1	Genome-wide signals of drift and local adaptation during rapid lineage divergence
2	in a songbird
3	
4	Running title: Drift and selection in a songbird radiation
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## 22 Abstract

23 The formation of independent evolutionary lineages involves neutral and selective factors, and understanding their relative roles in population divergence is a fundamental 24 25 goal of speciation research. Correlations between allele frequencies and environmental variability can reveal the role of selection, yet the relative contribution of drift can be 26 27 difficult to establish. Recently diversified systems such as that of the Oregon junco 28 (Aves: Emberizidae) of western North America provide ideal scenarios to apply genetic-environment association analyses (GEA) while controlling for population 29 structure. Genome-wide SNP loci analyses revealed marked genetic structure consisting 30 31 of differentiated populations in isolated, dry southern mountain ranges, and more 32 admixed recently expanded populations in humid northern latitudes. We used 33 correlations between genomic and environmental variance to test for three specific 34 modes of evolutionary divergence: (i) drift in geographic isolation, (ii) differentiation along continuous selective gradients, and (iii) isolation by adaptation. We found 35 36 evidence of strong drift in southern mountains, but also signals of local adaptation in several populations, driven by temperature, precipitation, elevation and vegetation, 37 especially when controlling for population history. We identified numerous variants 38 39 under selection scattered across the genome, suggesting that local adaptation can promote rapid differentiation over short periods when acting over multiple independent 40 41 loci. 42 Key words: local adaptation, isolation by adaptation, drift, selective gradients,

43 redundancy analysis, postglacial expansion

### 44 Introduction

45 Lineage diversification involves both selective and neutral factors, and elucidating their relative strengths and interactions in the process of evolutionary divergence is essential 46 47 to understand the mechanisms underlying the early stages of speciation (Covne and Orr 2004; Nosil 2012). Lineage differentiation can be driven by divergent natural selection 48 (Darwin 1859; Coyne and Orr 2004), a mechanism that is at the core of 'ecological 49 50 speciation' models, in which reproductive isolation arises as a by-product of cumulative, ecologically adaptive changes (Mayr 1947; Schluter 2000; Rundle and 51 Nosil 2005). In turn, accumulation of genetic differences caused by drift in geographic 52 53 isolation or in isolation-by-distance (IBD, Wright 1943; Wright 1946) has been proposed as a mode of divergence driven by neutral factors (Mayr 1954, 1963), a 54 55 mechanism that can be particularly strong in populations of small effective size (e.g. 56 Carson 1975; Templeton 1981; Uyeda et al. 2009). Selection and drift can also act jointly and even interact during evolutionary divergence in a number of ways. 57 58 Geographic distance usually implies environmental differences that may drive adaptation to local conditions and ecological differentiation, even if populations are 59 connected by moderate gene flow (Schluter 2000; Rundle and Nosil 2005). The 60 61 evolution of local adaptation can in turn result in isolation-by-adaptation (IBA), a mode of divergence where adaptive changes lead to intrinsic barriers to gene flow, enabling 62 genome-wide differentiation at both neutral and selected loci (Nosil et al. 2008; Funk et 63 64 al. 2011). Consequently, geographic distance and ecological divergence may promote 65 similar patterns of genetic diversity among populations, so that teasing apart the roles of 66 neutral evolution and ecological adaptation in evolutionary diversification requires approaches that account for both environmental heterogeneity and neutral population 67

structure (Wang and Bradburd 2014; Frichot et al. 2015; Rellstab et al. 2015; Forester et
al. 2016).

70

71 Our capacity to assess the relative roles of adaptation and neutral differentiation in driving population divergence has benefitted from our increased ability to survey 72 73 genome-wide variation thanks to the development of next generation sequencing (NGS) 74 techniques (McCormack et al. 2013; Faria et al. 2014). The increasingly large number of loci afforded by NGS provides improved resolution to detect neutral population 75 structure and patterns of gene flow among differentiated lineages. In addition, highly 76 77 differentiated loci identified as outliers in an  $F_{ST}$  distribution can be interpreted as potential targets of divergent selection standing out in a background of balanced or 78 79 neutrally maintained genomic variation (Faria et al. 2014; Rellstab et al. 2015). 80 However, methods of outlier detection relying solely on allele frequencies are sensitive to the confounding effects of historical factors, such as past sudden changes in 81 82 population size or strong drift in small populations that may result in high rates of false positives (Edmonds et al. 2004; Kawecki and Ebert 2004; Billiard et al. 2005; Christmas 83 et al. 2016). Moreover, changes in allele frequencies due to local adaptation can 84 85 sometimes go undetected by outlier analyses (Pritchard and Di Rienzo 2010; Bierne et al. 2011; Rellstab et al. 2015; Forester et al. 2017). Alternative approaches that integrate 86 environmental parameters by identifying allele frequencies that correlate with 87 ecological variability have proven useful to detect signals of adaptation, especially when 88 89 selective forces are weak (Frichot et al. 2015; Rellstab et al. 2015; Forester et al. 2017). 90 These methods, known as genetic-environment association (GEA, Hedrick et al. 1976; Mitton et al. 1977) analyses, have the potential to reveal genetic patterns of 91 92 differentiation due to local adaptation while testing for the role of multiple, specific

93	environmental variables as drivers of selection (Forester et al. 2017). Importantly, GEA
94	methods can correct for population history by controlling for general patterns of neutral
95	genomic variation (Rellstab et al. 2015; Forester et al. 2016), allowing us to separate the
96	respective effects of drift and selection in generating and maintaining variability. GEA
97	analyses have greatly benefitted from the development of high-throughput sequencing
98	techniques, resulting in a number of studies focusing on the genomic variability
99	associated with environmental parameters in groups as diverse as plants (Lasky et al.
100	2012; Jones et al. 2013; De Kort et al. 2014; Nadeau et al. 2016; Sork et al. 2016),
101	fungus (Ojeda Alayon et al. 2017), wolves (Forester et al. 2017) and birds (Manthey and
102	Moyle 2015; Safran et al. 2016; Szulkin et al. 2016; Termignoni-García et al. 2017).
103	
104	Recently diversified systems provide an ideal scenario for studying the relative roles of
105	selective and neutral factors in incipient divergence and speciation. Specifically, GEA
106	methods are particularly suitable when the system under study (i) is composed of
107	closely related populations, among which the signals of selection are still recent and
108	detectable; (ii) includes broad geographic distributions encompassing heterogeneous
109	habitats across ecological clines ( <i>i.e.</i> selective gradients) but also spatially
110	discontinuous habitats so that adaptive and neutral divergence can be assessed in
111	different spatial settings; (iii) shows large variability in the degree of geographical
112	isolation among populations, from extensive gene flow to total isolation; and (iv)
113	presents low variability in secondary sexual traits so that differential sexual selection
114	can be ruled out as a major driver of population divergence.
115	
116	The Oregon junco complex (Junco hyemalis oreganus) of western North America

117 provides a particularly well suited system to carry out genome-environment association

analysis. The complex originated recently as part of the postglacial radiation of dark-118 119 eved juncos across North America following a northward recolonization of the 120 continent as ice sheets retreated after the last glacial maximum, c.a. 18,000 years ago 121 (Milá et al. 2007; Friis et al. 2016; Milá et al. 2016). Among dark-eyed junco forms, the Oregon junco group presents the highest variability in terms of phenotype and 122 123 ecological range, encompassing a broad latitudinal range from Baja California to Alaska 124 (Fig. 1A). All forms of the Oregon junco share a characteristic dark hood, yet there is considerable population variation in plumage color, mainly of the hood, dorsum and 125 flanks, and the complex has been traditionally divided into at least 7 subspecific forms 126 127 (Dwight 1918; Miller 1941; Nolan et al. 2002), which include, from south to north: townsendi, from the San Pedro Mártir mountains in northern Baja California, Mexico; 128 pontilis, distributed just north of townsendi in the Sierra Juárez mountains, also in Baja 129 130 California, Mexico; thurberi, from the mountains of southern California and Sierra Nevada; pinosus, a coastal form from central California, predominantly distributed in 131 132 the Santa Cruz mountains; montanus, distributed across the interior of Oregon, 133 Washington and British Columbia; shufeldti, a more coastal form from Oregon and Washington; and oreganus from coastal British Columbia and southern Alaska (Miller 134 135 1941; Nolan et al. 2002; Fig. 1A, Fig. 2A, Table 1).

