1   # Dissociable Roles for Working Memory in Sensorimotor Learning

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ABSTRACT

The computations underlying cognitive strategies in sensorimotor learning are poorly understood. Here we investigate such strategies in a sensorimotor transformation task. We show that strategies assume two forms, reflecting distinct working memory representations: discrete response caching of stimulus-response contingencies (e.g., look-up table; RC), and time-consuming parametric computations (e.g. mental rotation; MR). Subjects' reaction times and errors suggest that both strategies are employed during learning, and trade off based on the progress of learning and the complexity of the task. Experiments using pressured preparation time support these working memory mechanisms: In discrete RC, time pressure elicits bimodal distributions of movements, in agreement with cached responses; in parametric MR, time pressure elicits a shifting distribution producing intermediate movements between visual targets and distal goals, consistent with analog re-computing of a movement plan. These results provide a specific model of working memory contributions to motor learning.

INTRODUCTION

When first learning a new motor skill, the process of selecting an appropriate action often involves a time-consuming, deliberative, and seemingly effortful process. Consider someone first learning to play the piano. Ideally, she would instantly be able to see the notes on the staff and immediately know the appropriate piano keys to play. However, as anyone who has ever attempted to play a new instrument knows, reading music is simple when the musical score only has a few notes in a small range, but as it gets more complicated things fall apart — if we just consider the number of piano keys in an octave, it easily exceeds our working memory capacity (Miller, 1956). A common strategy to overcome this limitation is to approach it parametrically: She can anchor her thumb on middle C and reference other notes on the lines of the musical staff relative to this piano key. While this strategy affords her the ability to play a more complex melody within a few minutes of practice, it becomes increasingly cumbersome the further a given note is from middle C — bearing remarkable similarity to the phenomena of mental scanning and mental rotation (Kosslyn et al., 1978; Shepard and Metzler, 1971). These two strategies, one discrete (caching each response-note relationship) and one parametric (computing the distance from C), offer two ways to approach learning a complex novel motor skill.

Even in a much simpler motor task, like the visuomotor rotation task (Krakauer, 2009), subjects appear to leverage strategies to rapidly improve performance, effectively "bootstrapping" the learning curve (Bond and Taylor, 2015; McDougle et al., 2017). These strategic processes appear to be related to higher reaction times (Haith et al., 2015a), improved task performance (Fernandez-Ruiz et al., 2011), and the direction of eye gaze (Rand and Rentsh, 2015; de Brouwer et al., 2018). Increased reaction time, and the fact that strategies are often verbalizable, suggests that they may reflect a form of deliberative, controlled processing rather than automatic processing (Cohen et al., 1990). Controlled processes appear to be reliant on working memory, making it one candidate system that may underlie the observed strategies in motor learning. Indeed, evidence suggests that spatial working memory ability correlates with performance in visuomotor learning (Seidler et al., 2012), and recruits similar neural circuits (Anguera et al., 2010). Moreover, spatial working memory ability specifically correlates with the use of explicit strategies in visuomotor rotation tasks (Christou et al., 2016). However, it remains unclear what operations of working memory directly drive strategies for motor learning.

Here we set out to directly characterize how known functions of working memory may manifest in a visuomotor rotation task (VMR), which requires subjects to adapt to sensory feedback that is rotated relative to their direction of movement. We hypothesized that deliberate working memory strategies could take two broad forms, either a discrete response caching (RC)

strategy or a parametric mental rotation (MR) strategy. A discrete RC strategy is here defined as

the maintenance of acquired associations between a set of stimuli and a set of responses

maintained in working memory (Collins and Frank, 2012), relying on processing in the prefrontal

cortex (Petrides 1985; Curtis & D'Esposito, 2003). As a form of discrete item-based working

memory, the efficacy of a discrete RC strategy should be subject to both load (e.g. the number of

items to be stored) and interference (e.g. the introduction of intervening items between trials).

Discrete RC should also be computationally inexpensive in simple learning contexts, but grow in

cost and decline in accuracy as the learning environment becomes more complex (Cowan, 2010;

Collins and Frank, 2012).

A parametric MR strategy, on the other hand, is the canonical example of an "analog"

computation (Shepard & Metzler, 1971; Cooper, 1976), in which an internal mental representation

is manipulated like an external physical object in visual working memory (Hyun & Luck, 2007). In

the motor domain, evidence from behavior and neurophysiology provides support for the use of

mental rotation in the planning of reaching movements: Reaction time (RT) during the mental

rotation of a reaching direction scales with the magnitude of required rotation, sharing remarkable

similarities to the mental rotation of visual percepts (Georgopoulos & Massey, 1987; Pellizzer &

Georgopoulos, 1993; Bhat & Sanes, 1998). Moreover, the decoded direction of neuronal

population vectors in motor cortex appears to "rotate" through direction space during the mental

rotation of a reach plan (Georgopoulos et al., 1989; but see Cisek & Scott, 1999). Critically, mental

rotation can compute rotations with arbitrary signs or magnitudes, and through multiple planes

(Shepard & Metzler, 1971). Mental rotation thus constitutes a flexible computation that appears

to rely on a straightforward heuristic that is formally equivalent to the application of a rotation

matrix to a mental representation. Similar heuristics appear to exist for other sensorimotor

transformations — for instance, mental-rotation-like RT effects are observed when humans

compute gain perturbations on reaching movements (Bhat & Sanes, 1998); here, the heuristic

would simply represent a scalar transformation rather than a rotation, though the same logic

applies.

We tested these ideas over three experiments: In *Experiment 1*, we provide evidence that

both discrete RC and parametric MR strategies are recruited in standard VMR tasks, where a

discrete RC strategy is recruited when the load is sufficiently low, and it is not sensitive to rotation

magnitude. A parametric MR strategy appears to be recruited in high load conditions, and it is

sensitive to rotation magnitude. In *Experiments 2* and *3*, we further elucidate the computations

underlying these two working memory strategies, showing how time pressure during the RT

reveals evidence for distinct within-trial parametric MR and discrete RC computations. Taken

together, these results reveal at least two roles for working memory in motor learning, with

implications for the psychological and neural mechanisms of skill acquisition.

RESULTS

Experiment 1: Dissociable working memory strategies in a visuomotor adaptation task

Subjects performed a standard visuomotor rotation task (see Methods; for a review see

Krakauer, 2009), where visual feedback was rotated relative to the reaching direction of the hand

(Fig. 1). Our hypothesis was that in low set size conditions (small number of targets), subjects

would primarily use a discrete response caching (RC) strategy (i.e., a look-up table), while in high

set size conditions they would use a parametric mental rotation (MR) strategy (i.e., rule-based

transformation algorithm). We grouped subjects via a 2 X 2 factorial design, crossing the factors

of set size and rotation magnitude. For set size, subjects were exposed to either 2 or 12 possible

target locations during the task: In the 2T condition, one of two targets, separated by 180°, was

presented on each trial (Fig. 1). The precise location of targets were randomized and

counterbalanced across subjects. In the 12T condition, target locations were pseudorandomly

presented at one of 12 possible locations (Fig. 1). For rotation magnitude, subjects either

Figure 1: *Experiment 1 task design*. Subjects performed a standard visuomotor rotation task, attempting to land a rotated visual cursor on a target. A between-subjects, 2 X 2 design was used, crossing the factors Rotation Magnitude (the size of the rotation in the rotation block; left) and Set Size (the number of possible target locations in the task; right).

experienced a 25˚ or 75˚ rotation during the rotation block (Fig. 1; varying rotation magnitude was necessary to test for the operation of a parametric MR strategy.

Subjects performed 36 baseline trials (no rotation), 300 rotation trials, and 36 washout trials (Fig. 1). Critically, feedback, in the form of a small visual cursor, was only provided at the end of the reach and was delayed by 2 s, a manipulation which has been shown to limit implicit motor adaptation (Kitazawa et al., 1995; Brudner et al., 2016; Schween et al., 2017) and thus better isolate strategic learning.

