Oscillatory Control of Cortical Space as a Computational Dimension

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Abstract

Flexible cognition depends on the ability to represent and apply relevant to the current task at hand. This allows the brain to interpret sensory input and guide behavior in a context-dependent manner. Recent work has proposed *Spatial Computing* as a mechanism for this flexibility, suggesting that task-related signals organize information processing through spatial patterns of oscillatory activity across the cortical surface. These patterns are proposed to act as "inhibitory stencils" that constrain where sensory-related information (the "content" of cognition) can be expressed in spiking activity. Here, we provide a comprehensive empirical test of Spatial Computing using multi-electrode recordings from the lateral prefrontal cortex in non-human primates performing a range of cognitive tasks (object working memory, sequence working memory, categorization). We found that alpha/beta oscillations encoded task-related information, were organized into spatial patterns that changed with task conditions, and inversely correlated with the spatial expression of sensory-related spiking activity. Furthermore, we found that alpha/beta oscillations reflected misattributions of task conditions and correlated with subjects' trial-by-trial decisions. These findings validate core predictions of Spatial Computing suggesting that oscillatory dynamics not only gate information in time but also shape where in the cortex cognitive content is represented. This framework offers a unifying principle for understanding how the brain flexibly coordinates cognition through structured population dynamics.

Introduction

A hallmark of cognition is its flexibility, the ability to adjust thoughts and actions according to current goals, rules, and contexts. This flexibility depends on more than just storing information. It requires the brain to control the expression of relevant information, when it should be accessed, and how it should be used. Central to this capacity is the use of learned information to shape how sensory inputs are interpreted and how behavior is guided.

Recent work has proposed a novel principle—*Spatial Computing*—as a potential mechanism for flexible context-dependent control [1,2]. The core idea is that the brain achieves flexible control by using oscillatory dynamics to control neural activity across the cortical surface. In essence, cortical space is utilized as a computational dimension. In this framework, internally-generated signals reflecting goals and context are carried by oscillatory activity, particularly alpha/beta-band oscillations (10–30 Hz). By exerting inhibition, these low-dimensional mesoscale patterns modulate local neuronal excitability and determine which high-dimensional populations of neurons on the microscale can be activated in a given context. This shapes where in the cortex sensory-related spiking activity can occur. Rather than relying solely on fixed anatomical pathways, the brain dynamically organizes information flow through spatially structured oscillatory patterns. In this way, Spatial Computing offers a mechanism for how sensory information is processed and organized in a context-dependent manner.

This framework provides a mechanistic explanation for flexible control of cognition. For example, it helps explain how low-dimensional subspaces encoding different task variables can emerge, allowing the brain to reduce interference and flexibly switch between representations [3,4]. It also accounts for how nonlinear mixed selectivity of neurons arises from the interaction of cognitive control and stimulus features [5–8]. By dynamically modulating local neuronal excitability, alpha/beta oscillatory patterns allow the same neural populations to support different cognitive operations depending on context.

Despite these advances, several critical questions about Spatial Computing remain unanswered. First, the theory proposes that different neural signals serve distinct roles. Oscillatory patterns are proposed to reflect information learned to complete any given goal-directed task, which we refer to as "task information". In contrast, spiking activity primarily represents bottom-up sensory inputs, the "*contents*" of cognition. However, little work has been done to directly compare the information carried by each signal type at the *population* level. Second, although Spatial Computing proposes that task information is reflected in alpha/beta oscillations, it is still unknown whether the spatial patterns of alpha/beta power reorganize with differing task conditions. It is also unknown whether any such reorganization between task conditions scales with cognitive demands, i.e. stronger cognitive demands result in larger task-based reorganization. Third, while the theory posits that alpha/beta regulates where sensory information is expressed, no direct test has shown that spiking-based sensory representation is spatially gated by alpha/beta patterns across the cortical space. Finally, although Spatial Computing is hypothesized to support flexible cognition and behavior, there is little direct

79 evidence that alpha/beta or gamma (30–150 Hz) oscillations predict trial-by-trial
80 decisions or reflect behavioral errors. Together, these open questions point to the need
81 for a more comprehensive empirical test of Spatial Computing Theory.
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Figure 1. Schematic overview of Spatial Computing and task designs.
(A). Conceptual framework of Spatial Computing Theory. Alpha/beta oscillations convey
cognitive control signals reflecting learned task information, and their spatial pattern gates
where sensory information ("content") can be expressed by spiking. Shaded grey areas indicate
stronger alpha/beta power where sensory information is suppressed while unshaded areas
indicate lower alpha/beta power where sensory information is expressed by spiking.
(B). "Recognition" rule version of Seq-image Task showing a sequence of sample images
followed by a delay. During the test phase, NHPs judged whether a probe sequence matched
the sample sequence.
(C). "Recall" rule version of Seq-image Task where, during the test phase, NHPs made
saccades to the remembered objects in their original order. The two rule versions alternated in
trial blocks (right).
(D). Single-image task where a single image was shown and, after a delay, NHPs made a
saccade to the remembered object.
(E). Category-match task where NHPs categorized exemplar dot patterns into two learned
categories.
(F). The exemplar dot patterns were generated by distorting two prototype patterns (left).
Distortion levels varied the similarity between exemplars and prototypes, controlling the degree
of abstractness (low vs. high).
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We analyzed multi-electrode recordings from the lateral prefrontal cortex as non-human
primates (NHPs) performed various cognitive tasks requiring working memory (WM),
sequencing, and/or categorization. We formalized five core predictions of Spatial
Computing and tested them across different cognitive demands. Our results show that
alpha/beta patterns encode task information, flexibly reorganize with task conditions,

spatially govern where sensory information is expressed in spiking activity, and closely
track behavioral performance. These results provide empirical support for Spatial
Computing Theory. They highlight its potential as a unifying framework for
understanding how the brain flexibly coordinates cognition through structured
population dynamics.
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Results
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Predictions of Spatial Computing Theory and Tasks
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We tested five basic predictions of *Spatial Computing Theory*: 1. Alpha/beta oscillations
primarily convey information reflecting cognitive control and task. 2. Spiking activity
conveys both sensory (e.g., stimulus identity) and task information. 3. Alpha/beta power
is spatially organized by task condition across the surface of cortex. 4. Sensory
information is spatially anti-correlated with shifting alpha/beta patterns reflecting task
information. 5. Alpha/beta oscillations reflect NHPs' behavioral performance and trial-
by-trial choices. As downstream effectors of cognitive control and structured by
alpha/beta, spiking and associated gamma are expected to correlate with behavior as
well. **Figure 1A** illustrates this framework, where alpha/beta patterns carry task
information and are dynamically shaped by the current task condition. These alpha/beta
patterns act as "inhibitory stencils" which constrain where higher levels of spiking occur.
This allows the stimulus-driven sensory information to be expressed only in regions
where alpha/beta is weak. In essence, cortical space becomes a computational
dimension. Alpha/beta patterns sculpt the spatial landscape of neural activity, thereby
supporting context-dependent information processing.
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To test these predictions, we analyzed multi-electrode neurophysiological recordings
(local field potentials (LFPs) and single/multi-unit spiking activity) from the lateral
prefrontal cortex (PFC) of NHPs (rhesus macaques) performing three cognitive tasks
(**Figure S1**). In one of the tasks, the Seq-image Task [9], the NHPs performed a
sequential working memory (WM) task involving two sample items. Two versions of the
task were used. Both required working memory for two sample items and their order of
appearance. They only differed in the rule that determined how NHPs reported these
memories. In the 'Recognition' version, the NHPs indicated whether a subsequent
sequence of test items was a match to the sample sequence (**Figure 1B**). In the 'Recall'
version, NHPs were offered three test items and asked to saccade to the two seen
earlier in the correct order (**Figure 1C**). The NHPs alternated between the two versions
across blocks of trials. Thus, NHPs needed to maintain two types of task information:
the order in which the two samples were presented on each trial, and the current rule
that needed to be applied to the stored samples and test items.
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The two other tasks involved working memory and categorization of individual images.
The Single-image task [10,11] was a standard delayed match-to-sample working memory
task (**Figure 1D**). In the Category Task [12,13], NHPs had to categorize dot-pattern stimuli
into two learned categories, each defined by random distortions around a prototype
(**Figure 1E1**). The level of distortion ("abstractness") was varied by manipulating the
similarity between the exemplars and their corresponding prototype (**Figure 1E2**). Thus,
in this task, the individual dot exemplars were the sensory inputs to cognitive
processing. Their inferred category was the learned task condition that determined

which behavioral response was appropriate for each exemplar. Additionally, manipulating the level of exemplar distortion across trials allowed us to test different degrees of cognitive demands.

