Ongoing, rational calibration of reward-driven perceptual biases

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Abstract

Decision-making is often interpreted in terms of normative computations that maximize a particular reward function for stable, average behaviors. Aberrations from the reward-maximizing solutions, either across subjects or across different sessions for the same subject, are often interpreted as reflecting poor learning or physical limitations. Here we show that such aberrations may instead reflect the involvement of additional satisficing and heuristic principles. For an asymmetric-reward perceptual decision-making task, three monkeys produced adaptive biases in response to changes in reward asymmetries and perceptual sensitivity. Their choices and response times were consistent with a normative accumulate-to-bound process. However, their context-dependent adjustments to this process deviated slightly but systematically from the reward-maximizing solutions. These adjustments were instead consistent with a rational process to find satisficing solutions based on the gradient of each monkey's reward-rate function. These results suggest new dimensions for assessing the rational and idiosyncratic aspects of flexible decision-making.

Introduction

Normative theory has played an important role in our understanding of how the brain forms decisions. For example, many perceptual, memory, and reward-based decisions show inherent trade-offs between speed and accuracy. These trade-offs are parsimoniously captured by a class of sequential-sampling models, such as the drift-diffusion model (DDM), that are based on the accumulation of noisy evidence over time to a pre-defined threshold value, or bound (Ratcliff, 1978; Gold and Shadlen, 2002; Bogacz et al., 2006; Krajbich et al., 2010). These models have close ties to statistical decision theory, particularly the sequential probability ratio test that can, under certain assumptions, maximize expected accuracy for a given number of samples or minimize the number of samples needed for a given level of accuracy (Barnard, 1946; Wald, 1947; Wald and Wolfowitz, 1948). However, even when these models provide good descriptions of the average behavior of groups of subjects, they may not capture the substantial variability in behavior that can occur across individuals and task conditions. The goal of this study was to better understand the principles that govern this variability, in particular how these principles relate to normative theory.

We focused on a perceptual decision-making task with asymmetric rewards. For this task, both human and animal subjects tend to make decisions that are biased towards the percept associated with the larger payoff (e.g. Maddox and Bohil, 1998; Voss et al., 2004; Diederich and Busemeyer, 2006; Liston and Stone, 2008; Serences, 2008; Feng et al., 2009; Simen et al., 2009; Nomoto et al., 2010; Summerfield and Koechlin, 2010; Teichert and Ferrera, 2010; Gao et al., 2011; Leite and Ratcliff, 2011; Mulder et al., 2012; Wang et al., 2013; White and Poldrack, 2014). These biases are roughly consistent with a rational strategy to maximize a particular reward function that depends on both the speed and accuracy of the decision process, such as the reward rate per trial or per unit time (Gold and Shadlen, 2002; Bogacz et al, 2006). This strategy can be accomplished via context-dependent adjustments in a DDM-like decision process along two primary dimensions (Fig. 1A): 1) the momentary sensory evidence, via the drift rate; and 2) the decision rule, via the relative bound heights that govern how much evidence is needed for each alternative. Subjects tend to make adjustments along one or both of these dimensions to produce overall biases that are consistent with normative theory, but with substantial individual variability (Voss et al., 2004; Cicmil et al., 2015; Bogacz et al., 2006; Simen et al., 2009; Summerfield and Koechlin, 2010; Leite and Ratcliff, 2011; Mulder et al., 2012; Goldfarb et al., 2014).

To better understand the principles that govern these kinds of idiosyncratic behavioral patterns, we trained three monkeys to perform a response-time (RT), asymmetric-reward decision task with mixed perceptual uncertainty (Fig. 1B). Like human subjects, the monkeys showed robust decision biases toward the large-reward option. These biases were sensitive to not just the reward asymmetry, as has been shown previously, but also to experience-dependent changes in perceptual sensitivity. These biases were consistent with adjustments to both the momentary evidence and decision rule in the

DDM. However, these two adjustments favored the large- and small-reward choice,
respectively, leading to nearly, but not exactly, maximal reward rates. We accounted for
these adjustments in terms of a satisficing, gradient-based learning model that calibrated
biases to balance the relative influence of perceptual and reward-based information on
the decision process. Together, the results imply the broad applicability of normative
theory, including sequential-sampling and heuristic-based strategies to understand how
the brain combines uncertain sensory input and internal preferences to form decisions
that can vary considerably across individuals and task conditions.

Results

We trained three monkeys to perform the asymmetric-reward random-dot motion
discrimination ("dots") task (Fig. 1B). All three monkeys were initially trained on a
symmetric-reward version of the task for which they were required to make fast eye
movements (saccades) in the direction congruent with the global motion of a random-dot
kinematogram to receive juice reward. They then performed the asymmetric-reward
versions that were the focus of this study. Specifically, in blocks of 30–50 trials, we
alternated direction-reward associations between a "LR-Right" reward context (the large
reward was paired with a correct rightward saccade and the small reward was paired
with a correct leftward saccade) and the opposite "LR-Left" reward context. We also
varied the ratio of large versus small reward magnitudes ("reward ratio") across sessions
for each monkey. Within a block, we randomly interleaved motion stimuli with different
directions and motion strengths (expressed as coherence, the fraction of dots moving in
the same direction). We monitored the monkey's choice (which saccade to make) and
RT (when to make the saccade) on each trial.

The monkeys' biases reflected changes in reward context and perceptual
sensitivity

For the asymmetric-reward task, all three monkeys tended to make more choices
towards the large-reward option, particularly when the sensory evidence was weak.
These choice biases corresponded to horizontal shifts in the psychometric function
describing the probability of making a rightward choice as a function of signed motion
coherence (negative for leftward motion, positive for rightward motion; Fig. 2A). However,
these biases differed in detail for the three monkeys. For example, bias magnitude
tended to be smallest in monkey F. Moreover, bias magnitude tended to decrease with
session number for all three monkeys, although this tendency was statistically significant
only for monkey C after accounting for co-variations with reward rate (Fig. 2B, middle).
Additionally, each monkey showed steady increases in perceptual sensitivity (steepness
of the psychometric function), which initially dropped relative to values from the
symmetric-reward task then tended to increase with more experience with asymmetric
rewards (Fig. 2B, top; H_0: partial Spearman's ρ of sensitivity versus session index after
accounting for session-specific reward ratios = 0, $p<0.01$ in all cases, except LR-Left for
monkey C, for which $p=0.57$). These improvements did not involve systematic changes
in lapse rates, which remained near zero across sessions (Fig. 2B, bottom), implying

that the monkeys knew how to perform the task. However, these increases in sensitivity
did correspond to systematic decreases in choice biases in some, but not all, cases (3
monkeys x 2 reward contexts; Fig. 2C), which is expected for an optimal process that
maximizes reward per trial (RTrial; Fig. S2). Compared to these optimal biases, the
monkey's biases tended to vary over sessions and overshoot for monkeys F and C (Fig.
2D).

To better understand the computational principles that governed these idiosyncratic
biases, while also taking into account systematic relationships between the choice and
RT data, we fit single-trial RT data (i.e., we modeled full RT distributions, not just mean
RTs) from individual sessions to a DDM. We used a hierarchical-DDM (HDDM) method
that assumes that parameters from individual sessions are samples from a group
distribution (Wiecki et al., 2013). The HDDM had six parameters for each reward context.
Four were from a basic DDM (Fig. 1A): a, the total bound height, representing the
distance between the two choice bounds; k, a scaling factor that converts sensory
evidence (motion strength and direction) to the drift rate; and t_0 and t_1, non-decision
times for leftward and rightward choices, respectively. The additional two parameters
provided biases that differed in terms of their effects on the full RT distributions
(Supplementary Fig. 2): me, which is additional momentary evidence that is added to the
motion evidence at each accumulating step and has asymmetric effects on the two
choices and on correct versus error trials (positive values favor the rightward choice);
and z, which determines the decision rules for the two choices and tends to have
asymmetric effects on the two choices but not on correct versus error trials (values >0.5
favor the rightward choice). A DDM variant with collapsing bounds provided qualitatively
similar results as the HDDM. Thus, subsequent analyses use the model with fixed
bounds, unless otherwise noted. Example fits for this model are shown in Fig. 3, and
summaries of best-fitting parameters and goodness-of-fit metrics are provided in
Supplementary Table 1.

