Simulation-based approaches to characterize the effect of sequencing effort on the quantity and quality

of metagenome-assembled genomes

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Abstract

We applied simulation-based approaches to characterize how sequencing effort influences the

properties of genomes identified in metagenomes assembled from short read sequences. An initial

analysis evaluated the quantity, completion, and contamination of complete-genome equivalents, a

bioinformatic-pipeline normalized metric for metagenome-assembled genomes (MAGs), as a function

of sequencing effort on four preexisting sequence read datasets taken from a maize soil, an estuarine

sediment, the surface ocean, and the human gut. These datasets were variously subsampled to simulate

the effect of sequencing effort on MAG retrieval. Modeling suggested that sequencing efforts beyond

what is typical in published experiments (1 to 10 Gbp) would generate diminishing returns in terms of

MAG binning. A second analysis explored the theoretical relationship between sequencing effort and

the proportion of available metagenomic DNA sequenced during a sequencing experiment as a function

of community richness, evenness, and genome size. Simulations from this analysis demonstrated that

while community richness and evenness influenced the amount of sequencing required to sequence a

community metagenome to exhaustion, the effort necessary to sequence an individual genome to a

target fraction of exhaustion was only dependent on the relative abundance and genome size. The

software tool, GRASE, was created to assist investigators further explore this relationship. Re-

evaluation of the relationship between sequencing effort and binning success in the context of the

relative abundance of genomes, as opposed to base pairs, provides a framework to design sequencing

experiments based on the relative abundance of microbes in an environment rather than arbitrary levels

of sequencing effort.

Importance

Short read sequencing with Illumina sequencing technology provides an accurate, high-

throughput method for characterizing the metabolic potential of microbial communities. Short read

sequences can be assembled and binned into metagenome-assembled genomes, thus shedding light on

the function of microbial ecosystems that are important for health, agriculture, and Earth system

processes. The work presented here provides an analytical framework to select sequencing effort as a

function of experimental goals in metagenome-assembled genomes creation projects. We hope that the

results presented here, as well as GRASE will be valuable to researchers planning sequencing

experiments.

Introduction

The reconstruction of high-accuracy short read sequences into metagenome-assembled genomes

(MAGs) is a recent approach to characterize microbial metabolisms within complex communities (1).

The recent creation of ~8,000 MAGs from largely uncultured organisms across the tree of life (2), the

spatial characterization of microbial metabolisms and ecology across Earth's oceans (3), and the

characterization of the potential impact that fermentation-based microbial metabolisms have on

biogeochemical cycling in subsurface sediment environments (4) provide a few examples of how

MAGs have helped to elucidate the relationships between microbial ecology, microbial metabolisms,

and biogeochemistry.

Sequencing experimental workflows involve selecting a target environment, sequencing effort,

bioinformatic pipeline software and parameters, metabolism characterization software (i.e., for gene

identification and similarity searches) and databases (Fig 1). At present, there is little information to

guide how much sequencing is appropriate to achieve scientific goals in such experiments (5). This gap

in knowledge is partly attributed to the unknown structure of target microbial communities. Further

complications arise in that the accuracy and efficiency of bioinformatic pipelines are often difficult to

characterize, and thus obscure the relationship between sequencing effort and MAG retrieval. For the

year 2017, estimates compiled by Quince et al. (5) suggest that up till now, metagenomic shotgun

sequencing experiments usually sequence between 1 Gbp and 10 Gbp DNA nucleotides. Researchers

require more precise guidance to select an appropriate shotgun sequencing effort in order to maximize

information and minimize cost for their specific experimental question.

Illumina sequencing technology is currently the most popular platform to generate

metagenomic shotgun sequences (5). Previous investigators established theoretical relationships

between contig formation rate (6) and single genome coverage (7) as a function of short read

sequencing effort. On the community level, heuristic approaches have been proposed for evaluating

community-level coverage to increases in sequencing effort. For example, it has been proposed to

utilize short read redundancies as a function of sequencing effort to estimate community level coverage

(8). Without *a priori* knowledge of the microbial community structure, practical application of these

methods to estimate MAG retrieval as a function of sequencing effort is hindered. This is particularly

true as heuristic methods will only be sensitive towards highly abundant genomes.

Here we present two distinct analyses which constrain the relationship between the quantity of

Illumina metagenomic shotgun sequences and the community-level sequence coverage. First, we

performed *in silico* experiments simulating the effect of sequencing effort on retrieved MAG properties

for Illumina sequence datasets. For these experiments, we propose a new metric, complete-genome

equivalents, to reduce sensitivity to the unknown accuracy and efficiency of our bioinformatic pipeline

(Fig 1). Second, we applied a theoretical model and numerical simulations to estimate the minimum

sequencing effort needed to sequence a metagenome to a target fraction of exhaustion. Our theoretical

model is unique compared to previous models (6, 7) in that we characterize sequencing effort in the

context of all physically-distinct DNA fragments in the metagenome. This is important, as complete

genome coverage does not necessarily translate into accurate read reconstruction due to interspecies

homologous DNA and intraspecies microdiversity (9). Coupling results from the two analyses provides

a framework for investigators to define sequencing experiments in the context of selecting a rarity and

fraction of exhaustion for a desired target genome when sequencing a community. Overall, the patterns

presented here can be used to guide sequencing effort decisions for future sequencing project when

MAG creation is a primary goal.