These three tasks allowed us to probe how sensory information (e.g., identity of sample items and exemplar dot patterns) and task information (e.g., rule, order, and category) are encoded by spiking and LFP dynamics. By examining the spatial distribution of these signals, we could directly test whether alpha/beta patterns are spatially organized, consistent with a "stencil" that organizes sensory information expressed by spiking. A complete description of the experimental design for each task can be found in the **Methods**. Trial-averaged power spectrograms for each dataset are provided in **Figure S2**, with additional figures and analyses available in the original publications that first reported these data [1,9–14]. The spectrograms showed clear spectral peaks within the 10–30 Hz band, indicating our results reflect true alpha/beta oscillations, rather than broadband aperiodic effects [15].

Figure 2. Alpha/beta power primarily carried task information, while spiking activity conveyed mixed sensory and task information.

(A). Sensory information (quantified by PEV) conveyed by each signal type in the Seq-image Task, for sample 1 (left) and sample 2 (right). Vertical dashed lines indicated the sample presentation periods (S1 and S2). Horizontal thin lines on top reflect times of significant PEV values for the corresponding signal type using the same color code (p<0.001, one-sample t-test). Error bar: mean ± SE, n = 48 (2 rules × 24 sessions).

(B). Task information conveyed by each signal type in the Seq-image Task. Left: PEV for rule (Recognition vs. Recall) during the second delay and test period. Right: PEV for sequence order (first vs. second order) during the sample presentation followed by delay. Error bar: mean ± SE, n = 24 sessions. Significance: p<0.001, one-sample t-test.

(C). Sensory information for each signal type in the Category Task. Left: PEV for dlPFC. Right: PEV for vlPFC. Error bar: mean ± SE, n = 51 sessions. Significance: p<0.001, one-sample t-test.

(D). Category information for each signal type in the Category Task. Left: PEV for dlPFC. Right: PEV for vlPFC. Error bar: mean ± SE, n = 51 sessions. Significance: p<0.001, one-sample t-test.

See Fig. S2 for corresponding paired comparisons between signal types.

Predictions 1 and 2: Alpha/beta Reflects Task Information, Spiking Reflects Both Sensory and Task Information

To test these predictions, we quantified the amount of information reflecting the sensory stimuli and task conditions carried by each signal type. Rather than analyzing single electrodes or individual units, we developed a method to compute the percentage of explained variance (PEV) by each factor across the entire population of units or LFPs (see **Methods**). For each factor, we computed an axis in population activity (spike rates or LFP band power) space that optimally separated contrasted conditions. We then projected population activity onto this axis and quantified its PEV. This approach allowed us to capture the information encoded by population-level LFP power and spiking.

We began with the sensory information in the Seq-image Task—the identity of the two sample items to be remembered (**Figure 1B**). In line with previous studies [16], we found that spiking activity carried the most sensory information. Information increased sharply during each sample presentation (**Figure 2A**). There was some sensory information in gamma, but it was much lower than in spiking activity (**Figure S3**). Sensory information was minimal in alpha/beta power, in line with Prediction 1.

We then assessed task condition encoding by measuring task information conveyed about the *rule* being performed (i.e., Recognition vs. Recall) during the delay period preceding test onset. Throughout the delay, all signals carried task information reflecting the current rule but alpha/beta and spiking carried more than gamma (**Figure 2B**, left; see **Figure S3** for paired comparison). Note the ramping up of alpha/beta information reflecting the rule. It peaked near the end of the delay and plateaued through the presentation of the test items when the rule needed to be applied. In addition to the rule, we also examined the sample sequence order (i.e., first vs. second sample item), which was another type of task factor in this paradigm. We computed information reflecting the sequence presentation order (i.e., sample/delay 1 vs. sample/delay2), while holding constant the sample image shown (see **Methods**). Similar to information reflecting the rule, all signals carried sequence order information, with alpha/beta and spiking carrying more than gamma (**Figure 2B**, right; see **Figure S3** for paired comparison). These results showed that, in the Seq-image Task, task information was reflected mainly in alpha/beta power and spiking, while WM content was reflected primarily in spiking activity, consistent with Predictions 1 and 2.

We then turned to the Category Task. The sensory stimuli in this task were the individual category exemplars, while their inferred category memberships were learned task conditions that determined the appropriate behavior for each exemplar. For this dataset, we have extensive sampling of both the dorsolateral (dlPFC) and ventrolateral (vlPFC) prefrontal subregions. In both subregions, spiking activity dominated sensory information, while alpha/beta and gamma in vlPFC also carried a modest amount (**Figure 2C**, see **Figure S3** for paired comparison). By contrast, for category information, we observed a clear distinction between the two regions, as we discuss next.

Previous work has shown that the dlPFC is more associated with top-down processing,
while vlPFC is relatively more driven by bottom-up sensory stimulation [13,17–19].
Consistent with this, alpha/beta oscillations in dlPFC were the primary carriers of
category information (**Figure 2D**, left; see **Figure S3** for paired comparison). By
contrast, in vlPFC, spiking activity dominated (**Figure 2D**, right; see **Figure S2** for paired
comparison). Importantly, these differences between dlPFC and vlPFC were not due to
differences in overall spiking rate or recording quality between these two regions.
Although vlPFC showed a brief (~100 ms) elevation in mean spiking rate relative to
dlPFC (**Figure S4**), their spiking rates were nearly identical through the rest of the trial.
Moreover, dlPFC and vlPFC carried comparable amounts of information reflecting the
behavioral response (choice of left vs. right test exemplar; **Figure S4**). This indicated
that the observed differences between dlPFC and vlPFC arose from their functional
distinctions. In dlPFC alpha/beta mainly reflected category, while in vlPFC spiking
conveyed both category and sensory inputs more prominently. This
dorsolateral/ventrolateral distinction suggested a gradient in the neural implementation
of top-down versus bottom-up processing, which we return to in the Discussion.
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For a further test of category, we examined the level of abstractness of the categories.
The Category Task manipulated the level of abstractness by different degrees of
distortions of the prototypes. For low distortions, the exemplars from the same category
were more similar in appearance to one another and to their prototypes. Thus, they
could be categorized by their bottom-up appearance, and required less abstraction. By
contrast, at high distortion levels, exemplars from different categories could appear
similar to each other. They thus had to be categorized by a task rule learned by the
NHPs [13].
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Figure 3. Category encoding varies with abstractness level.
(A). Category information (PEV) by alpha/beta (left), gamma (middle), and spiking (right) in dlPFC, shown separately for high- (solid lines) and low-abstractness (dashed lines) conditions. Horizontal grey bars on the x-axis represented significant PEV differences between high- and low-abstractness (p<0.05, two-sample t-test). Error bar: mean ± SE, n = 51 sessions.
(B). Cumulative difference in category PEV (high- minus low-abstractness) across time in dlPFC, shown for each signal type. *** p<0.001, **: p<0.01, n.s.: not significant, one-sample t-

