1 A genomic region containing *REC8* and *RNF212B* is associated with individual

2 recombination rate variation in a wild population of red deer (*Cervus elaphus*).

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6 **Abstract**

Recombination is a fundamental feature of sexual reproduction, ensuring proper disjunction,

preventing mutation accumulation and generating new allelic combinations upon which selec-

tion can act. However it is also mutagenic, and breaks up favourable allelic combinations pre-

viously built up by selection. Identifying the genetic drivers of recombination rate variation is a

key step in understanding the causes and consequences of this variation, how loci associated

with recombination are evolving and how they affect the potential of a population to respond to

selection. However, to date, few studies have examined the genetic architecture of recombina-

tion rate variation in natural populations. Here, we use pedigree data from ~2,600 individuals

genotyped at ~38,000 SNPs to investigate the genetic architecture of individual autosomal re-

combination rate in a wild population of red deer (*Cervus elaphus*). Female red deer exhibited

a higher mean and phenotypic variance in autosomal crossover counts (ACC). Animal models

fitting genomic relatedness matrices showed that ACC was heritable in females (h^2 = 0.12) but

not in males. A regional heritability mapping approach showed that almost all heritable varia-

tion in female ACC was explained by a genomic region on deer linkage group 12 containing the

candidate loci *REC8* and *RNF212B*, with an additional region on linkage group 32 containing

TOP2B approaching genome-wide significance. The *REC8/RNF212B* region and its paralogue

RNF212 have been associated with recombination in cattle, mice, humans and sheep. Our

findings suggest that mammalian recombination rates have a relatively conserved genetic ar-

chitecture in both domesticated and wild systems, and provide a foundation for understanding

the association between recombination loci and individual fitness within this population.

27 Introduction

Meiotic recombination (or crossing-over) is a fundamental feature of sexual reproduction and

an important driver of diversity in eukaryotic genomes (FELSENSTEIN, 1974; BARTON and

CHARLESWORTH, 1998). It has several benefits: it ensures the proper disjunction of homol-

ogous chromosomes during meiosis (HASSOLD and HUNT, 2001), prevents mutation accumu-

lation (MULLER, 1964) and generates novel haplotypes, increasing the genetic variance for fit-

ness and increasing the speed and degree to which populations respond to selection (HILL and

ROBERTSON, 1966; BATTAGIN et al., 2016). However, recombination can also come at a cost:

it requires the formation of DNA double strand breaks which increase the risk of local muta-

tion and chromosomal rearrangements (INOUE and LUPSKI, 2002; ARBEITHUBER et al., 2015);

it can also break up favourable allele combinations previously built up by selection, reducing

the mean fitness of successive generations (BARTON and CHARLESWORTH, 1998). Therefore,

as the relative costs and benefits of recombination vary within different selective contexts, it is

expected that recombination rates should vary within and between populations (BURT, 2000;

OTTO and LENORMAND, 2002). Indeed, recent studies have shown that recombination rates

can vary within and between chromosomes (i.e. recombination "hotspots"; MYERS et al. 2005),

individuals (KONG et al., 2004), populations (DUMONT et al., 2011) and species (STAPLEY et al.,

2017).

Genomic studies in humans, cattle, sheep and mice have shown that variation in recombination

rate is often heritable, and may have a conserved genetic architecture (KONG et al., 2014; MA

et al., 2015; JOHNSTON et al., 2016; PETIT et al., 2017). The loci *RNF212*, *REC8* and *HEI10*,

amongst others, have been identified as candidates driving variation in rate, with *PRDM9* driving

recombination hotspot positioning in mammals (BAUDAT et al., 2010; BAKER et al., 2017). This

oligogenic architecture suggests that recombination rates and landscapes have the potential to

evolve rapidly under different selective scenarios, in turn affecting the rate at which populations

respond to selection (BARTON and CHARLESWORTH, 1998; BURT, 2000; OTTO and BARTON,

2001; GONEN et al., 2017). However, it remains unclear how representative the above studies

are of recombination rate variation and its genetic architecture in natural populations. For exam-

ple, experimental and domesticated populations tend to be subject to strong selection and have

small effective population sizes, both of which have been shown theoretically to indirectly select

for increased recombination rates to escape Hill-Robertson interference (OTTO and BARTON

2001; OTTO and LENORMAND 2002; but see MUÑOZ-FUENTES *et al.* 2015). Therefore, it may

be that prolonged artificial selection results in different recombination dynamics and underlying

genetic architectures. As broad recombination patterns are characterised in greater numbers

of natural systems (JOHNSTON *et al.*, 2016, 2017; THEODOSIOU *et al.*, 2016; KAWAKAMI *et al.*,

2017), it is clear that broad and fine-scale recombination rates and landscapes can vary to a

large degree even within closely related taxa (STAPLEY *et al.*, 2017). Therefore, determining

the genetic architecture of recombination rate in non-model, natural systems are key to eluci-

dating the broad evolutionary drivers of recombination rate variation and quantifying its costs

and benefits at the level of the individual.

In this study, we investigate the genetic basis of recombination rate variation in a wild population

of Red deer (*Cervus elaphus*) on the island of Rùm, Scotland (CLUTTON-BROCK *et al.*, 1982).

This population has been subject to a long term study since the early 1970s, with extensive

pedigree and genotype information for $\sim$2,600 individuals at >38,000 SNPs (HUISMAN *et al.*,

2016; JOHNSTON *et al.*, 2017). We use this dataset to identify autosomal crossover rates and

their genetic architecture in >1,300 individuals. The aims of the study are to: (a) determine

which common environmental and individual effects, such as age, sex and birth year affect in-

dividual recombination rates; (b) determine if recombination rate is heritable; and (c) identify

genomic regions that are associated with recombination rate variation. Addressing these ob-

jectives will provide a foundation for future studies investigating the association between the

genetic architecture of recombination rate and individual fitness, to determine how this trait

evolves within contemporary natural populations.

79 Materials and Methods

80 Study population and genomic dataset.

The study population of red deer is situated in the North Block of the Isle of Rùm, Scotland

(57°02'N, 6°20'W) and has been subject to individual monitoring since 1971 (CLUTTON-BROCK

et al., 1982). Research was conducted following approval of the University of Edinburgh's An-

imal Welfare and Ethical Review Body and under appropriate UK Home Office licenses. DNA

was extracted from neonatal ear punches, cast antlers and post-mortem tissue (see HUISMAN

et al., 2016 for full details). DNA samples from 2880 individuals were genotyped at 50,541

SNP loci on the Cervine Illumina BeadChip (BRAUNING *et al.*, 2015) using an Illumina geno-

typing platform (Illumina Inc., San Diego, CA, USA). SNP genotypes were scored using Illu-

mina GenomeStudio software, and quality control was carried out using the *check.marker* func-

tion in GenABEL v1.8-0 (AULCHENKO *et al.*, 2007) in R v3.3.2, with the following thresholds:

SNP genotyping success >0.99, SNP minor allele frequency >0.01, and ID genotyping success

>0.99, with 38,541 SNPs and 2,631 IDs were retained. There were 126 pseudoautosomal SNPs

identified on the X chromosome (i.e. markers showing autosomal inheritance patterns; JOHN-

STON *et al.*, 2017). Heterozygous genotypes within males at non-pseudoautosomal X-linked

SNPs were scored as missing. A pedigree of 4,515 individuals has been constructed using

microsatellite and SNP data using the software Sequoia (see HUISMAN, 2017). The genomic

inbreeding coefficient ($\hat{F}_{III}$), was calculated for each deer in the software GCTA v1.24.3 (YANG

et al., 2011), using information for all autosomal SNP loci passing quality control. A linkage

map of 38,083 SNPs has previously been constructed, with marker orders and estimated base-

pair positions known for all 33 autosomes (CEL1 to CEL33) and the X chromosome (CEL34)

(JOHNSTON *et al.*, 2017 and data archive doi:10.6084/m9.figshare.5002562). All chromosomes

are acrocentric with the exception of one metacentric autosome (CEL5).

