A resurrection experiment reveals reduced adaptive potential in a common agricultural weed

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Abstract

Adaptive evolution is critical to population responses under accelerating anthropogenic global change. Although theory and some empirical work suggest that contemporary rates of environmental change can be met by equally rapid evolutionary shifts, mismatches between expected and realized evolution remain common in natural systems. Using a resurrection approach, we demonstrate that the structure of genetic covariances among pollination and mating-system traits in the common morning glory (*Ipomoea purpurea*) is increasingly constraining evolutionary trajectories. Pollinator-attracting traits evolved on a rapid time scale of less than ten years, but genetic covariances between traits limited the adaptive rate and evolutionary trajectory of flowering phenology, a trait with widely recognized importance in adapting to climatic shifts. Our results show that changing patterns of genetic constraint can substantially reduce adaptive potential even when selection favors rapid trait evolution.

Introduction

Understanding whether evolutionary adaptation can keep pace with rapid environmental shifts is critical to understanding current biodiversity declines and the continued resilience of ecosystems to global change (1, 2). Species can respond to rapid environmental changes by shifting their geographic range, plastic trait responses, and *in situ* adaptation. Geographic range shifts can be limited by large-scale habitat fragmentation and destruction (3, 4) whereas compensatory responses via plasticity have a limiting upper bound (5, 6). Taken together, these limits suggest that long-term, sustained responses to climatic changes rely heavily on *in situ* adaptation. A major unresolved question relating to adaptive evolution is how adaptive potential, defined here as genetic variation that can lead to a response to selection, may be changing in wild populations facing contemporary environmental pressures.

Given environmental change, a population's adaptive capacity depends on the presence and availability of genetic variation, as well as how efficiently selection acts on phenotypic variation. Trait-trait covariances can slow or prohibit evolutionary response if directional selection on genetically linked traits is misaligned with the sign of the trait covariance (7–9).The presence of trait covariance in wild populations (10, 11) as well as the possibility for these covariances to induce lags in adaptive rate in response to artificial selection (12–14) are both well documented. However, genetic correlations can change over time in response to novel environments (9, 15), with one of the strongest driving causes being conditions of fluctuating correlational selection (16, 17). Global change is exposing wild populations to multiple, simultaneous shifts in abiotic and biotic variables. Evidence suggests that historically coupled selective agents may be decoupling (18), creating a context in which misaligned selection or changing correlational selection is becoming increasingly probable. Characterizing how genetic correlations may change in the face of global change, however, remains a largely unmet challenge. As such, the causes and consequences of constraint are open questions that pertain directly to projections of adaptive rate and population viability in the face of global change.

The potential for constraints on climate change adaptation induced by among-trait genetic correlations has been assessed using a space-for-time approach (7), but spatial gradients of climatic change can be misleading proxies for climatic change over time or population response to that change (19). Non-climatic variables may decouple species response from climatic changes, or adaptation to local conditions can result in localized declines given climatic shifts rather than adaptation to spatially predicted trait means (20). Resurrection approaches mitigate these issues by comparing temporally sampled lineages from the same locations, thus removing spatial differences among populations altogether. While resurrection ecology is increasing in popularity for measuring phenotypic shifts, it has yet to be applied to assess the real-time evolution of constraints. Here, we combine a resurrection approach with analyses of the genetic variance–covariance structure underlying floral traits in wild morning glory (*Ipomoea purpurea*) populations and address the question of what drives the evolution of constraint by evaluating trait-specific trade-offs and pollinator visitation.

Approach

We conducted a resurrection experiment using replicate maternal lines of *Ipomoea purpurea* sampled at two time points (2003 and 2012; hereafter ancestral and descendant, respectively) as seed from seven wild populations (Fig. S1). Replicates of each maternal line were planted into field conditions at the Matthaei Botanical Gardens (Ann Arbor, MI) on June 4, 2022. During the experiment we recorded trait values for five floral traits (date of first flower, corolla width, corolla length, anther-stigma distance, and nectar sucrose content (°Bx)) along with plant size (leaf count) and fitness (total seed count). These measurements were used to calculate heritable genetic variation for each trait and a multivariate statistic, *R,* which compares the expected rate of adaptation, *W*, given a selection gradient, **β,** and variance-covariance matrix, *G*, to the rate of

