

1 The genomics of local adaptation in trees:
2 Are we out of the woods yet?

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11 **Abstract**

12 There is substantial interest in uncovering the genetic basis of the traits underlying adaptive
13 responses in tree species, as this information will ultimately aid conservation and industrial
14 endeavors across populations, generations, and environments. Fundamentally, the
15 characterization of such genetic bases is within the context of a genetic architecture, which
16 describes the multidimensional relationship between genotype and phenotype through the
17 identification of causative variants, their relative location within a genome, expression,
18 pleiotropic effect, environmental influence, and degree of dominance, epistasis, and additivity.
19 Here, we review theory related to polygenic local adaptation and contextualize these
20 expectations with methods often used to uncover the genetic basis of traits important to tree
21 conservation and industry. A broad literature survey suggests that most tree traits generally
22 exhibit considerable heritability, that underlying quantitative genetic variation (Q_{ST}) is structured
23 more so across populations than neutral expectations (F_{ST}) in 69% of comparisons across the
24 literature, and that single-locus associations often exhibit small estimated per-locus effects.
25 Together, these results suggest differential selection across populations often acts on tree
26 phenotypes underlain by polygenic architectures consisting of numerous small to moderate
27 effect loci. Using this synthesis, we highlight the limits of using solely single-locus approaches to
28 describe underlying genetic architectures and close by addressing hurdles and promising
29 alternatives towards such goals, remark upon the current state of tree genomics, and identify
30 future directions for this field. Importantly, we argue, the success of future endeavors should not
31 be predicated on the shortcomings of past studies and will instead be dependent upon the
32 application of theory to empiricism, standardized reporting, centralized open-access databases,
33 and continual input and review of the community's research.

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84 **Introduction**

85 Trees are plants with an arborescent habit, which is loosely defined as a tall-statured
86 growth form usually producing wood (reviewed by Petit & Hampe 2006). Approximately 15% to
87 25% of plant taxa are classified as trees (Oldfield *et al.* 1998; Grandtner 2005; Wortley &
88 Scotland 2004), with forested ecosystems accounting for approximately 30% of terrestrial
89 vegetation (Costanza *et al.* 1997) and providing habitat for terrestrial biodiversity. Indeed, trees
90 play important ecological roles in diverse communities across the globe, such as vertical
91 structural habitat, seeds for wildlife forage, forest cover, the production of oxygen, carbon
92 sequestration, air and water filtration, as well as the reduction of erosion, protracting snowmelt,
93 and desertification. Of these, biological roles are ultimately defined by a set of life history char-
94 acteristics common to most tree species (Petit & Hampe 2006. These include predominantly
95 outcrossing mating systems with high levels of gene flow and fecundities, as well as long
96 lifespans and generation times (Loehle 1988; Mitton & Williams 2006; Savolainen *et al.* 2007),
97 although these may differ in, for example, clades of tropical trees. As a result, tree species
98 typically have large effective population sizes, moderate to high levels of genetic diversity, and
99 frequent occurrences of locally adapted ecotypes (Savolainen *et al.* 2007; Alberto *et al.* 2013;
100 Sork *et al.* 2013; Boshier *et al.* 2015; Prunier *et al.* 2015; Holliday *et al.* 2017). Across species,
101 however, rates of morphological and molecular evolution tend to be slow (reviewed in De La
102 Torre *et al.* 2017). Additionally, genome size varies enormously across species of trees, ranging
103 from 0.4Gbp to 31Gbp (reviewed in Neale *et al.* 2017). Recent sequencing efforts in
104 gymnosperms, which represent the largest tree genomes, reveal that much of genome size
105 variation is due to transposable element dynamics and gene family evolution (Leitch & Leitch
106 2012; Morse *et al.* 2009; Nystedt *et al.* 2013; Prunier *et al.* 2015; Neale *et al.* 2017) where
107 duplication events of select gene families may contribute to the ability of trees to colonize
108 marginalized habitats (Leitch & Leitch 2012; Prunier *et al.* 2015; Neale *et al.* 2017).

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109 In trees, the general presence of large geographical ranges and extensive gene flow
110 also provides an ideal setting to disentangle neutral from selective evolutionary processes
111 (Neale & Kremer 2011). Indeed, their longevity and wide and heterogeneous geographical
112 distributions lend trees suitable for addressing several key evolutionary questions about the
113 importance of historical climatic fluctuations, and local adaptation involving shifts in allele
114 frequencies (Lotterhos & Whitlock 2014; Savolainen *et al.* 2007, 2013; Platt *et al.* 2015). As we
115 detail in subsequent sections, evidence consistent with local adaptation in trees is ubiquitous,
116 even across fine spatial scales where it had been hypothesized that gene flow may overcome
117 selection of locally favored alleles (e.g., Mitton *et al.* 1998; Budde *et al.* 2014; Csilléry *et al.*
118 2014; Vizcaíno-Palomar *et al.* 2014; Eckert *et al.* 2015; Holliday *et al.* 2016; Roschanksi *et al.*
119 2016; Lind *et al.* 2017).

120 Quantitative phenotypes are often used as a proxy for total lifetime fitness, which is com-
121 posed of two broad components: survival and reproduction. Since most quantitative traits are
122 related to some component of total lifetime fitness, they are often used to assess potential for
123 local adaptation. For many plant taxa, selection pressures are expected to be strongest for vari-
124 ation in survival during the juvenile stages of development (Donohue *et al.* 2010), particularly for
125 those taxa with high reproductive output, as is the case for many tree species. As such, juvenile
126 stages in plants have been found to contribute substantially to total lifetime fitness (Postma &
127 Agren 2016). Phenotypic traits associated with juvenile survival have thus received the majority
128 of genetic research focus in trees, particularly due to their long-lived nature. Such studies have
129 led to intriguing insights gained through a long history of common garden experimentation
130 (Langlet 1971; Morgenstern 1996). For example, traits such as growth (e.g., height and
131 diameter), form (e.g. specific gravity, straightness), phenology (e.g. bud flush, bud set), juvenile
132 performance (e.g. germination rate, seed traits) and physiology (e.g. cold hardiness, water-use
133 efficiency) have all been shown to be under moderate to high genetic control (reviewed in Corn-
134 elius 1994, Howe *et al.* 2003, Alberto *et al.* 2013; this review). Variation for these traits is also

135 often partitioned among populations (this review), despite the vast majority of neutral variation
136 remaining within populations (Howe et al. 2003; Neale & Savolainen 2004). With few exceptions
137 (e.g., major gene resistance in the white pine-blister rust pathosystem; Kinloch et al. 1970; Liu
138 et al. 2017), variation for these traits forms a continuum across individuals, thus implying that
139 the underlying genetic architecture is composed of a large number of small to moderate effect
140 loci (i.e., a polygenic architecture; concept reviewed in Savolainen et al. 2007, 2013; Gagnaire &
141 Gaggiotti 2016; Hoban et al. 2016; Timpson et al. 2017). There is some uncertainty, however,
142 concerning the properties of the effect size distributions comprising polygenic architectures
143 (*sensu* Fisher 1930, Kimura 1983, and Orr 1998), the relative importance of various forms of
144 gene actions (e.g., dominance, epistasis) in producing trait variation (Crow 2010, Hansen 2013),
145 how these interact to affect the evolution of polygenic architectures in natural populations
146 (Hansen 2006), and how these factors will ultimately influence evolutionary processes and out-
147 comes in forest trees (Savolainen et al. 2007; Sork et al. 2013; Prunier et al. 2015). Consider-
148 able strides, made in the past through genotype-phenotype-environment studies (*sensu* Sork et
149 al. 2013), have contributed intriguing insight into the genomic basis of local adaptation for tree
150 species. However, given the large genome size of many tree species, such methods have been
151 criticized as lacking in power and sufficient coverage needed to detect small effect loci, which is
152 further exacerbated by rapid decay of linkage disequilibrium (LD) in most forest trees (Mackay
153 2009; Savolainen et al. 2007). Despite these limitations, association studies have been
154 moderately successful in linking genotypes and phenotypes, including providing information for
155 making inferences about local adaptation.

156 In this review, we set out to summarize theory related to polygenic local adaptation and,
157 using these expectations, contextualize the progress of describing the genetic architectures
158 underlying traits important to conservation and industry in undomesticated tree species. We first
159 highlight the extensive evidence for local adaptation in trees by reviewing transplant designs
160 often used in investigations of quantitative genetic differentiation. Using an extensive literature

161 survey across both gymnosperm and angiosperm species, we provide an overview of these
162 transplant methods, give examples of each, and quantify the distribution of narrow sense herita-
163 bility and Q_{ST} estimates across various trait categories. We further use this survey to establish
164 patterns of comparative quantitative and neutral genetic differentiation (i.e., Q_{ST} - F_{ST} tests) which
165 until this review had not been suitably synthesized in trees. Before we transition into discussing
166 common methods used to uncover loci underlying adaptation, we establish expectations for the
167 genetic architecture of polygenic, fitness-related traits by reviewing the theory available to date.
168 We then provide an extensive review of genotype-phenotype associations in trees and provide
169 the distribution of the percent phenotypic variance explained by empirically associated loci.
170 Using this distribution, we underscore the limitations of using solely single-locus approaches to
171 uncover the loci underlying local adaptation in tree species. Given this synthesis, we highlight
172 exemplary genomic resources available to fill knowledge gaps, identify promising avenues of
173 future research, identify key benchmarks and necessary steps towards truly integrating studies
174 of trees into the genomic era, and address our primary question, “Are we out of the woods yet?”.

175 **Identifying heritable phenotypic variation**

176 Trees have evolved numerous adaptations as a result of their vast ecological breadth. As such,
177 it has long been the goal of forest scientists to understand the traits important to viability and
178 persistence. Among the most frequent designs used, common gardens and reciprocal
179 transplants have aimed at describing genetically based differentiation of measured phenotypes
180 among various source populations of varying sizes and across various geographic scales.
181 Across these designs, investigators seek to better understand the phenotypes relevant to local
182 adaptation and the selective pressures influencing these phenotypes. The exact design chosen,
183 however, is generally based on the questions driving the research endeavor and often by the
184 availability of resources (Morgenstern 1996; Blanquart et al. 2013; de Villemereuil et al. 2015).
185 In this section, we briefly review these designs, identify relevant questions and inferences,

186 highlight some of the important practical applications of these techniques, and discuss
187 examples of past investigations in various tree species.

188 There is a rich history of forest scientists using the common garden approach dating
189 back hundreds of years (Langlet 1971; Mátyás 1996). In a broad sense, a common garden
190 design is used to test for differentiation among genetically distinct groups in a homogeneous
191 environment. These groups can be clonal replicates or sibships (families) derived from species
192 or hybrids sampled from various populations, provenances, varieties, cultivars, or agricultural
193 accessions (Cheplick 2015). When individuals from various origins are grown together under the
194 same conditions, the observed phenotypic differentiation is expected to reflect underlying gen-
195 etic variation, especially when maternal effects are assumed or shown to be absent. Common
196 garden and provenance trial designs can also establish evidence that the phenotypes under
197 study are heritable, a prerequisite for an adaptive response to selective agents (Supplemental
198 Box S1), and that populations exhibit quantitative genetic differentiation (i.e., Q_{ST} ; Spitz 1993).
199 When driven by questions related to differentiation alone, a single common garden approach
200 can be used to describe levels of quantitative genetic variation within and among genetically dis-
201 tinct groups. In these cases, no environmental variables are manipulated, and thus, unequivocal
202 evidence for trait divergence among groups, and the contributing factors influencing this diver-
203 gence (e.g., neutral or selective processes), is often limited because conclusions must be based
204 on *post hoc* inferences about source environments for the materials established in the common
205 garden. Even so, single common garden approaches can be a powerful tool to demonstrate
206 evidence congruent with local adaptation. For instance, the white carob tree (*Prosopis alba*
207 Griseb., Leguminosae) growing in Argentina is an ideal multipurpose tree that has potential for
208 use in reforestation and afforestation applications in the region. However, this genus is known to
209 invade other regions, encroach on farmland and waterways, and has a thorny growth habit that
210 can cause sepsis in livestock. To better understand how forestry applications can balance the
211 benefits of production and forest protection, Bessega et al. (2015) used a single common

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212 garden representing eight provenances of *P. alba* to compare estimates of neutral genetic
213 patterns to the quantitative genetic variation of life history traits related to economic importance.
214 They found that for most traits there existed considerable underlying genetic variation ($\overline{Q_{ST}} =$
215 0.139). Additionally, source environments were often correlated with measured trait variation in
216 the common garden, suggesting that the observed differentiation was driven by temperature,
217 precipitation, wind speed, and sunshine fraction, with signals of divergent selection corroborated
218 across Q_{ST} - F_{ST} comparisons and tests for selection (e.g., S test, *sensu* Ovaskainen et al., 2011).
219 Bessega et al. (2015) concluded that the signal of non-neutral differentiation was indicative of
220 divergent phenotypic optima across populations, and that this variation could be used to direct
221 future breeding programs across the region.

222 When there is evidence that environmental differences among source populations may
223 be driving adaptive divergence, strong environmental candidates can be manipulated (artificially
224 or via site selection) in a multiple common garden design to further investigate hypotheses of differ-
225 entiation and adaptation. For instance, the sweet chestnut (*Castanea sativa* Mill., Fagaceae),
226 also known for its edible fruit, is distributed across much of Minor Asia and southern Europe and
227 is an ecologically important component of many Mediterranean systems. *Castanea sativa* exhibits
228 ecological, physiological, morphological, and genetic variability as the range overlays a climatic
229 transition from xeric Mediterranean conditions to wetter Euro-Siberian environments (see refs in
230 Lauteri et al., 2004). Previous common garden experiments carried out by Lauteri and col-
231 leagues have indicated that populations across this transition are further differentiated by water
232 use efficiency (the ratio of plant carbon gain to water loss) and carbon isotope discrimination, Δ .
233 To further explore variability of drought-related traits, Lauteri et al. (2004) used an *ex situ*
234 multiple common garden design using two water and temperature treatments in individual
235 climatic chambers to assess differentiation among six populations across Spain, Italy, and
236 Greece. They found *treatment* and *population x treatment* effects were significant, suggesting

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237 variation in drought adaptation across populations. Additionally, populations originating from dry
238 sites generally exhibited higher values of Δ , which was also composed of significant additive
239 genetic variation ($h^2 = 0.15-0.52$), and suggests that genetic and physiological mechanisms of
240 drought adaptation confer a capacity to colonize a wide arrange of environmental conditions,
241 while strong negative relationships between Δ and growth-related traits is suggestive of strong
242 evolutionary constraints at juvenile stages.

243 While *ex situ* common gardens approaches (e.g., Lauteri et al. 2004) can provide strong
244 evidence of adaptive divergence among populations, and in some cases corroborate putative
245 drivers of observed differentiation, these studies can often exclude key environmental factors,
246 possibly leading to confounding signals of adaptation (Kawecki & Ebert 2004). When *in situ*
247 experimentation is feasible, site selection can be used to test for environmental drivers of local
248 adaptation. For example, Evans et al. (2016) investigated traits related to growth and phenology
249 in juvenile narrowleaf cottonwood (*Populus angustifolia* James, Salicaceae) by planting families
250 from nine populations across the native range into three common gardens, one each at the
251 northern, southern, and interior extent of the range. Using $Q_{ST}-F_{ST}$ comparisons and clinal
252 analyses alongside the quantitative genetic analyses, Evans et al. (2016) concluded that climate
253 cues played a major role in structuring adaptive variation across the range of *P. angustifolia*,
254 and that future industrial and conservation applications should utilize this information to inform
255 source environments for optimal outcomes.

256 As both *in situ* and *ex situ* common garden trials can include multiple environmental
257 influences in their design, reciprocally transplanting to all source environments is not necessarily
258 a requirement to decompose genetic variation underlying adaptive traits or to provide evidence
259 for, or the drivers of, differentiation among populations. Thus, these designs may preclude
260 inferences regarding local adaptation *sensu stricto*. To produce such evidence, source
261 populations can be planted in a (full- or incomplete-factorial) reciprocal transplant design and

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262 allow for traits related to fitness to be assessed across native and non-native environments. If a
263 population is locally adapted, individuals exposed to their native environments should show
264 increased growth, survival, and reproduction relative to non-native genotypes (Kawecki & Ebert
265 2004; Leimu & Fischer 2008; Hereford 2009; Savolainen et al. 2013). For example, with the
266 goal of delineating conservation units based on molecular and quantitative trait differentiation,
267 Rodríguez-Quilón et al. (2016) used four reciprocally-transplanted common gardens to assess
268 height and survival of samples from 35 natural populations of maritime pine (*Pinus pinaster*
269 Aiton, Pinaceae). For both traits, Q_{ST} was consistently larger than F_{ST} across the four sites, a
270 pattern suggestive of divergent selection. Six distinct gene pools based on evolutionary history
271 of neutral markers were identified, and because high quantitative differentiation (Q_{ST}) was found
272 within these pools, hierarchical analyses were used to further identify ten adaptive population
273 groups for use in conservation and breeding approaches.

