Mapping fitness landscapes to interpret sex allocation in hermaphrodites

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9 **Summary**

Sex allocation theory predicts sex ratios of dioecious organisms, but it has been poor at

explaining sex allocation in hermaphrodites in which the assumed trade-off between male and

female functions is often obscure. Here, we apply sex-allocation theory to hermaphrodites by

mapping components of seasonal reproductive success onto a fitness landscape defined by

potentially independent measures of allocation to male and female functions on orthogonal axes.

We find that peaks of reproductive success in a perennial hermaphroditic plant reflect the

interactive effect of both male and female allocations on self-fertilization and the effects of

inbreeding depression. The rugged landscape corresponds well to the complex pattern of sex

allocation observed in natural populations in which individuals produce a mix of male and

bisexual flowers and express a type of gender diphasy. Our approach may help to interpret

common complexities of sex allocation in hermaphroditic plants and animals.

Keywords: alpine plant; andromonoecy; androdioecy; fitness surface; gain curve; gender

diphasy; modularity; paternity; sexual interference; sexual selection

Introduction

The theory of sex allocation has provided a simple and powerful framework for understanding the reproductive strategies of species with two different gamete sizes[1]. The theory explains when a species should be dioecious versus hermaphroditic[2], what sex ratio dioecious species should have under different scenarios of inbreeding, kin competition or conflict[3,4], whether sex should be determined by genes or by the environment[5], what proportion of reproductive resources hermaphrodites should allocate to their two sexual functions[2,6], whether hermaphrodites should allocate to both sexes simultaneously or sequentially[7], and whether (and how) their allocation strategy should depend on their size[8,9]. But whereas sex-allocation theory has been relatively successful in predicting the sex ratio and timing of allocation to sons versus daughters in species with two separate sexes[10], it has been notoriously difficult to test in populations that include hermaphrodites[11–16] (but see sex change in hermaphroditic animals[17]).

Applying sex-allocation theory to hermaphrodites has been unsuccessful for several reasons. First, the theory is based fundamentally on the assumption that individuals must divide a given reproductive resource between their male and female functions in a zero-sum game, with a simple allocation trade-off between the two sexes[1,2,6]. This assumption is straightforward in dioecious species in which the sex allocation is simply the progeny sex ratio[4,18]. However, empirical studies on hermaphroditic species have found it difficult to identify the resource that hermaphrodites must divide between their male and female functions – and even to define what ultimately constitutes allocation to each function, e.g., whether it should include costly ancillary structures that promote mating success[12,14,15]. Size-dependent sex-allocation theory recognizes a likely important dimension of individual variation in resource availability, but it does not tell us which resource to measure or precisely how the resource should vary with an individual's size[8]. Ignorance of the relation between size and resource availability may be one of the reasons for our failure, with few exceptions, to uncover male-female allocation trade-offs – even when size is statistically accounted for[19,20]. Moreover, sex-allocation trade-offs may also be obscured by trade-offs between reproduction and other allocation sinks such as growth or defense[21–24].

Second, testing sex-allocation theory for hermaphrodites requires an accurate estimate of the fitness gained through each sexual function, but we have few estimates of both male and female reproductive success in natural populations, especially in plants[25–27]. Estimates of the female reproductive success of hermaphrodites have been easier than those of male reproductive success, but they have usually been made in terms of the number of seeds or fertilized eggs without accounting for establishment probability or viability[12]. In self-compatible hermaphrodites, self-fertilization is known to strongly affect progeny establishment through inbreeding depression[28]. Yet although some sex-allocation models incorporate the effects of selfing and inbreeding depression[2,29], to our knowledge there are still no empirical tests of these models or evaluations of how fitness gains are affected by inbreeding.

Third, tests of sex-allocation theory have been largely limited to considering allocation at the individual level[30–32], yet in modular organisms such as plants, sex allocation often varies in complex ways among their modules[33,34]. While most flowering plants produce bisexual flowers with both male and female sexual organs, the individuals of many species vary in their sex allocation among flowers or reproductive modules. Such heteromorphic strategies include 'monoecy' (the simultaneous production of separate male and female flowers), and 'andromonoecy' (the production of bisexual and male flowers)[35]. In many functionally hermaphroditic species, individuals also vary their sex allocation over time, either within a season in terms of phenology (e.g., flowers may be protandrous or protogynous), or among seasons (e.g., species showing 'gender diphasy', in which individuals shift from being male to female or hermaphrodite over the course of their lives)[36,37]. Some models of sex allocation seek to predict the optimal allocation strategy for 'temporal' modules of an organism, i.e., how it should allocate its reproductive resources at different times or ages, often corresponding to different sizes[8,21,24,38], but we know of only one attempt to use sex allocation theory to understand the evolution of separate male and female modules on the same individuals[39], and no studies have attempted to test such ideas empirically.