adaptation with the same selection gradient but a variance-covariance matrix with all trait
covariances constrained to zero, W_{nc} [21]. Estimates of R < 1 demonstrate evolutionary
constraint such that the rate of adaptation is slower than expected without covariance.
Conversely, R > 1 would indicate evolutionary facilitation. We then investigated each trait
individually by comparing predicted evolutionary change with ($\Delta\bar{z}$) and without ($\Delta\bar{z}_{nc}$) trait
covariance to identify which floral traits influence the rate of adaptation (R). We further dissected
temporal patterns of selection and potential genetic correlations between traits to detect instances
where new or emerging pairwise trait correlations act in an antagonistic direction to selection,
indicating evolution of constraints or trade-offs between traits [22]. Finally, we used pollinator
visitation surveys and structural equation modeling to determine if fitness costs and benefits
associated with traits are mediated through pollinator visitation frequency and thus whether
adaptation in pollinator-attracting traits may constrain the response to other possible selective
agents.
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91 **Results**
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Trait Covariance is Causing a Rapid Decline in Adaptive Rate Over Time
In both the ancestral and descendant populations, across traits, the overall rate of adaptation is
lower when genetic covariances between traits are considered. Notably, however, evolutionary
constraint is much stronger in the descendant populations with R = 0.09 compared to R = 0.76 in
the ancestral populations (99.85% significance threshold; Fig. 1A). Thus, in the ancestral
populations, accounting for trait covariances results in an adaptive rate that is 76% as fast as
predictions without trait covariances, but this predicted rate drops precipitously to 9% in the
descendant populations, indicating that constraints have increased over time. An alteration to the
R statistic, R' (with values >1 indicating greater adaptive potential in ancestral populations)**,** that
compares the ancestral and descendant populations directly (see SI Methods) shows that
predicted rates of adaptation without covariance are largely sustained over time, though perhaps
marginally faster in ancestral populations, R_{nc}' = 1.81. However, when accounting for trait
covariance, R' = 23.61, demonstrating that, contrary to univariate predictions, the adaptive rate
was 23x faster in ancestral populations - *i.e.* over nine years the populations lost 96% of their
adaptive potential. These findings indicate that constraints among floral traits evolve on a rapid
time scale of less than ten years and are acting to limit adaptive potential, overall highlighting that
consideration of multi-trait dynamics is critical to understanding the impact of global change on
patterns of biodiversity.
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We next compared expected adaptation of individual traits with and without covariance ($\Delta\bar{z}$ to
$\Delta\bar{z}_{nc}$) to identify which traits underlie observed shifts in R. We found significant (>95%
significance threshold) constraint on the rate of adaptation in corolla width as well as a change in
the direction of evolution for corolla length in both the ancestral and descendant populations
(corolla width, ancestral: Δz_{cw} = 0.02, Δz_{nc_cw} = 0.13, descendent: Δz_{cw} = 0.01, Δz_{nc_cw} = 0.09;
corolla length, ancestral: Δz_{cl} = 0.01, Δz_{nc_cl} = -0.07, descendent: Δz_{cl} = 0.01, Δz_{nc_cl} = -0.07;
Table 1, Fig. 1B). Comparatively, for anther-stigma distance (ASD) and day of first flowering, we
find that genetic covariance between traits in ancestral populations results in adaptive facilitation
such that expected adaptation of smaller ASD (Δz_{ASD} = -0.07, Δz_{nc_ASD} = -0.002; 96.55%
significance threshold) and earlier flowering (Δz_{DoFF} = -0.14, Δz_{nc_DoFF} = -0.07; 99% significance
threshold) is faster than without covariance. Descendant populations, however, display no effect
of covariances on ASD, and show adaptive constraint on flowering phenology such that the
adaptive rate of flowering time is slowed by covariance with other traits (Δz_{DoFF} = -0.03,
Δz_{nc_DoFF} = -0.05; 92% significance threshold; Table 1, Fig. 1B). A comparison of Δz between
years shows that there is a decrease over time in expected rate of adaptation toward earlier