274 Available evidence suggests that many populations of tree species have substantial
275 heritable genetic variation, and that the quantitative traits under study often show signals of
276 divergent selection across both broad and fine spatial scales. But how broadly can we apply this
277 statement? Are there overall patterns of heritability and quantitative genetic structure across
278 tree species? Because estimates of heritability and Q_{ST} are often only applicable to a specific
279 set of populations, for a specific set of environments, at any specific point in time (e.g., see
280 Figure 2D), a large sample of these estimates is therefore necessary to synthesize the current
281 literature with regard to patterns across taxa. To accomplish this aim, we synthesized estimates
282 from 129 published studies with estimates of narrow sense heritability ($n = 114$) from replicated
283 progeny trials and/or estimates of quantitative genetic differentiation (Q_{ST} ; $n = 37$). However, we
284 excluded papers that have been cited for estimates of Q_{ST} or heritability that were calculated
285 *post hoc* from variance components (i.e., we only recorded estimates that were explicitly
286 reported as h^2 or Q_{ST} in the original publication). For comparison, we further grouped measured

287 traits into 14 broad categories: cold hardiness, disease resistance, drought hardiness, form,
288 growth, herbivore and insect resistance, leaf and needle properties, phenology, plant secondary
289 metabolites, reproduction, resource allocation, seed and early germination properties, survival,
290 and wood properties. Because sample size can influence the precision of both heritability and
291 Q_{ST} , for each trait category we used a weighted average where weights were equal to the
292 number of families used to estimate variance components for each estimate of h^2 and Q_{ST} .

293 In agreement with Cornelius (1994), our survey found that many of the traits important to
294 conservation and industry exhibit non-zero narrow sense heritability ($\overline{h^2} = 0.367$; File S1;
295 Figures S1-S4) and are thus amenable to selection. The mean weighted Q_{ST} across traits
296 groups from our survey (Table S1; File S1) was between 0.10-0.28, except for drought
297 hardiness (0.06) and disease resistance (0.04), with median values from the unweighted
298 distribution generally falling below the weighted average for each trait group (Figure 1). This
299 suggests that over various geographic and environmental distances, population histories, and
300 species, there is a general pattern of substantial genetic variation underlying measured traits.
301 Given our synthesis of Q_{ST} estimates in trees, we were curious of the evidence for adaptive
302 divergence among populations ($Q_{ST} > F_{ST}$). Of the 37 articles reporting Q_{ST} estimates in our
303 review, 23 compared Q_{ST} with F_{ST} or G_{ST} estimated from the same populations under study
304 (however, we excluded studies that used F_{ST} measurements taken from the literature, e.g., as in
305 McKay & Latta 2002; Alberto et al. 2013). Indeed, as pointed out by Crnokrak & Merilä (2002),
306 comparisons of Q_{ST} and F_{ST} estimated from different populations and/or at different time points
307 are uninformative. Of these 23 studies, 18 compared Q_{ST} and F_{ST} in a statistical framework
308 while the remaining five studies compared Q_{ST} and F_{ST} numerically. Across numerical and
309 statistical comparisons combined, 67% (254 of 381 traits) exhibited higher Q_{ST} than F_{ST} , with
310 69% (170 of 246 traits) exhibiting significantly higher Q_{ST} than F_{ST} . Although we did not tally
311 instances where Q_{ST} was reported to be less than F_{ST} (statistically or otherwise), as this was not

312 the focus of our review, there were some instances in which this was the case. For instance,
313 Lamy et al. (2011) found such patterns when quantifying population genetic differentiation of
314 cavitation resistance across the species range of maritime pine (*Pinus pinaster* Aiton,
315 Pinaceae), while Mahalovich et al. (2011) also found that $Q_{ST} < F_{ST}$ for traits related to white
316 pine-blister rust resistance in inoculated seedlings of whitebark pine (*Pinus albicaulis* Engelm.,
317 Pinaceae). While various explanations for such patterns were outlined by Lamy et al. (2011),
318 canalization was argued as the most likely process driving the observed patterns, while
319 Mahalovich et al. (2011) offered similar arguments for selection favoring the same genotype in
320 different environments (see Lamy et al. 2012 for more regarding this aspect).

321 Despite neutral genetic differentiation partitioned primarily within populations, adaptive
322 genetic variation seems to be structured to a greater degree across populations, more often
323 than not, for the various fitness-related traits reviewed here. Such a pattern is indeed consistent
324 with local adaptation, assuming that (among other considerations such as the recency of
325 selection) mutation rates are considerably lower than migration rates in these populations
326 (Whitlock 1999; Hendry 2002; Leinonen et al 2013). In any case, given an extensive literature
327 supporting the local adaptation hypothesis in trees, our results appear consistent with patterns
328 of selective forces acting on abundant, heritable genetic variation across populations, even in
329 the face of gene flow (discussed further in the next section).

330 **Expectations for the loci underlying quantitative traits**

331 The homogenous environments of the common garden and reciprocal transplant
332 designs are ideally suited to test hypotheses of local adaptation in trees (Sork et al. 2013).
333 However, uncovering the genetic basis and contributory influence of specific loci underlying
334 these adaptive traits is a sizable endeavor on its own, and the success of such pursuits will be
335 determined, in part, by the trait's underlying genetic architecture (i.e., the number, effect size,
336 type, location, expression, pleiotropic effect, environmental influence, and interaction of under-

337 lying loci), which is generally not known *a priori* (Stinchcombe & Hoekstra 2008; Rellstab et al.
338 2015; Savolainen et al. 2013; Hoban et al. 2016; Burghardt et al. 2017; Wadgyamar et al. 2017).
339 Much of our early understanding of the architectures of complex traits came shortly after
340 Nilsson-Ehle (1909) and East (1910) independently demonstrated evidence for multiple-factor
341 inheritance, where Fisher (1918) laid the groundwork for quantitative genetics by incorporating
342 the additive properties of variance to partition phenotypic variation into components tractable to
343 a model of Mendelian inheritance. It was this work, and that of Fisher's geometric model (1930),
344 which founded the basis for attributing continuous variation of phenotypes to a polygenic model
345 of many underlying heritable components of mainly small effect. From this model, Fisher (1930)
346 concluded that mutations of small effect were the main drivers of adaptation, suggesting large-
347 effect substitutions to contribute little to adaptation due to negative pleiotropic effects
348 constraining effect size. Therefore, the fate of a given locus would be conditioned on its
349 average, marginal effect on fitness calculated across the species, with non-additive deviations
350 from this linear model of inconsequential influence. This micro-mutationist view, to a large
351 extent, remained the dominant thought for nearly half a century (Orr 2005). It was then that
352 Kimura (1983) established that for an allele to contribute to adaptation, it would need to survive
353 the stochastic nature of drift. Thus, new mutations of low frequency and effect were less likely to
354 contribute substantially to adaptive evolution. Considering the adaptive contribution probability
355 of large and small effect loci, Kimura concluded that mutations of moderate effect would be the
356 most plausible. Years later, Orr (1998) showed that over the entire bout of selection via an
357 adaptive walk, the distribution of fixed substitutions resembles an exponential distribution, with
358 effect size decreasing with the proximity to the phenotypic optimum. In addition, the distribution
359 of fitness effects of beneficial mutations is also expected to be exponential (Orr 2003; for more
360 discussion on this aspect, see also Orr 2006; Eyre-Walker & Keightley 2007; Martin &
361 Lenormand 2008, Kopp & Hermisson 2009b; Keightley & Eyre-Walker 2010, Dittmar et al. 2016).
362 Despite major advances in theory and technology, there still remains substantial uncertainty

363 regarding the exact number of loci underlying many adaptive traits, the effect size distribution of
364 these loci, and how the number of underlying loci and effect distribution may change under
365 various evolutionary regimes (Orr 2001; Slate 2005; Hansen 2006; Mackay et al. 2009). In this
366 section, we describe how various factors can contribute to the (perhaps, effective) number of
367 causative loci, and the distribution of effects underlying continuously distributed adaptive traits,
368 beginning first with aspects of the architecture itself (gene action), and concluding with
369 explanations of how various processes (e.g., selection) play an influential role in the evolution of
370 underlying genetic architectures. Establishing these expectations is essential for assessing
371 common approaches and guiding future directions. In the next section we then compare these
372 expectations with methods used in, and results from, genotype-phenotype associations in trees.
373 While we discuss these examples in isolation, we highlight the fact that the underlying biological
374 processes are often not independent.

375 ***Gene action***

376 The classical genotype-phenotype map is largely one of additive effects, and is
377 represented by a statistical regression of the phenotype on genetic content, as developed by
378 Fisher (1918) and extended by others (e.g., Cockerham 1954; Kempthorne 1954). Indeed,
379 much of the work done in trees has relied on such additive effects to describe heritable and
380 quantitative genetic variation (see previous section). In this model, the phenotypic variance is
381 partitioned into orthogonal (i.e., independent) contributions from the genetic variance (σ_G),
382 environmental variance (σ_E), and the variance due to interaction between genotype and
383 environment ($\sigma_{G \times E}$; Figure 2; see Supplemental Box S1). Further, σ_G is also the sum of
384 orthogonal variance components, each term representing a different form of gene action. The
385 additive, dominance, and epistatic terms respectfully designate the associated variance
386 contribution of independent alleles, the non-additive contribution to variance of interactions
387 among alleles at the same locus, and the contribution to variance of non-additive interactions

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388 among alleles at different loci (the latter of which can take one of many forms such as additive-
389 by-additive, additive-by-dominance, etc.; Lynch & Walsh 1998). As a result, non-additive gene
390 action is minimized as non-linear contributions to the overall phenotype (Moreno 1994; Whitlock
391 et al. 1995) which contributes little to the distinction of the different forms of dominance and
392 epistasis (Cheverud & Routman 1995; Hansen & Wagner 2001; Hermisson et al. 2003; Hansen
393 2006; Mackay 2014) nor towards the inference of aspects of the underlying genetic architecture
394 in general (Nelson et al. 2013; Huang & Mackay 2016).

395 These statistical conveniences afforded by Fisher and others led to the notion that such
396 non-additive effects were transient (i.e., are due to LD, which will decay with the relaxation of
397 selection), or that trends of statistical epistasis were representative of functional epistasis in
398 general, and therefore epistasis was unimportant to evolutionary dynamics (e.g., Bulmer 1980;
399 Crow 2008, 2010; Hill et al. 2008). While minimized in a statistical regression, this does not
400 necessarily mean that epistasis and dominance will not have a profound impact on the genetic
401 architecture, or towards a given population or species' long-term evolutionary trajectory, even if
402 statistical epistatic or dominance variance is minimal (Goodnight 1988; Chevrud & Routman
403 1995; Hansen & Wagner 2001; Hansen 2013; Nelson et al. 2013; Griswold 2015; Paixão &
404 Barton 2016). Indeed, parameterizing a model in which the type I sums of squares is
405 determined by non-additive parameters, as opposed to additive variance in the conventional
406 regression model, the majority of genetic variation is still captured by the primary effect in the
407 model regardless of the underlying architecture (Huang & Mackay 2016). Given the prevalence
408 of evidence for non-additive contributions (e.g., Phillips 2008; de Visser et al. 2011; see also
409 references in Hansen 2013), it is likely that non-additive effects will play a role in evolutionary
410 outcomes. For instance, Huber et al. (2017) showed that the degree of dominance in
411 *Arabidopsis* is an outcome based upon functional importance and optimal expression level.
412 Further, Carter et al. (2005) show that, relative to a purely additive trait (or with non-directional
413 epistasis) under directional selection, positive and negative epistasis can respectfully increase

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414 or decrease the additive genetic variance, and thus increase or decrease the rate of phenotypic
415 response to selection (see also Le Rouzic & Álvarez-Castro 2016). As Jones et al. (2014) show,
416 for a two-trait phenotype controlled by pleiotropic and epistatic effects, epistasis in the presence
417 of selection can also affect the mutational architecture of complex traits, where the average
418 allelic effect evolves to be negatively correlated with the average epistatic coefficient, the
419 strength of which is greater in larger population sizes. Yet, as described by Barton et al. (2016),
420 and further discussed by Barton (2017) and Paixão & Barton (2016), the infinitesimal model can
421 be generalized to include epistatic effects, particularly when the number of underlying loci is
422 large and selection on individual loci is weak. In the case of non-systematic, weak pairwise
423 epistasis, and without mutation or environmental noise, the infinitesimal model holds to a good
424 approximation (Barton et al. 2016). In the case of sparse epistasis with selection and a large
425 number of loci, the change in the trait mean over 100 generations is greater than that under a
426 purely additive architecture, and the decrease in additive genetic variance exceeds, to an
427 extent, that of the neutral case after about 30 generations (which is exacerbated with simpler
428 architectures), with a reduction of the frequency of segregating alleles with positive effect on the
429 trait (Barton et al. 2016; Barton 2017).

430 Despite an ongoing debate within the literature (Wright 1932; Whitlock 1995; Crow 2008,
431 2010; Gibson 2012; Zuk et al. 2012; Hansen 2013; Hemani et al. 2013; Nelson et al. 2013;
432 Mäki-Tanila & Hill 2014; Ávila et al. 2014; Paixão & Barton 2016), and given that there seems to
433 be no general prevalence of either positive or negative epistatic interactions (Mackay 2014), the
434 infinitesimal model is likely to continue to contribute to our understanding of the evolution of
435 complex traits, as exemplified in its application towards breeding applications (Turelli & Barton
436 1994) and specifically those successfully applied to trees (Savolainen et al. 2007;
437 Thavamanikumar et al. 2013; Isik et al. 2015; Grattapaglia 2017). Ultimately, the success of
438 such models will be conditioned on the context, as well as the distinction between physiological
439 and statistical gene action. Here, (higher order) non-additive contributions to phenotypic

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440 variance will likely have minimal deviations from the limit of the infinitesimal model in the short-
441 term, particularly if this is primarily due to independent, low-order interactions, and should thus
442 be applied with this in mind. As such, while short-term evolutionary processes are likely to hold
443 in this limit, identifying the non-additive loci which underlie the trait, and their respective gene
444 action, may still need further inquiry (Grattapaglia 2017). Indeed, it is often argued that non-
445 additive gene action is too often neglected in studies of complex traits (e.g., Carlborg & Haley
446 2004), possibly due to the large sample sizes required to detect significant interactions, and lack
447 of statistical power incurred due to multiple hypothesis testing (Mackay 2014). Given the recent
448 reduced cost of sequencing technology and availability of novel computational and laboratory
449 tools, future studies incorporating investigations of epistasis and dominance (where appropriate
450 and feasible) would contribute to our understanding of genetic architectures, quantitative trait
451 evolution, and breeding applications in trees (Vitezica *et al.* 2017). For example, breeding
452 applications assessing hybridization across divergent backgrounds, as is also prevalent across
453 species in nature, have shown the importance of non-additive effects in phenotypic outcomes
454 (as in *Eucalyptus*, e.g. Tan *et al.* 2017, and *Pinus*, e.g. Dungey 2001). Even so, the additive
455 model is still a powerful tool to describe the loci underlying adaptive traits.

456 Pleiotropy is another considerable factor influencing the expectations of the genetic
457 architecture of quantitative traits, its evolution or evolvability, and indeed the genotype-pheno-
458 type map (Hansen 2003; Orr 2006; Chevin *et al.* 2010b; Tenallion 2014). While multiple defin-
459 itions exist across the literature (see Paaby & Rockman 2013), pleiotropy is generally identified
460 as a single locus influencing multiple phenotypic traits. Other than linkage disequilibrium,
461 pleiotropy is the fundamental cause of genetic covariance among phenotypes (Lande 1980).
462 Given that the number of independent traits under selection is likely limited (Barton 1990),
463 pleiotropy likely plays a substantial role in evolutionary dynamics. It is expected that as the
464 number of traits, n , influenced by a locus increases, the probability of a beneficial mutation will
465 decrease with the effect size of a mutation; where the effect size, r , relative to the distance to

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466 the phenotypic optimum, $d \cdot n^{-1/2}$, must be (much) less than d in order to be beneficial (Fisher
467 1930; the so-called ‘cost of complexity’: Orr 2000). Yet, empirical data seem to contradict this
468 hypothetical cost, as the effect size of mutations often do not scale with pleiotropy in this way,
469 and instead increase with the dimensionality of targeted traits (Wagner et al. 2008; Wang et al.
470 2010). Additionally, universal pleiotropy, where all mutations affect all phenotypes, and where
471 there is no net directionality of mutations (i.e., mutational isotropy; both aspects as in Fisher
472 1930), has also been challenged by findings which suggest that only a fraction of phenotypic
473 traits are affected by pleiotropic loci (Wagner et al. 2008; Wang et al. 2010). Relaxation of such
474 assumptions from Fisher’s geometric model have shown that the total number of traits affected
475 by pleiotropy has a relatively decreased effect on the rate of evolution in more general models
476 (e.g., Martin & Lenormand 2006; see also Simons et al. 2017, and references in Wagner &
477 Zhang 2011 and Tenailon 2014). It seems that if model organisms (e.g., Pickrell et al. 2016,
478 Smith 2016) are taken as a bellwether for expectations in trees, pleiotropy is likely a contributing
479 factor for many quantitative traits. Thus, the fraction of beneficial mutations is likely limited when
480 the number of traits influenced is large, suggesting that the cost of complexity (or, more
481 precisely, pleiotropy) may be generally robust (Welch & Waxman 2003), particularly when a
482 population is close to its phenotypic optimum where selection acts against dimensionality of
483 pleiotropic effects (Zhang 2012). Thus, the degree of pleiotropy for underlying loci, distance
484 from phenotypic optima, and covariance among traits under selection can have profound effects
485 on evolutionary outcomes. This is particularly true for the evolvability of architectures and
486 distribution of effect sizes, which further depends on the variational autonomy of the traits
487 affected by pleiotropy and the modularity of mutations, the former of which is ultimately
488 determined by the direction and size of effect among a set of pleiotropic loci across a set of
489 characters (see Arnold 1992; Wagner & Altenberg 1996; Hansen 2003, 2006; Wagner et al.
490 2007; Chevin et al. 2010b; Wagner & Zhang 2011; MacPherson et al. 2015).

