Evaluation of redundancy analysis to identify signatures of local adaptation

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ABSTRACT

Ordination is a common tool in ecology that aims at representing complex biological information in a reduced space. In landscape genetics, ordination methods such as principal component analysis (PCA) have been used to detect adaptive variation based on genomic data. Taking advantage of environmental data in addition to genotype data, redundancy analysis (RDA) is another ordination approach that is useful to detect adaptive variation. This paper aims at proposing a test statistic based on RDA to search for loci under selection. We compare redundancy analysis to pcadapt, which is a nonconstrained ordination method, and to a latent factor mixed model (LFMM), which is a univariate genotype-environment association method. Individual-based simulations identify evolutionary scenarios where RDA genome scans have a greater statistical power than genome scans based on PCA. By constraining the analysis with environmental variables, RDA performs better than PCA in identifying adaptive variation when selection gradients are weakly correlated with population structure. Additionally, we show that if RDA and LFMM have a similar power to identify genetic markers associated with environmental variables, the RDA-based procedure has the advantage to identify the main selective gradients as a combination of environmental variables. To give a concrete illustration of RDA in population genomics, we apply this method to the detection of outliers and selective gradients on an SNP data set of *Populus*

21 *trichocarpa* (Geraldes *et al.*, 2013). The RDA-based approach identifies the main

22 selective gradient contrasting southern and coastal populations to northern and

23 continental populations in the northwestern American coast.

25 **KEYWORDS:** Genome scans - Multivariate analysis - Redundancy analysis –

26 Biological adaptation – Selection – Environmental variables

INTRODUCTION

Natural selection results from environmental pressures. The environment acts

simultaneously on several characteristics, and these adaptive characteristics can be

determined by a large number of alleles of small effects (Pritchard and Di Rienzo,

2010). Patterns of adaptive variation within species are usually studied through

physiological, morphometric and fitness comparisons in common garden

experiments. However, this is a difficult and constraining task, which can even be

unrealistic to achieve in species with a long generation time. Recently, the

development of next-generation sequencing (NGS) technologies has opened the

possibility to access to a large amount of genetic variation across the genome.

Therefore, relationships between genetic polymorphism, phenotypic variation and

environmental variables can now be studied and quantified by in situ approaches

(Steane *et al.*, 2014, Vangestel *et al.*, 2018).

By looking for loci that are either excessively differentiated between

populations or significantly associated with environmental gradients, adaptive

variation can be detected (Foll and Gaggiotti 2008, Frichot *et al.*, 2013, Vatsiou *et al.*,

2015, Duforet-Frebourg *et al.*, 2016, Hoban *et al.*, 2016).

In many genome scan procedures, outlier loci are identified based on genetic

differentiation, assuming that adaptive alleles have increased genetic differentiation.

A common statistic to evaluate genetic differentiation is F_{st} (Lewontin and Krakauer,

1973) and many likelihood or Bayesian methods use F_{st}-related statistics to scan

genomes for local adaptation (Bazin *et al.*, 2010, De Villemereuil and Gaggiotti, 2015,

Foll and Gaggiotti, 2008, Whitlock and Lotterhos, 2015). One limitation of F_{st}-based

approaches is that they require defining discrete populations, whereas many species

are not structured in different populations but display continuous genetic gradients

(Martins *et al.* 2016). In these cases, defining populations can be challenging and can

directly interfere with the result of the analysis (Yang *et al.*, 2012).

A solution to avoid such an issue is to use individual-based multivariate

methods such as PCA that ascertain population structure. Duforet-Frebourg and

colleagues (2016) proposed a genome scan method based on PCA (pcadapt), where

outlier loci are the ones excessively correlated with one or more ordination axes

(Duforet-Frebourg *et al.*, 2016). Simulations show that pcadapt has a similar power

to model-based methods under island models and perform better when the

simulated demographic model drifts from the island model, for instance under

hierarchical population structure or isolation by distance patterns (Luu *et al.*, 2017).

Nevertheless, one conundrum of such approaches is the difficulty to interpret

ordination axes in terms of ecological meanings. These are usually tight to

geographical axes (latitudinal or longitudinal), but they are not necessarily linked to

an environmental variable such as temperature, drought, diet habit, etc. When

environmental information exists, it has to be used, a posteriori, as a means of

interpretation, but it is not involved in the inference process.

