Contact networks of small mammals highlight potential transmission foci of *Mammarenavirus lassaense*.

Authors

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Abstract

Lassa fever (*Mammarenavirus lassaense*; LASV), is an endemic zoonosis in West Africa. Human infections arise from rodent-to-human transmission, mainly from the synanthropic reservoir *Mastomys natalensis*. In Sierra Leone, small-mammal communities vary across land use gradients, shaping LASV transmission risk. How anthropogenic environments facilitate the rodent-rodent interactions remains poorly understood.

We sampled small mammals over 43,266 trap nights in Sierra Leone's LASV-endemic Eastern Province, detecting 684 rodents and shrews. To assess potential transmission, we constructed space-sharing networks from co-trapping events within species-specific home range radii. These networks approximated shared space use, allowing comparison of encounter patterns across habitats.

Network topology varied significantly by land use. Village networks were the most connected (highest average degree), whereas agricultural communities supported the most species (higher rarefied richness) and were the most fragmented (higher modularity). Notably, the probability of intraspecific space sharing among *M. natalensis* was highest in agriculture, suggesting land use modulates key intra-specific transmission pathways.

LASV seroprevalence was 5.7% community-wide, with antibodies in nine species. We found no statistically significant association between overall seroprevalence and land use or aggregate network structure (mean degree). However, predictive modeling for *M. natalensis* indicated that higher individual degree is associated with seropositivity, suggesting complex, scale-dependent relationships. These findings show that simple ecological drivers do not fully explain LASV exposure, highlighting the importance of species-specific behaviors (i.e., *M. natalensis* clustering in agriculture) and the multi-host serological landscape in assessing transmission risk.

Introduction

Lassa fever caused by *Mammarenavirus lassaense* (LASV) is a rodent-associated zoonosis endemic to West Africa, with an estimated 100,000-900,000 infections annually[1,2]. While outbreaks in Nigeria are routinely reported, cases in other endemic countries - Guinea, Liberia and Sierra Leone - are sporadically documented[3–5]. In Sierra Leone disease outbreaks frequently go undetected, consistent with findings of up to 80% LASV seropositivity among human communities in regions previously not considered endemic[6]. Human infections typically result from transmission from rodent hosts, with limited subsequent human-to-human transmission[7]. Understanding LASV transmission in endemic settings requires a detailed characterization of small-mammal community interactions, through which pathogen transmission occurs and is maintained.

The primary reservoir of LASV, *Mastomys natalensis* is a native synanthropic rodent species, widespread across sub-Saharan Africa. Pathogen challenge studies on *M. natalensis* colonies suggest that acute infection does not significantly alter rodent

behavior or cause clinical pathology[8–10]. LASV is transmitted through both direct contact
(e.g., superficial wounds caused by infected conspecifics) and indirect contact (e.g.,
exposure to contaminated environments), at low infectious doses[9]. Vertical transmission
(mother-to-pup) is also thought to be an important mechanism for viral persistence and
spread[10,11]. Infected adult rodents exhibit detectable viral RNA as early as 3 days post-
infection, with viral loads peaking within 1-2 weeks and resolving by 40 days[9]. Among
individuals infected within the first 2 weeks of life viral RNA is detectable up to 16
months post infection[10]. The transient nature of acute infection - outside of neonatal
infections - has led many studies to focus on LASV-specific antibody detection, rather
than markers of acute infection, such as viral RNA[12–14].

The antibody response dynamics to LASV in rodents are not yet well understood. Based
on data from a similar arenavirus, Morogoro virus, seroconversion is expected to occur
by 7 days post-infection, with antibodies (e.g., IgG) persisting beyond the decline of
detectable RNA[11]. LASV-infected rodents are presumed to develop lifelong immunity to
reinfection upon seroconversion, however, the efficacy of neutralizing antibodies is
unclear and the role of immune or partially immune individuals in transmission networks
is not known[9,10,15]. Although antibody-based studies have limitations, the higher
prevalence of LASV seropositivity compared to acute infections provides valuable
insights into viral dynamics within endemic regions. For example, a recent study in Bo
district, Sierra Leone, reported a 2.8% LASV IgG seroprevalence among rodents and
shrews, compared to a 0.3% prevalence of acute infection detected via PCR,
underscoring the challenges of identifying acute infections in small-mammal
populations[16].

While *M. natalensis* is considered the primary reservoir species of LASV, evidence of
LASV infection has been found in 15 other small-mammal species in endemic regions:
five identified with acute infection and ten showing previous infection based on
serological evidence.[14,17–19]. The contribution of these additional species to pathogen
transmission into human populations - and their role in viral transmission or
maintenance within their species communities - remains unclear. In species-rich
environments, both direct and indirect contact among small mammals may result in
incidental infections of non-reservoir species, which are subsequently detected through
surveillance. Incidental infections of non-primary reservoir species may have little
impact on viral transmission or maintenance[20]. Alternatively, these species could
facilitate the transfer of LASV across landscapes, linking geographically isolated *M.*
natalensis populations and causing reintroduction of the virus into reservoir species
populations[21,22]. To account for this uncertainty we refer to *M. natalensis* as the primary
reservoir species and other small-mammal species previously found to be infected as
hosts. Increasing recognition of multi-species host systems in zoonoses underscores
the importance of expanding surveillance efforts to the wider community in which the
host resides to better understand pathogen prevalence and dynamics[23,24].

Understanding the structure of small-mammal contact networks in LASV-endemic
regions may offer valuable insight into the structural drivers of transmission, even in the
absence of direct observations of infection. While prior studies have described rodent
diversity and abundance, these do not capture interaction patterns relevant for
pathogen spread[16,25,26]. Spatial proximity can be used as a proxy for potential contact

and, when analyzed as networks, can help identify structural features, such as
connectivity, inter- and intra-specific mixing, and network fragmentation that influence
pathogen transmission. In contexts where direct behavioral observation or tagging of
individuals is infeasible, such as post-Ebola Sierra Leone, proximity-based approaches
provide a pragmatic alternative to infer likely transmission pathways[27].