136

The diverse spatial configuration of populations and environmental variability across
the Oregon junco distribution are critical aspects that will affect our capacity to
disentangle the roles of adaptive and neutral factors in explaining genomic variance.
Here we use a conceptual framework to classify into three main settings the distinct
spatial scenarios observable in the Oregon junco with respect to gene flow and
environmental variation. These include (i) geographically isolated populations in similar

habitats, as in the case of the Baja California townsendi and pontilis forms, where low 143 144 levels of local adaptation and low rates of gene flow should result in limited adaptive divergence and high neutral divergence by drift (Fig. 2A); (ii) parapatric populations 145 146 under divergent ecological conditions, as exemplified by the *pinosus* and *thurberi* forms in California, where divergence is expected to increase due to local adaptation, while 147 148 geographic proximity and moderate gene flow should lead to intermediate levels of 149 neutral differentiation by drift (Fig. 2B); and (iii) populations found along a continuous environmental gradient, as in the northernmost forms of Oregon junco (thurberi, 150 151 shufeldti, montanus and oreganus) where neutral divergence is expected to be low due 152 to high levels of gene flow, while local adaptation along the gradient may result in a pattern of high differentiation in adaptive variation (Fig. 2C). 153 154 155 Here we use the Oregon junco complex to study how geographic isolation, population 156 history, and local ecological adaptation have driven population differentiation across the 157 range, using extensive population sampling and genome-wide SNPs obtained from 158 'genotyping by sequencing' (GBS, Elshire et al. 2011). A draft consensus genome of junco has also been sequenced and assembled to be used as a reference. We first assess 159 160 patterns of neutral genetic structure across the complex using selectively neutral SNPs, 161 and we then look for correlations between environmental variables and allele frequencies across the Oregon junco distribution using redundancy analysis (RDA). We 162 163 also use climatic variables to test for significant niche divergence while controlling for 164 spatial autocorrelation. Finally, we map GBS sequences harboring significant outlier SNP loci to the zebra finch (*Taeniopygia guttata*) reference genome in order to recover 165 the chromosomal position of polymorphic sites and explore how adaptive variation is 166 distributed across the genome. 167

168

### 169 Materials and methods

170

## 171 *Population sampling*

172 Oregon junco populations were sampled across the geographical range of the species

using mist nets in order to obtain biological samples for DNA extraction (Fig. 1A).

174 Each captured individual was aged, sexed, and marked with a numbered aluminum

band. A blood sample was collected by venipuncture of the sub-brachial vein and stored

in Queen's lysis buffer (Seutin 1991) or absolute ethanol at -80°C in the laboratory. All

177 sampling activities were conducted in compliance with Animal Care and Use Program

regulations at the University of California Los Angeles, and with state and federal

scientific collecting permits in the USA and Mexico. A high-quality tissue sample for

180 whole-genome sequencing was obtained from a red-backed junco (*J. hyemalis dorsalis*)

specimen (MLZ 69090) provided by the Moore Laboratory of Zoology, Occidental

182 College, that was collected at Dude Mountain, Coconino County, Arizona. Genomic

183 DNA was extracted from blood and tissue samples using a Qiagen DNeasy kit

184 (Qiagen<sup>TM</sup>, Valencia, CA).

185

186 Whole-genome sequencing and assembly

We assembled a draft dark-eyed junco genome by combining low-coverage genomes of eight different junco individuals, which we then used as a conspecific reference in the SNP calling process. Libraries for seven of the genomes were prepared with the Kapa Library Preparation Kit (Kapa Biosystems, Inc.) using TruSeq-style adapters (Faircloth and Glenn 2012). They were pooled after random shearing and individual barcoding and sequenced in a single lane of an Illumina HiSeq platform. The eighth genome was 193 sequenced at a higher coverage by means of two 101-bp paired-end shotgun libraries 194 and two 101-bp mate-paired libraries with insert sizes of 8 Kb in length at Macrogen Inc. The TruSeq Nano DNA Kit (Illumina) was used for the preparation of the shotgun 195 196 libraries, while the mate-paired libraries were prepared with Nextera Mate Pair Kit (Illumina). We used FASTQC (Andrews 2010) to evaluate the quality of the sequenced 197 data, and quality filtering was carried out with NextClip (Leggett et al. 2013) in the case 198 199 of the mate paired libraries. For the rest of them we used Trimmomatic (Bolger et al. 200 2014), applying a sliding window filtering approach with a size of 4 bp and a phred quality score threshold of 25. We also set a minimum length of 50 bp, below which 201 202 reads were filtered out after trimming. We used the software SOAPDENOVO2 (Luo et 203 al. 2012) to perform the assembly. The average insert size for each library was 204 estimated in a preliminary run, and we set a Kmer size of 27 and minimum edge 205 coverage of 2. Gaps that emerged during the scaffolding process were removed with the 206 GapCloser tool from SOAPDENOVO2. Finally, we filtered out all the scaffolds shorter 207 than 500 bp so the genome was functional as a mapping reference. The final assembled 208 genome had 37,904 scaffolds with an N50 of 147,816 bases, a L50 of 1,951 scaffolds, a 209 total size of 1.09 Gb, 17.5 Mb of missing sites, and an overall coverage of ~56X, as 210 computed with VCFTOOLS version 0.1.13 (Danecek et al. 2011).

211

## 212 Genotyping-by-sequencing

213 We used genotyping-by-sequencing (Elshire *et al.* 2011) to obtain individual genotypes

from 127 Oregon juncos belonging to the following subspecific taxa: *townsendi* (n=16),

215 *pontilis* (n=16), *thurberi* (n=35), *pinosus* (n=16), *montanus* (n=16), *shufeldti* (n=12),

216 *oreganus* (n=16) (Table 1). GBS libraries were prepared and sequenced at Cornell

217 University's Institute for Genomic Diversity, using the restriction enzyme PstI for

digestion. Sequencing of the 127 individually-barcoded libraries was carried out in three
different lanes (along with other 149 junco samples intended for other studies) of an
Illumina HiSeq 2000, resulting in an average of 229.2 million good single-end reads
100 bp in length per lane. Samples of the same form were distributed among at least two
of the lanes to avoid sequencing bias, except in the case of *shufeldti* individuals, which
were all introduced in the last set of sequenced samples.