491 In many investigations of local adaptation, the primary interest is in trait evolution and

492 thus the underlying genetic components. As such, environmental effects and interactions are
493 not often pursued, or perhaps even detected (Yoder & Tiffin 2017), particularly in studies of a
494 single common garden or environment, and are instead treated in much the same way as
495 epistatic interactions discussed above. Nonetheless, genotypic effects can evolve through
496 genotype-by-environment interactions with a changing environment just as is the case for the
497 evolution of non-additive interactions with a changing genetic background (Hansen 2006).
498 Indeed, it is likely that consistent fluctuations in the environment would select for
499 environmentally-perceptive responses, which seems to be the case across many tree species
500 (Li et al. 2017). The contribution to the effect size distribution from GxE interactions will be a
501 function of the variation in selection across the environments experienced by the interacting
502 allele(s) as well as the level of gene flow between environments and fitness differences among
503 various genetic backgrounds, but to our knowledge such information (to the extent of that for
504 e.g., selective sweeps) is lacking within the literature.

505 *Negative selection*

506 Negative selection acts against deleterious mutations that arise within populations. It is
507 one, but not the only, mechanism that underlies stabilizing selection, defined at the level of the
508 phenotype where deviations from an optimal value are selected against. Optima in this
509 framework can be thought of either globally (i.e., across all individuals) or locally (i.e., individuals
510 within a population), where the latter can have varying optima across populations. The nature of
511 the optima (i.e., being local or global) affects the detectable trait architecture. For example, trait
512 architecture should be composed of rare alleles with a negative relationship between effect size
513 and allele frequency (cf. Eyre-Walker 2010 and references therein), where this relationship can
514 also be confounded with degree of dominance and gene expression network connectivity
515 (Huber et al. 2017), under models of a single global optimum. From a population genetic
516 perspective, the ubiquity of negative selection is encapsulated in the name background

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517 selection, which has extensive reviews about its presence in natural systems (Charlesworth
518 2013), its importance for the neutral and nearly neutral theories of molecular evolution (Ohta
519 1992, 1996), and its contribution to observable patterns of hitchhiking (Stephan 2010). Important
520 for the study of polygenic adaptation and its architecture, however, is that loci identified using
521 GWAS may also include segregating deleterious variation (as argued and hinted at in Eckert et
522 al. 2013b; cf. Yang et al. 2017; Gazal et al. 2017) as this creates trait variance, with little known
523 about their prevalence (including differential prevalence across traits), differentiation in
524 frequencies across populations (but see Zhang et al. 2016), and effects on downstream
525 inferences about divergent selection pressures across populations. It is sets of GWAS loci,
526 though, that are currently analyzed for signatures of local adaptation via spatially divergent (i.e.,
527 locally positive) natural selection (e.g., Berg & Coop 2014).

528 Recent exemplary work with expression networks in *Populus tremula* L. (Salicaceae;
529 Mähler et al. 2017) and the herbaceous *Capsella grandiflora* Boiss. (Brassicaceae; Josephs et
530 al. 2015, 2017a) have revealed intriguing insight into the effects of negative selection on the
531 architecture of complex traits in plants, as well as the relationship between network connectivity
532 and the strength of negative selection. In *P. tremula*, genes with expression levels that were
533 significantly associated with sequence variation were found more often in the periphery of the
534 co-expression network (lower network connectivity) than within network module hubs (higher
535 connectivity), while expression-associated SNPs were negatively correlated with network
536 connectivity and effect size, a pattern also found between connectivity and expression variance,
537 and minor allele frequency and QTL effect size (Mähler et al. 2017). Genes associated with
538 sequence variation had less skewed site-frequency spectra (i.e., the frequency distribution of
539 allelic variants) and lower estimates of nonsynonymous to synonymous divergence (d_N/d_S) than
540 genes not associated with sequence variation, together suggesting that genes within the
541 periphery of co-expression networks are likely under less selective constraint than those genes
542 with high network connectivity which likely experience greater intensities of purifying selection.

543 These genes thus tend to have more segregating variation and may be those most likely to be
544 detected with current sample sizes utilized in GWAS, which has implications for estimation of
545 trait architecture and its ‘degree’ of polygenicity. Even so, while there is prevalent evidence of
546 negative selection in trees (e.g., Krutovsky & Neale 2005, Palmé et al. 2009, Eckert et al.
547 2013a,b; De La Torre et al. 2017), more inquiry is needed.

548 ***Positive selection***

549 The temporal and spatial heterogeneity of selection can impact the evolution of genetic
550 architectures underlying adaptation. These impacts are often thought of on a spectrum of trade-
551 offs, with one end being antagonistic pleiotropy where allelic effects vary between positive and
552 negative on fitness across populations, and conditional neutrality where allelic effects on fitness
553 are positive in one or more populations and nearly zero in others (Anderson et al. 2012,
554 Savolainen et al. 2013). For instance, alleles incorporated into a population after a shift in
555 environmental influence can increase from low to high frequency via positive selection. The
556 existence of such a beneficial allele can manifest in several ways: from new mutations,
557 introgression through gene flow, or molecular reorganization through novel recombination,
558 inversion, transposition, copy number variation, or insertion-deletion events. If there is strong
559 selection acting on this allele ($N_e s \gg 1$), it will sweep to high frequency creating a signature of
560 reduced polymorphism at neutral sites physically linked to the allele (‘genetic hitchhiking’,
561 Maynard Smith & Haigh 1974) resulting in a hard ‘selective sweep’ (Berry et al. 1991). However,
562 in structured populations with limited gene flow, this process can take significantly longer to
563 reach fixation, resulting in incomplete sweeps (Whitlock 2003). Additionally, Pavlidis et al.
564 (2012) found that, in congruence with Chevin & Hospital (2008), a multilocus genotype often
565 prevents the trajectories of individual alleles from sweeping to fixation, with an increasing
566 number of loci leading to decreasing probability of fixation, and as a result, an altered selective
567 signature at such loci (see also Jain & Stephan 2017). As such, hard selective sweeps in a

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568 polygenic architecture are expected to be rare (but not completely absent) under most
569 circumstances, particularly when the shift in environment causes a relatively small deviation
570 from the phenotypic optimum. Thus, hard sweeps most likely apply to loci with relatively large
571 effect above a calculated, context-dependent threshold value (Orr 2005; de Vladar & Barton
572 2014; Stephan 2015; see specifically Jain & Stephan 2015, 2017).

573 While early literature (Maynard Smith & Haigh 1974; Kaplan et al. 1989) focused on the
574 rapid sweep of an allele incorporated into a population after an environmental shift, research
575 within the last few decades have focused on ‘soft sweeps’ resulting from neutral or deleterious
576 mutations that are present in the standing genetic variation prior to the change in the selective
577 environment, wherein the selection coefficient changes with the environmental shift such that
578 the allele(s) become evolutionarily advantageous (reviewed in Hermisson & Pennings 2005,
579 Barret & Schluter 2008, Messer & Petrov 2013, and Hermisson & Pennings 2017; see also
580 Jensen 2014). These allele(s) could manifest via a single low-frequency variant, multiple
581 variants caused by parallel recurrent mutation/reorganization on multiple haplotypes, or multiple
582 unique alleles that arise independently within, perhaps multiple, populations. In such cases
583 where selection acts via soft sweeps, the rate of evolution at the phenotypic level is expected to
584 exceed those of hard sweeps because the alleles under selection have escaped the stochastic
585 nature of drift to a greater degree and are segregating within multiple individuals and genetic
586 backgrounds within the population. The extent to which soft sweeps alter the effect size
587 distributions underlying the genetic architecture is likely dependent upon both the strength of
588 selection and effect size before and after the environmental change (Messer & Petrov 2013;
589 Matuszewski et al. 2015; Jain & Stephan 2017), while the frequency before selection influences
590 the likelihood of subsequent detection (Innan & Kim 2004). Additionally, if multiple mutations are
591 segregating during the sweep, the probability of fixation for any given locus also decreases
592 (Pennings & Hermisson 2006a, 2006b; Chevin & Hospital 2008; Ralph & Coop 2010). Evidence
593 for hard sweeps in tree species exist within the literature, although they are rare (e.g., disease

594 response genes in *Pinus taeda* Ersoz et al. 2010; see also Table 2 in Siol et al. 2010). However,
595 for many species of trees, which often experience high gene flow and strong diversifying
596 selection across populations, adaptive divergence for polygenic traits is expected to result more
597 often from soft sweeps than hard sweeps, affecting phenotypes by subtle allele frequency
598 changes across populations, such that allele frequency differences of individual loci across
599 populations for neutral and selective sites will often be nearly indistinguishable (Latta 1998,
600 2003; Barton 1999; Le Corre & Kremer 2012; Stephan 2015; Yeaman 2015; Jain & Stephan
601 2015, 2017). Indeed, the large effective population sizes found in most tree species would
602 permit large effective mutation rates (or reorganization events) necessary for a soft selective
603 sweep from multiple unique variants, particularly when the phenotype is underlain by a large
604 mutational target. Even so, and as highlighted by Stephan (2015) and Bailey & Bataillon (2016),
605 the extent to which scientists can detect the influence of demographic processes on soft versus
606 hard sweeps, and vice versa, remains challenging (Jensen et al. 2005; Chevin & Hospital 2008;
607 Schrider et al. 2015, 2016; Schrider & Kern 2016; Hermisson & Pennings 2017).

608 While discrete directional selection events are likely to be a common evolutionary
609 influence across taxa, fluctuating or sustained directional selection (i.e., moving optima) are also
610 likely to be contributory factors influencing the genetic architecture of quantitative traits
611 (reviewed in Kopp & Matuszewski 2013; see also McCandlish & Stoltzfus 2014). For a
612 sustained moving optimum, the effect size distribution of beneficial alleles is expected to be
613 dependent upon the effect distribution of standing or *de novo* mutations as well as the strength
614 of selection: if the rate of change is dramatic, adaptation from new mutations is expected to
615 occur through intermediate to large-effect loci (Kopp & Hermisson 2009a; Matuszewski et al.
616 2014) or from small-effect loci when adaptation occurs via standing variation (particularly when
617 epistasis is considered, Matuszewski et al. 2015). Under lesser rates of environmental change,
618 adaptation is expected to proceed through mainly alleles of small-effect (Collins et al. 2007;
619 Kopp & Hermisson 2009a, 2009b) where intermediate effects will dominate the long-term

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620 distribution of effect sizes (Kopp & Hermisson 2009b). In the case of fluctuating environments,
621 outcomes often depend directly on the degree of temporal autocorrelation of the changing
622 environment. In such cases of stochastic fluctuation around a linear trend of environmental
623 change, extinction risk increases relative to that of the strictly linear trend (Bürger & Lynch 1995)
624 where local adaptation lags, to some degree, behind any given contemporaneous scenario. In
625 comparison, and similar in some ways, stochastic fluctuations around a constant mean are
626 expected to resemble the dramatic environmental change scenario described above,
627 characterized by strong selection pressures, maladaptation between generations, and a large
628 lag load (Bürger 1999; Chevin 2012; Kopp & Matuszewski 2013). In the case of autocorrelated
629 shifts, the ‘predictability’ of such fluctuations may decrease the possibility of extinction, increase
630 probability of local adaptation, and lead to similar scenarios as discussed for gradual changes in
631 the environment (Kopp & Matuszewski 2013).

632 *Gene flow*

633 Gene flow, to the extent that would be appreciable to that found in trees (reviewed in
634 Savolainen et al. 2007), is also an important component shaping quantitative expectations.
635 Indeed, since the early 1900s we have known that gene flow can disrupt adaptation if selection
636 is not strong enough to overcome the loss of beneficial alleles (Haldane 1930; Wright 1931; Slatkin
637 1987; reviewed in Felsenstein 1976, Lenormand 2002, Savolainen et al. 2007, 2013, Feder et
638 al. 2012a, and Tigano & Friesen 2016). Particularly when gene flow is asymmetric between core
639 and peripheral populations, adaptation can be inhibited in marginal habitats (Kirkpatrick &
640 Barton 1997; Kawecki 2008). Even so, there is abundant evidence that gene flow can promote
641 adaptation and maintain polymorphisms within populations, including white sand lizards
642 (Laurent et al. 2016), stick insects (Comeault et al. 2014, 2015), cichlid fishes (Meier et al.
643 2017), Darwin’s finches (Lamichhaney et al. 2015), and lodgepole pine (Yeaman & Jarvis 2006).

644 The magnitude of gene flow between populations can also impact the distribution of

645 effect sizes, for when gene flow falls below a critical threshold, and over many thousands of
646 generations, there is an increase in the probability of establishment and persistence times of
647 large-effect alleles, thus reducing the proportion of the polymorphism due to small-effect loci
648 (Yeaman and Otto 2011; Yeaman and Whitlock 2011). These dynamics are further influenced
649 by the susceptibility of alleles to 'swamping' (Slatkin 1975; Bürger & Akerman 2011; Lenormand
650 2002; Yeaman 2015; *sensu* Haldane 1930). For alleles that are prone to swamping, adaptive
651 phenotypic divergence depends on genetic variation and is driven by allelic covariance among
652 populations particularly when the underlying architecture is highly polygenic, the mutation rate is
653 high, and the number of loci underlying the trait exceeds the number needed to achieve the
654 local optimum phenotype (genetic redundancy; Yeaman 2015). Conversely, when there is little
655 genetic redundancy underlying the trait, limited divergence is observed unless the effect size of
656 a given swamping-prone allele exceeds the critical migration threshold. In these cases where
657 swamping-prone alleles contribute to adaptive divergence, the genetic architecture is transient
658 and any given locus contributes ephemerally to phenotypic divergence, even for loci of relatively
659 large effect (Yeaman 2015). In the case of swamping-resistant alleles, the evolved architecture
660 is enriched for large-effect loci and adaptive divergence can be maintained with little genetic
661 variation or input from mutation. Yet while the contribution from such loci can last many
662 thousands of generations, the architecture can again become transient as the genetic
663 redundancy or mutation rate increases (Yeaman and Whitlock 2011; Yeaman 2015).

664 Physical linkage and reduction of recombination between adaptive loci can also play a
665 considerable role in adaptive processes in the face of gene flow (Feder & Nosil 2010; Feder et
666 al. 2012a,b; Yeaman 2013; references therein). In such cases, loci that are tightly linked to other
667 loci already under selection will have an increased probability of contributing to local adaptation,
668 both because of physical linkage as well as by reducing the effective recombination among loci
669 within the sequence block. For instance, Yeaman & Whitlock (2011) showed that under
670 divergent selection with gene flow, the number of contributing loci decreases with increasing

671 recombination while small effect loci tend to cluster in groups that act as a single large effect
672 locus (see also Remington 2015), and strong selection can maintain these clusters of linked loci
673 over greater map distances than can weak selection. More recently, Yeaman (2013) employed
674 individual-based simulations to provide evidence that the clustering of alleles throughout a bout
675 of adaptation is unlikely to be driven mainly by divergence hitchhiking alone, and that instead
676 competition between genetic architectures and chromosomal rearrangements occurring
677 throughout adaptive processes under a range of environmental fluctuation scenarios can lead to
678 the evolution of tightly clustered adaptive loci which persist in the event of gene flow, unlike the
679 clusters identified by Yeaman & Whitlock (2011). Yeaman (2013) found that the level of
680 clustering was a function of the temporal fluctuation period, the rate of rearrangement itself is an
681 important determinant on the evolution of clustered architectures, and clusters can in some
682 cases be evolutionarily disadvantageous. Together, these results suggest that genomic
683 rearrangements (reviewed in Ortiz-Barrientos et al. 2016), including inversions (Kirkpatrick &
684 Barton 2006; reviewed in Hoffman & Rieseberg 2008), which decrease the effective rates of
685 gene flow among adaptive sequences can be an essential component of local adaptation, and
686 indeed some cases of speciation, in the face of gene flow.

687 ***Summary***

688 While we provided an overview of the factors that can influence the genetic architecture
689 of local adaptation, we acknowledge that it is far from exhaustive. Because the phenotypes
690 used in studies of local adaptation (particularly those assumed or corroborated to be a compo-
691 nent of total lifetime fitness) often have a continuous distribution, and are thus quantitative in
692 nature, the underlying genetic basis for these traits is likely polygenic and is predicted to be
693 underlain by multiple (often many) segregating loci, many of which may confer small phenotypic
694 effects (and are thus unlikely to be detected using single-locus approaches). Even so, a contin-
695 uum exists, where the true genetic architecture (the number of contributing loci, as well as their

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696 relative locations within the genome, phenotypic effects, and interactions) underlying a given
697 complex trait is itself determined by a combination of evolutionary forces that encompass an
698 interplay between the strength, timing, and direction of (background) selection against the
699 homogenizing effects of gene flow and recombination, disruptive effects of drift, linkage, trans-
700 position, inversion, and mutation, interactions between underlying loci as well as between these
701 loci and the environment, structural variation, relationship to gene expression networks, as well
702 as other factors related to life history. Consequently, the contemporary genetic architecture is a
703 result of past evolutionary processes, while the adaptive response to future evolutionary
704 dynamics is influenced in part by the contemporary architecture and genetic variance at hand.