Quantification of meiotic crossovers.

A standardised sub-pedigree approach was used to identify the positions of meiotic crossovers

(JOHNSTON *et al.*, 2016). The full pedigree was split as follows: for each focal individual (FID)

and offspring pair, a sub-pedigree was constructed that included the FID, its mate, parents and

offspring (Figure S1), where all five individuals were genotyped on the SNP chip. This pedi-

gree structure allows phasing of SNPs within the FID, characterising the crossovers occurring

in the gamete transferred from the FID to the offspring. All remaining analyses outlined in this

section were conducted in the software CRI-MAP v2.504a (GREEN *et al.*, 1990) within the R

package crimaptools v0.1 (JOHNSTON *et al.*, 2017) implemented in R v3.3.2. Mendelian incom-

patibilities within sub-pedigrees were identified using the *prepare* function and removed from

all affected individuals; sub-pedigrees containing more than 0.1% mismatching loci between

parents and offspring were discarded. The *chrompic* function was used to identify the grand-

parental phase of SNP alleles on chromosomes transmitted from the FID to the offspring, and

to provide a sex-averaged linkage map. Switches in phase indicated the position of a crossover

(Figure S1). Individuals with high numbers of crossovers per gamete (>60) were assumed to

have widespread phasing errors and were removed from the analysis; the maximum number of

crossovers for an individual was 45 in the remaining dataset.

Errors in determining allelic phase can lead to incorrect calling of double crossovers (i.e. $\geq$

2 crossovers occurring on the same chromosome) over short map distances. To reduce the

likelihood of calling false double crossover events, phased runs consisting of a single SNP were

recoded as missing (390 out of 7652 double crossovers; Figure S2) and *chrompic* was rerun.

Of the remaining double crossovers, those occurring over distances of $\leq$ 10cM (as measured

by the distance between markers immediately flanking the double crossover) were recoded as

missing (170 out of 6959 double crossovers; Figure S2). After this process, 1341 sub-pedigrees

passed quality control, characterising crossovers in gametes transmitted to 482 offspring from

81 unique males and 859 offspring from 256 unique females.

Genetic architecture of recombination rate variation.

Heritability and cross-sex genetic correlation. Autosomal crossover count (ACC) was mod-

elled as a trait of the FID. A restricted maximum-likelihood (REML) "animal model" approach

(HENDERSON, 1975) was used to partition phenotypic variance and examine the effect of fixed

effects on ACC; these were implemented in ASReml-R (BUTLER *et al.*, 2009) in R v3.3.2. The

additive genetic variance was calculated by fitting a genomic relatedness matrix (GRM) con-

structed for all autosomal markers in GCTA v1.24.3 (YANG *et al.*, 2011); the GRM was adjusted

assuming similar frequency spectra of genotyped and causal loci using the argument *--grm-adj*

0. There was no pruning of related individuals from the GRM (i.e. we did not use the *--grm-cutoff*

argument) as there is substantial relatedness within the population, and initial models included

parental effects and common environment which controls for effects of shared environments

between relatives. ACC was modelled first using a univariate model:

$$y = X\beta + Z_1 a + Z_r u_r + e$$

where y is a vector of ACC; X is an incidence matrix relating individual measures to a vector

of fixed effects, β; Z_1, and Z_r are incidence matrices relating individual measures with addi-

tive genetic and random effects, respectively; a and u_r are vectors of GRM additive genetic

and additional random effects, respectively; and e is a vector of residual effects. The narrow-

sense heritability h^2 was calculated as the ratio of the additive genetic variance to the sum of

variance components estimated for all random effects. Model structures were tested with sev-

eral fixed effects, including sex, $\hat{F}_{III}$ and FID age; random effects included individual identity

(i.e. permanent environment) to account for repeated measures in the same FID, maternal

and paternal identity, and common environment effects of FID birth year and offspring birth

year. The significance of fixed effects was tested with a Wald test, and the significance of ran-

dom effects was calculated using likelihood-ratio tests (LRT, defined as $2 \times$ the difference in

log-likelihoods between the two models, distributed as χ^2 with 1 degree of freedom) between

models with and without the focal random effect. $\hat{F}_{III}$ and individual identity were retained in all

models, irrespective of statistical significance, to account for possible underestimation of ACC

and pseudoreplication, respectively. As the variance in recombination rates differed between

the sexes, models were also run for each sex separately.

Bivariate models of male and female ACC were run to determine whether additive genetic vari-

ation was associated with sex-specific variation and the degree to which this was correlated

between the sexes. The additive genetic correlation r_A was determined using the CORGH

error-structure function in ASReml-R (correlation with heterogeneous variances) with r_A set to

be unconstrained. Model structure was otherwise the same as for univariate models. To deter-

mine whether genetic correlations were significantly different from 0 and 1, the unconstrained

model was compared with models where r_A was fixed at values of 0 or 0.999. Differences in

additive genetic variance in males and females were tested by constraining both to be equal

values using the CORGV error-structure function in ASReml-R. Models then were compared

using LRTs with 1 degree of freedom.

Genome-wide association study Genome-wide association studies (GWAS) of ACC were

conducted using the function *rGLS* in the R library RepeatABEL v1.1 (RÖNNEGÅRD *et al.*, 2016)

implemented in R v3.3.2. This function accounts for population structure by fitting the GRM as

a random effect, and allows fitting of repeated phenotypic measures per individual. Models

were run including sex and $\hat{F}_{III}$ as fixed effects; sex-specific models were also run. Associa-

tion statistics were corrected for inflation due to population stratification that was not captured

by the GRM, by dividing them by the genomic control parameter λ, which was calculated as

the observed median χ^2 statistic divided by the null expectation median χ^2 statistic (DEVLIN

et al., 1999). The significance threshold after multiple testing was calculated using a linkage

disequilibrium (LD) based approach in the software $K_{effective}$ (MOSKVINA and SCHMIDT, 2008)

specifying a sliding window of 50 SNPs. The effective number of tests was calculated as 35,264,

corresponding to a P value of 1.42×10^{-06} at α = 0.05. GWAS of ACC included the X chromo-

some and 458 SNP markers of unknown position. It is possible that some SNPs may show an

association with ACC if they are in LD with polymorphic recombination hotspots (i.e. associa-

tions in *cis*), rather than SNPs associated with recombination rate globally across the genome

(i.e. associations in *trans*). Therefore, we repeated the GWAS modelling *trans* variation only,

by examining associations between each SNP and ACC, minus the crossovers that occurred

on the same chromosome as the SNP. For example, if the focal SNP occurred on linkage group

1, association was tested with ACC summed over linkage groups 2-33. Marker positions are

known relative to the cattle genome vBTA_vUMD_3.1; in cases of significant associations with

recombination rate, gene annotations and positions were obtained from Ensembl (Cattle gene

build ID BTA_vUMD_3.1.89). LD was calculated between loci in significantly associated re-

gions using the allelic correlation r^2 in the R package LDheatmap v0.99-2 (SHIN *et al.*, 2006) in

R v3.3.2.

