**The macroecology of viral coinfection**

Cecilia A. Sánchez[1], Colin J. Carlson[1], Amy R. Sweeny[2*]

[1]Department of Epidemiology of Microbial Diseases, Yale School of Public Health, New Haven,

CT, USA

[2]Institute of Ecology and Evolution, University of Edinburgh, Edinburgh, UK

[*]Corresponding author: Amy R. Sweeny. Email: amyr.sweeny@gmail.com

Keywords: coinfection, disease ecology, wildlife surveillance, macroecology, viral dynamics

**Abstract**

Coinfection is common in wild animals, and can profoundly influence disease outcomes and

transmission. However, most coinfection research is based on laboratory experiments or a few

well-studied wildlife systems. Here, we use data from the PREDICT project – the largest

standardized wildlife disease surveillance project ever conducted – to describe patterns of viral

coinfection and evaluate their association with host and virus factors. Within the viruses

prioritized for testing, we find that coinfection is rare (detected in just 223 of 65,662 animals),

but still more common than expected by chance – especially coinfections of coronaviruses,

paramyxoviruses, and influenza A virus. We further find that coinfection is associated with host

age, but bats and rodents exhibit opposing relationships. We find that captive wild animals

(specifically rats and mallards) exhibit higher coinfection than free-ranging wild animals,

highlighting potential risks to wildlife and human health from the wildlife trade. Finally, we find

that biases in sampling and testing can shape observed patterns of coinfection. Our results

characterize associations among viruses in the largest relevant dataset currently available.

Factors associated with coinfection in this dataset such as host-level variation, virus-virus

interactions, and human interference should be further tested via surveillance and laboratory

approaches to better resolve the role of coinfection as an important and likely overlooked driver

of viral dynamics in nature.

**Introduction**

Parasites (including microparasites and macroparasites) exist in complex communities, where coinfection with multiple parasites at any time is the norm for wild animals (Cox 2001; Petney and Andrews 1998). Coinfections provide an opportunity for direct interactions between parasites, such as competition for resources (Budischak et al. 2017) or space (Knowles et al. 2013), or indirect interactions, such as those mediated by host immune responses (Ezenwa et al. 2021). These interactions can have a dramatic effect on outcomes at the host level, and transmission dynamics at the population level (Telfer et al. 2010; Gorsich et al. 2018; Pedersen and Fenton 2007; Graham 2008).

Several theoretical frameworks in disease ecology have incorporated the dynamics and consequences of coinfection (Pedersen and Fenton 2007; Fenton and Pedersen 2005; Graham 2008; Sweeny et al. 2021), but it can be difficult to test their predictions against empirical evidence. Within-host processes are easiest to study in laboratory settings (Fenton et al. 2014), but the controlled settings that make experimental approaches useful (e.g., laboratory organisms from a single host species, with known ancestry and no unknown parasites) also limit their similarity to real-world dynamics in natural communities. A few well-developed (and often long-term) field systems have also led to a number of insights about interactions within diverse host and parasite communities (Ezenwa et al. 2021), and some systems allow for natural experiments such as parasite removal or nutritional supplementation (Sweeny et al. 2020; Knowles et al. 2013; Ezenwa and Jolles 2015). However, neither field nor experimental case studies can provide insight into the full network of interactions occurring within parasite communities (or their consequences) at the ecosystem scale or above.

Whereas large-scale biodiversity data have reached volumes that allow for unprecedented

research and monitoring, wildlife disease data collection has historically been fragmented

between emerging infectious disease surveillance, wildlife management efforts, and ecological

research. Across all three fields, there are significant global gaps in surveillance (Carlson et al.

2025; Cohen et al. 2023), and – despite broader progress towards open science in ecology and

epidemiology – the sharing of spatially-explicit, host-level infection data remains rare

(Schwantes et al. 2025; Albery et al. 2022). Some studies have used data synthesis to develop

global infection datasets (Cohen et al. 2023, 2020; Heckley and Becker 2024; Warmuth et al.

2023), but in most cases, they have focused on single-host, single-parasite relationships, rather

than single-host, multi-parasite dynamics.

Here, we explore coinfection dynamics at the global scale for the first time, using the

single largest standardized study of wildlife disease ever conducted. The PREDICT project was a

decade-long, multi-country initiative funded by the U.S. Agency for International Development

(USAID) that conducted viral surveillance primarily in wild animals (PREDICT Consortium

2020). Focused on virus families of public health concern but capturing a sizable fraction of

wildlife virus diversity, the project produced one of the largest standardized multi-virus datasets

ever assembled and resulted in the discovery of new viruses (Anthony et al. 2017b; Goldstein et

al. 2018), an improved understanding of drivers of virus infection and diversity (Anthony et al.

2017a; Eskew et al. 2025), and insight into human and virus-related risk factors for zoonotic

spillover (Saylors et al. 2021). However, minimal work has been done using the PREDICT data

to test macroecological hypotheses. Using these data, we set out to explore three questions: (1)

How common is viral coinfection? (2) Is coinfection more common in certain host groups, life

stages, and settings? And finally: (3) Are there non-random patterns of coinfection driven by

specific virus taxa, or by associations between specific pairs of virus taxa? Through this work,

we also aimed to identify promising avenues for future research on viral coinfection.