flowering (Table 1, 99.85% significance threshold) and plant size (ancestral Δz_{size} = 0.19,
descendant Δz_{size} = 0.01; 99.95% significance threshold).
129
Constraints Result in New Trait Optima
Two commonly proposed mechanisms for lack of evolutionary change are an absence of
heritable genetic variation or trade-offs in the form of negative genetic correlations or covariances
(9); either, or both, of these mechanisms could explain the results on adaptive rate. However, we
found each of the floral traits to exhibit heritable genetic variation across collection years (Table
1). This means that lack of appreciable genetic variation within individual traits will not induce
evolutionary constraint, further corroborated by variation maintained in $\boldsymbol{W_{nc}}$ but not in $\boldsymbol{W}$ (Fig. 1A).
Analysis of genetic correlations showed that, while trade-offs between traits were present and
generally took the form of negative genetic correlations between floral traits and plant size, most
remained similar in sign and strength between the ancestral and descendent populations (Fig.
1C, S2). Two notable exceptions are the correlation between corolla width or corolla length and
date of first flower (corolla width and date of first flower: ancestral r = -0.005, p = 0.481,
descendant r = 0.34, p < 0.0001; corolla length and date of first flower: ancestral r = 0.02, p =
0.627, descendant r = 0.22, p = 0.003). The emergence of a positive correlation between corolla
size and flowering date in descendant populations indicates that a new trade-off has evolved
between larger flowers and earlier flowering.
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We likewise identified different trait optima in the descendent versus ancestral populations
favoring evolutionary investment in pollinator-attracting traits at the cost of evolvability in flowering
phenology. In the common garden setting, where the selection regime is shared, a difference in
selection coefficient between the ancestral and descendent populations means that there is a
different fitness response for a given trait in that shared environment. As such, we report
comparisons of total and direct selection on traits to show shifts in adaptive capacity.
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We found that total selection on the date of first flower differed between ancestral and
descendant populations (ANCOVA contrasting slopes: F = 5.39, p = 0.020, Table S2). In
ancestral populations, we uncovered higher fitness in earlier flowering plants (S = -0.20, p = 0.02;
Table S2) as well as a marginally-significant phenotypic shift to earlier flowering in descendant
populations by ~1 day (t-ratio = 1.78, p = 0.07; Fig. S3, Table S1, *also shown as a significant shift*
in greenhouse conditions, (23)). In contrast, in line with the constraint indicated by Δz_{DoFF} and a
trade-off with flower size, earlier flowering dates are no longer favored in descendant populations,
and instead highest fitness is found for intermediate flowering dates (γ = -0.12, p = 0.002; Table
S2). Descendent populations also displayed the emergence of marginally-significant, but novel
correlational selection between date of first flower and three other traits: corolla width (γ = -0.34, p
= 0.071), nectar sucrose content (γ = -0.14, p = 0.067), and corolla length (γ = 0.39, p = 0.045,
Fig. 2), all featuring a fitness ridge favoring intermediate values of flowering time.
166
Despite constraints on the rate of adaptation of corolla width, both corolla width and °Bx were
under selection toward greater investment in pollinator-attracting traits of larger floral display and
higher sugar content in nectar. Specifically, corolla width displays significant positive selection in
both the ancestral and descendent populations (S = 0.14, p < 0.001 ancestral; S = 0.14, p <
0.001 descendent; Table S2); further, the continued evolution of broader corollas is corroborated
by a significant increase in mean trait value from ancestral to descendant populations by ~1mm
(t-ratio = -2.53, p = 0.01; Fig S4, Table S1). Although total selection indicates positive selection
for corolla length (S = 0.11, p = 0.002 ancestral; S = 0.06, p = 0.031 descendant), analysis of
direct selection shows no selection in ancestral populations and negative selection on corolla
length in descendant populations (β = -0.37, p = 0.03; Table S3). Thus, the positive total selection
on corolla length likely arises from the strong correlation between corolla length and width (r =
0.92, p < 0.001) and the overall positive selection on corolla width. For nectar, ancestral
populations display a pattern of disruptive selection where both high and low levels of sucrose are
favored (γ =0.07, p = 0.02; Table S2), however, there is a marginally-significant change in the

