Prior learning shapes effort choice, revealing valence asymmetries and disruption in anhedonia

Calum Guinea[1,*], Bowen Xiao[1] and Rebecca P. Lawson[1,2]

[1] Department of Psychology, University of Cambridge, United Kingdom

[2] MRC Cognition and Brain Sciences Unit, United Kingdom

* Corresponding author

Word count (excluding abstract, methods, tables/figures and references): 5000

Author note

Correspondence concerning this manuscript should be addressed to Calum Guinea, Department of Psychology, Downing Place, University of Cambridge, CB2 3EB, United Kingdom. **Email:** calumlukeguinea@gmail.com. All authors declare no conflict of interest. This study was not preregistered. **Data availability:** The data from this study are available from the corresponding author upon reasonable request. **Funding statement:** CG is funded by the Harding Distinguished Postgraduate Scholarship. BX is funded by the Trinity Henry Barlow Studentship and the Cambridge International Scholarship. This work was supported by a Wellcome Trust Henry Dale Fellowship awarded to RPL. **Ethics:** This work was approved by the Cambridge Psychology Research Ethics Committee (CPREC), ethics number PRE.2019.110.

Author ORCIDs

Calum Guinea: https://orcid.org/0009-0006-9444-9838

Bowen Xiao: https://orcid.org/0000-0002-9627-3454

Rebecca P. Lawson: https://orcid.org/0000-0003-1228-1244

Abstract

Goal directed behaviour requires the integration of learned beliefs about the world with decisions about whether an outcome is worth the effort required to obtain it. Although value learning and effort-based choice are central to motivated behaviour, they are typically examined in isolation. Here we introduce a novel behavioural paradigm that directly links probabilistic associative learning with effortful decision making, for the first time allowing trial by trial assessment of how learned beliefs guide willingness to exert effort.

In a large sample of participants (n=252), individuals reliably acquired stimulus-outcome associations but showed substantial variability in learning fidelity. Effort allocation was not determined by objective reward probabilities, but was better predicted by participants subjective beliefs, demonstrating that effortful action depends on internal value representations rather than external contingencies alone. Reward and loss contexts showed asymmetric effects on effort motivation, revealing a valence specific dissociation in belief driven action.

Critically, higher levels of anhedonia were associated with reduced willingness to engage in effortful behaviour and a weaker coupling between learned value and action, indicating a disruption in the integration of belief and motivation. Together, these findings identify a mechanistic link between learning and effort expenditure and suggest that motivational impairments in psychiatric symptoms may arise from failures to translate learned value into action.

Introduction

The way we allocate effort in the present is inherently shaped by what we have learned about the value of available options in the past. Yet, despite this intrinsic connection, research on learning and effort-based decision-making has largely evolved in parallel, forming substantial but predominantly separate bodies of literature in cognitive neuroscience and related fields. To date, no existing work has systematically examined how prior learning influences subsequent effort-based decision-making.

Decision-making can be understood as a form of cost-benefit analysis, where people weigh potential rewards against associated costs. Principal among these costs is effort: be it physical or cognitive, each decision we make is likely to incur some effort cost. For example, visiting a new café might require walking two miles. While the absolute effort cost (i.e., the distance) is fixed, individuals differ in their sensitivity to that cost (i.e., how aversive they find the walk) and in how much they value the potential reward (i.e., how enjoyable they expect the café to be). Simply put, the decision to visit the café depends on whether the anticipated reward justifies the required effort.

Early studies established that humans generally seek to minimise effort whenever possible[1,2]. This aversion applies to both physical[3] and cognitive effort[4], but can be overridden when effort leads to reward. Notably, in some cognitive tasks, effort itself can be intrinsically rewarding or preferred over rest[5]. When access to rewards is effort-dependent, we conceptualise the reward's subjective value as being discounted based on the effort required to obtain it[6]. Importantly, this effort discounting occurs even when rewards are hypothetical, such as virtual apples[7] or virtual money[8]. While effort discounting is observed in the general population, it is steeper in individuals with psychiatric symptoms such as apathy and anhedonia[9,10].

Willingness to exert effort for reward is most commonly assessed using effort-based decision-making tasks[10–14]. Physical effort has been measured experimentally in several ways, including repeated pressing of keys on a keyboard[10,15], but most commonly, using grip force, where the effort requirements in the task are defined as a proportion of a person's maximum grip strength (Maximum Voluntary Contraction, MVC)[11–14]. Typically,

these tasks then ask participants to choose whether to exert different levels of effort for

different amounts of reward such that effort exertion is deterministically related to reward

(see [15,16] for exceptions). However, in naturalistic decision making, exerting effort is often

probabilistic and does not guarantee the receipt of reward (i.e. just because we exert

effort to visit the new café, it does not guarantee that our experience of the café will be

positive). Our tasks must therefore account for the *likelihood* of getting the desired

outcome or avoiding the undesired one. To reflect the probabilistic relationships between

effort and outcomes, some experimental designs choose to cue the probability of

receiving reward at the start of each trial[15–17]. While this enhances experimental control,

it comes at the cost of overlooking a crucial individual difference – namely, that people

may vary in the subjective beliefs they acquire about the uncertain relationship between

effort and its outcomes. These differences in subjective beliefs can arise from differences

in learning, which produce different experiences of uncertainty. Extending our example,

you may feel uncertain about a new café, as you have little prior knowledge about the

quality of its coffee. Yet uncertainty applies even to familiar choices – your usual café

might still disappoint despite past good experiences. How we learn about this uncertainty

and use what we have learned to guide our choices is therefore both central to

naturalistic effort choice behaviour and psychiatrically relevant[18–22].

Evident throughout the literature on learning is the idea that people learn differently, even

from identical experiences. Computational models highlight variability in learning rates -

some people update beliefs quickly, others slowly[18,23,24]. A simpler way to measure

people's learning is through explicit choices or preferences, having people learn

probabilistic relationships between stimuli and rewards, then selecting between the

stimuli in a "tournament". This approach provides a detailed measure of how accurately

they learned those initial probabilities[25–27]. Such differences matter for effort-based

decisions – if someone overestimates reward probability (e.g. believing that a 70%

chance of reward is actually 80%), they may be more willing to exert effort. Conversely,

someone who underestimates reward probability might be less willing to exert effort.

Learning accuracy is therefore important to effort-based choice, yet effort decision-

making studies often treat reward as deterministic, such that if the effort requirement on

a given trial is met, then reward is obtained[14]. In reality, effort decisions hinge on learning

accuracy, reward sensitivity and motivation. Without measuring and independently

manipulating these variables, we cannot pinpoint why people differ in effort exertion. This

may be particularly important in the context of mental health conditions where diagnoses

affecting learning and effort decision-making are highly co-occurring[28].

Finally, most of the research on effort-based decision-making has tended to focus on

effort for reward[3,10,11,13–15,29], though the vigour of effort to prevent negative outcomes has

been studied previously[30]. Our willingness to exert effort to avoid negative outcomes has

rarely been studied, despite the fact that avoidance and apathy are characteristic of

mood and anxiety disorders[6,31]. Valence-driven differences in learning from prediction

errors are well-documented[32], and loss aversion is a robust feature of economic

decision-making - losing £100 feels worse than gaining £100 feels good[33]. Some studies

have investigated willingness to exert effort for loss; in one example, participants were

required to exert either high or low effort for appetitive or aversive stimuli[34]. Here, effort

exertion is deterministically related to the outcomes; this exertion was not calibrated to

the individual (as is common in grip force paradigms[14,35] and participants can never

entirely avoid exerting effort. In a study of prosocial effort, the grip force task was used to

compare participants' willingness to exert effort to obtain reward or avoid loss for oneself

or for another person[35]. Again, effort exertion was deterministically related to the

outcomes participants received and motivation to obtain reward and avoid loss was not

compared. As with reward, these losses may be probabilistic and require the decision-

maker to track or infer this probability, alterations in the tracking and estimation of

uncertainty have regularly been associated with mood disorders[36]. Given the literature's

emphasis on effort for rewards, whether humans avoid effort to prevent losses in the

same way they pursue effort for gains remains an open question.