Genotype-environment association (GEA) methods are based on a different

principle than methods based on genetic differentiation; they assume that adaptive

loci are significantly associated with environmental variation. These methods aim at

detecting alleles that are associated with environmental variables (e.g., temperature

and drought) with the idea that these alleles may confer a selective advantage in

some environment (Coop *et al.*, 2010, Frichot *et al.*, 2013). Most of these methods

only use one predictor variable at a time, which can be problematic when the main

selective gradient is unknown or when we want to disentangle multifactorial

associations between genetic and environmental variation.

Some authors have proposed to use a constrained ordination procedure such

as RDA (redundancy analysis) or CAPs (canonical analysis of principal coordinates)

to combine the advantages of both multivariate methods and genotype-environment

association procedures (Lasky *et al.*, 2012, DeKort *et al.*, 2014; Steane *et al.*, 2014,

Hecht *et al.*, 2015). This family of methods subsequently aligns adaptive variation

with environmental data in order to be able to identify the environmental gradients

that are the most correlated with adaptive variation. In *Arabidopsis thaliana,* loci

involved in adaptation to climate have been found using RDA (Lasky *et al.*, 2012).

Outliers were identified as SNPs with the greatest squared loadings along the first

RDA axes (i.e., those in the 0.5% tail). RDA can also be used to derive an adaptive

index that predicts the performance of individuals in different environmental

conditions. Steane *et al.* (2014) applied a similar approach (using canonical analysis

of principal coordinates) to *Eucalyptus tricarpa* populations and showed that both

detection of adaptive variation with RDA and common garden experiments provided

evidence for local adaptation. Finally, in two recent studies, Forester and colleagues

(2015 and 2017) tested the power of an RDA-based method for detecting signatures

of local adaptation in a heterogeneous landscape (Forester *et al*, 2015) and in

comparison with other constrained ordination methods (Forester *et al*, 2017). Using

simulations, they found that an RDA-based method has a superior combination of

low false positive and high true positive rates when compared to generalized linear

models (GLM) or latent factor mixed models (LFMM).

In this paper, we build on these previous applications of RDA to develop a

statistical test for searching loci under selection. We use simulations and real data

sets to shed light on the conditions within which RDA is the most efficient and to

document the possibilities given by RDA when studying landscape genomics. We

also compare results of a genome scan based on RDA with genome scans obtained

with *pcadapt* and LFMM (Frichot et al. 2013, Luu *et al.,* 2017). This investigation

gives the opportunity to show the advantages of a constrained ordination method

such as RDA in comparison with nonconstrained ordination method (PCA) and

univariate genotype-environment association method (LFMM). Finally, to illustrate

the pertinence of such a method in a concrete example, we apply it to a *Populus*

trichocarpa SNP dataset obtained from Geraldes *et al.* (2013). We found a large

overlap between the outlier loci detected with RDA and the loci previously detected

with *Fst* based methods (Geraldes *et al.,* 2013). In addition, RDA revealed what the

main selective gradients are along the natural populations of *Populus trichocarpa* in

the northwestern coast of North America.

MATERIALS AND METHODS

1.   Genome scan using an RDA-based approach

Redundancy analysis (RDA) was first introduced by Rao (1964). It is the

extension of multiple regressions to the modeling of multivariate response data

(Legendre and Legendre 2012, section 11.1). The data are separated in two sets, a

response matrix Y of the variable to be explained (e.g., species abundance in a set of

sites; m sites and n species) and an explanatory matrix X (e.g., a set of environmental

variables within each site; m sites and p environment). In the following analysis, loci

replace species abundances and individuals replace sites. RDA seeks to project the

genetic variation between individuals that is explained by environmental data on a

reduced space. RDA assumes that there are linear relationships between the

response (Y) and explanatory (X) variables.