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This revealed category information in alpha/beta was stronger for the high-abstractness
exemplars (especially in dlPFC). By contrast, category information in spiking was
stronger for the low-abstractness exemplars (in vlPFC). In dlPFC, alpha/beta (**Figure
3A,B**, left) and gamma (**Figure 3A,B**, center) power encoded significantly more
category information under high-abstractness than under low-abstractness. These
effects were greater than the modulation of dlPFC spiking activity by abstractness,
which was itself not significant (**Figure 3A,B**, right). In vlPFC, by contrast, spiking
activity carried more category information under low-abstractness conditions than under
high abstractness (**Figure 3C,D**, right), while alpha/beta and gamma showed the
opposite direction of modulation by abstractness level (**Figure 3C,D**, left and center).
Thus, when exemplars were heavily distorted (high-abstractness), dlPFC alpha/beta
plays a dominant role in representing the relevant category through top-down
mechanisms. Conversely, when exemplars were easily recognized (low-abstractness),
vlPFC spiking could represent the category more directly through bottom-up
mechanisms. This is consistent with Predictions 1 and 2.
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Overall, these findings support Predictions 1 and 2 of Spatial Computing. They showed
that alpha/beta (especially in dlPFC) strongly reflected learned task information, while
spiking reflected both sensory and task information. Task information carried by
alpha/beta (in dlPFC) was greater when tasks required greater cognitive demands. The
distinction between dlPFC and vlPFC aligned with their respective bias toward top-down
vs. bottom-up processing.
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**Prediction 3: Alpha/beta Power is Spatially Organized by Task Condition Across
the Surface of Cortex**
Spatial Computing theory further predicts that alpha/beta oscillations form spatial
patterns across the cortical surface that reflect task information and reorganize with
changing task conditions. A critical component of this prediction is that, for each distinct
task condition, alpha/beta power should exhibit spatial clustering, rather than being
spatially unstructured. Because the Category Task and Single-image Task used multi-
electrode Utah arrays (see **Methods**), we could test this prediction directly with these
datasets.
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To quantify the spatial clustering of alpha/beta power within each array, we computed
Moran's I [20–22], a common measure of the overall strength of spatial autocorrelation (see
Methods). We compared the observed values to those obtained from spatially shuffled
versions of the same patterns. If alpha/beta power is spatially clustered, its spatial
autocorrelation should be significantly positive. By contrast, the shuffled patterns,
lacking spatial structure, should yield autocorrelation values near zero.
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We found that in both dlPFC and vlPFC, Moran's I values for the original alpha/beta
patterns were significantly greater than those of the shuffled patterns, which were

around zero (**Figure 4A,B**, left). This effect held for both low and high levels of
abstractness of the Category Task, as well as for the Single-Image Task (**Figure S5A,B**,
left). We confirmed this spatial clustering with an alternative metric by computing the
absolute power differences between neighboring electrodes (see **Methods**). This
yielded consistent results: shuffled patterns exhibited significantly larger power
differences between nearest neighbors compared to the actual observed alpha/beta
patterns (**Figure S5A,B**, right; **S5C,D**). These results indicated that alpha/beta power
was spatially organized across the cortical surface, as predicted by Spatial Computing
Theory.
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Figure 4. Alpha/beta power is spatially organized across the surface of cortex
(A). (Left). Spatial autocorrelation, measured by Moran's I, of the trial-averaged patterns of
alpha/beta power during sample presentation in dlPFC (solid bars), for low and high
abstractness of the Category Task. As a control, Moran's I for patterns randomly shuffled across
space is shown (hatched bars). (Right). Same as left but for alpha/beta power differences
*between category-1 and category-2. *** p<0.001, two-sample t-test. Error bar: mean ± SE, n =*
51 sessions.
(B). Same as (A) but for vlPFC.
(C). Example spatial patterns of alpha/beta power averaged across trials, shown for two
categories under low- and high-abstractness in one session from dlPFC.
(D). Same as (C) but for vlPFC.
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Spatial Computing theory predicts spatial organization not only of overall alpha/beta
power, but also of the task information it carries. That is, alpha/beta at nearby sites
should "prefer" the same category. To quantify this, we computed the signed power
difference between categories (Category-1 – Category-2) at each electrode. This is a
measure of category selectivity that reflects both its category preference (sign) and
strength (magnitude). We then quantified their spatial autocorrelation with Moran's I and
compared them against spatially shuffled controls. These category selectivity maps
showed significantly positive spatial autocorrelation, whereas shuffled maps were near
zero (**Figure 4A,B**, right). These results reflected contiguous clusters of electrodes that
consistently favor one category over the other. Thus, alpha/beta power formed patterns
that were spatially organized and category-selective. Moreover, further spatial
correlation analysis confirmed that these patterns reconfigured with category (i.e., lower
between-category correlation than within-category, **Figure S5E-H**). Example spatial

maps in **Figure 4C** and **4D** illustrated this spatially structured organization of alpha/beta patterns with increasing category differentiation under high-abstractness in dlPFC and vlPFC respectively. These results are consistent with Prediction 3.

Prediction 4: Sensory Information is Spatially Anti-correlated with Shifting Alpha/Beta Patterns Reflecting Task Information

Alpha/beta patterns are proposed to spatially regulate the expression of sensory information carried by spiking. This spatial regulation is expected to shift with changing task conditions. **Figure 5A** shows a schematic of this hypothesis: In each task condition, areas of high alpha/beta power exhibit lower sensory information carried by spiking, whereas areas with low alpha/beta power show higher spiking sensory information. Moreover, the spatial pattern of alpha/beta, and thus the location of spiking sensory expression, varies across task conditions.

First, we looked for evidence consistent with alpha/beta regulating where spiking sensory information is expressed. We analyzed data from the Single-image Task (**Figure 1D**), where multi-electrode Utah arrays were implanted in both dlPFC and vlPFC. We grouped LFPs/spiking by their positions in the electrode array. For each electrode, we computed the time-averaged alpha/beta power during sample presentation. We then split electrodes into two groups—those with higher or lower alpha/beta power across electrodes. Finally, we computed PEV reflecting sample object identity separately in spiking and gamma for each single electrode in order to construct spatial maps of sensory information (see **Methods**).

Figure 5. Alpha/beta patterns are spatially anti-correlated with the distribution of sensory information in cortical space.

(A). Schematic illustration of spatial regulation: in each task condition, regions with higher alpha/beta power suppress local spiking, thereby reducing sensory encoding and information.

(B). Sensory information (PEV for stimulus identity) carried by spiking activity at low- and high-alpha/beta electrodes in dlPFC (top) and vlPFC (bottom). Horizontal grey bars on the x-axis represented significant PEV differences between low- and high-alpha/beta power (p<0.05, two-sample t-test). Error bar: mean ± SE. Sample sizes (low/high alpha-beta): dlPFC spiking (n = 650/427), vlPFC spiking (n = 198/262).

(C). Same as (B) but for sensory PEV carried by gamma at low- and high-alpha/beta electrodes in dlPFC (top) and vlPFC (bottom). Error bar: mean ± SE. Sample sizes (low/high alpha-beta): dlPFC gamma (n = 672/730), vlPFC gamma (n = 365/909).

(D). Example spatial patterns from a single session in dlPFC (top) and vlPFC (bottom) showing time-averaged alpha/beta power, sensory PEV by gamma, and sensory PEV by spiking during the sample presentation period. Each pattern was normalized individually.

(E). Correlation between spatial patterns of alpha/beta power and sensory PEV of gamma and spiking in dlPFC (top) and vlPFC (bottom). *** p<0.001, *: p<0.05, n.s.: not significant, one-sample t-test. Error bar: mean ± SE. Sample sizes: dlPFC (gamma: n = 22, spiking: n = 15), vlPFC (gamma: n = 20, spiking: n = 18).

(F). Correlation between spatial patterns of alpha/beta power and sensory PEV of spiking for the same category (within-category) and different categories (between-category), for dlPFC (top) and vlPFC (bottom). *** p<0.001, *: p<0.05, n.s.: not significant, one-sample and two-sample t-test. Error bar: mean ± SE, n = 51 sessions.

As predicted, spiking activity on electrodes with lower alpha/beta power exhibited significantly greater sensory information than spiking activity on electrodes with higher power (**Figure 5B**). This differentiation was significant but weaker in gamma (**Figure 5C**). To visualize these spatial relationships, **Figure 5D** showed example maps of the alpha/beta power, sensory information in gamma, and sensory information in spiking from a representative session. It showed anti-correlation between the spatial patterns of alpha/beta power and of sensory information in spiking. Across all sessions, we found a significant negative spatial correlation between alpha/beta power and sensory information in spiking (**Figure 5E**). This supports the idea that alpha/beta regulates the spatial distribution of sensory information in spiking.

