Inter-individual gene expression variability implies stable regulation of brain-biased genes across

organs

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Abstract

Phenotypic variation among individuals plays a key role in evolution, since variation provides the material on

which natural selection can act. One important link between genetic and phenotypic variation is gene

expression. As for other phenotypes, the range of accessible expression variation is limited and biased by

different evolutionary and developmental constraints. Gene expression variability broadly refers to the

tendency of a gene to vary in expression (i.e., between individuals or cells) due to stochastic fluctuations or

differences in genetic, epigenetic, or environmental factors, separately from the differences between

conditions (e.g. organs), and is often estimated based on observed expression variance. The variance due to

biomolecular stochasticity (transcriptional 'noise') and cell-to-cell heterogeneity has been well-studied in

isogenic populations of unicellular organisms such as bacteria and yeasts. However, for more complex

organisms with multiple cells, tissues, and organs sharing the same genetic background, the interplay between

inter-individual expression variability, gene and organ function, and gene regulation remains an open

question. In this study, we used highly multiplexed 3'-end Bulk RNA Barcoding and sequencing (BRB-seq) to

generate transcriptome profiles spanning at least nine organs in outbred individuals of three ray-finned fish

species: zebrafish, Northern pike, and spotted gar. For each condition, we measured the expression variance

per gene independent of mean expression level as an estimate of its variability. We observed that lowly

variable genes are enriched in cellular housekeeping functions whereas highly variable genes are enriched in

stimulus-response functions. Furthermore, genes with highly variable expression between individuals evolve

under weaker purifying selection at the coding sequence level, indicating that intra-species gene expression

variability predicts inter-species protein sequence divergence. Genes that are broadly expressed across organs

tend to be both highly expressed and lowly variable between individuals, whereas organ-biased genes are

typically highly variable within their top organ of expression. For genes with organ-biased expression profiles,

we inferred differences in selective pressure on gene regulation depending on their top organ. We found that

genes with peak expression in the brain have low inter-individual expression variability across non-nervous

organs, suggesting stabilizing selection on regulatory evolution of brain-biased genes. Conversely, liver-biased

genes have highly variable expression across organs, implying weaker regulatory constraints. These patterns

show that gene regulatory mechanisms evolved differently based on constraints on the primary organ.

Introduction

Phenotypic variation among individuals is crucial for evolutionary processes, as it serves as the substrate for

natural selection. Gene expression is a key mediator of this phenotypic diversity. Similar to how phenotypic

variation is shaped and constrained by evolutionary and developmental factors, variation in gene expression

among individuals is also subject to comparable limitations and biases, with some genes having a wider

expression distribution than others. This gene expression variance is influenced by various 'intrinsic' (gene-

specific) and 'extrinsic' factors spanning biological scales, including transcriptional or translational

stochasticity [1], cellular heterogeneity [2], genetic variation [3], epigenetic modifications [2,4], as well as

micro- and macro-environmental cues [5]. The observed expression variance (or dispersion) is often used as a

proxy measure for expression *variability*, broadly defined as the *propensity* of a gene to vary its expression

level in response to stochastic, genetic, or environmental perturbations [6].

Focusing on transcriptional or translational stochasticity ('expression noise') in genetically identical

cells, many studies on isogenic populations of unicellular organisms (e.g., *Escherichia coli* [7–11] and

Saccharomyces [12–18]) have probed the relationship between expression noise and gene promoter features

[8–11], as well as properties of the gene regulatory network [19–21]. Evidence suggests that natural selection

can shape expression noise according to gene function [8,12,13,15,22] and that constraints on expression

noise are mediated by environmental context [23]. Furthermore, a correlation between gene expression noise

in isogenic yeast cells and broader expression variation in non-isogenic populations has been reported [24],

suggesting that some shared regulatory mechanisms underlie both intra-genotype expression noise and inter-

genotype expression variability [25]. The extent to which these findings in unicellular organisms can be

extrapolated to complex multicellular organisms, which have functionally specialized organs with distinct

expression profiles despite a common genetic background, is not well understood. While many studies have

advanced our understanding of genetic variants underlying variation in gene expression levels within a

population (e.g., eQTL analysis [26,27]), much remains to be known about the relationship among expression

variability, gene function, and gene regulation in animals. Although organ-specific enhancers exist to facilitate

spatiotemporal control of gene expression [28], we expect some level of pleiotropy of regulation of gene

expression in different organs. Thus selection on gene expression in different organs is not totally

independent, and the expression variability of a gene within any given organ may potentially be constrained

by antagonistic organ-specific expression optima.

Studies on gene expression variability in animals have primarily focused on select cell types or organs

in model organisms [29–31], with particular interests in the role of stochasticity in cell differentiation [32]

(including shifts in inter-cell [33,34] or inter-embryo expression variability over development [35]), as well as

expression variability in human organs and its association with disease [36–43]. (Of note, although most

studies compare expression in different organs, the terminology of 'tissues' and 'tissue-specificity' is

frequently used in the literature instead.) Although the functional relevance of expression variability is

increasingly recognized [44–46], how such variability is modulated by organ context remains largely

unexplored, even in studies that sample multiple organs [25,40,42,43]. Few studies have been published on

variability in outbred, non-model animal populations (but see [47–49]). This is partly due to heavy resource

requirements such as the need for many replicates for statistical significance, as well as the challenges

associated with disentangling different sources of expression variation [50], whether in bulk or at the single-

cell level.

Recently, the advent of low-cost high-throughput sequencing methods has made it increasingly

feasible to generate sufficient transcriptome data to measure within-organ gene expression variance across

several individuals and estimate within-organ expression *variability*. In this study, we present the first analysis

of inter-individual gene expression variability across multiple organs, using newly generated 3'-end Bulk RNA

Barcoding and sequencing (BRB-seq) [51] data in three ray-finned fish species. We confirmed that gene

expression variability between individuals is a functionally relevant trait. Genes that are broadly expressed

across organs have lower within-organ variability in expression between individuals compared to organ-biased

genes, reflecting pleiotropic constraints on expression. We further found that genes with organ-biased

expression profiles tend to be variably, and thus dynamically, expressed in their top ('focal') organ, and that

moreover their variability across other organs depends on their focal organ. Investigating gene expression

variability within and between animal organs provides a new framework for understanding gene regulation

and the evolution of gene function in complex multicellular organisms.

Results

Organ transcriptome profiles of ray-finned fishes

To investigate inter-individual gene expression variability within and between organs, we generated

transcriptomic data across multiple organs using 3'-end Bulk RNA Barcoding and sequencing (BRB-seq) [51]

for three species of ray-finned fishes (Actinopterygii): two teleost species (*Danio rerio* (zebrafish) and *Esox*

lucius (Northern pike)) and one non-teleost outgroup (*Lepisosteus oculatus* (spotted gar)). Zebrafish and

Northern pike represent two of the major teleost clades (Otomorpha and Euteleosteomorpha, respectively)

and neither species have undergone additional lineage-specific genome duplications after the teleost whole-

genome duplication [52], while the spotted gar is a slowly evolving outgroup to the teleosts [53]. For each

species, we sampled multiple outbred individuals spanning several organs in both sexes (when known)

(**Figure 1A**) and applied a two-step filter to remove low quality samples based on sequencing quality

(**Figures S1-S3**) and within-organ expression correlation (Materials and Methods). We observed that this

stringent two-step filtering protocol is sufficient to remove any strong batch effects in our dataset (**Figures S4**-

S6). Hierarchical clustering of zebrafish samples based on pairwise Pearson correlation shows that
transcriptome profiles cluster into five major groups by related organ systems (**Figure 1B**; see **Figures S7-S9**
for PCA plots). Reproductive organs (cluster 1) appear to have the most divergent expression profiles, while
the following clusters can be observed for non-reproductive organs: gastrointestinal and urinary systems
(kidney, intestine, and liver; cluster 2); nervous system (brain and eye; cluster 3); external organs (pectoral fin,
skin, and gills; cluster 4); and lastly, muscular system (heart and muscle; cluster 5). Apart from the gonads, we
do not observe strong clustering by sex except for the liver and pectoral fin. Although some differences can
be observed in Northern pike and spotted gar, clustering of organs of the nervous system, muscular system,
and external organs are generally consistent among species (**Figure S10**).

