

Establishment of locally adapted mutations under divergent selection

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Abstract

We study the establishment probabilities of locally adapted mutations using a multi-type branching process framework. We find a surprisingly simple and intuitive analytical approximation for the establishment probabilities in a symmetric two-deme model under the assumption of weak (positive) selection. This is the first analytical closed-form approximation for arbitrary migration rate to appear in the literature. We find that the establishment probability lies between the weak and the strong migration limits if we condition the origin of the mutation to the deme where it is advantageous. This is not the case when we condition the mutation to first occur in a deme where it is disadvantageous. In this case we find that an intermediate migration rate maximizes the probability of establishment. We extend our results to the cases of multiple demes, two demes with asymmetric rates of gene flow, and asymmetric carrying capacities. The latter case allows us to illustrate how density regulation can affect establishment probabilities. Finally we use our results to investigate the role of gene flow on the rate of local adaptation and identify cases in which intermediate amounts of gene flow facilitate the rate of local adaptation as compared to two populations without gene flow.

24 **Keywords:**

25 local adaptation; branching process; establishment probabilities

26 **Introduction**

27 Studying the maintenance of genetic variation under migration-selection balance has a long
28 tradition in population genetics. While most theoretical research on the establishment and
29 maintenance of local adaptation and population divergence has focused on deterministic models
30 (reviewed in [Felsenstein, 1976, Karlin, 1982, Lenormand, 2002, Nagylaki and Lou, 2008]; see
31 also [Nagylaki and Lou, 2007, Star et al., 2007, Bürger, 009a, Bürger, 009b, Nagylaki, 2009]),
32 considerably less work has been done on the probability of establishment of locally adapted
33 mutations. Even in infinitely large populations, new beneficial mutations experience genetic
34 drift while they are rare, and hence can get lost from the population despite their selective
35 advantage. The probability that a new beneficial mutation evades extinction due to stochas-
36 tic fluctuations has been called the invasion probability, establishment probability or fixation
37 probability, depending on the context. In the simplest case of a single panmictic popula-
38 tion of infinite size, Haldane's classical result states that the establishment probability of a
39 mutation with time- and frequency-independent selection coefficient s is approximately $2s$
40 [Haldane, 1927]. Since then, Haldane's result has been generalized and extended to several
41 scenarios (see [Patwa and Wahl, 2008] for a review about fixation probabilities of beneficial
42 mutations).

43 Traditionally, there are two main approaches to study establishment probabilities: branch-
44 ing processes and diffusion approximations. Branching processes often allow for the deriva-
45 tion of simple and intuitive results [Harris, 2002], but are restricted to beneficial mutations
46 in (infinitely large) populations. The diffusion approximation, first used by [Kimura, 1962]
47 in this context, is a powerful tool that allows the derivation of results for both beneficial
48 or deleterious mutations of arbitrary initial frequency in finite populations. The downside is
49 that the derivation of closed form solutions is often harder as compared to branching pro-
50 cess, and that the underlying assumptions are not always clear. Applications of establish-
51 ment or fixation probabilities include the quantification of the rate of adaption of populations

52 [Orr and Otto, 1994, Wilke, 2004, Desai and Fisher, 2007, Gonçalves et al., 2007], extinction
53 risk due to the accumulation of deleterious mutations [Lynch and Gabriel, 1990], or the rate of
54 emergence of drug resistance or evolutionary rescue [Carlson et al., 2014].
55 In the context of spatially structured populations, [Barton, 1987] extended the diffusion approx-
56 imation to account for spatially variation in fitness along a one-dimensional habitat and derived
57 analytical solutions for some special cases. [Kirkpatrick and Peischl, 2013] used a similar ap-
58 proach to study the contribution of new mutations to evolutionary rescue in environments that
59 change in space and time. [McPeck and Holt, 1992] studied the conditions at which a geno-
60 type can invade populations fixed for another genotype in environments varying spatially or
61 temporally in a two-patch model. [Tachida and Iizuka, 1991, Gavrillets and Gibson, 2002] and
62 [Whitlock and Gomulkiewicz, 2005] have explored the probability of a single mutant allele fix-
63 ing in both patches of a two-patch model using diffusion approximations. These studies present
64 the fixation probability as the solution of a system of two quadratic equations that can be solved
65 numerically but so far no closed form solution for the fixation probability has been derived.
66 [Yeaman and Otto, 2011] have used a heuristic "splicing" approach in which they combine the
67 leading eigenvalues of the transition matrix of a deterministic two-deme model with Kimura's
68 classical fixation probability formula. Their approach is surprisingly accurate and allowed them
69 to determine the probability of a locally beneficial mutation becoming permanently established
70 and the critical threshold migration rate above which the maintenance of polymorphism is un-
71 likely in finite populations. [Vuilleumier et al., 2008] studied the fixation of locally beneficial
72 alleles through simulations of a metapopulation in a spatially heterogeneous environment. Their
73 findings suggest that a mutation experiencing strong positive selection in parts of an otherwise
74 neutral environment has a higher chance of reaching fixation than a unconditionally beneficial
75 mutation with the same average selection coefficient. This illustrates that heterogeneity in
76 selection coefficients across space can have a large impact on the probability of fixation.
77 Despite these advances, several open questions remain. Perhaps most importantly, no closed
78 form approximation for the establishment probability in a spatially structured population is
79 available, even for the simplest cases of two-demes with heterogeneous selection (apart from the
80 heuristic formula obtained in [Yeaman and Otto, 2011]). Furthermore, [Vuilleumier et al., 2010]
81 showed that details on how migration is modeled can have a large influence on the outcome

