Neurobehavioural Correlates of Obesity are Largely Heritable

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Abstract

Recent molecular genetic studies have shown that the majority of genes associated with obesity
are expressed in the nervous system. In other studies, obesity has been associated with various
neurobehavioural factors like brain morphology, cognitive performance, and personality. Here,
we tested whether these neurobehavioural factors were associated with the heritable variance in
obesity measured by body mass index (BMI) in the Human Connectome Project (N = 895
siblings). We found that obesity was associated with various neurobehavioural factors: cortical
thickness findings supported Alonso-Alonso and Pascual-Leone's "right brain hypothesis" for
obesity. Namely, increased BMI associated with decreased cortical thickness in right frontal
lobes and increased thickness in the left frontal lobes. In addition, lower thickness and volume in
entorhinal-parahippocampal structures, and increased thickness in parietal-occipital structures in
obese subjects supported the role of visuospatial function in obesity. Morphometry results were
supported by cognitive tests, which outlined obesity's negative association with visuospatial
function, verbal episodic memory, and cognitive flexibility. Personality-obesity correlations were
small. We then aggregated the effects for each neurobehavioural factor for behavioural genetics
analysis and demonstrated their genetic overlap with obesity. Namely, all aforementioned
neurobehavioural factors had 0.22-0.52 genetic correlations with obesity and the phenotypic
correlations were 80%-100% explained by genetic factors. In a separate analysis, brain
morphology effects were partly represented by the aggregated personality and cognitive factors,
as they mediated 8-12% of brain morphology effects on obesity. In summary, the genetic
vulnerability to obesity is expressed in various neurobehavioural factors, which can inform
intervention strategies.

Introduction

Obesity is a widespread condition leading to increased mortality (1) and economic costs (2).

Twin and family studies have shown that individual differences in obesity are largely explained

by genetic variance (3). Gene enrichment patterns suggest that obesity-related genes are

preferentially expressed in the brain (4). While it is unclear how these brain-expressed genes lead

to obesity, several lines of research show that neural, cognitive, and personality differences have

a role in vulnerability to obesity(5, 6). Here we seek to test whether these neurobehavioral factors

could explain the genetic variance in obesity.

Diverse neurobehavioural effects have been found. In the personality literature, obesity is most

often negatively associated with Conscientiousness (exhibiting self-discipline and orderliness)

and positively with Neuroticism (a tendency towards negative affect) (7, 8). In the cognitive

domain, tests capturing executive function, inhibition, and attentional control have a negative

association with obesity (5, 6, 8–10). Neuroanatomically, obesity seems to have a negative

association with the grey matter volume of prefrontal cortex, and to a lesser extent the volume of

parietal and temporal lobes, as measured by voxel based morphometry (11). It has also been

suggested that structural and functional asymmetry of the prefrontal cortex might underlie

overeating and obesity (12). For genetic analysis, cortical thickness estimates of brain structure

have been preferred over volumetric measures (13). So far, cortical thickness patterns associated

with obesity have been inconsistent (14, 15). As a prerequisite to our goal of ascertaining the

heritability of brain-based vulnerability to obesity, we sought to extend these neurobehavioural

findings in a large multi-factor dataset from the Human Connectome Project. Since cortical

thickness measurement is limited to cortex only, we also included volumetric estimates of medial

temporal lobe and subcortical structures, which have been implicated in appetitive control (e.g.,

16).

The first goal was to assess whether the outlined obesity-neurobehavioural associations are of

genetic or environmental origin. Recent evidence from behavioural and molecular genetics

suggests that there is considerable genetic overlap between obesity, cognitive ability, and brain

imaging findings (17–19). However, the evidence so far is not comprehensive across all

neurobehavioural factors discussed. A recent paper has assessed the heritability of obesity-

associated brain structures themselves and found them varying (20). However, that study did not

analyse the heritability of the association between brain and obesity. The latter analysis is crucial

for understanding whether brain anatomy and obesity could have a genetic overlap, which would

suggest that the heritability of vulnerability to obesity is expressed in the brain.

The second goal is to test whether psychological measures from personality questionnaires and

cognitive tests mediate any brain morphology-obesity associations. Since variance in

psychological measures seems to originate from the brain (e.g., 21), performance in the cognitive

tests and questionnaires could be expected to mediate brain-obesity associations (6). This

hypothesis has rarely been tested. Within a normal population, only one study has suggested a

cognitive measure of inhibition to mediate the association between thickness of right superior

prefrontal cortex and obesity (22).

Establishing the mediation role of psychological measures in the associations between cortical

thickness and obesity would be beneficial in several ways. First, it would provide more concrete

meaning to the outlined brain-obesity associations, and support the theory that weight gain is

mostly caused by the behavioral and cognitive manifestations of brain variability. Second, if

psychological measures partly explain the effect of brain-obesity associations, the psychological

measures could be used as proxies for brain phenotypes (6)

Taken together, the goal of the current analysis is to use a large multifactor dataset to analyse the

heritability of the associations between obesity and brain/behavior. We further test whether

personality and cognitive phenotypes mediate the relation between brain and body mass index

(BMI).

