

1 Genetic structure of the grey side-gilled sea slug  
2 (*Pleurobranchaea maculata*) in coastal waters of New  
3 Zealand

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## 25 **Abstract**

26 *Pleurobranchaea maculata* is a rarely studied species of the Heterobranchia found  
27 throughout the south and western Pacific – and recently recorded in Argentina – whose  
28 population genetic structure is unknown. Interest in the species was sparked in New  
29 Zealand following a series of dog deaths caused by ingestions of slugs containing high  
30 levels of the neurotoxin tetrodotoxin. Here we describe the genetic structure and  
31 demographic history of *P. maculata* populations from five principle locations in New  
32 Zealand based on extensive analyses of 12 microsatellite loci and the *COI* and *CytB*  
33 regions of mitochondrial DNA (mtDNA). Microsatellite data showed significant  
34 differentiation between northern and southern populations with population structure  
35 being associated with previously described regional variations in tetrodotoxin  
36 concentrations. However, mtDNA sequence data did not support such structure,  
37 revealing a star-shaped haplotype network with estimates of expansion time suggesting  
38 a population expansion in the Pleistocene era. Inclusion of publicly available mtDNA  
39 sequence from Argentinian sea slugs did not alter the star-shaped network. We interpret  
40 our data as indicative of a single founding population that fragmented following  
41 geographical changes that brought about the present day north-south divide in New  
42 Zealand waters. Lack of evidence of cryptic species supports data indicating that  
43 differences in toxicity of individuals among regions are a consequence of differences in  
44 diet.

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46 **Keywords:** *Pleurobranchaea maculata*, tetrodotoxin, microsatellites, cytochrome c oxidase  
47 subunit 1, cytochrome b, population structure

48 **Short Title:** Genetic structure of *Pleurobranchaea maculata*

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## 50 **Introduction**

51 The grey side-gilled sea slug (*Pleurobranchaea maculata*) is an opportunistic carnivore  
52 that feeds on invertebrates including sea anemones, marine worms and other molluscs  
53 [1]. It is native to New Zealand (NZ), southeastern Australia, China, Sri Lanka and  
54 Japan where it is found in habitats ranging from sandy sediments to rocky reefs, and  
55 from shallow sub-tidal flats to depths of 300 m [1, 2]. Little is known about the life  
56 history of the species but studies of comparative development report the production of  
57 planktotrophic veligers that hatch within eight days and remain planktonic for three  
58 weeks before juveniles settle [1-3].

59 In late 2009 this otherwise little-known sea slug attracted attention after it was  
60 implicated in dog deaths on beaches in Auckland [4]. Forensic analyses revealed that  
61 deaths were a consequence of tetrodotoxin (TTX) poisoning associated with ingestion  
62 of *P. maculata* [4]. This was the first time that TTX had been reported in NZ and in a  
63 species of the taxonomic clade Heterobranchia [4]. *P. maculata* that have recently  
64 invaded coastal waters of Argentina also contain TTX [5, 6].

65 TTX is a potent neurotoxin found in numerous terrestrial and marine organisms, but  
66 neither the origin of TTX nor the causes of variation in TTX levels among species are  
67 understood. The structure of TTX suggests a microbial origin [7] and while certain  
68 microbes have been implicated in TTX production (reviewed in Magarlamov *et al*,  
69 2017), all such claims have been refuted [8, 9]. Nonetheless, while not excluding a  
70 microbial origin, there is recognition that TTX in animals is often acquired via diet. For  
71 example, variability in TTX levels found in puffer fish has been attributed to exposure  
72 to toxic food sources (reviewed in Noguchi and Arakawa, 2008). For *P. maculata*, there  
73 is mounting evidence that toxin accumulation occurs through feeding [10-13]. An

74 alternate possibility is that TTX arises from commensal or symbiotic microorganisms  
75 that are associated with *P. maculata* [13], but no TTX-producing bacteria have been  
76 found [14, 15].

77 Studies of individual and temporal differences in TTX concentration has established  
78 that *P. maculata* populations from northern regions of the North Island (Whangarei,  
79 Auckland, Tauranga) have high TTX concentrations (the highest average being 368.7  
80 mg kg<sup>-1</sup> per individual), while populations from the South Island (Nelson and Kaikoura)  
81 have trace amounts of TTX (<0.5 mg kg<sup>-1</sup>) or none at all [2, 4, 10, 13]. A recent study  
82 reported TTX concentrations as high as 487 mg kg<sup>-1</sup> [13]. Significant individual and  
83 seasonal variations have also been observed [2]. A single individual obtained from  
84 Wellington in the south of the North Island was found to have a low concentration of  
85 TTX (2.2 mg kg<sup>-1</sup>) supporting the notion of a geographical cline [2].

86 New Zealand (NZ) extends from the subtropical Kermadec Islands (29° S) to the sub  
87 Antarctic Campbell Island group (52° S), it comprises ~700 islands affording ~15,000  
88 km of coastline [16, 17]. Subtropical and sub-Antarctic waters converge at the juncture  
89 between the two major landmasses (the North and South Islands) creating a complex  
90 oceanography of currents and eddies.

91 Phylogenetic and population genetic studies of NZ marine organisms have shown that  
92 different species manifest a range of genetic structures (reviewed in Gardner *et al*, 2010)  
93 with a common pattern being a disjunction between northern and southern populations.  
94 Such disjunctions are evident in bivalves [19], teleosts [20], polyplacophores [21] ,  
95 echinoderms [22] and arthropods [23].

96 The genetic structure of *P. maculata* is unknown, but variation in the established  
97 differences in toxicity between northern and southern populations suggests that

98 geographic variability in TTX concentration correlates genetic structure – even the  
99 possibility that northern and southern populations define separate species. Here we test  
100 this hypothesis using a combination of microsatellite and mitochondrial DNA (mtDNA)  
101 sequence markers. Analysis of more rapidly evolving microsatellites showed evidence  
102 of a genetic break along the predicted north-south divide. However, no such structure  
103 was apparent from analysis of mtDNA data.

## 104 **Materials and methods**

### 105 **Sampling, DNA extraction and tetrodotoxin assay**

106 A total of 156 samples were collected from nine regions around New Zealand between  
107 2009 and 2013 (Fig 1 and S1 Table). DNA was extracted as described in Yıldırım *et al.*  
108 [24]. The Tauranga (TR) population included samples from Tauranga Harbour whereas  
109 the Auckland (AKL) population included samples from Tamaki Strait and Waitemata  
110 Harbour. Some samples were from the studies of Wood *et al.* [2] and Khor *et al.* [10].

#### 111 **Fig 1. Sampling locations for the *Pleurobranchaea maculata* individuals.**

112 The numbers within the circles indicate the sampling size of each region. The arrows  
113 show magnified maps of Auckland and Tauranga. Populations containing *P. maculata*  
114 individuals with high, and low and trace amounts of tetrodotoxin concentrations in red  
115 and blue colour, respectively.

116 At the outset of this study there was limited knowledge of the toxicity of *P. maculata*  
117 individuals from Wellington (WL) as only one individual had been previously tested  
118 [2]. To obtain a better understanding, the TTX concentration of eight (of eighteen)  
119 individuals collected from WL in October 2012 was determined as described in Khor et

120 al. [10]. The TTX assay was performed at the Cawthron Institute (Nelson) using a liquid  
121 chromatography-mass spectrophotometry method that is described in McNabb et al. [4].  
122 Population-level analyses were performed only for five populations which are Ti Point  
123 (TP), AKL, TR, Wellington (WL) and Nelson (NL) due to the small sample sizes of the  
124 other locations. TP, AKL and TR, which included highly toxic individuals [10] were  
125 designated as the “northern cluster”, whereas the WL and NL population, which  
126 contained slightly toxic and non-toxic individuals [2, 10, 11] were designated as the  
127 “southern cluster”.