Regional heritability analysis As a single locus approach, GWAS has reduced power to

detect variants with small effect sizes and/or low linkage disequilibrium with causal mutations

(YANG *et al.*, 2011). Partitioning additive genetic variance within specific genomic regions (i.e.

a regional heritability approach) incorporates haplotype effects and determines the proportion

of phenotypic variance explained by defined regions. The additive genetic variance was par-

titioned across all autosomes in sliding windows of 20 SNPs (with an overlap of 10 SNPs) as

follows (NAGAMINE *et al.*, 2012; BÉRÉNOS *et al.*, 2015):

$$y = X\beta + Z_1 v_i + Z_2 v_{ni} + Z_r u_r + e$$

where y is a vector of ACC; X is an incidence matrix relating individual measures to a vector

of fixed effects, β; v_i is a vector of additive genetic effects explained by autosomal genomic

region in window i; v_{ni} is the vector of the additive genetic effects explained by all autosomal

markers not in window i; Z_1, and Z_2 are incidence matrices relating individual measures with

additive genetic effects for the focal window and the rest of the genome, respectively; Z_r is

an incidence matrix relating individual measures with additional random effects, where u_r is a

vector of additional random effects; and e is a vector of residual effects. The mean window

size was 1.29 $\pm$ 0.32 Mb. Models were implemented in ASReml-R (BUTLER *et al.*, 2009) in R

v3.3.2. GRMs were constructed in the software GCTA v1.24.3 with the argument *--grm-adj 0*

(YANG *et al.*, 2011). The significance of additive genetic variance for window i was tested by

comparing models with and without the $Z_1 v_i$ term with LRT (χ^2_1). To correct for multiple testing,

a Bonferroni approach was used, taking the number of windows and dividing by 2 to account

for window overlap; the threshold P-value was calculated as 2.95×10^{-5} at α = 0.05. In the

most highly associated region, this analysis was repeated for windows of 20, 10 and 6 SNPs in

sliding windows overlapping by $n - 1$ SNPs in order to fine map the associated regions. This

was carried out from approximately 5MB before and after the significant region.

Accounting for sample size difference between males and females. Sample sizes within

this dataset are markedly different between males and females (see above and Table 1). A

consequence of this may be that there is lower power to detect associations with male recombi-

nation rate. We repeated the heritability and GWAS analyses in sampled datasets of the same

size within each sex. Briefly, 482 recombination rate measures (representing the total number

in males) were sampled with replacement within the male and female datasets, and the animal

model and GWAS analyses were repeated in the sampled dataset. This process was repeated

100 times, with sampling carried out in R v3.3.2. The observed and simulated heritabilities com-

pared to see how often a similar results would be obtained. This was repeated for association

at the most highly associated GWAS SNPs and regional heritability regions. The differences

between the mean simulated values in each sex were investigated using a Welch two-sample

t-test assuming unequal variances.

Haplotyping and effect size estimation. Haplotype construction was carried out to exam-

ine haplotype variation within regions significantly associated with recombination rate variation

in the regional heritability analysis. SNP data from deer linkage group 12 was phased using

SHAPEIT v2.r837 (DELANEAU *et al.*, 2012), specifying the linkage map positions and recombi-

nation rates for each locus. This analysis used pedigree information with the *--duohmm* flag to

allow the use of pedigree information in the phasing process (O'CONNELL *et al.*, 2014). Haplo-

types were then extracted for the most significant window from the regional heritability analysis.

Effect sizes on ACC for the top GWAS SNPs were estimated using animal models in ASReml-

R; SNP genotype was fit as a fixed factor, with pedigree relatedness fit as a random effect to

account for the remaining additive genetic variance. To determine the effect sizes on ACC for

the regional heritability analysis, animal models were run as follows: for a given haplotype, A,

its effect was estimated relative to all other haplotypes combined, i.e. treating them as a single

allele, B, by fitting genotypes A/A, A/B and B/B as a fixed factor. This was repeated for each

haplotype allele where more than 10 copies were present in the full dataset.

Data availability

Raw data are publicly archived at doi:10.6084/m9.figshare.5002562 (JOHNSTON *et al.*, 2017).

Code for the analysis is archived at https://github.com/susjoh/Deer_Recombination_GWAS.

243 **Results**

244 **Variation and heritability in autosomal crossover count.**

Autosomal crossover count (ACC) was significantly higher in females than in males, where fe-

males had 4.32 ± 0.41 more crossovers per gamete (Z = 10.57, P_{Wald} <0.001; Figure 1); there

was no effect of FID age or inbreeding on ACC (P >0.05, Table S1). Females had significantly

higher phenotypic variance in ACC than males (V_P = 32.02 and 15.33, respectively; Table 1).

ACC was heritable in both sexes combined (h^2 = 0.13, SE = 0.05, P = 0.002) and within females

only (h^2 = 0.11, SE = 0.06, P = 0.033), but was not heritable in males alone (P > 0.05; Table 1).

The remaining phenotypic variance was explained by the residual error term, and there was no

variance explained by the permanent environment effect, birth year, year of gamete transmis-

sion, or parental identities of the FID in any model (P >0.05). Bivariate models of ACC between

the sexes indicated that the genetic correlation (r_a) between males and females was 0.346, but

was not significantly different from zero or one (P_{LRT} >0.05). This may be due to the relatively

small sample size of this dataset resulting in a large standard error around the r_A estimate, or

the fact that ACC was not heritable in males. Sampling of 482 measures from each sex showed

no difference in heritability estimates between the sexes, indicating reduced power to quantify

heritable variation in the smaller male dataset (t = 0.242, P = 0.810, Figure S3).

Table 1: Data set information and animal model results for autosomal crossover count (ACC). Numbers in parentheses are the standard error, except for *Mean*, which is the standard deviation. N_{OBS}, N_{FID} and N_{xovers} are the number of ACC measures, the number of focal individuals (FIDS) and the total number of crossovers in the dataset. The mean ACC was calculated from the raw data. V_P and V_A are the phenotypic variance and additive genetic variance, respectively. h^2, pe^2 and e^2 are the narrow-sense heritability, the permanent environment effect, and the residual effect, respectively; all are calculated as the proportion of V_P that they explain. The additive genetic components were modelled using genomic relatedness matrices. $P(h^2)$ is the significance of the V_A term in the model as determined using a likelihood ratio test.

Analysis	N_{OBS}	N_{FID}	Mean	N_{xovers}	V_P	V_A	h^2	pe^2	e^2	$P(h^2)$
Both	1341	337	25.03 (5.49)	34911	26.42 (1.17)	3.46 (1.34)	0.13 (0.05)	0.05 (0.04)	0.82 (0.03)	0.002
Females	859	256	26.62 (5.62)	24025	32.02 (1.67)	3.46 (1.87)	0.11 (0.06)	0.05 (0.05)	0.84 (0.04)	0.033
Males	482	81	22.21 (3.88)	10886	15.33 (1.09)	1.03 (1.66)	0.07 (0.11)	0.06 (0.1)	0.87 (0.05)	0.554

Figure 1: Distribution of ACCs in the raw data for females and males.