The analysis starts using a classical RDA procedure performed with the

function *rda* of the R package *"vegan"* (Oksanen *et al.*, 2015). An individual genetic

dataset is used as the response matrix Y, and a set of environmental variables is used

as explanatory matrix X. We assume in the following that genotypes are coded using

the values 0,1, or 2 that correspond to homozygote for the most frequent allele,

heterozygote and homozygote for the less frequent allele. RDA amounts at

constructing a matrix Y' of fitted genetic values estimated from the regression of

each locus by the environmental variables and at performing principal component

analysis on the matrix Y' (Legendre and Legendre 2012, section 11.1). After this

constrained ordination step, we follow the methodology implemented in *pcadapt* to

find outlier loci (Luu *et al.,* 2017). First, we recover the loci loadings from the RDA

analysis. Only the loadings of the most informative ordination axis are kept for the

rest of the procedure. The number of axes used (K) is determined by looking at the

amount of information retained on the different axes of the RDA. A Mahalanobis

distance D is then computed for each locus to identify loci showing extreme D values

compared to the rest of the SNPs. A Mahalanobis distance is a multidimensional

generalization of the idea of measuring how many standard deviations is a point

from an average point. Computation of the Mahalanobis distance uses the *covRob*

function of the *"robust"* R package (Wang et al., 2014). Mahalanobis distances are

distributed as chi-squared distribution with K degrees of freedom after correcting

with the genomic inflation factor (Luu et al., 2017). Inflation factors are constant

values that are used to rescale chi-square statistics in order to limit inflation due to

diverse confounding factors (François et al 2016). We then adjust the resulting p-

values for the false discovery rate (FDR) by computing q-values with the "qvalue" R

package (Dabney and Storey, 2011). A locus is considered as an outlier if its q-value

is less than 10%, meaning that 10% or less of the identified outliers could be false

positives. The complete procedure is described in Fig. 1 and our R script is available

in the supplementary material (Script S1).

To evaluate the capacity of the RDA-based approach to detect loci underlying

environmental selective pressure, we first test the method with simulated datasets.

A total of 100 independent runs of simulations were used (see below for details). We

evaluate the power of the procedure based on RDA to detect positive loci (loci

simulated under selection) and to return a controlled number of false positives

(neutral loci detected as outliers by the analysis). Then, to emphasize the utility of

this method, we compared it to the PCA-based procedure implemented in *pcadapt*

that does not use environmental variables (Luu *et al.*, 2017). In the same vein, we

compare the multivariate RDA-based procedure to LFMM, which is an univariate

genotype-environment association method (Frichot *et al.*, 2013). This method

processes one environmental variable at a time, whereas we wanted to evaluate to

what extent different techniques can disentangle covariation between multiple

environmental gradients and multilocus genetic variation. To tackle this issue, we

performed the LFMM analysis with a set of composite environmental variables

corresponding to the first two PCs of a PCA performed on the environmental

simulated variables (Frichot *et al.*, 2013). It gave us orthogonal composite

environmental variables for which we tested covariation with genetic variability

through the LFMM procedure. The LFMM analysis was launched with only one latent

factor (K = 1), taking into account that simulations have been made without any

population divergence.

2. Simulated data

We used simulations performed with the simuPop python library (Peng *et al.*,

2005). A lattice of 8x8 populations is simulated (i.e., 64 populations in total). Each

population is initialized with 200 diploid individuals with random genotypes.

Migration is set to 0.1, resulting in an isolation-by-distance pattern across the

species range. Loci are assumed to be biallelic SNPs. The allele frequency of the

whole population is initialized at 0.5. 1000 loci are simulated in total and they are

separated in 200 chunks of 5 SNPs in physical linkage with the recombination rate

between adjacent loci fixed at 0.1. Three different quantitative traits are coded by a

group of 10 different loci (quantitative trait locus, QTL). The first quantitative trait is

coded by the loci 1, 11, 21,..., 91. The trait value is simply the sum of the genotype

values and therefore can take values between 0 and 20. We add to each trait a

random noise (nonheritable variation) drawn from a normal distribution $N(0,2)$. The

second quantitative trait is coded by loci 101, 111,..., 191 and the third is coded by

loci 201, 211,..., 291. Each quantitative trait is therefore coded by 10 independent