## 128 **Genotyping**

129 Twelve microsatellite markers (*Pm01, 02, 07, 08, 09, 10, 11, 13, 17, 19, 20* and *23*) [24]  
130 were genotyped for 149 samples. PCR amplification and genotyping procedures for the  
131 primers were as described in Yıldırım *et al.* [24] with some modifications (S2 Table).  
132 Details regarding amplification and genotyping processes are described in the  
133 Supporting Information.

134 A 1060 bp and 1153 bp region of mitochondrial cytochrome B (*CytB*) and cytochrome  
135 oxidase I (*COI*) genes, respectively, were amplified and sequenced in all 156 *P.*  
136 *maculata* individuals. For details regarding the primer pairs and amplification see  
137 Supporting Information and S3 Table. Geneious Pro 6.1.6 (Biomatters, New Zealand)  
138 was used to trim, assemble, align and concatenate the resulting DNA sequences.

## 139 **Statistical analysis**

### 140 **Genetic Diversity**

141 Microsatellite genotyping data were tested for scoring errors due to stuttering, null  
142 alleles, and large allele dropout using MICRO-CHECKER v.2.2 [25]. Departures from

143 Hardy-Weinberg equilibrium (HWE) were estimated using Nei's [26] inbreeding  
144 coefficient  $G_{IS}$  with 100,000 permutations using GenoDive v2.0b25 [27]. FreeNA [28]  
145 was used to estimate null allele frequency as in Dempster *et al.* [29]. Interference of  
146 putative null alleles on genetic differentiation between sampling sites was determined  
147 by calculating global and pair-wise  $F_{ST}$  values [30], either with or without "exclusion of  
148 null alleles" [28]. Linkage disequilibrium (LD) between pairs of loci was estimated in  
149 FSTAT v2.9.3.2 [31] and the significance of LD was determined by applying a  
150 Bonferroni correction. The total number of alleles ( $N_a$ ), allele frequencies, observed  
151 heterozygosity ( $H_o$ ), unbiased expected heterozygosity ( $H_e$ ) corrected for a small  
152 sampling size [26], and private alleles ( $PA$ ) per locus and population were calculated  
153 using GenAlex [32]. Allelic richness ( $A_R$ ) was calculated using the rarefaction method  
154 implemented in ADZE v.1.0 [33].  $H_e$  and  $A_R$  were used to compare the amount of  
155 genetic diversity among populations from different regions using one-way ANOVA  
156 (<http://vassarstats.net>).

157 For mtDNA, several estimates of genetic diversity, including the number of singletons  
158 ( $Sin$ ), haplotypes ( $Hap$ ) and segregating sites ( $S$ ), the average number of nucleotide  
159 differences between sequences ( $k$ ) [34], haplotype ( $h$ ) and nucleotide diversity ( $\pi$ ) [35]  
160 were calculated for the *CytB*, *COI* and concatenated sequences for each sampling  
161 location using DnaSP 5.10.0.1 [36].

## 162 **Population Structure**

163 For microsatellite data, global differentiation and pairwise differentiation between each  
164 pair of populations was investigated using various differentiation estimators, including a  
165 log-likelihood ratio (G)-based test [37], fixation index  $F_{ST}$  [30], standardised fixation  
166 index  $G''_{ST}$  [38], and Jost's [39] differentiation ( $D_{est}$ ). The statistical power to detect

167 true population differentiation and  $\alpha$ -error probability were assessed in POWSIM v4.1  
168 [40]. STRUCTURE v2.3.4 [41] was used to determine the probable number of distinct  
169 populations ( $K$ ) and individuals were assigned to populations using a Bayesian  
170 assignment approach. Parameters were set to 5,000,000 MCMC iterations with a burn-  
171 in of 500,000 values of  $K$  between one and ten, with a series of ten independent  
172 replicates for each  $K$  value, assuming an admixture model and correlated allele  
173 frequencies across the populations, both with and without introducing *a priori* sampling  
174 location. The most likely value of  $K$  was resolved using the  $\Delta K$  method [42] with the  
175 Structure Harvester v0.6.93 [43], and the results were introduced to the CLUMPP  
176 v1.1.2 software [44]. Destruct v1.1 [45] was used to visualize the results.

177 AMOVA [46] was performed in ARLEQUIN to determine the hierarchical genetic  
178 structure based on  $R_{ST}$  [47] and  $F_{ST}$  [30] for microsatellite data. Two separate  
179 calculations were performed: one with structure and one without structure. For the  
180 former, individuals from all regions were grouped; for the latter, the TP, AKL and TR  
181 populations were grouped into a northern cluster and the WL and NL populations were  
182 grouped into a southern cluster. This nested design was based on the results of  
183 population structure suggested by F-statistics, multivariate and STRUCTURE analyses  
184 of microsatellite data.

185 For the mtDNA sequences, POPART v1 (<http://www.leigh.net.nz/software.shtml>) was  
186 used to create a median joining haplotype network (MJN) [48]. We created an  
187 additional MJN for shorter COI sequences (624 bp) in order to accommodate four *P.*  
188 *maculata* samples obtained from individuals isolated in Argentina (Farias et al., 2016).  
189 A saturation test was performed in DAMBE v6.2.9 [49] using the test by Xia et al. [50].  
190 The proportion of invariant sites ( $P_{inv}$ ) was estimated by Jmodeltest v0.1.1 [51] with the



191 Akaike information criteria (AIC). The  $P_{inv}$  values (0.844 and 0.789 for *CytB* and *COI*,  
192 respectively) obtained from the most likely models suggested by the software (HKY+I  
193 and TrN+I for *CytB* and *COI*, respectively) and default settings for other parameters  
194 were used for the calculations. Haplotype-frequency-based  $F_{ST}$  [26] and distance-based  
195  $\Theta_{ST}$  [46] were calculated in ARLEQUIN to estimate population differentiation. For  $\Theta_{ST}$ ,  
196 the Tamura-Nei mutational model [52] was used for both genes as being the closest  
197 models to the ones suggested as most likely to explain mtDNA data by Jmodeltest. For  
198 concatenated sequences hierarchical structure was investigated by AMOVA in  
199 ARLEQUIN using both F-statistics and  $\Phi$ -statistics, and with the same structure pattern  
200 used for the microsatellite data analysis (TP, AKL and TR were grouped into a northern  
201 cluster, and WL and NL were grouped into a southern cluster).

202 Patterns of differentiation were also analysed using a multivariate approach. For  
203 microsatellite data, the Manhattan distance ( $DM$ ), which calculates the mean character  
204 differences between individuals, and clonal distances ( $DCL$ ), which assumes a stepwise  
205 mutational model (SMM) [46], were used. For mtDNA data, a distance matrix ( $D_{SEQ}$ )  
206 was calculated as a standardized bp difference between every pair of individuals.

207 Statistical analyses on resulting distance matrices were done using PRIMER v6 [53]  
208 with PERMANOVA+ [54]. Patterns of inter-sample distances were visualized using  
209 non-metric multi-dimensional scaling ordination (MDS) [55]. Permutational  
210 multivariate analysis of variance (PERMANOVA) [56, 57] was used to formally test for  
211 differences in genetic structures among different locations and canonical analysis of  
212 principal coordinates (CAP) [54, 58] was used to discriminate among specific  
213 populations and identify their distinctiveness, using leave-one-out allocation success.  
214 PERMDISP was used to test the null hypothesis of homogeneity of within-group  
215 dispersions among populations [59]. All permutation tests used 10,000 permutations.

216 A maximum likelihood (ML) tree using the Tamura-Nei mutational model [34] with  
217 default settings was reconstructed for 44 *P. maculata* individuals from this study and  
218 three Argentinian *P. maculata* individuals using COI sequences [5] (redundant  
219 sequences were removed) in MEGA7 [60]. COI sequences of five species from the  
220 same family (Pleurobranchidae, Genbank codes in brackets) including *Pleurobranchaea*  
221 *meckeli* (KU365727.1), *Pleurobranchaea nevaezelandiae*, *Pleurobranchus peronii*  
222 (KM521745.1), *Berthella ocellata* (KM521694.1) and *Berthellina citrina*  
223 (KM521694.1) were used as outgroups. The analysis involved 200 informative  
224 positions of 616. The phylogeny was tested with 1,000 bootstrap replicates.