**Methods**

*Data cleaning*

Analyses were performed using the R statistical environment v4.5.1 (R Core Team 2025). All

steps of the analysis pipeline (data cleaning, statistical modeling, figure generation) can be

reproduced using the code available at https://github.com/viralemergence/macroecology-

coinfection. Data from the PREDICT-2 project (2015-2019; the second half of the decade-long

initiative) were openly shared on the USAID Development Data Library prior to the takedown of

the USAID website. Our analyses relied on two PREDICT datasets: an animal sampling dataset

containing details such as species, sex, and age class and a second dataset containing virus test

results for individual specimens (which also included some basic host information). We cleaned

the test result dataset by excluding human results and fixing minor inconsistencies in animal host

taxonomy (e.g. host scientific name was mostly provided at the species level, but sometimes only

at the genus or family level). We modified the animal taxonomic groups originally provided by

joining "birds" and "poultry/other fowl" into one "birds" group, and separating a

"rodents/shrews" group into "rodents" and "shrews" groups. Two detections of the hantavirus

Thottapalayam virus were originally classified as belonging to the Bunyaviridae family; we

reclassified them to the Hantaviridae family to align with current virus taxonomy(Kuhn et al.

2024). Finally, we joined the two datasets so that each test result was paired with the relevant

animal data.

*Data overview*

Specimen collection ($n$ = 126,952 unique specimens from 65,662 unique animals) was

concentrated in countries in West and Central Africa as well as South and Southeast Asia, with >

10,000 specimens collected in Sierra Leone, Guinea, and Bangladesh each (Fig. S1). Most

specimens were collected from bats ($n$ = 77,334, 60.9%) and rodents ($n$ = 25,683, 20.2%). The

dataset included some farm and companion animals, including poultry, dogs, and swine. Virus

screening was conducted using consensus polymerase chain reaction (PCR) assays (714,954 total

tests), mostly targeting virus families of significance to human health (e.g., coronaviruses,

filoviruses, and paramyxoviruses), but in some cases testing for specific viruses of interest (e.g.,

influenza viruses and ebolaviruses). Further details on PREDICT can be found in the PREDICT

Legacy Book (https://ohi.vetmed.ucdavis.edu/programs-projects/predict-project/reports).

*Analyses*

Across all animals and within each animal taxonomic group (16 total groups; Table S1), we

calculated infection prevalence (number infected divided by number tested) and coinfection

prevalence (number coinfected divided by number tested). We considered an animal to be

infected when one or more test results for specimens collected from that animal were positive for

at least one virus, and to be coinfected when test results were positive for two or more unique

viruses. For bats, rodents, and birds (the three largest taxonomic groups), we compared the

observed coinfection prevalence to the expected pairwise prevalence based on the prevalence of

each virus, using a permutation testing approach. Specifically, we randomized infection across

individual animals while maintaining the observed prevalence of each virus within the taxonomic

group, then calculated the number of simulated coinfected animals. We repeated this procedure

1000 times for each taxonomic group, then plotted the simulated number of coinfected animals

compared to the number observed in the PREDICT data.

To explore patterns of coinfection among virus taxa, we created a coinfection matrix at

the virus family level, where each count within the matrix represents a coinfection of two virus

families occurring in the same animal. We included virus family coinfections that represented

detections of two viruses in the same family (e.g. two coronaviruses) or in different families (e.g.

one coronavirus and one paramyxovirus). For animals infected with three viruses, we counted

each coinfection separately (e.g., an animal infected with coronavirus X, coronavirus Y, and

paramyxovirus Z would count for two coronavirus-paramyxovirus coinfections and one

coronavirus-coronavirus coinfection). We used the *ggraph* package (Pedersen 2024) to visualize

the frequency of virus family coinfection. To test if certain virus families were co-detected more

often than expected by chance, we first calculated the expected coinfection prevalence of two

virus families as the product of each virus family's overall prevalence in the dataset. We then

performed a binomial test for each virus family–virus family combination to determine if the

observed coinfection prevalence was significantly greater than expected. To account for multiple

comparisons, we adjusted *P* values using a Bonferroni correction. Finally, we repeated the above

analyses for two subsets of the full data: bats and rodents. We also created a coinfection matrix at

the individual virus level and visually explored coinfection patterns; however, due to high

sparsity in this matrix, we did not analyze these data further.

For each infected animal *(n = 3,271)*, we quantified the number of unique viruses

detected in its test results and classified its coinfection status as either single infection (exactly

one virus detected) or coinfection (two or more viruses detected). We then used generalized

linear models to explore correlates of coinfection status. For all model datasets, we excluded

taxonomic groups with ≤ 10 animals (carnivores, cattle/buffalo, dogs, goats/sheep, other),

animals that were missing sex or age class data, and neonates and fetuses ($n$ = 809 total animals

excluded). **1)** For an all-animals model ($n$ = 2,462), we used the 'glm' function in the *stats*

package (R Core Team 2025) to model animal coinfection status as a function of geographic

region (Africa and West Asia; South, East, and Southeast Asia), taxonomic group (bats, birds,

rodents, shrews), sex (female, male), age class (adult, subadult, juvenile), captivity status (wild

free-ranging, owned domesticated, wild in captivity), and the number of virus families tested for

across all specimens associated with that animal. Given inconsistency in host identification (i.e.,

not all animals were identified to species level), we did not include any species-level traits as

model predictors. **2)** For a bats-only model *(n* = 1,519), we classified whether each bat species is

known to roost in caves using data from Willoughby *et al.* (Willoughby et al. 2017) and the

IUCN Red List. We excluded bats for which cave roosting behavior was not available and bats

not identified to at least family level. Using the 'glmmTMB' function in the *glmmTMB* package

(Brooks et al. 2017), we modeled coinfection status as a function of five fixed effects variables

(geographic region, sex, age class, cave roosting status, number of virus families tested), and one

random effects variable (host family). **3)** For a rodents-only model (n = 488), we modeled

coinfection status as a function of sex, age class, and number of virus families tested. Subadults

and juveniles were collapsed into one age class category due to sample size restrictions, while

host family and geographic region were not included as variables because nearly all rodents were

murids and were collected in South, East, and Southeast Asia.