82 on the effects of spatial structure on establishing locally adapted mutations. The same study
83 also identified cases where fixation probabilities lie outside of the range set by low- and high-
84 migration limits, in contrast to what is observed in simpler analytical models
85 [Tachida and Iizuka, 1991, Gavrillets and Gibson, 2002, Whitlock and Gomulkiewicz, 2005]. De-
86 riving a closed expression for the establishment probability in patchy environments is necessary
87 to better understand the role of habitat fragmentation and dispersal on adaptation in spatially
88 heterogeneous environments, for instance in models of evolutionary rescue or evolution of drug
89 resistance [Gomulkiewicz and Holt, 1995], [Uecker et al., 2013], studying the spatial origin of
90 mutations that cause range expansion [Behrman and Kirkpatrick, 2011], or studying the role
91 of gene-flow on the establishment of local adaptation [Seehausen, 2004].
92 Here we study the establishment of locally adapted mutations in a discrete migration-selection
93 model using the framework of multitype branching processes. We present a surprisingly simple
94 and intuitive analytical approximation for the probability that new mutations escape genetic
95 drift and become permanently established. Our results are valid in the limit of weak positive
96 selection in one of the selective habitats but allow for arbitrary migration rates and arbitrarily
97 strong negative selection in the other habitat. We recover previous results for special cases
98 such as weak or strong migration. Our results allow us to quantify the effects of migration on
99 the fate of mutations – depending on whether mutations first occur in individuals living in the
100 deme where the mutation is adapted to or in the deme where the mutations is selected against.
101 We apply our results to biologically interesting scenarios and derive simple results for the ef-
102 fect of asymmetric carrying capacities and density regulation, as well as asymmetric migration,
103 the rate of local adaptation and the contribution of different demes to local adaptation. In
104 particular, we derive conditions under which gene flow between demes facilitates the rate of
105 establishment of locally adapted alleles as compared to the case without gene flow. The latter
106 result is in contrast with common wisdom that gene flow tends to hinder local adaptation, and
107 could have interesting implications for the the role of hybridization during adaptive radiations.

108 **Results**

109 We start with a symmetric two-deme model to present our model and derive our main result.
110 We then generalize our results to asymmetric migration, multiple demes and unequal carrying

111 capacities.

112 **Two-deme model**

113 We consider an infinitely large population at demographic equilibrium. Generations are discrete
114 and non-overlapping. The population is structured into two demes that exchange migrants.
115 Population size is regulated independently in the two demes, that is, selection is soft (sensu
116 [Wallace, 1975]). Migration is homogenous and isotropic and the two demes have the same
117 carrying capacity. Migration is therefore conservative and the number of immigrants and emi-
118 grants are equal in each deme. At the focal locus, a resident allele is fixed in both demes and
119 a new mutation occurs in a single individual. Each copy of the mutant allele in the population
120 produces a random number of descendant copies in the following generation (“offspring”) that
121 is independent of the number produced by other mutant copies. The mean number of offspring
122 of a mutant copy is given by $1 + s_i$ in deme i .

123 In the following, we will focus on the case where selection is acting in opposing directions in
124 the two demes, that is $\text{sign}(s_1) \neq \text{sign}(s_2)$, but note that our derivations do not require this
125 assumption. For the remainder we set $s_1 > 0 > s_2$ without loss of generality. In a haploid
126 population, s_i is the relative fitness (dis)advantage of a mutant in deme i , while in a randomly
127 mating diploid population, it is the relative fitness (dis)advantage of a heterozygote. Because
128 the ultimate fate of the mutation is decided while mutant homozygotes are still rare, we can
129 ignore their fitness. After reproduction each mutant copy migrates to the other deme with
130 probability $m/2$ and remains in its current deme with probability $(1 - m/2)$. Thus, $m = 0$
131 corresponds to two demes without gene flow whereas $m = 1$ corresponds to a single panmictic
132 population with frequency dependent selection (i.e., the Levene model, [Levene, 1953]). There
133 are two possible outcomes to the above described process: the mutant allele dies out or it
134 becomes permanently established. Note that establishment does not necessarily imply fixation
135 in our model: alleles may become permanently established in a balanced polymorphism under
136 certain conditions [Bulmer, 1972, Nagylaki, 2009]. We denote by $p^{(i)}$ the probability of estab-
137 lishment of a mutation that initially appears in an individual in deme i .

138 We model the evolution of the number of copies of a mutant allele that first appears in a single
139 individual using a branching processes with two types of individuals. The type i ($i \in \{1, 2\}$)

140 corresponds to the deme in which an individual carrying a copy of the mutant allele resides.
 141 We assume that the number of offspring of each copy of the mutant allele is independent of
 142 the offspring numbers of the rest of the population. The number of mutant copies present in
 143 generation n can then be described by a vector $X(n) = (X_1(n), X_2(n))$, where $X_i(n)$ denotes
 144 the number of mutant copies in deme i at generation n . Each random variable $X_i(n)$ has two
 145 associated random variables $\xi_j^{(i)}$, describing the number of offspring of type j from a parent of
 146 type i .
 147 The theory of multi-type branching processes (e.g., [Harris, 2002]) tells us that the vector of
 148 extinction probabilities $(1 - p^{(1)}, 1 - p^{(2)})$ is given by

$$\lim_{n \rightarrow \infty} \mathbf{f}^{(n)}(0, 0), \quad (1)$$

149 where \mathbf{f} is the vector of probability-generating functions of the offspring distribution and $\mathbf{f}^{(n)}$
 150 denotes the n -fold application of \mathbf{f} . It can be shown that the establishment probabilities are
 151 given by the smallest positive solution of $(1 - p^{(1)}, 1 - p^{(2)}) = (f_1(1 - p^{(1)}), f_2(1 - p^{(2)}))$, with
 152 f_i the i -th element of vector $\mathbf{f}^{(n)}$ (see Supplemental Material, equations S1–S17, for details).