Our first analysis focused on reaction time (RT), a standard metric for quantifying the operation of controlled processes (Cohen et al., 1990). We hypothesized that in the high set size group (12T), RT would be modulated by rotation magnitude throughout learning, consistent with the idea that subjects were parametrically mentally rotating their motor plans on every trial. However, in the low set size group (2T), we expected this effect only early in learning while subjects discovered the solution to the task, after which we expected RTs to converge across rotation magnitudes due to the caching of discrete responses.

Consistent with these predictions, we observed that RT decreased over training for all conditions but it remained specifically elevated for the group of subjects training to counteract a 75° rotation at 12 target locations and somewhat less so for the 25˚ rotation at only 2 target locations (Fig. 2A). In contrast, subjects counteracting a 75° rotation at only 2 target locations showed a rapid drop in RT similar to groups counteracting a much smaller rotation. To quantify this, we examined the means of median RTs over 6 cycles between early learning (first 6 cycles) and late learning (last 6 cycles) across our four conditions. Note, we defined a cycle as 2 trials in the 2T group and 12 trials in the 12T group to control for inherent differences between set size conditions which affect the number of exposures at each target location. These RTs (Fig. 2C) were submitted to a 3-way mixed factorial ANOVA with a within-subjects factor of Time (early vs. late learning), and between-subject factors of Set Size (2T vs. 12T) and Rotation Magnitude (25° vs. 75°).

We found a significant main effect of Time ($F(76) = 67.97$, $p < 0.001$), reflecting the fact that all groups decreased their RT over training. A significant effect of Rotation Magnitude was also observed, suggesting that subjects may be employing a parametric MR strategy — larger rotations resulted in larger RTs ($F(76) = 8.32$, $p = 0.005$). While we did not find a main effect of Set Size ($F(76) = 0.90$, $p = 0.34$), we did find a significant two-way Time X Set Size interaction ($F(76) = 20.87$, $p < 0.001$) and, more importantly, a three-way Time X Set Size X Rotation interaction ($F(76) = 4.56$, $p = 0.036$). Two *post hoc* *t*-tests (Bonferroni-corrected) revealed an effect of Rotation Magnitude on late learning RT only in the 12T group ($t(38) = 2.45$; $p = 0.019$),

Figure 2: *Experiment 1 results*. (**A**) Time course of mean RTs, averaged over 2-trial and 12-trial cycles in the 2T and 12T groups. Vertical lines delineate the task block (baseline, rotation, and washout). (**B**) Movement angle across the groups, averaged over 12-trial bins. (**C**) Mean of median RTs averaged over the cycles of trials in the beginning (early) and end (late) of learning, and the full baseline and washout blocks (see Methods). Error bars = 1 s.e.m.

but not in the 2T group ($t(38) = 0.15$; $p = 0.88$). Taken together, these results are consistent with the hypothesis that all groups may have used a parametric MR strategy to find the task solution early in learning, and then either stayed with MR for the remainder of the task (12T group), or transitioned to a discrete RC strategy, where RT was not modulated by rotation size (2T group).

It is worth noting that we observed no significant differences in RT in either the baseline or washout blocks (Fig. 2C). The mean of median RTs over the baseline revealed no significant difference with respect to Set Size or Rotation Magnitude ($p = 0.63$ and $p = 0.97$, respectively). While we did not expect any differences as a function of Rotation Magnitude — as the rotation had yet to be imposed — there could have been differences in Set Size as subjects had less trials at each target location, or had to perform an additional visual search to locate the changing target location. Likewise, in the washout block we found no main effect of Rotation Magnitude on RT ($F(76) = 2.40$; $p = 0.13$) and only a marginal effect of Set Size ($F(76) = 3.35$; $p = 0.07$). Given this trend in washout block RTs (Fig. 2C), one speculation is that differences in RT at the end of

training carried over into the washout, perhaps due to habitual rather than decision-related factors (Wong et al., 2017). Overall, the baseline and washout RT results show that differences in RT were limited to the rotation block.

Furthermore, these differences in RT could not be attributed to differences in performance between conditions: All groups displayed similar asymptotic performance with some subtle differences in the rate of learning (Fig. 2B). To quantify this, we computed the average movement angular error (i.e., rotation - movement angle) over the first 6 cycles (early learning) and last 6 cycles (late learning) for all groups and submitted these values to a two-way repeated measures ANOVA. We found a significant main effect of Time on movement error ($F(76) = 74.70$, $p < 0.001$), reflecting learning. We found no significant main effects of Rotation ($F(76) = 0.23$, $p = 0.64$) or Set Size ($F(76) = 1.25$, $p = 0.27$), nor their interaction ($F(76) = 1.34$, $p = 0.25$). We also found no interaction between Rotation X Time ($F(76) = 1.18$, $p = 0.28$). We did, however, find a significant Time X Set Size interaction ($F(76) = 6.28$, $p < 0.014$), reflecting the slight learning advantage seen in the 2T groups. Taken together, the general lack of significant results suggests that neither rotation magnitude nor set size significantly affected the efficacy of employing strategies in VMR, consistent with previous results (Bond and Taylor, 2015; McDougle et al., 2017). Finally, there were no significant differences in movement error over the baseline block with respect to Set Size or Rotation Magnitude ($p = 0.96$ and $p = 0.66$, respectively).

We now consider the role of implicit motor adaptation in our task. First, delayed feedback appeared to be effective for limiting implicit adaptation, as average aftereffects across the sample were only 2.71˚ (Fig. 2B). This suggests that the vast majority of learning was driven by a strategy rather than implicit motor adaptation (Taylor et al., 2014). However, even though aftereffects were subtle, there were significant differences in movement angle in the washout block with respect to both Set Size and Rotation Magnitude ($F(76) = 16.10$, $p = 0.002$ and $F(76) = 9.95$, $p < 0.001$, respectively), though with no significant interaction ($F(76) = 1.67$; $p = 0.20$). Corrected t-tests revealed that only the 2T 25˚ group showed statistically reliable aftereffects ($t(19) = 3.49$; $p < 0.001$; all other p's > 0.09). Parsimony would suggest this effect was not due to implicit adaptation, as in that case significant aftereffects would be seen across all four conditions (Bond and Taylor, 2015). This result is, however, consistent with use-dependent learning, which describes a bias toward repeated movement directions: Use-dependent learning would be stronger in the 2T setting because only two responses are repeated and would thus create a stronger bias (Verstynen and Sabes, 2011).

Distributions of large reaching errors support strategy differences

Two secondary analyses also comport with our hypothesis of discrete RC and parametric
MR working memory strategies. The first analysis was a *post hoc* analysis of large errors: We
analyzed subjects' "sign" errors (i.e. trials where subjects reached in the wrong clockwise/counter-
clockwise direction relative to the correct response). Normatively, discrete RC and parametric MR
strategies make different predictions about sign errors: In discrete RC, sign errors should
represent trials where participants accidentally extract the wrong response held in working
memory (i.e. a working memory "swap" error; Bays, 2016). This would constitute a -155˚ error in
the 25˚ condition, and a -105˚ error in the 75˚ condition. In parametric MR, however, sign errors
should represent trials where participants accidentally "flip" the rotation angle, thus reaching with
the correct magnitude relative to the target but in the wrong direction. This would constitute a -25˚
error in the 25˚ condition, and a -75˚ error in the 75˚ condition.