**Results**

We found that both infection and coinfection were rare overall for viruses prioritised in

sampling. Out of 65,662 unique animals, 3,271 (4.98%) were infected with one or more viruses

(Fig. 1A); virus prevalence in birds (487/5,617; 8.67%) and bats (2,059/36,181; 5.69%) was

somewhat higher than average, while prevalence in rodents (549/12,876; 4.26%) was slightly

lower than average (Table S1). Animals with at least one virus detection were most often

captured in South, East, and Southeast Asia (Fig. S2). Within infected animals, there was uneven

geographic distribution of animal taxonomic groups; infected bats, rodents, and shrews were

widely distributed, while infected birds and swine were found only in South, East, and Southeast

Asia (Fig. 1B).

A single virus was detected in most infected animals ($n = 3,048$, 93.18%); coinfections of

two ($n = 211$, 6.45%) or three ($n = 12$, 0.37%) viruses in the same animal were infrequent. Of the

223 coinfected animals, over half were bats ($n = 128$, 57.4%) with the remainder being 50

rodents, 39 birds, 5 swine, and 1 shrew (Table S1). Within infected animals, coinfection rates

were highest in rodents (50/549, 9.11%), birds (39/487, 8.01%), and bats (128/2,059, 6.22%).

Roughly half of the triple coinfections (7/12; Table S2) were domesticated ducks and geese,

which were all infected with duck coronavirus, influenza A virus, and either Newcastle virus or

avian paramyxovirus 6; the remainder were bats infected with a mix of coronaviruses,

paramyxoviruses, influenza A virus, and a novel rhabdovirus.

*Coinfection is non-random across phylogeny and life stage*

Bats, rodents, and birds exhibited unique patterns of coinfection, even accounting for their over-

representation in the dataset. Coinfections were more common than expected based on the

observed prevalence of all viruses: for all three taxonomic groups, we applied a permutation

testing approach, and found that observed coinfection rates exceeded the maximum simulated

coinfection rates (across 1,000 simulations for each taxonomic group; Fig. S3). However, when

accounting for geographic region, sex, age class, captivity status, and number of virus families

tested in a generalized linear model, we found that both birds (odds ratio (OR) = 0.24, 95%

confidence interval (CI): 0.09-0.54, $p$ = 0.001) and rodents (OR = 0.55, 95% CI: 0.33-0.89, $p$ =

0.018) were significantly less likely to be coinfected than bats (Fig. 2A, Table S3). These

findings could support the widespread hypothesis (and growing evidence) that bats are able to

tolerate viral infection, and potentially coinfection, through a unique set of immune adaptations

(Letko et al. 2020; Irving et al. 2021; Becker et al. 2025). However, we note that coinfection

rates within the all-animals model dataset (i.e., animals with available sex and age data) differed

compared to the coinfection rates in the dataset of all 3,048 infected animals: while coinfection

was still high in rodents (46/488, 9.43%) and bats (112/1,628, 6.88%), it was noticeably lower in

birds (12/316, 3.80%).

Coinfection was associated with life history, but those relationships varied between

groups. Based on previous studies (Anthony et al. 2017a; Peel et al. 2025), we hypothesized that

infection prevalence – and therefore coinfection rates – would be higher in younger animals, due

to waning of maternal antibodies and less-developed immune systems (Goenka and Kollmann

2015). We found that younger rodents (juveniles and subadults) were more likely to be

coinfected than adults, though this result was not statistically significant (OR = 1.84, 95% CI:

0.95-3.51, $p$ = 0.066; Fig. 2E, Table S5). However, juvenile bats were less likely to be coinfected

than adults (OR = 0.13, 95% CI: 0.03-0.55, $p$ = 0.006; Fig. 2C, Table S4), and this effect was

strong enough that it dominated the all-animals model (OR = 0.32, 95% CI: 0.10-0.80, $p$ = 0.030;

Fig. 2A, Table S3). This pattern is aligned with the widespread hypothesis that bats' unique

immune adaptations allow them to tolerate persistent viral infections (Calisher et al. 2006;

Plowright et al. 2016), which might accumulate over the course of an individual's lifespan

(Aguillon et al. 2025). Alternatively, depending on the timing of sampling by the PREDICT

study, juvenile bats may have been captured at times when levels of maternal antibodies were

still high. Though logistically challenging, longitudinal sampling within populations can help

elucidate patterns of virus infection and coinfection in relation to host maturation

(Wacharapluesadee et al. 2018; Brook et al. 2019; Peel et al. 2025).

Finally, we found that host social structure could drive coinfection patterns, but found

divergent effects between natural and un-natural settings. For bats, cave-roosting behavior was

not a significant predictor of coinfection status (OR = 0.93, 95% CI: 0.53-1.63, $p$ = 0.796; Fig.

2C, Table S4). Cave-roosting bats can exist in large groups in close proximity, potentially

increasing infection rates, but other roosting strategies can also create large, close aggregations

of bats (e.g., tree-roosting *Pteropus* species that cluster together). Finer-scale contact patterns

that we could not examine here (e.g., segregation by age or sex, or centrality in social networks)

might lead to more correlated infection risk across viruses at the individual level. Sickness

behaviors (e.g., reduced social interactions) after an initial infection could also reduce an

individual's subsequent likelihood of infection with other viruses (Ripperger et al. 2020; Moreno

et al. 2021).

