Published estimates of group differences in multisensory integration are inflated

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Acknowledgments: This research was supported by NIH R01NS065395 to MSB.

Abstract

A common measure of multisensory integration is the McGurk effect, an illusion in which

incongruent auditory and visual speech are integrated to produce an entirely different

percept. Published studies report that participants who differ in age, gender, culture, native

language, or traits related to neurological or psychiatric disorders also differ in their

susceptibility to the McGurk effect. These group-level differences are used as evidence for

fundamental alterations in sensory processing between populations. However, there is high

variability in the McGurk effect *within* groups: some participants never report the illusion while

others always do. High *within-group* variability means that sample sizes much larger than those

in the published literature are necessary to accurately estimate *between-group* differences. Using

empirical data and simulations, we show that a typical study with a sample size of 15 participants

per group and a true group difference of 10% will report a group difference of 31%, an effect

inflation of three-fold. Reducing the effect inflation by half requires increasing the sample size

ten-fold, to 150 participants per group. The combination of within-group variability and small

sizes explains two prominent features of the McGurk literature. First, group differences are

inflated and highly variable across studies. Second, follow-up studies fail to replicate published

reports of large between-group susceptibility differences because initial effect-size estimates are

driven by statistical inflation rather than true group differences. These replication failures are

especially problematic in studies of clinical populations because they have real-world

consequences. Due to the high within-group variability, caution is essential when using the

McGurk effect to assess deficits in clinical groups or propose treatment plans, such as training to

improve sensory processing.

Introduction

Since different sensory modalities carry distinct sources of information about the world, integrating them provides a more reliable picture of the world. The underlying computations and behavioral manifestations of this multisensory integration have received increasing attention. Researchers have developed a variety of measures to assess multisensory integration, focused on a change in perception due to the addition of a second sensory modality. For instance, in speech perception, integrating auditory information from the talker's voice and visual information from the talker's face enhances speech recognition accuracy[1-3]. Perhaps the most common, single measure of multisensory integration is the McGurk effect, an illusion in which incongruent auditory and visual speech are integrated to produce a new percept that matches neither of the component modalities. The original report of the McGurk effect[4] has been cited thousands of times and the illusion is used as a textbook example of multisensory integration. As a behavioral assay, susceptibility to the McGurk effect has been used to argue for important differences in multisensory integration between genders[5], in atypical human development[4, 6], across the lifespan[7], in mental health disorders[8, 9], and between cultural/linguistic groups[10, 11]. Amidst the enthusiasm for using the McGurk effect to study such between-group differences, recent work has highlighted large within-group variability in susceptibility to the illusion using relatively homogeneous subject pools[12, 13]. Using empirical data and simulations, we show that high within-group variability, coupled with routine use of small sample sizes, leads to a proliferation of inflated estimates of group differences in multisensory integration. Our results explain why recent studies have failed to replicate the large reported differences between cultures[14], genders[12], and children with developmental disorders[15]. Studies of multisensory integration must increase sample sizes by an order of magnitude to produce

accurate and reliable estimates of group differences. Continued publication of inaccurate group

differences can lead to vague, unscientific theories(16) and, worse, unscientific treatments aimed

at ameliorating misunderstood clinical deficits.

To demonstrate how small studies can produce inflated effect estimates, we modeled the

consequences of studying group differences in multisensory integration with at a variety of

sample sizes. This modeling effort is made possible by a recent large-scale, in-person behavioral

study, that can accurately sample the true variability within a large population of homogeneous,

young healthy college undergraduates(12). For other situations, such as comparisons with

clinical populations, the within-group variance will be at least as large as for college

undergraduates. Therefore, these simulations provide a lower bound on the sample sizes

necessary to study group differences. Using empirical data rather than off-the-shelf power

analysis avoids assumptions, such as population normality, that are demonstrably false for the

McGurk effect(12).