The DDM fits provided a parsimonious account of both the choice and RT data.
Consistent with the results from the logistic analyses, the HDDM analyses showed that
the monkeys made systematic improvements in psychometric sensitivity (H_0: partial
Spearman's ρ of sensitivity versus session index after accounting for session-specific
reward ratios=0, $p<0.01$ in all cases except $p=0.06$ for LR-Left for monkey A). Moreover,
there was a negative correlation between psychometric sensitivity and choice bias (H_0:
partial Spearman's ρ of sensitivity versus total bias after accounting for session-specific
reward ratios=0, $p<0.001$ in all cases). These fits ascribed the choice biases to changes
in both the momentary evidence (me) and the decision rule (z) of the decision process,
as opposed to either parameter alone (Supplementary Table 2). These fits also indicated
context-dependent differences in non-decision times, which were smaller for all large-
reward choices for all three monkeys except in LR-Right context for monkey C and A (t-
test, $p<0.05$).

The monkeys' bias adjustments were adaptive with respect to optimal reward-rate
functions

To try to identify common principles that governed these monkey- and context-
dependent decision biases, we analyzed behavior with respect to optimal benchmarks
based on certain reward-rate functions. We focused on reward per unit time (RR) and
per trial (RTrial), which for this task are optimized in a DDM framework by adjusting
momentary-evidence (me) and decision-rule (z) biases, such that both favor the large-
reward choice. However, the magnitudes of these optimal adjustments depend on other
task parameters (a, k, t_0, and t_1, non-bias parameters from the DDM, plus the ratio of the
two reward sizes and inter-trial intervals; Supplementary Fig. 12) that can vary from
session to session. Thus, to determine the optimal adjustments, we performed DDM
simulations with the fitted HDDM parameters from each session, using different
combinations of me and z values (Fig. 4A). As reported previously (Bogacz et al., 2006;
Simen et al., 2009), when the large reward was paired with the leftward choice, the
optimal strategy used $z<0.5$ and $me<0$ (Fig. 4B, top panels, purple and orange circles for
RR and RTrial, respectively). Conversely, when the larger reward was paired with the
rightward choice, the optimal strategy used $z>0.5$ and $me>0$ (Fig. 4B, bottom panels).

The monkeys' adjustments of momentary-evidence (me) and decision-rule (z) biases
showed both differences and similarities with respect to these optimal predictions (Fig.
4B, black circles; similar results were obtained using fits from a model with collapsing
bounds, Supplementary Fig. 5). In the next section, we consider the differences, in
particular the apparent use of shifts in me in the adaptive direction (i.e., favoring the
large-reward choice) but of a magnitude that was larger than predicted, along with shifts
in z that tended to be in the non-adaptive direction (i.e., favoring the small-reward
choice). Here we focus on the similarities and show that the monkeys' decision biases
were adaptive with respect to the reward-rate function in four ways (RTrial provided
slightly better predictions of the data and thus are presented in the main figures; results
based on RR are presented in the Supplementary Figures).

First, the best-fitting me and z values from each monkey corresponded to near-maximal
reward rates (Fig. 5A). We compared the optimal values of reward per trial ($RTrial_{max}$) to
the values predicted from the monkeys' best-fitting me and z adjustments ($RTrial_{predict}$).
Both $RTrial_{predict}$ and $RTrial_{max}$ depended on the same non-bias parameters in the HDDM
fits that were determined per session (a, k, t_0, and t_1) and thus are directly comparable.
Their ratios tended to be nearly, but slightly less than, one (mean ratio: 0.977, 0.984,
and 0.982 for monkeys F, C, and A, respectively) and remained relatively constant
across sessions (H_0: slopes of linear regressions of these ratios versus session
number=0, $p>0.05$ for all three monkeys). Similar results were also obtained using the
monkeys' realized rewards, which closely matched $RTrial_{predict}$ (Spearman's $\rho=0.991$,
0.979, and 0.906 for monkeys F, C, and A, respectively, $p<0.0001$ in all three cases).
These results reflected the shallow plateau in the RTrial function near its peak (Fig. 5B),
such that the monkeys' actual adjustments of me and z were within the contours for 97%
$RTrial_{max}$ in most sessions (Fig. 5C; see Supplementary Fig. 6 for results using RR).
Thus, the monkeys' overall choice biases were consistent with strategies that lead to
nearly optimal reward outcomes.

Second, the across-session variability of each monkey's decision biases was predicted
by idiosyncratic features of the reward functions. The reward functions were, on average,
different for the two reward contexts and each of the three monkeys (Fig. 6A). These
differences included the size of the near-maximal plateau (red patch), which determined
the level of tolerance in RTrial for deviations from optimal adjustments in *me* and *z*. This
tolerance corresponded to the session-by-session variability in each monkey's *me* and *z*
adjustments (Fig. 6B). In general, monkey F had the smallest plateaus and tended to
use the narrowest range of *me* and *z* adjustments across sessions. In contrast, monkey
A had the largest plateaus and tended to use the widest range of *me* and *z* adjustments
(Pearson's ρ between the size of the 97% RTrial contour, in pixels, and the sum of the
across-session variances in each monkeys' *me* and *z* adjustments=0.83, *p*=0.041).
Similar analyses using the RR function produced qualitatively similar results
(Supplementary Fig. 8B).

Third, the session-by-session adjustments in both *me* and *z* corresponded to particular
features of each monkey's context-specific reward function. The shape of this function,
including the orientation of the plateau with respect to *z* and *me*, depended on the
monkey's perceptual sensitivity and the reward ratio for the given session. The monkeys'
me and *z* adjustments varied systematically with this orientation (Fig. 6C and D for
RTrial, Supplemental Fig. 8C and D for RR). This result was not an artifact of the fitting
procedure, which was able to recover appropriate, simulated bias parameter values
regardless of the values of non-bias parameters that determine the shape of the reward
function (Supplementary Fig. 9).

Fourth, the monkeys' *me* and *z* adjustments were correlated with the values that would
maximize RTrial, given the value of the other parameter for the given session and
reward context (Fig. 6E for RTrial, Supplemental Fig. 8E for RR). These correlations
were substantially weakened by shuffling the session-by-session reward functions
(Supplementary Fig. 7). Together, these results suggest that all three monkeys used
biases that were adaptively calibrated with respect to the reward information and
perceptual sensitivity of each session.

237 *The monkeys' adaptive adjustments were consistent with a satisficing, gradient-*
238 *based learning process*

Thus far, we showed that all three monkeys adjusted their decision strategies in a
manner that matched many features of the optimal predictions based on their
idiosyncratic, context-specific reward-rate functions. However, their biases did not match
the optimal predictions exactly. Specifically, all three monkeys used shifts in *me* favoring
the large-reward choice (adaptive direction) but of a magnitude that was larger than
predicted, along with shifts in *z* favoring the small-reward choice (non-adaptive direction).
We next consider if and how a consistent, adaptive process used by all three monkeys
could lead to these idiosyncratic and not-quite optimal patterns of adjustments.