Next, we assessed how this spatial anti-correlation shifted with changing task conditions, using data from the Category Task. Because alpha/beta patterns flexibly reorganize with category (**Figure S5E-H**), the spatial pattern of spiking sensory information should also reorganize accordingly. To examine this, we computed spatial correlations between alpha/beta power and spiking sensory PEV within the same category, and between different categories. Within the same category, spiking sensory PEV remained negatively correlated with the corresponding alpha/beta pattern (**Figure 5F**). By contrast, when comparing spiking sensory PEV from one category to alpha/beta patterns from the other, this anti-correlation weakened significantly and became more uncorrelated (**Figure 5F**). This suggested that the expression of spiking sensory information *remapped* with the spatial reorganization of alpha/beta power. It did so in a way that the alpha/beta pattern in one category better predicted the spatial distribution

427 of spiking sensory information within the same category than in the other. Together,
428 these results support Prediction 4.
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Prediction 5: Alpha/Beta Correlates with Behavior
431 If alpha/beta truly encodes task information in a way that influences the NHPs'
432 decisions, we would expect these signals also to differ between correct and incorrect
433 trials and to correlate with trial-by-trial decisions.
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Figure 6. Alpha/beta and spiking reflect correct versus incorrect decisions
about sequential order, but not about stimulus identity.
(A). Task information (PEV) about sample order in Seq-Image task for correct trials (solid lines)
and order-error trials (dashed lines), where NHPs chose the correct samples but in reversed
order. Results are shown separately for alpha/beta (left), gamma (middle), and spiking (right).
Horizontal grey bars on the x-axis indicate significant PEV differences between correct and error
trials (p<0.05, two-sample t-test). Error bar: mean ± SE, n = 24 sessions.
(B). Cumulative difference in order information (correct minus incorrect) across time for each
signal type. *** p<0.001, **: p<0.01, *: p<0.05, n.s.: not significant, one-sample t-test for each
signal type and Wilcoxon rank-sum test for across-signal comparisons. Error bar: mean ± SE,
n = 24 sessions.
(C). Same as in (A) but for sensory information about stimulus identity.
(D). Same as in (B) but for sensory information.
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We first tested this in the Recall version of Seq-image Task, where NHPs had to report
the two sample items in the correct order. This design allowed us to isolate task-related
errors. For many of the incorrect trials, NHPs chose the correct samples but in the wrong
order, indicating an error of reversed order. Thus, we compared the neural information
reflecting sample order between correct and order-error trials. We found that all signals
carried significantly more order information in correct trials than in order-error trials
(**Figure 6A**). This behavioral modulation was greatest for alpha/beta and spiking
(**Figure 6B**). By contrast, information about sensory stimuli (i.e., sample identity) did not
differ between correct and order-error trials for all signals (**Figure 6C** and **6D**). Similar
results were also found for the Category Task by comparing correct vs. incorrect
categorization trials (**Figure S6** and **S7**). This suggested that alpha/beta and spiking

were crucial for maintaining the correct sequence order and category that influenced behavioral decisions.

We next turned to the Category Task to determine whether these signals correlate with NHP's trial-by-trial decisions of category membership. We adapted a trial-by-trial analysis similar to Cohen & Maunsell, 2010 [23]. Specifically, we projected each trial's neural population data (time-averaged over sample presentation) onto an axis in population activity space that optimally separates the two learned categories on correct trials (see **Methods**). We then normalized projections onto this "category axis" so that a value of –1 corresponded to the average for Category-1 correct decisions and +1 to the average for Category-2 correct decisions. This provided a single-trial metric of population category coding strength. We then projected onto this axis the activity from the incorrect trials, which were not used to define the axis. If alpha/beta patterns are indeed relevant for categorization, we would expect the projections of the incorrect trials to be "shifted" toward the opposite, incorrect, category.

Figure 7. Alpha/beta patterns reflect category errors and predict trial-by-trial decisions.

(A). Histogram of single-trial projections of dlPFC alpha/beta activity onto the category axis for correct (upward bars) and incorrect (downward bars) trials in an example session, pooled across low- and high-abstractness conditions.

(B). Deviation (|Δ|) between mean projections of correct and incorrect trials under low- and high-abstractness conditions, shown for alpha/beta, gamma, and spiking in dlPFC (left) and vlPFC (right). *** p<0.001, n.s.: not significant, one-sample t-test for each signal type and Wilcoxon rank-sum test for across signals. Error bar: mean ± SE, n = 51.

(C). Correct choice probability as a function of projection values along the category axis for alpha/beta (left), gamma (middle), and spiking (right) in dlPFC. Solid line: trials where Category-1 is the correct category; dashed line: Category-2 is the correct category. Grey bars on the x-axis represent significant differences between the choice probability of category-1 and category-2 (p<0.001, Wilcoxon rank-sum test). Error bar: mean ± SE, n = 51.

(D). Same as (C) but for vlPFC.

We found the predicted shift in alpha/beta for error trials in which the incorrect category
was chosen. **Figure 7A** showed an example session in which alpha/beta projections for
incorrect decisions on Category-1 trials shifted toward the Category-2 distribution, and
vice versa. We quantified this effect by computing $|\Delta|$, the absolute deviation between
the mean projections of correct and incorrect category decisions (see **Methods**). Larger
$|\Delta|$ values reflect a stronger shift in population activity on error trials, on average. We
applied the same analysis to gamma and spiking (**Figure 7B**). We found the average
$|\Delta|$ deviations across sessions differed significantly from zero (one-sample t-test,
$p<0.001$) for all signals under both low- and high-abstractness conditions. However,
alpha/beta projections showed the greatest deviations $|\Delta|$ under high-abstractness
conditions as compared to gamma and spiking, in both dlPFC and vlPFC.

To further characterize the direction of the shift, we analyzed the signed projections of
incorrect trials. The shift toward the opposite category direction should reveal
projections that consistently reversed sign relative to the correct category. We found
this robust effect in alpha/beta projections across all conditions. In both dlPFC and
vlPFC, for both low- and high-abstractness levels (**Figure S8**, left). Gamma also
exhibited significant sign reversals, but only under high-abstractness conditions in both
regions (**Figure S8**, middle). Spiking showed a similar pattern, but only in dlPFC and
only for high-abstractness (**Figure S8**, right). These results suggested that while all
signals reflect categorization errors to some extent, the categorization errors were most
consistently and strongly associated with a systematic shift in alpha/beta toward
patterns reflecting the incorrect category. This underscored the critical role of alpha/beta
in representing task conditions used to guide decisions.

We also found a strong correlation with behavioral performance on a trial-by-trial basis.
Figure 7C and **7D** show the proportion of correct category decisions as a function of
the single-trial population activity projections. Trials with projections closer to the ideal
value for the given category (*x*-axis values of –1 for Category-1, +1 for Category-2)
resulted in the highest probabilities of correct decisions. Trials with projections toward
the opposite end of the category axis (*x*-axis values of +1 for Category-1, –1 for
Category-2) resulted in more incorrect decisions. In between these values, choice
probability varied in a graded manner. The slopes of the choice-probability curves were
similar across all signal types. This likely reflected the fact that, although alpha/beta
showed a larger mean separation between correct and incorrect decisions, it also had
greater trial-to-trial variability. In contrast, spiking and gamma showed smaller mean
shifts but also lower variance (**Figure S9**), resulting in comparable overall predictive
performance across signals.

All together, these results support Prediction 5. Alpha/beta patterns reflected sequence
order errors. They reflected category errors. They reflected the NHPs' trial-by-trial
category decisions. We also observed that spiking and gamma signals correlated with
behavioral performance. This is consistent with the theory: if alpha/beta oscillations
regulate where sensory-related spiking can emerge, then behaviorally relevant
information should also be reflected in spiking and associated gamma activity.

Discussion

Our results provide evidence for the central predictions of Spatial Computing Theory across multiple cognitive tasks. Alpha/beta power conveyed minimal sensory information (i.e., the items held in WM) but robustly carried information about task conditions (i.e., task rules, order, and categories). Alpha/beta power carried stronger task information with greater cognitive demands. By contrast, neuronal spiking activity encoded both task and sensory information (**Predictions 1 and 2**). Alpha/beta power was spatially organized into patterns reflecting task conditions (**Prediction 3**). Sensory information carried by spiking was spatially anti-correlated with alpha/beta patterns. Recording sites with lower alpha/beta power had significantly stronger sensory information. This spatial anti-correlation suggested alpha/beta's role in spatially organizing information processing in the cortex (**Prediction 4)**. Lastly, alpha/beta power closely tracked NHPs' behavioral performance. Alpha/beta power reflected misattributions of the current task condition, whereas sensory information remained unaffected. Alpha/beta patterns also correlated with NHPs' trial-by-trial decisions about task conditions, especially under higher cognitive demands (**Prediction 5**). Collectively, these findings validate predictions of Spatial Computing Theory [1], demonstrating that alpha/beta oscillations organize cortical activity spatially to support context-dependent cognitive processing.