## 225 **Migration**

226 The microsatellite data were analysed with GeneClass2 [61] to identify the first-  
227 generation migrants using the Bayesian criterion of Rannala and Mountain [62] and the  
228  $L_{\text{home}}/L_{\text{max}}$  likelihood, with a threshold  $p$ -value of 0.01 and a Monte-Carlo resampling  
229 algorithm [63].

## 230 **Neutrality tests**

231 BOTTLENECK v1.2 [64] was used to test the possibility of recent population reduction  
232 for microsatellite data assuming SMM and two-phase models (TPM) with default  
233 settings using a Wilcoxon signed rank test [64]. A possible sign of a recent bottleneck  
234 was investigated also by a mode-shift analysis [65].

235 Deviations from neutrality and demographic changes within and across the populations  
236 were calculated with Tajima's  $D$  [66], Fu's  $F_s$  [67] and mismatch distribution analysis  
237 in ARLEQUIN for the concatenated mtDNA sequences. The null hypothesis of  
238 expansion was statistically tested with the sum of squared deviations ( $SSD$ ) from the  
239 expected values [68] and Harpending's *raggedness* index [69]. Mismatch distribution

240 analysis also estimated tau ( $\tau$ ), which is the age of the population expansion. The  
241 approximate date of population expansion across all samples was calculated for *COI* as  
242 outlined by Schenekar and Weiss [70] using  
243 <http://www.unigraz.at/zoowww/mismatchcalc/index.php> by converting  $\tau$  to time since  
244 expansion ( $t$ ) in years using the formula  $t = \tau / 2\mu k$ , where  $\mu$  is the mutation rate per site  
245 per generation, and  $k$  is the sequence length [71]. One year of generation time and both  
246 5.3% divergence/mya (average  $\mu$  for marine invertebrate COI sequences) [72], a  
247 compatible mutation rate estimated for a planktotrophic heterobranchia species  
248 *Costasiella ocellifera* (Ellingson and Krug, 2015), were assumed. McDonald and  
249 Kreitman's [73] neutrality test was performed pooling all *P. maculata* COI sequences  
250 (1153 bp) in DnaSP using *P. meckeli* COI sequences as an outgroup species. Fisher's  
251 exact test (two-tailed) was used to identify significant deviations from neutrality.

## 252 **RESULTS**

### 253 **Tetrodotoxin levels in *P. maculata* from Wellington**

254 Previous analyses have established that northern WH, AKL, TR, and CR populations  
255 have high levels of TTX, marking these populations as “toxic”, while southern  
256 populations from NL and KK are recorded as containing either trace, or no TTX [2, 4,  
257 10, 13]. For WL populations, previous measurements existed for only one individual  
258 documented as having a low level of TTX (2.2 mg/kg) [2]. For this study, 18 slugs were  
259 obtained from WL of which eight randomly chosen individuals were subject to TTX  
260 assay. Three contained extremely low concentrations (0.12, 0.16 and 0.5 mg/kg) of  
261 TTX. No TTX was detected in the remaining five individuals. Accordingly, the WL, NL  
262 and KK samples (the southern cluster) were classified as “non-toxic”.

## 263 Genetic diversity

### 264 Microsatellite analyses

265 All loci were highly polymorphic with between five and 23 alleles for each locus  
266 (diversity statistics in Table 1 and S4 Table).  $H_e$  across populations ranged from 0.407  
267 to 0.843, with an average of 0.665. Rarefaction curves for  $A_R$  across each locus levelled  
268 off for each sampling location indicating that a reasonable portion of the available  
269 allelic diversity was sampled at each location (S1 Fig; allele frequencies are reported in  
270 S5 Table). Populations did not exhibit significant differences in genetic diversity for  
271 either  $A_R$  ( $F_{4,146}=0.0048$ ,  $P=1.000$ ) or  $H_e$  ( $F_{4,146}=1.102$ ,  $P=0.358$ ). No significant linkage  
272 disequilibrium was found after Bonferroni correction ( $P<0.05$ ) (S6 Table). Populations  
273 met Hardy-Weinberg expectations (MICROCHECKER, Table 1) with few exceptions  
274 (S4 Table). There was a low level of null allele frequency ( $\leq 10\%$ ) for all populations at  
275 all loci (FreeNA, S7 Table), and there were only slight changes in global and pairwise  
276  $F_{ST}$  values between populations for each locus after accounting for null alleles (S8  
277 Table). Original allele frequencies were therefore used in subsequent analyses.

278

279

280 **Table 1. Summary of the genetic diversity statistics at microsatellite loci across five**  
281 **locations.**

Locus	Na	Size (bp)	$H_o$	$H_e$	$G_{IS}$	$F_{ST}$	$G''_{ST}$	Dest
<i>Pm01</i>	23	108–208	0.842	0.843	0.000	0.057 <sup>c</sup>	0.328 <sup>c</sup>	0.291 <sup>c</sup>
<i>Pm02</i>	9	105–137	0.671	0.742	0.090 <sup>a</sup>	0.026	0.044	0.033
<i>Pm07</i>	10	128–164	0.737	0.720	-0.024	0.014	-0.007	-0.005
<i>Pm08</i>	6	141–165	0.710	0.660	-0.079	0.115 <sup>c</sup>	0.365 <sup>c</sup>	0.274 <sup>c</sup>
<i>Pm09</i>	16	91–142	0.733	0.736	0.004	0.045 <sup>c</sup>	0.142 <sup>c</sup>	0.108 <sup>c</sup>
<i>Pm10</i>	6	107–122	0.737	0.699	-0.055	0.071 <sup>c</sup>	0.233 <sup>c</sup>	0.175 <sup>c</sup>
<i>Pm11</i>	11	157–187	0.838	0.813	-0.032	0.035 <sup>b</sup>	0.137 <sup>b</sup>	0.114 <sup>b</sup>
<i>Pm13</i>	8	103–136	0.572	0.576	0.007	0.007	-0.024	-0.014

<b><i>Pm17</i></b>	8	155–187	0.457	0.452	-0.011	0.058 <sup>b</sup>	0.097 <sup>b</sup>	0.046 <sup>b</sup>
<b><i>Pm19</i></b>	5	169–181	0.523	0.519	-0.008	0.041 <sup>b</sup>	0.067 <sup>b</sup>	0.036 <sup>b</sup>
<b><i>Pm20</i></b>	5	114–132	0.407	0.407	-0.001	0.181 <sup>c</sup>	0.340 <sup>c</sup>	0.174 <sup>c</sup>
<b><i>Pm23</i></b>	14	156–184	0.703	0.692	-0.016	0.132 <sup>c</sup>	0.467 <sup>c</sup>	0.378 <sup>c</sup>
<b>Ave</b>	6.667		0.661	0.655	-0.009	0.064 <sup>c</sup>	0.175 <sup>c</sup>	0.122 <sup>c</sup>
<b>SE</b>	0.428		0.020	0.019		0.014	0.046	0.0352

282 <sup>a</sup> Significant deviation from HWE ( $P<0.05$ ). Significant genetic differentiation: <sup>b</sup> ( $P<0.01$ ), <sup>c</sup>  
283 ( $P<0.001$ ).

## 284 **Mitochondrial DNA analyses**

285 The basic diversity values for *COI* and *CytB* sequences are presented in Table 2. The  
286 total number of variable sites is 173 (*COI*: 105; *CytB*: 68), 98 of which are singleton  
287 mutations (*COI*: 59; *CytB*: 39); together these defined 130 distinct haplotypes for the  
288 concatenated sequences (*COI*: 103, *CytB*: 74). In contrast to the range of values  
289 obtained for  $h$  (0.922-0.989), the corresponding ranges for  $k$  (2.9-4.8) and  $\pi$  (0.254-  
290 0.414%) were low for both genes. The mean values for  $k$  and  $\pi$  are 7.30 (*COI*:  
291  $3.81\pm 0.018$ , *CytB*:  $2.75\pm 0.018$ ) and 0.330 % (*COI*: 0.381%, *CytB*: 0.275%), respectively  
292 for the concatenated sequences. Similar diversity was observed between sampling  
293 locations.

