Challenges in Replay Detection by TDLM in Post-Encoding Resting State

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Abstract:

Using temporally delayed linear modelling (TDLM) and magnetoencephalography (MEG), we investigated whether items associated with an underlying graph structure are replayed during a post-learning resting state. In these same data, we previously provided evidence for replay during on-line (non-rest) memory retrieval. Despite successful decoding of brain activity during a localizer task, and contrary to predictions, we found no evidence for replay during a post-learning resting state. To better understand this, we performed a hybrid simulation analysis in which we inserted synthetic replay events into a control resting state recorded prior to the actual experiment. This simulation revealed that replay detection using our current pipeline requires an extremely high replay density to reach significance (>1 replay sequence per second, with "replay" defined as a sequence of reactivations within a certain time lag). Furthermore, when scaling the number of replay events with a behavioural measure, we were unable to induce a strong correlation between sequenceness and this measure. We infer that even if replay was present at plausible rates in our resting state dataset, we would lack statistical power to detect it with TDLM. Finally, contrasting our novel hybrid simulation to existing purely synthetic simulations indicated that the latter approaches overestimate the sensitivity of TDLM. We discuss approaches that might optimize the analytic methodology, including identifying boundary conditions under which TDLM can be expected to detect replay. We conclude that solving these methodological constraints will be crucial for optimizing the non-invasive measurement of human replay using MEG.

Introduction

Reactivation of memory traces during offline states is considered a core mechanism supporting memory consolidation. Such memory reactivation has been investigated primarily in animal models, particularly in the context of hippocampal place cells (O'Keefe, 1976), wherein firing patterns that encode movements through a maze re-emerge later during rest periods (Pavlides & Winson, 1989), with their occurrence correlating with post-rest performance. Subsequently, memory reactivation has been linked to consolidation, i.e., the stabilization of long-term memory traces (Born & Wilhelm, 2012; Feld & Born, 2017, 2020; Klinzing et al., 2019, Diekelmann & Born, 2010; Findlay et al., 2020; Foster, 2017; Wilson & McNaughton, 1994). Here, "memory reactivation" refers to the re-emergence of activation patterns absent an eliciting sensory stimulus (Genzel et al., 2020). "Memory replay", on the other hand, emphasizes a sequential ordering of reactivations, where replay patterns follow a forward or backward trajectory (Diba & Buzsáki, 2007; Foster & Wilson, 2006). For example, backward replay is linked to reward processing during breaks in task performance (Ambrose et al., 2016). An additional feature of replay is its time-compression, such that its time-course is much faster than the temporal evolution of the original encoded activity (Nádasdy et al., 1999).

Replay has primarily been demonstrated using cellular recordings within a large variety of mammalian model organisms (Hoffman & McNaughton, 2002; Lee & Wilson, 2002; Pavlides & Winson, 1989). Replay studies in humans using intracranial microelectrode recordings are few, but include work demonstrating compressed replay of firing-pattern sequences in motor cortex during rest and sleep (Eichenlaub et al., 2020; Rubin et al., 2022) as well as single-unit replay of trial-specific cortical spiking sequences during episodic retrieval (Vaz et al., 2020). By contrast, studies using macroelectrodes mainly report stimulus-specific reinstatement or ripple-locked activity changes, as a correlate of replay (J. Zhang et al., 2025), without explicit demonstration of temporal sequential cellular replay (e.g., Axmacher et al., 2008; Staresina et al., 2015; H. Zhang et al., 2018). As these methods can only be applied under restricted clinical circumstances, such as during pre-operative neurosurgical assessments, this limits opportunities to investigate human replay. Therefore, this gives urgency to efforts aimed at developing novel methods to investigate human replay non-invasively.

Non-invasive neuroimaging techniques suffer a trade-off (Hall et al., 2014) – either their spatial precision is high but their temporal precision is lacking, as with functional magnetic resonance imaging (fMRI), or vice versa as with magneto- or electroencephalography (MEG/EEG). This has motivated the development of sophisticated analytical approaches to overcome these limitations. While measuring faster sequences is possible with fMRI (Schuck & Niv, 2019; Wittkuhn et al., 2024;

Wittkuhn & Schuck, 2021), most recent human replay research has focused on MEG- or EEG-based

analysis (Y. Liu et al., 2022).

Using advances in brain decoding that also exploit the temporal precision of MEG, an

emerging approach to quantifying replay involves examining lag differences between individually

decoded reactivation events (Kurth-Nelson et al., 2016; Y. Liu, Dolan, et al., 2021), a method referred

to as "Temporally delayed linear modelling" (TDLM). In essence, by computing a "sequenceness"

score, the method quantifies whether a predicted sequence of decoded brain activity appears more

often than expected by chance. Using this approach, a series of studies have provided evidence that

individual states are replayed in sequence not only during rest (Eldar et al., 2020; Q. Huang et al.,

2024; Y. Liu et al., 2019; Nour et al., 2021, 2023; Yu et al., 2023) but also in the context of active

cognition including planning/decision making (Eldar et al., 2020; W. Huang et al., 2020; Kurth-Nelson

et al., 2016; McFadyen et al., 2023; Schwartenbeck et al., 2023; Seow et al., 2024; Wimmer et al.,

2023; Wise et al., 2021), memory retrieval (Huang & Luo, 2023; Kern et al., 2024; Wimmer et al.,

2020) and following receipt of reward (Liu, Mattar, et al., 2021).

Relevant to the current study, off-line replay of task-relevant sequences has been reported

across several studies. Liu et al. (2019) reported on participants who first learned a rule-based

mapping of visually presented items, followed by a 5-minute recording of resting state activity, prior

to completion of a subsequent inference task. In one study iteration, value learning was applied to

one of two sequences of sensory items. During the resting session, states were replayed with an

approximate 40-60 milliseconds state to state lag, in a predominantly forward or backward direction,

depending on whether a specific sequence was rewarded or not. Eldar et al. (2020) found

sequenceness during resting state were associated with lessened flexibility to subsequent changes in

value-location pairings, consistent with replay supporting a form of model-free decision-making. In

the latter, backward replay was reported during rest at a 180-millisecond time lag, albeit using a

cross-correlation approach, a predecessor method to TDLM. Nour et al. (2021) found attenuated rest

period replay at a time lag of 40-50 milliseconds in patients with schizophrenia compared to healthy

controls. In the same data, Nour et al. (2023) found that patients showed reduced coupling of replay

to default-mode-network-activation, an impairment that correlated with memory retention. Yu et al.

(2023) used EEG recordings in a sequence-learning task and found that participants with high trait

anxiety showed reduced reverse replay of reward sequences during a resting state. Finally, Huang et

al. (2024), using simultaneous EEG-fMRI to examine the spatial and temporal dynamics of memory

replay in humans, showed that during mental simulation replay events were associated with

activation in the hippocampus and medial prefrontal cortex.

In the current study, our focus was on the consolidation of graph-based task structures during resting

state as captured in MEG data. Previous research indicates that recently acquired mnemonic

memories are replayed in subsequent rest or sleep sessions (Humiston et al., 2019; Schapiro et al.,

2018; Schreiner et al., 2021; Tambini et al., 2010; Wamsley, 2022), with replay density decreasing

over a time window of two hours (Eschenko et al., 2008a). Additionally, there is evidence that replay

at rest, and during sleep, preferentially consolidates items based on properties such as encoded

strength (Drosopoulos et al., 2007; Huelin Gorriz et al., 2023; Schapiro et al., 2018) or graph node

properties (Feld et al., 2022). Based on this, in our study we employed criterion learning to 80%

correct retrieval on a graph network, a criterion similar to the performance participants achieved on

a similar graph-based task which showed a behavioural sleep benefit (Feld et al., 2022). We assumed

this would achieve a level of memory acquisition that engenders subsequent sequence replay, as it

allowed for a further memory improvement, which we conjectured, depends on replay. Such

criterion learning is standard practice in studies on sleep-dependent memory consolidation (e.g. Feld

et al., 2013; Galer et al., 2015; Payne et al., 2012; Rasch et al., 2006).

Here we acknowledge that previous studies using TDLM focused primarily on structural inference

and decision making and not on mnemonic learning (Eldar et al., 2020; Y. Liu et al., 2019; Y. Liu,

Mattar, et al., 2021; Nour et al., 2021, 2023). Nevertheless, we considered it likely that there would

also be replay after mnemonic learning (Buhry et al., 2011; Findlay et al., 2020; Zhou et al., 2024)

and, if present, this would be amenable to quantification using TDLM. Despite this, we did not find

replay evidence in our post-learning resting state session leading us to conduct a second study

wherein we simulated replay by inserting neural events taken from the localizer data, at varying

densities, into a control resting state recorded before any stimulus exposure. This simulation allowed

us to quantify a lower bound of replay density necessary to detect replay. Additionally, to simulate a

previously demonstrated association of replay and performance (e.g. Schapiro et al., 2018; Wimmer

et al., 2020), we scaled simulated densities by participant's performance. Finally, we discuss existing

uncertainties regarding parameter choices and possible resulting confounders when using TDLM as

well as suggesting ways to improve the robustness of the method.

Results Study I: Resting State Analysis

In Study I, participants were first placed in an MEG scanner for acquisition of data during an 8-minute resting state (see Figure 1 for an overview of the procedure). Next, while subjects remained in the scanner, we performed a localizer task on which a classifier was trained. Participants then learned triplets of associated items according to a graph structure. Within this learning session, participants performed a maximum of six learning blocks, but the session was stopped if participants reached 80% memory performance (criterion learning, see Methods for details). Finally, participants were asked to remain in the scanner for a further 8-minute resting state measurement after which we tested memory retrieval. Mean learning performance of ~82% indicated that the applied criterion learning was successful.

In the post-task resting state, we found no evidence for sequenceness reflective of task structure. Behavioural results indicated that the experimental manipulation worked as intended, i.e., participants learned the graph to criterion, without overlearning which might have precluded the occurrence of replay (see Figure 2). Additionally, we tested for a correlation of memory performance and sequenceness strength, as group-level sequenceness thresholds might be too conservative alone, but found none. Note that data from the retrieval stage were published elsewhere (Kern et al., 2024). The study and analysis were preregistered at aspredicted.org (https://aspredicted.org/kx9xh.pdf, #68915).

Figure 1. Experimental procedure in the MEG. Control: Eight minutes of resting state activity were recorded before participants encountered any stimuli. Localizer task: Next, the ten individual items were repeatedly presented to participants auditorily and visually to extract multisensory activity patterns. Learning: Participants learned triplets of the ten items pseudo-randomly generated based on an underlying graph, by performing a three-alternatives-forced-choice. Participants were unaware of the graph layout. Consolidation: Eight minutes of resting state activity were recorded. Retrieval:

Participants' recall was tested by cueing triplets from the graph. The letters in the pictograms are placeholders for individual images. Figure and caption adapted from Kern et al., (2024). For details, see Methods section Figure 10. For an animated overview of the task and the experiment, see also our GitHub pages of the study: https://cimh-clinical-psychology.github.io/DeSMRRest-TDLM-Simulation/

Behavioural results

All thirty participants, except one, learned the graph consisting of images with an accuracy above

50% (preregistered as an exclusion criterion for "successful learning above chance"). After a

consolidation resting state, a marginal increase in performance was observed (76% to 82% accuracy,

t=-2.053, p=0.053, effect size r=0.22; Figure 2A), but note that this may not necessarily be attributed

to consolidation, as the last block also included further feedback. More details on behavioural results

are found in Kern et al. (2024).

*Figure 2. Behavioural results and localizer decoding accuracy. **A)** Memory performance of participants after completing the first block of learning, the last block of learning (block 2 to 6, depending on speed of learning), and retrieval performance after the resting state. Participants completed up to six blocks of learning trials. After reaching 80% in any block, no more learning blocks were performed (criterion learning). **B)** Decoding accuracy of the currently displayed item during the localizer task for participants with a decoding accuracy above 30% (n=21, 9 participants' data did not pass this criterion). The mean peak time point across all participants corresponded to 210 ms, with an average decoding peak accuracy of 42% (n=21). Note that the displayed graph combines accuracies across participants, where peak values were computed on an individual level and then averaged. Therefore, at a group level, the individual mean peak does not necessarily correspond to the average peak. **C)** Classifier transfer within the localizer when trained and tested at different time points as determined by cross validation, after ~200 ms, most training and testing time points generalize to each other (significant clusters indicated by cluster permutation testing, alpha<0.05).*

Decoder training

We trained classifiers on data from the localizer task, in which participants were presented with

images from the learning task in a pseudorandom order. In each trial, a word describing the stimulus

was played auditorily, after which the corresponding stimulus was shown. In ~11% of cases, there

was a mismatch between word and image (oddball trials), and these trials were excluded from the

localizer training. Decoding from brain activity was at a level comparable to previous reports as

determined in cross-validation (~42% across all participants, range 32–57%, chance level: 10%,

excluding participants with peak accuracy <30%). We used the average peak decoding accuracy time

point of 210 milliseconds across all participants (Figure 2B) to train final participant-specific decoders

to apply on resting state data.