As sign errors were relatively rare (especially in the 2T groups), we pooled the sign errors
over all subjects in each group. Sign errors were designated as trials where the subject reached
≤ -15˚ from the target location (i.e. in the direction opposite the correct response). As depicted in
the histograms in Figure 3A, swap errors were the more common sign error in the 2T groups,
consistent with a discrete, item-based working memory representation. Flip errors, however, were
the more common error in the 12T groups, consistent with a parametric computation that

Figure 3: *Sign errors and interference.* (**A**) Histograms of pooled "sign error" data in each group. Vertical dotted lines represent predicted locations of "swap" errors (generating the wrong response given two choices) and "flip" errors (reaching relative to the correct magnitude of rotation but the wrong sign). (**B**) Normalized absolute reach errors, grouped by the interval between the current trial and the previous trial in which the same target was presented. Error bars = 1 s.e.m. ** = $p < 0.01$

211   occasionally confuses the sign of rotation. We note however, that given the sparsity of these data
212   and the necessity of pooling, this is a qualitative analysis that should be interpreted with
213   appropriate caution.
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Evidence for interference effects in low set size but not high set size conditions

One hallmark of working memory is forgetting in a working memory trace due to
interference (Peterson and Peterson 1959; Jonides et al., 2008), which can here be
operationalized by the number of intervening presentations of a new stimulus between
presentations of a single learned item held in working memory (Gorgoraptis et al., 2011; Collins
and Frank, 2012). If the discrete RC strategy above reflects a form of "motor working memory,"
where the working memory trace represents a specific movement direction, interference would
suggest that movement precision should decrease as the number of intervening trials between
presentations of the same target increases. We hypothesized that we would not see this effect in
the case of a parametric strategy like mental rotation, as it involves a single rule that does not
require maintaining individual responses in working memory.

To assess the effect of interference on reach error, we first combined the groups based
on set size (i.e. the hypothesized strategy), and normalized each subject's absolute reach error
during the full rotation block between 0 and 1. We then classified errors based on the number of
trials elapsed since the current target was previously encountered (range: 1-3 trials for 2T groups,
2-23 trials for 12T groups), and performed regressions for each subject. For display purposes
(Fig. 3B), all longer intervals (≥ 3) in 12T groups were averaged and plotted in a single bin. As
predicted, we found a significant effect of delay in the 2T condition (Fig. 3B; t-test on regression
coefficients: $t(39) = 2.18$; $p = 0.007$), but not in the 12T condition ($t(39) = 0.70$; $p = 0.48$), consistent
with interference-based noise in an S-R working memory trace (Collins and Frank, 2012), but
relative stability in a parametric-MR strategy.

Experiment 2: Constraining RT reveals parametric transformation of movement plans

Experiment 1 provided evidence supporting our hypothesis that parametric MR and
discrete RC strategies comprise two deliberate action selection strategies in a visuomotor
adaptation task. In *Experiments 2* and *3*, we wanted to further confirm that these mechanisms
were valid descriptions of within-trial mental computations. To do this, we adopted a "forced
reaction time" task (Ghez et al., 1997), which limits the amount of time subjects have to prepare
their responses. This procedure has recently been used to show how different learning processes
progress in VMR (Haith et al., 2015a). Here, however, we wanted to test how rapidly-induced

movements could "read out" the state of cognitive processing (Cooper, 1976). Our specific

hypotheses were as follows: Interrupting a parametric MR strategy should induce intermediate

movements that represent partially-rotated movement plans (*Experiment 2*), but interrupting a

discrete RC strategy should induce bimodal movement distributions with modes at each cached

movement location, and negligible intermediate movements (*Experiment 3*).

To investigate mental rotation-like computations, we used a within-subject design with two

tasks: One task (FREE task) had no constraints on RT — this was essentially a control task

designed to replicate previous findings that, during putative mental rotation of a reaching

movement, RTs are a linear function of rotation magnitude (Georgopoulos & Massey, 1987;

Pellizzer & Georgopoulos, 1993; Bhat & Sanes, 1998). The second task (FORCED task) used a

forced-RT procedure to reveal the within-trial dynamics of a parametric MR strategy.

In the FREE task, subjects performed a series of trial pairs (Fig. 4A): In the first trial of

each pair, the "learning" trial (Fig. 4A, left), subjects were instructed to reach directly at the

displayed target and observe where the cursor landed. Feedback of the cursor was limited to

Figure 4: *Experiment 2 task design.* (**A**) FREE task: Subjects performed trial pairs consisting of learning trials (left) and execution trials (right). Thirteen rotation sizes were pseudorandomly presented (-90°:15°:90°). (**B**) FORCED task: Subjects were required to initiate their movement after target presentation and < 100 ms after the fourth tone. Targets could appear in either one of 12 locations (Exp. 2) or one of 2 locations (Exp. 3). The time of target appearance was titrated to induce subjects to react with a distribution of RTs.

endpoint feedback, which was presented after the subjects passed the invisible target ring, and
the appearance of the feedback was delayed by 500 ms to inhibit implicit adaptation, as in
Experiment 1. In the second trial of the pair, the "execution" trial (Fig. 4A, upper right), subjects
were told to apply what they learned about the relationship between their movement and the
resultant feedback and attempt to make the cursor terminate within the target (target positions
were different between the learning and execution trials; see Methods). Rotations applied in the
learning trials ranged from -90˚ to 90˚ by 15˚ intervals and were pseudo-randomly presented
throughout the task. We hypothesized that RT would be a linear function of subjects' reaching
angles as they counteracted rotations of varying magnitude (Georgopoulos & Massey, 1987; Bhat
& Sanes, 1998). Subjects performed the FREE task in a single block consisting of 140 trial pairs.
In the FREE task, subjects were generally accurate on execution trials, with reach angles
on the execution trials tracking the magnitude of the rotation imposed on the corresponding
learning trial (Fig. 5A; *t*-test on regression slopes: $t(31) = 30.57$, $p < 0.001$). As predicted, RT on
execution trials was linearly scaled by the angle of movement produced relative to the target (Fig.
5B; *t*-test on regression slopes: $t(31) = 6.59$, $p < 0.001$), consistent with a mental rotation (MR)
process. These results replicate the result that cognitive re-planning of a reaching vector
(Georgopoulos & Massey, 1987; Bhat & Sanes, 1998) also displays the psychophysical standard
applied to the mental rotation of visual percepts (Shepard and Metzler, 1971).
The FORCED task utilized a modified forced-RT paradigm to interrupt putative mental
rotation (Ghez et al., 1997; Haith et al., 2016; Fig. 4B), with the prediction that movement angle
would be a linear function of RT. On each trial, a countdown of four audio tones was played. The
tones were played 600 ms apart, and subjects were instructed to try and synchronize the initiation
of their reach with the fourth tone, effectively "replacing" that tone, allowing us to constrain their
RTs (see Methods). Targets could appear in one of 12 off-cardinal locations (10˚, 40˚, 70˚, 100˚,
130˚, 160˚, 190˚, 220˚, 250˚, 280˚, 310˚, 340˚). Feedback of the cursor was also delayed by 500
ms to prevent implicit adaptation. The moment of target appearance was titrated such that
subjects had varying amounts of time with which to compute the target location, plan, and execute
their movements. Subjects performed the FORCED task in three blocks, with 48 baseline trials,
624 rotation trials, and 36 washout trials. For the rotation trials, a fixed rotation of 90˚ (or -90˚, for
counterbalancing) was imposed on the cursor, and subjects were thoroughly educated about the
rotation before this block began.
We focused on trials where subjects reached on time in accordance with the fourth tone,
the "Go" signal ($\mu = 78.90\%$ of trials). Our first analysis involved measuring subjects' reach angles
as a function of RT. RTs were binned by 25 ms from 0 ms through 400 ms, with the final bin

including all RTs above 400 ms. To quantify MR, we first needed to identify the RT at which

subjects' movement angles "settled" on a direction that was reliably related to the presented target

location (at very short RTs, movements should be directed randomly since there is insufficient

time to process the target). To determine this "critical" RT, we looked at variance in reach

directions in each successive 25ms RT bin. Circular variance first significantly decreased ($t(31)$ =

2.28, p = 0.02; Supplementary Fig. 1) from the 7th to the 8th RT bin (150-175ms), suggesting that

at RTs over 150ms subjects, began to make non-random movements. We note that this result is

consistent with previous work (Haith et al., 2016).

After this critical bin, reaching angles linearly increased with RT (Fig. 5C; $t(31)$ = 14.17, p

< 0.001). Figures 5C and 5D depict results as, respectively, circular means, and pooled circular

means with vector lengths. This result closely echoes the observed rotation of a population vector

in the motor cortex, and suggests a mental rotation operation during the RT (Georgopoulos et al.,

Figure 5: *Experiment 2 results*. (**A**) Absolute movement angles as a function of absolute rotation magnitude in the FREE task. (**B**) Mean of median RTs as a function of movement angle (binned by 15˚). (**C**) Circular mean of movement angles as a function of RT (binned in 25 ms bins) in the FORCED task. Gray dots represent RT bins where movements were deemed as primarily random (see Methods). (**D**) Pooled vector representation of data in (**C**). Error bars = 1 s.e.m.

