Evolution of the rate, spectrum, and fitness effects of mutation under minimal selection in

Caenorhabditis elegans

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Abstract

Evolution of mutation rate is of longstanding interest. The classical explanation is a "cost of

fidelity", which posits a tradeoff between selection to reduce deleterious mutations and selection

to minimize resources devoted to genome surveillance. An alternative is the "drift barrier",

whereby the mutation rate is under directional selection downward; at equilibrium, the increase

in fitness from a further increase in replication fidelity is too small to overcome genetic drift.

We propagated *C. elegans* "mutation accumulation" (MA) lines under minimal selection

for ~150 generations, divided each line into two sublines and propagated them for another ~150

generations. The base-substitution rate did not change, but variation among lines increased, as

predicted by the CoF. Conversely, the insertion/deletion rate increased, as predicted by the DB.

The base-substitution spectrum did not change, whereas the spectra of insertions and deletions

did.

Also of interest is the nature of selection on deleterious mutations. If epistasis is

synergistic, several evolutionary quandaries could be resolved. To quantify mutational effects on

fitness, and to characterize the pattern of epistasis, we assayed competitive fitness at 150 and

300 generations. Mutations had a mean effect of -0.1%, and a model of constant mutational

effects (no epistasis) satisfactorily explains the results.

The evolution of mutation rate is a longstanding conundrum in evolutionary biology. Because

most mutations with effects on fitness are deleterious, why does the mutation rate not evolve to

zero[1,2]? The two most compelling hypotheses are (1) the Cost of Fidelity[3] (CoF) and (2) the

Drift Barrier[4] (DB). The CoF posits that the optimum mutation rate represents a tradeoff

between direct selection to minimize the input of deleterious mutations and indirect selection to

minimize the resources devoted to genome surveillance, resulting in stabilizing selection on

mutation rate (Figure 1A). The DB posits that directional selection acts to reduce the mutation

rate until the fitness increase resulting from a further reduction of the mutation rate is less than

the reciprocal of the effective population size (N_e), at which point selection is too weak to

counteract genetic drift and the (assumed) mutational bias toward increased mutation rate

(Figure 1B).

The DB predicts that a decrease in N_e should result in an increase in mutation rate, and

vice versa. The CoF makes an analogous prediction: if the mutation rate is at the optimum, a

change in N_e, should not result in a directional change in the mutation rate. However,

populations at the same optimum should experience a predictable change in the variance where

a decrease in N_e should result in increased variance around the optimum, and vice versa.

Stabilizing selection comes with a "variance drift barrier" because mutations with fitness effects

$s<1/N_e$ will accumulate in the population, irrespective of whether they increase or decrease trait

value. However, the prediction of no directional change under relaxed selection is predicated on

symmetry of the fitness function around the optimum (Figure 1A). If the fitness function is

asymmetric, a population at mutation-selection-drift equilibrium will exhibit a mutational bias,

such that more mutations cause the mutation rate to change in the direction of stronger

selection than in the direction of weaker selection (Supplementary Figure 1). Without a

mutational bias, the population would not be at equilibrium and would be under directional

selection.

A second feature of the mutational process of longstanding interest is the distribution of

fitness effects (DFE) of new mutations[1,5,6]. Of particular interest is how deleterious mutations

interact. Whether the interaction between deleterious mutations is multiplicative (no epistasis),

synergistic (negative epistasis) or diminishing-returns (positive epistasis) has important

implications for evolution, prominently including the evolution of sex and recombination[7,8].

To investigate these questions in a tractable experimental system, we employed a

"second order mutation accumulation" (O2MA) experiment[9-11] using the nematode *C. elegans*

(Figure 2A). We initially constructed a set of "first order MA" (O1MA) lines[12-14]. After ~150

generations, each O1MA line was subdivided into two replicate O2MA lines and propagated for

an additional ~150 generations, at which time the genomes of the O2MA lines were sequenced,

and competitive fitness of the O1MA and O2MA lines was assayed (Figure 2B). The sequence

data allow us to quantify the evolution of the rate and spectrum of mutation over ~300

generations of minimal selection. The fitness data allow us to quantify the evolution of the mean

and variance of fitness effects of the accumulated mutations.

Results

Mutation rate. Over the first ~150 generations we observed 2938 putative mutations (2430

SNVs, 348 deletions, and 160 insertions (Supplementary Tables 1 and 2; Supplementary Figure

2). Ignoring false inferences, the first-order per-nucleotide SNV rate $\mu_{O1,SNV}$=1.77$\pm$0.038$\times$10$^-$

^{9}/generation; the indel rate $\mu_{O1,Indel}$=0.36$\pm$0.019$\times10^{-9}$/generation (Table 1). The total first-order

mutation rate (μ_{O1}) is 2.14$\pm$0.042$\times10^{-9}$/generation. In the second ~150 generations of the

experiment we observed 6631 total variants (5084 SNVs, 1016 deletions, 763 insertions). Still

ignoring false inferences, the second-order SNV rate $\mu_{O2,SNV}$=1.90$\pm$0.034$\times10^{-9}$/generation. The

second-order indel rate $\mu_{O2,Indel}$=0.60$\pm$0.024$\times10^{-9}$/generation. The total second-order mutation

rate μ_{O2}= 2.48$\pm$0.046$\times10^{-9}$/generation.

False positives (FP) and false negatives (FN) affect the estimated mutation rates in

predictable ways (Supplementary Figure 3). FPs artificially inflate the inferred μ_{O2}, and false

negatives (FN) artificially deflate μ_{O2} on average. Ignoring FNs results in a conservative test of

the hypothesis that $\mu_{O2} > \mu_{O1}$. Because we have only a rough estimate of the number of FPs ($\approx$

2 SNVs and 0.33 indels per genome; Supplementary Table 3), we tested the null hypothesis that

$\mu_{O2} = \mu_{O1}$ by "titrating" the number of FPs by simulation and asking the question: what is the

fewest FPs per O2MA line that would lead us to falsely reject the null hypothesis $\mu_{O2} = \mu_{O1}$ (i.e.,

Type I error)? For SNVs, there must be fewer than 0.6 FPs for the observed difference $\Delta\mu_{SNV} =$

0.23$\times10^{-9}$/generation to not be sufficiently explained by FPs (Figure 3A), so we conclude that

there is no evidence that $\mu_{O2,SNV} > \mu_{O1,SNV}$. For indels, the same test shows the critical number

of FPs per line is 2.7 (Figure 3B), so we conclude that the evidence suggests that $\mu_{O2,Indel} >$

$\mu_{O1,Indel}$.

