Genomic analysis of European Drosophila populations reveals

longitudinal structure and continent-wide selection

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Abstract

Genetic variation is the fuel of evolution. However, analyzing evolutionary dynamics in natural populations is challenging, sequencing of entire populations remains costly and comprehensive sampling logistically difficult. To tackle this issue and to define relevant spatial and temporal scales of variation, we have founded the European *Drosophila* Population Genomics Consortium (*DrosEU*). Here we present the first analysis of 48 *D. melanogaster* population samples collected across Europe in 2014. Our analysis uncovers novel patterns of variation at multiple levels: genome-wide neutral SNPs, mtDNA haplotypes, inversions, and TEs showing previously cryptic longitudinal population structure; signatures of selective sweeps shared among populations; presumably adaptive clines in inversions; and geographic variation in TEs. Additionally, we document highly variable microbiota and identify several new *Drosophila* viruses. Our study reveals novel aspects of the population biology of *D. melanogaster* and illustrates the power of extensive sampling and pooled sequencing of populations on a continent-wide scale.

Keywords: *Drosophila*, population genomics, demography, selection, clines, SNPs, structural variants, symbionts.

Introduction

Genetic variation is the raw material for evolutionary change. Understanding the processes that create and maintain variation in natural populations remains a fundamental goal in evolutionary biology. The identification of patterns of genetic variation within and among

taxa (Dobzhansky 1970; Lewontin 1974; Kreitman 1983; Kimura 1984; Hudson *et al.* 1987;

McDonald & Kreitman 1991; e.g., Adrian & Comeron 2013) provides fundamental insights

into the action of various evolutionary forces. Historically, due to technological constraints,

studies of genetic variation were limited to single loci or small genomic regions and to

static sampling of small numbers of individuals from natural populations. The development

of population genomics has extended such analyses to patterns of variation on a genome-

wide scale (e.g., Black *et al.* 2001; Jorde *et al.* 2001; Luikart *et al.* 2003; Begun *et al.* 2007;

Sella *et al.* 2009; Charlesworth 2010; Casillas & Barbadilla 2017). This has resulted in

fundamental advances in our understanding of historical and contemporaneous

evolutionary dynamics in natural populations (e.g., Sella *et al.* 2009; Hohenlohe *et al.*

2010; Cheng *et al.* 2012; Fabian *et al.* 2012; Pool *et al.* 2012; Messer & Petrov 2013;

Ellegren 2014; Harpur *et al.* 2014; Kapun *et al.* 2014; Bergland *et al.* 2014; Charlesworth

2015; Zanini *et al.* 2015; Kapun *et al.* 2016a; Casillas & Barbadilla 2017).

However, large-scale sampling and genome sequencing of entire populations remains

largely prohibitive in terms of sequencing costs and labor-intensive sample collection,

limiting the number of populations that can be analyzed. Evolution is a highly dynamic

process across a variety of spatial scales in many taxa; thus, to generate a comprehensive

context for population genomic analyses, it is essential to define the appropriate spatial

scales of analysis, from meters to thousands of kilometers (Levins 1968; Endler 1977;

Richardson *et al.* 2014). Furthermore, one-time sampling of natural populations provides

only a static view of patterns of genetic variation. Allele frequency changes can be highly

dynamic even across very short timescales (e.g., Umina *et al.* 2005; Bergland *et al.* 2014;

Behrman *et al.* 2018), and theoretical work suggests that such temporal dynamics may be

an important yet largely understudied mechanism by which genetic variation is maintained

(Wittmann *et al.* 2017). It is thus essential to define the relevant spatio-temporal scales for

sampling and population genomic analyses accordingly.

To generate a population genomic framework that can deliver appropriate high-resolution

sampling and to provide a unique resource to the research community, we formed the

European *Drosophila* Population Genomics Consortium (*DrosEU*; https://droseu.net). Our

primary objective is to utilize the strengths of this consortium to extensively sample and

sequence European populations of *Drosophila melanogaster* on a continent-wide scale

and across distinct timescales. In close cooperation with a complementary effort focused

on North American populations, the *Drosophila* Real Time Evolution Consortium (*Dros-*

RTEC; http://web.sas.upenn.edu/paul-schmidt-lab/dros-rtec/), our long-term goal is to

define the appropriate spatio-temporal scales at which populations should be sampled and

analyzed and to gain novel insights into the dynamics of genetic variation.

D. melanogaster offers several advantages for such a concerted sampling and analysis

effort: a relatively small genome, a broad geographic range, a multivoltine life history that

allows sampling across generations over short timescales, ease of sampling natural

populations using standardized techniques, an extensive research community and a well-

developed context for population genomic analysis (Powell 1997; Keller 2007; Hales *et al.*

2015). The species is native to sub-Saharan Africa and has subsequently expanded its

range into novel habitats in Europe over the last 10,000-15,000 years and in North

America and Australia in the last several hundred years (e.g., Lachaise *et al.* 1988; David

& Capy 1988; Keller 2007). On both the North American and Australian continents, the

prevalence of latitudinal clines in frequencies of alleles (e.g., Schmidt & Paaby 2008;

Turner *et al.* 2008; Kolaczkowski *et al.* 2011b; Fabian *et al.* 2012; Bergland *et al.* 2014;

Machado *et al.* 2016; Kapun *et al.* 2016a), structural variants such as chromosomal

inversions (Mettler *et al.* 1977; Voelker *et al.* 1978; Knibb *et al.* 1981; Knibb 1982; 1986;

Anderson *et al.* 1991; Rako *et al.* 2006; Kapun *et al.* 2014; Rane *et al.* 2015; Kapun *et al.*

2016a) and transposable elements (TEs) (Boussy *et al.* 1998; González *et al.* 2008; 2010),

as well as complex phenotypes (de Jong & Bochdanovits 2003; Schmidt & Paaby 2008;

Schmidt *et al.* 2008; Flatt *et al.* 2013; Adrion *et al.* 2015 and references therein; Kapun *et*

al. 2016b; Behrman *et al.* 2018) have been interpreted to result from local adaptation to

environmental factors that co-vary with latitude or as the legacy of an out-of-Africa

dispersal history. However, sampling across these latitudinal gradients has not been

replicated outside of a single transect on the east coasts of both continents. The observed

latitudinal clines on the east coasts of North America and Australia may have been

generated, at least in part, by demography and differential colonization histories of

populations at high and low latitudes (Bergland *et al.* 2016). In North America, for example,

temperate populations appear to be largely of European origin, whereas low-latitude

populations show evidence of greater admixture from ancestral African populations and

the Caribbean (Caracristi & Schlötterer 2003; Yukilevich & True 2008a; b; Duchen *et al.*

2013; Kao *et al.* 2015; Bergland *et al.* 2016). More intensive sampling and analysis of both

African as well as European populations is thus essential to disentangling the relative

importance of local adaptation versus colonization history and demography in generating

the clinal patterns that have been widely observed. While there has been a great deal of

progress in the analysis of ancestral African populations (e.g., Begun & Aquadro 1993;

Corbett-Detig & Hartl 2012; Pool *et al.* 2012; Fabian *et al.* 2015; Lack *et al.* 2015; 2016),

the European continent remains largely uncharacterized at the population genomic level

(Božičević *et al.* 2016; Pool *et al.* 2016; Mateo *et al.* 2018).

Here, we present the first analysis of the *DrosEU* pool-sequencing data from a set of 48

European population samples collected in 2014. Our main focus is on describing spatial

variation across the European continent. A similar consortium has been organized mainly

in the United States, the *Dros-RTEC* consortium. While the two consortia share the

common goal of widespread and coordinate sampling, the *Dros-RTEC* consortium

concentrates on seasonal dynamics in North American populations (Machado *et al.* 2018).

We examine the 2014 *DrosEU* data at three levels: (1) patterns of variation at single-

nucleotide polymorphisms (SNPs) in the nuclear ($\sim$5.5 x 10^6 SNPs) and mitochondrial

(mtDNA) genomes; (2) variation in copy number of transposable elements (TEs); (3)

cosmopolitan chromosomal inversions previously associated with climate adaptation; and

(4) variation among populations in microbiota, including endosymbionts, bacteria, and

viruses (Figure 1).

Figure 1. The conceptual framework of the *DrosEU* consortium. By intensive spatio-temporal sampling of natural populations of *Drosophila melanogaster*, the European *Drosophila* Population Genomics Consortium (*DrosEU*; http://droseu.net/), aims to uncover the factors that shape the evolutionary dynamics of this exemplary model organism. Each of the repeatedly and consistently sampled *DrosEU* populations is subject to evolutionary forces ("Evolution axis", from neutral to adaptive evolution) in interaction with the environment ("Environment axis", from local aspects to global patterns, including spatial factors). In addition, there are several dimensions along which the fly genomes can be studied, from single SNPs and genes to structural variants and co-evolving genomes ("Genomics axis"), both over short and long timescales ("Timescale axis").

We find that European populations of *D. melanogaster* exhibit novel patterns of variation at all levels investigated: neutral SNPs in the nuclear genome and mtDNA haplotypes that

reveal previously unknown longitudinal population structure; genomic regions consistent

with selective sweeps that indicate selection on a continent-wide scale; new evidence for

inversion clines in Europe; and spatio-temporal variation in TEs frequencies. We also

identify four new DNA viruses and for the first time assemble the complete genome of a

fifth. These novel features are revealed by the comprehensive magnitude of our

coordinated sampling, thus demonstrating the utility of this approach.

Together with other large-scale genomic datasets for *D. melanogaster* (Casillas &

Barbadilla 2017) our data provide a rich and powerful community resource for studies of

molecular population genetics. Importantly, the *DrosEU* dataset represents the first

comprehensive characterization of genetic variation in *D. melanogaster* on the European

continent and might yield important insights into how this species has adapted to

temperate climates after its migration out of Africa.

Results

As part of the *DrosEU* effort, we collected and sequenced 48 population samples of *D.*

melanogaster from 32 geographical locations across Europe in 2014 (Table 1; Figure 2

and Figure 3A).

Figure 2. The geographic distribution of population samples. The map shows the geographic locations of all samples in the 2014 *DrosEU* dataset. The color of the circles indicates the sampling season for each location (see Table 1 and Supplemental Table 1). Note that some of the 12 Ukranian locations overlap in the map.

Figure 3. Sampling and data analysis pipeline. The schematic diagram shows the workflow of data collection and processing (A) followed by bioinformatic approaches used for quality assessment and read mapping (B) as well as the downstream analyses (C) conducted in this study (see Materials and Methods for further information).

While our analyses focus on spatial patterns, thirteen of the 32 locations were sampled

repeatedly during the year (at least twice, once in summer and once in fall), allowing a

first, crude analysis of seasonal changes in allele frequencies on a genome-wide level

(Figure 2). For an extensive analysis of temporal (seasonal) patterns in mainly North

American populations see the companion paper by Machado *et al.* (2018). All 48 samples

were sequenced to high coverage, with a mean coverage per population of >50x (Table

S1 and Figure 4).

Figure 4. Chromosome-wide average coverages. Barplot showing the distribution of chromosome-wide coverages for

X-chromosomes (black) and autosomes (grey). The error bars represent the standard deviation of coverages across all

autosomal arms (*2L, 2R, 3L, 3R*).

Using this high-quality dataset, we performed the first comprehensive, continent-wide

genomic analysis of European *D. melanogaster* populations (Figure 3). In addition to

nuclear SNPs, we also investigated variation in mtDNA, TE insertions, chromosomal

inversion polymorphisms, and the *Drosophila*-associated microbiome (Figure 3).

Most SNPs are widespread throughout Europe

We identified a total of 5,558,241 "high confidence" SNPs with frequencies > 0.1% across

all 48 samples (Figure 3B, Table S1 and S2). Of these, 17% (941,080) were shared

among all samples, whereas 62% were polymorphic in fewer than 50% of the samples

(Figure 5A).

Figure 5 with 2 supplements. **SNP sharing.** (A) Shared SNPs among *DrosEU* samples. Number and proportion of

SNPs in different samples, ranging from one specific sample to being shared among all 48 samples. (B) Shared SNPs

among three worldwide populations. Elliptic Venn diagram showing the number and proportion of SNPs overlapping

among the 5,361,256 biallelic SNPs in *DrosEU* (Europe), 3,953,804 biallelic SNPs in *DGRP* (North America) and

4,643,511 biallelic SNPs in Zambia (Africa) populations.

Figure 5 - figure supplement 1. SNP sharing among *DrosEU* samples. Number and proportion of SNPs in different

33 samples, each from one location. In the locations with more than one sample we took (A) the earliest sample or (B)

the latest of the season.