Psychiatric relevance of learning and effort-based choice

Both learning and effort-based decision-making are of established psychiatric and

neuropsychological interest. Anxiety has been associated with elevated or improper

adaptation of learning rates[18,37] and, different types of anxiety drive distinct biases in

safety versus threat learning[38]. Misestimating uncertainty can reinforce avoidance –

anxious individuals may overgeneralise from negative experiences, exerting excessive

effort to avoid potential threats. Core features of anxiety like rumination, are effortful
because of the processes they recruit, like mental replay of unpleasant experiences[39] and
persist despite the cost they incur, which suggests that anxiety may be associated with
an increased willingness to exert effort. While avoidance characteristically defines
anxiety disorders, evidence also links anxiety to increased willingness to exert cognitive
effort in reward-based tasks[40]. Recent literature has found that different types of anxiety
have different behavioural effects on learning[38] and effects separable from depression[41].
Given the heterogeneity of experience in mood disorders[42], its impact on physical effort
remains little understood.

Anhedonia and apathy also shape effort decisions[7,10,15,16,40]. Anhedonia, often tied to
depression but present across many psychiatric and neurological diagnoses[6], reduces
effort motivation and blunts sensitivity to reward probability and outcome
magnitude[15,16,40]. It has also been argued that anhedonia reflects impaired Pavlovian
biases (i.e. action to approach reward and suppress action to avoid) rather than altered
reward or punishment sensitivity[43]. Fatigue, a symptom that is likewise implicated across
psychiatric diagnoses, further complicates the specificity of anhedonia and apathy's
effects on effort decision-making. Fatigue frequently co-occurs with both, and all three
can manifest as lack of motivation and reduced energy levels. Computational modelling
has shown that fatigue affects people's ability to persist in effortful tasks, potentially
mimicking or amplifying the effects of apathy and anhedonia[14]. Distinguishing these
constructs is crucial for understanding how different psychiatric symptoms contribute to
effort-based decision-making.

Effort-related impairments likely arise from multiple mechanisms – limiting the utility of
diagnostic labels. This reflects a near-ubiquitous issue with current psychiatric nosology,
for which several alternatives have been proposed to simplify or improve upon it (e.g. p-
factor[44] and the RDoC matrix[45]). Transdiagnostic factor modelling, which takes a large
number of self-reported questionnaire responses about mental health symptoms and
identifies common structure amongst them, has been used to both simplify and identify
commonalities across psychiatric diagnoses[46]. Using this approach, the anxious-
depression factor and the social withdrawal factor are both associated with the
avoidance of mental effort[47]. Such dimensional approaches, then, might be fruitful in

disentangling the impact of blurred psychiatric constructs on learning and effort
decision-making.

Here we propose that across individuals, the fidelity of a person's learning, their
sensitivity to reward and their sensitivity to effort may have convergent effects on effort-
based decision-making. To arbitrate between these possible causes, we designed a novel
task that combines Pavlovian learning with effort-based decision-making. This allows us
to dissociate the contributions of learning, reward sensitivity and effort sensitivity on
effort decision-making, offering a clearer picture of how these factors shape decision-
making in anxiety, anhedonia and apathy. First, in the learning phase (**Figure 1A**),
participants (n=252) learned probabilistic associations between eight different aliens
with varying outcome probabilities (0.25, 0.75) and magnitudes of reward/loss (+ 40
points, +160 points, - 40 points, - 160 points), reporting objective estimates of the
outcome probability on each trial. We then assessed how well participants could
discriminate between pairs of these stimuli in a "tournament" phase (**Figure 1B**). Finally,
participants chose to 'accept' or 'reject' the opportunity to exert effort (keypresses) to
approach or avoid one of the eight aliens they had previously learned about (effort
decision-making phase; **Figure 1C**). Effort levels (low - 60%, medium - 75%, high - 90%)
were adjusted to each participant's maximum number of presses in three calibration
trials at the beginning of the experiment.

Results

Effort calibration

All participants, except one, comfortably exceeded the required minimum of 19 presses
in one of their three calibration trials. The average number of keypresses across the
sample was 45, with a standard deviation of 11.004 and a good range of individual
differences captured, min = 19, max = 78 (**Figure S1**).

Measures of learning

Several robust behavioural indicators suggest that participants were, on average, able to
learn the outcome probabilities of the eight different aliens.

Trial-by-trial estimates of stimulus outcome probability

In the learning phase of the experiment, participants' trial-by-trial estimates of each
alien's outcome probability trended towards the true outcome probability (**Figure 2A**).
However, there was significant variation across the sample. There were substantial
individual differences in learning accuracy, with some participants overestimating the
true outcome probability and others tending to underestimate the true outcome
probabilities (**Figure 2B and 2C**).

Post-task outcome probability ratings

At the end of the learning, tournament and effort-choice tasks, participants gave 'final'
explicit ratings of the outcome probability for all eight aliens. This provides a measure of
what was learned about the stimuli and whether it changed across time (**Figure S2**). After
initial learning, participants' average estimates of the outcome probabilities did not
significantly change across time points ($F_{(2, 502)} = 0.851$), Greenhouse-Geisser
corrected $p = 0.428$; see **Figure S2**), suggesting that learning remained stable. Comparing
participants' average estimates against the true average probabilities of the stimuli
indicated that, at the group level, they generally correctly approximated the true outcome
probabilities (**Figure 2C; Table S1**).

Tournament performance as a measure of learning

After the learning phase, participants completed a binary-choice tournament task which
enabled a more detailed assessment of participants' preferences amongst the learned
stimuli. Overall participants made accurate choices (mean = 78%, SD = 10.95) and were
good at approaching the best alien when it was present in a choice pair (mean = 85%, SD
= 17.57) as well as avoiding the worst stimulus when it was present in a pair (mean = 82%,
SD = 18.89). A Bayesian t-test provided support for the null hypothesis that there was no
difference in participants' approach and avoid accuracy ($BF_{10} = 0.304$). When choices
were more difficult, with pairs containing two stimuli of the same valence, participants
were less accurate for both reward (mean = 65%, SD = 20.19) and loss (mean = 62%, SD
= 17.84). Again, a Bayesian t-test indicated there was weak evidence for the alternative
hypothesis that reward and loss accuracy were different ($BF_{10} = 0.19$). See **Figure S3**.

Effort decision-making

In the effort decision making task, participants accepted a high percentage of effort offers overall (mean = 75%, SD = 18.81) and tended to accept higher effort offers (mean = 59%, SD = 28.82) less than medium effort offers (mean = 77%, SD = 22.41) and medium effort offers less than low effort offers (mean = 91%, SD = 14.55). Overall participants' effort choices were sensitive to effort level ($F_{(1, 251)}$ = 353.6, $p < 0.001$), outcome magnitude ($F_{(1, 251)}$ = 71.82, $p < 0.001$), outcome probability ($F_{(1, 251)}$ = 207.3, $p < 0.001$) and valence ($F_{(1, 251)}$ = 15.81, $p < 0.001$).

There was a significant effort by outcome probability interaction ($F_{(2, 502)}$ = 75.79, $p < 0.001$), indicating that participants accepted more effort for higher probability outcomes than for lower probability outcomes. A significant outcome magnitude by outcome probability interaction ($F_{(1, 251)}$ = 16.91, $p < 0.001$), suggesting participants accepted more offers for higher probability and higher outcome magnitude stimuli. There was a significant outcome magnitude by valence interaction ($F_{(1, 251)}$ = 29.4, $p < 0.001$), showing that participants accepted similar amounts of effort for low magnitude reward and loss stimuli but accepted more effort for larger magnitude rewards than losses.