A

	D. rerio		*E. lucius*		*L. oculatus*
	♀	♂	♀	♂	♀/♂
blood			2	3	6
brain	13	11	10	9	11
eye	11	11	6	5	10
gills	7	9	9	9	6
gonads	11	9	7	5	5
heart	1	2	10	9	9
intestine	11	12	8	8	15
kidney	4	6	9	8	
liver	10	9	4	10	8
muscle	1	1	7	8	7
pectoral fin	10	12	10	10	8
skin	10	11			12
swim bladder			3	5	7
Σ	89	93	85	89	104

B

Figure 1. Study design for multi-organ analysis of inter-individual gene expression variability in fishes. (A) Summary of BRB-seq dataset. Table shows the number of individuals per organ (rows) and sex (columns) within each species, after removal of low-quality samples. There is no information on sex for spotted gar (*L. oculatus*). Organ-sex conditions with fewer than 4 replicates (light gray) were excluded from expression variability analyses. **(B)** Hierarchical clustering heatmap of organ transcriptomes (*n* = 182) in zebrafish (*D. rerio*). Samples cluster into 5 groups by related organ systems: reproductive system (1), gastrointestinal and urinary system (2), nervous system (3), external organs (4), and muscular system (5). Pairwise distances between samples are computed based on Pearson correlation of log-transformed, TMM-normalized counts per million (log$_2$ TMM-CPM) of protein-coding genes (*n* = 19,466). The color keys for organ (inner color bars) and sex (outer color bars) are shown in (A).

Gene expression variability is estimated independently of expression level

Here, we use 'variation' to refer to a measurable statistical feature of an ensemble (i.e., statistical dispersion), and 'variability' to refer to a dispositional property inferred from such measurements [6]. Measuring expression variation as a proxy for quantifying variability has not been standardized in the literature, as it is not well-studied beyond unicellular organisms or model cell lines. Different measures of statistical dispersion (e.g., standard deviation, variance, median absolute deviation, or coefficient of variation (CV)) are known to be strongly anti-correlated with mRNA abundance (**Figure 2A, Figure S11**) [54,55]. To correct for this mean-dispersion relationship, many studies use regression to model variation as a function of expression level [55] and use either the ratio [30,35,42] or difference [25,41] (i.e., model residual) of observed to expected variation as an adjusted estimate of expression variability. Although this removes the global correlation between expression mean and variation, this does not correct for the non-homogenous distribution of variation across the range of expression levels [29,40]. To address this, other studies apply percentile rank normalization (e.g., 'local coefficient of variation' (LCV) [40,47] or 'relative noise rank' [29]), in which the effect of mean expression is removed by ranking the expression variation of each gene relative to a window of genes with similar expression levels. Here, we used the local (residual) coefficient of variation as our estimate of expression variability because it is globally uncorrelated with mean expression level, as in regression-based approaches, and it also imposes a uniform expression variability distribution across expression levels. Although the latter is a very strong correction, by forcing a uniform distribution we maximally remove mean expression as an explanatory factor for downstream analyses.

Since actual expression variation may be skewed by samples with more divergent expression profiles, we computed the LCV for each gene within an organ (for unsexed spotted gar) or organ-sex condition (for sexed Northern pike and zebrafish) using a jackknife resampling or 'leave-one-out' procedure to mitigate the effect of outliers on expression variability estimates. For each subset of *n* - 1 replicates, we first computed the log-transformed, squared coefficient of variation (log(CV)^2, where CV is the ratio of standard deviation to the mean) and fitted a local polynomial regression (LOESS) to model log(CV)^2 as a function of mean expression (**Figure 2B**). Second, we ranked the residual log(CV)^2 ('residual CV') of each gene within sliding windows of 100 neighboring genes sorted by expression level, such that genes with rank 0.00 have the lowest dispersion whereas genes with rank 1.00 have the highest dispersion within a local gene neighborhood (**Figure 2C**). Lastly, we calculated the mean LCV over all *n* - 1 subsets within a condition and used this metric, which we call 'variability rank', as our estimate of gene expression variability (**Figures S12**-**S14**).

Because variability ranks are computed relative to the set of genes expressed within a specific
condition, it is not an absolute measure of variability that is directly comparable between conditions. A gene
with a high variability rank in the brain is considered highly variable only with respect to other genes expressed
in the brain (e.g., a gene with a variability rank of 0.80 in the brain is not necessarily more variable, in the
absolute sense, than a gene with a variability rank of 0.60 in the liver). By its construction, variability ranks are
thus generally organ-specific, with weak to moderate correlation across organs (**Figures S15-S17**).

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Figure 2. Measures of expression variation as a function of expression level. Data from female zebrafish brain (n = 13 samples). All
values are means estimated from jackknife (n - 1) resampling. Mean expression refers to log-transformed, TMM-normalized counts
per million ($\log_2$ TMM-CPM). **(A)** Log-transformed (squared) coefficient of variation (CV, defined as the ratio of standard deviation to
the mean). **(B)** Residuals of $\log_2$ (squared) CV. **(C)** Local percentile rank of residual $\log(CV)^2$ ('variability rank') within sliding windows of
100 neighboring genes sorted by expression level. Blue lines are LOESS regression curves and R is Pearson's correlation coefficient.

One limitation of the variability rank metric is that it does not describe the modality of the underlying
expression distribution. Overdispersion may be due to either a broader range of expression centered at a
single mean (i.e., a unimodal distribution) or the presence of two or more expression peaks (a multimodal
distribution), the latter of which may be indicative of hidden substructures within the population [46]. To
investigate this possibility, we assessed whether more bimodal expression profiles can be observed for sets of
increasingly more variable genes (**Figure S18**) (Supplementary Results; Supplementary Methods). Overall, we
find weak evidence for bimodality in our dataset, only in some conditions, and only at very high variability
ranks. Thus, we infer that higher variability is likely primarily driven by a wider range of expression centered
around a single peak, rather than by bimodality of expression.

Gene expression variability is associated with gene function
After applying a strong correction to remove the effect of mean expression, we verified that our estimate of
gene expression variability retains a functionally relevant signal. For this, we performed a Gene Ontology (GO)
enrichment analysis of GO biological process terms repeatedly represented ($n \geq 3$ conditions) in genes with
lowly or highly variable expression. We observed that sets of lowly (respectively highly) variable genes show a
substantial overlap in enriched GO biological process categories across several conditions. Underdispersed
genes (variability rank $\leq$ 0.20) are consistently enriched in core cellular (housekeeping) functions such as

protein localization and transport, protein modification, RNA processing, and organelle organization (**Figure 3A**). Conversely, overdispersed genes (variability rank ≥ 0.80) are enriched in functions such as signaling, chemotaxis, response to stimuli, and immune response (**Figure 3B**). Functional categories that are under-represented in lowly variable genes also tend to be over-represented in highly variable genes, and vice versa (**Figure S25**). We also performed a GO enrichment analysis for moderately dispersed genes (0.40 ≤ variability rank ≤ 0.60) but observed virtually no over-represented terms, and under-representation of terms related to defense response, immune response, and muscle movement (**Figure S26**). These findings are consistent with previous studies in humans [42,43] and *Drosophila* [25] which have shown that lowly and highly variable genes are enriched in distinct functional categories.