95   **Results**

Background

We analysed data from 895 participants from the Human Connectome Project (23), consisting of

111 pairs of monozygotic twins and 189 pairs of dizygotic twins and siblings. Univariate

heritability using an AE (A: additive genetics and E: unique environment) model on twin and

sibling data revealed that BMI heritability was A = 70% [95% CI: 59%;77%], which is close to

the published meta-analytic estimate (A = 75% [74%; 76%] 3). Most control variables, such as

age, gender, race, ethnicity, education, income, as well as family structure had considerable

influence on BMI (SI Results: Control variables, Figure S2). Therefore, the effect of control

variables is accounted for in all analyses reported below.

Cognitive and Personality Factors

Log-transformed and residualized BMI was negatively correlated with several cognitive tests

capturing visuospatial ability, verbal memory and cognitive flexibility (Figure 1A). Tests

measuring reading abilities and lack of impulsivity (Delay Discounting) also showed a negative

trend, which was not significant at the false discovery rate (FDR) corrected level. No personality

test surfaced when FDR correction was applied (Figure 1B). Trendwise, there were effects of

Agreeableness, Aggression and Fear: somatic symptoms.

Brain Morphology

Cortical thickness was estimated from each T1-weighted MRI using CIVET 2.0 software (24).

Parcel-based analysis identified negative associations with BMI in right inferior lateral frontal

cortex, and bilateral entorhinal-parahippocampal cortex (Figures 2A & 3A). Positive associations

were found with the left superior frontal cortex, left inferior lateral frontal cortex, and bilateral

parietal cortex parcels.

We formally tested the effect of frontal lobe asymmetry on BMI. We averaged all frontal parcels

(Frontal section in Figure 2A) into left and right frontal lobes. Cortical thickness of the left

frontal lobe correlated positively with BMI ($r = 0.17$, $p<0.001$), and right frontal lobe thickness

correlated negatively with BMI ($r = -0.10$, $p = 0.021$). This correlation difference was highly

significant ($z = 4.27$, $p<0.001$)(25). The areas demonstrating the greatest asymmetry effect were

in the inferior frontal gyrus (Figure 2A & Figure 3A).

Brain volume was measured for estimation of brain morphology-obesity associations in brain

structures not covered by the cortical thickness algorithm. Namely, medial temporal lobe and

subcortical volumes were individually segmented and measured by registering each brain to a

labelled atlas using ANIMAL software (26). Volumetric results supported the role of entorhinal

cortex and also highlighted a positive association of left amygdala volume with BMI (Figures 2B

& 3B). No subcortical region had a significant association with BMI.

Creating poly-phenotype scores

We performed dimension reduction for heritability and mediation analyses to reduce

measurement noise and avoid multiple testing with redundant measures. We used the weights of

each individual neurobehavioural factor (Figures 1 & 2) to create an aggregate BMI risk score or

poly-phenotype score (PPS). This approach is similar to polygenic approach in genetics, where

the small effects of each polymorphism are aggregated to yield a total effect score (e.g., 27). We

used the correlation values as weights to multiply each participant's scaled measurements, and

aggregated the results into a single composite variable, the PPS. The PPS reflects the total

association of a given factor with BMI. PPS-s were created separately for cognition (correlation

with BMI: $r = 0.18$, $p<0.001$), personality ($r = 0.17$, $p<0.001$), cortical thickness (r = 0.33,

$p<0.001$), medial temporal brain volume (r = 0.28, $p<0.001$), and subcortical brain volume ($r =$

0.07, $p = 0.034$). Ten-fold cross-validation showed that in out-of-sample data, the effect sizes

dropped but stayed significant if the PPS has significant individual phenotype weights (cognition

r = 0.12, $p<0.001$; cortical thickness r = 0.26, $p<0.001$; medial temporal lobe volume r = 0.19,

$p<0.001$, see SI Results: Overview of PPS-obesity associations). While the PPS-obesity effect

sizes are likely overestimated, it is less of a concern in the current analysis, as we focus on the

proportion of the PPS-obesity association explained by genetic factors, and by mediation models.

For the heritability analysis, we used PPSs not residualized for family structure, education, or

income, as this information is needed for modelling heritability.

Heritability

In the bivariate heritability analysis, we preferred the AE model that does not model the shared

environment (C) because it had lower Akaike Information Criteria (AIC) values than the classic

ACE model (Figure 4A, SI Table S4). All PPSs were found to be highly heritable, with the A

component explaining 37-80% of the variance (Figure 4B, SI Table S5). Genetic correlations (r_g)

between BMI and neurobehavioural factors ranged from $r_g = 0.25$ to $r_g = 0.51$, suggesting

considerable genetic overlap between the heritable variances of BMI and each factor, with the

exception of subcortical structures ($r_g = 0.06$, Figure 4C, SI Table S6). As expected from high

heritability of the traits and high genetic correlations, the phenotypic BMI-PPS correlations

described in the previous sections were 86-100% explained by genetic factors (Figure 4D, SI

Table S5).

