Lifespan Normative Modeling of Brain Microstructure

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Abstract

Normative models of brain metrics based on large populations could be extremely valuable for detecting brain abnormalities in patients with a variety of disorders, including degenerative, psychiatric and neurodevelopmental conditions, but no such models exist for the brain's white matter (WM) microstructure. Here we present the first large-scale normative model of brain WM microstructure– based on 19 international diffusion MRI datasets covering almost the entire lifespan (totaling N=54,583 individuals; age: 4-91 years). We extracted regional diffusion tensor imaging (DTI) metrics using a standardized analysis and quality control protocol and used hierarchical Bayesian regression (HBR) to model the statistical distribution of derived WM metrics as a function of age and sex. We extracted the average lifespan trajectories and corresponding centile curves for each WM region. We illustrate the utility of the method by applying it to detect and visualize profiles of WM microstructural deviations in a variety of contexts: in mild cognitive impairment, Alzheimer's disease, and 22q11.2 deletion syndrome – a neurogenetic condition that markedly increases risk for schizophrenia. The resulting large-scale model provides a common reference to identify disease effects on the brain's microstructure in

[1] Data used in preparation of this article were obtained from the Alzheimer's Disease Neuroimaging Initiative (ADNI) database (adni.loni.usc.edu). As such, the investigators within the ADNI contributed to the design and implementation of ADNI and/or provided data but most of them did not participate in analysis or writing of this report. A complete listing of ADNI investigators may be found at: http://adni.loni.usc.edu/wp-content/uploads/how_to_apply/ADNI_Acknowledgement_List.pdf

individuals or groups, and to compare disorders, and discover factors affecting WM abnormalities. The derived normative models are a new resource publicly available to the community, adaptable and extendable to future datasets as the available data expands.

1. Introduction

Large-scale international initiatives have greatly increased the availability of brain imaging data, driving the development of advanced statistical tools such as normative[1] and generative models[2] to study brain measures across life, health, and disease. Normative modeling (NM) is a statistical technique that estimates the normative distribution, i.e., the centiles of variation, of a biological measure in a reference population, given explanatory or clinical variables such as age and sex. The resulting model can be used to gauge and track brain abnormalities and factors that influence them. NM is sensitive to individual differences and offers multivariate metrics of abnormality, extending the standard case-control maps of group differences. This has sparked broader interest in NM for comprehensive studies of neurological and psychiatric conditions, and to support personalized medicine[3].

NM has previously been applied to anatomical brain scans to create lifespan trajectories for structural brain metrics. Rutherford et al.[4] used the Predictive Clinical Neuroscience toolkit[1] (https://pcntoolkit.readthedocs.io) to chart the lifespan trajectories for a range of structural brain measures, available at the PCNPortal[5] (https://pcnportal.dccn.nl/). Bethlehem et al.[6] aggregated structural MRI scans from over 100 studies, creating lifespan 'brain charts' from 101,457 human participants from birth to 100 years of age. Ge et al.[7] applied algorithms including multivariable fractional polynomial regression, warped Bayesian linear regression and hierarchical Bayesian regression to model brain morphometry in 37,407 healthy individuals from 86 international sites as part of the ENIGMA-Lifespan project, creating CentileBrain (https://centilebrain.org). Similar approaches have been extended to infant brain data, most recently by the ENIGMA Origins group[8].

So far, we lack lifespan normative models of the brain's white matter (WM) microstructure, which is sensitive to complex maturation, aging, and disease processes not detectable with standard structural MRI. Diffusion magnetic resonance imaging (dMRI)[9,10] probes the local microstructural environment of brain tissue, yielding a rich set of metrics sensitive to a broad range of brain diseases[11]. Creating a universal normative model for dMRI metrics is challenging due to the influence of scanning protocol parameters, such as voxel size, the number of diffusion gradient directions, *b*-values (diffusion weightings), and field strengths[12–15]. Here we address this challenge using hierarchical Bayesian regression (HBR)[16], an approach designed to model site, scanner and cohort effects in multi-site data. To incorporate sufficient dMRI data for NM, we pooled dMRI data from diverse international studies, accommodating data collected with different acquisition protocols. As some datasets focus on development and others on aging, we merged data from different age ranges to build a normative model covering the lifespan (4-91 years).

We built normative models for the most widely used dMRI brain metrics - fractional anisotropy (FA), and mean, radial, and axial diffusivity (MD/RD/AD) - based on 54,583 individuals scanned in 19 public neuroimaging studies. We then leveraged the estimated normative models for 21 deep WM regions to identify WM abnormalities in two age-related conditions: dementia, mild cognitive impairment (MCI) for senescence phase and a neurogenetic disorder, 22q11.2 deletion syndrome (22qDel) for the

development phase. In this anomaly detection scenario, we first show how the deviation scores from the reference population norm can identify abnormal microstructure in mild cognitive impairment (MCI), dementia, and 22qDel.

Lastly, to better understand the lifespan charts, we tested two hypotheses proposed by the theory of *retrogenesis*[17,18]. First, we tested the so-called "*last-in first-out"* hypothesis, which suggests that the early-maturing cerebral structures are somewhat spared from neurodegeneration in old age and tend to degenerate last in late life, relative to late-maturing structures. By contrast, the "*gain-predicts-loss"* hypothesis argues that the neural tissues tend to show correlations between their rate of tissue maturation in early life and degeneration in later life, with fast developing structures showing the most rapid decline.

2. Results

Table 1.

List of full names for each of the regions of interest of the white matter according to the JHU-ENIGMA-DTI template space. There are 21 white matter regions and the "Global-WM" (including all WM in a single metric for a total of 22 total regions of interest included in the analyses.

Acronym	Region of Interest (ROI)
Global-WM	Mean across the entire ENIGMA-DTI white matter skeleton.
ACR	Anterior corona radiata
ALIC	Anterior limb of internal capsule
BCC	Body of corpus callosum
CGC	Cingulum of the cingulate gyrus
CGH	Cingulum of the hippocampus
CST	Cortico-spinal tract
EC	External/extreme capsule
FX	Fornix
FXST	Crus of the fornix/stria terminalis
GCC	Genu of the corpus callosum
PCR	Posterior corona radiata
PLIC	Posterior limb of internal capsule
PTR	Posterior thalamic radiation
RLIC	Retrolenticular part of internal capsule
SCC	Splenium of the corpus callosum
SCR	Superior corona radiata
SFO	Superior fronto-occipital fasciculus
SLF	Superior longitudinal fasciculus

SS	Sagittal stratum
TAP	Tapetum of the corpus callosum
UNC	Uncinate fasciculus

Table 2.

Demographics (age and sex) of the datasets used for hierarchical Bayesian regression (HBR) model estimation and model adaptation (NIMHANS and UCLA).

Study/Dataset	Controls/ Cases	Age Range	Median Age (SD)	N (Males %)	Country
ABCD	-	8 - 13	11.2 (1.1)	8669 (52)	United States
AOMIC	-	18 - 26	22.5 (1.7)	1267 (46)	The Netherlands
CAMCAN	-	18 - 88	55 (18.5)	629 (49)	United Kingdom
CHBMP	-	18 - 46	31 (7.7)	139 (89)	Cuba
CHCP	-	18 - 71	23 (17.7)	302 (44)	China
HCP-A	-	36 - 89	56.3 (13.9)	638 (42)	United States
HCP-D	-	8 - 21	13.4 (3.6)	580 (48)	United States
HCP-YA	-	22 - 36	29 (3.6)	1025 (45)	United States
PedsDTI	-	4 - 21	12.2 (4.2)	233 (48)	United States
PING	-	3 - 21	12.3 (4.8)	852 (52)	United States
PNC	-	8 - 22	15.2 (3.3)	839 (45)	United States
QTAB	-	9 - 14	11 (1.2)	392 (51)	Australia
QTIM	-	15 - 30	21.8 (2.9)	756 (37)	Australia
SLIM	-	17 - 24	20 (1.0)	424 (44)	China
UKBB	-	46 - 81	64.6 (7.5)	36864 (46)	United Kingdom
ADNI3	Controls	55 - 95	71.9 (7.5)	447 (41)	United States
	MCI	55 - 93	75.4 (8.0)	214 (57)	
	Dementia	55 - 95	78 (8.8)	69 (59)	
HBN	Controls	5 - 19	10.3 (3.6)	211 (53)	United States
OASIS3	Controls	42 - 92	71.4 (8.4)	232 (48)	United States
	MCI	62 - 87	76.8 (7.5)	11 (36)	
	Dementia	65 - 82	76.1 (5.1)	12 (41)	
PPMI	Controls	30 - 80	63.1 (10.5)	84 (63)	United States
NIMHANS	Controls	51 - 81	65.0 (7.5)	118 (50)	India
	MCI	52 - 88	68.0 (7.4)	89 (65)	

	Dementia	51 - 86	70.5 (7.4)	90 (58)	
UCLA	Controls	7 - 31	14.2 (5.2)	67 (47)	United States
	22q11.2 Del. S.	6 - 27	14.0 (6.1)	80 (46)	

Footnote: The N for each group reflects the participants after quality control and removal of outliers, as explained in the methods section. The hyphen (-) in the Controls/Cases column indicates that the study was population-based. (SD=standard deviation).