Some types of mutations occur independently of genome replication and accrue at an

approximately constant rate in chronological time. If the chronological mutation rate remains

constant but the generation time increases, the result will be an increase in the per-generation

mutation rate. There was little difference in the average number of generations between the O1

and O2 lines over the course of the experiment (O1 mean = 145.6 generations; O2 mean =

144.2 generations), and the conclusions drawn from mutation rates scaled per-generation apply

similarly to mutation rates scaled per-day (Table 1). The number of indels was positively

correlated with generations of MA in both O1 ($r_{O1,indel}$=0.29, p<0.004) and O2 lines ($r_{O2,indel}$=0.25,

p<0.001). In the O1 lines, the number of SNVs was uncorrelated with the number of

generations ($r_{O1,SNV}$=0.10, p>0.32), whereas in the O2 lines the number of SNVs was positively

correlated with the number of generations ($r_{O2,SNV}$=0.46, p<<0.0001). The lack of correlation of

number of SNVs in the O1MA lines is likely due to lack of power resulting from the smaller

sample size and the lack of variation in generations of MA. Taken together, we reject the

hypothesis that mutation rate is independent of generation time, but we cannot rule out a

contribution of replication-independent mutation.

Mutation spectrum. Consistent with the observed lack of evolution of μ_{SNV}, the SNV mutation

spectra did not differ between O1MA and O2MA lines (Monte Carlo Fisher's Exact Test, p >

0.12) (Figure 4A,B). G:C→A:T transitions (28.6%) and A:T→T:A transversions (28.1%) were the

most common classes of base substitutions, as is typical in *C. elegans* MA experiments[15,16].

The indel spectrum, in contrast, was significantly different between O1MA and O2MA

lines (Figure 4C,D). In the O1MA lines, deletions outnumbered insertions by nearly 2.2:1,

whereas the deletion bias in the O2MA lines was significantly reduced to ~1.6:1 (Fisher's Exact

Test (FET), p < 0.0001). Considered separately, the deletion spectrum differed significantly

between O1 and O2 lines (FET, p < 0.003), with an increase in the shortest (1 bp) and longest

(>10 bp) categories and concomitant decreases in the intermediate categories. The insertion

spectrum also differed between O1 and O2 (FET, p < 0.0001), also with increases in the 1 bp

and >10 bp categories and decreases in the intermediate categories.

Variance in mutation rate. To begin, we ask: can the observed variance in mutation rate among

O1MA lines ($V_{\mu O1}$) be sufficiently explained by sampling variance of a uniform mutational

process? For SNVs the answer is yes; $V_{\mu O1,SNV}$ is sufficiently explained by sampling if mutation

conforms to a uniform Poisson process (p>0.09, Supplementary Figure 4A). For indels,

however, $V_{\mu O1,Indel}$ is too great to be explained by a uniform Poisson process (p<0.02,

Supplementary Figure 4B).

To extend the analysis to the O2MA lines, FPs must be considered. We assume a

liberal average of 3 SNV and 1 indel FPs per sample, leading to a conservative test of the

hypothesis of a uniform mutational process. For both SNVs and indels there is more variance

among O2MA lines than is consistent with a uniform mutational process (Supplementary Figure

5).

A different approach is to ask: how much heritable variation in μ_{O2} exists among O1MA

families? Over a given time period, a trait may not be heritable either because it evolves too

slowly for significant variation to accumulate, or because it evolves so rapidly that the signal of

ancestor-descendant relationship is erased.

Neither $\mu_{O2,SNV}$ nor $\mu_{O2,Indel}$ is significantly heritable ($h^2_{\mu_{O2,SNV}} = 0.037$, p>0.71; $h^2_{\mu_{O2,Indel}} =$

0.089, p>0.38), nor is total mutation rate ($h^2_{\mu_{O2}} = 0.046$, p>0.65). For SNVs, the lack of

heritability is consistent with inferred uniformity of the O1MA process. For indels, the simplest

explanation is that (*i*) heritability is inherently low, and (*ii*) FPs add additional noise, so 96 O1

families provide insufficient power to detect heritability of ~10% over 150 generations. A more

speculative possibility is that in the absence of purifying selection, indel mutation rate is an

accelerating function of mutation load[17,18].

Mutational effects on fitness. Because all mutations in an MA line are in complete linkage

disequilibrium, the effects of individual mutations are unidentifiable. However, knowing the

number of mutations in each MA line allows a straightforward estimate of the first two moments

of the DFE, because the sum of the effects of a large number of individual mutations within a

line can be well-approximated by a normal distribution regardless of the underlying DFE,

according to the Central Limit Theorem.

Competitive fitness (*W*) decreased by about 0.2% per generation over the first 150

generations of MA and by about half that over the next 150 generations of MA (Supplementary

Figure 5, Supplementary Tables 4 and 5). There is a weak negative association between

mutation rate and fitness (Supplementary Figure 6).

Given that the number of mutations is similar between O1MA and O2MA lines, the

implication is that the mean effect on competitive fitness of O1 mutations (u_{O1}) is greater than

that of O2 mutations (u_{O2}). To test that hypothesis, we compared a model in which the mean

mutational effect u is constrained to remain constant over the entire 300 generations ("uniform

effects") to a "two-stage" model in which u_{O1} and u_{O2} are allowed to differ. Models were

compared using the Widely Applicable Information Criterion (WAIC) and leave-one-out cross-

validation (LOO) to assess model fit and predictive performance[19]. The differences between

models are substantial (Δ_{WAIC} = 13.52; Δ_{LOO} = 12.69) and favor the uniform effects model. The

point estimate of the mean mutational effect u = -0.0139, with variance σ_g^2 = 1.22×10⁻⁴ (Figure

5).

Discussion

Our primary goal was to test empirically the predictions of the two prominent alternative

theoretical models of the evolution of mutation rate. The Drift Barrier (DB) model predicts that

the mutation rate should increase when selection is relaxed. Explicit in the DB model is the

existence of a mutational bias in the direction of increased mutation rate. The Cost of Fidelity

(CoF) model posits that mutation rate is under apparent[20] stabilizing selection, whereby direct

selection to reduce the input of deleterious mutations is counterbalanced by indirect selection to

minimize the resources devoted to genome surveillance. The "strong" form of the CoF is

predicated on symmetry of the fitness function around the optimum and predicts that relaxation

of selection will not lead to a change in the mean mutation rate, but that the variance in mutation

rate will increase. The "weak" form of the CoF, which relaxes the assumption of symmetry of

the fitness function, also predicts an increase in the variance in mutation rate upon relaxation of

selection. It allows for a change in the mean mutation rate, provided that the mutational bias is

in the direction of stronger selection.