Lastly, there was a significant four-way interaction between effort, outcome probability, outcome magnitude and valence ($F_{(1, 502)}$ = 13.56, $p < 0.001$). Across all effort levels, participants accepted more effort for reward than for loss. Generally, they discriminated more precisely between high and low probability as well as high and low magnitude reward stimuli than loss stimuli. For example, participants accepted effort for a similar amount of low and high magnitude stimuli on low effort trials, regardless of outcome probability. Overall, these results suggest that the levels of effort, probability, outcome magnitude and valence were well calibrated in this part of the experiment (see **Figure 3** for 3D bar plots representing how the probability of accepting changes across these variables), with main effects and interactions as expected based on prior literature (Treadway et al., 2009; Treadway et al., 2012).

Effort choices are sensitive to prior learning

After establishing the key behavioural features of the learning, tournament, and effort-based decision-making phases of our experiment, we now examine our primary hypothesis: that individual differences in learning the outcome probabilities of each stimulus influence participants' subsequent effort choices. First, we fit a logistic regression (M0) that assumes all participants learn equally and perfectly about the probability of the aliens they encountered. Full details of all predictors in the model are outlined in the methods section below (see also **Table S2**). Overall, the model was a significant improvement over the null (X^2 = 4221.8, p < 0.001, McFadden R^2 = 0.156). As expected, effort level (β = -7.004, p < 0.001) and objective outcome probability (β = 2.58, p < 0.001) were significant predictors of effort choice but in opposite directions; with increasing effort level reducing the odds of accepting an effort offer and increasing outcome probability increasing the odds of accepting. Likewise, outcome magnitude was a significant predictor of effort choice (β = 0.774, p < 0.001); the odds of accepting increased with increased outcome magnitude. We also observed small but significant effects of learning order (β = -0.102, p = 0.002), where learning about the loss aliens first was associated with decreased odds of accepting an effort offer, and an effect of valence (β = 0.224, p < 0.001), where reward aliens were associated with increased odds of accepting. Of the demographic variables, only gender was a significant predictor (p < 0.001), with small but significant differences largely driven by non-binary participants and those not reporting a gender identity (see **Table S2**).

As M0 assumes participants learn accurately and identically to one another, we next compared a suite of models with different individualised measures of participants' learning as predictors of effort-choice. Full details of all models considered can be found in **Supplementary Table 3**. Crucially, all models containing individual signatures of learning significantly improved the prediction of participants' effort-choice over and above M0 (see **Supplementary Table 3** for a comparison of models containing individualised learning measures and **Supplementary Figure 4** for a correlation matrix of the relationships amongst these learning measures), but the winning model substituted

participants' *post-tournament probability estimates* in place of the objective outcome

probabilities.

Post-tournament probability estimates were a significant predictor (β = 6.01, SE = 0.37,

Wald stat. < 0.001) of effort choice, for every unit of increase in participants' probability

estimates, the log odds of accepting an effort offer increased by 6.01. Outcome

magnitude was a small, positive and significantly non-zero predictor of effort choice, with

the log odds of accepting an effort offer increasing by 0.34 (SE = 0.16, Wald stat. p = 0.035)

for every unit of increase in outcome magnitude. Stimulus valence was also a positive

and significantly non-zero predictor of effort choice, with the log odds of accepting an

effort offer increasing by 0.39 (SE = 0.09, Wald stat. p < 0.001) for reward stimuli. By

contrast, effort level (β = -12.78, SE = 0.65, Wald stat. p < 0.001) was a negative and

significantly non-zero predictor, for every unit of increase in effort level the log odds of

accepting an effort offer decreased by -12.78. See **Figure 4** for a visualisation of the fixed

effects and slope estimates from the best-fitting model M1.

The effect of self-report mental health measures on effort choice

Over and above demonstrating that a model accounting for the individual signatures of

learning improves the prediction of participants' choices, we sought to understand

whether transdiagnostic and self-report measures of mental health symptoms also

predicted participants' effort choices. **Figure S5** shows the distribution of scores on the

SHAPS, FAS and the transdiagnostic factor dimensions across the sample. Accordingly,

we fit three additional binomial generalised linear mixed models (GLMMs) by adding the

transdiagnostic and self-report mental health measures as fixed effects to the best-fitting

model (M1). Model 2 (M2) added the three transdiagnostic symptom measures, Model 3

(M3) added the sum score of the SHAPS and Model 4 (M4) added the FAS. AIC scores

suggest that Model 3 (M3), which included the SHAPS sum score, was the best

performing of these models (**Table S4**) and crucially, also outperformed Model 1 (M1).

Participants post-tournament probability estimates (β = 5.98, SE = 0.37, Wald stat. p <

0.001), stimulus outcome magnitude (β = 0.34, SE = 0.16, Wald stat. p = 0.032), effort

level (β = -12.8, SE = 0.65, Wald stat. p < 0.001) and valence (β = 0.4, SE = 0.1, Wald stat.

p < 0.001) remain significant in this model. In addition to the significant predictors of the

previously reported model, participants' sum scores on the SHAPS were a significantly non-zero predictor of their effort choices (β = -0.34, SE = 0.15, Wald stat. p = 0.026). The log odds of accepting decreased by 0.34 for every unit of increase in SHAPS sum score (**Figure 5**). In comparison, none of the three transdiagnostic factors or the FAS were significant fixed effects predictors of effort choice in their respective models (see **Table S4** for a comparison of these models containing mental health measures against Model 0 and Model 1). As we identified an effect of the SHAPS on trial-by-trial effort choice, we ran partial correlations controlling for the effects of age and gender. SHAPS scores were significantly negatively correlated with the percentage of medium (τ = -0.086, p = 0.043) and high effort offers accepted (τ = -0.083, p = 0.049). Kendall's tau was used here because the percentage of effort offers accepted is not normally distributed, and there were a number of participants with the same percentage of offers accepted (e.g. 100% of low effort offers accepted).

The interaction of anhedonia and individualised measures of learning

As our analysis of participants' trial-by-trial effort choices showed that participants' scores on the SHAPS were a significant fixed effect predictor of their effort choices, we conducted some additional exploratory analyses to assess whether SHAPS scores significantly interact with any of the decision variables included in our winning model (M3; see **Supplementary Materials** for full details of these exploratory models). Only the interaction between participants' post-tournament probability estimates and SHAPS scores was significant (β = -0.71, SE = 0.33, p = 0.0303; **Supplementary Materials EM1**). This suggests that the effort choices of participants with higher anhedonia are less sensitive to their own estimates of the outcome probabilities. Specifically, compared to participants reporting no or low anhedonia, those with high anhedonia are less willing to accept effort as their estimates of the stimulus outcome probability increase (**Figure 6;** see also **Figure S6** for a visualisation of how the probability of accepting varies with SHAPS scores across effort levels).

When participants estimate that the outcome probability of a stimulus is high, all participants are equally likely to accept effort offers. However, as participants' estimates of the outcome probability decrease, participants scoring high on the SHAPS are less

346   likely to accept effort offers. All other predictors in the model remained significant except

347   the fixed effect of SHAPS scores (β= 0.01, SE = 0.22, p > 0.05)

Discussion

Learning and effort-decision making are both central to motivated behaviour, yet they are

typically studied in isolation. By integrating these processes into a single paradigm, we

capture a more ecologically valid model of goal-directed behaviour. Our novel task

combined probabilistic associative learning with effort decision-making to understand

how these two behaviours interact. Analysis of learning and tournament behaviour

clearly demonstrates that participants were able to learn the outcome associations of

the stimuli and then use that learning to accurately discriminate amongst them. However,

we also observe significant variability in learning across participants. Crucially, trial-by-

trial effort decisions were best predicted by participants' *own* estimates of stimulus-

outcome probabilities, rather than the objective probabilities. Moreover, individuals with

higher anhedonia scores were less inclined to exert effort and exhibited a weaker

alignment between their learned beliefs and their decisions—suggesting a disruption in

the translation of expected value into motivated action.