2.1. Lifespan trajectories of brain microstructural metrics

We fitted normative models to four diffusion tensor imaging (DTI) metrics (FA, MD, AD, RD) from 21 deep WM regions and the "Global-WM" (a single metric including the entire WM) based on 19 worldwide cohorts after rigorous quality control (see **Table 1**). Normative models were estimated - for 22 WM regions - for each DTI metric using a generative model that uses 10 repetitions of 80%-20% train-test splits with different random seeds -of all datasets but excluding the clinical population-. **Figures 1 and 2** show the results of these sampling procedures for the Global-WM. The posterior distributions for the parameters of the HBR model are estimated from the training data using the No-U-Turn Sampler (NUTS)[19]. For inference on the test data, we sample from the posterior predictive distribution, which represents the distribution over new observations given the estimated parameters. To derive the normative ranges, we sampled from the posterior predictive distribution across the full age range for both males and females (see Methods). In data from 19 sites, we see that each site has a different intercept (**Fig. 1B**). This is expected, as the mean values of DTI metrics depend on the scanning parameters, such as voxel size and *b*-value, yielding a different distribution of diffusivity and anisotropy values per site. In **Figure 1D** we show the 2nd, 10th, 25th, 75th and 90th and 98th percentile curves, corresponding to $Z=\pm0.68$, ±1.29 and ±2, respectively. For the Global-WM, the FA showed an inverted U-shaped trajectory, with a peak at 29 years of age. The three diffusivity measures (MD, AD, and RD) show a U-shaped trajectory, with minima later in life: at 37 years for RD, 43 years for MD, and 53 years for AD (**Figure 2**). **Figures S1**, **S2**, **S3** and **S4** show the lifespan trajectories and percentile curves for each DTI metric for the 21 regions besides the Global-WM.

For those regions that had a U-shape trajectory, FA peaked between 16 years (GCC) and 39 years (FX), MD reached minima between age 30 (PTR) and 51 years (UNC), while RD showed a pattern similar to that of MD with minima as early as 27 years (PTR) to 44 years (UNC). AD showed the greatest variability across the regions among all metrics; 14 out of 22 regions showed a U-shaped pattern for which minimum AD values occurred between 42 and 55 years. For AD, some regions had a quasi-linear trajectory from ages 4 to 80 years (CGH, FXST), and three showed an initial increase up to age ~20 to reach a plateau or later increase again at a later age (ALIC, CGC, CST). The average age of maturation (peak/minimum) of the DTI metrics differed. Pairwise comparisons of peak/minimum ages across regions were significantly different: FA to MD, FA to AD, FA to RD, MD to AD, MD to RD, AD to RD. On average, FA reached maturity at the youngest age (trimmed mean=29.2 years), followed in order by RD (trimmed mean=35.8 years), MD (trimmed mean=41.7 years) and AD (trimmed mean=48.2 years). **Figure 6** shows the peak/minimum ages for FA, MD, AD, and RD for 21 WM regions.

A key question for the practical utility of normative models is how well the models fit the data. To test model accuracy, we ran two diagnostic metrics for all 22 regions: the standardized mean squared error (SMSE) that assesses the goodness of fit and central tendency and the mean standardized log loss (MSLL)[20] that evaluates both the mean and standard deviation of the predicted distribution. An SMSE value closer to 0 indicates a better fit, and a more negative MSLL value indicates better performance. Additionally, to test the reliability of the estimated centiles we calculated Shapiro-Wilk's *W*, the kurtosis and the skewness of the estimated Z-scores (**Figure 3**). We aimed to verify how well the estimated centiles followed the normal distribution. Median SMSE scores across all regions and across all DTI metrics were below 0.67 and median MSLL scores were below -0.24, reflecting reasonable goodness of fit. AD had on average the best SMSE and MSLL scores, denoting a good fit for the bulk of the distribution, followed, in order, by FA, MD and RD. The Global-WM had the best SMSE and MSLL scores for FA and AD, whereas the FX had the best scores for MD and RD. Since all *W* scores across all four DTI metrics were above 0.97–where 1.0 is a perfect score and indicative of normality–we further investigated the distribution of the estimated Z-scores. For all DTI metrics we found, in general, platykurtic[2] distributions of the Z-scores, with some notable exceptions, such as the FX (for FA, AD and RD), the TAP (MD, AD, RD) showing higher values beyond 3.0 (leptokurtic[3]), especially for MD and RD. Interestingly, the FX showed the highest skewness across all DTI metrics, negative for FA, positive for the MD, RD, AD, despite showing the best goodness of fit metrics (SMSE and MSLL). The TAP showed high positive skewness for MD, RD, AD, following the FX. A possible explanation for this is that the FX and the TAP are close to the ventricles and may be affected by partial voluming of the WM with cerebrospinal fluid in the same image voxels (see **Figures S5, S6, S7, S8**).

[2] Platykurtic refers to a statistical distribution in which the excess kurtosis value is negative. Such a distribution will have thinner tails than a normal distribution will, resulting in fewer extreme positive or negative values.

[3] Leptokurtic refers to distributions with a positive kurtosis larger than that of a normal distribution. Leptokurtic distributions have heavier tails or a higher probability of extreme outlier values.

Figure 1.

Normative models of brain microstructure. A) Age distribution of the 19 datasets used in the experiments. Median ages are shown above the boxplots, and total numbers of subjects are listed to the right. For clinical samples (HBN, ADNI3, OASIS3 and PPMI), only the control subjects are shown (as they were the only ones used to define the normative model). The color code for the 19 datasets is the same for panels A, B and C. **B)** Shows the lifespan trajectory for the global white matter fractional anisotropy metric (Global-WM FA) as a mean across males and females (these are separated in D). Each curve shows the result of generating synthetic predictions for each of the 19 scanning protocols by sampling from the predictive posterior distribution. Thus, each curve has a different intercept (random intercept option used for the hierarchical Bayesian Regression). **Note:** HBN_SI, ABCD_Siemens, PedsDTI_Siemens, PING_Philips, AOMIC_ID1000 and ADNI3_S31 are subsamples of HBN, ABCD, PedsDTI, PING, AOMIC and ADNI3, respectively. **C)** shows the raw data points of FA for the Global-WM, where each cohort is shown in a different color. **D)** shows the generated predictions for the site "CAMCAN" for the Global-WM FA metric, for each of the 10 HBR iterations. The reference charts can be made separately for women and for men by modeling sex as a batch effect. To the left, the lifespan trajectory of FA for the Global-WM after averaging the 10 iterations for each sex separately. It also shows centile curves for Z=±0.68, ±1.29 and ±2. Note: When generating data from one site and for all 10 repetitions, we noticed that for ages greater than 91 years and under 4 years there is instability across the repetitions, likely caused by the scarcity of subjects in those age ranges. We show the main trajectory of FA between ages 4 and 91.

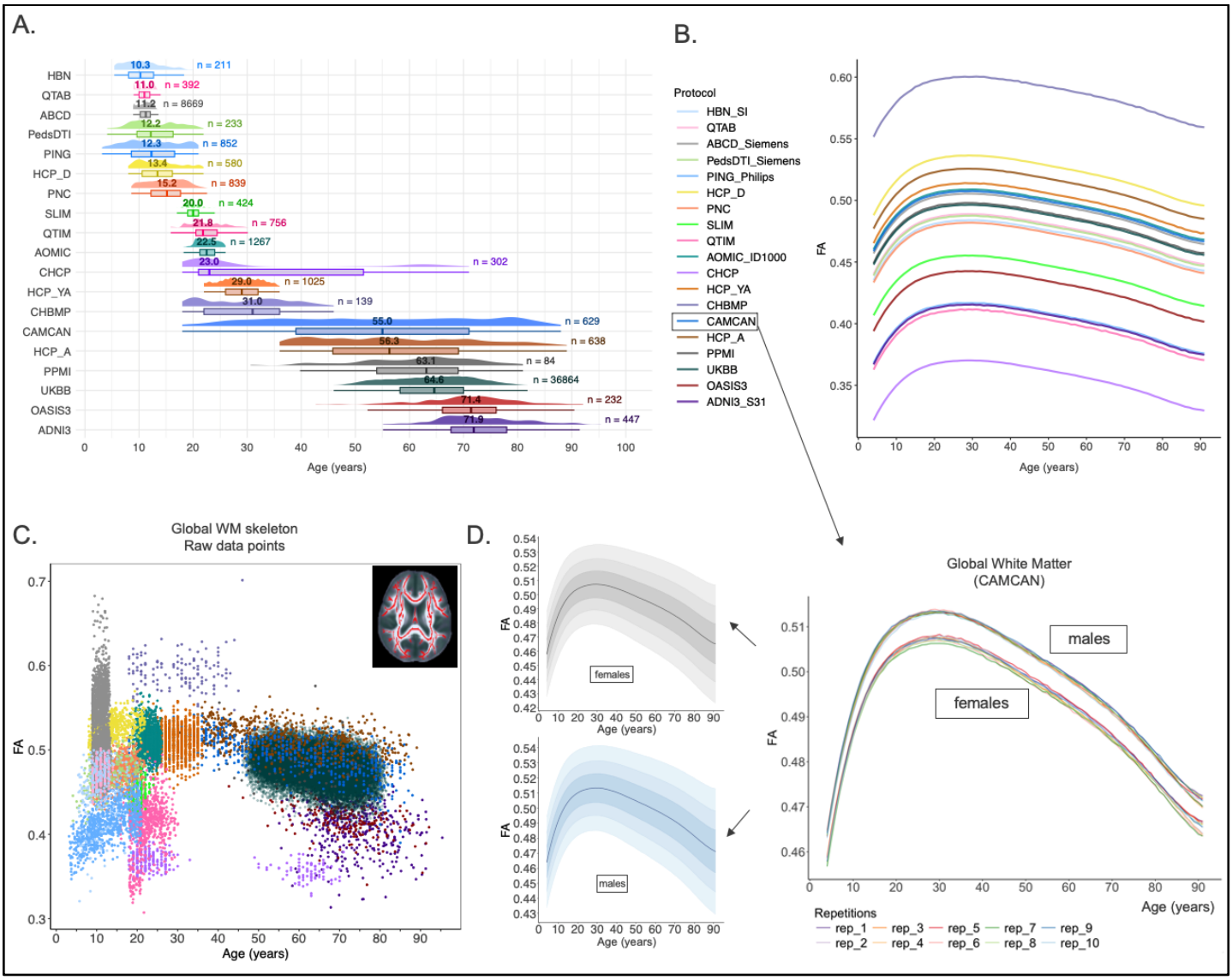


Figure 2.