We propose that a gradient-based satisficing model can account for these patterns of

adjustments. The intuition for the model is shown in Fig. 7. The lines on the RTrial
heatmap represent the trajectories of a gradient-tracking procedure that adjusts *me* and
z values to increase RTrial until reaching 97% of the maximum possible value. Gradient
lines are color-coded based on how *me* and *z* values at the end points relate to the
optimal *me* and *z* values. For example, consider adjusting *me* and *z* by following all of
the magenta gradient lines until their endpoints. This procedure would result in *me* shifts
in the adaptive direction with magnitudes larger than the optimal *me*, plus *z* shifts in the
non-adaptive direction. In other words, as long as the initial *me* and *z* values fall within
the area covered by the magenta lines, the positive gradient-tracking procedure would
lead to a good-enough solution with over-shifted *me* and non-adaptive *z* values similar to
what we found in the monkeys' data.

We simulated this process with the reward function derived for each reward context,
session, and monkey, using different starting points and a termination rule
corresponding to achieving the estimated RTrial$_{predict}$ from those conditions (see above).
This process is illustrated for LR-Left blocks in an example session from monkey C (Fig.
8A). We estimated the unbiased *me* and *z* values as the midpoints between their values
for LR-Left and LR-Right blocks (square). At this point, the RTrial gradient is larger along
the *me* dimension than the *z* dimension, reflecting the tilt of the reward function. We set
the initial point at baseline *z* and a very negative value of *me* (-90% of the highest
coherence used in the session; overshoot in the adaptive direction) and referred to this
setting as the "over-*me*" model. The *me* and *z* values were then updated in a step-wise
fashion (magenta trace), according to the RTrial gradient, until the monkey's RTrial$_{predict}$
or better was achieved (magenta circle). The endpoint of this updating process was very
close to monkey C's actual adjustment (gray circle). For comparison, three alternative
models are illustrated. The "over-*z*" model selects *z* as the initial dimension and
assumes updating from the baseline *me* and over-adjusted *z* values (blue, initial *z* set as
0.1 for the LR-Left context and 0.9 for the LR-Right context). The "over-both" model
assumes updating from the over-adjusted *me* and *z* values (green). The "neutral" model
assumes the same updating process but from the baseline *me* and baseline *z* (black).
The endpoints from these alternative models deviated considerably from the monkey's
actual adjustment.

The "over-*me*" model produced better predictions than the other three alternative models
for all three monkeys. Of the four models, only the "over-*me*" model captured the
monkeys' tendency to bias *me* toward the large-reward choice (positive Δme) and bias *z*
toward the small-reward choice (negative Δz; Fig. 8B). In contrast, the "over-*z*" model
predicted small adjustments in *me* and large adjustments in *z* favoring the large-reward
choice*;* the "over-both" model predicted relatively large, symmetric *me* and *z*
adjustments favoring the large-reward choice; and the "neutral" model predicted
relatively small, symmetric adjustments in both *me* and *z* favoring the large-reward
choice. Accordingly, for each monkey, the predicted and actual values of both Δme and
Δz were most strongly positively correlated for predictions from the "over-*me*" model
compared to the other models (Fig. 8C). The "over-*me*" model was also the only one of

the models we tested that recapitulated the measured relationships between both *me*-
and *z*-dependent biases and session-by-session changes in the orientation of the RTrial
function (Fig. 8D). Qualitatively similar results were obtained with a different criterion that
stopped updating when RTrial reached a fixed 98% of RTrial$_{max}$, suggesting that the
ability of the "over-*me*" model to mimic monkeys' performance did not simply result from
setting session-specific final RTrial values.

Discussion

We analyzed the behavior of three monkeys performing a decision task that encouraged
the use of both uncertain visual motion evidence and the reward context. All three
monkeys made choices that were sensitive to the strength of the sensory evidence and
were biased toward the larger-reward choice, which is roughly consistent with previous
studies of humans and monkeys performing similar tasks (Maddox and Bohil, 1998;
Voss et al., 2004; Diederich and Busemeyer, 2006; Liston and Stone, 2008; Serences,
2008; Feng et al., 2009; Simen et al., 2009; Nomoto et al., 2010; Summerfield and
Koechlin, 2010; Teichert and Ferrera, 2010; Gao et al., 2011; Leite and Ratcliff, 2011;
Mulder et al., 2012; Wang et al., 2013; White and Poldrack, 2014).However, we also
found that these adjustments differed considerably in detail for the three monkeys, in
terms of overall magnitude, dependence on perceptual sensitivity and offered rewards,
and relationship to RTs. We quantified these effects with a logistic analysis and a
commonly used model of decision-making, the drift-diffusion model (DDM), which
allowed us to compare the underlying decision-related computations to hypothetical
benchmarks that would maximize reward. We found that all three monkeys made reward
context-dependent adjustments with two basic components: 1) an over-adjustment of the
momentary evidence provided by the sensory stimulus (*me*) in favor of the large-reward
option; and 2) an adjustment to the decision rule that governs the total evidence needed
for each choice (*z*), but in the opposite direction (i.e., towards the small-reward option).
These adjustments tended to provide nearly, but not exactly, maximal reward intake. We
proposed a common heuristic strategy based on the monkeys' individual reward
functions to account for the idiosyncratic adjustments across monkeys and across
sessions within the same monkey.

Considerations for assessing optimality and rationality

Assessing decision optimality requires a model of the underlying computations. In this
study, we chose the DDM for several reasons. First, it provided a parsimonious account
of both the choice and RT data (Palmer et al., 2005; Ratcliff et al., 1999). Second, as
discussed in more detail below, the DDM and related accumulate-to-bound models have
provided useful guidance for identifying neural substrates of the decision process
(Roitman and Shadlen, 2002; Ding and Gold, 2010; Ding and Gold, 2012; Hanks et al.,
2011; Ratcliff et al., 2003; Rorie et al., 2010; Mulder et al., 2012; Summerfield and
Koechlin, 2010; Frank et al., 2015). Third, these models are closely linked to normative
theory, including under certain assumptions matching the statistical procedure known as
the sequential probability ratio test that can optimally balance the speed and accuracy of

uncertain decisions (Barnard, 1946; Wald, 1947; Wald and Wolfowitz, 1948). These
normative links were central to our ability to use the DDM to relate the monkeys'
behavior to different forms of reward optimization.

Assessing optimality also requires an appropriate definition of the optimization goal. In
our study, we mainly focused on the goal of maximizing reward rate (per trial or per unit
of time). Based on this definition, the monkeys showed suboptimal reward-context-
dependent adjustments. It is possible that the monkeys' were optimizing for a different
goal, such as accuracy or a competition between reward and accuracy ("COBRA,"
Maddox and Bohil, 1998). However, the monkeys' behavior was not consistent with
optimizing for these goals, either. Specifically, none of these goals would predict optimal
z adjustment that favors the small reward choice: accuracy maximization would require
unbiased decisions: me=0 and z=0.5, whereas COBRA would require z values with
smaller magnitude (between 0.5 and those predicted for reward maximization alone), but
still in the adaptive direction. Therefore, the monkeys' strategies were not consistent with
simply maximizing commonly considered reward functions.

Deviations from optimal behavior are often ascribed to a lack of effort or poor learning.
However, these explanations seem unlikely to be primary sources of suboptimality in our
study. For example, lapse rates, representing the overall ability to attend to and perform
the task, were consistently near zero for all three monkeys. Moreover, the monkeys'
reward outcomes (RTrial or RR with respect to optimal values) did not change
systematically with experience but instead stayed close to the optimal values. These
results imply that the monkeys understood the task demands and performed consistently
well over the course of our study. More importantly, the monkeys' reward-context-
dependent adjustments were adapted to changes in the reward function due to session-
specific reward ratios and motion coherences and the monkeys' daily variations of
perceptual sensitivity and speed-accuracy trade-offs (Fig. 6E, Fig. S8). Based on these
observations, we reasoned that the seemingly sub-optimal behaviors may instead reflect
a common, adaptive, rational strategy that aimed to attain good-enough outcomes.