In contrast, across all animals, we found that wild animals in captivity were significantly

more likely to be coinfected than free-ranging wild animals (OR = 2.12, 95% CI: 1.16-3.82, $p$ =

0.013; Fig. 2A, Table S3). These results support the hypothesis that captivity can increase

parasite transmission through both crowding, stress, and unusual patterns of interspecific contact.

This hypothesis has been strongly supported by previous analyses of coronavirus testing

conducted by the PREDICT project (Anthony et al. 2017b; Huong et al. 2020), and is a particular

point of concern, given the potential for coinfection to lead to recombinant coronaviruses (Wells

et al. 2023), and the high risk of spillover in wildlife farms and markets (Sánchez et al. 2021).

However, we note that of the 163 animals classified as "wild animals in captivity" in the all-

animals model dataset, 131 (80.4%) were rats (*Rattus* spp.) and 28 (17.2%) were mallards (*Anas*

*platyrhynchos*). Further work quantifying coinfection rates in captive counterparts of a broader

range of species is needed to better understand if these results are generalisable, and to examine

the role of potential mechanisms in driving coinfection in captivity.

*Sampling and testing biases can shape observed coinfection patterns*

Animals sampled in South, East, and Southeast Asia were more likely to be coinfected than those

sampled in Africa and West Asia in both the all-animals model (OR = 2.42, 95% CI: 1.63-3.61, *p*

< 0.001; Fig. 2A, Table S3) and the bats-only model (OR = 2.79, 95% CI: 1.78-4.39, *p* < 0.001;

Fig. 2C, Table S4). Rather than reflecting any characteristic of one region versus another, this

result might be driven by sampling artifacts. For instance, flying foxes (*Pteropus* spp.) had a high

coinfection rate (all-animals model dataset: 37/245; 15.1%) and were sampled only in South,

East, and Southeast Asia. Rats (*Rattus* spp.) were sampled in both geographic regions, but

infected animals were almost entirely found in South, East, and Southeast Asia, and coinfection

was high within this group (44/403; 10.9%). Virus family testing effort was also positively

correlated with coinfection in the all-animals model (OR = 1.11, 95% CI: 1.00-1.24, *p* = 0.042;

Fig. 2A, Table S3) and the rodents model (OR = 1.57, 95% CI: 1.29-1.94, *p* < 0.001; Fig. 2E,

Table S5).

*Evidence of virus-virus associations*

Coinfection networks revealed that certain viruses–namely, coronaviruses (CoVs),

paramyxoviruses (PMVs), and orthomyxoviruses (specifically, influenza A virus)–were most

frequently detected together (Fig. 2B, 2D, 2F). Notably, there were 108 CoV-CoV coinfections

(57 in bats, 48 in rodents), 69 CoV-PMV coinfections (44 in bats, 2 in rodents), 45 CoV-

influenza coinfections (23 in bats), and 10 PMV-influenza coinfections (1 in bats). Binomial

tests indicated that CoV-PMV, CoV-influenza, and PMV-rhabdovirus coinfections occurred

significantly more than expected in all animals and in bats, while coinfections with multiple

coronaviruses occurred significantly more than expected in rodents (Tables S6-S8). All other

virus family coinfections occurred < 10 times, and no coinfections involving arenaviruses,

filoviruses, or hantaviruses were detected.

Examining coinfection patterns at the level of individual viruses, we found that viruses

that are transmitted by poultry and other livestock (influenza A virus, Newcastle disease virus,

and infectious bronchitis virus) had the highest degree centrality in the network (Fig. 3).

However, a handful of commonly occurring betacoronaviruses, including murine coronavirus,

PREDICT CoV-16 and PREDICT CoV-17, and Bat coronavirus HKU9, were also well

connected within the network.

**Discussion**

Here, we conducted the first analysis of viral coinfection patterns across multiple continents and

host taxonomic groups. We found that within the virus groups sampled, coinfection is rare –

detected in just 223 of 65,662 animals – but occurs more frequently than expected, particularly

among specific virus groups. Our findings suggest that widely-occurring viruses could be

important to account for when researching other pathogens. For example, influenza A virus was

often detected in combination with other viruses; one possible explanation for this finding is that

influenza A virus might alter host susceptibility to secondary infections. Future work could

explore how specific viral coinfections influence factors such as the magnitude and duration of

shedding, particularly as it could relate to zoonotic spillover risk (Schuh et al. 2025). We also

found coinfection dynamics that suggest further testing would be beneficial to resolve whether

these associations are modulated by broad differences in host immunology and/or fine-scale host

contact processes, including in anthropogenic settings like farms and the wildlife trade. These

findings are also of potential importance for public health, as they highlight the need for both

basic research on host-parasite interactions, and surveillance of environments where coinfection

can produce new reassortant and recombinant viruses with epidemic or pandemic potential

(Anthony et al. 2017a).

Although our study is the largest macroecological study of coinfection to date, our

findings should be interpreted in light of the limitations of our data. Most importantly, the

PREDICT project was not designed with ecological research as the primary objective, and so the

testing data have uneven coverage both across and within taxonomic groups. For instance, the

PREDICT project was almost single-handedly responsible for bats becoming the best-studied

wildlife hosts of viruses (Gibb et al. 2022). Given that over 60% of specimens were collected

from bats, these data probably drove many of our broad-scale findings. However, only half of all

bat families (10 of 21) were sampled, and an emerging body of work highlights that while bats

are unique viral reservoirs, there is still considerable important (and interesting) variation within

the group (Becker et al. 2025). We also note that data collection for PREDICT was cross-

sectional in nature, and therefore interpret our results as associations that warrant further

mechanistic testing, for example using longitudinal or experimental approaches to test causality

of interactions. Finally, the surveillance nature of the data should not be viewed as entirely

representative of underlying populations. More complete sampling agnostic to host and virus

taxonomy is therefore required for more confident understanding of coinfection dynamics across

and within host groups. However, we note that more complete, unbiased datasets do not yet

exist, and leveraging existing sampling efforts provide a foundation for future sampling

priorities.

