TITLE: Balances: a new perspective for microbiome analysis

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ABSTRACT

High-throughput sequencing technologies have revolutionized microbiome research by

allowing the relative quantification of microbiome composition and function in different

environments. One of the main goals in microbiome analysis is the identification of

microbial species that are differentially abundant among groups of samples, or whose

abundance is associated with a variable of interest. Most available methods for microbiome

abundance testing perform univariate tests for each microbial species or taxa separately,

ignoring the compositional nature of microbiome data.

We propose an alternative approach for microbiome abundance testing that consists on the

identification of two groups of taxa whose relative abundance, or balance, is associated

with the response variable of interest. This approach is appealing, since it has direct

translation to the biological concept of ecological balance between species in an ecosystem.

In this work, we present *selbal*, a greedy stepwise algorithm for balance selection.

We illustrate the algorithm with 16s abundance data from an HIV-microbiome study and a

Crohn-microbiome study.

Importance

A more meaningful approach for microbiome abundance testing is presented. Instead of

testing each taxon separately we propose to explore abundance balances among groups of

taxa. This approach acknowledges the compositional nature of microbiome data.

INTRODUCTION

Human microbiome research, focused on understanding the role in health and disease of

microbes living in the human body, has experienced significant growth in the last few years.

High-throughput sequencing technologies have revolutionized this field by allowing the

quantification of microbiome composition and function in different environments. Large

scale projects, like the Human Microbiome Project (1),(2) or MetaHIT (Metagenomics of the

Human Intestinal Tract), have established standardized protocols for creating, processing

and interpreting metagenomic data (3). However, the analysis of microbiome data for

differential abundance or association with sample metadata is still challenging.

Typically, after DNA sequencing, bioinformatics preprocessing and quality control of the

sequences, an abundance table with the number of sequences (*reads*) per sample for

different microbial species (taxa) is obtained. Total number of sequences for each sample

is highly variable, and depends on laboratory sample preparation. Indeed, raw abundances

and the total number of reads per sample are non-informative since they depend on

physical and technical mechanisms when sequencing the DNA. In order to mitigate the

problem of different sampling depth, microbiome data are often normalized previous to

differential abundance testing (4),(5).

Working with proportions, that is, relative abundances instead of raw abundances, does not

solve the problem since there is a dependence structure in the data that may lead to

misleading results such as spurious correlations or incoherent distances (6)·(7).

Rarefaction, which consists of random sampling of the same number of sequences for each

sample, is similar to working with relative abundances. Though rarefaction might be

convenient for richness and diversity analyses and avoids the problem of different sample

depth, it supposes a loss of information and the increase of Type I error for differential

abundance analyses (4). Other normalization alternatives, developed for RNA-Seq, are also

applied in microbiome analysis for dealing with the problem of different number of counts

per sample through variance stabilizing transformations (5). However, these RNA-Seq

proposals also present problems with the false discovery rate when library sizes are very

different among samples (8).

An alternative approach to rarefaction and normalization methods for microbiome analysis

is to acknowledge the compositional nature of microbiome data and to use the

mathematical theory available for compositional data (CoDa). Compositional data is defined

as a vector of strictly positive real numbers carrying relative information. Relative

information refers to the fact that the information of interest is contained in the ratios

between the components of the composition and the numerical value of each component

by itself is irrelevant (9).

As mentioned before, raw microbiome abundances are by itself non-informative since they

depend on technical artifacts such as sequencing depth. Thus, microbiome data fits the

definition of compositional data except for the fact that microbiome abundance tables

contain many zeros. Assuming that observed zeros are rounded zeros, meaning that they

correspond to values below the detection limit, they can be replaced by a positive value or

pseudo count (10) so that CoDa analysis in terms of relative abundances between groups of

microorganisms can be applied.