The gradient-based, satisficing model we proposed was based on these considerations
to account for our results. We do not yet know how well this model generalizes to other
tasks and conditions, but it exemplifies an additional set of general principles for
assessing the rationality of decision-making behavior: an alternative satisficing (good-
enough) goal, potential heuristic strategies based on the properties of the utility function,
and flexible adaptation to changes in the external and internal conditions. Below we
discuss the assumptions of our proposed model in more detail.

Adjustments with respect to the gradient of the utility function

In general, finding rational solutions through trial-and-error or stepwise updates requires
a sufficient gradient in the utility function to drive learning (Sutton and Barto, 1998). Our
proposed scheme couples a standard gradient-following algorithm with principles that
have been used to explain and facilitate decisions with high uncertainties, time

pressures, and/or complexity to achieve a satisficing solution (Simon, 1966; Wierzbicki,
1982; Gigerenzer and Goldstein, 1996; Nosofsky and Palmeri, 1997; Goodrich et al.,
1998; Sakawa and Yauchi, 2001; Goldstein and Gigerenzer, 2002; Stirling, 2003;
Gigerenzer, 2010; Oh et al., 2016). This scheme complements but differs from a
previously proposed satisficing strategy to account for human subjects' suboptimal
calibration of the speed-accuracy trade-off via adjustments of the decision bounds of a
DDM that favor robust solutions given uncertainties about the inter-trial interval
(Zacksenhouse et al., 2010), In contrast, our proposed strategy focuses on reward-
biased behaviors for a given speed-accuracy tradeoff and operates on reward per trial,
which is, by definition, independent of inter-trial-interval. Our scheme was based on
three key assumptions, as follows.

Our first key assumption was that the starting point for gradient following was not the
unbiased state (i.e., me=0 and z=0.5) but an over-biased state. Notably, in many cases
the monkeys could have performed as well or better than they did, in terms of optimizing
reward rate, by making unbiased decisions. The fact that none did so prompted our
assumption that their session-by-session adjustments tended to reduce, not inflate,
biases. Specifically, we assumed that the initial experience of the asymmetric reward
prompted an over-reaction to bias choices towards the large-reward alternative. In
general, such an initial over-reaction is not uncommon, as other studies have shown
excessive, initial biases that are reduced or eliminated with training (Gold et al., 2008;
Jones, et al., 2015; Nikolaev et al., 2016). The over-reaction is also rational because the
penalty is larger for an under reaction than for an overreaction. For example, in the
average RTrial heatmaps for our task (Fig. 6A), the gradient dropped faster in the under-
biased side than in the over-biased side. This pattern is generally true for tasks with
sigmoid-like psychometric functions (for example, the curves in Supplementary Fig. 2).
Our model further suggests that the nature of this initial reaction, which may be driven by
individually tuned features of the reward function that can remain largely consistent even
for equal-reward tasks (Supplementary Fig. 12) and then constrain the end-points of a
gradient-based adjustment process (Fig. 8), may help account for the extensive
individual variability in biases that has been reported for reward-biased perceptual tasks
(Voss et al., 2004; Summerfield and Koechlin, 2010; Leite and Ratcliff, 2011; Cicmil et
al., 2015)).

Our second key assumption was that from this initial, over-biased state, the monkeys
made adjustments to both the momentary evidence (me) and decision rule (z) that
generally followed the gradient of the reward function. The proposed step-wise
adjustments occurred too quickly to be evident in behavior; e.g., the estimated biases
were similar for the early and late halves in a block (data not shown). Instead, our
primary support for this scheme was that the steady-state biases measured in each
session were tightly coupled to the shape of the reward function for that session. It would
be interesting to design tasks that might allow for more direct measurements of the
updating process itself, for example, by manipulating both the initial biases and relevant
reward gradient that might promote a longer adjustment process.

Our third key assumption was that the shallowness of the utility of the function around
the peak supported satisficing solutions. Specifically, gradient-based adjustments,
particularly those that use rapid updates based on implicit knowledge of the utility
function, may only be sensitive to relatively large gradients. For our task, the gradients
were much smaller around the peak, implying that there were large ranges of parameter
values that provided such similar outcomes that further adjustments were not used. In
principle, it is possible to change the task conditions to test if and how subjects might
optimize with respect to steeper functions around the peak. For example, for RTrial, the
most effective way to increase the gradient magnitude near the peak (i.e., reducing the
area of the dark red patch) is to increase sensory sensitivity (k) and cautiousness
(Supplementary Fig. 12; a; i.e., emphasizing accuracy over speed). For RR, the gradient
can also be increased by increasing the time-out penalty (not shown). Despite some
practical concerns about these manipulations (e.g., increasing time-out penalties can
decrease motivation), it would be interesting to study their effects on performance in
more detail. Such studies would help inform how constraints of the utility function govern
decision strategies, including when to use satisficing or "good enough" instead of optimal
ones like in this study (Simon, 1956; Simon, 1982).

Possible neural mechanisms

Our study benefitted greatly from the ability of each of the three monkeys to perform this
complex decision task nearly optimally and with few lapses of attention over many
sessions. It also benefitted from an RT task design that provided a richer set of
constraints for inferring characteristics of the underlying decision process than choice
data alone (Feng et al., 2009; Nomoto et al., 2010; Teichert and Ferrera, 2010).
Together, these benefits establish a strong foundation for planned, future studies of the
underlying neural mechanisms, which are informed by our present results.

Specifically, the DDM framework has been used effectively to identify and interpret
neural substrates of key computational components of the decision process for
symmetric-reward versions of the motion-discrimination task. For example, the
momentary evidence is represented, in part, by motion-sensitive neurons in the middle
temporal (MT) area of extrastriate visual cortex (Newsome et al., 1989). When the
decision is indicated with an eye movement, the accumulation of momentary evidence
into the decision variable is represented in several brain areas involved in preparing and
executing the eye-movement response, including parts of the parietal and prefrontal
cortex, the basal ganglia, and the superior colliculus (e.g. Shadlen and Newsome, 2001;
Ding and Gold, 2012; Ding and Gold, 2010; Horwitz and Newsome, 2001). In broad
strokes, the *me* adjustment that we found may be implemented as reward modulation of
MT output and/or synaptic weights for the connections between different MT
subpopulations and decision areas (Cicmil, et al., 2015). In contrast, the *z* adjustment
may be implemented as reward context-dependent baseline changes in neural
representations of the decision variable and/or the rule that determines the final choice
(Lo and Wang, 2006; Rao, 2010; Lo et al., 2015; Wei et al., 2015). Finally, modulation of

non-decision times may be implemented as reward modulation in higher-order motor-
planning regions and/or lower-level motor generation circuits.

To summarize, our results suggest that the monkeys' idiosyncratic decision-making
adjustments under changing contexts may reflect a rational heuristic strategy operating
on the gradient of their individual reward functions. This new view extends the more
conventional reward maximization-based test of rationality and may help to account for
individual variations in decision-making strategies. It also opens up new questions about
how the brain acquires, retains, and uses the knowledge needed to make such flexible
adjustments. Answers to these questions will add to our growing understanding of
increasingly complex forms of decision-making.