Normative models of the overall white matter (Global-WM) and peak ages per ROI. A) Age-dependent charts of the four DTI metrics (FA, MD, AD, RD) across the lifespan (age 4-91 years). In order to fit them in the same plot we adjusted the y-axis as a percentage of the maximum value for each metric. The dashed vertical lines indicate the age of maximum FA and the minimum age for MD, AD, RD. **B)** We show the centile curves for FA, MD, AD and RD for the Global-WM (females). The bottom row shows the trajectories for males and females separately.

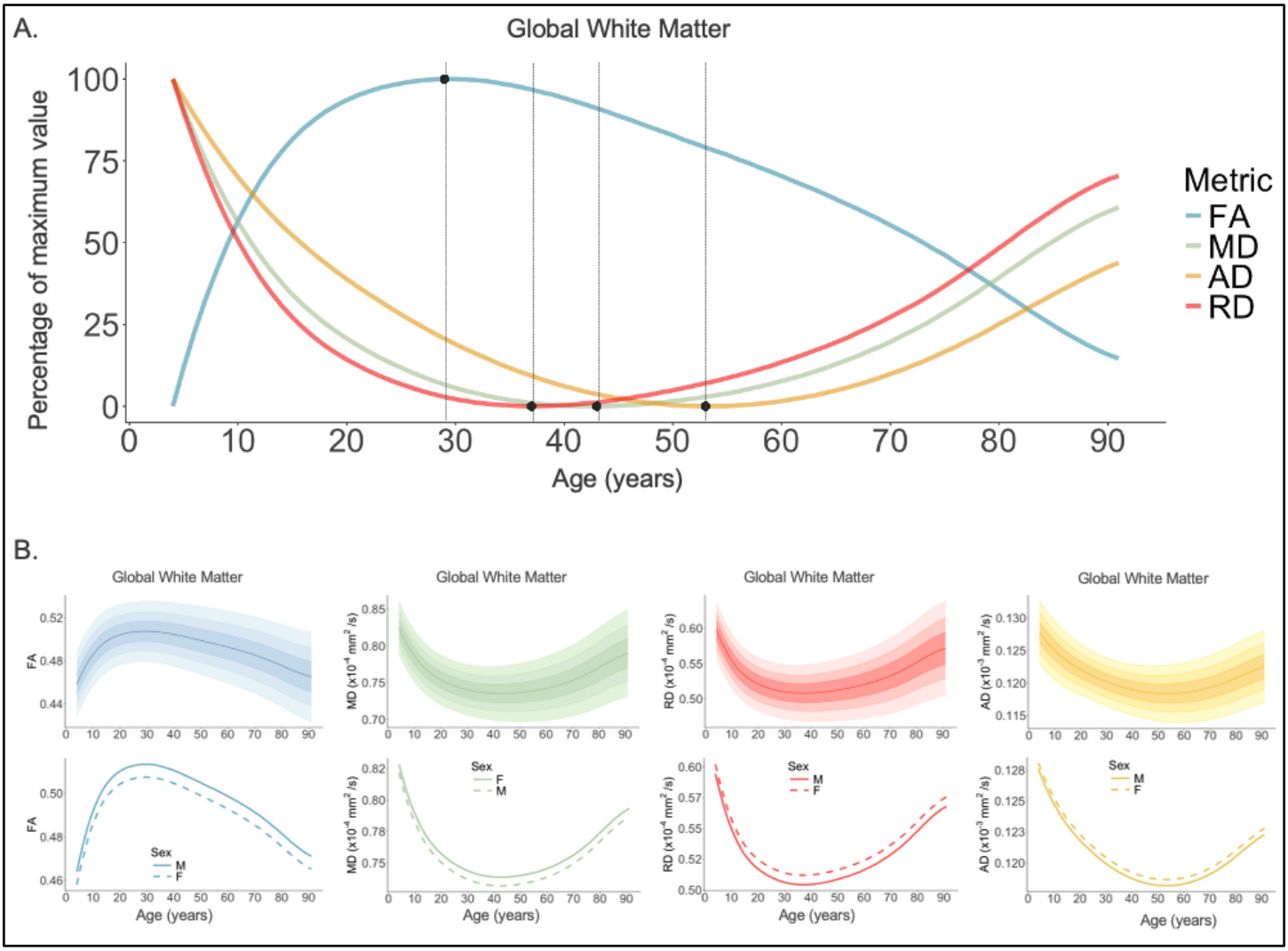


Figure 3.

Hierarchical Bayesian Regression: model diagnostics. Here we show how well the models fit the data, as a function of age. **A)** We show Shapiro-Wilk's *W* statistic, the standardized mean squared error (SMSE) and mean standardized log loss (MSLL) for all four DTI metrics. All WM regions are lumped into one density plot for each DTI metric. The line inside the density plot indicates the median. Axial diffusivity (AD) had on average the lowest SMSE (median=0.47) and MSLL (median=-0.39) of all DTI metrics. RD had on average the highest SMSE (median=0.67) and MSLL (median=-0.24). FA had on average the highest *W* (median=0.98) and MD the lowest (median=0.97). **B)** We show kurtosis, skewness, SMSE, MSLL per ROI for fractional anisotropy (FA). Skewness and kurtosis are shown to contextualize the results of the Shapiro-Wilk normality test. We show that most regions had a platykurtic distribution (kurtosis < 1.0) with the exception of the FX and the SCC. Most regions had a skewness between -0.25 and 0.25, except for the SCR (positively skewed) and the SCC, BCC, GCC and FX (negatively skewed). The regions with the best scores for SMSE and MSLL are the Global-WM, the FXST and EC. ROI-wise plots for the other metrics (MD, AD, RD) are found in the supplements (**Fig. S6, S7, S8**).