No sequential replay during resting state

By applying trained decoders to the resting state session, we used probability estimates of the ten

classes as input to TDLM. In brief, TDLM quantifies time-lagged reactivation ("sequenceness") of

states according to expected transitions, e.g., quantify how strongly state A->B was reactivated with

a temporal distance of N milliseconds. As there was no specific sub-hypothesis about which

transitions would be replayed, we tested for all allowable graph transitions. We found no evidence

for sequential replay measurable during the post-learning resting state. More specifically, the GLM

did not exceed significance thresholds (conservative, i.e., the maximum sequenceness across all

permutations and timepoints or liberal criteria, i.e., the 95% percentile of aforementioned

sequenceness) for any individual time lag, neither in forward nor backward direction (Figure 3A).

Additionally, we calculated the difference between the pre-resting state and post-resting state

sequenceness values, to remove potential systemic biases on sequenceness fluctuation present in

both sessions. We calculated an FDR corrected paired t-test for all 30 time lags individually, for

forward and backward sequenceness between pre- and post-resting state respectively. All time lags

were non-significant (p_{FDR}>0.05; without correction, three time points were significant: forward 100

ms, t=-2.43, p=0.024 and 240 ms, t=2.13, p=0.045, backward 200 ms time lag, t=-2.11, p=0.047).

However, as noted in Y. Liu, Dolan, et al., (2021), correct statistical analysis of sequenceness scores is

non-trivial and FDR-correction might be too conservative. The permutation test shuffles state labels

and measures a fixed effect, making it susceptible to outliers on a group level, which we addressed

by z-scoring the sequenceness results. We also included a newly developed method to take random-

effects at a group level into account, specifically employing a sign-flip permutation test that randomly

inverts a subset of participants' sequences 10000 times and calculates a one-sided t-test against null

for each time lag per permutation. The maximum t-value per permutation is used to construct an

empirical null distribution. This approach circumvents the multiple comparison problem inherent in

the previous method (see Methods section for details). Neither the control nor the post-learning

resting state surpassed sign-flip significance thresholds (Figure 3C control: forward p=0.065, t=2.63

backward p=0.641, t=1.61, post-learning: forward p=0.082, t=2.62, backward p=0.566, t=1.731).

Additionally, as time lags are dependent on each other, we also calculated a one-sample-cluster-

permutation-test of difference values with 10000 permutations that considers the temporal

correlation of time lag bins. This revealed a single forward non-significant cluster at 130-140 ms

(p=0.039) ms (p=0.44). For a visualization of the reported differences see Figure 3B. Finally, we report

forward minus backward sequenceness as well as our motivation for an across-session post-pre

comparison, instead of within-session forward-backward in Supplement Figure 10.

Additionally, we ascertained whether there was an interaction between memory retrieval

performance and individual sequenceness strength by calculating the time lag of peak sequenceness

across all participants (control forward = 100 ms, control backward: 10 milliseconds, post-learning

forward: 280 ms, backward: 240 ms) and correlating individual sequenceness values with

participants' performance. There was no significant correlation between peak sequenceness scores

and memory performance at post-learning resting state (forward r=0.0, p=0.99, backward: r=0.02,

p=0.935, see Figure 3D), nor with the performance difference between post-rest retrieval minus the

performance of the last block before rest (see Supplement Figure 9)

Lastly, similar to Liu et al. (2019), we split the 8-minute resting state into discrete time bins

and assessed replay strength for each segment separately (Figure 4). In none of these bins was a

maximum permutation threshold exceeded. A more lenient 95% permutation threshold was

exceeded in one segment of the control resting state for minute 4 (100 ms backward sequenceness).

However, here the sequenceness value was negative suggesting, if anything, suppression of replay.

Individual sequenceness curves for all segments can be found in Supplement Figure 4.

Figure 3. Sequenceness result of the resting state sessions. **A)** *Forward and backward sequenceness values on the y-axis at different time lags (x-axis) for control resting state (light coloured) and post-learning resting state (dark coloured). Two significance thresholds are indicated, the more lenient 95% quantile of the maximum sequenceness of all permutations (lightly dotted line) and the maximum mean peak sequenceness across all permutations of all subjects (bold dotted line). None of the thresholds were surpassed. The peak mean sequenceness has been marked by a vertical line respectively.* **B)** *Difference between pre-learning and post-learning resting states across time lags. Using FDR-corrected paired t-tests for each time lag revealed no significant difference between conditions (without correction two sequenceness differences were significant: forward 140 ms, p=0.039 and 220 ms, p=0.035). An additionally performed cluster-permutation test found a non-significant cluster at 130-140 ms forward direction (p=0.43). Bands indicate two standard errors of the mean of the difference scores.* **C)** *T-value distribution of a 10000 iterations-sign-flip permutation test. Per permutation, a random subset of subjects' sequenceness values were flipped and the maximum t-value across lags used to construct the empirical distribution. The true maximum t-value indicated in red (nonsignificant in all cases)* **D)** *Individual participant's retrieval memory performance relative to their sequenceness values at the peak mean sequenceness (indicated in A). Regression lines and 95% confidence intervals are fitted. For correlation with performance deltas between post and pre resting state retrievals see Supplement Figure 9. NB, however, as our simulation shows, correlations of sequenceness with behavioural markers are likely to be underpowered and occur only with very high replay rates or much higher sample size. See our simulation discussion for a more detailed explanation on how correlations may be inherently biased, where fluctuations in baseline sequenceness overshadow individual scaling with behavioural markers. For individual sequenceness curves, see Supplement Figure 3. For within-session forward minus backward sequenceness, see Supplement Figure 10.*

Figure 4. Sequenceness map of different segments. Each of the eight minutes of resting state are on the y-axis, and different time lags on the x-axis. For each minute, we took either the first or second half of 30 seconds for the final calculation and the other 30-second segment for the exploration data set. Sequenceness crossed the 95% permutation threshold in the control resting state in minute 4 (forward 100 millisecond time lag) alone. However, the sequenceness score was negative, rendering its interpretation non-trivial. Note that we did not apply any control for multiple comparisons across the segments. For sequenceness curve per segment, see Supplement Figure 4.

Interim discussion

Overall, in resting state data, recorded directly after a learning task block, we did not find evidence for sequenceness using TDLM. This involved implementing previously described protocols in terms of localizer recording, pre-processing and decoder settings, albeit within a different cognitive protocol compared to those previously implemented within resting states (Y. Liu et al., 2019; Nour et al., 2021, 2023; Wimmer et al., 2020). Even though we found a single time lag with putative replay during one of the thirty-second segments, the evidence here was weak and its validity is questionable. Firstly, multiple comparison correction would have removed the effect. Second, interpretation of negative sequenceness values is unclear. As forward and backward sequenceness are estimated separately, in our interpretation, negative forward sequenceness suggests an expected sequenceness is replayed *less* than expected, which could be interpreted as replay suppression. One possible way this could occur is via a repetition suppression process (Schwartenbeck et al., 2023), if one assumes that

activation of one item occurs simultaneously with the next in line and subsequently that pattern of

neuronal activity is suppressed. However, this appears to us, at best, a convoluted argument.

Additionally, recent reports suggest replay avoids self-repetition in addition to avoidance of replay

for just experienced episodes (Mallory et al., 2025), which would show as negative sequenceness.

However, this effect has so far been shown to occur only at a time-scale of seconds and not minutes.

Previous reports involving TDLM analysis of resting state data has involved paradigms with a

dominant inference component (Y. Liu et al., 2019; Y. Liu, Mattar, et al., 2021; Nour et al., 2021).

Notably, in our study, we used a declarative graph-based memory paradigm that lacked this

inferential component. We note previous research in humans has shown reactivation and replay

during rest in the context of non-inferential paradigms using fMRI (Ogawa et al., 2024; Schapiro et

al., 2018; Tambini & Davachi, 2013), and especially in the context of declarative learning (using

fMRI&EEG: Deuker et al., 2013; using intracortical electrodes: Eichenlaub et al., 2020; using fMRI:

Schlichting & Preston, 2014; using intracranial EEG: Zhang et al., 2018). On the other hand, it has

been reported that replay is more likely to be expressed for weakly encoded memories (Denis et al.,

2020; Schapiro et al., 2018) and hence it might be absent if encoding strength is too strong. While we

employed a learning criterion of 80% for this reason, as is usual in the broader context of sleep and

memory research, we cannot be certain if this represents a level of encoding strength that lessens a

requirement for task-related replay in our study. Anecdotal evidence for replay occurrence might

come from many participants reporting involuntarily seeing segments of the sequence flashing in

front of their eyes during the resting state, though we caution that there is mixed evidence for a link

between anecdotal reports of conscious experience and the occurrence of replay (Komsky, 2015;

Wittkuhn et al., 2024).

Additionally, given that the stimuli were presented in combined triplets, participants may have

formed a singular representation of associated items and subsequently replayed these (e.g., AB→C),

instead of replaying item-by-item transitions (A→B→C). Under such a scenario, a classifier trained on

individual items would not detect these newly formed representations, particularly if they diverge

strongly from single-item patterns. In our previous study, where we address retrieval (Kern et al.,

2024), we found that states were to varying extent co-reactivated, yet classifiers trained on single

items retained sensitivity to detect these combined reactivation events. Consistently, prior work

suggests that unified representations retain overlap with their constituent item representations

(Dennis et al., 2024; Liang et al., 2020). However, there is also evidence that different brain regions

are involved if unitization occurs (Staresina & Davachi, 2010), potentially confusing classifiers.

Therefore, we cannot exclude that rest-related consolidation replays engendered unitized

representations that were insufficiently captured by our single-item classifiers.

As we did not know whether absence of replay was inherent to our paradigm, in a second study we
explored potential explanations for this null finding. Specifically, we conducted a hybrid simulation
study that combined localizer data and data from a control resting state. In the context of our
paradigm, this simulation allows quantification of the strength of replay necessary for it to be
detectable using the TDLM methodology.

274 Results Study II: Simulation

In a hybrid simulation study, we enriched the control resting state data of participants with simulated
replay events. That means, we inserted sequences of neural patterns extracted from the localizer
(taken from the time point of peak decoding accuracy) into the control resting state with a state-to-
state lag of 80 milliseconds. We chose this timepoint (80 millisecond state-to-state lag) as its
sequenceness value was close to zero in the baseline condition as well as it being distant to notable
alpha rhythms of participants (which varied between ~9-11 Hz). Additionally, we subtracted the
mean sequenceness value of the baseline at 80 millisecond lag such that any simulation effects
would, on average, start at zero sequenceness.

We varied the density of replay, traditionally denoted as "replay events per minute" (min^{-1}) and
looked at the resulting sequenceness strengths. Additionally, we scaled the replay density inversely
with the participants' retrieval memory performance to simulate a putative increase of replay for less
stable memories. For example, a participant with lowest memory performance (50%) would have a
factor 1.0 applied to their replay amount, while a participant with highest memory performance
(100%) would show reduced replay and have a scaling factor of 0.5 applied to the number of inserted
replays. We used different scaling factors to explore how they contribute to relationships between
measures of sequenceness and memory performance. While we acknowledge the relation between
replay and memory performance is probably more complex (Kern et al., 2024; Schapiro et al., 2018;
Wimmer et al., 2020), we applied this simplified scaling to assess whether TDLM can detect
behavioural correlations with sequenceness strength, as has shown previously (Kern et al., 2024; Y.
Liu et al., 2019; Nour et al., 2021; Wimmer et al., 2020; Yu et al., 2023). Our results indicate that
replay and correlations with behaviour were only detectable at very high - putatively biologically
implausible - densities.

Figure 5. Schematic of procedure for inserting simulated replay into the control resting state. First, neural patterns are extracted, for each stimulus, from the peak of decodability during the localizer task. Participant-specific patterns are normalized by subtracting the average sensor values of seeing any other stimulus from the sensor values of seeing only the stimulus of interest. Next, these subtle patterns are inserted into the control resting state at a fixed interval, following a one-step transition pattern according to the task sequence. A refractory period is retained between each replay event to prevent overlaps. For details see Methods section.

Sequenceness only detectable at very high replay density

Our simulation analysis found that a replay density of 130 min^{-1} and a density of ~80 min^{-1} were

necessary in order to reach the maximum and 95% quantile permutation threshold, respectively (see

Figure 6A). We also performed the above-mentioned sign-flip permutation test on the simulated

data, reaching comparable results, with significance reached at around 80 min^{-1} (see Figure 6B).

Figure 6C shows sequenceness curves for selected densities at 0 min^{-1} (baseline), 40 min^{-1}, 80 min^{-1}

and 120 min^{-1}. At 40 min^{-1} only a small increase of sequenceness above baseline is detectable. Note

the significant drop of sequenceness at 100 milliseconds, which is due to the alpha control

introduced at 10 Hz by inserting repeating patterns of the decoded items (Y. Liu, Dolan, et al., 2021).

