The contributions of host population size and maternal transmission rate to fluctuating *Wolbachia* frequencies

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Keywords: host-microbe interaction, endosymbiosis, vertical transmission, *w*Mel, *Wolbachia*, *Drosophila*

Running Head: Fluctuating *Wolbachia* frequencies

ABSTRACT

Many insects and other animals host maternally transmitted endosymbionts that alter host

fitness and reproduction. The facultative endosymbiont *Wolbachia* is found in roughly half of all

insect species, but population frequencies fluctuate across time and geography for unknown

reasons. This is especially true for *w*Mel *Wolbachia* in *Drosophila melanogaster* and related

"*w*Mel-like" strains that fluctuate over timescales as short as a month. Recent work suggests

these fluctuations may be due to temperature effects that disrupt maternal transmission in some

or all female hosts under certain conditions. We simulated temporal *Wolbachia* dynamics in the

new R package *wlbcmodeler* to test whether fluctuations can be solely explained by finite-host

population stochasticity or, alternatively, by perturbation of maternal transmission rates. Across a

range of biologically plausible parameters, we show that declining transmission rates increase

variance and decrease mean *Wolbachia* frequencies over time, and these effects outweigh the

impact of *Wolbachia* effects on host fitness (i.e., fecundity) and *Wolbachia*-induced cytoplasmic

incompatibility. Large temporal fluctuations can arise due to stochasticity or changes to maternal

transmission rates that alter *Wolbachia* equilibria. These simulated dynamics may explain

fluctuations of *w*Mel-like *Wolbachia* over years-long periods, but additional factors (e.g., insect

demography) likely contribute to fluctuations over shorter timescales.

INTRODUCTION

Animal life is characterized by interactions with microbes that span a continuum from beneficial to antagonistic. Many insects and other animals carry heritable endosymbionts that reside within their own cells. These microorganisms can alter basic aspects of host fitness, including reproduction, immune function, and nutrient acquisition (McFall-Ngai et al. 2013; Shropshire et al. 2020; Perreau and Moran 2022; Bennett et al. 2024; Hoffmann and Cooper 2024). Heritable endosymbionts often form close associations with female ovaries and the germline to ensure that microbes are maternally transmitted from one host generation to the next (Russell et al. 2019; Hague et al. 2024). This process of vertical-maternal transmission is a key determinant of the spread, persistence, and long-term evolutionary dynamics of endosymbiosis. Strict vertical maternal transmission aligns the fitness interests of both hosts and endosymbionts, which may lead to long-term, stable relationships that can approach the status of organelles (Wade 2007; McCutcheon and Moran 2012; Fisher et al. 2017; McCutcheon et al. 2019; Perreau and Moran 2022).

Ultimately, the ecological and evolutionary consequences of endosymbioses are determined by whether or not endosymbionts are prevalent in host populations. Many endosymbionts are facultative from the host perspective, such that a proportion of host individuals in the population do not carry endosymbionts due to imperfect maternal transmission (Mateos et al. 2006; Carrington et al. 2011; Hamm et al. 2014; Corbin et al. 2017; Hague et al. 2020b). The frequency of facultative endosymbionts in host populations can vary widely across host systems, geography, and time (Hamm et al. 2014; Kriesner et al. 2016; Cooper et al. 2017; Hague et al. 2020b), suggesting that variable maternal transmission rates contribute to variation in endosymbiont prevalence. For instance, if maternal transmission is perturbed by

environmental conditions (e.g., temperature), high rates of imperfect maternal transmission may

cause endosymbiont frequencies to decline in certain host populations (Hague et al. 2020b, 2022;

Ross et al. 2020). Nonetheless, the relative contribution of imperfect maternal transmission to

variation in endosymbiont prevalence is poorly understood, especially in the context of other

important contributing factors, such as endosymbiont effects on host fitness and stochastic

dynamics associated with host population size (Caspari and Watson 1959; Hoffmann et al.

1990).

Maternally transmitted *Wolbachia* are the most common endosymbionts on earth,

infecting roughly half of all insect species, as well as other arthropods and nematodes (Ferri et al.

2011; Zug and Hammerstein 2012; Weinert et al. 2015). *Wolbachia* are vertically-maternally

transmitted through the female germline (Russell et al. 2019; Porter and Sullivan 2023;

Radousky et al. 2023; Hague et al. 2024); although, phylogenetic evidence indicates that

horizontal and introgressive transfers occur frequently on evolutionary timescales (O'Neill et al.

1992; Raychoudhury et al. 2009; Conner et al. 2017; Gerth and Bleidorn 2017; Turelli et al.

2018; Cooper et al. 2019; Vancaester and Blaxter 2023; Shropshire et al. 2024). *Wolbachia*

generally form facultative relationships with insects (from the host perspective), as the

endosymbionts are usually found in only a proportion of host individuals due to imperfect

maternal transmission. *Wolbachia* frequencies have been shown to vary considerably among and

within host systems (Hoffmann et al. 1990; Kriesner et al. 2013, 2016; Hamm et al. 2014;

Cooper et al. 2017; Hague et al. 2020b, 2022; Turelli et al. 2022); however, the relative

contribution of imperfect maternal transmission to this frequency variation is unclear.

Population frequencies and equilibria dynamics of *Wolbachia* can be approximated by a

discrete generation, deterministic model that incorporates three parameters: (1) the proportion of

uninfected ova produced by *Wolbachia*-positive females (μ; i.e., imperfect maternal

transmission), (2) the fitness of *Wolbachia*-positive females relative to *Wolbachia*-negative

females (F; i.e., components of host fitness like fecundity), and (3) the relative egg hatch of

uninfected eggs fertilized by *Wolbachia*-positive males (s_h) due to *Wolbachia*-induced

cytoplasmic incompatibility (CI) (Hoffmann et al. 1990). The initial spread and persistence of

Wolbachia in natural host populations is generally considered to be driven by beneficial effects

on host fitness ($F > 1$), which are balanced by imperfect maternal transmission ($\mu > 0$)

(Hoffmann et al. 1990; Turelli and Hoffmann 1991; Kriesner et al. 2013; Kriesner and Hoffmann

2018; Ross et al. 2019*b*; Turelli et al. 2022). Although beneficial *Wolbachia* effects on host

fitness are poorly understood in nature, recent work suggests *Wolbachia* block viruses in their

native *Drosophila* hosts (Hedges et al. 2008; Teixeira et al. 2008; Osborne et al. 2009; Martinez

et al. 2014; Cogni et al. 2021; Bruner-Montero and Jiggins 2023), in addition to a potential role

of nutrient provisioning (Brownlie et al. 2009; Nikoh et al. 2014; Newton and Rice 2020). Some

Wolbachia strains also cause CI ($s_h > 0$), a crossing incompatibility that generates a frequency

dependent benefit for *Wolbachia*-positive females (Shropshire et al. 2020; Turelli et al. 2022).

Once a CI-causing strain exceeds a critical frequency threshold, *Wolbachia* will rapidly spread to

high frequencies, which are balanced by imperfect maternal transmission (Hoffmann and Turelli

1988; Hoffmann et al. 1990; Turelli and Hoffmann 1995; Carrington et al. 2011; Kriesner et al.

2013; Turelli et al. 2022).

In host systems where *Wolbachia* do not cause strong CI, infection frequencies tend to

fluctuate across time and geographic space (Kriesner et al. 2013, 2016; Hamm et al. 2014;

Cooper et al. 2017; Hague et al. 2020*b*). For instance, *w*Mel in *Drosophila melanogaster* and the

closely related ("*w*Mel-like") *w*San and *w*Yak strains in *D. yakuba*-clade hosts (diverged from

*w*Mel ~30,000 years ago) generally cause no or weak CI. The frequencies of these *w*Mel-like

Wolbachia can fluctuate dramatically over time periods as short as a month at a given location

(Figure 1) (Hoffmann et al. 1998; Reynolds and Hoffmann 2002; Cooper et al. 2017, 2019). For

example, at Gold Coast in subtropical eastern Australia, the frequency (p) of *w*Mel in *D.*

melanogaster dropped from $p = 0.919$ (binomial confidence interval: 0.787, 0.972) in March

1995 down to $p = 0.25$ (0.12, 0.449) just a month later in April and then rebounded back up to p

$= 0.615$ (0.355, 0.823) in September (Hoffmann et al. 1998). Similarly, on the island of São

Tomé off the coast of west Africa, the frequency of *w*San in host *D. santomea* dropped from $p =$

0.772 (0.648, 0.862) in 2009 to $p = 0.368$ (0.234, 0.527) in 2015 and then back up to $p = 0.795$

(0.692, 0.87) in 2018 (Cooper et al. 2017; Hague et al. 2020*b*). The frequencies of *w*Mel and *D.*

yakuba-clade *Wolbachia* have also been shown to vary geographically among host populations

by latitude and altitude (Kriesner et al. 2016; Cooper et al. 2017; Hague et al. 2020*b*). These

variable patterns contrast those of strong CI-causing *Wolbachia* strains (e.g., *w*Ri in *D.*

simulans), which tend to persist at high, stable frequencies due to the frequency dependent

advantage of CI (Hoffmann et al. 1990; Turelli and Hoffmann 1991, 1995; Rousset and Solignac

1995; Bourtzis et al. 1996; James and Ballard 2000; Ballard 2004; Carrington et al. 2011; Choi et

al. 2015; Turelli et al. 2018).