Small-mammal communities in LASV-endemic regions are structured along
anthropogenic land use gradients[14,28]. As such, the risk of Lassa fever outbreaks in
human populations is expected to correlate with these gradients[6,29,30]. Within human-
dominated land use types, the prevalence of typically synanthropic rodent hosts of
LASV is anticipated to be higher due to increased food availability, shelter, and reduced
predation pressure[31–33]. These factors influence rodent abundance and population
dynamics which in turn may promote greater pathogen persistence, as observed in
several other rodent-associated zoonosis systems[34–37]. Understanding how small-
mammal contact networks, alongside small-mammal occurrence and abundance, vary
along these anthropogenic gradients could reveal potentially distinct pathogen
transmission networks in different habitats. We hypothesize that both contact frequency
and overall network connectivity increase in human-dominated habitats where
resources are concentrated, leading to enhanced potential for pathogen transmission.

In this study, we define "contact" as the inferred co-occurrence within a spatial and
temporal window (i.e., proximate capture locations within four trap-nights), consistent
with prior spatial proximity models[38,39]. While this does not imply physical contact or
indirect contact suitable for transmission, it serves as a proxy for shared habitat use and
potential indirect or environmental contact relevant to LASV transmission. Interaction
networks derived from wildlife data have previously been employed to study pathogen
transmission. These networks are particularly valuable for investigating the role of
community structure and the impact of contact rate heterogeneity between species in
multi-host pathogen systems[40–42].

We leverage rodent and shrew trapping data collected over three years in the Lassa
fever-endemic region of the Eastern Province, Sierra Leone. We characterize potential
interactions - both direct and indirect - within these communities as a network, where
the nodes represent rodents or shrews, and the connections (or edges) between them
represent potential interactions. We further hypothesize that the spatial clustering of
conspecifics and the increased abundance of commensal species in anthropogenically
dominated environments will lead to higher intra-specific contact rates compared to
inter-specific contact rates within these communities. We use network analysis to
explore how contact patterns vary along an anthropogenic land use gradient, focusing
on *M. natalensis* while also assessing the roles of other potentially important, highly
connected species ("hubs"). Finally, we report the prevalence of antibodies against
LASV among individual small mammals in the study region and investigate the
association between contact rates with seropositivity.

Methods

Study area

Rodent trapping surveys were conducted between October 2020-April 2023 within and around four village study sites (Baiama, Lalehun, Lambayama, and Seilama) in the Lassa fever endemic zone of the Eastern Province of Sierra Leone (Figure 1A). Surveys were conducted within trapping grids along a land use gradient of anthropogenic disturbance comprising forest, agriculture (including fallow and currently in-use areas), and villages (within and outside of permanent structures) (see Supplementary Information for study site and trap grid locations). One study site (Lambayama) was more developed than the other villages due to its proximity to the province capital (Kenema), we refer to this site as peri-urban with the others termed rural. Trapping survey sessions occurred four times annually with two trapping surveys in each of the rainy and dry seasons (May to November and December to April, respectively), producing up to a total of 10 trapping sessions in each village over the study period.

*Figure 1: A) Location of village study sites (circles) in the Eastern Province of Sierra
Leone. Kenema, the largest city in the province, and the national capital, Freetown, are
also shown (+). The inset map highlights the location of Sierra Leone within West Africa.
B) An example of a rodent contact network derived from trapping data during visit 5 in
village land use. Each colored node represents an individual small mammal, with lines
(edges) indicating inferred contacts between individuals. The number of edges
connected to a node represents its degree. Betweenness reflects the importance of a
node in connecting different parts of the network. Pale nodes indicate unobserved*

individuals, for whom contacts (edges) were not recorded. Species listed in the legend
without colors were never detected in the shown land use type (village).

Alt text: [A two-panel figure. Panel A is a map of Sierra Leone, marking the four village
study sites in the Eastern Province, with an inset showing Sierra Leone's location in
West Africa. Panel B is a network diagram showing potential contacts between
individual small mammals, where nodes colored by species represent animals and lines
represent shared space use.]

Study sites were selected to represent the range of land use in the Eastern Province of
Sierra Leone, considering both accessibility throughout the year and acceptability of the
study protocol to the village communities (Supplementary Information). At each trapping
grid 49 Sherman traps (7.62cm x 8.89cm x 22.86cm) (H.B. Sherman Traps,
Tallahassee, USA), were arranged in a 7 x 7 grid, with traps placed 7 meters apart in a
grid conforming to the local landscape (median trapping grid area = 1,672 m^2). In
permanent structures, trap placement deviated from the grid structure. At each visit,
permanent structures were randomly selected from a grid projected over the village
area, with four traps placed within each structure. The location of each trap within the
grid was geolocated. Traps were baited with a locally produced mixture of oats, palm oil
and dried fish. Each morning, traps were checked, closed for the day, and re-baited in
the evening. Each trapping survey session consisted of four consecutive trap-nights
(TN) at each trapping grid within the village study site. Rodents and shrews were
associated with the coordinates of the trap they were detected. Geospatial processing
was performed using the sf package in R (version 4.1.2)[43,44].

All small mammals were handled by trained researchers using appropriate personal
protective equipment. Animals were sedated using halothane and euthanized according
to established protocols[45]. Morphological measurements and samples of blood and
tissue (skin, liver and spleen) were collected. The study was approved by the Clinical
Research Ethical Review Board and Animal Welfare Ethical Review Board of the Royal
Veterinary College, UK (URN: 2019 1949-3), and the ethics committee of Njala
University, Sierra Leone. The study adhered to national and institutional ethical
guidelines for the humane treatment of animals and safe handling of zoonotic
pathogens. Trapping and sampling protocols were designed to minimize stress to the
animals and reduce the risk of pathogen exposure to researchers and communities.
Community engagement sessions were conducted to ensure understanding and
acceptance of the study objectives, and all work complied with the principles outlined in
the ARRIVE guidelines (v2)[46]. All carcasses were incinerated to mitigate pathogen
transmission risks.