SNV mutations are consistent with the predictions of the strong CoF model: the mean

base substitution mutation rate μ_{SNV} did not change significantly over the course of the

experiment, whereas the variance in μ_{SNV} increased significantly. Indels fulfill the predictions of

the DB model but not the strong CoF. The indel mutation rate increased significantly over the

course of the experiment; the point estimate of μ_{Indel} increased by about 2/3 from O1 to O2 and

the number of false positive indels would have to be much greater than the estimated number to

explain the observed increase in mutation rate. However, the data do not allow us to rule out

the weak CoF, which posits that if a mutational bias exists, then the bias should be in the

direction of stronger selection (Supplementary Figure 3). The trend for a negative association

between competitive fitness and mutation rate (Supplementary Figure 6) is consistent with both

the DB and the weak CoF, given the mutational bias for increased indel mutation rate.

The findings of this study reinforce and extend previous findings from *C. elegans* and

other organisms, in which various types of physiological stress – exogenous and endogenous –

increase the indel mutation rate, but have little if any effect on the base substitution mutation

rate. Rajaei et al.[15] performed an MA experiment with an N2-derived strain of *C. elegans*

carrying a mutation (*Mev-1*) that causes increased cellular oxidative stress. In that experiment,

the indel (and specifically, the insertion) rate increased by about 50% over wild-type, whereas

the SNV mutation rate decreased slightly. Similarly, MA lines of *Drosophila melanogaster*

initiated from strains engineered to carry large-effect deleterious mutations experience a higher

rate of short deletions than their wild-type controls, whereas the SNV mutation rates do not

differ[21]. *Arabidopsis thaliana* MA lines propagated under benign and stressful conditions of

salinity[22] and temperature[23] show large and highly significant increases in indel mutation rates

compared to much smaller increases in SNV rates. A similar increase in indel frequency was

reported for the green alga *Chlamydomonas reinhardtii* MA lines maintained under salinity

stress[24].

Our second-order MA experiment has three antecedents. First, Ávila et al.[9] initiated an

O2MA experiment from a single high-fitness MA line of *D. melanogaster*. Their study predated

economical whole-genome sequencing, but they found that the decline in fitness in the O2MA

experiment was greater than in the O1MA experiment. They attributed the increased rate of

decline in fitness to an increase in mutation rate, although they could not formally rule out

greater mutational effects in the O2MA experiment. The subsequent findings of Sharp and

Agrawal[21] add credence to their conclusion. Second, Saxena et al.[10] did an O2MA experiment

with *C. elegans*, designed to test the hypothesis that genomes that carry a high mutation load

mutate faster than less-loaded genomes[17,18]. The indel rate (specifically the insertion rate)

increased significantly in the O2MA lines; contrary to expectation, the O2MA indel rate was

greater in the high-fitness (= less-loaded) lines. The SNV rate increased on average, but not

significantly. Two of the ten O1MA families (n=5 O2MA sub-lines per O1MA family) evolved

SNV O2 mutation rates that were significantly different from the ancestor; one family increased

and the other decreased.

The third antecedent is the study of Liu and Zhang[11], who brought the power of yeast

genetics to bear on the question of mutation rate evolutiion. They first constructed a set of

O1MA lines in which a mismatch repair gene (*MSH-2*) was deleted by CRISPR-Cas9

engineering to increase the mutation rate. They then constructed a set of O2MA lines by re-

inserting the *MSH-2* gene into the O1MA lines and allowing mutations to accumulate further.

They found (*i*) that the O2MA mutation rate increased on average, as predicted by the DB, but

also (*ii*) 12/48 O2MA lines evolved a decreased mutation rate. They interpreted the high

frequency of lines evolving decreased mutation rate as inconsistent with the DB hypothesis and

consistent with the yeast mutation rate being maintained by apparent stabilizing selection (i.e.,

the weak CoF). Moreover, they identified a gene with antagonistic pleiotropic effects on cellular

physiology (autophagy) and mutation rate, as predicted by the CoF. Their results, if not their

conclusions, are in agreement with those of Saxena et al.[10] in *C. elegans*.

In 1995, Kondrashov[25] famously posed the question "Why have we not died 100 times

over?". Given the sizes of genomes (large) and effective population sizes (much smaller), the

continued existence of humans seemed incompatible with the inferred genetic load (~100 lethal

equivalents) if the effects of deleterious mutations are independent. The simplest explanation is

that the effects of deleterious mutations are not independent, but rather they interact

synergistically, i.e., epistasis is negative. However, it is not immediately obvious why

deleterious mutations should interact synergistically on average. One possible explanation is

that selection is "soft"[26], wherein population absolute fitness is established by resource

availability and relative fitness results from competition among individuals for resources[27,28]. If

so, it implies that epistasis is more likely to be detected under competitive conditions.

This study provides no support for the possibility that deleterious mutations interact

synergistically, and in fact the point estimate of the effect of second-order mutations (u_{O2}) is

about half that of the first-order effect, suggestive of diminishing returns epistasis. However, we

can confidently reject the two-stage model of mutational effects in favor of the one-stage model

of constant mutational effects ($\Delta_{WAIC} \approx 12$). There are obviously many caveats associated with

that inference, including environmental context and statistical power. We note that our findings

recapitulate those of Peters and Keightley[29] and Katju et al.[30] who looked for epistatic effects of

on non-competitive fitness in *C. elegans* and did not find them.

Conclusions. The results of this study are clear: there is an upward mutational bias for indels,

whereas there is no such bias for SNV mutations, although the variance in μ_{SNV} increased over

the course of the experiment. The discrepancy between evolution of the indel and SNV

mutational processes, in this study and others, is intriguing. One possibility is that the genomic

mutational target for the control of the indel process is simply larger than that of SNVs, and that

if selection was relaxed for long enough that the SNV mutation rate would begin to drift upward.

We suspect this is probably true. However, it is worth considering the possibility that the indel

and SNV mutational processes are subject to different patterns of selection, and that the DB or

weak CoF model applies to indels, whereas the strong CoF model applies to SNVs.