The temporal fluctuations of *w*Mel-like *Wolbachia* could be strictly due to stochasticity

associated with a finite host population. Alternatively, changes in the maternal transmission rate,

host fitness effects, and/or CI between host generations could generate deterministic changes to

Wolbachia equilibria over time. We first explore the hypothesis that *Wolbachia* frequencies

fluctuate over time as a result of finite-population stochasticity (Jansen et al. 2008; Kriesner and

Hoffmann 2018; Turelli and Barton 2022; Turelli et al. 2022). The deterministic model of

Wolbachia equilibria can be adapted to approximate the stochasticity induced by finite

population size using a stochastic transition matrix analogous to a haploid Wright-Fisher model

of genetic drift (Crow and Kimura 1970; Turelli and Barton 2022; Turelli et al. 2022). The model

uses an effective population size of N female hosts (because *Wolbachia* are maternally

transmitted) and binomial sampling to model the stochastic effect of finite population size and

their interactions with maternal transmission (μ), host fitness effects (F), and CI (s_h). Following

Turelli et al. (2022) and Turelli and Barton (2022), we evaluate how host population size

influences the stochasticity of *Wolbachia* frequencies over time, measured as the standard

deviation of *Wolbachia* frequencies (p_{SD}), in simulated host populations across a wide range of

biologically plausible values for *w*Mel-like *Wolbachia*. We also explored how different values of

μ, in conjunction with F and s_h, impact the magnitude of stochastic fluctuations over time.

Second, we investigate the hypothesis that *Wolbachia* frequencies may change

deterministically over time as a result of temporal changes in the maternal transmission rate, host

fitness effects, or CI strength between host generations (Turelli and Hoffmann 1995; Carrington

et al. 2011; Hague et al. 2020*b*). Recent work on *w*Mel and *w*Yak showed that rearing hosts at a

relatively cool temperature (20ºC) reduces maternal transmission rates compared to standard

conditions (25ºC) (Hague et al. 2020*b*, 2022). Maternal transmission estimates from field-

collected females in these systems also indicate that a proportion of *Wolbachia*-positive females

may be "low transmitters" with especially high rates of imperfect maternal transmission ($\mu > 0.7$)

(Hoffmann et al. 1998; Carrington et al. 2011; Hamm et al. 2014; Hague et al. 2020*b*). This

phenomenon has been observed in wild-caught females transmitting *w*Mel (Hoffmann et al.

1998), *w*San, and *w*Yak (Hague et al. 2020*b*), as well as other *Wolbachia* strains (Carrington et

al. 2011; Hamm et al. 2014). Thus, it is conceivable that rapid changes in environmental

conditions (e.g., a bout of cold weather) could perturb maternal transmission for all hosts or a

subpopulation of low transmitters, causing deterministic changes to *Wolbachia* frequencies in

subsequent host generations. Changes in the strength of host fitness effects or CI could similarly

generate deterministic *Wolbachia* fluctuations, because these factors can also be context-

dependent, for example, depending on environmental conditions, host age, and the *Wolbachia*

and host genomes (Hoffmann et al. 1990; Clancy and Hoffmann 1998; Olsen et al. 2001;

Reynolds and Hoffmann 2002; Bordenstein and Bordenstein 2011; Versace et al. 2014; Kriesner

et al. 2016; Cooper et al. 2017; Cao et al. 2019; Shropshire et al. 2021). We used the stochastic

model to evaluate how biologically plausible values of μ, F, and s_h influence *Wolbachia*

equilibria, measured as the mean *Wolbachia* frequency ($\bar{p}$) from each simulated host population.

We implemented the simulations in the new publicly available R package *wlbcmodeler*.

The package simulates temporal *Wolbachia* dynamics within a host population with discrete

generations and a given effective population size. We used the R package to run 491,400

simulations across 3,276 unique parameter combinations to explore how biologically plausible

values of host population size ($10^3 \leq N \leq 10^6$), imperfect maternal transmission ($0.001 \leq \mu \leq 0.5$),

host fitness effects ($1 \leq F \leq 1.5$), and CI ($s_h = 0$ or 0.1) influence the temporal dynamics of

*w*Mel-like *Wolbachia*. The package treats μ as a random binomial variable and allows for the

inclusion of a subpopulation of low transmitter females (e.g., 10% of *Wolbachia*-positive

females with $\mu = 0.8$). We also allow for fluctuating host fitness effects by treating F as a log-

normal random variable (Turelli et al. 2022). We began each simulation at an arbitrary

intermediate *Wolbachia* frequency of $p_0 = 0.4$ and ran the population for 10,000 host generations.

After removing the initial 500 host generations as burn-in, we calculated the standard deviation

(p_{SD}) and mean ($\bar{p}$) *Wolbachia* frequency from generations 501-10,000 for each simulation.

Broadly, our results demonstrate that strains that cause weak or no CI, such as *w*Mel-like

Wolbachia, tend to generate larger stochastic fluctuations over time compared to strong-CI

causing strains like *w*Ri in *D. simulans*. While the simulated stochastic and/or deterministic

dynamics in our model may explain the observed fluctuations of *w*San on São Tomé over years-

long periods, other factors likely contribute to the fluctuations of *w*Mel at Gold Coast over much

shorter month-long periods.

THEORETICAL FRAMEWORK

Deterministic analysis of Wolbachia frequencies

167   The deterministic model of *Wolbachia* population frequencies and equilibria dynamics

incorporates imperfect maternal transmission (μ), host fitness effects modeled as differential

fecundity (F), and CI (s_h) (Hoffmann et al. 1990). We assume that uninfected ova produced by

Wolbachia-positive females (due to imperfect maternal transmission) are susceptible to CI, just

as uninfected ova from *Wolbachia*-negative females (Carrington et al. 2011). Embryos produced

by fertilizations of *Wolbachia*-negative ova with sperm from *Wolbachia*-positive males hatch

with frequency $H = 1 - s_h$ relative to the other three possible fertilizations, which are all

considered equally compatible (Cooper et al. 2017). Thus, s_h represents the severity of CI, or the

frequency of unhatched eggs in pairings between *Wolbachia*-negative ova and *Wolbachia*-

positive sperm.

177   We first ignore CI and consider a stable equilibrium ($\hat{p}$) balanced by positive host fitness

effects ($F > 1$) and imperfect maternal transmission ($\mu > 0$) (Figure 2). When initially rare,

Wolbachia must generate $F(1 - \mu) > 1$ to spread deterministically from low frequency, regardless

of whether they cause CI. Here, we assume $F(1 - \mu) > 1$ given the spread and persistence of non-

CI causing *Wolbachia* strains in nature (Hoffmann and Turelli 1997; Kriesner et al. 2013, 2016;

Hamm et al. 2014; Kriesner and Hoffmann 2018; Meany et al. 2019). In contrast, when $F(1 - \mu)$

< 1, rare infections will be deterministically lost (Hoffmann et al. 1990). A value of $F(1 - \mu) > 1$

also provides an explanation for the low frequency spread of strong CI-causing strains,

particularly *w*Ri (Turelli and Hoffmann 1991, 1995; Bakovic et al. 2018). When $F(1 - \mu) > 1$, the

stable equilibrium frequency is

$$\hat{p} = 1 - \frac{\mu F}{F - 1}. \tag{1}$$

In this case, $\hat{p}$ increases from 0 to $1 - \mu$ as F increases from $1/(1 - \mu)$.

We explored a range of plausible μ values based on previous estimates of maternal

transmission in the field, as well as under varying conditions in the lab. Empirical estimates of

mean μ values ($\pm BC_a$ confidence intervals) for *w*Mel generated from field-collected *D.*

melanogaster have ranged from low ($\mu = 0.026$ [0.008, 0.059]; Hoffmann et al. 1998) up to

moderate levels of imperfect transmission ($\mu = 0.11$ [0.07, 0.17]; Olsen et al. 2001). Estimates

generated in the lab range as high as $\mu = 0.415$ (0.332, 0.493) when females are reared in the

cold (20ºC) relative to a standard temperature (25ºC) (Hague et al. 2022, 2024). Estimates of

mean μ values for *w*San from field collected *D. santomea* on the island of São Tomé were

moderate ($\mu = 0.068$ [0.027, 0.154]), whereas *w*Yak in *D. yakuba* ranged from low at low

altitude ($\mu = 0.038$ [0.003, 0.184]) to moderate at high altitude ($\mu = 0.20$ [0.087, 0.364]) (Hague

et al. 2020*b*). Estimates of μ generated under lab conditions for *w*Yak in *D. yakuba* ranged as

high as $\mu = 0.15$ (0.109 0.196) when females were reared in the cold (Hague et al. 2020*b*). We

thus decided to explore a broad range of μ values from $0.001 \leq \mu \leq 0.5$.

We know far less about *Wolbachia* effects on host fitness (F) in natural populations due

to the challenges of estimating fitness of *Wolbachia*-positive and -negative individuals in the

wild (Ross et al. 2019*b*). *Wolbachia* cells are found throughout host somatic tissue (Pietri et al.