Several recent works acknowledge the compositional nature of microbiome abundance

data and propose their analysis accordingly (11,12). Most of these approaches consider the

centered log-ratio transformation (clr) and perform relative abundance testing for each clr

transformed component, which is given by the logarithm of the component divided by the

geometric mean of all the components in the sample. This allows the identification of clr

transformed components that are associated with a specific characteristic of interest.

However, the interpretation of such association is not straightforward because the clr

transformation involves the abundances of all the taxa in the sample.

Instead, we propose to perform microbiome relative abundance testing by identifying two

groups of taxa whose relative abundance is associated with the phenotype of interest. For

this we use the notion of balance between two groups of components of a composition,

which is a central concept in CoDa analysis.

Mathematically, a balance is defined as follows. Let $X = (X_1, X_2, \ldots, X_k)$ be a composition

of the number of counts for k different microbial species or taxa. Given two disjoint subsets

of components in X, denoted by X_+ and X_-, indexed by I_+ and I_-, and composed by k_+ and

k_- taxa, respectively, the balance between X_+ and X_- is defined as:

$$B = \sqrt{\frac{k_+ \cdot k_-}{k_+ + k_-}} \log \frac{\left(\prod_{i \in I_+} X_i\right)^{1/k_+}}{\left(\prod_{j \in I_-} X_j\right)^{1/k_-}} \; .$$

Expanding the logarithm, the balance is proportional to

$$B \; \propto \; \frac{1}{k_+} \sum_{i \in I_+} \log X_i - \frac{1}{k_-} \sum_{j \in I_-} \log X_j \; ,$$

which is a more familiar expression corresponding to the difference in means of the log-

transformed abundances between the two groups.

Balances are in compositional data analysis a key element in the construction of new

coordinates through the so called *isometric log-ratio transformation* (ilr) (13) .

The concept of balance, as proposed in the compositional data theory, provides a new and

interesting perspective for microbiome data analysis, since this mathematical concept is

related to the biological concept of ecological balance in ecosystems.

Recently, some authors have proposed the use of CoDa approaches for microbiome analysis

with different objectives such as the differential abundance between groups (14),

differentiation of niches (15), or the inclusion of phylogenetic associations between the

components included in the study (16).

In this work, we propose an algorithm for the identification of balances between groups of

taxa that are associated with a dependent component of interest. This approach provides a

new perspective to differential abundance and microbiome association studies. Starting

with the balance composed by only two taxa that is most associated with the response, the

algorithm performs a forward selection process and, at each step, a new taxon is added to

the existing balance so that the specified association criterion is maximized. The algorithm

stops when none of the possible additions improves the current association.

The paper is organized as follows. In the Results and Discussion section, the proposed

algorithm is applied to an HIV-microbiome study and to a Crohn's disease-microbiome

study. Then these results are analyzed and both the advantages and technical issues of the

algorithm when applied to microbiome data sets are discussed. Finally, in Material and

Methods we present a detailed explanation of the algorithm.

RESULTS

We illustrate the proposed methodology with a dataset from a cross – sectional HIV –

microbiome study conducted in Barcelona (Spain) including both HIV – infected subjects

and HIV – negative controls (17). Microbiome information is derived from MiSeq^{TM} 16SrRNA

sequence and bioinformatically processed with Mothur. After applying abundance filters

and agglomerating taxa to genus level, microbiome abundance is summarized in a matrix of

raw abundances for 155 samples and 60 different genera (Bioproject accession number:

PRJNA307231, SRA accession number: SRP068240). Below, we present the results for two

different analyses, the association of microbiome abundance with HIV status and with the

inflammation parameter, sCD14. In the first case, the component of interest is dichotomous

while in the second case it is continuous.

We also present the results of a Crohn's disease study (18). Only patients with Crohn's

disease (n = 662) and those without any symptom (n = 313) were analyzed. The information

was obtained from MiSeq^{TM} 16SrRNA sequence, agglomerated to the genus level, resulting

in a matrix with information of 48 genus for 975 samples. In this case, the goal is to identify

groups of taxa whose abundance balance is associated with Crohn's disease.