Here, we overcome some of the limitations of applying sex-allocation theory to hermaphrodites by

using a different approach that sidesteps the trade-off assumption (whether it is upheld or not).

Our approach involves mapping male and female fitness components on orthogonal axes of a

two-dimensional fitness landscape. Although mapping reproductive success to genotypes with

continuous traits in a fitness landscape has been common in evolutionary biology[40–42], to our

knowledge this approach has not previously been applied in empirical studies of sex allocation in

hermaphrodites. We illustrate our approach by using it to interpret the complex hermaphroditic

allocation strategy of the self-compatible perennial herb, *Pulsatilla alpina* (L.) Delarbre

(Ranunculaceae), which displays both andromonoecy and a type of gender diphasy[43,44]. This

variation demands a functional explanation, but it also presents an opportunity to consider the

fitness implications of a wide range of allocation phenotypes at the flower level. While sex

allocation is already highly variable among *P. alpina* flowers[45], we further enhanced this variation

by removing stamens from a subset of flowers (Figure 1). This manipulation allowed us to infer

the fitness implications of producing female flowers, which do not occur naturally in the species

(and to offer an explanation for why). By focusing on a population of *P. alpina* comprising mostly

single-flowered individuals, we could use genetic markers to attribute paternity to individual

flowers and their phenotypes, which is not possible for multi-flowered individuals.

Although fitness is ultimately a property of individuals, by assessing the contribution to fitness

through seed production and siring success at the modular, floral, level, we could begin to ask

why a plant might produce both male and bisexual flowers with strongly different sex allocation.

How, for instance, does the production of a male flower contribute to an individual's marginal

fitness gain compared to the production of a hermaphroditic (or female) flower in its place? Such

an evaluation is presumably the general basis of allocation decisions taken by plants with

heteromorphisms such as andromonoecy or monoecy, with each new module receiving

resources for its male versus female functions depending on the context of, and prospective

fitness contribution made by, that module and other modules of the same individial[46] (we

comment on limitations of using individuals with only single flowers in the Discussion).

We showcase our approach by modelling fitness components for flowers of *P. alpina* at the flower

scale in terms of male and female allocations under two scenarios of inbreeding depression, both

on their own as single variables and jointly. We first modelled estimates of male and female

fitness in terms of components of sex allocation to characterize 'fitness gain curves' that relate

fitness through each sexual function to allocation to that same function using univariate analysis.

We then used multivariate selection gradient analysis to test how female and male allocations

jointly affect fitness components. Finally, we map fitness components onto a two-dimensional

fitness landscape. This mapping reveals fitness peaks that correspond closely to the complex

sex-allocation strategy adopted by *P. alpina* in the wild, reflecting the interactive effects of male

and female allocation beyond any simple trade-off. Our analyses provide new insights into the

evolution of the complex reproductive strategies of a perennial and modular organism that may

reinvigorate the empirical application of a conceptually powerful theory to hermaphrodites.

Results

Variation in floral sex allocation and components of reproductive success

We measured the male and female components of sex allocation for a total of 175 flowers, with

104 individuals producing a single flower and 31 individuals producing more than one flower

(Figure S1). Prior to the stamen removal manipulations, 46 and 129 flowers were male and

bisexual, respectively. Note that experimental stamen removal did not alter the flowering duration

of the flowers[45]. We then estimated female, male, and total reproductive success for all

individuals based on a paternity analysis using seeds sampled from all of the bisexual flowers in

the population (N = 104 seed families; 25 bisexual flowers were aborted or missing). Each seed

family produced 90.5 ± 55.4 mature seeds (mean ± SD; N = 104). We could assign paternity for

854 seeds to a single most likely father under a relaxed confidence interval (80%), corresponding

to a 96% successful assignment rate for what amounted to around 9% of all mature seeds

produced in the population. Our inferences are based on a sampling intensity greater than that of

most recent studies of the mating system[47] and phenotypic selection in hermaphroditic plants

[26,27,32,48–51]. Using selfing-rate estimates for seed families[45] and adopting a value of inbreeding

depression of δ = 0.95, which we estimated on the basis of comparisons of inbreeding

coefficients between adults and seed progeny[52], we inferred the mean female, male, and total

reproductive success of single-flowered individuals to be 41.4 ± 48.6, 33.6 ± 41.5, and 75 ± 61.3,

respectively (N = 87; flowers of 17 individuals were aborted or missing; Figure S2). These values

compare with corresponding estimates of 65.8 ± 58.4, 57.9 ± 51.9, and 123.7 ± 88.3,

respectively, under the assumption of a hypothetical scenario of δ = 0 (no difference in fitness

between selfed and outcrossed progeny). Comparing fitness estimates between contrasting

scenarios of inbreeding depression reveals that inbreeding depression may strongly affect

inferences of selection on sex allocation and likely other traits.