260 Genetic architecture of autosomal crossover count.

Genome-wide association study. No SNPs were significantly associated with ACC at the

genome-wide level (Figure 2, Tables 2 and S2). The most highly associated SNP in both sexes

was *cela1_red_10_26005249* on deer linkage group 12 (CEL12), corresponding to position

26,005,249 on cattle chromosome 10 (BTA10). This marker was also the most highly asso-

ciated SNP when considering recombination in *trans*, indicating that this region affects ACC

across the genome (Table S2). The observed association was primarily driven by female ACC

(Table 2, Figure 2). In females, the most highly associated SNP was *cela1_red_10_25661750*

on CEL12, corresponding to posiiton 25,661,750 on BTA10. For both SNPs, sampling of 482

measures from each sex showed that the observed associations were significantly higher in

females than in males when considering the same sample size (*cela1_red_10_25661750*: t

= 18.60, P < 0.001; *cela1_red_10_26005249*: t = 4.89, P < 0.001; Figure S4). Based on its

position relative to the cattle genome, *cela1_red_10_26005249* was ∼600bp upstream of an

olfactory receptor *OR5AU1* and ∼24kb downstream from a gene of unknown function (ENSB-

TAG00000011396). There were four candidate genes within 1Mb of both loci, including *TOX4*,

CHD8, *SUPT16H* and *CCNB1IP1* (Figure 4; see Discussion). Similar results were obtained

when considering recombination rate on all chromosomes excluding that on which the SNP

occurred, indicating that all associations affect recombination rate variation in *trans* across the

genome (Table S2).

Figure 2: Manhattan plot of genome-wide association of autosomal crossover count (ACC) for (A) all deer, (B) females only and (C) males only. The dashed line is the genome-wide significance threshold equivalent to P <0.05. The left-hand plots show association relative to the estimated genomic positions on deer linkage groups from JOHNSTON *et al.* (2017). Points have been colour coded by chromosome. The right-hand plots show the distribution of observed $-log_{10}P$ values against those under the null expectation. Association statistics have been corrected for the genomic control inflation factor λ. Underlying data are provided in Table S2 and sample sizes are given in Table 1.

Table 2: The top five most significant hits from a genome-wide association study of ACC in (A) Both sexes, (B) Females only and (C) Males only. No SNPs reached the genome-wide significance of P = 1.42×10^{-06}. The SNP locus names indicate the position of the SNPs relative to the cattle genome assembly vBTA_vUMD_3.1 (indicated by *Chromosome_Position*). Linkage groups and map positions (in centiMorgans, cM) are from JOHNSTON *et al.* (2017). A and B are the reference alleles. Effect B is the estimated effect and standard error of the B allele as estimated in RepeatABEL (RÖNNEGÅRD *et al.*, 2016). P-values have been corrected for the genomic inflation parameter λ. Full results are available in Table S2.

Sex	SNP Locus	Deer Linkage Group	Map Position (cM)	A	B	Effect B	(SE)	χ_1^2	P	MAF
A. Both	cela1_red_10_26005249	12	36.4	G	A	1.53	0.28	22.73	1.87e-06	0.33
	cela1_red_8_100681301	16	43.5	A	G	6.42	1.19	21.87	2.91e-06	0.02
	cela1_red_10_25661750	12	35.6	A	G	2.18	0.42	20.6	5.67e-06	0.1
	cela1_red_1_35423049	31	46.2	A	G	1.4	0.29	17.31	3.18e-05	0.25
	cela1_red_10_21372438	12	34.5	A	G	1.22	0.26	16.2	5.69e-05	0.42
B. Females	cela1_red_10_25661750	12	35.6	A	G	2.81	0.56	21.07	4.44e-06	0.1
	cela1_red_10_26005249	12	36.4	G	A	1.58	0.35	16.84	4.07e-05	0.33
	cela1_red_8_100681301	16	43.5	A	G	6.25	1.4	16.72	4.34e-05	0.02
	cela1_red_1_35423049	31	46.2	A	G	1.62	0.37	16.37	5.22e-05	0.25
	cela1_red_11_91378678	11	86.5	A	G	13.61	3.22	15.03	1.06e-04	0.02
C. Males	cela1_red_10_49732924	12	52.6	G	A	-2.9	0.66	18.25	1.94e-05	0.14
	cela1_red_1_128593904	19	13.5	G	A	-1.99	0.46	17.32	3.15e-05	0.18
	cela1_red_15_6941417	1	8.9	A	G	-1.93	0.46	16.74	4.28e-05	0.21
	cela1_red_2_101879999	8	35.2	G	A	1.51	0.39	14.28	1.58e-04	0.44
	cela1_red_15_7417500	1	9.2	A	C	-1.93	0.5	13.9	1.93e-04	0.17

Regional heritability analysis. The genome-wide regional heritability analysis of ACC showed a significant association in both sexes and in females only with a ~2.94Mb region on CEL12 (Figure 3, Table 3). The most highly associated window (~1.36 Mb) within this region contained 42 genes, including REC8 meiotic recombination protein (*REC8*; 20,810,610 - 20,817,662 bp on BTA10). Detailed examination of this region in sliding windows of 6, 10 and 20 SNPs found the highest association at a 10 SNP window of ~463kb containing 36 genes, including *REC8* (Table 3). This region explained all heritable variation in recombination rate, with regional heritability estimates of 0.143 (SE = 0.053) and 0.146 (SE = 0.045) for all deer and females only, respectively. The sex-specific effect was supported by sampling of 482 measures, where females had consistently higher associations than in males (t = 19.03, P < 0.001, Figure S5). The total significant region after detailed examination was ~3.01Mb wide, flanked by SNPs *cela1_red_10_18871213* and *cela1_red_10_21878407* (Figure 4 & Table S4) and containing ~87 genes. This wider region contained the protein coding region for ring finger protein 212B

(*RNF212B*; 21,466,337 - 21,494,696 bp on BTA10), a homologue of *RNF212*, which has been

directly implicated in synapsis and crossing-over during meiosis in mice (REYNOLDS *et al.*,

2013). Genetic variants at both *RNF212B* and *RNF212* have been associated with recombina-

tion rate variation in humans, cattle and sheep (KONG *et al.*, 2008; MA *et al.*, 2015; JOHNSTON

et al., 2016; PETIT *et al.*, 2017). Whilst this region was close to the most highly associated

SNPs from the genome-wide association study, there was no overlap between the two anal-

yses, with the mostly highly associated regions separated by an estimated ∼5.5Mb (Figure

4). The mean r^2 LD between the top regional heritability window and the top GWAS SNPs

was 0.258 for *cela1_red_10_25661750* and 0.276 for *cela1_red_10_26005249*, with the top r^2

of 0.665 observed between the SNPs *cela1_red_10_21807996* and *cela1_red_10_26005249*

(Figure 4).

A second region on linkage group 32 almost reached genome-wide significance in the regional

heritability analysis, corresponding to the region ∼38.7 - 41.3Mb on cattle chromosome 27.

This region contained the locus topoisomerase (DNA) II beta (*TOP2B*); inhibitors of this gene

lead to defects in chromosome segregation and heterochromatin condensation during meiosis

I in mice, *Drosophila melanogaster* and *Caenorhabditis elegans* (LI *et al.* 2013; GÓMEZ *et al.*

2014; HUGHES and HAWLEY 2014; JARAMILLO-LAMBERT *et al.* 2016; Figure 3, Tables 3 and

S3). Full results for the regional heritability analyses are provided in Tables S3 and S4.

Table 3: The most significant hits from a regional heritability analysis of ACC in (A) Both sexes, (B) Females only and (C) Males only. Sliding windows were 20 SNPs wide with an overlap of 10 SNPs. Lines in italics are the most highly associated regions from detailed examination of significant regions - in each case these are for 10 SNP windows. The χ^2 and P values are for likelihood ratio test comparisons between models with and without a genomic relatedness matrix for that window; values in bold type are significant the the genome-wide level. The SNP locus names indicate the position of the SNPs relative to the cattle genome assembly vBTA_vUMD_3.1 (indicated by *Chromosome_Position*). Full results are available in Tables S3 & S4.