Methods

Subjects

We used three rhesus macaques (*Macaca mulatta*), two male and one female, to study
behavior on an asymmetric-reward reaction-time random-dot motion discrimination task
(Fig. 1B, see below). Prior to this study, monkeys F and C had been trained extensively
on the equal-reward RT version of the task (Ding and Gold, 2010, 2012b, a). Monkey A
had been trained extensively on non-RT dots tasks (Connolly et al., 2009; Bennur and
Gold, 2011), followed by >130 sessions of training on the equal-reward RT dots task. All
training and experimental procedures were in accordance with the National Institutes of
Health Guide for the Care and Use of Laboratory Animals and were approved by the
University of Pennsylvania Institutional Animal Care and Use Committee.

Behavioral task

Our task (Fig. 1B) was based on the widely used random-dot motion discrimination task
that typically has symmetric rewards (Roitman and Shadlen, 2002; Ding and Gold, 2010).
Briefly, a trial started with presentation of a fixation point at the center of a computer
screen in front of a monkey. Two choice targets appeared 0.5 s after the monkey
acquired fixation. After a delay, the fixation point was dimmed and a random-dot
kinematogram (speed: 6 °/s) was shown in a 5° aperture centered on the fixation point.
For monkeys F and C, the delay duration was drawn from a truncated exponential
distribution with mean=0.7 s, max=2.5 s, min=0.4 s. For monkey A, the delay was set as
0.75 s. The monkey was required to report the perceived global motion direction by
making a saccade to the corresponding choice target at a self-determined time (a 50-ms
minimum latency was imposed to discourage fast guesses). The stimulus was
immediately turned off when the monkeys' gaze left the fixation window (4, 4, and 3°
square windows for monkey F, C, and A, respectively). Correct choices (i.e., saccades to
the target congruent with actual motion direction) were rewarded with juice. Error
choices were not rewarded and instead penalized with a timeout before the next trial
began (timeout duration: 3 s, 0.5-2 s, and 2.5 s, for monkeys F, C, and A, respectively).

On each trial, the motion direction was randomly selected toward one of the choice
targets along the horizontal axis. The motion strength of the kinematogram was
controlled as the fraction of dots moving coherently to one direction (coherence). On
each trial, coherence was randomly selected from 0.032, 0.064, 0.128, 0.256, and 0.512
for monkeys F and C, and from 0.128, 0.256, 0.512, and 0.75 for monkey A. In a subset
of sessions, coherence levels of 0.064, 0.09, 0.35, and/or 0.6 were also used for monkey
A.

We imposed two types of reward context on the basic task. For the "LR-Left" reward
context, correct leftward saccades were rewarded with a larger amount of juice than
correct rightward saccades. For the "LR-Right" reward context, correct leftward
saccades were rewarded with a smaller amount of juice than correct rightward saccades.
The large:small reward ratio was on average 1.34, 1.91, and 2.45 for monkeys F, C, and
A, respectively. Reward context was alternated between blocks and constant within a
block. Block changes were signaled to the monkey with an inter-block interval of 5 s. The
reward context for the current block was signaled to the monkey in two ways: 1) in the
first trial after a block change, the two choice targets were presented in blue and green
colors, for small and large rewards, respectively (this trial was not included for analysis);
and 2) only the highest coherence level (near 100% accuracy) was used for the first two
trials after a block change to ensure that the monkey physically experienced the
difference in reward outcome for the two choices. For the rest of the block, choice
targets were presented in the same color and motion directions and coherence levels
were randomly interleaved.

Basic characterization of behavioral performance

Eye position was monitored using a video-based system (ASL) sampled at 240 Hz. RT
was measured as the time from stimulus onset to saccade onset, the latter identified
offline with respect to velocity ($> 40°/s$) and acceleration ($> 8000°/s^2$). Performance was
quantified with psychometric and chronometric functions (Figs. 2 and 3), which describe
the relationship of motion strength (signed coherence, *Coh*, which was the proportion of
the dots moving in the same direction, positive for rightward motion, negative for leftward
motion) with choice and RT, respectively. Psychometric functions were fitted to a logistic
function (Equation (1)), in which $\boldsymbol{\lambda}$ is the error rate, or lapse rate, independent of the
motion information; $\boldsymbol{\alpha}_0$ and $(\boldsymbol{\alpha}_0 + \boldsymbol{\alpha}_{rew})$ are the bias terms, which measures the coherence
at which the performance was at chance level in the LR-Right and LR-Left reward
contexts, respectively. $\boldsymbol{\beta}_0$ and $(\boldsymbol{\beta}_0 + \boldsymbol{\beta}_{rew})$ are the perceptual sensitivities in the LR-Right
and LR-Left reward contexts, respectively.

$$P_{rightward\ choice} = \lambda + (1 - 2\,\lambda) \times \frac{1}{e^{-Sensitivity(Coh-Bias)}} \quad (1)$$

Reward-biased drift-diffusion model

To infer the computational strategies employed by the monkeys, we adopted the widely
used accumulation-to-bound framework, the drift-diffusion model (DDM; Fig. 1A). In the

standard DDM, motion evidence is modeled as a random variable following a Gaussian
distribution with a mean linearly proportional to the signed coherence and a fixed
variance. The decision variable (DV) is modeled as temporal accumulation (integral) of
the evidence, drifting between two decision bounds. Once the DV crosses a bound,
evidence accumulation is terminated, the identity of the decision is determined by which
bound is crossed, and the decision time is determined by the accumulation time. RT is
modeled as the sum of decision time and saccade-specific non-decision times, the latter
accounting for the contributions of evidence-independent sensory and motor processes.

To model the observed influences of motion stimulus and reward context on monkeys'
choice and RT behavior, we introduced two reward context-dependent terms: z specifies
the relative bound heights for the two choices and me specifies the equivalent
momentary evidence that is added to the motion evidence at each accumulating step.
Thus, for each reward context, six parameters were used to specify the decision
performance: a: total bound height; k: proportional scaling factor converting evidence to
the drift rate; t_0 and t_1: non-decision times for leftward and rightward choices,
respectively; and z and me. Similar approaches have been used in studies of human
and animal decision making under unequal payoff structure and/or prior probabilities
(Voss et al., 2004; Bogacz et al., 2006; Diederich and Busemeyer, 2006; Summerfield
and Koechlin, 2010; Hanks et al., 2011; Mulder et al., 2012).

To fit the monkeys' data, we implemented hierarchical DDM fitting using an open-source
package in Python, which performs Bayesian estimates of DDM parameters based on
single-trial RTs (Wiecki et al., 2013). This method assumes that parameters from
individual sessions are samples from a group distribution. The initial prior distribution of
a given parameter is determined from previous reports of human perceptual
performance and is generally consistent with monkey performance on equal reward
motion discrimination tasks (Ding and Gold, 2010; Matzke and Wagenmakers, 2009).
The posterior distributions of the session- and group-level parameters are estimated with
Markov chain Monte Carlo sampling.

For each dataset, we performed 5 chains of sampling with a minimum of 10000 total
samples (range: 10000-20000; burn-in: 5000 samples) and inspected the trace,
autocorrelation and marginal posterior histogram of the group-level parameters to detect
signs of poor convergence. To ensure similar level of convergence across models, we
computed the Gelman-Rubin statistic (R-hat) and only accepted fits with R-hat<1.01.

To assess whether reward context modulation of both z and me was necessary to
account for monkeys' behavioral data, we compared fitting performance between the
model with both terms ("full") and reduced models with only one term ("z-only" and "me-
only"). Model selection was based on the deviance information criterion (DIC), with a
smaller DIC value indicating a preferred model. Because DIC tends to favor more
complex models, we bootstrapped the expected ΔDIC values, assuming the reduced
models were the ground truth, using trial-matched simulations. For each session, we
generated simulated data using the DDM, with single-session parameters fitted by me-

only or *z*-only HDDM models and with the number of trials for each direction ×
coherence × reward context combination matched to the monkey's data for that session.
These simulated data were then re-fitted by all three models to estimate the predicted
∆DIC, assuming the reduced model as the generative model.

