Longitudinal studies at birth and age 7 reveal strong effects of genetic variation on ancestry-

associated DNA methylation patterns in blood cells from ethnically admixed children

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Abstract

The epigenetic architecture in humans is influenced by genetic factors, exposure histories and

biological factors such as age, but little is known about their relative contribution or their

longitudinal dynamics. Here, we studied DNA methylation levels at over 750,000 CpG sites in

mononuclear blood cells collected at birth and age 7 from 196 children of primarily self-reported

Black and Hispanic ethnicities to study age- and ancestry-related patterns in DNA methylation.

We first identified a large set of age-related CpGs at birth and age 7 and showed that gestational

age related CpGs at birth are the same as the age-related CpGs at age 7. We then developed a

novel Bayesian inference method for longitudinal data and showed that the vast majority of the

ancestry-associated methylation patterns present at birth are also present at age 7. A large

proportion of ancestry-associated CpGs (59%) had a nearby methylation quantitative trait locus

(meQTL) and we show that at least 13% of the ancestry-associated methylation patterns were

mediated through local genotype. These combined results indicate that ancestry-associated

methylation patterns in blood are in large part genetically determined. Our results, further

suggest that DNA methylation patterns in blood cells are robust to environmental exposures, at

least during the first 7 years of life.

Introduction

Epigenetic patterning in human genomes reflects the contributions of genetic variation[1; 2]

exposure histories[3-8], and biological factors, such as age[9-18], ethnicity[19-24] and disease status[25-28],

among others. However, little work has been done to elucidate the relative contributions or

longitudinal dynamics of each on epigenetic patterning.

To directly examine the relationship between age, ethnicity, genetic variation, early life

exposures and allergic phenotypes and an epigenetic mark, we studied global DNA methylation

patterns at over 750,000 CpG sites on the EPIC array in cord blood mononuclear cells (CBMCs)

collected at birth and in peripheral blood mononuclear cells (PBMCs) collected at 7 years of age

from 196 children participating in the Urban Environment and Childhood Asthma (URECA)

birth cohort study[29; 30]. This cohort is part of the NIAID-funded Inner-City Asthma Consortium

and is comprised of children primarily of Black and Hispanic self-reported ethnicity, with a

mother and/or father with a history of at least one allergic disease living in poor urban areas (see

Gern et al.[30] for details of enrollment criteria). Mothers of children in the URECA study were

enrolled during pregnancy and children were followed from birth through at least 7 years of age.

The longitudinal design of the URECA study provided us with the resolution to partition

genetic from non-genetic effects on ancestry-associated DNA methylation patterns, and yielded

new insight into the factors affecting DNA methylation patterns at CpG sites in mononuclear

(immune) cells during early life in ethnically admixed children. Using a novel statistical

inference method that provides a general framework for analyzing longitudinal genetic and

epigenetic data, we show that ancestry-dependent methylation patterns are conserved over the

first 7 years of life and that these patterns are strongly influenced, and often mediated, by local

genotype. Further, chronological age, but not measured exposures during pre- or post-natal

71 periods or disease status by age 7, was associated with methylation patterns in this sample of

72 children. Finally, we show that the "epigenetic clock"[10] present during pregnancy is the same

73 clock that is present after birth. Together with published data, we suggest that genetic ancestry

74 contributes to the epigenetic prediction of age at all stages of the life cycle, beginning *in utero*.

Materials and methods

Sample composition

URECA is a birth cohort study initiated in 2005 in Baltimore, Boston, New York City and St.

Louis under the NIAID-funded Inner City Asthma Consortium[29]. Pregnant women were

recruited. Either they or the father of their unborn child had a history of asthma, allergic rhinitis,

or eczema, and deliveries prior to 34 weeks gestation were excluded (see Gern et al.[29] for full

entry criteria). Informed consent was obtained from the women at enrollment and from the

parent or legal guardian of the infant after birth.

Maternal questionnaires were administered prenatally and child health questionnaires

administered to a parent or caregiver every 3 months through age 7 years. Gestational age at

birth and obstetric history were obtained from medical records. Additional details on study

design are described in Gern et al.[29] and in the Supplement. Frozen paired cord blood

mononuclear cells (CBMCs) and peripheral blood mononuclear cells (PBMCs) at age 7, were

available for 196 of the 560 URECA children after completing other studies. After QC (see

below), DNA methylation data were available for 194 children at birth, 195 children at age 7,

and 193 children at both time points; genotype data were available in 193 children (194 at birth;

195 at age 7) (Supplementary Table 1).