2016), which is likely to have complex effects on components of host fitness in nature

(Hoffmann et al. 1998; Weeks et al. 2007; Kriesner et al. 2016; Hague et al. 2020*a*, 2021; Cogni

et al. 2021; Bruner-Montero and Jiggins 2023). Prior work found no evidence for fecundity

effects of *w*Mel in *D. melanogaster* (Hoffmann et al. 1994) or *Wolbachia* in *D. yakuba*-clade

hosts (Cooper et al. 2017); however, fecundity represents only one aspect of host fitness.

Wolbachia effects on other measures of host fitness in the lab have ranged from negative

(Hoffmann and Turelli 1988; Hoffmann et al. 1990; Kriesner et al. 2016) to positive (Hedges et

al. 2008; Teixeira et al. 2008; Brownlie et al. 2009; Russell et al. 2023) to context-dependent

(Olsen et al. 2001; Fry et al. 2004). Beneficial *Wolbachia* effects of $F > \sim 1.2$ (roughly a 20%

relative fitness advantage for *Wolbachia*-positive females) have not been documented in any

system (Weeks et al. 2007; Hamm et al. 2014; Meany et al. 2019; Hague et al. 2020*b*). We

generally expect $F > 1$ based of the spread and persistence of *Wolbachia* strains in natural host

populations (described above); however, we expect values of $F > 1.5$ to be biologically

unrealistic (Weeks et al. 2007; Hamm et al. 2014; Cooper et al. 2017; Meany et al. 2019; Hague

et al. 2020*b*, 2022). Therefore, we chose to explore F values ranging from $1 \le F \le 1.5$.

Next, we considered equilibria that incorporate CI ($s_h > 0$) (Figure 2). Like non-CI

causing strains, CI-causing strains must still increase host fitness to spread deterministically from

low frequency such that $F(1 - \mu) > 1$. CI-causing infections that generate $F(1 - \mu) > 1$ and $F\mu <$

1 produce a single stable equilibrium between 0 and 1 given by

$$\hat{p} = \frac{s_h + 1 - F + \sqrt{(s_h + 1 - F)^2 + 4s_h[F(1 - \mu) - 1](1 - F\mu)}}{2s_h(1 - F\mu)}.$$
[2]

Prior estimates of CI using field-collected *D. melanogaster* males from central and northern

Australia suggest that *w*Mel causes weak CI on the order of $s_h = 0.05$ (Hoffmann et al. 1998;

Kriesner et al. 2016). CI strength declines rapidly with male age (Shropshire et al. 2021) and

Reynolds and Hoffmann (2002) found that one-day-old males derived from wild-collected larvae

and pupae produce hatch rates of 0.39 in incompatible crosses, with large confidence intervals

(0.15, 0.64). Given this information, Kriesner et al. (2016) conjectured that s_h values for wMel of

0 to 0.1 are plausible, but $s_h > 0.1$ is unlikely. We are not aware of CI estimates for *D. yakuba-*

clade *Wolbachia* using field-collected males; however, estimates in the laboratory indicate that

these strains reduce egg-to-adult viability in CI crosses by about 10-20% (wSan: $s_h = 0.15$ [0.12,

0.18]; wYak: $s_h = 0.16$ [0.13, 0.20]; wTei: $s_h = 0.2$ [0.17, 0.23]), although the presence and

intensity of CI can also vary depending on host and *Wolbachia* genotype (Cooper et al. 2017).

Based on these previous data, we decided to explore the effects of weak CI ($s_h = 0.1$) on

Wolbachia equilibria and temporal dynamics. We also considered strong CI ($s_h = 0.45$) that is

characteristic of field-collected male *D. simulans* carrying wRi (Turelli and Hoffmann 1995;

Carrington et al. 2011); although, this level of CI has not been observed for the wMel-like

Wolbachia described here.

Assuming discrete host generations, the total population fraction of offspring in the next

generation can be divided with respect to whether mated males and females are carrying

Wolbachia:

1. *Wolbachia-positive male and female.* In this case, there is a population of $F(1 - \mu)p_t^2$

and also a population of $F\mu(1 - s_h)p_t^2$ when ova become uninfected due to imperfect

maternal transmission, resulting in a population fraction of $F(1 - \mu s_h)p_t^2$.

2. *Wolbachia-negative male and female.* The population fraction is simply $(1 - p_t)^2$.

3. *Wolbachia-negative male, Wolbachia-positive female.* Accounting for fecundity in

Wolbachia-positive females, the population fraction is then $Fp_t(1 - p_t)$.

4. *Wolbachia-positive male, Wolbachia-negative female*. The population from generation

t is $p_t(1 - p_t)$. A fraction $1 - s_h$ of this population survives from this incompatible

pairing, so the population fraction for generation $t + 1$ is $p_t(1 - p_t)(1 - s_h)$.

Assuming equal infection frequencies in males and females, the adult infection frequency in

generation t, denoted p_t, changes between generations as follows:

$$p_{t+1} = \frac{p_t F(1 - \mu)}{1 + p_t(F - 1 - s_h) + p_t^2 s_h(1 - \mu F)} \qquad [3]$$

253 (Hoffmann et al. 1990). In the absence of CI ($s_h = 0$), this model becomes

$$p_{t+1} = \frac{p_t F(1 - \mu)}{1 + p_t(F - 1)}. \qquad [4]$$

For a stability analysis of [4], see Appendix 1.

Simulations to approximate temporal Wolbachia dynamics under finite population size

Following Turelli and Barton (2022) and Turelli et al. (2022), we approximated the

stochasticity produced by finite population size using a stochastic transition matrix analogous to

a haploid Wright-Fisher model of genetic drift, which assumes that generations are discrete and

the new generation is a random sample of constant population size drawn from the total number

of offspring produced by the previous generation (Crow and Kimura 1970). The model uses an

effective population size of N females. The finite-population stochasticity, modeled as binomial

sampling, is superimposed on the deterministic equilibria dynamics described by [3]. We

explored a range of biologically plausible host population sizes ranging from 10^3 to 10^6

(McKenzie 1980; McInnis et al. 1982; Powell 1997; Gravot et al. 2004; Bergland et al. 2014).

Estimates of local population sizes of *Drosophila* made from mark-release-recapture methods

report local census sizes on the order of 10^4 to 10^5, with considerable variation among seasons,

years, and localities (McKenzie 1980; McInnis et al. 1982; Powell 1997; Gravot et al. 2004;

Bergland et al. 2014).

We now derive a model for modeling infection frequency p_t of *Wolbachia*-positive hosts

which includes randomness effects from a finite female population of size N, $I_t = Np_t$ of which

are infected. We assume females produce the same number of eggs with each mating and that N

does not change over generations. Starting with the current adult (female) infection frequency p_t,

the infection frequency among viable gametes in the next generation is determined by [3]. Then,

the infection frequency in the next generation of N adult females is obtained from binomial

sampling of this deterministic projection.

First, we compute the total number of *Wolbachia*-positive adult offspring in generation

$t + 1$ in the absence of CI ($s_h = 0$). For *Wolbachia*-positive females, we note that there are FI_t

ova produced. Each ovum has a probability $1 - \mu$ of being infected, so the total number of

infected offspring can be represented by a binomial random variable

$$X_t \sim \text{Bin}\big(FI_t, (1 - \mu)\big). \tag{5}$$

281   For the total number of offspring produced, we observe that all *Wolbachia*-negative females

282   produce *Wolbachia*-negative ova. This gives a total of $N - I$ *Wolbachia*-negative offspring for

283   the next generation. For *Wolbachia*-positive females, we noted that FI_t offspring are produced

284   from matings. This gives a total population of $N + (F - 1)I_t$ offspring. The proportion of

285   *Wolbachia*-positive adults for generation $t + 1$ is then

$$p_{t+1} = \frac{X_t}{FI_t + N - I_t} = \frac{X_t/N}{p_t(F - 1) + 1}. \tag{6}$$

286   From [6], the expected value and variance of p_{t+1}, given knowledge of p_t, are readily computed

287   as

$$\mathbb{E}[p_{t+1}|p_t] = \frac{F(1-\mu)p_t}{p_t(F-1)+1}, \quad \text{Var}(p_{t+1}|p_t) = \frac{F\mu(1-\mu)p_t}{N(p_t(F-1)+1)^2}. \tag{7}$$

From the law of large numbers, as the population $N \to \infty$, the variance in [7] tends to 0, and the

random recursion [6] converges to the deterministic model [4].

Now, allowing for CI ($s_h > 0$), we can derive a random recursion for p_t. For each of the

stages of mating, ova production, and the generation of adult offspring, we record total

populations in Table S1. The total *Wolbachia*-positive adult offspring resulting from *Wolbachia*-

positive male and female matings is given by $X_1 \sim \text{Bin}(FpI, (1-\mu))$. For *Wolbachia*-negative

males and -positive females, total *Wolbachia*-positive offspring is $X_2 \sim \text{Bin}(F(1-p)I, (1-$

$\mu))$. The total number of offspring is found by summing the four terms in the row labeled "adult

offspring after CI" in Table S1. Denoting the normalized fractions $\tilde{X}_i = X_i/N$, the fraction of

Wolbachia-positive offspring over the total population is then

$$p_{t+1} = \frac{\tilde{X}_1 + \tilde{X}_2}{s_h(1-F)p_t^2 + (F-s_h-1)p_t + s_h\tilde{X}_1 + 1}. \tag{8}$$

As $N \to \infty$, from the law of large numbers [8] approaches the deterministic model [3].