Effect inflation measured with a known true effect

To examine how sample size can influence experimental results, we used bootstrapped

data from a large behavioral study and simulated population differences in McGurk

susceptibility, defined as the mean percentage of fusion responses to McGurk stimuli across all

possible stimuli and subjects (see *Methods* for details). In the first example, we created two

populations: population A has susceptibility 45% while population B has susceptibility 55%, and

thus a mean difference of 10% (Figure 1A). How likely is it for a given experiment to accurately

estimate this group difference? Figure 1B shows an example experiment in which 150 subjects

are sampled from each population for a total sample size of 300. With this large sample size,

Figure 1. Small sample sizes lead to inflated estimates of population differences. **A**. Two example populations with McGurk susceptibility of 45% (Population A, orange) and 55% (Population B, blue), producing a true difference in McGurk susceptibility of 10%. **B**. A representative study using a sample of 300 subjects, 150 from each population. The estimated population difference is 12%, close to the true difference. **C**. A representative study using a typical sample size of 30 subjects (15 from each population). The estimated population difference is 31%, far from the true difference. **D**. Long run distribution of population difference estimates from studies with $N = 300$. Across all studies, the mean effect estimate is accurate (mean of 10%, gray bars). Considering only studies with significant results (blue bars) inflates the mean estimate to 13% because only 67% of the studies are considered (power = 67%). **E**. Long run distribution of population difference estimates from studies with $N = 30$. Considering only studies with significant results inflates the mean estimate to 30%, three times higher than the true difference on 10%.

susceptibilities are estimated at 48% ± 3% (standard error of the mean) for population A and

60% ± 3% for population B, resulting in a difference estimate of 12% ± 9% (95% confidence

interval), a reasonable approximation to the true population difference of 10%.

However, a sample size of N = 300 is much larger than that used in published studies of

multisensory integration. Figure 1C shows an example experiment with a more typical sample

size of N = 30. This experiment results in estimated susceptibilities of 36% ± 9% for population

A and 67% ± 8% for population B. The estimated population difference in this experiment is

31% ± 24%, greatly inflated from the actual difference of 10%.

This effect inflation occurs when experimenters only consider experiments that result in a

significant population difference, usually defined as a between-groups t-test producing $p < 0.05$;

this criterion can be instantiated in many ways in science, most often in the manuscript

preparation and submission phase in which only significant results are included(17-19). Figure

1D shows the result of implementing this significance filter on thousands of simulated

experiments with a sample size of 300. Of these experiments, 67% result in significant

differences between populations and are "published"; the other 33% are discarded. The discarded

studies have population difference estimates that are always *smaller* than the true difference

(mean of 5%; 100% of estimates less than 10%; gray bars in Fig 1D) while the significant studies

have population difference estimates that are usually *greater* than the true difference (mean of

13%; 77% of estimates greater than 10%; blue bars in Figure 1E). A weighted average of the

significant and discarded experiments recovers the true population difference of 10% (13% ×

0.67 + 5% × 0.32= 10%), but examining only the significant experiments biases the population

difference estimate upwards by a factor of 1.3.

Figure 1E shows the results of implementing a significance filter on thousands of

simulated experiments with sample size of 30. Only 13% of experiments result in significant

differences between populations; the remaining 87% are discarded. The discarded studies have

population estimates that are close to the true difference (mean of 7%; gray bars in Figure 1E),

while the included studies have population difference estimates that are much greater than the

true difference (mean of 30%; blue bars in Figure 1E). Considering only significant experiments

inflates the population difference estimates by a factor of 3.

A key point is that small sample sizes not only inflate the mean effect estimate, but also

increase effect estimate variance from study to study. For any particular $N = 30$ study with a true

population difference of 10%, the difference estimate will vary from -16% to 36% (two standard

deviations around the true difference; gray distribution in Figure 1E). This variance in effect

estimation is large enough that in 3% of statistically significant studies, the population difference

will be in the wrong direction: experiments will incorrectly conclude that McGurk susceptibility

is higher in group A, rather than the true effect of higher susceptibility in group B. In contrast,

across all $N = 300$ studies, every statistically significant study estimates the population difference

in the correct direction.

Experimental manipulations to counter effect inflation

The first example demonstrated that a true difference between populations of 10% will be

vastly overestimated using procedures that are common in the literature: sample sizes of 30 and a

statistical significance filter. Since true effect sizes are rarely known in advance, and are often a

motivator for performing the experiment, we next examined effect size inflation at a range of

true population differences. As shown in Figure 2A, effect size inflation is most extreme when

population differences are small. For a true population difference of 5%, effect size inflation