Table 1: **The mean number of individual of type 1 and 2 for parents of type 1 and 2**

Type of the parent	Type of the offspring	
	1	2
1	$(1 + s_1)(1 - m/2)$	$(1 + s_2)m/2$
2	$(1 + s_1)m/2$	$(1 + s_2)(1 - m/2)$

153 Under a Wright-Fisher model of selection, offspring numbers are determined via binomial
 154 sampling from the parental generation, which can be approximated by Poisson-distributed
 155 offspring in large populations. The mean number of offspring for each type are summarized in
 156 table 1. Assuming Poisson distributed offspring, the establishment probabilities are then given
 157 by the smallest positive solution of

$$1 - p^{(1)} = e^{-(1 - \frac{m}{2})(1 + s_1)p^{(1)} - \frac{m}{2}(1 + s_2)p^{(2)}}, \quad (2)$$

$$1 - p^{(2)} = e^{-\frac{m}{2}(1 + s_1)p^{(1)} - (1 - \frac{m}{2})(1 + s_2)p^{(2)}}. \quad (3)$$

158 Equations (2) and (3) are transcendental equations for which one can in general not obtain

159 exact solutions and we resort to approximation. We first introduce new parameters that describe
160 the strength of the evolutionary forces relative to the strength of selection for the mutant allele
161 in deme 1: $\zeta = s_2/s_1$ and $\chi = m/s_1$. Assuming $s_1 > 0 > s_2$, and taking the limit of weak
162 selection, that is ignoring second- and higher-order terms in s_1 , in the Supplemental Material
163 (derivation of equations S17) we show that the establishment probabilities can be written as

$$p^{(1)} = \max[s_1(1 + \sigma) + s_2\mu, 0] \quad (4)$$

$$p^{(2)} = \max[s_1\mu + s_2(1 - \sigma), 0] \quad (5)$$

164 where

$$\sigma = \frac{s_1 - s_2}{\sqrt{m^2 + (s_1 - s_2)^2}} \quad (6)$$

165 and

$$\mu = \frac{m}{\sqrt{m^2 + (s_1 - s_2)^2}} \quad (7)$$

166 are scaled measures of the heterogeneity in selection and the migration rate, respectively. We
167 note that $\sigma, \mu \in [0, 1]$. Equations (4) and (5) show that the probability of establishment can
168 be written as a weighted sum of the strength of selection in the two demes. Figure 1 shows
169 establishment probabilities for various combinations of selection intensities.

170 The weak selection approximation requires that the establishment probability is small, that is,
171 $s_1 \ll 1$ and is valid even if the mutation is, on average, deleterious, *i.e.*, $s_1 + s_2 < 0$. Figure 2
172 shows comparison between the analytical approximation and exact solutions of equations (2)
173 and (3), obtained by numerical iteration of the probability-generating function. We find that
174 the approximation is very accurate with respect to the solutions of equations (2) and (3).

175 The establishment probabilities are positive if $m < \frac{s_1 s_2}{s_1 + s_2}$ or if $s_1 + s_2 > 0$. Note that this con-
176 dition is equivalent with the invasion conditions derived in deterministic models [Bulmer, 1972].
177 Because σ is monotonically decreasing in m and μ is monotonically increasing in m it follows
178 immediately that $p^{(1)}$ is monotonically decreasing in m (see figure 1). For $p^{(2)}$ the dependence
179 in m is more complicated. If $m = 0$, it is clear that $p^{(2)} = 0$ because $s_2 < 0$. Because $p^{(2)} > 0$ if
180 $m < \frac{s_1 s_2}{s_1 + s_2}$ or $s_1 + s_2 > 0$, $p^{(2)}$ is always maximized for some positive migration rate if $s_1 < 0 < s_2$
181 (figure 1). Straightforward calculations yield that either the maximum of $p^{(2)}$ is attained at
182 $m = \frac{s_1}{s_2}(s_2 - s_1) > 0$ or $p^{(2)}$ is monotonically increasing in m .