A striking finding was the asymmetry in participants' willingness to exert effort for

potential reward versus avoiding equivalent losses. On average, participants were more

inclined to work for a possible gain than to prevent a comparable loss, despite equal

expected value. This asymmetry was robust and did not vary with anhedonia, suggesting

it reflects a fundamental motivational bias rather than individual differences in hedonic

capacity. To date, few studies have investigated how individuals make effort choices to

avoid negative outcomes, in these studies successful completion of the effort always

leads to reward[34,35]. Here we combine features from existing paradigms, namely: the

calibration of effort to the individual and providing participants with the opportunity to

avoid effort entirely and extend this work by requiring participants to learn the

probabilistic outcomes of effort. Our design combines two major advantages that more

closely mirror real-world effort decision-making, first allowing participants to opt out of

effort so individuals can choose rest over exertion. Second, by making effort

probabilistically related to outcomes, individuals must consider that there is uncertainty

about whether their effort exertion will have the expected outcome. These features
enhance sensitivity to symptoms like anhedonia, as their expression in naturalistic
behaviour may be to opt out of effortful behaviour altogether, yet this opting out could
result from reduced motivation, reduced outcome sensitivity or inaccurate learning
about this uncertainty. While prior research has overwhelmingly focused on effort in
service of obtaining reward, our study broadens the concept of goal-directed behaviour
to include the avoidance of loss. Crucially, we find that reward and loss stimuli with equal
absolute expected value do not motivate effort exertion equally, here participants were
more willing to exert effort to obtain reward than to avoid loss.

Like apathy, anhedonia has been associated with reduced willingness to exert effort for
reward[15,16,48]. Consistent with this, we observed that participants who scored higher on
self-reported anhedonia were less likely to accept effortful offers on a trial-by-trial basis.
Importantly, exploratory analyses revealed a significant interaction between anhedonia
and participants' own estimates of stimulus outcome probability, such that higher
reported anhedonia was associated with weaker coupling between their choices and
their probability estimates. In other words, compared to those with lower anhedonia,
their likelihood of accepting a given effort offer increased less steeply as a function of
their perceived probability of an outcome. Accordingly, we report weak but significant
negative correlations between self-reported anhedonia and the percentage of both
medium and high effort offers. These findings suggest that anhedonia may disrupt the
ability to integrate expected value and outcome probability when deciding whether to
exert effort, potentially leading to less adaptive motivational behaviour.

To our knowledge, this is the first study to examine the relationship between
transdiagnostic mental health factors[49] and physical effort choice. These factors
putatively capture a range of symptoms that are likely related to choice (e.g. items from
the Obsessive-Compulsive Inventory)[50] and effort (e.g. items from the Apathy Evaluation
Scale; AES)[51]. Despite this, no previous work has assessed whether such transdiagnostic
constructs relate to physical effort choice. One prior study previously reported
associations between the anxious-depression and social-withdrawal factors and
avoidance of mental effort[47]. Though the anxious-depression factor comprises
symptoms related to apathy (Apathy Evaluation Scale)[51], social anxiety. (Liebowitz Social

Anxiety Scale)[52], depression (Self-rating Depression Scale)[53] and state and trait anxiety

(State-Trait Anxiety Inventory)[54], it does not necessarily capture anhedonia and here we

did not find evidence of similar associations between anxious-depression and *physical*

effort.

One possibility is that the heterogeneous symptoms subsumed by these factors exert

opposing influences on physical effort-based decision-making, leading to overall null

effects. Indeed, recent work has taken a more targeted approach, using factor analysis

across a battery of questionnaires specifically designed to deeply phenotype anhedonia

and apathy[55] - symptoms with clearer theoretical links to motivational impairments.

Moreover, the impact of symptoms on effort decision-making may be both symptom-

specific and domain-specific. Supporting this, recent work assessing participants'

sensitivity to both physical and mental effort costs found that apathy and anhedonia have

a domain-specific effect on physical effort sensitivity, while anxiety specifically

influenced sensitivity to cognitive effort[40]. These findings underscore the importance of

examining specific symptom dimensions, rather than broad transdiagnostic factors, and

of distinguishing between different types of effort when investigating motivational

impairments in mental health.

Cognitive and physical effort are distinct forms of exertion and may have distinct

relationships to psychiatric symptoms[40]. Anxiety has been linked to reduced sensitivity to

cognitive effort, potentially due to the high costs associated with symptoms like worry

and rumination, which involve excessive mental replay[39]. In contrast, depression-related

symptoms are typically studied in relation to physical effort[15], where they are associated

with increased effort discounting[10]. While the cognitive and physical domains may

involve distinct mechanisms, they may converge through their accumulation of fatigue –

a potential shared cost signal. Recent computational work has further distinguished

between two types of fatigue: one that is recoverable through rest, and one that is not[14].

These latent components of fatigue have been shown to closely track participants'

subjective estimates of effort exertion and tiredness[56]. Fatigue thus appears central to

understanding effort decision-making. However, in our study we were unable to apply

these modelling approaches to estimate individual fatigue parameters, as they rely on

patterns in trial-by-trial decision behaviour that were not elicited in our design. Moreover,

in our study participants likely experienced and recovered from varying degrees of fatigue

across the learning and tournament phases prior to the effort-choice task. While self-

reported fatigue was measured, we found no evidence that it influenced effort decisions

in this sample, unlike in recent work which has reported similar effects of fatigue and

apathy on effort decision-making[9]. Future studies using dynamic fatigue modelling may

reveal how individual differences in fatigue shape effort-based decisions across domains.

Though we observed substantial individual variation in learning measures, these

differences were not associated with any indicators of mental health. At first this might

seem surprising, given that alterations in learning have been reported in association with

a range of different psychiatric diagnoses and symptoms[43]. However, the learning task

used here was deliberately simple—featuring fixed outcome probabilities, no

opportunities for choice, and an emphasis on tracking probabilities rather than

magnitudes. This simplicity was important to ensure that participants could reliably

acquire the outcome contingencies, as this learning served as a foundation for

subsequent effort-based choice. As a result, this learning task was not able to investigate

the associations between specific mental health symptoms and learning behaviour seen

in the literature, including: the relationship between anhedonia and choice temperature[43]

as participants did not make choices during learning; the improper adaptation of learning

rate in anxiety[18] or internalising psychopathology more generally[57] as the outcome

contingencies were stable for all learned stimuli and, anhedonia's relationship to

reduced reward sensitivity[58] as the requirement to provide trial-by-trial estimates of

stimulus outcome probability likely shifted participants attention away from outcome

magnitude. Future studies should test whether such associations might emerge under

more complex or uncertain learning conditions, where greater demands are placed on

flexibility, valuation, or motivational processes.

By integrating probabilistic associative learning and effort-based decision-making in a

single task, our results reveal that participants can learn equally well but differ in how

learning informs action, particularly in the context of psychiatric symptoms. This work

bridges two previously separate literatures and suggests that anhedonia-specific

mechanisms may underlie impairments in goal-directed behaviour.

Methods

Participants

In total, 299 participants were recruited using the online recruitment platform Prolific
using the following inclusion/exclusion criteria: aged 18-65, fluent English speaker, not
colourblind and a Prolific study approval rating above 90%, refer to **Table 1** for a summary
of the sample's demographic information. All these participants provided informed
consent at the start of the study. After applying further exclusion criteria, 252 participants
were included in the final analyses. 42 participants were excluded for repeatedly failing
at least one of the instruction quizzes designed to test participants' understanding of the
learning, tournament and effort decision-making tasks. Five participants were removed
because their overall choice accuracy in the tournament task was 50% or lower,
indicating random responding. Participants who completed all tasks and questionnaires
received £10.50. This study was approved by the Cambridge Psychology Research Ethics
Committee (CPREC; Ethics number PRE.2019.110).