What remains to be determined is the nature of the fitness function with respect to

mutation rate. Liu and Zhang[11] reported that mutation rate of their yeast O2MA lines is

uncorrelated with growth rate. Visual inspection of their plot of growth rate against mutation rate

(their Supplementary Figure 2E) reveals a weak positive relationship between mutation rate and

growth rate, contrary to the prediction that lines with high mutation rates (mutators) should have

lower than average fitness, as expected if mutation rate is under stabilizing selection in the

presence of a mutational bias toward increased mutation rate. However, growth rate was

measured in only one environment and both mutation rate and fitness in yeast are known to be

context dependent[31], so it remains entirely possible that mutation rate in yeast is under global

stabilizing selection. What Liu and Zhang clearly did show is (*i*) antimutator mutations are not

trivially rare in yeast, but also (*ii*) on average, mutation rate drifts upward when selection is

relaxed.

The idea that epistasis among deleterious mutations should be synergistic on average

has been as influential among evolutionary geneticists as the supporting evidence for it is

weak[32]. Beyond the obvious possibility that 50 years of inconclusive evidence is the result of

underpowered experiments intended to quantify an inherently noisy trait, the intriguing

possibility exists that epistasis itself is not only evolvable[33,34], but that the direction of epistasis

favored by selection is dependent on idiosyncratic properties of the population including

effective population size and mutation rate[35,36].

Methods

Mutation Accumulation (MA) Experiment

We obtained a stock of the EG8072 strain of *C. elegans* (https://cgc.umn.edu/strain/EG8072)

from the Caenorhabditis Genetics Center (CGC) at the University of Minnesota in September

2018. EG8072 is a derivative of the standard N2 strain, with green (GFP) and red (mCherry)

fluorescent reporters stably integrated into the genome. We propagated the founder stock for

seven generations by transfer of a single immature hermaphrodite at 4-day intervals, at which

point the population is expected to have reached mutation-drift equilibrium ($6N_e$ generations[37]).

The population was expanded to large size (thousands of mixed-stage individuals) for 2-3

generations and cryopreserved.

In November 2018 we thawed an aliquot of the cryopreserved stock, cleaned it by

bleaching[38] and initiated 148 replicate populations, each initiated from a single immature

hermaphrodite. 48 replicates were expanded to large size (N ≈ a few thousand) for an

additional 2-3 generations and cryopreserved in 20 replicate PCR plates. These are the G0

ancestral "pseudolines". The remaining 100 populations – the first-order (O1) MA lines - were

propagated by transfer of a single immature hermaphrodite at four-day intervals for an additional

600 days (Gmax=150 generations). MA lines were maintained under standard conditions: on 35

mm NGMA agar plates spotted with 50 µl of an overnight culture of *E. coli* OP50, kept in a

Percival incubator set to a constant 20° C. Plates were kept on cafeteria trays, in 6 × 8 arrays,

48 plates per tray. Even numbered and odd numbered replicates were kept on different trays.

After the leading-generation worm was transferred, the plate it was taken from was

moved to a different incubator set to 10° C; this is the (first) backup plate. Worms do not

reproduce (or reproduce only very slowly) at 10° C. If at the time of transfer a worm of the

leading generation failed to reproduce or was not visibly gravid, it was replaced by a sibling

taken from the previous generation plate. If there were offspring on the leading generation plate

but no worm had reached at least the L2 stage, the plate was "held over" without transferring a

worm. This protocol assures that generation number of a line is known with (near) certainty.

Demographic data on backups and holdovers is given in Supplementary Table 6.

All lines were cryopreserved at 50 generation intervals. After 150 transfers (600 days,

Gmax=150), each O1MA line was replicated into two second-order (O2) MA lines; the two

O2MA lines derived from an O1MA line are an "O1MA family". The O2MA lines were

propagated as described for an additional ~150 generations (Gmax=306), at which time they

were cryopreserved analogously to the G0 pseudolines as described above.

Genome Sequencing and Variant Calling

(i) <u>DNA Extraction and library preparation</u>. Cryopreserved tubes of worms were thawed and

grown for several days on 60 mm NGMA agar plates containing Streptomycin and Nystatin and

seeded with 110 µl of overnight culture of the HB101 strain of *E. coli*. When food was nearly

exhausted, a small piece (~2 cm^2) of the agar plate was transferred onto a 100 mm NGMA plate

and worms were allowed to grow for another 2-3 days. When food was nearly exhausted,

worms were collected into 15 ml centrifuge tubes on ice and allowed to settle. Approximately

100 µl of settled worms were transferred into a 1.5 ml microcentrifuge tube and stored at -80°C

until all samples were ready for DNA extraction, at which time samples were thawed, and

genomic DNA was extracted with the DNeasy Kit (Qiagen) following the manufacturer's protocol.

Extracted DNA was stored at -20°C prior to library preparation.

Frozen DNA samples were thawed and diluted with diH$_2$O to a concentration of 0.2

ng/µl. Tagmentation was performed per manufacturer's instructions (Illumina, Nextera DNA

Sample preparation kit, FC-121-1030). Following tagmentation, amplification was performed

using custom barcoded IDT primers. Following amplification, a 96-pooled well sample was

generated. Once pooled, 170µL of sample was combined with 30µL of 6X loading dye and run

on a 2% agarose gel. DNA segments of 300-500 bp were excised from gel. The QIAquick Gel

Extraction Kit (QIAquick, 28704) was used to clean and elute DNA from the gel, following

manufacturer's instructions. Pooled library samples were quantified using the Qubit HS kit

(Qubit, Q33230) and sequenced to ~25X average depth of coverage with 150 bp paired-end

Illumina sequencing on a NovaSeq 6000 by Novogene, Inc.

(ii) <u>Genome Assembly.</u> Raw sequence reads were filtered for quality (MQ≥3) and adapter

sequence trimmed with fastp[39]. Reads were then aligned to the *C. elegans* N2 strain reference

genome (WS292) using bowtie2[40]. The resulting BAM files were sorted and filtered with

Samtools[41] and bamTools[42] for downstream processing. Read group information was added

with Picard (http://broadinstitute.github.io/picard). Duplicate reads were identified and removed

using MarkDuplicates in GATK[43]. Sequence quality was assessed with FastQC[44] and MultiQC[45]

after trimming and alignment and prior to variant calling.