1. establishment of localization in ce...
2. mRNA splicing, via spliceosome
3. RNA processing
4. organelle organization
5. peptide transport
6. retrograde transport, endosome to G...
7. vesicle fusion
8. cellular protein catabolic process
9. protein localization
10. modification-dependent protein cata...
11. protein localization to membrane
12. protein K63-linked ubiquitination
13. cellular catabolic process
14. protein transport
15. DNA repair
16. visual perception
17. Golgi to plasma membrane transport
18. cellular aromatic compound metaboli...
19. GPI anchor biosynthetic process
20. positive regulation of DNA-template...

1. immune response
2. skeletal muscle contraction
3. chemokine-mediated signaling pathwa...
4. multicellular organismal movement
5. negative regulation of endopeptidas...
6. lymphocyte chemotaxis
7. complement activation
8. response to chemokine
9. alcohol biosynthetic process
10. response to xenobiotic stimulus
11. exogenous drug catabolic process
12. biological process involved in inte...
13. cell-cell recognition
14. cellular response to estrogen stimu...
15. steroid biosynthetic process
16. response to oxygen-containing compo...
17. heart process
18. negative regulation of chromatin si...
19. neutrophil chemotaxis
20. regulation of immune system process

Figure 3. Semantic similarity bubble plots of enriched Gene Ontology (GO) biological process terms for (A) underdispersed (variability rank ≤ 0.20) and (B) overdispersed (variability rank ≥ 0.80) genes in zebrafish organs. Over-representation analysis of GO terms was performed separately for the set of under- and overdispersed genes in each organ-sex condition using the conditional hypergeometric test in *GOstats* **[56]**. Terms that are significant (unadjusted p-value < 0.01) in at least three organ-sex conditions were retained. Redundancy reduction and visualization of terms was performed using *GO-Figure!* **[57]**. Each circle represents a group of similar GO terms collapsed into a representative term. Representative terms are plotted in a semantic similarity space such that more similar terms are positioned closer together. The top 20 representative terms are listed and sorted by the number of organ-sex conditions in which the representative is significant. Circle color indicates the number of conditions in which a representative term is significant while circle size is scaled by the number of related terms summarized in a cluster.

Variably expressed genes evolve under weaker purifying selection

By testing for an association between inter-individual gene expression variability and dN/dS, we investigated whether variability evolves under selection pressures that run in parallel to those acting on the protein-coding sequence. For this, precomputed selection statistics were retrieved from Selectome [58,59], a database of conservative branch-site likelihood tests for selection conducted on vertebrate gene trees (Materials and Methods). For each species, the most underdispersed genes (median variability rank ≤ 0.2 averaged across conditions) have evolved under stronger negative selection (dN/dS) compared to the most overdispersed genes (median variability rank > 0.8 across conditions). As there is little variance in dN/dS on the bulk of genes, the global inverse correlation is significant but relatively weak (**Figure 4**; **Figure S20**).

Next, we compared evidence for positive selection for genes with consistently low, moderate, or high expression dispersion across multiple conditions within a species. We observed a statistically significant association between estimated variability and the proportion of branches with evidence for positive selection in zebrafish (Pearson's chi-squared test, χ^2 = 21.866, p = 1.786 x 10^{-5}) and Northern pike (χ^2 = 48.122, p = 3.552 x 10^{-11}), but not in spotted gar (χ^2 = 1.4127, p = 0.4935). Specifically, we found depletion of branches under positive selection for frequently underdispersed genes (variability rank ≤ 0.20 in at least half of available conditions) in both zebrafish and pike, as well as enrichment of positive selection for frequently overdispersed genes (variability rank ≥ 0.80 at least half of conditions) in pike (**Table S4**). Thus, higher gene expression variability between individuals is associated with faster protein-coding sequence evolution between species, both due to relaxed negative selection and to increased positive selection.

Figure 4. Purifying selection of zebrafish genes grouped by median expression variability across conditions. Selection on coding sequences computed from conservative branch-site likelihood tests were retrieved from the Selectome database [58,59]. (A) Species phylogeny with tested internal branches highlighted (turquoise). Tested branches included in all three species are: Euteleostomi, Actinopterygii, Neopterygii; for teleosts, i.e. zebrafish and pike, only: Osteoglossocephalai, Clupeocephala; for zebrafish only: Otomorpha, Otophysi; and for pike only: Euteleosteomorpha, Protacanthopterygii. Icons retrieved from PhyloPic: *Danio rerio* by Jake Warner (CC0 1.0), *Esox lucius* by Timothy Knepp and T. Michael Keesey (PDM 1.0), *Lepisosteus oculatus* by Ingo Braasch (CC0 1.0). Edge lengths are not scaled by divergence time. (B) Zebrafish protein-coding genes were grouped into 10 bins of size 0.1 ([0,0.1], (0.1,0.2], …, (0.8,0.9], (0.9,1.0]) based on their median variability rank across organ-sex conditions. Boxplots show the distribution of ratios of nonsynonymous substitutions (dN) to synonymous substitutions (dS) of sites that have evolved under purifying selection (ω_0 ratio, for dN/dS < 1). The median ω_0 is indicated within each boxplot and the number of genes included are annotated below each. *R* is the Spearman correlation coefficient between ω_0 and the unbinned median variability across conditions, and its corresponding *p*-value. The Wilcoxon *p*-value corresponds to a Wilcoxon rank-sum test comparing the least variable genes (variability rank ≤ 0.2) against the most variable genes (variability rank > 0.8). See **Figure S27** for equivalent analyses in Northern pike and spotted gar.

Brain-biased genes have low expression variability across non-nervous organs

We have shown that genes which are lowly variable across individuals in multiple conditions are enriched in housekeeping functions, while highly variable genes are enriched in stimulus-sensitive functions (**Figure 3**). Since housekeeping genes are, by definition, ubiquitously expressed across conditions (but see [60] for a discussion), this suggests that the relationship works in both directions: genes expressed broadly across organs should also show low within-organ variability across individuals, consistent with stronger pleiotropic constraints. To test this hypothesis, we quantified expression breadth per gene using the expression specificity metric tau (τ), with range [0,1], in which genes with τ close to 0 are uniformly expressed across organs whereas genes with τ close to 1 are highly organ-specific [61,62] (Materials and Methods). Importantly, τ only considers mean expression differences between conditions and does not take within-condition expression dispersion into account, thus it is independent of our estimate of variability. Across conditions in each species, results consistently show a positive association between variability rank and organ specificity (**Figures S28**-**S30**). Based on the lower quartile of the τ distribution averaged across species, we designated genes with $\tau \leq 0.3$ as 'broadly expressed' and genes with $\tau > 0.3$ as 'organ-biased' (**Figure S31**). Note that we explicitly use the term 'organ-biased' rather than 'organ-specific' because our relaxed τ cutoff necessarily includes genes that are preferentially expressed in one organ but are still expressed in other organs at lower levels. Per organ, broadly expressed genes are generally underdispersed relative to organ-biased genes (**Figure 5**, **Figures S32-S34**, **Figures S35-S37**, **Figures S38-S40**). Similar results were obtained with a more stringent τ cutoff ($\tau > 0.5$ as organ-biased) (**Figures S41-S43**). These results are consistent with previous studies suggesting that expression variability is associated with both tissue and cell-type-specific expression [49,63].