705 **The genomics of local adaptation in trees**

706 *Common approaches used to identify adaptive loci*

707 Across taxa, and specifically in trees, the predominant association and outlier methods
708 for uncovering sets of loci underlying local adaptation have relied upon single-locus population
709 genetic approaches. Putatively adaptive loci are often identified by elevated allele frequency
710 differences among populations relative to patterns genome-wide. Yet, as revealed in the
711 previous section, loci underlying polygenic traits will often be indistinguishable from non-
712 causative sites in this way. Further, outlier tests based on F_{ST} (*sensu* Lewontin & Krakaur 1973)
713 do not incorporate information regarding putative phenotypic targets of selection nor
714 environmental drivers of differentiation, often do not correct for neutral population structure (but
715 see Lotterhos & Whitlock 2015), and will inevitably isolate a biased set of candidate loci
716 (Hermisson 2009; Cruickshank & Hahn 2014). In the case of single-locus genotype-environment
717 associations (reviewed in Rellstab et al. 2015; see also De Mita et al. 2013), information about
718 possible environmental drivers is incorporated by assessing the association between allele
719 frequencies and environmental heterogeneity, yet without information regarding traits
720 hypothesized to be influenced by selection (Schoville et al. 2012). Single-locus genome wide

721 association studies (see next section; Supplemental Box S2) and quantitative trait loci (QTL)
722 experiments (reviewed in Ritland et al. 2011, Hall et al. 2016) have also been used in trees,
723 quantifying the differential effects of typed alleles on a given phenotype. Despite the
724 shortcomings of these methods, such studies provide candidate loci that can be investigated in
725 further detail (Tiffin & Ross-Ibarra 2014), which is particularly advantageous when resources are
726 limited. Indeed, as discussed below, these approaches dominate the methods used to uncover
727 complex traits (adaptive or otherwise) in trees.

728 *Current progress in trees*

729 In light of the expectations outlined above for the architecture of quantitative traits under various
730 evolutionary regimes, and the methods commonly used to detect these loci, we reviewed the lit-
731 erature of single-locus genotype-phenotype associations (GPAs, which included associations to
732 gene expression levels) from studies in forest trees. In doing so, we identified 52 articles across
733 10 genera and 24 species with a total of 2113 GPAs (Supplemental Table S2, Supplemental
734 File F2). Because most studies in trees do not report phenotypic effect sizes of individual loci
735 (i.e., regression coefficients), we report r^2 values which can be used to quantify the percent
736 phenotypic variance explained by the associated locus. In cases where multiple SNPs from a
737 given locus (e.g., a gene or scaffold) were associated to a trait, we averaged the r^2 values for
738 that locus. As with our review of trait heritability and Q_{ST} , we grouped phenotypic traits used in
739 associations into twelve broad categories (in this case, no phenotypes fell into Survival or Seed
740 and Seedling Properties groups). If traits important to tree conservation and industry are often of
741 a polygenic basis, we would expect small to moderate effects from loci empirically associated to
742 phenotype. Indeed, across the trait groups considered here, the mean r^2 was 0.039, where
743 80.79% ($n = 1707$) of recorded estimates had r^2 values less than 0.05, 18.78% ($n = 397$) of r^2
744 values falling between [0.05,0.22], and nine values of r^2 greater than 0.22, which were all
745 related to *Cronartium ribicola* resistance in *Pinus monticola* Douglas ex. D. Don (Figure 3).

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746 Of the twelve trait groups, all but those traits relating to both reproduction and herbivore
747 and insect resistance had r^2 estimates greater than 0.10, with traits relating to disease
748 resistance, growth, leaf and needle properties, phenology, and wood properties each
749 contributing over 10% of these outliers. These small effects tend to also not account for much of
750 the observed heritability, but can explain sizeable fractions in some instances (e.g., primary
751 metabolites in Eckert et al. 2012). Of the loci associated with expression levels, r^2 estimates
752 were between 0.05 and 0.152 in all but one case ($n = 54$). We also assessed the propensity of
753 individual loci to be associated to more than one phenotype or expression level across our
754 literature review. Without correcting for the multiple associations of a locus to yearly phenotypes
755 (e.g., bud flush 2009, bud flush 2010), we found that the average number of loci associated to
756 multiple phenotypes per study was 6.94, while after correcting for multiple years the average
757 number decreased to 5.59. The median number of SNPs utilized for association per study was
758 206, where 75% (39/52) of studies used less than 1,000 SNPs, eight studies using between
759 1,000-10,000 SNPs, four studies using between 29,000-35,000 SNPs, and one study utilizing
760 2,822,609 SNPs for association (all studies with greater than 10,000 SNPs were from either
761 *Pinus* or *Populus* species).

762 *Are we out of the woods yet?*

763 From insight gained from the literature review of genotype-phenotype associations it seems that
764 the vast majority of the genetic architecture of local adaptation and complex traits in trees remains
765 largely unexplained using common GWAS methods (see also Box 1), a consistent pattern across
766 the past decade of research in trees (Neale & Savolainen 2004; Savolainen et al. 2007; Čalić et
767 al. 2015; Hall et al. 2017). Furthermore, it is likely that the estimates for percent variance
768 explained are inflated due to a combination of QTLs that break down into smaller effect loci
769 (Remington 2015), the Beavis effect (Beavis 1994; Xu 2003), and the Winner's Curse (Görning
770 et al. 2001; Zöllner & Pritchard 2007) where locus effects are inflated by using the same data for

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771 both gene identification and phenotypic prediction (see Box 1 in Josephs et al. 2017b for a
772 detailed synopsis of these biases). Such a pattern suggests that, indeed, many of the traits im-
773 portant to evolutionary, breeding, and conservation insight in trees are likely of a polygenic basis
774 and that future studies must take this into account when seeking to identify the underlying loci.

775 Even within studies of model organisms, missing heritability is nothing new. Across taxa,
776 missing heritability is less frequent within phenotypes of mono- to oligogenic bases (as seen for
777 the *Cr2* major-gene resistant locus in *Pinus monticola*, Liu et al. 2017), as would be expected,
778 and is a recurrent, pervasive shortcoming from genotype-phenotype associations of complex
779 traits, particularly those maintaining single-locus perspectives. A number of explanations have
780 been put forth to explain the missing heritability, such as epistasis (Hemani et al. 2013) and its
781 inflationary effect on heritability estimates (Zuk et al. 2012), environmental or epigenetic inter-
782 actions (Feldman & Lewontin 1975) as well as their inflationary effect on heritability estimates
783 (Zuk et al. 2012), (unmeasured) low-frequency variants of large effect (Dickson et al. 2010),
784 genetic or variance heterogeneity of individual alleles (Leiserson et al. 2013; *cf.* Box 1 in Nelson
785 et al. 2013), or common variants with effect size below detection thresholds (Yang et al. 2010).
786 As such, here we avoid supporting one causative hypothesis over another, particularly given the
787 ongoing discussion within the literature, for which strengths and weakness for any viewpoint are
788 apparent (e.g., Gibson et al. 2010), and because of the progress yet to be made in trees.

789 Indeed, the dissection of the genetic architectures underlying complex traits in trees is
790 still in its nascency compared to the progress of model organisms (for which missing heritability
791 is still an issue), and beyond issues of coverage, genomic saturation, and genomic resources
792 (discussed below in The Path Forward), we must approach this issue with all possibilities in
793 mind. Given the unique properties of the life histories, genome size and organization of many
794 tree species, and the limited numbers of studies with large sets of molecular markers, causative
795 sources of the missing heritability should be ruled out, or supported, as with any other
796 hypothesis, particularly as we gain information from contemporary studies of trees that address

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797 shortcomings of those in the past. Further, we must keep in mind differences between functional
798 and statistical gene action (Álvarez-Castro *et al.* 2007; Nelson *et al.* 2013; Huang & Mackay
799 2016; Huber *et al.* 2017). In any case, it seems that sample sizes of single-locus approaches
800 will need to be increased (Hall *et al.* 2016), albeit with diminishing returns (Boyle *et al.* 2017;
801 Simons *et al.* 2017), to discover a higher proportion of the underlying loci in trees due to small to
802 moderate additive effects. Alongside suggestions outlined in *The Path Forward*, incorporating
803 investigations into such aspects of epistasis, dominance, pleiotropy, GxE effects, and network
804 analyses (when appropriate), may be a worthwhile complement (e.g., Lotterhos *et al.* 2017,
805 Mähler *et al.* 2017, Mizrahi *et al.* 2017; Tan *et al.* 2017).

806 While the infinitesimal model will continue to prove to be immensely useful for breeding
807 programs and for short-term evolutionary predictions, and we may find that the missing
808 heritability in trees is truly due to consequences of the infinitesimal regime (as is often cited to
809 be the majority consensus across taxa for missing heritability), it has been argued that the
810 analysis paradigm for such studies is near its limits in describing the functional genetic
811 architecture of quantitative traits, and that it is therefore necessary to move beyond single-locus
812 perspectives and reconsider common practices (Pritchard & Di Rienzo 2010; Nelson *et al.* 2013;
813 Sork *et al.* 2013; Tiffin & Ross-Ibarra 2014; Wadgyamar *et al.* 2017). At this stage, it seems that
814 we investigators seeking to describe the genetic architecture of quantitative traits in trees have
815 some ways yet to go before we are truly out of the woods. In the next section, we describe the
816 path forward to describing genetic architectures from a polygenic and functional perspective,
817 identify resources available to advance our knowledge and fill knowledge gaps, as well as future
818 directions for this research area.

819 **The Path Forward**

820 As we have outlined, there is still ample room for improvement in our description and
821 understanding of the genetic architecture of quantitative traits in trees (see Table 1 and Box 1).

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822 Importantly, methods used to uncover causative loci should take into consideration the expected
823 degree of polygenicity, the relative contributions of various forms of gene action, as well as how
824 past evolutionary phenomena has likely shaped current adaptive expectations. In this section,
825 we orient our path forward by first highlighting utilities available to, and underused within, the
826 forest genetics community to describe the genetic architecture of complex traits. We then outline
827 several suggestions to facilitate further progress and advocate for prospective perspectives in
828 future studies such that information and data may continue to be used easily in subsequent
829 syntheses across pathways, environments, species, and towards insight to identify future
830 needed resources as our understanding progresses. While our recommendations are specific to
831 the tree community, we also acknowledge other valuable recommendations from recent reviews
832 (e.g., Savolainen et al. 2013; Tiffin & Ross-Ibarra 2014; Lotterhos & Whitlock 2015; Gagnaire &
833 Gaggiotti 2016; Hoban et al. 2016; Wellenreuther et al. 2016; Burghardt et al. 2017; Wadgymar
834 et al. 2017).

835 *Stepping off the path – what’s in our pack?*

836 The genetic architecture underlying local adaptation and complex traits likely has a
837 polygenic basis composed of many loci of relatively weak effect yet many of the common
838 association or outlier methods will often fail to detect many of the causative loci of small to
839 moderate influence. Such investigations have so far led to an incomplete description of studied
840 architectures, and, in many cases, have limited our understanding of complex traits in trees to a
841 handful of loci. While we do not advocate that such single-locus methods be avoided in future
842 studies (considered further in the next section), here we outline underused and promising
843 approaches to identify and describe underlying loci that explicitly take into account the polygenic
844 basis of such traits and may help advance our understanding in future studies, including some
845 of the questions we have outlined in Table 1. Multivariate, multiple regression, and machine
846 learning techniques are three such examples, and differ from univariate analyses by analyzing

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847 patterns among multiple loci simultaneously.

848 The Bayesian sparse linear mixed model (BSLMM), for instance, such as that deployed
849 in the software package `GEMMA` (Zhou et al. 2013), is developed for both genomic prediction (see
850 also Box 1) and mapping of complex traits that offers considerable advantages over single-locus
851 genotype-phenotype approaches (Guan & Stephens 2011; Ehret et al. 2012; Zhou et al. 2013;
852 Moser et al. 2015). This analysis has gained in popularity recently, being used across diverse
853 taxa such as stick insects (Comeault et al. 2015, Riesch et al. 2017), butterflies (Gompert et al.
854 2015), Darwin's finches (Chaves et al. 2016), and trees (Lind et al. 2017). BSLMM is a hybrid of
855 LMM and Bayesian variable regression that extends the Lande & Arnold (1983) multiple
856 regression approach in an attempt to address the sparsity of common data sets used in
857 genotype associations, where the number of model parameters (loci) is often much greater than
858 the number of observations (sampled individuals; Zhou et al. 2013; Gompert et al. 2016).
859 Specifically, the model takes into account relatedness among individuals and provides a means
860 to summarize estimates of selection across the genome such as the proportion of phenotypic
861 variation explained (PVE) across genotyped markers by estimating the combined influence of
862 markers with either polygenic (infinitesimal) or measurable (moderate to large) effect, the
863 proportion of PVE explained by genetic loci with measurable effects (PGE), and the number of
864 loci with measurable effects that underlie the trait (for more details see Guan & Stephens 2011;
865 Zhou et al. 2014; Gompert et al. 2016). Additionally, `GEMMA` returns the posterior inclusion
866 probability for each marker providing evidence for association with the phenotype. While the
867 approach remains promising considering its performance in the context of genomic prediction
868 and inference of PVE (e.g., Zhou et al. 2013, Speed & Balding 2014), there has been no
869 attempts, to our knowledge, to assess the approach under various demographic histories,
870 genetic architectures, and sampling designs. A close approximation to this comes from analyses
871 carried out by Gompert et al. (2016), in which `GEMMA` was evaluated for PVE estimation,

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872 estimated effects of causative loci, and the estimated number of underlying SNPs based on
873 various author-specified numbers of causal loci, underlying heritability ranges, and numbers of
874 sampled individuals. In short, the authors convey that GEMMA is promising, but that there are
875 important limitations to consider (Gompert et al. 2016). However, because the authors simulated
876 architectures by randomly assigning effects to loci from an empirically-derived sequence data
877 set, and while they were thorough in their data exploration, we encourage these results be
878 replicated *in silico* through full modeling of genomic loci across various demographic, LD,
879 sampling, and architecture scenarios to ensure underlying allele frequencies among populations
880 and LD (within and among populations) reflect realistic patterns which may have an effect on
881 model performance. Such additional analyses will also allow for more specific insight into model
882 performance based on *a priori* biological insight available to investigators, allowing more
883 informed decisions when choosing an appropriate genotype-phenotype association method
884 such as BSLMM.

885 Random Forests (Breiman et al. 2001) is a machine learning algorithm used to identify
886 patterns in highly dimensional data sets to further generate predictive models. Alongside uses
887 outside of evolutionary biology, the Random Forests algorithm has gained popularity in
888 association studies across taxa as well as in trees such as that of genotype-phenotype
889 associations in Sitka spruce (*Picea sitchensis*; Holliday et al. 2012) and genotype-environment
890 associations in white spruce (*P. glauca*; Hornoy et al. 2015). Random Forests is based upon
891 classification (for discrete variables, e.g., soil type) and regression (continuous variables; e.g.,
892 temperature or phenotypic measurements) trees (so-called CART models). During its
893 implementation, Random Forests creates these decision trees using two layers of stochasticity:
894 the first layer is used to grow each tree by using a bootstrap sample of observations
895 (environmental or phenotypic) while the second uses a random subset of predictors (marker
896 loci) to create a node which is then split based on the best split of the observations across
897 permutations of predictors using the residual mean square error (see Figure 2 in Hornoy et al.

898 2015). The observations that were not used as training data to create the model are then used
899 to estimate model accuracy, which can be further used to assess variable importance (Holliday
900 et al. 2012; Hornoy et al. 2015; Forester et al. 2017).

901 While creating a promising alternative to univariate approaches, until recently the
902 Random Forests algorithm has not been fully explored to assess model performance for use in
903 association studies. Forester et al. (2017) provide a thorough analytical assessment using
904 simulated data to remark on performance for use in genotype-environment association studies
905 (GEA). In their analysis, they used published simulations of multilocus selection (Lotterhos &
906 Whitlock 2014, 2015) of various demographic histories and selection intensities across 100
907 causative (with 9900 neutral) loci to compare the Random Forests algorithm to the multivariate
908 approaches of constrained ordination (redundancy analysis, RDA, and distance-based RDA,
909 dbRDA - both of which are mechanistically described in Legendre & Legendre 2012, but are
910 multivariate analogs of multiple regression on raw or distance-based data) and to the univariate
911 latent factor mixed model (LFMM). In short, Forester et al. (2017) found that LFMM performed
912 better than Random Forests as a GEA, while constrained ordinations resulted in relatively lower
913 false positive and higher true positive rates across levels of selection than both Random Forests
914 and LFMM. Additionally, the authors found that correction for population structure had little
915 influence on true and false positive rates of ordination methods, but considerably reduced true
916 positive rates of Random Forests. They also note that further testing is needed across various
917 evolutionary scenarios. Even so, constrained ordination provides an effective means by which to
918 detect loci under a range of both strong and weak selection (Forester et al. 2017). While
919 promising under a GEA framework, future analyses may provide evidence that such methods
920 also perform well in genotype-phenotype associations as well. Empirically, it has been used in
921 trees to explore multivariate relationships between phenotypes, genotypes, and environments
922 (e.g., Sork et al. 2016). Additionally, there have been many extensions of the original Random
923 Forests model, such that extensions with purportedly better performance should be assessed

924 alongside other popular association methods in the future.