Species identification

Species identification was performed in the field based on external morphological
characteristics, including body length, tail length, ear length, and pelage coloration,
following the taxonomic keys of Happold and Kingdon and Monadjem *et al.*[47,48]. Field
identification was supplemented by molecular methods to confirm species identity for
individuals identified as *Mastomys sp.*, *Mus sp.*, *Rattus sp.* and *Crocidura sp.* alongside
a random subset of remaining individuals (50% of remaining samples).

All samples remained in Sierra Leone and were stored at -20°C until processing.
Genomic DNA was extracted using QIAGEN DNeasy kits as per the manufacturers
instructions[49] (Supplementary Information). DNA extracts were amplified using platinum
Taq polymerase (Invitrogen) and *cytochrome B* primers[16]. DNA amplification was
assessed through gel electrophoresis, and successful amplification products were
Sanger sequenced (performed by Eurofins Genomics). The sequences were attributed
to rodent species using the BLAST program, comparing the obtained sequences to
cytochrome B records in the NCBI database (accessed 2023-06-30)[50].

Serological Analysis

Serological status of trapped rodents and shrews was determined using the
BLACKBOX® LASV IgG ELISA Kit (Diagnostics Development Laboratory, Bernhard
Nocht Institute for Tropical Medicine), which has been validated for rodent samples[51,52].
The protocol is reproduced in Supplementary Information. Briefly, 1µL of whole blood
was inactivated and mixed with the provided sample dilution buffer (1:50). Where whole
blood was unavailable (21 samples, 3%), blood was eluted from dried blood spots
stored on filter paper by incubating with phosphate-buffered saline containing 0.08%
Sodium Azide and 0.05% Tween-20[53]. Samples, alongside negative and positive
controls, were incubated on the ELISA kit plates for 24 hours at 4-8 °C in a wet
chamber. Following incubation, the plates were washed and incubated for a further hour
with 1:10,000 diluted HRP-labelled streptavidin. A final wash was performed prior to the
addition of 100µL of 3,3',5,5'-Tetramethylbenzidine (TMB) substrate to wells, with
incubation for 10 minutes. The colorimetric reaction was stopped by adding 100µL of a
stop solution.

A deviation from the kit protocol occurred due to local ELISA plate reader limitations.
We measured the optical density (OD) at 450nm and 630nm, as opposed to 450nm and
620nm but this was not expected to have an important effect on absorbance patterns,
as advised by the manufacturer. The index value was calculated by subtracting OD_{630}
from OD_{450} and dividing by the cut-off value (the mean values of the negative controls +
0.150). Samples were classified as positive if the index value was greater than or equal
to 1.1, negative if the index value was less than or equal to 0.9, and inconclusive if the
index value was between 0.9 and 1.1. Inconclusive results were retested.

The prevalence of seropositive individuals is reported aggregated by species. A
Bayesian logistic regression model was constructed, using the `brms` package, to
estimate the Odds Ratio (OR) of seropositivity for each species compared to *M.*
natalensis, which served as the reference species[54]. Specifically, a Bernoulli regression
with normal, uninformative priors for population-level effects was used, with a binary
dependent variable for seropositivity and an independent variable of small mammal
species. Only species with more than 10 individuals assayed for antibodies to LASV
were included in this model. Posterior distributions are presented in graphical format,
alongside the posterior mean and 95% Credible Interval (CrI). Unlike frequentist
approaches, Bayesian inference does not rely on *p*-values; rather, statistical support for
an effect is assessed based on the posterior distribution, with associations interpreted in
terms of the central tendency (e.g., OR) and whether the CrI excludes the null value
(OR = 1).

In addition to the species-level seroprevalence analysis, we conducted *post-hoc*
exploratory analyses to assess differences in LASV seroprevalence by village and land
use type. These analyses employed Bayesian logistic regression models analogous to
those used in species-level comparisons, with results interpreted cautiously given their
non-pre-specified nature.

Community Diversity Analysis

To compare small-mammal diversity across land use types, it was necessary to account
for the different number of animals captured in each habitat. We therefore used
individual-based rarefaction (IBR), a statistical technique that standardizes species
richness estimates as if an equal number of individuals had been sampled in each
habitat. This approach allows for a robust comparison of alpha diversity by controlling
for unequal sampling effort.

Using the `mobr` package in R, we calculated the expected species richness for each
habitat rarefied to a common sample size of 10 individuals, which was the minimum
number of captures in any single habitat[55,56]. To further investigate the ecological
processes shaping these communities, we compared these rarefied richness values to
the expectations of a null model. This comparison allowed us to assess whether the
observed species abundance patterns deviated from random community assembly, for
instance, to detect potential dominance effects where one or a few species reduce
overall community evenness and richness.

Contact Network Construction

Species space-sharing networks were reconstructed from the trapping data. Capture-
mark-recapture (CMR) methods have previously been used to identify space-sharing by
individuals[39,57,58]. In our study system, a CMR design was not feasible due to study
communities concerns around the risk of releasing an infected animal. Therefore, we
considered that spatial and temporal co-occurrence proxies direct or indirect contacts
with other small mammals[38]. We assumed these potential contacts were sufficient to
transmit LASV if they were trapped within a species specific buffer zone centered on the
location of the trap during the same 4 trap night session. A key assumption underlying
this approach is that an individual was trapped at the center of their home range[58].