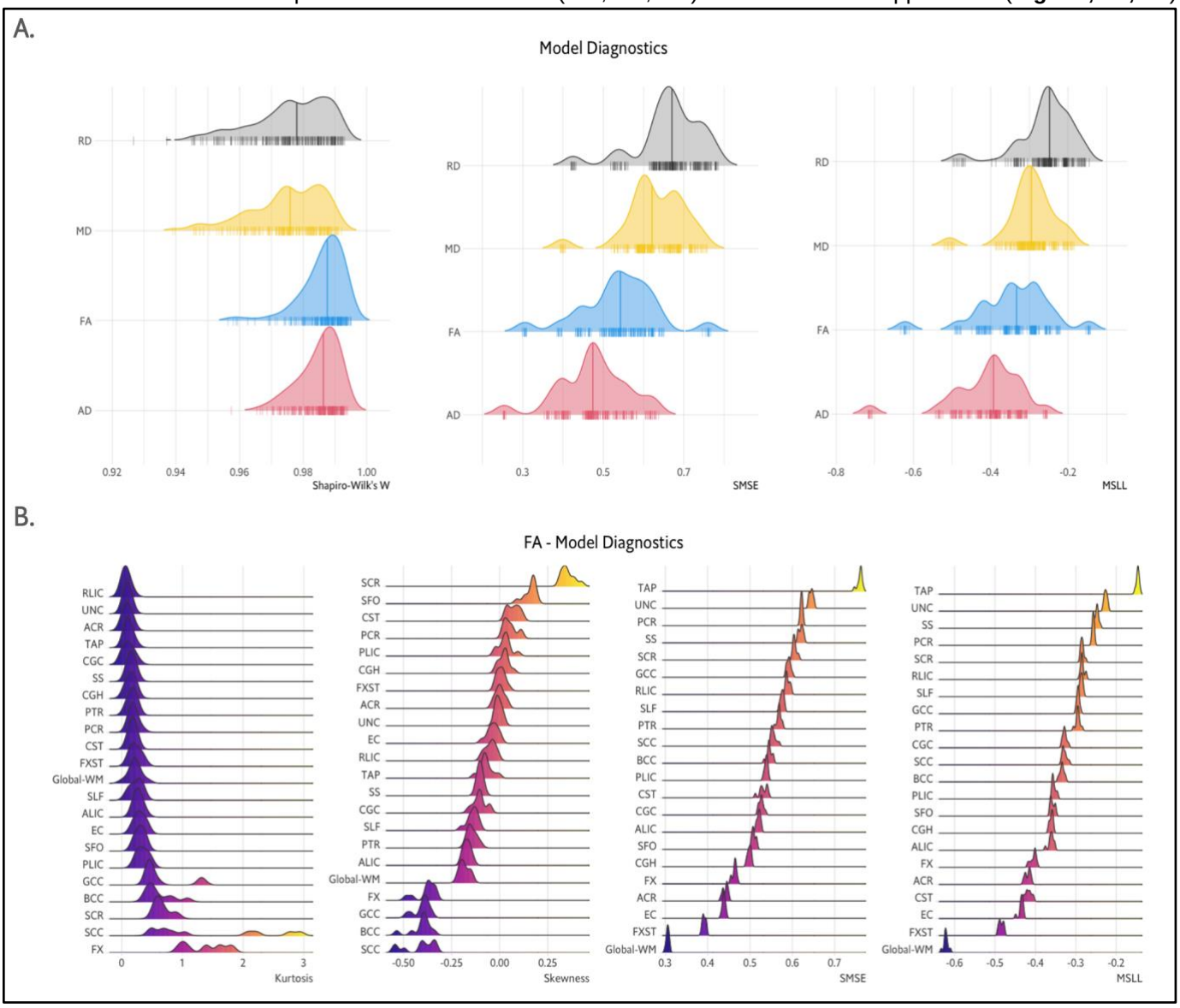
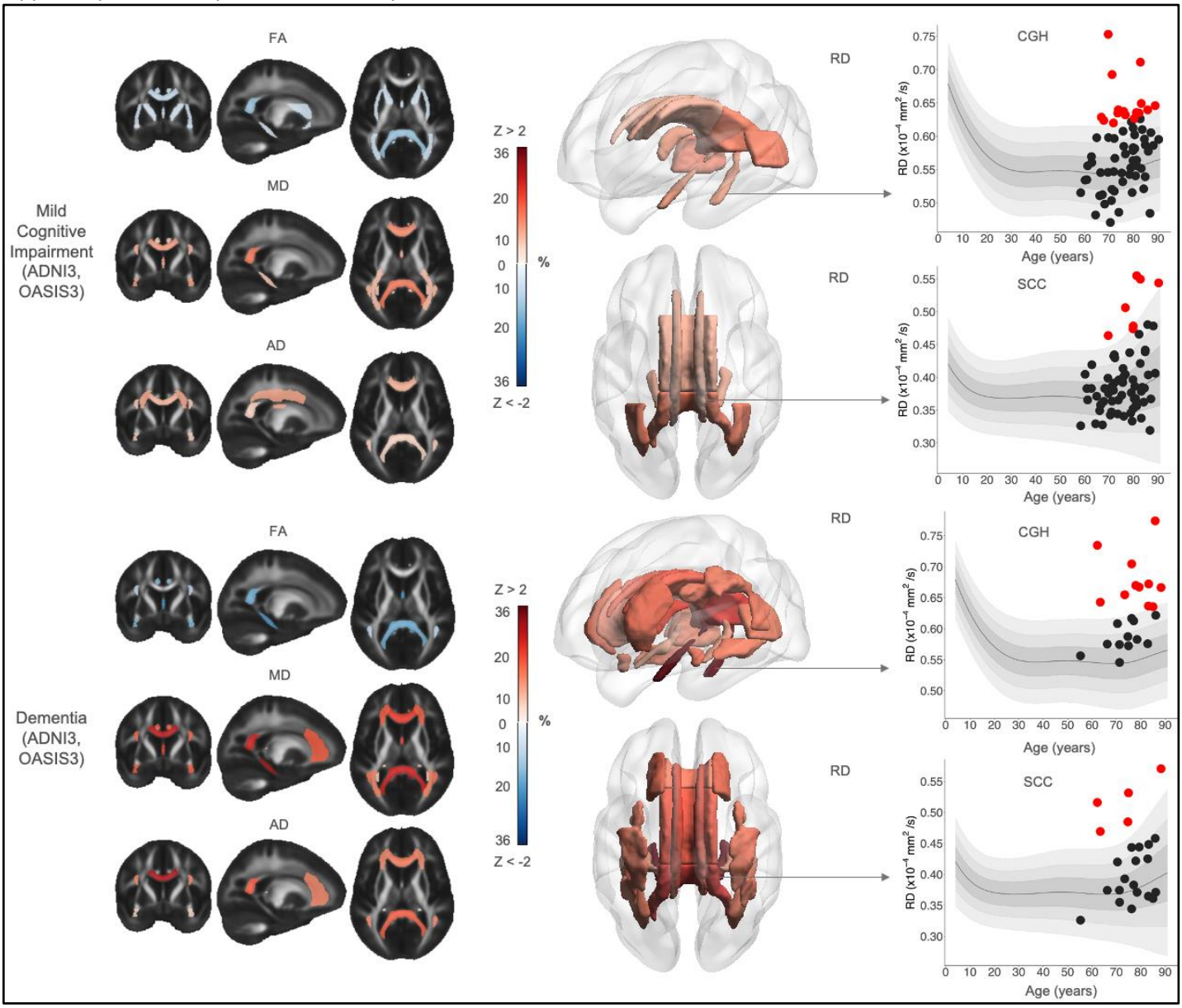


Figure 4.

Anomaly detection in elderly individuals with mild cognitive impairment (MCI) and dementia (ADNI3 and OASIS3 datasets). The WM regions shown have AUC > 0.5 (i.e., the models in these regions significantly discriminate the cases from controls) and passed multiple comparisons testing while controlling the false discovery rate (FDR). The extreme deviations (|Z|>2) are shown for FA, MD, AD in 2D, and for RD in 3D. The regions are color-coded based on the percentage of participants with extreme deviations --blue scale for negative deviations, red scale for positive deviations. The metric with the greatest proportion of extreme deviations is RD, with up to 36% of positive extreme deviations for the **CGH** in dementia patients. As expected, the extreme deviations for FA are in the opposite direction than for diffusivity metrics, i.e., MD, AD, RD. The centile curves on the right are specifically for the ADNI3-S55 dataset. Dots displayed on the centile curves denote the ADNI3-S55 patients. The dots in red are those with Z>2. **CGH**=cingulum of the hippocampus; **SCC**=splenium of the corpus callosum.



2.1.1. Retrogenesis hypotheses.

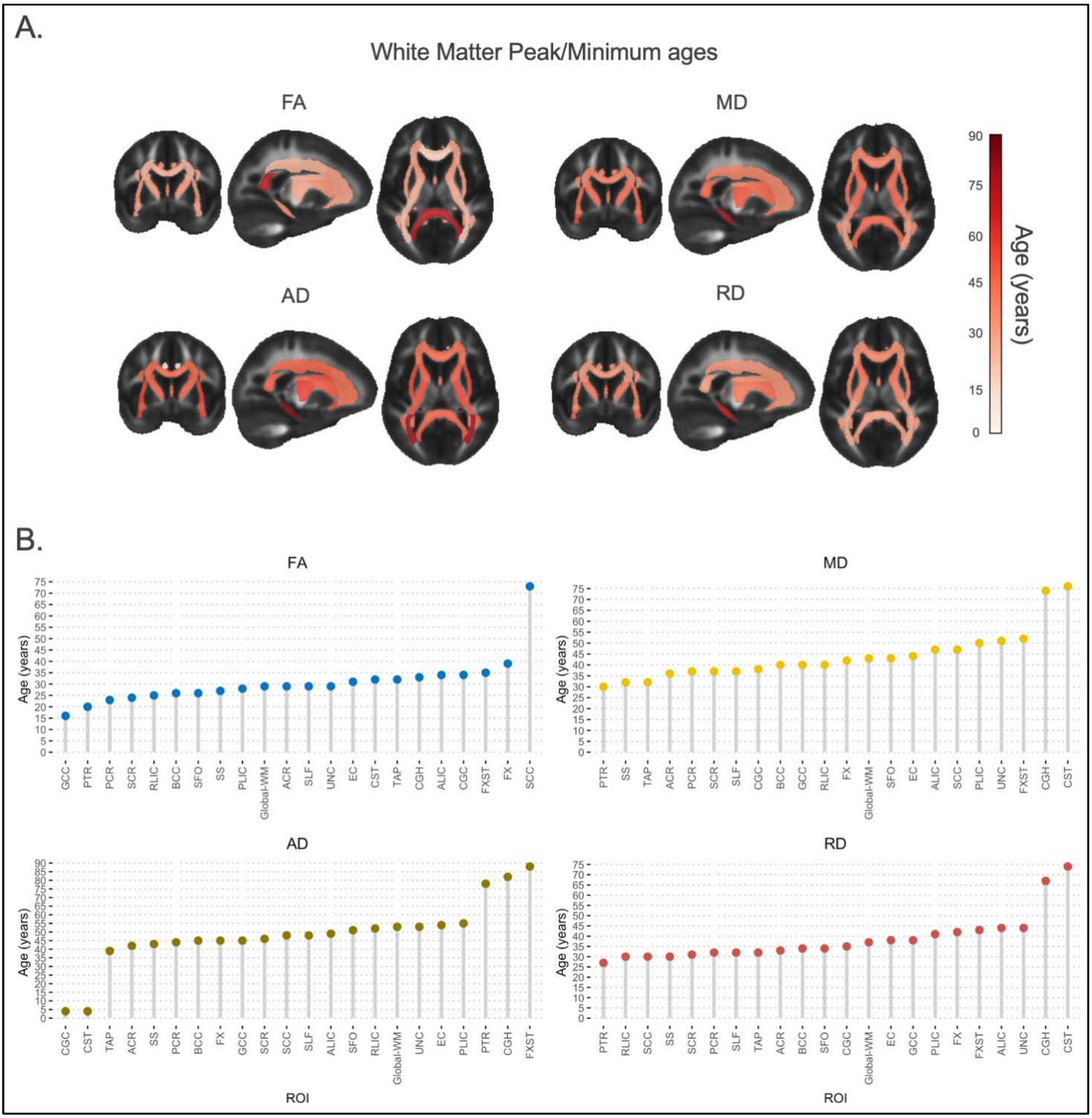
We did not find evidence to support the *'gain-predicts-loss'* hypothesis (**Figures S9 to S12**). We found some evidence supporting the *'last-in first-out'* hypothesis, but only for MD (*Rho*=0.47, *p-value* = 0.03); that is, the greater the percent change from age 4 to 14 years (see **Figure 5 & S14**) – indicating a faster decrease of MD – the younger the age of maturity, or minimum MD. This might also indicate that the regions with a slower maturation reach a minimum MD value later in life which in turn might indicate steeper increases of MD -faster degeneration- from the minimum to age 91 years. No other DTI metrics showed evidence for the *'last-in first-out'* hypothesis (**Figures S13** to **S16)**. Additionally, we did not find differences in the age of peak/minimum among the three groups of WM regions: commissural, association and projection fibers.