Microbiome and HIV status

The main goal of this analysis is to find a microbiome balance associated with HIV-status,

that is, a microbiome balance that is able to discriminate between HIV-positive and HIV-

negative individuals. As exposed in Noguera–Julian et al. (17), the HIV risk factor MSM (Men

who have Sex with Men) vs non-MSM should be considered as a possible confounder in any

HIV - microbiome study. The proposed algorithm implements a regression model which

allows adjustment for other variables. Thus, we applied the algorithm to Y=HIV-status and

X=microbiome abundance at genus level, adjusted by Z=MSM factor.

According to the cross-validation (cv) procedure implemented with function *selbal.cv*, the

optimal number of components to be included in the balance is 2 (Figure 1). The balance

we identified as the most associated with HIV-status is given by X_+, a taxon of the family

Erysipelotrichaceae and unknown genus and X_-, a taxon of the family *Ruminococcaceae*

and unknown genus (Figure 2). HIV-positive status is associated with higher balance scores,

that is, larger relative abundances of *Erysipelotrichaceae* with respect to *Ruminococcaceae*.

The discrimination accuracy of this balance is moderate, with an AUC of 0.786 on the whole

sample and a cross-validation AUC of 0.674. As can be observed in the boxplot in Figure 2,

HIV-negative individuals are associated with lower balance values, most of them negative,

while HIV-positive individuals have heterogeneous balance values. Figure 3 shows the result

of the cross – validation procedure. The balance identified with the whole dataset is the

most frequently identified in the cross-validation procedure, appearing 44% of the times,

an indicator of robustness for the proposed global balance.

Microbiome and sCD14 inflammation parameter

Acute and chronic inflammations typically occur after HIV infection. Even patients under

antiretroviral medications and undetectable viral load present chronic inflammation, which

may cause tissue damage and is associated with many chronic diseases. In this context,

there is a great interest in defining possible interventions involving modifications of the gut

bacterial environment, which may reduce inflammation in HIV patients. This requires a good

understanding of the association between gut microbial composition and several

inflammation parameters. In this case, we focus on an immune–marker related to the

chronic inflammation: the levels of soluble CD14 (sCD14), which was measured for a subset

of samples (n = 151). The optimal number of components to be included in the model is

four, according to the cv-MSE (Figure 4). The balance that is identified as the most

associated with sCD14 is composed by two taxa in the numerator, $X_+ =$

{*g_Subdoligranulum, f_Lachnospiraceae_g_unclassified*} and two in the denominator $X_- =$

{*f_Lachnospiraceae_g_Intertae_Sedis, g_Collinsella*}. The association is moderate, with R =

0.53. Figure 5 provides a scatter plot of the balance values and sCD14 values, indicating

that higher balance scores are associated with higher sCD14 values. The robustness of the

selected balance can be evaluated through the results of the cv-procedure (Figure 6). We

see that the proposed global balance is also the one that has been more frequently selected

in the cv, 34% of the times. The four taxa defining the global balance correspond to the top

4 most frequently selected in the cross - validation. These results emphasize the robustness

of the selected global balance.

Crohn's disease

Crohn's disease is an inflammatory bowel disease (IBD) linked to microbial alterations in the

gut (18)[,](19). We ran *selbal.cv* algorithm with the goal of identifying groups of taxa whose

abundance balance can discriminate between individuals with Crohn's disease from those

without the disease.

The optimal number of components in the balance is twelve according to the MSE criterion

(Figure 7). The groups defining the balance are X_+= {*g_Roseburia, o_Clostridiales_g_,*

*g_Bacteroides, f_Peptostreptococcaceae_g_*} and X_- = {*g_Dialister, g_Dorea,*

o_Lactobacillales_g_, g_Eggerthella, g_Aggregatibacter, g_Adlercreutzia, g_Streptococcus,

g_Oscillospira}. Cases with Crohn's disease have lower balance scores than controls (Figure

8) which means lower relative abundances of X_+ with respect to X_-. The discrimination

value of the identified balance is important, with an AUC = 0.838 and a cv-AUC = 0.819.