Mediation

We used a mediation model to test whether the PPSs of cognition and personality mediated the

brain morphology effects on BMI. For cortical thickness (Figure 5A), personality and cognition

in total mediated 7.5% (p = 0.002) of the effects on BMI (Figure 5D). On their own, cognition

mediated 4.5% (p = 0.012), and personality 3% (p = 0.06) of the effect. For medial temporal lobe,

personality and cognition in total mediated 11.7% (p<0.001), with personality mediating 6.3% (p

= 0.009) and cognitive score mediating 5.4% (p = 0.017) of the effect (Figure 5B, Figure 5D).

The association between subcortical volume and BMI was not mediated (Figure 5C, Figure 5D).

We were unable to replicate a previously suggested mediation model where inhibition task

performance mediates the effect of right Superior Frontal Gyrus on BMI (22)(SI Methods:

Mediation).

171   **Discussion**

Cortical thickness, medial temporal lobe volume, and cognitive and personality measures all had

covariation with BMI, and their effect on BMI was almost entirely heritable. The effects of

cortical thickness and medial temporal lobe volume on BMI were 8-12% explained by variation

in cognitive abilities and personality traits. Together, our results from a large sample support the

role of brain and psychological constructs in explaining genetic variance in obesity.

BMI correlated with increased cortical thickness in the left prefrontal cortex and decreased

thickness in the right prefrontal cortex, supporting the right brain hypothesis for obesity (12). The

effect was most prominent in the inferior frontal gyrus (Figures 2A and 3A). Only preliminary

support for the right brain hypothesis has been previously available (15). Right prefrontal cortex

has been implicated in inhibitory control (21) and possibly bodily awareness (12). Many

neuromodulation tasks aimed at increasing self-regulation capacity often target right prefrontal

cortex. On the other hand, effects have also been demonstrated in studies targeting left prefrontal

cortex (28).

Cortical thickness results also highlighted the role of temporo-parietal perceptual structures in

obesity. Namely, obesity was associated with bilaterally decreased thickness of the

parahippocampal and entorhinal cortices, and with mostly right-lateralised increased thickness of

parietal and occipital lobes. Volumetric results within the medial temporal lobe supported the role

of entorhinal cortex and also suggested that obesity is positively associated with the volume of

left amygdala. Emergence of the effects of the right parietal structures together with right

prefrontal structures hint at the role of ventral frontoparietal network (29), whereas the

parahippocampal cortex is associated with episodic memory and context mediation (30). The

amygdala is implicated in emotional and appetitive responses to sensory stimuli, including food

cues (31).

Integrating these findings, one could envision a model where obesity is associated with a certain

cognitive profile (32). The model starts with a hyperactive visual attention system attributing

heightened salience to food stimuli, implicating the ventral visual stream and amygdala. These

signals are then less optimally tied into relevant context by the parahippocampal structures, and

less well moderated (or filtered) by the prefrontal executive system. This could result in

consummatory behavior more likely driven by the presence of appetitive food signals that are

ubiquitous in the obesogenic environment. An impaired response inhibition and salience

attribution model of obesity has been suggested based on the functional neuroimaging literature

(reviewed in 5, 31, 33). Functional MRI studies have consistently identified obesity to associate

with heightened salience response to food cues, coupled with reduced activation in prefrontal and

executive systems involved in self-regulation and top-down attentional control. A similar

conclusion emerged from a recent resting state network analysis of a subset of the HCP data (34),

in which obesity was associated with alterations in perceptual networks and decreased activity of

default mode and central executive networks.

This brain morphology-derived model has some support from cognitive tests. The role of

prefrontal executive control is outlined by obesity's negative association with cognitive

flexibility. Surprisingly, there was no effect of motor inhibition as measured by the Flanker

inhibitiory task. A relation between obesity and reduced motor inhibition, while often mentioned,

has been inconsistent even across meta analyses (8, 9). More intriguingly, obesity was found to

be negatively associated with spatial orientation and verbal episodic memory, tasks that tap into

the key functions associated with entorhinal and parahippocampal regions implicated in our study

(30). Therefore, both cognitive tests and brain morphology measures propose that the increased

salience of food stimuli could be facilitated by dysregulated context representation and executive

control in obesity.

Regarding personality, we were unable to find any questionnaire-specific effects, notably with

respect to Neuroticism and Conscientiousness, both often thought to be associated with obesity

(5, 6). There are potential explanations for this discrepancy. First, the meta-analytical effects of

Neuroticism and Conscientiousness are small ($r \approx 0.1$, 8), for which we might have been

underpowered after p-value correction. Second, controlling for family structure likely further

reduced the effect sizes (35). Third, the personality-obesity associations tend to pertain to more

specific facets and nuances than broad factors (36), therefore, further analysis with more detailed

and eating-specific personality measures is needed. Still, using the PPS approach, we were able to

uncover a general personality risk score that aggregated the multiple small effects, which we used

in subsequent analyses.