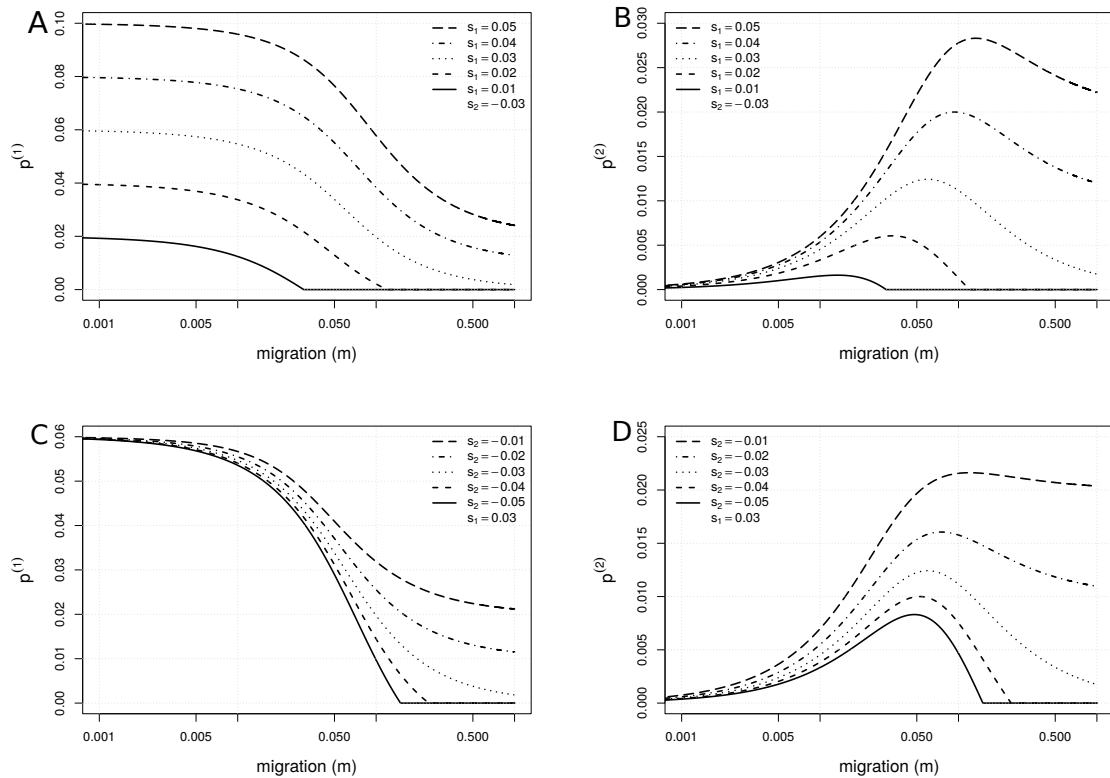


Figure 1: Establishment probabilities as a function of migration rate for various combinations of selection intensities. (A) $p^{(1)}$ with fixed $s_2 = -0.03$. (B) $p^{(2)}$ with fixed $s_2 = -0.03$. (C) $p^{(1)}$ with fixed $s_1 = 0.03$. (D) $p^{(2)}$ with fixed $s_1 = 0.03$.

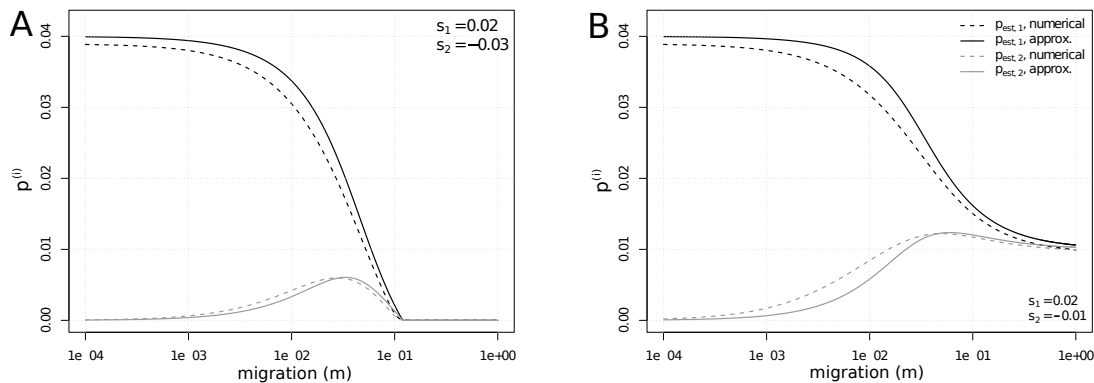


Figure 2: Comparison between exact solution of (2) and (3) and our approximation from equations (4) and (5). The exact solution is obtained numerically after 10'000 iterations of (2) and (3) (see equation (1)). (A) Probabilities of establishment for a scenario where $s_1 + s_2 < 0$. (B) Probabilities of establishment for a scenario where $s_1 + s_2 > 0$. The limit for very high migration is $p^{(1)} = p^{(2)} = s_1 + s_2$ (see main text).

183 Comparison with previous results

184 Equations (2) and (3) recover several previous results for establishment probabilities. In the
 185 absence of migration, we have that $\sigma = 1$ and $\mu = 0$ and we get $p^{(i)} = \max[2s_i, 0]$, in

186 agreement with Haldane’s classical result for a single panmictic population [Haldane, 1927].
187 In the limit of strong migration, we get $p^{(1)} = p^{(2)} = \max[s_1 + s_2, 0]$ which means that
188 the establishment probability is determined by the average selection coefficient across demes
189 [Nagylaki, 1980]. In the limit of weak migration, we get $p^{(1)} = 2s_1 - \frac{ms_2}{|s_1 - s_2|}$ and $p^{(2)} = \frac{ms_1}{|s_1 - s_2|}$.
190 [Gavrilets and Gibson, 2002] used a diffusion approximation to compute fixation probabilities
191 in a biallelic one-locus two-deme model similar to ours. The key difference between our and
192 their approach that we calculate establishment rather than fixation probabilities, which makes it
193 hard to directly compare our results. Furthermore, no closed-form solution is available for their
194 model. However, in the case where establishment implies fixation, our results are in very good
195 agreement (see Supplemental Material, figure S1). [Yeaman and Otto, 2011] extended classical
196 deterministic two-deme models for diploid individuals (such as [Bulmer, 1972]) to derive a
197 heuristic approximation for the establishment probability of new mutations. They calculate
198 the rate of increase in frequency of a rare locally beneficial mutation and use this initial growth
199 rate as a the selection coefficient in Kimura’s classical equation for fixation probabilities in finite
200 populations [Kimura, 1962]. Numerical comparison reveals a good fit between their results and
201 our equations (4) and (5) (see Supplemental Material, figure S2).