Let us also consider the case where a subpopulation of "low transmitter" females with

high rates of imperfect maternal transmission are included in the host population. Here, the host

population consists of M groups with group fractions $(r_1, \dots, r_M)$. Rather than a single μ value,

the ith subpopulation assumes a value μ_i for $i = 1, \dots, M$. To compute *Wolbachia*-positive

offspring, we are now summing M separate binomial distributions. However, most computations

from above carry over. For each group i, we compute $X_i \sim \text{Bin}(FIr_i, 1-\mu_i)$. The number of

Wolbachia-positive female eggs then becomes $S^M = \sum_{i=1}^{M} X_i$. For no CI, the total population is

still $FI + N - I$, and

$$p_{t+1} = \frac{S^M_{t+1}}{FI_t + N - I_t} = \frac{S^M_{t+1}/N}{p_t(F-1)+1}.$$ [9]

307 When $\mu_i \equiv \mu$ for all $i = 1, ..., M$, then $S^M \sim \text{Bin}(1-\mu, FI)$ and we recover [6]. Expected values

308 are also straightforward to compute, with

$$\mathbb{E}[p_{t+1}|p_t] = \frac{F(1-\sum_{i=1}^M \mu_i r_i)p_t}{p_t(F-1)+1}.$$ [10]

309 Each of the *Wolbachia*-positive populations X_i for $i ..., M$ terms are independent, so by additivity

310 of variance,

$$\text{Var}(p_{t+1}|p_t) = \frac{F(\sum_{i=1}^M \mu_i(1-\mu_i)r_i)p_t}{N(p_t(F-1)+1)^2}.$$ [11]

311 The stable equilibrium, where $\hat{p} = \mathbb{E}[p_{t+1}|p_t = \hat{p}]$, is very similar to [1]. Accounting for

312 multiple μ_i values, we have

$$\hat{p} = 1 - \frac{F}{F-1}\sum_i \mu_i r_i.$$ [12]

313 Finally, to generalize [8], we replace X_1 and X_2 with sums of independent binomials $\{Y_i\}_{1\leq i\leq M}$

314 and $\{Z_i\}_{1\leq i\leq M}$, with

$$X_1 = \sum_{i=1}^M Y_i, \qquad Y_i \sim \text{Bin}(Fpr_iI, 1-\mu_i),$$ [13]

$$X_2 = \sum_{i=1}^M Z_i, \qquad Z_i \sim \text{Bin}(F(1-p)r_iI, 1-\mu_i).$$ [14]

Previous estimates of μ from a large sample of field-collected female *D. santomea* ($N =$

62) and *D. yakuba* ($N = 71$) on São Tomé in 2018 suggest the majority of females have perfect or

near perfect transmission rates, but a small proportion of low transmitters have high values of μ.

For *w*San, the total sample of *Wolbachia*-positive females had a mean value of $\mu = 0.068$ (0.027,

0.154), but 8% of these females had a μ value greater than 0.5 with a mean of $\mu = 0.728$.

Similarly, for *w*Yak, the total sample of *Wolbachia*-positive females had a mean value of $\mu =$

0.126 (0.062, 0.238), but approximately 11% of the females had a μ value greater than 0.5 with a

mean of $\mu = 0.803$ (Hague et al. 2020a). Previous work using field-collected females found

similar evidence of low transmitters for wMel in *D. melanogaster* (Hoffmann et al. 1998), wSuz

in *D. suzukii* (Hamm et al. 2014; see Figure 3), wInn in *D. innubila* (Unckless et al. 2009), and

wRi in *D. simulans* (Carrington et al. 2011; see Figure 6). Hamm et al. (2014) termed females

with high rates of imperfect wSuz transmission as "low transmitters" and we follow the same

terminology here. Using the 2018 *D. yakuba*-clade *Wolbachia* data as a guide, we explored the

effect of treating a 10% subpopulation of *Wolbachia*-positive females as low transmitters with a

random binomial μ variable set to $\mu = 0.8$. The rest of the *Wolbachia*-positive females had μ

values set as described above ($0.001 \leq \mu \leq 0.5$).

While we initial assumed a fixed F value, we also considered the case where host fitness

effects fluctuate from one host generation to the next as independent, identically distributed log-

normal random variables (Turelli et al. 2022). We assume that each generation F is chosen

independently from a log-normal distribution such that $F = e^X$, where X is a normal random

variable with mean μ_X and variance σ^2 (Turelli et al. 2022). This implies that F has median $m = $

e^{μ_x} and a squared coefficient of variation $CV^2 = \text{Var}(F)/[\text{E}(F)^2] = e^{\sigma^2} - 1$. To produce a

particular median, m, and CV for F, we set $\mu_x = \ln(m)$ and $\sigma^2 = \ln(CV^2 + 1)$. While it is

generally unknown whether host fitness effects fluctuate in natural host populations (Turelli et al.

2022), there is evidence that the fitness effects of wMel on *D. melanogaster* fecundity (Kriesner

et al. 2016) and virus-blocking (Chrostek et al. 2021) can depend on the environmental context.

We explored scenarios with weakly ($CV = 0.01$) and strongly ($CV = 0.1$) fluctuating host fitness

effects. To illustrate how these values correspond to fluctuating F values in our simulations,

when median(F) = 1.05, a CV = 0.01 value corresponds to 2.5 and 97.5 percentiles of 1.03 and

1.071, respectively, and a CV = 0.1 value corresponds to 0.863 and 1.271, respectively.

We ran a set of 25 simulations for every unique combination of parameters. Each

individual simulation began with the arbitrary intermediate *Wolbachia* frequency of p_0 = 0.4 and

ran for 10,000 host generations (similar p_0 values do not qualitatively alter the results). We

removed an initial burn-in of 500 host generations as *Wolbachia* spread from p_0 = 0.4 to a stable

equilibrium, and then calculated the standard deviation (p_{SD}) and mean ($\bar{p}$) *Wolbachia* frequency

from generations 501-10,000 for each simulation. To quantify the influence of the different

parameter values on the outcome of simulations, we modeled the p_{SD} and $\bar{p}$ values with linear

effects for N, μ, the presence/absence of low transmitters, F, the presence/absence of fluctuating

F values (i.e., treating F as fixed or a random variable), and s_h. Since the frequency is always

between zero and one, we modeled the distribution of the p_{SD} and $\bar{p}$ values as beta random

variables. Models were implemented in R version 4.4.2 using the *mgcv* package version 1.9.1

(Wood 2017). We found that a small handful of simulations had infections that persisted at $p > 0$

for the full 10,000 host generations for combinations of parameter values that do not predict a

stable equilibrium of $\hat{p} > 0$ (e.g., F = 1.05, μ = 0.05, s_h = 0). We removed these simulations from

our regression analysis, because they are predicted to be deterministically lost over time and we

did not want the p_{SD} and $\bar{p}$ values to bias our analyses of stable equilibria. Finally, we used the

"plot_predictions" function from the *marginaleffects* package version 0.25 (Arel-Bundock et al.

2024) to generate model predictions for p_{SD} and $\bar{p}$ as a function of the μ value modulated by the

presence/absence of low transmitters, F, and s_h.

RESULTS

We ran 491,400 simulations to examine how different parameter combinations of host population size, imperfect maternal transmission, host fitness effects, and CI contribute to stochastic fluctuations (measured as p_{SD}) and *Wolbachia* equilibria (measured as $\bar{p}$) over time in simulated host populations (Figures S1, S2). Multiple regression was used to summarize how N, μ, F, and s_h impacted p_{SD} (Table 1, Figure 3) and $\bar{p}$ (Table 2, Figure 4) in our simulations. We first explore how different parameters values influence the size of stochastic fluctuations over time (p_{SD}) and then examine how temporal changes to parameter values, particularly the maternal transmission rate, between host generations alter deterministic *Wolbachia* equilibria ($\bar{p}$). For both scenarios, we consider whether the dynamics in our simulations can explain empirical examples of temporally fluctuating *w*Mel-like *Wolbachia* in Australia and São Tomé.

