous environments (Gillespie & Turelli, 1989; Ji et al., 2023; Jin et al., 2023; Kelly, 2019).

Tobacco (*Nicotiana tabacum L.*) is a model organism in plant genetics and synthetic biology. Tobacco leaf exhibits rich diversity and plasticity in morphology, and metabolic traits (Honarnejad & Shoai-Deylami, 2004; Legg & Collins, 1975; White, Pandeya, & Dirks, 1979), making it an ideal model organism for botanical and plant synthetic biology research. Therefore, genetic regulation of tobacco leaf morphological traits has attracted many attention during the past two decades. With the development of sequencing technology, the cost of sequencing has gradually decreased, the development of molecular markers has increased (Edwards et al., 2017; Khan et al., 2019; C. Wang et al., 2022), and genome-wide association analysis have been widely used in leaf morphological development, and many genes related to leaf number, width and length have been identified seat (Ahmed et al., 2019; Liu et al., 2022; Safdar et al., 2020; Zuo et al., 2022). Liu et al performed GWAS with a multiparent advanced generation inter-cross (MAGIC) population which consisting of 600 individuals and locating a tobacco plant height-related QTL at 140 cM on chromosome 13 (Liu et al.,

2022). However, little is known about the plasticity of leaf morphological development. While these studies provided insights on the genetic basis of leaf morphological traits, little investigation was performed on the genetic basis of leaf morphological traits plasticity.

In this study, we utilized a MAGIC population of 600 lines derived from eight parents to dissect the genetic basis of leaf width mean and leaf width plasticity across four environments. A total of 45 QTL associated with leaf width mean and plasticity were detected. We found that the dynamics change in QTL effects across environments was associated variation of leaf width plasticity, indicating that QTL-by-environment interactions may drive leaf width plasticity variation. Through integrating genetic diversity, environmental variation, and their interactions, we developed a model for site-specific predictions and improved accuracy considerably. Overall, this study reveals complex genetic basis involving multiple alleles, and genotype interactions underlying variations of leaf width mean and plasticity. These findings contribute to assessing the role of plasticity in responding to climate or other environmental changes.

**Results**

**Tobacco leaf width mean variation within each of the four environments**

We explored the genetic basis of leaf width mean and plasticity using 600 tobacco lines derived from a multi-parent advanced generation intercross （MAGIC）with genotypes using the 430K tobacco SNP array with 432,362 markers and leaf width (LW) measured under 4 different environments in Guiyang, Hunan (25°27′N 112°13′E), and Zhucheng, Shandong (35°59′N 119°24′E) in 2019 and 2020. In this study, the studied environment was abbreviated as: Hunan Guiyang in 2019 (GY19), Hunan Guiyang in 2020 (GY20), Shandong Zhucheng in 2019 (ZC19), Shandong Zhucheng in 2020 (ZC20). The leaf width mean and plasticity were abbreviated as LWm and LWp, respectively. LW displayed continuous variation, being the shortest at 13.5 cm and the longest at 37.5 cm. Changes in environment altered the mean and phenotypic variance of LWm (Vp; Fig

107 1B) across 4 environments. There was widespread reaction norm for LWm. For

108 example, when the environment changed from ZC19 to GY19, the leaf width at

109 Guiyang (LWm = 27.8cm) showed drastic difference in leaf width at Zhucheng (LWm

110 = 21.8cm), leading to variation in LWp (Fig. 1B). This result indicates that

111 environmental changes have an impact on both the LWm and LWp.

**Fig. 1A)** Geographical illustration of two Filed plantations, Zhucheng and Guiyang,
where the MAGIC population was phenotyped. **B)** Boxplot of LWm measured at four
environments.