202 **Island model with multiple demes**

203 Here we extend our results to an island model with multiple demes and two selective habitats.
204 Let s_1 and s_2 denote the selection coefficients of the mutation in habitat 1 and 2, respec-
205 tively. We assume that migration occurs at rate m between all demes (*i.e.*, the island model,
206 [Wright, 1931]). This model can readily be reduced to a two-deme model with asymmetric
207 migration rates [Whitlock and Gomulkiewicz, 2005]. Let m_{ij} denote the migration rate from
208 selective habitat i to selective habitat j . We assume that all demes are of the same size and that
209 selective habitats 1 and 2 contain k_1 and $k_2 = n - k_1$ demes, respectively. The migration rates
210 between the two selective habitats are then given by $m_{12} = mk_2/n$ and $m_{21} = m(n - k_2)/n$.
211 Note that migration is conservative in this scenario since each deme sends and receives the
212 same number of migrants. In the Supplemental Material (equations S23), we show that the

213 establishment probabilities for the island model are given by

$$p_{\text{IM}}^{(1)} = \max[s_1(1 + \sigma - \Delta\mu) + s_2 \cdot \mu_{12}, 0] \quad (8)$$

$$p_{\text{IM}}^{(2)} = \max[s_1 \cdot \mu_{21} + s_2(1 - \sigma + \Delta\mu), 0] \quad (9)$$

214 where $\Delta\mu = (\mu_{12} - \mu_{21})/2$ (with $\mu_{ij} = m_{ij}/\lambda$), $\sigma = (s_1 - s_2)/\lambda$. In these last definitions, we
 215 used $\lambda = \sqrt{\bar{m}^2 + (s_1 - s_2)^2 - (m_{12} - m_{21})(s_1 - s_2)}$, with $\bar{m} = (m_{12} + m_{21})/2$. An $m_{\text{crit}} > 0$
 216 always exists, such that $p^{(i)}$ is positive if $m \in [0, m_{\text{crit}}]$; also, $p_{\text{IM}}^{(1)}$ is monotonically decreasing.
 217 The theoretical maxima of the functions can be easily calculated (see Supplemental Material,
 218 derivation of equation S23).

219 If the selective habitat where the mutation is beneficial is larger than the habitat where the
 220 mutation is detrimental, we have more individuals migrating from habitat 2 to habitat 1 and
 221 $\Delta\mu < 0$. As a consequence the contribution of selection in habitat 1 is amplified as compared
 222 to the case with symmetric migration (see eqs. (8) and (9)), and the chances of establishment
 223 generally increase with respect to the symmetric migration model (Figure 3). We note that
 224 even though we introduced this scenario as a multiple-deme island model, the result (8) and
 (9) are also valid in a 2-deme model with asymmetric migration between demes.

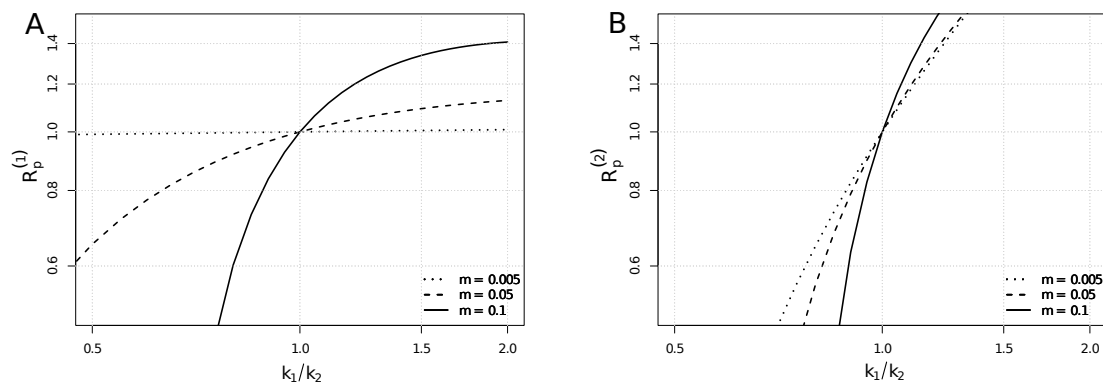


Figure 3: Ratio between probabilities of establishment computed within the model with symmetric migration or within the island model, defined as $R_p^{(i)} = p_{\text{IM}}^{(i)}/p^{(i)}$, as a function of k_1/k_2 . (A) $i = 1$ and (B) $i = 2$. In both cases $s_1 = 0.02$ and $s_2 = -0.03$.