925 Once a set of candidate loci have been identified to putatively underlie a phenotype or
926 environment of interest, these loci can be used to further test the hypothesis of polygenic local
927 adaptation. For instance, Berg & Coop (2014) use the significant hits from GWAS data sets to
928 estimate within-population additive genetic values by calculating the frequency-weighted sum of
929 effects across these loci. These values are then compared to a null model of genetic drift that
930 accounts for population structure to test for an excess of variance among populations, ultimately
931 identifying the populations most strongly contributing to this signal. The excess variance statistic
932 (Q_x) is analogous to Q_{ST} and is composed of two quantities – an F_{ST} -like component describing
933 allele frequency differentiation across populations and a LD-like component describing
934 coordinated and subtle allele frequency shifts across populations. This method thus allows
935 explicit hypothesis tests related to the expected polygenic architecture of local adaptation
936 across populations of trees. It is also noteworthy in that it combines aspects of the genotype-
937 environment-phenotypic spectrum that underlies local adaptation within a single methodological
938 framework (cf. Sork et al. 2013). Prior attempts take a pairwise approach examining each
939 pairwise combination of the genotype-environment-phenotype spectrum (e.g., Eckert et al.
940 2015). Despite the promising insight from this method, it has not been used widely outside of
941 model organisms. Future applications in trees should consider the number of causal loci
942 identified to be associated with quantitative phenotypes (driven somewhat by the number of loci
943 used in mapping studies), the number of populations needed to increase power, especially in
944 the correlation of genetic values to environmental data, and the ability to reliably estimate
945 genotypic effects.

946 ***At the trail junction – where to next?***

947 While we have outlined methods above that have not yet realized their full potential in
948 describing genetic architecture of complex traits in trees, there are several matters that we, as a

949 field, must keep in mind such that we can continue to progress our understanding in the most
950 efficient manner. Here we believe the path forward lies in three critical areas which we discuss
951 in further detail below: 1) needed data, 2) standardized data reporting, and 3) empirical studies
952 in trees designed to test theoretical expectations of genetic architectures.

953 *Needed data*

954 While the common garden approach can facilitate understanding of evolutionary
955 processes without specifically identifying underlying loci (Rausher & Delph 2015), identifying
956 features of the genetic architecture will ultimately inform breeding applications important to
957 management, conservation, and industry, and thus requires knowledge about underlying loci.
958 Consequently, we have not yet had sufficient sampling of both marker densities and studies
959 amenable to replication across systems to truly exhaust the use of single-locus approaches,
960 particularly as the sample size of markers, individuals, and populations increase in the near
961 future. Indeed, Hall et al. (2016) estimated that the number of causative loci underlying
962 quantitative traits in trees is likely in the several hundreds, and to capture 50% of the heritable
963 genetic variation using single-locus approaches, population sizes of about 200 will be needed
964 for mapping disease traits, and about 25,000 for traits such as growth. Even so, we recommend
965 that such single-locus associations should not be used as the sole method of architecture
966 description as we carry out future studies unless justified *a priori* based on biological principles,
967 knowledge of the expected architecture, and/or for testing specific hypotheses. While the limits
968 of such methods should be considered, these approaches can be used alongside other lines of
969 evidence to either support or spur further testing of underlying loci (*sensu* Sork et al. 2013). For
970 instance, there is little downside to performing both a single-locus association and a multivariate
971 analysis in the same study, even if some or all of the results for a given technique are excluded
972 to the supplement (e.g., Sork et al. 2016). Further, contextualizing genotype-phenotype and
973 genotype-environment relationships with results that describe local adaptation (e.g., phenotype-

974 environment, Q_{ST} - F_{ST} comparisons) can also stimulate further understanding particularly for data
975 that is made publically available for future synthesis. Specifically, studies which do so within the
976 context of comparisons within and across species (e.g., Yeaman et al. 2016) or environments
977 (Holliday et al. 2016), offer unique circumstances under which to advance our understanding of
978 complex traits in trees (Table 1; Lotterhos & Whitlock 2015; Čalić et al. 2016; Hoban et al. 2016;
979 Ingvarsson et al. 2016; Mahler et al. 2017).

980 Isozymes (Adams & Joly 1980), restriction fragment length polymorphisms (Devey et al.
981 1994), randomly amplified DNA (Grattapaglia & Sederoff 1994), and expressed sequence tag
982 polymorphisms (Temesgen et al. 2001) were among the first used to test evolutionary
983 hypotheses in trees related to genome organization and the mapping of complex traits
984 (discussed in Eckert et al. 2009). Marker technology has progressed considerably since this
985 time (dozens of markers) to include markers capable of more densely sampling tree genomes
986 (up to millions of markers). For example, array-based designs (Silva-Junior et al. 2015) and
987 exome capture (Suren et al. 2016) allow for hundreds to tens of thousands of markers (which
988 can be dwarfed by the number of subsequently called SNPs) whereas RADseq (reviewed in
989 Parchman et al. in review) is in the range of tens- to hundreds of thousands of markers (e.g.,
990 Parchman et al. 2012) and whole genome sequencing in the range of millions (e.g., Stölting et
991 al. 2015). However, while the continual advent of sequencing technology will likely allow for
992 more SNPs and longer sequences, it is ultimately the concordance between polygenic
993 expectations and analytical methods of marker data that will determine the success of such
994 endeavors. With this in mind, future studies aimed at answering outstanding questions (Table 1)
995 will benefit from a diverse set of markers that represent both functional proteins (genic regions)
996 as well as those which control aspects of their expression or post-transcriptional regulation. If
997 one lesson is to be gained from the recent discussion of the applicability of reduced
998 representation techniques (Lowry et al. 2016, 2017; Catchen et al. 2017; McKinney et al. 2017),
999 it is that genomic resources are paramount to advancement of knowledge, especially when

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1000 developed with knowledge of patterns of linkage disequilibrium or, if not with this knowledge,
1001 with goal of quantifying it. However, RADseq remains one of the most cost-effective approaches
1002 available to trees and should thus be assessed in the specific context of tree species,
1003 particularly when strengths and limitations are understood and addressed (as reviewed in
1004 Parchman et al. in review). No matter the approach used for association, some aspect of the
1005 architecture is likely to be missed in trees. For example, RADseq-based markers developed
1006 within large genomes are not enriched within genic regions where structural changes to proteins
1007 are expected to affect phenotypes, although choice of enzymes can affect the relative
1008 proportion of genic regions in tree genomes, as evidenced from *in silico* digestions of reference
1009 genomes from *Populus*, *Eucalyptus*, *Amborella*, *Pseudotsuga*, and *Pinus* species (Parchman et
1010 al. in review). In contrast, exome based approaches are anchored within coding regions thus
1011 excluding putative regulatory elements outside of the exomic regions used to develop probes.
1012 Recent marker development approaches, such as RAPTURE (Ali et al. 2016), however, have
1013 blurred the lines between RADseq and exome based approaches and may offer a promising,
1014 cost-effective path forward that explicitly avoids biased assumptions about the importance of
1015 exomic versus intergenic loci comprising the architecture of local adaptation.

1016 Beyond dense genetic linkage maps (e.g., Friedline et al. 2015) and reference genomes,
1017 which undoubtedly should be among our top priorities, other techniques outside of traditional
1018 genomics, such as transcriptomics, have the potential to complement genomic studies in many
1019 ways without great need for existing species-specific resources (reviewed in Romero et al.
1020 2012, Strickler et al. 2012; Vialette-Guiraud et al. 2016). For instance, comparative
1021 transcriptomic techniques in trees can be used to identify putatively orthologous sets of markers
1022 (e.g., Wachowiak et al. 2015; Yeaman et al. 2016) that can be used to describe the evolution of
1023 architecture (e.g., shared orthologs versus paralogs across species) or for comparative linkage
1024 mapping (Ritland et al. 2011) across systems. Additionally, with the appropriate study design,
1025 transcriptomics can be implemented in tree species to describe various aspects of differential

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1026 expression (Cohen et al. 2010; Carrasco et al. 2017; Cronn et al. 2017), selective constraint
1027 (Mähler et al. 2017), prevailing selective forces (Hodgins et al. 2016), mapping of disease
1028 resistance (Liu et al. 2016; Liu et al. 2017), and regulatory networks (Zinkgraf et al. 2017). The
1029 multilocus paradigm of transcriptomics is amenable to identifying and testing hypotheses of the
1030 genetic architecture of complex traits in a network framework (Jansen et al. 2009; Leiserson et
1031 al. 2013; Civelek & Lusk 2014) and will no doubt provide valuable contributions for tree evolu-
1032 tionary biologists. Other areas amenable to network description such as metabolomics and prot-
1033 eomics would also be a complement (see Cowen et al. 2017), particularly if genetic studies con-
1034 textualize results with findings from such approaches and vice versa. Ultimately the goal is to
1035 use *a priori* knowledge synthesized across past studies, techniques, and perspectives to guide
1036 further hypotheses about underlying architecture, as exemplified by Mizrachi et al. (2017) and
1037 Lotterhos et al. (2017). Finally, high-throughput phenotyping as well as environmental measures
1038 at fine spatial scales below square-kilometers will also facilitate and advance our understanding
1039 of complex traits in trees (Sork et al. 2013; Rellstab et al. 2015; Leempoel et al. 2017).

1040 ***Standardized data reporting***

1041 As we continue to accrue genotype-phenotype, genotype-environment, and phenotype-
1042 environment relationships within and across tree species, authors should consider how their
1043 results can most effectively be used in further studies and syntheses, both for the purpose of
1044 validation or comparison as well as novel insights yet to be seen. Here we outline a few
1045 suggestions that can be broken down into reporting within manuscripts and metadata. For
1046 instance, in our survey of common garden studies used to estimate h^2 and Q_{ST} , in many cases
1047 the exact design of the study could not be replicated with the information from the manuscript
1048 alone. While an abbreviated design may be suitable for the main text, authors can provide much
1049 more detail in supplemental materials that can facilitate replication and comparison across
1050 studies (e.g., total individuals per garden, family, or block – as opposed to averages or ranges),

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1051 which will ultimately facilitate syntheses regarding future directions. Further, future studies
1052 would benefit from estimating relatedness using marker data which will ultimately improve the
1053 precision of h^2 , Q_{ST} , and missing heritability estimates (de Villemereuil et al. 2016) including
1054 those estimates made in the field (Castellanos et al. 2015). For cases in which estimating
1055 relatedness from markers is not appropriate or feasible, the field would benefit by authors
1056 exploring a range of underlying sibships (e.g., Eckert et al. 2015), which are often assumed to
1057 be half-sib relationships. While some studies in our survey assumed a mixed sibship
1058 relationship for open-pollinated sources, ultimately such assumptions without data exploration
1059 will affect the outcome or conclusions for any given study. A recently released R package by
1060 Gilbert and Whitlock (2014) allows for such an exploration of effects of mixed sibships on
1061 inference of Q_{ST} and its magnitude relative to F_{ST} . Inclusion of such exploration, even in the
1062 supplement, will help contextualize such studies as they are published. For studies estimating
1063 causality for genotype to phenotype, it would be worthwhile to include the regression
1064 coefficients or other estimates of effect size (e.g., odds ratios) in addition to PVE (r^2).
1065 Importantly, the units of the effect size must be explicitly reported (e.g., Julian days versus
1066 phenotypic standard deviations), with the standard deviation also reported. For all association
1067 studies, supplemental tab- or comma-delimited text files (outside of a word processing
1068 document) easily analyzed with programming languages would also facilitate synthesis (even if
1069 providing redundant information from the main text), particularly if such files are well described
1070 with a README file and contained data regarding marker position, putative orthogroups, hits to
1071 reference genomes, effect size, PVE, genotypes by individual identifiers, individual population
1072 assignments, and if the sequence or marker was significantly associated to phenotype or
1073 environment. Such an operating procedure may work well in the short term, however in the long
1074 term such information will need to be easily accessible from one or a central hub of repositories.

1075 Data standardization, the inclusion of meta-information, and compilation of these data
1076 specific to trees into a database with common terminology will be crucial to future inquiries with

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1077 the purpose of synthesizing evidence for underlying architectures across species and
1078 environmental systems (e.g., as for human GWAS data: <https://www.ebi.ac.uk/gwas/>). If the
1079 data generated by tree biologists is disparate and housed across databases and journal
1080 supplements this impedes synthesis first by forcing scientists to collate information across
1081 sources, which may be further impeded by data redundancies or inconsistencies in data format
1082 and utilized nomenclature (Wegrzyn et al. 2012). While many journals have required submission
1083 of sequence data to repositories such as NCBI, such databases are lacking with regard to
1084 information pertaining to phenotypic, environmental, and geographic information upon which
1085 much of the foundation of our field is built. Submissions to Dryad somewhat overcome this, but
1086 there is no standardization within the community for content for such submissions and important
1087 information may be lacking. Currently, this information is often appended in supplemental files
1088 that cannot be readily accessed, compared, or queried in an efficient manner. Hierarchical
1089 ontologies can be used to ease this burden. Gene Ontology is likely the most recognizable to
1090 evolutionary biologists, but there also exist Plant Ontologies for organismal structure and
1091 developmental stages, Environmental Ontologies for habitat categorization, and Phenotypic,
1092 Attribute, and Trait Ontologies for the annotation of phenotypes. Such ontologies not only
1093 standardize nomenclature, but also assist in database queries. The utilization of such databases
1094 will no doubt encourage comparative studies and syntheses, as infrastructure and data
1095 accessibility are essential to the comparative approach (Neale et al. 2013; Ingvarsson et al.
1096 2016; Plomion et al. 2016). Luckily, such a database exists for the broader tree genetics
1097 community. The open-source genomics and phenomics database, called TreeGenes
1098 (treegenesdb.org), is part of a central hub of repositories, including the Hardwood Genomics
1099 Project (hardwoodgenomics.org) and the Genome Database for Roseaceae (rosaceae.org), that
1100 communicate and integrate data from each other. Unlike many other repositories for tree
1101 genomic data, TreeGenes is not project or institution specific. The data and metadata for
1102 roughly 1700 species housed on TreeGenes can be accessed, queried, and visualized through

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1103 DiversiTree, a web-based, desktop-style interface (Wegrzyn et al. 2008). DiversiTree connects
1104 to the geographical interface CartograTree (Vasquez-Gross et al. 2013) to encourage
1105 comparative synthesis by providing technology to filter and visualize geo-referenced biotic and
1106 abiotic data housed on TreeGenes. As promising as such database hubs are, they are only as
1107 useful as the data that is deposited to them. While TreeGenes will regularly import and enhance
1108 data from public repositories (through e.g., sequence alignment to published genomes, or data
1109 from Genbank, Phytozome, PLAZA, etc), often pertinent metadata necessary for comparative
1110 synthesis is lacking (Wegrzn et al. 2008, 2012). Indeed, from our survey of published GPA since
1111 the release of the database in 2008, less than 13% (6/48) of the studies submitted their data
1112 directly to TreeGenes. To better prepare for future synthesis, we advocate that authors submit
1113 their data to the TreeGenes database and that reviewers and editors enforce this habit, as
1114 currently implemented for linkage maps published in *Tree Genetics & Genomes*. Consolidated,
1115 open-source resources will be crucial to the advancement of this field (Neale et al. 2013), and
1116 will no doubt spur knowledge that would not have been recognized otherwise. Prime examples
1117 of advancement to knowledge because of these types of resources and community-wide efforts
1118 come from the human GWAS literature where such resources provide crucial information
1119 necessary to study polygenic adaptation (e.g., Berg & Coop 2014).

1120 *Empirical tests of theory*

1121 In combination with the development of truly genome-wide public resources, there is
1122 need to use these resources to validate and better characterize foundational ideas and
1123 assumptions in the theory of polygenic adaptation relative to the life history strategies of tree
1124 species. For example, Gagnaire & Gaggiotti (2016) highlight that the degree of polygenicity can
1125 be tested as a function of the number of GWAS hits relative to the length of contigs or
1126 chromosomes containing these markers. Simple models of polygenicity predict that there should
1127 be a positive correlation between these quantities. Thus, rather than assuming some functional

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1128 form of a polygenic architecture (i.e., an approximate infinitesimal model) during analysis,
1129 researchers can strive to characterize, or at least exclude some forms of, the underlying genetic
1130 architecture prior to interpretation. In a related fashion, publically available data sets would spur
1131 comparisons across species and study systems to test hypotheses about polygenic
1132 architectures (e.g., the modularity of genetic architectures as in Lotterhos *et al.* 2017, or
1133 perhaps genomic organization or effect size distribution) due to the relative timing of selection,
1134 degree of environmental contrast (e.g., diversifying selection and changes to the strength of
1135 negative selection), selection strength, and level of gene flow across diverging lineages. As an
1136 example, much of the theory of polygenic adaptation requires assumptions about simplistic
1137 demographics (where violations have consequences for standing levels of non-neutral diversity,
1138 e.g., Wang *et al.* 2017) and the equilibration among co-acting evolutionary forces over a large
1139 number of generations (Brandvain & Wright 2016). Indeed, differing architectures are expected
1140 as a function of the timing for the onset of selection (Le Corre & Kremer 2003; Kremer & Le
1141 Corre 2012), with subtle allele frequency shifts across populations dominating architectures
1142 near the onset of selection and larger allele frequency shifts much later in time. While there is
1143 need for empirical validation of this theory, there is also a need to characterize the prevalence of
1144 its predicted patterns across differing clades of tree species. In other words, researchers could
1145 imagine testing the theory itself in natural populations (e.g., as begun by Le Corre & Kremer
1146 2012) or assuming its validity and characterizing the circumstances under which to expect large
1147 shifts in allele frequencies across tree species with differing life history strategies. Little of any of
1148 this (Table 1), however, will be possible without development of needed data and its deposition
1149 into publically available, standardized databases.