Results

The number of 'complete-genome equivalents' (equation 1) as a function of high-quality bases

empirically fit the Gompertz equation (equation 2; Fig 2B; parameters in Table 1). Complete-genome

equivalents represent the sum of MAGs retrieved, weighted by each MAG's checkM completeness

score (10). This metric is less sensitive to whether bioinformatic pipelines tend to lump contigs into

fewer, more-complete MAGs versus split them into fewer, less-complete MAGs.

For each environment, the number of complete-genome equivalents as a function of simulated

sequencing effort fit the Gompertz equation. Here, A, μ, and λ correspond to the maximum number of

complete-genome equivalents assembled with the pipeline, the maximum rate which complete-genome

equivalents form as with more sequencing, and the "lag bases," or the bases which must be sequenced

prior to a sufficient number of sequence reads existing to generate overlap and form contigs (6). The

predicted maximum complete-genome equivalents varied from ~17 in the gut community to ~97 in the

surface ocean, the predicted maximum rate that complete-genome equivalents increased varied from

~1.4 to ~5.8, and the minimum threshold of sequencing necessary prior to seeing complete-genome

equivalents varied from ~0.6 to ~6.7. The *Tara Oceans* dataset, where complete-genome equivalents

decreased at sequencing effort >20 Gbp, was an exception. For the estuary, maize, and human gut

datasets, the complete-genome equivalents yield began to asymptote with increasing sequencing

efforts. The *Tara Oceans* dataset followed a similar pattern at <25 Gbp. However, when the number of

sequenced bases was >25Gbp, the number of complete-genome equivalents decreased and became

insensitive to sequencing effort. The total number of MAGs retrieved was considerably higher than the

number of complete-genome equivalents.

Mean MAG completeness also increased towards an asymptote with increasing sequencing

effort (Fig 2C). Completeness was highest for the human gut dataset, with a maximum of 23.9%, and

increased continuously as sequencing effort increased. The mean MAG completeness reached an

asymptote of ~10-15% for the other three datasets with sequenced bases >10 Gbp. Note that when >10

Gbp were sequenced, the number of complete-genome equivalents created still increased as new

sequences were added. For all datasets, mean MAG contamination was <2% (Fig 2D) and did not

depend strongly on sequencing effort.

SIMULATION EXPERIMENTS

Using equation 8, we calculated the number of k-length sequence reads required to sequence all

physically-distinct DNA sequences of length, k (k-mers), in four hypothetical metagenomes, a perfectly

even, moderately uneven, highly uneven, and a lognormally-distributed structure (Fig 3A-D). Note that

equation 8 treats homologous DNA uniquely from one another. The expectation value of the log

number of sequences required to fully sequence metagenomes of the hypothetical communities was

linear with respect to log-transformed size of the metagenome (i.e., physically-distinct k-mers in the

population, or unique base pairs in a metagenome); this suggests a power-law relationship between

metagenome size and expectation value of sequence reads required to sequence the metagenome to

exhaustion (Fig 3E). For all community structures, the slope of the relationship between log-

transformed sequenced reads and log-transformed physically-distinct number of sequenced reads was

within 1% of 1.06. The structure of the population strongly influenced the number of reads required

such that more-even community structures required far fewer reads than less-even structures.

As equation 8 only estimates the number of reads to sequence a metagenome to exhaustion, we

used a numerical simulation to estimate the number of k-sized reads to sequence a metagenome to a

target fraction of exhaustion. The numerical simulation and equation 8 agreed on the expected

sequences reads to sequence 100% of a given metagenome (Fig 4); this supported the use of this

simulation. The log-transformation of both total physically-distinct k-sized reads ($|K_{MG}|$ in equation 7)

and sequenced reads showed a linear response for all target fractions and all community structures. The

amount of sequences required to achieve a given target of physically-distinct k-sized reads was variable

for the different communities shown in Fig 3A. For instance, the lognormally-distributed community,

when compared to other communities, required the most sequencing effort to sequence a metagenome

to exhaustion but required similar sequencing effort to sequence a metagenome to a target fraction of

50%.