Since organ-biased genes have likely evolved under selection for an organ-specific function, one prediction is that these genes would have a narrower expression range within their top (focal) organ of expression, indicating more fine-tuned regulation for its putative primary function. On the contrary, we observed significant overdispersion of organ-biased genes within their focal organ (**Figure 5A-C**, **5E-G**, **5I-K; Figures S32-S34**, **Figures S35-S37**, **Figures S38-S40**). For example, genes that are most highly expressed in the pectoral fin are overdispersed within the pectoral fin, relative both to broadly expressed genes as well as to other organ-biased genes expressed in this organ (**Figure 5A, 5E, 5I**). This is consistent with a previous claim that highly variable genes within a cell type population are typically indicative of cell type function [64]. This pattern is lost when the top organ is assigned randomly, such that focal-biased genes have dispersion levels that are indistinguishable from the background of other organ-biased genes (**Figures S45-S47; Figure S48**). We infer that expression of organ-biased genes within their focal organ tends to be highly plastic in response to genetic or environmental cues, and that context-dependent upregulation of expression of organ-biased genes may be mediated by *trans*-acting factors that amplify variation.

Patterns of expression dispersion across organs can reveal the extent to which the inter-individual expression variability of a gene is independently regulated per condition [43]. On the one hand, if control of variability is organ-specific, dispersion should differ across organ contexts. On the other hand, if variability is

controlled by a shared underlying mechanism, dispersion should appear consistent across organs. The second
scenario entails differences in the evolution of regulatory architecture depending on the top organ of a gene.
Interestingly, we observe a consistent pattern across organ contexts that is dependent on the focal organ. For
example, we find that brain-biased genes have consistently underdispersed expression across non-nervous
system organs relative to other organ-biased genes, regardless of differences in cell type composition per
organ (**Figure 5D, 5H, 5L; Figures S32-S34**). (See **Figures S41-S44** for the same analysis given a cutoff of $\tau >$
0.5.) By contrast, liver-biased genes, which evolve under weaker purifying selection at the protein-coding
sequence level (**Figure S49**) [65], tend to be overdispersed across multiple organs (**Figures S35-S37**), implying
weaker global expression constraints. Furthermore, we found that genes preferentially expressed in the
gonads, especially the ovary, tend to be underdispersed across organs (**Figure 5D, 5H, 5L**) (**Figures S38-S40**).
Taken together, these results are consistent with the expected pattern if gene-specific expression variability
is regulated by a shared mechanism across organ contexts [43] that is dependent on the top organ of a gene.
We observed qualitatively similar organ-dependent patterns of expression dispersion after controlling for
differences in transcriptome complexity (i.e., number of genes expressed) across organs [66–68] (**Figure S50**),
with particularly robust results in zebrafish (**Figure S51D**), although the effect is weaker in Northern pike
(**Figure S51H**) and spotted gar (**Figure S51L**). No pattern is recovered in any of the species when the top organ
is assigned randomly (**Figures S52-S54**), confirming that these results are biological and not due to biases in
data generation or processing. These results suggest that organ-biased genes have evolved particular *cis*-
regulatory mechanisms to facilitate either highly stable expression downregulation across other organs (for
brain- and gonads-biased genes) or more variable expression (for liver-biased genes). Indeed, we observed
that zebrafish promoter regions for each set of organ-biased genes can be distinguished from both matched
random promoter sequences (**Figure S55**) and from each other (**Figure S56**) (Methods). Promoter regions from
zebrafish brain-, ovary-, and testis-biased genes yield better model performance, consistent with stronger
promoter sequence conservation for lowly variable genes. By contrast, promoters regions from liver-, pectoral
fin-, gill-, and intestine-biased genes have somewhat lower model performance than sets of random
promoters, suggesting weaker sequence conservation (**Figure S57**).

Suboptimal expression can be more or less deleterious in different organs, due to differences in
physiological function, cellular composition and turnover, and sensitivity to perturbations [69]. In particular,
it has been suggested that neural tissues are highly sensitive to protein misfolding, and that this sensitivity
contributes to strong evolutionary constraints of genes expressed in the brain [70]. We confirm that protein-
coding sequences of brain-biased genes have evolved under stronger purifying selection not only compared
to other organ-biased genes, but also to broadly expressed genes (**Figure S42**), which is consistent with reports
of strong negative selection on the mammalian brain relative to other organs [71–73,65,74]. We observed
moderate rates of purifying selection on both testis- and ovary-biased genes (**Figure S42**), though highly testis-
specific genes have been shown to be faster-evolving compared to other organ-specific genes in mammals
[65]. Overall, our findings show a generally positive association between inter-species protein sequence

divergence and intra-species expression variability (**Figure 4**), suggesting that evolutionary constraints at the
protein sequence level co-vary with organ-dependent patterns of expression variability. One interesting
exception involves eye-biased genes, which evolve under strong purifying selection across the three fish
species (**Figure S49**) but which tend to have overdispersed expression in pike and spotted gar (**Figure 5H, 5L**).
We suggest that the differential variability of eye-biased genes may be an interesting subject for future
research, particularly given the remarkable diversity of visual opsin gene repertoires in fishes and its relation
to differences in light environment and ecology [75].

From these results, we infer that (1) organ-biased genes are highly variable within their focal organ,
likely due to context-sensitive expression, but that (2) their global variability reflects organ-dependent
evolutionary constraints on both protein-coding and *cis*-regulatory sequences. We found that these organ-
dependent patterns of expression variability are largely consistent in all three sampled species, thus
constituting an evolutionarily conserved signal across at least 320 million years of evolution.

Figure 5. Relationship between expression variability and organ-biased expression for (A-D) zebrafish, (E-H) Northern pike, and (I-L) spotted gar. (A) Boxplot of expression variability ranks of genes expressed in female ('F') zebrafish pectoral fin, for sets of genes classified by organ bias and with an emphasis on brain-biased genes. Broadly expressed genes ('broad') have a measured organ expression specificity index (tau) of $\tau \leq 0.3$; genes with $\tau > 0.3$ are considered organ-biased. Among the set of organ-biased genes, genes are further classified into three categories: focal-biased genes ('focal') are preferentially expressed in the focal organ (i.e., pectoral fin), brain-biased genes ('brain') are preferentially expressed in the brain, and the remaining ('other') are preferentially expressed in other organs. The number of genes included are annotated for each category. Comparisons between broadly expressed and focal-biased genes, and between brain-biased and other organ-biased genes were performed using Wilcoxon rank-sum tests (p-value significance levels: '****': $p \leq 0.0001$, '***': $p \leq 0.001$, '**': $p \leq 0.01$, '*': $p \leq 0.05$, 'ns': $p > 0.05$). The dashed line indicates a variability rank of 0.5. **(B)** Boxplot of variability ranks of genes expressed in female zebrafish intestine, with genes classified by organ bias and with an emphasis on liver-biased genes. **(C)** Boxplot of variability ranks of genes expressed in female zebrafish eye, with genes classified by organ bias and with an emphasis on ovary-biased genes. **(D)** Heatmap summarizing pairwise comparisons of expression variability ranks between subsets of organ-biased genes for female ('F', left) and male ('M', right) zebrafish. Rows correspond to focal organ whereas columns correspond to subsets of organ-biased genes. Pairwise comparisons were performed using Wilcoxon rank-sum tests with Benjamini-Hochberg ('BH') correction for multiple testing (adjusted p-value < 0.05). Effect size was computed using Glass's delta (δ), with color scale shown in (L). Cells in gray are not included in the dataset. Three cells are annotated, corresponding to the following comparisons in boxplots (A-C): (A) brain-biased vs. other organ-biased genes expressed in female zebrafish pectoral fin, (B) liver-biased vs. other organ-biased biased genes expressed in female zebrafish intestine, and (C) ovary-biased vs. other organ-biased genes expressed in female zebrafish eye. A negative effect size entails lower expression variability relative to other organ-biased genes (A,C) (turquoise), whereas a positive effect size entails higher variability (B) (salmon). **(E-H)** Corresponding plots for Northern pike, with cells annotated corresponding to comparisons in boxplots (E-G). **(I-L)** Corresponding plots for unsexed spotted gar, with cells annotated corresponding to comparisons in boxplots (I-K).