224

### 225 Alignment and variant calling

We evaluated GBS read quality using FASTQC after sorting them by individual with 226 227 FASTXTOOLKIT (Gordon and Hannon 2010), and performed the trimming and 228 filtering treatment using PRINSEQ (Schmieder and Edwards 2011). All resulting reads were 69 bp long and had a mean genotyping phred quality score of at least 30, with no 229 230 positions below 20. The reads were then mapped against the assembled junco genome 231 using the mem algorithm in the Burrows-Wheeler Aligner (BWA, Li and Durbin 2009). Two samples were excluded at this step due to too low read mapping (1 thurberi and 1 232 233 montanus). We used the Genome Analysis Toolkit (GATK, McKenna et al. 2010) 234 version 3.6-0 to realign reads around indels using the IndelRealigner tool and then we 235 applied the HaplotypeCaller tool to call the individual genotypes. We finally used the 236 GenotypeGVCFs tool to gather all the per-sample GVCFs files generated in the previous step and produce a set of joint-called SNPs and indels in the variant call format 237 238 (vcf) (GATK Best Practices, DePristo et al. 2011; Auwera et al. 2013). Because GBS 239 data do not provide enough coverage for base quality score recalibration, we used VCFTOOLS to implement a 'hard filtering' process, customized for each of the 240 downstream analyses (Table 2). 241

242

### 243 *Genetic structure analyses*

244 To explore genome-wide population structure among all Oregon junco forms, we ran a principal components analysis (PCA) based on SNP data. Using VCFTOOLS we 245 246 retained all samples with less than 25% missing data after a 'soft filtering' (coverage range between 2 and 100, minimum phred quality score of 40), resulting in a dataset of 247 248 88 samples, and between 8 and 12 individuals per population (24 in the case of 249 thurberi). We filtered out all the sites with any non-genotyped individuals or a minor 250 allele frequency (MAF) below 0.02. We also applied a threshold for SNPs showing highly significant deviations from Hardy-Weinberg equilibrium (HWE) with a p-value 251 of 10<sup>-4</sup> to filter out false variants arisen by the alignment of paralogous loci, resulting in 252 253 a matrix of 11,261 variants. We then excluded SNPs putatively under selection using 254 BayeScan (Foll and Gaggiotti 2008). BayeScan computes per-SNP  $F_{ST}$  scores and 255 decomposes them into a population-specific component shared by all loci that 256 approximates population related effects, and a locus-specific component shared by all populations, which accounts for selection. The program compares two models of 257 258 divergence, with and without selection, and assumes a departure from neutrality when the locus-specific component is necessary to explain a given diversity pattern (Foll 259 260 2012). We used BayeScan with default settings and a thinning interval size to 100 to 261 ensure convergence. For each SNP we obtained the posterior probability for the selection model and the  $F_{ST}$  coefficient averaged over populations. For outlier detection 262 263 and exclusion, we implemented a false discovery rate of 0.3. To filter out the SNPs 264 under linkage disequilibrium (LD) we used the function snpgdsLDpruning from the 265 SNPrelate package (Zheng 2012) in R Studio (R Studio Team 2015) version 1.0.136 266 with R (R Core Team 2013) version 3.2.2. We applied the correlation coefficient 267 method with a threshold of 0.2 (method ="corr", ld.threshold=0.2), resulting in a final

data matrix of 9,436 SNP loci (Table 2). We then used the function snpgdsPCA also
available in SNPrelate to perform the PCA and obtain the eigenvectors to be plotted.

271 We examined population structure with STRUCTURE (Pritchard et al. 2000), using a smaller, more heavily filtered SNP data matrix to reduce the computational time of the 272 273 analysis. Using VCFTOOLS, we retained the eight samples of each population (16 in 274 the case of *thurberi*) with the lowest proportion of missing sites for a final number of 64 samples. We constructed a data matrix of biallelic SNPs excluding those out of a range 275 of coverage between 2 and 100, or with a genotyping phred quality score below 70. 276 277 Positions with less than 90% of individuals genotyped were removed from the data 278 matrix, along with those presenting a MAF below 0.02. Once again, we implemented a 279 threshold for SNPs showing highly significant deviations from Hardy-Weinberg equilibrium (HWE) with a p-value of  $10^{-4}$ , and performed the filtering for non-neutral 280 positions and linkage disequilibrium exactly as done for the PCA, to obtain a final data 281 282 matrix of 34,367 SNP loci (Table 2). We converted the vcf file to STRUCTURE format using PGDspider (Lischer and Excoffier 2012) version 2.0.5.1. Bash scripts to perform 283 the analyses were created with STRAUTO (Chhatre and Emerson 2016) and we ran the 284 program five times per K, with K ranging from 1 to 10 after running a preliminary 285 analysis to infer the lambda value. The burn-in was set to 50K iterations and the 286 287 analysis ran for an additional 100K iterations. Similarity scores among runs and 288 graphics were computed with CLUMPAK (Kopelman et al. 2015). 289

290 *Redundancy analysis and variance partition* 

291 When applying GEA methods there are two main potentially confounding effects

292 related to neutral factors: (i) structure among populations derived from strong drift in

293	isolation may result in genetic patterns similar to those related to adaptive divergence;
294	and (ii) demographic expansions along latitudinal axes may create gradients of allele
295	frequencies at neutral loci correlated with latitude, that in turn would correlate with any
296	environmental variable that changes with latitude, mimicking a pattern of selective
297	sweep and local adaptation (Excoffier and Ray 2008; Excoffier et al. 2009; Rellstab et
298	al. 2015; Forester et al. 2016). Redundancy analysis (Van Den Wollenberg 1977;
299	Legendre and Legendre 1998) is a canonical ordination method that allows computing
300	the variance of a set of response variables explained by a number of constraining or
301	explanatory variables. In addition, partial RDA enables computation of this shared
302	variance between two sets of variables while conditioning or holding constant the
303	effects of a third set of covariables. Here we used RDA and partial RDA as
304	implemented in the R package vegan (Oksanen et al. 2016) to explore the associations
305	between genetic variability and environmental data. Ecological data were obtained from
306	7 of the 19 variables available in the BioClim database (Hijmans et al. 2005),
307	specifically chosen in accordance to their relevance to junco ecology (Miller 1941;
308	Nolan et al. 2002). They measured mean temperature and precipitation over the year
309	(BIO1 and BIO12); mean temperature and precipitation over the warmest quarter
310	(BIO10 and BIO18), which corresponds to the birds' breeding season; isothermality,
311	referring to how the range of day-to-night temperature differs from the range of
312	summer-to-winter, where a value of 100 indicates equality between them; and
313	seasonality of temperature and precipitation (BIO4 and BIO15). We also included three
314	vegetation variables from the Moderate Resolution Imaging Spectroradiometer
315	(MODIS) satellites as available in <u>https://modis.gsfc.nasa.gov</u> : percent tree cover
316	(TREE), Normalized Difference Vegetation Index (NDVI, a measure of canopy
317	greenness), and NDVI's annual standard deviation (std_NDVI). Finally, we included

318 the high-quality elevation data provided by the NASA Shuttle Radar Topographic Mission (ELEV), downloadable from http://www2.jpl.nasa.gov/srtm for a total of 319 320 eleven variables (Table 3). All ecological variables were centered and standardized. 321 Following Blanchet et al. (2008), we implemented a forward selection method using the forward.sel function from the R-package packfor (Dray et al. 2009) to reduce the 322 number of variables in the model. This procedure applies two stopping criteria: a 323 significance level for each tested variable, which we set at 0.01; and a maximum limit 324 for global adjusted  $R^2$ , equal to the adjusted  $R^2$  of the RDA model including all initial 325 variables. In doing so we prevent inflation of the overall type I error and of the amount 326 327 of explained variance. After this, we excluded those retained variables with a variance 328 inflation factor (VIF) over 10 (Borcard et al. 2011) to avoid high collinearity. Despite 329 signs of low orthogonality observable among variables in the partial RDA (especially 330 among BIO18, TREE and NDVI, see Results) we chose not to exclude more variables 331 or to apply dimension-reduction treatments like PCA to the environmental space of 332 variables so as to assess their specific and relative contributions to differentiation 333 patterns (McCormack et al. 2010) and discuss signals of adaptation with higher confidence. The final selected ecological variables were used as explanatory variables in 334 335 two RDAs, with and without subtracting the effects of neutral processes, which were approximated by the first two principal components of the PCA of population structure 336 based on selectively neutral loci (see above). As response variables, we used the same 337 SNP dataset used for the PCA, but excluding LD and neutrality filters. SNP data were 338 339 coded as counts of the alternative allele for each position (*i.e.*, 0, 1 or 2 copies) with VCFTOOLS and transformed following Patterson et al. (2006). Statistical significance 340 was obtained using a permutation-based procedure with 10,000 permutations, assuming 341 342  $\alpha = 0.01$ . We also used variance partitioning as implemented in the vegan R-package to