The alpha correction likely over-corrected and introduced negative sequenceness at 10 Hz seen in

some participants. Therefore, we include the analysis without alpha-correction in the Supplement

Figure 5. NB the strong negative sequenceness at 100 millisecond time lag was not present in our

exploratory data set and only appeared in data held back for the final calculation (as preregistered,

we used 50% of the resting state data for exploration and parameter search and report here on the

unseen 50% after parameters were chosen, see Methods section for details).

Figure 6. Sequenceness results of the hybrid simulation. **A)** *Simulation results for inserting replay at different densities. On the y-axis, sequenceness at 80-millisecond state-to-state time lag, at different replay densities on the x-axis with standard error as shaded orange. The 95% confidence interval significance threshold is reached at a density of ~80 min⁻¹, while the maximum permutation threshold is reached only around 130 min⁻¹.* **B)** *Observed t-values and significance thresholds of a sign-flip permutation test, in which the maximum t-value across time lags for sign-flip permutated sequenceness curves are used to construct a non-parametric null distribution (see Methods for details). Significance (p<0.05) is reached similarly at around 80 min⁻¹.* **C)** *Sequenceness across time lags for replay at baseline (0 min⁻¹) and simulated for a density of 40, 80 and 120 min⁻¹ with a simulated time lag of 80 milliseconds. In this simulation, we assumed a temporally uniform distribution of replay throughout the resting state, as TDLM is relatively invariant to temporal distribution (see Appendix A.1 and Appendix A Figure 1)*

316 ## Correlation with behaviour only present at high replay speeds

The correlation of behaviour with sequenceness at 80 milliseconds time lag was non-significant, even at the maximum replay densities of 200 min⁻¹ (see Figure 7A). Here we assumed a realistic scenario that linearly downscales replay density with increasing performance from 100% to down to 50% for increased memory performance. To amplify the correlation effect, in a secondary analysis, we

normalized the performance scaling between 100-0% (i.e. 100% performance would lead to 0% of replay inserted). In this maximum-scaling scenario, the correlation of sequenceness with performance was markedly boosted, yet did not reach significance even at a density of 200 events min^{-1}. When further investigating this lack of influence of our scaling, we found that even in the baseline condition without simulated replay (0 min^{-1}), the correlation with performance seems to oscillate, depending on which time lag it is computed on (Figure 7B), indicating fluctuations of sequenceness between participants that are non-random. If we chose a different time lag for our simulation, added effects of baseline plus simulation would have probably reached significance more easily (e.g. at the 240-millisecond time lag). This baseline variation becomes apparent when comparing sequenceness of participants at baseline with sequenceness values boosted via simulated replay at 80 millisecond time lag. Even in the baseline condition (0 events min^{-1}, Figure 7D), some participants express higher sequenceness values than others with simulated replay of 160 min^{-1}. This indicates that the effect that we wanted to introduce via scaling the inserted events with memory performance is overshadowed by some form of bias present at baseline, that induces marked variations in sequenceness without any replay present, even after a recommended GLM-based alpha control (Y. Liu, Dolan, et al., 2021).

The resulting correlation is of moderate size even at higher densities (Pearson's r of around 0.3), indicating our paradigm was likely underpowered for this specific analysis. Therefore, we performed a bootstrapped power analysis by oversampling existing participants, with 10,000 repetitions per sample size. We found that more than 160 participants would have been necessary to detect a correlation between performance and sequenceness in the simulated density of 100 min^{-1} condition (Figure 7C) with 80% power where the above-mentioned maximum scaling applied. Nevertheless, it is worth bearing in mind that this might also lead to increase power to detect nuisance effects in the baseline condition, if the effect is added to a pre-existing correlation in the baseline condition.

Figure 7. Correlation with behaviour is underpowered and strongly affected by systematic baseline fluctuations. **A)** *Relationship (Pearson's r) between memory performance (behaviour) and sequenceness at 80 ms time lag on the y axis, with replay density on the x-axis for two scaling rules: Realistic scaling: Replay density was inversely scaled with participant's performance to between 50 and 100% (i.e., the worst performer had 100% while the best performer only had 50% of replay density). Maximal scaling: Replay density was scaled in from 0-100% (i.e., the best performer had no replay inserted). Significance of the correlation was not reached even at the maximum density. However, note, that the baseline correlation started at r=0.1, and simulated effect without this offset would have reached significance at a density of around 200 min⁻¹. **B)** Correlation of sequenceness computed across different time lags in the control condition (i.e., without inserted replay). Even when no replay is present, the correlation fluctuates significantly depending on which time lag it has been computed, possibly due to baseline variations that fluctuate differentially between participants. **C)** Bootstrap power analysis showing how many participants would be necessary to detect a significant correlation between replay and performance in the maximally scaled scenario with a replay density of 130 min⁻¹. 80% power to detect a correlation is reached around a sample size of around 160. **D)** Sequenceness at 80 milliseconds time lag in the control condition (left), in the realistic scenario with replay scaling from 50-100% (middle) and the maximal scenario with scaling from 0-100% (right). Participant's performance is indicated by increasing saturation of blue. While all participants show an increase in sequenceness with increased scaling, some participants have a higher sequenceness value and an ordering according to performance is barely visible in the maximum scaling case.*

Previous Synthetic Simulation Overestimates TDLM Sensitivity

Next, we compared our hybrid simulation to a purely synthetic TDLM simulation, closely following the demonstration code in Y. Liu, Dolan, et al., (2021). We show that in the absence of endogenous oscillatory patterns, and using synthetic patterns, significance thresholds are easily surpassed at around 10-20 min^{-1} (Figure 8A). In this setting, synthetic resting state data is generated autoregressively by sampling values from a multivariate normal distribution given a sensor covariance matrix. The class-specific stimulus patterns are created by first drawing a random sensor pattern common to all classes (akin to an ERP) and adding a random class-specific pattern for each class, plus noise for each trial of that class. Null data is created by drawing from a normal distribution. For the full procedure see Methods Study II. Note, for consistency, we used the same insertion logic and code as for the hybrid simulation, which deviates slightly from the MATLAB implementation and simulates with an unrealistic 2000 min^{-1} (see Supplement Figure 6).

The maximum permutation threshold is reached already at a density of 19 min^{-1} (Figure 8A), similarly to the original MATLAB code (see Supplement Figure 6, note the slightly increased threshold is explained by our scaling with performance, which puts an effective density mean slightly lower than 19 min^{-1}). The 95% of permutation sequenceness maxima is reached at 12 min^{-1}. Similarly, the sign-flip permutation threshold is reached at 10 min^{-1}, showing again an increased sensitivity of this new permutation approach. Individual sequenceness plots (Figure 8B) show that the purely synthetic setup produces robust sequenceness effects at substantially lower replay densities than evident in our hybrid simulation, surpassing thresholds manifold even for lower sequenceness. However, importantly, when examining the baseline condition, a notable lack of oscillatory behaviour is observed compared to empirical data (Figure 8B upper left).

To quantify why sequenceness emerged more readily in the synthetic case, we compared decoder output distributions at timepoints where a true pattern is present. Therefore, we extracted classifier probabilities given to the true class and all other classes. Figure 9A shows histograms of log-scaled classifier probability estimates for the peak-decoding time point in the localizer (left, cross-validated) and during a simulated reactivation events in both the hybrid simulation (middle) and synthetic simulation. Classifier probabilities of the target class in the synthetic simulation were substantially more extreme, indicating stronger separation between target and non-target patterns than in the empirical localizer. Next, we quantified the difference between the target and other probability estimate distributions using the Wasserstein distance (Villani, 2009) per participant (Figure 9B), as a higher distance should result in better discriminability. Pattern discriminability was lowest in the localizer trial (W_1=0.11), intermediate in the hybrid simulation (W_1=0.14) (consistent with the fact that inserted patterns were derived from empirical data) and markedly increased for

352   the synthetic simulation (W_1=0.38). Notably, both simulations overestimate pattern separability, as

353   we hypothesize that the pattern discriminability during a real reactivation event is less pronounced

354   than in the localizer during the strong signal induced by visual image processing.

*Figure 8. Results of the fully synthetic simulation, akin to the one presented in Liu et al. (2021). **A)** Mean sequenceness at the simulated time lag of 80 milliseconds. Thresholds are reached already at a much lower density of a rate of ~19 events/minute, in line with the original MATLAB code (Supplement Figure 6). Note that due to scaling with performance, real average events per minute are slightly below the indicated density, explaining why the thresholds are reached at a slightly higher density (i.e. ~19 vs ~14). **B)** Four exemplar plots for sequenceness at a density of 0 (baseline), 40, 80 and 120 events per minute. Oscillatory behaviour of the sequenceness curves is very minimal in the baseline condition. Maximum permutation thresholds are reached easily in the 40 events per minute case and surpassed several times with higher densities. However, exactly the lack of these spurious confounders and a too strong pattern discriminability make it possible for TDLM to work well in a*

synthetic dataset. See Supplement Figure 6 to compare result to the original MATLAB simulation code.

Next, we tested what would happen if we increased the strength of the inserted pattern.

Therefore, we scaled the synthetic pattern by a factor of 0 to 2 before insertion and measured the

Wasserstein distance of resulting decoder output. Unsurprisingly, the stronger the magnitude of the

inserted pattern, the better the discriminability between the target and other class (Figure 9C). This

allowed us a rough estimate of the pattern strength during the localizer and hybrid simulation by

mapping their Wasserstein distances to the strength in the synthetic scaled simulation. Next, we

compared how increased pattern discriminability would affect sequenceness (Figure 9D). We

simulated synthetic replay as above at a density of 80 min^{-1}. As expected, the higher the

discriminability between the target class and other classes, the higher the sequenceness. Thus,

greater discriminability will lead to more extreme classifier outputs, increasing the beta weights of

the first-level GLM of TDLM.

Overall, our results indicate that the purely synthetic simulation produces substantially

higher pattern discriminability than observed in the empirical localizer and the hybrid simulation.

This enhanced discriminability is accompanied by a strong increase in sequenceness magnitude,

suggesting that previous synthetic models may provide an over-optimistic estimate of TDLM's

sensitivity. Additionally, by omitting endogenous oscillatory confounders, synthetic simulations fail to

account for spurious baseline sequenceness that confound empirical detection.

*Figure 9. Pattern discriminability between the target class and other classes impact sequenceness sensitivity. **A)** Histograms of log-scaled classifier probability estimates during the peak decodability of localizer (left, cross-validated), for reactivation events in the hybrid simulation (middle) and synthetic simulation (right). **B)** KDE-plot of the Wasserstein distance of participants' classifier probabilities. The hybrid simulation likely over-estimates how well the decoder finds a reactivation event, while the synthetic simulation largely over-estimates. **C)** Pattern discriminability as a function of the inserted pattern strength (akin signal-to-noise-ratio). We scaled the magnitude of the inserted pattern in the synthetic simulation. Dotted lines indicate estimated pattern strength of the empirical data and hybrid simulation as given by their mean Wasserstein distance **D)** TDLM Sequenceness increases with better pattern discriminability. However, it is questionable if the pattern discriminability during a resting state reactivation event is higher than during the localizer.*

Discussion

Using temporally delayed linear modelling (TDLM), we investigated whether sequential replay of learned graph-structured associations could be detected during a post-learning resting state. Contrary to expectation based on previous findings, we did not find evidence for replay of well-learned sequences during a post-learning rest period. This was despite using a similar methodology to that used in prior replay investigations of resting state, albeit these involved different paradigms (Eldar et al., 2020; Q. Huang et al., 2024; Y. Liu et al., 2019; Nour et al., 2021; Yu et al., 2023). However, due to differences from previously used paradigms, we cannot conclude that lack of evidence for replay reflects the sensitivity of our detection method (TDLM) or reflects that our task was not eliciting sufficiently strong resting state replay to enable replay detection using TDLM.

Based on the above, we designed a simulation study to explore the strength of replay being necessary for successful detection. This simulation study also speaks to the potential sensitivity of

TDLM as an explanation of our null findings. By inserting synthetic replay events into pre-task resting

state data, we show that replay is only be detectable under an assumption of very high replay rates—

with more than 80 replay events per minute being necessary to reach significance thresholds. Here,

one "replay event" is defined as a train of sequential reactivations with a certain state-to-state time

lag, i.e., 80 ms, meaning more than one replay event per second on average throughout the entire

resting state. Furthermore, we find that already at baseline (i.e. without any replay) sequenceness

values are biased by, yet unknown, confounders that introduce absolute offsets in baseline

sequenceness values between participants. We contrast our novel hybrid approach to a purely

synthetic simulation and show that previously proposed approaches are likely to overestimate

pattern discriminability and endogenous neural patterns as oscillations, overestimating TDLM's

sensitivity.

Our hybrid simulation indicates that when an increase in sequenceness by actual replay is

introduced, the baseline offset between participants is larger than the increase at the group level.

Therefore, correlating behavioural measures with individual sequenceness values should be regarded

with caution unless this baseline offset is controlled for. Our findings highlight a potential need for

further methodological refinements in both experimental paradigms and analysis techniques to

reliably detect human replay, using non-invasive measures, during extended offline periods (such as

resting state or sleep). Finally, we make recommendations for future studies and suggest ways to

enhance the reliability of replay detection methods.