Sex	Deer Linkage Group	χ^2_1	P	First SNP	Last SNP	Region h^2	SE
A. Both	*12*	*32.30*	*1.32e-08*	*cela1_red_10_20476277*	*cela1_red_10_20939342*	*0.143*	*0.053*
	12	28.62	8.81e-08	cela1_red_10_19617695	cela1_red_10_20977030	0.080	0.043
	12	25.11	5.41e-07	cela1_red_10_20519507	cela1_red_10_21807996	0.080	0.045
	12	22.91	1.70e-06	cela1_red_10_18871213	cela1_red_10_20476277	0.105	0.055
	32	16.55	4.73e-05	cela1_red_27_38731584	cela1_red_27_40264086	0.056	0.034
	32	15.76	7.21e-05	cela1_red_27_39821973	cela1_red_27_41274975	0.071	0.045
B. Females	*12*	*28.14*	*1.13e-07*	*cela1_red_10_20476277*	*cela1_red_10_20939342*	*0.146*	*0.045*
	12	24.34	8.06e-07	cela1_red_10_19617695	cela1_red_10_20977030	0.089	0.048
	12	23.5	1.25e-06	cela1_red_10_20519507	cela1_red_10_21807996	0.102	0.056
	12	20.03	7.61e-06	cela1_red_10_18871213	cela1_red_10_20476277	0.133	0.068
	12	13.72	2.12e-04	cela1_red_10_21000545	cela1_red_10_22450693	0.089	0.054
	12	12.32	4.49e-04	cela1_red_10_21878407	cela1_red_10_26041475	0.177	0.087
C. Males	5	14.07	1.76e-04	cela1_red_19_15289588	cela1_red_19_16108226	0.133	0.052
	5	12.61	3.84e-04	cela1_red_19_15753501	cela1_red_19_16923111	0.137	0.058
	20	8.77	3.06e-03	cela1_red_3_110763634	cela1_red_3_112123206	0.142	0.085
	32	8.51	3.52e-03	cela1_red_27_38731584	cela1_red_27_40264086	0.119	0.076
	1	8.21	4.17e-03	cela1_red_15_6354196	cela1_red_15_7482634	0.123	0.056

Figure 3: Regional heritability plot of association of autosomal crossover count for (A) all deer, (B) females only and (C) males only. Each point represents a sliding window of 20 SNPs with an overlap of 10 SNPs. The dashed line is the genome-wide significance threshold equivalent to P <0.05 as calculated using Bonferroni. Lines have been colour coded by chromosome. Underlying data are provided in Table 3.

Figure 4: Detailed figure of genes, association statistics and linkage disequilibrium patterns at the most highly associated region on on CEL12 (homologous to BTA10) for all deer of both sexes. All X-axis positions are given relative to the cattle genome vBTA_vUMD_3.1. The top panel shows protein coding regions, with annotation for candidate loci. The central panel shows the results for the regional heritability analysis (where lines represents a sliding windows of 6, 10 and 20 SNPs with an overlap of n-1 SNPs) and the genome-wide association study (where points indicate single SNP associations). The dashed lines are the genome-wide significance thresholds (green = regional heritability, black = genome-wide association). The checked shaded area shows the position of the T cell receptor alpha/delta locus (see Discussion). Underlying data are provided in Table S4. The lower panel shows linkage disequilibrium between each loci using allelic correlations (r^2).

Effect size estimation. At most highly associated GWAS SNP, *cela1_red_10_26005249*, car-

rying one or two copies of the G allele conferred 3.3 to 3.9 fewer crossovers per gamete in

females (Wald P < 0.001) and 1.8 - 2.8 fewer crossovers per gamete in males (P = 0.009; Table

4). The most highly associated SNP in females, *cela1_red_10_25661750*, had a significant

effect on ACC in females (P < 0.001) but not in males (P > 0.05; Table 4). This locus conferred

2.03 more crossovers in A/G females and 13.68 more in G/G females; however, the latter cat-

egory contained 7 unique measures in only two individuals, and so this estimate is likely to be

subject to large sampling error.

A total of 17 haplotypes in the 10 SNP region spanning *cela1_red_10_20476277* and

cela1_red_10_20939342 had more than ten copies in unique individual females (Table S5). Of

these, two haplotypes, AGGAGAGAAG and AGAGAAGAGA, had a significant effect on ACC

relative to all other haplotypes (Tables 4 and S5, Figure S6). Haplotype AGGAGAGAAG in-

creased female ACC by 2.4 crossovers per gamete in heterozygotes (P < 0.001); homozygotes

for the haplotype were rare (13 measures in 4 individuals) and so the large effect size estimate

was again likely to be subject to large sampling effects (Table 4). The haplotype AGAGAAGAGA

reduced female ACC by 2.2 crossovers per gamete in heterozygous individuals (P < 0.05; Table

4). The r^2 LD between haplotype AGGAGAGAAG and the two most highly associated GWAS

SNPs was 0.464 and 0.885 for *cela1_red_10_26005249* and *cela1_red_10_25661750*, respec-

tively; for haplotype AGAGAAGAGA, it was 0.229 and 0.036 for *cela1_red_10_26005249* and

cela1_red_10_25661750, respectively.

Table 4: Effect sizes for the most highly associated GWAS SNPs and for the AGGAGAGAAG haplotype at the most highly associated regional heritability region. Models were run for each sex separately and included a pedigree relatedness as a random effect. Count and ID Count indicate the number of ACC measures and the number of unique individuals for each genotype, respectively. Wald.P indicates the P-value for a Wald test of genotype as a fixed effect.

Locus	Sex	Genotype	Count	ID Count	Solution	S.E.	Z Ratio	Wald.P
cela1_red_10_26005249	Female	A/A (Intercept)	98	28	29.575	0.732	40.43	3.43e-06
		A/G	388	114	-3.269	0.733	-4.46	
		G/G	377	116	-3.888	0.786	-4.944	
	Male	A/A (Intercept)	27	6	24.405	0.964	25.327	8.90e-03
		A/G	248	40	-1.813	0.993	-1.826	
		G/G	207	35	-2.863	1.025	-2.793	
cela1_red_10_25661750	Female	A/A (Intercept)	688	208	25.979	0.36	72.114	6.34e-10
		A/G	168	48	2.026	0.56	3.619	
		G/G	7	2	13.684	2.386	5.736	
	Male	A/A (Intercept)	411	65	22.13	0.367	60.345	0.399
		A/G	57	14	0.934	0.69	1.353	
		G/G	14	2	0.345	1.512	0.228	
Haplotype AGGAGAGAAG	Female	A/A (Intercept)	690	208	25.905	0.364	71.125	7.29e-09
		A/B	160	46	2.387	0.573	4.166	
		B/B	13	4	8.701	1.73	5.029	
	Male	A/A (Intercept)	406	66	22.017	0.36	61.153	0.037
		A/B	62	13	1.72	0.669	2.571	
		B/B	14	2	0.506	1.492	0.339	
Haplotype AGAGAAGAGA	Female	A/A (Intercept)	795	242	26.591	0.451	58.928	0.026
		A/B	68	16	-2.244	1.005	-2.233	
	Male	A/A (Intercept)	481	80	22.279	0.351	63.403	0.775
		A/B	1	1	-1.122	3.925	-0.286	

330 **Discussion**

In this study, we have shown that autosomal crossover count (ACC) is 1.2× higher in red deer-

females than in males, with females exhibiting higher phenotypic and additive genetic variance

for this trait; ACC was not significantly heritable in males. Almost all genetic variation in females

was explained by a ~7Mb region on deer linkage group 12. This region contained several

candidate genes, including *RNF212B* and *REC8*, which have previously been implicated in re-

combination rate variation in other mammal species, including humans, mice, cattle and sheep

(KONG *et al.*, 2008; REYNOLDS *et al.*, 2013; MA *et al.*, 2015; JOHNSTON *et al.*, 2016; PETIT *et al.*,

2017). Here, we discuss in detail the genetic architecture of individual recombination rate, can-

didate genes underlying heritable variation, sexual-dimorphism in this trait and its architecture,

and the conclusions and implications of our findings for other studies of recombination in the

wild.