SNPs (QTL), resulting in a total of 30 causal SNPs among the 1000. Selection can

have an effect on linked loci, for instance, loci 2, 3, 4 and 5 can be impacted by

selection on locus 1. However, recombination is high enough (0.1) to expect a limited

linkage effect. We have defined 10 different environmental variables. The first one

determines the selective pressure on trait 1, the second one on trait 2 and the third

one on trait 3. The first environment variable is a quadratic gradient coded by

function $env1 = -(\cos(\theta)*(i\text{-}3.5))^2 -(\sin\theta)*(j\text{-}3.5))^2 + 18, \theta = \pi/2$, with i and j being

the population indicator on the 8x8 lattice. The second one is a linear plan gradient

coded by function $env2$ = h*cos(θ)*(i-1) + h*sin(θ)*(j-1) + k with h = 2, θ = $\pi/4$ and

k=3. The third environment variable simulates a coarse environment with value

$env3$ = 2 for all populations except population (i,j) = (2,2), (2,3), (3,2), (3,3), (6,2),

(6,3), (7,2), (7,3), (2,6), (2,7), (3,6), (3,7), (6,6), (6,7), (7,6), (7,7), for which $env3$ = 18.

Env4, env5 and env6 have exactly the same equation as env1, env2 and env3,

respectively, except for a noise term explained below. The remaining 4 environment

variables are similar to env2 but with different values of h and θ. Env7 has h = 2, θ =

0 and k = 3. Env8 has h = 2, θ = $\pi/4$ and k = 0. Env9 has h = 1, θ = $\pi/4$ and k=4. Env10

has h=0.5, θ = $\pi/4$ and k = 8. Graphical representation of the mean environmental

value is given in Fig. 2. Environmental variables 4, 5 and 6 have, respectively, the

same mean value spatial distribution as environmental variables 1, 2 and 3. The

environmental equation gives a mean value of the environment variable. To avoid

collinearity between environmental variables, we added noise by drawing an

environment value within a normal distribution N(μ = env, σ = 1). The fitness for

each trait is set to be $-\exp(x-env)^2/(2*\omega^2)$, with x being the quantitative trait

value, env the environmental value and ω is the defining selection strength and has

been set to 20, which was found to be enough for loci to be often detected. To get the

overall fitness for a given individual, the fitness associated with each trait is

multiplied. Fitness values are used to determine the number of offspring during the

simulations, which have been launched for 500 generations. At the end of the

simulation, we sampled 10 individuals per population resulting in 640 individuals

with 1000 SNP-like loci.

3.   **Real dataset**

The *Populus trichocarpa* dataset is a sample of 424 individuals genotyped on

33,070 SNPs from 25 drainages (i.e., topographic units separated by watershed

barriers) (Geraldes *et al.*, 2014). Genotyping of each accession was performed with a

34K Populus SNP array targeting 34,131 SNPs mostly within (plus 2 kb upstream

and downstream) 3543 genes (Geraldes *et al.*, 2013). A total of 21 climatic variables

are available on each sampling site (Table S1). From the 33,070 SNPs, we removed

the SNPs with missing values and a minor allele frequency below 5% resulting in

17,224 SNPs.

We explored the adaptive genetic variability of the *Populus trichocarpa*

dataset using PCA-based (*pcadapt*) and RDA-based method using the 21

environmental variables as explanatory variables (Table S2). Similar to the

simulations above, we created two environmental composite variables by keeping

the two first axes of a PCA on the 21 environmental variables. We then used the

latent factor mixed model (LFMM) procedure on these composite variables and with

one latent factor (K=1). We finally compared the list of outliers found with these

three approaches to the ones found using *Fst* based methods (Geraldes *et al.*, 2014).

We then performed a second RDA with only the loci previously found as

outliers (q-values < 0.1). This set of outliers gives an "adaptively enriched genetic

space" (Steane *et al.*, 2014) and using a second RDA on these specific loci, we can

identify the environmental variables that are the most correlated with "putatively"

adaptive variation. Using only outlier loci avoids potential interferences brought by

the "neutral" genetic variation during the alignment between genetic and

environmental variation performed by RDA. We then used the scores of the different

Populus trichocarpa populations along the first two RDA axes to build composite

indexes, which correspond to the main environmental pressures driving adaptive

genetic variation. The individuals' scores on the axes of the adaptively enriched

genetic RDA can be interpreted as cursors of individual genetic adaptation to the

environmental variables associated with these axes. We thus represented the mean

RDA1 and RDA2 individual score for each of the 133 populations on a map. It allows

visualizing the adaptive landscape across the sampling area.