Future work could build on initial steps here by designing surveillance to more explicitly

search for and report coinfection. A recent systematic review found that coinfections were

common in bats, but were incidentally reported within study findings (Jones et al. 2023). More

deliberate focus on sampling to capture coinfection dynamics may also require specific sampling

strategies: for example, reliably detecting coinfection in an individual host requires capture of

animals in-hand (versus, e.g., pooled or environmental sampling). Similarly, different parasites

are associated with different organs and sites in the body, and so prioritizing particular samples

or assays could be necessary, depending on study aims. Although broad sample and parasite

coverage presents substantial cost considerations, several recent developments offer possible

solutions. Multiplex, metagenomic, and metatranscriptomic methods all provide a fuller view of

the host virome, and are especially useful when sample sizes are small due to biological

limitations (e.g., collecting blood from small animals) (Boyd et al. 2015; Raghwani et al. 2023;

Chang et al. 2020). Longitudinal surveillance also improves the odds of viral detection, and

allows researchers to test hypotheses about asynchronous interactions between parasites, host

immunity, and external environmental factors (Plowright et al. 2019). Finally, host-level data on

wildlife disease are also shared infrequently (Schwantes et al. 2025), preventing the kinds of re-

analyses we conducted here, or approaches that synthesize findings through systematic reviews

and meta-analyses (Jones et al. 2023). Developing a better global view of wildlife disease

dynamics is an essential step to conducting more macroecological research on the consequences

of coinfection.

**Acknowledgements**

This work was supported by NSF DBI 2515340. We thank our colleagues on the Verena team

for helpful feedback and brainstorming, and the countless members of the PREDICT consortium

who contributed to one of the largest and most unique ecological datasets ever collected.

**Conflict of Interest Statement**

We declare no conflicts of interest.

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     **Figure captions**

     **Figure 1. A.** Number of unique animals infected with one or more viruses, separated by

     taxonomic group. Groups in which only single infections were observed are not displayed (dogs:

     $n = 10$; goats/sheep: $n = 4$; cattle/buffalo: $n = 2$; carnivores: $n = 1$; other: $n = 1$). **B.** Distribution

     of animals infected with one or more viruses, colored by taxonomic group. Point size

     corresponds to the number of animals captured at a given location; coordinates were truncated so

     that nearby locations could be grouped together. Points are partially transparent and jittered for

     improved visibility.

     **Figure 2.** Predictors of virus coinfection (**A**, **C**, **E**) and virus family coinfection networks (**B**, **D**,

     **E**). **Left column:** generalized linear model outputs (odds ratios and 95% confidence intervals

     (CIs)) for predictors of coinfection are shown for all animals (**A**), bats only (**C**), and rodents only

     (**E**). CIs that do not cross the vertical line at 1 (i.e., statistically significant) are purple. Reference

     levels for categorical predictors are Africa and West Asia (geographic region), bats (taxonomic

     group), female (sex), adult (age class), wild free-ranging (captivity status), and non-cave roosting

     (roosting status). SES Asia = South, East, and Southeast Asia. **Right column:** networks of virus

     family coinfection are shown for all animals (**B**), bats only (**D**), and rodents only (**F**). Nodes

     represent virus families, while edges represent the frequency of virus family coinfection.

     Asterisks indicate coinfections that occurred significantly more than expected.

526  **Figure 3.** Virus coinfection network. Nodes represent individual viruses, while edges represent

527  the frequency of virus coinfection. Only nodes with degree centrality $\geq 4$ are labeled, and nodes

528  not involved in coinfections are not displayed.

530     **Figure 1**

**Figure 2**

**Figure 3.**

**Supporting Information for:**

**The macroecology of viral coinfection**

Cecilia A. Sánchez[1], Colin J. Carlson[1], Amy R. Sweeny[2*]

This Supporting Information contains Figures S1-3 and Tables S1-8.

**Figure S1.** Map of sampling effort in the PREDICT-2 dataset; color corresponds to the number

of unique specimens collected in a country. Note that the color legend is on a log scale.

**Figure S2.** Geographic distribution of unique animals infected with one or more viruses. Points

are partially transparent and jittered for improved visibility.

**Figure S3.** Density plots of 1000 coinfection simulations each for bats ($n$ = 36,181), rodents ($n$ =

12,876), and birds ($n$ = 5,617). The observed number of coinfected animals (in the PREDICT-2

data) for each taxonomic group is plotted with a dashed vertical line.