All the variables in our regression model had a significant effect on stochastic fluctuations measured as p_{SD}; however, some parameters were more influential than others (Table 1, Figures 3, S1). Notably, host population size had only a small effect (β = -4.14 x 10^{-7}, P < 0.001), although p_{SD} generally declined as population size increased and *Wolbachia* frequencies hovered closer to predicted equilibria. Given that host population size had only a minor effect on p_{SD}, we focus here on our results from a single biologically plausible population size of $N = 10^4$ as a simple case to illustrate how other parameters contribute to p_{SD} (Figures 5, S3). Broadly, our results illustrate an intuitive pattern where the parameter values that yield high mean infection frequencies—high rates of maternal transmission, strong fitness benefits, and CI—tend to generate lower p_{SD} values and smaller fluctuations. In contrast, the parameters that yield low infection frequencies ($\bar{p}$ < 0.2) typically generate the largest *Wolbachia* fluctuations over time. Within the biologically plausible range of μ, F, and s_h values that we explored, μ had

the strongest relative effect on p_{SD} ($\beta = 8.51$, $P < 0.001$), as lower maternal transmission rates

(i.e., larger μ values) are associated with lower infection frequencies and larger fluctuations. To a

lesser extent, the inclusion of a subpopulation (10%) of low transmitter females ($\mu = 0.8$) also

generally increased p_{SD} values ($\beta = 1.12$, $P < 0.001$). Here, we also found an interaction effect

between the value of μ and the presence of low transmitters ($\beta = -4.42$, $P < 0.001$), such that the

rate of increase in p_{SD} as a function of μ is smaller when low transmitters are present. The value

of F had a smaller relative effect on p_{SD} than μ ($\beta = -2.04$, $P < 0.001$), wherein stronger host

fitness benefits are associated with higher infection frequencies and smaller fluctuations. If we

instead treat F as a log-normal random variable and assume fluctuating host effects, weak

fluctuations ($CV = 0.01$) only had a minor effect on increasing p_{SD} values ($\beta = 0.423$, $P < 0.001$),

whereas strong fluctuations predictably had a larger effect ($\beta = 1.78$, $P < 0.001$). Finally, the

addition of weak CI at a biologically plausible level for wMel-like *Wolbachia* ($s_h = 0.1$) had a

relatively small effect on p_{SD} values ($\beta = -9.85 \times 10^{-2}$, $P < 0.001$). CI generally causes higher

infection frequencies and smaller fluctuations. Strong CI characteristic of strains like wRi ($s_h = $

0.45) had a larger effect on p_{SD} ($\beta = -0.783$, $P < 0.001$), as strong CI is associated with high,

stable infection frequencies (Figure S3).

405   Can these stochastic dynamics explain empirical observations of temporally fluctuating

wMel-like *Wolbachia*? At Gold Coast in Australia, wMel frequencies dropped from $p = 0.919$

(0.787, 0.972) in March 1995 down to $p = 0.25$ (0.12, 0.449) in April and then back up to $p = $

0.615 (0.355, 0.823) in September, representing a maximum absolute change in frequencies (p_{max}

- p_{min}) of 0.669 occurring over just a month-long period (Figure 1). On São Tomé, wSan

frequencies dropped from $p = 0.772$ (0.648, 0.862) in 2009 to $p = 0.368$ (0.234, 0.527) in 2015

and then back up to $p = 0.795$ (0.692, 0.87) in 2018, representing a maximum frequency change

of 0.427 over the course of a six-year period (Figure 1). Generally, the simulations assuming

fixed or weakly fluctuating host effects ($CV = 0.01$) produced small frequency fluctuations that

cannot explain wMel or wSan (Figures 6, S4). Assuming a host population size of $N = 10^4$, we

found that the parameter combination of $\mu = 0.1$ (without low transmitters), a fixed value of $F =$

1.125, and $s_h = 0.1$ produced the largest frequency fluctuations, with an absolute difference

between the maximum and minimum frequency of 0.07 in an individual simulation spanning

10,000 host generations. Adding weakly fluctuating host fitness effects ($CV = 0.01$) produced

only a slightly larger maximum absolute difference of 0.12 from the parameters $\mu = 0.05$

(without low transmitters), $F = 1.05$, and $s_h = 0.1$. Only when we assume strongly fluctuating

fitness effects ($CV = 0.1$) do we find large fluctuations that can potentially account for empirical

observations of wMel and wSan. The parameters $\mu = 0.01$ (without low transmitters), $F = 1.025$,

and $s_h = 0$ produced the largest absolute frequency difference of 0.855 between p_{max} and p_{min}

over the course of a single simulation (Figure S5).

425   These results highlight a pattern where the specific combination of parameters that

generate the largest p_{SD} values does not necessarily align strictly with the general trends outlined

above in the regression analysis. In the absence of strong fluctuating host effects, simulations

with CI ($s_h = 0.1$) and intermediate values of μ ($0.05 \leq \mu \leq 0.15$) generated the highest p_{SD} values

(Figure 5A). With the addition of strong fluctuating host effects, simulations with no CI ($s_h = 0$)

and small μ values ($0.005 \leq \mu \leq 0.01$) combined with small F values ($F = 1.025$) generated the

highest p_{SD} values (Figure 5B). In this region of parameter space, no CI and the presence of

strongly fluctuating F values with a median slightly greater than one cause *Wolbachia* to

fluctuate between strongly favored ($F[1 - \mu] > 1$) and disfavored ($F[1 - \mu] < 1$) in the host

population (Figure S5). With a median of $F = 1.025$ and $CV = 0.1$, the 2.5 and 97.5 percentiles

435 are 0.843 and 1.251, respectively. Only the simulations with no CI and strong fluctuating host

436 effects ($CV = 0.1$) produced absolute frequency differences ≥ 0.669 in individual simulations that

437 were on the scale of wMel fluctuations observed at Gold Coast. Because the wSan fluctuations at

438 São Tomé were not quite as large, both the no CI and weak CI simulations with strongly

439 fluctuating host effects produced maximum absolute frequency differences ≥ 0.427 that are

440 comparable to wSan. Without CI, the presence of a low transmitter subpopulation generally

441 reduced maximum observed frequency differences (Figure 6) and low transmitters cannot be

442 present in order to generate maximum frequency differences ≥ 0.669 comparable to wMel. In

443 fact, the values of $0.005 \leq \mu \leq 0.01$ (without low transmitters), $F = 1.025$ ($CV = 0.1$), and $s_h = 0$

444 were the only combinations of parameters that generated absolute frequency differences ≥ 0.669

445 as observed for wMel at Gold Coast.

446   We also examine how the values of μ, F, and s_h influenced *Wolbachia* equilibria as

447 measured by $\bar{p}$ (Figures 4, S2) to explore the alternative possibility that changes to these

448 parameter values between host generations could generate deterministic increases/decreases in

449 *Wolbachia* frequencies over time that explain wMel-like *Wolbachia* dynamics. Again, we used

450 multiple regression to summarize how different values of N, μ, F, and s_h impacted $\bar{p}$ in our

451 simulations (Table 2). Host population size had a very small relative effect on $\bar{p}$ values ($\beta = -1.58$

452 x 10^{-8}, $P = 0.009$). As expected, our results otherwise align tightly with equilibria predictions

453 from the strictly deterministic model (Figure 2), which shows that high rates of transmission

454 (smaller μ values), strong beneficial host effects (larger F values), and strong CI (larger s_h

455 values) yield high equilibrium frequencies. Within the biologically plausible range of parameters

456 we explored, the value of μ had the strongest relative effect on $\bar{p}$ ($\beta = -22.4$, $P < 0.001$). To a

457 lesser extent, the inclusion of a subpopulation of low transmitter females was also associated

with lower $\bar{p}$ values ($\beta = -2.43$, $P < 0.001$). We also found an interaction effect between the μ

value and the presence of low transmitters ($\beta = 8.26$, $P < 0.001$), such that the rate of decrease in

$\bar{p}$ as a function of μ is slightly less when low transmitters are present. The value of F had a

smaller relative effect on $\bar{p}$ compared to μ ($\beta = 3.77$, $P < 0.001$) and the addition of fluctuating

host fitness effects had only a minor effect on $\bar{p}$ ($CV = 0.01$: $\beta = 1.30$ x 10^{-4}, $P = 0.961$; $CV = 0.1$:

$\beta = 9.68$ x 10^{-3}, $P < 0.001$). Finally, the addition of weak CI had a small effect on $\bar{p}$ ($\beta = 0.427$, P

< 0.001), whereas strong CI had a much stronger effect ($\beta = 1.92$, $P < 0.001$), as the frequency

dependent benefit of CI drives *Wolbachia* to high equilibrium frequencies.

466   While including a subpopulation of low transmitters ($\mu = 0.8$) generally reduced $\bar{p}$ values,

we note that this effect was strongest for *Wolbachia* strains that do not cause CI (Figures 4, S2).

Again, we focus on the simple case of $N = 10^4$ to illustrate how different parameter values

contribute to $\bar{p}$ (Figures 5, S3). Here, with no CI ($s_h = 0$), the median $\bar{p}$ value dropped from 0.807

for all our simulations with a strictly binomial μ value to 0.472 for the simulations that included

a low transmitter subpopulation. In contrast, the median $\bar{p}$ value dropped from 0.875 to 0.591

with low transmitters for weak CI ($s_h = 0.1$) and from 0.945 to 0.858 for strong CI ($s_h = 0.45$).

The impact of low transmitters on $\bar{p}$ is especially strong for relatively small values of μ (≤ 0.01)

and F (≤ 1.15), which are otherwise predicted to generate high equilibrium frequencies (Figure

2). For instance, the values of $\mu = 0.001$, $F = 1.1$, and $s_h = 0$ generated a pooled mean of $\bar{\bar{p}} =$

0.989 across the set of 25 simulations, but the addition of low transmitters dropped the mean

down to $\bar{\bar{p}} = 0.11$.