Task design

Effort calibration

At the start of the experiment, to establish each individual's maximum capacity for key
pressing, participants were asked to press the spacebar on their keyboard using their
right and left index fingers as fast as possible in a six second window. With every keypress
the participant added a star graphic to the screen and their task was to "fill the sky with
stars" to help light the path home for an astronaut. The maximum number of keypresses
is a measure comparable to maximum voluntary contraction from grip force tasks[11–14].
This was repeated three times, with their highest number of key presses from the three
runs taken forward into the effort decision-making. After each attempt, participants were

Summary Table 1

Sample size	252
Demographics	
Age, mean (SD)	38.93 (11.92)
Gender, number (%)	
Male	125 (49.6)
Female	125 (49.6)
Non-binary	1 (0.4)
Rather not say	1 (0.4)
Ethnicity, number (%)	
Asian	15 (5.95)
Black	13 (4.76)
Mixed	6 (2.38)
Other	2 (0.79)
White	217 (86.11)
Socioeconomic status, mean (SD)	5.05 (1.7)
Transdiagnostic factor loadings	
Anxious-depression, mean (SD)	-0.6 (0.76)
Compulsivity and intrusive thought, mean (SD)	0.14 (0.79)
Social withdrawal, mean (SD)	0.1 (0.79)
Questionnaire measures	
SHAPS, mean sum score (SD)	10.17 (6.93)
FAS, mean sum score (SD)	25.73 (9.6)

Table 1: **Demographic information of the sample**. Participant loadings onto the three transdiagnostic factors and scores on self-report measures of anhedonia and fatigue.

told how many stars they added to the sky and were encouraged to try and beat it on the next attempt. To prevent participants from minimising their effort, participants need to press at least 19 times to proceed beyond the calibration stage of the experiment.

Pavlovian Learning

Participants learned about eight different stimuli (aliens) and their probabilistic association with reward or loss outcomes. Four of the aliens were probabilistically associated with reward and four were probabilistically associated with loss. The outcome probabilities and magnitudes were arranged such that each stimulus was associated with a unique combination of probability and magnitude (see **Supplementary Table 5**). The reward and loss aliens were presented in separate blocks and the trial order in each block was randomised once, such that all participants experienced the same order of trials. Each alien was presented a total of 16 times so participants completed 128 trials across the two blocks. On each trial, participants were shown one of the eight aliens and

given an outcome (here points that were not related to participants' payment for participation in the study). They were then asked to use a slider to estimate the outcome probability that this alien gives you or takes away points (i.e. the reward or loss probability of the alien; see **Figure 1 Panel A** for a schematic of the learning trials). This permitted the explicit measure of trial-by-trial learning across encounters with a given alien, without the need for instrumental responses that might confound subsequent measures of choice in later stages of the experiment. As participants encountered reward and loss aliens in separate blocks, the order of these blocks was counterbalanced across the sample, half of participants learned about the reward aliens first and the other half of participants learned about the loss aliens first.

At the end of the final learning block, participants were asked to provide a final estimate of the reward probability of all eight stimuli.

Tournament

After the learning phase, participants were asked to complete a 'tournament' in which they made binary decisions between pairs of the eight aliens. All possible pair combinations of the stimuli were presented five times in a randomised order, with all participants experiencing the same trial order. Participants were instructed to choose between the alien pairs based on which one they felt was 'best' from the learning phase. Participants had three seconds to respond to each pair using the left and right arrow keys to choose the left and right stimuli respectively. If participants did not respond within the three second limit, this was recorded as no response and the next pair would appear on the screen immediately (see **Figure 1 Panel B**).

At the end of the tournament, participants were again asked to provide a final estimate of the reward probability of all eight stimuli.

Effort-based decision-making

Then participants completed an effort-based decision-making task (**Figure 1 Panel C**). On each trial, participants chose to 'accept' or 'reject' the opportunity to exert effort to approach or avoid one of the eight aliens they had previously learned about. Prior to

starting the effort-based decision-making, participants are reminded to press the spacebar with their right and left index fingers as they did during the calibration task. Here effort is defined as the number of keypresses required to move a spaceship to reach a 'trade' (approach) or 'safety' (avoid) zone (see screen two in **Figure 1 Panel C**). Participants were presented with one of three effort levels (60%, 75% and 90% of each participant's maximum number of key presses), represented on a horizontal bar in the middle of the screen showing the distance they would be required to travel. For approach trials, the alien was presented next to the 'trade' zone, indicating this was an approach trial. Participants were instructed that accepting an effort offer on this trial meant agreeing to exert effort for the chance to trade with the alien located in the 'trade' zone. Whereas for avoid trials, the learned stimulus was presented on the left side of the effort bar, indicating that it was an avoid trial. On these trials, participants were instructed that accepting an effort offer here was like taking the opportunity to avoid having points stolen by the alien. Critically, no information was presented about the reward/loss probability or magnitude associated with each alien, requiring participants to make choices based on the effort level and *what they learned about the stimuli previously*.

When participants accepted an effort offer, they progressed to a second screen and had six seconds to meet the effort requirement, the six second countdown started on their first button press. Participants could see their spaceship progress toward the 'trade' or 'safety' zone and therefore, determine whether they had met the effort requirement based on whether their spaceship had passed the green marker on the effort bar (see screen two in **Figure 1 Panel C**). In both approach and avoid trials, participants were only told whether they had met the effort requirement or not. This prevented participants from continuing to learn about the reward/loss probabilities of the aliens. On trials where participants rejected the effort offer, participants had the opportunity to rest for a duration equivalent to the effort exertion window (six seconds). Participants completed 96 effort choices in total, which ensured that each combination of effort level and stimulus was presented four times in total in a pre-randomised order and the same order for all participants.

At the end of the effort-choice task, participants were again asked to provide a final estimate of the reward probability of all eight stimuli.

 Transdiagnostic and self-report mental health measures

 At the end of the experiment participants completed a shortened version of the

 transdiagnostic three factor questionnaire[59] which contains questions relating to apathy,

 the Snaith-Hamilton Pleasure Scale (SHAPS)[60] , a measure of anhedonia which we

 calculate as a sum score to increase its fidelity as a predictor[61] and the Fatigue

 Assessment Scale (FAS)[62] to measure fatigue.

 Analysis

 Measuring learning

 With the inclusion of trial-by-trial probability estimates, final post-task probability

 estimates and tournament performance, there are several ways to quantify how

 participants learned the outcome probabilities associated with each alien. Analyses

 were conducted across a range of measures (outlined below) to confirm that, on average,

 participants were capable of learning the stimulus-outcome probabilities in this task.

 Mean trial-by-trial probability estimates: here we calculate the mean of a participant's

 probability ratings for each alien (16 trials) of the learning task.

 Final learning probability estimates: here we take the last estimate participants gave for

 each stimulus during the learning task. The trial order of the learning was randomised for

 the reward and loss blocks, meaning the final learning trials for each stimulus were not

 the last eight trials of the learning.

 Post-task ratings: here we took the explicit rating provided for each stimulus at the end of

 a given task (e.g. post-learning, post-tournament and post-effort choice). This allowed us

 to examine whether learning persisted or changed across different phases of the

 experiment.