## Environmental perturbation buffers/releases additive variation for LWm by altering the underlying genetic architecture

There was a moderate variation (0.13-0.61; Fig 2A) in the estimated narrow-sense

heritability ($h^2$) across different environmental conditions. Given that phenotypic

variation ($V_p$) varied significantly among the four environments, the amount of additive

genetic variation ($V_a$), estimated as $V_p$ times $h^2$, changed considerably from 4.1 to 11.2

(Fig 2B). These fluctuations in $V_a$ were likely caused by either altered genetic effects

from one or multiple QTL or turning on/off of a completely new set of QTL, indicating

the genetic architecture of LWm was dynamic across the four environments. This

hypothesis was supported by the variation of pairwise genetic correlation from

multivariate mixed model analysis (Fig 2C). The estimated genetic correlation between

environments was as low as 0.4, suggesting considerable changes in the polygenic

scores/breeding values across individuals. Collectively, these results illustrated that the

genetic variance of LWm changed across these environments due to alterations in

genetic architecture. In the following sections, we will quantify change and dissect the

genetic basis underlying LWm variation across environments.

**Fig.2 A)** Estimated narrow-sense heritability ($h^2$) for LWm at each environment. Each
dot represents a point estimate of $h^2$ indicated on the y-axis and the corresponding
environment was labelled in the x-axis. Standard errors were plotted as error bars. **B)**
Estimated additive variance at each environment. **C)** Upper triangular: Pairwise
Spearman rank correlation among LWm measured under four environments. **C)** Lower
triangular: Pairwise genetic correlation among LWm measured under four
environments.

**Polygenic architecture underlying tobacco LWm variation and LWp variation**

There was a widespread reaction norm for LWm. At population level, the median of

LWm in Zhucheng is significantly smaller than that in Guiyang (Fig. 1B), resulting in

variation in LWp. We, therefore, quantified LWp using two types of metrics. The first

type involved two measures including Finlay–Wilkinson Regression (FWR), and across

environment variance in rank transformed LWm (VarR) quantifying the overall

plasticity (OP) of LW across all environments. The second metric involved assessing

pairwise LWm differences for all possible pairs of environments, quantifying specific

changes across paired environments and resulting in a total of 6 specific plasticity (SP)

measures. Almost all the plasticity measurements were continually distributed with

intermediate to high narrow sense heritability (0.02-0.61).

A total of 45 QTL were identified for LWm and LWp using genome-wide association

analysis (GWA) (Fig. 3A), including 14 QTL for LWm, 9 QTL for leaf width BLUP

and 43 QTL for LWp with variable degrees of overlaps, indicating a shared genetic

basis between LWm, LWp and BLUP (Fig. 3C).

By contrasting genetic effects of QTL across sites, two types of QTL, whose effects

changed in magnitude, were associated with LWp For example, qLW1-3 showed a

significant effect for $LW_{ZC19}$ ($3.6 \pm 0.3$cm; $P = 4.8 \times 10^{-16}$; Fig. 3D), $LW_{ZC20}$ ($2.7 \pm 0.4$cm;

$P = 4.4 \times 10^{-9}$) as well as on the specific plasticity measures $LW_{GY19-ZC20}$ ($P = 5.9 \times 10^{-8}$

; Fig. 3D) and $LW_{ZC19-GY19}$ ($P = 2.2 \times 10^{-14}$; Fig. 3D), indicating changes in the

magnitude of genetic effects contributed to the variation of LW plasticity. In contrast,

qLW9-1 (Fig. 3E), was exclusively detected for several LWp but not for any of the

LWm and $LW_{BLUP}$ measurements. Altogether, these results suggested that changes in

magnitude of genetic effects across sites caused variation in plasticity. The changing

genetic effects highlighted the role of QTL by environment interaction in the variation

of LWm and LWp.

**Fig. 3 A)** Manhattan plots overlaying genome-wide association analysis results for the
mean, plasticity measurements and BLUP for LW. The black horizontal dashed line

indicated the Bonferroni-corrected genome-wide significance threshold derived as 0.05/Me (Me is the effective number of independent markers; Methods) **B)** QTL associated with the LWm and LWp measured at four environments and the LW$_{BLUP}$ (y-axis). Each dot represents a SNP, and the size of the dot is proportional to its -Log10 p value, as indicated in the legend on the right. Loci with a p value above the genome-wide significance threshold are colored in blue. **C)** Venn diagram illustrating the overlap of QTL detected for the 3 types of LW measurements. **D) and E)** Genotype-to-phenotype maps for LWm and LW$_{SP}$ at Chr 1: 38450718bp, Chr 9: 55539734bp.