The genetic architecture of individual recombination rate. Using complementary trait mapping approaches, we identified a ~7Mb region on deer linkage group 12 (homologous to cattle chromosome 10) associated with ACC. The most highly associated GWAS region occurred between ~25.6 and 26Mb (relative to the cattle genome position), although this association was not significant at the genome-wide level. The most highly associated regional heritability region occurred between ~20.5 and 20.9MB, around 5Mb away from the top GWAS hits (Figure 4); association at this region was significant at the genome-wide level and explained almost all of the heritable variation in ACC in both sexes and in females only. Most variation in mean ACC was attributed to two haplotypes within this region (Tables 4 and S5; Figure S6).

At present, it is not clear why the results of the two analyses occur in close vicinity, yet do not overlap. Assuming homology with humans, cattle, sheep and mice (Ensembl release 91, ZERBINO *et al.* 2018), the two regions are separated by the highly repetitive T-cell receptor alpha/delta variable (TRAV/DV) locus, which may contain up to 400 TRAV/DV genes in cattle (REININK and VAN RHIJN 2009; Figure 4). This region is of an unknown size in deer; relative to the cattle genome, these regions are separated by 4.72Mb, but the deer linkage map distance is estimated as 1.86 centiMorgans (cM). The sex-averaged genome-wide recombination rate in deer is ~1.04cM/Mb, suggesting this genomic region may be shorter in deer (JOHNSTON *et al.*, 2017) and that these two regions are in closer vicnity. This is supported by both the linkage map distance and patterns of linkage disequilibrium between the associated loci, particularly at the associated haplotypes (see Results & Figure 4). In addition, the small sample size used in the current study may result in increased sensitivity to sampling effects and bias in the estimation of the relative contribution of SNPs to the trait mean (GWAS) or variance (Regional heritability). Further investigation with higher samples sizes, whole genome sequencing approaches and improved genome assembly may allow more accurate determination of the most likely candidate genes and potential causal mutations (coding or regulatory) within this species.

Candidate genes for recombination rate variation.
Regional heritability analysis. The most highly associated region in the regional heritability analysis contained the gene *REC8*, the protein of which is required for the separation of sister chromatids during meiosis (PARISI *et al.*, 1999). It also contained *RNF212B*, a paralogue of *RNF212*. *RNF212* has been associated with recombination rate variation in humans, cattle and sheep (KONG *et al.*, 2008; SANDOR *et al.*, 2012; JOHNSTON *et al.*, 2016; PETIT *et al.*,

2017); the *REC8/RNF212B* region is associated with recombination rate in cattle, and has

a large effect size on ACC phenotype than *RNF212* in this species (SANDOR *et al.*, 2012;

MA *et al.*, 2015). A second region on deer linkage group 32 almost reached genome-wide

significance (Figure 3, Tables 3 and S3). This region was relatively gene-poor, but contained

∼6 genes, including the candidate topoisomerase (DNA) II beta (*TOP2B*): inhibitors of this gene

lead to defects in chromosome segregation and heterochromatin condensation during meiosis

I in mice, *Drosophila melanogaster* and *Caenorhabditis elegans* (LI *et al.*, 2013; GÓMEZ *et al.*,

2014; HUGHES and HAWLEY, 2014; JARAMILLO-LAMBERT *et al.*, 2016). No association was

observed at the region homologous to *RNF212* (predicted to be at position ∼109.2Mb on cattle

chromosome 6, corresponding to ∼57.576cM on deer linkage group 6) for the GWAS or regional

heritability analysis.

Genome-wide association study (GWAS). Examination of annotated regions within 500kb

of either side of the most significant GWAS SNPs identified three genes, TOX High Mobility

Group Box Family Member 4 (*TOX4*), Chromodomain Helicase DNA Binding Protein 8 (*CHD8*)

and SPT16 Homologue Facilitates Chromatin Remodelling Subunit (*SUPT16H*). These genes

are involved in with chromatin binding and structure (*SP16H, TOX4*), histone binding (*CHD8,*

SUPT16H), nucleosome organisation (*SP16H*) and cell cycle transition (*TOX4*). One of these

genes, *SUPT16H*, interacts with NIMA related kinase 9 (*NEK9*), which is involved with mei-

otic spindle organisation, chromosome alignment and cell cycle progression in mice (YANG

et al., 2012) and is a strong candidate locus for crossover interference in cattle (WANG *et al.*,

2016). The SNP *cela1_red_10_26005249* was ∼825kb from Cyclin B1 Interacting Protein 1

(*CCNB1IP1*), also known as Human Enhancer Of Invasion 10 (*HEI10*), which interacts with

RNF212 to allow recombination to progress into crossing-over in mice (QIAO *et al.*, 2014) and

Arabidopsis (CHELYSHEVA *et al.*, 2012); this locus is also associated with recombination rate

variation in humans (KONG *et al.*, 2014).

Sexual dimorphism in genetic architecture of recombination rate. The results of this anal-

ysis suggest that there is sexual dimorphism in the genetic architecture of recombination rate

variation in deer. Higher female recombination rates are typical in mammals (BRANDVAIN and

COOP, 2012) and in this system is driven by higher female recombination rates near centromeric

regions (JOHNSTON *et al.*, 2017). Male ACC was not significantly heritable, although we could

not rule out that this was a consequence of the smaller sample size relative to females (Figure

S3). No regions of the genome were significantly associated with male ACC in the regional her-

itability and GWAS analyses, but sampling did indicate that differences observed between male

and female genomic associations were statistically significant (Figures S4 & S5). Investigation

of genetic correlations between males and females was inconclusive, as the r_A of ACC was not

significantly different from 0 or 1. The observed sex differences are consistent with previous

studies of the genetic architecture of ACC in mammals, where a sexually-dimorphic architec-

ture has been observed at the paralogous *RNF212* region in humans and sheep (KONG *et al.*,

2014; JOHNSTON *et al.*, 2016). Nevertheless, some observed associations were stronger when

considering both male and female deer in the same analysis, for example at the most highly

associated GWAS SNP, and the amplified signal for the regional heritability analysis on linkage

group 33 (Figures 2 & 3), suggesting that there may be some degree of shared architecture

within these regions.

Conclusions and implications for studies of recombination in the wild. We have shown

that recombination rate is heritable in female red deer, and that it has a sexually dimorphic

genetic architecture. Genomic regions associated with recombination rate in red deer are as-

sociated with this trait in other mammal species, supporting the idea that recombination rate

variation has a conserved genetic architecture across distantly related taxa. A key motivation

for this study is to compare how recombination rate and its genetic architecture is similar or

different to that of model species that have experienced strong selection in their recent history,

such as humans, cattle, mice and sheep. The heritability of recombination rate in deer was

lower than that observed in other mammal systems (DUMONT *et al.*, 2009; KONG *et al.*, 2014;

MA *et al.*, 2015; JOHNSTON *et al.*, 2016), with no observed heritable variation present in male

deer. Whilst we were able to test their effects, we found no contribution of contribution of indi-

vidual and common environmental effects on recombination rate (i.e. age, year of birth, year

of gamete transmission); indeed, most phenotypic variance in recombination was attributed to

residual effects. This suggests that despite some underlying genetic variation, recombination

rate is mostly driven by stochastic effects, or otherwise unmeasured effects.