All the associations discussed were largely due to shared genetic variance between

neurobehavioural factors and BMI. This is in accordance with recent molecular genetics evidence

that 75% of obesity related genes express preferentially in the brain (4). The current evidence

further supports the brain-gene association with obesity vulnerability.

A possible interpretation here is pleiotropy – the existence of a common set of genes that

independently influence variance in both obesity and brain function. However, our results could

also support a causal relationship – that the genetic correlation is due to a persistent effect of

heritable brain factors on over-eating and hence BMI. For instance, we could hypothesize that the

heritable obesity-related cognitive profile introduces overeating when high-calorie food is

available. As high-calorie food is abundant, the cognitive profile could lead to repeated

overeating providing an opportunity for genetic obesity-proneness to express. Such longitudinal

environmental effects of a trait need not to be large, they just have to be consistent (37, see

discussion in 38). Of course, a reverse scenario is also possible – obesity leads to alterations in

cortical morphology due to the consequences of cardiometabolic complications, including

inflammation, hypertension, and vascular disease (reviewed in 11, 39). However, we find this

less plausible in our study for two reasons. First, the reviews outline that reported consequence of

chronic visceral adipose tissue accumulation is often global brain atrophy. In contrast, current

data suggests focal brain effects – vascular or inflammatory atrophy would not explain the left-

right asymmetry in the prefrontal cortex. Second, the same reviews suggest that global brain

atrophy is mostly seen in older participants, whereas in current sample had a mean age of 29.

Young adults often experience "healthy or transitional obesity", where clinical inflammation

levels (40) and other cardiometabolic comorbidities have not yet developed (41). Further, a recent

mendelian randomisation study suggested that low genetic potential for education (proxy for

cognitive ability, Figure S3B) leads to high risk for cardiovascular disease, whereas high genetic

risk for cardiovascular disease does not lead to low education (42). Therefore, this indirect

evidence suggests that the genetic neurobehavioural factors described here lead to higher risk of

obesity.

The mediation models provide evidence that the brain structure associations with BMI are partly

explained by cognitive abilities and personality traits. While we were unable to replicate a

previous specific effect implicating poor inhibitory control and the lateral prefrontal cortex (22),

our PPS-based analysis suggested that the general principle is true – brain structure effects were

partly mediated by cognitive and personality factors. Granted, the effects were small so cognitive

tests cannot fully replace brain structure measures yet. While a more focused analysis on

psychological measures and brain regions that had strong effects in current analysis might have

yielded higher mediation effects, we abstained from such analysis to avoid "double dipping" (43).

From a practical point of view, current evidence suggests that evidence from psychology and

neuroscience can be used to design intervention strategies for people with higher genetic risk for

obesity. One way would be modifying neurobehavioural factors, e.g. with cognitive training, to

improve people's ability resist the obesogenic environment (28, 32). Another path could be

changing the immediate environment to be less obesogenic (e.g., 44) so that individual

differences in neurobehavioural factors would be less likely to manifest. In any case, current

evidence highlights that obesity interventions should not only focus on diet, but also take into

account that obesity is genetically intertwined with variation in neurobehavioural profiles.

The current analysis has limitations. Due to the cross-sectional nature of the dataset, mediation

models are only suggestive – longitudinal designs would enable better insight into the causal

associations between brain morphology, psychological measures, and BMI or weight gain. While

the measurement of cognitive abilities and personality was exhaustive, it lacked some common

behavioural tasks like the stop-signal task, or common questionnaires measuring self-control,

impulsivity, and eating-specific behaviours (5, 6), that have been previously associated with body

weight. Particularly, the common eating-specific behaviours such as uncontrolled eating (45) are

likely better candidates for explaining brain morphology-BMI associations as they are more

directly related to the hypothesized underlying behaviour. Ideally, the PPSs should be applied not

to same data but to new data for unbiased estimates of PPS-obesity effect sizes. This is feasible,

as cross-validation analysis showed that the weights obtained are mostly generalizable to new

data, albeit with slightly smaller effect sizes. However, as the current study does not address the

effect size of PPS-obesity associations, but the relative proportion of effect explained by genetic

factors or mediation models, the possible overestimation of effect size is not a major concern.

One has to be careful in translating individual differences in cortical thickness in normal

population to underlying neural mechanisms. Diverse biological processes have been suggested

to influence MRI-based cortical thickness measures, ranging from synaptic density to apparent

thinning due to synaptic pruning and myelination (summarised in 46, 47). A definitive model of

the underlying mechanism that links normal variations in cortical thickness to differences in brain

function cannot be given, as cortical thickness has not been mapped with both MRI and histology

in humans (47). Similarly, BMI is a crude proxy for actual eating behaviours or health status. At

the same time, the associations between brain morphology and BMI were also present in out-of-

sample cross-validation, suggesting that the pattern is robust. Our conceptual interpretation has

support from measures of both brain structure and cognitive function. Still, future hypothesis-

driven studies have to elucidate the neurobehavioural mechanisms behind obesity proneness.