**Table S1.** Number of infected animals and coinfected animals within the PREDICT-2 dataset,

separated by taxonomic group. NA = not applicable (if no animals were infected, then none

could be coinfected).

| Animal taxonomic group | Total number tested | Number infected (≥ 1 viruses detected) | Number coinfected (≥ 2 viruses detected) |
|---|---|---|---|
| Bats | 36181 | 2059 | 128 |
| Birds | 5617 | 487 | 39 |
| Camels | 356 | 0 | NA |
| Carnivores | 406 | 1 | 0 |
| Cats | 62 | 0 | NA |
| Cattle/Buffalo | 517 | 2 | 0 |
| Dogs | 961 | 10 | 0 |
| Goats/Sheep | 441 | 4 | 0 |
| Horses | 1 | 0 | NA |
| Non-human primates | 6196 | 0 | NA |
| Pangolins | 332 | 0 | NA |
| Rodents | 12876 | 549 | 50 |
| Shrews | 1050 | 33 | 1 |
| Swine | 611 | 125 | 5 |
| Ungulates | 32 | 0 | NA |
| Other | 23 | 1 | 0 |

568    **Table S2.** Characteristics of the 12 animals that were infected with three viruses. Note that host

569    and virus names are provided as originally described in the PREDICT dataset, and may not

570    conform to current taxonomies.

| PREDICT sample ID | Country | Taxonomic group | Species | Virus family | Virus |
|---|---|---|---|---|---|
| 56344 | Egypt | Bats | Rousettus aegyptiacus | Coronaviridae | Bat coronavirus HKU9 |
| | | | | Orthomyxoviridae | Influenza A |
| | | | | Paramyxoviridae | PREDICT_PMV-116 |
| 108315 | Thailand | Bats | Pteropus lylei | Coronaviridae | PREDICT_CoV-17 |
| | | | | Paramyxoviridae | PREDICT_PMV-6 |
| | | | | Paramyxoviridae | PREDICT_PMV-128 |
| 108328 | Thailand | Bats | Pteropus giganteus | Coronaviridae | PREDICT_CoV-16 |
| | | | | Coronaviridae | PREDICT_CoV-68 |
| | | | | Paramyxoviridae | PREDICT_PMV-125 |
| 133237 | Cameroon | Bats | Rhinolophus cf. alcyone | Coronaviridae | Kenya bat coronavirus/BtKY83/59/58 |
| | | | | Coronaviridae | PREDICT_CoV-70 |
| | | | | Coronaviridae | PREDICT_CoV-109 |
| 145203 | Cambodia | Bats | Scotophilus kuhlii | Coronaviridae | Bat coronavirus 512/2005 |
| | | | | Paramyxoviridae | PREDICT_PMV-66 |

| | | | | Rhabdoviridae | PREDICT_RbdV-21 |
|---|---|---|---|---|---|
| 54944 | Bangladesh | Birds | Anas platyrhynchos domesticus | Coronaviridae | Duck Coronavirus |
| | | | | Orthomyxoviridae | Influenza A |
| | | | | Paramyxoviridae | Newcastle Disease Virus |
| 108865 | Bangladesh | Birds | Anas platyrhynchos domesticus | Coronaviridae | Duck Coronavirus |
| | | | | Orthomyxoviridae | Influenza A |
| | | | | Paramyxoviridae | Newcastle Disease Virus |
| 126366 | Bangladesh | Birds | Cairina moschata | Coronaviridae | Infectious bronchitis virus (IBV) |
| | | | | Coronaviridae | Duck Coronavirus |
| | | | | Orthomyxoviridae | Influenza A |
| 126464 | Bangladesh | Birds | Anas platyrhynchos domesticus | Coronaviridae | Duck Coronavirus |
| | | | | Orthomyxoviridae | Influenza A |
| | | | | Paramyxoviridae | Newcastle Disease Virus |
| 126519 | Bangladesh | Birds | Anas platyrhynchos domesticus | Coronaviridae | Duck Coronavirus |
| | | | | Orthomyxoviridae | Influenza A |
| | | | | Paramyxoviridae | Avian Paramyxovirus 6 |
| 126536 | Bangladesh | Birds | Anser cygnoides domesticus | Coronaviridae | Duck Coronavirus |
| | | | | Orthomyxoviridae | Influenza A |
| | | | | Paramyxoviridae | Avian Paramyxovirus 6 |
| 126561 | Bangladesh | Birds | Anas | Coronaviridae | Duck Coronavirus |

| | | | platyrhynchos domesticus | Orthomyxoviridae | Influenza A |
|---|---|---|---|---|---|
| | | | | Paramyxoviridae | Avian Paramyxovirus 6 |

573 **Table S3.** Model coefficients for a generalized linear model to examine effects of taxonomy,

574 sex, age, captivity status, and testing effort on virus coinfection status (binary: single infection or

575 coinfection) in bats, birds, rodents, and shrews. *p* values < 0.05 are in bold.

| Variable | | Estimate | SE | z | *p* | Odds ratio (95% CI) |
|---|---|---|---|---|---|---|
| Intercept | | -3.51 | 0.31 | -11.20 | **< 0.001** | 0.03 (0.02-0.05) |
| Geographic region | Africa and West Asia | – | – | – | – | – |
| | South, East, and Southeast Asia | 0.88 | 0.20 | 4.35 | **< 0.001** | 2.42 (1.63-3.61) |
| Taxonomic group | bats | – | – | – | – | – |
| | birds | -1.42 | 0.45 | -3.18 | **0.001** | 0.24 (0.09-0.55) |
| | rodents | -0.60 | 0.25 | -2.36 | 0.018 | 0.55 (0.33-0.89) |
| | shrews | -0.90 | 1.03 | -0.88 | 0.382 | 0.41 (0.02-1.96) |
| Sex | female | – | – | – | – | – |
| | male | -0.06 | 0.16 | -0.39 | 0.697 | 0.94 (0.68-1.29) |
| Age class | adult | – | – | – | – | – |
| | subadult | 0.12 | 0.20 | 0.58 | 0.563 | 1.13 (0.74-1.66) |
| | juvenile | -1.14 | 0.53 | -2.17 | **0.030** | 0.32 (0.10-0.80) |
| Captivity status | free-ranging wild animal | – | – | – | – | – |
| | owned domesticated | 0.27 | 0.60 | 0.44 | 0.658 | 1.31 (0.39-4.38) |

| | animal | | | | | |
|---|---|---|---|---|---|---|
| | wild animal in captivity | 0.75 | 0.30 | 2.48 | **0.013** | 2.12 (1.16-3.82) |
| Number of virus families tested | | 0.11 | 0.05 | 2.04 | **0.042** | 1.11 (1.00-1.24) |