High replay density necessary to detect sequenceness

In our study, more than one replay event per second (80 min^{-1}) would have been necessary to detect

replay at a group level. Sharp-wave-ripples (SWR) are traditionally considered substrates for

representational reactivation and hence strong replay expression (Eschenko et al., 2008b; Joo &

Frank, 2018; Roumis & Frank, 2015), with each SWR containing several sequential item reactivations.

While every SWR might not contain task-relevant activity (Schreiner et al., 2024), their occurrence

rate is considered a reasonable proxy for estimating *in vivo* replay density. Previous research in

rodents has reported that ripple rate significantly varies between tasks and states, being most

prominent in offline states such as NREM and post-task rest with a replay density of around 20-25

events per min^{-1} (Eschenko et al., 2008b; Ponomarenko et al., 2003; Wiegand et al., 2016). In

humans, evidence from intracranial recordings is sparse, but studies estimate a ripple density of

between 1-15 min^{-1} during NREM sleep (X. Jiang et al., 2020; Ngo et al., 2020) and ~1.5-4 min^{-1} during

resting state (Y. Y. Chen et al., 2021). However, it should be noted that reported ripple rates depend

on an applied detection criteria and so might underestimate their actual true occurrence (A. A. Liu et

al., 2022). While *in vitro* studies have reported that CA1/CA3 neurons can maintain a (experimentally

induced) density of up 50-200 min^{-1} (Buzsáki, 2015; H. Jiang et al., 2018; Kanak et al., 2013), to our

knowledge, such densities have not been reported *in vivo*.

As replay was detectable at around 80 min^{-1} (~1.33 replay events per second), with more strict

thresholds only being reached at around 130 min^{-1} (~2.5 replay events per second). Based on

previous research it might be considered implausible that these densities would be maintained

throughout an 8-minute resting state. Nevertheless, as it is known that SWRs occur in bursts

(Buzsáki, 2015), and it is conceivable that higher ripple rates are maintained during shorter time

windows, as for example during short intervals of recall, where they can reach more than 40 min^{-1}

(Norman et al., 2019).

Interestingly, Higgins et al. (2021) report that replay events detected by TDLM were not

evenly distributed throughout the resting state and, even though not explicitly reported, varied

between 9-40 min^{-1} for different segments of an exemplar participant. When replay occurs in bursts,

certain time windows might reach necessary density thresholds, while averaging over larger windows

might be inconclusive. However, such high-density windows would need to be selected on a

principled basis to avoid inflating false positive rates. Nevertheless, in our post-hoc analogue sub-

analysis of 30-second individual segments of the resting state we found no clear evidence for replay.

Currently, in the context of MEG replay studies it is not reported what replay rates (events min^{-1}) are

assumed. Therefore, future studies might consider providing estimates for replay rates under which

reported findings are plausible, analogous to SWR rates reported in animal studies, e.g. via

simulations.

However, while we use indices of SWRs as a proxy for replay density estimation, the

relationship between hippocampal replay and replay detected by TDLM remains uncertain. Current

decoding approaches measure replay-like phenomena over cortical sites, with previous reports

highlighting a spectral power increase in hippocampal areas coinciding with TDLM detected replay

episodes (e.g., Y. Liu et al., 2019; Wimmer et al., 2023). Nevertheless, it is conceivable that cortical

replay detected by TDLM might occur independent of hippocampal replay and SWRs, as well as being

generated by different mechanisms. Furthermore, some TDLM-studies find a replay state-to-state

time lag of above 100 ms (Q. Huang & Luo, 2023; Y. Liu, Mattar, et al., 2021; Wimmer et al., 2020,

2023), which is much slower than previously reported place cell replay. One obvious conclusion is

that we know too little about the characteristics of replay as detected by TDLM, specifically with

respect to its relationship to SWRs. Future studies might employ simultaneous intracranial and

cortical surface recordings to throw light on the precise relationship between hippocampal replay

and replay as detected by TDLM

Similarly, in our simulation, for simplicity and to maintain consistency with previous stimulations, we

restricted replay events to those spanning two reactivation events. While the characteristics of

replay as measured by TDLM are unknown, it is conceivable that several event steps can be replayed

within one replay event. We show that the vanilla version of TDLM is fundamentally sensitive to the

number of single-step transitions alone, and disregards if these are either replayed chained or

chunked (Appendix A.2 and Appendix A Figure 2). Nevertheless, if the number of reactivation events

chained within a replay event increases, TDLMs sensitivity is increased relative to the replay density

and thresholds are reached earlier (see Appendix A Figure 4). See Appendix A.4 for a simulation of

multi-step replay events and our discussion of the caveats.

Correlation with behaviour is overshadowed by baseline fluctuation of sequenceness

The occurrence of replay, in the absence of ground truth, can to some extent be validated by a

correlation between a replay measure and behavioural performance (Y. Liu, Mattar, et al., 2021;

Wimmer et al., 2020). In our simulation, we could not induce a strong relationship of sequenceness

to participants' performance at realistic replay densities. Even in an unrealistic, maximal-scaling

scenario, with no replay for the best-performing and 100% replay for the worst-performing

participant, correlations were only significant at implausibly high replay densities of around 200 min^{-1}

^{1}. A power analysis revealed that around 160 participants would be necessary to establish a

correlation in the maximum scaling case at a density of 80 min^{-1}. We find that the magnitude of

baseline sequenceness values across participants was higher than the induced effects and was hence

not affected strongly by inserting synthetic replay (see Figure 7D). Sample sizes in previous TDLM

literature usually range between 20 to 40 participants. A bootstrap power analysis reveals that even

at those sample sizes and beyond, power would remain low unless unrealistically high replay rates

are assumed (Supplement Figure 11). Additionally, the correlation analysis between sequenceness

and behaviour would be greatly underpowered for higher sample sizes, even under an assumption of

high replay densities (see *Figure 7*C).

Spurious sequenceness in the control condition

In data without any expected replay sequenceness should be close to zero, or else be attributed to

noise or bias. While random noise should average out with increasingly large number of participants,

it remains possible that some systematic bias is shared among participants. For example, this might

be the case with alpha oscillations, which occur at roughly the same frequency and tend to increase

in strength from pre to post resting state due to participant fatigue. In our study, non-significant

fluctuations in sequenceness already occurred in the control condition, a phenomenon not unique to

this study (and reported in supplementary material of Nour et al., 2021, 2023 and Liu et al., 2019).

Thus, it seems as if a yet unknown bias introduces spurious systematic sequenceness effects that, in

our simulation, may serve to overshadow actual sequenceness at low densities. While, as previously

discussed (Liu, Dolan, et al., 2021; Liu et al., 2019), endogenous oscillations such as eyes-closed-

alpha can impact TDLM, the spurious (albeit non-significant) sequenceness in our case was not

sufficiently controlled by alpha correction to make it disappear completely (see Supplement Figure 5

for analyses without the 10 Hz correction method). Interestingly, the magnitude of the correlation of

performance with sequenceness, as discussed above, fluctuates approximately at a 10 Hz rhythm.

This leads us to speculate that alpha oscillations may impact baseline sequenceness in a manner not

entirely removed by controlling for alpha using recommended methods (Liu, Dolan, et al., 2021).

Furthermore, we speculate that other factors might contribute to spurious sequence detection, for

example oscillatory inter-trial-coherence during localizer training (Hoffman et al., 2013; Kragel et al.,

2020; van Diepen & Mazaheri, 2018), (phase)-coupling of two or more oscillations (Fell & Axmacher,

2011; Watrous et al., 2015) or phase-differences in travelling waves (Hughes, 1995; Muller et al.,

2018). For example, if two classifiers' probabilities oscillate at 10 Hz but at a different phase, a

spurious time lag can be found reflecting this phase shift. We speculate that more complex

interactions between classifiers oscillating at different phases are also conceivable. Exploring these

possibilities is beyond the scope of the current study and should be evaluated systematically in

future experiments.

Finally, yet importantly, it is important to consider that endogenous oscillations might play a

functional role. Thus, it has been proposed that certain theta phases support encoding -and

consolidation (Hasselmo et al., 2002; Kerrén et al., 2018, 2022), or that oscillations provide a window

for reactivation (Higgins et al., 2021; Staresina et al., 2015). Such a scenario poses a challenge for

differentiating between a theta-phase-bias, not attributable to memory processing, from actual

reactivation present only during a specific phase of an oscillation.

Probing the conditions under which TDLM works requires realistic simulations

Previous studies performed sequenceness analyses on simulated replay, but replay density has rarely

been simulated under realistic conditions. For example, in the sample code of Y. Liu, Dolan, et al.

(2021), replay was simulated with a density of 2000 min^{-1}. Similarly, Y. Liu, Mattar, et al. (2021) a

density of 2400 min^{-1} was used. Re-running these simulation scripts, with arguably a more realistic

replay density (i.e. 15 min^{-1} for NREM, as found in X. Jiang et al. (2020), sequenceness failed to reach

significance, even with purely synthetic data (Supplement Figure 6B). Our recreation of the purely

synthetic simulation, closely following Y. Liu, et al. (2021), reproduced an impression of high

sensitivity, with significance thresholds exceeded at densities on the order of ~10–20 events·min^{-1}

(Figure 8A). However, in contrast to empirical resting state, the synthetic baseline showed little of

the oscillatory and non-stationary structure that drives spurious sequenceness fluctuations in real

data (Figure 8B upper left). Moreover, synthetic class patterns were substantially more separable

than empirical patterns, yielding markedly more extreme classifier probability estimates during

reactivation time points than observed in the localizer or in our hybrid simulation (Figure 9A). It is

especially doubtful whether pattern discriminability would be higher during a reactivation event than

during the localizer at the timepoint of peak decodability. Together, these results indicate that purely

synthetic simulations are liable to substantially overestimate TDLM sensitivity unless calibrated to

match both empirical nuisance structure and empirically observed decoder-pattern separability. A

hybrid approach naturally provides both aspects.

TDLM at its core is a parameter-free method. Nevertheless, a large space of decoding

modalities exists that can influence an analysis outcome. In this paper we primarily applied default

choices as previously reported (L1 Logistic Regression trained on peak decoding time point as in Kern

et al., 2024; Kurth-Nelson et al., 2016; Y. Liu et al., 2019; Nour et al., 2021; Wimmer et al., 2020,

2023). Nevertheless, we note many different parameter combinations have been applied in other

studies and their combined effect remains unclear (Eldar et al., 2020; Q. Huang & Luo, 2023; Wimmer

et al., 2020; Wise et al., 2021). For example, here we observed that mean baseline probabilities

produced by the classifier vary per participant (Supplement Figure 1). While probabilities increased

more or less linearly across all participants, we opted to normalize these by dividing the mean

probability per class per participant. Similarly, in an exploration dataset, we observed that

sequenceness effects were driven by a few participants who produced high magnitudes of

sequenceness scores. To mitigate these effects, we z-scored sequenceness for each trial, as reported

previously (Wimmer et al., 2020; Yu et al., 2023). However, this approach also has its drawbacks, as it

could potentially deflate true events in participants with good decodability and so inflate false

positives, and vice versa.

Furthermore, many different classifier regimes have been used in the context of TDLM (e.g. Lasso

Logistic Regression vs SVM, sensor-based decoding vs. temporal embedding with PCA, normalization

of sequenceness scores via z-scoring vs no normalization, using classifier raw decision values vs

probability estimates, binary-class training vs multi-class training to only name a few). It remains

unclear how these, and especially their combinations, impact TDLM's ability to differentiate true

sequenceness results from a systemic bias due to, for example, to endogenous oscillatory

confounders. To improve methodological certainty, further studies could usefully investigate this

question in form of a multiverse re-analysis of previously published data (Chambers, 2020; Dafflon et

al., 2022).

Improving TDLMs sensitivity

TDLM is a complex and sophisticated approach for detecting human replay. However, in our study,

TDLM was not sufficiently sensitive to detect replay during offline periods, under an assumption that

replay was indeed present in this 8-minute rest-period. Nonetheless, we suggest several factors

could potentially be improved to increase its power and sensitivity. First, the success of the approach

hinges on the quality of the decoders. Improving their specificity as well as their ability to transfer

from the localizer to a condition of interest should be assessed more closely. While most decoding

approaches simply assume that transfer is seamless, previous research has indicated that this

depends heavily on target modality, such as the way the decoders are trained as well as the cognitive

processes involved (Bezsudnova et al., 2024; Dijkstra et al., 2019, 2021). Further research should

optimize, and validate, that decoders trained primarily on sensory evoked activity can be flexibly be

applied to replay detection, for example by introducing a control that demonstrates classification

accuracy in relation to another sensory modality. If transfer is not possible from visual sensory to, for

example auditory or imagined stimuli, it might indicate failure of representational transfer that is

likely to be a necessity for detection of reactivation during replay. Relatedly, focusing on the peak of

the decodability probably introduces a reliance on sensory processing that does not necessarily

transfer well to other, sensory and non-sensory, conditions. Thus, training decoders independently,

or as an ensemble of the time dimension as, for example, performed by Eldar et al. (2020) or Wise et

al., (2021) could help circumvent this problem. Additionally, a wide variety of oddballs has been used

(e.g. upside-down, scrambled, or mismatched images, cues presented visually, as words, auditorily,

etc), and at this time it is unclear if these affect the representations that the classifier learns.