Female and male floral fitness gain curves

We estimated the shape of the male and female gain curves with a power function for individual

flowers on the basis of reproductive success estimates for the 87 parents in our sample that

produced a single flower (Figure 2). Regression of our estimate of female fitness on pistil

production within flowers points to an accelerating female fitness gain curve (exponent b = 1.91 ±

0.3; mean ± se; significantly > 1.0; P < 0.01; Figure 2A and Table S1), confirming the result

reported by Chen and Pannell[44]. Whereas under the hypothetical scenario that assumes no

inbreeding depression, we infer a slightly saturating gain curve for the female function (b = 0.85 ±

0.16; Figure 2C), though the curve does not differ statistically from linearity (b is not significantly

different from 1.0; P > 0.05; Table S1). In contrast to the inferred accelerating female fitness gain

curve, our data point to saturating gain curves for male function, whether incorporating our

estimate of inbreeding depression (b = 0.47 ± 0.38; Figure 2B) or not (b = 0.52 ± 0.27; Figure 2D)

– though again neither of these two gain curve estimates differed significantly from linearity (P >

0.05; Table S1). Note that the high variance around the fitted male gain curves likely reflects the

potential effects of cross-sex interactions not captured by univariate analysis or due to greater

uncertainty around the estimates of male fitness.

Linear, quadratic, and correlational selection gradients on female and male allocation

We further performed a multivariate selection gradient analysis[53] on the single-flowered

individuals (N = 87), which fully confirmed the above results: female reproductive success was a

clearly accelerating function of allocation to pistils when incorporating inbreeding depression in

our fitness estimates, but linear when inbreeding depression was ignored (Figure 3A). Similarly,

male reproductive success was a mostly saturating function of allocation, though saturation was

only significant for the hypothetical scenario of no inbreeding depression (Figure 3E and

Supplementary Information Table S2). Our selection-gradient analysis for female reproductive

success also points to mostly negative directional selection on male allocation when fitness

estimates incorporate inbreeding depression (Figure 3D and Supplementary Information Table

S2), as a result of ovule discounting[54], because flowers with more stamens had a higher selfing

rate[44]. For male reproductive success, our results point to stabilizing selection on female

allocation, but only when inbreeding depression is ignored (Figure 3B and Supplementary

Information Table S2). Total reproductive success depended only on female allocation when

inbreeding depression was accounted for, but on both female and male allocation when ignoring

it (Figure 3C and F; and Table S2). We could detect no correlational selection on female and

male allocations (correlational selection gradients did not differ significantly from zero; Table S2).

Two-dimensional fitness landscapes for female, male, and total reproductive success

Maps of female, male, and total reproductive success on a two-dimensional sex-allocation

landscape are presented in Figures 4A, B, and C, respectively. The two fitness peaks in the figure

for total reproductive success correspond qualitatively closely to phenotypes of unisexual male

and bisexual flowers, respectively (Figure 4C and S5). Standardized linear, quadratic, and

correlational selection gradients, estimated using the *GSG* package for R[55], are presented in

Table S3. Note that we detected evidence for negative correlational selection gradients for total

reproductive success when incorporating inbreeding depression in our fitness estimates (Table

S3). Variance explained by fitted models for female, male, and total reproductive success were

79.9%, 58.2%, and 53.4%, respectively, when considering the influence of inbreeding depression,

and 85.7%, 63.3%, and 67.5% when inbreeding depression was ignored.

Interpretation on gender-diphasy from the fitness landscapes

To aid interpretation of gender-diphasy in *P. alpina,* we drew hypothetical lines on the predicted

fitness landscapes for strategies reflecting individuals of low, medium, and high resource status

(when δ = 0.95); here, we assumed a linear sex-allocation trade-off with an arbitrarily chosen

negative slope (Figure 4A, B, and C). We then plotted male, female and total fitness gain curves

along each of these three hypothetical trade-off relations (following classic sex-allocation theory),

with relative sex allocation computed as stamen number over the number of stamens and pistils

together (Figures 5). Adopting either a unisexual male or a bisexual strategy yielded the highest

total fitness returns for individuals with low (dotted line) and high (solid line) resource status,

respectively (Figure 5C).