**qLW14 associated with LWm and LWp measurements**

We detected a QTL，qLW14 located at chromosome 14: 116-123 Mb, associated with LWm at Zhucheng (LW$_{ZC19}$, LW$_{ZC20}$) and multiple LWp measurements (SP-LW$_{GY19-ZC20}$, SP-LW$_{ZC19-GY19}$, SP-LW$_{ZC19-GY20}$, SP-LW$_{ZC20-GY20}$, Fig. 4A). Among the eight parental lines, haplotype analysis revealed two main haplotypes, with one parent *Vam* harbouring a deletion of approximately 3Mb at the qLW14 (121-124 Mb), while no deletions were observed in the remaining seven parents (Fig. 4B). Linkage disequilibrium analysis revealed a high level of linkage disequilibrium at this region, which overlapped with the deleted fragment in *Vam* (Fig. 4C). Phenotypic analysis showed that lines carrying this deletion exhibited a 20% increase in leaf width at Zhucheng in comparison to that for lines without the deletion (Fig. 4D). In contrast, lines carrying this deletion did not significantly increase leaf width at Guiyang. Overall these results highlighted a general contribution from qLW14 by environment interaction to variation of LWp (Fig. 4E).

**Fig. 4A)** Manhattan plot for plasticity measurement $LW_{ZC19-GY20}$. The black horizontal dashed line represents the Bonferroni-corrected genome-wide significance threshold of 0.05/Me (Me is the effective number of independent markers; Methods). **B)** Haplotype analysis from re-sequenced eight parents for MAGIC population (*Vam*, *Beinhart1000-1*, *Florida301*, *Samsun*, *Basma*, HD: *honghuadajianyuan*, TP: *tangpeng*, XHQ: *xiaohuaqing*) revealed a 3Mb deletion in *Vam*. **C)** LD heatmap showing the region surrounding a strong peak (chromosome 14: 116 Mb–124 Mb). Difference in leaf width of the two haplotype lines qLW14 in the year of 2019 and 2020 at Zhucheng, Shandong **D)** and Guiyang, Hunan **E)**.

**Accounting for dynamics in genetic architecture improved LWm prediction across environment**

Here, we evaluated the potential of integrating genetic diversity, environmental variation, and their interaction in complex trait prediction by jointly modelling

genotype, environment, and their interaction (referred to as GEAI model, Methods).

Two cross-validation schemes were considered. First, we explored the ability to predict

untested lines (Newly developed lines) at the four environments using five-fold cross-

validation. In this scenario, 80% of the lines with phenotype measured from all the four

environments were randomly selected as training set and the remaining 20% lines,

whose phenotypes were masked as NA, were used as test set (Fig. 5A, 5B). Within site

Pearson correlation (r) between predicted and measured phenotype were used to

evaluate the accuracy of predicted phenotype at each of the four environments.

Compared with the GBLUP, with a universal prediction for all sites, our model not only

provided site-specific predictions but also increased prediction accuracy. The averaged

prediction accuracy for LW increased by 8.7%.

In the second scenario, we performed prediction for novel sites without any tested lines

by utilizing phenotypic data from three out of the four environments as the training set,

while the phenotype data from the remaining environment served as the test set.

Compared to the GBLUP model, our GEAI model exhibited an average improvement

of 7.2% in prediction accuracy. This substantial enhancement underscores the

effectiveness of our GEAI model in capturing and leveraging genotype-environment

interactions to enhance prediction accuracy across different environmental contexts.

Overall, our study highlighted the potential of intergrading QTL by environment

interaction in understanding complex traits variation and predictions.

**Fig 5 A)** Use 5-fold cross-validation to predict the phenotypes of the remaining 20% lines in each environment. **B)** Use the phenotypes of all strains in three environments in the four environments as training set for LW, and the remaining environments as a test set. **C)** Schematic illustration of the prediction design for **A)**. **D)** Schematic illustration of the prediction design for **B)**, with $LW_{ZC20}$, $LW_{GY20}$, and $LW_{GY19}$ as training set and $LW_{ZC19}$ as the tested set.

## Discussion

Consistent with studies in maize and rice (Guo et al., 2020; Ji et al., 2023), we found a considerable proportion of QTL (33%) associated with LWp were shared with those associated with LWm, suggesting a shared genetic basis between LWm and LWp. However, due to a lack of environmental factor data, we were not able to definitively determine which environmental factors influence the plasticity of leaf width. Further investigation is needed to explore the specific roles of environmental factors in the variation of LWm and LWp.