We applied the simulation to demonstrate the effect of community evenness on the number of

reads required to sequence a community to a target fraction of completion. These communities ranged

from perfectly even (a=0, equation 10) to more uneven (a=0.02, Fig 5A). Evenness was quantified

using the Pielou evenness index, which expresses Shannon diversity relative to the diversity of a

perfectly even community (11). The sequencing effort required to characterize genomes depended

strongly on both the evenness and the target fraction of completion (Fig 5B). Again, less-even

communities required more sequence reads than more-even communities. The strength of this

relationship also depended on the target fraction of completion. A community with Pielou evenness of

0.97 required 3 orders of magnitude more sequence reads to sequence a metagenome to a target

fraction of exhaustion than a perfectly even community while the same community only required about

42% more reads to sequence 50% of the metagenome.

The minimum sequencing effort required to sequence a microbe genome given a combination

of target fraction, genome size, and fraction of the metagenome community was modeled with a

generalized additive model (GAM; Fig 6). The GAM shows expected sequencing effort required for

microbial genome sizes 0.5, 2, 5, 10, and 20 Mbp, target genome completeness fractions from 0.5 to

1.0, and genome relative abundances from 1 to 0.0001. The smooth dimensions for target fraction,

genome size, and fraction of the metagenome community were 50, 6, and 29. To normalize for

different sequence read length, sequence reads were converted to bases and ranged from 1×10^7 to

1×10^{13} total bases. More bases were required to sequence microorganisms 1) when the genome was

relatively rarer in the community, 2) to achieve better coverage of the genome, and 3) when average

genome sizes were larger.

Using the relationship shown in Fig 6, we can convert sequencing effort to an abundance cutoff.

Fig 7 shows complete-genome equivalents retrieved for a genome of a given level of abundance that is

sequenced to a target depth for the human gut, maize soil, estuarine sediment, and the surface ocean

microbiomes analyzed earlier. Correlation coefficients for the regressions used to relate log-

transformed sequencing effort (base pairs) to genome relative abundance were $R=1$ for all three

genome sizes evaluated (1 Mbp, 5 Mbp, and 20 Mbp). Evaluation of 1 Mbp and 20 Mbp define the

range of uncertainty in predicting complete-genome equivalents as the true size of genomes is

unknown. Unlike the asymptotic response of complete-genome equivalents to sequencing effort (Fig

2B), complete-genome equivalents increase exponentially as the abundance cutoff decreases (note that

the abundance cutoff is on a log-scale). The target genome completeness fraction was held at a constant

of 0.5 for all regressions and the sensitivity of complete-genome equivalents will change with respect

to the genome abundance cutoff with different values of target genome completeness fraction.

Discussion

We sought to establish evidence-based guidelines for selecting a sequencing effort during

shotgun metagenomic sequencing experiments. Random subsamples of existing short read datasets

(individually sampled, assembled, and binned) simulated the effect of creating MAGs from datasets of

different sizes and environments. The datasets analyzed here are representative of both the sequencing

effort (1 to 10 Gbp) (5) and the types of target environments that microbial ecologists often investigate

(12). A variety of software is available for all steps of MAG creation pipelines, and the quantity/quality

of MAGs will depend on software selection, software configuration, and sequenced environment (Fig

1; (5)). Furthermore, it is best-practice to manually curate algorithmically-created MAG bins (13). We

do not argue that the pipeline used here is objectively optimal for generating high-quality (high

completeness and low contamination) MAGs. Rather, our pipeline was configured to minimize

contamination (Fig 2D) associated with retrieved MAGs at the expense of reduced completeness. Thus,

MAG quantity was not directly reported but expressed as complete-genome equivalents (equation 1).

Reporting complete-genome equivalents reduces bias associated with binning software.

As sequencing effort increased, there was an initial lag in complete-genome equivalents

followed by a rapid increase in complete-genome equivalents, and then diminishing returns at higher

sequencing efforts. Previous investigators modeled the response of 16S RNA gene (14–16), Hill's

number diversity (17), taxon-resolved abundance (18), and gene abundance (18) as a function of

sequencing effort with rarefaction curves, or collector's curves. The complete-genome equivalents as a

function of sequencing effort did not match a traditional collector's curve, which lacks an initial lag.

Rather, the data appear sigmoidal. We modeled the data using the Gompertz function (equation 8),

because its parameters can be interpreted in terms of quantities that are familiar from microbial growth

curves (lag time, growth rate, and maximum density) (19). The three Gompertz parameters, A, μ, and λ

are interpreted as the maximum possible number of complete-genome equivalents at infinite

sequencing effort, maximum rate that complete-genome equivalents increased with increasing

sequencing effort, and a minimum sequencing threshold prior to rapid complete-genome equivalents

retrieval (Fig 2A), respectively. Consequently, the Gompertz equation fit parameters may provide

means for inter-comparisons between different metagenomic datasets.

The fits to the Gompertz function illustrate that there may be an optimal sequencing effort for

MAG creation efforts corresponding to the upper shoulder of the Gompertz curve ("late log phase").