DNA Methylation

DNA for methylation studies was extracted from thawed CBMCs and PBMCs using the Qiagen

AllPrep kit (QIAGEN, Valencia, CA). Genome-wide DNA methylation was assessed using the

Illumina Infinium MethylationEPIC BeadChip (Illumina, San Diego, CA) at the University of

Chicago Functional Genomics Facility (UC-FGF). Birth and 7-year samples from the same child

were assayed on the same chip and the data were processed using Minfi[31]; Infinium type I and

type II probe bias were corrected using SWAN[32]. Raw probe values were corrected for color

imbalance and background by control normalization. Three out of the 392 samples (two at birth

and one at age 7) were removed as outliers following normalization. We removed 82,352 probes

that mapped either to the sex chromosomes or to more than one location in a bisulfite-converted

genome, had detection P values greater than 0.01% in 25% or more of the samples, or

overlapped with known SNPs with minor allele frequency of at least 5% in African, American or

European populations. After processing, 784,484 probes were retained and M-values were used

for all downstream analyses, which were computed as $\log_2$ (methylated intensity + 100) - $\log_2$

(unmethylated intensity + 100). The offset of 100 was recommended in Du et al.[33].

Genotyping

DNA from the 196 URECA children was genotyped with the Illumina Infinium

CoreExome+Custom array. Of the 532,992 autosomal SNPs on the array, 531,755 passed

Quality control (QC) (excluding SNPs with call rate <95%, Hardy-Weinberg P values <10^{-5}, and

heterozygosity outliers). We conducted all analyses in 293,696 autosomal SNPs with a minor

allele frequency ≥5%. Genotypes for three children failed QC and were excluded from

subsequent analysis that involved genotypes, including methylation quantitative locus (meQTL)

mapping, inferred genetic ancestry, or used genetic ancestry PC1 as a covariate (see below).

These three children were included in all other analyses.

Estimating inferred genetic ancestry

Ancestral principal component analysis (PCA) was performed using a set of 801 ancestry

informative markers (AIMs) from Tandon et al.[34] that were genotyped in both the URECA

children and in HapMap[35] release 23. Because PC1 captured the majority of variation in genetic

ancestry (Figure 1), we refer to PC1 as inferred genetic ancestry and consider it as a surrogate

measure for percent African ancestry.

Statistical analysis

To determine the effect of gestational age on methylation in CBMCs, we used standard linear

regression models with the child's gender, recruitment city, inferred genetic ancestry and

methylation plate number as covariates in our model. We also estimated cell composition and

other unobserved confounding factors using a method described in McKennan et al.[36]. We then

computed a gestational age P value for each CpG site and used q-values[37] to control the false

discovery rate at a nominal level. We took the same approach to determine CpGs whose

methylation changed from birth to age 7, except the response was measured as the difference in

methylation at birth and age 7. In this analysis, we included the child's gender, gestational age at

birth, inferred genetic ancestry and recruitment site as covariates. Because all paired samples

were on the same plate, we did not include plate number as a covariate in this analysis. We also

estimated unobserved factors that influence differences in methylation at birth and age 7 using

McKennan et al.[36] and included these latent factors in our linear model. See models (S1) and

(S2) in the Supplement for more detail.

We used data from the self-reported Hispanic and Black individuals with methylation

measured at both time points to analyze the effect of ancestry (either inferred genetic ancestry or

self-reported race) on methylation at birth and age 7 jointly using a Bayesian model. We did not

include the 10 individuals of other reported races in this analysis because we did not want our

estimates to be influenced by the groups with small samples sizes. We included age (birth or age

7), recruitment site, gestational age at birth, gender and methylation plate number as covariates

in our model, and estimated additional unobserved covariates (including cell composition) using

a method specifically designed for correlated data[36]. Once we estimated the relevant hyper-

parameters, we extended the sign rate paradigm developed in Stephens[38] to perform inference in

longitudinal data. This is discussed in more detail in the context of the specific questions we

present in the results section; the mathematical formulation is presented in the Supplement.