Figure 5.

White matter changes across the lifespan. A) Percent change of DTI metrics every 10 years across the lifespan (4-91 years). regions are grouped by type: association, projection, and commissural fibers and Global-WM.

Figure 6.

White matter peak/minimum ages. A) Peak ages for the 21 white matter regions of interest for all four DTI metrics. **B)** Age at peak or minimum for each DTI metric, organized from youngest to oldest age of maturation.



2.2. Anomaly detection

As detailed in the Methods, we next set out to find significant deviations from the reference model for each ROI per metric across 10 experimental repetitions. We did this for two datasets, for which 80% of healthy controls were used in the model estimation, i.e., ADNI3 and OASIS3. The anomaly detection focused on the participants with dementia or MCI. The Z-scores were first converted to abnormality scores, i.e., the probability of each metric being normal, based on a modified normal cumulative distribution function[16]. To determine the discriminative ability of the brain metrics for each condition, we calculated the area under the ROC curve (AUC) for classification of the patients (AlzD or MCI) versus the 20% controls in the test set for each of the 10 experimental repetitions. Here, we report the regions per metric with significantly above the chance performance (AUC>0.5) that passed multiple comparisons correction (across metrics) in 9 out of 10 repetitions to consider deviations significant and stable; for dementia, we found the highest classification accuracy for the CGH for RD (AUC=0.74), for MD (AUC=0.7) and FA (AUC=0.65). For axial diffusivity in dementia, the BCC had the highest accuracy (AUC=0.67). For MCI, the AUCs were lower, as expected, than in dementia: values ranged from 0.55 to 0.6 (**Table S4 & S5**). The best performance was observed for MD in the SCC (AUC=0.6), for AD in the BCC (AUC=0.67), and in the FX for FA and MD (AUC=0.58/0.59, respectively). Extreme deviations (|Z|>2) were mostly in the same direction for dementia and MCI; most subjects showed positive deviations (Z>2) for the diffusivity measures (MD, RD, AD) and negative deviations for FA (**Tables S2 & S3**). In MCI and dementia, the metric with the largest number of extreme deviations was RD (**Figure 4**). 37% of dementia participants had positive extreme deviations for RD in the CGH, and 25% in the SCC; 17% of subjects with MCI had positive extreme deviations in the SCC and 11% of subjects in the CGH. For FA, negative extreme deviations were mostly found in the CGH for dementia (20%) and the FX (21%) and in the SCC for MCI (13%).

2.3. Model adaptation

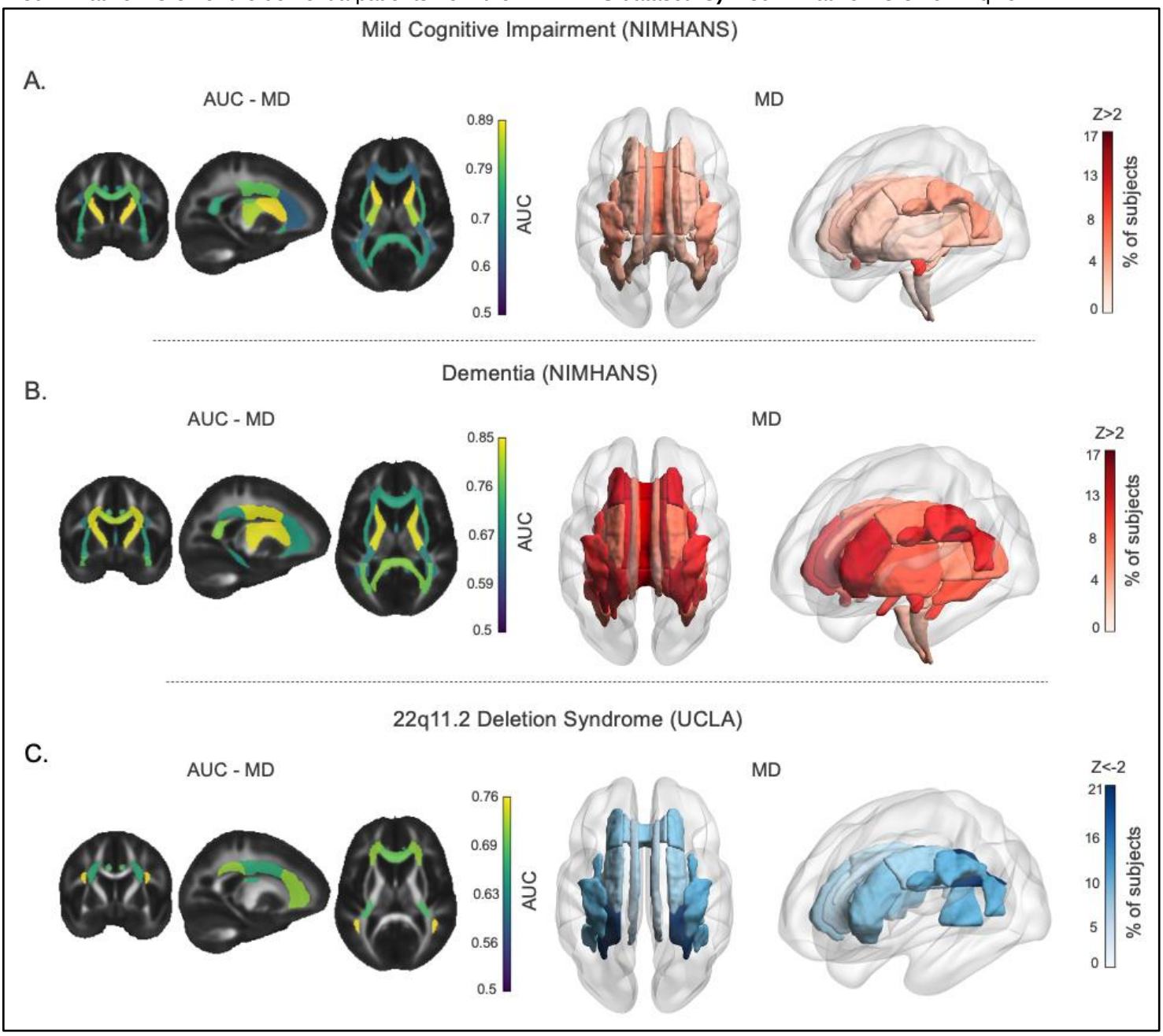
Next, we adapted the large-scale trained reference model to healthy controls from two new independent datasets. We assume that these two datasets were not available during the development phase of the reference model to simulate a real-world scenario where the model needs to be adapted to local data in the deployment stage. The NIMHANS elderly cohort has MCI and dementia participants and age-matched healthy controls; and the UCLA neurogenetic dataset has 22qDel participants and age-matched healthy controls (**Table 2**). For each dataset we used 50% of the healthy controls for recalibrating the model. Thereafter, the remaining 50% of healthy controls and the patients were used for anomaly detection, as described in the Methods (sections 9 and 10). For the NIMHANS cohort, we found a greater severity -assessed by the percentage of subjects with extreme deviations- for MCI compared to dementia (**Figure 7A & 7B**). Additionally, for MCI we found subjects with extreme positive deviations for MD, AD and RD, whereas for dementia, all four DTI metrics showed regions with significant extreme deviations; negative deviations for FA and positive deviations for diffusivities (**Tables S6, S7, S8, S9).**

For 22qDel, we found significant effects in multiple regions for MD and AD. MD showed more regions with extreme deviations (9 of 21 regions) than AD (5 of 21 regions). Interestingly, the entire *corona radiata* showed negative extreme deviations (ACR, SCR, PCR). For both DTI metrics, extreme

deviations were negative, except for the PLIC in AD, where the extreme deviations were positive. (**Figure 7C; Tables S10 & S11**).

Figure 7.

Results of the model adaptation experiment. Anomaly detection for patients with Mild Cognitive Impairment (MCI) and dementia in the NIMHANS dataset and patients with 22q11.2 Deletion Syndrome (22qDel) in the UCLA dataset. The colored ROIs are those that passed FDR correction, with AUC>0.5 when tested against the test sample of healthy controls (50%) from the same dataset (UCLA or NIMHANS). We show the AUCs per ROI and the percentage of extreme deviations for MD (Z>2 & Z<-2). **A)** Discriminative results for the MCI patients from the NIMHANS dataset. **B)** Discriminative ROIs for the dementia patients from the NIMHANS dataset. **C)** Discriminative ROIs for 22qDel.



3. Discussion

Here, we used HBR to create normative models of the brain's WM microstructure, based on a diverse, large-scale international dataset of diffusion MRI scans. In addition to revealing lifespan trends for the most widely used metrics of brain microstructure, we showed how the model is useful for investigating neurological, and neurogenetic conditions. Reference models were created to quantify the distributional properties of FA, MD, AD, and RD throughout life. By mapping their normal range of variation, and projecting clinical data into this normative range, we detected patterns of deviation from this range for dementia, MCI, and 22qDel. Such large-scale normative models offer transdiagnostic clinical value; unlike group differences in case-control analyses, normative modeling yields individual profiles of anomalies, which are agnostic to the clinical labels and accommodate the well-supported premise that not all individuals with the same disease deviate in the same brain regions and in the same direction[1,3]. Mapping the extent and scope of individual-level deviations can complement the aggregated group difference effects reported in large-scale case-control studies, but may also flag anomaly patterns that would be washed out in group comparisons of cases and controls.