Haplotype analysis on qLW14 revealed the presence of multiple duplicate MADS-box

transcription factors. The candidate gene MADS-box has been confirmed to be involved

in leaf development in *Arabidopsis* and *tomato*. In *Arabidopsis*, a MADS-box member

*FRUITFULL(FUL)* affects leaf development by affecting vascular bundle development

(Bemer et al., 2017); in *tomato*, a MADS-box member *SlMBP21* affects leaf

development by inhibiting genes involved in determining seed development (Y.-s.

WANG et al., 2021). Although the candidate interval of qLW14 is still large in our

results, with the advancement of multi-gene editing technologies such as

CRISPR/LbCpf1, it is worthwhile to conduct fine mapping and functional verification

of potential candidate genes for this QTL.

We constructed a genomic prediction model that considers genotype-environment

interactions. When using three of the environments as training set to predict phenotypes

in the remaining environment, this model outperformed Genomic Best Linear Unbiased

Prediction (GBLUP). This superiority of the GEAI model is attributed to its inclusion

of interactions between detectable QTL and environmental conditions. Due to a lack of

various environmental factor data, we cannot confirm the interactions between QTL

and specific environmental factors, which requires further exploration. Despite its

imperfections, we recommend using the GEAI model for cross-environment trait

prediction, as it can further enhance prediction accuracy. By establishing a prediction

model that responds to environmental changes, we can predict leaf width phenotypes

under different environments, laying the foundation for improving the accuracy of

genome-wide prediction of agronomic traits under multiple environments.

**Methods**

***Phenotype and Genotype***

A total of eight different tobacco varieties were utilized as parents to develop the

MAGIC population. These included *Beinhart1000-1* and *Florida301*, which is cigar

tobacco varieties; *Basma* and *Samsun*, belonging to oriental tobacco; *Xiaohuaqing* (XHQ) and *Tangpeng* (TP) from sun-cured tobacco; *Vam*, a variety of Virginia tobacco; and finally, the flue-cured tobacco variety *Honghuadajinyuan* (HD). Initially, these eight parents were crossed in pairs, resulting in four two-way crosses. Subsequently, the offspring from these crosses were subjected to incomplete diallel crossing to generate four-way crosses. Further crossing led to the production of eight-way crosses. These eight-way crosses were then subjected to random mating, yielding a total of 320 eight-way crossbreds. From these 320 eight-way crossbreds, three individual plants were selected per crossbred, and then propagated for five generations through the single seed descent (SSD) method, resulting in the establishment of over 800 eight-way tobacco MAGIC inbred lines.

In 2019 and 2020, the MAGIC population, along with the eight parents, were planted separately in Shandong Zhucheng, China, and Guiyang, Hunan, China, respectively. Randomized complete block design was employed with two replicates. Replicate of each line comprised two rows of 10-plants each, with row length of 10 m, row spacing of 1.2 m, and plant spacing of 0.5 m.

During each experiment, five randomly selected plants were selected for phenotypic evaluation, and the average value was analyzed. Leaf width was defined by measuring the width of the blade perpendicular to the main vein for the middle leaf .

Genotyping was conducted using a tobacco SNP chip with 432,362 markers from the Zhengzhou Tobacco Research Institute of the China Tobacco Corporation. Detailed information regarding genotyping and SNP calling procedures can be found in a previous study (Yuan et al., 2022).

***Calculation of kinship heritability and genetic correlation***

Kinship heritability was estimated using a mixed linear model:

$$\overline{y} = \mu + Zb + e \quad (1)$$

Here, $\overline{y}$ is LWm of 600 lines. $\mu$ is the intercept representing the population mean. $Z$

is the corresponding design matrix that satisfies $ZZ^T = G$, where $G$ is the genomic

kinship matrix estimated from GCTA (Yang, Lee, Goddard, & Visscher, 2011).

Therefore, $b$ follows a normal distribution $N(0, I\sigma_g^2)$. $e$ is the residual with

$e \sim N(0, I\sigma_e^2)$. Kinship heritability is calculated as interclass correlation $h^2 = \frac{\sigma_g^2}{\sigma_g^2 + \sigma_e^2}$.