To test an alternative model, we also fitted monkeys' data to a DDM with collapsing
bounds (Zylberberg et al., 2016). This DDM was constructed as the expected first-
stopping-time distribution given a set of parameters, using the PyMC module (version
2.3.6) in Python (version 3.5.2). The three model variants, "full", "*me*-only" and "*z*-only",
and their associated parameters were the same as in HDDM, except that the total bound
distance decreases with time. The distance between the two choice bounds is set as
$a/(1 + e^{\beta(t-d)})$, where a is the initial bound distance, β determines the rate of collapsing,
and d determines the onset of the collapse. Fitting was performed by computing the
maximum *a posteriori* estimates, followed by Markov chain Monte Carlo sampling, of
DDM parameters given the experimental RT data.
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Optimality analysis

To examine the level of optimality of the monkeys' performance, we focused on two
reward functions: reward rate (RR, defined as the average reward per second) and
reward per trial (RTrial, defined as the average reward per trial) for a given reward
context for each session. To estimate the reward functions in relation to *me* and *z*
adjustments for a given reward context, we numerically obtained choice and RT values
for different combinations of *z* (ranging from 0 to 1) and *me* (ranging from -0.6 to 0.6
coherence unless otherwise specified), given *a, k* and non-decision time values fitted by
the full model. We then calculated RR and RTrial, using trial-matched parameters,
including the actual ITI, timeout, and large:small reward ratio. RR_{max} and $RTrial_{max}$ were
identified as the maximal values given the sampled *me-z* combinations, using 1000 trials
for each coherence × direction condition. Optimal *me* and *z* adjustments were defined as
the *me* and *z* values corresponding to RR_{max} or $RTrial_{max}$. $RR_{predict}$ and $RTrial_{predict}$ were
calculated with the fitted *me* and *z* values in the full model.

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601   **Competing interests**

608   The authors declare that no competing interests exist.

609   **References**
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 Figure legends

 Figure 1. Theoretical framework and task design.

A, Schematics of the drift-diffusion model (DDM). Motion evidence is modeled as
samples from a unit-variance Gaussian distribution (mean: signed coherence, *Coh*).
Effective evidence is modeled as the sum of motion evidence and an internal
momentary-evidence bias (*me*). The decision variable starts at value a×*z*, where *z*
governs decision-rule bias, and accumulates effective evidence over time with a
proportional scaling factor (*k*). A decision is made when the decision variable reaches
either bound. Reaction time (RT) is assumed to be the sum of the decision time and a
saccade-specific non-decision time.

B, Reaction-time (RT) random-dot visual motion direction discrimination task with
asymmetric rewards. A monkey makes a saccade decision based on the perceived
global motion of a random-dot kinematogram. Reward is delivered on correct trials and
with a magnitude that depends on reward context. Two reward contexts (LR-Left and
LR-Right) were alternated in blocks of trials with signaled block changes. Motion
directions and strengths were randomly interleaved within blocks.

Figure 2. Relationships between sensitivity and bias from logistic fits to choice
data.

A, For each monkey, the probability of making a rightward choice is plotted as a function
of signed coherence (–/+ indicate left/right motion) from all sessions, separately for the
two reward contexts, as indicated. Lines are logistic fits.

B, Top row: Motion sensitivity (steepness of the logistic function) in each context as a
function of session index (colors as in A). Solid lines indicate significant positive partial
Spearman correlation after accounting for changes in reward ratio across sessions
(*p*<0.05). Black dashed lines indicate each monkey's motion sensitivity in the same task
with equal rewards before training on this asymmetric reward task. Middle row: ΔBias
(horizontal shift between the two psychometric functions in the two reward contexts at
chance level) as a function of session index. Solid line indicates significant negative
partial Spearman correlation after accounting for changes in reward ratio across
sessions (*p*<0.05). Bottom row: Lapse rate as a function of session index (median=0 for
all three monkeys).

C, ΔBias as a function of motion sensitivity for each reward context (colors as in A). Solid
line indicates a significant negative partial Spearman correlation after accounting for
changes in reward ratio across sessions (*p*<0.05).

D, Optimal versus fitted Δbias. Optimal Δbias was computed as the difference in the
horizontal shift in the psychometric functions in each reward context that would have
resulted in the maximum reward per trial, given each monkey's fitted motion sensitivity

and experienced values of reward ratio and coherences from each session (see Fig.
S2a). Solid lines indicate significant negative Spearman correlation (p<0.01). Partial
Spearman correlation after accounting for changes in reward ratio across sessions are
also significant for money F and C ($p<0.01$).

Figure 3. Comparison of choice and RT data to HDDM fits with both momentary-
evidence (*me*) and decision-rule (*z*) biases.

A, Psychometric data (points as in Fig. 2A) shown with predictions based on HDDM fits
to both choice and RT data.

B, RT data (circles) and HDDM-predicted RT distributions (lines). Both sets of RT data
were plotted as values corresponding to the 20th, 40th, 60th, and 80th percentiles of the
full distribution. Top row: Trials in which monkey chose the left target. Bottom row: Trials
in which monkeys chose the right target. Columns correspond to each monkey (as in A),
divided into choices in the large- (left column) or small- (right column) reward direction
(correct/error choices are as indicated in the left-most columns; note that no reward was
given on error trials). The HDDM-predicted RT distributions were generated with 50 runs
of simulations, each run using the number of trials per condition (motion direction $\times$
coherence $\times$ reward context $\times$ session) matched to experimental data and using the
best-fitting HDDM parameters for that monkey.

Figure 4. Actual versus optimal adjustments of momentary-evidence (*me*) and
decision-rule (*z*) biases.

A, Schematic of the comparison procedure. Choice and RT data from the two reward
contexts in a given session were fitted separately using the HDDM. These context- and
session-specific best-fitting *me* and *z* values are plotted as the monkey's data (black
circles in *B* and *C*). Optimal values were determined by taking the best-fitting values of
parameters *a, k,* and non-decision times from the HDDM fits, then sampling *me* and *z*
uniformly over broad ranges to predict choice and RT behavior. For each *me* and *z*
combination, the predicted probability of left/right choice and RTs were used with actual
task information (inter-trial interval, error timeout and reward sizes) to calculate the
expected reward rate (RR) and average reward per trial (RTrial). Optimal *me*/*z*
adjustments were then identified to maximize RR (purple) or RTrial (orange).

B, Scatterplots of the monkeys' *me*/*z* adjustments (black), predicted optimal adjustments
for maximal RR (purple), and predicted optimal adjustments for maximal RTrial (orange),
for the two reward contexts in all sessions (each data point was from a single session).
Values of *me*>0 or z>0.5 produce biases favoring rightward choices.

C, Scatterplots of the differences in *me* (abscissa) and *z* (ordinate) between the two
reward contexts for monkeys (black), for maximizing RR (purple), and for maximizing

RTrial (orange). Positive Δme and Δz values produce biases favoring large-reward
choices.

Figure 5. Predicted versus optimal reward per trial (RTrial).

A, Scatterplots of $RTrial_{predict}$:$RTrial_{max}$ ratio as a function of session index. Each session
was represented by two ratios, one for each reward context. Mean ratio across contexts
and sessions: 0.977 for monkey F, 0.984 for monkey C, and 0.983 for monkey A.

B, 97% $RTrial_{max}$ contours for all sessions, computed using the best-fitting HDDM
parameters and experienced coherences and reward ratios from each session. Light
grey: LR-Left blocks; Dark grey: LR-Right blocks.