In summary, the current analysis provides comprehensive evidence that the obesity-related

differences in brain structure, cognitive abilities, and personality traits are largely due to shared

genetic factors. Mediation analysis using aggregate scores for each factor outlined that, in

principle, the brain morphology-obesity associations are partly mediated by personality traits and

cognitive abilities. We hope that increasingly larger longitudinal data sets, dedicated studies, and

more sophisticated methods will help to outline more specific neurobehavioural mechanisms for

a better understanding of obesity, and providing a basis for designing informed interventions.

Methods

Data were provided by the Human Connectome Project (23). Certain people were excluded due

to missing data or not fulfilling typical criteria. Exclusion details, demographics and family

structure are summarized in SI Methods: Participants and Table S1. See Table S2 for a complete

list of 22 questionnaires and 18 cognitive tests used in our analysis. Software pipelines for

obtaining cortical thickness, quality control, and brain parcellation are described in SI Methods:

Measures. Analysis scripts are available on request from the first author.

Figure S1 provides a schematic pipeline for data analysis. Details of each data analysis step are

outlined in SI Methods: Data analysis. We discuss missing data handling, and how cross-

validation was applied to test the robustness of PPS-based results. We further describe the main

principles of twin and sibling based heritability analysis and model selection. Finally, the

software and packages used are listed.

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Figures

Figure 1. Associations between body mass index (BMI), cognitive ability (A), and personality
traits (B). Error bars mark 95% confidence intervals. See Table S2 for explanation of cognitive
test names. Numerical values are reported in Table S3. EF = executive function; FFM = Five-
Factor Model; FDR = false discovery rate; Imp = (lack of) impulsivity; Lang = language; Mem =
memory; Neg = negative affect; Perc = perception; PWB = psychological well-being; Soc =
social Relationships; SSE = stress and self efficacy; WM = working memory

Figure 2. Associations between body mass index (BMI), cortical thickness (A) and regional brain

volume (B). Error bars mark 95% confidence intervals. Numerical values are reported in SI Table

S2. FDR = false discovery rate; Fro = frontal, Ins = insula; Occ = occipital; Par = parietal; Tem =

temporal; MTL = medial temporal lobe; SC = subcortical.

Figure 3. Associations between body mass index (BMI), cortical thickness (A) and brain volume
(B) on a brain template. Values are the same as in Figure 2. Colour bar applies to both sub-plots.

Figure 4. Heritability analysis of the association between poly-phenotype scores (PPS) and body mass index (BMI). (A) Akaike Information Criterion values for different model decompositions. The AE model was chosen for subsequent analyses. (B) Heritability of each trait. BMI has multiple estimates, since it was entered into a bivariate analysis with each other PPS separately. The effect of unique environment (E) is not shown, since E = 100-A. (C) Genetic correlations between BMI and each PPS. The genetic correlations are positive, because the PPSs are designed to positively predict BMI. None of the environmental correlations were significant and therefore not shown. (D) Heritability of the phenotypic correlation between BMI and PPS. Horizontal lines depict 95% confidence intervals. Confidence intervals were not estimated for subcortical structures, because the genetic and phenotypic correlations with BMI were very small. AIC = Akaike Information Criterion; Cogn = PPS of cognitive tests; CT = PPS of cortical thickness; MTL = PPS of medial temporal lobe volume; Pers = PPS of personality tests; SC = PPS of subcortical structure volumes

Figure 5. Summary of mediation models. (A-C) The role of cognition and personality in mediating brain effects on body mass index (BMI). Path strengths greater or equal to 0.07 have p <0.05. Total effect (c) is in brackets. (D) Summary of % mediated effect size [(a+b)/c]. Error bars mark 95% confidence intervals. Cogn = PPS of cognitive tests; CT = PPS of cortical thickness; MTL = PPS of medial temporal lobe volume; Pers = PPS of personality tests; SC = PPS of subcortical structures' volume; Total = personality + cognition effect.

Supporting information

SI Methods

Participants

Data were provided by the Human Connectome Project (23) WU-Minn Consortium (Principal

Investigators: David Van Essen and Kamil Ugurbil; 1U54MH091657, RRID:SCR_008749)

funded by the 16 NIH Institutes and Centers that support the NIH Blueprint for Neuroscience

Research; and by the McDonnell Center for Systems Neuroscience at Washington University,

Background, personality and cognitive data were analysed from 970 participants (S900 data

release). Participant relatedness information was updated from the S1200 data release. For the

current analysis, we applied the following exclusion criteria, as these they might confound brain-

obesity associations: people with missing values on crucial variables, such as age, BMI,

education, income, gender, race, and ethnicity (n = 6), hypo/hyper thyroidism (n = 4), other

endocrine problems (n = 16), underweight (BMI < = 18, n = 9), women who had recently given

birth (n = 9). In addition, as we used family information control for participants' relatedness, we

excluded participants that were half-siblings to other participants (n = 31).