Discussion

Estimation of male and female gain curves in a natural population of flowers

Our study provides a rare estimate of the shape of fitness gain curves for a natural plant

population. It has hitherto been difficult to estimate the shape of fitness gain curves for several

reasons, not least because most species do not display a sufficiently wide range of phenotypic

variation in natural populations over which to estimate fitness. By removing some or all stamens

of a sample of flowers of our study species *P. alpina*, we extended the already wide range of sex

allocation represented in the population to include not only male and bisexual flowers, which

occur naturally, but also fully female flowers. Because the range of sex allocation in the study

population covered the full plane of allocations, we could estimate male and female gain curves

relatively independently of one another[56].

Our analyses indicate that the male fitness gain curve for *P. alpina* tended towards saturation in

single-flowered individuals. Although there have been very few empirical estimates of the shape

of the male gain curve for animal-pollinated plants[12,16], those that do exist have also found

evidence for saturating male gain curves[57–59] (but see reference[60]). Saturating fitness gain curves

are expected if pollen accumulation saturates on pollinators' bodies[12], or if pollen from a given

individual is delivered to a small number of receptive stigmas, causing local mate

competition[6,61,62] or local sperm competition[15]. *Pulsatilla alpina* is pollinated almost exclusively by

flies that cause substantial within-flower self-pollination[44] and that disperse pollen among

individuals over short distances (mean = 3.16 m)[63], conditions that should give rise to local mate

competition and thus to saturating male fitness gain curves.

In contrast with the tendency towards saturation for male gain curve, the female gain curve for

single-flowered individuals of *P. alpina* was strongly accelerating, likely as a result of sexual

interference between male and female allocation[64]. The female gain curve in plants is often

suspected to be saturating rather than accelerating, because competition among progeny

dispersed into a limited seed shadow ('local resource competition') compromises the seed

producer's inclusive fitness[12,65,66]. While we did not measure possible effects of local resource

competition on the female gain curve, we believe this possibility is negligible in *P. alpina* because

the single-seeded fruits (achenes) are furnished with parachute-like structures that aid their

dispersal by wind from elongated floral stalks[67,68]. Accordingly, although pollen is dispersed over

very short distances, there is almost no genetic structure in populations of *P. alpina*[63], consistent

with a scenario of well-dispersed seeds[69]. The accelerating fitness gain curve estimated for *P.*

alpina flowers likely derives from the fact that flowers with more pistils and/or fewer stamens have

a lower within-flower selfing rate as a result of an increased separation between pistils and

stamens within flowers[44] and thus a lower rate of seed discounting[54]. This explanation is

confirmed by analysis that ignores the effects of inbreeding depression, which predicted a slightly

(albeit not statistically significant) saturating female gain curve for flowers rather than the

accelerating curve predicted by analysis that incorporates inbreeding depression into estimates of

floral contributions to fitness.

A landscape approach for interpreting complex sex allocation strategies

We found that individuals of *P. alpina* occupy a large part of the area on the male/female

allocation plane, indicating that sex allocation is clearly not constrained to lie on a single

straightforward trade-off diagonal assumed by sex-allocation theory (see Figure 1). Such a

scenario is likely to be common in perennial plants and animals whose resource status varies

with size or for other, more cryptic, reasons[15,70]. A corollary of our finding here is that the female

component of reproductive success in *P. alpina* likely depends not only on female allocation, but

also on male allocation in ways that go beyond linear sex-allocation trade-offs (and vice versa for

male reproductive success). Our results indicate that the selfing rate and a flower's contribution to

female reproductive success was indeed a function not only of the number of pistils, but also of

the number of stamens for a given pistil number[44]. Classic sex-allocation theory is unable to

account for such patterns in a straightforward manner, though indirect effects of allocation to one

sex on fitness through the other may be incorporated in a framework of sexual selection[71].

Indeed, although we found that the contribution to male reproductive success in *P. alpina* was

independent of the number of pistils in single-flowered individuals, a previous study on the same

species, but focusing only on bisexual flowers, found evidence for a cross-sex effect of this sort,

consistent with the within-flower sexual conflict[45].

To appreciate the value of mapping fitness components on a two-dimensional landscape of

empirically measurable components of sex allocation when resource status obscures potentially

underlying trade-offs, consider the diagonal lines traced on the fitness maps inferred for *P. alpina*

in Figure 4A-C. Each of the lines represents a hypothetical linear sex-allocation trade-off, and a

walk over the fitness landscape following these paths (Figure 5) illustrates how the shape of

fitness gain curves may vary with plant resource status (i.e., small, medium and large). This

depiction suggests that small individuals of *P. alpina* should allocate most of their reproductive

resources to their male function, while larger individuals (with more resources) should allocate

substantially to both male and female functions (Figure 5). With growth, individuals should thus

shift from an all-male to a hermaphroditic allocation strategy, i.e., they should display a type of

gender diphasy[21,36,72], as indeed observed in wild populations of *P. alpina*[43] and many other

perennial plants[73–75] and animals[76].