C, The monkeys' adjustments (blue in LR-Left blocks, red in LR-Right blocks) were
largely within the 97% $RTrial_{max}$ contours for all sessions and tended to cluster in the *me*
over-biased, *z* under-biased quadrants (except Monkey F in the LR-Right blocks). The
contours and monkeys' adjustments are centered at the optimal adjustments for each
session.

Figure 6. Relationships between adjustments of momentary-evidence (*me*) and
decision-rule (*z*) biases and RTrial function properties.

A, Mean RTrial as a function of *me* and *z* adjustments for the LR-Left (top) and LR-Right
(bottom) blocks. RTrial was normalized to $RTrial_{max}$ for each session and then averaged
across sessions.

B, Scatterplot of the total variance in *me* and *z* adjustments across sessions (ordinate)
and the area of ≥97% max of the average RTrial patch (abscissa). Variance and patch
areas were measured separately for the two reward blocks (circles for LR-Left blocks,
squares for LR-Right blocks).

C, D, The monkeys' session- and context-specific values of *me* (C) and *z* (D) co-varied
with the orientation of the >97% heatmap patch (same as the contours in Fig. 5B).
Orientation is measured as the angle of the tilt from vertical. Circles: data from LR-Left
block; squares: data from LR-Right block; lines: significant correlation between *me* (or *z*)
and patch orientations across monkeys ($p<0.05$). Colors indicate different monkeys (see
legend in B).

E, Scatterplots of conditionally optimal versus fitted Δme (top row) and Δz (bottom row).
For each reward context, the conditionally optimal *me (z)* value was identified given the
monkey's best-fitting z (*me*) values. The conditionally optimal Δme (Δz) was the
difference between the two conditional optimal *me (z)* values for the two reward contexts.
Grey lines indicate the range of conditional Δme (Δz) values corresponding to the 97%

maximal RTrial given the monkeys' fitted z (*me*) values.

Figure 7. Relationships between starting and ending values of the satisficing, reward function gradient-based updating process.

Example gradient lines of the average RTrial maps for the three monkeys are color
coded based on the end point of gradient-based *me* and *z* adjustments in the following
ways: 1) *me* biases to large reward whereas *z* biases to small reward (magenta); 2) *z*
biases to large reward whereas *me* biases to small reward (blue); 3) *me* and *z* both bias
to large reward (green), and 4) *me* and *z* both bias to small reward (yellow). The gradient
lines ended on the 97% $RTrial_{max}$ contours. Top row: LR-Left block; bottom row: LR-
Right block.

Figure 8. The satisficing reward function gradient-based model.

A, Illustration of the procedure for predicting a monkey's *me* and *z* values for a given
RTrial function. For better visibility, RTrial for the LR-Left reward context in an example
session is shown as a heatmap in greyscale. Gradient lines are shown as black lines.
The square indicates the unbiased *me* and *z* combination (average values across the
two reward contexts). The four trajectories represent gradient-based searches based on
four alternative assumptions of initial values (see table on the right). Open circles
indicate the end values. Grey filled circle indicates the monkey's actual *me* and z. Note
that the end points differ among the four assumptions, with the magenta circle being the
closest to the monkey's fitted *me* and *z* of that session.

B, Scatterplots of the predicted and actual Δ*me* and Δz between reward contexts. Grey
circles here are the same as the black circles in Fig. 4C. Colors indicate model identity,
as in A.

C, Average regression coefficients between each monkey's Δ*me* (left four bars) and Δ*z*
(right four bars) values and predicted values for each of the four models. Filled bars: *t-*
test*, p* < 0.05*.*

D, Covariation of *me* (top) and *z* (bottom) with the orientation of the >97% RTrial
heatmap patch for monkeys and predictions of the four models. Blue: data from LR-Left
blocks, red: data from LR-Right blocks. Data in the "Monkey" column are the same as in
Fig 6C and D. Note that predictions of the "over-*me*" model best matched the monkey
data than the other models.

 Supplementary Figure legends

 Figure S1. Logistic regression fits.

A, Choice data from example sessions for each of the three monkeys, plotted as in Fig.
2A. Lines: fits with logistic regressions.

B, Scatterplots of mean likelihood per trial, assuming binomial errors. Red circles
indicate the sessions shown in A.

Figure S2.

A, Identification of the optimal Δbias for an example session using logistic fits. For each
reward context (blue for LR-Left and red for LR-Right), RTrial was computed as a
function of bias values sampled uniformly over a broad range, given the session-specific
sensitivities, lapse rate, coherences and large:small reward ratio. The optimal Δbias was
defined as the difference between the bias values with the maximal RTrial for the two
reward contexts. The fitted Δbias was defined as the difference between the fitted bias
values for the two reward contexts.

B, The optimal bias decreases with increasing sensitivity. The example heatmap shows
normalized RTrial as a function of sensitivity and bias values in the LR-Right blocks,
assuming the same coherence levels as used for the monkeys and a large:small reward
ratio of 2.3. The black curve indicates the optimal bias values for a given sensitivity value.

C, Scatterplots of optimal Δbiases obtained via the procedure described above as a
function of sensitivity for each of the two reward contexts. Same format as Figure 3B.
Solid lines indicate significant partial Spearman correlation after accounting for changes
in reward ratio across sessions ($p<0.05$). Note that the scatterplots of the monkeys'
Δbiases and sensitivities in Fig. 2C also show negative correlations, similar to this
pattern.

Figure S3. Qualitative comparison between the monkeys' RT distribution and DDM
predictions.

A, RT distributions as predicted by a DDM with no bias in decision rule (z) or momentary
evidence (me; left), with $me>0$ (middle), and with $z>0.5$ (right). RT distributions are
shown separately for correct (red) and error (black) trials and using values
corresponding to 20th, 40th, 60th, and 80th percentiles. Positive/negative coh values
indicate rightward/leftward saccades. The values of me and z were chosen to induce
similar choice biases (~0.075 in coherence units). Note that the me bias induces large
asymmetries in RT both between the two choices and between correct and error trials,
whereas the z bias induces a large asymmetry in RT for the two choices, but with little

asymmetry between correct and error trials.

B, The monkeys' mean RTs for four quantiles for the LR-Right (*top*) and LR-Left (*bottom*)
reward contexts, respectively (same convention as in A). Note the presence of
substantial asymmetries between correct and error trials for all three monkeys.

Figure S4. Fits to a DDM with collapsing bounds.

A, B, A DDM with collapsing bounds and both momentary evidence (*me*) and decision-
rule (*z*) biases fit to each monkey's RT data. Same format as Fig. 3.

C, The model that included both *me* and *z* adjustments ("full") had smaller Akaike
Information Criterion (AIC) values than reduced models ("me" or "z" only) across
sessions. Note also the different ranges of ΔAIC for the full–*me* and full–z comparisons.
The mean ΔAIC$_{full-me}$ and ΔAIC$_{full-z}$ values are significantly different from zero (Wilcoxon
signed rank test, *p*=0.0007 for Monkey F's full–*me* comparison and *p*<0.0001 for all
others).

D, RT distributions as predicted by the DDM with collapsing bounds, using no bias in *z*
or *me* (*left*), *me*>0 (*middle*), or *z*>0.5 (*right*). Same format as Fig. S2A.

1002

Figure S5. Estimates of momentary-evidence (*me*) and decision-rule (*z*) biases
using the collapsing-bound DDM fits. Same format as Fig. 4B and C, except here
1005 only showing fits to the monkeys' data. As with the model without collapsing bounds, the
1006 adjustments in *me* tended to favor the large reward but the adjustments in *z* tended to
1007 favor the small reward.

1008

Figure S6. Predicted versus optimal reward rate (RR). Same format as Fig. 5. Mean
1010 RR$_{predict}$:RR$_{max}$ ratio across sessions=0.970 for monkey F, 0.980 for monkey C, and
1011 0.979 for monkey A.