quadratic relationship between fitness and °Bx in descendant populations (ANCOVA contrasting
slopes: F = 3.40, p = 0.066, Table S2) such that lower sucrose values are no longer favored and
instead there is selection for high values of °Bx (S = 0.07, p = 0.04, Table S2).
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Overall, these data suggest that the evolution of earlier flowering times in descendant populations
is limited by selection for increased corolla width (Fig. 1B & 1C, Tables S2 & S3), with a
concomitant increase in nectar sucrose. These changes are additionally accompanied by a
decrease in relative fitness of descendant populations compared to ancestral (ancestral = 1.06,
descendant = 0.93; p = 0.001), indicating that, at least in some environments, reduced adaptive
capacity is also resulting in lowering fitness. Beyond the rapid evolution of trade-offs, our results
show that constraints due to changes in genetic architecture have evolved within the nine-year
period studied, and that the covariance structure—particularly between floral traits and the timing
of flowering—may remain unstable as global changes continue.
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Pollinators as Agents of Selection
We next examined the potential that pollinators impose selection on both corolla width and nectar
sucrose content in *Ipomoea purpurea* using structural equation modeling. Our best supported
model (Table S4, S5) indicates that increases in corolla width and nectar sucrose content
between ancestral and descendant populations are driven by pollinator preference, whereas
changes in flowering phenology are not responsive to pollinator foraging frequency, but rather
some other environmental factor. Specifically, we find a positive effect of corolla width on
frequency of pollinator floral approaches (r = 0.242, p = 0.02) and pollinator visits (r = 0.286, p =
0.004) as well as a positive relationship between nectar sucrose content and pollinator
approaches (r = 0.149, p = 0.002), meaning that large corolla widths and high sucrose content in
nectar attracts more pollinators to a flower (Fig. 3, Table S6). An additional high conversion of
pollinator approach to foraging (r = 0.366, p < 0.0001) and positive relationship between pollinator
foraging and plant fitness (r = 0.173, p = 0.032) results in an indirect relationship between corolla
width and nectar on plant fitness where high values of each are favored by pollinators and result
in higher plant fitness. Conversely, earlier dates of first flower have a positive effect on plant
fitness (r = 0.321, p = 0.005), but do not impact pollinator foraging behavior, meaning that
selection on flowering phenology is driven by something other than pollinators, an observation
corroborated by numerous studies showing that flowering phenology is core to adaptation to
global change factors such as temperature (24, 25) and precipitation (26). Overall, and in
conjunction with the development of constraints on flowering phenology due to trade-offs with
corolla width and first flower, our results highlight an adaptive pathway that is driven by
investment in pollinator attraction at the cost of evolvability in flowering phenology, a shift that has
occurred within a nine-year period.
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219 **Discussion**
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A central question in evolutionary biology is whether adaptive evolution can keep pace with rapid,
multi-dimensional environmental change. Although models often predict rapid evolutionary
responses, empirical studies repeatedly uncover lags between expected and realized adaptation
in the wild (2). Using a resurrection framework, we demonstrate that genetic constraints among
floral traits in *Ipomoea purpurea* have strengthened markedly over just nine years, resulting in a
96% reduction in predicted adaptive rate. These results provide a mechanistic explanation for
why adaptation may stall even in populations with abundant genetic variation: rapid evolution of
trait covariances can fundamentally reorganize the genetic architecture that determines how
efficiently selection translates into evolutionary change.
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Our findings reveal that constraints themselves evolve and do so on ecologically relevant
timescales. Ancestral populations showed moderate constraint, whereas descendant populations
exhibited strong antagonistic covariances—particularly involving corolla width and flowering
phenology—that greatly reduced predicted evolutionary response. Importantly, heritable variation

was present for all focal traits, indicating that limits to adaptation arise not from the loss of genetic
variance, but from the reorganization of variance across traits. Strengthened correlations between
floral size and flowering time in descendant populations mark the evolution of a new trade-off
linking investment in pollinator attraction to delays in flowering, thereby shifting multivariate
adaptive trajectories.
240
Pollinators appear to be central to this shift. Structural equation models show that larger corollas
and higher nectar sucrose content increased pollinator visits and, indirectly, plant fitness,
whereas pollinators did not mediate the fitness effects of flowering time. This asymmetric
influence means that selection by pollinators pushes floral size upward but simultaneously
creates antagonistic covariance with flowering phenology. In effect, adaptation toward greater
pollinator attraction reshapes the covariance structure in ways that constrain the evolution of
earlier flowering, a trait strongly tied to climatic adaptation in many plant species. These findings
align with emerging evidence that global change factors may not act independently, but instead
impose conflicting (18) or interacting (27) selective pressures that channel adaptation in opposing
directions.
251
Several considerations provide context for these inferences. First, distinguishing drift from
selection as the mechanism behind changes in covariance structure is challenging. However, if
drift were the major agent underlying the changes in covariance structure that we document, we
would expect to see proportional change in variances and covariances across traits (28, 29).
However, we observed changes in genetic correlations that were uneven across the traits (Fig.
1C) and that the evenness of **G** (E_λ) declines from ancestral to descendant populations—
indicating that genetic variance becomes less evenly distributed and that drift is an unlikely
mechanism ($E_{\lambda_{ancestral}}$= 0.62, $E_{\lambda_{descendant}}$= 0.49, 21). Second, we used selfed maternal lines in our
common garden experiment, which allowed us to reduce genetic heterogeneity and more directly
attribute differences between sampling years to evolutionary rather than mating-system variation.
We note that this approach may not fully capture covariance structure as expressed in naturally
outcrossed populations; however, given that *I. purpurea* freely self-pollinates across much of its
range (30), selfed lines offer a biologically appropriate and experimentally tractable means of
minimizing within-line variation and isolating temporal genetic differences. Finally, previous work
indicates high genetic connectivity among *I. purpurea* populations (31), so we standardized traits
within populations and analyzed them collectively to estimate species-wide responses to an
average pattern of selection rather than focus on population-specific differences. As a diagnostic,
we also calculated ***R*** and **Δz̄** for each population (Table S9), but models fitted separately had
limited power to estimate higher-order interaction terms. Pooling populations therefore provides a
more robust estimate of the typical strength and sign of these interactions. Ultimately, these
caveats do not diminish the core conclusion: the covariance structure underlying floral traits has
evolved rapidly, with major consequences for adaptive potential.
274
Prior work has emphasized shifts toward increased selfing (*Viola praemorsa*, (32); *Viola arvensis*,
(25)), investment in pollinator attraction due to reduced pollinator abundance (*Lobelia siphilitica*
(33)), or climatically-induced shifts in flowering phenology (34). Together, our results identify
another pathway by which plants may respond to global change: the evolution of genetic
constraints themselves. We show that selection favoring pollinator-attracting traits may indirectly
restrict the evolution of flowering time, a trait critical for climatic adaptation. This has important
implications for population persistence.
282
Recent decades have seen significant reductions in pollinator populations (35–37), with
unprecedentedly high declines of bumblebee species in North America specifically between the
years of 2000-2014 (38) when these experimental *Ipomoea purpurea* populations were collected.
While the evolution of flowering time is well established as a key mechanism for flowering plants
to adapt to climatic changes in temperature and precipitation (39–42), our research suggests that
declining pollinator populations could undermine resilience to climate change through adaptive