Figure 5 - figure supplement 2. SNP sharing with other populations by frequency bins. Elliptic Venn-diagram

showing the proportion of SNPs shared and specific to each population, *DrosEU* (Europe), *DGRP* (North America) and

Zambia (Africa) for different frequency bins: (A) frequency ≥ 0.5, (B) frequency ≥ 0.1 and < 0.5 and (C) frequency < 0.1.

For proportions <1%, the number of SNPs is not depicted; *DrosEU*-Zambia: 4,088 SNPs, *DrosEU-DGRP*: 24,049 SNPs,

DrosEU-DGRP-Zambia: 3,627 SNPs.

Due to our filtering scheme, SNPs that are private or nearly private to a sample will be

recovered only if they are at a substantial frequency in that sample (~5%). In fact, only a

small proportion of SNPs (1% = 3,645) was found in fewer than 10% of the samples, and

only 0.004% (210) were specific to a single sample (Figure 5A). To avoid an excess

contribution of SNPs from populations with multiple (seasonal) sampling, we repeated the

analysis by considering only the earliest (Figure 5 - figure supplement 1A) or the latest

(Figure 5 - figure supplement 1B) sample from populations with seasonal data. We

observed similar patterns across the three analyses: (i) a very small number of sample-

specific, private SNPs (210, 527 and 455, respectively), (ii) a majority of SNPs shared

among 20% to 40% of the samples (53%, 52% and 52%, respectively), and (iii) a

substantial proportion shared among all samples (17%, 20% and 19%, respectively; Figure

5A and Figure 5 - figure supplement 1). These results suggest that most SNPs are

geographically widespread in Europe and that genetic differentiation among populations is

moderate, consistent with high levels of gene flow across the European continent.

Derived European and North American populations share more SNPs with each

other than they do with an ancestral African population

D. melanogaster originated in sub-Saharan Africa, migrated to Europe ~10,000-15,000

years ago, and subsequently colonized the rest of the world, including North America and

Australia ~150 years ago (Lachaise *et al.* 1988; David & Capy 1988; Keller 2007). To

search for genetic signatures of this shared history, we investigated the amount of allele

sharing between African, European, and North American populations. We compared our

SNP set to two published datasets, one from Zambia in sub-Saharan Africa (*DPGP3*; Lack

et al. 2015) and one from North Carolina in North America (*DGRP*; Huang *et al.* 2014).

Populations from Zambia inhabit the ancestral geographical range of *D. melanogaster*

(Pool *et al.* 2012; Lack *et al.* 2015); North American populations are thought to be derived

from European populations, with some degree of admixture from African populations,

particularly in the southern United States and the Caribbean (Caracristi & Schlötterer

2003; Yukilevich & True 2008a; b; Yukilevich *et al.* 2010; Duchen *et al.* 2013; Kao *et al.*

2015; Pool 2015; Bergland *et al.* 2016). The population from North Carolina exhibits

primarily European ancestry, with ~15% admixture from Africa (Bergland *et al.* 2016).

Approximately 10% of the SNPs (~1 million) were shared among all three datasets (Figure

5B). Since the out-of-Africa range expansion and the subsequent colonization of the North

America continent by European (and to a lesser degree African) ancestors was likely

accompanied by founder effects, leading to a loss of African alleles, and adaptation to

temperate climates (Mettler *et al.* 1977), we predicted that a relatively high proportion of

SNPs would be shared between Europe and North America. As expected, the proportion

of shared SNPs was higher between Europe and North America (22%) than between

either Europe or North America and Zambia (11% and 13%, respectively; Figure 5B).

When we analyzed SNPs in variant frequency bins, the proportion of SNPs shared across

at least two continents increased from 26% to 41% for SNPs, with variant frequencies

larger than 50% (Figure 5 - figure supplement 2A). In contrast, only 6% of the SNPs at low

frequency (<10%; Figure 5 - figure supplement 2C) were shared. These results are

consistent with the loss of low-frequency variants during the colonization of the European

continent; they suggest that intermediate frequency alleles are more likely to be ancestral

and thus shared across broad geographic scales. Interestingly, as compared to Africa and

North America, we identified nearly 3 million private SNPs that are specific to Europe

(Figure 5B). Given that North American and Australian populations are – at least partly – of

European ancestry (see Lemeunier & Aulard 1992 for more details), future analysis of our

data may be able to shed light on the demography and adaptation of these derived

populations.

European and other derived populations exhibit similar amounts of genetic variation

Next, we estimated genome-wide levels of nucleotide diversity within the European

population samples using population genetic summary statistics. Pairwise nucleotide

diversity (π and Watterson's θ), corrected for pooling (Stalker 1976; Mettler *et al.* 1977;

Voelker *et al.* 1978; Stalker 1980; Sezgin *et al.* 2004), ranged from 0.0047 to 0.0057 and

from 0.0045 to 0.0064, respectively (Figure 6 and Figure 7), with our estimates being

qualitatively similar to those from non-African *D. melanogaster* populations sequenced as

individuals (Knibb *et al.* 1981; Knibb 1982; Anderson *et al.* 1987) or as pools (Inoue &

Watanabe 1979; Inoue *et al.* 1984).

Figure 6. Chromosome-wide average Tajima's π. Barplot showing the distribution of chromosome-wide estimates of

Tajima's *π* for *X-chromosomes (*black) and autosomes (grey). The error bars represent the standard deviation of π

across all autosomal arms (*2L, 2R, 3L, 3R*).

Figure 7. Chromosome-wide average Watterson's *θ*. Barplot showing the distribution of chromosome-wide estimates

of Watterson's *θ* for *X-c*hromosomes (black) and autosomes (grey). The error bars represent the standard deviation of *θ*

across all autosomal arms (*2L, 2R, 3L, 3R*).

Figure 8. Genetic variation in worldwide samples. Barplot showing the distribution of genome-wide estimates of Tajima's π of the *DrosEU* and other genomic datasets (see Materials and Methods for more details) The error bars in the *DrosEU* dataset represent the standard deviation of π across all 48 population samples.

Estimates of π were slightly lower than, but in close agreement with, estimates of θ, leading to a slightly negative average of Tajima's *D* (Das & Singh 1990; 1991; Singh & Das 1992; Singh 2018). Due to our SNP calling approach (see Materials and Methods), we found a deficiency of alleles with frequencies ≤ 0.01, both in the sample-wise site frequency spectra (SFS) as well as in the combined SFS by SNP type, with the sample-wise SFS being skewed towards low frequency variants (Figure 9A).

Figure 9. Site frequency spectra (SFS). (A) Chromosome-wise SFS: The five panels show the means (dark blue line) and standard deviations (light blue area) of the folded SFS across all 48 samples. The red dashed line indicates the expected SFS based on a θ of 10^{-4}. (B) SFS by SNP type: The two panels show the means (dark line) and standard deviations (light area) of the folded SFS for different SNP types (intergenic, intronic, non-synonymous and synonymous), combining SNPs across the autosomes (left) and the *X*-chromosome (right).

In addition, we observed an excess of low-frequency SNPs at non-synonymous sites as compared to other types of sites, which is consistent with purifying selection eliminating

deleterious non-synonymous mutations (Endler 1977).

Overall, we detected only minor differences in the amount of genetic variation among

populations. Specifically, genome-wide π ranged from 0.005 (Yalta, Ukraine) to 0.006

(Chalet à Gobet, Switzerland) for autosomes, and from 0.003 (Odesa, Ukraine) to 0.0035

(Chalet à Gobet, Switzerland) for the X chromosome (Table S1 and Figure 6). When

testing for associations between geographic variables and genome-wide average levels of

genetic variation, we found that both π and θ were strongly negatively correlated with

altitude, but neither was correlated with latitude or longitude (Table 2). There were no

correlations between the season in which the samples were collected and levels of

average genome-wide genetic variation as measured by π and θ (Table 2).

The X chromosome showed markedly lower genetic variation than the autosomes, with the

ratio of X-linked to autosomal variation (π_X/π_A) ranging from 0.53 to 0.66. These values

are well below the ratio of 0.75 (one-sample Wilcoxon rank test, p < 0.001) expected under

standard neutrality and equal sex ratio, but are consistent with previous findings for

European populations of *D. melanogaster* and can be attributed to either selection (Knibb

et al. 1981) or changes in population size (Knibb *et al.* 1981). This pattern is consistent

with previous estimates of relatively low X-linked diversity for European (Kapun *et al.*

2014) and other non-African populations (Kapun *et al.* 2016a). Interestingly, the ratio π_X/π_A

was significantly, albeit weakly, positively correlated with latitude (Spearman's r = 0.315, p

= 0.0289), with northern populations having slightly higher X/A ratios than southern

populations. This is at odds with the prediction of periodically bottlenecked populations

leading to a lower X/A ratio in the north and perhaps reflects more complex demographic

scenarios (Mettler *et al.* 1977; Voelker *et al.* 1978; Knibb *et al.* 1981; Knibb 1982; Das &

Singh 1991; Van 't Land *et al.* 2000; de Jong & Bochdanovits 2003; Anderson *et al.* 2005;

Umina *et al.* 2005; Rako *et al.* 2006; Kapun *et al.* 2014; 2016a).

In contrast to π and θ, we observed major differences in the genome-wide averages of

Tajima's *D* among samples (Figure 10).

Figure 10. Chromosome-wide average Tajima's *D*. Barplot showing the distribution of chromosome-wide estimates of

Tajima's *D* for *X*-chromosomes (black) and autosomes (grey). The error bars represent the standard deviation of *D*

across all autosomal arms (*2L, 2R, 3L, 3R*).

The chromosome-wide Tajima's *D* was negative in approximately half of all samples and

close to zero or slightly positive in the remaining samples, possibly due to heterogeneity in

the proportion of sequencing errors among the multiplexed sequencing runs. However,

models that included sequence run as a covariate did not explain more of the observed

variance than models without the covariate, suggesting that associations of π and θ with

geographic variables were not confounded by sequencing heterogeneity (see Supporting

Information; Table S4). Moreover, our results for *π*, *θ* and *D* are unlikely to be confounded

by spatio-temporal autocorrelations: after accounting for similarity among spatial neighbors

(Moran's *I* ≈ 0, *p* > 0.05 for all tests), there were no significant residual autocorrelations

among samples for these estimators.

Genetic variation was not distributed homogeneously across the genome. Both *π* and *θ*

were markedly reduced close to centromeric and telomeric regions (Figure 11), which is in

good agreement with previous studies reporting systematic reductions in genetic variation

in regions with reduced recombination (Kennison 2008).

Figure 11 with 1 supplement. **Genome-wide estimates of genetic diversity and recombination rates**. The distribution

of Tajima's *π*, Watterson's *θ* and Tajima's *D* (from top to bottom) in 200 kb non-overlapping windows plotted for each

chromosomal arm separately. Bold black lines depict statistics which were averaged across all 48 samples and the

upper and lower grey area show the corresponding standard deviations for each window. Red dashed lines highlight the

vertical position of a zero value. The bottom row shows log-transformed recombination rates (*r*) in 100kb non-overlapping

windows as obtained from Comeron et al. (2010).

Figure 11 - figure supplement 1. Correlation between recombination and genetic diversity. Smooth local regression (LOESS) between Comeron et al. (2012) recombination rate in cM/Mb and the average of the 48 samples genetic diversity (π) in 100kb non-overlapping windows by chromosome arm.

Consistent with this, we detected strong correlations with estimates of recombination rates

based on the data of Comeron *et al.* (2012) (linear regression, $p < 0.001$; not accounting

for autocorrelation), suggesting that the distribution of genome-wide genetic variation is

strongly influenced by the recombination landscape (Table S5). For autosomes, fine-scale

recombination rates explained 41-47% of the variation in π, whereas broad-scale

recombination rates (Roberts 1998; Pimpinelli *et al.* 2010) explained 50-56% of the

variation in diversity. We obtained similar results for *X*-chromosomes, with recombination

rates explaining 31-38% (Dobzhansky & Sturtevant 1938; Kunze-Mühl & Müller 1957;

Ashburner & Lemeunier 1976) or 24-33% (Wesley & Eanes 1994; Andolfatto *et al.* 1999;

Matzkin *et al.* 2005; Corbett-Detig *et al.* 2012) of the variation (Figure 11, Table S5, Figure

11 - figure supplement 1).

We also observed variation in Tajima's D with respect to genomic position (Figure 11).