The final dataset analysed consisted of 895 participants, demographics of which are summarised

in Table 1. The sample had good gender balance and reasonable variation in BMI and income. As

limitations, the sample was relatively young and well educated. Most people were white and non-

Hispanic, however other races-ethnicities were also represented. The participants were nested

into 383 families, typically having 1-3 siblings in the dataset.

Table S1. Descriptive statistics of sample analysed.

Variable	Summary
Age	$\bar{x} = 28.83$ (SD = 3.67) years
BMI	$\bar{x} = 27.28$ (SD = 5.77) kg/m^2
Education	$\bar{x} = 14.85$ (SD = 1.82) years

Ethnicity	Not Hispanic/unknown: 819 (91.51%) Hispanic/Latino: 76 (8.49%)
Family structure	381 families. 1 sibling: 82 (21.41%) 2 siblings: 112 (29.24%) 3 siblings: 150 (39.16%) 4 siblings: 35 (9.14%) 5 siblings: 4 (1.04%)
Gender:	Men: 420 (45.85%) Women with birth control: 146 (15.94%) Women without birth control: 350 (38.21%)
Handedness	Laterality score: $\bar{x}$ = 64.84, SD = 45.39
Income	<$10,000: 66 (7.37%) 10K-19,999: 79 (8.83%) 20K-29,999: 115 (12.85%) 30K-39,999: 104 (11.62%) 40K-49,999: 99 (11.06%) 50K-74,999: 180 (20.11%) 75K-99,999: 119 (13.3%) > = 100,000: 133 (14.86%)
Race	White: 665 (74.3%) Asian, native Hawaiian, or other Pacific islander: 45 (5.03%) Black or African American: 144 (16.09%) American or Alaskan native, more than one race, or other/unknown: 41 (4.58%)

For the heritability analysis between each neurocognitive factor and BMI, we randomly chose
one sibling pair per family, ensuring that the pair had complete data. Non-twin sibling pairs were
considered equivalent to dizygotic twin pairs with respect to heritability analyses once data was
residualized for age and gender. If multiple sibling pairs within a family had complete data, we
prioritized choosing monozygotic twin pairs and dizygotic twin pairs over non-twin sibling pairs.
Depending on the neurocognitive factor, the heritability analysis was conducted on 64-111 pairs
of monozygotic twins (median = 105) and 65-189 pairs of dizygotic twins and siblings (median =
160).

Measures

Psychological measures.

Participants completed an extensive set of questionnaires and cognitive tests (see 48, 49 for an

overview). In the current analysis, we included 22 questionnaires and 18 cognitive tests (see

Figure 2 and Table S2 for complete list). We named the set of questionnaires personality

variables, as personality encompasses various patterns of what people want, say, do, feel, or

believe (50). Based on our previous review (6) we chose cognitive tests capturing aspects of

executive function, memory, and language.

Cortical thickness.

All T1-weighted MRI images were processed using the CIVET pipeline (version 2.0) (24, 51,

52). Processing was executed on the Canadian Brain Imaging Network (CBRAIN) High

Performance Computing platform for collaborative sharing and distributed processing of large

MRI datasets (53). Briefly, native T1-weighted MRI scans were corrected for non-uniformity

using the N3 algorithm (54). The corrected volumes were masked and registered into stereotaxic

space, and then segmented into gray matter (GM), white matter (WM), cerebrospinal fluid (CSF)

and background using a neural net classifier (55). The white matter and gray matter surfaces were

extracted using the Constrained Laplacian-based Automated Segmentation with Proximities

algorithm (56, 57). The resulting surfaces were resampled to a stereotaxic surface template to

provide vertex based measures of cortical thickness (58). All resulting images were visually

inspected for motion artefacts by experienced personnel and then subsequently processed through

a stringent quality control protocol, which only 559 of the 879 participants in our initial cohort

passed. For those participants who passed, cortical thickness was then measured in native space

using the linked distance between the two surfaces across 81924 vertices and a 20mm surface

smoothing kernel was applied to the data (59). The Desikan–Killiany–Tourville (DKT) atlas was

used to parcellate the surface into 64 cortical regions (60). Cortical thickness was averaged over

all vertices in each region of interest for each subject (61) and the effect of mean cortical

thickness was regressed to allow for regional analysis (62).

Volumetric estimates.

Because our cortical thickness method does not cover all medial temporal and subcortical

structures, we used volumetric estimates for these brain regions. For subcortical volumetric

estimation, T1-weighted scans of the subjects were pre-processed through a computerized

pipeline (n = 897). Image denoising (63), intensity non-uniformity correction (54), and image

intensity normalization into range (0-100) using histogram matching were performed. After

preprocessing, all images were first linearly (using a 9-parameter rigid registration) and then

nonlinearly registered to an average template (MNI ICBM152) as part of the ANIMAL software

(26, 64). The subcortical structures, i.e., thalamus, putamen, caudate, and globus pallidus were

segmented using ANIMAL by warping segmentations from ICBM152 back to each subject using

the obtained nonlinear transformations. The medial temporal lobe structures, i.e. hippocampi,

amygdala, temporal pole, and parahippocampal, entorhinal and perirhinal cortices, were

segmented using an automated patch-based label-fusion technique (65). The method selects the

most similar templates from a library of labelled MRI template images, and combines them with

a majority voting scheme to assign the highest weighted label to every voxel to generate a

discrete segmentation. Quality control was performed on the individual registered images as well

as the automated structure segmentations by visual inspection, and inaccurate results were

discarded. The final results included data from n = 827, 8 parcels per subjects for the subcortical

structures, and n = 591, 12 parcels for the medial temporal lobe structures.