constraint–specifically, the diminished adaptive potential in flowering phenology is at odds with
projected climatic shifts that continue to alter temperature and precipitation patterns (43). These
findings thus highlight the necessity of considering multiple dimensions of global change, and
emphasize that adaptation to one factor, such as reduced pollinator populations, could conflict
with the adaptive responses required for other factors, like climate-induced shifts in temperature
and precipitation.
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297   **Materials and Methods**
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Study System and Sampling Design
Ipomoea purpurea (Convolvulaceae) is an annual, mixed-mating vine native to the subtropical
Americas and common across disturbed eastern and midwestern U.S. landscapes. Its large,
showy flowers exhibit substantial natural variation in morphology and reward traits, and the
species is visited by a broad assemblage of generalist pollinators (44). These features make *I.
purpurea* well suited for examining contemporary evolution in traits influencing pollination and
mating system dynamics.
306
To assess evolutionary change over time, we conducted a resurrection experiment using seeds
collected from seven wild populations in South Carolina, North Carolina, and Tennessee at two
sampling points: 2003 (ancestral) and 2012 (descendant) (Fig S1). For each seed cohort, two
greenhouse refresher generations were produced to reduce maternal environmental effects.
Given the species' mixed mating system (population outcrossing rates 0.2–0.8, 30), selfed
descendants remain representative of natural genotypes while enabling replicated evaluation of
line means. Three replicates of each line were grown in 4-inch pots in a growth room with daily
watering and a 12-hr light cycle. Plants were allowed to naturally self-pollinate, and all seeds
were collected.
316
In May 2022, we established a common garden at the University of Michigan Matthaei Botanical
Gardens. We planted 1,536 scarified seeds representing 3–8 maternal lines per population and
five replicate plants per line within each of three spatial blocks (2 years x 7 populations x 3-8
maternal lines x 5 replicates x 3 blocks). Ancestral and descendant cohorts were fully
randomized within blocks. Plants were staked, watered as needed, and monitored throughout the
growing season. Germination exceeded 98%; however, vole herbivory reduced the final sample
size to 839 individuals.
324
325   We quantified five traits associated with pollinator attraction, mating system, and phenology:
326   corolla width, corolla length, anther–stigma distance (ASD), nectar sucrose content (°Bx), and
327   date of first flowering (Julian date). Leaf count at flowering was recorded as an index of plant size
328   and included as a covariate in selection models.
329
330   For each plant, four flowers sampled across the season were measured for floral morphology and
331   nectar reward. Caliper measurements were recorded to ±0.01 mm. Nectar sucrose was estimated
332   using standardized nectar extraction followed by refractometry. Flowers were bagged the night
333   before sampling to prevent nectar removal. Total seed set per plant served as the estimate of
334   fitness. Full measurement protocols are provided in SI Methods.
335
Pollinator Observations
Pollinator visitation was surveyed weekly from August through early October. Observations
occurred between 10:00–13:00 under suitable weather conditions and followed a standardized
20-min focal plot protocol. Across nine subplots (28–33 plants each), 277 focal plants were
observed, generating 1,116 recorded visits. Pollinators were identified to order and
morphospecies based on body size to maintain consistency across observers. We scored