RESULTS

1.   Genome scan on simulated dataset using RDA, pcadapt and LFMM

methods

The procedure was first performed on one simulation (sim1) with pcadapt,

RDA and LFMM approaches. To perform RDA, we considered environmental

variables 1-10 as explanatory variables. For pcadapt and RDA, we retained the first

four ordination axes to compute Mahalanobis distances as indicated by the scree

plots (Fig. S1). In parallel, we independently launched LFMM analysis with the first

two PCs resulting from a PCA applied to environmental variables (Fig. S2). We then

estimate that a locus is considered as an outlier by LFMM if there is a significant

correlation with at least one of the two composite variables.

The software package *pcadapt* is successful at detecting QTL2 SNPs (Fig. 3

topleft) but fails entirely at detecting QTL1 and QTL3 SNPs. The RDA-based

approach also detects QTL2 SNPS and most of the QTL1 SNPs in contrast to *pcadapt*

(Fig. 3 topright). The LFMM method is also able to detect most of the QTL1 and all

the QTL2 SNPs. These SNPs are detected when respectively regressing the genetic

dataset with the first and the second environmental composite variables (Fig. 3

bottom). For the SNPs associated with QTL3, the three methods have negligible

power to detect them.

When looking at the projection of each SNP in the RDA space, we can detect

links between QTL SNPs and environmental variables (Fig. 4). RDA1 is strongly

correlated to env2 and env5 and the QTL2 SNP show extreme scores on this axis

compared to the other SNPs, especially the neutral ones (Fig. 4A). RDA2 is also

associated with env7, env8, env9 and env10, which had not been used as direct

selective drivers during the simulations. RDA3 is correlated to env1 and env4

corresponding to QTL1 SNPs and there is also a weak association between RDA4

(env3 and env6) and the SNPs associated with QTL3 (Fig. 4B & 3C).

When looking at the percentage of environmental variance explained by the

RDA and PCA axes, we find that the first four RDA axes explain more variance of each

environmental variable than the first four axes of the PCA (Fig. 5). The difference is

especially true for env1 and env4, which are strongly associated with the first four

axes of the RDA, whereas PCA axes show an almost null correlation. The variables

env7, env8, env9 and env10 are also correlated to the first four axes of the RDA,

whereas PCs show far less correlation with these environments. Finally, RDA axes

show more correlation with the env3 and env6 variables than PCs (R^2 = ~0.30 vs R^2

= ~0). Furthermore, each RDA axis, from the first one to the fourth one, is associated

with more than one environmental variable, whereas only one PC (PC1) is

associated with env2 and env5 (Fig. S3).

Finally, over the 100 simulations, we measured the average FDR and power

for *pcadapt,* LFMM and RDA (Fig. 6). The three methods correctly identify almost all

of the QTL2 SNPs (Fig. 6A). RDA-based analysis and LFMM are more powerful than

pcadapt to identify QTL1 SNPs with, respectively, 75%, 48% and less than 10% of

true identification. For QTL3, *pcadapt* is not able to find any of the SNPs under

selection, whereas RDA reaches 24% of true identification and LFMM reaches 22%.

The three methods properly control for the false discovery rate because the

proportion of false discoveries obtained with *pcadapt*, LFMM and the RDA-based

methods are, respectively, 5%, 2.3% and 2.3% when controlling for a nominal error

rate of 5%.

2.   *Populus trichocarpa*

For the RDA and *pcadapt* analysis on the *P. trichocarpa* dataset, we retained

respectively $K=3$ and $K=6$ axes as indicated by the scree plots (Fig. S4) and we used

the two first axes of a PCA on environmental variables as composite variables for the

LFMM procedure (Fig. S5). Analysis with an FDR of 0.1 provides a list of 151 outliers

for RDA, 117 for *pcadapt* and 36 for LFMM. Interestingly, the three methods have

only 9 outliers in common and 117, 77 and 11 SNPs are outliers, respectively,

specific to RDA, *pcadapt* and LFMM.