343 estimate (i) the total proportion of genomic variation explained by ecological variables 344 alone; (ii) by neutral structure alone; and (iii) the effects of both sets of variables. 345 Finally, we repeated the whole RDA treatment for a subset of 87 SNPs identified as 346 selectively divergent by BayeScan, with no conditional treatment. The analyses were conducted in R version 3.2.2 (code included in Appendix I, Supplementary Material). 347 348 349 Niche divergence tests 350 To further explore patterns of ecological divergence in the Oregon Junco, we tested for niche divergence applying the method developed by McCormack et al. (2010), a method 351 352 that allows us to examine each environmental variable separately. To avoid a loss of 353 statistical power due to multiple analyses, we conducted three specific comparisons of 354 forms presenting different patterns of genetic divergence and geographical settings, 355 including: (i) townsendi and southern thurberi forms, in order to estimate niche 356 divergence between geographically isolated, genetically differentiated forms; (ii) the 357 ecologically divergent *pinosus* with the parapatric northern *thurberi* form, to further test 358 a possible case of isolation-by-adaptation; and (iii) northern and southern populations of thurberi, as conspecific extremes of a potential adaptive gradient. We used occurrence 359 360 points from our own georreferenced field sampling records, and this set of occurrence 361 records was further revised to avoid spatial autocorrelation and to match the spatial resolution of environmental variables (1-km grid). Our final dataset comprised 80 362 363 localities: *pinosus* (n = 14), *thurberi* north (n = 26), *thurberi* south (n = 19), and 364 *townsendi* (n = 21). We decided to improve quality (geographic accuracy) vs. quantity 365 (number of occurrence records), by using fewer data but with higher spatial accuracy 366 (Engler et al. 2004). To generate a background dataset for each population, we drew 367 1000 random points from a background representing the geographic range of each junco

368 population. In order to select an appropriately sized area for the niche divergence tests, 369 we included accessible habitats according to the dispersal ability of each population (Soberon and Peterson 2005). We generated background samples from a 100-km 370 371 "buffer zone" around known occurrences (Warren et al. 2008). For populations with small ranges or small dispersal ability (thurberi south, pontilis and townsendi) we 372 373 restricted the buffer zone to 10 km to reduce spatial inaccuracies in the null distribution 374 (Barve et al. 2011), after testing different buffer sizes to test the robustness in delimiting accessible areas for juncos. Next, we extracted the environmental data (same as the data 375 used for the RDA, see above) for both occurrence points and random background points 376 377 from within the geographic range of each junco form. Niche divergence and 378 conservatism were tested by comparing the observed environmental differences among 379 forms against a null model of background divergence (generated by calculating the 380 difference between background points using a bootstrapping approach and 1000 resamples) for each environmental variable using a two-tailed test. We conducted all the 381 382 analyses in R 3.2.2. 383

384 *Genome scans* 

We performed genome scans for different Oregon junco forms using BayeScan.

In order to obtain the chromosomic positions of the SNPs, we mapped the GBS reads

387 against the zebra finch (*Taeniopygia guttata*) genome v87 available in Ensembl (Yates

et al. 2016), applying the same set of tools and parameters as for mapping against the

- junco genome. Using the same set of samples as in the PCA and the RDA, we
- 390 conducted the analysis for all forms together; for *townsendi* against *pontilis*; and for
- 391 townsendi against all thurberi (see Table 2 for final dataset sizes). For each of these
- matrices, we retained only biallelic SNPs with coverage between 2 and 100 and a

393	genotyping phred quality score over 40. Positions with less than 25% of the individuals
394	genotyped were removed from each data matrix, along with those presenting a MAF
395	below 0.05. Once again, we implemented a p-value threshold for HWE of $10^{-4}$ to filter
396	out false variants arisen by the alignment of paralogous loci. We ran BayeScan with the
397	same settings used for filtering out SNPs under selection in population structure
398	analyses, but implemented a more conservative 10% FDR for outlier detection. Genome
399	scan plots were conducted in R 3.2.2 using the package qqman (Turner 2014).
400	
401	
402	Results
403	
404	Neutral genetic structure
405	The plot of the first two principal components from the PCA revealed four distinct
406	clusters in the Oregon junco group. The most differentiated groups were townsendi and
407	pontilis from Baja California, which formed two highly divergent clusters with respect
408	to each other and to other populations. A third, highly differentiated group corresponded
409	to pinosus from coastal California, showing less differentiation than the Baja California
410	forms with respect to a fourth cluster, which included all the remaining forms in the
411	PCA (Fig. 1B). Within this fourth cluster, southern thurberi individuals presented
412	certain degree of differentiation from the rest of the forms, a pattern more conspicuous
413	when plotting the third and fourth components, which also revealed a slight signal of
414	divergence in the oreganus form (Fig. S1 in Supplementary Information).
415	
416	The STRUCTURE results were generally congruent with the PCA. The $K = 2$ plot
417	recovered townsendi as an independent population, which shared a considerable amount

418	of variance with <i>pontilis</i> . In the analysis for $K = 3$ , <i>pontilis</i> separated, and $K = 4$
419	identified the same four main clusters seen in the PC1 vs. PC2 plot (Fig. 1B), revealing
420	<i>pinosus</i> as an independent genetic cluster. The plot for $K = 5$ clearly captured the
421	differentiation of the southern <i>thurberi</i> form in a fifth cluster, and in K=6, <i>oreganus</i>
422	appears as an independent northern group with all individuals from northern thurberi
423	and especially montanus and shufeldti forms showing an increasing probability of
424	assignment to the oreganus cluster from south to north (Fig. 2).
425	
426	Patterns of divergence in Nei's distances and $F_{ST}$ values among forms were highly
427	congruent with previous results, with townsendi, pontilis and pinosus showing the
428	highest values for both indices, and northern forms showing lower levels of
429	differentiation, especially between southern and northern thurberi individuals (Table
430	S1).
431	
432	Forward selection of explanatory variables
433	Out of eleven potentially relevant ecological variables for juncos, six were retained after
434	the forward selection method intended for excluding non-significant effects. Retained
435	variables included isothermality (BIO3), mean temperature of the warmest quarter
436	(BIO10), mean precipitation of the warmest quarter (BIO18), vegetation cover (TREE),
437	greenness (NDVI), and elevation (ELEV) (Table 3). None of these variables was

438 excluded due to excessive correlation as VIF values were below 10 (maximum

439 recovered VIF = 5.75).