We obtained species specific buffer radii for the primary analysis using a hierarchical
approach that prioritized the most robust data available for each species. Our primary
source was systematically compiled data within the `HomeRange` R package (version
1.0.2), from which we extracted median radii at the species level where possible,
including for *M. natalensis* (median home range radii = 10.6m), *Mus musculus* (9.6m),
Lemnisomys striatus (8m), *Lophuromys sikapusi* (8.4m), and *Rattus rattus* (29.3m)[59].
For detected species with no species level home range data, *Praomys rostratus* (7.2m),
Mus setulosus (9.6m), *Crocidura spp.* (7.2m) and *Hylomyscus simus* (7.7m) we used
the genus level median home range. Where no data were available we obtained home
range radii from single-study estimates in the literature. This provided values for
Malacyoms edwardsi (35m), *Gerbilliscus guineae* and *Hybomys planifrons* (both
22m)[60–62]. For *Dasymys rufulus*, which lacked any of the above data, we assigned the

median of all previously determined species-level radii (11 m). To assess the impact of
these assumptions, we performed pre-specified sensitivity analyses by applying uniform
buffer radii of 15, 30, and 50 meters to all species.

Networks were constructed from observed animals (nodes) and the presence or
absence of contact between them (edges). Data were aggregated by land use type and
sampling visit, producing a potential 32 distinct networks from 201 trapping grid, village
and visit combinations (see Figure 1B for an example network). However, as no rodents
were detected in three networks derived from forest sites, only 29 networks were used
in subsequent analysis.

Network Structure and Mixing Pattern Analysis

We analyzed the structure of the space-sharing networks at three distinct levels: 1) the
overall network within each land use type, 2) the individual animals (nodes) within those
networks, and 3) species-level roles, including connectivity patterns (degree distribution,
hub identification) and mixing dynamics (inter- vs. intra-specific contacts).

First, to characterize the overall networks, we calculated global metrics for each land
use type (village, agriculture, and forest). These included the total number of nodes
(animals) and edges (inferred contacts), as well as the network-wide mean degree and
betweenness centrality. We then used a Kruskal-Wallis test to formally compare the
distributions of these metrics across the three land use types. For metrics with a
significant overall difference, we performed a Dunn's *post-hoc* test with Bonferroni
correction to identify which specific pairs of habitats differed.

Second, to evaluate the roles of individual animals, we calculated two key node-level
centrality measures: degree centrality, the number of connections for each animal, as a
direct measure of its contact frequency, and betweenness centrality, to identify
individuals acting as critical bridges within the network. We then investigated species-
specific interaction patterns by analyzing the degree distribution for each species to
identify key hub species (i.e., those with notably high connectivity) in each habitat
(Figure 3). To formally test for differences in connectivity across land uses, we used a
Wilcoxon rank-sum test to compare the network-level mean degree for key abundant
species.

To quantify mixing patterns, we also calculated the proportion of each species' contacts
that were intraspecific (with its own species) versus interspecific (with other species),
providing a measure of network homophily across the land use gradient.

Modeling Contact Probability in *Mastomys natalensis*

To examine the association between land use and species with the probability of a
contact between two individuals, we modeled these contacts using Exponential-Family
Random Graph Models (ERGMs)[63]. The analysis was limited to *M. natalensis*, the
primary rodent host of LASV. Estimation of ERGM parameters provide an OR for the
probability of an edge in a network – conditional on the rest of the network - based on
network properties included in the model and nodal attributes. Within our trapping grids,
only a subset of all individuals are detected in traps. Including unobserved individuals -

and thus, unobserved contacts - enhances the interpretability and generalizability of the
network models. This approach allows for a more accurate estimation of the total
population size by accounting for missing data, thereby making the network models
more representative of the entire population from which the analytic sample was
derived.

Previous analysis of our study system suggests that the probability of detecting a rodent
at each trap is less than 10% for 4 trap nights, provided that the species is present in
the trapping grid[28]. To estimate the abundance of individuals of each species within a
trapping grid, we modeled abundance (i.e., total population size) from repeated count
data using an N-mixture model implemented in the `unmarked` R package (version
1.2.5)[64,65]. The latent abundance distribution was modeled using Poisson, negative
binomial or zero-inflated Poisson random variables. The abundance model included the
number of trap nights and season as replicate-dependent detection covariates, as well
as location (rural vs. peri-urban setting) and land use type (forest, agriculture or village)
as occurrence covariates.

To select the most appropriate model for each species, we compared the Akaike
Information Criterion (AIC) of the Poisson, negative binomial, and zero-inflated Poisson
models. The best-fitting model was then used to derive the estimated abundance. The
median estimated abundance from the distribution produced for each trapping grid was
used to estimate the number of unobserved individuals in each network, aggregated by
land use type (Supplementary Figures 2.1-2.12). The number of observed individuals
was subtracted from the predicted abundance to derive the number of unobserved
individuals for each species. These unobserved individuals were explicitly set to have
missing (i.e., unobserved) edge values.

Finally, the constructed adjacency matrices were converted to networks using the
`network` R package (version 1.13.0.1) for subsequent ERGM modelling[66] (Figure 1B)
and Supplementary Figures 3.1-3.3).

ERGMs were specified for each of our inferred contact networks to compare the
probabilities of edges forming based on rodent characteristics (i.e., species). The
general model is shown in Equation 1:

$$P(Y = y) \propto \exp\left(\theta_1 g_1(y) + \theta_2 g_2(y) + \ldots + \theta_p g_p(y)\right) \ (1)$$

Where p is the number of terms in the model, and the values of the coefficients θ
represent the size and direction of the effects of the covariates $g(y)$ on the overall
probability of an edge being present in the network. At the edge level the expression for
the probability of the entire graph can be re-expressed as the conditional log-odds of a
single edge between two nodes (a contact between two rodents) as shown in Equation
2.

$$\text{logit}\left(P\left(Y_{ij} = 1 | y_{ij}^c\right)\right) = \theta' \delta\left(y_{ij}\right) \ (2)$$

Where Y_{ij} is the random variable for the state of the node pair ij and $y_{i,j}^c$ signifies all
dyads in the network other than $y_{i,j}$. θ' is the coefficient for the change production of an
edge between the two nodes conditional on all other dyads remaining the same
$(\delta(y_{ij}))$.