We performed meQTL mapping in the 145 genotyped, self-reported Black children using

the set of 269,622 SNPs with 100% genotype call rate in this subset. We restricted ourselves to

this subset of samples to minimize heterogeneity in effect sizes. To identify CpG-SNP pairs, we

considered SNPs within 5kb of each CpG, as this region has been previously shown to contain

the majority of genetic variability in DNA methylation[1], and computed a P value for the effect of

the genotype at a single SNP on methylation at the corresponding CpG with ordinary least

squares. We then defined the meQTL for each CpG site as the SNP with the lowest P value. In

addition to genotype, we included inferred genetic ancestry (i.e., ancestry PC1), gestational age

at birth, gender, recruitment city and methylation plate number in the linear model, along with

the first nine principal components of the residual methylation data matrix after regressing out

the intercept and the five additional covariates. We then tested the null hypothesis that a CpG did

not have an meQTL in the 10kb region by using the minimum marginal P value in the region as

the test statistic and computed its significance via bootstrap. Lastly, we used q-values to control

the false discovery rate.

Results

Our study included 196 children participants in the Urban Environment and Childhood Asthma

(URECA) cohort who had stored CBMCs and PBMCs collected at birth and age 7,

respectively[29], and passed quality control (QC) checks as described in Methods. The URECA

children were classified by parent- or guardian-reported race into one of the following

categories: Black, n =147; Hispanic, n = 39; White, n = 1; Mixed race n = 7, and Other, n = 2. A

description of the study population is shown in Table 1 and in Supplementary Materials.

Ancestry, assessed using ancestral PCs, revealed varying proportions of African and European

ancestry along PC1, which we refer to as inferred genetic ancestry (IGA); there is little

separation along PC2 (Figure 1). The reported races (RR) of the children are also shown in

Figure 1. The means and ranges of gestational age stratified by reported race are shown in Table

1; the distribution of gestational age at birth is shown in Figure S1 in the supplement.

The "epigenetic clock" *in utero* is the same "epigenetic clock" after birth

DNA methylation patterns change at many CpG sites across the genome as humans age[9; 13; 14; 39]

and an individual's epigenetic age may be more associated with their health and disease states

than chronological age[11; 15]. Two epigenetic age predictors have been proposed: one after birth[10]

and one prior to birth[13]. However, it is not known how the "epigenetic clock" present *in utero*

(measured at birth) is related to the epigenetic clock present later in life, and we therefore sought

to determine whether or not the same CpGs determine one's epigenetic age *in utero* and at age 7.

Among the 784,484 CpG sites in CBMCs that passed QC, the methylation levels at 2,836 CpGs

were significantly correlated with gestational age at an FDR of 5%. Methylation at 62% of these

sites decreased with increasing gestational age (P value for trend $< 10^{-10}$). We did not find any

significant associations between gestational age and methylation at age 7, indicating that

gestational age within the ranges available has little to no effect on methylation patterns later in

life.

To determine whether methylation levels at gestational age-associated CpG sites

continued to change between birth and age 7, we next identified a set of 16,172 CpG sites whose

methylation levels changed between these ages at an FDR of 5% (see model (S2) in the

Supplement for more detail). The 16,172 age-related CpGs included 1,986 (70%) of the 2,836

gestational age-associated CpGs. Thus, the gestational age-associated CpGs are largely a subset

of the CpGs that changed from birth to age 7. Furthermore, the estimates of age effects among

the CpGs that changed from birth to age 7 showed the same direction of change as the estimated

gestational age effects in 97% of the 16,172 CpGs, which included 14,186 gestational age-

associated effects that were not significant at the 5% FDR threshold. This concordance in

direction of effect is unlikely to occur by chance ($P < 10^{-10}$, see Supplement for calculation) and

provides strong evidence that CpGs whose methylation decreases with increasing gestational

ages *in utero* continue to decrease after birth (and vice-versa if the methylation increases *in*

utero).

To further validate that the same CpGs determine one's epigenetic age both *in utero* and

after birth, we investigated whether or not our gestational age- and age-related CpGs were

enriched for the 353 CpGs reported by Horvath[10] to predict age after birth using DNA

methylation patterns across many human tissues, as well as for the 148 CpGs used by Knight et

al.[13] to predict gestational age from cord blood DNA methylation. Even though only 6 CpGs

overlapped between the sets of 353 and 148 CpGs in each study, our gestational age and age-

related CpGs are highly enriched for both sets of CpGs from the Knight et al. and Horvath

studies (Figure 2).