440

441 *Redundancy analysis and variance partition* 

All six ecological variables retained in the forward selection method were included as 442 443 explanatory variables in the RDA and the partial RDA models. RDA computes, in successive order, a series of axes that are linear combinations of the explanatory 444 445 variables, and that best explain the variation in the matrix of response variables (Borcard et al. 2011). Six RDA axes (named RDA1 to RDA6 hereafter, ordered by the 446 amount of variance explained by each one, as reflected by the adjusted  $R^2$ ) explained 447 448 6.26% of the total genetic variance in the non-conditioned model, and 1.18% when removing the effects of neutral genetic structure (Table 4). The amount of explained 449 variance increased to 36.61% when using only BayeScan outliers as response variables 450 451 (Table 4, Table S2). The permutation tests for the RDA models yielded a p-value below 0.001 in all three analyses, confirming the significance of the constraining variable 452 453 effects. 454

455 Loadings of ecological explanatory variables on each of the axes varied across the three 456 different RDA models (Fig 4, Table 4, Table S2). In the non-conditioned RDA, the 457 RDA1 axis had a large negative contribution of TREE and NDVI, and loaded positively on elevation (Fig. 4A). The RDA2 axis loaded mostly on BIO3 and BIO10. The plot of 458 459 per-individual projections on these two axes revealed a pattern generally similar to the 460 PCA. The forms townsendi, pontilis and to a lesser extent pinosus, showed distinctive high values of correlation with both axes. The remaining forms showed similar 461 462 correlation patterns with respect to RDA1, with southern *thurberi* individuals showing a 463 clear association with RDA2 (Fig. 4A). 464

In the partial RDA, RDA1 had a large contribution of BIO3, while RDA2 loaded mostly

466 on BIO10 and to a lesser extent, BIO18 and TREE (Fig. 4B). Plotting these first two

RDA axes revealed patterns of genetic correlation especially related to the first RDA 467 468 axis for *pinosus*, which consequently presented the strongest association with isothermality. The individuals from the southern population of *thurberi* showed a more 469 470 pronounced negative association with RDA2 than in the non-conditioned RDA, further evidence of a positive correlation with mean temperature of the warmest quarter. 471 472 Northern forms of the Oregon junco (northern thurberi, shufeldti, montanus and 473 oreganus) also separated along the second RDA axis, with the northernmost oreganus showing the strongest association with the particularly conspicuous mean precipitation 474 of the warmest quarter gradient, a pattern that was not visible in the non-conditioned 475 476 RDA. Townsendi from Baja California occupied positions closer to the origin of coordinates, suggesting a lower association between environmental and genetic variance 477 478 (Fig. 4B). 479 480 In the RDA based on outlier loci potentially under selection, the first axis showed

481 moderate negative contributions from TREE and NDVI, and also a positive contribution 482 of ELEV (Fig. 4C). Variance in BIO18 was almost entirely captured by RDA2, which also had a relatively high, negative contribution from BIO3. The plot showed a pattern 483 484 of correlation between *pontilis* and *townsendi* along RDA1, while genetic variance in oreganus appeared strongly associated with the gradient of mean precipitation of the 485 warmest quarter along RDA2. The rest of Oregon junco forms (and one atypical 486 oreganus individual) were distributed in an opposite fashion, with small differences in 487 488 their patterns of correlation with environmental variability captured in the second axis of the RDA (Fig. 4C). 489

490

491	The variance partition analysis showed that climate and neutral structure together
492	explained 7.17% of the total genetic variability (fractions A+B+C, Fig. 5). Since
493	variable sets are not orthogonal, a 5.08% of variation was explained jointly by the
494	environmental data and the first two components of the PCA based on neutral genetic
495	positions (fraction B, Fig. 5). As recovered in the partial RDA, environmental variables
496	alone explained 1.18% of the total variance (fraction A, Fig. 5), while the non-
497	overlapping fraction of neutral genetic structure explained 0.91% of the variability in
498	the SNP dataset (fraction C, Fig. 5). The p-value computed through the 10,000-step
499	permutation test for each individual fraction was below 0.001 in all cases, thus
500	confirming the significant effects of both variable sets.
501	
502	Niche divergence tests
503	We tested for niche divergence and conservatism on each of the environmental
504	variables. We found significant niche divergence between pinosus and northern thurberi
505	for three of the six environmental variables analyzed (isothermality, mean precipitation
506	of the warmest quarter and elevation; Table 5). When considering northern <i>thurberi</i> vs.
507	southern thurberi, we found significant divergence for mean temperature of the warmest
508	quarter and conservatism for isothermality (Table 5). NDVI was the only variable that
509	exhibited significant divergence between townsendi and thurberi south (Table 5).
510	
511	Genome scans
512	The BayeScan survey comparing all Oregon junco forms together detected 32 SNPs
513	potentially under divergent selection, and 5 significant SNPs potentially under
514	balancing selection. In the two pairwise comparisons townsendi vs. pontilis and
515	townsendi vs. thurberi, 20 and 30 significant SNPs with signs of having diverged under

516	selection were detected, respectively, and no significant SNPs under balancing selection
517	were found in either case. SNPs potentially differentiated under divergent selection
518	appeared distributed across the genome in all comparisons, without obvious signs of
519	heterogeneity among regions (Fig. 6). Chromosomes 1B and 16 harbored no SNPs so
520	they are not shown in the plot.
521	
522	
523	Discussion
524	
525	Neutral population structure and local adaptation explains genomic variance among
526	Oregon junco forms
527	Our results reveal that both neutral and selective factors have played a role in driving
528	divergence among Oregon junco populations, and that the relative contributions of
529	geographic isolation and environment-driven selection are not uniform across the
530	distribution range of the complex. Environmental variables explained 1.18% of genomic
531	variation when controlling for population structure, and environment and neutral
532	structure together accounted for 7.17% of the variability in the 11,256 SNP matrix. The
533	remaining 92.83% of the variance corresponds potentially to loci under balancing
534	selection or selective pressures not represented in our ecological variables, and to
535	neutral variation shared by all Oregon junco forms because of their close relatedness
536	and/or gene flow among them. The amount of variance explained solely by
537	environmental variables in our study was comparable to the values reported in studies
538	applying RDA to detect specific correlations between genomic variation and a given set
539	of potentially correlated variables, as shown in plants (e.g. Lasky et al. 2012; Vincent et
540	al. 2013; De Kort et al. 2014) or other avian species (Safran et al. 2016; Szulkin et al.

541 2016). Previous studies on birds have used simple spatial variables such as geographic 542 distance to control for the effects of spatial autocorrelation (Safran et al. 2016; Szulkin et al. 2016). Here, we controlled for genome-wide patterns of neutral variation by 543 544 subtracting the variance captured by the first two PCs of a PCA based on neutral genome-wide SNPs, a method which should better account for population history and 545 546 structure, including changes in effective population size, geographic isolation and 547 related effects (Forester et al. 2016). Since spatial autocorrelation is usually reflected by neutral genetic structure, we did not include spatial covariates to avoid over-548 549 conditioning the model (Rellstab et al. 2015). Given that only a small fraction of the 550 surveyed genome is expected to be related to genes coding for climatic adaptation or 551 linked to them (Meirmans 2015), a significant 1.18% of association between genomic 552 variation and environmental variability in the conditioned (partial) RDA over only 553 11,261 genome-wide distributed SNPs is a compelling signal of local adaptation. 554

555 Genetic-environment association patterns in the diversification of the Oregon junco 556 The RDA revealed a number of strikingly different patterns of covariation between genetic variance and ecological variables likely to have played a role in Oregon junco 557 558 diversification, especially when the effects of population history were removed. The 559 forms pontilis and townsendi from Baja California, markedly isolated in terms of 560 geography and neutral genetic variability, presented a low genetic-environment 561 association when controlling for population history. This suggests that the 562 differentiation between townsendi and pontilis is due largely to geographic isolation, in 563 this case caused by unsuitable desert habitat in the lowlands surrounding their 564 respective mountain ranges, a pattern also known as isolation by resistance (IBR, McRae and Beier 2007). In this scenario, our results suggest that differentiation is 565

566 caused by drift under conditions of small population size and reduced gene flow, rather 567 than divergent selection due to local adaptation, fitting the classic allopatric speciation 568 model (Mayr 1942, 1963; Coyne and Orr 2004). This hypothesis is consistent with the 569 niche divergence test comparing *townsendi* with southern *thurberi*, for which all tested 570 variables but NDVI showed no signal of divergence beyond expectations based on 571 background divergence.