Notably, Tajima's D was markedly lower than the corresponding chromosome-wide

average in the proximity of telomeric and centromeric regions on all chromosomal arms.

These patterns possibly reflect purifying selection or selective sweeps close to

heterochromatic regions (Navarro & Faria 2014; Kapun *et al.* 2014; 2016a), or might

alternatively be a result of sequencing errors having a stronger effect on the SFS in low

SNP density regions.

Localized reductions in Tajima's D are consistent with selective sweeps

We identified 144 genomic locations on the autosomes with non-zero recombination,

reduced genetic variation, and a local reduction in Tajima's D (see Methods, Table S6),

which jointly may be indicative of selective sweeps. Although we cannot rule out that these

patterns are the result of non-selective demographic effects (e.g., bottlenecks), two

observations suggest that at least some of these regions are affected by positive selection.

First, bottlenecks are typically expected to cause genome-wide, non-localized reductions

in Tajima's D. Second, several of the genomic regions in our data coincide with previously

identified, well-supported selective sweeps in the proximity of *Hen1*, *Cyp6g1* (Andolfatto *et*

al. 1999; Corbett-Detig & Hartl 2012), *wapl* (Andolfatto *et al.* 1999; Matzkin *et al.* 2005;

Kennington *et al.* 2007; Corbett-Detig & Hartl 2012; Kennington & Hoffmann 2013; Kapun

et al. 2014; 2016a), *HDAC6* (Begun 2015; Lavington & Kern 2017), and around the

chimeric gene *CR18217* (Kirkpatrick 2010).

Figure 12 with 3 supplements. **Signals of selective sweeps.** The central figure shows the distribution of Tajima's *D* in

50 kb sliding windows with 40 kb overlap. The red and green dashed lines show Tajima's *D* = 0 and Tajima's *D* = -1,

respectively. The top panel magnifies a genomic region on chromosomal arm *2R* that harbors well-known candidate loci

for pesticide resistance, *Cyp6g1* and *Hen1* (highlighted in red), where strong selection resulted in a selective sweep. This

sweep is characterized by an excess of low-frequency SNP variants, indicated by an overall negative Tajima's *D* in all

samples. The colored solid lines depict Tajima's *D* for each sample separately, whereas the black dashed line shows

Tajima's *D* averaged across all samples. (A legend for the color codes of the samples can be found in the Supporting

Information file in Figure 12 - figure supplement 3). The bottom figure shows a genomic region on *3L* which has not been

previously identified as a potential target of selection but shows Tajima's *D* patterns similar to the top figure. Notably,

both regions are also characterized by a strong reduction of genetic variation (Figure 12 - figure supplement 1).

Figure 12 - figure supplement 1: Genetic variation in regions of putative selective sweeps. This figure is equivalent

to Figure 12 in the main text but shows the distribution of genetic variation *(π)* in regions with depressed Tajima's *D*

around the well-studied *Cyp6g1* locus (A) and a previously unknown candidate region on *3L* (B). Similar to Tajima's *D, π*

was calculated in 50 kb sliding windows with 40 kb overlap. See Table S6 for more examples. A legend for the color

codes of the samples can be found in Figure 12 - figure supplement 3.

Figure 12 - figure supplement 2. Signals of selective sweeps in local populations. This figure is equivalent to Figure

12 in the main text but shows depressions in Tajima's *D* in single population samples. Such regions were identified by

486 enriching for samples with window-wise Tajima's D values < -0.9, which were smaller than 2 times the standard deviation

487 of Tajima's D for all samples in the corresponding genomic region. The top and bottom panels show two such examples

488 where samples with reduced Tajima's D are highlighted by thick lines and the corresponding sample names. See Table

489 S6 for additional details. A legend for the color codes of the samples can be found in Figure 12 - figure supplement 3.

● AT_14_Mau_1	● UA_14_Ode_19	● ES_14_Lle_35
● AT_14_Mau_2	● UA_14_Ode_20	● FI_14_Aka_36
● TR_14_Yes_3	● UA_14_Ode_21	● FI_14_Aka_37
● TR_14_Yes_4	● UA_14_Ode_22	● FI_14_Ves_38
● FR_14_Vil_5	● UA_14_Kyi_23	● DK_14_Kar_39
● FR_14_Vil_7	● UA_14_Kyi_24	● DK_14_Kar_41
● FR_14_Got_8	● UA_14_Var_25	● CH_14_Cha_42
● UK_14_She_9	● UA_14_Pyr_26	● CH_14_Cha_43
● UK_14_Sou_10	● UA_14_Dro_27	● AT_14_See_44
● CY_14_Nic_11	● UA_14_Cho_28	● UA_14_Kha_45
● UK_14_Mar_12	● UA_14_Cho_29	● UA_14_Kha_46
● UK_14_Lut_13	● SE_14_Lun_30	● UA_14_Cho_47
● DE_14_Bro_14	● DE_14_Mun_31	● UA_14_Cho_48
● DE_14_Bro_15	● DE_14_Mun_32	● UA_14_Kyi_49
● UA_14_Yal_16	● PT_14_Rec_33	● UA_14_Uma_50
● UA_14_Yal_18	● ES_14_Lle_34	● RU_14_Vald_51

Figure 12 - figure supplement 3. *Legend for color code in Figure 12, Figure 12 - figure supplement 1 and Figure 12 -*

figure supplement 2.

However, some regions, such as those around *wapl* or *HDAC6,* are characterized by low

recombination rates (< 0.5 cM/Mb; Table S5), which can itself lead to reduced variation

and Tajima's D (see also Nolte *et al.* 2013). Our screen also uncovered several regions

that have not previously been described as harboring sweeps (Table S6). These represent

promising candidate regions containing putative targets of positive selection. For several of

these candidate regions, patterns of variation were highly similar across the majority of

European samples, suggesting the existence of continent-wide selective sweeps that

either predate the colonization of Europe (e.g., Beisswanger *et al.* 2006) or that have

swept across all European populations more recently. In contrast, some candidate regions

were restricted to only a few populations and characterized by highly negative values of

Tajima's D, i.e. deviating from the among-population average by more than two standard

deviations, thus possibly hinting at cases of local, population-specific adaptation (Figure 12

- figure supplement 2 and Table S6 for examples).

European populations are strongly structured along an east-west gradient

We next investigated patterns of genetic differentiation due to demographic substructure.

Overall, pairwise differentiation as measured by F_{ST} was relatively low, though markedly

higher for X-chromosomes (0.043–0.076) than for autosomes (0.013–0.059; Student's t-

test; $p < 0.001$; Figure 13), possibly reflecting differences in effective population size

between the X chromosome and the autosomes (Hutter *et al.* 2007). One population, from

Sheffield (UK), showed an unusually high amount of differentiation on the X-chromosome

as compared to other populations (Figure 13).

Figure 13. Chromosome-wide average F_{ST}. Bar-plot showing the distribution of chromosome-wide F_{ST} averaged

across all possible pairwise comparisons for a given sample. The black bars show observed F_{ST} values and the dark

grey bars expected F_{ST} values (based on autosomal values; see Materials and Methods) for the X chromosome. The

light grey bars show autosomal F_{ST} values and error bars represent the standard deviation of average chromosome-wide

F_{ST} across all autosomal arms (*2L, 2R, 3L, 3R*).

Despite these overall low levels of among-population differentiation, European populations

showed some evidence of geographic substructure. To analyze this pattern in more detail,

we focused on a set of SNPs located in short introns (< 60 bp), as these sites are relatively

unaffected by selection (Haddrill *et al.* 2005; Singh *et al.* 2009; Parsch *et al.* 2010;

Clemente & Vogl 2012; Lawrie *et al.* 2013). We analyzed the extent of isolation by

distance (IBD) within Europe by correlating genetic and geographic distance and using

pairwise F_{ST} between populations as a measure of genetic isolation. F_{ST} was overall low

but significantly correlated with distance across the continent, indicating weak but

significant IBD (Mantel test; *p* < 0.001; max. F_{ST} ~ 0.05; Figure 14A). We also examined

those populations that were most and least separated by genetic differentiation, estimated

by pairwise F_{ST} (Figure 14B). In general, longitude had a stronger effect on isolation than

latitude, with populations showing the strongest differentiation separated along an east-

west, rather than a north-south, axis (Figure 14B). This pattern remained unchanged when

the number of populations sampled from Ukraine was reduced to avoid over-

representation (Figure 14 - figure supplement 1).

Figure 14 with 2 supplements. **Genetic differentiation among European populations**. (A) isolation by distance estimated by average genetic differentiation (F_{ST}) of 21,008 SNPs located in short introns (<60 bp) plotted against geographic distance. Mantel tests and linear regression (red dashed line and statistics in upper left box) indicate

significance. (B) Average neutral F_{ST} among populations. The center plot shows the distribution of average neutral F_{ST}

values for all 1,128 pairwise combinations. Mean neutral F_{ST} values were calculated by averaging individual F_{ST} values

from 20,008 genome-wide intronic SNPs for each pairwise comparison. The plots on the left and the right show

population pairs in the lower (blue) and upper (red) 5% tails of the F_{ST} distribution. (C) Population structure of all *DrosEU*

samples as determined by PCA of allele frequencies of 20,008 SNPs located in short introns (< 60 bp). The optimal

number of five clusters was estimated by hierarchical model fitting using the first four principal components. Cluster

assignment of each population, which was estimated by k-means clustering, is indicated by color.

Figure 14 - figure supplement 1: Genetic differentiation among European populations. Similar to Figure 14, the top plot

shows the population structure of all *DrosEU* samples as determined by PCA of allele frequencies of 20,008 SNPs

located in short introns (< 60 bp) but only on a reduced dataset with one randomly drawn Ukrainian sample to test if the

longitudinal pattern found in the full dataset was an artifact of the excessive sampling in the Ukraine. The bottom left plot

shows the cumulative variance explained by each of the principal components (PC) and the bottom right barplot depicts

the Eigenvalues of each PC.

Figure 14 - figure supplement 2: PCA of neutrally evolving SNPs in the 48 DrosEU samples. The top plot shows the cumulative variance explained by each of the principal components (PC) from a PCA based on the allele frequencies of 20,008 putatively neutrally evolving SNPs located in short (<60bp) introns distant from chromosomal inversions. The bottom barplot depicts the Eigenvalues of each PC.

To further explore these patterns, we performed a principal component analysis (PCA) on

the allele frequencies of SNPs in short introns. The first three principal components (PC)

explained more than 25% of the total variance (PC1: 16.3%, PC2: 5.4%, PC3: 4.8%,

eigenvalues = 599.2, 199.1, and 178.5 respectively; Figure 14C and Figure 14 - figure

supplement 2). As expected, PC1 was strongly correlated with longitude. Despite

significant signals of autocorrelation, as indicated by Moran's test on residuals from linear

regressions with PC1, the association with longitude was not due to spatial

autocorrelation, since a spatial error model also resulted in a significant association. PC2

was similarly, but to a lesser extent, correlated with longitude and also with altitude. PC3,

by contrast, was not associated with any variable examined (Table 2). None of the major

PC axes were correlated with season, indicating that there were no shared seasonal

differences across samples in our dataset. Hierarchical model fitting based on the first

three PC axes resulted in five distinct clusters (Figure 14C) that were oriented along the

axis of PC1, supporting the notion of strong longitudinal differentiation among European

populations. To the best of our knowledge, such a pronounced longitudinal signature of

differentiation has not previously been reported in European *D. melanogaster*.

Remarkably, this pattern is qualitatively similar to that observed for human populations

(Cavalli-Sforza 1966; Xiao *et al.* 2004; Francalacci & Sanna 2008), perhaps consistent

with co-migration of this commensal species.

Mitochondrial haplotypes also exhibit longitudinal population structure

Our finding that European populations are longitudinally structured is also supported by an

analysis of mitochondrial haplotypes. We identified two main mitochondrial haplotypes in

Europe, separated by at least 41 mutations (between G1.2 and G2.1; Figure 15A). Our

findings are consistent with similar analyses of mitochondrial haplotypes from a North

American *D. melanogaster* population (Cooper *et al.* 2015) as well as from worldwide

samples (Wolff *et al.* 2016), revealing varying degrees of differentiation among haplotypes,

ranging from only a few to hundreds of substitutions. The two G1 subtypes (G1.1 and

G1.2) are separated by only four mutations, and the three G2 subtypes are separated by a

maximum of four mutations (between G2.1 and G2.3). The estimated frequency of these

haplotypes varied greatly among populations (Figure 15B). Qualitatively, three types of

European populations can be distinguished based on these haplotypes, namely those with

(1) a high frequency (> 60%) of the G1 haplotypes, characteristic of central European

samples, (2) a low frequency (< 40%) of G1 haplotypes, a pattern common for Eastern

European populations in summer, and (3) a combined frequency of G1 haplotypes

between 40-60%, which is typical of samples from the Iberian Peninsula and from Eastern

Europe in fall (Figure 15 - figure supplement 1).