1989). Moreover, the observed linear trend of intermediate movement directions (Fig. 5C, D) may

reflect intermediate states of mental rotation, which confirm a fundamental prediction of

parametric computations (Shepard & Metzler, 1971; Cooper, 1976).

Mental rotation paces are correlated between FREE and FORCED tasks

Our conclusions assume that the FREE and FORCED tasks recruit the same parametric

MR strategy. To test this, we derived a mental rotation pace parameter (ms/deg) from each task.

For the FREE task, this was done by regressing each subjects' execution trial data (Fig. 5B), and,

for the FORCED task, by regressing each subject's full data set above the critical RT (RT X

Movement Angle), as well as fitting a sigmoid to those data (Supplementary Fig. 2). The resulting

regression slopes from the FORCED task were then inverted to match the units of the FREE task

(ms/deg). Linear MR pace was strikingly similar between tasks ($t(31) = 0.19$, $p = 0.85$; Bayes

Factor = 7.81 in favor of the null). Moreover, a significant correlation was found in subjects' mental

rotation paces across tasks ($R^{pearson} = 0.46$; $p = 0.008$; $R^{spearman} = 0.41$; $p = 0.02$), and this

correlation held when using sigmoidal fits on the FORCED data (Supplementary Fig. 2). This

result suggests that the same mental rotation computation is operative in both free and

constrained RT contexts.

We note that average RTs in the FREE task (Fig. 5B) were well above those that produced

equivalent movement directions in the FORCED task (Fig. 5C). We believe this extra computation

time is likely the result of a relatively low sense of urgency in the FREE task compared to the

FORCED task: First, we note that the range of RTs observed in the FREE task is comparable to

the range observed in a similar study (Georgopoulos & Massey, 1987). Second, the regression

analysis on the FREE task (Fig. 5B) revealed no correlation across subjects between the intercept

of the regression, which reflects putative non-rotation RT, and its slope, which reflects the mental

rotation pace ($R^{pearson} = 0.03$; $p = 0.87$; $R^{spearman} = 0.09$; $p = 0.64$). This suggests that the excess

RT in the FREE task is the product of processing unrelated to MR.

Control analyses confirm predictions of a parametric-MR strategy

We now address three alternative explanations for the observed linear rise in mean reach

directions over RT in the FORCED task (Figs. 5C, D): First, because subjects made many random

(i.e. uniformly distributed) movements at low RTs and gradually made correct movements at

higher RTs, a linear trend could appear as an artifact of averaging without any truly purposeful

intermediate movements. However, subjects' most frequent (mode) reach directions also

displayed intermediate values, gradually increasing with RT (Supplementary Fig. 3) and showing

a significant linear trend (p = 0.002). This argues against this averaging confound.

Second, subjects' non-random reaches could be limited to the 0˚ target direction (i.e. a

"prepotent" response; Stroop, 1935) and the 90˚ solution direction, with the linear trend merely

representing a change in the relative frequency of each as RT increases (another form of

averaging artifact). First, we note the mode analysis also argues against this explanation

(Supplementary Fig. 3). Second, the full distribution of reach directions plotted in Figure 6A shows

no clear evidence for a mode at the 0˚ location. A subtle degree of bimodality was indeed

observed — subjects occasionally reached with an approximately correct magnitude of rotation

but with the wrong sign. We interpret these errors as sign flips, which were found in *Experiment*

1 (Fig. 3A) and are predicted by a parametric MR strategy.

Figure 6: *FORCED task full data distribution and model fits*. (**A**) Full distribution of movement angles across all RTs in the FORCED task. (**B**) Fitted mean parameters from the "Free-μ" mixture model (starting at the 150 ms bin) for directed reaches with a correct sign (Von Mises +, purple) and a flipped sign (Von Mises -, green). See Methods for model details. (**C**) PDFs of model fits in each RT bin. Vertical lines represent the fitted mean parameter (Von Mises +) for each corresponding RT bin.

To further examine the distributions of reach angle over the RT bins, we fit the data with
two different mixture models that accounted for random reaches (i.e. modeled with a uniform
distribution) and "directed" reaches (i.e. modeled with two Von Mises distributions, one with a
positive mean and one with a negative mean; see Methods for details). The first model, a "Free-
μ" model, allowed two mean parameters, one positive and one negative, to vary freely over RT
bins (thus capturing mental rotation and sign flips). The second model, a "Fixed-μ" model, fixed
the mean parameters at +90° and -90°, which would be consistent with the *a priori* predictions of
a discrete RC strategy (and swap errors).

As predicted, the mean parameters of the Free-μ model showed evidence for MR (Fig.
6B): Both the positive (purple) and negative (green) mean parameters gradually approached,
respectively, +90° and -90° as RT increased. Linear regressions on the fit mean parameters
revealed a significant positive linear trend for the positive μ parameter ($p < 0.001$), consistent with
MR. Regression on the negative μ parameter, however, revealed no significant trend ($p = 0.38$),
though we reasoned this result may be driven by the two deviant points (both > 1.5 sd from the
mean) in the first two RT bins (Fig. 6B, green). Figure 6C shows the full probability density
functions of the best fit models in each RT bin, with vertical lines representing the positive Von
Mises mean in each bin. We emphasize that by accounting for random movements (with a uniform
distribution), the mixture model offers further support for the mental rotation of directed
movements. Moreover, the apparent rotation of the negative Von Mises mean (Fig. 6B, green)
suggests that sign flips themselves follow MR, albeit it in the erroneous direction. Critically, when
the quality of fit of the Free-μ model is compared with the Fixed-μ model, we find that the Free-μ
provides a far superior fit to the data (ΔAIC = 1189); this large difference comports with MR, as
the Free-μ model can more accurately capture intermediate movements.

Finally, apparent mental rotation could appear as the result of "response substitution"
between a prepotent response directed at the target location gradually becoming inhibited, while
the rotated response gets excited (Cisek & Scott, 1999). In this interpretation, intermediate
movements could occur via the averaging of simultaneously active motor plans (Cisek & Kalaska,
2005), with one plan directed at the target and one at the solution. Critically, this model of
response substitution and our model of MR make distinct predictions concerning movement
speed: Response substitution predicts a slowdown of movement speed on trials where responses
are averaged (Aslin and Shea, 1987), whereas MR does not. Importantly, movements were fast
throughout the FORCED task (μ MT = 128.13 ms; Supplementary Fig. 4), and, in fact, speed

subtly and monotonically increased with RT (Supplementary Fig. 5; *t*-test on regression slopes:
t(31) = 4.28, *p* < 0.001), arguing against response substitution.

The different speed predictions of response substitution versus MR can be captured in the
length of a neural population vector, which is correlated with movement speed (Moran &
Schwartz, 1999). A follow-up computational modeling analysis using two neural models to
formalize different movement speed predictions provides further evidence that the speed data are
compatible with MR but not response substitution (Supplementary Fig. 5). Finally, we note here
that recent research has questioned the premise that intermediate movements represent
involuntary averaging of parallel motor plans (see Discussion; Haith et al., 2015b; Wong and
Haith, 2017).

Experiment 3: Constrained RT in low set size condition reveals a discrete working memory
strategy

Experiment 3 was designed to provide further evidence for a discrete RC strategy: We
used the identical forced-RT task as *Experiment 2,* but with only 2 possible target locations instead
of 12, matching the set size manipulation of *Experiment 1* (Fig. 1). We hypothesized that subjects'
behavior in this context would reveal a bimodal distribution of reaches reflecting cached
responses, but would not show evidence for mental rotation (i.e. via intermediate movements).
(Note, we we did not include a FREE task for the 2-target condition in *Experiment 2,* given that
the FREE task itself requires a large variety of rotation magnitudes and signs to test for a linear
relationship between rotation and RT; under these conditions, a discrete RC strategy is not
plausible given working memory constraints. If we constrained it to a single rotation magnitude
and sign, then it degenerates into the 2T condition of *Experiment 1*.)