The existence of two fitness peaks on the sex allocation landscape for *P. alpina* flowers (Figure 4)

also points to andromonoecy as a successful strategy, with some flowers adopting a fully male

and others a bisexual strategy. Again, this pattern corresponds qualitatively to the strategy

adopted by *P. alpina* (Figure S3). By abandoning their female function and producing fully male

flowers, individuals of *P. alpina* may capitalize on opportunities for mating in the early season

when most of the other bisexual protogynous flowers are in their female stage[45]. It is plausible

that individuals with substantial reproductive resources in general might set aside as much as

they can to produce bisexual flowers and use whatever remains for one or more male flowers at

the beginning of the season (note that all *P. alpina* floral buds are formed below ground at the

end of the previous growing season[77]). In andromonoecious species of *Solanum*, for instance,

there is a strong correlation between the size of the fruit and the fraction of male flowers,

supporting the notion that male flowers in andromonoecious species in general may often serve

to balance sex allocation in this way[78].

Finally, the topography of the fitness landscape provides a clear indication of a saturating (rather

than the otherwise accelerating) female fitness curve for flowers of *P. alpina* from which stamens

had experimentally been removed, i.e., for a phenotype that does not occur in nature (because *P.*

alpina flowers always bear stamens; Figures 4A and S4). *Pulsatilla alpina* stamens function not

only to sire ovules but also to attract pollinators to flowers that do not produce nectar, so that the

lower female reproductive success of flowers without stamens may have been a result of

pollinator limitation – though this conjecture needs testing. Nonetheless, this possibility further

exemplifies the utility of our approach in being able to reveal cross-sex effects when they exist.

The function of stamens as both the source of pollen for siring offspring as well as a reward for

pollinators might further help to explain why the flowers of *P. alpina* always produce stamens and

thus why the species is andromonoecious (with unisexual male and bisexual flowers) and not

monoecious (with unisexual female and male flowers), likely resembling the cases of

andromonoecy in *Solanum*[79–81].

We have interpreted strategies at the floral level using a fitness landscape inferred from a

population of mostly single-flowered individuals in a species that often produces multiple flowers.

This needs to be borne in mind when interpreting a multi-flower strategy. First, inferences of

positive marginal effects on fitness through the production of male flowers will need to be

tempered if the flowers negatively affect the fitness contribution of other flowers with a female

function, e.g., through increased geitonogamous selfing[82] (but see[25]). This seems an unlikely

scenario in *P. alpina*, because its male flowers tend to be open at the beginning of the flowering

season and thus often function alone even for individuals that produce bisexual flowers[43].

Second, the inference of an accelerating fitness gain curve for the female function of individuals

with only one flower would be misleading for individuals that disperse seeds from many flowers

into a small seed shadow that then suffer from local resource competition[12,65,66]. As discussed

above, we believe that local resource competition is unlikely to be important in *P. alpina*[67,68].

Finally, the fitness landscape we have inferred for single-flowered *P. alpina* individuals might

seem to point to androdioecy (with distinct male and hermaphrodite individuals), as much as to

andromonoecy. But multi-flowered males would produce their male flowers over the course of the

reproductive season, yet the andromonoecious strategy of *P. alpina* involves producing one or a

few male flowers only at the beginning of the season when the species' strong protogyny ensures

high mating opportunities[43,45]. It is well established that androdioecy is difficult to maintain when

males must compete with hermaphrodites for siring opportunities, especially in partially selfing

species[83,84], as is the case of *P. alpina*[44].

Conclusion

Sex allocation theory has been successful in helping to explain variation in sex ratios in dioecious

species, particularly animals, but its quantitative application to explain the complexity of sex-

allocation strategies in hermaphrodites has had limited success. Our study has overcome key

operational difficulties by successfully estimating both the male and female contributions to

fitness by floral modules across their full potential range of allocations. It also showcases a simple

but potentially useful approach for interpreting complex patterns of sex allocation in terms of the

shape of a fitness landscape defined by independent measures of male and female allocation on

orthogonal axes. This approach may provide an empirically accessible rescue-line for a body of

theory that has been increasingly criticized for being too difficult to apply to the messy world of

hermaphroditic reproduction[12,15,56,85,86].

Resource availability

Lead contact

Further information and requests for resources should be directed to the lead contact, Kai-hsiu

Chen (kai-hsiu.chen@unil.ch).

Materials availability

There are no new materials associated with this paper.

Data and code availability

Data and codes used in this study can be accessed in Zenodo (DOI: 10.5281/zenodo.14289403).