The identified global balance is very robust as the results of the cv reveal (Figure 9). The

global balance obtained with the whole dataset is also the most frequently identified

balance in the cv-procedure, namely 36% of the times. Moreover, the components defining

the global balance are also the ones more frequently selected in the cv procedure. The

balance identifies *Bacteroides* and *Clostridiales* as part of the denominator of the balance,

which have also been identified previously as less abundant in Crohn's disease individuals

than in controls (18).

DISCUSSION

The identification of individual microbial species, or taxa, that are differentially abundant

among groups of samples is challenging because the change in relative abundance of one

taxon affects the relative abundances of the other taxa. As an alternative, we propose the

analysis of relative abundances among groups of taxa instead of analyzing each taxon

separately. In this work, we present *selbal*, a greedy stepwise algorithm for balance

selection that takes into account the compositional nature of microbiome abundance data.

The algorithm identifies two groups of taxa whose relative abundance, or balance, is

associated with the response variable of interest.

selbal overcomes the problem of differences in sample size that is usually treated with

different methods based on count-normalization, rarefaction or transformation into

proportions. The only way in which data is altered in *selbal* is at the zero imputation stage

required because of the use of logarithms and ratios in the definition of balances. This

replacement of zeros by positive numbers is performed under the assumption that

observed zeros are rounded zeros, that is, all taxa are present in all the samples but some

of them are not detected because of low abundance and insufficient sample depth.

However, it is not clear how the imputation method and the presence of structural zeros

(absence of the taxa in the sample) may influence the results. Future research will be

focused on the treatment of zeros with the aim of more precisely evaluating if zeros are

rounded or structural and on selecting the best replacement method.

Due to the computational cost, *selbal* does not explore the whole balance space and the

method for selecting the optimal balance is suboptimal and may be improved. Thus,

exploring for alternative approaches in the search of the optimal balance is another topic

of future research.

In order to improve classification or prediction accuracy of the variable of interest a

prediction model with several balances can be obtained by applying *selbal* algorithm

sequentially. This sequential search of balances can be performed similarly to partial least

squares approach: when the first balance B1 is identified, all variables are deflated by the

first balance, that is, each variables is adjusted for the first balance, by regressing the

variable on B1 and taking residuals. Then, the second balance is searched on the new

orthogonalized data.

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Endorsed by the compositional treatment of microbiome abundance data, *selbal* can also

be useful for comparing different microbial studies. Since balances are based on relative

abundances among groups of taxa, this relative information is likely to remove the noise

and biases of each particular study.

MATERIALS AND METHODS

Let $X = (X_1, X_2, \ldots, X_k)$ be a composition, that is, a vector of strictly positive real numbers.

Given two disjoint subsets of components in X, denoted by X_+ and X_-, indexed by I_+

and I_-, composed by k_+ and k_- components respectively, the balance between X_+ and X_-

is defined as:

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$$B(X_+, X_-) = \sqrt{\frac{k_+ \cdot k_-}{k_+ + k_-}} \log \frac{\left(\prod_{i \in I_+} X_i\right)^{1/k_+}}{\left(\prod_{j \in I_-} X_j\right)^{1/k_-}} .$$

279 Expanding the logarithm, the balance is proportional to

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$$B(X_+, X_-) \propto \frac{1}{k_+} \sum_{i \in I_+} \log X_i - \frac{1}{k_-} \sum_{j \in I_-} \log X_j = M_+ - M_- ,$$

which corresponds to the difference of the arithmetic means of the log-transformed initial

components in the two groups that we denote by M_+ and M_-, respectively. This second

expression is preferable from a computational point of view and is the one implemented in

the proposed algorithm.