572

The form *pinosus* showed considerable neutral genetic structure and a conspicuous 573 pattern of genetic-environment association in both non-conditioned and partial RDA. 574 575 When controlling for population structure, *pinosus* individuals showed high positive 576 correlation values with isothermality, while correlating negatively with elevation. 577 Indeed, *pinosus* presents the highest isothermality values, and the second lowest 578 elevation after *oreganus*, reflecting a tolerance for low elevation conditions that are 579 absent in neighboring *thurberi*, a pattern also recovered in the niche divergence test. Unlike other differentiated forms like pontilis and townsendi, pinosus does not show 580 high geographic isolation, and zones of intergradation with *thurberi* have been 581 described (Miller 1941). Neutral population divergence despite the absence of 582 583 geographic barriers to gene flow along with signs of local adaptation is a pattern 584 consistent with isolation-by-adaptation, where barriers to gene flow may have arisen as 585 individuals adapted to the distinct habitat of the coastal mountains of California. Niche 586 distinctiveness and the genetic-environment association pattern of this form is thus 587 congruent with a combination of warm latitude, low elevation and coastal influence that 588 has seemingly resulted in the adaptive differentiation of *pinosus* from the rest of the 589 Oregon junco taxa. As a result, differentiation by drift may have led to positive 590 correlations between adaptive and neutral genetic divergence (Nosil et al. 2008).

591

592	The southern thurberi individuals from Mont Laguna showed high overlap in terms of
593	neutral genetic structure with northern thurberi and other boreal forms, and only slight
594	differences in their genetic-environment association patterns when no controls for
595	confounding factors were implemented. The Mount Laguna site represents the
596	southernmost tip of the thurberi range in Southern California, which extends northward
597	and reaches Oregon, forming a relatively continuous distribution (Miller 1941; Nolan et
598	al. 2002), suggesting potentially high gene flow. However, the partial RDA revealed a
599	distinctive pattern of high correlation with the mean temperature of the warmest quarter
600	for the southern <i>thurberi</i> juncos, differentiating them from the rest of Oregon forms.
601	They also correlated negatively with the mean precipitation of the warmest quarter. This
602	pattern seems congruent with the habitat of Mount Laguna, and in general with the
603	southern inland range of Oregon juncos, quite arid during summer but subject to
604	snowfalls in winter due to the high elevations (Miller 1941), contrasting sharply with
605	the more climatically moderate coastal and northern populations. The limited neutral
606	genetic structure between <i>thurberi</i> range extremes but considerable differentiation in the
607	genetic-environment association patterns is consistent with a process of local adaptation
608	across a selective gradient (Forester et al. 2016), in which selection is the prominent
609	evolutionary force driving differentiation (Haldane 1948; Slatkin 1973; Nagylaki 1975;
610	Felsenstein 1976) while neutral alleles may move freely across space. The niche
611	divergence test comparing southern and northern thurberi populations was significant
612	for mean temperature of the warmest quarter.
613	

613

The boreal Oregon junco forms including *oreganus*, *montanus*, *shufeldti*, and *thurberi*individuals from Tahoe, California, presented a more conspicuous pattern of local

616 adaptation along a shallow selective gradient. These forms showed very low neutral 617 genetic structure or differences in ecological covariances in the non-conditioned RDA, yet showed an increasing signal of association following their latitudinal distribution in 618 619 the partial RDA. A strong association pattern emerged especially for mean precipitation 620 of the warmest quarter and for correlated environmental variables of tree cover and 621 greenness, matching quite precisely their latitudinal distribution along a gradient of 622 increasing humidity and vegetation cover. The ecology-related differences in genetic 623 variance, consistent with the taxonomic classification of these forms, is especially 624 relevant considering the relative phenotypic similarity of these taxa, and their apparent 625 intergradation (Miller 1941; Nolan et al. 2002). 626 627 GEA methods present a number of limitations, including potentially high rates of type I 628 error (see Supplementary Materials for details). Here, rather than detecting specific loci 629 under selection, we aimed to explore how selection and neutral processes shape the 630 variability in Oregon juncos, but the risk of finding false significant associations between genetic variance and ecological parameters persists. To further test the 631 632 environmental associations revealed in this study, we implemented a highly 633 conservative approach using only BayeScan outliers as response variables in the model. 634 The non-conditioned RDA based on 87 SNP loci identified by BayeScan as potential targets of divergent selection yielded relatively lower resolution than the partial RDA. 635 636 BayeScan has been shown to produce relatively few false positives, but it is also a 637 conservative approach, the sensitivity of which decreases with selection strength 638 (Narum and Hess 2011). The RDA suggests that BayeScan correctly identified outliers related to low temperatures and high precipitation for *oreganus* samples, a pattern 639 640 congruent with the habitat and with the outcomes of previous analyses for this form. It

641 also detected highly differentiated positions in *pontilis* and *townsendi* that correlate with 642 RDA1, but in this case associations with specific environmental variables were lower, 643 and the pattern disappears in the RDA based on the entire SNP dataset when correcting 644 for population structure. This may suggest that BayeScan failed to exclude the effects of demographic history, or in turn, that controlling for the genetic variance captured by the 645 646 PCA was overly conservative. BayeScan was also less successful in detecting adaptive 647 divergence in *pinosus*, and especially in northern Oregon junco forms, where selection 648 may be weaker or have acted during a shorter period. Nevertheless, the outlier SNP dataset explained 36.61% of the total climatic variability, a considerable amount 649 650 compared with the full SNP data RDA models, indicating a good fit of the retained 651 outliers to the linear regression on the environmental parameters. 652 653 Interactions among environment, geography and demographic processes result in three 654 different modes of divergence within the Oregon junco lineage 655 The Oregon junco is one of the six phenotypically and genetically differentiated dark-656 eyed junco taxa evolved during a northward expansion from Central America after the last glacial maximum, c.a. 18,000 years ago (Milá et al. 2007; Friis et al. 2016). Similar 657 658 postglacial expansions have been reported for many other bird species (Seutin et al. 659 1995; Milá et al. 2000; 2006; Hansson et al. 2008; Malpica and Ornelas 2014; Alvarez 660 et al. 2015). However, the population structure documented in this study reveals a 661 variety of different spatial, selective and demographic factors not previously 662 documented in other avian taxa. In light of our genetic-environment association 663 analyses, the patterns recovered by the PCA and STRUCTURE analysis reveal at least 664 three different effects of geography and demographic history interacting to varying 665 degrees with selection in the process of Oregon junco diversification. First, the IBR

666 pattern of differentiation presented by *pontilis* and *townsendi* may suggest that these 667 forms are peripheral remnants of an original, broader distribution of the Oregon juncos, thereafter isolated in 'sky islands' of Baja California and diverging predominantly by 668 669 drift. Indeed, in his thorough monograph on the geographic variation in juncos, Alden Miller (1941) had perceptively suggested early on that the habitat of Oregon juncos 670 671 from Baja California did not seem to account for their phenotypic differentiation from 672 Californian forms, and that their distinctive traits appeared to be predominantly 673 "historical" (p. 306). The spatial configuration and recovered patterns of neutral and 674 adaptive divergence for *pontilis* and *townsendi* fit the first model of neutral divergence 675 of isolated population in approximately similar habitats (scenario A in Fig. 2). Second, 676 the IBA pattern found in *pinosus* suggests that the current area of intergradation 677 corresponds to a secondary contact zone that formed after diverging in relative isolation, 678 maybe linked to the ancient coastal closed-cone pine forest that has allegedly 679 diminished since the Pleistocene (Miller 1941). The mode of divergence between 680 parapatric populations in different habitats (scenario B in Fig. 2) has been only partially fulfilled, since local adaptation seems to have resulted in reduced levels of gene flow, 681 682 leading to increased neutral genome-wide differentiation (Nosil et al. 2008; Funk et al. 683 2011; Flaxman et al. 2014). Third, the geographic continuum represented by *thurberi*, shufeldti, montanus and oreganus is also captured in the STRUCTURE analysis by a 684 gradual signal of differentiation following a latitudinal distribution, suggesting that 685 686 ongoing gene flow may occur among forms. Combined with the signal of increasing 687 environmental association recovered in the partial RDA, these outcomes are consistent 688 with a process of differentiation driven by local adaptation along a selective gradient in the direction of the northward expansion, fulfilling the third hypothesized mode of 689 690 divergence for these forms of Oregon junco (scenario C in Fig. 2). Interestingly,

ecological association approaches have been shown to perform better along clines of
selection where demographic expansions align with the gradient of ecological variables,
usually related to latitude (Frichot et al. 2015), as is the case in the *Junco* system.