Figure 15 with 1 supplement. **Mitochondrial haplotypes.** (A) TCS network showing the relationship of 5 common

mitochondrial haplotypes; (B) estimated frequency of each mitochondrial haplotype in 48 European samples.

Figure 15 - figure supplement 1. Mitochondrial haplotypes. (A) Graphical summary of the combined frequency of G1

haplotypes in Europe. Summer and Fall are represented at the top and bottom of the circles, respectively. White – no

information; green, yellow and red represent a combined frequency of G1 haplotypes lower than 40%, in between 40%

and 60% and higher than 60%, respectively. (B) Correlations between the combined frequency of G1 haplotypes and

longitude (red diamonds for populations below 20° and red circles for populations above 20°).

We observed a significant shift in the relative frequencies of the two haplotype classes

between summer and fall samples in only two of the nine possible comparisons among

haplotypes. While there was no correlation between latitude and the combined frequency

of G1 haplotypes, we found a weak but significant negative correlation between G1

haplotypes and longitude ($r^2 = 0.10$; $p < 0.05$), which is consistent with the longitudinal

east-west population structure observed for intronic SNPs. In a subsequent analysis, we

divided the dataset at 20º longitude into an eastern and a western subset since in northern

Europe 20º longitude corresponds to the division of two major climatic zones, namely C

(temperate) and D (cold), according to the Köppen-Geiger climate classification (Peel *et al.*

2007). When splitting the populations in a western (longitude < 20º E) and an eastern

group (longitude > 20° E), we found a clear correlation between longitude and the

combined frequency of G1 haplotypes, explaining as much as 50% of the variation in the

western group (Figure 15 - figure supplement 1B). Similarly, in the eastern populations

longitude and the combined frequency of G1 haplotypes were correlated, explaining

approximately 20% of the variance (Figure 15 - figure supplement 1B). Thus, our data on

mitochondrial haplotypes clearly confirm the existence of pronounced east-west population

structure and differentiation in European *D. melanogaster*. While this might be due to

climatic selection, as recently found for clinal mitochondrial haplotypes in Australia (Camus

et al. 2017), we can presently not rule out an effect of demography.

The majority of TEs vary with longitude and altitude

To examine the population genetics of structural variants in our data, we first focused on

transposable elements (TEs). The repetitive content of the 48 samples analyzed ranged

from 16% to 21% with respect to nuclear genome size (Figure 16). The vast majority of

detected repeats were TEs, mostly represented by long terminal repeats (LTR) and long

interspersed nuclear elements (LINE; Class I), as well as a few DNA elements (Class II).

LTR content best explained total TE content (LINE+LTR+DNA) (Pearson's r = 0.87, p <

0.01, *vs.* DNA r = 0.58, p = 0.0117, and LINE r = 0.36, p < 0.01 and Figure S16A).

Figure 16 with 1 supplement. **Transposable elements.** Relative abundances of repeats among samples**.** Proportion of

647   each repeat class was estimated from sampled reads with dnaPipeTE (2 samples per run, 0.1X coverage per sample).

Figure 16 - figure supplement 1. **Transposable Elements.** (A) shows the contribution of each of the main TE classes to variation of the total TE content. Correlations (Pearson's correlation tests) between each of the three main TE classes (LTR, LINE and DNA) and the total TE content of each pool (LTR+LINE+DNA) in Kb. (B) The site frequency spectrum of TE frequencies per chromosome arm. Each dot represents the proportion of TEs in each bin per sample and a smoother geometric line had been added to highlight the trend. Lower panel is a zoom in of the above panel.

We next estimated population-wise frequencies of 1,630 TE insertions annotated in the *D. melanogaster* reference genome *v.6.04* using *T-lex2* (Table S7, Fiston-Lavier *et al.* 2010). On average, 56% of the TEs annotated in the reference genome were fixed in all samples. The remaining polymorphic TEs usually segregated at low frequency in all samples (Figure 16 - figure supplement 1A), potentially due to the effect of purifying selection (González *et al.* 2008; Petrov *et al.* 2011; Kofler *et al.* 2012; Cridland *et al.* 2013; Blumenstiel *et al.* 2014). However, we also observed 142 TE insertions present at intermediate (>10% and <95%) frequencies (Figure 16 - figure supplement 1B), which might be consistent with transposition-selection balance (Charlesworth *et al.* 1994).

In each of the 48 samples TE frequency and recombination rate were negatively correlated on a genome-wide level (Spearman rank sum test; $p < 0.01$), as previously reported (Bartolomé *et al.* 2002; Petrov *et al.* 2011; Kofler *et al.* 2012). This pattern still holds when only polymorphic TEs (population frequency <95%) are analyzed, although it becomes statistically non-significant for some chromosomes and populations (Table S8). In either case, the correlation is more negative when using broad-scale, rather than fine-scale, recombination rate estimates (Materials and methods, Tables S8B, S8D). This indicates that broad-scale recombination patterns may best capture long-term population recombination patterns.

We further tested whether the distribution of TE frequencies among samples could be explained by geographical or temporal variables. We focused on the 141 TE insertions that showed frequency variability among samples (interquartile range, (IQR) > 10; see Materials and Methods). Of these, 73 TEs showed significant associations with geographical or temporal variables after multiple testing correction (Table S9). Note that we used a conservative p-value threshold (< 0.001), and we did not find significant residual spatio-temporal autocorrelation among samples for any TE tested (Moran's *I* > 0.05 for all tests; Table S9). 16 out of 73 TEs were located in regions of very low recombination (0 cM/Mb for either of the two recombination measures used). Among the 57 significant TEs located in high recombination regions, we observed significant correlations of 13 TEs with longitude, 13 with altitude, 5 with latitude, and 3 with season (Table S9). In addition, the frequencies of the other 23 insertions were significantly correlated with more than one of the above-mentioned variables (Table S9). These significant TEs were scattered along the main five chromosome arms (Table S9). Among

the 57 significant TEs located in high recombination regions two TE families were enriched

(χ^2 p-values after Yate's correction < 0.05): the LTR *297* family with 11 copies, and the

DNA *pogo* family with 5 copies (Table S10). We also checked the genomic localization of

the 57 TEs. Most of them (42) were located inside genes: two in 5'UTR, four in 3'UTR, 18

in the first intron, and 18 TEs in subsequent introns. Additionally, 7 TEs are <1 kb from the

nearest gene, indicating that these might potentially affect the regulation of nearby genes

(Table S9). Interestingly, 14 of these 57 TEs coincide with previously identified candidate

adaptive TEs (Table S9), suggesting that our dataset might be enriched for adaptive

insertions. However, further analyses are needed to discard the effect of non-selective

forces on the patterns observed.

Inversion polymorphisms in Europe exhibit latitudinal and longitudinal clines

Chromosomal inversions are another class of important and common structural genomic

variants, often exhibiting frequency clines on multiple continents, some of which have been

shown to be adaptive (e.g. Knibb 1982; Umina *et al.* 2005; Kapun *et al.* 2014; 2016a).

However, little is known yet about the spatial distribution and clinality of inversions in

Europe. We used a panel of inversion-specific marker SNPs (Kapun *et al.* 2014) to

examine the presence and frequency of six cosmopolitan inversion polymorphisms

(*In(2L)t, In(2R)NS, In(3L)P, In(3R)C, In(3R)Mo, In(3R)Payne*) in the 48 samples. All

populations were polymorphic for one or more inversions (Figure 17). However, only

In(2L)t segregated at substantial frequencies in most populations (average frequency =

20.2%). All other inversions were either absent or occurred at low frequencies (average

frequencies: *In(2R)NS* = 6.2%, *In(3L)P* = 4%, *In(3R)C* = 3.1%, *In(3R)Mo* =2.2%,

In(3R)Payne = 5.7%).

Figure 17 with 1 supplement. **Distribution of inversion frequencies.** Cumulative bar plots showing the absolute

frequencies of six cosmopolitan inversions (*In(2L)t, In(2R)NS, In(3L)P, In(3R)C, In(3R)Mo, In(3R)Payne*) in all 48

population samples of the *DrosEU* dataset.

Figure 17 - figure supplement 1. Clinal variation of the inversion *In(3R)Payne* across continents. Parallel

Despite their overall low frequencies, several inversions exhibited clinal patterns across

space (Table 3). We observed significant latitudinal clines for *In(3L)P, In(3R)C* and

In(3R)Payne. Although they differed in overall frequencies, *In(3L)P* and *In(3R)Payne*

showed latitudinal clines in Europe that are qualitatively similar to the clines previously

observed along the North American and Australian east coasts (Figure S17 and Table

S11, Kapun *et al.* 2016a). For the first time, we also detected a longitudinal cline for *In(2L)t*

and *In(2R)NS*, with both inversions decreasing in frequency from east to west, a result that

is consistent with our finding of strong longitudinal among-population differentiation in

Europe. *In(2L)t* also increased in frequency with altitude (Table 3). Except for *In(3R)C,* we

did not find significant residual spatio-temporal autocorrelation among samples for any

inversion tested (Moran's $I \approx 0$, $p > 0.05$ for all tests; Table 3), suggesting that our analysis

was not confounded by spatial autocorrelation for most of the inversions. It will clearly be

interesting to examine the extent to which clines in inversions (and other genomic variants)

across Europe are shaped by selection and/or demography in future work.

European *Drosophila* microbiomes contain trypanosomatids and novel viruses

We were also interested in determining the abundance of microbiota associated with *D.*

melanogaster from the Pool-Seq data – these endosymbionts often have crucial functions

in affecting the life history, immunity, hormonal physiology, and metabolic homeostasis of

their fly hosts (e.g., Trinder *et al.* 2017; Martino *et al.* 2017). The taxonomic origin of a total

of 262 million non-*Drosophila* reads was inferred using MGRAST, which identifies and

counts short protein motifs ('features') within reads (Meyer *et al.* 2008). The largest fraction

of protein features was assigned to *Wolbachia* (on average 53.7%; Figure 18)*,* a well-

known endosymbiont of *Drosophila* (Werren *et al.* 2008). The relative abundance of

Wolbachia protein features varied strongly between samples ranging from 8.8% in a

sample from the UK to almost 100% in samples from Spain, Portugal, Turkey and Russia

(Table 1). Similarly, *Wolbachia* loads varied 100x between samples if we use the ratio of

Wolbachia protein features divided by the number of *Drosophila* sequences retrieved for

that sample as a proxy for relative micro-organismal load (for a full table of micro-

organismal loads standardized by *Drosophila* genome coverage see Table S12).

Figure 18: Microbiome. Relative abundance of *Drosophila*-associated microbes as assessed by MGRAST classified

shotgun sequences. Microbes had to reach at least 3% relative abundance in one of the samples to be presented

Acetic acid bacteria of the genera *Gluconobacter*, *Gluconacetobacter*, and *Acetobacter*

were the second largest group, with an average relative abundance of 34.4%.

Furthermore, we found evidence for the presence of several genera of Enterobacteria

(*Serratia, Yersinia, Klebsiella, Pantoea, Escherichia, Enterobacter, Salmonella*, and

Pectobacterium). *Serratia* occurs only at low frequencies or is absent from most of our

samples, but reaches a very high relative abundance in the Nicosia summer collection

(54.5%). This high relative abundance was accompanied by an 80x increase in *Serratia*

bacterial load. We detected several eukaryotic microorganisms, although they were less

abundant than the bacteria. The fraction of fungal protein features is larger than 3% in only

three of our samples from Finland, Austria and Turkey (Table 1). Interestingly, we detected

the presence of trypanosomatids in 16 of our samples, consistent with other recent

evidence that *Drosophila* can host these organisms (Wilfert *et al.* 2011; Chandler & James

2013; Hamilton *et al.* 2015).

Our data also allowed us to detect the presence of five different DNA viruses (Table S13).