Given Y, a response variable, which can be either numeric or dichotomous, a composition

$X = (X_1, X_2, \ldots, X_k)$ and additional covariates $Z = (Z_1, Z_2, \ldots, Z_r)$, the goal of the algorithm

is to determine the sub-compositions of X, X_+ and X_-, indexed by I_+ and I_-, respectively,

so that the balance B between X_+ and X_- is highly associated with Y after adjustment for

covariates Z. Depending on the nature of the dependent variable, the association can be

defined in several ways.

For a continuous variable Y, the optimization criterion is defined as maximization of the

coefficient of determination of the linear regression model:

$$Y = \beta_0 + \beta_1 B + \gamma' Z \, .$$

For a dichotomous variable Y, we fit the logistic regression model

$$logit(Y) = \beta_0 + \beta_1 B + \gamma' Z \, ,$$

and, in this case, we consider three possible optimization criteria: the area under the ROC

curve (default option), the maximization of the explained variance (20) or the discrimination

coefficient (21).

The main function of the proposed algorithm to detect the most associated balance is called

selbal and follows these steps:

STEP 0: *Zero replacement*

The initial matrix of counts in a microbiome study, denoted by $\widetilde{X}$, typically contains zeros.

In order to apply the mathematical theory of compositional data, the observed zeros are

assumed to be non-structural zeros but a consequence of under detection limit. They are

replaced by a positive value using a Bayesian-Multiplicative replacement (10) of count zeros

as implemented in function *cmultRep()* of the R package *zCompositions* (22). It is important

to remark that this transformation keeps the information contained in the ratios between

non-zero components. The resulting matrix without zeros is denoted by X and coincides

with $\widetilde{X}$ only if the latter has no null value.

STEP 1: *Optimal balance between two components*

The algorithm evaluates exhaustively the optimization criterion for all possible balances

composed by only two components; that is, all the balances of the form:

$$B = \sqrt{\frac{1}{2}}\left(\log(X_i) - \log(X_j)\right)$$

326   for $i, j \in \{1, \ldots, k\}$ $i \neq j$. We denote by $B^{(1)}$ the optimal two- component balance in

327   terms of maximization of the association value.

328   For each pair of components (X_i, X_j) there are two options when defining a balance:

$$\sqrt{\frac{1}{2}}\left(\log(X_i) - \log(X_j)\right) \qquad \text{and} \qquad \sqrt{\frac{1}{2}}\left(\log(X_j) - \log(X_i)\right)$$

330   differenced only by their sign. For dichotomous variables, they will provide the same AUC

331   value; nevertheless *selbal* returns the balance whose coefficient in the regression model is

332   positive.

335 **STEP s:** *Optimal balance adding a new* component

337 For $s > 1$ and until the stop criterion is fulfilled, let $B^{(s-1)}$ be

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$$B^{(s-1)} \propto M_+^{(s-1)} - M_-^{(s-1)} = \frac{1}{k_+^{(s-1)}} \sum_{i \in I_+^{(s-1)}} \log(X_i) - \frac{1}{k_-^{(s-1)}} \sum_{j \in I_-^{(s-1)}} \log(X_j)$$

where $I_+^{(s-1)}$ and $I_-^{(s-1)}$ are two disjoint subsets of indices in $\{1, \ldots, k\}$, with $k_+^{(s-1)}$ and

$k_-^{(s-1)}$ elements, respectively.