694

695 A relevant aspect of the marked population structure found among Oregon junco forms 696 is that it is based on a relatively small subset of genome-wide SNPs randomly sampled 697 from across the genome, representing a genomic fraction not greater than 0.2% of the 698 total of 1.2 Gb. The clear signal of divergence mediated by environmental factors recovered also in the RDA indicates that divergence may have taken place at the level 699 700 of the entire genome, suggesting the role of multiple selective pressures during local 701 adaptation along the latitudinally broad and heterogeneous distribution of the Oregon 702 juncos. The presence of outliers potentially under positive selection scattered across the 703 genome seems to support this hypothesis of selection-driven genome-wide divergence, 704 rather than widespread drift among isolated populations. Other examples of such 705 patterns of genomic differentiation due to divergent selection at early (e.g. Parchman et 706 al. 2013; Brawand et al. 2014; Egan et al. 2015) and intermediate (e.g. Riesch et al. 707 2017) stages of speciation have been reported recently, contrasting with proposed 708 models of speciation initiated by divergent selection in a few, localized genes involved 709 in reproductive isolation (e.g. Nadeau et al. 2012; Poelstra et al. 2014). 710

711

### 712 Conclusion

Our analyses reveal the role of both local adaptation and demography in driving rapid
diversification during the northward recolonization of western North America by the

715 Oregon junco. The combined effects of a demographic expansion along a selective

716 gradient with a heterogeneous landscape of environmental variability have resulted in a 717 striking array of divergence modes within a single lineage, from isolated forms in Baja California that have differentiated largely by drift in isolated 'sky islands', to adaptive 718 719 diversification along selective gradients with no obvious geographic barriers to gene flow. There is also a compelling example of isolation by adaptation in the case of 720 721 *pinosus*, where ecological barriers to gene flow seem to maintain its divergence with 722 respect to nearby forms. Genome-wide patterns of divergence indicate that Oregon junco diversification has been driven by multiple ecological factors acting on many loci 723 724 across the genome, and suggests that selection may promote local adaptation in short 725 periods of time, highlighting the role of adaptive divergence in the early stages of the 726 speciation process. Future analyses of dense sequencing and functional gene 727 characterization will be necessary to further identify adaptive changes promoting 728 barriers to gene flow and reveal the genomic architecture of rapid diversification.

729

730

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- 741 Ministry of Science and Innovation to BM.
- 742

## 743 Data Accessibility

- 744 Genomic data will be deposited in Dryad in short.
- 745

## 746 Author Contributions

- 747 G. Friis and BM designed the study and carried out field sampling; G. Friis, JM, and BF
- 748 generated and analyzed genomic data; G. Friis, G. Fandos and AZ generated and
- analyzed environmental data; G. Friis and BM wrote the manuscript with input from all
- 750 co-authors.
- 751
- 752

- **Table 1.** Oregon junco forms and number of genotyped individuals per locality. State
- abbreviations are the following: British Columbia (BC) in Canada; Oregon (OR),
- 755 California (CA) in the USA; Baja California Norte (BCN) in Mexico.

756

Form	State	Localities	Sequenced	
oreganus	BC	Banks Island, Porcher Island, Susan Island	16	
shufeldti	OR	Willamette N.F.	12	
montanus	OR	Wallowa N.F.	16	
northern thurberi	CA	Tahoe	18	
southern thurberi	CA	Mount Laguna	17	
pinosus	CA	Santa Cruz Mountains	16	
pontilis	BCN	Sierra Juárez	16	
townsendi	BCN	Sierra San Pedro Mártir	16	
Total			127	

757

# **Table 2.** SNP data matrices used in each analyses. General filters included a depth

range from 2 to 100 and a p-value for Hardy-Weinberg deviation of  $10^{-4}$ .

## 761

Analysis	Number of	Min. phred	MAF	Number	Allowed
	samples	score	threshold	of SNPs	missing data
STRUCTURE	64	70	0.02	34,367	10%
PCA	88	40	0.02	9,436	0%
RDA:					
On all <i>loci</i>	88	40	0.02	11,261	0%
On BayeScan outliers	88	40	0.02	87	0%
Genome Scans:					
All lineages	88	40	0.02	29,868	75%
townsendi vs. pontilis	24	40	0.05	22,773	75%
townsendi vs. thurberi	24	40	0.05	22,516	75%

762

- **Table 3.** Set of environmental variables included in the initial stepforward selection
- method. Significant variables retained by the method are shown in bold.
- 766

Environmental	Description
variable	
BIO1	Annual mean temperature
BIO3	Isothermality
BIO4	Temperature seasonality (standard deviation *100)
BIO10	Mean temperature of the warmest quarter
BIO12	Annual precipitation
BIO15	Precipitation seasonality (coefficient of variation)
BIO18	Precipitation of the warmest quarter
NDVI	Normalized Difference Vegetation Index (greenness)
NDVI SD	Annual NDVI standard deviation (greenness seasonality)
TREE	Percent tree cover
ELEV	Elevation from the NASA Shuttle Radar Topographic Mission

768	Table 4. RDA loads of the constraining variables in the first two axes and their
769	explained variance for each one of the RDA models. The total variance explained by the
770	full model (adjusted $R^2$ for the resultant six axes) is shown for each analysis. In all of
771	the three analyses, the p-value for the full models was below 0.001. See Table 3 for
772	variable definitions.

773

	Non conditioned RDA		Partial RDA		RDA of BayeScan outliers	
Total explained variance (adjusted R <sup>2</sup> )	6.26%		1.18%		36.61%	
Variable	RDA1	RDA2	RDA1	RDA2	RDA1	RDA2
BIO3	0.124	0.239	0.839	-0.427	0.214	0.763
BIO10	0.108	-0.610	0.359	-0.679	0.221	0.374
BIO18	0.052	0.021	-0.452	0.501	-0.031	-0.994
ELEV	0.646	-0.031	-0.324	-0.301	0.629	0.420
NDVI	-0.641	0.588	-0.320	0.199	-0.731	-0.277
TREE	-0.710	0.320	-0.270	0.326	-0.776	-0.423
Variance						
Eigenvalue	1304.262	567.705	481.320	350.714	63.655	12.657
Proportion explained	0.025	0.011	0.003	0.002	0.250	0.050
Cumulative proportion	0.025	0.037	0.003	0.005	0.250	0.300

774

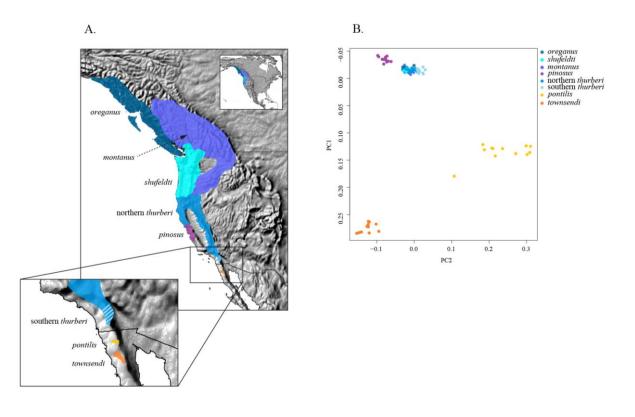
- **Table 5.** Results from the niche divergence test for *pinosus* vs. northern *thurberi*,
- northern vs. southern *thurberi* and southern *thurberi* vs. *townsendi*. Variables showing
- significant divergence (Diverged) or conservatism (Conserved) are shown in bold (p-
- value < 0.05). e.d.b.b.: expected divergence based on background.
- 780