ERGMs are implemented using the `ergm` package (version 4.3.2) in R[67]. Three terms
were included in the final ERGM to model the probability of the formation of ties
(Equation 3). The first term (edges), describes the density of the network, representing
the probability of a tie being observed in the network. The second term (species)
represents the conditional probability of a tie forming, conditional on the species of the
nodes. The third term (species homophily) accounts for intraspecific tie formation
among rodent individuals (i.e., the conditional probability of two individuals of the same
species forming a tie).

$$P(Y = y) \propto \exp\left(\theta_{edges}g_{edges}(y) + \theta_{species}g_{species}(y) + \theta_{homophily}g_{homophily}(y)\right) \text{ (3)}$$

ERGMs were implemented on the individual networks for each land use type at each
visit. The effect sizes from each model were pooled through random-effects meta-
analysis, stratified by land use, to produce a land use-specific summary effect size for
each coefficient[68]. Inclusion in meta-analysis was restricted to ERGMs that produced
stable estimates for each of the model terms (i.e., sufficient detections of *M. natalensis*
within the network). Random-effects models were conducted using the `metafor` package
(version 4.0.0) in R[69]. Heterogeneity across the models was assessed using the Q-test
and the restricted maximum-likelihood estimator (τ^2) with a prediction interval for the
true outcomes produced[68,70]. The Q-test assesses whether there is greater variability
among effect sizes than expected by chance, with a significant result indicating
substantial heterogeneity. The τ^2 statistic estimates the between-study variance,
quantifying the degree of heterogeneity rather than just testing for its presence. Weights
for each network included in meta-analysis were assigned using inverse-variance
weights[71].

The presence of influential networks was assessed using Cook's distance, for models
including influential networks leave-one-out sensitivity analysis were performed[72].
Forest plots were generated to visualize the summary OR of the probability of a tie for
each model term, stratified by land use type.

Models with unstable estimates for the species homophily term were not included in the
random-effects meta-analysis. No contact networks from forest land use contributed to
meta-analysis as no *M. natalensis* were detected in these settings. Five models from
agricultural settings and six from village settings were included in meta-analysis.

Analysis of Network Centrality and Serostatus

To investigate pathogen transmission within our networks, using seropositivity as a
proxy for prior exposure to LASV, we first report the small-mammal species found to
contain individuals that were seropositive for LASV. We then compared the nodal
degree of seropositive and seronegative individuals using a Wilcoxon rank-sum test with

continuity correction[73]. This analysis was repeated stratified by species to assess
whether contact rates were associated with an individual being seropositive. Finally, we
compared the node-level betweenness of seropositive and seronegative individuals to
determine whether an individual's position within a structured contact network was
associated with prior exposure to LASV.

Modelling Seropositivity Risk Factors in *M. natalensis*

We next tested whether an individual *M. natalensis*'s risk of being LASV-seropositive
was influenced by its position in the contact network. Specifically, we investigated the
effects of two key properties: its total number of contacts (degree) and its tendency to
connect with conspecifics (homophily). Because our trapping data provides only one
realization of a dynamic social system, we used a simulation-based approach to
account for this structural uncertainty. Based on the network formation rules derived
from the ERGMs described previously, we generated 50 plausible network simulations
for each trapping grid that contained at least one seropositive *M. natalensis*.

Using these simulated networks, we fitted a Bayesian generalized linear mixed-effects
model (GLMM) to assess the association between network position and serostatus. The
GLMM predicted LASV serostatus as a function of degree, homophily, and their
interaction, with random intercepts for each original trapping grid and simulation
replicate to account for the nested data structure. The model was specified using the
following formula:

$$\text{logit}(p_i) = \beta_0 + \beta_1(Homophily_i) + \beta_2(Degree_i) + \beta_3\big(Homophily_i * Degree_i\big) + u_j + v_k \ (4)$$

Where u_j (the random effect for each network) is drawn from a Normal distribution with
a mean of 0 and a variance $\sigma^2_{network}$, and v_k (the random effect for each simulation) is
drawn from a Normal distribution with a mean of 0 and a variance σ^2_{sim}. For each *M.*
natalensis individual in each simulation, we calculated its degree and a node-level
homophily score (defined as the proportion of its direct contacts that were with other *M.*
natalensis). Evidence for an association was assessed using the posterior distributions
of the model coefficients.

Results

A total of 684 small mammals were captured over 43,266 trap-nights, representing 17
species (13 rodent and 4 shrew species). *M. natalensis* was the most commonly
detected species (N = 113, 16.5%), followed by *Crocidura olivieri* (N = 105, 15.3%) and
Praomys rostratus (N = 102, 15%) (Table 1). Rarefied species richness (*S*),
standardized to 10 individuals, was 5.86 in agricultural habitats, 5.13 in forests, and
4.36 in village settings. Observed species richness was 9 in agricultural sites, 6 in
villages and 3 in forests. In village habitats, rarefied richness values were consistently
lower than the expectations of null models of species abundance distributions.

LASV Seroprevalence Across Species and Habitats

Antibodies to LASV were identified in 39 rodents and shrews (39/684, 5.7%) from 9 species, including *M. natalensis* (11/39, 28%), *C. olivieri* (8/39, 21%), *Lophuromys sikapusi* (8/39, 21%) and *Rattus rattus* (4/39, 10%) (Table 1).

In a Bayesian model with *M. natalensis* as the reference, the OR for LASV seropositivity was highest in *L. sikapusi* (OR = 2.11, 95% Credible Interval (CrI) = 0.89-4.88). However, like most other species, its credible interval included 1.0, indicating substantial uncertainty. In contrast, *M. musculus* had the lowest odds of being seropositive (OR = 0.22, 95% CrI = 0.06–0.74), and its CrI did not include 1.0, suggesting stronger evidence for reduced odds compared to *M. natalensis*.

Table 1: The number of individuals detected and antibodies to LASV among those individuals.