When sequencing effort is too close to λ, MAGs bin poorly; when sequencing effort is too great, the

number and quality of MAGs per unit sequencing effort is low. We speculate that our choice of

pipeline, and specifically the fact that we discarded contigs <3kb, caused poor performance at higher

sequencing effort for the *Tara Oceans* dataset. Species-level microdiversity and interspecies

homologous DNA can cause "bubbles" which impair assembly in larger data sets (9, 20). Improved

assembly would likely have yielded more complete-genome equivalents for our assembly of the largest

subsets of the *Tara Oceans* data.

As mean MAG completeness converged to an asymptote well below 100% (Fig 2B), complete-

genome equivalent yields, or the ratio of the maximum observed complete-genome equivalents to the

theoretical maximum complete-genome equivalents (*A*) (Table 1) were close to 100%. The asymptote

in mean MAG completeness (Fig 2C) and negligible values of mean MAG contamination (Fig 2D)

demonstrates that the retrieved complete-genome equivalents at higher sequencing efforts represent

true signal opposed to false positive signal. If the complete-genome equivalents were false signal, then

either 1) the mean MAG completeness would decrease with the addition of false positive marker genes

which fail to cluster (i.e., more MAGs with low completeness) or 2) false positive signal would cluster

with preexisting MAGs leading to an increase in mean MAG contamination. Neither of these scenarios

occurred. More so, the fact that mean MAG completeness was at steady state suggests more

sequenceable DNA exists in these communities. Thus, we asked how much sequencing was necessary

to sequence a community to exhaustion.

The theoretical probability model (equation 8) demonstrated that the sequencing effort to

sequence a metagenome to exhaustion was predictable, regardless of community structure (Fig 3). One

limitation to equation 8 is predicting sequencing effort to sequence a metagenome to a fraction of

exhaustion. As such, a numerical simulation was used for predicting the sequencing effort necessary to

sequence an individual metagenome to a fraction of exhaustion. The numerical simulation agreed with

the theoretical model when predicting the sequencing effort to sequence a metagenome to exhaustion

(Fig 4). A further analysis explored the semi-quantitative relationship between community evenness

and richness and the necessary sequencing effort to achieve a target fraction of exhaustion for a

metagenome (Fig 5).

In practice, information about a target community structure may not be available for estimating

sequencing effort. The GAM built here predicts the minimum number of sequences necessary to

sequence a given fraction of a target genome as a function of average genome size and the relative

abundance of the target genome in the community (Fig 6). Even without knowledge of a target

community's structure, the GAM provides a useful constraint for designing whole-genome shotgun

metagenomic sequencing experiments.

Metagenomic shotgun sequencing experimental designs should be rationally designed such that

sequencing effort is selected to capture a desired fraction of a target microbial metagenome.

Investigators should be cognizant of the rarest microbial metagenome desired to be characterized as

well as the degree of characterization of that microbial metagenome when designing a sequencing

experiment. To that end, we have built a tool, `Genome Relative Abundance to`

`Sequencing Effort (GRASE)`, to report estimated sequencing effort required to capture a

defined fraction of a genome as a function of the relative abundance of the corresponding

microorganism in the community and average genome size. This R-based GUI app can be accessed

online at http://adsteen.shinyapps.io/grase and is archived at http://github.com/adsteen/GRASE, from

which it can be downloaded and run locally.

When the sequence read datasets analyzed here (Table 2; Fig 2) are re-evaluated in the context

of the relative abundance of a microbial metagenome (g_{MG}) sequenced to a target fraction (0.5),

complete-genome equivalents increase appreciably in response to minor increases in deeper

characterization of the community metagenome (Fig 7). This observation contrasts the complete-

genome equivalents response to sequencing effort (in base pairs), where substantial increases in

sequencing effort (by contemporary standards) leads to diminishing returns in complete-genome

equivalents. It is important to note that abundance cutoff and sequencing effort are interchangeable;

however, complete-genome equivalents response to changes in the respective predictor (i.e., base pairs

versus abundance cutoff) alters the optics of the collector's curve. Modest increases in sequencing

effort contribute minor amounts to extending abundance cutoff. A substantial amount of genomic and

metabolic data can be gained from targeting rarer microbes (metagenomic abundances <0.005), with

the caveat that whole-genome shotgun sequencing technology (as well as computational power)

requires significant increases in either the number or length of reads generated per run.

Materials and Methods

SEQUENCE DATA SOURCES

All sequence data were downloaded from NCBI's Sequence Read Archive (SRA) using the

SRA Toolkit (fastq-dump –split-files) (21). Exact duplicate reads for both forward and reverse reads

were removed using PRINSEQ (-derep 1; v0.20.4) (22). All sequencing datasets were limited to

Illumina shotgun metagenomic paired-end reads. Four datasets were analyzed for this analysis. The

first dataset was from oceanic surface water collected at 5m depth in the Caribbean Sea as a part of the

Tara Oceans expedition (23). The second dataset was from sediment from a depth of 8-10 cm below

the surface (sulfate-rich zone) and collected at the White Oak River Estuary, Station H, North Carolina,

USA (4). The third dataset was collected from maize soil (24). The last dataset was collected from

human fecal samples and represented a human gut microbiome (25). All datasets analyzed in this study

are summarized in Table 1.