Inferred genetic ancestry effects on DNA methylation patterns are conserved in magnitude

and direction between birth and age 7

Previous cross-sectional studies have revealed associations between ancestry and DNA

methylation at birth[19; 23] and later in life[20-22; 24; 25]. These correlations were generally attributed to

the combined effects of genetic variation and environmental exposures[20-23]. However, because of

the cross-sectional nature of these studies, it is not known if the association between ancestry and

methylation patterns present at birth persist (or change) in childhood. Moreover, because

ancestry is typically confounded with environmental exposures[40], it has been proposed that

ancestry effects on methylation levels may reflect the effects of exposure histories, which also

may vary by race or ethnicity[20]. Alternatively, ancestry effects on DNA methylation patterns

could also be due to genetic differences. In these cases, we would expect ancestry-associated

methylation patterns to be conserved from birth to later childhood. Using the longitudinal data

from the URECA cohort, we tested this hypothesis by addressing three questions. What is the

correlation between ancestry and methylation levels at individual CpG sites at birth and age 7? Is

the direction and magnitude of the correlation between ancestry and methylation levels

conserved between birth and age 7? Are there any CpGs for which the correlation between

methylation and ancestry at birth is significantly different from the correlation between

methylation and ancestry at age 7?

Standard hypothesis testing can be used to answer the first question but is not appropriate

for answering the second or third because failure to reject the null hypothesis that the effects are

equal at birth and age 7 does not imply the null hypothesis is true. Additionally, because our

studies were conducted in cord blood cells at birth and peripheral blood cells at age 7, effects at

birth and age 7 may differ due to differences in cell composition[41]. To circumvent these issues,

we built a Bayesian model (see model (S3) in the Supplement) and let the data determine both

the strength of the correlation between inferred genetic ancestry or reported race and

methylation, and how similar the correlations were at birth and age 7. We then answered the

first, second and third questions by defining and estimating the correct (*cor*), conserved (*con*)

and discordant (*dis*) sign rates for each CpG $g = 1,\ldots, 784{,}484$:

$cor_g^{(0)}$ = Posterior probability that the estimate for the direction of CpG *g*'s ancestry effect

245   at birth was correct

246        $cor_g^{(0,7)}$ = Posterior probability that the directions of CpG *g*'s ancestry effects at birth and

247   age 7 were the same AND the directions were estimated correctly

248        $dis_g^{(0)}$ = Posterior probability that the ancestry effect for CpG *g* at birth was non-zero and

249   was 0 or in the opposite direction at age 7 AND both directions were estimated

250   correctly

The correct and discordant sign rates at age 7 ($cor_g^{(7)}$, $dis_g^{(7)}$) were defined as follows. The

correct sign rate at birth and age 7 is always at least as large as the conserved sign rate. That is, if

a CpG has a significant conserved sign rate, the correct sign rates at birth and age 7 will also be

significant. We therefore say that the birth and age 7 effects for CpG g overlap if its conserved

sign rate ($con_g^{(0,7)}$) is above a designated threshold. Supplemental Figures S2 and S3 provide

insight into how the correct and conserved sign rates compare with standard univariate P values.

We refer the reader to "Joint modeling of methylation at birth and seven" in the Supplement for a

detailed description of our model and estimation procedure.

After fitting the relevant parameters in the model to the data, we were able to estimate the

fraction of CpGs with non-zero effects at both ages that fell into one of four possible bins: the

two effects were completely unrelated ($\rho = 0$), moderately similar ($\rho = 1/3$), very similar ($\rho =$

2/3), or identical ($\rho = 1$). Note that if a non-trivial fraction of CpG sites had effects that were

negatively related, they would be assigned to the first bin ($\rho = 0$). In fact, less than 1% of the

CpGs with non-zero inferred genetic ancestry effects at both ages had unrelated or moderately

similar IGA effects ($\rho = 0$ or 1/3), whereas 34% fell in the very similar bin and 66% had

identical inferred genetic ancestry effects at birth and age 7 (Supplemental Figure S4). This

indicates that when inferred genetic ancestry effects on methylation are non-zero at both birth

and age 7, they tend to be very similar or exactly the same with respect to both direction and

magnitude.

We then estimated the correct and conserved sign rates for all 784,484 probes, and

identified 2,873 inferred genetic ancestry-associated CpGs (IGA-CpGs) in CBMCs ($cor_g^{(0)} \geq$

0.95), 3,834 in PBMCs at age 7 ($cor_g^{(7)} \geq 0.95$), and 2,659 whose effects were conserved in sign

($cor_g^{(0,7)} \geq 0.95$). Methylation tended to increase with increasing African ancestry at 1,494 of the

2,659 conserved CpGs (56%), suggesting that individuals with more African ancestry tend to

have more methylation ($P < 10^{-10}$). This is consistent with the study of Moen et al.[22], which used

the Illumina 450K array to quantify the differences in methylation between European and

African populations. Supplemental Figure S5 shows the IGA-CpG locations in the genome and

Figure 3a illustrates the overlap between IGA-CpGs at birth and age 7. This strong overlap

corroborates our above observations and answers the second question in the affirmative: if

methylation is strongly correlated with inferred genetic ancestry at birth, the magnitude and

direction of the correlation is conserved at age 7.