As expected, for most regions, we found an inverted U-shaped lifespan trajectory for FA, increasing during maturation and declining after a peak. Diffusivity metrics exhibited a U-shape lifespan trajectory, with an initial decrease during maturation and an increase later in life. FA typically reached maturity (i.e., a maximum) at a younger age, followed by RD and MD, whereas AD reached maturity (i.e., a minimum) at age ~40 or ~50 years, depending on the ROI. This is in line with prior cross-sectional and longitudinal studies of age-dependent changes in WM DTI metrics[21–27]. Even though most regions in this study showed a U-shaped or inverted U-shaped trajectory across the lifespan, particularly for FA, MD, and RD, this was not the case for all WM regions. AD was the metric with the most regions deviating from this U-shaped trajectory–six out of 22 regions–with some regions showing a generally linear trend, followed by FA, whereas RD and MD were the most consistent in terms of displaying the inverted U-shape. Future normative modeling efforts focusing on these few WM regions may consider fitting age with linear models.

Relatively few studies have tested the retrogenesis hypothesis on WM DTI metrics[17,18,25,28–31], but such an analysis is now feasible as large dMRI datasets are available from childhood to senescence. Retrogenesis was first proposed in the context of Alzheimer's disease, where the course of functional and cognitive decline followed a pattern that was to some extent the reverse of the normal human developmental sequence[32]. This hypothesis has been termed '*last-in first-out'*. It was later hypothesized that structural and functional neuroimaging phenotypes might follow the same *'last-in first-out'* trend[33], with some evidence supporting this for WM DTI metrics in Alzheimer's disease[34,35]. The underlying hypothesis for the white matter is that the early myelinating fibers, such as the internal capsules, cerebral peduncles, and corticospinal tracts, may be less vulnerable to degeneration than the late myelinating fibers, such as the inferior and superior longitudinal fasciculi, anterior centrum semiovale, corpus callosum, and hippocampus. Nonetheless, the empirical evidence for this theory for DTI metrics has been rather moderate[30] or absent[17,18,22]. Even so, despite our tests in a large and diverse population, we did not find a significant relationship between the age at maturation and the percentage of change post-maturation. However, we found a positive association between the percent change of MD from age 4 to 14 years and the age of peak (i.e., the minimum MD) across regions. This finding may suggest that the fastest and earliest maturing regions may have a less pronounced slope from

“peak” to age 91 years and thus effectively degenerate slower. Nevertheless, approaches using piecewise linear fits with three defined sections of the lifespan: maturation, plateau and degeneration[17], may be of interest for future experiments.

Another related retrogenesis theory is the ‘*gain-predicts-loss’* hypothesis. It suggests that the rate of tissue maturation during development is inversely correlated with the rate of degeneration at an older age. This hypothesis has been tested on lifespan samples using either quadratic fits or by correlating the pre-maturation slope to the post-maturation slope, assuming some degree of symmetry on either side of the maturation peak. Even so, some studies found no such symmetry in DTI lifespan trajectories[17,31]. Others have found moderate or strong evidence[25,29], especially for FA and RD. Even so, we did not find sufficient evidence for the ‘*gain-predicts-loss’* hypothesis for any of the DTI metrics.

The post-analysis of the anomaly detection experiments identified extreme deviations ($|Z|>2$) in both directions, positive and negative–a hallmark of heterogeneity within the clinical group. In MCI and dementia, most deviations for MD and RD were positive, indicating higher WM diffusivity; however, for some regions, a smaller set of participants showed negative deviations (lower diffusivity). This was more prevalent in MCI than in dementia, and more prevalent for FA and AD than for MD and RD. FA may indeed be higher in the early stages of MCI than in healthy controls, and may even reflect recent evidence of Alzheimer’s disease subtypes: typical, limbic-predominant, and hippocampal-sparing Alzheimer’s disease[36]. In the future, normative models may help to identify the disease subtype to which an individual belongs to, given the availability of other biomarkers such as tau, amyloid, and APOE, although this is beyond the scope of the current manuscript.

Our anomaly detection framework was also able to gauge disease severity in new subjects by adapting the normative models to an unseen sample. For the NIMHANS dataset, we found some WM regions with extreme deviations in MCI, whereas all regions had significantly anomalous deviations for dementia, especially for MD. We found some differences between the ADNI3/OASIS3 and NIMHANS samples for MCI and dementia, perhaps due to differences in sample sizes and study recruitment inclusion criteria. Ancestry might play a role as well, as ADNI3/OASIS3 consist primarily of white Caucasian North Americans and NIMHANS evaluates Indian ancestry participants resident in India.

We also found a similar phenomenon for 22qDel: for some regions, a proportion of 22qDel participants also had deviations in the opposite direction to the majority of subjects; this aspect would be missed using approaches that averaged multi-subject data prior to group comparisons. In particular, for 22qDel, even though the largest case-control studies show lower diffusivity for almost the entire WM[37], we identified some regions for which a small proportion of extreme deviations was in the opposite direction–that is, higher diffusivity (MD and AD). This may reflect biological differences within the sample of cases, such as the length of the 22q11.2 locus deletion. Another reason may be that many factors influence DTI, such as neurite density, axonal dispersion, and crossing fibers; higher-order diffusion models and advanced data acquisitions may be needed to disentangle these effects.

We observed a peculiar lifespan trajectory for a few WM regions. This aspect is characterized by a ‘double peak’, around age 50 to 60 years. For instance, the SCC for FA had a second peak at age 73 years of age. Similarly, the PLIC, RLIC, SCR, and SFO also showed similar trajectories for FA, and the CST, FX, CGH, and PTR for diffusivity metrics. This has been reported before for DTI FA: Zhu et al.

reported a similar pattern for FA in SCC[28]. Bouhrara et al.[21] reported that FA displayed a more heterogeneous pattern–where some regions, such as the SCR, had higher values with age. Burzynska et al.[38] also found no decrease in FA for the splenium with age. Westlye et al. reported that cingulum bundles showed a less clear-cut or 'non-U-shaped' trajectory[22]. Sullivan et al.[39] did not find significant linear age effects for CGH and the splenium in men. Similar results have been reported in longitudinal studies: in Storsve et al.[40], the *forceps major*, a tract running through the SCC, showed no longitudinally significant changes for FA, similar to the left cingulum angular bundle and left ILF. However, the magnitude of this double bump trajectory is much lower than the estimated variance in the sample. That is to say, that the overall, global variance of the estimated model is larger than the local variance of a specific age range or even for the variance explaining the sex batch effects. This can be observed in Supplementary **Figures S1-S4.** Finally, from the dMRI modeling perspective, DTI is a non-specific microstructural model. FA is a metric sensitive to a combination of multiple tissue properties, including axonal density, myelination, axonal orientation dispersion, crossing fibers, and free-water. Free-water may predominate in many voxels as atrophy and tissue shrinkage occur after age ~50 years and could cause these local effects on the age trajectory of the WM region. Other microstructural properties, such as dispersion, cannot be ruled out - a key topic for future work.

One limitation of this study is that the ENIGMA-DTI protocol has not been optimized for very early childhood (3-5 years), so models based on infant data[8] may be more appropriate for this age range. Some regions, such as the fornix, may be contaminated with cerebrospinal fluid due to partial volume effects caused by misregistration to the ENIGMA template. Tractometry techniques may also more accurately segment WM tracts close to regions with tissue partial voluming. Future studies will include a richer set of microstructural measures, including Standard Model Imaging[41] and Neurite Orientation Dispersion and Density Imaging[42]. Although our results presented are based on DTI, the reference models are still highly sensitive to WM microstructural deviations from the norm in psychiatric and neurological diseases. Additionally, the single-shell scanning protocols are, to this day, the most commonly available on most MRI scanners. More advanced multi-compartmental models beyond DTI will require multi-shell dMRI data, which is still limited in clinical populations, but worthy of examination in future. Finally, more advanced distributional models may also be beneficial for normative modeling, such as high-order non-Gaussian distributions modeled using nonlinear probability warping[43], SHASH function models[44] and high-dimensional variational autoencoders that can model the joint distribution of multiple features[45,46].

Methods

1. Datasets analyzed

Table S1 summarizes the diffusion MRI acquisition protocols for all datasets. We included the following public datasets with various age ranges to cover the lifespan: the Adolescent Brain Cognitive Development (ABCD)[47], the Amsterdam Open MRI Collection (AOMIC)[48], the Cambridge Centre for Ageing and Neuroscience (CAMCAN)[49], the Cuban Human Brain Mapping Project (CHBMP)[50], the Chinese Human Connectome Project (CHCP)[51], the Human Connectome Project Aging & Development (HCP_A & HCP_D)[52], the Human Connectome Project Young Adult (HCP_YA)[53], the DTI component of the NIH MRI Study of Normal Brain Development (PedsDTI)[54], the Pediatric Imaging, Neurocognition, and Genetics data repository (PING)[55], the Philadelphia Neurodevelopmental Cohort (PNC)[56], the Queensland Twin Adolescent Brain (QTAB)[57], the Queensland Twin Imaging dataset (QTIM)[58], the Southwest University Longitudinal Imaging Multimodal (SLIM) study[59], and the UK BioBank (UKBB)[60].

Several clinical datasets were also included in the analyses: the third phase of the Alzheimer's Disease Neuroimaging Initiative (ADNI3)[61,62], the third release of the Open Access Series of Imaging Studies (OASIS3)[63], the Parkinson's Progression Markers Initiative (PPMI)[64], and the Healthy Brain Network (HBN)[65]. Importantly, only the healthy controls (HC) of the latter four datasets were added to the pool of training data to build the normative models, yielding a total of 19 public datasets. **Figure 1A** summarizes the age distribution for the 19 datasets used for training. In contrast, the patients in ADNI3 and OASIS3 were included as test data for the anomaly detection experiment. ADNI3 and OASIS3 included patients with mild cognitive impairment (MCI) and dementia. Two additional datasets were used for the model adaptation experiments: the National Institute of Mental Health and Neuro Sciences (NIMHANS) dataset[66] and the 22q11.2 copy number variant (CNV) dataset from the University of California Los Angeles (UCLA)[67,68]. **Table 2** summarizes the demographics of all datasets analyzed.