All the analysis were performed in GCTA. Pairwise genetic correlation among LWm

measured under the four environments was estimated by a multivariant mixed linear

model in ASReml.

***Quantification of plasticity measurements***

Since all the MAGIC lines were phenotyped for LW across four environments, we

quantified and studied the genetics of LW plasticity in response to environmental

variation. Here, the phenotypic plasticity was classified into two categorie. The first

category is overall plasticity describing plasticity across all the studied environments,

while the second category-specific plasticity is more unique to certain pairs of sites,

which only captures the plasticity across two environments. The motivation underlying

such classification is that some individuals are more robust across most of the studied

sites except only one or a few sites, while other individuals are plastic among most of

the sites. One metric, pairwise difference in phenotypic value between two sites, was

used to quantify the specific plasticity. In addition, two additional approaches were used

to quantify the overall plasticity. First the across environment variance (VarR) (Vanous

et al., 2019) of the rank transformed phenotype proposed in Vanous et al., 2019 was

used, was also used to account for the mean difference. The fourth score for the overall

plasticity (FWR) (Finlay & Wilkinson, 1963; Lian & de Los Campos, 2016) applies

Finlay-Wilkinson Regression to partition the phenotype into two components, one is

constant across environments and another responds dynamically to environmental

changes. Using the linear mixed model, the phenotype of each line is partitioned into

these two components and the plasticity component is used as a measurement of

plasticity. In total, the described approaches resulted in eight measurements of

phenotypic plasticity (abbreviated as SP, VarR and FWR).

***Genome-wide association analysis for LW mean/BLUP and plasticity measurements***

Genome-wide association analysis was conducted using a linear mixed model as

implemented in GCTA. We applied a significance threshold of $p < 0.05/Me$ for SNPs,

where Me represents the number of independent SNPs estimated following methods

described by Li et al. (Li, Yeung, Cherny, & Sham, 2012). The effect sizes of identified

quantitative trait loci (QTL) were obtained from the GCTA output. The effect sizes for

each SNP were estimated using the aforementioned mixed linear model.

***Predicting the site-specific performance of the LW***

We fitted the following model to predict the performance of each site for LW.

$$Y = X_1\beta_1 + Zu + e \quad (2)$$

$$Y = X_2\beta_2 + Zu + e \quad (3)$$

Here, Y is a vector of length n*p (n = 600, the number of individuals; p = 4, the number

of sites; n*p = 2400), representing the trait means measured at four sites. u is a vector

of length n*p, representing the breeding value of the n maize line, and e is the randomly

distributed residual with length n*p. The Z matrix satisfies ZZ'=G $\otimes$ I, where G is the

identity by state (IBS) matrix and I is a diagonal matrix of pxp. $X_1$ is a design matrix

with one column of 1 representing column mean and additional 3 columns representing

the environmental effects from the remaining 3 sites, and $\beta_1$ is a vector of

corresponding effect sizes. $X_2$ includes all the columns from $X_1$ and additional

columns with genotypes of the k QTL associated with the mean and plasticity measures

of the tested trait, and additional columns representing the interaction between the k

QTL and the five sites, capturing the effects from QTL by environment factor

interaction. The fitted values from model 3 were referred to as GBLUP predictions

while the fitted values from model 4 were referred to as GEAI predictions. These models were fitted using rrBLUP (Endelman, 2011) package in R (https://www.R-project.org/). In the first case, 80% of the lines with phenotype measured from all the five sites were used as training set and the remaining 20% lines, whose phenotypes were masked as NA, were used as test set. Within site $r^2$ between predicted and measured phenotype were used to evaluate the accuracy of predicted phenotype at each of the four sites. In the second case, we randomly selected all lines with measured phenotypes from three sites to be used as the training set, and the remaining one site was used as the test set and masked to NA. Accuracy was estimated as the regression r2 between the measured and predicted phenotype within each locus. All the lines with measured phenotype were used as training set and lines masked as NA was used as test set. Accuracies were estimated as the regression $r^2$ between measured and predicted phenotypes within each site.

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**Author Contributions**

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398     **The authors have no conflicts of interest to declare.**

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401     **Reference**
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