"approach" (inspection without contact) and "visit" (contact with anthers or stigma). Explicit
procedures and morphogroup definitions appear in SI Methods.
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345 **Genetic Variation and Potential for Adaptation**
346 To determine whether traits possessed genetic variance necessary for adaptive evolution, we fit
347 random effect models for each trait within each year:
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349 Trait ~ Pop:ML + Block
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where Pop:ML represents maternal line nested within population, and Block accounts for spatial
heterogeneity across the garden. Significance of the maternal-line effect was assessed via a
likelihood-ratio χ^2 test. Broad-sense heritability (H^2) was estimated as the proportion of trait
variance attributable to maternal line relative to total phenotypic variance. We repeated analyses
within each population to confirm within-population variation. One population lacking variation in
corolla width was excluded from subsequent analyses, as our analyses and approach require
genetic variance in the focal traits. Additional model diagnostics are in SI Methods.
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359
360 **Cross-Trait Constraints**
361 To measure the impact of genetic covariances on the rate of adaptation across traits, we follow a
362 procedure described in Agrawal and Stinchcombe, 2009 ([21]). We first estimated two variance-
363 covariance matrices using trait values standardized such that maternal line variance is unity. The
364 first, **G,** is calculated using a multivariate mixed effects model
365
366 Measurement ~ Trait Type + Maternal Line + Block
367
where 'Measurement' is a single measurement of a given trait, 'Trait Type' is one of the six
measured phenotypic traits, and maternal line and block are both included as random effects.
The second matrix, $\mathbf{G}_{nc}$, has genetic covariances manually constrained to zero such that the
diagonal of the matrix contains trait variances, but all other terms in the matrix are set to zero.
Vectors of predicted response to selection, $\mathbf{\Delta\bar{z}}$ and $\mathbf{\Delta\bar{z}}_{nc}$ were obtained by multiplying **G** and $\mathbf{G}_{nc}$
by a vector of selection coefficients, $\mathbf{\beta}$ (for information on $\mathbf{\beta}$ estimation see *Correlational Selection*
section of SI Methods). Defining the rate of adaptation as the rate of increase of fitness of the
mean phenotype, and approximating fitness by a quadratic function, we use the following
equation for change in fitness of the mean phenotype [(21)]:
377
$$\Delta W(\bar{z}) = \Delta z^{-T}\beta + 0.5\Delta z^{-T}\gamma\Delta\bar{z}$$
379
We then use $\Delta\mathbf{W(\bar{z})}$ and $\Delta\mathbf{W}_{nc}\mathbf{(\bar{z})}$ to compute a single, statistic, $\mathbf{R} = \frac{\Delta W(\bar{z})}{\Delta W_{nc}(\bar{z})}$ showing response to
selection. This provides an intuitive metric for quantifying the rate of adaptation with covariances
relative to the expected rate without them. If $\mathbf{R}$ = 0, then there is an absolute constraint on
adaptation. If $\mathbf{R}$ < 1, then evolutionary constraint is present; for example, if $\mathbf{R}$ = 0.5 the fitness of
the mean phenotype is increasing only 50% as quickly as if the traits were genetically
independent. Conversely, if $\mathbf{R}$ > 1, then genetic covariances are facilitating evolution and
accelerating the pace of adaptation.
387
We used bootstrap resampling to test whether the difference between evolutionary response with
and without accounting for the covariance structure was significant. For each of 2000 iterations,
we resampled individuals and then estimated standardized $\mathbf{\beta}$, **G**, and $\mathbf{G}_{nc}$ for each bootstrap
sample and calculated $\mathbf{\Delta\bar{z}}$, $\mathbf{\Delta\bar{z}}_{nc}$, and $\mathbf{R}$. If 95% or more of the calculated $\mathbf{R}$ values were <1, we
concluded that there is significant evolutionary constraint acting among traits. Each step of this
analysis was performed separately for plants collected from ancestral and descendant
populations. $\mathbf{R}_{ancestral}$ and $\mathbf{R}_{descendant}$ are then compared such that if $\mathbf{R}_{ancestral}$ - $\mathbf{R}_{descendant}$ > 0 for