 Tournament performance: To assess choice accuracy, we calculated the expected value

 (EV) of each alien by multiplying its outcome magnitude (points) by its outcome

 probability. The correct choice for a given pair was the alien with the higher EV. We also

 examined approach accuracy (choosing the highest-EV alien when present) and avoid

593 accuracy (avoiding the lowest-EV alien when present)[25]. Additionally, we measured
594 accuracy in choices between only reward or only loss stimuli ('within-valence'
595 comparisons).

Logistic regression predicting trial-by-trial choice

To first establish that participants' effort choices were sensitive to the effects of effort,
outcome probability and outcome magnitude, a baseline logistic regression was used to
predict participants' trial-by-trial choice data with effort level, outcome probability and
outcome magnitude, learning order, stimulus valence, age, gender and socioeconomic
status as predictors. Learning order was a binary predictor included to measure a
possible effect of block order on effort choice and valence was also a binary predictor
included as this design presents a unique opportunity to compare how people make
effort choices to approach reward and avoid loss when outcome probabilities and
absolute outcome magnitudes are equivalent. Gender uses male (including trans male)
as a reference category, the remaining levels were: female (including trans female), non-
binary and gender not specified. These models were fit using R (R version 4.4.1; R Core
Team, 2024; https://www.R-project.org/) and *stats* package (version 4.4.1; R Core Team,
2024).

$$\mathbf{\textit{M}0}: \; \textit{choice} \sim \textit{outcome probability} + \textit{outcome magnitude} + \textit{effort level}$$
$$+ \textit{learning order} + \textit{stimulus valence} + \textit{age} + \textit{gender} + \textit{SES}$$

Generalised linear mixed effects modelling

The simple logistic regression model above assumes that all participants learn perfectly
(i.e. they learn the exact outcome probabilities of each stimulus) and identically to one
another (i.e. the logistic regression treats all participants as one participant). To capture
individual differences in learning, a set of models which included different individualised
measures of participants' learning were also fit to determine whether these individual
measures of learning better explain participants' effort choices (refer to **Supplementary**
Table 3 for a full list of models and see the **Supplemental Materials** for details of the
measures tested). We allowed formal model comparison to arbitrate between different
individual learning signatures as predictors of effort-choice, and report details of the

winning model in the results section. Here we used AIC which punishes less for model

complexity than BIC to identify winning models but AIC scores, along with BIC scores, log

likelihood and deviance are also reported in the **Supplementary Materials**. Bold text

indicates where models have changed from the previous model:

$\textbf{\textit{M1}}$: $choice \sim \textbf{\textit{individual learning measure}} + outcome\ magnitude$

627           $+ effort\ level + stimulus\ valence + age + gender + (1$

628           $+ \textbf{\textit{individual learning measure}} + outcome\ magnitude$

629           $+ effort\ level + stimulus\ valence\ |\ participant)$

Lastly, in this general population sample, we were interested in whether participants'

effort choices were related to a set of relevant psychiatric dimensions and clinical

symptoms (e.g. three transdiagnostic factor scores comprising anxious-depression,

compulsivity and intrusive thought and social withdrawal; in addition to anhedonia and

fatigue). We therefore fit the following suite of models (bold text indicates where models

have changed from the previous model):

$\textbf{\textit{M2}}$: $choice \sim individual\ learning\ measure\ + outcome\ magnitude + effort\ level$

637           $+ stimulus\ valence + age + gender + \textbf{\textit{AD}} + \textbf{\textit{Compul}} + \textbf{\textit{SW}} + (1$

638           $+ individual\ learning\ measure\ + outcome\ magnitude + effort\ level$

639           $+ stimulus\ valence\ |\ participant)$

640 $\textbf{\textit{M3}}$: $choice \sim individual\ learning\ measure\ + outcome\ magnitude + effort\ level$

641           $+ stimulus\ valence + age + gender + \textbf{\textit{SHAPS}} + (1$

642           $+ individual\ learning\ measure\ + outcome\ magnitude + effort\ level$

643           $+ stimulus\ valence\ |\ participant)$

644 $\textbf{\textit{M4}}$: $choice \sim individual\ learning\ measure\ + outcome\ magnitude + effort\ level$

645           $+ stimulus\ valence + age + gender + \textbf{\textit{FAS}} + (1$

646           $+ individual\ learning\ measure\ + outcome\ magnitude + effort\ level$

647           $+ stimulus\ valence\ |\ participant)$

648 To improve mixed model fitting and to aid interpretation and comparison of estimates

649 generated by these models, questionnaire scores were rescaled and zero-centred. As all

650 values of effort level and outcome probability ranged between 0 and 1, outcome

magnitude was min-max normalised with a minimum of 0.2 and a maximum of 0.8, such that the lowest absolute outcome magnitude became 0.2 and the highest absolute outcome magnitude became 0.8 thus preserving the relationship between the original absolute magnitudes of 40 and 160 (i.e. the large outcome magnitude being four times the small outcome magnitude).

References

1. Hull, C. L. *Principles of Behavior, an Introduction to Behavior Theory*. (London: Appleton-Century Company, New York, NY, 1943).

2. Zipf, G. K. *Human Behavior and the Principle of Least Effort: An Introduction to Human Ecology*. (Ravenio Books, 1949).

3. Pessiglione, M., Vinckier, F., Bouret, S., Daunizeau, J. & Le Bouc, R. Why not try harder? Computational approach to motivation deficits in neuro-psychiatric diseases. *Brain* **141**, 629–650 (2018).

4. Westbrook, A., Kester, D. & Braver, T. S. What Is the Subjective Cost of Cognitive Effort? Load, Trait, and Aging Effects Revealed by Economic Preference. *PLOS ONE* **8**, e68210 (2013).

5. Wu, R., Ferguson, A. M. & Inzlicht, M. Do humans prefer cognitive effort over doing nothing? *Journal of Experimental Psychology: General* **152**, 1069–1079 (2023).

6. Husain, M. & Roiser, J. P. Neuroscience of apathy and anhedonia: a transdiagnostic approach. *Nat Rev Neurosci* **19**, 470–484 (2018).

7. Bonnelle, V. *et al*. Characterization of reward and effort mechanisms in apathy. *Journal of Physiology-Paris* **109**, 16–26 (2015).

8. Schmidt, L., Palminteri, S., Lafargue, G. & Pessiglione, M. Splitting Motivation: Unilateral Effects of Subliminal Incentives. *Psychol Sci* **21**, 977–983 (2010).

9. Jurgelis, M. *et al*. Heightened effort discounting is a common feature of both apathy and fatigue. *Sci Rep* **11**, 22283 (2021).

10. Mehrhof, S. Z. & Nord, C. L. A common alteration in effort-based decision-making in apathy, anhedonia, and late circadian rhythm. *eLife* **13**, (2024).

11. Cléry-Melin, M.-L. *et al.* Why Don't You Try Harder? An Investigation of Effort

Production in Major Depression. *PLOS ONE* **6**, e23178 (2011).

12. Klein-Flügge, M. C., Kennerley, S. W., Saraiva, A. C., Penny, W. D. & Bestmann, S.

Behavioral Modeling of Human Choices Reveals Dissociable Effects of Physical Effort

and Temporal Delay on Reward Devaluation. *PLOS Computational Biology* **11**,

e1004116 (2015).

13. Meyniel, F. *et al.* A specific role for serotonin in overcoming effort cost. *eLife* **5**, e17282

(2016).

14. Müller, T., Klein-Flügge, M. C., Manohar, S. G., Husain, M. & Apps, M. A. J. Neural and

computational mechanisms of momentary fatigue and persistence in effort-based

choice. *Nat Commun* **12**, 4593 (2021).

15. Treadway, M. T., Buckholtz, J. W., Schwartzman, A. N., Lambert, W. E. & Zald, D. H.

Worth the 'EEfRT'? The Effort Expenditure for Rewards Task as an Objective Measure

of Motivation and Anhedonia. *PLOS ONE* **4**, e6598 (2009).