More broadly, our results are consistent with prior work showing that alpha/beta-band power and coherence in PFC reflects task rules and rule shifts [24,25]. Our observations of increased context information in alpha/beta with increased cognitive demands parallels results from Wutz, et al., 2018 [13], who reported increased dlPFC alpha/beta power under similar conditions. Our findings go further by directly quantifying the amount of context information in alpha/beta and showing how it varies with task demands. Alpha/beta effects have also been observed across different task periods of simpler working memory tasks [16,26–28], putatively reflecting their differing requirements for task-related usage of information in working memory (i.e. encoding vs. maintenance vs. readout from working memory). These findings support the broader idea that low-frequency oscillations flexibly adapt to cognitive demands and help coordinate cortical processing.

In addition to these functional distinctions between signal types (alpha/beta vs gamma and spiking), we also observed overlaid broad differences between PFC subregions. We found stronger overall alpha/beta and stronger correlations with cognitive factors in dlPFC. We observed stronger spiking activity and stronger correlations with bottom-up sensory information in vlPFC. This is consistent with prior studies showing that dlPFC is more involved in maintaining abstract rules and goals, whereas vlPFC is more engaged in stimulus-driven encoding and maintenance of object features [13,17–19]. It's also consistent with the stronger connectivity between sensory cortex and vlPFC [29,30], and the stronger connectivity between dlPFC and areas involved in decision-making and motor output [31,32]. Thus, these two areas may reflect a functional gradient, where top-down cognitive functions (and their associated signals) are more strongly

represented in dlPFC and sensory-related functions (and their associated signals) are more strongly represented in vlPFC.

Previous studies have linked alpha/beta oscillations with behavioral performance [33–35], as we have in this study. We found that behavioral correlates in the alpha/beta signals were comparably strong to those observed in spiking activity. This suggests that the spatial patterns observed in alpha/beta oscillations are not merely reflective of internal state, but likely play an active role in shaping the neural processes that lead to behavior. They actively structure the neural landscape in ways that influence downstream decision-making.

Our findings align closely with conceptual advances from human EEG and MEG research. In humans, alpha oscillations have been widely implicated in the allocation of neural resources via the inhibition of task-irrelevant brain areas, particularly under top-down control [36–38]. While most human studies focus on posterior cortical areas, modulations of frontal alpha activity have also been reported during goal-directed tasks, including language and working memory [39,40]. Our data extend this framework by showing in non-human primates that alpha/beta oscillations in prefrontal cortex not only reflect top-down control but also shape where spiking-based sensory representations emerge. Our results further show that this control can be exerted in spatial patterns *within* a cortical area, and that these patterns show correlations with behavior. By linking alpha/beta patterns to both local excitability and behavioral performance, our results bridge a critical gap between non-human primate and human research. This suggests that low-frequency oscillations implement a shared control strategy across species and cortical areas.

Our results support the view that alpha/beta oscillations exert inhibitory control over local neural populations [41,42]. Prior studies have shown that alpha/beta power decreases in cortical areas engaged by a task and increases in disengaged areas [43–45]. In macaque cortex, increased alpha/beta power has been associated with reduced local spiking activity, suggesting a suppressive gating function [28,35,46–48]. Consistent with this, we observed a strong inverse relationship between alpha/beta power and content-related spiking. This correlational relationship supports an inhibitory role for alpha/beta. Future work, including causal analyses and modeling, will be needed to clarify the directionality of this relationship and provide a mechanistic account of how alpha/beta might regulate spiking activity.

One possible neural circuit mechanism underlying this spatial regulation involves local inhibitory interneurons. Recent evidence, both experimental [49–52] and computational [53–55], suggests that alpha/beta oscillations may be generated by somatostatin-expressing (SST) interneurons. These interneurons are well-positioned to mediate top-down modulation [56,57]. Top-down inputs could locally recruit SST cells, leading to an increase in the local alpha/beta power and targeted suppression of nearby excitatory neurons and spiking. In this way, the spatial structure of alpha/beta activity could emerge from selective activation of inhibitory circuits, which in turn sculpt where content-related spiking is permitted or suppressed.

Our findings build on prior work showing that working memory is implemented through
interactions between alpha/beta oscillations and brief spiking/gamma bursts [16,28]. More
recently, Lundqvist et al. (2023) [1] formalized these ideas into Spatial Computing theory,
proposing that cortical space is used as a computational dimension for cognitive control.
They showed that low-dimensional oscillatory activity carries control-related signals.
While that study established key principles of Spatial Computing, several important
questions remained open. This includes how spatial patterns of neural activity
reorganize with context and scale with cognitive demand, and whether these dynamics
relate to behavior. Our results extend this framework by providing direct empirical tests
of these predictions. Our results suggest that alpha/beta oscillations not only gate
information in time but also sculpt the spatial landscape of cortical activity, determining
where spiking-based sensory representations emerge. This offers strong support for the
idea that cortical space functions as an active computational resource, with alpha/beta
rhythms regulating both the location and timing of information expression.

This spatial regulation offers a mechanistic basis for how the brain flexibly reconfigures
its representational geometry with task demands. This is closely aligned with theories
of subspace coding [3,4] and nonlinear mixed selectivity [5-8]. Subspace coding proposes
that different task contexts activate orthogonal patterns of population activity, minimizing
interference between representations. Mixed selectivity refers to neurons that respond
to specific combinations of stimuli and task context. It is essential for flexible cognition[6].
While these properties have been observed across the brain, the underlying mechanism
has remained unclear. Our results suggest that alpha/beta oscillations provide that
mechanism by dynamically regulating where in cortical space certain subspaces and
selective combinations can emerge. By spatially suppressing spiking activity, alpha/beta
patterns may reduce the overlap between neuronal populations recruited in different
contexts, leading to more orthogonal population subspaces. This selective recruitment
could also support the emergence of nonlinear mixed selectivity, as neurons participate
in different combinations depending on where alpha/beta inhibition is released. In doing
so, Spatial Computing explains how a single cortical circuit can express a wide range
of task-dependent representations without structural rewiring.

Gamma is typically associated with bottom-up sensory processing and closely tied to
local spiking activity [58-60]. We found gamma carried some sensory information, though
to a much lesser extent than spiking. This likely reflects its lower spatial and temporal
resolution relative to spiking activity. Gamma also carried some task information and
was correlated with behavioral choices. These findings suggest that while gamma is not
traditionally considered a carrier of top-down signals, it may inherit context-related
structure through its close coupling with spiking activity that is spatially regulated by
alpha/beta oscillations. In this sense, gamma may act as an intermediary, reflecting both
the bottom-up content and the spatial constraints imposed by top-down control. Prior
work has shown that gamma synchronization is modulated by attention and correlates
with behavioral measures, including perceptual accuracy and reaction time [61,62].

Taken together, our results highlight Spatial Computing as a framework for
understanding how the brain flexibly controls cognitive processes. By showing that
alpha/beta oscillations dynamically shape when and where content can be expressed,
we report a spatial mechanism for cognitive control. This shifts the focus from fixed
anatomical circuits to dynamic, spatially structured population activity as the substrate
of cognition. Spatial Computing offers a unifying principle that explains how diverse
cognitive operations, such as working memory, rule application, categorization, and
decision-making, can emerge from a shared circuitry guided by oscillatory control. More
broadly, it suggests that cortical space itself is a computational dimension, repurposed
in real time to serve changing behavioral goals. This view opens new avenues for
understanding the flexible and adaptive nature of neural computation, and for building
models of cognition grounded in the analog dynamics of brain activity.

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Data and materials availability: All data needed to evaluate the conclusions in the paper are available upon reasonable request from the Corresponding Author.

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Methods

Non-human primate (NHP) subjects

We analyzed data from three previous studies. The Seq-image Task was from [9,14], the Single-image Task was from [1,10,11], and the Category Task was from [12,13,63–65]. For details of experimental design and data collection, please see those publications.