When we compared RDA genome scan results with the outliers found by

Geraldes *et al.* (2014) based on Fdist, BayeScan and bayenv methods, we found a

substantial overlap between them (Fig. 7). Among the 151 outliers found with the

RDA-based method, 69 have already been picked up by *Fst*-based methods (see

Table S3). Interestingly, the loci showing the lowest p-values with the RDA analysis

are also pointed out as outliers by the *Fst* and *pcadapt* methods, but not by the

LFMM procedure (Fig. 7). However, the RDA-based genome scan found 72 SNPs that

no other methods, including *pcadapt*, detected as outliers.

When we performed a new RDA on the set of 151 outlier loci as detected by

the RDA-based genome scan method (adaptive enriched genetic space), we found

that the two first axes keep most (respectively 40% and 18%) of the adaptive

genetic variance among the populations (Fig. 8A). These two axes correspond to the

two principal selective pressures in the sampling area. RDA1 is correlated with

several environmental variables used in the analysis (Fig. 8A), including DD_0, DD18,

MAT, MCMT, and EMT, which correspond to temperature indexes. RDA1 is also

correlated to moisture variables (e.g., AHM or MAP), continentality (TD) and

flowering period length (bFFP and eFFP). This axis summarizes a gradient of global

climatic variation encountered in the sampling area with southern and coastal

populations opposed to northern and continental populations (Fig. 8B). On the other

side, RDA2 also correlates with moisture and temperature variables such as AHM,

MAP, TD, MCMT, and eFFP and contrasts individuals from coastal locations to

individuals from inland areas (Fig. 8C).

DISCUSSION

Genome scans based on genetic differentiation, as implemented in *pcadapt*,

detect true positives when environmental gradients related to adaption are

correlated to population structure. For instance, QTL2 SNPs are adaptive along an

environmental gradient that is correlated with population structure (see Fig. 2) and

they are detected by *pcadapt* (Fig. 6). By contrast, when the environmental gradient

is not correlated to population structure (QTL3 and QTL1 to a lesser extent) it fails

to detect adaptive SNPs. In this case, PCA ordination is not capable of orienting

genetic variation into the direction of environment 1 or 3. Fig. 6 shows that RDA has

a larger statistical power than *pcadapt* to detect QTL1 SNPs by taking advantage of

information from environmental local conditions. It can be attributed to a better

alignment between the genetic space and the environmental variable improving the

power to detect true positives. This better alignment is easily visible when looking at

the PCA and RDA correlation with environmental variables (Fig. 5 & Fig. S3). PCA is

only capable of identifying axes, which correspond to maximum genetic

differentiation (PC1 and PC2), whereas RDA is also able to capture the genetic

variability associated with other environmental variables, such as env1 (RDA3) and

to a lesser extent env3 (RDA4).

Genotype-environment approaches (GEA) explicitly take into account the

relationship between genotypes and the environment (Joost *et al.* 2007, de

Villemereuil *et al.* 2014, Lotterhos and Whitlock 2015). We showed that univariate

GEA, implemented in LFMM, and multivariate GEA, based on RDA, can show similar

results in identifying adaptive variation although RDA is more powerful than LFMM

for QTL1 SNPs (Fig. 6). As suggested in the literature, we limited the number of

environmental predictors during the LFMM procedure by using environmental PCA

axes as composite variables (Frichot *et al.*, 2013). We then only used the two first

PCs, but adding the third and fourth ones could potentially increase the capacity of

LFMM to detect QTL1 and QTL3 SNPs. It points out one of the advantages of a

multivariate procedure (e.g., RDA) compared to univariate GEA because multivariate

methods take all environmental variation into account at the same time and can

simultaneously detect associations between different sets of loci and different sets of

environmental variables (Forester *et al.,* 2017).