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Figures

Figure 1. Morphological space of stamen and pistil number of the single-flowered individuals after

stamen removal treatments. Green, orange, and blue points represent all-stamen-removed, half-

stamen-removed, and intact flowers, respectively (N = 8, 18, and 61 respectively). Data points

were jittered along the y-axis.

Figure 2. Fitness gain curves of female (upper panels) and male functions (bottom panels) at the floral level under the condition when inbreeding depression is taken into account (left-hand panels) or not (right-hand panels). Each point represents one individual with one flower ($N = 87$). The points were jittered to avoid overlapping. The shape of the gain curves was estimated by fitting exponential curves (see materials and methods for details) and the exponent b is shown in the figure with the standard error. An asterisk denotes that the curve was significantly non-linear.

Figure 3. The dependence of female (upper panels), male (middle panels), and total (bottom panels) reproductive success (RS) on female and male allocation under the condition when inbreeding depression is taken into account (green lines) or ignored (yellow lines), estimated by selection gradient analyses (N = 87, see materials and methods for details; Table S2). The shaded ribbon indicates the range of one standard error of the partial regression curves.

Regression lines of non-significant and marginally non-significant dependency of reproductive

success on the sex function are shown in dotted and dashed lines, respectively. Raw data points

were jittered to avoid overlapping. Panel **A** and **D** fully confirmed the results of a previous study

based on the reproductive success of bisexual flowers only [44].

Figure 4. Representations of the fitness landscape for female, male, and total reproductive

success (RS) as a function of pistil and stamen number in a flower under the condition when

inbreeding depression is taken into account (left-hand panels) or ignored (right-hand panels),

predicted by generalized additive models (*gam*) using 87 individuals with a single flower (grey

circles). The colour gradient from white to red represents low to high predicted reproductive

success in terms of the number of seeds. Hypothetical linear trade-off lines between male and

female functions were depicted by dotted (line a), dashed (line b), and solid lines (line c) for

individuals of low, medium, and high resource status, respectively (Panels **A**-**C**). Note that the

slope of the trade-off lines is conceptual, because we do not know the linearity and the actual

trade-off ratio relating female to male units of allocation. Individuals with a given amount of

resource are only able to explore the left and bottom part of the trade-off line on the fitness

landscape.

Figure 5. Conceptual figures demonstrating how female (**A**), male (**B**), and total (**C**) fitness gain

curves depend on the resource status of an individual based on the study of *P. alpina*.

Relationship of the projected reproductive success and sex allocation to the male function along

the hypothetical trade-off lines a, b, and c for individuals of low, medium, and high resource

status, respectively, were extracted from Figure 4 **A** - **C**. Sex allocation in terms of maleness was

calculated by dividing the stamen number by the sum of stamen and pistil numbers. (**C**) Inferred

optimal sex allocation strategies changed from predominantly male-biased toward highly female-

biased for individuals with low and high resource status, respectively. Note that the slope of the

trade-off lines is conceptual, because we do not know the linearity and the actual trade-off ratio

relating female to male units of allocation.

Materials and Methods

Experimental model and study participant details

We studied a population of *Pulsatilla alpina* (L.) Delarbre (Ranunculaceae) at Solalex in the pre-Alps of Vaud canton, Switzerland ('Population S1+'; latitude: 46°17′42″N, longitude: 7°09′09″E; elevation: 1758 a.s.l.) in the flowering season of 2022. The species grows in sub-alpine to alpine habitats in central Europe, with longevity likely exceeding 30 years[87]. Each spring, several vegetative and/or reproductive shoots emerge from a rhizome soon after the snowmelt, with single flowers on separate flowering shoots. Small plants often produce a single male or bisexual flower, while larger plants could produce up to about 20 usually bisexual flowers, i.e., the species displays gender diphasy as a result of size-dependent sex allocation[43]. Bisexual flowers have a similar number of stamens to that of male flowers and up to 400 single-ovulated pistils (Figure 1)[45]. Flowers are predominantly visited by flies[68]. Ripe fruits (technically achenes) with elongated pappus hairs are dispersed by wind in early autumn[67]. *Pulsatilla alpina* is self-compatible and has a selfing rate of about 0.47 (95% *CI*: 0.38 – 0.54; based on 1,000 bootstraps), with the rate of selfing being a positive function of within-flower male allocation[44]. Inbreeding depression was estimated to be 0.95 (0.92 – 0.96 under the 95% *CI* of mean selfing rate)[44] based on comparisons of inbreeding coefficients (F_{IS}) between adults and seed progeny in the study population[52]. As in most studies of inbreeding depression in plants, our estimate of inbreeding depression ignores the effect of inbreeding on seed development from the zygote and is thus somewhat conservative. The study population comprised a total of 104 individuals producing a single flower and 31 individuals producing more than one flower in 2022 (N = 135 flowering individuals, and a total of 175 flowers; Figure S1). The population was situated on an open slope of sub-alpine grassland that we fenced to exclude the most predominant herbivores, i.e., cattle and other browsers. We removed all floral buds from the few unsampled individuals outside the plot at the beginning of the flowering season to prevent them from siring progeny in the plot.