Data Analysis

A schematic pipeline of the analysis is displayed in Figure S1. Data from all neurocognitive

factors were first residualized for control variables (age, education, ethnicity, gender, handedness,

income, race) using a linear mixed model, adding a random intercept for family (Figure S1).

Using a random intercept for family resulted in very similar results, as compared to using the

theoretical genetic relatedness matrix to account for family effects. As BMI was skewed (long-

tail at the upper end of the scale), it was log-transformed to achieve a normal-like distribution.

Handedness was reversed and also log normalised.

As some psychological measures had missing responses, but psychological tasks often correlate,

missing data were replaced using multiple imputation, taking the mean of 100 simulated datasets.

For each factor category (cognition, personality, cortical thickness, medial temporal volume,

subcortical volume), factor-BMI relationships were assessed using univariate correlation between

each brain parcel or test score and BMI. We initially also tried using a partial least squares (PLS)

correlation approach, which is a multivariate technique suited to handling correlated predictors

(66, 67). However, the PLS estimates were extremely close to univariate correlations, and

univariate correlations were preferred for simplicity. As a result, we received an estimate of the

relative contribution (weight) of each predictor within a given factor.

Estimates used in this study are presented in Table S3. In that table, type enables filtering a

certain factor, (eg "ct" for cortical thickness), and control variables used. "_reg" stands for

regular control variables (e.g. all mentioned variables). Other types include weights without

controlling for education and income ("_neiyf"), and without controlling for education, income,

and family structure ("_neinf"). Finally, we also ran the correlation in 10 different combinations

of 90% subsets of the data and took the average of these estimates. Estimates obtained this way

were very close to full sample estimates and had little variation (Table S3, columns

"rmean_cv90" and "rsd_cv").

Using these weights, we created an aggregate BMI risk score or *poly-phenotype score (PPS)* for

each neurocognitive factor. This approach is inspired by polygenic risk score approach, where the

effects of single-nucleotide polymorphisms are added up to a total genetic score (27).

Specifically, we used the correlation-derived weights to multiply each participant's measured

values, and aggregated the results into a single composite variable for a given factor, the PPS. A

PPS would reflect the total association that a given factor has with BMI. Similarly to polygenic

scores, we do not apply a p-value threshold (68).

To get a sense of robustness of the PPSs, we ran ten-fold cross-validation on the BMI-PPS

correlations. Namely, we estimated the correlations on 90% of the data, and used the correlation

weights obtained to build PPSs both in the same 90% of data and in the 10% out of sample data.

The estimates obtained from 10 folds of 90% data were averaged. The estimates obtained from

10 folds of 10% data were put together, and a global estimate for 10% folds was found. When

creating the 10% folds, we created folds for each factor separately, as each factor has a different

number of available data points, ensuring that folds were as equal in size as possible. We also

ensured siblings from the same family were in the same fold. Therefore, no data from family

members were used in calculating both the correlation weights and performing out of sample

predictions.

In the heritability analysis, a typical behavioural genetics decomposition uses relatedness

assumptions between individuals to divide variance in a trait to the following components:

genetic variance (A, additive and interactive effects), shared environmental (C, family and shared

school effects), and unique environmental (E, unique experience and measurement error). The

assumptions are:100% of genetic variance shared between monozygotic twins, 50% of genetic

variance shared between dizygotic twins and sex-and gender residualized siblings, 100% of

family environment shared by all siblings, 0% unique variance shared between siblings. Such

decomposition is called univariate heritability.

Besides establishing univariate heritability, one can also conduct heritability analysis on the

covariance between two traits. For instance, a genetic correlation is the correlation between A

components of trait 1 and trait 2. A bivariate heritability analysis decomposes the phenotypic

correlation between trait 1 and trait 2 into A, C, and E components.

Heritability analysis was conducted on PPS scores not residualized for education, income and

family structure, as this information is used in heritability modelling. We first assessed the fit of

different decomposition scenarios (ACE, AE, CE, E), and chose the one with the lowest Akaike

Information Criterion (AIC). As AE had the lowest or close to the lowest AIC in all models, we

opted for this decomposition (Table S4). We then ran bivariate heritability analysis separately

between each PPS and BMI, which provided univariate heritability estimates of the PPSs and

BMI, genetic and environmental correlations between the univariate estimates of PPSs and BMI,

and bivariate decomposition of the phenotypic correlation between each PPS and BMI. We report

only standardized A estimates in the main results, as in the univariate and bivariate analysis of the

AE model, E = 100-A. Also, no environmental correlations were significant. All standardized and

unstandardized estimates are reported in the supplementary materials (Tables S5-S6).