Discussion

Until recently, most gene expression studies have focused on testing for differences in average expression levels between groups of interest, largely to the exclusion of expression variance and variability. Yet a more complex picture is now emerging: gene expression variability is increasingly implicated in disease [36–39] and drug efficacy [40], aging [41], circadian rhythm [76], cellular differentiation [32–34], embryogenesis [35,77], developmental canalization [47], and other biological processes. Here, we apply an evolutionary perspective based on a comparative analysis of inter-individual gene expression variability across multiple adult organs in three species of ray-finned fishes. We confirm that genes which are lowly variable among individuals are typically enriched in housekeeping functions, whereas highly variable genes are enriched in stimulus-sensitive functions, as reported previously in other species [25,42,43,63]. This biologically intuitive result reaffirms that expression variability is a functionally relevant trait independent of average expression. There is a growing body of literature that recognizes the functional relevance of expression variability [44–46] and how this variability is associated with promoter characteristics such as sequence, core promoter motifs, and transcriptional start site architecture [25,63,78]. However, there have been few attempts to characterize how gene expression variability in animals is systematically modulated by the interplay between gene function and organ context, though some studies have sampled diverse organs [40,43]. To our knowledge, this is one of the first studies to investigate how gene expression variability is influenced both by expression specificity and organ context.

Our cross-organ approach helps clarify the connection between two previously reported findings. First, although lowly and highly variable genes show distinct functional signatures, some studies report that gene expression variability tends to be tissue-specific (i.e., organ-specific) [43], with little overlap in the set of

most overdispersed genes between organs [41,42] or cell types [64] (but see [25]). Second, organ- or cell-type-

specific genes are found to have highly variable expression [49,63,64]. We verify that organ-biased genes are

more variably expressed relative to broadly expressed genes, but in addition, we have also found that organ-

biased genes are typically highly variable within their primary (focal) organ. Since the most variable genes

within an organ are typically genes which are biased or specific to it, we think it is unsurprising to find few

highly variable genes shared between multiple conditions. We infer that plastic expression within a focal organ

is a functional feature of more specialized genes, and that expression is dynamically induced by context-

specific regulatory cues, such as organ-specific transcription factor binding to enhancer elements or enhancer-

promoter interactions. Promoters of specialized genes (e.g., developmental genes) have been shown to

interact with different classes of enhancers compared to those of housekeeping genes [79]. In addition, such

specialized genes tend to have more enhancer-promoter interactions than housekeeping genes [28,80], and

the number of interacting enhancers has been found to have an additive effect on gene expression levels

[28,81,82]. Thus, to augment transcription of organ-biased genes in the relevant context, we expect an

increase in regulatory inputs, which introduces more sources of expression variation [11]. Since gene-specific

expression variability is influenced by a gene's regulatory inputs [11] and its position in the gene regulatory

network [83], we expect subsets of genes to have co-regulated patterns of variability, analogous to co-

regulation of gene expression levels. An analysis of the co-regulation of gene expression variability would

require even larger sample sizes in order to estimate variability across clusters of individuals (e.g., from

different populations).

Outside of their focal organ, organ-biased genes have basal expression levels that are relatively more

stable among individuals. We find striking evidence across species that the degree to which biased genes are

stably downregulated across non-focal organs reflects organ-dependent evolutionary constraints. Notably, we

observe consistently underdispersed expression of brain-biased genes across non-nervous organs, despite

these organs having very different transcriptional repertoires, cell-type compositions, and local micro-

environments. Low variability within a condition suggests stabilizing selection on expression, which is

consistent with strong purifying selection on *cis*-regulatory sequences, and/or with buffering of expression

variation by compensatory *cis-trans* effects [84]. Given that we observe consistent underdispersion across

multiple organs having very different local micro-environments, we reason that stable basal expression of

brain-biased genes is likely to be predominantly driven by strong purifying selection at the *cis*-regulatory level

rather than by context-dependent *trans* effects. Our finding that organ-biased genes have distinguishable

promoter sequences (**Figure S55**) further supports the view that inter-individual variability across organ

contexts is at least partly under *cis*-regulatory control. Moreover, the results support stronger selection on

promoters that regulate lower variability in non-focal organs.

We recognize that comparing gene expression variability between organs is challenging because any

measured changes in dispersion may simply be due to technical variation (e.g., sampling design, sequencing

depth or RNA quality [40]) or 'extrinsic' sources of biological variation (e.g., cell type composition [49,50],

transcriptome complexity [66,67]). Indeed, a limitation of bulk RNA-seq is that it does not allow us to distinguish between multiple sources of variation within an organ, whether it is expression variation within a cell type population, expression variation between cell type populations, or heterogeneity in cell type composition between individuals. Observing a consistent signal of underdispersion across conditions with heterogeneous backgrounds is thus particularly informative and suggests that our method captures a signal of stable *cis*-mediated downregulation that is independent of organ-context. We infer that brain-biased genes have evolved under strong purifying selection, not only at the protein-coding sequence level, but also at the *cis*-regulatory level. From this, we hypothesize that genes selected for function in the brain may have co-evolved with more stringent 'enhancer grammar' (e.g., number, type, order, affinity, spacing, or orientation of transcription factor binding sites) [85,86] relative to other organ-biased genes. Our results highlight how analysis of gene expression variability across organs can yield additional insights into gene regulatory strategies in animals. Extending on these findings, we propose that intra-specific expression variability can potentially be informative of broader macro-evolutionary trends, such as patterns of duplicate gene retention and functional divergence after whole-genome duplication.

Materials and Methods

Sample preparation

Sample collection

<u>Zebrafish</u>

Thirteen female and twelve male zebrafish (*Danio rerio*), strain AB, were raised at the University of Oregon Aquatic Animal Care Services. Animals were raised in E5 medium [87] and fed rotifers from 4-10dpf [88]. After 10dpf, fish were moved to a circulating fish water system and fed Zeigler Adult Zebrafish Diet. Three-month old adults were euthanized by cold shock [87], and organs were dissected and immediately stored in RNAlater. Procedures were approved by the University of Oregon IACUC protocol #18-13.

<u>Northern pike</u>

Ten female and ten male one-year-old Northern pike (*Esox lucius*) were sourced from a commercial hatchery (Aqua 2B, Pouancé, France). The fish were euthanized using a lethal dose of tricaine methanesulfonate (MS 222, 400 mg/L), supplemented with 150 mg/L of sodium bicarbonate. Following euthanasia, organs were immediately dissected and preserved in RNAlater. All animal procedures were conducted in full compliance with French and European legislation (French decree 2013-118 and European Directive 2010-63-UE), governing the ethical use and care of laboratory animals in scientific research.