Species	Order	Individuals (N)	LASV Antibody detected (%)	Percentage of all positive individuals	OR (95% CrI)
Mastomys natalensis	Rodentia	113	11 (9.7%)	28.2%	ref
Crocidura olivieri	Eulipotyphla	105	8 (7.6%)	20.5%	1.15 (0.51-2.65)
Praomys rostratus	Rodentia	102	2 (2%)	5.1%	0.4 (0.13-1.14)
Mus musculus	Rodentia	90	0 (0%)	0%	0.22 (0.05-0.74)
Rattus rattus	Rodentia	88	4 (4.5%)	10.3%	0.71 (0.25-1.92)
Lophuromys sikapusi	Rodentia	57	8 (14%)	20.5%	2.11 (0.89-4.88)
Mus setulosus	Rodentia	43	3 (7%)	7.7%	1.01 (0.32-2.85)
Crocidura buettikoferi	Eulipotyphla	23	0 (0%)	0%	0.44 (0.09-1.83)
Crocidura grandiceps	Eulipotyphla	15	0 (0%)	0%	0.52 (0.11-2.23)
Malacomys edwardsi	Rodentia	11	1 (9.1%)	2.6%	1.12 (0.24-4.53)
Lemniscomys striatus	Rodentia	11	0 (0%)	0%	0.58 (0.11-2.8)
Hylomyscus simus	Rodentia	9	0 (0%)	0%	
Hybomys planifrons	Rodentia	7	1 (14.3%)	2.6%	
Mastomys erythroleucus	Rodentia	4	1 (25%)	2.6%	
Crocidura theresae	Eulipotyphla	3	0 (0%)	0%	
Gerbilliscus guineae	Rodentia	2	0 (0%)	0%	
Dasymys rufulus	Rodentia	1	0 (0%)	0%	

Figure 2: Odds Ratios of seropositivity to LASV among small-mammal species,
compared to Mastomys natalensis. Only species with more than 10 individuals assayed
for antibodies to LASV were included in this analysis.

Alt text:[A forest plot comparing the odds ratios of LASV seropositivity for several small
mammal species against the reference species, *Mastomys natalensis*. Each species
has a point estimate and a 95% credible interval. *Lophuromys sikapusi* shows a higher
odds of being seropositive (OR 2.09), while *Mus musculus* shows lower odds (OR 0.22).
The credible intervals for most species are wide and cross the null value of 1, indicating
uncertainty.]

Antibody positive small mammals were detected in three of the study villages, Lalehun
(N = 18, 46%), Seilama (N = 12, 31%) and Baiama (N = 9, 23%). No positive individuals
were detected in Lambayama, the most urbanized village study site. Antibody positive
individuals were detected during all study visits except visit 9 (2023-February). The
proportion of antibody positive among all captures was 6.3% (24/379) in agricultural
sites, 5% (13/261) in villages, and 4.5% (2/44) in forests.

In exploratory *post-hoc* models, the OR for seropositivity was 2.57 in Lalehun (95% CrI
= 1.28-5.15), 1.55 in Baiama (95% CrI = 0.68-3.44), and 0.21 in Lambayama (95% CrI =
0.05-0.67), relative to Seilama as the reference village. The OR for agricultural land use
was 1.27 (95% CrI = 0.68-2.42), and 0.85 (95% CrI = 0.23-2.59) in forests compared to
villages. As the 95% CrIs for several comparisons (specifically, Baiama village, and
agricultural and forest land uses) included 1.0, we did not find strong evidence for a
difference in seroprevalence in these settings compared to their respective references.

Structure of Small-Mammal Contact Networks

The number of individuals (nodes) was highest in agricultural settings (n = 379)
compared to villages (n = 261) and forests (n = 44). While mean network degree was
numerically highest in village settings (mean = 3.39), followed by forest (mean = 2.23)

and agricultural settings (mean = 1.84), these differences were not statistically
significant (Kruskal-Wallis, H = 3.05, p = 0.22). Similarly, we found no significant
difference in mean betweenness centrality across land use types (H = 2.64, p = 0.27).

In contrast, network modularity differed significantly among habitats (H = 15.9, p <
0.001). *Post-hoc* tests revealed that agricultural networks were significantly more
modular than both village networks (p = 0.02) and forest networks (p < 0.001). The
difference in modularity between village and forest networks was not statistically
significant. This indicates that small-mammal communities in agricultural settings form
more fragmented, clustered interaction networks.

At the species level, degree centrality varied by habitat, with distinct hub species
emerging in different settings (Figure 3). In villages, the synanthropic species *M.*
musculus and *R. rattus* were highly connected, reaching maximum degrees of 14 and
11 with high mean degrees (4.1 and 4.7, respectively). In forests, the native species
Malacomys edwardsi was a key hub, exhibiting the highest mean degree (5.4) and the
maximum individual degree (9) for that habitat. By contrast, the principal reservoir *M.*
natalensis, though abundant, showed more moderate connectivity with a lower
maximum degree and lower mean degrees across habitats (mean = 2.3, max = 8 in
villages and mean = 1.74, max = 6 in agriculture).

Figure 3: The degree (number of contacts) of individual small mammals stratified by species and land use type. Boxes contain the median and inter-quartile range of the degree distribution. Whiskers include the upper and lower quartile with outliers shown as points.

While the mean degree for several synanthropic species appeared numerically elevated in villages, formal comparisons of network-level mean degrees found no statistically significant differences in the degree distributions for species, including *M. natalensis* and *R. rattus*, across land use types (Wilcoxon rank-sum tests, *p* > 0.05 for all comparisons).

Inter- and Intra-specific Mixing Patterns

Species with more detected individuals generally exhibited a greater number of inter-specific contacts (Pearson correlation, $r = 0.73$, $p < 0.001$, $n = 15$). For instance, the frequently detected species, *M. natalensis*, *P. rostratus* and *R. rattus* each had contacts with more than 11 other species. An exception to this trend was *M. musculus*, which, despite being the fourth most frequently detected species, only had observed contacts with four other species (Figure 4 and Supplementary Figures 4.1 and 4.2).