Similarly, performing localizers on categorical stimuli instead of specific stimuli, as well as training on

multi-sensory localizers, might prove fruitful, as it reduces overreliance on specific visual features

that might not central to the reactivated memory representation. In summary, we would expect a

multimodal categorical localizer, and a classifier that is not trained on a specific timepoint, to

generalize best.

Second, as our simulation shows, detecting replay in shorter windows of time should be

possible, as here a high replay density needed has previously been reported (Norman et al., 2019).

Therefore, restricting the analysis to time 'regions of interest' wherein a high replay frequency is

expected could drastically improve detection power. As it is known that replay occurs in bursts, such

windows might exceed necessary replay densities. At this stage we cannot offer a general

recommendation for window sizes as they are likely to depend on details of the research paradigm.

However, intracranial recordings can be used as proxy to estimate the duration of replay bursts, for

example as reported in Norman et al. (2019) increased SWRs were seen up to 1500 ms after retrieval

cue onset. While studies researching planning or retrieval naturally offer such restricted analysis

windows, this becomes more difficult for longer recordings such as resting state and sleep analysis.

Furthermore, to prevent false positives and "double-dipping" (Kriegeskorte et al., 2009), such time-

windows would need to be chosen beforehand or include an appropriate control for multiple-

comparisons.

Detection might also be enhanced through use of contrasts of specified oscillation phases (e.g. slow

oscillations during sleep or theta phases during resting state) during which it is known that replay

typically occurs (Xiao et al., 2024). Another approach would be contrasting windows after targeted

memory reactivation cues and respective control phases of control stimulation to validate the replay

detection (Z. Chen et al., 2024). Last but not least, search windows could also be artificially narrowed

by using a sliding window similar to the approach implemented by Wise et al., (2021). However, in

this case multiple comparison problems between different windows are introduced and circular

testing would need to be avoided (e.g., by using exploration and validation datasets). Nevertheless,

as the considerable branching factor poses a threat of increased false-positive findings we opt to

focus the current simulations on previously published pipelines and parameters. Future studies

should systematically evaluate parameter choices on TDLM under different conditions, something

beyond the remit of the current study.

Lastly, there are certain assumptions that TDLM makes that might not hold (see Methods

Study II). Current implementations look for a fixed time lag that is the same across all participants

and between all reactivation events. If time lags differ across participants, TDLM will fail to find them.

Similarly, TDLM assumes a fixed sequence order and is not robust against slight within-sequence

permutations or in-sequence-missing reactivation events. However, from other data sources., such

as hippocampal place cell recordings, it is known that such permutations can occur where some

states are skipped or fail to decode during replay. Similarly, it is assumed that each reactivation event

lasts between 10-30 milliseconds, but the true temporal evolution of reactivation measured by TDLM

is currently unknown. Future method development might focus on improving invariance to these

assumptions.

Furthermore, in this paper, in collaboration with some of the original authors of TDLM, we

introduced a novel way to test for significance that is more robust and sensitive than that previously

used state-permutation maximum. The newly proposed sign-flip permutation test has markedly

increased power and is more robust, as can be observed from the analysis presented in Supplement

Figure 11, and its use is strongly recommended for future studies using TDLM.

Finally, what do our simulations imply for the broader MEG replay literature? Our

implementation successfully detects replay when boundary conditions are met, as shown in the

simulation. But sensitivity depends critically on high fidelity between the analysis window and the

density of replay events. A systematic evaluation of these conditions as they apply to prior studies

remains beyond the scope of the current paper. Instead, our focus is on delineating boundary

conditions that we hope will motivate conduct of power analyses in future work as well as inclusion

of simulations that approximate realistic experimental conditions.

In summary, there are several areas where TDLM might be improved, including a restriction

in its search space, improvement in classifiers, a validation of localizer representation transfer to

other domains (e.g. memory representations), and the extension of TDLM to render it more robust

against violations of its core assumptions.

Limitations

In an effort to explain null-findings, we conducted a simulation study. Despite close attention to the

bounds of this simulation, it is possible that the reported results are unique to our analysis pipeline

or recording device (e.g. application of MaxFilter, see Methods section), even though we rigorously

followed best practices of the field. While we are reasonably confident the proposed approach

approximates expected veridical replay situations, certain assumptions of our analysis may be too

strong and not model actual replay well (e.g. the way we chose to extract and insert neural patterns).

Furthermore, it is possible that TDLM behaves differently with different MEG systems, different

paradigms, pre-processing or different analysis approaches and that we simply did not find the

correct parameters to make it work. Nevertheless, we closely followed previous pipelines applied on

the same dataset (Kern et al., 2024) that did find effects using TDLM and also tried many plausible

alternative pipelines without success. Additionally, while the analysis we propose arguably simulates

replay under more realistic conditions than previous work, there are still ways to increase realism

even further, that were out-of-scope for this paper. Future studies could for example simulate

nesting of replay with oscillations such as gamma or theta or look how clustered replay affects

sequenceness. Finally, while initially planning for thirty participants, due to exclusion criteria, our

study featured fewer participants than most previous studies using TDLM (i.e. usually 25-40, but 21

in our study). While we are confident that our simulation results hold under these sample sizes, as

sample sizes of other studies show comparable power to our (Supplement Figure 11), we cannot rule

out a possibility that our null-findings are explained by a lack in power alone.

Conclusion

We were unable to detect replay during an eight-minute rest epoch in our data set using TDLM.

Conducting a simulation, we first show that even if replay was present, implausibly high replay

densities would be necessary to achieve significance at our sample size. Second, despite our best

efforts, we were unable to induce a correlation of sequenceness with behaviour at the individual

level, and uncovered large baseline fluctuations in sequenceness across participants. Taken together,

our results suggest that future studies should refine and validate decoder training paradigms and

limit the search space for successful replay detection during resting states via TDLM. Additionally,

future studies should calculate expected replay densities that would be necessary for reported

results to demonstrate plausibility. Finally, our results highlight the overall importance of delineating

the limits of replay detection for the interpretation of null effects.

Methods Study I: Resting State Analysis

Participants

Note that parts of this methods section are taken from Kern et al. (2024) for convenience of the reader. We recruited thirty participants (15 male and 15 female-identifying), between 19 and 32 years old (mean age 24.7 years). Inclusion criteria were right-handedness, no claustrophobic tendencies, no current or previous diagnosed mental disorder, non-smoker, fluency in German or English, age between 18 and 35 and normal or corrected-to-normal vision. Caffeine intake was requested to be restricted for four hours before the experiment. Participants were recruited through the institute's website and mailing list and various local Facebook groups. A single participant was excluded due to a corrupted data file and replaced with another participant. We acquired written informed consent from all participants, including consent to share anonymized raw and processed data in an open access online repository. The study protocol was approved by the ethics committee of the Medical Faculty Mannheim of Heidelberg University (ID: 2020-609). The experiment and analysis were preregistered at aspredicted.org (https://aspredicted.org/kx9xh.pdf, #68915).

Procedure

 Participants came to the laboratory for a single study session of approximately 2.5 hours. After filling out a questionnaire about their general health, their subjective sleepiness (Stanford Sleepiness Scale, Hoddes et al., 1973) and mood (PANAS, Watson et al., 1988), participants performed five separate tasks while in the MEG scanner. First, before any stimulus exposure, an eight-minute eyes-closed resting state was recorded. This was followed by a localizer task (~30 minutes), in which all 10 items were presented 54 times in pseudo-randomized order, using auditory and visual stimuli. Next, participants learned a sequence of the 10 visual items embedded into a graph structure until they achieved 80% accuracy or reached a maximum of six blocks (7 to 20 minutes). Following this, we recorded another eight-minute eyes-closed resting state to allow for initial consolidation and, finally, a cued retrieval session (four minutes), of which the results have previously been reported in Kern et al., (2024). For an overview, see Figure 1. For an animated version of the task, see GitHub pages at https://cimh-clinical-psychology.github.io/DeSMRRest-TDLM-Simulation/. Note: For brevity, we have only included information that is relevant for the understanding of the current study. Further details on the localizer, learning and retrieval task as well as the stimulus material can be found in Kern et al., (2024).

A Localizer Trial

B Sequence

C Learning Trial

D Retrieval Trial

*Figure 10. Detailed overview of the experiment procedure. **A)** During the localizer task, a word describing the stimulus was played via headphones and the corresponding visual item was then shown to the participant. In 11% of trials, the audio and visual cue did not match and in this case, participants were instructed to press a button on detection (to sustain attention to the stimuli and check for inattention). **B)** Graph layout of the task. Two elements could appear in two different triplets. The graph was directed such that each tuple had exactly one successor (e.g., apple→zebra could only be followed by cake and not mug), but individual items could have different successors (zebra alone could be followed by mug or cake). Participants never saw the illustrated birds-eye-view. **C)** During learning, in each trial one node was randomly chosen as the current node. First, its predecessor node was shown, followed by the current node with the participant then given a choice of three items. They were then required to choose the node that followed the displayed cue tuple. Feedback was then provided to the participant. This process was repeated until the participant reached 80% accuracy for any block or reached a maximum of six blocks of learning. **D)** The retrieval followed the same structure as the learning task, except that no feedback was given.*

Localizer task

Ten stimuli were shown to participants repeatedly in a pseudo-randomized order with 54

presentations per image. They were images drawn from Rossion & Pourtois, (2001), selected to be

diverse in colour, shape and category and for having short descriptive words (one or two syllables)

both in German and English. In each trial, a fixation cross was shown and after 0.75-1.25 seconds a

word describing the current stimulus was played via headphones. Auditory stimuli were created

using the Google text-to-speech API, availing of the default male voice (SsmlVoiceGender.NEUTRAL)

with the image description labels, either in German or English, based on the participants' language

preference. Auditory stimulus length ranged from 0.66 to 0.95 s. Next, the stimulus image was shown

for 1 second. As a distractor, in ~ 11% of the trials the audio and image were incongruent, which

participants should indicate with a button press. These oddball trials were excluded from all further

analysis and decoder training. Image transitions were balanced such that each image was followed by

each other image the same number of times.

Graph Learning

The ten images were randomly assigned to nodes of the graph, as shown in Figure 10B. Participants

were instructed to learn a randomized sequence of elements with the goal of reaching 80%

performance within six blocks of learning. During each block, participants were presented with each

of the twelve edges of the graph exactly once.

After a fixation cross of 3.0 seconds, a first image was shown on the left of the screen. After 1.5

seconds, the second image appeared in the middle of the screen. After another 1.5 seconds, three

possible choices were displayed in vertical order to the right of the two other images. One of the

three choice options was the correct successor of the cued edge. Of the two distractor stimuli, one

was chosen from a distal location on the graph (five to eight positions away from the current item),

and one was chosen from a close location (two to four positions away from the current item) and

neither were connected to the currently onscreen item on the graph. Feedback was shown for 3

seconds. No audio was played during learning. After completing the task, participants were told that

their knowledge would be prompted once more a few minutes later (see retrieval).

Resting State

Two resting state recordings were recorded in the MEG. One before any stimulus exposure at the

beginning of the experiment (pre-RS), as well as a second one (post-RS) after graph learning, before

the retrieval part. Each resting state session lasted eight minutes. Participants were instructed to

close their eyes and to "not think of anything particular".

Retrieval

After the resting state, we presented participants with a single retrieval session block, which followed

the exact layout of the learning task with the exception that no feedback was provided as to whether

entered choices were correct or incorrect (Figure 10D). Analysis of the retrieval period and more

methodological details can be found in (Kern et al., 2024).

MEG Acquisition and Pre-Processing

MEG was recorded in a passively shielded room with a MEGIN TRIUX (MEGIN Oy, Helsinki, Finland) with 306 sensors (204 planar gradiometers and 102 magnetometers) at 1000 Hz with a 0.1–330 Hz band-pass acquisition filter at the ZIPP facility of the Central Institute for Mental Health in Mannheim, Germany. Before each recording, empty room measurements ensured no ill-functioning sensors were present. Head movement was recorded using five head positioning coils. Bipolar vertical and horizontal electrooculography (EOG) as well as electrocardiography (ECG) was recorded. After recording, the MEGIN proprietary MaxFilter algorithm (version 2.2.14) was run using temporally extended signal space separation (tSSS) and movement correction with the MaxFilter default parameters (Taulu & Simola, 2006, raw data buffer length of 10 s, and a subspace correlation limit of .98. Bad channels were automatically detected at a detection limit of 7; none had to be excluded. The head movement correction algorithm used 200 ms windows and steps of 10 ms. The HPI coil fit accept limits were set at an error of 5 mm and a g-value of .98). Using the head movement correction algorithm, the signals were virtually re-positioned to the mean head position during the initial localizer task to ensure compatibility of sensor-level analysis across the recording blocks. The systematic trigger delay of our presentation system was measured and visual stimuli appeared consistently 19 milliseconds after their trigger value was written to the stimulus channel, however, to keep consistency with previous studies that do not report trigger delay, timings in this publication are reported uncorrected (i.e., 'as is', not corrected for this delay).