Inferred genetic ancestry is more correlated with methylation than is self-reported race

The observed correlations between ancestry and methylation levels may reflect differences in

environmental exposures[20; 22], due to associations between race or ethnicity with socio-cultural,

nutritional, and geographic exposures, among others[40]. In fact, Galanter et al.[20] showed in a

cross-sectional study that self-reported ethnicity explained a substantial portion of the variability

in whole blood DNA methylation patterns from Latino children of diverse ethnicities, even more

so than inferred genetic ancestry. They concluded that ethnicity captures genetic, as well as the

socio-cultural and environmental differences that influence methylation levels. If this were the

case in the URECA children, we should observe just as large, if not larger, an effect as we did for

inferred genetic ancestry when we substitute reported race for inferred genetic ancestry in the

analyses presented in the previous section. However, using reported race, we identified only 457

CpGs at birth and 709 CpGs at age 7 whose correct sign rate was at least 0.95, and 424 whose

reported effects were conserved from birth to age 7 ($cor_{\mathrm{g}}^{(0,7)} \geq 0.95$), far fewer than the 3,991

CpGs whose methylation was significantly correlated with inferred genetic ancestry at birth or at

age 7.

To explore this further, we examined the overlap between IGA-CpGs and reported race-

associated CpGs (RR-CpGs) in CBMCs at birth and in PBMCs at 7 (Figures 3c-d). Because

reported race is an estimate of inferred genetic ancestry, there is a still substantial overlap

between IGA-CpGs and RR-CpGs. In fact, almost all of the RR-CpGs are among the IGA-CpGs,

but the opposite is not true. This indicates that while IGA-CpGs include most RR-CpGs, reported

race does not capture most of the variation in methylation attributable to genetic ancestry in these

children.

The observed correlations between DNA methylation and ancestry are primarily genetic

Both age and ancestry effects on methylation could be due to genetic variation or to

environmental exposures that differ throughout life or by ethnicity. To address this question, we

used local genetic variation (within 5kb of a CpG site) and DNA methylation data at birth and

age 7 in the 147 self-reported Black children in our study. Of the 519,622 CpGs within 5kb of a

SNP, 65,068 and 70,898 had at least one meQTL in CBMCs at birth and in PBMCs at age 7,

respectively, at an FDR of 5%. In addition, 59% of IGA-CpGs with at least one SNP in the ±5kb

window had at least one meQTL at birth or age 7 at an FDR of 5%, indicating IGA-CpGs were

enriched for CpGs with meQTLs (Figure 4a-b).

To provide further evidence that local genotype mediates the effect of inferred genetic

ancestry on methylation, we used logistic regression to regress the genotype of each of the

269,622 SNPs in our study set onto inferred genetic ancestry. The goal was to determine the

fraction of IGA-CpGs that were mediated through local genotype, i.e. IGA-CpGs with both

edges a and c in Figure 4a. Our analysis first revealed that the genotypes at meQTLs whose

target CpGs were IGA-CpGs in either CBMCs at birth or PBMCs at age 7 (IGA-meQTLs) were

significantly more correlated with inferred genetic ancestry than the genotype of non-IGA-

meQTLs (Figure 4c). Moreover, approximately 13% of the IGA-CpGs with at least one SNP in

their ±5kb windows had an inferred genetic ancestry effect that was mediated through local

genotype (i.e. had edges a and c, see Supplement for calculation details), which is likely an

underestimate of the true number IGA-CpGs mediated through local genotype because our

sample size was relatively small[42]. Nonetheless, this is striking compared to the 0.1% of non-

IGA-CpGs whose corresponding SNP had edge c at a 20% FDR.

Lastly, we used DNA methylation data on 573 ethnically diverse U.S. Latino children

ages 9 to 16 years old from the Galanter et al. study[20] to further explore the effect of ancestry on

DNA methylation in whole blood. Children and teenagers in the Galanter study were classified

as Mexican, Puerto Rican, mixed Latino, or other Latino. They reported 916 CpG sites whose

methylation was significantly associated with reported race, 773 of which were also significantly

associated with estimated percent European, Native American and African ancestry at a

Bonferroni P value threshold of 1.6 x 10^{-7}. A total of 726 of their 916 ethnic-associated CpGs

were also in the set of 784,484 probes CpG sites in our study. Our set of IGA-CpGs at birth or

age 7 contained a significant fraction of the 726 ethnic-associated CpGs from the cross-sectional

study, but there was considerably less overlap with our RR-CpGs (Figure 5a-b). If the correlation

we observed between ancestry and methylation was largely due to responses to environmental

exposures, as suggested in the Galanter study, then the overlap with reported race should have

been at least as large as the overlap with inferred genetic ancestry. That was not the case, further

suggesting that the inferred genetic ancestry effects on methylation in the URECA cohort are

primarily genetic in origin.