Figure 7: *Experiment 3 results*. (**A**) Circular mean of movement angles as a function of RT (binned in 25 ms bins) in the 2-target FORCED task. (**B**) Pooled vector representation of data in (A). Error bars = 1 s.e.m.

With only 2 possible target locations, subjects could fully counteract the rotation with very

short RTs (Fig. 7). Examining the circular means of the movement angles (Fig. 7A) reveals an

abrupt jump from highly variable movements at short RTs (< 225ms) to consistent movements

directed at the rotation solution (90°). Indeed, the pooled vector plot (Fig. 7B) does not show the

several highly-variable intermediate values seen in the circular means (Fig. 7A), suggesting that

those are likely the result of averaging of extreme values (e.g. -90˚ and +90˚; see below for model

fitting). Furthermore, the full histogram of reaching directions (Fig. 8A) is consistent with a bimodal

response distribution: Subjects' responses were concentrated near the solution (+90˚) and its

opposite (-90˚).

Figure 8: *2-target FORCED task full data distribution and model fits*. (**A**) Full distribution of movement angles across all RTs in the 2-target variant of the FORCED task. (**B**) Fitted mean parameters from the "Free-*μ*" mixture model (starting at the 150 ms bin) for directed reaches with a correct sign (Von Mises +, purple) and a flipped sign (Von Mises -, green). See Methods for model details. (**C**) PDFs of model fits in each RT bin.

For our next analysis, we analyzed the mode of pooled movement angles across the RT

bins. Unlike the 12-target condition in *Experiment 2*, we did not find a linear trend in the 2-target

condition (p = 0.20; Supplementary Fig. 3). Next, to better quantify the distribution of reaches in

the 2-target condition, we fit the Free-μ and Fixed-μ mixture models to the pooled data. Critically,

the Von Mises mean parameters for the Free-μ model appeared to saturate immediately, not

displaying the signature of mental rotation (Fig. 8B). In fact, linear regression on the positive mean

parameter revealed a slight negative slope (p = 0.004), which is in the wrong direction for mental

rotation, and regression on the negative mean parameter revealed no significant trend (p = 0.36).

When the quality of fit of the Free-μ model is compared with the Fixed-μ model, we find that the

Fixed-μ model provides a far superior fit to the data (ΔAIC = 1941); this large difference comports

with a discrete RC strategy, as there is no need for shifting mean free parameters in the absence

of intermediate movements.

Finally, we sought to directly compare between the 12-target (*Experiment 2*) and 2-target

(*Experiment 3*) forced-RT conditions. First, we performed serial *t*-tests between circular mean

reaching directions at each RT bin after the critical bin. Means were significantly different at each

bin in the 10th - 13th RT bins (all *p's* < 0.05), spanning the range of RTs between 200 and 300

ms. Next, we performed within-condition *t*-tests at each RT bin to look for significant differences

between the group mean movement angles and the solution at 90˚. In the 12-target condition, all

bins after and including the 7th bin were significantly different from 90˚ (all *p's* < 0.01). In the 2-

target condition, movement angles were not significantly different from 90˚ starting at the 8th bin

onward (all *p's* > 0.05). Finally, we performed a more granular model comparison analysis: Recall

we found that the 12-target condition (*Experiment 2*) was better fit by the Free-μ model, consistent

with using a parametric MR strategy, and the 2-target condition (*Experiment 3*) was better fit by

Fixed-μ model, consistent with a discrete RC strategy. However, for these analyses we summed

AIC values across all model fits after and including the critical (7th) RT bin. Here, we computed

ΔAIC between models at each of these RT bins (Fig. 9). The Free-μ model fit the 12-target group

better at every RT bin (ΔAIC range: 6.92 - 276.13). In contrast, the Fixed-μ model fit the 2-target

group better at every RT bin (ΔAIC range: 15.69 - 345.46). Thus, consistent with our predictions,

subjects in the 12-target condition were likely using a parametric MR strategy whereas subjects

in the 2-target condition were likely using a discrete RC strategy.

Figure 9: *Model comparison*. ΔAIC values over each RT bin between the Free-*μ* and Fixed-*μ* models for the 2-target (grey) and 12-target (black) FORCED-RT experiments. Negative values constitute better fits for the Free-*μ* model, and positive values constitute better fits for the Fixed-*μ* model,

Mental rotation pace from FREE task predicts RT differences during standard visuomotor learning

By intermixing different rotation sizes, the FREE task in *Experiment 2* provided us with a precise estimate of subjects' mental rotation paces when RT was unconstrained (Fig. 5B). If our hypothesis is correct, the average mental rotation pace from the FREE task should correspond to observed RT differences between the 75˚ and 25˚ conditions from *Experiment 1*. Moreover, this correspondence should hold for the 12T conditions at both the beginning and end of learning, but only the beginning of learning in the 2T conditions in *Experiment 1* (Fig 2C). As shown in Figure 10, the regression line predicted by the FREE task was consistent with MR occurring in all subjects early in learning (Fig. 10A; slope difference between FREE task distribution of slopes and early learning regression slope for 12T groups: $t(31) = 1.37$, $p = 0.18$; 2T groups: $t(31) = 1.55$, $p = 0.13$), but only the 12T groups late in learning (Fig. 10B; FREE task slope difference from 12T groups: $t(31) = 0.78$, $p = 0.44$; 2T groups: $t(31) = 6.16$, $p < 0.001$). This result is consistent with a

    set-size-dependent shift during learning from a parametric working memory strategy to a discrete

459   one (Fig. 10C).

Figure 10: *Mental rotation pace from FREE task matches RT data in Experiment 1*. (**A**) Mean of median RTs in early learning from *Experiment 1* (triangle markers) with a regression line, for visualization, on all four groups' data (red), and the regression line predicted by the mental rotation pace gleaned from the FREE task using the same intercept (dashed black). (**B**) Same as (A) on late learning data, with separate regression lines for the 12T groups (red) and 2T groups (blue). (**C**) Schematic describing a parametric to discrete learning shift, which is mediated by set size. Error bars = 1 s.e.m.

DISCUSSION

The role of working memory in motor learning is not well understood, though it is clear that

controlled, deliberative processes are important for learning (Taylor et al., 2014). Here we provide

evidence for at least two working memory strategies that contribute to learning in a visuomotor

rotation task — discrete response caching (RC) and parametric mental rotation (MR). A

parametric MR strategy manifests as a linear relationship between rotation magnitude and RT,

consistent with classic mental rotation (Shepard and Metzler, 1971; Georgopoulos and Massey,

1987; Bhat and Sanes, 1998). In contrast, a discrete RC strategy manifests as something akin to

a look-up table of S-R relationships (Petrides, 1985; Collins and Frank, 2012), consistent with

capacity-limited working memory (Cowan, 2010; Luck & Vogel, 2013).

One key question concerns the relationship between parametric versus discrete working

memory and long term motor skill acquisition. One useful analogy here could be the dissociation

of model-based and model-free reinforcement learning (Daw et al., 2005), where the former relies

on an explicit model of transition probabilities between responses and sensory states, and the

latter merely reinforces directly rewarded actions. One speculation could be that model-based

computations, which we consider to be analogous to parametric motor learning strategies, are

themselves made automatic over time (Economides et al., 2015). This would suggest that the

transition from parametric to discrete strategies we observed (Fig. 10) could represent mental

rotation itself becoming automatic. However, if a complex computation like mental rotation

becomes "cached," it is hard to discern if the computation can still be said to be operative — it

could be that the responses that usually reflect the computation have been reinforced and cached,

and the computation itself is thus no longer needed (Haith and Krakauer, 2018). This kind of

caching could represent an intermediate form of processing that lies somewhere between model-

based planning and model-free habit.

One idea could be that during learning, there is a first transition from model-based

planning (e.g. leveraging a parametric model of the world) to discrete, working memory S-R (Fig.