Briefly, each study involved two Rhesus macaques (*Macaca mulatta*) that were trained on a behavioral task using positive rewards (juice), until their performance indicated a solid understanding of all the task requirements. The Seq-image Task included NHPs S (female, 6 years, 6.5 kg) and A (male, 4 years, 10 kg). The Single-image Task included NHPs T (male, 17 years, 13 kg) and D (male, 8 years, 10 kg). The Category Task included NHPs P (female, 8 years, 9 kg) and G (male, 9 years, 13 kg). All NHPs were experimentally naive prior to the experiments analyzed here, except for NHP T, who had previously participated in studies of categorization [63] and working memory [65]; NHP P, who had participated in a previous study of category learning [66]; and NHP G, who had participated in a study of rule-based categorization [64]. NHPs were pair-housed, on a 12-hr day/night cycle, and in a temperature-controlled environment (80° F). All procedures followed the guidelines of the Massachusetts Institute of Technology Committee on Animal Care and the National Institutes of Health.

Datasets analyzed

For the Seq-image Task, NHPs were trained on two versions of a sequence working memory task. Blocks of each task version—Recognition and Recall—were interleaved, 100–250 trials per block. In both versions, a sequence of two sample images was presented (500 ms each, 1 s blank delay between them). The identity and order of the images then had to be maintained in working memory over a brief delay (1 s). In the Recognition version, a test sequence of two images was then presented, which required a bar-release response during the presentation of the second image if the test images matched the samples in both identity and order. Otherwise, subjects withheld their response until a subsequent matching sequence was shown. In the Recall version, an array of three test images was shown simultaneously, and the subject was required to saccade to the two images that matched the samples in the same order as initially presented. There was no explicit cue to indicate that the task had switched, as it was obvious from context during the first trial after the block switched. NHPs typically performed 2–4 blocks of each task during each recording session. Performance was well above chance on both versions of the task. An infrared-based eye-tracking system (ISCAN) continuously monitored eye position at 60 kHz. In each session, eight tungsten microelectrodes (FHC) were lowered through custom grids using custom-made independently moveable microdrives (Crist). LFPs were referenced to ground, band-pass filtered, and recorded at 1 kHz (Plexon Multichannel Acquisition Processor). Spikes were detected from high-frequency data using an online manually-set amplitude threshold, and manually sorted into isolated single units offline (Offline Sorter, Plexon). For NHP S, a DC-coupled headstage, LFP filtering from 3.3–88 Hz, and spike filtering from 154–8800 Hz was used. For NHP A, an AC-coupled headstage, LFP filtering from 0.7–170 Hz, and spike filtering from 250–8000 Hz was used.

For the Single-image Task, NHPs were trained on a delayed match-to-sample WM task. Each trial began with fixation for 500 ms, followed by a sample object, blank memory delay, and then a free-viewing two-choice test. During the test period, NHPs were required to saccade to the remembered sample and hold fixation there. An infrared-based eye-tracking system (Eyelink 1000 Plus, SR-Research, Ontario, CA) continuously monitored eye position at 1 kHz. Neural signals were recorded with four chronically implanted 64-electrode (8x8) "Utah" arrays (1.0 mm length, 400 µm spacing, iridium-oxide tips; Blackrock Microsystems, Salt Lake City, UT). Arrays were implanted in both hemispheres of dorsolateral PFC and ventrolateral PFC. LFPs were referenced to low-impedance subdural reference wires, AC-coupled, band-pass filtered (0.05–300 Hz), and recorded at 1 kHz (Blackrock Cerebus). Spikes were detected from high-frequency data (250–5000 Hz) using a manually-set amplitude threshold and spike-sorted online using time-amplitude windows.

For the Category Task, NHPs were trained to categorize two categories of dot-pattern category exemplar stimuli, each of which was defined by random distortions around a prototype. Each prototype consisted of 7 pseudo-randomly placed dots. The exemplars were selected by randomly sampling locations from a normal distribution centered at each prototype dot location. The session was split into blocks of trials, with exponentially increasing numbers of category exemplars and increasing levels of distortion around each prototype in successive blocks. The first block featured two exemplars of each category, and each subsequent block doubled the number of exemplars, cumulatively keeping all exemplars from previous blocks and adding an equal number of new exemplars. Each block, the variance of the normal distributions around the prototype dots, which was used to generate each exemplar (i.e. the distortion), progressively increased. NHPs moved from one block to the next when a learning criterion was achieved. Neural signals were recorded with two chronically implanted 64-electrode (8x8) "Utah" arrays (1.0 mm length, 400 µm spacing, iridium-oxide tips; Blackrock Microsystems, Salt Lake City, UT). Arrays were implanted unilaterally in dorsolateral PFC and ventrolateral PFC. LFPs were referenced to ground, AC-coupled, band-pass filtered (0.75–300 Hz), and recorded at 1 kHz (Blackrock Cerebus). Multi-unit spiking activity was obtained by offline high-pass filtering and thresholding of the 30 kHz-sampled raw broadband signals. A channel was defined as recording spikes (multi-unit (MUA) channel) when its spike rate averaged over trials and sample presentation period (0–1 sec) was larger than 1 Hz.

Note that any global fluctuations that might have been introduced via the reference wire (Single-Image Task) or ground (Seq-Image Task, Category Task) would be expected to, if anything, work against our primary results. Our results rely on the fact that alpha/beta oscillations exhibit *specific spatial patterns*. Since global fluctuations are, by definition, consistent across all electrodes, they would actually mask any such patterns. Further, to the extent that global fluctuations reflect noise or other non-task-related factors, they would also work against our results showing that alpha/beta patterns differ between task conditions.

Signal processing

In all datasets, local field potentials (LFPs) were first preprocessed by removing the DC offset, and removing the 60 Hz line noise and its harmonics using an adaptive sinusoid fit method. Subsequently, signals were band-pass filtered between 0.5 and 300 Hz.

For spectral analyses of the LFPs, we applied Morlet wavelet analysis [67] after subtracting the evoked potentials phase-locked to trial events, estimated using the grand mean signal across all trials. The Morlet wavelets had a fixed width of 6 cycles, with center frequencies ranging from 2^1 to $2^{7.5}$ Hz in increments of 0.2 (on a logarithmic scale). Resulting spectrograms were then normalized to correct for the ~1/frequency baseline distribution of power by multiplying power at each frequency by the frequency itself [68,69]. To obtain band-specific oscillatory power, we averaged the spectral data within the alpha/beta (10–30 Hz) and gamma (30–150 Hz) frequency bands. Note that, without the above correction for the 1/frequency distribution of power, these band-pooled power estimates would be heavily biased toward the lower-frequency end of each band. Oscillatory power was calculated with a temporal resolution of 1 ms.

Neural spiking data across all datasets were smoothed using a Gaussian kernel with a width of 50 ms. The smoothed spiking activity is referred to as the spike rate, with a temporal resolution of 1 ms. Both LFP band power and spike rate data were baseline-corrected by subtracting the mean activity (averaged across trials and time points) in the time interval preceding stimulus onset: [–1, 0] s relative to sample onset for Seq-image Task and Category Task and [–0.5, 0] s for Single-image Task.

In addition to these standardized preprocessing steps, dataset-specific procedures from the original studies were applied as necessary; further details are provided in those publications. Briefly, for the Seq-image Task, sessions containing high-power, broadband frequency artifacts were excluded from analyses. In the Single-image Task, channels identified as having poor electrode contact or significantly elevated noise amplitudes, based on visual inspection, were removed prior to analysis. Unless specified otherwise, all computations below were performed on a session-by-session basis.

Information quantification using PEV

To quantify the amount of information carried by each neural signal type (alpha/beta power, gamma power, and spike rate), we computed the bias-corrected percentage of explained variance (PEV) [70] across trials under different task conditions. For different analyses, this quantification was performed at both the population level and the level of individual electrodes or units.

At the population level, we first defined an optimal one-dimensional readout axis based on all correct trials. At each time point, neural population activity (LFP power or spike rate) formed distinct clusters in a high-dimensional space, where each dimension represented one electrode or unit. Each data point corresponded to the neural response from a single trial, and clusters represented two distinct experimental conditions (i.e., a binary classification). For task information, these conditions included Category-1 vs.

Category-2 (Category task), first vs. second presentation order (Seq-image task), or
Recognition vs. Recall rule versions (Seq-image task). For order information in the Seq-
image Task, we held the image shown at both sequence orders constant to reduce the
effect of stimulus identity on the variance. Each image shown gave us two separate
clusters corresponding to first and second sequence order. For sensory information
(e.g., stimulus identity), each stimulus pair formed two separate clusters.