MAG ASSEMBLY PIPELINE

The pipeline developed here followed similar pipelines described by other authors (3, 26). All

sequence datasets were analyzed as follows. Trimmomatic (v0.36) (27) removed adapters and trimmed

low-quality bases from individual reads. Read leading and trailing quality scores were required to be

>3. The sliding window was set to 4 base pairs and filtered base pair windows with a mean score <15.

Quality controlled reads were assembled into contigs using MEGAHIT (v1.1.2; --presets meta-large)

(28). Due to RAM limitations, assembled contigs <3000 bp long were excluded from the analysis.

Redundant contigs were removed using CD-HIT (v4.6.8; cd-hi-est -c 0.99 -n 10) (29). Remaining

contigs with high intra-contig similarity based on intra-contig sequence alignments (Minimus2 -D

OVERLAP=100 MINID=95) were removed. The quality-controlled reads (i.e., after using

Trimmomatic) were mapped to the remaining contigs using Bowtie 2 (v2.3.3) (30) to generate a

coverage score for individual contigs.

Resultant contigs were iteratively clustered into MAGs using the unsupervised, clustering

algorithm Binsanity (v0.2.6) (26). Similar to Tully et al. (3), six initial clustering iterations were

performed with the parameter, *preference* (-p), set to -10 (iteration 1), -5 (iteration 2), -3 (iteration 3-6).

Between iterations, a refinement step (Binsanity-refine) was performed on the putative MAGs with

constant *preference* (-p) of -25. The refined putative MAGs were evaluated for contamination and

completeness using the software CheckM (v1.0.6) (10), which uses HMMER (v3.1) and Prodigal

(v2.6.3) (31). Contigs associated with putative MAGs meeting one of the following criteria: 1) had a

completeness > 90% and contamination < 10%, 2) had a completeness > 80% and contamination < 5%,

or 3) had a completeness > 50% and contamination < 5% were treated as high-quality. All other MAGs

were considered low-quality MAGs. MAGs defined as high-quality were not modified any further.

Contigs associated with the high-quality MAGs were not used in the subsequent reclustering and

refinement steps. The contigs associated with low-quality MAGs were pooled together and reclustered

during the next iteration of Binsanity clustering. After the sixth iteration, the remaining MAGs which

did not fall into one of the three categories underwent additional refinement using Binsanity-refine.

During this step, MAGs were iteratively refined with *preference* set to -10 (iteration 1), -3 (iteration 2),

and -1 (iteration 3). Between each refinement step, metrics of contamination and completeness were

evaluated using CheckM. Again, MAGs meeting the criteria of a high-quality category described above

were not further modified. The respective contigs associated with putative MAGs were not used in

proceeding refinement steps. After the last iteration of refinement, all MAGs were reevaluated for

completeness, contamination, and a taxonomic rank using CheckM. Completeness values for MAGs

with a minimum resolved taxonomic rank of phylum were summed together. Complete-genome

equivalents, g_{eq}, were defined as:

$g_{eq} = \sum_{i=1}^{n} f_i$ (1

where *n* is the total number of retrieved MAGs and f_i is the corresponding CheckM completeness (%).

Common practice when reporting MAGs involves discarding MAGs not meeting a criterion of

completeness and contamination. In the context of characterizing MAG response to sequencing effort a

MAG pipelines often fails at generating high-quality MAGs. One must discern whether this is a failure

of the bioinformatic pipeline versus insufficient sequencing effort (Fig 1). We argue that MAGs with

low completeness still provide information regarding the community level coverage when generating

collector's curves. More so, a community has finite number of unique marker genes, and an infinite

sequencing effort would result in sequencing all unique marker genes, regardless of whether or not

these genes cluster to one another.

SUBSAMPLING SEQUENCE READ DATASETS

The effect of decreased sequencing effort was simulated by subsampling the initial sequence read

datasets described above. Downloaded sequence read datasets were randomly sampled at set fractions

of 1%, 10%, 20%, 40%, 60%, 80%, 90%, 95%, and 100%. Each fraction was resampled, assembled,

and binned in triplicate. All triplicates were analyzed using the MAG assembly pipeline described

above.

MODELING MAG RESPONSE TO SEQUENCING EFFORT

Complete-genome equivalents as a function of sequencing effort was modeled for environmental

sequence datasets using the Gompertz equation, as reformulated by Zweitering et al. (19) for use with

microbial growth curves:

$$g_{eq}(A, \mu, \lambda, b) = A \times e^{-e^{\frac{\mu \times e}{A}(\lambda - b) + 1}} \tag{2}$$

where A, μ, and λ are fit coefficients and b is high-quality bases (Gbp).