(iii) <u>Variant Calling Pipeline.</u> Variants were called with the HaplotypeCaller tool in GATK[43], run in

GVCF mode. HaplotypeCaller identifies both single nucleotide variants (SNVs) and indels by

local de novo assembly of haplotypes in regions with evidence of variation, improving accuracy

by reconstructing the sequence context in those regions. GVCFs from all samples were merged

with GenomicsDBImport, and joint genotyping was performed using GenotypeGVCFs. We used

bcftools[46] to identify homozygous variants unique to a line, which were counted as putative

mutations.

Variants that met the following three criteria were considered as putative mutations: (1)

they were called homozygous; (2) they were present in one and only one O1MA family, and (3)

the ancestor genotype is homozygous wild-type. If the variant is present in both O2MA sublines,

the mutation is inferred to have occurred in the first 150 generations and be a first-order (O1)

mutation. If the variant is present in only one of the two O2MA sublines, it is inferred to have

occurred subsequent to the split of the two O2MA sublines and be a second-order (O2)

mutation. Comparisons of O1 and O2 mutation rates were restricted to the additional criteria (4)

a site must be called in both members of an O1MA family and (5) the site must be called in at

least 90% of O1MA families.

One O2MA line from O1MA family 9 was not sequenced. The identities of two O2MA

lines appear to have been swapped, one from O1MA family 34 and one from family 46. One

O1MA line (family 26) evolved an indel mutator phenotype, as evidenced by the >5X greater

indel rates of the two O2MA sublines than of the O2MA line with the next highest indel rate

(Supplementary Table 1). Except where noted, those families were omitted from further

analysis.

In the O2MA lines, over 97% of the ~100.2 Mb genome was covered at the liberal 3X

coverage criterion (mean= 97.6%, median = 97.8%; Supplementary Table 1). At 3X coverage,

the mean and median fractions of the genome included in the O1MA analysis are 97.3% and

97.5% respectively. At 10X coverage the O2MA mean and median are 93.5% and 95.8%; the

O1MA mean and median are 89.8% and 93.2%.

Assessment of falsely inferred variants

(*i*) <u>False positives</u>. EG8072 is derived from the N2 strain, which is the source of the *C. elegans*

reference genome[47]. Any difference between our EG8072 ancestor and the reference genome

is either (1) a mutation that fixed subsequent to the divergence of EG8072 and the N2 reference

strain, (2) a false negative in the N2 reference, or (3) FP in our EG8072 ancestor. In the first

two cases, all MA lines derived from the EG8072 ancestor should also have the variant allele.

However, if the variant is FP in our EG8072 ancestor, the derived MA lines should all have the

reference allele. The EG8072 G0 ancestor was sequenced from three independent Illumina

libraries. We estimated the frequency of FPs from each of the three samples of the ancestor.

At the liberal (3X) coverage criterion we identified an average of 1662 fixed variants

between our G0 ancestor and the reference genome (1676, 1647, and 1664 in the three

replicates, respectively; Supplementary Table 3), of which approximately 42% were single

nucleotide variants (SNVs) and 58% were insertions or deletions (indels). In the three replicates

we identified three (1 deletion, 2 SNVs), three (3 SNVs) and one (1 SNV) FPs respectively. One

of the SNVs (a G:C→C:G transversion) was common to all three replicates; the others were

only found in one replicate. Of the seven FPs detected at the liberal 3X criterion, six were also

found at the stringent 10X criterion. By way of comparison, the same analysis with a different

N2 ancestor found one FP (an SNV) out of 1439 variants with 68 MA lines at the 3X criterion[15].

(*ii*) Failure to recall (FtR) and false negatives (FN). A mutation can fail to be detected ("failure to

recall", FtR) for two reasons. First, a site with a true mutant may not be called (i.e., missing

data). Second, a true mutant allele may be mistakenly identified as an ancestral allele; that is a

false negative (FN). FNs are a subset of FtR and cause the mutation rate to be underestimated.

FtR due to missing data may or may not affect estimation of the mutation rate, depending on

whether sites with true mutations are more or less likely to be called than sites without

mutations. All else equal, sites with true mutations are less likely to be called than sites without

mutations. The reason is that if a sequence with n or more mismatches exceeds the cutoff for

mapping, a sequence with a mutation has one more mismatch than the same sequence without

the mutation. All else also equal, FNs are equally likely to be true O1 mutations or true O2

mutations. Since there are twice as many opportunities for O2 as O1 mutations (Supplementary

Figure 3), FNs introduce a downward bias in the O2 mutation rate relative to the O1 mutation

rate.

To assess our ability to recall mutations, we introduced simulated "dummy" mutations at

random into the reference genome and analyzed the data using our standard variant-calling

pipeline, but with the simulated genome as the reference[15]. The changed site in the reference

should appear as a variant in our sequenced sample. For example, suppose the true reference

allele at a position is an A and we change it to a C. The site should be called variant (A/A in this

example) in our data. If the site is either not covered in the data or is called anything other than

a variant homozygote (A/A), we have failed to recall the mutation. If the site is covered in the

data but not called a variant homozygote, that is FN.

We constructed a pseudo-reference genome by inserting 1200 single-nucleotide variants

(SNVs) and 12,000 indels (1–20 bp, size chosen at random) into the *C. elegans* N2 (WS292)

reference genome (Supplementary Table 7). First, we randomly inserted "dummy" SNVs and

indels into the reference genome in random positions using the R packages GenomicRanges,

IRange, and Biostrings and compiled them into a VCF file. The VCF was then sorted and

indexed with samtools, and the pseudo-reference genome was produced with GATK

FastaAlternateReferenceMaker. The genome assembly and variant calling pipelines were rerun

with all the MA lines and the G0 ancestor using this pseudo-reference genome. Line-specific

simulated FtR and FN rates and counts are given in Supplementary Table 8.

To validate the consistency of the procedure, the original variants (i.e., putative

mutations) were re-mapped against the pseudo-reference genome. All sites identified as variant

in the original analysis were also called as variant against the pseudo-reference. We repeated

the simulation twice independently and obtained consistent results across runs (Supplementary

Table 9).

At the liberal 3X criterion, the average FtR rate was ~2.2% for SNVs and ~2.5% for

indels. The FN rates were 0.21% and 0.53% for SNVs and indels. At the stringent 10X criterion

the failure to recall rates were 7.8% and 7.7% for SNVs and indels respectively; the FN rates

were 0.084% and 0.12%. All analyses are based on the 3X coverage data; 10X data are

reported in the Supplementary Tables.