These included approximately two million reads from *Kallithea* nudivirus (Webster *et al.*

2015), allowing us to assemble the complete *Kallithea* genome for the first time (>300-fold

coverage in the Ukrainian sample UA_Kha_14_46; Genbank accession KX130344). We

also identified around 1,000 reads from a novel nudivirus that is closely related to *Kallithea*

virus and to *Drosophila innubila* nudivirus (Unckless 2011) in sample DK_Kar_14_41 from

Karensminde, Denmark (Table 1). These sequences permitted us to identify a publicly

available dataset (SRR3939042: 27 male *D. melanogaster* from Esparto, California;

Machado *et al.* 2016) that contained sufficient reads to complete the genome (provisionally

named "*Esparto* Virus"; KY608910). We further identified two novel Densoviruses

(*Parvoviridae*), which we have provisionally named "*Viltain* virus", a relative of *Culex*

pipiens densovirus found at 94-fold coverage in sample FR_Vil_14_07 (Viltain; KX648535)

and "*Linvill Road* virus", a relative of *Dendrolimus punctatus* densovirus that was

represented by only 300 reads here, but which has previously been found to have a high

coverage in dataset SRR2396966 from a North American sample of *D. simulans*

(KX648536; Machado *et al.* 2016). In addition, we detected a novel member of the

Bidnaviridae family,"*Vesanto* virus", a bidensovirus related to *Bombyx mori* densovirus 3

with approximately 900-fold coverage in sample FI_Ves_14_38 (Vesanto; KX648533 and

KX648534), Using a detection threshold of >0.1% of the *Drosophila* genome copy number,

the most commonly detected viruses were *Kallithea* virus (30/48 of the pools) and *Vesanto*

virus (25/48), followed by *Linvill Road* virus (7/48) and *Viltain* virus (5/48), with *Esparto*

virus being the rarest (2/48). In some samples, the viruses reached strikingly high titers: on

13 occasions the virus genome copy number in the pool exceeded the host genome copy

number, reaching a maximum of nearly 20-fold in Vesanto.

This continent-wide analysis of the microbiota associated with fruit flies suggests that

natural populations of European *D. melanogaster* differ greatly in the composition and

relative abundance of microbes and viruses.

Discussion

In recent years, large-scale population resequencing projects have shed light on the

biology of both model (Mackay *et al.* 2012; Langley *et al.* 2012; Consortium 2015; Lack *et*

al. 2015; Alonso-Blanco *et al.* 2016; Lack *et al.* 2016) and non-model organisms (e.g.,

Hohenlohe *et al.* 2010; Wolf *et al.* 2010). Such massive datasets contribute greatly to our

growing understanding of the processes that create and maintain genetic variation in

natural populations. However, the relevant spatio-temporal scales for population genomic

analyses remain largely unknown. Here we have applied, for the first time, a

comprehensive sampling and sequencing strategy to European populations of *D.*

melanogaster, allowing us to uncover previously unknown aspects of this species'

population biology.

A main result from our analyses of SNPs located in short introns and presumably evolving

neutrally (Parsch *et al.* 2010) is that European *D. melanogaster* populations exhibit very

pronounced longitudinal differentiation, a pattern that – to the best of our knowledge – has

not been observed before for the European continent (for patterns of longitudinal

differentiation in Africa see e.g. Michalakis & Veuille 1996; Aulard *et al.* 2002; Fabian *et al.*

2015). Genetic differentiation was greatest between populations from eastern and western

Europe (Figure 14). The eastern populations included those from the Ukraine, Russia, and

Turkey, as well as one from eastern Austria, suggesting that there may be a region of

restricted gene flow in south-central Europe. However, populations from Finland and

Cyprus are more similar to western populations than to eastern populations, possibly as a

result of migration along shipping routes in the Baltic and Mediterranean seas. More data

from populations in the unsampled, intermediate regions are needed to better delineate

the geographic limits of the eastern and western population groups. Consistent with the

strong differentiation between eastern and western populations, our PCA analysis

revealed that longitude was the major factor associated with among-population

divergence, with no significant effect of latitude (Figure 14C; Table 2). Thus, the patterns

of neutral genetic differentiation in Europe contrast with those previously reported for North

America, where latitude impacts neutral differentiation (Machado *et al.* 2016; Kapun *et al.*

2016a). However, our present analysis does not exclude the existence of clinally varying

polymorphisms in European populations outside short introns: for example, we detected

latitudinal frequency clines both for TEs and inversion polymorphisms. A detailed analysis

of genome-wide patterns of clinal variation in the 2014 *DrosEU data* is beyond the scope

of this paper and currently under way.

The mitochondrial genome and several chromosomal inversions and TEs showed similar

patterns of differentiation as the rest of the genome, with the main axis of differentiation

being longitudinal. Uncovering the extent to which this pattern is driven by demography

and/or selection, and identifying the underlying environmental correlates (including any

potential role of co-migration with human populations), will be an important task for future

analyses. Due to the high density of samples and the large number of SNP markers

examined, our results reveal that European populations of *D. melanogaster* exhibit much

more differentiation and structure than previously thought (e.g., Baudry *et al.* 2004;

Dieringer *et al.* 2005; Schlötterer *et al.* 2006; Nunes *et al.* 2008; Mateo *et al.* 2018).

Within the eastern and western population groups there was a low – but detectable – level

of genetic differentiation among populations, including those that are geographically close

(Figure 14C). These population differences persisted over a timespan of at least 2–3

months, as there was less genetic differentiation between the summer and fall samples of

the 13 locations sampled at multiple time points than between neighboring populations

(Figure 14C). Thus, while the weak but significant signal of IBD suggests homogenizing

gene flow across geography, there is seasonally stable differentiation among populations.

The season in which samples were collected did not show a significant association with

genetic differentiation, except when considered in conjunction with longitude or altitude

(Table 2). However, the data analyzed here are from a single year only: demonstrating

recurrent shifts in SNP frequencies due to temporally varying selection will require analysis

of additional annual samples. For an extensive analysis of patterns of seasonal variation

across a broad geographic scale see Machado *et al.* (2018)

Our Pool-Seq data also allowed us to characterize geographic patterns in both inversions

and TEs. In marked contrast to putatively neutral SNPs, the frequencies of several

chromosomal inversions, including *In(3L)P, In(3R)C,* and *In(3R)Payne*, showed a

significant correlation with latitude (Table 3). For *In(3L)P* and *In(3R)Payne*, the latitudinal

clines were in qualitative agreement with parallel clines reported from North America and

Australia, with the inversions decreasing in frequency as distance from the equator

increases (Mettler *et al.* 1977; Knibb *et al.* 1981; Fabian *et al.* 2012; Kapun *et al.* 2014;

Rane *et al.* 2015; Kapun *et al.* 2016a). This suggests that these inversions may contain

genetic variants that are better adapted to warmer environments than to temperate

climates. The overall frequencies of these inversions are, however, low in Europe (<5%),

indicating that they might play only a minor role in local adaptation to European habitats.

Some euchromatic TE insertions also showed geographic or seasonal patterns of variation

(Table S7), indicating that they might play a role in local adaptation, particularly as many of

them are located in regions where they could affect gene regulation. Importantly, several

inversions and TEs also showed longitudinal frequency gradients, thus supporting the

notion that European populations exhibit marked longitudinal differentiation.

We also examined signatures of selective sweeps in our dataset. We found 144 genomic

regions that showed signatures of hard sweeps in regions of normal recombination (cM/Mb

≥ 0.5), and with reduced variation and negative Tajima's D ($D \leq$ -0.8) in all European

populations (Figure 12, Table S6). Four of these regions were identified in previous studies

as potential targets for positive selection.

The first region, at the center of chromosome arm *2R* (Figure 12A, Table S6), was

previously found to be strongly differentiated between African and North American

populations (Langley *et al.* 2012) and contains two genes, *Cyp6g1* and *Hen-1*, that are

associated with recent, strong selection. The cytochrome P450 gene *Cyp6g1* has been

linked to insecticide resistance (Daborn *et al.* 2002; Schmidt *et al.* 2010), shows evidence

for recent selection independently in both *D. melanogaster* and *D. simulans* (Schlenke &

Begun 2003; Catania *et al.* 2004), and is associated with a large differentiated region in

the Australian latitudinal cline (Kolaczkowski *et al.* 2011a). *Hen-1*, a methyltransferase

involved in maturation of small RNAs involved in virus and TE suppression, showed

marginally non-significant evidence for selective sweeps in North American and African

populations of *D. melanogaster* (Kolaczkowski *et al.* 2011b).

The second region previously implicated in a selective sweep is located on chromosome

arm *3L* (Figure 12B, Table S6) and centered around the chimeric gene *CR18217*, which

formed from the fusion of a gene encoding a DNA-repair enzyme (*CG4098*) and a

centriole gene (spd-2; Rogers & Hartl 2012). *CR18217* appears to be unique to *D.*

melanogaster, but – in spite of its recent origin – segregates at frequencies of around 90%

(Rogers & Hartl 2012), consistent with a recent strong sweep in this region of the genome.

This putative sweep region also spans *Prosbeta6*, which (like HDAC) encodes a gene

involved in proteolysis (Flybase v. FB2017_05; Gramates *et al.* 2017). *Prosbeta6* also

shows homology to genes involved in immune function (Lyne *et al.* 2007; Handu *et al.*

2015), which might explain why it has been a target of positive selection.

The third previously characterized sweep region, surrounding the *wapl* gene on the *X*

chromosome (Table S6), was identified as showing evidence of strong selective sweeps in

both African and European *D. melanogaster* populations (Beisswanger *et al.* 2006; Boitard

et al. 2012). The genic targets of selection in this region are unclear, but most likely are

ph-p in Europe and *ph-p* or *ph-d* in Africa (Beisswanger *et al.* 2006). These genes are

tandem duplicates involved in the Polycomb response pathway, which functions as an

epigenetic repressor of transcription (reviewed in Kassis *et al.* 2017).

The fourth previously observed sweep region, originally identified in African populations of

D. melanogaster, is also located on the *X* chromosome (Table S6), but 30 cM closer to the

telomere and thus not implicating the *wapl* region (Beisswanger *et al.* 2006; Boitard *et al.*

2012). Selection in this region has been attributed to the *HDAC6* gene (Svetec *et al.*

2009). *HDAC6*, although nominally a histone deacetylase, actually functions as a central

player in managing cytotoxic assaults, including in transport and degradation of misfolded

protein aggregates (reviewed in Matthias *et al.* 2008; Svetec *et al.* 2009).

Our data support the widespread occurrence of these previously identified sweeps in many

populations in Europe. Notably, practically all European populations examined showed

reduced variation and negative Tajima's *D* in these sweep regions. This is consistent with

the sweeps either pre-dating the colonization of Europe (e.g., Beisswanger *et al.* 2006) or

having swept across Europe more recently (also see Stephan 2010 for discussion). In

addition, we also uncovered several novel genomic regions with tentative evidence for

hard sweeps (Table S6) – these regions represent a valuable source for future analyses of

signals of adaptive evolution in European *Drosophila*.

Finally, we used our Pool-Seq data to identify microbes and viruses and to quantify their

presence in natural populations of *D. melanogaster* across the European continent.

Wolbachia was the most abundant bacterial genus associated with the flies, but its relative

abundance and load varied greatly among samples (Figure 18). The second most

abundant bacterial taxon was acetic acid bacteria (*Acetobacteraceae*), a group previously

found among the most abundant bacteria in natural *D. melanogaster* isolates (Chandler *et*

al. 2011; Staubach *et al.* 2013). Other microbes were highly variable abundance in relative

abundance. For example, *Serratia* abundance was low in most populations, but very high

in the Nicosia sample, which might reflect that there are individuals in the Nicosia sample

that carry a systemic *Serratia* infection generating high bacterial loads. Future sampling

may shed light on the temporal stability and/or population specificity of these patterns.

Contrary to expectation, we found relatively few yeast sequences. This is a bit surprising

because yeasts are commonly found on rotting fruit, the main food substrate of *D.*

melanogaster, and have been found in association with *Drosophila* before (Barata *et al.*

2012; Chandler *et al.* 2012). This suggests that, although yeasts can attract flies and play

a role in food choice (Becher *et al.* 2012; Buser *et al.* 2014), they might not be highly

prevalent in or on *D*. *melanogaster* bodies. While trypanosomatids have been reported in

association with *Drosophila* before (Wilfert *et al.* 2011; Chandler & James 2013; Hamilton

et al. 2015), our study provides the first systematic detection across a wide geographic

range in *D*. *melanogaster*. Despite being host to a wide diversity of RNA viruses (Huszar &

Imler 2008; Webster *et al.* 2015), only three DNA viruses have previously been reported in

association with *Drosophilidae*, and only one from *D. melanogaster* (Unckless 2011;

Webster *et al*. 2015; 2016). Here, we have discovered four new DNA viruses in *D.*

melanogaster. Although it is not possible to directly estimate viral prevalence from pooled

sequencing data, we found that the DNA viruses of *D. melanogaster* can be very

widespread, with *Kallithea* virus detectable at a low level in most populations.