Both RDA and LFMM fail to detect QTL3 outliers in, respectively, 75% and

80% of the simulations (Fig. 6), possibly because of the additional difficulty in

detecting adaptive SNPs on a coarse environment compared to a smooth

environmental gradient as in environment 1 and 2. LFMM and the RDA-based

genome scan rely on linear association between the environment and allelic

abundance (0, 1 or 2). Patchy selective pressure, such as env3, is difficult to identify

with a correlative approach. Forester and colleagues (2015) already found that

heterogeneous environmental selection decreases the capacity of RDA to identify the

loci under selection. The strength of selection and the level of allelic dispersion

during the simulations strongly influence the statistical power of RDA; weak

selective strength and strong dispersion are the less favorable conditions (Forester

et al., 2015). Nonetheless, our simulations plead in favor of using a constrained

ordination method when relevant environmental variables are available in order to

both orientate the ordination axis in the direction of environmental gradients and to

account for most of the environmental variation. This result confirms that

constrained ordination shows a superior combination of low false positive and high

true positives rates across all levels of selection (Forester *et al.*, 2017).

During our RDA-based procedure, we use Mahalanobis distances as a test

statistic to identify genetic variation significantly associated with environmental

gradients. This distance is computed based on the scores of more than one RDA axis

and can thus condense the information of several RDA axes in one statistic. Some

other studies have used different RDA-based statistics to identify the loci under

putative selection. They often used a value estimated independently for the different

RDA axes, such as the loadings of the loci along the axis (e.g., extreme squared

loadings for Latsky *et al.*, 2012) and considering a threshold such as ± 3 SD (Forester

et al., 2017). Mahalanobis distances can capture complex selection patterns where

allelic variation depends on more than one environmental gradient. This statistic is

appropriate to identify selective pressures on natural populations, which usually

encompass multifactorial gradients of selection (Salmela, 2014).

One application of ordination methods is to detect selective gradients. The

RDA-based approach succeeded very well in decomposing the QTL SNPs-

environments associations along the different axes, at least in our simulations. In our

example, we succeeded in dissociating the outliers linked to the different

environmental variables, and thus to precisely identify which set of variables drives

each part of the adaptive genetic variability (Fig. 4). RDA1 is strongly associated with

env2, RDA2 with env7 (collinear with isolation by distance pattern in the grid),

RDA3 with env1 and RDA4 with env3. As expected, the correlated environments are

also strongly associated with these respective axes. This is reflecting the fact that it

is difficult to distinguish among several collinear variables which environmental

variable has a causal effect on the individual fitness. However, it is often sufficient to

identify the combination of environment variables having a strong association with

adaptive variation without precisely knowing the underlying mechanical process.

These simulations serve as a proof of concept to show that ordination axes can

provide holistic measures of genomic adaptation (Steane *et al.,* 2014). Under this

idea, scores of individuals on one RDA axis reflect their degree of adaptation along

this axis. RDA produces useful visualization of gradients of genetic adaptation (Fig.

8).

From the analysis of *Populus trichocarpa*, RDA found numerous outlier loci

(72) that are not detected by any other genome scan method, including *pcadapt* and

LFMM (Fig. 7 and Table S3). These results support the simulation analysis

conclusions that the RDA-based genome scan is able to pick up more genetic

markers in relation to environmental gradients than *pcadapt*. It is, however,

noticeable that a proportion of outliers are shared between PCA and RDA-based

methods (29), highlighting the fact that both approaches are partially similar. The

difficulty in interpreting the PCA axes is compensated by the fact that it is a blind

method in regard to environmental data, so it can pick up genes that we will miss

using RDA because some crucial environmental data are lacking. In the same vein,

half of the outlier loci found with the LFMM procedure (14 over 33 loci) are common

with RDA results (Fig. 7 and Table S3). The two methods rely on linear regression

between environmental and genetic variability, and by using PCs as composite

environmental variables in the LFMM procedure, we even add some similarity

between the two approaches.

After focusing the RDA analysis on the loci showing significant association

with environmental variables (151 outliers), we identified the predominant

environmental gradients structuring the genetic adaptation of *Populus trichocarpa.*

The main one is a composite climatic variation (RDA1 in Fig. 8), gathering

precipitation, temperature and seasonality variations. It clearly contrasts the

populations coming from the extreme northern part of the sampling area (Alaska)

and the population coming from the inland locations (higher altitudes) to the

populations close to the coast in British Columbia (Canada) and the southern

populations, near Portland, Oregon (USA). This result is in accord with the

conclusions of Geraldes *et al.* (2014). In their study, they point out the need of

adaptation to seasonality, photoperiod and frost events for *Populus trichocarpa*

along its continuous distribution from 44 to almost 60 degrees of latitude. RDA2 is

associated with moisture regimes and points out a gradient of adaptation between

coastal locations, strongly watered, and inland populations, receiving less

precipitation (Fig. 8). These two environmental gradients are the strongest selective

pressures driving *Populus trichocarpa* climatic adaptation in northwest America,

considering the set of environmental variables available and the scale of the analysis.