DEFINING THE MICROBIAL METAGENOME AND SEQUENCING PROBABILITY

Here we draw on set theory to provide a theoretical grounding for our *in silico* simulations described

below. The expected number of sequences to sequence a fraction of an individual microbe's genome

can be modeled with probability theory by defining a community metagenome with set theory. Fig 8A-

E provides a cartoon illustrating the application of this set theory on a hypothetical microbial

population, *G*. *G* contains unique genomes (*g*) with finite abundances (*n*). The definition of microbial

species is somewhat contentious (32). Here we define *g* as the average genome (length and

composition) for all individuals genomes associated with the rank, species, based on an arbitrary

definition. As such, the species richness (*s* or unique *g*) of *G* depends on the definition of *g*. In the

example community (Fig 8A-E), $s=6$ and $n=13$. Thus, *G* is represented as (Fig 8A):

$$G = \{n_1 g_1, n_2 g_2 \dots n_s g_s | n \in \mathbb{N}\} \tag{3}$$

where *s* is the species richness. When characterizing *G* via shotgun metagenomics, the ith genome, g_i,

can be sequenced at *K* unique sections given a characteristic read length, *k*, and average genome size, *l*,

341 in number of base pairs (Fig 8B). Thus, the number of physically-distinct k-sized reads, K, associated

342 with the ith genome, g_i, within G is equal to:

343 $K_{g_i} = l(g_i) - k + 1$                  (4

344 Note that equation 4 considers homologous DNA as unique fragments. From equation 4, the

345 metagenome, g_{MG}, for g_i is defined as the set of all physically-distinct k-sized reads (Fig 8C) or:

346 $g_{MG,i} = \{g_{i,1,1+k}, g_{i,2,2+k} \cdots, g_{i,Kg_i,Kg_i+k}\}$         (5

347 where the subscripts for g_i represent a given k-sized read spanning from an arbitrary starting base pair

348 to the arbitrary starting base pair plus k. By substituting $g_{MG,i}$ into all g for equation 3 (Fig 8D), the

349 metagenome for a microbial community, G_{MG}, is derived to be:

350 $G_{MG} = \{n_1 g_{MG,1}, n_2 g_{MG,2} \dots n_s g_{MG,s} | n \in \mathbb{N}\}$        (6

351   while the population of physically-distinct k-sized reads in the metagenome, G_{MG} (Fig 8E), is

352 represented as:

353 $K_{MG} = \{g_{MG,1}, g_{MG,2} \cdots g_{MG,s}\}$            (7

354   From equation 5, one can determine the cardinality, or the total number, of physically-distinct

355 k-sized reads in associated with G_{MG} (expressed as $|K_{MG}|$). When attempting to fully sequence G_{MG}

356 using shotgun metagenomics, we assume that sampling events (sequence reads) are independent and

357 are sampled with replacement. In fact, Illumina sequencing technology sequences reads in parallel via

358 the individual DNA fragments binding to individual clusters. Furthermore, the fragmented DNA cannot

359 be sequenced twice as the sequencing process is destructive (33). Nonetheless, the mass of DNA

360 extracted from a target environment represents a negligible fraction of the total DNA existing in that

361 environment. As the relative abundance of the k-sized reads in K_{MG} does not change when DNA is

362 extracted from an environment, sampling events can be treated as independent and thus, DNA sampling

363 reduces to sampling with replacement. If the proportion DNA mass extracted had a significant impact

on the remaining mass of DNA in the environment, then one would be more suited to sequence all the

DNA versus a smaller proportion of the DNA. The sequencer should have no impact on sampling

assuming no sequencing errors due to misreading or spatial sampling issues (i.e., clonal density issues).

These issues do exist, but for the sake of a first order approximation, these biases can be ignored.

The probability of sequencing all elements in G_{MG} reduces to a coupon collectors problem (34)

by making the above assumptions. Using the general functional form for calculating expected samples

for sampling all unique elements in a set (equation 13b in (8)), one can predict the number of sequences

necessary to sequence all elements in K_{MG}, such that the expected number of sequences, $\text{E}(G_{MG})$, is:

$$\text{E}(G_{MG}) = \int_0^\infty \left(1 - \prod_{j \in K_{MG}} (1 - e^{-p_j t})\right) dt \qquad (8$$

where j is a given element within K_{MG}, t is the number of sampling events, and p_j is equal to the

proportion of the jth k-sized read within a given population of k-sized reads. p_j can be expressed as

follows:

$$p_j = \frac{n_i \times j \in K_{MG}}{|G_{MG}|} \qquad (9$$

where n_i is the respective abundance for the species whose MAG contains the jth k-sized read within

K_{MG}, and $|G_{MG}|$ is the cardinality of G_{MG}, or the total number of k-sized reads in the metagenome, G_{MG}.