294 **Table 2. Summary of the genetic diversity statistics and bottleneck analysis for individuals sampled from five locations.**

Pop	N	Microsatellite					COI					CytB						
		Na	AR	Ho	He	GIS	S	Sin	Hap	h±SD	k	π±SD%	S	Sin	Hap	h±SD	k	π±SD%
<b>Ti Point</b>	20	6.08	6.02	0.679	0.680	0.001	27	16	17	0.984±0.020	4.8	0.414±0.048	17	9	16	0.957±0.021	3.1	0.294±0.041
<b>Auckland</b>	30	6.67	5.88	0.681	0.667	-0.021	39	31	26	0.989±0.013	4.1	0.352±0.033	23	15	18	0.922±0.034	2.7	0.254±0.043
<b>Tauranga</b>	33	6.88	5.89	0.697	0.676	-0.032	39	27	27	0.981±0.015	4.5	0.390±0.045	29	20	21	0.936±0.032	2.8	0.263±0.041
<b>Wellington</b>	18	6.00	6.00	0.606	0.636	0.047	25	17	16	0.987±0.023	4.2	0.363±0.033	18	10	13	0.954±0.034	3.4	0.317±0.046
<b>Nelson</b>	45	7.75	5.93	0.641	0.616	-0.041	44	33	30	0.957±0.018	4.0	0.348±0.035	35	23	29	0.941±0.063	2.9	0.276±0.032
<b>Total</b>	146	121	-	0.661	0.655	-0.009	105	59	103	0.980±0.005	4.4	0.381±0.018	68	39	74	0.945±0.011	2.9	0.275±0.018

295 N: sample size, Na number of alleles, AR: mean allelic richness based on 18 diploid individuals, Ho observed heterozygosity, He: unbiased expected  
 296 heterozygosity, GIS: inbreeding coefficient. S: number of segregating sites, Sin: singleton mutations, Hap: number of haplotypes detected, h: haplotype  
 297 diversity, k: number of pairwise nucleotide differences, π: nucleotide diversity, SD: standard deviation.

## 298 **Genetic Structure**

### 299 **Microsatellite analyses**

300 Global genetic differentiation, estimated using  $F_{ST}$ ,  $G''_{ST}$  and  $Dest$ , was low  
301 to moderate: 0.064, 0.175 and 0.122, respectively, and highly significant  
302 ( $P \leq 0.0001$ ). Differentiation was significant for most loci (Table 2). Location  
303 had a significant effect on population structure (PERMANOVA,  $DM$ :  
304  $F_{4,146}=7.3914$ ,  $DCL$ :  $F_{4,146}=9.8256$ ,  $P=0.0001$ ). Pairwise comparisons of  
305 genic ( $\chi^2=\infty$ , d.f.=24,  $P<0.0001$ ) and genetic differentiation ( $F_{ST}$ :  
306 0.074-0.122,  $G''_{ST}$ : 0.216-0.337,  $Dest$ : 0.153-0.246) as well as  
307 PERMANOVA tests ( $t=2.54-4.69$ ) showed that southern populations (WL  
308 and NL) were significantly differentiated from northern populations (TP,  
309 AKL and TR) ( $P<0.0001$ , Table 3). Most loci supported this pairwise  
310 differentiation pattern among populations ( $F_{ST}$  and  $Dest$  values in S8 and S9  
311 Tables, respectively). None of the estimators suggested significant  
312 differentiation among the northern populations of TP, AKL and TR.  $F_{ST}$  and  
313 PERMANOVA comparisons suggest weak but significant differentiation  
314 between the WL and NL populations ( $P<0.046$ ). Analysis of statistical  
315 power by POWSIM showed a 100% probability of detecting population  
316 differentiation at an  $F_{ST}$  value of 0.01. The probability of  $\alpha$  error was  $\sim 0.05$   
317 ( $P=0.04$  and  $0.057$  based on chi-square and Fisher methods, respectively),  
318 suggesting a low risk for Type I error. The differentiation pattern was  
319 therefore considered real and reinforced by  $F_{ST}$  values between the

320 significantly differentiated populations being for the most part greater than  
 321 0.01. Additionally, PERMDISP analysis showed no significant differences  
 322 in dispersion for either *DM* ( $F_{4,141}=0.1243$ ,  $P=0.1243$ ) or *DCL* ( $F_{4,141}=$   
 323  $0.4856$ ,  $P=0.4856$ ), meaning that average nucleotide distances from  
 324 individuals to their own population centroid did not differ among the groups  
 325 (i.e. that within-group genetic variability did not differ among populations).

326 **Table 3. Pairwise population differentiation estimates and associated**  
 327 **tests across five populations.**

Groups	<i>Microsatellite data</i>						<i>COI data</i>			<i>CytB data*</i>	
	Genic	$F_{ST}$	$G''_{ST}$	<i>Dest</i>	<i>t-DM</i>	<i>t-DCL</i>	<i>t</i>	$F_{ST}$	$\Theta_{ST}$	$F_{ST}$	$\Theta_{ST}$
TP-AKL	17.26	-0.009	-0.027	-0.019	0.699	0.766	1.585 <sup>a</sup>	0.014 <sup>b</sup>	0.024	0.050 <sup>c</sup>	0.028
TP-TR	29.67	-0.004	-0.012	-0.008	0.727	0.79	1.359	0.017 <sup>b</sup>	0.018	0.043 <sup>b</sup>	0.023
TP-WL	Inf <sup>c</sup>	0.087 <sup>c</sup>	0.256 <sup>c</sup>	0.185 <sup>c</sup>	2.851 <sup>c</sup>	2.711 <sup>c</sup>	1.233	0.014 <sup>a</sup>	0.008	0.033 <sup>c</sup>	-0.026
TP-NL	Inf <sup>c</sup>	0.118 <sup>c</sup>	0.330 <sup>c</sup>	0.241 <sup>c</sup>	4.160 <sup>c</sup>	4.083 <sup>c</sup>	1.494	0.030 <sup>b</sup>	0.029	0.040 <sup>b</sup>	0.000
AKL-TR	20.48	-0.004	-0.011	-0.007	0.983	0.655	1.712 <sup>a</sup>	0.015 <sup>c</sup>	0.020	0.071 <sup>c</sup>	0.019
AKL-WL	Inf <sup>c</sup>	0.095 <sup>c</sup>	0.274 <sup>c</sup>	0.198 <sup>c</sup>	2.548 <sup>c</sup>	2.958 <sup>c</sup>	1.202	0.012 <sup>a</sup>	0.014	0.062 <sup>c</sup>	-0.003
AKL-NL	Inf <sup>c</sup>	0.122 <sup>c</sup>	0.337 <sup>c</sup>	0.246 <sup>c</sup>	3.677 <sup>c</sup>	4.693 <sup>c</sup>	1.283	0.028 <sup>c</sup>	0.005	0.068 <sup>c</sup>	-0.001
TR-WL	Inf <sup>c</sup>	0.074 <sup>c</sup>	0.216 <sup>c</sup>	0.153 <sup>c</sup>	2.541 <sup>c</sup>	3.035 <sup>c</sup>	1.403	0.016 <sup>a</sup>	0.017	0.056 <sup>c</sup>	0.012
TR-NL	Inf <sup>c</sup>	0.097 <sup>c</sup>	0.272 <sup>c</sup>	0.194 <sup>c</sup>	3.725 <sup>c</sup>	4.661 <sup>c</sup>	1.076	0.031 <sup>c</sup>	0.007	0.061 <sup>c</sup>	0.026 <sup>a</sup>
WL-NL	53.20 <sup>b</sup>	0.008 <sup>a</sup>	0.021	0.013	1.235 <sup>a</sup>	1.593 <sup>a</sup>	1.326	0.029 <sup>b</sup>	0.020	0.053 <sup>c</sup>	-0.014