2. Preprocessing of diffusion MRI scans

Three studies provided preprocessed dMRI images through their official project websites: HCP_YA, ABCD and AOMIC. Two studies provided precomputed DTI maps (PedsDTI and UKBB). The remaining datasets were preprocessed in-house. In all cases, preprocessing included correction for eddy currents, movement, and EPI-induced susceptibility distortions, even though the protocols were not fully aligned. A detailed description of the in-house preprocessing performed on each dataset may be found in the Supplementary Methods section. Importantly, tensor fitting was performed on the datasets with single-shell dMRI data or on the lowest non-zero shell available, if multi-shell data were provided. DTI maps (FA, MD, RD and AD) were computed with DiPy's tensor fitting function[69] for the following datasets: AOMIC, CAMCAN, CHBMP, CHCP, HCP_A, HCP_D, HCP_YA, NIMHANS, PNC, QTAB and UCLA. FSL's *dtifit*[70] was used to fit the diffusion tensor for ABCD, ADNI3, HBN, OASIS3, PING, PPMI, QTIM, SLIM. All subjects' FA maps were linearly registered with FSL's *flirt* and nonlinearly registered to the ENIGMA-FA template with ANTs' diffeomorphic non-linear registration algorithm[70,71]. For QTIM, non-linear registration was performed using the 3DMI algorithm[72]. The combined linear and non-linear transformations were applied to the MD, RD and AD maps. Mean DTI metrics were extracted from 21 bilateral regions of interest (regions) from the Johns Hopkins University WM atlas (JHU-WM)[73] using

the ENIGMA-DTI protocol[74]. The mean for each DTI metric across the whole skeleton (Global-WM) was also extracted for each participant, yielding a total of 22 regions.

3. Quality control

All study participants had information on age and sex at the time of the imaging session along with a usable dMRI scan. ADNI3, OASIS3, PPMI, CAMCAN and HBN were manually checked for head movement, FOV (field of view) related artifacts and other dMRI acquisition artifacts and preprocessing errors caused by poorly tuned software tools. For all other datasets used for training, outliers were detected with an anomaly detection algorithm called *isolation forests*[75] implemented in *Scikit-learn*[76], as detailed in the Supplementary Methods. Briefly, isolation forests were run on each ROI for each DTI metric, separately, to identify a subset of "outlier subjects" for each DTI metric. Subsequently, outlier subjects that were duplicated in at least two DTI metrics were removed from the sample. Aggregating the subjects left after this procedure plus the subjects from the datasets that were manually checked, a total of 54,583 subjects remained for training.

4. White matter regions assessed

Twenty-two JHU-WM regions of interest were assessed in this study (**Table 1**), including the following: PCR=posterior *corona radiata*, CGH=cingulum of the hippocampus, CGC=cingulum of the cingulate gyrus, UNC=uncinate fasciculus, RLIC=retrolenticular part of internal capsule, SCR=superior *corona radiata*, ACR=anterior *corona radiata*, EC=external/extreme capsules, PLIC=posterior limb of internal capsule, GCC=*genu* of the corpus callosum, SS=sagittal stratum, ALIC=anterior limb of internal capsule, FXST=*crus* of the fornix/*stria terminalis*, BCC=body of corpus callosum, TAP=tapetum of the corpus callosum, SLF=superior longitudinal fasciculus, SFO=superior fronto-occipital fasciculus, SCC=splenium of the corpus callosum, FX=fornix, PTR=posterior thalamic radiation, Global-WM=mean across the entire white matter skeleton of the ENIGMA-DTI template.

5. Normative Modeling with HBR

The HBR NM framework is thoroughly described by Kia et al.[16], but here we summarize the basic principles.

Let $\mathbf{X} \in \mathrm{R}^{n \times p}$ be a matrix of n the number of participants and p covariates. We added age and sex as the covariates for the present study and we modeled the non-linear effect of age with a cubic *b*-spline basis set expansion with five evenly spaced knots. Here, we denote the dependent variable as $y \in \mathbb{R}$. In its simplest form, NM assumes a Gaussian distribution over y, i.e., $y \sim \mathcal{N}\left(\mu, \sigma^2\right)$, and it aims to find a parametric or non-parametric form for μ and σ given $\mathbf{X}$**.** In its parametric form, μ and σ are respectively parameterized on $f_\mu\left(\mathbf{X}, \theta_\mu\right)$ and $f_\sigma^+\left(\mathbf{X}, \theta_\sigma\right)$, where θ_μ and θ_σ are the parameters of f_μ and f_σ^+. f_σ^+ is a non-negative function used to fit the standard deviation of the noise ϵ_i. Consequently, in a multi-site scenario, a separate set of model parameters could be estimated for each site, or batch, i, as follows:

$$y_i = f_{\mu_i}(\mathbf{X}, \theta_{\mu_i}) + \varepsilon_i \qquad i \in \{1, ..., m\}$$

However, an assumption in HBR is that θ_{μ_i} (and θ_{σ_i}) across different batches come from the same joint prior distribution that functions as a regularizer and prevents overfitting on small batches. HBR is

described as a *partial pooling* strategy in the multi-site scenario, i.e., similar to the *no-pooling* scenario, in HBR, the parameters for f_{μ_i} and $f^+_{\sigma_i}$ are estimated separately for each batch. Then, in the NM framework, the deviations from the norm can be quantified as Z-scores for each subject in the i_{th}batch:

$$Z_i = \frac{y_i - f_{\mu_i}(\mathbf{X}_i, \theta_{\mu_i})}{f^+_{\sigma_i}(\mathbf{X}_i, \theta_{\sigma_i})}$$

The harmonization of Z-scores happens at this stage where the Z-scores are adjusted for additive and multiplicative batch effects using the estimated f_{μ_i} and $f^+_{\sigma_i}$ for each batch. In fact, the model does not yield a set of "harmonized" data, i.e., with the batch variability removed. Instead, it preserves the sources of biological variability that may correlate with the batch effects. Consequently, the frequent statistical dependence in multi-site neuroimaging projects that occurs between age and site (see **Figure 1A**), caused by different inclusion criteria across cohorts, may be better tackled with HBR. When removing the site effect with harmonization methods such as ComBat[77] some of the biological variance correlated with age and site may be inadvertently removed[16].

6. Regression models

We identified 37 dMRI acquisition protocols across the 19 datasets (see **Table S1**). Some of the 19 datasets had subsamples acquired at multiple locations, often at different time points, at scanners from different vendors, ultimately modifying the acquisition protocols and leading to differences in the numbers of gradient directions, number of *b*-zero volumes, *b*-values, voxel sizes, etc. Essentially, these subsamples constituted a slightly different protocol within each dataset. Consequently, we used the *dMRI protocol* as the batch effect when training the normative model with HBR. HBR models these site-dependent parameters in a hyperprior, and the model hyperparameters are adapted to incoming data from each site. Subsequently, we ran an HBR model for each DTI metric across all regions that included sex as a second grouping variable in the HBR setting, and age as a fixed independent variable. We ran the training and testing of the HBR normative models using the PCNtoolkit package (v29) and Python 3.9 (Details in Supplementary Methods).

7. Model Diagnostics

We evaluated the goodness of fit of the HBR models, using the standardized mean squared error (SMSE), the mean standardized log loss (MSLL) and Shapiro-Wilk's *W* for all four DTI metrics. The SMSE assesses the goodness of fit and central tendency, for which a value closer to 0 is better. We also used the mean standardized log loss (MSLL)[20] which evaluates both the mean and standard deviation of the predicted distribution; a more negative value indicates a better performance. We also aimed to evaluate the calibration of the model; that is, if the estimated centiles of the distribution matched the observed centiles. First, we tested the normality of the estimated Z-scores of the test data using the Shapiro-Wilk test[78]. We used the test statistic (*W*) to quantify the degree to which the observed distribution deviates from the assumed normal distribution. Values less than one (1) indicate deviation from normality. Secondly, we also calculated the skewness and the kurtosis of the estimated Z-scores to further investigate their distributions. We ran these fit validation metrics (i.e., SMSE, MSLL, *W*, skewness and kurtosis) for each ROI across all 10 experimental iterations. The results are shown in **Figure 3** and **Figures S4, S5, S6**.

8. Plotting of the age trajectories

It is customary for the distribution estimated in a Bayesian model to be computed by sampling from the posterior predictive distribution. To generate samples from the posterior predictive distribution, we use the posterior samples $\{\theta_i\}$ and generate new data points $\tilde{y}_i$ from the likelihood function $p(y \mid \theta_i)$. To this end, for each posterior sample θ_i, a corresponding sample $\tilde{y}_i$ from the predictive distribution conditioned on θ_i is drawn. We generated these samples for the 10 repetitions of the HBR experiment (each with a different 80%-20% train-test split). To model the trajectory of each of the DTI metrics for each WM ROI, with respect to age and for males and females separately, we used the generative features of the hierarchical Bayesian model. For each DTI metric per ROI we generated samples for the age range 4 to 91 years and for both sexes 20 times and for one site only. We picked CAMCAN as the site for the simulations of the Global-WM (**Figures 1D and 2**). After generating these samples, we averaged the predicted data across the 20 iterations to generate a smooth curve. In **Figure 1B** we generated samples for 19 sites for the Global-WM to display the site differences.