10C). This form of S-R could then itself be proceduralized into habitual model-free S-R, which

would no longer be dependent on working memory. In this framework, rich representational

models are not transformed or discarded during skill learning, but rather come to be skipped over

during response preparation. To return to the piano learning example proffered in the Introduction,

novice piano students could first learn the parametric strategy (i.e. measuring distances from

middle C) to bootstrap performance quickly. Then, as they acquire the skill, the CNS may

construct a rich look-up table of discrete responses to the variety of notes on the page, and the

model-based strategy could eventually be omitted. Future studies using extended training and

generalization tests could test these hypotheses.

Presumably, it is important that a model-based parametric strategy can still be flexibly

recruited as needed after extensive training, especially in the event that the environment changes.

For example, expert riflemen adeptly use "Kentucky windage" to adapt to the direction and force

of the wind. Curiously, several studies on long-term visual mental rotation (using the task

described in Shepard and Metzler, 1971) find that although practicing mental rotation for weeks

leads to an exponential reduction in overall RT, the mental rotation effect (i.e. RT as a linear

function of rotation magnitude) decreases only subtly and does not go away (Leone et al., 1993;

Wright et al., 2008). Critically, in the longest version of this task (Leone et al., 1993), a set size of

12 different objects and many rotation magnitudes was used, thus making a discrete RC (i.e.,

lookup table) strategy onerous. In a different study with long-term practice of visual mental

rotation, only three stimuli and four rotation angles were used, and mental rotation effects

significantly decreased after extensive training (Tarr and Pinker, 1989). Intriguingly, when novel

orientations were introduced at the end of learning, the mental rotation effects returned,

suggesting a "dual process" model where subjects either retrieve an over-learned object from rote

memory or perform a parametric computation (Tarr and Pinker, 1989; Provost et al., 2013). These

results in the perceptual learning domain overlap considerably with our results in the motor

learning domain, suggesting homologous mechanisms.

We note that the results of our 12-target forced reaction time task (Fig. 5) also have
implications for the study of mental imagery. Our results provide novel evidence supporting a
critical prediction of analog cognitive computations, namely that mental rotation should predictably
proceed through intermediate states (Shepard and Metzler, 1971). Previously, intermediate states
have been implied from RT measures (Cooper, 1976) or neural population statistics
(Georgopoulos et al., 1989), but have not been rendered in overt motor behavior. Our results
suggest that a mental rotation-like operation drives volitional re-planning of a movement goal via
a continuous sweep through direction space at a constrained pace. We note that this mechanism
doesn't require that cells are necessarily directionally-tuned, but could instead represent a low-
dimensional projection of a high-dimensional representation onto the 2D angular space
prescribed by our task (Churchland et al. 2012). Importantly, the aforementioned results in motor
cortex (Georgopoulos et al., 1989) were challenged by an alternative explanation, which posits
that the gradual "averaging" of an initial motor plan at $0°$ with a second plan at $90°$ could give the
appearance of mental rotation (Cisek & Scott, 1999). This response substitution account has been
supported by the behavioral finding that subjects appear to involuntarily average co-active motor
plans in "go-before-you-know" paradigms (Stewart et al., 2014) or "target-jump" saccade tasks
(Aslin and Shea, 1987), where goals are cued late in a trial.
This averaging hypothesis has been challenged by recent findings suggesting that
nominally averaged plans may actually be the product of a decision-making process optimizing a
single movement plan to account for multiple goals (Haith et al., 2015b; Wong and Haith, 2017).
In a recent study, Wong & Haith (2017) used a go-before-you-know task where subjects had to
initiate a movement between two targets before one was cued as the goal. Critically, when moving
relatively slowly (average movement times ~ 600 ms), subjects often reached at an angle between
two competing targets until the goal target was cued, after which they swerved to the correct
choice. However, when moving quickly (average movement times ~ 300 ms), subjects often
moved in a straight line to one of the two options, making no corrections after the cue. These and
other similar results (Haith et al., 2015b) suggest that when making reaching movements between
two competing options, subjects' movement directions may reflect an optimal decision process
that can buy itself time to make the correct decision under uncertainty. Critically, not only did our
FORCED task have a single goal (a $90°$ reach) instead of multiple goals, but, more importantly,
subjects elicited rapid reaching movements (average movement times ~ 130 ms) well below the
movement time threshold for plan averaging (Wong & Haith, 2017). Taken together, our task
design and results suggest that neither obligatory averaging of co-active motor plans (Stewart et

al., 2014) nor strategic plan averaging (Wong & Haith, 2017) were at play in our task. We maintain

that parametric rotation of the motor plan is the most parsimonious explanation.

One trend in our data is the tendency for subjects to "under-rotate" when they were

putatively using a parametric MR strategy. This can be observed in *Experiment 1* (12T groups;

Fig. 2B), and both the FREE and FORCED tasks of *Experiment 2* (Fig. 5). In contrast, when

performing discrete RC, subjects appear to rotate all the way to the solution (2T groups Fig. 2B;

Fig. 7). We do not have a clear result that speaks to this issue. One speculation is that due to the

extra computation time needed for mental rotation, a latent "urgency" signal may drive subjects

to initiate their movement before they're finished rotating, reflecting a kind of within-trial speed-

accuracy trade off. Another detail in our data is that MR in the 12-target FORCED task

(*Experiment 2*) appears to start with a rapid jump followed by a linear rotation; this is most clearly

shown when looking at the model fit (Fig.6B). This suggests that the rotation process may proceed

as a sequence of several discrete steps with varying magnitudes.

In terms of the neural substrates of parametric MR and discrete RC strategies, we discuss

several speculations. For parametric MR, parietal areas — the putative locus of mental rotation

and similar sensorimotor transformations (Kosslyn et al.,1998; Harris & Miniussim, 2003; Bueno

& Andersen, 2006; Zacks, 2008) — could continually feed a shifting movement goal to motor

cortex, which would explain the observed intermediate movements when RT was cut short.

Consistent with this hypothesis, Anguera and colleagues (2010) found overlapping activation in

inferior parietal and dorso-lateral prefrontal cortices during both a visuomotor adaptation task and

a mental rotation task. Discrete RC, on the other hand, being a form of capacity-limited item-

based working memory, could rely on the maintenance of S-R relationships in the dorso-lateral

prefrontal cortex (Curtis & D'Esposito, 2003). Future studies could test these hypotheses using

brain imaging or stimulation techniques.

The dissociation of parametric and discrete memory representations is a theme in several

learning and memory domains. One example is human mathematical reasoning, where there is

a strong behavioral and neural dissociation between numerical computations learned by rote and

then recalled from semantic memory ("two times four = eight"), versus flexible parametric

operations performed on a "mental number line," represented in fronto-pareital spatial working

memory regions (Deheane and Cohen, 1997; Hubbard et al., 2005). Here we present an

analogous dissociation, showing evidence for both the maintenance of discrete motor responses,

and the parametric manipulation of such responses, in working memory. We suggest that working

memory may operate on internally generated movement plans much like it operates on any

number of sensory and abstract representations, with implications for how novel motor skills are
learned.
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METHOD
Participants
122 right-handed subjects (age range 18-34, 57 women) were recruited from the research
participation pool maintained by Princeton University in exchange for course credit or monetary
compensation. Handedness was verified using the Edinburgh handedness inventory (Oldfield,
1971). All subjects participated in accordance with the university's IRB and provided written,
informed consent. In *Experiment 1* (n = 80), 20 subjects were used per condition. A single subject
was excluded for disregarding the key task instruction of attempting to land the cursor on the
target (asymptotic movement error > 6 sd from the mean). We note that the sample size was not
determined by a power analysis, although it is consistent with sample sizes in other studies using
similar tasks (Stewart et al., 2014; Wong and Haith, 2017; McDougle et al., 2017). In *Experiment*
2, a power analysis was used (alpha = 0.95) and revealed that a sample of 19 subjects would
replicate the effect size of a relevant correlation result (d = 0.66; correlation of mental rotation
paces in Pellizzer and Georgopoulos, 1993). To be conservative, we sought to approximately
double it, so we recruited a sample size of 32 (given counterbalancing). This is also in the range
of a previous study that used a similar within-subject design and a somewhat similar task, and it
had a sample size of 26 (Pellizzer & Georgopoulos, 1993). A power analysis was used to
determine the sample size in *Experiment 3*, which aimed to test the regression effect seen in
Experiment 2 (d = 2.60; regression on movement angle X RT). The necessary sample size was
5, but, again to be conservative and counterbalance rotation signs, we recruited a sample of 10.
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General Experimental Procedures and Analysis
In all experiments*,* subjects made rapid, center-out, open-loop reaching movements to
visual targets (5.0 mm radius) using a digitizing tablet, holding on to a digital pen with their hand
in a power grip position and sliding the pen across the tablet (Intuous Pro; Wacom). The task was
controlled by custom software written in MATLAB (Mathworks, Natick, MA; Psychophysics
Toolbox). Hand position was sampled at 140 Hz. Stimuli were shown on a 21.5-inch LCD
computer monitor (Planar), mounted horizontally 25 cm above the tablet, occluding vision of the
hand. A small cursor (2.5 mm radius) provided endpoint feedback after each reach terminated.
Analyses were conducted in MATLAB and R (GNU).