MODELING EXPECTED SEQUENCES

Equation 8 provides an estimate for the total number of sequences to sequence all K_{MG}. The

influence of increasing species richness (i.e., s in equation 3) on the expected number of sequences was

tested for four hypothetical communities. The first community had an even structure such that all

physically-distinct DNA fragments were equally distributed across all K_{MG}. In the second community,

90% of the physically-distinct DNA fragments were equally distributed in 50% of K_{MG}, and the remaining

10% of the physically-distinct DNA fragments were distributed equally across the remaining 50% of

K_{MG}. This community represented a community with relatively moderate species evenness. In the third

community, 90% of the physically-distinct DNA fragments were equally distributed across 10% of K_{MG},

and the remaining 10% of the physically-distinct DNA fragments were distributed equally across the

remaining 90% of K_{MG}. This community represented a community with relatively low species evenness.

The last community had 10 equally-sized groups. The abundance of the physically-distinct DNA

fragments in each group was based on the function form of a lognormal community (35) which has been

observed in microbial populations (e.g., (17, 36)), such that:

$S(R) = S_0 e^{-a^2 R^2}$ (10

where S_0 was treated as the maximum relative of abundance ($S_0 = 1$), a was the inverse width of the

distribution, R was treated as the positive octave range spanning 0 to 9, and $S(R)$ represented the

abundance for a given group. For the lognormal abundance distribution in Fig 3D, a was set to a value

of 0.2. Each hypothetical community started with $|K_{MG}| = 1 \times 10^2$. $|K_{MG}|$ incrementally increased at 10

equally-spaced, linear steps to a maximum of $|K_{MG}| = 1 \times 10^6$. As $|K_{MG}|$ increased, all community

structures remained constant. Graphical representation of rank abundance in Fig 3A was normalized by

a given $|K_{MG}|$ to reflect that populations retained the same structure even as population size varied. We

defined a normalized rank abundance r_n such that

$r_n = \frac{r}{s}$ (11

where r and s are untransformed rank abundance and richness, respectively. For each community, at

each step, the expected number of sequences was calculated using equation 8. The expected number of

sequences as a function of $|K_{MG}|$ were modeled with linear regressions.

Equation 8 gives the expected number of sequences required to sequence any sized community

to exhaustion. Numerical sequencing simulations were performed to determine the number of

sequences necessary to sequence a subset of all physically-distinct DNA (K_{MG}). These numerical

sequencing simulations were applied to four hypothetical community structures described above.

Numerical simulations were performed such that $|K_{MG}| = 3 \times 10^7,\ 4 \times 10^7, 5 \times 10^7, 7 \times 10^7, 9 \times$

10^7, and 1×10^8. During each of these simulations, the parameters read length (k) and average

genome size (l) were set to 100 and 1×10^6, respectively, for all g. Random elements from K_{MG} were

selected with replacement to simulate sequencing events. Numerical simulations were performed until

the percentage of $|K_{MG}|$ sequenced was 50%, 70%, 90%, 95%, 99%, or 100%. A weight distribution

was applied to elements in a given K_{MG}. The weight distribution biased sequencing to reflect the

relative abundances of the four hypothetical communities described above. The percentage of $|K_{MG}|$

sequenced was evaluated every 1×10^7 sequences. Numerical simulations were performed in triplicate

for all $|K_{MG}|$ and all target fractions of $|K_{MG}|$.

We explored the influence of community evenness on required sequencing effort by performing

sequencing simulations on 6 different lognormally-distributed communities. The numerical sequencing

simulations followed the simulations described above. The 6 lognormal communities were modeled

such that each community had S_0 =1, R=10, and $|K_{MG}| = 1 \times 10^7$. Values of a for the 6 lognormal

distributions were a=0, a=0.005, a=0.008, a=0.01, a=0.015, and a=0.02. Evenness was represented

using the Pielou evenness index (11), or the ratio of the Shannon diversity index (37) for an observed

community to an even community of equal richness. Shannon diversity was calculated in the context of

a metagenomes such that:

$H_{MG} = \sum_{j \in K_{MG}} -p_j log(p_j)$ (12

where p_j is the proportion that the jth k-sized read represents among all physically-distinct DNA sequences

in the metagenome. Thus, the Pielou evenness index (11) was calculated such that:

$J = \dfrac{H_{MG}{'}}{H_{MG,max}}$ (13

where J was the Pielou evenness index, $H_{MG}{'}$ was the metagenome Shannon diversity index, and

$H_{MG,max}$ represented the metagenome Shannon diversity index when all p_j were equal (a=0).