Age-related CpGs overlap with ancestry-associated CpGs that are also genetically

regulated

Recent work has suggested that race and ethnicity may play a role in determining one's

epigenetic age[13; 14]. However, it has not been determined if there is a genetic basis for this

phenomenon or if differences arise solely due different environmental exposures that are

confounded with ethnicity. To investigate this, we examined the overlap between age-related

CpGs identified in our study and the 263 ethnic-associated CpGs identified in the Galanter study

that also had an meQTL at a 5% FDR. Figure 5c-d illustrates that these 263 ethnic-associated

meQTLs are enriched for age-related CpGs in our study, suggesting that the previously observed

discrepancy in epigenetic age between ethnic groups may in part be due to genetic differences

between the groups.

Non-genetic factors influence the observed ancestry-methylation correlation more at age 7

than at birth

Although most of the variation in methylation levels at ancestry-associated CpGs can be

attributed to genetic variation in the URECA children, a small proportion may be due to non-

genetic (environmental) factors. To explore this possibility, we further hypothesized that non-

genetic effects on methylation levels at ancestry-associated CpGs would be greater at age 7 than

at birth, due to accumulated exposures over the first 7 years of life. We note that none of the

direct or indirect measures of exposures that were available in this cohort were associated with

methylation levels at either age, including maternal asthma, maternal infections during

pregnancy, pet ownership, bedroom allergens, mother stress, anxiety and depression metrics,

maternal cotinine levels during pregnancy, number of smokers in the household, number of

siblings, number of previous live births, daycare attendance, number of colds at age 2 or 3, and

allergic sensitization or asthma in the child (see Supplement for details). Therefore, if exposure

histories influenced methylation levels in these children, it is not likely due to any of these

measured variables.

To directly test the hypothesis that non-genetic factors tend to have larger effects on

methylation levels at age 7 than at birth, we used our Bayesian model to estimate the proportion

of CpGs in our study whose methylation was not associated with ancestry at birth but associated

at age seven and the proportion that were associated with ancestry at birth but not at age 7. The

former was greater than 14% while the latter was less than 1.5% using either inferred genetic

ancestry or reported race as a measure of ancestry. Even though over 14% of all CpGs in our

study had ancestry effects present at age 7 but not at birth, we were only able to identify 18

discordant IGA-CpGs and 4 discordant RR-CpGs at age 7 using a liberal threshold of $dis_g^{(7)} \geq$

0.80. That is, for almost all CpGs that are associated with ancestry at age 7 but not at birth, the

expected ancestry effect sizes were quite small relative to the statistical error, making it

impossible to assign the direction of effect on methylation changes with confidence (Figure S6).

Therefore, while some CpG sites may be influenced by exposures that are correlated with

ancestry at age 7 but not at birth, their effects were far too small to estimate in this sample size.

Discussion

The relationships between DNA methylation, chronological age, and ancestry have the potential

to shed light on disease etiology and may help determine the relative genetic and environmental

contributions to the observed inter-individual variability of the epigenome[9-15; 19-24]. For example,

it is well known that the methylome undergoes global changes as humans age[9; 10; 12], and

accelerated aging (the difference in epigenetic age and chronological age) is related to a variety

of disease phenotypes in adults[11; 15; 39], and with pregnancy complications, as well as with

developmental and stress-related phenotypes and birthweight in infants[42-44]. However, it has not

previously been shown that the same CpGs influence methylation age both prior to and after

birth. Our results suggest that not only is the collection of age-related CpGs similar pre- and

post-natally, but that the direction of age-related change in methylation levels is consistent before

and after birth.

Our discovery of the strict monotonicity of the change in methylation levels from birth to

age 7 is particularly interesting because it suggests the biological mechanism(s) involved with

advancing methylation age begin(s) *in utero* and remain(s) active at least to age 7. This provides

further support for and mechanistic insights into hypotheses on the fetal origins of adult

diseases[45; 46]. Moreover, because methylation age is an aggregate measure of methylation levels

across the genome, using this metric could reveal associations with phenotypes of interest even

when correlations with individual CpG sites are weak. Our results also suggest that including

CpGs that also predict epigenetic age *in utero* may provide more accurate measurements of

accelerated aging in adults.