Competitive Fitness Assay

(*i*) <u>Focal worms</u>. 50 G150 O1MA lines and 20 G0 pseudolines (PS lines) were chosen at

random for inclusion in assay block 1. Block 2 included one G300 O2MA replicate from each of

those same 50 lines and the same 20 PS lines; Block 3 included the other O2MA replicate,

along with the same 20 PS lines. At the beginning of each assay (*Day 1*) a cryopreserved

aliquot from each line (MA and PS) was thawed, and ~150 µl pipetted onto a 60 mm NGMA

plate seeded with the 100 µl of an overnight culture of the OP50 strain of *E. coli* (the thaw plate;

Figure 2B) Two days later (*Day 3*), seven immature worms (L3/L4 stage) from each line were

picked singly to a 35 mm NGMA plate seeded with 50 µl of OP50. These plates are the parental

(P) generation. Each P plate was identified with a random number and arranged in numerical

order on cafeteria trays, 48 plates/tray, and placed in a Percival incubator maintained at 20˚C

and ambient humidity. Thaw plates were incubated at 10˚C to prevent further reproduction.

Four days later (*Day 7*), a single immature (L3/L4) hermaphrodite was picked to a new plate and

maintained as before. These plates are the F1 generation. If a P generation worm failed to

reproduce, it was replaced in the assay with a sibling from the thaw plate. Four days later (*Day

11*) the process was repeated for the F2 generation. If an F1 worm failed to reproduce, the

replicate was removed from the assay and treated as missing data. Three days later (*Day 14*) a

single L1 stage worm was picked from each F2 plate onto the assay plate, along with an L1

competitor strain worm.

(*ii*) Competitor worms. On *Day 1* of the assay, an aliquot of N2 strain worms was thawed and

pipetted onto a seeded 35 mm plate and allowed to reproduce. When the food was exhausted

(around *Day 5*), a ~2 cm chunk was transferred to a new 100 mm NGMA plate seeded with 400

µl of OP50 and allowed to reproduce. When food was exhausted (around *Day 9*), a chunk was

transferred to a new plate. On *Day 13*, worms were washed from the plate and bleached to

synchronize the population. Bleached embryos were pipetted in equal aliquots onto four 100

mm NGMA plates supplemented with streptomycin and the antifungal agent nystatin and

seeded with 400 µl of the HB101 strain of *E. coli*.

(*iii*) Assay plates. A single L1 focal worm (MA or G0 pseudoline) and a single L1 competitor

worm were picked to a 35 mm NGMA plate supplemented with streptomycin and nystatin and

seeded with 50 µl of HB101. Assay plates were identified by random number, placed in random

order on trays (48 plates/tray) and incubated at 20°C for eight days (*Day 21*; approximately two

overlapping generations), at which point nearly all plates were depleted of bacterial food.

Worms were washed from the assay plate in 2 ml of M9 buffer and transferred to a 2 ml deep-

well 96-well plate. The plate was centrifuged for 1 min at 1,000 g, and the supernatant was

aspirated with an 8-channel strip aspirator set to leave 100 µl in each well. The worms were

washed once more by adding 1.4 ml of M9 buffer, centrifuging, and aspirating as before,

followed by the addition of 1.4 ml of M9. Worms were mixed by gently pipetting with a 12-

channel pipette and a 110 µl sample was transferred into a well in a 96-well, clear-bottom tissue

culture plate. 5 mM levamisole was added to each well to immobilize worms prior to imaging.

(iv) Imaging. Bright-field and GFP fluorescence images (470±20 nm excitation, 525±25 nm

emission) were captured at 20x or 30x magnification for each well using an automated

epifluorescence microscope (IX-70, Olympus, Pittsburgh, PA, USA) fitted with a CCD camera

(Retiga-2000R, QImaging, Surrey, BC, Canada), XYZ stage and focus motors (Prior Scientific,

Cambridge, UK), and controlled by Image Pro-plus software (Media Cybernetics, Rockville, MD,

USA). 96-well plates were imaged first under bright field (BF) and then under fluorescence

(GFP). Images of the same well (BF and GFP) were paired and renamed after checking by eye

that the images matched.

(v) Counting. Worms were counted from stacked images using ImageJ software. We

summarize the protocol here; a detailed tutorial of the ImageJ analysis is included as

Supplementary Appendix 1.

Step 1. Paired images were stacked using the "Stacks" option under the "Image" tab.

Step 2. Images were merged and colorized using the "Merge Channels" tool under the "Image"

tab, with fluorescent worms colored differently from non-fluorescent ones.

Step 3. For best effect, brightness and contrast of GFP images were adjusted by setting

"minimum" and "maximum" light low and "contrast" high under the "Image" tab.

Step 4. If an image was too crowded with worms, the ROI selection tool in the "Analyze" tab was

used to crop a sample portion of the stacked images to count. The stacked images were

synchronized using Analyze →Tools →Synchronize windows → Synchronize All commands.

Step 5. Worms were counted by eye using the "Multipoint" tool to mark individual worms, with

different marking colors for fluorescent and non-fluorescent worms.

Data Analysis

i) Mutation Rate. We calculated the per-generation mutation rate, μ_i, for each MA line *i* as $\mu_i =$

$m_i/n_i t_i$, where *m* is the number of mutations, *n* is the number of callable sites, and *t* is the number

of generations of MA[48]. First-order and second-order mutation rates were calculated separately.

Our first question of interest is: does the mutation rate change over time? To test that

hypothesis, we can estimate the mutation rates of O1MA lines (μ_{O1}) and O2MA lines (μ_{O2}),

calculate the difference $\Delta\mu = \mu_{O2} - \mu_{O1}$, and test the hypothesis that the average difference is

significantly different from zero. There is a complication, however, because false inferences

affect estimates of μ_{O1} and μ_{O2} in different ways (Supplementary Figure 3).

To proceed, we employed a parametric bootstrap approach. We assume that mutation

can be modeled as a Poisson process. We further assume that there is a uniform mutation rate

$\mu = \mu_{O1} = \mu_{O2}$, i.e., the mutation rate does not change over the course of the experiment.