9. Anomaly detection on the reference cohort

Training and testing data sets were created with an 80% to 20% sample split, stratified by site, based on the total number of 54,583 participants. To achieve stability, we repeated the same procedure 10 times by using *Scikit-learn's* '*train_test_split'* function with different random seeds, thus ensuring a different split each time. To ensure robustness against data perturbations, we ran HBR 10 times on each of the 80% sample splits (n=43,666). Z-scores were re-calculated on each of the 10 test sets (20% sample split; n=10,917). Z-scores were also calculated for the Alzheimer's dementia and MCI cases from ADNI3 and OASIS3. We calculated the probabilities of abnormality $P(z)$, i.e., the abnormality scores, from the Z-scores for both controls (20% sample split) and clinical samples (dementia and MCI) using the following formula[16]:

$$P(Z) = 2\left(1 - \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{|Z|} e^{\frac{-t^2}{2}} dt\right)$$

ROI-wise areas under the ROC curves (AUCs) were calculated to determine the classification accuracy of the computed deviations. Subsequently, we performed permutation tests with 1,000 random samples to derive permutation *p*-values for each DTI metric per ROI, and applied a false discovery rate (FDR) correction on each DTI metric separately across regions to identify those that showed significant group differences. The regions that showed significance in 9 out of the 10 experimental iterations were retained[4,16]. For each clinical group, i.e., dementia and MCI, we summarized the individual deviations within each group by first separating them into positive and negative deviations, counting how many subjects had an extreme deviation (defined as Z>2 for positive deviations, Z<-2 for negative deviations) at a given ROI, and then dividing by the group size (see **Figure 4**).

10. Model adaptation

Following the development phase, we aimed to deploy the trained normative models by transferring the estimated hyperparameters to new datasets. This deployment strategy, called *model adaptation*, allows us to re-infer the parameters of the model at local sites. We used the HBR model adaptation function

called '*transfer'* available in the PCNtoolkit v29[16]. Here, we took 50% of the healthy control subjects from the new sites to adapt the parameters of the reference model to the local model. Then the Z-scores for the remaining 50% healthy controls and the cases of the new incorporated sites are computed as described above in Methods section 5. We tested the model adaptation scenario on two cohorts with pathologies affecting the brain at two different stages of the lifespan, i.e., UCLA and NIMHANS (see **Table 2**). The UCLA dataset includes people with a rare neurogenetic condition who carry chromosome 22q11.2 CNVs (copy number variants; 22qDel) that affect brain development and have a higher risk for future psychiatric conditions such as neurodevelopmental delay, autism spectrum disorder, attention deficit hyperactivity disorder and schizophrenia. The NIMHANS dataset has patients with MCI and Alzheimer's dementia, both affecting the brain in later stages of life. Once the normative models were adapted to these new sites, we summarized the individual deviations within each group (22qDel, MCI and dementia) by first separating them into positive and negative deviations, counting how many subjects had extreme deviations at a given ROI, and then dividing by the group size (see **Figure 6**).

11. Retrogenesis of WM Regions

We tested two retrogenesis hypotheses that have been proposed for drawing general principles from the lifespan charts of brain metrics. Here we use the term "peak" for all DTI metrics, that is, for the maximum of FA and for the minimum of MD, AD, RD. To test the "*gain-predicts-loss"* hypothesis, we ran *Spearman* correlations between the percent change from age 4 years to the peak and the percent change between the peak and age 91 years. By doing so, we are testing correlations between metrics derived from the two sides of the curve, the maturation and the decline during aging. According to the "*gain-predicts-loss"* hypothesis both sides of the curve should have slopes that are correlated, i.e., the rate of senescence should depend on the rate of maturation. We assessed statistical significance with permutation tests (R package *coin*[79]). The other hypothesis, the "*last-in, first-out"* hypothesis, assumes that the cortical regions of higher complexity (supporting higher order functions) have a slower maturation rate and may be more sensitive to degeneration after reaching their maturational peak. By extension, the WM tracts connecting these more complex cortical regions should also have a slower maturation rate. We first calculated the percent change for each of DTI metrics every 10 years across the lifespan for each ROI (4-14,15-24, 25-34, 35-44, 45-54, 55-64, 65-74, 75-84, 85-91 years; **Figure 5**). We then assessed associations between the percent change from age 4-14 years, 75-85 years and 85-91 years with the age of peak. If the "*last in, first-out"* were true, a higher percent change during development should be correlated with a younger age of peak, and a higher percent change in senescence should be correlated with an older age of peak. Finally, we compared the average age of peak across regions between DTI metrics. For this, we performed pairwise comparisons with a repeated measures ANOVA design using a percentile bootstrap method on trimmed means (R package *WRS*[80]). Multiple comparisons were corrected with Hochberg's method. We also grouped the regions into three groups: commissural (BCC, GCC, SCC, TAP), association (CGC, CGH, EC, FX, FXST, SFO, SLF, SS, UNC), and projection tracts (ACR, ALIC, CST, PCR, PLIC, PTR, RLIC, SCR) and compared their peak ages per DTI metric using the same ANOVA percentile bootstrap method as above.

Data Availability

The estimated normative models will be available for the public at the PCNPortal[5] (https://pcnportal.dccn.nl/).

Author contributions

J.E.V.R. contributed with Conceptualization, Data Curation, Formal Analysis, Investigation, Writing – Original Draft Preparation, Investigation, Software, Visualization
A.H.Z. contributed with Writing – Review & Editing, Data Curation, Project Administration
S.B. contributed with Data Curation
C.A.M. contributed with Writing – Review & Editing
Y.F. contributed with Data Curation
T.C. contributed with Data Curation
L.N. contributed with Data Curation
L.K. contributed with Data Curation, Project Administration, Project Administration
C.E.B. contributed with Funding Acquisition, Writing – Review & Editing, Resources
J.P.J. contributed with Funding Acquisition, Resources
H.J. contributed with Data Curation
S.I.T. contributed with Data Curation, Project Administration
K.E.L. contributed with Data Curation, Writing – Review & Editing
T.M.N. contributed with Data Curation, Writing – Review & Editing
N.J. contributed with Funding Acquisition, Writing – Review & Editing, Resources, Supervision
S.M.K. contributed with Methodology, Software, Writing – Review & Editing, Supervision, Software
A.F.M. contributed with Methodology, Writing – Review & Editing
P.M.T. contributed with Conceptualization, Funding Acquisition, Supervision, Writing – Review & Editing, Resources

Competing interests

There are no personal or financial interests of any of the authors that could be considered a conflict of interest in relation to this manuscript.

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ABCD. Some of the data used in preparing this article were obtained from the Adolescent Brain Cognitive Development SM (ABCD) Study (https://abcdstudy.org), held in the NIMH Data Archive (NDA). This is a multisite, longitudinal study designed to recruit more than 10,000 children aged 9-10 and follow them over 10 years into early adulthood. The ABCD Study® is

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ADNI. Data used in the preparation of this article were obtained from the Alzheimer's Disease Neuroimaging Initiative (ADNI) database (adni.loni.usc.edu). The ADNI was launched in 2003 as a public-private partnership, led by Principal Investigator Michael W. Weiner, MD. The primary goal of ADNI has been to test whether serial magnetic resonance imaging (MRI), positron emission tomography (PET), other biological markers, and clinical and neuropsychological assessment can be combined to measure the progression of mild cognitive impairment (MCI) and early Alzheimer's disease (AD). For up-to-date information, see www.adni-info.org. Data collection and sharing for the ADNI project was funded by the Alzheimer's Disease Neuroimaging Initiative (ADNI) (National Institutes of Health Grant U01 AG024904) and DOD ADNI (Department of Defense award number W81XWH-12-2-0012). ADNI is funded by the National Institute on Aging, the National Institute of Biomedical Imaging and Bioengineering, and through generous contributions from the following: AbbVie, Alzheimer's Association; Alzheimer's Drug Discovery Foundation; Araclon Biotech; BioClinica, Inc.; Biogen; Bristol-Myers Squibb Company; CereSpir, Inc.; Cogstate; Eisai Inc.; Elan Pharmaceuticals, Inc.; Eli Lilly and Company; EuroImmun; F. Hoffmann-La Roche Ltd and its affiliated company Genentech, Inc.; Fujirebio; GE Healthcare; IXICO Ltd.; Janssen Alzheimer Immunotherapy Research & Development, LLC.; Johnson & Johnson Pharmaceutical Research & Development LLC.; Lumosity; Lundbeck; Merck & Co., Inc.;Meso Scale Diagnostics, LLC.; NeuroRx Research; Neurotrack Technologies; Novartis Pharmaceuticals Corporation; Pfizer Inc.; Piramal Imaging; Servier; Takeda Pharmaceutical Company; and Transition Therapeutics. The Canadian Institutes of Health Research is providing funds to support ADNI clinical sites in Canada. Private sector contributions are facilitated by the Foundation for the National Institutes of Health (www.fnih.org). The grantee organization is the Northern California Institute for Research and Education, and the study is coordinated by the Alzheimer's Therapeutic Research Institute at the University of Southern California. ADNI data are disseminated by the Laboratory for Neuro Imaging at the University of Southern California.

AOMIC. The Amsterdam Open MRI Collection (AOMIC) is a collection of three datasets with multimodal (3T) MRI data. All raw/preprocessed data is publicly available from the Openneuro data sharing platform:
ID1000: https://openneuro.org/datasets/ds003097
PIOP1: https://openneuro.org/datasets/ds002785
PIOP2: https://openneuro.org/datasets/ds002790

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CHBMP. The Cuban Human Brain Mapping Project (CHBMP) repository is an open multimodal neuroimaging and cognitive dataset available for registered users on the LORIS repository which is part of the MNI neuroinformatics ecosystem at: https://chbmp-open.loris.ca/

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PNC. The Philadelphia Neurodevelopmental Cohort (PNC) data are available at:
https://www.med.upenn.edu/bbl/philadelphianeurodevelopmentalcohort.html

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