Figure 2. Factors contributing to effect inflation and statistical power in studies of population differences in McGurk susceptibility. **A**. Effect inflation decreases as sample size increases for a fixed population difference. Doubling the known population difference (5%, 10%, 20%, 30%; blue lines) halves effect inflation at a given sample size. **B**. Effect inflation decreases with increasing number of stimuli (1, 2, 4, 8, 16; red lines), shown for a fixed population difference of 10%. **C**. Increasing the number of trials (10, 100, 1000; green lines) has little impact on effect inflation. **D**. Statistical power (ability to detect a non-zero population difference) increases with increasing sample size and increasing population difference. **E**. Increasing the number of stimuli produces diminishing returns for increasing statistical power. **F**. Increasing the number of trials has minimal effect on statistical power.

reaches over 7-fold for $N = 20$. In other words, on average, experiments with $N = 20$ (10 in each

group) that select subjects from populations with true difference of 5% will publish effect

estimates of over 35%.

In general, increasing the sample size and increasing the true population effect decrease

effect inflation severity. However, for small population differences even very large sample sizes

do not eliminate effect inflation: for a 5% population difference, effect inflation was more than

two-fold even for sample sizes of 300. Conversely, for small N even a large population

difference does not eliminate effect inflation. For a very large population difference of 20%, a

sample size of 20 still results in two-fold effect inflation.

What about other experimental manipulations? There is considerable procedural variation

in the number of stimuli used to assess the McGurk effect, with some using a single stimulus,

and others many more(20, 21). To examine how stimulus count influences effect inflation, we

modelled the effect of increasing stimuli for a true population difference of 10% (Figure 2B). In

general, increasing the number of stimuli reduces effect inflation. The effect is most pronounced

when increasing the number of stimuli from 1 to 2, which reduces the effect inflation from 1.8 to

1.6 at $N = 100$. Further increases in the number of stimuli produce diminishing returns, to an

effect inflation of 1.2 for 16 stimuli.

Another experimental approach is to present only a single stimulus but increase the

number of trials. A typical McGurk experiment may present 10 repetitions of a single stimulus to

a subject. As shown in Figure 2C, even a vast increase in the number of trials, from 10 to 1000

results in little or no decrease in effect inflation, regardless of the number of subjects tested.

Relationship between effect inflation and statistical power

The statistical power of an experiment is defined as the probability that an experiment

with a known true effect size and given sample size will result in a significant *p*-value. Power

calculations implicitly incorporate a statistical significance filter, as they assume that the

experimenter's goal is to detect a difference between populations with a significance of $p < 0.05$,

rather than to accurately estimate the group difference in the long run. A power of 80% is often

used as a benchmark in the literature(22), as it ensures that only one in five experiments will be

discarded for not being significant (assuming that the true population difference is known).

As shown in Figure 2D, large sample sizes are required for adequate statistical power

across the range of population differences. Even for a very large difference between-population

difference of 20% an N of 110 is required to achieve 80% power. For a population difference of

5% and a sample size of N of 300, power is only 23%.

Increasing the number of stimuli can increase power (Figure 2E), but this increase is not

enough to fully overcome small sample sizes or a small population differences. For a population

difference of 10%, the sample size required to achieve 80% power decreases from $N = 450$ with

one stimulus, to $N = 300$ with 2 stimuli, to $N = 205$ with 16 stimuli.

Changing the number of trials has minimal effect on statistical power (Figure 2F). For a

single stimulus and a population difference of 10%, 450 subjects are required to achieve 80%

power with 10 trials; increasing to 1000 trials only reduces the required number of subjects to

405.

Relevance to the published literature on multisensory integration

Autism spectrum disorder (ASD) is frequently referred to as a disorder in which patients

have impaired ability to integrate information across modalities(23). This consensus is based on

behavioral studies comparing people with ASD and healthy controls. Figure 3 shows a summary

of 13 of these studies, with the difference in McGurk susceptibility between ASD and healthy

controls plotted against the sample size of the studies. There is wide variation across studies,

Figure 3. Reported population difference in McGurk susceptibility between children with autism spectrum disorder (ASD) and children with typical development (TD). Each symbol represents a single published study. Statistically significant ($p < 0.05$) difference are colored in blue, non-significant gray. Positive values indicate more McGurk susceptibility in the TD group *vs*. the ASD group. Mean across significant studies is 27% (blue dashed line).