478   Can changes to the maternal transmission rate or the presence of low transmitters

between host generations generate large deterministic drops in *Wolbachia* frequencies over time

on the scale observed for *w*Mel-like *Wolbachia*? We further examined the $\bar{\bar{p}}$ values from our

simulations (Figure 5C) to infer how changes to μ, F, or s_h between host generation could alter

equilibrium frequencies over time. We found that a range of parameter values could plausibly

explain the drop in wMel frequencies from $p = 0.919$ in March down to $p = 0.25$ in April of 1995

at Gold Coast if we assume that changes to μ values or the presence of low transmitters occur

between host generations (Figures S6). For instance, the values of $\mu = 0.005$, $F = 1.075$, and $s_h =$

0 produced a value of $\bar{\bar{p}} = 0.928$, which falls within the confidence interval of p for wMel in

March. Then, assuming a drop in the transmission rate to $\mu = 0.05$ in subsequent generations

(while all else remains equal), $Wolbachia$ frequencies drop to $\bar{\bar{p}} = 0.283$, which falls within the

confidence interval of p in April (Figure S7). Or, assuming the values of $\mu = 0.005$, $F = 1.125$,

and $s_h = 0$ in March ($\bar{\bar{p}} = 0.955$), and then adding a subpopulation of low transmitters in April ($\bar{\bar{p}}$

$= 0.239$) could also explain the drop. Of course, subtle changes in multiple parameters, including

F and s_h, could also explain the temporal drop in frequencies. For example, the values of $\mu =$

0.01, $F = 1.025$, and $s_h = 0.1$ could explain March ($\bar{\bar{p}} = 0.920$) and then a slight decrease in the

maternal transmission rate, increase in host fitness effects, and the loss of CI ($\mu = 0.05$, $F =$

1.075, and $s_h = 0$) could explain April ($\bar{\bar{p}} = 0.283$).

We found that a similar range of parameter values could plausibly explain the drop in

wSan from $p = 0.772$ in 2009 to $p = 0.368$ in 2015 on São Tomé (Figure S8). For instance, the

values of $\mu = 0.01$, $F = 1.075$, and $s_h = 0$ produced a value of $\bar{\bar{p}} = 0.857$, which falls within the

confidence interval of p for wSan in 2009. Then, assuming a drop in the transmission rate to $\mu =$

0.05 (while all else remains equal), $Wolbachia$ frequencies drop to $\bar{\bar{p}} = 0.283$, falling within the

2015 p estimate (Figure S9). Assuming the addition of low transmitters in 2015 could also

plausibly explain the drop in $Wolbachia$ frequencies. For example, the values of $\mu = 0.05$, $F =$

1.175, and $s_h = 0.1$ produced $\bar{\bar{p}} = 0.782$, and could explain 2009, and then adding a subpopulation

504 of low transmitters drops *Wolbachia* frequencies to $\bar{\bar{p}} = 0.284$, falling within the 2015 p estimate.

505 Subtle changes in multiple parameters could also plausibly explain the drop in frequencies. For

506 instance, the values of $\mu = 0.005$, $F = 1.025$, and $s_h = 0$ produced $\bar{\bar{p}} = 0.795$ and could explain

507 2009, whereas a slight drop in the maternal transmission rate and increase in host fitness effects

508 ($\mu = 0.05$, $F = 1.075$, and $s_h = 0$) produced $\bar{\bar{p}} = 0.283$ and could explain 2015.

510 **DISCUSSION**

511   We explored how finite host population size in conjunction with maternal transmission

512 rates, host fitness effects, and CI contribute to the temporal dynamics of *w*Mel-like *Wolbachia*.

513 Broadly, we find that strains like *w*Mel and *w*San that cause no or weak CI are more likely to

514 generate large stochastic fluctuations over time than strains like *w*Ri that cause strong CI. The

515 combination of no CI, host effects that fluctuate between mutualistic and deleterious, and a

516 modest maternal transmission rate generated the absolute largest stochastic fluctuations over

517 time in our model, regardless of population size. Changes between host generations to the

518 maternal transmission rate or the presence/absence of low transmitters can also generate large

519 deterministic changes to *Wolbachia* frequencies, as can between-generation changes to host

520 fitness effects and CI. The results suggest that stochastic dynamics and/or between-generation

521 changes to parameter values can provide plausible explanations for the temporal fluctuations of

522 *w*San on São Tomé over years-long periods, but the fluctuations of *w*Mel at Gold Coast are more

523 difficult to explain in the context of our model.

524   We first evaluated whether finite-population stochasticity alone can explain the temporal

525 dynamics of *w*Mel-like *Wolbachia*. Assuming the values of μ, F, and s_h remain fixed over time,

526 we found that stochasticity in our model generates only minor fluctuations over time, regardless

of host population size. Even small host population sizes (10^3) were generally outweighed by

deterministic dynamics, as *Wolbachia* frequencies tend to fluctuate closely around predicted

equilibria. Notably, across all the biologically plausible parameter values that we explored, poor

rates of maternal transmission (i.e., increasing values of μ) had the strongest relative effect on

increasing *Wolbachia* fluctuations in our simulations (Table 1). However, very specific

conditions were required to generate the absolute largest frequency fluctuations. Assuming fixed

host fitness effects, weak-CI causing *Wolbachia* ($s_h = 0.1$) with moderate fitness benefits (1.125

$\leq F \leq 1.2$) and relatively poor rates of maternal transmission ($0.1 \leq \mu \leq 0.2$) generated that

largest p_{SD} values (Figure 5A). Adding a subpopulation of low transmitters generally increased

the size of fluctuations; however, low transmitters were not present in the simulations that

produced the largest p_{SD} values. Only the addition of strongly fluctuating host fitness effects (*CV*

$= 0.1$) under a limited set of conditions can replicate fluctuations on the scale of *w*Mel-like

Wolbachia, particularly for *w*Mel at Gold Coast in Australia. Here, non-CI causing *Wolbachia*

with median F values close to one ($1.025 \leq F \leq 1.075$) combined with a high fidelity of

transmission ($0.005 \leq \mu \leq 0.05$) generated the largest p_{SD} values (Figure 5B), as *Wolbachia*

fluctuate between mutualistic ($F[1 - \mu] > 1$) and deleterious ($F[1 - \mu] < 1$) (Figure S5). In

contrast, conditions that yield equilibrium frequencies near one—strong CI, near perfect

maternal transmission, and large host fitness benefits—generate smaller stochastic fluctuations

over time.

These results are generally consistent with empirical observations of *Wolbachia*

frequency fluctuations in natural populations. *w*Mel-like *Wolbachia* that cause no or weak CI and

are imperfectly transmitted in nature tend to show evidence of temporally and/or spatially

fluctuating *Wolbachia* frequencies (Hoffmann et al. 1994; Kriesner et al. 2016; Cooper et al.

2017). Estimates of CI for wMel using crosses between field-collected male *D. melanogaster*

from Australia and *Wolbachia*-negative laboratory females are consistent with $s_h \approx 0$ (unless

males are very young) (Hoffmann et al. 1998; Reynolds and Hoffmann 2002; Kriesner et al.

2016; Shropshire et al. 2021). Estimates of wMel maternal transmission rates from field-

collected females have been recorded as high $\mu = 0.11$ (0.07, 0.17) (Olsen et al. 2001). Similar

levels of imperfect maternal transmission have also been recorded for wSan from field-collected

D. santomea ($\mu = 0.120$, [0.032, 0.346]) and for wYak from *D. yakuba* ($\mu = 0.20$ [0.087, 0.364])

on São Tomé (Hague et al. 2020b). Importantly, these patterns are not strictly characteristic of all

"wMel-like" *Wolbachia*, as some wMel-like strains cause strong CI in their native hosts (e.g.,

wSeg in *D. seguyi*; Shropshire et al. 2024) and other *Wolbachia* such as the wRi-like strain wSuz

in *D. suzukii* do not cause CI, show imperfect transmission in the field ($\mu = 0.14$ [0.04, 0.27]),

and fluctuate across geographic space (Hamm et al. 2014; Turelli et al. 2018). The patterns for

non-CI causing *Wolbachia* contrast empirical observations of the wRi strain in *D. simulans*,

which causes strong CI ($s_h \approx 0.45$) and tends to persist at high, stable infection frequencies. wRi

also seems to have relatively high rates of maternal transmission in the field (e.g., $\mu = 0.026$

[0.004, 0.057]) and in the lab at different temperatures (Hoffmann et al. 1990; Turelli and

Hoffmann 1995; Carrington et al. 2011; Kriesner et al. 2013; Hague et al. 2024). Similarly, the

CI-causing strains wHa (in *D. simulans*) and wAna (in *D. ananassae*) persist at high, stable

infection frequencies (Rousset and Solignac 1995; Bourtzis et al. 1996; James and Ballard 2000;

Ballard 2004; Choi et al. 2015) and have thermally stable maternal transmission rates in the lab

(Hague et al. 2024).