This represents one of the smallest datasets in which recombination rate has been investigated,

and so it may be that the observed effects are underestimated due to the small sample size,

sampling effects, or perhaps that other genetic variants present in this species do not segre-

gate in the Rùm deer population. Nevertheless, identification of clear candidate genes and their

effects on phenotype represents a valuable contribution to understanding the genetic architec-
ture of recombination more broadly. Ultimately, our findings allow future investigation of the
fitness consequences of variation in recombination rate and the relationship between identified
variants and individual life-history variation, to address questions on the maintenance of ge-
netic variation for recombination rates, and the relative roles of selection, sexually antagonistic
effects and stochastic processes in contemporary natural populations.

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452 Author Contributions

453 S.E.J and J.M.P. conceived the study. J.M.P organised the collection of samples. J.H. con-
454 ducted DNA sample extraction and genotyping and constructed the pedigree. S.E.J. analysed
455 the data and wrote the paper. All authors contributed to revisions.

456 References

ARBEITHUBER, B., A. J. BETANCOURT, T. EBNER, and I. TIEMANN-BOEGE, 2015 Crossovers
are associated with mutation and biased gene conversion at recombination hotspots. Proc.

Natl. Acad. Sci. U.S.A. **112**: 2109–2114.

AULCHENKO, Y. S., S. RIPKE, A. ISAACS, and C. M. VAN DUIJN, 2007 GenABEL: an R library

for genome-wide association analysis. Bioinformatics **23**: 1294–1296.

BAKER, Z., M. SCHUMER, Y. HABA, L. BASHKIROVA, C. HOLLAND, *et al.*, 2017 Repeated losses

of PRDM9-directed recombination despite the conservation of PRDM9 across vertebrates.

eLife **6**: e24133.

BARTON, N. H., and B. CHARLESWORTH, 1998 Why sex and recombination? Science **74**:

187–195.

BATTAGIN, M., G. GORJANC, A.-M. FAUX, S. E. JOHNSTON, and J. M. HICKEY, 2016 Effect of

manipulating recombination rates on response to selection in livestock breeding programs.

Genet. Select. Evol. **48**: 44.

BAUDAT, F., J. BUARD, C. GREY, A. FLEDEL-ALON, C. OBER, *et al.*, 2010 PRDM9 is a major

determinant of meiotic recombination hotspots in humans and mice. Science **327**: 836–840.

BÉRÉNOS, C., P. A. ELLIS, J. G. PILKINGTON, S. H. LEE, J. GRATTEN, *et al.*, 2015 Hetero-

geneity of genetic architecture of body size traits in a free-living population. Mol. Ecol. **24**:

1810–1830.

BRANDVAIN, Y., and G. COOP, 2012 Scrambling eggs: meiotic drive and the evolution of female

recombination rates. Genetics **190**: 709–23.

BRAUNING, R., P. J. FISHER, A. F. MCCULLOCH, R. J. SMITHIES, J. F. WARD, *et al.*, 2015

Utilization of high throughput genome sequencing technology for large scale single nucleotide

polymorphism discovery in red deer and canadian elk. bioRxiv **doi:10.1101/027318**.

BURT, A., 2000 Sex, Recombination, and the Efficacy of Selection - was Weismann Right?

Evolution **54**: 337–351.

BUTLER, D. G., B. R. CULLIS, A. R. GILMOUR, and B. J. GOGEL, 2009 Mixed Models for S

language Environments: ASReml-R reference manual.

CHELYSHEVA, L., D. VEZON, A. CHAMBON, G. GENDROT, L. PEREIRA, *et al.*, 2012 The Ara-

bidopsis HEI10 Is a New ZMM Protein Related to Zip3. PLoS Genet. **8**: e1002799.

CLUTTON-BROCK, T., F. GUINNESS, and S. ALBON, 1982 *Red Deer. Behaviour and Ecology of*

Two Sexes.. University of Chicago Press.

DELANEAU, O., J.-F. ZAGURY, and J. MARCHINI, 2012 Improved whole-chromosome phasing
for disease and population genetic studies. Nat. Methods **10**: 5–6.

DEVLIN, A. B., K. ROEDER, and B. DEVLIN, 1999 Genomic Control for Association. Biometrics
55: 997–1004.

DUMONT, B. L., K. W. BROMAN, and B. A. PAYSEUR, 2009 Variation in genomic recombination
rates among heterogeneous stock mice. Genetics **182**: 1345–9.

DUMONT, B. L., M. A. WHITE, B. STEFFY, T. WILTSHIRE, and B. A. PAYSEUR, 2011 Exten-
sive recombination rate variation in the house mouse species complex inferred from genetic
linkage maps. Genome Res. **21**: 114–125.

FELSENSTEIN, J., 1974 The evolutionary advantage of recombination. Genetics **78**: 737–756.

GÓMEZ, R., A. VIERA, I. BERENGUER, E. LLANO, A. M. PENDÁS, *et al.*, 2014 Cohesin removal
precedes topoisomerase IIα-dependent decatenation at centromeres in male mammalian
meiosis II. Chromosoma **123**: 129–146.

GONEN, S., M. BATTAGIN, S. JOHNSTON, G. GORJANC, and J. HICKEY, 2017 The potential of
shifting recombination hotspots to increase genetic gain in livestock breeding. Genet. Select.
Evol. **49**: 55.

GREEN, P., K. FALLS, and S. CROOKS, 1990 *Documentation for CRIMAP, version 2.4.* Wash-
ington University School of Medicine.

HASSOLD, T., and P. HUNT, 2001 To err (meiotically) is human: the genesis of human aneu-
ploidy. Nat. Rev. Genet. **2**: 280–291.

HENDERSON, C. R., 1975 Best linear unbiased estimation and prediction under a selection
model. Biometrics **31**: 423–447.

HILL, W. G., and A. ROBERTSON, 1966 The effect of linkage on limits to artificial selection.
Genet. Res. **8**: 269–294.

HUGHES, S. E., and R. S. HAWLEY, 2014 Topoisomerase II Is Required for the Proper Sepa-
ration of Heterochromatic Regions during Drosophila melanogaster Female Meiosis. PLoS
Genet. **10**: e1004650.

HUISMAN, J., 2017 Pedigree reconstruction from SNP data: parentage assignment, sibship
clustering and beyond. Mol. Ecol. Resour. **17**: 1009–1024.

HUISMAN, J., L. E. B. KRUUK, P. A. ELLIS, T. CLUTTON-BROCK, and J. M. PEMBERTON, 2016

Inbreeding depression across the lifespan in a wild mammal population. Proc. Natl. Acad.

Sci. U.S.A. **113**: 3585–3590.

INOUE, K., and J. R. LUPSKI, 2002 Molecular mechanisms for genomic disorders. Ann. Rev.

Genomics Hum. Genet. **3**: 199–242.

JARAMILLO-LAMBERT, A., A. S. FABRITIUS, T. J. HANSEN, H. E. SMITH, and A. GOLDEN, 2016

The Identification of a Novel Mutant Allele of topoisomerase II in Caenorhabditis elegans

Reveals a Unique Role in Chromosome Segregation During Spermatogenesis. Genetics

204: 1407–1422.