72   ranging from 45% *less* integration in ASD ($N = 34$;(24)) to 10% *more* integration in ASD ($N = 36$;(25)). These differences across studies have been attributed to interactions between clinical diagnosis and other factors, including stimulus, gender, temporal processing abilities, or participant age. For instance, one group reported a population difference for younger children but not older children(26) while another group reported the exact opposite effect(27). We suggest that a more likely explanation than a variety of *post hoc* moderators is that small sample sizes inevitably result in high variability in population effect estimates, causing wild swings in their magnitude and sign across studies(28, 29). This contention is reinforced by noting that the ASD study with the largest sample size ($N = 76$) also reported the smallest effect, a population difference of only 3%(15). Assuming that this large study is a more accurate measure of the true effect size, we can ask what the expected effect size will be for studies with a more typical sample size of $N = 36$ (mean across 13 ASD studies). Our simulations show that applying a significance-filter to studies of this size will inflate the population difference around nine-fold, resulting in an effect estimate of 28%, similar to the observed significance-filtered across-study mean of 27%.

A similar pattern is observed in studies of cultural differences in multisensory integration,

which compare the prevalence of the McGurk effect between different linguistic or cultural

populations. For instance, a number of studies have considered differences in McGurk

susceptibility between speakers of tonal (*e.g.*, Mandarin) *vs.* non-tonal (*e.g.*, English) languages.

Using small samples sizes, effect estimates have varied from +36% (*more* McGurk for non-tonal

speakers; $N = 24$; (30)) to +17% ($N = 48$; (10)) to -8% (*less* McGurk for non-tonal speakers; $N =$

40; (31)). In contrast, a study with a large sample size found the smallest estimated difference: -

4% (less McGurk for non-tonal speakers; $N = 376$; (14)).

Although these examples are taken from studies of population-level differences, the same

problems arise in studies in which otherwise identical subjects receive different experimental

manipulations, such as viewing different stimuli. A recent review(32) of 27 studies of the

McGurk Effect found a median sample size of $N = 25$.

Is this problem restricted to studies of multisensory integration?

Because a hallmark of many cognitive functions, including multisensory integration, is

individual variation(33), large sample sizes are necessary to accurately measure a population

difference or the effect of an experimental manipulation. Inadequate sample sizes are not unique

to studies of multisensory integration: both the neuroscience(34) and psychology(35) literatures

suffer from the same problem. Studies with small sample sizes over-estimate true effects, leading

investigators in fruitless pursuit of the source of "large" effects despite failures to replicate(36).

While replication failures are often attributed to *post hoc* moderators or to the inevitable

experimental differences between studies, our results show that a proliferation of small studies

cannot resolve conflicting results. Instead, highly-powered studies aimed at accurate estimation

are the only way to provide rigorous, reliable and reproducible studies of human behavior.

Method

Bootstrap procedure to estimate effect inflation and statistical power

We used a bootstrapping procedure to create hypothetical replication datasets based on a

large behavioral dataset ($N = 165$) collected in-person from Rice University undergraduate

students and described previously(12). The goal of the simulation was to determine how

experimental design choices impacted statistical power (ability to reject the null hypothesis) and

effect estimation (accuracy of mean estimates from studies that reject the null hypothesis). We

conducted the simulations using R(37); source code is available on the authors' website:

http://www.openwetware.org/wiki/Beauchamp:DataSharing.

The simulations proceeded in a series of 7 steps:

1. Simulation parameters were set. N: number of participants (20, 30, 40, 50, 100, 150, 200, or

300), E: size of the population difference (5%, 10%, 20%, and 30%), S: number of stimuli (1,

2, 4, 8 or 16), T: the number of trials for each stimulus (10, 100, or 1000).

2. McGurk perception rates, pM, for N participants at S stimuli were sampled with replacement

from the empirical data and stored in a new dataset, D. This procedure ensures realistic

within-subject correlations across stimuli.

3. Participants in D were randomly assigned to group A or B, ensuring equal group sizes (group

size = $N / 2$).

4. To produce the population effect, the values of pM was shifted by E for all participants

assigned to group B. Because of ceiling effects, the actual size of the shift needed to be

greater than the desired mean population difference. We determined these values via

simulation using sample sizes of 35,000 and stimulus count of 200. Actual shift values:

9.45%, 17.25%, 33%, and 55% produce population differences of 5%, 10%, 20%, and 30%,

respectively. To ensure the simulation remained stochastic, *pM* was truncated to [5%, 95%]

for all participants in *D*.