Because FPs necessarily inflate μ_{O2} relative to μ_{O1} and FNs do the opposite on average,

ignoring FNs leads to a conservative test of the prior hypothesis that $\mu_{O2} > \mu_{O1}$. To begin, we

set $\mu=$ to the point estimate of μ_{O1}. At each generation, an MA line accumulates mutations in its

genome drawn from a Poisson distribution with parameter $\lambda_m = \mu L$, where μ is the per-

generation mutation rate and *L* is the size of the *C. elegans* genome. The total number of

mutations in line *i*, $m_i = \lambda_m t_i$, where t_i is the number of generations of MA for line *i*. At

Gmax=150, the O1MA line is split into the two O2MA sublines, and the mutational process is

continued for the remaining Gmax=150 generations. Each mutation was accepted ("called") or

rejected with binomial probability equal to the proportion of the genome covered in the line

(Supplementary Table 1). The number of generations for each MA line was held constant to the

observed number for the line. Upon reaching Gmax=300, n FPs were added to each O2MA

line, drawn from a Poisson distribution with parameter λ_{FP}= the expected number of FPs per

line. We next calculate μ_{O1} and μ_{O2} from the simulated data and subtract μ_{O1} from μ_{O2} to find

$\Delta\mu$. This procedure was repeated 10,000 times to generate a null distribution of $\Delta\mu$. The value of

λ_{FP} for which the observed value of $\Delta\mu$ exceeds 95% of simulated values is the maximum

number of FPs per O2MA line that is consistent (at p>0.05) with $\mu_{O2} = \mu_{O1}$, i.e., $\Delta\mu > 0$ cannot

be explained by FPs if $\mu_{O2} = \mu_{O1}$.

ii) Variance in mutation rate. Also of interest is the question: does the variance in mutation rate

(V_μ) change over time? We address this question in two ways. First, from the same simulations

described in the preceding section. For each simulation run we calculate the variance among

O1MA lines $(V_{\mu O1})$ and O2MA lines $(V_{\mu O2})$ to generate the relevant null distributions. We then

compare the observed variances to the null distributions to test the hypothesis that there is more

variance in mutation rate than expected if the mutation rate is uniform[10].

Second, the design of the experiment permits a simple test of the heritability of mutation

rate over the time frame of the first-order MA ($\approx$ 150 generations). The phenotypic correlation

between relatives $\rho_{x,y} = r_{x,y}h^2$, where $r_{x,y}$ is the coefficient of relatedness between individuals x

and y and h^2 is the narrow-sense heritability[49] (Equation 6.4 in Ref. 49). The two O2MA

sublines of an O1MA family are identical twins at the time of "birth" (Gmax=150), and $r_{x,y}$ = 1.

Thus, the heritability of mutation rate is simply the correlation between the mutation rates of the

O2MA sublines.

iii) Fitness effects. Competitive fitness W is defined as the logarithm of the ratio of the

frequency of focal strain worms (p) to competitor strain worms (1-p) on an assay plate, i.e., W =

log[p/(1-p)].   To quantify how mutations contribute to competitive fitness, we applied a fully

Bayesian mixed-effects modeling framework[50]. We compared two Bayesian models: a one-

stage ("uniform effects") model that consolidates all mutations into a single genetic effect

structure using a unified genetic covariance matrix, and a two-stage model that explicitly

partitions mutations into first-order (O1) and second-order (O2) effects, each with its own

genetic parameters and separate covariance matrices. Each value of W for observation i is

modeled as the sum of a global intercept (c), a block-specific deviation ($b_{block[i]}$), genetic effects

(g_i), and a residual error (ε_i):

$$W_i = c + b_{block[i]} + g_i + \varepsilon_i, \quad \varepsilon_i \sim N(0,\, \sigma_e^2).$$

The block effects (b_j) account for systematic batch-to-batch variability arising from separate

experimental runs, modeled as independent deviations from a normal distribution:

$$b_j \sim N(0,\, \sigma_b^2).$$

The genetic effects (g_i) capture heritable variation in W explained by genetic similarity among

strains. Following the Central Limit Theorem, we assume that the cumulative effect of multiple

mutations can be approximated by a normal distribution. For each individual i, the genetic effect

is decomposed into two additive components:

$$g_i = g_{1,i} + g_{2,i},$$

where $g_{1,i}$ represents the genetic effect from first-order mutations (accumulated from

generations 0 to 150) and $g_{2,i}$ represents the genetic effect from second-order mutations

(accumulated from generations 150 to 300). Each genetic effect component is assumed to

follow a normal distribution with mean and variance proportional to the number of mutations:

$$g_{1,i} \sim N(n_{1,i}\mu_1,\, n_{1,i}\sigma_1^2) \quad and \quad g_{2,i} \sim N(n_{2,i}\mu_2,\, n_{2,i}\sigma_2^2),$$

where $n_{1,i}$ and $n_{2,i}$ represent the number of mutations for individual i in the first-order and

second-order experiments, respectively. The parameters μ_1 and μ_2 represent the mean fitness

effect per mutation, while σ_1^2 and σ_2^2 represent the variance in fitness effects per mutation.

In the one-stage ("uniform effects") model, we model the total genetic effect for each individual

using a single genetic covariance matrix (G) that combines both O1 and O2 mutations:

$$g_i \sim N(n_{total,i}\mu_g,\, n_{total,i}\sigma_g^2), \quad n_{total,i} = n_{1,i} + n_{2,i},$$

where μ_g and σ_g represent the overall genetic mean and scale parameters, and the covariance

structure is captured by a unified G matrix incorporating all mutations. In the two-stage model,

we model the genetic effects separately using independent covariance matrices (G_1 for O1

mutations and G_2 for O2 mutations):

574               $$g_{1,i} \sim N(n_{1,i}\mu_{g1},\ n_{1,i}\sigma_{g1}^2) \quad and \quad g_{2,i} \sim N(n_{2,i}\mu_{g2},\ n_{2,i}\sigma_{g2}^2),$$

where μ_{g1}, μ_{g2}, σ_{g1}^2, and σ_{g2}^2 represent the separate genetic parameters for first-order and

second-order mutations, and the covariance structure is partitioned

between G_1 and G_2 matrices.

We implemented both models using PyMC[51] with weakly informative priors to regularize

estimation. Intercepts and genetic means were assigned $N(0,10)$ priors, while standard

deviations received HalfNormal(5) priors. Model parameters were estimated using the No-U-

Turn Sampler (NUTS) with 2,000 tuning and 20,000 posterior draws.