225

226 Asymmetric carrying capacities and density regulation

227 So far we ignored the effects of density regulation because we assumed demographic equilibrium
 228 and that migration is conservative, i.e., gene flow does not change the number of individuals in

229 each deme. We next relax the assumption of equal deme sizes and introduce a simple model
230 of deme-independent density regulation (soft selection, sensu Wallace 1953). Let κ_1 and κ_2
231 denote the carrying capacities of deme 1 and 2, respectively. If migration is isotropic the larger
232 deme acts as a source, that is, it sends out more migrants than it receives. Here we assume
233 that density regulation acts after migration and brings each deme back to its carrying capacity
234 instantaneously. Let κ_i denote the carrying capacity of deme i . Initially both demes are at
235 carrying capacity. The number of individuals in deme i after migration but before density
236 regulation are denoted N'_i and are given by

$$N'_1 = \kappa_1(1 - m) + \kappa_2 m, \quad \text{and} \quad N'_2 = \kappa_1 m + \kappa_2(1 - m). \quad (10)$$

237 Density regulation will then change the number of individuals in each deme by a factor

$$\delta_i = \frac{\kappa_i}{\kappa_i(1 - m) + \kappa_j m}; \quad i, j \in \{1, 2\}, \quad i \neq j. \quad (11)$$

238 This modifies the absolute fitness of individuals in deme i (see table 1) to $w_i = (1 + s_i)\delta_i$.

239 The establishment probabilities for this case are explicitly calculated in the Supplemental Ma-
240 terial (see equation (S28)).

241 If $\kappa_1 < \kappa_2$, deme 1 receives more migrants than it sends out and density regulation will cull the
242 population size back to carrying capacity. Thus, in a certain sense deme 1 is behaving like a
243 shrinking population, which should reduce the establishment probability of mutations that are
244 beneficial in that deme [Otto and Whitlock, 1997]. Our results confirms this intuition (figure
245 4) and show that the establishment probability increases if the deme where the allele is bene-
246 ficial has a smaller carrying capacity. Likewise, if $\kappa_2 < \kappa_1$, deme 1 is growing after migration,
247 which reduces drift and increases the establishment probability of mutations that are adapted
248 to deme.

249 **Comparison with simulations**

250 A comparison with simulations yields good results (see figure 5) for the case of symmetric demes,
251 asymmetric migration and also both cases of asymmetric demes (*i.e.* $\kappa_1 > \kappa_2$ and $\kappa_1 < \kappa_2$).
252 Simulations were performed assuming that the population dynamics is represented through a

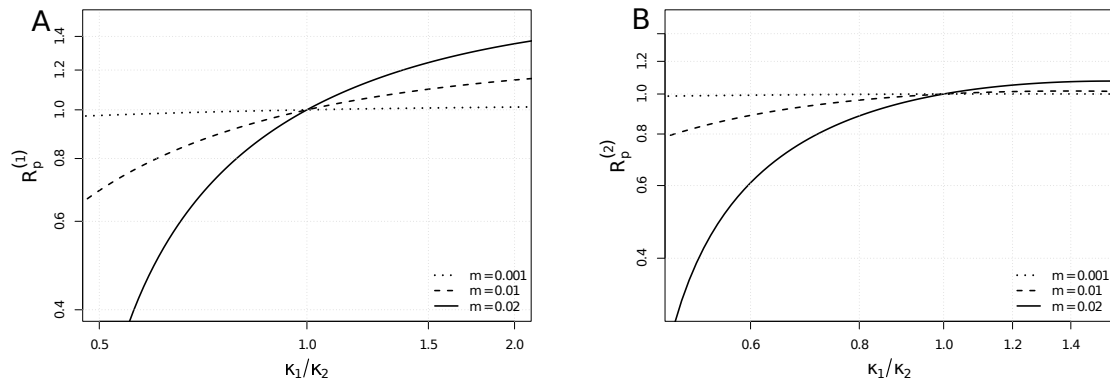


Figure 4: Ratio between probabilities of establishment computed within the model with symmetric migration or within the model with asymmetric carrying capacities (see equations for $p_{\text{dens}}^{(i)}$, Supplemental Material, equation (S28)), defined as $R_p^{(i)} = p_{\text{dens}}^{(i)}/p^{(i)}$, as a function of κ_1/κ_2 . (A) $i = 1$ and (B) $i = 2$. In both cases $s_1 = 0.02$ and $s_2 = -0.03$.

253 logistic growth, and inserting one mutant in either deme 1 or 2 and letting the system evolve
 254 for 20000 generations. 50000 replications are done for each simulated scenario. Simulations
 255 were done with small populations (300). Our approximation tends to overestimate $p^{(1)}$ and
 256 underestimate $p^{(2)}$. This is expected, since the same behavior can be seen when compared the
 257 approximation to the exact solution (see figure 2).

258 Global rate of establishment of locally adapted alleles

259 It is commonly assumed that gene flow hampers or even prevents local adaptation [Lenormand, 2002].
 260 Here we use our results to quantify the effect of migration on the overall establishment proba-
 261 bility, defined as

$$P = cp_{\text{est}}^{(1)} + (1 - c)p_{\text{est}}^{(2)} \quad (12)$$

262 where c and $1 - c$ denote the relative sizes of deme 1 and deme 2, respectively.