5. For each participant in *D*, we calculated an *observed* McGurk perception rate *pF* for each

stimulus in *S* by sampling from a binomial distribution with *T* trials and true proportion *pM*.

6. A hypothesis test was conducted to compare the *pF* between groups A and B. For single-

stimulus experiments, we used a two-sample, equal-variance *t*-test. For multi-stimulus

experiment, we first averaged across stimuli within each subject, and then used a *t*-test. We

used a *t*-test rather than a linear mixed-effects model here for computational efficiency.

Because of how the population effect was created (a fixed shift for all participants and all

stimuli), results for the LME and *t*-test were similar.

7. To obtain long-run behavior, we repeated these steps 35000 times for each parameter

combination.

To summarize the results of the simulations we used two summary measures: Statistical

Power: the proportion of simulations that rejected the null hypothesis, and Effect Inflation: the

mean ratio of the estimated effect magnitude (absolute value of the mean difference between

groups) and the true effect magnitude, for the subset of the simulations that rejected the null

hypothesis, also called the expected Type M error or exaggeration ratio(28, 29). Effect inflation

provides a quantitative measure of the impact of the statistical significance filter on population

effect estimates.

Summary of population differences in children with autism spectrum disorder

To assess the impact of small sample sizes in an important area of multisensory

integration research, we reviewed studies that compared McGurk susceptibility between children

with autism spectrum disorder (ASD) and children with typical development (TD). We included

all studies for which group means and sample sizes were available. For studies with multiple

group-level comparisons, we present them as separate data points. For graphing purposes, the

points are jittered (uniform distribution between -1 to 1) in each direction to prevent overlap. A

list of the studies used and the data collected are available from the authors' website:

openwetware.org/Beauchamp:DataSharing.

References

1.   Sumby WH & Pollack I (1954) Visual contribution to speech intelligibility in noise. *J*

Acoust Soc Am 26(2):212-215.

2.   Grant KW & Seitz PF (2000) The use of visible speech cues for improving auditory

detection of spoken sentences. *J Acoust Soc Am* 108(3 Pt 1):1197-1208.

3.   Ma WJ, Zhou X, Ross LA, Foxe JJ, & Parra LC (2009) Lip-reading aids word

recognition most in moderate noise: a Bayesian explanation using high-dimensional

feature space. *PLoS One* 4(3):e4638.

4.   McGurk H & MacDonald J (1976) Hearing lips and seeing voices. *Nature*

264(5588):746-748.

5.   Traunmüller H & Öhrström N (2007) Audiovisual perception of openness and lip

rounding in front vowels. *Journal of Phonetics* 35(2):244-258.

6.   de Gelder B, Vroomen J, & Van der Heide L (1991) Face recognition and lip-reading in

autism. *European Journal of Cognitive Psychology* 3(1):69-86.

7.   Tremblay C*, et al.* (2007) Speech and non-speech audio-visual illusions: a developmental

study. *PLoS One* 2(1):e742.

8.   de Gelder B, Vroomen J, Annen L, Masthof E, & Hodiamont P (2003) Audio-visual

integration in schizophrenia. *Schizophr Res* 59(2-3):211-218.

9.   Romero YR*, et al.* (2016) Alpha-band oscillations reflect altered multisensory processing

of the McGurk illusion in schizophrenia. *Frontiers in human neuroscience* 10.

10.   Burnham D & Lau S (1998) The effect of tonal information on auditory reliance in the

McGurk effect. *AVSP'98 International Conference on Auditory-Visual Speech*

Processing.

11.   Sekiyama K & Tohkura Y (1991) McGurk effect in non-English listeners: few visual

294   effects for Japanese subjects hearing Japanese syllables of high auditory intelligibility. *J*

295   *Acoust Soc Am* 90(4 Pt 1):1797-1805.

296   12.   Basu Mallick D, Magnotti JF, & Beauchamp MS (2015) Variability and stability in the

297   McGurk effect: contributions of participants, stimuli, time, and response type.

298   *Psychonomic bulletin & review*:1-9.

299   13.   Strand J, Cooperman A, Rowe J, & Simenstad A (2014) Individual Differences in

300   Susceptibility to the McGurk Effect: Links With Lipreading and Detecting Audiovisual

301   Incongruity. *Journal of Speech, Language, and Hearing Research* 57(6):2322-2331.

302   14.   Magnotti JF*, et al.* (2015) Similar frequency of the McGurk effect in large samples of

303   native Mandarin Chinese and American English speakers. *Exp Brain Res* 233(9):2581-

304   2586.