It is important to note that strongly fluctuating host fitness effects are required as a

condition to generate stochastic fluctuations on the scale of wMel at Gold Coast or wSan at São

Tomé (Figure 6). Unfortunately, we have a poor understanding of *Wolbachia* effects on host

fitness in nature, although the endosymbionts are generally considered to be beneficial based on

the spread and persistence of *Wolbachia* in nature (Weeks et al. 2007; Hamm et al. 2014; Cooper

et al. 2017; Meany et al. 2019; Ross et al. 2019*b*; Hague et al. 2020*b*, 2022; Hoffmann and

Cooper 2024). We do not know if host fitness effects fluctuate in nature (Turelli et al. 2022), but

there is evidence that the effects of *w*Mel on *D. melanogaster* fecundity (Kriesner et al. 2016)

and virus-blocking (Chrostek et al. 2021) can depend on the environmental context. The non-CI

causing strain *w*Au in *D. simulans* has also been shown to have context-dependent host effects

depending on where host breeding occurs (Cao et al. 2019). Given that temperature has pervasive

effects on *Wolbachia* densities in host tissues (Hague et al. 2020*a*, 2024), temperature-induced

changes to *Wolbachia* densities could in turn alter *Wolbachia* effects on various host fitness

components. The strongly fluctuating host effects we modeled here may also help explain

altitudinal variation in *w*Yak frequencies in *D. yakuba* observed on São Tomé in 2018, which

otherwise requires invoking unrealistic values of F and s_h (Figure S10; see Hague et al. 2020 for

further discussion).

Our simulations build on previous work focused on the role of finite-population

stochasticity in the initial establishment of *Wolbachia* from low frequencies in host populations

(Jansen et al. 2008). At low initial frequencies, *Wolbachia* that generate $F(1 - \mu) > 1$ may be

stochastically lost and, conversely, CI-causing strains that generate $F(1 - \mu) < 1$ may become

established; however, these probabilities decreases as host population size increases (Rigaud and

Rousset 1996; Egas et al. 2002; Jansen et al. 2008). *Wolbachia* that generate $F(1 - \mu) < 1$ are

disfavored by selection; however, strong-CI causing *Wolbachia* may still become established if

stochastic fluctuations carry *Wolbachia* above the critical threshold frequency at which CI favors

infection spread. This is especially relevant for biocontrol programs using transinfected *Aedes*

aegypti mosquitoes, because *w*Mel is deleterious but also generates strong CI in the mosquito

background (Ross et al. 2019*b*). Stochastic dynamics can also cause deleterious infections to be

lost if *Wolbachia* fluctuate back below the threshold frequency and they are no longer favored by

CI (Jansen et al. 2008). Our simulations here reflect the fact that *Wolbachia* infections in natural

hosts are now generally thought to generate $F(1 - \mu) > 1$ (Hoffmann and Turelli 1997; Kriesner

et al. 2013, 2016; Hamm et al. 2014; Kriesner and Hoffmann 2018; Meany et al. 2019).

We also considered the hypothesis that changes to maternal transmission rates, host

fitness effects, and/or CI between host generations could cause deterministic changes to

Wolbachia frequencies over time that resemble *w*Mel-like *Wolbachia* in nature. Predictably, the

$\bar{p}$ values from our simulations (Figures 5C) closely align with equilibria from the deterministic

model (Figure 2) (Hoffmann et al. 1990); although, we also considered the addition of low

transmitters to the population. Generally, moderate changes to maternal transmission rates (e.g.,

$\mu = 0.005$ to 0.05) and/or the addition of low transmitters between host generations can plausibly

explain drops in the frequencies of *w*Mel-like *Wolbachia*. Moderate changes to host fitness

effects (e.g., $F = 1.025$ to 1.075) and CI strength (e.g., $s_h = 0.1$ to 0) can also cause similar

changes to infection frequencies over time. These parameter values generally fall within a

biologically plausible range based on aforementioned estimates from field-collected *Drosophila*

hosts (Hoffmann et al. 1998, 1998; Olsen et al. 2001; Reynolds and Hoffmann 2002; Kriesner et

al. 2016; Shropshire et al. 2021), especially considering that maternal transmission rates of *w*Mel

and *D. yakuba*-clade *Wolbachia* can change depending on thermal conditions (Hague et al.

2020*b*, 2022, 2024). Of course, it is important to note that our simplified model treats μ, F, and s_h

as discrete values from one host generation to the next, whereas these parameters are much more

likely to exhibit quantitative variation on a continuous scale in natural populations with

overlapping host generations.

By incorporating a subpopulation of low transmitters into our simulations, we considered

the possibility that maternal transmission rates vary among hosts individuals within a given host

generation. This phenomenon has been observed for a number of *Wolbachia* strains in natural

host populations, including *w*Mel, *w*San, and *w*Yak (Hoffmann et al. 1998; Carrington et al.

2011; Hamm et al. 2014; Hague et al. 2020*b*). The presence of low transmitters caused the

largest drops in *Wolbachia* frequencies for strains that do not cause CI (Figure 4). These results

imply that heterogeneity of transmission rates among females and changes to the

presence/absence of low transmitters between host generations may provide another plausible

explanation for *w*Mel-like *Wolbachia* fluctuations over time.

Why might some hosts be low transmitters? Heterogeneity in transmission rates could

arise within a generation if host individuals experience different thermal conditions that impact

maternal transmission rates and *Wolbachia* densities (Ulrich et al. 2016; Ross et al. 2017, 2019*a*;

Foo et al. 2019; Hague et al. 2020*b*, 2020*a*, 2022, 2024; Chrostek et al. 2021). Insects are

ectotherms that inhabit complex, heterogeneous environments where thermal conditions can vary

on the scale of individual fruits or leaves (Roberts and Feder 2000; Potter et al. 2009; Saudreau

et al. 2009; Pincebourde and Woods 2012; Woods et al. 2015). Insects thermoregulate

behaviorally, but the immobile egg and pupal stages of development leave hosts more

susceptible to local microclimatic conditions that could influence maternal transmission or

Wolbachia densities (May 1979; Heinrich 1993; Dillon et al. 2009; Rajpurohit and Schmidt

2016). Heterogeneity of *Wolbachia* densities in hosts could also arise for other reasons, for

example, due to individual variation in diet (Serbus et al. 2015; Christensen et al. 2019), age

(Reynolds and Hoffmann 2002; Shropshire et al. 2021), and the host and *Wolbachia* genomes

(Gu et al. 2022; Hague et al. 2022). For instance, the density of *w*Inn, a strain related to *w*Mel

and *w*San, varied by 20,000 fold among host individuals in natural populations of *D. innubila* in

Arizona (Dyer and Jaenike 2004; Unckless et al. 2009; Hill et al. 2019).

We must point out that our discrete generation model generally requires a number of host

generations to achieve stochastic or deterministic changes to *Wolbachia* frequencies over time on

the order ≥ 0.669 for *w*Mel and ≥ 0.427 for *w*San (Figures S5, S7, S9). These timescales may be

plausible for *w*San on São Tomé, where frequencies fluctuated on the scale of years; however, it

is implausible for *w*Mel at Gold Coast, where frequencies fluctuated over just a month-long

period. This period from March to April corresponds to roughly a single host generation in the

field (Turelli and Hoffmann 1995; Hoffmann et al. 1998). Conservatively, the maximum

stochastic change in *Wolbachia* frequences (p_{max} - p_{min}) we observed over five consecutive host

generations in our simulations was 0.209 using the parameters that generated the largest p_{SD}

values ($\mu = 0.01$, $F = 1.025$, $CV = 0.1$, $s_h = 0$) (see Appendix 2 for a detailed analysis of

stochasticity over many host generations). Even if we assume an implausibly small host

population of $N = 100$, the maximum absolute change in *Wolbachia* frequencies that we observe

over five consecutive generations is 0.224 in a set of 25 simulations using the aforementioned

parameters. If we instead assume that μ, F, and/or s_h values change between host generations, it

takes roughly 50 host generations for deterministic changes in equilibria to occur (Figures S7,

S9). These results suggest that between-generation changes to maternal transmission, host fitness

effects, or CI cannot yield large deterministic changes to *Wolbachia* frequencies over a single

host generation according to the assumptions of our model.

What else could potentially explain large temporal fluctuations in *Wolbachia* frequencies

over just one or two host generations? Importantly, our model assumes that hosts have discrete

generations and a fixed population size. Accounting for aspects of insect population dynamics

and ecology, such as overlapping generations, density-dependent population regulation, or

seasonal variation in population size could also contribute to fluctuating *Wolbachia* frequencies

over time (Rasgon and Scott 2004; Farkas and Hinow 2010; Turelli 2010; Hancock et al. 2011;

Turelli and Barton 2022). Host migration among structured populations with different *Wolbachia*

frequencies could also plausibly alter frequencies within a given focal population (Hoffmann et

al. 1998; Flor et al. 2007; Jansen et al. 2008; Engelstädter and Telschow 2009; Haygood and

Turelli 2009; Hancock et al. 2011). *Wolbachia* frequencies could potentially vary on small

geographic scales due to the effects of microclimatic variation on μ, F, and s_h. At Gold Coast, *D.*

melanogaster hosts were sampled from 20-25 banana mash buckets along transects of at least 3

km. Perhaps microclimatic differences along the transect generated complex variation in

Wolbachia frequencies (Roberts and Feder 2000; Potter et al. 2009; Saudreau et al. 2009;

Pincebourde and Woods 2012; Woods et al. 2015). It is also conceivable that environmental

conditions could cause *Wolbachia* frequencies to change within a single host generation if there

are extreme effects on *Wolbachia* densities (e.g., a heat wave) that effectively clear *Wolbachia*

from some hosts (Ross et al. 2019*a*, 2020).