16. Treadway, M. T., Bossaller, N. A., Shelton, R. C. & Zald, D. H. Effort-based decision-

making in major depressive disorder: A translational model of motivational

anhedonia. *Journal of Abnormal Psychology* **121**, 553–558 (2012).

17. Kuhn, M. *et al.* Computational Phenotyping of Effort-Based Decision Making in

Unmedicated Adults with Remitted Depression. *Biological Psychiatry: Cognitive*

Neuroscience and Neuroimaging (2025) doi:10.1016/j.bpsc.2025.02.006.

18. Browning, M., Behrens, T. E., Jocham, G., O'Reilly, J. X. & Bishop, S. J. Anxious

individuals have difficulty learning the causal statistics of aversive environments. *Nat*

Neurosci **18**, 590–596 (2015).

19. Dubois, M. & Hauser, T. U. Value-free random exploration is linked to impulsivity. *Nat*

Commun **13**, 4542 (2022).

20. Gibbs-Dean, T. *et al.* Belief updating in psychosis, depression and anxiety disorders:

A systematic review across computational modelling approaches. *Neuroscience &*

Biobehavioral Reviews **147**, 105087 (2023).

21. Pulcu, E. & Browning, M. The Misestimation of Uncertainty in Affective Disorders.

Trends in Cognitive Sciences **23**, 865–875 (2019).

22. Sandhu, T. R., Xiao, B. & Lawson, R. P. Transdiagnostic computations of uncertainty:

towards a new lens on intolerance of uncertainty. *Neuroscience & Biobehavioral*

Reviews **148**, 105123 (2023).

23. Rouhani, N., Norman, K. A. & Niv, Y. Dissociable effects of surprising rewards on

learning and memory. *J Exp Psychol Learn Mem Cogn* **44**, 1430–1443 (2018).

24. Rouhani, N. & Niv, Y. Signed and unsigned reward prediction errors dynamically

enhance learning and memory. *eLife* **10**, e61077 (2021).

25. Frank, M. J., Seeberger, L. C. & O'Reilly, R. C. By Carrot or by Stick: Cognitive

Reinforcement Learning in Parkinsonism. *Science* **306**, 1940–1943 (2004).

26. Lawson, R. P. *et al.* The habenula encodes negative motivational value associated

with primary punishment in humans. *Proceedings of the National Academy of*

Sciences **111**, 11858–11863 (2014).

27. Lawson, R. P. *et al.* Disrupted habenula function in major depression. *Mol Psychiatry*

22, 202–208 (2017).

28. Kessler, R. C. *et al.* Anxious and non-anxious major depressive disorder in the World

Health Organization World Mental Health Surveys. *Epidemiology and Psychiatric*

Sciences **24**, 210–226 (2015).

29. Hezemans, F. H., Wolpe, N. & Rowe, J. B. Apathy is associated with reduced precision

of prior beliefs about action outcomes. *Journal of Experimental Psychology: General*

149, 1767 (2020).

30. Nord, C. L. *et al.* Vigour in active avoidance. *Sci Rep* **7**, 60 (2017).

31. Stein, M. B. & Paulus, M. P. Imbalance of Approach and Avoidance: The Yin and Yang

of Anxiety Disorders. *Biological Psychiatry* **66**, 1072–1074 (2009).

32. Palminteri, S., Lefebvre, G., Kilford, E. J. & Blakemore, S.-J. Confirmation bias in

human reinforcement learning: Evidence from counterfactual feedback processing.

PLOS Computational Biology **13**, e1005684 (2017).

33. Kahneman, D. & Tversky, A. Prospect Theory: An Analysis of Decision Under Risk. in

World Scientific Handbook in Financial Economics Series vol. 4 99–127 (WORLD

SCIENTIFIC, 2013).

34. Rzepa, E., Fisk, J. & McCabe, C. Blunted neural response to anticipation, effort and

consummation of reward and aversion in adolescents with depression

symptomatology. *J Psychopharmacol* **31**, 303–311 (2017).

35. Lockwood, P. L. *et al.* Prosocial apathy for helping others when effort is required. *Nat*

Hum Behav **1**, 0131 (2017).

36. Bishop, S. J. & Gagne, C. Anxiety, Depression, and Decision Making: A Computational

Perspective. *Annual Review of Neuroscience* **41**, 371–388 (2018).

37. Huang, H., Thompson, W. & Paulus, M. P. Computational Dysfunctions in Anxiety:

Failure to Differentiate Signal From Noise. *Biological Psychiatry* **82**, 440–446 (2017).

38. Wise, T. & Dolan, R. J. Associations between aversive learning processes and

transdiagnostic psychiatric symptoms in a general population sample. *Nat Commun*

11, 4179 (2020).

39. Bedder, R. L., Pisupati, S. & Niv, Y. Modelling Rumination as a State-Inference Process.

Proceedings of the Annual Meeting of the Cognitive Science Society **45**, (2023).

40. Bustamante, L. A. *et al.* Major depression symptom severity associations with

willingness to exert effort and patch foraging strategy. *Psychological Medicine* **54**,

4396–4407 (2024).

41. Gagne, C., Agai, S., Ramiro, C., Dayan, P. & Bishop, S. Biased belief priors versus

biased belief updating: Differential correlates of depression and anxiety. *PLOS*

Computational Biology **18**, e1010176 (2022).

42. Lynch, C. J., Gunning, F. M. & Liston, C. Causes and Consequences of Diagnostic

Heterogeneity in Depression: Paths to Discovering Novel Biological Depression

Subtypes. *Biological Psychiatry* **88**, 83–94 (2020).

43. Robinson, O. J. & Chase, H. W. Learning and Choice in Mood Disorders: Searching for

the Computational Parameters of Anhedonia. *Computational Psychiatry* **1**, 208–233

(2017).

44. Caspi, A. *et al.* The p Factor: One General Psychopathology Factor in the Structure of

Psychiatric Disorders? *Clinical Psychological Science* **2**, 119–137 (2014).

45. Cuthbert, B. N. & Insel, T. R. Toward the future of psychiatric diagnosis: the seven

pillars of RDoC. *BMC Med* **11**, 126 (2013).

46. Gillan, C. M., Kosinski, M., Whelan, R., Phelps, E. A. & Daw, N. D. Characterizing a

psychiatric symptom dimension related to deficits in goal-directed control. *eLife* **5**,

e11305 (2016).

47. Patzelt, E. H., Kool, W., Millner, A. J. & Gershman, S. J. The transdiagnostic structure

of mental effort avoidance. *Sci Rep* **9**, 1689 (2019).

48. Ang, Y.-S., Gelda, S. E. & Pizzagalli, D. A. Cognitive effort-based decision-making in

major depressive disorder. *Psychological Medicine* **53**, 4228–4235 (2023).

49. Wise, T., Robinson, O. J. & Gillan, C. M. Identifying Transdiagnostic Mechanisms in

Mental Health Using Computational Factor Modeling. *Biological Psychiatry* **93**, 690–

703 (2023).

50. Foa, E. B. *et al.* The Obsessive-Compulsive Inventory: Development and validation of

a short version. *Psychological Assessment* **14**, 485–496 (2002).

51. Marin, R. S., Biedrzycki, R. C. & Firinciogullari, S. Reliability and validity of the apathy

evaluation scale. *Psychiatry Research* **38**, 143–162 (1991).

52. Liebowitz, M. R. Social Phobia. (1987) doi:10.1159/000414022.

53. Zung, W. W. K. A Self-Rating Depression Scale. *Archives of General Psychiatry* **12**, 63–

70 (1965).

54. Spielberger, C., Gorsuch, R., Lushene, R., Vagg, P. & Jacobs, G. *Manual for the State-*

Trait Anxiety Inventory. (Consulting Psychologists Press., 1983).