579    **Table S4.** Model coefficients (fixed effects only) for a generalized linear mixed model to

580    examine effects of sex, age, cave roosting behavior, and testing effort on virus coinfection status

581    (binary: single infection or coinfection) in bats. *p* values < 0.05 are in bold.

| Variable | | Estimate | SE | z | *p* | Odds ratio (95% CI) |
|---|---|---|---|---|---|---|
| Intercept | | -2.97 | 0.64 | -4.61 | **< 0.001** | 0.05 (0.01-0.18) |
| Geographic region | Africa and West Asia | – | – | – | – | – |
| | South, East, and Southeast Asia | 1.03 | 0.23 | 4.44 | **< 0.001** | 2.79 (1.78-4.39) |
| Sex | female | – | – | – | – | – |
| | male | -0.08 | 0.21 | -0.40 | 0.686 | 0.92 (0.61-1.38) |
| Age class | adult | – | – | – | – | – |
| | subadult | -0.53 | 0.32 | -1.64 | 0.102 | 0.59 (0.32-1.11) |
| | juvenile | -2.03 | 0.73 | -2.77 | **0.006** | 0.13 (0.03-0.55) |
| Cave roosting | no | – | – | – | – | – |
| | yes | -0.07 | 0.29 | -0.26 | 0.796 | 0.93 (0.53-1.63) |
| Number of virus families tested | | -0.03 | 0.09 | -0.33 | 0.741 | 0.97 (0.81-1.16) |

584 **Table S5.** Model coefficients for a generalized linear model to examine effects of sex, age, and

585 testing effort on virus coinfection status (binary: single infection or coinfection) in rodents. *p*

586 values < 0.05 are in bold.

| Variable | | Estimate | SE | z | *p* | Odds ratio (95% CI) |
|---|---|---|---|---|---|---|
| Intercept | | -5.79 | 0.86 | -6.76 | **< 0.001** | 0.00 (0.00-0.01) |
| Sex | female | – | – | – | | – |
| | male | 0.13 | 0.32 | 0.42 | 0.676 | 1.14 (0.61-2.17) |
| Age class | adult | – | – | – | | – |
| | subadult/juvenile | 0.61 | 0.33 | 1.84 | 0.066 | 1.84 (0.95-3.51) |
| Number of virus families tested | | 0.45 | 0.10 | 4.36 | **< 0.001** | 1.57 (1.29-1.94) |

589  **Table S6.** Expected versus observed coinfection prevalence for virus families, for all animals. *P*

590  values were calculated from binomial tests and adjusted using a Bonferroni correction. $P_{adj}$

591  values < 0.05 are in bold.

| Virus family combination | Expected coinfection prevalence | Observed coinfection prevalence | Observed > expected | $P_{adj}$ |
|---|---|---|---|---|
| Arenaviridae-Arenaviridae | 5.80e-9 | 0 | No | 1 |
| Arenaviridae-Coronaviridae | 2.76e-6 | 0 | No | 1 |
| Arenaviridae-Filoviridae | 1.39e-8 | 0 | No | 1 |
| Arenaviridae-Flaviviridae | 5.80e-9 | 0 | No | 1 |
| Arenaviridae-Hantaviridae | 4.64e-9 | 0 | No | 1 |
| Arenaviridae-Orthomyxoviridae | 5.49e-7 | 0 | No | 1 |
| Arenaviridae-Paramyxoviridae | 5.42e-7 | 0 | No | 1 |
| Arenaviridae-Rhabdoviridae | 5.68e-8 | 0 | No | 1 |
| Coronaviridae-Coronaviridae | 1.31e-3 | 1.64-3 | Yes | 0.841 |
| Coronaviridae-Filoviridae | 6.62e-6 | 0 | No | 1 |
| Coronaviridae-Flaviviridae | 2.76e-6 | 0 | No | 1 |
| Coronaviridae-Hantaviridae | 2.21e-6 | 0 | No | 1 |
| Coronaviridae-Orthomyxoviridae | 2.61e-4 | 6.85e-4 | Yes | **< 0.001** |
| Coronaviridae-Paramyxoviridae | 2.58e-4 | 1.05e-3 | Yes | **< 0.001** |
| Coronaviridae-Rhabdoviridae | 2.70e-5 | 4.57e-5 | Yes | 1 |
| Filoviridae-Filoviridae | 3.34-8 | 0 | No | 1 |
| Filoviridae-Flaviviridae | 1.39e-8 | 0 | No | 1 |
| Filoviridae-Hantaviridae | 1.11e-8 | 0 | No | 1 |
| Filoviridae-Orthomyxoviridae | 1.32e-6 | 0 | No | 1 |