95% or more of the resampled populations, then we conclude that genetic covariances constrain
evolution less in ancestral populations.
397
A comparison of R across years provides a metric for changes in how strongly evolutionary
constraints impact adaptive rate (*i.e.* the discrepancy in response to selection with and without
covariance), however it does not directly measure whether the absolute adaptive rate changes
between ancestral and descendant populations. To do this, we implement a modified R statistic,
R' where $R'_u = \frac{\Delta W_{nc}(\bar{z})_{ancestral}}{\Delta W_{nc}(\bar{z})_{descendant}}$ represents a predicted difference in response between ancestral
and descendant populations to selection without constraints (*i.e.* a **u**nivariate analysis of how
adaptive potential changes over time). Similarly, $R'_m = \frac{\Delta W(\bar{z})_{ancestral}}{\Delta W(\bar{z})_{descendant}}$ represents a predicted
difference in response between ancestral and descendant populations to selection with
constraints (*i.e.* a **m**ultivariate analysis of how adaptive potential changes over time). If $R' < 1$ for
95% or more of the resampled populations, we conclude that the adaptive rate is lower in
ancestral populations, meaning that adaptive capacity is sustained or even increasing over time.
Conversely, if $R' > 1$ for 95% or more of the resampled populations, we conclude that the
adaptive rate is greater in ancestral populations, meaning that adaptive capacity is decreasing
over time.
412
Trait-Specific Constraints
The existence of evolutionary constraint across a suite of traits does not guarantee every trait is
impacted in the same manner. To assess which traits underlie changes in R and R', we compare
$\Delta\bar{z}_{ancestral}$ to $\Delta\bar{z}_{descendant}$ for each trait and $\Delta\bar{z}$ to $\Delta\bar{z}_{nc}$ for each trait and year independently. If 95%
or more values of Δz_i from the 2000 bootstrap samples described above are less than Δz_{i_nc}, we
conclude that evolution in that trait in that year is subject to adaptive constraint. If there is no
significant difference between Δz_i and Δz_{i_nc}, then there is no constraint or facilitation on that trait,
and if $\Delta z_i > \Delta z_{i\_nc}$, then the trait is subject to adaptive facilitation. If 95% or more values of
$\Delta\bar{z}_{descendant} < \Delta\bar{z}_{ancestral}$, then we conclude that predicted adaptive rate with covariance has
decreased over time, whereas if $\Delta\bar{z}_{descendant} > \Delta\bar{z}_{ancestral}$, then predicted adaptive rate has
increased and no difference between the two indicates a stable adaptive rate.
424
425   **Trait Correlations & Conflicting Selection**
426   We quantified Pearson correlations among maternal-line means for all trait pairs and tested for
427   temporal shifts in correlations via 10,000 bootstrap replicates. Directional agreement or conflict
428   between direct selection (**β**) and trait correlations was evaluated to determine whether emerging
429   trade-offs aligned with observed covariance evolution.
430
To assess total selection on traits as a metric for adaptive capacity, we estimated selection
differentials (*S*) from univariate regressions of relative fitness on each trait, with traits
standardized to variance of one and block as a random effect. Total balancing or disruptive
selection were assessed with a near-identical model including both the linear and quadratic terms
for a given trait. Selection gradients (β, γ) were obtained from a multivariate regression including
all traits simultaneously. Traits were standardized such that each maternal-line variance equaled
one. Significance was assessed using Type III ANOVA. Extended correlation tests and selection
models provided in SI Methods.
439
Pollinators as Agents of Selection: Structural Equation Models
To test whether evolution may be mediated by pollinator behavior, we evaluated seven *a priori*
structural equation models representing alternative causal structures linking floral traits, pollinator
visitation (approach and visit), and fitness in the ancestral populations. The hypotheses tested
include five nested hypotheses regarding the extent to which pollinators mediate the impact of
floral traits on fitness, and two non-nested hypotheses - one testing whether relationships are
correlated but not causal and another testing whether all trait variation is determined by plant
size. Analyses were conducted using piecewiseSEM in R, which accommodates mixed-model

components and non-Gaussian data types such as count (45, 46).
449

Model fit was assessed using a log-likelihood χ^2 statistic comparing model-implied and observed
variance-covariance relationships; non-significant ($p > 0.05$) χ^2 indicates adequate fit. One model
lacked degrees of freedom and was assessed via AIC comparisons. Path coefficients and indirect
effects were estimated for only the best-fit model and statistical power was assessed using
pwrSEM (47). All *a priori* hypotheses, model structures and fit statistics are included in SI
Methods and Tables S4–S7.
456

457   **Realized Phenotypic Evolution**
458   To evaluate phenotypic change between years, we fit mixed models with year, population, and
459   their interaction as fixed effects and block as a random effect:
460

461   Trait ~ Year*Population + Block
462

463   Traits were log-transformed to improve normality. Post hoc comparisons of estimated marginal
464   means were conducted using the emmeans package in *R* (45, 48). Full model outputs appear in
465   SI Methods.
466

467   **Acknowledgments**
468

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R.S.B. conceived and designed the study with R.S.B. acting as advisor. S.B. collected and
analyzed the data and J.S. provided critical review and analysis recommendations. S.B. and
R.S.B drafted the initial manuscript with all authors contributing to the final version. **Competing
Interest Statement:** The authors have no competing interests to declare.
483

485   **Data and materials availability:** Primary data and R code used in these analyses are made
486   available through the following digital dryad repository that will be made public upon acceptance:
487   https://datadryad.org/dataset/doi: 10.5061/dryad.2z34tmpz8