Method details

Flowering phenology

We recorded the location and flowering state of all individuals in the population every three or four

days from late May to late June, 2022 (in total twelve census times), noting the number and

sexual phenotype of their flowers. For each flower and sampling date, we recorded each flower's

sexual stage in terms of seven and five ordinal categories for bisexual and male flowers,

respectively (detailed description of the categories in Chen & Pannell[45]). These categories

allowed us to manipulate flowers at the same developmental stage, i.e., before anther

dehiscence.

<u>Manipulation of floral sex allocation</u>

Although flowers of *P. alpina* vary widely in their sex allocation, natural variation does not include

flowers with few or no stamens. To estimate the potential fitness contributions of female or nearly

fully female flowers, and thus to evaluate the full two-dimensional fitness landscape, we removed

stamens of a sample of flowers at times throughout the flowering season, as has been done in

other studies[32,56,88]. Prior to the stamen removal manipulations, 46 and 129 flowers were male

and bisexual, respectively (Figure S1). At each observation time point, we randomly selected

about a quarter of the bisexual flowers in their early female stage and used tweezers to remove

100% or 50% of their stamens. Similarly, we removed 50% of the stamens of about a quarter of

male flowers in their early male stage[45]. Stamen removal did not alter the flowering duration of the

flowers[45], so its effect on components of fitness manifested only through changes in the number

of stamens. Because stamen-removal occurred before anthers opened, there was no accidental

intra-flower selfing.

<u>Estimates of floral sex allocation</u>

We quantified absolute floral sex allocation to the two sex functions as the number of stamens

and pistils produced by each flower. For male allocation, we photographed all bisexual and male

flowers at the late female stage and the mid.male stage (detailed description of the stages in

Chen & Pannell[45]), respectively, and later counted the number of stamens based on the

photographs calibrated against measures for a sample of 15 fresh flowers[43]. Three weeks after

flowering ended, flowers with developing achenes were enclosed in individual paper bags to

prevent seed dispersal, and seeds were later collected for counting. We quantified female

allocation by counting the total number of achenes in each flower; each pistil, which contains a

single ovule, develops into an achene and remains on the floral receptacle until the end of the

season, whether fertilized or not[68].

Paternity assignment and estimates of floral selfing rates

We first counted the number of mature seeds from a total of 22,612 achenes belonging to 104

seed families (19 and 6 seed families from 129 bisexual flowers were aborted or missing,

respectively). Each seed family produced 90.5 ± 55.4 mature seeds ($N = 104$; mean $\pm$ SD). The

seeds were well mixed and randomly selected from each seed family for genotyping. We

extracted DNA (BioSprint 96 DNA Plant Kit), conducted PCR (HotStarTaq Master Mix Kit), and

aimed to assign paternity to each of ten seeds sampled randomly from each flower in our sample,

using variation at ten microsatellite loci (manually checked in GeneMapper v 6. 0) and the

software Cervus v 3.0.7 set with a confidence level of 80% and an error rate of 0.018 (see details

of the genetic markers in Chen & Pannell[45]). The ten genetic markers were highly polymorphic in

the study population, with a mean of 8.7 alleles per locus and a mean polymorphic information

content (PIC) of 0.65. We included all the flowering individuals (i.e., all single- and multiple-

flowered individuals) as potential sires in the paternity analysis, none of which had an identical

genotype. Furthermore, the average kinship coefficient F_r was -0.0001 over all the flowering

individuals[63], reassuring the accuracy of our paternity analysis[89]. In total, 892 of 1054 sampled

seeds could be genotyped for at least five loci and used for paternity analysis (PCR was

unsuccessful for the remaining seeds and were not included). The selfing rate was calculated by

dividing the number of selfed seeds by the total number of paternity-assigned seeds for each

flower.

Estimates of female reproductive success

We calculated components of reproductive success assuming both the estimated value of

inbreeding depression (δ = 0.95), as well as assuming no inbreeding depression (δ = 0), with

female reproductive success (RS_F) computed as the inferred number of mature outcrossed seeds

plus (1 − δ) times the inferred number of mature seeds produced by selfing, based on the number

of mature seeds (N_T) and the selfing rate (S) estimated for each bisexual flower with the following

formula: $RS_\mathrm{F} = N_\mathrm{T} \times (1 - S) + N_\mathrm{T} \times S\,(1 - \delta)$.