For each index $p \notin (I_+^{(s-1)} \cup I_-^{(s-1)})$, the algorithm evaluates the optimization criterion

of the balance that is obtained by adding $\log(X_p)$ to $B^{(s-1)}$ including p either in $I_+^{(s-1)}$ or

in $I_-^{(s-1)}$. That is, the algorithm evaluates the optimization criterion for both, $B^{(s+)}$

and $B^{(s-)}$, defined as:

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$$B^{(s+)} = \sqrt{\frac{(k_+^{(s-1)} + 1) \cdot k_-^{(s-1)}}{k_+^{(s-1)} + k_-^{(s-1)} + 1}} \left(\frac{k_+^{(s-1)} * M_+^{(s-1)} + \log(X_p)}{k_+^{(s-1)} + 1} - M_-^{(s-1)} \right),$$

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$$B^{(s-)} = \sqrt{\frac{k_+^{(s-1)} * (k_-^{(s-1)} + 1)}{k_+^{(s-1)} + k_-^{(s-1)} + 1}} \left(M_+^{(s-1)} - \frac{k_-^{(s-1)} * M_-^{(s-1)} + \log(X_p)}{k_-^{(s-1)} + 1} \right),$$

and selects as $B^{(s)}$ the one that maximizes the optimization criterion. If $B^{(s)} = B^{(s+)}$, the

sets of new indices are defined as, $I_+^{(s)} = I_+^{(s-1)} \cup \{p\}$ and $I_-^{(s)} = I_-^{(s-1)}$, and similarly

for $B^{(s)} = B^{(s-)}$.

STOP criterion. *selbal* function has two parameters to decide the stopping criterion:

-   *th.imp,* threshold improvement (default 0). The algorithm stops the iteration

process when the improvement in association is lower than the specified threshold

improvement.

-   *maxV*, maximum number of components. The algorithm stops when the specified

maximum number of components has been included in the balance.

Cross-validation: *selbal.cv*

We perform a cross-validation procedure with two goals: (1) to identify the optimal number

of components to be included in the balance and (2) to explore the robustness of the global

balance identified with the whole dataset.

The cv procedure is implemented in the *selbal.cv* function.

For each cv process, the dataset is divided into K folds (default value, K = 5). K-1 folds are

used to obtain the balance (with *th.imp* = 0 as the stop rule) and the remaining fold is used

to test the result. The process is repeated M times (default value, M = 10)

Optimal number of components

For each combination of K and M we perform the *selbal* function on the training dataset

and find the optimal balance with c components, $c \in \{2, \ldots, C\}$ (default value, C = 20) and

evaluate the mean squared error (MSE) of the model on the test dataset. For each c we

obtain $\overline{MSE}_c$, the mean MSE of the different models with c components and the

corresponding standard error. The optimal number of components is defined with the 1se

rule, as the minimum number of components whose mean MSE is below the minimum $\overline{MSE}$

plus its standard error.

For dichotomous components, the MSE is computed in the same way codifying the two

groups as 0 and 1.

Robustness of the result

Once the optimal number of components k_{opt} has been chosen, all the balances obtained

in the cv procedure are reduced to k_{opt} components. Then, a frequency table is built both

for balances and for individual components. This information, available in the output of

selbal.cv, is summarized in a table as those shown in Figure 3, Figure 6 and Figure 9*.*

The cv process also provides the association or discrimination value for each balance in the

cv which can be used as a more accurate measure of association or discrimination of the

global model.

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Figure 1: Mean squared error (MSE) as a function of the number of components included in the balance. The optimal number of components is highlighted with a vertical dashed line.

Figure 2: The components defining the selected balance are specified on top of the boxplot that represents the distribution of the balance score for each of the groups. The right part of the figure contains the ROC–curve with its AUC value (0.786) and the density curve for each group.

Figure 3: Cross–validation (cv) results: first column contains the names of the taxa appearing in the most frequently selected balances in the cv procedure, the second column provides the frequency of selection (in percentage), the third column corresponds to the global balance, that is, the balance obtained using all the samples. Columns 4 to 6 represent the most frequent balances identified in the cv procedure. Colored rectangles indicate if the component is in the numerator of the balance (*red*), in the denominator (*blue*) or not included (*white*). The last row provides the proportion of times the balance has been selected as optimal in the cv procedure.

Figure 4: Mean squared error (MSE) as a function of the number of components included in the balance. The optimal number of components is highlighted with a vertical dashed line.