	Variable	Result	p-value	Observed mean	Background mean
				difference	difference
<i>pinosus</i> vs.	TREE	e.d.b.b.	0.636	3.95	7.17
northern thurberi	BIO10	e.d.b.b.	0.718	20.40	17.34
	BIO3	Diverged	0.000	16.78	9.45
	BIO18	Diverged	0.000	41.92	27.24
	ELEV	Diverged	0.032	1329.81	978.26
	NDVI	e.d.b.b.	0.550	192.67	428.18
northern vs.	TREE	e.d.b.b.	0.580	2.11	4.52
southern <i>thurberi</i>	BIO10	Diverged	0.030	41.85	23.87
	BIO3	Conserved	0.002	1.37	2.87
	BIO18	e.d.b.b.	0.860	15.02	14.25
	ELEV	e.d.b.b.	0.474	202.60	140.29
	NDVI	e.d.b.b.	0.144	258.89	1019.97
southern thurberi	TREE	e.d.b.b.	0.536	15.01	12.25
vs. t <i>ownsendi</i>	BIO10	e.d.b.b.	0.412	47.88	43.03
	BIO3	e.d.b.b.	0.472	2.47	2.89
	BIO18	e.d.b.b.	0.188	74.77	67.26
	ELEV	e.d.b.b.	0.786	750.22	717.21
	NDVI	Diverged	0.012	1597.62	574.28

781

782

- **Figure 1.** Geographic distribution and neutral genetic structure of the Oregon junco
- forms. (A) Breeding ranges of all the Oregon junco forms. (B) Genetic structure of
- 786 Oregon junco forms based on a principal components analysis of independent,
- 787 selectively neutral genome-wide SNPs.

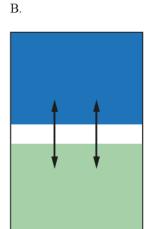


788

789

- 791 Figure 2. Expectations for neutral and adaptive divergence under different
- environmental and spatial configurations found across the Oregon junco distribution. 792
- 793 (A) Geographically isolated populations in similar habitats. (B) Parapatric populations
- in ecologically divergent habitats. (C) Population continuum across a selective gradient. 794

А.





Habitat types: Gene flow: Neutral divergence: Adaptive divergence:

ONE LOW HIGH LOW



TWO

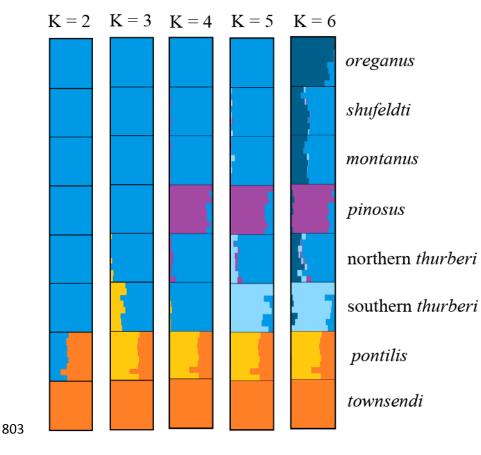
HIGH

GRADIENT HIGH LOW HIGH

795

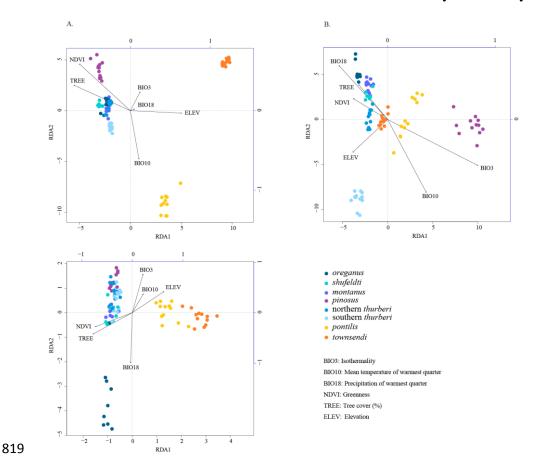
- **Figure 3.** Genetic structure of the Oregon junco forms based on 34,367 selectively
- neutral genome-wide SNPs using the program STRUCTURE. Each horizontal bar
- corresponds to an individual, with colors corresponding to posterior assignment
- 800 probabilities to each of a number of genetic clusters (K). Colors correspond
- approximately to those in Fig. 2A.

802



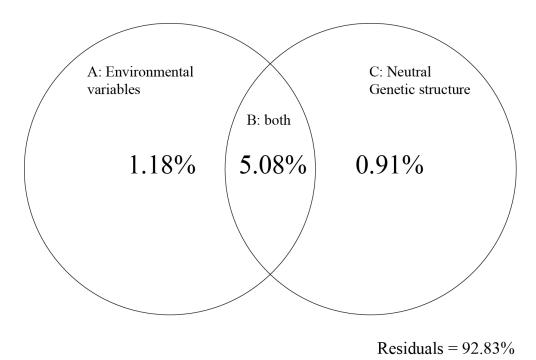
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806 Figure 4. Genetic-environment association analyses in the Oregon junco. Points 807 represent the projection of individual genotypes on the first two RDA axes. Marker colors correspond to those on the range map on Fig. 2A. The explanatory variables are 808 809 shown within the space defined by RDA1 and RDA2 by labeled vectors. Their contribution to each axis is represented by the length of their orthogonal projections 810 811 over the scale bars along top and right sides of the graphs. Arrows indicate the direction 812 of the gradient of variation for the corresponding environmental parameter. The value for each sample point on each explanatory variable can be obtained by an orthogonal 813 projection on the corresponding plotted vector. (A) First two RDA axes of a non-814 815 conditioned RDA based on 11,261 SNPs. (B) First two RDA axes of a partial RDA based on 11,261 SNPs conditioned by neutral genetic structure, approximated by the 816 817 first two PCs of a PCA based on neutral markers. (C) First two RDA axes of a non-818 conditioned RDA based on 87 SNP outliers identified in a BayeScan analysis.



- **Figure 5.** Plot of the fractions of the genetic variability in Oregon samples explained by
- 821 (A) environmental variables alone; (B) the overlap of both environmental variables and
- 822 neutral structure; and (C) neutral genetic structure alone; and the unexplained genetic
- variability (residuals). P-values computed through a 1000-step permutation test for the
- fractions A, B and C, were below 0.001 in all cases.

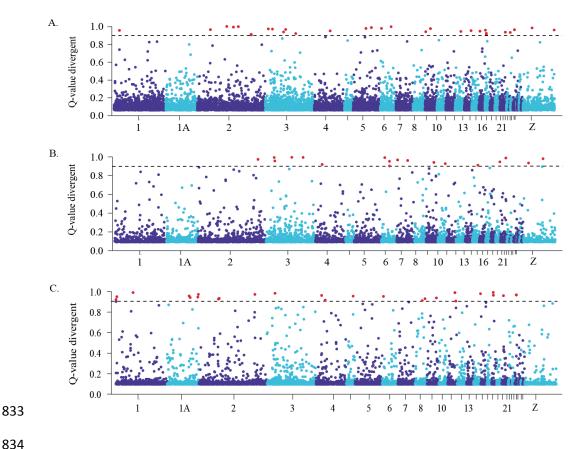
825



- Figure 6. Plot of per-SNP posterior probability of divergence mediated by selection 828
- 829 (shown as 1 – Q-value as computed by BayeScan) in (A) all Oregon junco forms

together; (B) townsendi against pontilis and (C) townsendi against all thurberi. Loci 830

- above the dotted line (in red) are those below a false discovery rate of 10%. 831
- 832



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