1075

For each binary contrast, we computed the linear discriminant axis, which reflects the
optimal linear decoder that maximizes the separation between classes [71,72] and the
Fisher information [73]. The optimal axis was calculated as the product of the pooled
inverse covariance matrix (Σ) of the two clusters and the vector connecting their cluster
means: $\Sigma^{-1}(\mu_1 - \mu_2)$. After averaging this optimal axis across time (and across
stimulus pairs for stimulus identity, or across the images shown for order), we projected
the original high-dimensional data onto this one-dimensional axis. This step effectively
reduced the dimensionality of the neural data, facilitating the computation of PEV
analogous to analyses conducted at the single electrode/unit level.

1085

For the resultant one-dimensional projected population data, as well as for single
electrode/unit data, we applied a one-way ANOVA, grouping trials according to different
experimental conditions (e.g., stimulus identity, task categories, task rules). Only correct
trials were included. A bias correction was applied to PEV estimation (ω^2 formula) to
address potential biases associated with small sample sizes [74]:

$$\omega^2 = \frac{SS_{between\ groups} - df \times MSE}{SS_{total} + MSE}$$

where $SS_{total} = \sum_{i=1}^{N} (x_i - \underline{x})^2$, $SS_{between\ groups} = \sum_{group}^{G} n_{group}(\underline{x}_{group} - \underline{x})^2$, df is
the degrees of freedom (i.e., the number of groups G − 1) and MSE is the mean squared
error $MSE = \sum_{i=1}^{N} (x_i - \underline{x}_{group})^2$. Consequently, bias-corrected PEV ω^2 provided a
robust measure of the amount of information encoded by modulations in neural spiking
rates or oscillatory power at both the population and single-electrode/unit levels.

1097

For the Seq-image Task, the sensory PEV was computed based on the trials of the
Recognition and Recall versions separately for each session. They were aggregated
across all the sessions of the two NHPs. The task PEV was computed by all the trials
together with the two groups given by "Recognition" and "Recall ". For the order
information, we separated each trial into two temporal segments: sample 1 + delay and
sample 2 + delay. These two segments from each trial gave the two groups for PEV.
The order PEV was computed for Recall trials only so it could be compared with the
PEV of order-error trials, which are most easily identifiable in the Recall version.

1106

1107 For the Category Task, the sensory PEV was computed only over trials from the first
1108 block, where the number of exemplar dot patterns was small. This minimized the effects
1109 of category learning on sensory representation. The category information was computed
1110 for the blocks later than block three, by which NHPs typically learned the category [12].

1111

1112 To estimate the information during error trials, we projected incorrect single-trial neural
1113 responses onto the optimal one-dimensional axis derived from the correct trials. The

1114 same one-way ANOVA procedure described above was subsequently applied to these
1115 projections.
1116
Low- and high-abstractness based on exemplar distortion
For the Category Task, we classified exemplars into low- and high-abstractness groups
using the same criterion as described previously by [13]. Specifically, the dataset from
each recording session was divided into two abstractness levels based on a critical
distortion threshold of 1.1 degrees visual angle (DVA), corresponding to the inflection
point in behavioral performance. Further methodological details are provided in [13].
Analysis based on abstractness levels, for example, category PEV, was performed on
the trials of low- or high-abstractness separately. For example, for the category
information under low- and high-abstractness conditions, we computed the one-
dimensional axis from the low- and high-abstractness trials separately and projected the
data separately.
1128
Spatial clustering analysis of alpha/beta patterns
To characterize the spatial clustering and organization of alpha/beta power, we
computed Moran's I, a commonly used statistic to quantify the overall strength of spatial
autocorrelation [20–22]. It is essentially a distance-weighted correlation between all pairs
of electrodes. As with other correlation measures, its value ranges from −1 (i.e.,
maximally spatially dispersed/anti-correlated) to +1 (i.e., maximally spatially
clustered/autocorrelated). A value of 0 indicates a lack of correlation. We computed it
separately on the baseline-subtracted alpha/beta power and on a measure of alpha/beta
power selectivity (signed difference in power between task conditions). Note that per-
electrode baseline subtraction removes any global power gradients [27], so any observed
effects are unlikely to be explained by nonspecific spatial structure.
1140
1141 Moran's I is defined as:

$$I = \frac{N}{W} \frac{\sum_{i=1}^{N} \sum_{j=1}^{N} w_{ij} (x_i - \bar{x})(x_j - \bar{x})}{\sum_{i=1}^{N} (x_i - \bar{x})^2}$$

1142

where N is the number of electrodes indexed by i and j, x_i is the observed alpha/beta
power (or selectivity) at the i-th electrode, $\bar{x}$ is the averaged power (or selectivity) across
all electrodes, w_{ij} are the elements of a matrix of spatial weights with zeroes on the
diagonal, and W is the sum of all w_{ij} (i.e., $W = \sum_{i,j=1}^{N} w_{ij}$). The spatial weights we chose
were binary weights where nearest neighbors (i.e., the electrodes at a minimal distance
from a given electrode) were set to 1 and all other weights were set to 0. We computed
Moran's I for the actual observed alpha/beta patterns and patterns spatially shuffled
across each recording array. As an alternative measure of spatial clustering, we also
computed the absolute differences of baseline-subtracted alpha/beta power between
nearest-neighbor electrodes, and also compared this to values for spatially-shuffled
data (**Figure S5**). Both metrics gave consistent results.
1154
1155 **Spatial correlation between alpha/beta power and sensory PEV maps**
1156 To quantify the spatial relationship between alpha/beta power and the sensory
1157 information (PEV) encoded by gamma power and spiking, we first constructed spatial
1158 maps representing both power and sensory PEV.

1159

For alpha/beta power, we first subtracted the baseline (pre-stimulus) values as
described above. Then, for each electrode, power was averaged across all correct trials
and across a specific time window during stimulus presentation (approximately 0.1–0.4
s for the Single-image Task and 0.2–0.6 s for the Category Task). This resulted in a
vector of 64 alpha/beta power values recorded in each session. These values were then
mapped onto the 8-by-8 electrode grid locations to build the power map.

1166

For the sensory PEV maps, we first computed the PEV of stimulus identity separately
for each electrode/unit at each time point. Note that, unlike the population PEV
described above, this was computed separately for each individual electrode/unit,
without any projection onto a population coding axis. We then averaged the resulting
PEV time series within the same time window used for alpha/beta power averaging,
producing a sensory PEV vector with dimension equal to the number of electrodes/units.
For gamma activity, the resulting PEV vector was mapped directly to the grid locations
of each electrode. For spiking activity, the approach varied slightly: for unsorted multi-
unit activity (MUA) recorded in the Category Task, electrode locations were directly
mapped. However, for sorted single-unit activity recorded in the Single-image Task,
electrodes could have multiple units or none. When multiple units were recorded on the
same electrode, their PEV values were averaged to yield a single electrode-level PEV
value. Electrodes with no recorded units were excluded from spatial correlation
computations.

1181

Spatial correlations were computed using Pearson's correlation coefficient. For each
recording session, we correlated the alpha/beta power vector with the corresponding
PEV vector, including only electrode locations containing valid data points (i.e.,
excluding electrodes with missing values because they recorded no units). Note that the
observed anti-correlation between alpha/beta power and sensory PEV is unlikely to be
attributed to variations in the 1/f spectral slope between trial epochs or experimental
conditions [15,75]. First, power spectra were corrected for 1/f scaling by multiplying power
at each frequency by its corresponding frequency prior to any further analysis. Second,
all spatial correlation analyses were performed within the same trial epoch and
experimental condition (e.g., for a single memorized object identity), ensuring that any
across-epoch or across-condition differences in 1/f slope could not account for the
reported effects.