While there was substantial change in blood methylation levels over time among children

in this cohort, inferred genetic ancestry and self-reported race effects on methylation are

overwhelmingly conserved in both direction and magnitude from birth to age 7, indicating that

variation in methylation patterns in blood cells are robust to changes in environment and are

consistent through early childhood. This result is interesting in and of itself because it provides

an example of fetal epigenetic variation that persists later in life, and more generally an example

of a persistent effect on DNA methylation levels, which has been cited as a critical area of future

epigenetic research[47; 48]. Furthermore, the novel statistical framework we used to infer effects

that are conserved versus those that vary over time can be easily applied to other longitudinal

DNA methylation data, as a way to avoid the spurious logic often used in applications of

frequentist hypothesis testing that failing to reject the null hypothesis implies the null is true.

The observation that inferred genetic ancestry and reported race effects are conserved

from birth to age 7 does not necessarily mean that the observed effects are primarily genetic in

nature because genetic and environmental components, or their interactions, could determine

ancestry-related methylation *in utero* and those patterns could persist as the child ages. To

further explore this, we showed that the IGA-CpGs were enriched among CpGs with meQTLs,

and that methylation levels at many of the IGA-CpGs are mediated by local genotype, indicating

that much of the ancestry-methylation correlation could be attributed to genetic variation.

Moreover, the RR-CpGs were only a small subset of IGA-CpGs in our study. This is opposite to

the findings of Galanter et al. [20], who argued that ancestry-dependent methylation patterns in

admixed populations are in large part determined by differences in exposure histories. Because

their data were cross-sectional they could not evaluate whether the observed patterns arose

during childhood or were also present at birth. While our results only provide evidence for

genetics accounting for most of the correlation between methylation and ancestry, they imply

that the genetic contribution to variability in blood methylation is substantial.

Our observations in support of strong genetically – and weak environmentally –

determined ancestry-associated methylation patterns in blood may seem paradoxical to the

plethora of studies showing that DNA methylation levels in blood cells are associated with

environmental exposures, such as cadmium, arsenic and smoking, to name a few[5-8; 20; 49-53].

Whereas the estimated genetic effect sizes in our study are substantially larger than many of the

environmentally-associated effects on methylation patterns previously reported, the effects of

environmental exposures on methylation in blood is probably too small to estimate with even

moderate to large sample sizes[45]. For example, it was only by performing a meta-analysis in

6,685 individuals that Joubert et al.[5] were able to identify 6,000 CpGs whose DNA methylation

levels in blood from infants and adolescents were associated maternal smoking exposure. In one

sense, we were able to corroborate previous observations of small non-genetic effects on

methylation in blood by showing that while methylation patterns at an estimated 14% of all

CpGs in our study were not correlated with ancestry at birth but correlated with ancestry at age 7,

the correlation at individual CpGs at age 7 was too small to be identified as statistically

significant. However, we were not able to find any statistically significant correlations between

methylation at birth or at age 7 and any of the environmental exposure variables measured in

blood cells from this cohort. We note that cord blood cotinine levels, a measure of in-utero

tobacco smoke exposure, were above the level of detection in only 34 of the 196 mothers in our

study.

The combined results of our study and those cited above suggest that DNA methylation

levels in blood cells are fairly robust to environmental perturbations. A better understanding of

tissue-specific methylation responses to environmental exposures could inform the design of

future studies and provide insights into the mechanisms through which exposures and gene-

environment interactions influence health and disease.

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631   **Table 1**: Covariates for the 196 URECA children in our study, stratified by self-reported race.
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	Black	**Hispanic**	**White**	**Mixed**	**Other**
Sample Size (N)	147	39	1	7	2
Males (%)	71 (48%)	25 (64%)	0 (0%)	4 (57%)	0 (0%)
Asthma diagnosis at age 7 (%)	38 (26%)	12 (31%)	0 (0%)	2 (29%)	0 (0%)
Gestational age at birth, in weeks (mean [range])	39.0 [34,42]	38.9 [35,41]	36.0	39.1 [37,40]	39.0 [38,40]
Recruitment city					
Baltimore (%)	64 (44%)	1 (3%)	1 (100%)	3 (43%)	2 (100%)
Boston (%)	17 (12%)	5 (13%)	0 (0%)	3 (29%)	0 (0%)
New York (%)	23 (16%)	32 (82%)	0 (0%)	1 (14%)	0 (0%)
St. Louis (%)	43 (29%)	1 (3%)	0 (0%)	1 (14%)	0 (0%)

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 Figure Legends

Figure 1: Estimated ancestry principal components (PCs) 1 and 2. Nearly all the variation in

ancestry separates along PC1 in the URECA sample. Filled triangles represent the 196

URECA children in this study, with their self-reported race shown in different colors.