328  $\chi^2$  values for the homogeneity of allele frequencies in pairwise comparisons tested  
 329 with the exact *G*-test (Genic). Pseudo *t* statistic (*t*) from pairwise PERMANOVA  
 330 tests on the basis of either Manhattan (*DM*) or Clonal distances (*DCL*) for  
 331 microsatellite data, or standardized nucleotide differences for mtDNA data. Other  
 332 estimators used were fixation index  $F_{ST}$ , standardised fixation index  $G''_{ST}$ , Jost's  
 333 differentiation (*Dest*) and distance-based  $\Theta_{ST}$ .  
 334 TP: Ti Point, AKL: Auckland, TR: Tauranga, WL: Wellington, NL: Nelson.  
 335 Significant differentiation: <sup>a</sup>  $P<0.05$ , <sup>b</sup>  $P<0.0005$ , <sup>c</sup>  $P\leq 0.0001$ . \* Pairwise  
 336 comparisons using pseudo *t* were not done for *CytB* data because the overall test  
 337 was not statistically significant.  
 338



339 A north-south differentiation is evident from MDS ordinations of the *DM*  
340 and *DCL* matrices (Figs 2A and B). Note that although stress is relatively  
341 high (0.23 and 0.24) for the 2-dimensional MDS ordinations, the north-  
342 south distinction was also very clearly apparent in the corresponding 3-  
343 dimensional MDS ordinations (not shown here), which had lower stress  
344 (0.18 in both cases). Discriminant analysis (CAP, S2 Fig) for the two-group  
345 north-south split was highly significant ( $P < 0.001$ ), with the CAP model  
346 showing a leave-one-out allocation success of 97.26% (142/146 for *DCL*  
347 and also for *DM*). In contrast, the five-group CAP models (TP, AKL, TR,  
348 WL and NL) were poor, having overall allocation success of only 43.15%  
349 (63/146 for either *DCL* or *DM*). Furthermore, there was no discrimination  
350 among the three northern locations (TP, AKL and TR) for either *DCL* or  
351 *DM* (CAP,  $P > 0.76$  in both cases), justifying their pooling into a single  
352 group. Similarly, there was no basis for discriminating the two southern  
353 populations (WL and NL) using CAP models ( $P > 0.11$  for *DCL* and *DM*).  
354 Bayesian clustering of individuals based on allele frequencies as  
355 implemented by STRUCTURE showed a  $\Delta K$  value and mean log  
356 probabilities of data ( $\ln P(x/K)$ ) that were maximal at  $K=2$  (Fig 2C), again  
357 supporting the same two distinct north-south clusters (Figs 2D and E). This  
358 finding was not affected by inclusion of sampling locations as priors (data  
359 not shown).

360 **Fig 2. Visualization of genetic structure among *Pleurobranchaea***  
361 ***maculata* populations based on microsatellite data.**

362 All individuals from eight sampling locations were included. MDS  
363 ordination of pairwise (A) Manhattan (*DM*) and (B) Clonal (*DCL*) distances  
364 between individuals. Bayesian clustering analysis where the sampling  
365 location was not introduced for the calculations, (C) Plot of  $\Delta K$  versus  $K$   
366 indicating that data are best explained by  $K = 2$  clusters, (D) Population  
367 structure at  $K=2$ . Each individual is represented by a vertical line divided  
368 into two segments, which indicates proportional membership in the two  
369 clusters; (E) Group assignments, indicating proportional membership in  
370  $K=2$  clusters.

371 AMOVA results are shown in S10 Table. Analysis performed assuming no  
372 structure revealed that the highest proportion of the total variation stemmed  
373 from variation among individuals (88.3%) with a fixation index ( $R_{IT}$ ) value  
374 of 0.117. Variation between populations ( $R_{ST}=0.137$ ) constituted 8.02% of  
375 the total variation with high significance ( $P=0.0000$ ). Variation between  
376 individuals within populations did not contribute to genetic variation (-  
377 1.16%) ( $R_{IS}=-0.020$ ,  $P=0.700$ ). AMOVA analysis was repeated by  
378 clustering the populations into northern and southern clusters. The results  
379 show that the variation between the northern and southern clusters explains  
380 20.0% ( $R_{CT}=0.201$ ) of the total variation with high significance ( $P<0.0001$ ).  
381 Variation between populations within clusters explained only 0.1% of the

382 total variation ( $R_{SC}=0.001$ ,  $P=0.3418$ ). AMOVA analysis based of  $F_{ST}$   
383 yielded similar results (data not shown). These results show that total  
384 variation for the NZ *P. maculata* populations is best explained by  
385 differences in the distribution of variance between northern and southern  
386 clusters.

### 387 **Mitochondrial DNA analyses**

388 The sequence data for 156 individuals obtained from *COI* and *CytB* were  
389 submitted to GenBank (accession numbers: KY094153 - KY094309 for  
390 *COI* and KY094310 - KY094466 for *CytB*). MJN analysis of *CytB* (Fig 3)  
391 and *COI* (S3 Fig ) genes resulted in highly similar patterns. Thus, for  
392 simplicity, the *CytB* network was used to infer evolutionary relationships  
393 among individuals. The network shows a lack of noticeable geographical  
394 structure; common haplotypes are shared across populations. The two most  
395 common *CytB* haplotypes are shared by 25 (16.0% of the total dataset) and  
396 22 (14.1% of the total dataset) individuals (frequencies in S11 Table). The  
397 network is complex and reticulated with a star-like topology; many private  
398 haplotypes descend from central shared nodes with mostly one to two base  
399 pair distances [74]. Private haplotypes were detected in all sampling  
400 locations. Characteristics of the network showed little change on inclusion  
401 of four recently published Argentinian samples (Farias *et al*, 2015; Farias *et*  
402 *al*, 2016) (based on a slightly shorter *COI* fragment (see Materials and  
403 Methods and S3B Fig). Argentinian samples are closely related to NZ

404 samples with just a single base pair distance from a commonly shared  
405 haplotype. The index of substitution saturation (ISS) was used to test for  
406 homoplasy due to multiple substitutions [50]. For both symmetrical and  
407 asymmetrical tree topology models and for both genes, the observed ISS  
408 values are significantly larger than the critical *ISS* (*ISS.c*) values (S12  
409 Table), which indicate that the paired partitions are not saturated, and the  
410 degree of homoplasy is low.

411 **Fig 3. Median joining network of the *CytB* haplotypes.**

412 The network is coloured according to the sampling locations. The diameter  
413 of the circles reflects relative haplotype frequencies. The hashes indicate the  
414 mutational steps between the haplotypes. The black nodes represent the  
415 imaginary haplotypes necessary to create a bridge between the present  
416 haplotypes.