Experiment 1

        Subjects (N = 80) made rapid, 7 cm movements to targets in a blocked reaching task using a 2 X 2 between-subjects design, crossing the factors rotation magnitude and number of targets ("set size"; Fig. 1). For the rotation, two magnitudes were used, ±25˚ and ±75˚, with the sign of the rotation counterbalanced within each condition. For set size, subjects were exposed to either two targets (2T) or 12 targets (12T). In the 12T condition, target locations were pseudorandomly presented at 12 possible locations (0˚, 30˚, 60˚, 90˚, 120˚, 150˚, 180˚, 210˚, 240˚, 270˚, 300˚, 330˚), where the same location was never repeated in consecutive trials. In the 2T condition, two targets were presented, and the locations were randomized across subjects to include one position from the 12 possible locations above and its opposite 180˚ away (Fig. 1). Within a subject, targets were presented pseudorandomly, where the same target was not repeated more than twice in a row. At the start of every trial, to position their hand in a central "start" position (5 mm radius) subjects used a dynamic visual ring that was scaled based on their distance from the start. After holding the start for 500 ms, a visual target appeared and cursor feedback was removed.

        In a baseline block (36 trials), subjects reached to the targets with veridical endpoint feedback. At the start of the subsequent rotation block, the rotation was abruptly applied and was maintained for 300 trials, all with endpoint feedback. Finally, in the washout block, subjects were told to cease any strategy they may have adopted to counter the rotation and reach directly for each target. At the start of the experiment, subjects were told to do their best to land the cursor on the target. If the cursor hit the target, a pleasant chime was sounded, otherwise an aversive buzzer was sounded. In the washout block a neutral sound was played after each reach. Critically, to limit implicit sensorimotor adaptation and better isolate strategic explicit learning, all feedback was delayed by 2 s on every trial (Kitazawa et al., 1995; Brudner et al., 2016; Schween et al., 2017).

        Movement angles were computed as the angle of the hand, relative to the target, when it crossed the invisible target ring of 7cm. All movement angles were rotated to a common reference frame about the 0˚ axis of the unit circle. In all experiments, reaction time (RT) was measured as the time elapsed from target appearance to the hand leaving the start circle (i.e. 5mm from the center of the tablet), and movement time (MT) was measured as the duration of the movement from the end of the RT to the time at which the hand crossed the invisible workspace ring. In analyses with normalized absolute errors, angular errors were scaled between 0 and 1 for each participant. Lastly, trials where RTs exceeded three standard deviations above or below a subject's mean RT (1.30% of trials), or trials where movement angles exceeded three standard

deviations above or below a subject's mean movement angle (0.87% of trials) were excluded from

analyses (the latter exclusions were not performed for the sign error analysis, Fig. 3A).

Experiment 2

Subjects (N = 32) performed two tasks (FREE and FORCED; Fig. 4), with the order of

completion counterbalanced, as well as the rotation sign used in the FORCED task (-90˚ versus

90˚). The FREE task was designed to verify the classic signature of mental rotation, and the

FORCED task was designed to interrupt mental rotation.

FREE Task

At the start of every trial, subjects used continuous cursor feedback to position their hand

in a central "start" position (5 mm radius). After holding for 400 ms, a visual target appeared and

cursor feedback was removed. Participants were instructed to make a rapid, straight shooting

movement to the target. The trial concluded when the hand crossed an invisible ring (8 cm radius),

at which point feedback of the cursor was provided at the location where the hand crossed the

ring. Slow movement times (MT) were discouraged: If MT exceeded a 700ms limit, a "too slow"

warning was delivered visually to the subject by the task software. In line with the instructions,

subjects moved straight and rapidly, with a mean MT of 208.32 ± 8.54 ms (standard error of the

mean).

Subjects were instructed that they would be performing a number of trial "pairs" (Fig. 4A):

In the first trial of each pair, the "learning" trial (Fig. 4A), subjects were instructed to reach directly

at the displayed target and observe where the feedback cursor landed. In learning trials, the target

was blue and appeared in one of four off-cardinal locations (10˚, 100˚, 190˚, 280˚). In the second

trial of the pair, the "execution" trial (Fig. 4A), subjects were told to apply what they learned about

the relationship between their movement and the resultant feedback and attempt to make the

cursor terminate within the target. In execution trials, the target was red and appeared in one of

the three locations in which the learning target did not appear. Target locations were pseudo-

randomized within and between trial pairs. Subjects performed 140 trial pairs. This task was

modeled after a previous study (Pellizzer & Georgopoulos, 1993), but with the distinction that

subjects were not provided with an explicit symbolic cue regarding the exact solution to the

rotation; instead, subjects had to determine the rotation's size and sign themselves, so that our

task echoed canonical visuomotor adaptation paradigms (Krakauer, 2009).

Rotations used in the learning trials ranged from -90˚ to 90˚ by 15˚ intervals. Rotations

were pseudo-randomized throughout the task, and each rotation was seen on 10 different trial

pairs, except for the 0° null rotation, which was seen on 20 trial pairs. Each subject received an
individualized schedule of rotation magnitudes and target locations.

Analyses were conducted on execution trials only, limited to trials where there was an
imposed rotation. Movement angles were computed as the angle of the hand, relative to the
target, when it crossed the invisible target ring. All movement angles were rotated to the same 0°
target axis and matched to a single sign for analysis (see *experiment 1*). The first analysis was a
regression on subjects' reaching angles, relative to the imposed rotation, on execution trials. We
chose to use the absolute reaching angle since on a significant number of trials (8.71%), subjects
approximated the magnitude of the rotation but misinterpreted the sign (i.e. were within 15° of the
"flipped" solution). No significant difference in mean RTs was found between correct and "flipped"
execution trials ($t(31)$ = -0.99, p = 0.33), suggesting that subjects did not hesitate and change
their mind on flipped trials, but simply misremembered the proper clockwise/counterclockwise
direction at the onset of the execution trial (see Results and Discussion for a further discussion of
sign errors).

A one-sample *t*-test was performed on the fitted slopes resulting from the regression of
rotation magnitude and movement angle to test for the presence of a reliable trend. The second
analysis was a regression of each subject's realized movement angles onto corresponding RTs.
A one-sample *t*-test was performed on the fitted slope values to test for the presence of a reliable
trend. Moreover, each subject's slope derived from this regression analysis served as their mental
rotation "pace", in the units of milliseconds per degree (Pellizzer & Georgopoulos, 1993). We also
performed a similar linear regression using the imposed rotation angles as the predictor variable
to confirm that it echoed the regression using the actually-realized movement angles; both
analyses yielded comparable results, though the movement angle regression is likely to be a
better estimate of a subject's true mental rotation pace (Georgopoulos and Massey, 1987).