Intra-specific contacts were common for most species, but notable differences emerged across land use types. In agricultural settings, 46% of all contacts involving *M. natalensis* were intra-specific, even while interacting with 12 other species (Figure 4). In village settings, this proportion of intra-specific contact decreased to 28% (Supplementary Figure 4.2). Other species showed a higher degree of inter-specific mixing (e.g., *L. sikapusi* and *M. setulosus*).

Figure 4: The proportion of contacts between individual small mammals in agricultural land use. Darker colors indicate increasing proportions of observed contacts to a species (Contact to) from named species (Contact from). Numbers in the cells correspond to the proportion of contacts to a species from a named species. For example, 45% of all contacts to Mastomys natalensis are from other M. natalensis while 9% of contacts are from Lophuromys sikapusi. Percentages sum to 100% in the `Contact to` *axis, while they may exceed 100% in Contact from. Species are ordered by the total number detected in this study with M. natalensis (N = 113) in the bottom left.*

Modeling Contact Probability in *Mastomys natalensis*

Focusing on the primary reservoir, *M. natalensis*, we used a random-effects meta-analysis to pool the results from 11 individual ERGM models (6 village networks and 5 agricultural networks).

The overall probability of an inferred contact (the `edges` term) was low in both agricultural (OR = 0.04, 95% Confidence Interval = 0.03-0.07, *p* < 0.001) and village settings (OR = 0.15, 95% C.I. = 0.09-0.23, *p* < 0.001), with substantial heterogeneity

observed between individual networks in both land use types ($\hat{\tau}^2_{\text{agriculture}}$ = 0.21, Q = 38,
p < 0.001 and $\hat{\tau}^2_{\text{village}}$ = 0.29, Q = 35, p < 0.001) (Figure 5A).

For inter-specific inferred contacts, *M. natalensis* had significantly reduced odds of
forming a connection with another species in village settings (OR = 0.49, 95% C.I. =
0.33-0.73, p < 0.001). In agricultural settings the odds were also reduced, though this
effect was not statistically significant (OR = 0.61, 95% C.I. = 0.33-1.11, p = 0.1) (Figure
5B). There was no substantial heterogeneity in inter-specific contact odds between
networks ($\hat{\tau}^2_{\text{agriculture}}$ = 0.22, Q = 8.5, p = 0.07 and $\hat{\tau}^2_{\text{village}}$ = 0.09, Q = 8, p = 0.15).

Conversely, *M. natalensis* showed a strong and statistically significantly increase in the
odds of intra-specific inferred contacts in agricultural settings (OR = 5.05, 95% C.I. =
2.14-12, p < 0.001), but not in village settings (OR = 1.96, 95% C.I. = 0.79-4.81, p =
0.15) (Figure 5C). Heterogeneity in intra-specific contact odds was low in both land use
types ($\hat{\tau}^2_{\text{agriculture}}$ = 0, Q = 2.6, p = 0.63 and $\hat{\tau}^2_{\text{village}}$ = 0.14, Q = 6, p = 0.3).

Sensitivity analyses revealed no changes in the direction of effect sizes when altering
the contact radius although the magnitude of the effect sizes varied (Supplementary
Figures 5.1-3). Additionally, leave-one-out sensitivity analyses for influential networks
did not indicate meaningful changes in the effect size magnitude or direction. These
results support the robustness of the findings across space-sharing buffer area
assumptions and community composition changes across visits.

Figure 5: Random effects meta-analysis of ERGM network models reporting the odds of
a contact being observed for M natalensis. A) The odds ratio of a contact being
observed for M. natalensis in Agricultural or Village land use types. B) The odds ratio of
a contact being observed between M. natalensis and an individual of a different rodent

Observed Association Between Network Centrality and Serostatus

We first examined the observed association between an individual's network centrality and its LASV serostatus. Overall, we found no significant difference in the mean degree of seropositive individuals (mean = 2, SD = 2.5) compared to seronegative individuals (mean = 2.5, SD = 3) (Wilcoxon rank-sum test, $W = 11218$, $p = 0.29$).

In species-specific analysis, this lack of a significant association held for most species, including the principal reservoir, *M. natalensis* ($p = 0.43$). However, for the invasive synanthrope *R. rattus*, seropositive individuals had a significantly higher mean degree than their seronegative counterparts (mean 7.75 vs. 4.24, $W = 273.5$, $p = 0.034$). This finding should be interpreted with caution, as it is based on a small number of seropositive individuals (n = 4). We found no significant difference in betweenness centrality between seropositive and seronegative individuals, either for the overall community or within any individual species ($p > 0.05$ for all tests).

Modelling Predictors of Seropositivity Risks in *M. natalensis*

To move beyond correlation and to model the factors that predict seropositivity risk, we fitted a Bayesian generalized linear mixed-effects model for the primary reservoir, *M. natalensis*. This model, which included an interaction term, was favored over simpler models (Leave One Out comparison, ELPD difference >65).

In the final model, an individual's odds of being LASV seropositive increased significantly with its total number of contacts (degree; OR = 2.36, 95% CrI: 1.98–2.81) and its proportion of conspecific contacts (homophily; OR = 10.22, 95% CrI: 5.39–19.71). We also identified a strong negative interaction between degree and homophily (OR = 0.19, 95% CrI: 0.14–0.26).

Discussion

In the Eastern Province of Sierra Leone, we found that the structure of small-mammal contact networks varied by habitat, though not as initially hypothesized. While we found no statistically significant difference in mean degree (connectivity) across land use types, key hub species differed distinctly. Invasive synanthropes like *R. rattus* and *M. musculus* were highly connected in villages, whereas the native species *M. edwardsi* emerged as a central hub in forests. The primary reservoir *M. natalensis* consistently showed a high probability of intra-specific contact in agricultural settings, suggesting these areas may host distinct transmission dynamics. Our most complex findings, however, relate to the link between network position and serostatus.