Most analysis was conducted in R (69), using packages Amelia, car, cocor, cowplot, corrplot,

ggplot2, kinship2, lme4, MuMIn, pbkrtest, plyr, psych, R.matlab, rmatio, tidyr (70–84). PLS was

applied using PLS Applications (85) in MATLAB (86). Heritability analysis was conducted using

OpenMX (87), adapting scripts provided by the Colorado International Twin Workshop (88).

Mediation analysis and multiple regression analysis was conducted with the lavaan package (89)

using full information maximum likelihood for missing data.

619   **SI Results**

Control variables

Age, education, and race related to BMI, demonstrating the need for residualizing (Figure S2).

Marginal R^2 explaining only fixed effects was 0.09, and conditional R^2 explaining both fixed and

random effects was 0.35, highlighting the effect of family structure. Further, controlling for

family structure in a nested model as random intercept improved model fit (AIC dropped from

5610 to 5539), suggesting that family nesting needs to be taken into account.

Overview of PPS-obesity associations

PPS-obesity associations are presented in Figure S4. The association between the cognition PPS

and BMI was $r = 0.18$, $p<0.001$ ($R^2 = 0.03$) and the personality PPS-BMI association was $r =$

0.17, $p<0.001$ ($R^2 = 0.02$), cortical thickness PPS – BMI association was $r = 0.33$, $p<0.001$, ($R^2 =$

0.11), medial temporal lobe PPS – BMI association was $r = 0.28$, $p<0.001$ ($R^2 = 0.08$). The

associations with subcortical structures were smaller (subcortical PPS – BMI: $r = 0.07$, $p = 0.034$,

$R^2 = 0.01$). Leaving out random subsets of data did not change these correlations (Figure S4:

CV90%). The PPS-obesity associations are likely overestimated because the PPS weights were

created on and applied to same participants in current data. This was evidenced by out-of-sample

cross validation estimates being smaller (Figure S4, CV 10%). PPSs with statistically significant

single weights (Figures 1 and 2 in main text) seemed to be better generalizable to unseen data in

cross-validation (Figure S4, CV 10%). All neurobehavioural PPSs had independent effects on

BMI in a multiple regression (Figure S4, mult regr), suggesting that each factor has an

independent effect on BMI (Figure S4). The R^2 for multiple regression was 0.21. For the

heritability analysis, we used PPSs not residualized for family structure, education, or income, as

this information is needed for modelling heritability. Not residualizing for these variables

increased the PPSs association with BMI (Figure S4, Heritability) (see also 35), and also with

each other (compare Figure S3 left with S3 right).

Mediation

We sought to replicate the specific mediation model of Lavagnino et al (22), in which reduced

inhibitory control mediated the effect of lateral prefrontal cortex function on BMI. Without

control variables, the Flanker Inhibitory Control Task score had a slight negative association with

BMI (r = -0.06, p = 0.061), and Right Superior Frontal Gyrus had a positive association (r = 0.08,

p = 0.073). In the mediation model, where brain effects on BMI would be mediated by inhibition

performance, the mediation effect was not significant. When we used variables residualized for

control variables, no direct or mediated associations emerged.

<div align="center">**SI Figures**</div>

Figure S1. A schematic diagram of the analysis pipeline. All steps were conducted on all

neurocognitive factors separately. BMI = body mass index; CV = cross-validation; MTL =

medial temporal lobe; MR = multiple regression; MRI = magnet resonance image; PPS = poly-

phenotype score; SC = subcortical; SI = supporting information.

Figure S2. Regression weights of multilevel linear model nested for family. Lines mark standard 95% confidence intervals. Intercept is 27.37 (standard error: 2.16). For interpretability, regular BMI is unscaled here. Reference groups: gender – male, Race: white, Ethnicity: not Hispanic/unknown. Am. = American; BC = birth control; Is. = Islander

Figure S3. Correlations between BMI and modality poly-phenotype scores (PPS). (Left) correcting for all control variables. (Right) not correcting for education, income, and family nestedness. Correlations with $p>0.01$ are crossed out. Cogn = PPS of cognitive tests; CT = PPS of cortical thickness; MTL = PPS of medial temporal lobe volume; Pers = PPS of personality tests;

SC = PPS of subcortical structures.

Figure S4. Standardized effect sizes of poly-phenotype scores' (PPS) associations with body

mass index (BMI). Error bars mark 95% confidence intervals. Cogn = PPS of cognitive tests; CT

= PPS of cortical thickness; CV = cross-validation; Heritability = correlations using variables not

residualized for education, income, and family structure; MTL = PPS of medial temporal lobe

volume; Mult regr = multiple regression; Pers = PPS of personality tests; Regular = regular

correlations; SC = PPS of subcortical structures.

678   **SI References**

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