Trial data were pre-processed using Python-MNE (version 1.7, RRID:SCR_005972, Gramfort, 2013). Data were down-sampled to 100 Hz using the MNE function 'resample' (with default settings, which applies an anti-aliasing filter before resampling with a brick-wall filter at the Nyquist frequency in the frequency domain) and ICA applied using the 'picard' algorithm (Ablin et al., 2018) on a 1 Hz high-pass filtered copy of the signal using 50 components. As recommended, ICA was set to ignore segments that were marked as bad by Autoreject (Jas et al., 2017, RRID:SCR_022515) on two-second segments. Components belonging to EOG or ECG and muscle artifacts were identified and removed automatically using MNE functions 'find_bads_eog', 'find_bads_ecg' and 'find_bads_emg', using the EOG and ECG as reference signals. Finally, to reduce noise and drift, data were filtered with a high-pass filter of 0.5 Hz using the MNE filter default settings (hamming window FIR filter, -6 dB cutoff at 0.25 Hz, 53 dB stop-band attenuation, filter length 6.6 seconds). To improve numerical stability, signals were re-scaled to similar ranges by multiplying values from gradiometers by 1e10 and from magnetometers by 2e11. These values were chosen empirically by matching histograms for both channel types. As outlier values can have significant influence on the computations, after re-scaling, values that were still above 1 or below -1 were "cut off" and transformed to smaller values by multiplying with 1e-2.

Trials from the localizer task were created from -0.1 to 0.5 seconds relative to visual stimulus

onset to train decoders. We realized that the initial pre-registered window of 0 to 0.3 seconds was

rather short and we deviated from the pre-registration by extending the analysis window by 100

milliseconds before and 200 milliseconds after stimulus onset. However, the deviation was only to

improve interpretability and to include baseline data and did not alter the results in any way. No

baseline correction was applied. To detect artifacts, Autoreject was applied using default settings,

which repaired segments by interpolation in case artifacts were present in only a limited number of

channels and rejected trials otherwise. Anonymized and maxfiltered raw data are openly available at

Zenodo (RRID:SCR_004129), the localizer data has been previously published as part of a previous

publication (v1 https://doi.org/10.5281/zenodo.8001755), the resting state data can be found at v2

(https://doi.org/10.5281/zenodo.15629081), code and behavioural data will be made public on

GitHub (RRID:SCR_002630) at https://github.com/CIMH-Clinical-Psychology/DeSMRRest-TDLM-

Simulation.

From the resting state, segments were taken between the onset and offset markers, which

were approximately 8 minutes apart and in which the participant had the eyes closed. `Autoreject` is

not well tested on longer segment data, therefore, unlike for the localizer trials, we did not apply it to

the resting state data. However, a supplementary analysis revealed that both ICA and Autoreject did

not have any significant effect on decoding accuracy, a result also reported in other studies

(Delorme, 2022).

Split in exploration and validation dataset

As stated in the pre-registration, we used 50% of the resting state data for exploration to tune

classifiers and try out different approaches. To mitigate time effects on replay, we used interleaved

30 second-segments of the resting state data to divide into exploration and validation sets. To

counterbalance ordering effects, for the first 15 participants we concatenated the uneven segments

as exploration, while for the latter participants switched the assignment. After settling for the

parameters reported in this section, we used the other 50% of data to validate the results. All results

described in Study I are from the validation segments.

Decoding framework and training

We closely followed the approaches applied in previous research (Kurth-Nelson et al., 2016; Y. Liu et

al., 2019; Wimmer et al., 2020) and using the same data set as reported in our previous study (Kern

et al., 2024). We used Lasso regularized logistic regression on sensor-level data of localizer trials

using the Python package Scikit-Learn (Pedregosa et al., 2011, RRID:SCR_002577). Decoders were

trained separately for each participant, and each stimulus, using liblinear as a solver with 1000

maximum iterations and a L1 regularization of C=9.1, which was determined by cross-validating the

decoding accuracy in the time span of 150-250 ms after stimulus onset across participants using

values for C of a log-spaced interval from 10^{-2} to 10^{2} (see Supplement Figure *7* for details). To

circumvent class imbalance due to trials removed by Autoreject, localizer trials were stratified such

that they contained an equal number of trials from each stimulus presentation by randomly

removing trials from over-represented classes. Using a cross-validation schema (leaving one trial out

for each stimulus per fold, i.e., 10 trials left out per fold), for each participant the decoding accuracy

was determined across time (Figure 2B). During cross validation, for each fold, decoders were trained

on data of each 10 milliseconds time step and tested on left out data from the same time step.

Therefore, decoding accuracy reflects the separability of the stimulus classes by the sensor values for

each time step independently. Decoders were trained using a one-vs-all approach, which means that

for each class, a separate classifier was trained using positive examples (target class) and negative

examples (all other classes) plus null examples (data from before stimulus presentation, see below).

This approach allows the decoder to provide independent estimates of detected events for each

class.

For each participant, a final set of decoders (i.e., 10 decoders per participant, for each stimulus one

decoder) were trained at 210 milliseconds after stimulus onset, a time point reflecting the average

peak decoding time point computed for all participants. For the final decoders, data from before the

auditory stimulus onset was added as a negative class with a ratio of 1:2, based upon results from

previous publications reaching better sensitivity with higher negative class ratio (Y. Liu, Dolan, et al.,

2021). Adding null data allows decoders to report low probabilities for all classes simultaneously in

absence of a matching signal, and reduces false positives while retaining relative probabilities

between true classes. Together with the use of a sparsity constraint on the logistic regression

coefficients, this increases the sensitivity of sequence detection by reducing spatial correlations of

decoder weights (see also Liu et al., 2021). For a visualization of relevant sensor positions, see

Supplementary section of Kern et al., (2024). Decoders were then applied to the two-resting state

session data. This resulted in ten probability vectors across the resting state sessions, one for each

item, in steps of 10 milliseconds. These probabilities indicate the similarity of the current sensor-level

activity to the activity pattern elicited by exposure to the stimulus and can therefore be used as a

proxy for detecting active representations, akin to a representational pattern analysis approach (RSA,

Grootswagers et al., 2017).

Sequential replay analysis

To test whether individual items were reactivated in sequence at a particular time lag, we applied

temporally delayed linear modelling (TDLM, Liu et al., 2021) to the resting state segments, using a

Python (RRID:SCR_008394) implementation of the algorithm (TDLM-Python, Kern, 2024). In brief, the

TDLM method approximates a time lagged cross-correlation of reactivation strength in the context of

a particular transition pattern, quantifying evidence for a certain activity transition pattern

distributed in time. As input for the sequential analysis, we used the probabilities of the ten

classifiers corresponding to the stimuli, which we normalized by dividing each class probability by the

mean of all classes.

Using a linear model, the method first estimates evidence for sequential activation of the decoded

item representations at different time lags. For each item i at each time lag Δt up to 300 milliseconds

it estimates a linear model of form:

$Y\_i = Y(\Delta t) \times \beta\_i (\Delta t)$

where Y_i contains the decoded probability output of the classifier of item i and $Y(\Delta t)$ is simply Y time

lagged by Δt. When solving this equation for $\beta\_i (\Delta t)$ the predictive strength of $Y(\Delta t)$ for the

occurrence of Y_i at each time lag Δt can be estimated. Calculated for each stimulus i, an empirical

transition matrix $T\_e (\Delta t)$ that indexes evidence for a transition of any item j to item i at time lag Δt

can be created (i.e., a 10x10 transition matrix per time lag, each column j contains the predictive

strength of j for each item i at time lag Δt). These matrices are then combined with a ground truth

transition matrix T (encoding the valid sequence transitions of interest) by taking the Frobenius inner

product. This returns a single value $Z\_{\Delta t}$ for each time lag, indicating how strongly the detected

transitions in the empirical data follow expected task transitions, which is termed "sequenceness".

Using different transition matrices to depict forward (T_f) and backward (T_b) replay, evidence for

replay at different time lags for each trial is estimated separately. To control for auto-correlations of

the signal, caused by large brain oscillations such as occipital alpha (Y. Liu, Dolan, et al., 2021), we

added time-shifted versions of the decoded probabilities to the GLM in 10 Hz intervals. This process

is applied to each trial individually, and resulting sequenceness values are averaged to provide a final

sequenceness value per participant for each time lag Δt. To test for statistical significance, a baseline

distribution is sampled by permuting the rows of the transition matrix 1000 times (creating transition

matrices with random transitions; identity-based permutation (Liu et al., 2021) and calculate

sequenceness across all time lags for each permutation. The null distribution is then constructed by

taking the peak sequenceness across all time lags for each permutation. To minimize possible

participant-specific biases on the magnitude of sequenceness values, we z-scored along the time lag

axis for each permutation, for each participant, similar as previously proposed (Wimmer et al., 2020).

Sign-flip-permutation test

As previously noted (Y. Liu, Dolan, et al., 2021) the standard approach of taking the maximum or 95th

percentile of sequenceness from state-shuffled permutations controls for multiple comparisons very

conservatively. Its method implements a fixed-effects model, as it averages data across participants

to establish a threshold without accounting for inter-subject variability. Consequently, the results are

susceptible to outliers, which we partially tried to mitigate by z-scoring sequenceness values. To

implement a random-effects inference framework that accounts for between-participant variance, a

sign-flip permutation test can be employed. This approach is significantly more robust to outliers and

follows established non-parametric methods for multiple comparison correction in neuroimaging

(Nichols & Holmes, 2002). In this procedure, the sign of the entire sequenceness curve for a random

subset of participants is inverted for each permutation. A one-sided t-test against zero is then

calculated at every time lag for each permutation. By taking the maximum t-value across all time lags

for each permutation, an empirical null distribution of the maximum statistic is constructed. The

observed t-values are then compared against this distribution, providing Family-Wise Error Rate

(FWER) control across the temporal axis. This method offers higher sensitivity than state-shuffling

because it preserves the temporal structure of the individual data while explicitly penalizing effects

that are inconsistent across the cohort. Furthermore, this framework naturally extends to more

complex experimental designs, such as multiverse analyses involving varying classifier parameters or

regularization strengths.

Resting state analysis

To assess whether previously learned sequences are replayed during the resting state, we tested for

significant sequenceness using the above-mentioned permutation test for forward and backward

sequenceness. We tested for all allowable directed graph transitions. Additionally, we compared the

sequenceness results to the control resting state by subtracting the sequenceness values on the

participant level. Additionally to the maximum across all permutations, we report the 95%

confidence interval of the maxima across each permutation (Y. Liu, Dolan, et al., 2021). Finally, to

assess if sequential replay was correlated with the participant's performance, we calculated a

Pearson correlation between peak sequenceness across all time lags and the retrieval performance

(as in Kern et al., 2024).

Parameter exploration

For completeness and transparency, we report what parameters we explored during the exploration

phase on the first half of the data set (final results were calculated on a held-out portion of the

resting state data). For classifiers, we tried Random Forest Classifiers, Support-Vector-Machines as

well as LDA. We applied different normalizations of the MEG data by z-scoring the raw data along

channels, times as well as across the entire recording. These were all inferior in terms of localizer

decoding accuracy compared to the current approach. We trained on single time points and averaged time windows, including PCA. We employed z-scoring on the sequenceness values, on the probability estimates as well as using raw classifier outputs instead of probability estimates or normalizing each class by its mean probability. Nevertheless, after considering the vast space of possible parameters, we did not find alternative parameters that robustly showed a predicted replay pattern and thus, to avoid overfitting and remain comparable to the literature, we report this study with parameters most often used in previous research in terms of normalization, classifier choice (Logistic Regression) as well as training regimen. Crucially, these parameters were successfully applied by us to identify replay effects in our previous study using the retrieval data from this data set (Kern et al., 2024).

Exclusions

Replay analysis relies on a successive detection of stimuli, where the chance of detection exponentially decreases with each step (e.g., detecting two successive stimuli with a chance of 30% leaves a 9% chance of detecting a replay event). However, one needs to bear in mind that accuracy is a "winner-takes-all" metric indicating whether the top choice also has the highest probability, disregarding subtle, relative changes in assigned probability. As the methods used in this analysis are performed on probability estimates, and not class labels, one can expect that 30% is a rough lower bound and that the actual sensitivity within the analysis is higher. Additionally, based on pilot data, we found that attentive participants were able to reach 30% decodability, allowing its use as a data quality check. Therefore, we decided a priori that participants with a peak decoding accuracy of below 30% would be excluded from the analysis (nine participants in total), as obtained from the cross-validation of localizer trials. Additionally, as successful above-chance learning was necessary for the paradigm, we ensured all remaining participants had a retrieval performance of at least 50% (one participant had to be excluded, but was already excluded due to low decoding performance). No participant had more than 25% missed button presses, so no exclusion was necessary on this pre-registered criterion.