55. Valton, V. *et al.* A computational approach to understanding effort-based decision-

making in depression. *bioRxiv* 2024.06.17.599286 (2024)

doi:10.1101/2024.06.17.599286.

56. Müller, T., Milton, J., Husain, M. & Apps, M. A. Computational signatures of exertion

and rest underlie moment-to-moment dynamics of subjective perceptions of effort

and fatigue. (2025).

57. Gagne, C., Zika, O., Dayan, P. & Bishop, S. J. Impaired adaptation of learning to

contingency volatility in internalizing psychopathology. *eLife* **9**, e61387 (2020).

58. Huys, Q. J., Pizzagalli, D. A., Bogdan, R. & Dayan, P. Mapping anhedonia onto

reinforcement learning: a behavioural meta-analysis. *Biol Mood Anxiety Disord* **3**, 12

(2013).

59. Hopkins, A. K., Gillan, C., Roiser, J., Wise, T. & Sidarus, N. Optimising the

measurement of anxious-depressive, compulsivity and intrusive thought and social

withdrawal transdiagnostic symptom dimensions. Preprint at

https://doi.org/10.31234/osf.io/q83sh (2022).

60. Snaith, R. P. *et al*. A scale for the assessment of hedonic tone the Snaith-Hamilton

Pleasure Scale. *Br J Psychiatry* **167**, 99–103 (1995).

61. Franken, I. H. A., Rassin, E. & Muris, P. The assessment of anhedonia in clinical and

non-clinical populations: Further validation of the Snaith–Hamilton Pleasure Scale

(SHAPS). *Journal of Affective Disorders* **99**, 83–89 (2007).

62. Shahid, A., Wilkinson, K., Marcu, S. & Shapiro, C. M. Fatigue Assessment Scale (FAS).

in *STOP, THAT and One Hundred Other Sleep Scales* (eds Shahid, A., Wilkinson, K.,

Marcu, S. & Shapiro, C. M.) 161–162 (Springer New York, New York, NY, 2011).

doi:10.1007/978-1-4419-9893-4_33.

Learning phase
A

Tournament phase
B

Effort decision-making phase
C

Figure 1: Learning, tournament and effort-decision making phases of the experiment (A) An example

trial from the learning phase of the experiment. Participants first view a fixation cross, then one of the eight

stimuli is shown on screen before the outcome of the trial appears on the next screen. For reward stimuli,

when the alien gives points, a positive tone is played when the amount appears on screen and for loss trials,

a negative tone is played when the alien takes points. For both alien types, a neutral tone is played when

the outcome is zero. At the end of each trial, participants are asked to estimate the percentage of trials that

the alien is giving or taking points from them, this estimate is reported using a slider. **(B)** An example trial

from the tournament. Participants see a brief fixation cross and are then presented with two of the eight

learned stimuli and asked to make a choice between them in 3000ms using the left and right arrow keys on

their keyboard. If no choice is made this is recorded as no response. If participants make a choice in under

3000ms the next pair of stimuli appear on the screen after a brief delay. **(C)** An example trial from the effort

decision-making phase where the participant accepts the offer and meets/exceeds the effort requirement.

On each effort trial, participants are presented with one of the eight stimuli they learned about previously

and an effort amount visualised as the distance from a Trade/Safety zone on a horizontal bar. Participants

are required to accept or reject based on the effort required and what they remember about the outcome

probability and magnitude associated with the given stimulus. When accepting the offer, participants are

taken to a screen which includes a countdown timer that counts down to zero from 6000ms as soon as the

participant makes the first spacebar press. Rejecting the effort offer takes participants to a waiting screen

that also last for 6000ms.

Figure 2: Summary of performance in the learning phase of the task (A) The average of all participants' probability estimates across the learning trials. Here we take the mean rating of all participants' estimates for each stimulus on each trial (coloured lines) to visualise the trajectory of participants' learning and how, across trials, their ratings tend toward the true outcome probability (dashed black line). **(B)** Box and scatter plots of participants' final learning estimate for each stimulus. The stimuli have been ordered to group the low probability (p = 0.25) and high probability (p = 0.75) stimuli. Green indicates reward stimuli and red indicates loss stimuli. Participants tended towards the objective outcome probability but there was also notable variability in participants' estimates. **(C)** Here we take the mean post-tournament probability estimate of each participant and subtract the true average outcome probability from it (0.5), such that participants who on average, overestimate the true outcome probabilities are above the yellow line. This average error indicates there is significant variability in learning fidelity across participants.

Figure 3: Probability of accepting as a function of effort level, outcome magnitude and valence-probability. **(A)** A 3D bar plot showing how the probability of accepting effort varies as a function of effort level and the amount of points associated with a stimulus. The probability of accepting was highest for the highest outcome magnitude stimuli combined with the lowest effort and lowest for the low magnitude loss stimuli combined with the highest effort level. For each level of outcome magnitude, p(accept) decreased as effort level increased. **(B)** A 3D bar plot showing how the probability of accepting effort varies as a function of effort level and outcome probability for both reward and loss aliens. The probability of accepting was highest for the high probability reward stimuli combined with the lowest effort level and lowest for the low magnitude loss stimuli combined with the highest effort level. For each stimulus valence and outcome probability combination, p(accept) decreased as effort level increased.

Figure 4: Bar plot of fixed effect estimates and histograms of the random effects from M1. (A) Bar plot of the fixed effect estimates from Model 1 (M1), the winning model amongst those fit to compare the performance of individualised learning measures against the objective outcome probabilities. Each bar represents the fixed effect estimate for a predictor from the winning model (M1). Error bars represent the 95% confidence interval estimates for the predictor. Unlike in Model 0 (M0), in the GLMMs, none of the demographic predictors were significantly non-zero predictors of trial-by-trial effort choice. Post-tournament rating, outcome magnitude, effort level and stimulus valence were all significantly non-zero predictors ($p < 0.05$). **(B)** Histograms of the random slope estimates for each of the predictors in the winning model, Model 1 (M1). For example, where the fixed effect estimate for effort level is zero, each participant's slope estimate for the effect of effort level on choice is expressed as a difference from zero. Therefore, participants with a positive slope estimate for effort level are more sensitive to effort level when making effort choices. The slopes for each random effect are approximately normally distributed.

Figure 5: Bar plot of the fixed effects and histograms of the random effects from M3. **(A)** Bar plot of the fixed effect estimates from the overall best-fitting model, which included participants' SHAPS sum scores as a fixed effect predictor (Model 3). Each bar represents the fixed effect estimate for a predictor from the winning GLME model (Model 3), error bars are 95% confidence intervals. All predictors are significant ($p < 0.05$). **(B)** Histograms of the random slopes estimated for each participant from the winning model (Model 4). The slopes fit for each participant are expressed as a difference from zero, where zero is the fixed effect estimate for the predictor. So, for post-tournament rating, a participant with a positive slope estimate suggests that the participant's effort choices are more strongly influenced by their own probability estimate of the stimulus. For each predictor, the slope estimates are approximately normally distributed.

Figure 6: The interaction between SHAPS scores and participants' post-tournament probability
estimates. (A) Predictions from our winning interaction model (EM1) that visualise the effect of the
interaction between participants' post-tournament probability estimates and SHAPS scores on the
probability of accepting effort offers. Compared to participants scoring lower on the SHAPS (cream colour),
high SHAPS scorers' (red colour) effort choices are more weakly related to their estimates of the stimuli's
outcome probability. Lower scorers on the SHAPS show a sharper increase in probability of accepting as
their estimate of the outcome probability increases than participants who score higher on the SHAPS. **(B)**
A bar plot of the interaction effect estimates for high, low and mean scorers on the SHAPS. These estimates
were calculated by adding and subtracting the estimate for the interaction effect from the fixed effect
estimate for participants' post-tournament estimates from the winning interaction model (EM1).