A striking qualitative pattern in our microbiome data is the high level of variability among

populations in the composition and relative amounts of different microbiota and viruses.

Thus, an interesting open question is to what extent geographic differences in microbiota

might contribute to phenotypic differences and local adaptation among fly populations,

especially given that there might be tight and presumably local co-evolutionary interactions

between fly hosts and their endosymbionts (e.g., Haselkorn *et al.* 2009; Richardson *et al.*

2012; Staubach *et al.* 2013; Kriesner *et al.* 2016).

In conclusion, our study demonstrates that extensive sampling on a continent-wide scale

and pooled sequencing of natural populations can reveal new aspects of population

biology, even for a well-studied species such as *D. melanogaster*. Such extensive

population sampling is feasible due to the close cooperation and synergism within our

international consortium. Our efforts in Europe are paralleled in North America by the

Drosophila Real Time Evolution Consortium (*Dros-RTEC*), with whom we are currently

collaborating to compare population genomic data across continents. In future years, our

consortia will continue to sample and sequence European and North American *Drosophila*

populations in order to study these populations with increasing spatial and temporal

resolution and to provide an unprecedented resource for the *Drosophila* and population

genetics communities.

Materials and Methods

The 2014 *DrosEU* dataset analyzed here consists of 48 samples of *D. melanogaster*

collected from 32 geographical locations at different time-points across the European

continent, through a joint effort of 18 European research groups (see Figure 2, Table 1).

Field collections were performed with baited traps using a standardized protocol (see

Supplementary file for details). Up to 40 males from each collection were pooled, and DNA

extracted from each pool, using a standard phenol-chloroform based protocol. Each

sample was processed in a single pool (Pool-Seq; Schlötterer *et al.* 2014), with each pool

consisting of at least 33 wild-caught individuals. To exclude morphologically similar and

co-occurring species, such as *D. simulans*, as potential contaminants from the samples,

we only used wild-caught males and distinguished among species by examining genital

morphology. Despite this precaution, we identified a low level of *D. simulans* contamination

in our samples, and further steps were thus taken to exclude *D. simulans* sequences from

our analysis (see below). The 2014 *DrosEU* dataset represents the most comprehensive

spatio-temporal sampling of European *D. melanogaster* populations available to date

(Table 1, Figure 3).

DNA extraction, library preparation and sequencing

DNA was extracted from pools of 33–40 males per sample after joint homogenization with

bead beating and standard phenol/chloroform extraction. A detailed extraction protocol can

be found in the Supporting Information file. In brief, 500 ng of DNA in a final volume of 55.5

µl were sheared with a Covaris instrument (Duty cycle 10, intensity 5, cycles/burst 200,

time 30) for each sample separately. Library preparation was performed using NEBNext

Ultra DNA Lib Prep-24 and NebNext Multiplex Oligos for Illumina-24 following the

manufacturer's instructions. Each pool was sequenced as paired-end fragments on a

Illumina NextSeq 500 sequencer at the Genomics Core Facility of Pompeu Fabra

University (UPF; Barcelona, Spain). Samples were multiplexed in five batches of 10

samples each, except for one batch that contained only 8 samples (see Supplementary

Table S1 for further information). Each multiplexed batch was sequenced on four lanes to

obtain an approximate 50x raw coverage for each sample. Reads were sequenced to a

length of 151 bp with a median insert size of 348 bp (ranging from 209 to 454 bp).

1012

Mapping pipeline and variant calling

1013

Prior to mapping, we trimmed and filtered raw FASTQ reads to remove low-quality bases

(minimum base PHRED quality = 18; minimum sequence length = 75 bp) and sequencing

adaptors using cutadapt (v. 1.8.3; Martin 2011). We only retained read pairs for which both

reads fulfilled our quality criteria after trimming. FastQC analyses of trimmed and quality

filtered reads showed overall high base-qualities (median ranging from 29 to 35 in all 48

samples) and indicated a loss of ~1.36% of all bases after trimming relative to the raw

data. We used bwa mem (v. 0.7.15; Li 2013) with default parameters to map trimmed

reads against a compound reference genome consisting of the genomes from *D.*

melanogaster (v.6.12) and genomes from common commensals and pathogens, including

Saccharomyces cerevisiae (GCF_000146045.2), *Wolbachia pipientis* (NC_002978.6),

Pseudomonas entomophila (NC_008027.1), *Commensalibacter intestine*

(NZ_AGFR00000000.1), *Acetobacter pomorum* (NZ_AEUP00000000.1), *Gluconobacter*

morbifer (NZ_AGQV00000000.1), *Providencia burhodogranariea* (NZ_AKKL00000000.1),

Providencia alcalifaciens (NZ_AKKM01000049.1), *Providencia rettgeri*

(NZ_AJSB00000000.1), *Enterococcus faecalis* (NC_004668.1), *Lactobacillus brevis*

(NC_008497.1), and *Lactobacillus plantarum* (NC_004567.2), to avoid paralogous

mapping. We used Picard (v.1.109; http://picard.sourceforge.net) to remove duplicate

reads and reads with a mapping quality below 20. In addition, we re-aligned sequences

flanking insertions-deletions (indels) with GATK (v3.4-46; McKenna *et al.* 2010).

1033

1034 After mapping, Pool-Seq samples were tested for DNA contamination from *D. simulans*.

1035 To do this, we used a set of SNPs known to be divergent between *D. simulans* and *D.*

1036 *melanogaster* and assessed the frequencies of *D. simulans*-specific alleles following the

approach of Bastide *et al.* (2013). We combined the genomes of *D. melanogaster* (v.6.12)

and *D. simulans* (Hu *et al.* 2013) and separated species-specific reads for samples with a

contamination level > 1% via competitive mapping against the combined references using

the pipeline described above. Custom software was used to remove reads uniquely

mapping to *D. simulans*. In 9 samples, we identified contamination with *D. simulans*,

ranging between 1.2 % and 8.7% (Table S1). After applying our decontamination pipeline,

contamination levels dropped below 0.4 % in all 9 samples.

We used *Qualimap* (v. 2.2., Okonechnikov *et al.* 2016) to evaluate average mapping

qualities per population and chromosome, which ranged from 58.3 to 58.8 (Table S1). We

found heterogeneous sequencing depths among the 48 samples, ranging from 34x to

115x for autosomes and from 17x to 59x for *X*-chromosomes (Figure S1, Table S1). We

then combined individual BAM files from all samples into a single *mpileup* file using

samtools (v. 1.3; Li & Durbin 2009). Due to the large number of Pool-Seq datasets

analyzed in parallel, we had to implement quality control criteria for all libraries jointly in

order to call SNPs. To accomplish this, we implemented a novel custom SNP calling

software to call SNPs with stringent heuristic parameters (PoolSNP; see Supplementary

Information), available at Dryad (doi: https://doi.org/10.5061/dryad.rj1gn54). A site was

considered polymorphic if (1) the minimum coverage from all samples was greater than

10x, (2) the maximum coverage from all samples was less than the 95th coverage

percentile for a given chromosome and sample (to avoid paralogous regions duplicated in

the sample but not in the reference), (3) the minimum read count for a given allele was

greater than 20x across all samples pooled, and (4) the minimum read frequency of a

given allele was greater than 0.001 across all samples pooled. The above threshold

parameters were optimized based on simulated Pool-Seq data in order to maximize true

positives and minimize false positives (see Figure S18 and Supporting Information).

Additionally, we excluded SNPs (1) for which more than 20% of all samples did not fulfill

the above-mentioned coverage thresholds, (2) which were located within 5 bp of an indel

with a minimum count larger than 10x in all samples pooled, and (3) which were located

within known transposable elements (TE) based on the *D. melanogaster* TE library v.6.10.

We further annotated our final set of SNPs with SNPeff (v.4.2; Cingolani *et al.* 2012) using

the Ensembl genome annotation version BDGP6.82 (Figure 3).

1069

Combined and population-specific site frequency spectra (SFS)

We quantified the amount of allelic variation with respect to different SNP classes. For this,

we first combined the full dataset across all 48 samples and used the SNPeff annotation

(see above) to classify the SNPs into four classes (intergenic, intronic, non-synonymous,

and synonymous). For each class, we calculated the site frequency spectrum (SFS) based

on minor allele frequencies for the *X*-chromosome and the autosomes, as well as for each

sample and chromosomal arm separately, by counting alleles in 50 frequency bins of size

0.01.

1078

Genetic variation in Europe

We characterized patterns of genetic variation among the 48 samples by estimating three

standard population genetic parameters: π, Watterson's θ and Tajima's D (Watterson

1975; Nei 1987; Tajima 1989). We focused on SNPs located on the five major

chromosomal arms (*X*, *2L*, *2R*, *3L*, *3R*) and calculated sample-wise π, θ and Tajima's D

with corrections for Pool-Seq data (Kofler *et al.* 2011). Since PoPoolation, the most

commonly used software for population genetics inference from Pool-Seq data, does not

allow using predefined SNPs (which was desirable for our analyses), we implemented

corrected population genetic estimators described in Kofler *et al.* (2011) in Python

(PoolGen; available at Dryad; doi: https://doi.org/10.5061/dryad.rj1gn54). Before

calculating the estimators, we subsampled the data to an even coverage of 40x for the

autosomes and 20x for the *X*-chromosome to control for the sensitivity to coverage

variation of Watterson's θ and Tajima's *D* (Korneliussen *et al.* 2013). At sites with greater

than 40x coverage, we randomly subsampled reads to 40x without replacement; at sites

with below 40x coverage, we sampled reads 40 times with replacement. Using *R* (R

Development Core Team 2009), we calculated sample-wise chromosome-wide averages

for autosomes and *X* chromosomes separately and tested for correlations of π, θ and

Tajima's *D* with latitude, longitude, altitude, and season using a linear regression model of

the following form: y_i = *Lat + Lon + Alt +Season* + ε_i, where y_i is either π, θ and *D*. Here,

latitude, longitude, and altitude are continuous predictors (Table 1), while 'season' is a

categorical factor with two levels *S* ("summer") and *F* ("fall"), corresponding to collection

dates before and after September 1st, respectively. We chose this arbitrary threshold for

consistency with previous studies (Bergland *et al.* 2014; Kapun *et al.* 2016a). To further

test for residual spatio-temporal autocorrelation among the samples (Kühn & Dormann

2012), we calculated Moran's *I* (Moran 1950) with the *R* package *spdep* (v.06-15., Bivand

& Piras 2015). To do this, we used the residuals of the above-mentioned models, as well

as matrices defining pairs of samples as neighbors weighted by geographical distances

between the locations (samples within 10° latitude/longitude were considered neighbors).

Whenever these tests revealed significant autocorrelation (indicating non-independence of

the samples), we repeated the above-mentioned regressions using spatial error models as

implemented in the *R* package *spdep*, which incorporate spatial effects through weighted

error terms, as described above.

1111

1112   To test for confounding effects of variation in sequencing errors between runs, we

1113   extended the above-mentioned linear models including the run ID as a random factor

1114   using the *R* package *lme4* (v.1.1-14; see Supporting Information). Preliminary analyses

1115   showed that this model was not significantly better than simpler models, so we did not

1116   include sequencing run in the final analysis (see Supporting information and Table S4).

1117

To investigate genome-wide patterns of variation, we averaged π, θ, and D in 200 kb non-

overlapping windows for each sample and chromosomal arm separately and plotted the

distributions in *R.* In addition, we calculated Tajima's D in 50 kb sliding windows with a

step size of 10 kb to investigate fine-scale deviations from neutral expectations. We

applied heuristic parameters to identify genomic regions harboring potential candidates for

selective sweeps. To identify candidate regions with sweep patterns across most of the 48

samples, we searched for windows with log-transformed recombination rates ≥ 0.5,

pairwise nucleotide diversity ($\pi \leq 0.004$), and average Tajima's D across all populations $\leq$ -

0.8 (5% percentile). To identify potential selective sweeps restricted to a few population

samples only, we searched for regions characterized as above but allowing one or more

samples with Tajima's D being more than two standard deviations smaller than the

window-wise average. To account for the effects of strong purifying selection in gene-rich

genomic regions which can result in local negative Tajima's D (Tajima 1989) and thus

confound the detection of selective sweeps, we repeated the analysis based on silent sites

(4-fold degenerate sites, SNPs in short introns of $\leq$ 60 bp lengths and SNPs in intergenic

regions in $\geq$ 2000 bp distance to the closest gene) only. Despite of the reduction in

polymorphic sites available for this analysis, we found highly consistent sweep regions and

therefore proceeded with the full SNP datasets, which provided better resolution (results not

shown).