263 We next take the derivative of P with respect to m at $m = 0$. If this derivative is positive,
 264 the rate of adaptation increases when we introduce some gene flow as compared to the case
 265 without gene flow between demes. In other words, we derived the condition under which
 266 the unconditional establishment probability of locally adapted mutations increases when we
 267 introduce small amounts of migration. We find that this is the case in all our models if

$$N_1|s_1| > N_2|s_2| \quad (13)$$

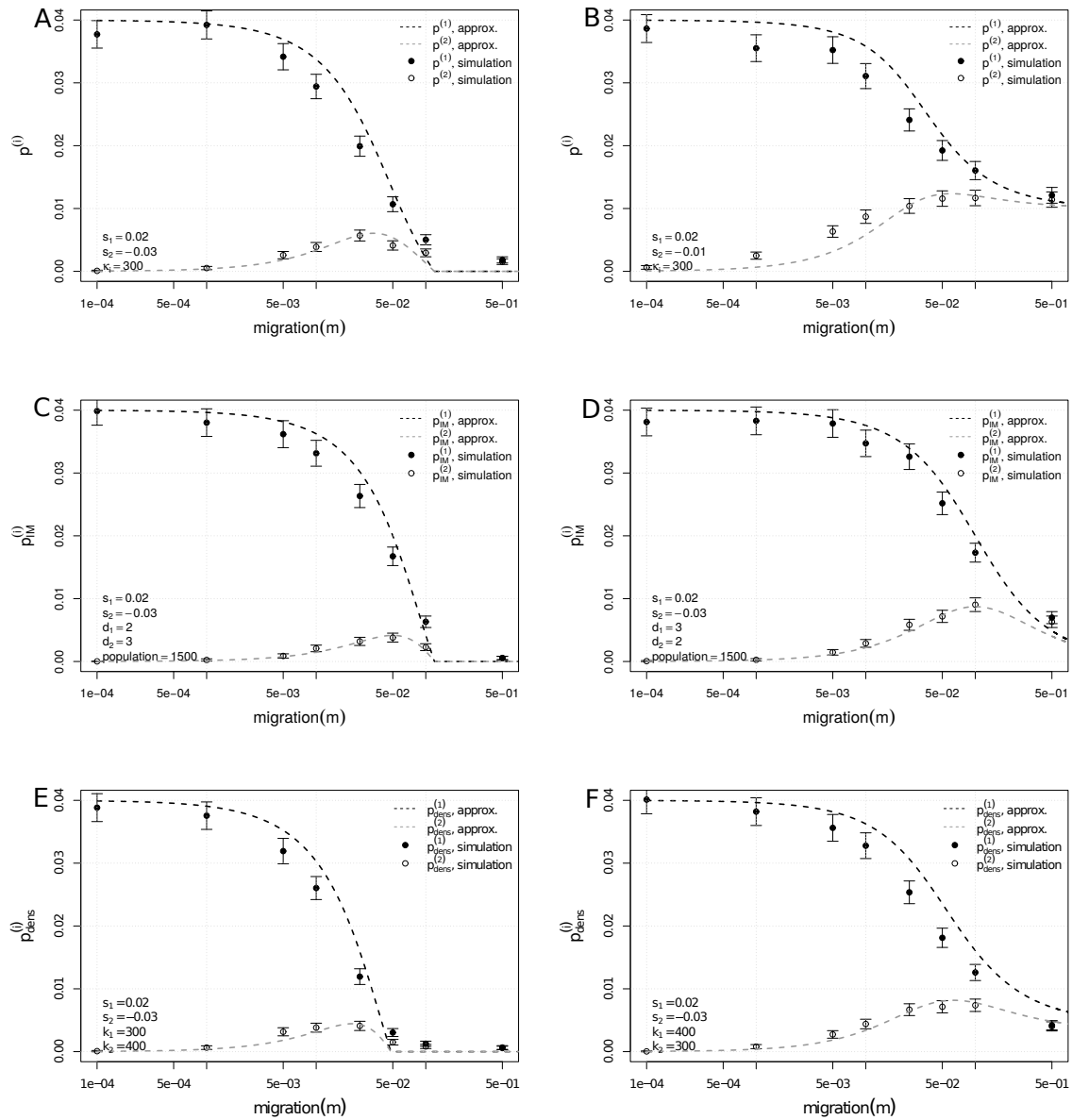


Figure 5: Comparison of simulations and analytical approximations. Our model is represented by the dashed line. (A) Model with symmetric migration and symmetric demes, $s_1 = 0.02$, $s_2 = -0.03$. (B) Model with symmetric migration and symmetric demes, $s_1 = 0.02$, $s_2 = -0.01$. (C) Island model with two demes of type 1 ($d_1 = 2$, $s_1 = 0.02$) and three demes of type 2 ($d_2 = 3$, $s_2 = -0.03$). (D) Island model with three demes of type 1 ($d_1 = 3$, $s_1 = 0.02$) and two demes of type 2 ($d_2 = 2$, $s_2 = -0.03$). (E) Model with different carrying capacities, with carrying capacities $\kappa_1 = 300$ and $\kappa_2 = 400$. (F) Model with different carrying capacities, with carrying capacities $\kappa_1 = 400$ and $\kappa_2 = 300$. The form of $p_{\text{dens}}^{(i)}$, shown in images (E) and (F), is described in the Supplementary Information (equations S28).

268 where N_1 and N_2 are the number of individuals in deme 1 and 2 respectively. Hence, in the
269 symmetric model where the probability of establishment is given by (4) and (5), we find that
270 gene flow increases the chances of establishment when $|s_1| > |s_2|$. Therefore, if the selective
271 advantage in one deme is larger than the selective disadvantage in the other deme, some gene
272 flow can facilitate the establishment of local adaptation as compared to completely isolated
273 demes. This can also be seen directly via the weak migration approximation $p_{\text{est}}^{(1)} = 2s_1 - \frac{ms_2}{|s_1 - s_2|}$
274 and $p_{\text{est}}^{(2)} = \frac{ms_1}{|s_1 - s_2|}$ derived above.
275 If we consider the model with asymmetric migration in the context of a multi-deme model (see
276 above), the condition becomes

$$m_{21}|s_1| > m_{12}|s_2|. \quad (14)$$

277 Using the definitions of m_{12} and m_{21} that we defined in the section above, for the multi-deme
278 model, equations (13) and (14) are identical.