490   **References**
491

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612

    Figures and Tables

615
Figure 1. Adaptive potential decreases over time due to a trade-off between corolla size
and flowering phenology. (A) Bootstrap estimates of **R**, plotted as the two component parts of
the ratio ∆Wz and ∆Wnc(z). Black points indicate resampled **R** estimates for ancestral
populations, and grey points indicate resampled **R** estimates for descendant populations. Ellipses
correspond to confidence intervals of 10, 34, 50, 68, 90, 95, and 98% with the respective color
gradient ranging from red to light blue. Actual **R** estimates are reported in the top right. (**B**) The
expected response to selection with (Δz_i) and without (Δz_{nci}) trait covariances (black and orange
bars, respectively) for each trait. Traits with an asterisk had a significant difference between Δz_i
and Δz_{nci} based on bootstrap resampling. **C)** Pairwise correlations between traits for both
ancestral and descendant populations calculated as Pearson's correlation coefficient. Traits
displaying a significant change in correlation over time based on bootstrap resampling are shown
with a "-" indicating that the direction of joint selection is antagonistic to the direction of correlation
or a "+" indicating that the direction of selection is reinforcing to the direction of trait correlation.
Asterisks next to each bar indicate a significant correlation within year ($p < 0.1\sim$, $p < 0.05*$, $p <$
$0.01**$, and $p < 0.001***$).
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632
Figure 2. Correlational selection is present in descendent but not ancestral populations.
Fitness surface for correlational selection in ancestral and descendant populations acting upon
date of first flower and three other floral traits tested using selection gradients including all linear,
non-linear, and interacting terms (multivariate selection). Relative fitness is depicted by the color
gradient with red being the highest fitness, yellow intermediate fitness, and blue lowest fitness. In
ancestral populations, no indication of correlational selection was present. All traits were included
in the selection gradient, only those with selection gradients indicative of correlational selection in
descendant populations are shown: corolla width (γ = -0.34, p = 0.071), corolla length (γ = 0.39, p
= 0.045), and °Brix (γ = -0.14, p = 0.067).

643
644 **Figure 3. Pollinator foraging behavior mediates fitness effects of floral size and quality**
645 **traits.** Best-fit SEM showing the direct and indirect effects determining pollinator visitation rate
646 and fitness. Standardized path coefficients are shown for each significant path where * indicates
647 a p-value < 0.05, ** indicates p < 0.01, and *** indicates p < 0.001. See Table S6 for a full listing
648 of direct, indirect, and total effect sizes of functional traits on fitness.
649
650

Table 1. The adaptive potentials of floral traits are limited by trait-trait covariances. For each trait, the mean and standard error for Δz_i and Δz_{nci}, with those displaying significant evolutionary constraint marked in bold. Significance is defined as the percent of bootstrap samples where Δz_i - Δz_{nci} is either above or below zero, depending on the expected difference calculated from the actual data. Δz is calculated with standardized traits such that variance(maternal line) = 1. Significant broad sense heritabilities (H^2) for each trait are also shown in bold. Estimated marginal mean trait values are shown for corolla width, length and ASD in cm, nectar sucrose in °Bx, flowering phenology as Julian Date, plant size as leaf count, and relative fitness as seed count. Ancestral and descendent populations are noted as A or D, respectively. Bolded values represent traits with a significant change in trait value between ancestral and descendant populations (Table S1).

	Year	Δz_i	SE	Δz_{nci}	SE	H^2	Trait Mean
Corolla Width (mm)	A	**0.02**	**1.86E-05**	**0.13**	**2.23E-05**	**0.11**	**48.0**
	D	**0.01**	**5.37E-06**	**0.09**	**1.54E-05**	**0.11**	**48.7**
Corolla Length (mm)	A	**0.01**	**1.70E-05**	**-0.07**	**1.36E-05**	**0.09**	53.4
	D	**0.01**	**4.50E-06**	**-0.07**	**1.12E-05**	**0.05**	53.4
ASD (mm)	A	-0.07	2.11E-05	-0.002	3.16E-07	**0.08**	0.37
	D	-0.02	7.14E-06	-0.01	2.13E-06	**0.09**	0.55
Nectar Sucrose (°Bx)	A	-0.006	1.88E-05	-0.04	8.77E-06	**0.12**	3
	D	0.004	4.76E-06	-0.005	1.07E-06	**0.11**	3
Flowering Phenology (Julian Date)	A	-0.14	1.91E-05	-0.07	8.33E-06	**0.35**	**218**
	D	**-0.03**	**7.11E-06**	**-0.05**	**4.02E-06**	**0.40**	**217.3**
Plant Size (Leaf Count)	A	0.19	3.32E-05	0.16	3.01E-05	**0.09**	**37.2**
	D	0.01	6.05E-06	-0.001	1.24E-07	**0.04**	**33**