In summary, we find that stochasticity or between-generation changes to the maternal

transmission rate (as well as host effects and CI) can provide plausible explanations for the

temporal fluctuations of *w*Mel-like *Wolbachia* over years-long periods, but not shorter month-

long timescales. Our results highlight the need for future work to explore the scale of temporal

and geographic variation of *w*Mel-like *Wolbachia* frequencies in nature, in conjunction with

insect population dynamics (e.g., census population size, migration). It is unclear, for example,

whether *Wolbachia* frequencies vary temporally on a scale of less than a month or

geographically on the scale of kilometers. Interestingly, the temporal dynamics of *w*Mel at Gold

Coast seem to be quite variable, but on a broader geographic scale, *w*Mel frequencies have

persisted as a stable latitudinal cline for decades along the coast of eastern Australia (Hoffmann

et al. 1994, 1998; Kriesner et al. 2016). In the future, field estimates of maternal transmission

rates, fecundity, and CI using wild-caught flies sampled at different temporal and geographic

scales can be used to model *Wolbachia* dynamics using the publicly available R package

wlbcmodeler. Uncovering sources of temporal and geographic variation in *Wolbachia* prevalence

is ultimately critical for the understanding the long-term evolutionary dynamics of these host-

endosymbiont relationships.

DATA ACCESSIBILITY STATEMENT

The *wlbcmodeler* R package and scripts to recreate figures will be made publicly

available upon publication.

COMPETING INTERESTS STATEMENT

The authors declare no competing interests.

AUTHOR CONTRIBUTIONS

JMG: funding acquisition (equal), methodology (equal), formal analysis (lead), software

(lead), visualization (lead), writing – original draft (supporting), writing – review and editing

(supporting). JK: methodology (equal), formal analysis (supporting), writing – original draft

(supporting), writing – review and editing (supporting). MTJH: funding acquisition (equal),

conceptualization (lead), supervision (lead), writing – original draft (lead), writing – review and

editing (lead).

ACKNOWLEDGEMENTS

This work was supported by the University of Scranton through a Faculty Internal

Research Award to JMG and MTJH and a startup package awarded to MTJH from the College of

Arts and Sciences Dean's Office.

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	TABLES

999	**Table 1.** Regression analysis of the effects of model parameters on the standard deviation of

1000	*Wolbachia* frequencies (p_{SD}).

		Coefficient	t value	*P* value
	Host population size	-4.14×10^{-7}	-1.13×10^{2}	< 0.001
Imperfect maternal transmission (μ)	μ **value ($0.001 \leq \mu \leq 0.5$)**	8.51	4.03×10^{2}	< 0.001
	Low transmitters	1.12	2.88×10^{2}	< 0.001
	μ **value * low transmitters**	-4.42	-1.39×10^{2}	< 0.001
Host fitness effects (F)	F **value ($1 \leq F \leq 1.5$)**	-2.04	-1.68×10^{2}	< 0.001
	Fluctuating F ($CV = 0.01$)	4.23×10^{-1}	1.01×10^{2}	< 0.001
	Fluctuating F ($CV = 0.1$)	1.78	4.49×10^{2}	< 0.001
Cytoplasmic incompatibility (s_h)	**Weak CI ($s_h = 0.1$)**	-9.85×10^{-2}	-3.20×10^{1}	< 0.001
	Strong CI ($s_h = 0.45$)	-7.83×10^{-1}	-2.17×10^{2}	< 0.001

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1002	**Table 2.** Regression analysis of the effects of model parameters on mean *Wolbachia* frequencies

1003	($\bar{p}$).

		Coefficient	t value	*P* value
	Host population size	-1.58×10^{-8}	-6.09	< 0.001
Imperfect maternal transmission (μ)	μ **value ($0.001 \leq \mu \leq 0.5$)**	-2.24×10^{1}	-1.02×10^{3}	< 0.001
	Low transmitters	-2.43	-7.33×10^{2}	< 0.001
	μ **value * low transmitters**	8.26	2.76×10^{2}	< 0.001
Host fitness effects (F)	F **value ($1 \leq F \leq 1.5$)**	3.77	4.01×10^{2}	< 0.001
	Fluctuating F ($CV = 0.01$)	1.30×10^{-4}	4.93×10^{-2}	9.61×10^{-1}
	Fluctuating F ($CV = 0.1$)	9.68×10^{-3}	3.62	< 0.001
Cytoplasmic incompatibility (s_h)	**Weak CI ($s_h = 0.1$)**	4.27×10^{-1}	1.63×10^{2}	< 0.001
	Strong CI ($s_h = 0.45$)	1.92	6.82×10^{2}	< 0.001

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 FIGURES

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Figure 1. Temporal fluctuations of *Wolbachia* frequencies for **(A)** *w*Mel in *D. melanogaster* and

(B) *w*San in *D. santomea*. Error bars represent 95% binomial confidence intervals. **(A)** *w*Mel

frequencies in *D. melanogaster* at Gold Coast in subtropical eastern Australia from the winter of

1995 to the summer of 1996. Data recreated from Hoffmann et al. (1998). The authors reported

no obvious seasonal effect on *w*Mel frequencies. **(B)** *w*San frequencies in *D. santomea* in the

high-altitude mist forests on the volcanic island of São Tomé estimated from 2005 to 2018. Data

recreated from Hague et al. (2020*b*). *w*Yak frequencies in *D. yakuba* also fluctuate temporally on

São Tomé; however, this variation is complicated by altitudinal variation in *Wolbachia*

frequencies and transmission rates (Hague et al. 2020*b*).

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1018 **Figure 2**. Equilibrium *Wolbachia* frequencies ($\hat{p}$) approximated by a discrete generation

1019 deterministic model incorporating imperfect maternal transmission (μ), *Wolbachia* effects on

1020 host fitness (F), and cytoplasmic incompatibility (s_h). Equilibrium frequencies are shown across

1021 a range of biologically plausible parameter values for no, weak, and strong CI.

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Figure 3. Predicted standard deviation (p_{SD}) values of *Wolbachia* frequencies over time from the

regression analysis as a function of imperfect maternal transmission (μ) and conditioned by the

presence/absence of low transmitters, strength of host fitness effects (F), and cytoplasmic

incompatibility (s_h). Note that prediction lines average the effects of host population size (N) and

fixed/fluctuating host effects ($CV = 0.01$ and 0.1).

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1031   **Figure 4.** Predicted mean *Wolbachia* frequency ($\bar{p}$) values from the regression analysis as a

1032   function of imperfect maternal transmission (μ) and conditioned by the presence/absence of low

1033   transmitters, strength of host fitness effects (F), and cytoplasmic incompatibility (s_h). Note that

1034   prediction lines average the effects of host population size (N) and fixed/fluctuating host effects

1035   ($CV = 0.01$ and 0.1).

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Figure 5. Summary of results from simulations exploring the contributions of μ, F, and s_h to

standard deviations (p_{SD}) and mean *Wolbachia* frequencies ($\bar{p}$) with a finite host population size

of $N = 10^4$. **(A)** Each cell represents the mean standard deviation ($\overline{p_{\text{SD}}}$) of *Wolbachia* frequencies

from 25 simulations with a given μ value and a fixed F value. Note that $\overline{p_{\text{SD}}}$ values are plotted on

a log scale. **(B)** Each cell represents $\overline{p_{\text{SD}}}$ from 25 simulations with a given μ value and a strongly

fluctuating F value ($CV = 0.1$). **(C)** Each cell represents the pooled mean of mean *Wolbachia*

frequencies ($\bar{\bar{p}}$) from a set of 25 simulations with a given μ value and a fixed F value. Similar

data are not shown for the simulations with fluctuating F values, because fluctuating host effects

had only a small effect on $\bar{p}$. Results for strong CI (considered implausible for *w*Mel and *w*San)

are shown in Figure S3.

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Figure 6. Results from individual simulations exploring the contributions of μ, F, and s_h to the

absolute difference between maximum and minimum *Wolbachia* frequencies (p_{max} - p_{min})

observed over 10,000 host generations with a finite population size of $N = 10^4$. Individual points

represent the maximum absolute change in *Wolbachia* frequencies observed in a given

simulation with various combinations of μ ($0.001 \leq \mu \leq 0.5$) and F ($1 \leq F \leq 1.5$). Datasets are

further separated based on whether low transmitters and fluctuating host effects ($CV = 0.01$ or

0.1) are present. Individual points are color-coded to indicate each simulation's mean *Wolbachia*

frequency ($\bar{p}$). The full set of results are shown in Figure S4. The first 500 generations were

removed from each simulation before calculating p_{max} - p_{min}.

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