305   15.   Stevenson RA*, et al.* (2017) The cascading influence of multisensory processing on

306   speech perception in autism. *Autism*:1362361317704413.

307   16.   Ferguson CJ & Heene M (2012) A vast graveyard of undead theories publication bias and

308   psychological science's aversion to the null. *Perspectives on Psychological Science*

309   7(6):555-561.

310   17.   Sterling TD (1959) Publication decisions and their possible effects on inferences drawn

311   from tests of significance—or vice versa. *Journal of the American statistical association*

312   54(285):30-34.

313   18.   Rosenthal R (1979) The file drawer problem and tolerance for null results. *Psychol Bull*

314   86(3):638.

19.   Gelman A & Weakliem D (2009) Of beauty, sex and power: Too little attention has been

paid to the statistical challenges in estimating small effects. *Am Sci* 97(4):310-316.

20.   Stropahl M, Schellhardt S, & Debener S (2017) McGurk stimuli for the investigation of

multisensory integration in cochlear implant users: The Oldenburg Audio Visual Speech

Stimuli (OLAVS). *Psychon Bull Rev* 24(3):863-872.

21.   Magnotti JF & Beauchamp MS (2015) The noisy encoding of disparity model of the

McGurk effect. *Psychonomic Bulletin & Review* 22(3):701-709.

22.   Cohen J (1992) A power primer. *Psychol Bull* 112(1):155-159.

23.   Rosenberg A, Patterson JS, & Angelaki DE (2015) A computational perspective on

autism. *Proc Natl Acad Sci U S A* 112(30):9158-9165.

24.   Bebko JM, Schroeder JH, & Weiss JA (2014) The McGurk effect in children with autism

and Asperger syndrome. *Autism Res* 7(1):50-59.

25.   Woynaroski TG*, et al.* (2013) Multisensory Speech Perception in Children with Autism

Spectrum Disorders. *J Autism Dev Disord*.

26.   Taylor N, Isaac C, & Milne E (2010) A comparison of the development of audiovisual

integration in children with autism spectrum disorders and typically developing children.

Journal of autism and developmental disorders 40(11):1403-1411.

27.   Stevenson RA*, et al.* (2014) Brief report: Arrested development of audiovisual speech

perception in autism spectrum disorders. *J Autism Dev Disord* 44(6):1470-1477.

28.   Gelman A & Carlin J (2014) Beyond Power Calculations: Assessing Type S (Sign) and

Type M (Magnitude) Errors. *Perspect Psychol Sci* 9(6):641-651.

29.   Gelman A & Tuerlinckx F (2000) Type S error rates for classical and Bayesian single and

multiple comparison procedures. *Computational Statistics* 15(3):373-390.

30.   Sekiyama K (1997) Cultural and linguistic factors in audiovisual speech processing: the

339   McGurk effect in Chinese subjects. *Percept Psychophys* 59(1):73-80.

340   31.   Chen Y & Hazan V (2007) Developmental factor in auditory-visual speech perception-

341   the Mcgurk effect in Mandarin-Chinese and English speakers. *AVSP'07 International*

342   *Conference on Auditory-Visual Speech Processing*.

343   32.   Alsius A, Paré M, & Munhall KG (2017) Forty Years After Hearing Lips and Seeing

344   Voices: The McGurk Effect Revisited. *Multisensory Research*.

345   33.   Vogel EK & Awh E (2008) How to exploit diversity for scientific gain: Using individual

346   differences to constrain cognitive theory. *Current Directions in Psychological Science*

347   17(2):171-176.

348   34.   Button KS*, et al.* (2013) Power failure: why small sample size undermines the reliability

349   of neuroscience. *Nat Rev Neurosci* 14(5):365-376.

350   35.   Bakker M, van Dijk A, & Wicherts JM (2012) The Rules of the Game Called

351   Psychological Science. *Perspect Psychol Sci* 7(6):543-554.

352   36.   Open Science C (2015) PSYCHOLOGY. Estimating the reproducibility of psychological

353   science. *Science* 349(6251):aac4716.

354   37.   R Core Team (2017) R: A language and environment for statistical computing (R

355   Foundation for Statistical Computing, Vienna, Austria).

356