To go further, these composite indexes could also serve to predict (i) provenances

that would perform well in common gardens; (ii) patterns of adaptation to local

climate across the geographic range; or (iii) future adaptive landscapes in the

context of climate change.

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beloved friend Eric Bazin (1977-2017) who passed away too soon.

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DATA ACCESSIBILITY STATEMENT

- Simulations datasets are available at the Dryad Digital Repository at the following address:
  https://datadryad.org/review?doi=doi:10.5061/dryad.1s7v5

- Scripts used for simulations and *Populus trichocarpa* analysis are available on Github: https://github.com/Capblancq/RDA-genome-scan

- *Populus trichocarpa* SNP data are available on Dryad:
  http://datadryad.org/resource/doi:10.5061/dryad.1051d

AUTHOR CONTRIBUTIONS

E. Bazin designed the study. E.B. and T. Capblancq performed the analysis and treatments. M. Blum, E.B. and T.C. wrote the manuscript, and all authors contributed substantially to revisions.

TABLES AND FIGURES

Figure 1: Schematic representation of the RDA-based procedure for searching loci under selection.

Figure 2: Graphical representations of mean environmental values for environments 1 to 10. Equation determining the environment is given in the main manuscript. Mean environmental values for environments 4, 5 and 6 are, respectively, equivalent to environments 1, 2 and 3.

Figure 3: Manhattan plots of the results obtained with *pcadapt* (upper left panel), RDA (upper right panel), and LFMM (lower panels) for one simulated dataset.

Figure 4: Projection of SNPs and environmental variables into the RDA space.

Figure 5: Proportion of environmental variance explained by the first four ordination axes using RDA or PCA. The values correspond to the r-squared of the regression between the environmental variable and the first four axes of the multivariate analysis.

Figure 6: Statistical power obtained with RDA, *pcadapt*, and LFMM. Estimate of power corresponds to an average over 100 simulated datasets (error bars are displayed). Power (A) is given separately for loci coding for quantitative traits 1, 2 and 3. False Discovery Rate (FDR) is given for all methods (B).

Figure 7: Scatter plot of the p-values returned by RDA and *pcadapt* procedures (A) and the p-values returned by RDA and LFMM procedure (B) for the *Populus trichocarpa* dataset. Black dots correspond to loci identified as outliers by *Fst* methods in Geraldes *et al*. (2014). The dashed lines indicate a 10-3 p-value with RDA and *pcadapt* or LFMM.

Figure 8: Projection of SNPs and environmental variables in the RDA performed on the adaptively enriched genetic space (151 loci) of *P. trichocarpa* dataset (A). The first two axes of the RDA projection represent 40% (RDA1) and 18% (RDA2) of the explained variance. (B) and (C) show respective spatial representations of RDA1 and RDA2 axes scores in the sampling area. The points represent the 133 sampled populations and their color depends on the mean of individual RDA1 (B) and RDA2 (C) scores in the populations.

SUPPORTING INFORMATION

Script S1: Detailed script of the RDA-based genome scan analysis.

Table S1: Environmental variables used in the *Populus trichocarpa* analysis.

Table S2: Climatic variable values for the 133 *Populus trichocarpa* sampled
populations.

Table S3: Comparison between RDA, *pcadapt*, *fdist* and *bayescan* outlier loci.

Figure S1: Scree plot for PCA and RDA analysis of the simulation 1 dataset.

Figure S2: PCA and scree plot of PCA on environmental variables of the first
simulation.

Figure S3: Heatmap of the R^2 resulting from linear regression between the four first
axes of PCA and RDA analysis and the environmental variables.

Figure S4: Scree plot for PCA and RDA analysis of *Populus trichocarpa* dataset.

Figure S5: PCA and scree plot of PCA on environmental variables of the *Populus*
trichocarpa dataset.