417 Analysis based on  $F_{ST}$  values showed a moderate but highly significant  
418 global genetic differentiation for both *COI* ( $F_{ST}=0.022$ ,  $P=0.0001$ ) and *CytB*  
419 ( $F_{ST}=0.057$ ,  $P=0.0001$ ). However  $\Phi_{ST}$  values were significant for *COI* ( $\Phi_{ST}$   
420 = 0.01552,  $P = 0.022$ ), but not for *CytB* ( $\Phi_{ST}=0.00845$ ,  $P=0.12$ ). Pairwise  
421  $F_{ST}$  values were low (0.033 – 0.074), yet significant for all the comparisons  
422 ( $P < 0.03$ ) (Table 3). Differences in the distribution of allele frequencies  
423 were observable for *CytB* (Fig 4A, *COI*: S4 Fig) and pairwise  $\Theta_{ST}$  values  
424 were low for all the comparisons for both *COI* and *CytB* (0.007-0.029); a

425 weak significant differentiation was observed only between TR and NL for  
426 *CytB* ( $\Theta_{ST}=0.026$ ,  $P=0.0226$ , Table 3). MDS ordination of the  $D_{SEQ}$  matrices  
427 revealed no observable structure for either gene (Fig 4B). However,  
428 PERMANOVA analysis suggests that sampling location has a significant  
429 effect on the population structure for *COI* ( $F_{4, 145}=1.8791$ ,  $P=0.0173$ ), but  
430 not for *CytB* ( $F_{4, 145}=1.3386$ ,  $P=0.1806$ ), which was consistent with the  
431 results of  $\Phi_{ST}$  analysis. Pairwise PERMANOVA tests showed significant  
432 but weak differentiation between the AKL and TP ( $P=0.0276$ ), and AKL  
433 and TR populations ( $P=0.0218$ ) (Table 3). PERMDISP revealed no  
434 significant differences in dispersion among populations for either *COI*  
435 ( $F_{4,141}=0.3793$ ,  $P=0.847$ ) or *CytB* ( $F_{4, 141}=0.358$ ,  $P=0.833$ ). CAP analysis  
436 was performed for *COI* but not for *CytB* because PERMANOVA did not  
437 reveal significant differences between groups for the latter. A CAP analysis  
438 to discriminate the three northern populations was statistically significant  
439 (trace=0.293,  $P=0.0038$ , S5 Fig), but the populations varied in their degree  
440 of distinctiveness under the CAP model. Specifically, leave-one-out cross-  
441 validation had the highest allocation success for TR (69.7%), followed by  
442 AKL (50%), but there was little or no identifiability of samples from Ti  
443 Point (only 25% correct allocation). Overall, the CAP results were  
444 consistent with the structures revealed by PERMANOVA: among the five  
445 populations, AKL is weakly distant from TP and TR.

446 **Fig 4. Graphical visualization of the results of population structure and**  
447 **demographic analysis for mtDNA data.**

448 (A) The frequencies of *CytB* haplotypes at each location ( $N=156$ ). The pie  
449 segment represents the relative haplotype frequencies. Each colour  
450 corresponds to a different haplotype. The patterned segment represents  
451 private haplotypes. The sizes of the circles are proportional to the sample  
452 size. (B) Non-metric MDS ordination of distances obtained from the  
453 standard nucleotide differences between individuals for the *COI* data. Each  
454 symbol also represents a different sampling location. (C) Mismatch  
455 distributions of pairwise base pair differences between the concatenated  
456 *COI* and *CytB* haplotypes.

457 PERMANOVA analysis was repeated by contrasting the TP, AKL and TR  
458 populations (the northern population cluster) against WL and NL (the  
459 southern population cluster) to test whether the north-south differentiation  
460 identified with microsatellite data could be observed with the mtDNA data.  
461 The results indicate no significant differentiation between the clusters (*COI*:  
462  $F_{1, 145}=0.819$ ,  $P=0.5268$ ; *CytB*:  $F_{1, 145}=1.2325$ ,  $P=0.2982$ ) and do not support  
463 the north-south differentiation identified by microsatellite data.

464 Hierarchical AMOVA analysis of concatenated sequences performed in the  
465 absence of population structure showed that differentiation among sampling  
466 sites explains only 0.69% and 1.26% of the total variation based on  $\Theta_{ST}$  and

467  $F_{ST}$  statistics, respectively, with high significance ( $\Theta_{ST}=0.01260$ ,  
468  $P=0.03082$ ;  $F_{ST}=0.00686$ ,  $P=0.0001$ ), whereas variation within populations  
469 explained 98.74% ( $\Theta_{ST}$ ) or 99.31% ( $F_{ST}$ ) of the total variation (S13 Table).  
470 Grouping of populations into northern and southern clusters explained a  
471 maximum 0.16% of the total variation, but it was not significant ( $\Theta_{CT}=-$   
472  $0.0062$ ,  $P=0.8994$ ;  $F_{CT} = 0.00161$ ,  $P=0.10029$ ). These analyses show that  
473 mtDNA data do not support the north-south differentiation evident in  
474 microsatellite data.

475 A phylogenetic analysis of *P. maculata* shows that *P. maculata* samples  
476 from NZ form a single clade (S6 Fig). Inclusion of the Argentinian  
477 sequences did not change the topology of the tree. Phylogenetic  
478 relationships between *P. maculata* individuals were unsupported (bootstrap  
479 values <50%), whereas there was good bootstrap support (between 74-  
480 100%) for five other species in the Pleurobranchidae family.

## 481 **Migration and demographic changes**

### 482 **Microsatellite analyses**

483 Migration analysis with GeneClass2 detected four first generation migrants  
484 ( $P=0.01$ ): one individual sampled from TP is a migrant from AKL, one  
485 individual sampled from TR is from AKL, and two individuals sampled  
486 from NL are migrants from WL (S14 Table). Individuals from each cluster  
487 were more likely to belong to populations from the same cluster. A lack of

488 first generation migrants between the clusters shows that these clusters are  
 489 genetically not well connected. However, CAP and STRUCTURE show  
 490 low connectivity between the clusters and admixed individuals in both  
 491 northern and southern clusters (Figs 2D and E). The highest  
 492 misclassifications between the clusters detected by CAP analysis and  
 493 highest admixture proportions detected by STRUCTURE were noted in TR  
 494 and WL. This suggests that the TR and WL populations are bridges between  
 495 the clusters. Admixture in the WL population may also explain weak  
 496 differentiation between the WL and NL populations that was found with  $F_{ST}$   
 497 and PERMANOVA tests.

498 The Wilcoxon test did not detect recent bottlenecks in any population under  
 499 either TPM or SMM models (Table 4). In addition, analysis of mode-shift in  
 500 the distribution of allele frequencies suggests that all the populations exhibit  
 501 a normal L-shaped pattern indicating no mode-shift in the frequency  
 502 distribution of alleles. Taken together these data suggest that none of the  
 503 populations experienced a recent or sudden bottleneck.

504 **Table 4. Results of the neutrality and demographic tests using either**  
 505 **microsatellite or mtDNA data.**

Population	Microsatellite		mtDNA – concatenated sequences				
	Bottleneck		Tajima's $D$	Fu's $F_s$	$\tau$	Mismatch Distribution	
	TPM	SMM				$SSD (P_{SSD})$	$Raggedness (P_r)$
Ti Point	0.741	0.945	-1.45	-14.57 <sup>c</sup>	8.783	0.010 (0.368)	0.033 (0.184)
Auckland	0.715	0.993	-2.14 <sup>b</sup>	-23.57 <sup>c</sup>	5.254	0.017 (0.028)	0.008 (0.980)
Tauranga	0.689	0.995	-2.10 <sup>b</sup>	-24.93 <sup>c</sup>	7.590	0.001 (0.873)	0.006 (0.968)
Wellington	0.633	0.052	-1.62 <sup>a</sup>	-9.47 <sup>c</sup>	7.691	0.001 (0.944)	0.011 (0.852)



<b>Nelson</b>	0.954	0.999	-2.20 <sup>b</sup>	-25.08 <sup>c</sup>	7.994	0.001 (0.870)	0.006 (0.953)
<b>Total</b>	-	-	-2.45 <sup>c</sup>	-24.74 <sup>c</sup>	7.797	0.001 (0.756)	0.004 (0.929)

506 Total values for Tajima's  $D$ , Fu's  $F_s$ ,  $\tau$  and the parameters of the mismatch  
507 distribution analysis were calculated pooling all available individuals into a  
508 single pool.