JOHNSTON, S. E., C. BÉRÉNOS, J. SLATE, and J. M. PEMBERTON, 2016 Conserved genetic

architecture underlying individual recombination rate variation in a wild population of soay

sheep (ovis aries). Genetics **203**: 583–598.

JOHNSTON, S. E., J. HUISMAN, P. A. ELLIS, and J. M. PEMBERTON, 2017 A High Density Link-

age Map Reveals Sexual Dimorphism in Recombination Landscapes in Red Deer (*Cervus*

elaphus). G3 (Bethesda) **7**: 2859–2870.

KAWAKAMI, T., C. F. MUGAL, A. SUH, A. NATER, R. BURRI, *et al.*, 2017 Whole-genome patterns

of linkage disequilibrium across flycatcher populations clarify the causes and consequences

of fine-scale recombination rate variation in birds. Mol. Ecol. **26**: 4158–4172.

KONG, A., J. BARNARD, D. F. GUDBJARTSSON, G. THORLEIFSSON, G. JONSDOTTIR, *et al.*,

2004 Recombination rate and reproductive success in humans. Nat. Genet. **36**: 1203–1206.

KONG, A., G. THORLEIFSSON, M. L. FRIGGE, G. MASSON, D. F. GUDBJARTSSON, *et al.*, 2014

Common and low-frequency variants associated with genome-wide recombination rate. Nat.

Genet. **46**: 11–6.

KONG, A., G. THORLEIFSSON, H. STEFANSSON, G. MASSON, A. HELGASON, *et al.*, 2008

Sequence variants in the RNF212 gene associate with genome-wide recombination rate.

Science **319**: 1398–1401.

LI, X.-M., C. YU, Z.-W. WANG, Y.-L. ZHANG, X.-M. LIU, *et al.*, 2013 DNA Topoisomerase

II Is Dispensable for Oocyte Meiotic Resumption but Is Essential for Meiotic Chromosome

Condensation and Separation in Mice. Biol. Reprod. **89**: 118, 1–11.

MA, L., J. R. O'CONNELL, P. M. VANRADEN, B. SHEN, A. PADHI, *et al.*, 2015 Cattle Sex-
Specific Recombination and Genetic Control from a Large Pedigree Analysis. PLoS Genet.
11: e1005387.

MOSKVINA, V., and K. M. SCHMIDT, 2008 On multiple-testing correction in genome-wide asso-
ciation studies. Genet. Epidemiol. **32**: 567–573.

MULLER, H., 1964 The relation of recombination to mutational advance. Mutat. Res. **1**: 2–9.

MUÑOZ-FUENTES, V., M. MARCET-ORTEGA, G. ALKORTA-ARANBURU, C. LINDE FORSBERG,
J. M. MORRELL, *et al.*, 2015 Strong Artificial Selection in Domestic Mammals Did Not Result
in an Increased Recombination Rate. Mol. Biol. Evol. **32**: 510–523.

MYERS, S., L. BOTTOLO, C. FREEMAN, G. MCVEAN, and P. DONNELLY, 2005 A fine-scale map
of recombination rates and hotspots across the human genome. Science **310**: 321–4.

NAGAMINE, Y., R. PONG-WONG, P. NAVARRO, V. VITART, C. HAYWARD, *et al.*, 2012 Localis-
ing Loci underlying Complex Trait Variation Using Regional Genomic Relationship Mapping.
PLoS ONE **7**: e46501.

O'CONNELL, J., D. GURDASANI, O. DELANEAU, N. PIRASTU, S. ULIVI, *et al.*, 2014 A General
Approach for Haplotype Phasing across the Full Spectrum of Relatedness. PLoS Genet. **10**:
e1004234.

OTTO, S. P., and N. H. BARTON, 2001 Selection for recombination in small populations. Evolu-
tion **55**: 1921–1931.

OTTO, S. P., and T. LENORMAND, 2002 Resolving the paradox of sex and recombination. Nat.
Rev. Genet. **3**: 252–261.

PARISI, S., M. J. MCKAY, M. MOLNAR, M. A. THOMPSON, P. J. VAN DER SPEK, *et al.*, 1999
Rec8p, a Meiotic Recombination and Sister Chromatid Cohesion Phosphoprotein of the
Rad21p Family Conserved from Fission Yeast to Humans. Mol. Cell. Biol. **19**: 3515–3528.

PETIT, M., J.-M. ASTRUC, J. SARRY, L. DROUILHET, S. FABRE, *et al.*, 2017 Variation in Recom-
bination Rate and Its Genetic Determinism in Sheep Populations. Genetics **207**: 767–784.

QIAO, H., H. B. D. P. RAO, Y. YANG, J. H. FONG, J. M. CLOUTIER, *et al.*, 2014 Antagonistic
roles of ubiquitin ligase HEI10 and SUMO ligase RNF212 regulate meiotic recombination.
Nat. Genet. **46**: 194–199.

REININK, P., and I. VAN RHIJN, 2009 The bovine T cell receptor alpha/delta locus contains over
400 V genes and encodes V genes without CDR2. Immunogenetics **61**: 541–549.

REYNOLDS, A., H. QIAO, Y. YANG, J. K. CHEN, N. JACKSON, *et al.*, 2013 RNF212 is a dosage-
sensitive regulator of crossing-over during mammalian meiosis. Nat. Genet. **45**: 269–78.

RÖNNEGÅRD, L., S. E. MCFARLANE, A. HUSBY, T. KAWAKAMI, H. ELLEGREN, *et al.*, 2016 In-
creasing the power of genome wide association studies in natural populations using repeated
measures: evaluation and implementation. Methods Ecol. Evol. **7**: 792–799.

SANDOR, C., W. LI, W. COPPIETERS, T. DRUET, C. CHARLIER, *et al.*, 2012 Genetic variants in
REC8, *RNF212*, and *PRDM9* influence male recombination in cattle. PLoS Genet. **8**: 1–13.

SHIN, J.-H., S. BLAY, B. MCNENEY, and J. GRAHAM, 2006 Ldheatmap: An r function for
graphical display of pairwise linkage disequilibria between single nucleotide polymorphisms.
J. Stat. Soft. **16**: Code Snippet 3.

STAPLEY, J., P. G. D. FEULNER, S. E. JOHNSTON, A. W. SANTURE, and C. M. SMADJA,
2017 Variation in recombination frequency and distribution across Eukaryotes: patterns and
processes. Philos. Trans. R. Soc. Lond., B, Biol. Sci. **372**: 20160455.

THEODOSIOU, L., W. O. MCMILLAN, and O. PUEBLA, 2016 Recombination in the eggs and
sperm in a simultaneously hermaphroditic vertebrate. Philos. Trans. R. Soc. Lond., B, Biol.
Sci. **283**: 20161821.

WANG, Z., B. SHEN, J. JIANG, J. LI, and L. MA, 2016 Effect of sex, age and genetics on
crossover interference in cattle. Sci. Rep. **6**: 37698.

YANG, J., S. H. LEE, M. E. GODDARD, and P. M. VISSCHER, 2011 GCTA: a tool for genome-
wide complex trait analysis. Am. J. Hum. Genet. **88**: 76–82.

YANG, S.-W., C. GAO, L. CHEN, Y.-L. SONG, J.-L. ZHU, *et al.*, 2012 NEK9 regulates spindle
organization and cell cycle progression during mouse oocyte meiosis and its location in early
embryo mitosis. Cell Cycle **11**: 4366–4377.

ZERBINO, D. R., P. ACHUTHAN, W. AKANNI, M. AMODE, D. BARRELL, *et al.*, 2018 Ensembl
2018. Nucleic Acids Res. **46**: D754–D761.