1194

Within- and between-category spatial correlation of alpha/beta power maps
For the Category Task, we assessed spatial correlations of alpha/beta power maps both
within and between categories. Trials from each recording session were segmented into
non-overlapping windows of 50 trials. Within each set, we computed spatial correlations
separately and subsequently averaged the results across the trial windows. For each
set of 50 trials, we constructed alpha/beta power maps by separately averaging trials
from Category-1 and Category-2, both across trials and across the stimulus
presentation temporal window. The between-category correlation was calculated
directly between these two resulting spatial maps. For within-category correlations, trials
from each category were randomly partitioned into two halves, and correlations were

computed between spatial maps averaged from these halves. This random partitioning
procedure was repeated 20 times, and the average correlation value was taken as the
within-category correlation estimate. Finally, spatial correlation values obtained from the
trial windows after the third block were averaged to yield the overall within- and between-
category correlation estimates for each recording session.
1210
Single-trial behavioral performance analysis
Single-trial behavioral performance analyses were conducted following the procedure
described by [23]. For each recording session, neural responses (LFP power or spike rate)
were first averaged across the stimulus presentation window for each individual trial,
resulting in trial-specific population response vectors. Using only correct trials, we then
calculated an optimal one-dimensional discriminant axis by multiplying the inverse
covariance matrix of two response clusters (corresponding to Category-1 and Category-
2) with the vector connecting their cluster means. Single-trial population responses were
projected onto this discriminant axis, and the resultant scalar projections were
normalized so that projections of -1 and +1 corresponded to the mean responses of
correct Category-1 and Category-2 trials, respectively.
1222
Note that we used all correct trials to estimate the decision axis to obtain a more stable
and robust estimate. This was particularly critical when the data were further subdivided
by category and abstractness level. This could, in principle, bias the axis toward better
separation of correct trials than of error trials. However, this cannot account for the
systematic shift of error trials to the *opposite* side of the axis relative to their correct
category (**Figure S8**). Thus, our results are unlikely to be explained by overfitting to
correct trials, but is consistent with genuine category mislabeling in error trials.
1230
Incorrect trials exhibited a projection distribution shifted toward the opposite category:
incorrect Category-1 trials shifted toward +1, while incorrect Category-2 trials shifted
toward -1. For each session, we used the one-sided t-test to see whether this shift in
the incorrect distribution was significant ($p<0.05$) or not. For those sessions with
significant shifts, we quantified these shifts by computing the mean projection values
separately for incorrect Category-1 and Category-2 distribution and then calculated the
absolute differences (denoted as $|\Delta|$) between these incorrect-trial means and their
respective correct means (-1 for Category-1, +1 for Category-2).
1239
To estimate the probability of correct choices as a function of scalar projection, we
binned the projection values from all trials within each recording session and counted
the number of correct and incorrect trials per bin. Given that the projections were
normalized for each session, bin alignments were consistent across sessions, with -1
representing the average correct Category-1 response and +1 the average correct
Category-2 response. Finally, we aggregated these binned trial counts across all
sessions and both NHPs.
1247
Statistical methods
Statistical tests for significance were performed with a two-sided Wilcoxon rank-sum
test (ranksum function in MATLAB) when samples were independent and with a two-

sided t-test (ttest function in MATLAB) for paired samples or single samples.
Correlations between two variables were computed as the Pearson correlation.
Statistical significance was defined by a p value <0.05. The statistical details (correlation
coefficient, p value, sample size n) are provided in the figures, figure legends, or the
text of the Results section. The specific meaning of the sample size n is clarified when
used.

Supplemental information

Oscillatory Control of Cortical Space as a Computational Dimension

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1264

Figure S1. Recording sites in the lateral PFC across tasks

(A). A brain schematic showing the location of lateral prefrontal cortex (PFC) where all the recordings were performed. (B). Anatomical locations of recording sites for the Seq-image Task, for both NHPs. (C). Utah array locations in the dlPFC and vlPFC for the Single-image Task and Category Task.

1269

Figure S2. Trial-averaged power spectrograms for each dataset show band-limited alpha/beta oscillations

(A). Trial-averaged power spectrogram for the Seq-image Task. The alpha/beta band is in-between the black dashed lines. The red dashed lines indicate sample presentation. (B). Same as (A) but for Single-image Task. (C). Same as (A) but for Category Task.

1275

Figure S3. Paired significance test for sensory and task information

(A-B). Sensory (A) and task (B) information by each signal type in the Seq-image Task. Error bar: mean ± SE, n = 2×24 sessions for (A) and n = 24 sessions for (B). Horizontal black lines on top represent significant PEV values for pairwise comparison between signal types (p<0.001, Wilcoxon rank-sum). (C-D). Sensory (C) and task (D) information by each signal type in the Category Task. Left: PEV for dlPFC. Right: PEV for vlPFC. Error bar: mean ± SE, n = 51 sessions. Horizontal black lines on top represent significant PEV values for pairwise comparison between signal types (p<0.001, Wilcoxon rank-sum).

1283

Figure S4. Firing rate and motor information for Category Task

(A). Firing rate (spikes per second) averaged over all dlPFC and vlPFC MUA channels (MUA channel

= average spike rate between 0 and 1 s > 1). (B). Information quantified by PEV of the left vs. right

saccade for dlPFC (left) and vlPFC (right). The vertical dash line indicates the time of test onset. Error

bar: mean ± SE, n = 51 sessions.

1289

Figure S5. Spatial regulation of sensory information by alpha/beta patterns

(A). The spatial clustering of alpha/beta power measured by Moran's I (left) and absolute power difference (right) in dlPFC recorded from Single-image Task. The values are compared with randomly shuffled patterns (bar with stripes). *** p<0.001, **: p<0.01, two-sample t-test. Error bar: mean ± SE. Sample sizes: n = 22. (B). Same as (A) but for vlPFC of Single-image Task. Error bar: mean ± SE. Sample sizes: n = 20. (C). The spatial clustering of alpha/beta power measured by absolute power difference in dlPFC recorded from Category Task. *** p<0.001, **: p<0.01, two-sample t-test. Error bar: mean ± SE. Sample sizes: n = 51. (D). Same as (C) but for vlPFC of Category Task. (E). Spatial correlation of alpha/beta power computed for the same category (within) and for different categories (between) in dlPFC. *** p<0.001, **: p<0.01, two-sample t-test. Error bar: mean ± SE. Sample sizes: n = 51. (F). Same as (E) but for vlPFC. (G). Spatial correlation differences (between- minus within-category) for high- and low-abstractness conditions in dlPFC. *** p<0.001, **: p<0.01, n.s.: not significant, one- and two-sample t-test. Error bar: mean ± SE, n = 51. (H). Same as (G) but for vlPFC. (I). Example alpha/beta power maps with the corresponding sensory PEV carried by spiking for both categories under low and high-abstractness conditions.

1305

Figure S6. Category information (PEV) in dlPFC for correct and incorrect trials

(A). Category information (PEV) for correct trials and incorrect trials for alpha/beta (left), gamma (middle), and spiking (right) under low-abstractness. Horizontal grey bars on the x-axis represented significant PEV differences between correct- and incorrect conditions (p<0.05, two-sample t-test). Error bar: mean ± SE, n = 51. (B). Cumulative difference in category information (correct minus incorrect) across time for each signal type. *** p<0.001, n.s.: not significant, one-sample t-test for each signal type and two-sample t-test for across signals. Error bar: mean ± SE, n = 51. (C-D). Same as (A-B) but under high-abstractness.

1313

Figure S7. Category information (PEV) in vlPFC for correct and incorrect trials

(A). Category information (PEV) for correct trials and incorrect trials for alpha/beta (left), gamma
(middle), and spiking (right) under low-abstractness. Horizontal grey bars on the x-axis represented
significant PEV differences between correct- and incorrect conditions (p<0.05, two-sample t-test). Error
bar: mean ± SE, n = 51. (B). Cumulative difference in category information (correct minus incorrect)
across time for each signal type. *** p<0.001, n.s.: not significant, one-sample t-test for each signal type
and two-sample t-test for across signals. Error bar: mean ± SE, n = 51. (C-D). Same as (A-B) but under
high-abstractness.

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Figure S8. Signed projections of incorrect trials onto the decision axis for Category Task

(A). Signed projections of incorrect trials on the category axis of each category under low- and high-abstractness conditions, shown for alpha/beta (left), gamma (middle), and spiking (right) in dlPFC. *** p<0.001, n.s.: not significant, one-sample t-test for each signal type and Wilcoxon rank-sum test for across signals. Error bar: mean ± SE, n = 51. (B). Same as (A) but for vlPFC.

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Figure S9. Standard deviation of the projections from incorrect trials for Category Task

(A). Standard deviation of the incorrect-trial projections for all signal types in dlPFC. *** p<0.001, **:
p<0.01, *: p<0.05, one-sample t-test for each signal type and Wilcoxon rank-sum test for across signals.
Error bar: mean ± SE, n = 51. (B). Same as (A) but for vlPFC. (C).