Methods Study II: Simulation

To better understand our null-finding, we conducted a simulation study. We applied a hybrid simulation approach, using resting state data recorded before any stimulus exposure as a sequence-free baseline, into which we inserted neural patterns taken from the localizer. By inserting these patterns at a specific time lag, we can create a ground truth of replay-like events that should be well-decodable. The advantage of using real-world MEG data over synthetic data is that it includes all nuisance confounders that might be present under realistic conditions, yet leaves room to control

the variable of interest. Finally, by varying the density and length of inserted events, we can assess an

upper bound of sensitivity of TDLM under optimal, yet realistic conditions.

Simulating replay by inserting localizer data into the control resting state

TDLM makes assumptions (implicit and explicit) about the underlying structure of replay. While it has

yet to be shown these assumptions are based on the characteristics of actual replay, all methods

pose other constraints (Cliff et al., 2023). Here, we aimed to create an optimal scenario tailored to

TDLM. Therefore, we adhere to the following assumptions:

1. Individual items are reactivated and can be decoded throughout the recording

2. Individual item reactivations are brief, usually only of ~10-30 ms[1] duration

3. Item reactivations are non-random and adhere to a specific sequential pattern

4. A replay event comprises two or more reactivation events in close temporal proximity

5. There is a fixed time lag between all reactivations across all replay events

Additionally, we include the following constraints for the simulation

1. Classifiers trained on the peak time point of decodability of sensors from a localizer transfer
953   well to reactivation events

954   2. Between each replay event (two or more stimulus reactivations), there is a refractory period
955   of at least the time lag of interest in the analysis (i.e. 150 milliseconds)

3. Replay is correlated with memory performance during the retrieval session.

4. The replay density is stable across the recording.

Based on these assumptions, we inserted replay events into the control resting state of each

participant by first extracting stimulus-specific patterns of sensor activity and then added these

patterns at sequential positions in the resting state. We created patterns by taking the mean sensor

activation at the peak of decodability during the localizer of a specific class (as defined above, i.e. at

210 milliseconds). To isolate the stimulus-specific pattern from activity elicited by general sensory

processing, we then subtracted the mean sensor values across all other classes (akin to an event

related potential of seeing any stimulus except the current target stimulus), for a visualization of the

mean patterns across participants, see Supplement Figure 8. Next, we determined N replay onset

points *t* throughout the resting state and inserted the pattern of a randomly chosen class by simply

adding the pattern to the sensor values at time t_i. After a fixed time, lag of 80 milliseconds after t_i, we

inserted the pattern of the class that would follow according to the underlying sequence structure.

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[1] As for example shown in the replay example visualization in Figure 2 in Liu et al., (2019) or Figure 3 in Wimmer et al., (2020)

This time lag (80 msec) was chosen in order to isolate precisely an effect of the experimentally

inserted sequenceness. Thus, we chose a lag at which the mean baseline sequenceness was close to

zero and where the correlation with behaviour was low. Additionally, we subtracted the mean

sequenceness value (at 80 milliseconds) at baseline from the specific lag recorded for each

participant, such that simulation effects would be initialized at zero sequenceness on average

enabling any effects to be attributed purely to inserted replay. Additionally, we excluded time lags

too close to the alpha rhythms of participants (which varied between ~9-11 Hz) or lags which would

have a harmonic with the rhythm. We did not allow 150 ms after and before each replay event from

containing any other replay events, as this would have introduced interactions between two replay

events within the used search window of 300 ms. For simplicity, we only used one-step replay event

transitions, i.e. containing two reactivation events.

Figure 11 shows the decoded probabilities at the timepoints *t* before and after addition of the

stimulus-specific pattern. The target class is now well decodable, while other classes remain at a low

baseline probability (Figure 11). Using this approach, we can simulate replay events while for the

most part preserving other characteristics of the resting state in terms of endogenous oscillations

that might introduce spurious sequenceness. However, note that this simulation approach is likely to

slightly over-estimate decodability of reactivation events (which reflects in an average decoding

accuracy of 43% after insertion which is higher than the decoding accuracy during the localizer).

Additionally, as the classifier follows a one-versus-all classification approach it estimates probabilities

for each class separately. Therefore, a small increase in probability can be seen even for non-target

classes, which most probably stems from remaining patterns from general sensory processing and

the slight increase in sensor magnitude after addition. Sensor values still follow the distribution of a

naturalistic MEG signal after inserting patterns, i.e. are centred around zero and have a comparable

standard deviation (see Supplement Figure 2).

Figure 11. Normalized probability estimates of time points in the control resting state for the target class of which the pattern has been added and non-target class. Left: Time points at which no pattern has been inserted, both classes have a similar probability estimate at baseline. Right: At time points where the stimulus-specific pattern was added, the decoded probability estimate is significantly increased. Note that also the probability estimate of non-target classes is slightly increased at the timepoint of pattern insertion, which is due to slight shifts in sensor values that heightens all probabilities by a small amount.

Replay density

Unlike intracranial or cell recordings, TDLM does not per se detect individual replay events,

but given a specific transition pattern ascertains evidence for sequenceness over a period of time.

Each in-sequence reactivation of two or more items increases the specific time lag's sequenceness

value. Therefore, the sequenceness magnitude strongly relies on the underlying density of decoded

replay events, which is usually denoted in events per minute or min^{-1}. To assess the sensitivity of

TDLM, we incrementally increased the density up to a theoretical maximum of around 200 min^{-1}, as

given by the time lag of 160 milliseconds plus the above-defined refractory period of 150

milliseconds. Additionally, as replay is seen as instrumental to memory consolidation, the density

was linearly scaled with the participant's retrieval memory performance (e.g. a participant with 50%

memory performance had at maximum 80 min^{-1} = 160 min^{-1} * 50% events inserted). To assess how

performance might affect replay in our specific dataset, we chose to use the original participants'

performance values instead of uniformly sampling the performance space (which ranged from 50 to

100%). However, for the correlation analysis, we additionally added a "best-case" scenario, in which

we scale replay from 0 to 1, an approach that is statistically equivalent to scaling values to the full

space of possible performance (0 to 100%) (see Results Study II: Simulation). Note that we

deliberately overestimated the link between replay and memory performance which has been

estimated to be as low as r=0.2 in humans (Schapiro et al., 2018).

Additionally, the timepoints of inserting replay within the resting state are sampled from a uniform

distribution. Even though TDLM tracks reactivation events over time, at a macro-scale, the algorithm

is invariant to the temporal distribution. At each time step, the GLM regresses onto a future time

step up to the maximum time lag of interest, yielding a predictor per lag. However, these predictors

within the GLM are independently assessed, and hence, TDLM is, outside of the time lag window,

relatively invariant to the temporal distribution of replay. To demonstrate our claim, we simulated

uniform replay vs "bursty" replay that only occurs in some parts of the resting state, both yield

equivalent sequenceness results (see Appendix A.1 and Appendix A Figure 1).

Reactivation event length

It is unclear how long is the span of individual decodable reactivation events in the context of MEG

decoding. In primates, sharp wave ripples (SWR), as a detectable proxy for a reactivation event, are

usually searched for in the range of 20 to 200 milliseconds (A. A. Liu et al., 2022), while mostly found

to range from 50 to 100 milliseconds (Buzsáki, 2015), hence each individual reactivation is much

shorter. There is no report of decoded reactivation event lengths in the context of MEG, however,

spans can be estimated from example illustrations given by previous studies, where they mostly span

10-30 milliseconds. Nevertheless, even if reactivation events are brief, there seems to be a waxing

period, a peak and a waning period. Therefore, we simulated this process by extracting the item-

specific pattern around the decoding accuracy peak and weighting it with a gaussian (sigma=1, i.e.

item-specific pattern at 210 milliseconds +- 20 milliseconds was inserted into 50 milliseconds centred

around the reactivation event marker in the resting state with the following weighting: [0.058, 0.24,

1, 0.24, 0.058].).

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Synthetic Simulation

In our paper we argue that a hybrid simulation is superior to previously proposed purely synthetic

simulation approaches, mainly for overestimated pattern discriminability of synthetic data as well as

a lack of oscillatory confounds. We have conducted a purely synthetic simulation that is based upon

the sample code proposed in Y. Liu, Dolan, et al. (2021). Following their sample code, we simulated

baseline resting state MEG by first sampling a symmetric covariance matrix to simulate sensor

dependencies. Using this covariance matrix, we generated spatially correlated noise and introduced a

temporal dependency by ensuring each time point was strongly correlated with the previous one

(correlation of 0.95, AR(1) process). This yielded a continuous time-series that preserved both the

spatial structure across sensors and the natural smoothness of the signal over time. For the classifier

patterns, same as in the example code, we first created a common pattern (i.e. akin a visual

processing ERP) by drawing from a normal distribution for each sensor. Then we created ten

stimulus-specific patterns by drawing from a normal distribution once again and adding to the

common pattern for each class. Then we simulated eighteen noisy trials per class, by once again

adding noise to the components for each trial with a ratio of 4:1, on which we trained a classifier.

Additionally, half of the patterns had additional normal noise added to them. Null data was created

by sampling from a normal distribution and adding with a ratio of 1:1. We used a logistic regression

with default settings and L1 regularization (C=1). To insert the events into the synthetic resting state,

we used the same procedure as in our main simulation. We inserted the stimulus-specific patterns in

varying densities to the data, with each event simulating one transition from the graph. Density was

again inversely scaled by sampling from the real performances of our participants. Diverging from the

original MATLAB code, we did not add a jitter to the time lag and only simulated up to a density of

200 min^{-1}, as the density of 2000 min^{-1} used in the original code necessarily introduces a lot of

overlap of replay events. Finally, as in the previous synthetic simulation, we induced a temporal

autocorrelation in the predicted resting state probabilities by adding a noise vector with a temporal

dependency of 0.95. Note, as there was no inter-individual difference in magnitude of sequenceness

scores (as they were simulated at the same scale), we did not apply z-scoring to the sequenceness

values (in line with the original simulation).

To assess discriminability of the inserted patterns, we extracted classifier probabilities at

time points at which we knew a true pattern was present. This was naturally the case during the peak

decodability of the localizer and at time points during the simulations in which we had inserted a

pattern. For the localizer, we used cross-validation to extract probability estimate for the target class

and all other classes of left-out trials. These can be seen as an empirical upper bound of how well

patterns are discriminable, as we hypothesize that patterns during a reactivation event will be less

pronounced than during visual processing. To assess how far apart the probability distributions of

target classes and other classes are, we used the Wasserstein distance ((Ramdas et al., 2017; Villani,

2009)). The Wasserstein distance measures how much "work" would be necessary to transform a

distribution into another, also called "earth-movers-metric". A larger Wasserstein distance W should

reflect a better pattern discriminability of the classifier, with 0 denoting no discriminability and

higher values reflecting more separate distributions. However, note this approach only yields a rough

estimate and has the caveat that pattern separability is effectively acting on a trial basis (i.e. a trial

with low probabilities could still be separable even if the winner class is only slightly higher), while we

compare probability distributions across all trials/timepoints. We use it nevertheless in favour of, for

example, taking the ratio within a trial as it allows for a scale-invariant approach of comparing

distributions of unequal sample size. Finally, to simulate increased pattern strength, we simply

multiplied the inserted patterns by a specific scalar and assessed the Wasserstein distance of

probabilities after insertion.

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Data availability

MaxFiltered and anonymized MEG raw data are available at Zenodo in two parts: Localizer data is

available as v1 of https://doi.org/10.5281/zenodo.8001755 (Kern, 2023). Resting state data is

available as v2 of https://doi.org/10.5281/zenodo.15629081 (Kern, 2025).

Code availability

The code of the analysis as well as the experiment paradigm and the stimulus material are available

at https://github.com/CIMH-Clinical-Psychology/DeSMRRest-TDLM-Simulation.

Supplement figures

For supplementary figures of this draft, see after references.

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Supplement

Which pattern to extract from the localizer

To extract a suitable pattern that can be added to the control resting state, some constraints need to be addressed. First, the pattern should elicit an increase in probability specific to the target class when added to the resting state data. Consequently, all non-target classes should not increase significantly compared to baseline. Because, if the baseline probability is increased, spurious sequenceness will occur simply due to the baseline increase, as is the case with oscillatory confounders (Y. Liu, Dolan, et al., 2021). Additionally, classifiers will also react to any event-locked visual activity, slightly increasing the probability of the target class, even if seeing an unrelated image. We initially elicited these patterns by taking the mean ERP per class adding the sensor values of the peak decoding time point to resting state positions. However, this shifted the distribution of sensor values in such a way that the baseline probability of non-target-classes also increased (Figure 11). Therefore, we opted to refine the pattern by taking the ERP of all other classes and subtracting it from the ERP of the target class. This has the advantage that it removes most of the background visual activity and only keeps activity specific to the target class.