<u>Spotted gar</u>

Spotted gars (*Lepisosteus oculatus*) were obtained as embryos from hormone-induced spawns of wild-caught broodstock from bayous near Thibodaux, Louisiana and raised at Michigan State University for 23 or 31 months

(April 2018 - March/November 2020) in 150 to 300 gallon tanks on a 12h light : 12h dark cycle and a diet starting from *Artemia* brine shrimp to feeder fish (mostly fathead minnows). Twenty-two (22) unsexed, pre-mature individuals of 20.1 - 23.3 cm standard length (tip of snout to caudal peduncle) and 24.8 - 29.0cm total length (tip of snout to tip of caudal fin) were euthanized in 300 mg/L MS-222 (Sigma). Tissues were dissected and stored in RNAlater (Invitrogen) at -80℃ until shipped frozen to the University of Lausanne. All spotted gar experiments were approved by the Institutional Animal Care & Use Committee (IACUC) of Michigan State University (animal use protocol PROTO201900309).

RNA extraction

Tissue samples were thawed on ice, removed from RNA*later* RNA Stabilization Solution (Invitrogen, catalog no. AM7024) and transferred to new Eppendorf tubes containing 1 mL TRIzol Reagent (Invitrogen, ref. 15596018) and about 10 - 15 SiLi beads (Type ZS, 1.4 - 1.6 mm) (Sigmund Lindner, ref. 9315-33). Excluding blood, samples were homogenized at 4°C using PreCellys Evolution (Bertin Technologies, ref. P000062 _PEV0) for at least 2 cycles at 6600 rpm for 30 s. Phase separation of homogenized samples was performed following the TRIzol Reagent user guide (MAN0001271). For Northern pike and spotted gar, RNA extraction with TRIzol Reagent was carried out following the same manual. After washing and resuspending RNA, samples were treated with RNase-free DNase I (Thermo Scientific, catalog no. EN0525) to remove DNA in accordance with the user guide (MAN0012000). For zebrafish, column-based RNA extraction with TRIzol was performed using the PureLink RNA Mini Kit (Invitrogen, ref. 12183018A), following manufacturer's instructions (MAN0000406). DNA digestion during on-column RNA purification was done using the PureLink DNase Set (Invitrogen, ref. 12185010). All samples were eluted in 50 - 60 µl RNase-free water. Following extraction, RNA concentration was quantified with the QuantiFluor RNA System (Promega, ref. E3310). A subset of samples was also assessed for RNA quality using the 5200 Fragment Analyzer System (Agilent Technologies, M5310AA) with PROSize 3.0.1.6 data analysis software (Agilent Technologies). Samples with low RNA concentration or signs of degradation were removed from subsequent steps. RNA preparations were diluted to a minimum concentration of 50 ng/µl, transferred to low-binding 96-well plates, and stored in -80°C prior to sequencing.

Data generation

Bulk RNA barcoding and sequencing (BRB-seq)

Samples were sent to Alithea Genomics (Lausanne, CH) for library preparation and sequencing using highly multiplexed 3'-end bulk RNA barcoding and sequencing (BRB-seq) [51] over three phases (corresponding to projects AG0012, AMP0027, and AMP0020) (**Table S1**) (Supplementary Data). BRB-seq libraries were prepared using the MERCURIUS BRB-seq library preparation kit for Illumina and following the manufacturer's manual (Alithea Genomics, #18013). For project AG0012 (Northern pike, spotted gar), libraries were sequenced using Illumina NovaSeq 6000, generating multiplexed, asymmetric paired-end reads, where read R1 (25 bp) contains the sample barcode and UMI and read R2 (80 bp) contains the cDNA fragment sequence to be aligned to the

reference genome. For projects AMP0027 (spotted gar) and AMP0020 (zebrafish), libraries were paired-end

sequenced using Illumina NextSeq 550 (R1 = 21 bp, R2 = 55 bp).

Read alignment and quantification

Read trimming, alignment, and quantification steps were performed by Alithea Genomics using STARsolo

(preprint) [89] from STAR 2.7.9a [90]. R2 reads were trimmed ("--clipAdapterType CellRanger4") and aligned

to the following reference genomes: for spotted gar, LepOcu1 (Ensembl 103 [91]); zebrafish, GRCz11 (Ensembl

101 [91]); Northern pike, Eluc_v4 (Ensembl 101). To generate both raw and UMI-deduplicated counts, the

following parameter was used: "--soloUMIdedup NoDedup 1MM_Directional". UMI-deduplicated count

matrices were used for downstream analyses.

Data analysis

Excluding the promoter sequence analysis, data analyses were performed using R 4.1.0 [92], Bioconductor

3.13 [93], and RStudio 2023.03.0+386 [94] unless specified otherwise.

Sample filtering

Sample libraries were filtered in two main steps independently for each species. First, samples with low

sequencing quality were removed based on a minimum number of uniquely mapped reads or number of genes

with nonzero counts ("detected genes"), following Liu et al. [35]. For example, in zebrafish, samples with less

than 500,000 uniquely mapped reads or less than 10,000 detected genes were excluded from downstream

analysis. To account for differences in reference genome annotation quality, we selected species-specific

minimum uniquely mapped reads and detected genes. (See **Table S2** for filtering criteria.) For each species,

we visually inspected the plot of detected genes against uniquely mapped reads to estimate the point in which

the relationship between the two variables becomes nonlinear and set this point as the minimum threshold

(**Figure S1A**, **Figure S2A**, **Figure S3A**). We verified that samples which do not meet specified minimum criteria

are visible outliers on a principal component analysis (PCA) (**Figure S1B-C**, **Figure S2B-C**, **Figure S3B-C**).

Second, samples that were poorly correlated with other samples within its annotated organ were

removed. We performed iterative removal of lowly correlated samples based on a protocol described by

Fukushima & Pollock [95]. For each iteration, we first generated the all-against-all pairwise Pearson correlation

matrix of retained samples based on gene-filtered and normalized expression data in the form of log-

transformed, TMM-normalized counts per million ($\log_2$ TMM-CPM) (Gene filtering and normalization). Then,

for each sample, we computed its mean Pearson correlation with each organ. We removed any sample which

is more highly correlated with another organ than the one in which it is annotated to. This process was

repeated until no more samples fit removal criteria. Considering the set of protein-coding genes in each

species, we verified that all retained samples cluster by organ based on hierarchical clustering of the pairwise

Pearson correlation matrix (**Figure 1B, Figure S10**) and PCA of gene-filtered and normalized expression (**Figures S7-S9**). Gene biotype information was retrieved for each species from Ensembl 105 [96].

For computing expression specificity (Organ expression specificity), we required a minimum of 1 replicate per organ-sex for Northern pike and zebrafish, and 2 replicates per organ for spotted gar (**Table S2**). For estimating expression variability (Gene expression variability) and testing its association with gene function (Gene Ontology (GO) statistical over- and under-representation analysis) and coding-sequence selection (Selection analysis), we required a minimum of 4 replicates per organ (for spotted gar) or organ-sex condition. Any conditions which did not meet the minimum number of replicates after quality filtering were excluded from downstream analysis (**Table S3**). See **Tables S2**-**S3** for samples retained after each step of the sample filtering pipeline and **Figure 1A** for a complete breakdown of the number of replicates retained per condition.

Gene filtering and normalization

At the start of each step of the sample filtering pipeline (Sample Filtering), we removed non-expressed genes from the full raw count matrix by retaining only genes with both mean and median counts per million (CPM) greater than 1.0 in at least one organ. From the gene-filtered raw count matrix, we performed library size normalization over all conditions using the trimmed mean of M-values (TMM) method in *edgeR* (3.34.1) [97], then re-converted counts to CPM and transformed CPM to $\log_2$-scale ($\log_2$ TMM-CPM). To remove negative values, $\log_2$ TMM-CPM was vertically offset by the absolute value of the global expression minimum such that the minimum value in the normalized expression matrix is always 0.