279 Discussion

280 We used multi-type branching processes to study the establishment of locally beneficial mu-
281 tations in a spatially heterogeneous environment with two selective habitats. Our main result
282 is a simple and analytical closed-form approximation for the probability of establishment of a
283 locally beneficial mutation in a two-deme model with divergent selection and symmetric mi-
284 gration between demes (equations (4) and (5)). By establishment we mean that a mutation
285 permanently establishes in the meta-population, either by going to fixation or by maintenance
286 as a balanced polymorphism. To our knowledge this is the first closed-form analytical approx-
287 imation for an establishment probability in this context that is valid for arbitrary migration
288 rates (but see [Yeaman and Otto, 2011] for a heuristic approach). The resulting formula is
289 intriguingly simple and intuitive: the probability of establishment is simply a weighted average
290 over selection coefficients in the two demes, where the weights are determined by the relative
291 contributions of migration and spatial variation in selection. We extended our main result to
292 asymmetric migration between and a multi-deme island model with two selective habitats, and
293 studied the impact of variation in carrying capacities and density regulation on the establish-
294 ment of locally adapted mutations. We show that establishment probabilities can fall outside

295 the range spanned by the weak or strong migration limits, and provide conditions for when this
296 is the case. In particular, we identify conditions under which small amounts of migration can
297 facilitate the build-up of adaptive divergence as compared to two demes without gene-flow.
298 Our derivation assumes an (infinitely) large population and that selection for the beneficial
299 mutation is sufficiently weak. Let $s_1 > 0$ denote the selection coefficient in the deme where the
300 mutation is beneficial. Our approximation should hold as long as the product of population
301 size and selection coefficient s_1 is larger than 1 ($Ns_1 > 1$, fig.5) and selection is so weak that
302 we ignore second- and higher-order terms in s . Worthy of note, our approximation remains
303 valid if the mutation is strongly deleterious (or even lethal) in one of the demes (see figure S3
304 in the Supplementary Material). Furthermore, we have modeled a haploid population to avoid
305 the intricacies of dominance. However, our results should remain valid in diploid populations if
306 we use the fitness of the mutant allele in our model as the fitness of heterozygotes in a diploid
307 model. The ultimate fate of a mutation will be determined while it is rare so that we can ignore
308 the fitness of homozygotes.

309 We have been discussing single mutations in isolation and neglected genetic events that may
310 interfere with the establishment process (e.g., clonal interference [Gerrish and Lenski, 1998]).
311 Our results should therefore hold in sexually reproducing species with strong recombination.
312 Our approximation is less plausible in organisms that reproduce with little or no recombina-
313 tion, such as most microbes, or for mutations in genomic regions with low recombination rates.
314 Competition between simultaneously spreading beneficial mutations (clonal interference) can
315 have severe impacts on each other's establishment [Gerrish and Lenski, 1998, Orr, 2000]. This
316 effect is difficult to account for [Wilke, 2004] and is most important in populations with little or
317 no recombination, because recombination can break associations between mutations that occur
318 in different parts of a genome [Muller, 1932, Hill and Robertson, 1966]. A second effect that
319 we did not account for is the genomic background on which the mutation falls via additive or
320 epistatic interactions between mutations. This can be sidestepped by interpreting the selection
321 coefficient as referring to the focal individuals fitness rather than to the effect of the mutation.
322 Previous work based on individual based simulations has shown that variation in densities
323 across demes can affect the establishment of new mutations [Vuilleumier et al., 2010]. Our
324 results confirm this and show selection is more efficient in source-demes. We show that the

325 combination of density regulation and asymmetric migration mimics the effects of a growing
326 population, which increases the absolute fitness of individuals and leads to more efficient posi-
327 tive selection [Otto and Whitlock, 1997].

328 The solution that we presented in this paper also assumes that migration rates remain fixed
329 in time. We know, however, that spatially varying selection can lead to evolution of dispersal
330 [Ronce, 2007]. Our model is therefore plausible if migration is mainly determined by geograph-
331 ical features of the environment, or if there is little or no genetic variability for traits related
332 to dispersal. Furthermore, variation in density can introduce non-random movement between
333 demes. We have accounted for this in eqs. (8) and (9) where we allowed migration rates to
334 be asymmetric. In equations (8) and (9), the carrying capacity of each type of habitat is not
335 explicitly taken into account. A combination of the result for asymmetric migration rates with
336 the scenario with density regulation (see equations S28) may be necessary to study extreme
337 cases in which asymmetry of migration is related with asymmetry in carrying capacities. While
338 this should be possible in a multi-type branching process framework, the derivations are com-
339 plicated and are beyond the scope of this work.

340 We have presented here a mathematically rigorous approximation of establishment probabilities
341 in a spatial framework using the theory of multi-type branching processes. It would be very
342 interesting to generalize our approach to more than two different types of individuals. While
343 the theoretical foundation is laid out, finding actual solutions for establishment probabilities
344 in higher-dimensional system poses algebraic challenges that might be difficult to overcome.
345 Nevertheless, the simple and intuitive form of our solution suggests that this approach can be
346 exploited further and that our results can be generalized and extend to various scenarios that
347 include more than two types of individuals.

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