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Supplement Figure 1. Mean classifier probabilities before and after normalization, per participant for time points selected for replay pattern insertion. Left: Variance in baseline probability estimates can already be seen before pattern insertion, i.e. for some participants a higher baseline is present that for others. This higher baseline probability values (or rather their stronger absolute fluctuation) will lead to falsely increased sequenceness values in the GLM betas. Right: After normalization, baseline probabilities are similar across all participants. Blue shading indicates the participant's performance with lighter shading at 50% and darkest shading at 100% memory performance.

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Supplement Figure 2. Distribution of raw sensor values before and after insertion of the simulation pattern. As MEG measures differences in magnetic field strength, the signal is centred on zero. Adding a pattern that is extracted from a zero-centred distribution should not affect the distributions mean and hence still produce a valid MEG signal, which is shown in the figure. However, a very small difference in spread between the two distribution can be seen, indicating that sensor values are not exactly following the same distribution characteristics after adding the stimulus-specific pattern, which might explain the small increase in baseline probability for non-target classes in Figure 11.

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Supplement Figure 3. Individual sequenceness curves for all participants in the control resting state (left) and the post-learning resting state (right), for forward (upper), and backward sequenceness (bottom). The participant's performance is indicated in shades of blue, from the lightest shade at 50% to the darkest shade at 100% retrieval performance. Note that there is an oscillatory component present in all sequenceness curves, however, their phase and exact amplitude seem to be shifted.

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*Supplement Figure 4. Individual sequenceness curves for different segments. **A)** Control resting state and **B)** post-learning resting state. Each segment is taken from minute N of the respective resting state. For each participant, from each segment either the first or second 30-second interval was taken. Two significance thresholds are indicated, the more lenient 95% quantile of the maximum sequenceness of all permutations (lightly dotted line) and the maximum mean peak sequenceness across all permutations of all subjects (boldly dotted line). Sequenceness crossed the 95% permutation threshold in the control resting state in minute 4 (forward 100 millisecond time lag). Note that sequenceness scores are negative for the latter, which would indicate a replay-suppression of these specific sequences.*

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Supplement Figure 5. Sequenceness curves when not correcting for alpha oscillations. An increase in oscillatory behaviour of the sequenceness curves can be observed. Alpha correction is usually applied to mitigate nuisance effects that are introduced by strong eyes-closed occipital alpha by adding probability time lines in 100ms steps to the GLM (Y. Liu, Dolan, et al., 2021).

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*Supplement Figure 6. Rerun of original MATLAB simulation code. **A)** Original TDLM (Y. Liu, Dolan, et al., 2021) simulation with a replay density of 2000 min⁻¹. **B)** Re-run of the code with realistic replay density of 15 min⁻¹. Using this density, significance thresholds are barely reached. **C)** Baseline sequenceness with no inserted sequences, some participants reach significance even in absence of replay, as can be seen by the confidence intervals **D)** Left: Time points at which no pattern has been inserted, both classes have a similar probability estimate at baseline. Right: At time points where the stimulus-specific pattern was added, the decoded probability estimate is significantly increased, cf. Figure 11. **E)** Probability estimates before and after insertion of the replay pattern, visualised as in Supplement Figure 1. The probability estimates at baseline are close to zero while increased by several orders of magnitude. Using synthetic data probably overestimates probability estimates of true events. Note that different from our simulation, the replay time lag has here been sampled from*

a gamma prior with mean 40 milliseconds, such that the majority of events is simulated at 40 milliseconds, with some events getting assigned a lag of 30 or 50 milliseconds.

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Supplement Figure 7. Cross-validation accuracy during the localizer across different regularization values. For each regularization value, a cross-validation was performed per participant as

documented in the methods section, evaluated at 150 to 250 milliseconds after stimulus onset. The best regularization value was found around C=9.1

Supplement Figure 8. Visualization of the sensor patterns that were inserted into the resting state per class, averaged across all participants. Patterns were created on a per-participant-basis by taking the mean sensor value at the peak time point of decodability during the localizer of all trials of the specific class. From these patterns, the mean across all trials (the ERP) at the peak timepoint was subtracted, to remove activity related to general visual processing. Saturation indicates value magnitude (a.U.), red is positive sensor values, blue negative sensor value.

Supplement Figure 9. Individual participant's retrieval performance delta of the post-rest retrieval to the pre-rest retrieval relative to their sequenceness values at the peak mean sequenceness (indicated in A). Regression lines and 95% confidence intervals are fitted. While we found a non-significant relation between a memory performance enhancement and post-learning forward sequenceness we are cautious not to overinterpret these results. As in the section "Correlation with behaviour only present at high replay speeds" the noted correlational measure oscillates heavily with baseline sequenceness fluctuations, and any true replay effect is likely to be overshadowed by such fluctuations.

Supplement Figure 10. Forward minus backward sequenceness within each resting state session. Previous papers often report subtraction of backward from forward sequenceness (fwd-bkw) as a means to remove oscillatory confounds that impact both sequenceness directions in synchrony. While required in early cross-correlation approaches (Kurth-Nelson et al., 2016), its validity in GLM-based frameworks depends on an assumption that confounds are global and in-phase across sensors. We observed this assumption is violated in our baseline data, where spurious sequenceness occasionally diverges in opposite directions at specific time lags (e.g., ~90 ms). In such instances, subtraction would increase the false-positive rate rather than suppress noise. In Figure 3B, we prioritized the comparison of pre-task versus post-task sequenceness within the same direction, as oscillatory confounds appeared more stable across time within a single direction, as opposed to across directions within a single session. However, we consider both approaches are valid. We now provide the fwd-bkw plots for completeness and comparison with previous literature. A) forward minus backwards sequenceness for Control (left) and Post-Learning resting state (right). B) T-value distribution of the sign-flip permutation test for Control (left) and Post-Learning resting state (right).

Supplement Figure 11. Power analysis of sequenceness significance for bootstrapped samples sizes.
A) *Powermap for state-permutation thresholds. However, here the bootstrap approach suffers from a conceptional problem: significance thresholds are defined by the permutation maximum and/or 95-percentile of the maximums across all sequence-permutations across participants. If we resample bootstrap-participants from our existing pool, the maximum thresholds computed will remain relatively stable across resampled participants, as it only compares against the mean and disregards the standard deviation.* ***B)*** *The newly presented statistical approach is significantly more sensitive at higher sample sizes. Note that even then, 80% power is only reached with replay density of higher than 50 min^{-1} at a sample size of 60 participants. Additionally, the sign-flip permutation test assumes that the mean is at zero. As we observed a non-zero mean due to spurious oscillations, we subtracted the mean sequenceness of the baseline condition from each participant before permuting to achieve a null distribution with mean zero, as otherwise, we would have found significant replay effects in the baseline condition at increasing sample size. Nevertheless, due to the higher sensitivity, the new sign-flip test is recommended over the previous sequence-permutation-based test. Colours indicate the power from 0 to 1 for different bootstrapped sample sizes and densities. 80% power thresholds are outlined in black*

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Supplement Appendix A: Replay characteristic and TDLM behaviour

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1530 In this section, we clarify the behaviour of TDLM under three specific conditions. We utilize
1531 simulations to demonstrate that the TDLM sequenceness metric is primarily a function of the total
1532 number of transitions detected within an analysis window.

1533 1) Bursty vs uniformly distributed replay: Invariance to temporal distribution of replay

Previous research has indicated that replay occurs in bursts and that the temporal distribution within
the resting state might not be uniform. Here, we simulated replay either with uniform temporal
distribution for each replay event or a more "bursty" distribution, which focuses replay in specific
areas of the recording. The same number of transitions has been inserted in both cases (taking into
account the refractory period between individual replay events). As can be seen in Appendix A Figure
1, TDLM is blind to temporal distribution of transitions within the analysis window. NB this only holds
true if no overlap between events occurs, i.e. the refractory period is regarded. If there is overlap of
different replay events, e.g. during a burst, a reduction in sensitivity is expected. Additionally,
restricting the analysis window to short period at which a burst is to be expected would increase the
sensitivity as the time window decreases.

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Appendix A Figure 1. TDLM is invariant to the macro-distribution of sequences within the recording.
***A)** forward and backward sequenceness of replay events that were uniformly distributed (light
orange/green) or that were "bursty", i.e., temporally constrained to specific areas (dark
orange/green). Shaded areas denote the standard error across participants. **B)** Probability
distribution from which the starting points for each replay events were drawn. **C)** Histogram of
empirical replay starting positions across all participants.*

1545 2) Tripled replay vs doublet replay: Invariance towards length of replay

We intentionally designed our study to encourage replay of triplets. However, this begs the question
as to whether it matters if triplets or individual chunks of a sequence are replayed at different time
points? Here, we simulated two scenarios. In one, we inserted replay of single transitions alone with
a refractory period, e.g. A->B and separate B->C transitions. In a second scenario, we simulate replay
of chained triplets, e.g. A->B->C, with a distance of 80 milliseconds each. Importantly, we kept the
number of transitions constant (i.e., A->B, … B->C and where A->B->C would both have 2 transitions.
This creates a context wherein a four-minute resting state would have ~100 events of A->B->C
inserted and ~200 events of A->B or B->C, such that in both cases this results in the same number of
single step transitions. We found both are equivalent, with TDLM agnostic to the length of sequence
trains, i.e., it does not matter if replay is chunked or chained under the assumption that the number
of transitions remains fixed, as can be seen in Appendix A Figure 2.

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Appendix A Figure 2. TDLM is invariant to the length of sequence replay trains under an assumption
that the number of target transitions (e.g. single steps) is fixed. We simulated replay either as two
temporally separate A->B, B->C events (light orange/green) or as a single A->B->C event (dark
orange/green), both yielding equivalent sequenceness. Shaded areas denote the standard error
across participants.

3) Full vs actual remembered transitions: Invariance to which specific states are replayed

From literature we know that replay is increased after learning and that less stable memories are
replayed more often. We simulated this effect by scaling our replay density inversely with
performance. However, for simplicity, in our simulation, we inserted sampled transitions from all
valid transitions given by the graph structure, i.e., the following transitions were valid: However, this
meant that some participants would have transitions inserted that they didn't actually remember. To
show that this would not change results, we simulated two scenarios: In the full sequence scenario,
all valid graph transitions are inserted (i.e. all participant's replay is sampled from 'A->B, B->C, C->D,
D->E, E->F, F->G, G->E, E->H, H->I, I->B, B->J, J->A'). In the second scenario (memorized transitions)
we only replayed transitions that the participant actually retrieved correctly during the post-resting
state testing sessions (i.e. a participant's replay would have been sampled from 'A->B, B->C, G->E, E-
>H, H->I', if those were the ones he remembered). In both scenarios, the number of events is kept
constant. The results are equivalent as can be seen in Appendix A Figure 3. NB this only holds under
the assumptions that classifiers are equally good at decoding each class.

Appendix A Figure 3. TDLM is insensitive towards which transitions are replayed and only sensitive to how many transitions are detected in total. Here we simulate transitions either sampled from the full graph (light orange/green) or participant-specific transitions of trials that participants correctly remembered (dark orange/green). Shaded areas denote the standard error across participants.

4) Simulating multi-step replay: Density increases while number of transitions stays similar

From rodent literature we know that a replay event often consists of several reactivation events. However, in our study, to keep the simulation as basic as possible and to keep consistent with previous simulations, we intentionally only simulated single transitions. As shown above, TDLM is invariant to the sequence length that is replayed and only sensitive to the number of transitions. Our definition of density scales inversely to the number of transitions in a replay event, giving the impression of an increased sensitivity. Here, we simulate how TDLM behaves under multi-step replay events. We simulated replay with increasing number of transitions while still preventing overlap by expanding the refractory period accordingly. As the number of transitions per replay events scales with the step size, the replay density to find replay is linearly decreasing as well (see *Appendix A Figure 4*). Nevertheless, TDLM's sensitivity is primarily governed by the total number of target transitions present in the data, with all three cases becoming significant at around 70-80 single-step transitions found. Additionally, multi-step sequences pose another problem: If we assume that a classifier has a decoding accuracy of 35%, successfully detect two events in a row would have an upper bound chance of ~12% and decreases exponentially for longer step size. As shown above, our simulation likely over-estimates pattern discriminability and therefore gives a biased view of how multi-step sequences would be detected by TDLM.

Appendix A Figure 4. Sequenceness for longer multi-step replay. Here, we inserted sequence trains of longer replay (e.g. A->B->C->D->E for the 4-step case) without keeping the number of transitions fixed. To prevent overlap, an increase in length of replay is accompanied by a longer refractory window, which leads to a reduced number of maximum replay events per minute. TDLM becomes more sensitive under the assumption that (i) multiple sequence steps are reactivated and (ii) all sequence steps are decoded accurately. Notably, as mentioned before, only the number of single step transitions is relevant for TDLM's sensitivity. Indicated line is sequenceness at the simulated 80-millisecond time lag across different time lags, with participants' standard error in the shaded area.