1150 **Concluding Remarks**

1151 The path forward provides a means by which we can most efficiently describe the
1152 underlying genetic architectures of traits important to management, conservation, and industry,

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1153 which can ultimately be used to expedite breeding projects (Box 1). The past evolutionary
1154 history will have a profound effect on the underlying genetic architecture of such traits, and thus
1155 strengths and weakness of the data and methods used to uncover such architecture should be
1156 specifically addressed in the future, particularly in how utilized methods perform across various
1157 demographic and architecture scenarios. Insights gained from empirically testing theory will also
1158 contribute to the advancement of this field and will ultimately quantify the variation in archi-
1159 tecture across environments and species and inform effective management. Importantly, the
1160 success of future genotype-phenotype efforts should not be predicated on past studies using
1161 single-locus approaches and small numbers of markers, and instead on overcoming such
1162 shortcomings by applying theoretical expectations to empirical inquiry. Even so, until
1163 sequencing technologies allow for cost-effective whole genome sequencing of individual trees,
1164 most genotype-phenotype studies (GS included) will be carried out via reduced representation
1165 techniques (i.e., a subset of all sites within the genome). Therefore, it is essential that
1166 processed data be uploaded to a repository that, in addition to raw sequences, includes
1167 genotypic, environmental, and spatial data, facilitates user-friendly queries, and allows for future
1168 meta-analysis. The future is bright, but we are not yet out of the woods. As such, efficient
1169 advancement in this field relies on community efforts, standardized reporting, centralized open-
1170 access databases, and continual input and review within the community's research.

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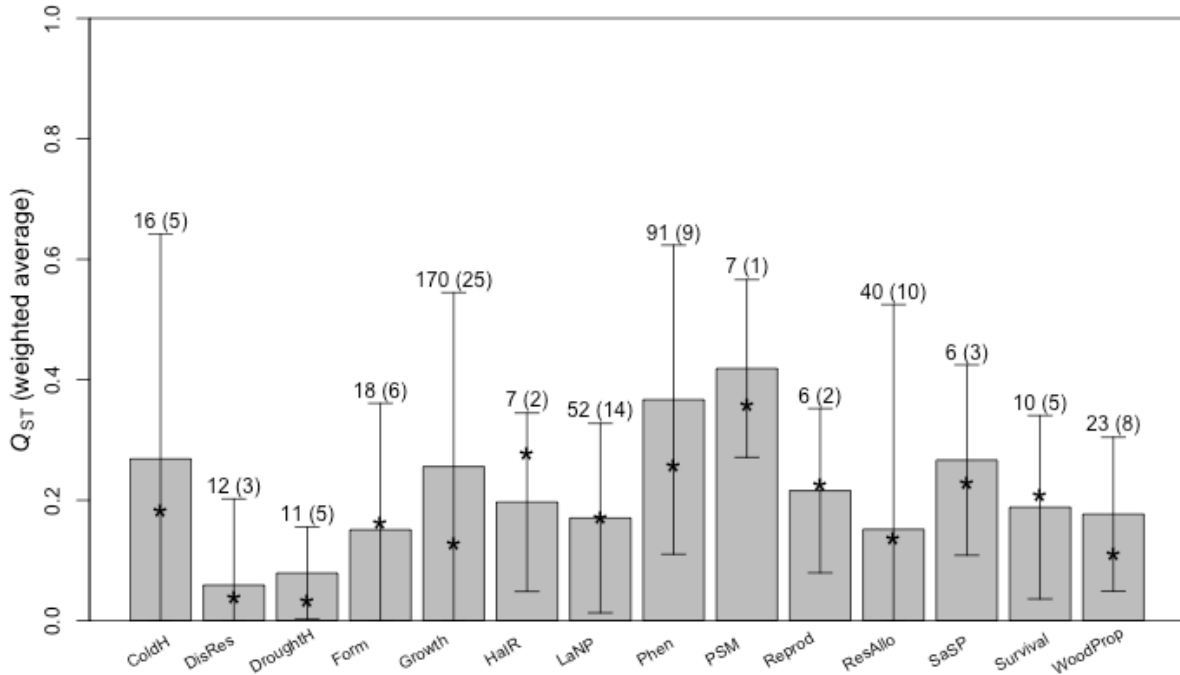
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2073 **Author Contributions**

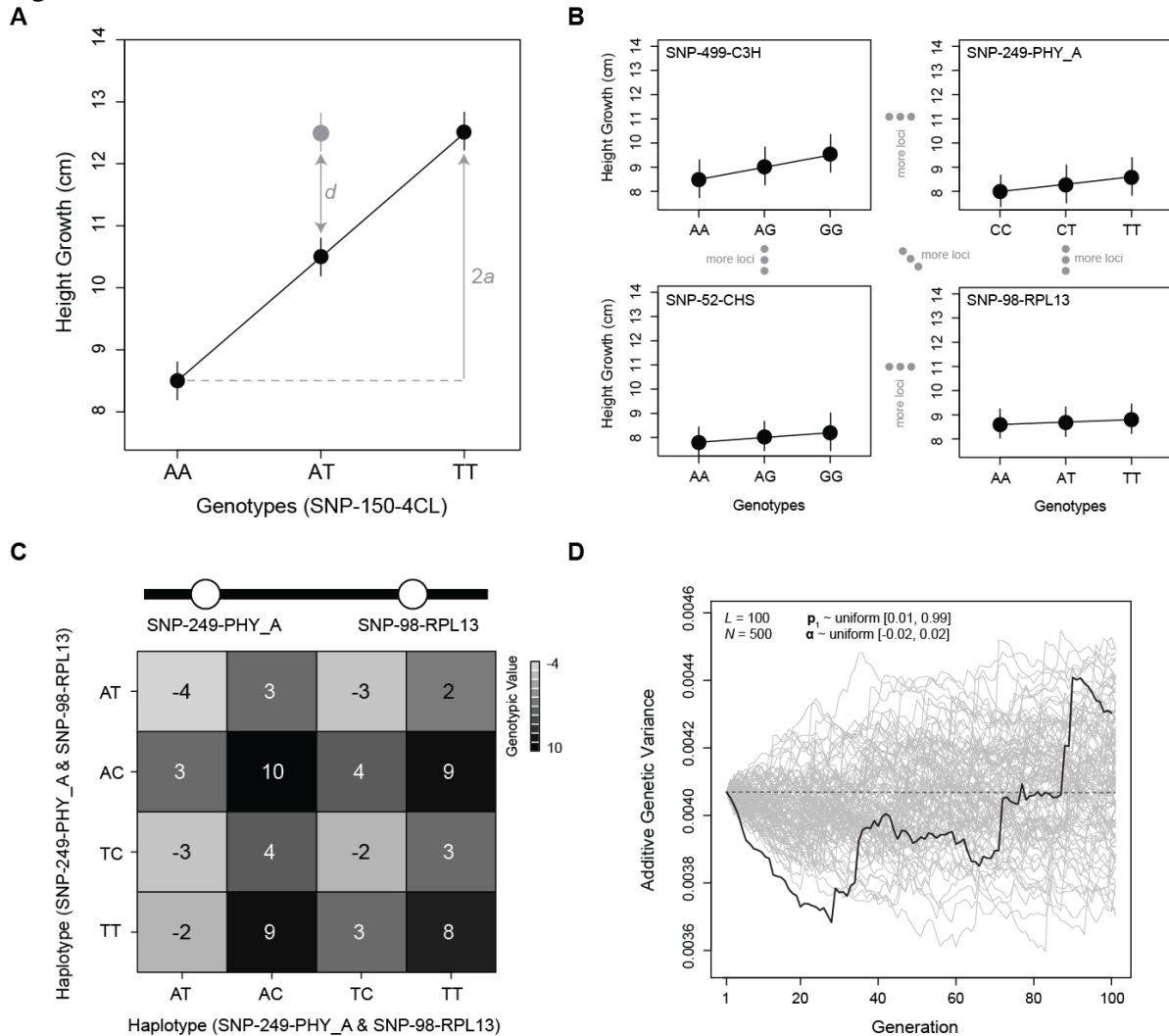
2074 BML and AJE conceived the review, with contributions from MM, CEB, and TMF. BML, MM,
2075 CEB, and TMF contributed to the literature search and survey which was analyzed by BML.
2076 CEB summarized Q_{ST} and F_{ST} comparisons. BML wrote the manuscript with contributions from
2077 MM and AJE. All authors contributed to the editing of the manuscript.

2078 **Figures**
2079 **Figure 1**



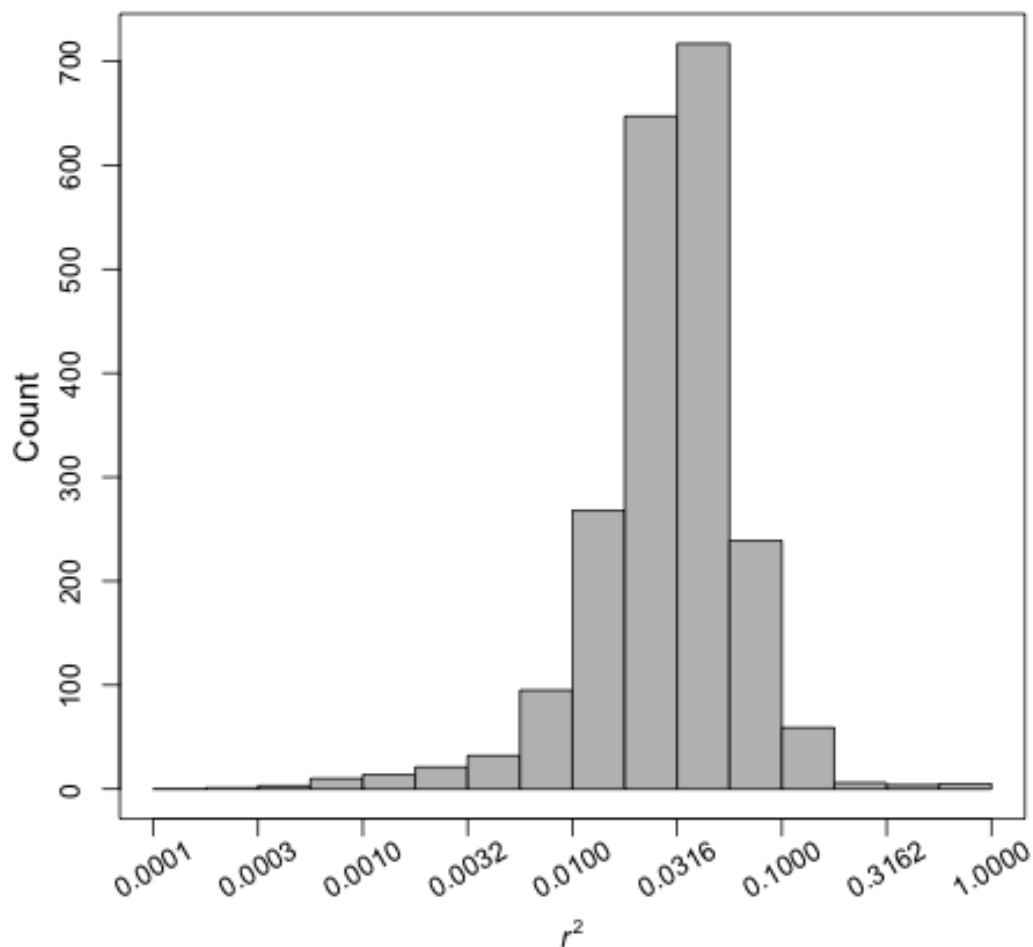
2080
2081 **Figure 1.** Average Q_{ST} for each of 14 trait categories from literature review calculated by
2082 weighting each estimate by the number of families used in the estimation. Error bars represent
2083 the standard deviation of the weighted averages. Numbers above error bars represent total
2084 number of estimates, with total number of unique species in parentheses. Asterisks indicate
2085 median values of the unweighted Q_{ST} distribution. ColdH = cold hardiness, DisRes = disease
2086 resistance, DroughtH = drought hardiness, HaIR = herbivore and insect resistance, LaNP = leaf
2087 and needle properties, Phen = phenology, PSM = plant secondary metabolites, Reprod =
2088 reproduction, ResAllo = resource allocation, SaSP = seed and seedling properties, WoodProp =
2089 wood properties.

2090 **Figure 2**



2091
 2092 **Figure 2.** Relevant quantitative genetic concepts are needed to understand the evolution of
 2093 polygenic traits. (A) Additive and non-additive effects at a single locus, where a is defined as the
 2094 additive effect (also known as the average effect of allelic substitution [α] when there is no
 2095 dominance) and d is defined as the dominance deviation. With dominance, $\alpha = a[1 + k(p - q)]$, where
 2096 k is the degree of dominance ($k = 0$: additive, $k = 1$: dominance, $k > 1$: over-dominance, see Lynch &
 2097 Walsh 1998). (B) Polygenic traits are determined by multiple genes, each with additive (shown) and
 2098 non-additive (not shown) effects. The total additive effect is the sum of the additive effects at all
 2099 causative loci. (C) Additive-by-additive epistasis, where the additive effect of an allele at the PHY_A
 2100 SNP depends on what allele it is paired with at the RPL13 SNP. In this case, the effects can be
 2101 thought of as dependent in the following manner using the four possible haplotypes at the PHY_A
 2102 (A/T SNP) and RPL13 (C/T SNP) SNPs – AC: +5, AT: -2, TC: -1, TT: 4. (D) The effect of genetic drift
 2103 on the additive genetic variance as determined by 100 independent, causative loci. Each line
 2104 represents a simulation of genetic drift in a constant sized population ($n = 500$ diploids) conditioned
 2105 on initial allele frequencies across loci (p_i) and effect sizes (α). The expected mean across all 100
 2106 simulations is given by the dashed black line. Any given simulation can deviate strongly from this
 2107 expectation (solid black line). Thus, when the elements of p change over time, in this case due to
 2108 genetic drift, so does the additive genetic variance. See also Supplemental Box S1.

2109 **Figure 3A**

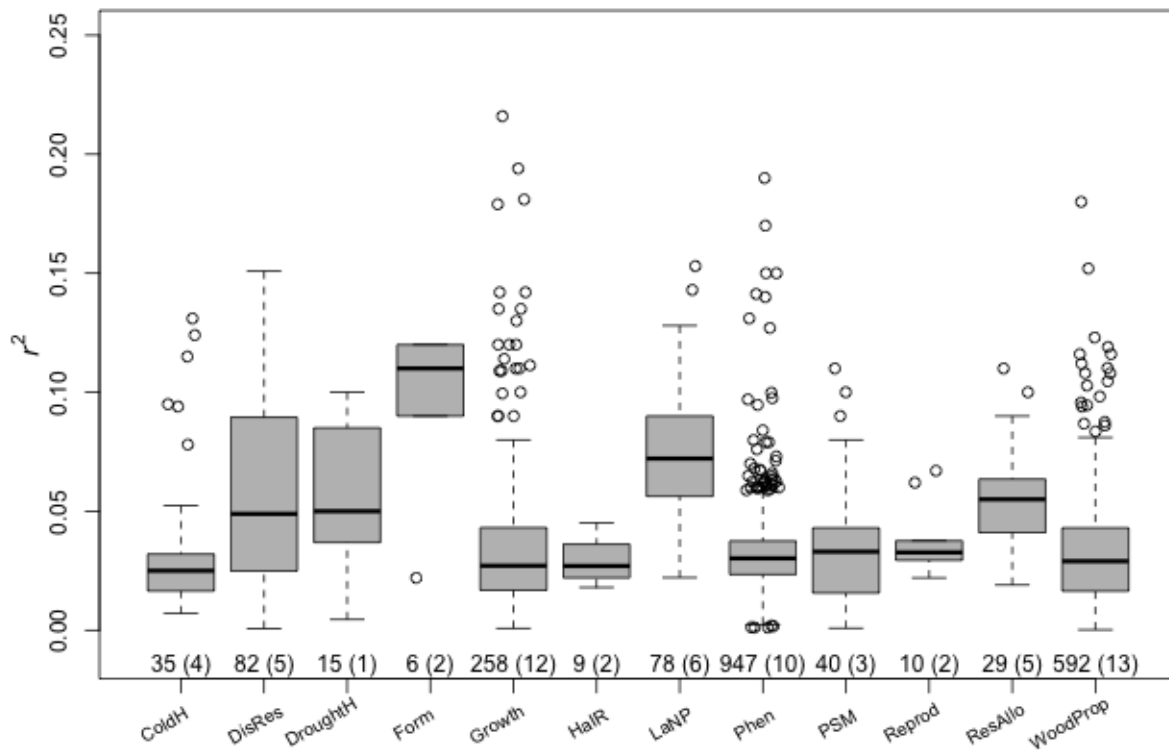


2110

2111 **Figure 3A.** Counts of per-locus percent variance explained (r^2) estimates from single-locus

2112 genotype-phenotype associations from literature review. Note logarithmic x-axis.