509 TPM: Two-phase mutational model, SMM: stepwise mutational model,  
510 SSD: sum of squared deviations and statistical significance,  $P_{SSD}$ : for the  
511 validity of the sudden expansion model,  $\tau$ : time passed since population  
512 expansion, *raggedness* and  $P_r$ : Harpending's raggedness index and its  
513 probability' respectively, for the null hypothesis test of goodness-of-fit.  $P$ -  
514 values: <sup>a</sup> <0.05, <sup>b</sup> <0.01, <sup>c</sup> <0.001

### 515 **Mitochondrial DNA analyses**

516 Overall, neutrality tests of Tajima's  $D$  and Fu's  $F_s$  revealed significant  
517 negative values in pooled samples (Table 4) suggesting a recent population  
518 expansion or purifying selection. This was also suggested by the uni-modal  
519 mismatch distributions of pairwise base pair differences for *COI* and *CytB*  
520 haplotypes (Fig 4C and S7 Fig, respectively), non-significant *SSD* and  
521 *raggedness* patterns (Table 4). Furthermore, the McDonald-Kreitman test  
522 found no evidence of positive selection: the ratio of nonsynonymous to  
523 synonymous substitutions within *P. maculata* ( $P_n/P_s=5/136$ ) and between  
524 species ( $D_n/D_s=5/100$ ) was statistically similar (neutrality index=0.400,  
525  $P=0.1104$ ).

526 The demographic parameter tau ( $\tau$ ), which represents the mutational time  
527 since expansion, was calculated after pooling all five populations (pooling  
528 of populations was justified because all exhibited signs of population  
529 expansion):  $\tau$  was estimated at 2.703 (CI 95%: 1.488-3.320) and 2.754 (CI  
530 95%: 1.344-3.615) for the *COI* and *CytB* data, respectively. The date of

531 expansion was estimated at 44.2 kya (CI 95%: 24.3-54.3 kya) based on *COI*  
532 data.

## 533 **DISCUSSION**

534 This study marks the first attempt to describe and account for patterns of  
535 genetic diversity in *P. maculata*. Overall, we observed marked signals of  
536 population structure; however, the population structure suggested by  
537 microsatellite versus mtDNA data differed. The nuclear data exhibit patterns  
538 of diversity indicative of a north-south disjunction. The northern samples  
539 formed one group and southern (WL and NL) samples formed another (with  
540 few examples of migration). However this disjunction was not supported by  
541  $F_{ST}$  analysis of mtDNA data, which indicated divergence among all  
542 populations.

543 Discordance between results obtained from nuclear versus mitochondrial  
544 markers is not uncommon, with explanations ranging from variation in  
545 lineage sorting to differences in rates of nuclear versus mitochondrial  
546 evolution [reviewed in 75, 76]. Taken together, we interpret our data as  
547 indicative of a single founding population that subsequently became  
548 fragmented following geographical and oceanographical changes that led to  
549 the present north-south divide in NZ waters.

550 According to this scenario, *P. maculata* was inhabiting the NZ waters  
551 before the end of the last glacial maximum when sea levels were low and  
552 NZ was a single land mass [77, 78]. Shallow-water species that inhabit  
553 intertidal or shallow subtidal areas were primarily affected as decreased sea  
554 levels destroyed available habitat [79, 80]. At some later time, most likely  
555 following the last glacial maximum (~22,000 years ago), large benthic  
556 habitats became available [79], and this may have facilitated population  
557 expansion and fragmentation aided by warming temperatures and rising sea  
558 levels [81].

559 In support of this model is the haplotype network based on mtDNA data  
560 showing star-like structure and high haplotype diversity, indicative of a  
561 population expansion arising from a small initial population [74]. Additional  
562 support comes from the unimodal mismatch distribution pattern of pairwise  
563 base pair differences for *COI* and *CytB* haplotypes (Fig 4C and S6 Fig,  
564 respectively), non-significant *SSD* and *raggedness* patterns (Table 4), and  
565 significant negative values for Tajima's *D* and Fu's *F<sub>s</sub>*. The approximate  
566 date of population expansion was estimated at 44.2 kya (CI: 24.3–54.3kya).  
567 This estimate assumes a 5.3% divergence rate for the *COI* gene, which  
568 while an estimate, nonetheless coincides with a periodic of cyclic climatic  
569 oscillation defined by the late Pleistocene era (~110–15 kya) [82].  
570 Glaciation has been suggested as the possible cause of demographic changes

571 in other NZ marine organisms, including triplefin species [20] and red alga  
572 *Bostrychia intricata* [83].

573 This scenario is also supported by the analysis of more rapidly evolving  
574 microsatellites, which are useful for uncovering recent barriers to gene flow  
575 [84-86], especially following periods of demographic expansion [84, 86].  
576 As sea levels rose at the end of the last glacial cycle [ $\sim$ 13,000 years ago;  
577 78], geographical factors and associated oceanography established barriers  
578 to gene flow for marine organisms. In particular, confluence of the East  
579 Cape current with the Wairarapa Eddy off the east coast of the North Island  
580 (between 37–39°S) created an oceanographic barrier [87]. A barrier was  
581 also formed by waters separating North and South Islands (the Cook Strait)  
582 [87, 88].

583 A north-south genetic differentiation has been observed in other marine  
584 organisms from NZ (reviewed in Gardner et al, 2010; Ross et al, 2012).  
585 Confluence of the East Cape current with the Wairarapa Eddy is thought to  
586 be responsible for population differentiation in organisms such as the  
587 amphipods *Paracorophium excavatum* and *P. lucasi* [87] and the gastropod  
588 *Diloma subrostrata* [89]. The Cook Strait barrier is thought to have  
589 underpinned north-south differentiation in organisms such as the green shell  
590 mussel (*Perna canaliculus*) [19], blackfoot paua (*Haliotis iris*) [21] and the  
591 Ornate limpet (*Cellana ornate*) [90].

592 One additional factor that has likely promoted recent population subdivision  
593 in *P. maculata* is the distribution of invasive species [91] that constitute a  
594 food source for *P. maculata*. The Asian date mussel, *Arcuatula senhousia*  
595 has been established in the Auckland region since the 1970s, forming large  
596 transient beds in sub- and inter-tidal areas of the Hauraki Gulf, Manukau  
597 Harbour and Whangarei Harbour [92]. Expansion of *A. senhousia* beds in  
598 the Hauraki Gulf appears to have preceded increases in the density of *P.*  
599 *maculata* populations. Interestingly, subsequent decline of near-shore beds  
600 of *A. senhousia* post 2010 was followed by rapid decline in the density of *P.*  
601 *maculata* populations [93]. Further evidence that range expansion of *P.*  
602 *maculata* may have been facilitated by availability of prey species comes  
603 from off-shore mussel farms in Tasman Bay (Nelson, NZ), where culture of  
604 the green shell and blue mussels have created new habitats for *P. maculata*,  
605 with high-density populations being found beneath mussel farms [93].  
606 Recently, *P. maculata* was identified in Argentinean waters with the species  
607 rapidly spreading along the Atlantic coast [5, 6]. The minor difference  
608 between mtDNA sequence in Argentinian versus NZ raises the possibility  
609 that NZ maybe the source of the recently discovered Argentinian  
610 population.

611 Life history traits such as the nature of egg and larval stages are of  
612 understandable importance in shaping population structure of the species.  
613 Species with benthic eggs tend to have more structured populations than

614 ones with pelagic eggs [94, 95] and an inverse relationship between pelagic  
615 larval duration (PLD) and genetic structure has been found [96, 97]: in a  
616 comparative analysis of NZ pelagic marine species, Ross et al. [96] showed  
617 a significant negative correlation between PLD and genetic differentiation.  
618 However, in the same meta-study, when NZ-wide sampling regimes were  
619 considered, NZ organisms with PLD durations similar to *P. maculata* (2–4  
620 weeks) exhibit structural patterns of diverse types ranging from no structure,  
621 to a north-south disjunction, IBD and differentiation within and between  
622 sampling locations [96].

623 Our prediction that the previously recorded cline in TTX might be explained  
624 by genetic structure holds only for microsatellite markers. Had this also held  
625 for mtDNA markers a case may have been made that *P. maculata* is a  
626 complex of two cryptic species, but no such evidence exists. Our  
627 phylogenetic analysis indicates that all *P. maculata* populations – including  
628 samples from Argentinian waters [5] – are conspecific. Short branches with  
629 no or low bootstrap support is also indicative of lack of genetic  
630 differentiation among *P. maculata*. Similar lack of differentiation between  
631 toxic and non-toxic populations has been shown for *Taricha granulosa*  
632 newts from various localities in western North America [98, 99] and the  
633 red-spotted newt *Notophthalmus viridescens* [100].