1137

For statistical analysis, the diversity statistics were log-transformed to normalize the data.

We then tested for correlations between π and recombination rate using *R* in 100 kb non-

overlapping windows and plotted these data using the *ggplot2* package (v.2.2.1., Wickham

2016). We used two different recombination rate measurements: (i) a fine-scale, high

resolution genomic recombination rate map based on millions of SNPs in a small number

of strains (Comeron *et al.* 2012), and (ii) the broad-scale Recombination Rate Calculator

based on Marey maps generated by laboratory cross data fitting genetic and physical

positions of 644 markers to a third-order polynomial curve for each chromosome arm

(Fiston-Lavier *et al.* 2010). Both measurements were converted to version 6 of the *D.*

melanogaster reference genome to match the genomic position of π estimates (see

above).

1149

SNP counts and overlap with other datasets

We used the panel of SNPs identified in the *DrosEU* dataset (available at Dryad; doi:

https://doi.org/10.5061/dryad.rj1gn54) to describe the overlap in SNP calls with other

published *D. melanogaster* population data: the *Drosophila* Population Genomics Project 3

(DPGP3) from Siavonga, Zambia (69 non-admixed lines; Lack *et al.* 2015; 2016) and the

Drosophila Genetic Reference Panel (DGRP) from Raleigh, North Carolina, USA (205

inbred lines; Mackay *et al.* 2012; Huang *et al.* 2014). For these comparisons, we focused

on biallelic SNPs on the 5 major chromosome arms. We used *bwa mem* for mapping and

a custom pipeline for heuristic SNP calling (PoolSNP; Figure 3). To make the data from

the 69 non-admixed lines from Zambia (Lack *et al.* 2015; 2016) comparable to our data,

we reanalyzed these data using our pipeline for mapping and variant calling (Figure 3).

The VCF file of the DGRP data was downloaded from http://dgrp2.gnets.ncsu.edu/ and

converted to coordinates according to the *D. melanogaster* reference genome v.6. We

depicted the overlap of SNPs called in the three different populations using elliptic Venn

diagrams with *eulerAPE* software (v3 3.0.0., Micallef & Rodgers 2014). While the *DrosEU*

data were generated from sequencing pools of wild-caught individuals, both the DGRP

and DPGP3 data are based on individual sequencing of inbreed lines and haploid

individuals, respectively.

1168

Genetic differentiation and population structure in European populations

To estimate genome-wide pairwise genetic differences, we used custom software to

calculate SNP-wise F_{ST} using the approach of Weir and Cockerham (1984). We estimated

SNP-wise F_{ST} for all possible pairwise combinations among samples. For each sample, we

then averaged F_{ST} across all SNPs for all pairwise combinations that include this particular

sample and finally ranked the 48 population samples by overall differentiation.

1175

We inferred demographic patterns in European populations by focusing on 20,008

putatively neutrally evolving SNPs located in small introns (less than 60 bp length; Haddrill

et al. 2005; Singh *et al.* 2009; Parsch *et al.* 2010; Clemente & Vogl 2012; Lawrie *et al.*

2013) that were at least 200 kb distant from the major chromosomal inversions (see

below). To assess isolation by distance (IBD), we averaged F_{ST} values for each sample

pair across all neutral markers and calculated geographic distances between samples

using the haversine formula (Green & Smart 1985) which takes the spherical curvature of

the planet into account. We tested for correlations between genetic differentiation and

geographic distance using Mantel tests using the *R* package *ade4* (v.1.7-8., Dray & Dufour

2007) with 1,000,000 iterations. In addition, we plotted the 5% smallest and 5% largest F_{ST}

values from all 1,128 pairwise comparisons among the 48 population samples onto a map

to visualize geographic patterns of genetic differentiation. From these putatively neutral

SNPs, we used observed F_{ST} on the autosomes (F_{aut}) to calculate the expected F_{ST} on X

chromosomes (F_X) as in Machado *et al*. (2016) using the equation

1190

1191 $$F_X = 1 - \left[\frac{9(z+1)*(1-F_{aut})}{8(2z+1)-(1-F_{aut})*(7z-1)}\right]$$

1192

1193   where z is the ratio of effective population sizes of males (N_m) and females (N_f), N_m/N_f

1194   (Ramachandran *et al.* 2004). For the purposes of this study we assume $z = 1$.

1195

We further investigated genetic variation in our dataset by principal component analysis

(PCA) based on allele frequencies of the neutral marker SNPs described above. We used

the *R* package *LEA* (v. 1.2.0., Frichot *et al.* 2013) and performed PCA on unscaled allele

frequencies as suggested by Menozzi *et al*. (1978) and Novembre and Stephens (2008).

We focused on the first three principal components (PCs) and employed a model-based

approach as implemented in the *R* package *mclust* (v. 5.2., Fraley & Raftery 2012) to

identify the most likely number of clusters based on maximum likelihood and assigned

population samples to clusters by *k*-means clustering in *R* (R Development Core Team

2009). Finally, we examined the first three PCs for correlations with latitude, longitude,

altitude, and season using general linear models and tested for spatial autocorrelation as

described above. A Bonferroni-corrected α threshold ($\alpha' = 0.05/3 = 0.017$) was used to

account for multiple testing.

1208

Mitochondrial DNA

To obtain consensus mitochondrial sequences for each of the 48 European populations,

reads from individual FASTQ files were aligned and minor variants replaced by the major

variant using *Coral* (Salmela & Schröder 2011). This way, ambiguities that might prevent

the growth of contigs from reads during the assembly process can be eliminated. For each

population, a genome assembly was obtained using *SPAdes* using standard parameters

and *k*-mers of size 21, 33, 55, and 77 (Bankevich *et al.* 2012) and the corrected FASTQ

files. Mitochondrial contigs were retrieved by blastn, using the *D. melanogaster* NC

024511 sequence as a query and each genome assembly as the database. To avoid

nuclear mitochondrial DNA segments (numts), we ensured that only contigs with a much

higher coverage than the average coverage of the genome were retrieved. When multiple

contigs were available for the same region, the one with the highest coverage was

selected. Possible contamination with *D. simulans* was assessed by looking for two or

more consecutive sites that show the same variant as *D. simulans* and looking for

alternative contigs for that region with similar coverage. As an additional quality control

measure, we also examined the presence of pairs of sites showing four gametic types

using *DNAsp* 6 (Rozas *et al.* 2017) – given that there is no recombination in mitochondrial

DNA no such sites are expected. The very few sites presenting such features were

rechecked by looking for alternative contigs for that region and were corrected if needed.

The uncorrected raw reads for each population were mapped on top of the different

consensus haplotypes using *Express* as implemented in *Trinity* (Grabherr *et al.* 2011). If

most reads for a given population mapped to the consensus sequence derived for that

population the consensus sequence was retained, otherwise it was discarded as a

possible chimera between different mitochondrial haplotypes. The repetitive mitochondrial

hypervariable region is difficult to assemble and was therefore not used; the mitochondrial

region was thus analyzed as in Cooper *et al*. (2015). Mitochondrial genealogy was

estimated using statistical parsimony (TCS network; Clement *et al.* 2000), as implemented

in *PopArt* (http://popart.otago.ac.nz), and the surviving mitochondrial haplotypes.

Frequencies of the different mitochondrial haplotypes were estimated from FPKM values

using the surviving mitochondrial haplotypes and expressed as implemented in *Trinity*

(Grabherr *et al.* 2011).

1240

Transposable elements

To quantify the transposable element (TE) abundance in each sample, we assembled and

quantified the repeats from unassembled sequenced reads using *dnaPipeTE* (v.1.2. ,

Goubert *et al.* 2015). The vast majority of high-quality trimmed reads were longer than 135

bp. We thus discarded reads less than 135 bp before sampling. Reads matching mtDNA

were filtered out by mapping to the *D. melanogaster* reference mitochondrial genome

(NC_024511.2. 1) with *bowtie2* (v. 2.1.0., Langmead & Salzberg 2012). Prokaryotic

sequences, including reads from symbiotic bacteria such as *Wolbachia*, were filtered out

from the reads using the implementation of *blastx* (translated nucleic vs. protein database)

vs. the non-redundant protein database (nr) using *DIAMOND* (v. 0.8.7., Buchfink *et al.*

2015). To quantify TE content, we subsampled a proportion of the raw reads (after

filtering) corresponding to a genome coverage of 0.1X (assuming a genome size of 175

MB), and then assembled these reads with *Trinity* assembler (Grabherr *et al.* 2011). Due

to the low coverage of the genome obtained with the subsampled reads, only repetitive

DNA present in multiple copies should be fully assembled (Goubert *et al.* 2015). We

repeated this process with three iterations per sample, as recommended by the program

guidelines, to assess the repeatability of the estimates.

1258

We further estimated frequencies of previously characterized TEs present in the reference

genome with *T-lex2* (v. 2.2.2., Fiston-Lavier *et al.* 2015), using all annotated TEs (5,416

TEs) in version 6.04 of the *D. melanogaster* genome from flybase.org (Gramates *et al.*

2017). For 108 of these TEs, we used the corrected coordinates as described in Fiston-

Lavier *et al*. (2015), based on the identification of target site duplications at the site of the

insertion. We excluded TEs nested or flanked by other TEs (<100 bp on each side of the

TE), and TEs which are part of segmental duplications, since *T-lex2* does not provide

accurate frequency estimates in complex regions (Fiston-Lavier *et al.* 2015). We

additionally excluded the INE-1 TE family, as this TE family is ancient, with thousands of

insertions in the reference genome, which appear to be mostly fixed (2,234 TEs; Kapitonov

& Jurka 2003).

1270

After applying these filters, we were able to estimate frequencies of 1,630 TE insertions

from 113 families from the three main orders, LTR, non-LTR, and DNA across all *DrosEU*

samples. *T-lex2* contains three main modules: (i) the presence detection module, (ii) the

absence detection module, and (iii) the combine module, which joins the results from the

former two detection modules. In the presence module, *T-lex2* uses *Maq* (v. 0.7.1., Li *et al.*

2008) for the mapping of reads. As *Maq* only accepts reads 127 bp or shorter, we cut the

trimmed reads following the general pipeline (Figure 3) and then used *Trimmomatic* (v.

0.35; Bolger *et al.* 2014) to cut trimmed reads longer than 100 bp into two equally sized

fragments using CROP and HEADCROP parameters. Only the presence module was run

with the cut reads.

1281

To avoid inaccurate TE frequency estimates due to very low numbers of reads, we only

considered frequency estimates based on at least 3 reads. Despite the stringency of *T-*

lex2 to select only high-quality reads, we additionally discarded frequency estimates

supported by more than 90 reads, i.e. 3 times the average coverage of the sample with the

lowest coverage (CH_Cha_14_43, Table 1), in order to avoid non-unique mapping reads.

This filtering allows to estimate TE frequencies for ~96% (92.9% to 97.8%) of the TEs in

each population. For 85% of the TEs, we were able to estimate their frequencies in more

than 44 out of 48 *DrosEU* samples.

1290

We tested for correlations between TE insertion frequencies and recombination rates

using Spearman's rank correlations as implemented in *R*. For SNPs, we used

recombination rates from Comeron *et al.* (2012) and from the Recombination Rate

Calculator (Fiston-Lavier *et al.* 2010) in non-overlapping 100 kb windows, and assigned to

each TE insertion the recombination rate of the corresponding 100 kb genomic window.

To test for spatio-temporal variation of TE insertions, we excluded TEs with an interquartile

range (IQR) < 10. There were 141 TE insertions with variable population frequencies

among the *DrosEU* samples. We tested the population frequencies of these insertions for

correlations with latitude, longitude, altitude, and season using generalized linear models

(ANCOVA) following the method used for SNPs but with a binomial error structure in *R*.

We also tested for residual spatio-temporal autocorrelations, with Moran's *I* test (Moran

1950; Kühn & Dormann 2012). We used Bonferroni corrections to account for multiple

testing (α'= 0.05/141 = 0.00035) and only considered Bonferroni-corrected p-values <

0.001 to be significant. TEs with a recombination rate that differed from 0 cM/Mb according

to both used measures (see above) were considered as high recombination regions. To

test TE family enrichment among the significant TEs we performed a χ^2 test and applied

Yate's correction to account for the low number of some of the cells.