2113 **Figure 3B**



2114
2115 **Figure 3B.** Distribution of per-locus percent variance explained (r^2) values for trait groups
2116 within genotype-phenotype literature review. Values along x-axis are total number of estimates
2117 and number of species across estimates. Not shown are nine outliers for disease resistance to
2118 *Cronartium ribicola* in *Pinus monticola* (range = [0.402, 1.0]) from Lui *et al.* 2017.
2119 Abbreviations as in Figure 1.

2120 **Tables**

2121 **Table 1.** Where to next? The Path Forward identifies meaningful ways in which we can progress
2122 our understanding of the architecture underlying complex traits in trees. Here we outline some
2123 questions that can be used to guide future inquiry as the number of markers and sequence
2124 length increase, and annotation becomes more precise and specific to tree biology.

- 2125
- 2126 1) Composition and evolution of architectures
 - 2127 a. How prevalent are non-additive contributions to underlying genetic architectures?
2128 Are there patterns across similar phenotypes or regulatory networks? Is there
2129 evidence that such non-additive effects have constrained or facilitated local
2130 adaptation?
 - 2131 b. Are adaptive loci most prevalent in areas of low recombination or repetitive
2132 sequences (e.g., retrotransposons, clustered gene families)? Are loci of similar
2133 effect sizes or expression profiles clustered within the genome?
 - 2134 c. At what frequency does local adaptation result in fitness tradeoffs across
2135 environments (Tiffin & Ross-Ibarra 2014; Wadgyman et al. 2017)? And does this
2136 interact with demographic history in trees?
 - 2137 d. Does pleiotropy play a substantial role in underlying tree architectures?
 - 2138 e. Which aspects of genetic architectures in trees are likely to exhibit deleterious
2139 variation? And how much of this signal are we capturing in genotype-phenotype
2140 applications?
 - 2141 2) Repetitiveness of architectural organization
 - 2142 a. Should we expect genome organization to vary across populations?
 - 2143 b. Under what situations in trees are we likely to observe genomic reorganization
2144 (e.g., physical linkage or dispersion) due to selection pressures (Lotterhos et al.
2145 2017)? Will reference genomes be suitable to assess this question across
2146 species or diverged populations, or can long-read sequencing technologies
2147 (reviewed in Jiao & Schneeberger 2017) offer appropriate resources?
 - 2148 c. At what level of the genetic architecture do we see patterns of convergence,
2149 parallelism, and divergence? Within core hubs, or perhaps within aspects of the
2150 periphery? What does the comparison of topologies from such architectures tell
2151 us about influential evolutionary processes?
 - 2152 d. What is the prevalence of convergent and parallel adaptation within polygenic
2153 architectures across populations and species?
 - 2154 e. How often are architectures influenced by variation in expression levels rather
2155 than structural variation in proteins? Do architectures differ in predictable ways
2156 with the prevalence of one or the other?

2157 **Boxes**

2158 **Box 1: A step in the right direction: Synergism between GWAS and Genomic Selection**

2159

2160 Early simulations showcased the promise of predicting breeding values from marker data to
2161 accelerate domestication and breeding of plants and animals (Meuwissen et al. 2001; Bernardo &
2162 Yu 2007; Heffner et al. 2009; Zhong et al. 2009), and particularly under the framework of
2163 genomic selection (GS) in trees (Wong & Bernardo 2008; Grattapaglia & Resende 2011; Iwata et
2164 al. 2011; defined and reviewed by Grattapaglia 2017). Much of the early exploration into the
2165 applicability of GS in trees discounted the utility of marker-assisted selection (MAS) because of
2166 the small estimated effects for the few loci significantly associated via single-locus approaches at
2167 the time, as well as having concerns related to replication because of the identification of
2168 markers across limited parental (genetic) backgrounds (Grattapaglia & Resende 2011; Iwata et
2169 al. 2011; Resende et al. 2012a, 2012b). Based on these arguments and results from simulations,
2170 genomic selection was identified as a more promising endeavor than MAS, particularly if the
2171 breeding cycle can be reduced via efforts such as grafting (Grattapaglia & Resende 2011) or
2172 somatic embryogenesis (Resende et al. 2012a).

2173

2174 While GS techniques often can explain a considerable proportion of narrow sense heritability,
2175 current implementation of GS in trees is often on par with, or marginally better than, traditional
2176 phenotypic selection when evaluating potential within the same generation and environment (see
2177 Table 9.1 in Grattapaglia 2017). Further, the predictive accuracy of various models are a function
2178 of underlying architecture (Resende et al. 2012c; Grattapaglia 2017). As pointed out by
2179 Grattapaglia (2017), current marker densities have produced satisfactory results due to the
2180 capture of relatedness between training and validation populations. Here, this success is likely
2181 due to the ability of markers to reasonably represent large haplotype blocks (and thus cumulative
2182 action of causative effects) due to the high level of relatedness between training and validation
2183 populations. Even so, Grattapaglia (2017) recommends higher marker densities so that markers
2184 also capture true marker-QTL LD and thus sustain long-term accuracies across generations and
2185 environments. We also believe GWAS applications (*sensu lato*) in trees will also see
2186 improvements through increased marker densities, the results of which can then be used to
2187 further test specific hypotheses regarding underlying architectures and to increase predictive
2188 accuracies of GS as well. Incorporating markers that putatively underlie the trait of interest into
2189 model prediction may spur opportunities that do not require high degrees of relatedness between
2190 training and validation populations, perhaps to the extent of incorporating material from outbred
2191 stands using predictive approaches (*sensu* Bérénos et al. 2014; Bontemps 2016) and heritability
2192 validation (*sensu* Castellanos et al. 2015) in the field.

2193

2194 In the end, the realized progress of our understanding regarding the genomics of complex traits
2195 in trees will therefore be enhanced by the deposition of data from both GS and GWAS
2196 approaches into a centralized open-access database hub such as TreeGenes (treegenesdb.org).
2197 Future meta-analyses can then synthesize past inquiry to summarize our current understanding of
2198 underlying genetic architectures, ultimately incorporating this knowledge towards future
2199 applications in industry and conservation (see The Path Forward; Table 1).

2200

Supplemental Information

2201

The genomics of local adaptation in trees:

2202

Are we out of the woods yet?

2203

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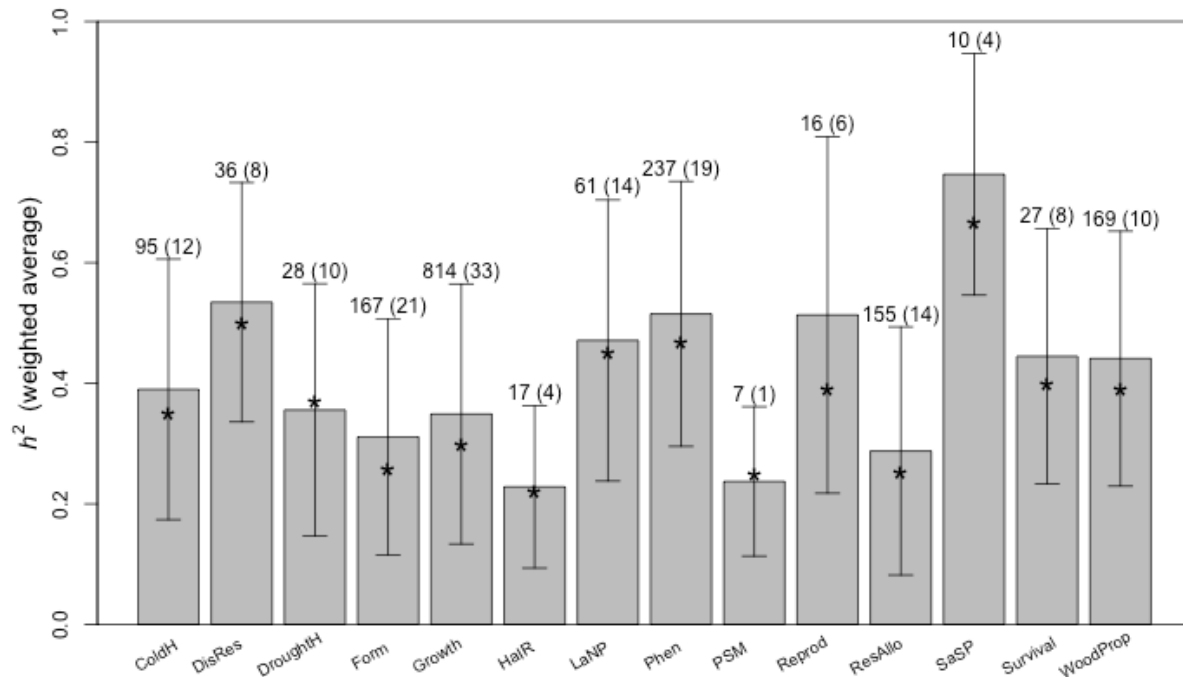
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2207

Running Title: Are we out of the woods yet?

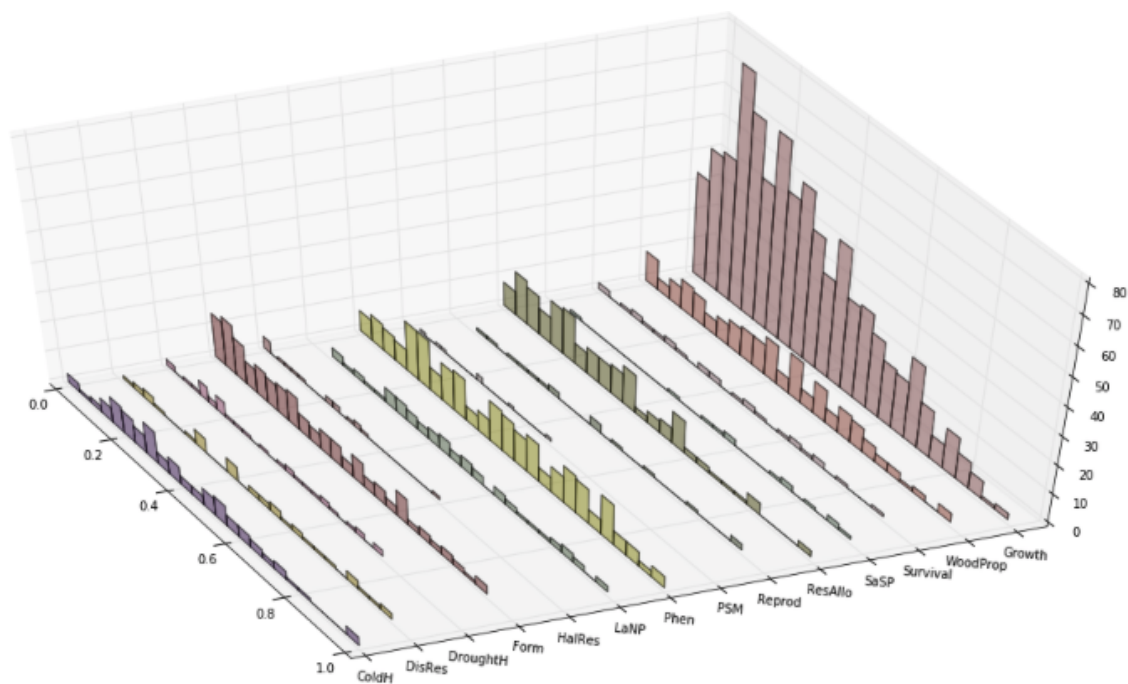
2208 **Supplemental Figures**

2209 **Figure S1**

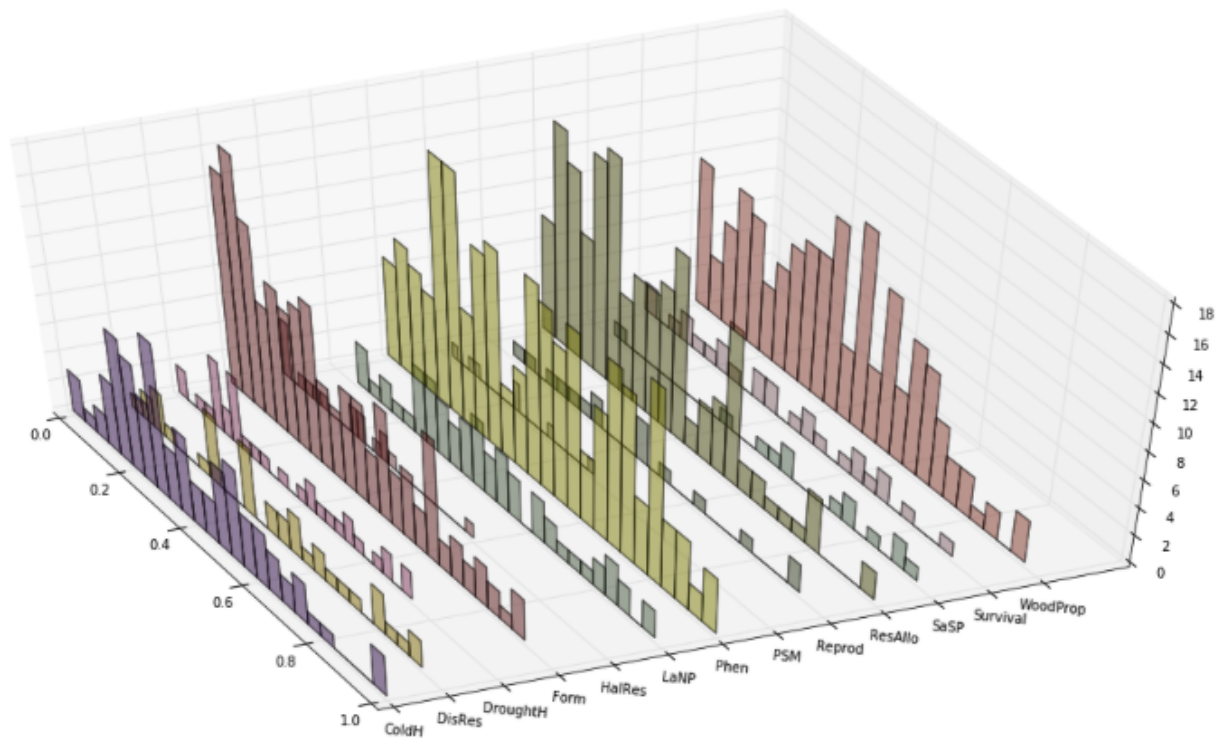


2210 **Figure S1.** Averages of narrow sense heritability calculated by weighting the number of families
2211 used in each estimate of heritability. Error bars represent the standard deviation of the weighted
2212 averages. Note that genetic variances in juvenile traits may be inflated due to instances of
2213 maternal effects, which we did not control for in our literature survey. Abbreviations as in Figure
2214 1 of the main text.
2215

2216 **Figure S2**

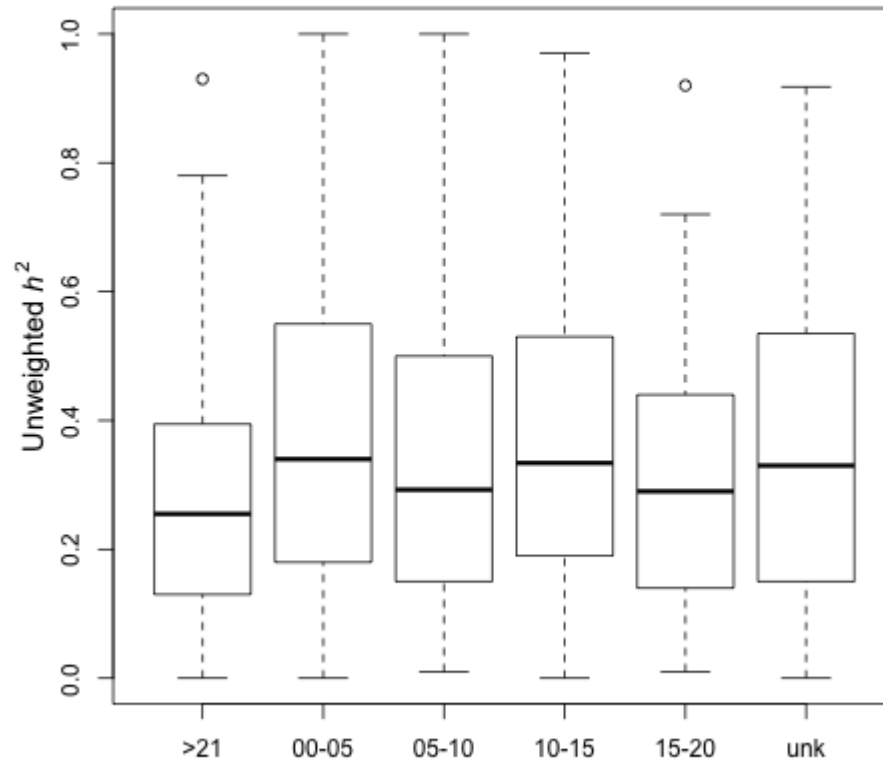


2217
2218



2219 **Figure S1.** Distributions of unweighted narrow sense heritability with (A) and without (B)
2220 inclusion of the Growth distribution. Trait abbreviations as in Figure 1 of the main text.
2221