Open circles are reference control samples from HapMap; red = Utah residents from

northern and western Europe (CEU), yellow = east Asian (Chinese and Japanese); dark

blue = Africans from Nigeria (Yoruban).

Figure 2: Distribution of *P* values for age (birth to age 7) (panels a and c) and gestational age

(GA) (panels b and d). Upper panels, stratified by whether or not the CpG was one of the

353 CpGs used to build the linear predictor of age in Horvath 2013[10] (red dots). The blue

dots are all other CpGs; the 10 enlarged red circles are for visual aid. Lower panels,

stratified by whether or not the CpG was one of the 148 CpGs used to build the linear

predictor of age in Knight et al.[13] (violet dots). The blue dots are all of the other CpGs; the

ten enlarged violet circles are just for visual aid.

Figure 3: Overlapping ancestry CpGs at birth and at age 7. (a): IGA-CpGs in CBMCs at birth

($\hat{cor}^{(0)} \geq 0.95$) and PBMCs at age 7 ($\hat{cor}^{(7)} \geq 0.95$). Overlapping CpGs (violet) have

$\hat{con}^{(0,7)} \geq 0.95$. (b): RR-CpGs at in CBMCs at birth ($\hat{cor}^{(0)} \geq 0.95$) and PBMCs at age 7 (

$\hat{cor}^{(7)} \geq 0.95$). Overlapping CpGs (violet) have $\hat{con}^{(0,7)} \geq 0.95$. (c): IGA-CpGs and RR-

CpGs ($\hat{cor}^{(0)} \geq 0.95$) in CBMCs at birth. (d): IGA-CpGs and RR-CpGs ($\hat{cor}^{(0)} \geq 0.95$) in

PBMCs at age 7.

Figure 4: IGA-CpGs are enriched for CpGs with meQTLs. (a) Illustration of the causal

relationship between the methylation (M) at a CpG site, the genotype (G) at the SNP

within ±5kb of the CpG that had the smallest meQTL *P* value and inferred genetic

660  ancestry (IGA). Each graph corresponds to a unique CpG. (b) Plots of the meQTL P value

661  for edge a in CBMCs at birth, where CpGs were stratified by whether or not it was an

662  IGA-CpG at birth or age 7 $\left(\max\left(c\hat{o}r_g^{(0)}, c\hat{o}r_g^{(7)}\right) \geq 0.95\right)$. The ten enlarged red circles are

663  just for visual aid. (c) Plots of the logistic regression P value for edge c (Genotype $\sim$ IGA

664  + Ancestry PC2 (see Figure 1)), stratified by whether or not the SNP was an IGA-meQTL.

665  **Figure** 5: Inferred genetic ancestry effects on methylation are primarily genetic in origin.

666  Upper panel. Histograms of $\max\left(c\hat{o}r^{(0)}, c\hat{o}r^{(7)}\right)$ in the IGA (a) and the RR (b) analyses. If

667  $\max\left(c\hat{o}r_g^{(0)}, c\hat{o}r_g^{(7)}\right)$ is close to 1 in the IGA or RR analysis, CpG g is an IGA-CpG or RR-

668  CpG, respectively, in CBMCs at birth or PBMCs at 7. The red histogram is created with

669  the set of 726 ethnicity-associated CpGs identified in Galanter et al.[20] that were also

670  among the 784,484 CpGs in our study and the blue are the remaining 783,758 CpGs.

671  Lower panel. P values for age (0 to 7) and gestational age (GA). The red points are the 263

672  ethnic-associated CpGs identified in Galanter et al., 2017[20] that had an meQTL at a 5%

673  FDR identified in our study. The blue points are the remaining 179,385 CpGs that were

674  within 5kB of a SNP in our data and among the CpGs studied in Galanter et al.[20]. The

675  large circles are shown for visual aid.

Figure 1

Figure 2

Figure 3

(a)

(b)   **(c)**

Figure 4

(a)　　　　　　　　　　　　**(b)**

(c)　　　　　　　　　　　　**(d)**

Figure 5