A key finding of our study is the divergence between our descriptive and modelling analyses of seropositivity. While our predictive model identified a higher degree as a

significant risk factor for seropositivity in *M. natalensis*, our descriptive, non-parametric
analysis found no significant overall difference in the mean degree of seropositive and
seronegative animals. This divergence highlights the challenge of detecting clear
disease signals in static, aggregated network data and suggests the underlying
association is likely obscured by confounding factors that a simple correlational test
cannot account for. The null result in our descriptive analysis is likely due to several
mechanisms. Survival bias may play a key role, where highly-connected individuals face
higher mortality or removal, thus masking the association between high degree and risk
among the surviving population[74]. Most plausibly, the signal is obscured by the temporal
dynamics of LASV. The virus likely persists through metapopulation dynamics and
episodic local fadeout, and our serological data, aggregated over three years, smooths
over these dynamics. In contrast, our model, which accounts for network structure, was
able to detect the underlying signal that higher connectivity does indeed increase
infection risk, a signal that is masked in a simple correlational analysis.

Our predictive model for *M. natalensis* also revealed a strong negative interaction
between degree and homophily. This indicates that while having more contacts
increased the odds of being seropositive overall, this effect was significantly weaker for
individuals whose contacts were primarily with conspecifics. In other words, for highly
connected animals, contacts with other species were more strongly associated with
seropositivity than contacts with their own species. This finding could challenge the
paradigm of *M. natalensis* as the sole driver of transmission, pointing towards a more
complex, community-level maintenance system where inter-species spillover events are
particularly important.

Analyzing the network structures highlights significant ecological heterogeneity across
the land use gradient. The upper tail of the degree distribution was skewed towards
individuals in villages and agriculture, underscoring the importance of individual-level
variation and a few highly connected hub individuals not captured by aggregated
metrics[75,76]. Furthermore, agricultural networks were more modular. While this suggests
these communities may be fragmented into distinct clusters, a structure that could
facilitate intense local transmission while slowing landscape-level spread, this finding
should be interpreted with caution. The observed modularity may be partially influenced
by the spatial arrangement of trapping grids in heterogeneous agricultural environments
rather than reflecting social clustering alone. The high species richness and greater
proportion of inter-specific space-sharing in agricultural areas also likely reflects edge
effects that facilitate interactions between synanthropic and sylvatic species, creating
opportunities for inter-species spillover[77,78]. Conversely, the high connectivity of invasive
synanthropes in villages may contribute to the lower-than-expected species richness
observed in those habitats through competitive exclusion or other mechanisms[79].

The structure of *M. natalensis* interactions also varied by land use. In agricultural
settings, *M. natalensis* exhibited significantly higher odds of intra-specific clustering, a
pattern consistent with the species' weak territoriality[80–82]. In villages, this strong
tendency for intra-specific contact was not statistically significant. This dynamic could
amplify intra-specific transmission chains in agricultural landscapes while potentially
diluting transmission pathways in villages where inter-specific encounters may be

proportionally higher. Movement between habitats by individuals driven by resource
availability may further modulate these dynamics[83].

The seroprevalence of LASV (5.7%), was consistent with prior estimates from Sierra
Leone[16]. Our study included forest sites further from human habitation, yet the
proportion of *M. natalensis* individuals testing positive was similar (~9%). Our detection
of LASV antibodies in nine distinct species, with *M. natalensis* comprising only 28% of
seropositive individuals, confirms previous reports of LASV exposure across a range of
small-mammal species in West Africa[12,18]. While seropositivity indicates past infection, it
does not confirm reservoir competence or define a species' role in active transmission.
Nonetheless, this finding underscores the importance of considering a multi-host
community perspective when studying LASV ecology[84].

These complex ecological findings have direct public health implications. The distinct
network structures suggest different human risk profiles by land use: the high
intraspecific connectivity of *M. natalensis* in agricultural settings may amplify the virus
within the primary reservoir, posing a direct spillover risk to the members of these
communities that most utilize agricultural landscapes. In villages, the multi-host
seropositivity and denser networks suggest a more diffuse risk to residents from a wider
range of species. Furthermore, our finding that seropositive animals do not, on average,
have higher connectivity implies that surveillance strategies focused only on the most
socially central animals may be insufficient. A more comprehensive sampling approach
is likely necessary to accurately assess community-level prevalence. Notably, we
detected no seropositive animals in the most urbanized study site (Lambayama),
suggesting factors associated with increased urbanicity, potentially including shifts in
community composition towards invasive species or altered habitat structure, might
disrupt local transmission cycles, though further investigation is needed. Ultimately, the
confirmation of a multi-host serological landscape, despite low absolute numbers of
seropositive individuals for many species, suggests that interventions targeting solely
the primary reservoir may be insufficient or yield counterintuitive outcomes in complex
ecological systems[85].

Our study's methodology rests on a series of assumptions that introduce cumulative
uncertainty. Contacts were inferred from spatio-temporal co-occurrence, which assumes
that proximity is a reliable proxy for interactions capable of transmission[38]. Our
approach also assumes that a capture location represents a central point within an
individual's short-term activity space[58]. Furthermore, removal trapping may have
created temporary spatial vacuums, and unobserved behavioral differences between
captured and uncaptured animals could bias inferences[86,87]. While our study design (≥3
months between sessions), sensitivity analyses, use of species-specific home range
radii, and leave-one-out meta-analysis sought to mitigate these issues, we acknowledge
our results represent a model of a complex system. Ultimately, replicating the study
across a greater number of sites would be valuable for evaluating the generalizability of
these patterns.

In conclusion, this study highlights the variability in small-mammal contact networks
across a land-use gradient in a Lassa fever-endemic region. While we found no direct
association between land use and seroprevalence, our results reveal that the ecological

drivers of transmission risk are complex and scale-dependent. The divergence between
our descriptive and modeling results suggests that static snapshots may not fully
capture dynamic disease processes. The serological evidence suggesting multi-host
LASV exposure, alongside the habitat-specific roles of hub species underscore the
importance of tailoring surveillance and control strategies to local ecological contexts to
mitigate Lassa fever risks effectively.

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