FORCED Task

The FORCED task utilized a modified forced-response-time paradigm to interrupt putative
mental rotation (Ghez et al., 1997; Haith et al., 2016; Fig. 4B). Like the FREE task, at the start of
every trial subjects used continuous cursor feedback to position their hand in a central "start"
position. Once subjects were positioned in the start, a countdown of four tones was emitted from
the computer speakers (Logitech). The tones were played 600 ms apart, and subjects were
instructed to try and synchronize the initiation of their reach with the fourth tone, effectively
"replacing" that tone (Fig. 4B). The experiment and instructions were deliberately designed to
emphasize early movements over late ones: If the subject initiated their movement > 100 ms after

the fourth tone, the screen was blanked and a "Too Late" message was displayed. Subjects could initiate their movement any time after target appearance, to encourage early RTs. However, if they moved before the target appeared the screen was blanked and a "Too Early" message was displayed. Similar versions of this forced-response-time task have been previously used to study the effect of restricted RTs on visuomotor learning and movement preparation (Ghez et al., 1997; Haith et al., 2016).

Subjects were instructed to prioritize reaching on time, and, secondarily, to try and land the cursor on the displayed target. On each trial, the targets could appear in one of 12 off-cardinal locations (10˚, 40˚, 70˚, 100˚, 130˚, 160˚, 190˚, 220˚, 250˚, 280˚, 310˚, 340˚). Endpoint cursor feedback was presented after the subjects passed the invisible target ring, and the appearance of the feedback was delayed by 500 ms. Like *experiment* 1, this brief delay was added to inhibit implicit adaptation to the imposed rotation and thus isolate the cognitive re-aiming process (Kitazawa et al., 1995; Brudner et al., 2016; Schween et al., 2017). This manipulation was successful, yielding subtle, though significant aftereffects (μ = 1.98˚; $t(31)$ = 2.06, p = 0.048).

The moment of target appearance was titrated such that subjects had varying amounts of time with which to compute the target location, plan, and execute their movements. These windows reflected the time elapsed between target appearance and the fixed moment of the fourth tone. Seventeen test RT windows were used, ranging from 200 - 600 ms by 25 ms intervals. An asymptotic RT window was also used, giving subjects up to 1200 ms to react. We reiterate that because subjects could move early in our variant of this task (i.e. any time after target appearance), the windows acted as "guides" rather than being perfectly predictive of subjects' realized RTs (i.e. subjects tended to move before they needed to; see Results). Finally, catch trials of 0 ms were used to ensure that subjects stayed on task and executed movements on time even if they had not perceived the target yet.

Subjects performed the FORCED task in three blocks. To get accustomed to the task, in the first block subjects performed 64 baseline trials with veridical endpoint cursor feedback, with pseudo-randomized RT windows and target locations. In the subsequent rotation block, subjects performed 624 trials, with 35 trials at each of the test RT windows, 14 trials at the asymptotic RT window, and 15 catch trials. For the rotation block, a fixed rotation of 90˚ (or -90˚, for counterbalancing) was imposed on the cursor. Subjects were thoroughly educated about the rotation before this block began. Subjects were told to try and counter the rotation and land the cursor on the target every trial. Due to the difficulty of the task, subjects were encouraged by a monetary bonus of up to $5 based on their performance in the rotation block, and were informed that missed trials (reaching too late or too slowly) would count against their performance score.

Finally, in the aftereffect block, subjects performed 48 trials with a 1200 ms RT window, pseudo-
randomized target locations, and no cursor feedback. Subjects were told to reach directly for the
target on every trial of the aftereffect block. These data were later compared to the baseline data
to get an estimate of any implicit adaptation.

All analyses were conducted on trials where subjects reached on time (78.90% of trials).
Like *Experiment 1* and the FREE task, movement angles were computed as the angle of the hand
relative to the target when it crossed the invisible target ring. Importantly, subjects followed the
instructions and made straight shooting movements: Movement times were rapid (μ = 128.13 ms)
and no feedback was provided during the reach, which helped to ensure movements were
straight. Movement speed was computed by taking the average of the derivative of hand position
from the time subjects left the start circle to the time they crossed the invisible target ring. (Due to
the rapid "shooting" movement required by the task, average speed was used instead of peak
speed to reduce noise in the estimate as subjects often reached peak speed after passing the
target; we note that this particular approximation of speed did not influence the main results of
the analyses related to movement speed).

Our first analysis involved investigating subjects' reach angles as a function of RT. RT
bins were taken every 25 ms, from 0 ms through 400 ms, with the final bin including all RTs above
400 ms. Similar to previous results (Haith et al., 2016), in catch trials (0 ms RT window) subjects
prioritized the timing demands of the task when no target appeared and moved relatively randomly
around the circle, with some slight biases. A single "critical RT bin" was computed, after which
reaches were determined to be primarily non-random (Supplementary Fig. 1). Linear regressions
were performed on subjects' full distribution (i.e. no binning) of reach angles and RTs after the
critical RT bin, and a *t*-test was performed on the resulting slopes to test for the presence of a
significant trend. These slopes were used as the mental rotation "pace" parameters in further
analyses.

Mental rotation paces were compared between the FREE and FORCED tasks in four
ways: First, a one-sample *t*-test was performed to show that the null could not be rejected.
Second, a Bayes Factor was computed on the resulting one-sample *t*-value using the JZS method
(Rouder et al., 2009) to quantify evidence for the null. Third, both parametric (Pearson) and non-
parametric (Spearman) correlations were conducted between the values to test for a significant
relationship that is robust to outliers (both metrics were used for all correlations in this study). To
confirm the robustness of this correlation, a secondary analysis that involved fitting a sigmoid to
the data was used as an alternative method for extracting the pace parameter in the FORCED
task. Lastly, two supplementary modeling analyses were conducted to test alternative

interpretations, involving both a mixture model analysis (see below; Fig. 6), and a neural model

which modeled movement speed and direction using cosine-tuned units (Supplementary Fig. 5).

Experiment 3

Subjects (N = 10) performed the FORCED task used in *Experiment 2* (Fig. 4B), with one

critical difference: Only two target locations were used, where one was drawn from one of the 12

possible locations of *Experiment 2* (FORCED), and the other target was its 180˚ counterpart. The

particular pair of targets used, and the sign of the 90˚ rotation, were counterbalanced across

subjects. All task instructions and analyses matched those described in *Experiment 2*. MTs were

similarly rapid (μ = 111.25 ms), and aftereffects were similarly small (μ = 3.17˚) but did not reach

significance ($t(9)$ = 2.12, p = 0.06). For comparison purposes, the same "critical" RT bin was used

in this experiment as in *Experiment 2* (RTs > 150 ms; see above).

Mixture Model

A mixture model was used to characterize data in the FORCED tasks from *Experiments*

2 and 3. In this analysis, we modeled reach data as a mixture of two circular normal (Von Mises)

distributions representing both positive (correct) and negative (flipped) directed reaches, and a

single uniform distribution that represented random reaches. The model had probability density

function,

$$P = w_1 \frac{e^{\kappa \cos(x - \mu_1)}}{2\pi I_0(\kappa)} + w_2 \frac{e^{\kappa \cos(x - \mu_2)}}{2\pi I_0(\kappa)} + (1 - w_1 - w_2)U(-\pi, \pi)$$

where I_0 is the modified Bessel function of order 0, w is the weight given to each distribution (w_1 +

$w_2 \le 1$), κ is the concentration parameter (which we fit as a single parameter between the two Von

Mises pdfs), and $U(-\pi, \pi)$ is the uniform probability density function of the unit circle. The mean

parameter, μ, represents the mean of a given Von Mises, including one mean for positively signed

reaches and one for negatively signed reaches (μ_1 > 0, $\mu_2 \le 0$). Data were pooled across subjects

in each RT bin and parameters were optimized separately in each bin using maximum likelihood

estimation, minimizing the negative log likelihood with the MATLAB function *fmincon*. Fifty

randomized starting parameter values were used for each fit to avoid local minima, and Aikake

Information Criterion values (AICs) were wither summed over all RT bins after the critical 7th bin

or compared separately at each bin for model comparisons. Two models were compared: In the

"Free μ" model, both μ parameters could vary freely. In the "Fixed μ" model, μ parameters were

fixed at our *a priori* prediction of μ_1 = +90˚ and μ_2 = -90˚. Thus, the Fixed *µ* model had two fewer

free parameters than the Free *µ* model.

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