Gene expression variability estimation

To estimate gene expression variability within an organ (for spotted gar) or organ-sex condition (for Northern pike and zebrafish), we considered only the set of conditions with at least 4 replicates after sample filtering. Gene filtering and normalization steps were performed as described in the previous section, but with each condition treated separately. As an additional filter to remove lowly expressed genes specific to a condition, we removed genes with mean or median $\log_2$ TMM-CPM < 1.0. Since actual expression variation may be sensitive to outliers, we used a jackknife resampling ('leave-one-out') procedure to compute expression dispersion statistics and generate a less biased estimate of gene expression variability. That is, for a given condition with n replicates, we formed all n possible subsets of n - 1 replicates and computed expression dispersion metrics for each subset, then computed the mean for each statistic over all subsets. We initially considered different measures of expression dispersion (Supplementary Methods), with the aim to correct for the strong covariation between expression variation and average expression (**Figure S11**). From the metrics we considered, we selected the local (residual) coefficient of variation (LCV) [29,40] as our estimate of expression variability, as it maximally removes the association between mean expression level and expression dispersion (**Figure 2C; Figures S12**-**S14**). The LCV was computed in two steps then averaged over all n - 1 subsets, as described below.

<u>Residual coefficient of variation</u>

The coefficient of variation (CV) is the ratio of the standard deviation to the mean. First, we computed the squared CV and applied a $\log_2$ transformation ($\log_2(CV)^2$) for each gene, following Faure et al. [29]. Next, using the 'loess' function from the *stats* package in R, we fit a LOESS regression curve (with span = 0.6 and default parameters) to model $\log_2(CV)^2$ as a function of mean expression level. The difference between observed and predicted $\log_2(CV)^2$ is the residual $\log_2(CV)^2$ ('residual CV', for brevity) after adjusting for expression level. Genes with residual CV > 0 overdispersed, while genes genes with residual CV < 0 are underdispersed.

<u>Local (residual) coefficient of variation (LCV)</u>

Given residual CVs for each gene, we performed a percentile rank normalization step following the 'local coefficient of variation' [40,47] or 'relative noise rank' [29] metric recommended in previous publications. First, genes were sorted by mean expression level. Next, we defined a sliding window of 100 neighboring genes and a step size of 1, with a focal gene at the center of each window. At each iteration of the sliding window, the residual CV of the focal gene was ranked against the residual CVs of genes within the window, on a percentile rank scale of 0.00 (lowest dispersion) to 1.00 (highest dispersion). We computed the LCV statistic only for complete windows; as such, for each *n* - 1 subset, we do not have variability estimates for the 50 most lowly and 50 most highly expressed genes.

<u>Summarizing jackknife results</u>

Since transcriptome-level expression distributions are positively skewed, with the most highly expressed genes spanning a wide expression range, expression variability tends to be poorly estimated for outlier genes with very high expression. As such, for each condition, we removed the top 5% most highly expressed genes based on jackknifed mean expression level. Then, we computed the mean LCV over all *n* - 1 subsets and used this metric, which we call 'variability rank', as our estimate of expression variability.

Gene Ontology (GO) statistical over- and under-representation analysis

For each organ-sex condition in zebrafish, we retrieved three sets of genes categorized by expression variability: genes with variability rank ≥ 0.80 (highly variable), genes with rank ≤ 0.20 (lowly variable), and genes with 0.40 ≤ rank ≤ 0.60 (moderately variable). For each variability category per condition, we used *GOstats* (2.58.0) [56] to conduct conditional hypergeometric tests for statistical over- and under-representation of GO biological process terms, based on the org.Dr.eg.db (3.13.0) zebrafish genome annotation [98]. Per test, the background gene set was defined as the set of genes which have expression variability estimates within the organ-sex condition. For each variability category, we retained GO terms that were significantly over-represented (or under-represented) (unadjusted *p*-value < 0.01) in at least three

conditions. Redundancy reduction and visualization of GO terms was done using the Python package *GO-*

Figure! (1.0.0) [57] (**Figure 3, Figures S25**-**S26**).

Coding-sequence selection analysis

Given a list of protein-coding genes (Ensembl 105 [96]) and taxonomic branches for each species, we retrieved

precomputed selection statistics from Selectome (https://selectome.org/), a database of evidence for positive

selection based on a conservative branch-site likelihood test applied to *Euteleostomi* gene trees (Ensembl 98)

[58,59]. For each gene - internal branch (i.e., taxonomic clade), the Selectome dataset provides a measure of

selective pressure for sites under purifying selection (ω_0 ratio for sites with dN/dS < 1), as well as a likelihood-

ratio test statistic for positive selection. The tested internal branches for each species are as follows. For

spotted gar: Euteleostomi, Actinopterygii, Neopterygii. For zebrafish: Euteleostomi, Actinopterygii,

Neopterygii, Osteoglossocephalai, Clupeocephala, Otomorpha, Otophysi. For Northern pike: Euteleostomi,

Actinopterygii, Neopterygii, Osteoglossocephalai, Clupeocephala, Euteleosteomorpha, Protacanthopterygii

(**Figure 4A**). We included only distinct gene - internal branch entries in our downstream analysis, as duplicate

gene - internal branch entries may be a result of uncertainties in the gene tree. For each species, the

background gene set consisted of protein-coding genes that are included in the Selectome dataset and that

have expression variability data in at least half of the available conditions.

First, to test for an association between expression variability and purifying selection, we grouped

genes into 10 bins of size 0.1 based on their median variability rank computed across available conditions (per

species) and compared the ω_0 distribution across bins (**Figure 4**; **Figure S27**). Second, to check for an

association between variability and positive selection, we computed the proportion of gene - internal

branches with evidence for positive selection based on a *q*-value cutoff of *q* < 0.05 on the branch-site likelihood

test (**Table S4**). For this, genes that were frequently lowly, moderately, or highly variable (i.e., the same

classification in at least half of available conditions) were grouped into three large categories based on the

following thresholds: rank ≤ 0.20 for lowly variable genes, rank ≥ 0.80 for highly variable genes, and 0.20 <

rank < 0.80 for moderately variable genes. To test for significance, we conducted a randomization test by

permuting the *q*-values on the background gene list (*n* = 2000 permutations), in addition to a Pearson's chi-

squared test.

Organ expression specificity

Data preprocessing

To compute organ expression specificity within a species, we required at least 2 individuals per organ for

spotted gar, and at least 1 individual per organ-sex condition for Northern pike and zebrafish. Furthermore, to

make expression specificity measurements comparable between species, we considered only the set of organs

with sufficient samples in every species, i.e., brain, eye, intestine, liver, heart, muscle, gills, pectoral fin, and

gonads. After excluding organs with missing data (blood, kidney, skin, swim bladder), we repeated our gene

filtering and normalization step starting from raw read counts (Gene filtering and normalization).

<u>Expression specificity index</u>

For each retained gene per species, organ expression specificity was measured using the tau (τ) index, which

is defined as:

$$\tau = \left(\frac{1}{n-1}\right) \sum_{i=1}^{n} \left(1 - \frac{\bar{x}_i}{\bar{x}_{max}}\right)$$

where n is the number of organs, $\bar{x}_i$ is the mean expression of gene x in organ i, and $\bar{x}_{max}$ is the maximum

mean expression of gene x within an organ $1 \le i \le n$ [61,62]. The tau index gives a single measure of organ

specificity per gene, ranging from broadly expressed across all organs ($\tau \to 0$) to specifically expressed in a

single organ ($\tau \to 1$). For species with differentiated sexes (Northern pike and zebrafish), the ovary and testis

were considered separate anatomical entities. However, for non-reproductive organs in the same species, we

did not consider each sex separately.