| | | | | |
|---|---|---|---|---|
| Filoviridae-Paramyxoviridae | 1.30e-6 | 0 | No | 1 |
| Filoviridae-Rhabdoviridae | 1.36e-7 | 0 | No | 1 |
| Flaviviridae-Flaviviridae | 5.80e-9 | 0 | No | 1 |
| Flaviviridae-Hantaviridae | 4.64e-9 | 0 | No | 1 |
| Flaviviridae-Orthomyxoviridae | 5.49e-7 | 1.52e-5 | Yes | 1 |
| Flaviviridae-Paramyxoviridae | 5.42e-7 | 0 | No | 1 |
| Flaviviridae-Rhabdoviridae | 5.68e-8 | 0 | No | 1 |
| Hantaviridae-Hantaviridae | 3.71e-9 | 0 | No | 1 |
| Hantaviridae-Orthomyxoviridae | 4.39e-7 | 0 | No | 1 |
| Hantaviridae-Paramyxoviridae | 4.33e-7 | 0 | No | 1 |
| Hantaviridae-Rhabdoviridae | 4.55e-8 | 0 | No | 1 |
| Orthomyxoviridae-Orthomyxoviridae | 5.19e-5 | 0 | No | 1 |
| Orthomyxoviridae-Paramyxoviridae | 5.12e-5 | 1.52e-4 | Yes | 0.090 |
| Orthomyxoviridae-Rhabdoviridae | 5.38e-6 | 0 | No | 1 |
| Paramyxoviridae-Paramyxoviridae | 5.06e-5 | 7.61e-5 | Yes | 1 |
| Paramyxoviridae-Rhabdoviridae | 5.31e-6 | 9.14e-5 | Yes | **< 0.001** |
| Rhabdoviridae-Rhabdoviridae | 5.57e-7 | 0 | No | 1 |

594    **Table S7.** Expected versus observed coinfection prevalence for virus families, for bats only. *P*

595    values were calculated from binomial tests and adjusted using a Bonferroni correction. $P_{adj}$

596    values < 0.05 are in bold.

| Virus family combination | Expected coinfection prevalence | Observed coinfection prevalence | Observed > expected | $P_{adj}$ |
|---|---|---|---|---|
| Coronaviridae-Coronaviridae | 1.80e-3 | 1.58e-3 | No | 1 |
| Coronaviridae-Filoviridae | 1.41e-5 | 0 | No | 1 |
| Coronaviridae-Flaviviridae | 1.17e-6 | 0 | No | 1 |
| Coronaviridae-Orthomyxoviridae | 2.37e-4 | 6.36e-4 | Yes | **< 0.001** |
| Coronaviridae-Paramyxoviridae | 3.89e-4 | 1.22e-3 | Yes | **< 0.001** |
| Coronaviridae-Rhabdoviridae | 5.75e-5 | 8.29e-5 | Yes | 1 |
| Filoviridae-Filoviridae | 1.10e-7 | 0 | No | 1 |
| Filoviridae-Flaviviridae | 9.17e-9 | 0 | No | 1 |
| Filoviridae-Orthomyxoviridae | 1.85e-6 | 0 | No | 1 |
| Filoviridae-Paramyxoviridae | 3.03e-6 | 0 | No | 1 |
| Filoviridae-Rhabdoviridae | 4.49e-7 | 0 | No | 1 |
| Flaviviridae-Flaviviridae | 7.64e-10 | 0 | No | 1 |
| Flaviviridae-Orthomyxoviridae | 1.54e-7 | 0 | No | 1 |
| Flaviviridae-Paramyxoviridae | 2.53e-7 | 0 | No | 1 |
| Flaviviridae-Rhabdoviridae | 3.74e-8 | 0 | No | 1 |
| Orthomyxoviridae-Orthomyxoviridae | 3.12e-5 | 0 | No | 1 |
| Orthomyxoviridae-Paramyxoviridae | 5.11e-5 | 2.76e-5 | No | 1 |
| Orthomyxoviridae-Rhabdoviridae | 7.56e-6 | 0 | No | 1 |
| Paramyxoviridae-Paramyxoviridae | 8.37e-5 | 1.11e-4 | Yes | 1 |

| Paramyxoviridae-Rhabdoviridae | 1.24e-5 | 1.66e-4 | Yes | **< 0.001** |
|---|---|---|---|---|
| Rhabdoviridae-Rhabdoviridae | 1.83e-6 | 0 | No | 1 |

600    **Table S8.** Expected versus observed coinfection prevalence for virus families, for rodents only.

601    *P* values were calculated from binomial tests and adjusted using a Bonferroni correction. $P_{adj}$

602    values < 0.05 are in bold.

| Virus family combination | Expected coinfection prevalence | Observed coinfection prevalence | Observed > expected | $P_{adj}$ |
|---|---|---|---|---|
| Arenaviridae-Arenaviridae | 1.51e-7 | 0 | No | 1 |
| Arenaviridae-Coronaviridae | 1.47e-5 | 0 | No | 1 |
| Arenaviridae-Hantaviridae | 6.03e-8 | 0 | No | 1 |
| Arenaviridae-Orthomyxoviridae | 6.03e-8 | 0 | No | 1 |
| Arenaviridae-Paramyxoviridae | 1.60e-6 | 0 | No | 1 |
| Coronaviridae-Coronaviridae | 1.44e-3 | 3.73e-3 | Yes | **< 0.001** |
| Coronaviridae-Hantaviridae | 5.90e-6 | 0 | No | 1 |
| Coronaviridae-Orthomyxoviridae | 5.90e-6 | 0 | No | 1 |
| Coronaviridae-Paramyxoviridae | 1.56e-4 | 1.55e-4 | No | 1 |
| Hantaviridae-Hantaviridae | 2.41e-8 | 0 | No | 1 |
| Hantaviridae-Orthomyxoviridae | 2.41e-8 | 0 | No | 1 |
| Hantaviridae-Paramyxoviridae | 6.39e-7 | 0 | No | 1 |
| Orthomyxoviridae-Orthomyxoviridae | 2.41e-8 | 0 | No | 1 |
| Orthomyxoviridae-Paramyxoviridae | 6.39e-7 | 0 | No | 1 |
| Paramyxoviridae-Paramyxoviridae | 1.69e-5 | 0 | No | 1 |