1308

Inversion polymorphisms

Since Pool-Seq data precludes a direct assessment of the presence and frequencies of

chromosomal inversions, we indirectly estimated inversion frequencies using a panel of

approximately 400 inversion-specific marker SNPs (Kapun *et al.* 2014) for six

cosmopolitan inversions (*In(2L)t, In(2R)NS, In(3L)P, In(3R)C, In(3R)Mo, In(3R)Payne*). We

averaged allele frequencies of these markers in each sample separately. To test for clinal

variation in the frequencies of inversions, we tested for correlations with latitude, longitude,

altitude and season using generalized linear models with a binomial error structure in *R* to

account for the biallelic nature of karyotype frequencies. In addition, we tested for residual

spatio-temporal autocorrelations as described above and Bonferroni-corrected the α

threshold (α'= 0.05/7 = 0.007) to account for multiple testing.

1320

Microbiome

Raw sequences were trimmed and quality filtered as described for the genomic data

analysis. The remaining high quality sequences were mapped against the *D. melanogaster*

genome (v.6.04) including mitochondria using bbmap (v. 35; Bushnell 2016) with standard

settings. The unmapped sequences were submitted to the online classification tool,

MGRAST (Meyer *et al.* 2008) for annotation. Taxonomy information was downloaded and

analyzed in *R* (v. 3.2.3; R Development Core Team 2009) using the matR (v. 0.9;

Braithwaite & Keegan) and RJSONIO (v. 1.3; Lang) packages. Metazoan sequence

features were removed. For microbial load comparisons, the number of protein features

identified by MGRAST for each taxon and sample was divided by the number of

sequences that mapped to *D. melanogaster* chromosomes *X, Y, 2L, 2R, 3L, 3R* and *4*.

1332

We also surveyed the datasets for the presence of novel DNA viruses by performing *de*

novo assembly of the non-fly reads using SPAdes 3.9.0 (Bankevich *et al.* 2012), and using

conceptual translations to query virus proteins from Genbank using DIAMOND 'blastp'

(Buchfink *et al.* 2015). In three cases (*Kallithea* virus, *Vesanto* virus, *Viltain* virus), reads

from a single sample pool were sufficient to assemble a (near) complete genome. In two

other cases, fragmentary assemblies allowed us to identify additional publicly available

datasets that contained sufficient reads to complete the genomes (*Linvill Road* virus,

Esparto virus; completed using SRA datasets SRR2396966 and SRR3939042,

respectively). Novel viruses were provisionally named based on the localities where they

were first detected, and the corresponding novel genome sequences were submitted to

Genbank (KX130344, KY608910, KY457233, KX648533-KX648536). To assess the

relative amount of viral DNA, unmapped (non-fly) reads from each sample pool were

mapped to repeat-masked *Drosophila* DNA virus genomes using bowtie2, and coverage

normalized relative to virus genome length and the number of mapped *Drosophila* reads.

1347

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1355

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1409

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1411

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1946

1947   **Tables**

1948

1949   **Table 1. Sample information for all populations in the *DrosEU* dataset.** The table shows the origin, collection data and season and

1950   sample size (number of chromosomes: *n*) of the 48 samples in the *DrosEU* dataset. Additional information can be found in the supporting

1951   information in Table S1.

ID	Country	Location	Coll. Date	Number ID	Lat (°)	Lon (°)	Alt (m)	Season	n	Coll. name
AT_Mau_14_01	Austria	Mauternbach	2014-07-20	1	48.38	15.56	572	S	80	Andrea J. Betancourt
AT_Mau_14_02	Austria	Mauternbach	2014-10-19	2	48.38	15.56	572	F	80	Andrea J. Betancourt
TR_Yes_14_03	Turkey	Yesiloz	2014-08-31	3	40.23	32.26	680	S	80	Banu Sebnem Onder
TR_Yes_14_04	Turkey	Yesiloz	2014-10-23	4	40.23	32.26	680	F	80	Banu Sebnem Onder
FR_Vil_14_05	France	Viltain	2014-08-18	5	48.75	2.16	153	S	80	Catherine Montchamp-Moreau
FR_Vil_14_07	France	Viltain	2014-10-27	7	48.75	2.16	153	F	80	Catherine Montchamp-Moreau
FR_Got_14_08	France	Gotheron	2014-07-08	8	44.98	4.93	181	S	80	Cristina Vieira
UK_She_14_09	United Kingdom	Sheffield	2014-08-25	9	53.39	-1.52	100	S	80	Damiano Porcelli
UK_Sou_14_10	United Kingdom	South Queensferry	2014-07-14	10	55.97	-3.35	19	S	80	Darren Obbard
CY_Nic_14_11	Cyprus	Nicosia	2014-08-10	11	35.07	33.32	263	S	80	Eliza Argyridou
UK_Mar_14_12	United Kingdom	Market Harborough	2014-10-20	12	52.48	-0.92	80	F	80	Eran Tauber
UK_Lut_14_13	United Kingdom	Lutterworth	2014-10-20	13	52.43	-1.10	126	F	80	Eran Tauber

DE_Bro_14_14	Germany	Broggingen	2014-06-26	14	48.22	7.82	173	S	80	Fabian Staubach
DE_Bro_14_15	Germany	Broggingen	2014-10-15	15	48.22	7.82	173	F	80	Fabian Staubach
UA_Yal_14_16	Ukraine	Yalta	2014-06-20	16	44.50	34.17	72	S	80	Iryna Kozeretska
UA_Yal_14_18	Ukraine	Yalta	2014-08-27	18	44.50	34.17	72	S	80	Iryna Kozeretska
UA_Ode_14_19	Ukraine	Odesa	2014-07-03	19	46.44	30.77	54	S	80	Iryna Kozeretska
UA_Ode_14_20	Ukraine	Odesa	2014-07-22	20	46.44	30.77	54	S	80	Iryna Kozeretska
UA_Ode_14_21	Ukraine	Odesa	2014-08-29	21	46.44	30.77	54	S	80	Iryna Kozeretska
UA_Ode_14_22	Ukraine	Odesa	2014-10-10	22	46.44	30.77	54	F	80	Iryna Kozeretska
UA_Kyi_14_23	Ukraine	Kyiv	2014-08-09	23	50.34	30.49	179	S	80	Iryna Kozeretska
UA_Kyi_14_24	Ukraine	Kyiv	2014-09-08	24	50.34	30.49	179	F	80	Iryna Kozeretska
UA_Var_14_25	Ukraine	Varva	2014-08-18	25	50.48	32.71	125	S	80	Oleksandra Protsenko
UA_Pyr_14_26	Ukraine	Pyriatyn	2014-08-20	26	50.25	32.52	114	S	80	Oleksandra Protsenko
UA_Dro_14_27	Ukraine	Drogobych	2014-08-24	27	49.33	23.50	275	S	80	Iryna Kozeretska
UA_Cho_14_28	Ukraine	Chornobyl	2014-09-13	28	51.37	30.14	121	F	80	Iryna Kozeretska
UA_Cho_14_29	Ukraine	Chornobyl Yaniv	2014-09-13	29	51.39	30.07	121	F	80	Iryna Kozeretska
SE_Lun_14_30	Sweden	Lund	2014-07-31	30	55.69	13.20	51	S	80	Jessica Abbott
DE_Mun_14_31	Germany	Munich	2014-06-19	31	48.18	11.61	520	S	80	John Parsch
DE_Mun_14_32	Germany	Munich	2014-09-03	32	48.18	11.61	520	F	80	John Parsch
PT_Rec_14_33	Portugal	Recarei	2014-09-26	33	41.15	-8.41	175	F	80	Jorge Vieira

ES_Gim_14_34	Spain	Gimenells (Lleida)	2014-10-20	34	41.62	0.62	173	F	80	Lain Guio
ES_Gim_14_35	Spain	Gimenells (Lleida)	2014-08-13	35	41.62	0.62	173	S	80	Lain Guio
FI_Aka_14_36	Finland	Akaa	2014-07-25	36	61.10	23.52	88	S	80	Maaria Kankare
FI_Aka_14_37	Finland	Akaa	2014-08-27	37	61.10	23.52	88	S	80	Maaria Kankare
FI_Ves_14_38	Finland	Vesanto	2014-07-26	38	62.55	26.24	121	S	66	Maaria Kankare
DK_Kar_14_39	Denmark	Karensminde	2014-09-01	39	55.95	10.21	15	F	80	Mads Fristrup Schou
DK_Kar_14_41	Denmark	Karensminde	2014-11-25	41	55.95	10.21	15	F	80	Mads Fristrup Schou
CH_Cha_14_42	Switzerland	Chalet à Gobet	2014-07-24	42	46.57	6.70	872	S	80	Martin Kapun
CH_Cha_14_43	Switzerland	Chalet à Gobet	2014-10-05	43	46.57	6.70	872	F	80	Martin Kapun
AT_See_14_44	Austria	Seeboden	2014-08-17	44	46.81	13.51	591	S	80	Martin Kapun
UA_Kha_14_45	Ukraine	Kharkiv	2014-07-26	45	49.82	36.05	141	S	80	Svitlana Serga
UA_Kha_14_46	Ukraine	Kharkiv	2014-09-14	46	49.82	36.05	141	F	80	Svitlana Serga
UA_Cho_14_47	Ukraine	Chornobyl Applegarden	2014-09-13	47	51.27	30.22	121	F	80	Svitlana Serga
UA_Cho_14_48	Ukraine	Chornobyl Polisske	2014-09-13	48	51.28	29.39	121	F	70	Svitlana Serga
UA_Kyi_14_49	Ukraine	Kyiv	2014-10-11	49	50.34	30.49	179	F	80	Svitlana Serga
UA_Uma_14_50	Ukraine	Uman	2014-10-01	50	48.75	30.21	214	F	80	Svitlana Serga
RU_Val_14_51	Russia	Valday	2014-08-17	51	57.98	33.24	217	S	80	Elena Pasyukova

1952

1953

1954 **Table 2**. **Clinality of genetic variation and population structure.** Effects of geographic variables and/or seasonality on genome-wide

1955 average levels of diversity (π, θ and Tajima's D; top rows) and on the first three axes of a PCA based on allele frequencies at neutrally

1956 evolving sites (bottom rows). The values represent F-ratios from general linear models. Bold type indicates F-ratios that are significant

1957 after Bonferroni correction (adjusted α'=0.0055). Asterisks in parentheses indicate significance when accounting for spatial

1958 autocorrelation by spatial error models. These models were only calculated when Moran's I test, as shown in the last column, was

1959 significant. *$p < 0.05$; **$p < 0.01$; ***$p < 0.001$.

1960

Factor	Latitude	Longitude	Altitude	Season	Moran's I
$\pi_{(X)}$	4.11*	1.62	**15.23***	1.65	0.86
$\pi_{(Aut)}$	0.91	2.54	**27.18***	0.16	-0.86
$\theta_{(X)}$	2.65	1.31	**15.54***	2.22	0.24
$\theta_{(Aut)}$	0.48	1.44	**13.66***	0.37	-1.13
$D_{(X)}$	0.02	0.38	5.93*	3.26	-2.08
$D_{(Aut)}$	0.09	0.76	5.33*	0.71	-1.45
PC1	0.06	**120.72***(***)**	5.35*(*)	2.53	4.15***
PC2	3.5	**10.22***	**15.21***	1.97	-1.96

		PC3	0.14	0.11		0.01	1.29	0.22

1961

1962

1963 **Table 3**. **Clinality and/or seasonality of chromosomal inversions.** The values represent *F*-ratios from generalized linear models with

1964 a binomial error structure to account for frequency data. Bold type indicates deviance values that were significant after Bonferroni

1965 correction (adjusted $α'=0.0071$). Stars in parentheses indicate significance when accounting for spatial autocorrelation by spatial error

1966 models. These models were only calculated when Moran's *I* test, as shown in the last column, was significant. *$p < 0.05$; **$p < 0.01$; ***p

1967 < 0.001

1968

Factor	*Latitude*	*Longitude*	*Altitude*	*Season*	*Moran's I*
In(2L)t	2.2	**10.09****	**43.94*****	0.89	-0.92
In(2R)NS	0.25	**14.43*****	2.88	2.43	1.25
In(3L)P	**21.78*****	2.82	0.62	3.6	-1.61
In(3R)C	**18.5***(***)**	0.75	1.42	0.04	2.79**
In(3R)Mo	0.3	0.09	0.35	0.03	-0.9
In(3R)Payne	**43.47*****	0.66	1.69	1.55	-0.89

1969

1970