Organ-biased expression and gene expression variability

Per species, genes with tau (τ) ≤ 0.3 were classified as broadly expressed, while genes with tau above this

threshold were classified as organ-biased. (See **Figures S41**-**S44** for an alternative cutoff of $\tau \le 0.5$.) For genes

which are organ-biased, the 'focal organ' is defined as the organ in which the gene is maximally expressed.

<u>Multiple pairwise comparisons</u>

For each species (spotted gar) or species-sex (Northern pike and zebrafish), let S_1 be the set of m organs which

(a) have been used in expression specificity computations, and (b) have gene expression variability estimates

(**Figure 1**):

$$S_1 = \{1, 2, \ldots, m\}.$$

650 Let S_2 be the set of $n \ge m$ organs which satisfy (a) but not (b):

$$S_2 = \{1, 2, \ldots, m, \ldots, n\}.$$

652 For each organ $k \in S_1$, we compared the variability rank distribution within k for broadly expressed genes (τ

653 ≤ 0.3) versus genes which are organ-biased ($\tau > 0.3$) and with top expression in focal organ k. For example,

654 given the pectoral fin as the focal organ, we compared the variability ranks of broadly expressed genes against

655 pectoral-fin-biased genes (**Figure 5A, 5E, 5I**).

656 Next, still given k, we iterated through the set of other (non-focal) organs in S_1 which are not k:

$$\{i \in S_1 \mid i \neq k\}.$$

658 For each iteration of $i \neq k$, we defined a set of i-biased genes, which are organ-biased and with top expression

659 in non-focal organ i, but still expressed in focal organ k. Then, we compared the variability rank distribution

within k for genes which are i-biased versus other organ-biased genes with top expression in the set J of (non-focal) organs in S_2 which are neither k nor i:

$$J = \{j \in S_2 \mid j \neq k, j \neq i\}.$$

As an example, for expression variability in the pectoral fin, we compared the variability ranks of brain-biased (or liver-biased, ovary-biased, etc.) genes expressed in the pectoral fin against the variability ranks of the set of other organ-biased genes that are neither pectoral-fin-biased nor brain-biased (or liver-biased, ovary-biased, etc.), and are also expressed in the pectoral fin (**Figure 5A, 5E, 5I**).

For each pairwise comparison of variability ranks of i-biased versus (other-biased) J-biased genes, iterated for each organ $k \in S_1$, we tested for a difference in group means using Wilcoxon-rank sum tests, and applied a Benjamini-Hochberg ('BH') correction for multiple testing. In addition, we calculated the effect size using Glass's δ, using the set of J-biased genes as the reference group:

$$\delta_{k,i} = \frac{\bar{v}_{k,i} - \bar{v}_{k,J}}{s_{k,J}}$$

where $\bar{v}_{k,i}$ is the mean variability rank of i-biased genes in focal organ k, $\bar{v}_{k,J}$ is the mean variability rank of J-biased genes, and $s_{k,J}$ is the sample standard deviation of the variability ranks of J-biased genes. We visualized the distribution of effect sizes using a heatmap, where rows correspond to focal organ k and the columns correspond to sets of i-biased genes (**Figure 5D, 5H, 5L**). We showed pairwise comparisons that are considered statistically significant after BH multiple testing correction (adjusted p-value < 0.05).

Analysis on randomized data

Given the gene x organ mean expression matrix of a species, we permuted the values of the mean expression vector per gene so that the focal organ of each gene is assigned randomly without changing tau. Since tau is unchanged, the same set of genes are classified as broadly expressed ($\tau \leq 0.3$) and organ-biased, but the subclassifications of organ-biased genes (e.g., brain-biased, liver-biased, etc.) differ. We pre-set two seeds (12345, 67890) for our sample randomizations. Multiple pairwise comparisons of expression variability ranks for subsets of organ-biased genes were performed as previously described (**Figure S48**).

Analysis on nonzero expression matrix

To control for differences in number of genes expressed among organs (**Figure S50**), we repeated the same analysis (i.e., gene expression normalization, expression variability estimation, comparisons of expression variability between subsets of genes) as previously described, considering only genes that have nonzero counts across all samples within a species (**Figure S51**).

Organ-biased expression and coding-sequence selection

Per species, we retrieved expression specificity measurements for each gene (Organ expression specificity)

and filtered selection statistics for each gene - internal branch (Coding-sequence selection analysis), as

described in previous sections. We annotated each gene from the Selectome dataset with its top (focal) organ

of expression (i.e., organ bias). If a gene has $\tau \leq 0.3$, its organ bias is labeled 'broad'. We compared the ω_0

distribution of genes across organ bias categories to check for any association between focal organ and

strength of purifying selection at the coding-sequence level (**Figure S49**).

Promoter sequence analysis

To test whether organ-biased promoters can be distinguished from random sequences and from each other

based on variability patterns, we trained gapped k-mer SVM models (gkmSVM) on promoter sequences.

Analyses were first performed on each organ-biased promoter set, using subsamples of up to 1,000 genes to

avoid sample-size bias **(Figure S55)**. Organ-biased promoter sets were also tested against promoters from the

other organ-biased sets to evaluate whether they could be distinguished from one another **(Figure S49)**. We

then grouped zebrafish organ-biased genes into two categories: those typically lowly variable in other organs

(brain, ovary, testis, eye) and those highly variable in the others (liver, pectoral fin, gills, intestine), with up to

5,000 genes per set. In addition, three random gene sets were generated **(Figure S57)**. Promoter regions were

defined as the 2-kb sequence upstream of the canonical TSS (Ensembl Biomart release 114) [99], and exonic

regions were removed with bedtools subtract [100].

For each promoter set, a matched negative set of genomic sequences was drawn from the DanRer11

assembly, matched for length, repeat content, and GC content. This was done by building a BSgenome object

with UCSC annotations [101] and sampling with the genNullSeqs function from the gkmSVM R package [102].

Models were trained with gkmtrain (LS-GKM) [103] using default parameters, and performance was assessed

by 5-fold cross-validation using the -x option and receiver operating characteristic (ROC) curves.

Author Contributions

M.R.-R., I.B., Y.G., and J.H.P. conceived the experimental design. C.F.B. and M.R.-R. conceived the design of

data analysis. A.W.T., B.L.R., and I.B. sourced and dissected the spotted gar samples; C.A.W. and J.H.P., the

zebrafish samples; J.B. and Y.G., the Northern pike samples. C.A. performed RNA extraction. S.M. provided

Selectome data for positive selection analysis. A.L. and A.T. performed the promoter sequence analysis. C.F.B.

performed the data analysis and generated the figures with input from M.R.-R. C.F.B. wrote the original draft

of the manuscript with input from M.R.-R; C.F.B., A.L., A.W.T., J.B., I.B., Y.G., J.H.P., and M.R.-R wrote the final

version of the manuscript. C.F.B., A.L., A.W.T., J.B., I.B., Y.G., J.H.P., and M.R.-R. contributed to the

interpretation and discussion of results.

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Competing Interests

The authors declare no competing interests.

Data Availability

Demultiplexed FASTQ files generated via BRB-seq for zebrafish (192 samples), Northern pike (252 samples), and spotted gar (177 samples) are available on the NCBI Sequence Read Archive (SRA) under accession number PRJNA1195359.

Data analysis scripts are available on GitHub: https://github.com/christabel-bucao/fish-variability-across-organs

Additionally, scripts and supplementary data files are available on Zenodo: https://doi.org/10.5281/zenodo.14063787

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