2222 **Figure S3**

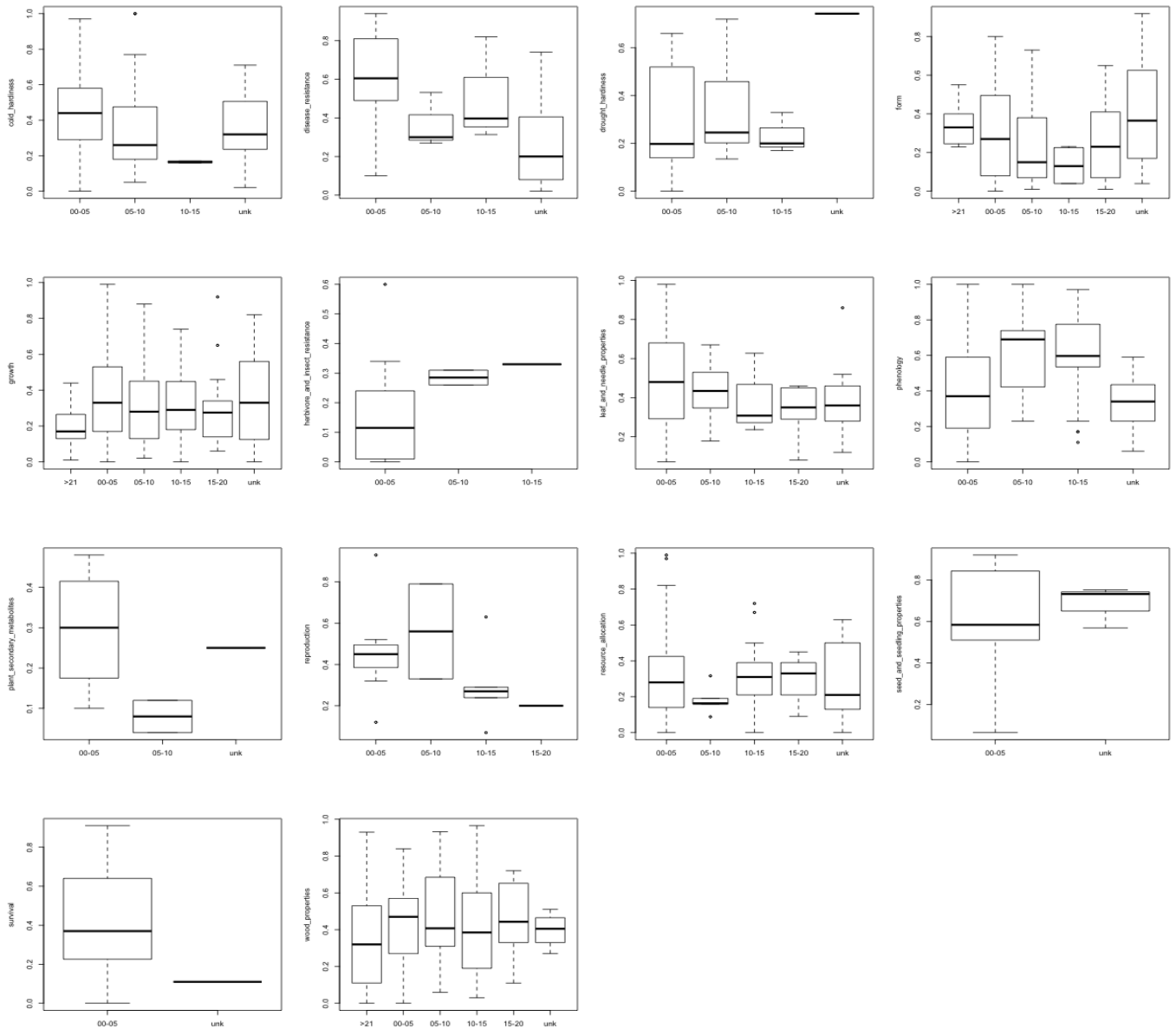


2223

2224 **Figure S2.** Unweighted narrow sense heritability distributions by age (years). Unk = unknown

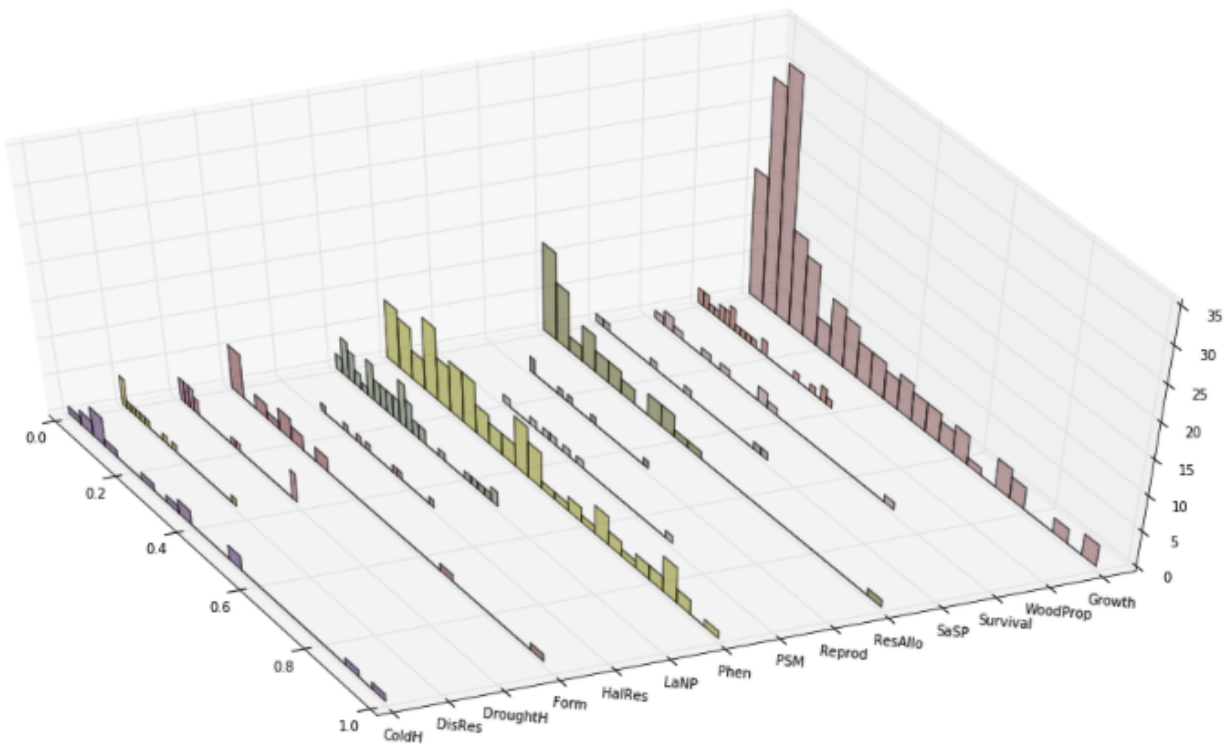
2225 age (i.e., not specified by article).

2226 **Figure S4**



2227
2228 **Figure S3.** Unweighted narrow sense heritability distributions by age (years) and by trait
2229 category. Unk = unknown (i.e., not specified by article)

2230 **Figure S5**



2231
2232 **Figure S4.** Distributions of unweighted QST estimates from literature survey. Abbreviations as
2233 in Figure 1 of the main text.

2234 Supplemental Tables

2235 Table S1

Trait Group	Total measurements	Total species	Angiosperm meas.	Gymnosperm meas.	Eucalypt meas.	Pine meas.	Populus meas.
Cold hardiness	16	5	3	13	0	10	2
Disease resistance	12	3	3	9	0	9	3
Drought hardiness	11	5	5	6	0	3	4
Form	18	6	15	3	1	3	12
Growth	170	25	73	97	13	74	44
Herbivore and insect resistance	7	2	7	0	6	0	0
Leaf and needle properties	52	14	44	8	11	5	12
Phenology	91	9	63	28	0	16	53
Plant secondary metabolites	7	1	7	0	7	0	0
Reproduction	6	2	2	4	2	4	0
Resource allocation	40	10	16	24	3	22	10
Seed and seedling properties	6	3	2	4	0	0	0
Survival	10	5	2	8	2	8	0
Wood properties	23	8	14	9	8	4	5

2236 **Table S1.** Summary of total and per-species measurements used in literature review of
 2237 differentiation of quantitative genetic variation (Q_{ST}). The Total Species column is the number of
 2238 unique species in our survey, whereas the remaining columns provide the total number of
 2239 measurements per category.

2240 **Table S2**

trait Group	Total measurements	Total species	Angiosperm measurements	Gymnosperm measurements	Eucalypt meas.	Pine meas.	Populus meas.
old hardiness	35	4	2	33	0	0	2
disease resistance	82	5	31	51	0	51	30
drought hardiness	15	1	0	15	0	15	0
form	10	3	4	6	0	5	4
growth	258	12	205	53	44	17	152
herbivore and insect resistance	9	2	9	0	6	0	3
leaf and needle properties	78	6	58	20	0	5	45
phenology	947	10	886	61	0	0	846
plant secondary metabolites	52	3	32	20	29	20	3
reproduction	10	2	0	10	0	9	0
resource allocation	29	5	19	10	4	8	15
wood properties	588	12	410	178	94	136	312

2241 **Table S2.** Summary of total and per-species measurements used in literature review of percent
 2242 phenotypic variance explained by associated markers (r^2). The Total Species column is the
 2243 number of unique species in our survey, whereas the remaining columns provide the total
 2244 number of measurements per category.

2245 Boxes

2246 Supplemental Box S1: Basic Concepts from Quantitative Genetics

2247 We follow the traditional decomposition of phenotypes into genetic and environmental components, which
2248 forms the basis of quantitative genetics (Fisher 1918, Lynch & Walsh 1998, Charlesworth & Charlesworth
2249 2010, reviewed by Hill 2010). The phenotype of an individual (P) can be decomposed into effects from its
2250 genotype (G), its environment (E), and the interaction between its genotype and environment (GxE).
2251 Typically, this is thought of as deviations from the population mean, with each causative locus having two
2252 alleles. Using this framework, phenotypic variance (σ^2_P) can be decomposed into genotypic variance
2253 (σ^2_G), environmental variance (σ^2_E) and the variance due to the interaction between genotypes and
2254 environments (σ^2_{GxE}):

$$\sigma^2_P = \sigma^2_G + \sigma^2_E + \sigma^2_{GxE}$$

2255 For a single locus, σ^2_G can be decomposed into variances arising from additive (σ^2_A) and dominance (σ^2_D)
2256 effects (Fig. 1). For multiple loci, σ^2_G can be decomposed into variances arising from additive, dominance,
2257 and epistatic (σ^2_I) effects, with the total additive effect across loci being the summation of the effects at
2258 each of the causative loci. Dominance and epistatic effects are jointly termed non-additive effects. Thus,
2259 the previous equation can be expanded to the following:

$$\sigma^2_P = \sigma^2_A + \sigma^2_D + \sigma^2_I + \sigma^2_E + \sigma^2_{GxE}$$

2260 The decomposition of σ^2_G into different types of effects provides a way of estimating narrow-sense
2261 heritability (h^2), which is defined as the ratio of additive genetic variance (σ^2_A) to total phenotypic variance
2262 (σ^2_P). For tree populations, this is often accomplished through variance partitioning techniques
2263 (Namkoong 1979) using half-sib designs in common gardens (White *et al.* 2007) or using molecular
2264 markers to estimate relatedness in the field (cf. Ritland & Ritland 1996). In the case of half-sib designs, if
2265 the assumptions of free recombination and little epistasis among causative loci, random mating, and lack
2266 of environmental covariance among sibs are satisfied, σ^2_A is given by (Lynch & Walsh 1998):

$$\sigma^2_A = 4\sigma^2_F$$

2267 where σ^2_F is the variance due to family (e.g., as extracted from a linear mixed model). Hence, for a half-
2268 sib design, $h^2 = 4\sigma^2_F/\sigma^2_P$. Other sibling designs are possible, with the 4 in the previous equation replaced
2269 by $1/r$, where r is the coefficient of relationship (e.g., Whitlock & Gilbert 2012). Clonal and controlled
2270 mating designs are also often used for estimation of heritability, often broad-sense heritability (Namkoong
2271 1979; White *et al.* 2007). When families are nested into populations, and an estimate of the among
2272 population variance component is made, these are the components also used to estimate Q_{ST} (Spize
2273 1993; Prout & Barker 1993). When compared against estimates of F_{ST} using a similar variance
2274 decomposition procedure (e.g., Yang 1998) and a method suitable to account for the substantial variance
2275 associated with these components (e.g., Whitlock & Guillaume 2009) conclusions about local adaptation
2276 can be reached.

2277 Heritability estimates are population, environment, and time specific, as evidenced by the relationship
2278 between σ^2_A and allele frequencies within populations (Lynch & Walsh 1998; e.g. Berg & Coop 2014):

$$\sigma^2_A = 2 \sum_{i=1}^L \alpha_i^2 p_i (1 - p_i)$$

2291 where the summation is over the number of causative loci (L), α is the average effect of allelic substitution
2292 at each locus (Fig. 1), and p_i is the frequency of one of the alleles at each of the causative loci. Thus, any
2293 evolutionary force altering p at some or all of the causative loci will change σ^2_A (cf. Box 3.7 in
2294 Charlesworth & Charlesworth 2010). Heritability is also uninformative about the underlying architecture
2295 itself, as are the relative magnitudes of the different variance components themselves (Huang & Mackay
2296 2016), and can often be misleading about evolutionary potential (Hansen *et al.* 2011).

2299 **Supplemental Box S2:**

2300 *Brief introduction to methods for single-locus genetic association analysis*

2301
2302 Detecting associations between genetic markers and complex trait variation relies on fitting and
2303 evaluating linear models, typically of the form:

$$\mathbf{y} = \mathbf{X}\boldsymbol{\beta} + \mathbf{Z}\mathbf{u} + \mathbf{e},$$

2304
2305
2306 where \mathbf{y} is a vector of observed or inferred phenotypic values, $\boldsymbol{\beta}$ and \mathbf{u} are vectors of random and fixed
2307 effects, respectively, \mathbf{X} and \mathbf{Z} are design matrices associated with $\boldsymbol{\beta}$ and \mathbf{u} , and \mathbf{e} is a vector of residuals
2308 (Yu *et al.* 2005). In the simplest model, the phenotype (\mathbf{y}) is modeled as a function of genetic effects at a
2309 single locus, represented by marker genotypes for the samples comprising values in \mathbf{y} , and covariates
2310 describing relatedness among sampled trees and the structure among populations from which those trees
2311 were sampled. Genetic effects are encoded based on *a priori* assumptions about the underlying
2312 architecture of the phenotypic trait under consideration, with the most frequent encoding being that for
2313 additive effects (e.g., counts of a reference allele) considered as either fixed or random effects (Goddard
2314 *et al.* 2009). Phenotypic values are often estimates derived through analysis of materials established
2315 within common gardens, either from clones or sibships, from which estimates of the genetic values of
2316 unmeasured trees (e.g., maternal trees for which markers have been genotyped) are made using the
2317 theory of Best Linear Unbiased Predictors (BLUPs; Henderson 1975; Searle *et al.* 1992; Piepho *et al.*
2318 2008). Inclusion of only fixed effects results in a General Linear Model (GLM), whereas a mixture of fixed
2319 and random effects results in a Mixed Linear Model (MLM or LMM). The use of covariates is necessary to
2320 avoid identification of false positive associations arising from the confounding between neutral genetic
2321 and phenotypic variation due to demographic history and the analysis of relatives (Devlin & Roeder 1999;
2322 Yu *et al.* 2005; Price *et al.* 2006).

2323
2324 Models as described above are typically fitted and evaluated using restricted maximum likelihood (REML,
2325 Patterson & Thompson 1971), although Bayesian methods are available and have the advantage of
2326 specifying *a priori* assumptions more clearly, remove the distinction between fixed and random effects,
2327 and are more applicable to testing biologically realistic models (Stephens & Balding 2009). Output from
2328 these models include estimates of effect sizes for markers (e.g. r^2 , coefficients for random effects,
2329 genotypic trait means) and, when used in a frequentist framework, probability values (p -values) of
2330 observing test statistics under a null model. Bayesian methods, in contrast, provide strength of evidence
2331 measures such as Bayes Factors for the association of each marker to the phenotype of interest. The
2332 ability to discover and correctly quantify effect sizes of true positives (i.e. causative markers or indirect
2333 associations resulting from linkage to causative markers) is dependent upon experimental design,
2334 including design of genotyping assays, and sample sizes (Long & Langley 1999; Zöllner and Pritchard
2335 2007; Spencer *et al.* 2009), as well as genome-wide patterns of linkage disequilibrium relative to the
2336 density of markers in the genome, the genetic distance between the indirectly associated marker and the
2337 causative locus, and the true underlying genetic architecture of the phenotypic trait under consideration
2338 (Platt *et al.* 2010; Prichard *et al.* 2010; Caballero *et al.* 2015).

2339
2340 One model is typically fitted and evaluated per marker-phenotypic trait combination (but see e.g. Wegrzyn
2341 *et al.* 2010 for haplotype analysis). Even without the issue of confounding described above, this increases
2342 the likelihood of false positives arising solely from performing many statistical tests. A variety of methods
2343 exist to deal with multiple testing, with the most popular methods being those based on the false
2344 discovery rate (Storey & Tibshirani 2003) and permutation (Hirschhorn & Daly 2005). Additional methods
2345 exist for situations where the multiple tests are not independent from one another (e.g. linkage
2346 disequilibrium among markers, see Johnson *et al.* 2010) or when permutation analysis is problematic
2347 (Joo *et al.* 2016).

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