To account for potential measurement errors in mutation counts, we implemented a

bootstrapping approach that simulates both false positives and false negatives. False positives,

which only affect O2 mutations, were modeled using a Poisson distribution with expectation $\lambda =$

3. False negatives, which in this context include all mutations missed due to failure to recall,

were also modeled using a Poisson distribution with rate

587               $$\lambda_{missed} = \mu_{mutation} \times n_{generations} \times genome\_length \times failure\_rate,$$

where $\mu_{mutation}$ represents the per-generation mutation rate (2.13×10^{-9} for O1, 2.47×10^{-9} for

O2), $n_{generations} = 150$, genome_length $= 10^8$ bp (the approximate size of the $C.\ elegans$ genome),

and failure_rate $= 0.025$. For each bootstrap iteration, we resampled mutation counts by

subtracting false positives from observed O2 counts, adding missed mutations to both O1 and

O2 counts, and fitting the Bayesian model with the resampled mutation counts. This approach

allows us to propagate uncertainty in mutation counts through to our parameter estimates while

maintaining the mathematical structure of our genetic effects model.

Data Availability

Raw sequence data are deposited in the NCBI Sequence Read Archive BioProject

#PRJNA1089216. Competitive fitness data are deposited in Dryad UPON ACCEPTANCE.

Code and unprocessed image files are available at https://github.com/Baer-Group/O2MAno2.

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Figure Legends

Figure 1. Theoretical models of mutation rate evolution. (**A**) Cost of Fidelity. (**B**) Drift Barrier.

Figure 2. (**A**) Schematic depiction of the MA experiment. Short black lines depict ancestral G0

"pseudolines" (PS), blue lines depict O1MA lines, orange lines depict O2MA lines. (**B**)

Schematic depiction of the competitive fitness assay. Label "LX" refers to "Line X" (MA or G0),

label LX.1,…LX.7 refers to replicates 1-7 of a line in the assay. See Methods for details.

Figure 3. Distribution of simulated values of $\Delta\mu=\mu_{O2}-\mu_{O1}$ with increasing number of false

positives. Red dashed line is the observed value. "FP cutoff" is the minimum number of false

positives necessary for the data to be consistent with the hypothesis $\Delta\mu=0$. (**A**) SNVs (**B**)

Indels.

Figure 4. Type-specific mutation rates and spectra. (**A**) SNV mutation rates (**B**) SNV frequency

spectrum (**C**) Indel mutation rates (**D**) Indel frequency spectrum.

Figure 5. Distribution of posterior probabilities of mutational effects u on competitive fitness. (**A**)

Mean u (**B**) Standard deviation σ_G.

Supplementary Figure 1. Cost of Fidelity with mutational bias. Curves depict the relationship

between fitness (*W*) and mutation rate (μ). Blue horizontal arrows represent the input of

mutational variance for mutation rate; red vertical arrows represent the strength of stabilizing

selection against (anti)mutator variants; $w(\mu) = \exp\left(-\frac{\mu^2}{2\sigma^2}\right)$. Thickness of the arrows depicts the

magnitude of the effect.

Supplementary Figure 2. Distribution of mutations along chromosomes. Top panel, SNVs;

bottom panel, indels.

Supplementary Figure 3. Schematic depiction of the effects of false inferences. Stem branch

is the O1MA line, bifurcating terminal branches are the two O2MA lines derived from the O1MA

line. Mutations are shown represented as arrows on the lineage. True sequence inferences are

shown in blue text at sequence position 6; false inferences are shown in red text.

Supplementary Figure 4. Simulated distributions of variance in line mean mutation rates under

the assumption of uniform mutation rate and the observed average number of false positives

(FP) per line. Vertical dashed lines show the observed values. (**A**) SNVs, 2 FPs/line (**B**) Indels,

0.33 FPs/line.

Supplementary Figure 5. Distribution of line mean competitive fitness *W* by treatment (MA vs

G0) and assay block. (**A**) G0 pseudolines (**B**) G150 MA lines (**C**) G300 MA lines (**D**) Comparison

across assay blocks.

Supplementary Figure 6. Plots of line mean fitness against mutation rate. (**A-C**) All data; (**E-G**)

generation 300 data only.

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Table 1. Summary statistics (SEM in parentheses). N_{Lines} is the number of MA lines, $\bar{t}$ is the average number of generations of MA

(top row) and number of days of MA (bottom row). Values of $\bar{\mu}$ are mutation rates ($\times 10^9$) calculated from the observed number of

mutations without correction for false inferences. Values of $\hat{\mu}$ are estimates of mutation rates ($\times 10^9$) corrected for false positives (FP)

and mutations that were failed to recall (FtR). We assume the average number of FPs (2 SNVs and 0.33 indels) per genome and the

average proportion of simulated mutations that were FtR at 3X coverage (2.2% SNVs, 2.5% indels). Estimated mutation rates ($\hat{\mu}$) are

calculated as $\dfrac{n-\overline{FP}+\overline{FtR}}{L \times t}$, where n is the observed number of variants and L (100,286,401 bp) is the size of the *C. elegans* N2

reference genome.

	N_{Lines}	%Genome	$\bar{t}$	$\bar{\mu}_{SNV}$	$\bar{\mu}_{DEL}$	$\bar{\mu}_{INS}$	$\bar{\mu}_{TOTAL}$	$\hat{\mu}_{SNV}$	$\hat{\mu}_{DEL}$	$\hat{\mu}_{INS}$	$\hat{\mu}_{TOTAL}$
O1	96	97.25 (0.051)	145.9 (0.76)	1.77 (0.027)	0.25 (0.011)	0.113 (0.0071)	2.13 (0.0296)	1.76 (0.027)	0.25 (0.010)	0.11 (0.007)	2.00 (0.0296)
			600 (0.0062)	0.427 (0.0062)	0.061 (0.0027)	0.028 (0.0017)	0.52 (0.0072)	0.43 (0.0062)	0.0605 (0.0027)	0.028 (0.0017)	0.452 (0.0064)
O2	192	97.25 (0.051)	144.75 (1.55)	1.88 (0.032)	0.368 (0.016)	0.23 (0.013)	2.47 (0.046)	1.72 (0.03)	0.34 (0.016)	0.21 (0.018)	2.06 (0.0378)
			620 (0.28)	0.431 (0.0075)	0.084 (0.0034)	0.053 (0.0026)	0.567 (0.0098)	0.40 (0.0074)	0.079 (0.0034)	0.047 (0.0026)	0.443 (0.0083)

[A]

[B]