1012

Figure S7. Correlation between fitted and conditionally optimal adjustments was
stronger for matched sessions than for unmatched sessions. A, C: Momentary-
evidence (Δ*me*) adjustments. B, D: Decision-rule (Δ*z*) adjustments. A, B: optimal values
obtained with the RTrial function. C, D: optimal values obtained with the RR function.
Red lines indicate the partial Spearman correlation coefficients between the fitted and
optimal Δ*me* or Δ*z* (obtained in the same way as data in Fig. 6E) for matched sessions.
Bars represent the histograms of partial correlation for unmatched sessions, which were
obtained by 100 random shuffles of the sessions). Note that the correlation values for
unmatched sessions were lower than those for matched sessions (Wilcoxson rank-sum

test, $p<0.001$ for all three monkeys and both Δme and Δz).

1023

1024   **Figure S8. The monkeys' momentary-evidence (*me*) and decision-rule (*z*)**
1025   **adjustments reflected RR function properties.** Same format as Fig. 6, but using RR
1026   instead of RTrial.

1027

Figure S9: The HDDM model fitting procedure does not introduce spurious
correlations between patch orientation and *me* value. Artificial sessions were
simulated with fixed *me* values (±0.1 for the two reward contexts) and different *k* values.
A: Recovered *k* values from HDDM fitting closely matched *k* values used for the
simulations. B: Recovered *me* values from HDDM fitting closely matched *me* values
used for simulation and did not correlate with RTrial patch orientation.

1034

1035   **Figure S10. RR gradient trajectories color-coded by the end points of the *me/z***
1036   **patterns.** Same format as Fig. 7 but using gradients based on RR instead of RTrial.

1037

1038   **Figure S11. Predictions of a RR gradient-based model.** Same format as Fig. 8 but
1039   using gradients based on RR instead of RTrial.

1040

Figure S12. Dependence of the orientation and area of the near-optimal RTrial
patch on parameters reflecting internal decision process and external task
specifications.

The top four rows show the RTrial heatmaps with two values of a single parameter
indicated above, while keeping the other parameters fixed at the baseline values. The
third and fourth rows show the estimated orientation (the slope value from a linear
regression analysis) and area (in pixels), respectively, of the image patches
corresponding to ≥97% of $RTrial_{max}$. The baseline values of the parameters are: *a*=1.5,
k=6, non-decision times=0.3 sec for both choices, *ITI*=4 sec, *Timeout*=8 sec, *large-*
reward (LR): small-reward (SR) ratio=2.
1051

1052 **Supplementary Table 1.**

1053 Best-fitting parameters of HDDM.

| | Monkey F (26079 trials) | | | | Monkey C (37161 trials) | | | | Monkey A (21089 trials) | | | |
| | LR-Left | | LR-Right | | LR-Left | | LR-Right | | LR-Left | | LR-Right | |
	Mean	Std	Mean	Std	Mean	Std	Mean	Std	Mean	Std	Mean	Std
a	1.67	0.16	1.43	0.12	1.77	0.09	1.53	0.13	1.33	0.13	1.36	0.09
k	10.22	1.87	9.91	2.11	6.58	0.51	5.08	0.92	4.04	0.33	3.45	0.46
t_1	0.31	0.03	0.29	0.03	0.35	0.04	0.33	0.05	0.29	0.04	0.27	0.04
t_0	0.28	0.04	0.31	0.05	0.33	0.04	0.31	0.03	0.21	0.08	0.26	0.04
z	0.6	0.03	0.57	0.04	0.62	0.03	0.2	0.04	0.57	0.06	0.39	0.04
me	-0.06	0.04	0.08	0.05	-0.14	0.04	0.21	0.06	-0.22	0.05	0.27	0.09

1054

Supplementary Table 2.

The difference in deviance information criterion (ΔDIC) between the full model (i.e., the
model that includes both *me* and *z*) and either reduced model (*me*-only or *z*-only), for
experimental data and data simulated using each reduced model. Negative values favor
the full model. Positive values favor the reduced models. Note that the ΔDIC for the
experimental data deviated substantially - toward the negative direction - from the ΔDIC
for the simulated data.

1062

	Experimental Data				Simu: *me* model		Simu: *z* model	
	ΔDIC: full–*me*		ΔDIC: full–*z*		ΔDIC: full–*me*		ΔDIC: full–*z*	
	Mean	Std	Mean	Std	Mean	Std	Mean	Std
Monkey F	-124.6	2.7	-2560.4	3.9	3.1	8.7	0.2	10.5
Monkey C	-1700.4	2.4	-6937.9	0.9	17.5	10.1	1.8	1.2
Monkey A	-793.6	2.7	-2225.7	3.1	25.4	8.0	1.2	3.1

1063

Figure 1

A

Evidence

me

Coh

0 → t

Decision time — Right

Decision Variable

a

Slope = k(Coh + me)

z x a

t

Left

B

Dots

RT

Reward

Saccade

Dots

Target

Fixation

Time

+

+

Left   Right

LR-Left blocks: Small Rew   Large Rew
LR-Right blocks: Large Rew   Small Rew

Figure 2

Logistic Fits

A

Monkey F   Monkey C   Monkey A

Prob. Right

LR-Right
LR-Left

Left ← Coh → Right

B

Sensitivity

ΔBias

Lapse Rate

Session   Session   Session

C

ΔBias

Sensitivity   Sensitivity   Sensitivity

D

Optimal Bias

Fit Bias   Fit Bias   Fit Bias

Figure 3

HDDM Fits

Figure 4

A

B

Figure 5

Figure 6

Figure 7

Trajectory end-point	me	z
———	Over biased	Under biased
———	Under biased	Over biased
———	Over biased	Over biased
———	Under biased	Under biased

Figure 8

Figure S1

A

Monkey F   Monkey C   Monkey A

Prob. right

0.5

−0.5 0 0.5 −0.5 0 0.5 −0.5 0 0.5

Left ← Coh → Right

B

Likelihood per trial

0.5

0   50   0   20   40   0   20   40

Early ← Session → Late

Figure S2

A

RTrial normalized

Optimal
ΔBias

LR-Right

Fitted
ΔBias

LR-Left

−0.5   0   0.5
Left ←   Bias   → Right

B

LR−Right

Bias → LR
SR ←

0.5

−0.5

Low ← Sensitivity → High

5   10   15

C

Monkey F

Optimal ΔBias

0.6

0.4

0.2

20   40
Sensitivity

Monkey C

5   10   15
Sensitivity

Monkey A

0   5   10
Sensitivity

Figure S3

A

No bias   ΔME bias   Δz bias

Quantile RT (s)

SR   LR   Correct / Error

Left Saccade ←Coh→ Right Saccade

B

Monkey F   Monkey C   Monkey A

LR-Left Quantile RT (s)

LR   SR

SR-Right Quantile RT (s)

SR   LR

Left Saccade ←Coh→ Right Saccade

Figure S4

Collapsing-bound Fits

Figure S5

Figure S6

Figure S7

Figure S8

Figure S9

Figure S10

Trajectory end-point	me	z
——— (magenta)	Over biased	Under biased
——— (blue)	Under biased	Over biased
——— (green)	Over biased	Over biased
——— (yellow)	Under biased	Under biased

Figure S11

A

Model	Initial Dimension	Overshoot?
neutral	Both	No
over-me	me	Yes
over-z	z	Yes
over-both	Both	Yes

- ● monkey
- ○ model end point
- □ baseline

B

Monkey F   Monkey C   Monkey A

C

D

Monkey Neutral Over-me Over-z Over-both

Patch orientation (degree from verticle)

Figure S12