634 Having called into question substantive genetic differences between north  
635 and south populations, differences in TTX levels are thus likely attributable  
636 to exogenous factors, such as differences in associated bacteria, exposure, or  
637 diet. Work to date is strongly suggestive of diet as the major source of TTX,  
638 with *P. maculata* accumulating TTX via feeding [10], while offspring from  
639 TTX positive individuals raised in a TTX-free environment become free of  
640 TTX [12]. Recent work studying cultured bacteria from *P. maculata* found  
641 no evidence for a bacterial origin of the toxin [15], but TTX has been  
642 reported in certain prey, including a Platyhelminthes *Stylochoplana* species,  
643 that co-occur with TTX-containing *P. maculata* [11].

## 644 **ACKNOWLEDGEMENTS**

645 We acknowledge Susanna Wood and David Taylor (Cawthron Institute);  
646 Lauren Salvitti, Dudley Bell and Warrick Powrie (The University of  
647 Waikato); Mike McMurtry (Auckland Regional Council); Severine Hannam  
648 and Wilma Blom (Auckland Museum); Shane Genage (Victoria University  
649 of Wellington); Steve Journee (The Dive Guys, Wellington); Richard  
650 Huges, Paul Caiger and Bakhti Patel (The Leigh Marine Laboratory at the  
651 University of Auckland); and Don Morrissey, Matthew Smiths and Stephen  
652 Brown (The National Institute of Water and Atmospheric Research). We  
653 thank Paul McNabb (Cawthron Institute) and Serena Khor (The University  
654 of Waikato) for conducting tetrodotoxin assays.

655 Supporting Information accompanies this paper on the website.

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## 958 **Supporting Information Captions**

959 **Supporting Information Text: Supplementary Methodology and**

960 **Results**

961 **S1 Table. Sampling locations and abbreviations for the New Zealand**

962 ***Pleurobranchaea maculata* populations.**

963 **S2 Table. PCR annealing temperature and elongation time for**

964 **microsatellite markers.**

965 **S1 Table. Primers used to amplify the mtDNA genes.**

966 **S2 Table. Summary genetic diversity statistics for the 12 microsatellites**

967 **and 5 populations of *Pleurobranchaea maculata*.**

968 **S3 Table. Allele frequencies at 12 microsatellite loci by population.**

969 **S4 Table. Results of the linkage equilibrium test for 12 microsatellite**

970 **markers.** The calculations were performed for each population and across

971 the populations. *P* values were estimated by 6600 permutations. The *P*-

972 value for the 5% nominal level is 0.000152 after Bonferroni correction.

973 **S5 Table. Estimating the null allele frequency using the ENA algorithm.**

974 **S6 Table. The estimation of the possible effect of null alleles on**  
975 **population structure analysis.** Estimated (A) global  $F_{ST}$  of Weir (1996) for  
976 each microsatellite locus and (B) pairwise  $F_{ST}$  for each pair of populations  
977 with and without using the ENA correction [28].

978 **S9 Table. Pairwise population matrix of the *Dest* values at each locus.**  
979 ***Dest* values below the diagonal.** The  $P$  values calculated with 9999  
980 permutations are shown above the diagonal. TP: Ti Point, AKL: Auckland,  
981 TR: Tauranga, WL: Wellington, NL: Nelson.

982 **S10 Table. AMOVA results for the *Pleurobranchaea maculata***  
983 **populations based on the microsatellite data.** Analysis was performed  
984 with SMM based on  $R_{ST}$ -statistics. The hierarchical distribution of variation  
985 was determined between the sampling locations without introducing any  
986 structure, and nesting the Ti Point-Auckland and Tauranga populations in  
987 the northern cluster and the Wellington and Nelson populations in the  
988 southern cluster.

989 **S3 Fig. Median joining network of the *COI* haplotypes.** (A) 1153 bp of  
990 *COI* sequences from 156 New Zealand individuals. (B) 624 bp of *COI*  
991 sequences from 156 New Zealand and four Argentinian individuals. The  
992 network was coloured according to the sampling locations. The diameters of  
993 the circles are proportional to the frequency of the haplotypes. The hashes  
994 indicate the mutational steps between the haplotypes. The black nodes

995 represent the imaginary haplotypes necessary to create a bridge between the  
996 present haplotypes.

997 **S11 Table. Haplotype frequencies for the mtDNA sequences by mere**  
998 **counting in populations.** The IDs are individual names. <sup>+</sup>Population codes.  
999 \*Sample size.

1000 **S12 Table. The results of the saturation test for *COI* and *CytB***  
1001 **sequences.** The results suggest only little saturation both genes while  
1002 assuming both symmetrical and asymmetrical topology as significant  
1003 difference was observed between *Iss* and *Iss.c* values, and  $Iss < Iss.c$ .

1004 **S13 Table. AMOVA results of concatenated *COI* and *CytB* sequences**  
1005 **based on F- and  $\Theta$ -statistics.**

1006 **S14 Table. Detection of first generation migrants.** Likelihood ratio:  
1007  $L_{\text{home}}/L_{\text{max}}$ . The number of individuals with a probability below the  
1008 threshold value is 4. The potential migrants are labelled in red ( $P < 0.01$ )  
1009 and the most likely population in green. ID: ID of the individuals.

1010 **S1 Fig. Allele rarefaction curves for the microsatellite data.**

1011 **S2 Fig. CAP analyses of the multivariate genetic distances between the**  
1012 ***P. maculata* individuals obtained from 12 microsatellite markers.** (A)  
1013 *DM* (the number of axes -  $m=13$ ) and (B) *DCL* ( $m=6$ ) distances. The axes



1014 represent the amount of variation between the populations that is explained  
1015 by the two axes. Each symbol represents a different sampling location.

1016 **S3 Fig. Median joining network of the *COI* haplotypes.** (A) 1153 bp of  
1017 *COI* sequences from 156 New Zealand individuals. (B) 624 bp of *COI*  
1018 sequences from 156 New Zealand and four Argentinian individuals. The  
1019 network was coloured according to the sampling locations. The diameters of  
1020 the circles are proportional to the frequency of the haplotypes. The hashes  
1021 indicate the mutational steps between the haplotypes. The black nodes  
1022 represent the imaginary haplotypes necessary to create a bridge between the  
1023 present haplotypes.

1024 **S4 Fig. The graphical representation of haplotypes frequencies for *COI***  
1025 **data at each sampling location.** The pie segment represents the relative  
1026 frequencies of the haplotypes. Each colour corresponds to a different  
1027 haplotype. Private haplotypes specific to each location are represented with  
1028 a patterned segment.

1029 **S5 Fig. Graphical visualization of the standard nucleotide differences**  
1030 **between individuals based on the *COI* data.** Canonical correlation  
1031 ordination analysis at  $m=4$  was performed.

1032 **S6 Fig. Molecular phylogenetic analysis of *COI* sequences by maximum**  
1033 **likelihood method.** The tree with the highest log likelihood (-2685.86 is  
1034 displayed. Initial tree(s) for the heuristic search were obtained automatically

1035 by applying Neighbor-Join and BioNJ algorithms to a matrix of pairwise  
1036 distances estimated using the Maximum Composite Likelihood (MCL)  
1037 approach, and then selecting the topology with superior log likelihood  
1038 value. The analysis involved 49 nucleotide sequences out of a total of 616.  
1039 The nodes are coloured based on the sampling locations. The numbers on  
1040 the nodes represent the % bootstrap support (values <50% not presented).

1041 **S7 Fig. Mismatch distribution of pairwise base pair differences between**  
1042 **the concatenated *COI* and *CytB* haplotypes.**