<u>Estimates of male reproductive success</u>

We calculated male reproductive success for each flower as the number of outcrossed seeds

sired on other individuals in the population (seed families from all the flowering individuals; N =

104) plus the number of seeds sired by selfing multiplied by (1 − δ), again for both δ = 0.95 and

0.0, as described before. Because we successfully assigned paternity for about eight seeds for all

flowers (total N = 854), irrespective of the total number of seeds produced by the flower, we

estimated the male reproductive success of a given potential sire by multiplying the fraction of the

seeds in the flower it sired by the total number of seeds in that flower (mean correlated paternity[90]

was 0.19 ± 0.27 based on 88 seed families that had at least two outcrossed seeds). It is worth

noting that although it would always be useful to increase the sample size of seeds from one

seed family to increase the precision of the estimates of the selfing rate and paternity share, the

power of our analysis lies in the replication of phenotypes varying in their sex allocation in the

population (N = 87 individuals in total as the experimental units in the following analyses). A

supplementary analysis using five seeds subsampled per family (total N = 497 seeds; all seeds

were used if the family produced less than five seeds) yielded an almost identical fitness

landscape, indicating that our fitness inferences based on estimated selfing rates and paternity

shares are robust (Figure S5).

Quantification and statistical analysis

All the statistical analyses were conducted in R studio (v2024.12.1+563) with R (4.3.1)[91].

To investigate the relationship between components of fitness to floral sex allocation, we focused

on female, male, and total reproductive success for all individuals that had only one flower ($N =$

87; 17 individuals with its flower aborted or missing were excluded). Because we wished to

measure the total strength of selection (direct and indirect effects via correlation with other traits)

on female and male sex allocation[41], we excluded floral traits from our model, e.g., stalk height,

petal length, and flowering date. Nonetheless, including the floral traits as covariates for

estimating the strength of direct selection on sex allocation yielded qualitatively highly consistent

results (see Figure S6), presumably because reproductive success is primarily determined by sex

allocation in *P. alpina*. Moreover, a previous study[45] found that stamen removal did not alter the

flowering duration of *P. alpina* flowers, with the treatment effect on components of fitness

manifested only through changes in stamen number.

We related prospective reproductive success to sex allocation using nonlinear univariate least

square models with a Gaussian error distribution (*nls* function in R *stats*[91]) to evaluate the shape

of fitness gain curves for female and male functions at the flower level, assuming both $\delta = 0$ or δ

$= 0.95$. The gain curve y was modelled as $y = ax^b$, where y is the female or male reproductive

success relative to the mean of that sex function for single-flowered individuals, x is the number

of pistils or stamens for the female and male functions, respectively, a is a constant, and $b < 1$

and $b > 1$ correspond to a saturating or accelerating dependence of fitness on sex allocation[32,92].

P-values were calculated using a one-sample *t*-test.

We also used conventional selection gradient analysis with multivariate regression to evaluate the

dependency of reproductive success on both the female and male components of allocation (i.e.,

as two phenotypic traits[93]), including second-order polynomial and interaction terms under the two

scenarios of inbreeding depression at the flower level[53]. We fitted relative female, male, and total

reproductive success (for both scenarios of inbreeding depression) as a response variable in six

separate models, each with a Gaussian error distribution (*lm* function in R *stats*[91]). We

standardized pistil number and stamen number to a mean of zero and a standard deviation of

one, setting linear, quadratic, and interaction terms for the two traits in each model to evaluate

linear and non-linear (i.e., quadratic and correlational) selection gradients on female and male

allocation[53,94]. For all quadratic gradients, we multiplied the regression coefficients by two to

obtain the correct estimate of stabilizing or disruptive selection[95]. *P*-values of the estimates were

calculated using *t*-tests in six multivariate regression models.

Finally, we used nonparametric regression with smoothing functions to characterize the fitness

landscapes for reproductive success in terms of female and male allocations under the two

inbreeding depression scenarios[55,96]. Here, we used generalized additive models (*gam* function in

R package *mgcv*[97]) for the dependency of components of reproductive success (as the numbers

of seeds produced and/or sired) on female and male allocations on pistil and stamen number,

assuming a Poisson error distribution; assuming negative-binomial or zero-inflated Poisson

distributions yielded qualitatively identical results but explained less variance (Figure S6). We

applied the *gam.gradients* function in the *R* package *GSG* to extract standardized linear,

quadratic, and correlational selection gradients from the fitted models and calculated the standard

errors and *P* values based on 1,000 bootstraps[55]. To plot the fitness surface, we used a

smoothing   term   with   thin   plate   splines[97].

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