Figure 5. Representation of the balance obtained (*X* axis) for the sCD14 immune-marker

values (*Y* axis), the bacteria groups composing it (top of the figure) and the corresponding

regression line (*blue*).

Figure 6: Cross – validation (cv) results: first column contains the names of the taxa

included in the most frequently selected balances in the cv procedure, the second column

provides the frequency of selection (in percentage), the third column corresponds to the

global balance, that is, the balance obtained using the whole sample. Columns 4 to 6

represent the most frequent balances identified in the cv procedure. Colored rectangles

indicate if the component is in the numerator of the balance (*red*), in the denominator

(*blue*) or not included (*white*). The last row provides the proportion of times the balance

has been the selected in the cv procedure.

Figure 7: Mean squared error (MSE) as a function of the number of components included

in the balance. The optimal number of components is highlighted with a vertical dashed

line.

Figure 8: The components defining the selected balance are specified on top of the boxplot

which represents the distribution of the balance score for each of the groups. The right part

of the figure contains the ROC – curve with its AUC value (0.838) and the density curve for

each group.

Figure 9: Cross – validation (cv) results: first column contains the names of the taxa

appearing in the most frequently selected balances in the cv procedure, the second column

provides the frequency of selection (in percentage), the third column corresponds to the

global balance, that is, the balance obtained using the whole sample. Columns 4 to 6

represent the most frequent balances identified in the cv procedure. Colored rectangles

indicate if the component is in the numerator of the balance (*red*), in the denominator (*blue*)

or not included (*white*). The last row provides the proportion of times the balance has been

the selected in the cv procedure.

Figure 1

Figure 2

Figure 3

	%	Global	BAL 1	BAL 2	BAL 3
f_Ruminococcaceae_g_Incertae_Sedis	76	🟦	🟦	🟦	🟦
f_Erysipelotrichaceae_g_unclassified	50	🟥	🟥		
g_Bacteroides	30			🟥	
g_Phascolarctobacterium	12				🟥
FREQ	–	–	0.44	0.24	0.06

Figure 4

Figure 5

DENOMINATOR

f_Lachnospiraceae_g_unclassified

g_Collinsella

NUMERATOR

g_Subdoligranulum

f_Lachnospiraceae_g_Incertae_Sedis

Figure 6

	%	Global	BAL 1	BAL 2	BAL 3
f_Lachnospiraceae_g_unclassified	94	🟦	🟦	🟦	🟦
g_Collinsella	76	🟦	🟦	🟦	
g_Subdoligranulum	72	🟥	🟥	🟥	🟥
f_Lachnospiraceae_g_Incertae_Sedis	54	🟥	🟥		🟥
g_Thalassospira	50			🟥	
g_Bifidobacterium	14				🟦
FREQ	–	–	0.34	0.12	0.08

Figure 7

Figure 8

Figure 9

	%	Global	BAL 1	BAL 2	BAL 3
g__Dialister	100	(blue)	(blue)	(blue)	(blue)
g__Roseburia	100	(red)	(red)	(red)	(red)
o__Clostridiales_g__	98	(red)	(red)	(red)	(red)
g__Bacteroides	98	(red)	(red)	(red)	(red)
g__Dorea	96	(blue)	(blue)	(blue)	(blue)
o__Lactobacillales_g__	94	(blue)	(blue)	(blue)	(blue)
g__Eggerthella	92	(blue)	(blue)	(blue)	(blue)
g__Aggregatibacter	92	(blue)	(blue)	(blue)	(blue)
g__Adlercreutzia	90	(blue)	(blue)	(blue)	(blue)
f__Peptostreptococcaceae_g__	86	(red)	(red)	(red)	(red)
g__Streptococcus	76	(blue)	(blue)		(blue)
g__Oscillospira	72	(blue)	(blue)	(blue)	
g__Actinomyces	26			(blue)	
g__Blautia	24				(blue)
FREQ	–	–	0.36	0.1	0.1