Lastly, numerical simulations were performed to determine the sequencing effort necessary to

achieve a target fraction for an individual metagenome (g_{MG}). Target fractions increased from 0.5 to 1

at 100 linearly-spaced intervals. The fraction of the metagenome community (G_{MG}) that g_{MG}

represented varied from 1% to 100% in 30 lognormally-spaced intervals. The target genome sizes

varied such that $l=0.5\times10^6$, $l=1\times10^6$, $l=2\times10^6$, $l=3\times10^6$, $l=5\times10^6$, $l=10\times10^6$, $l=15\times10^6$, and $l=20\times10^6$.

The sequencing effort for a given combination of target fraction, genome size, and fraction of the

metagenome community was modeled using the gam function (mgcv R package; (38)). For modeling

purposes, target fraction was raised to the 6th power and both genome size and sequences were log-

transformed. The number of smooth dimensions for fraction of community, genome size, and target

fraction were heuristically varied till the resulting fit demonstrated residuals with a normal distribution.

Note that the objective here was not build a predictive model but simply a first order approximation for

simulations performed here. The generalized additive model was extrapolated for genome relative

abundances spanning 0.01 to 0.0001 at all sizes and target fractions. The extrapolation was necessary

due to the computational limitation of analyzing genome relative abundances <0.01.

RELATING MAG RESPONSE TO THE THEORETICAL SEQUENCING MODEL

Sequencing effort (in base pairs) used for predicting complete-genome equivalents for the sequence

read datasets maize soil, an estuarine sediment, the surface ocean, and the human gut, were related to a

genome relative abundance utilizing the generalized additive model presented in the previous section.

This was accomplished by setting a constant genome size and target fraction (0.5) and performing a

linear regression between genome relative abundance (a genome's fraction of a community) and log-

transformed base pairs. The linear regression was performed with genome sizes of 1 Mbp, 5 Mbp, and

20 Mbp. Complete-genome equivalents were predicted as a function of genome relative abundance

sequenced to a target fraction of 0.5.

Data Availability

All simulations and codes used for modeling sequencing effort are freely available on Github at:

https://github.com/taylorroyalty/sequence_simulation_code. The GRASE GUI application is available

at: http://shinyapps.io/adsteen/GRASE.

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Tables

Table 1. Estimates of fit coefficients for the Gompertz equation (equation 2) for the complete-genome equivalents as a function of sequencing depth in published datasets from ocean surface water, estuarine sediment, maize soil, and the human gut. p values for all coefficients were <<0.05.

Environment	A (±SE)	μ (±SE)	λ (±SE)	Yield*
Ocean Surface Water	97.67 (4.15)	5.84 (0.31)	1.16 (0.33)	0.88
Estuary Sediment	26.25 (2.11)	1.63 (0.06)	3.70 (0.24)	0.86
Maize Soil	43.65 (1.98)	1.43 (0.07)	6.13 (0.60)	0.70
Human Gut	17.49 (1.02)	5.01 (0.46)	0.67 (0.14)	0.90

*Calculated as the ratio of the maximum complete-genome equivalents experimentally observed to maximum complete-genome equivalents (A)

Table 2. Summary of sequence datasets analyzed with the MAG pipeline.

Environment	NCBI SRA Accession	Sequencing Platform	Total Reads*	High Quality Bases*	General Notes	Citation
Ocean Surface Water	ERR599029	Illumina HiSeq 2000	3.372×10^8	3.340×10^{10}	Caribbean Sea (5 mbsl)	(23)
Estuary Sediment	SRR5248164	Illumina HiSeq 2000	1.137×10^8	1.589×10^{10}	Sulfate Zone (8-10 cmbsf)	(4)
Maize Soil	SRR351473	Illumina HiSeq 2000	4.727×10^8	3.824×10^{10}	Surface Soil	(24)
Human Gut	SRR5127631	Illumina HiSeq 2000	5.095×10^7	4.847×10^9	--	(25)

* Combined forward and reverse paired-end reads.

Figures

Fig 1. A flow diagram illustrating the workflow for sequencing experiments.

Fig 2. The influence that the parameters A, μ, and λ had on the Gompertz equation (A). The property of the Gompertz equation that each parameter influences is colored red. Complete-genome equivalents (B), mean MAG completeness (C), mean MAG contamination (D) as a function of simulated sequencing effort (Gbp) for sequence datasets of the human gut, maize soil, estuarian sediment, and surface ocean microbiomes, using the pipeline described in the methods section. Translucent lines in (B) correspond to nonlinear least squares fits of the Gompertz equation to the respective environmental dataset.

Fig 3. Average expected sequences required to fully sequence four different community structures, one with relatively high community evenness (A), relatively moderate community evenness (B), relatively low community evenness (C), and one with a lognormal community structure (D), were predicted using linear regressions (E) and the log of $|K_{MG}|$ from equation 7 as a predictor.

Fig 4. Sequences necessary to reach variable target sequencing depths (colors) for four different community structures, one with relatively high community evenness (A), relatively moderate community evenness (B), relatively low community evenness (C), and one with a lognormal community structure (D). Red translucent lines correspond with linear regression curves for the respective community in Fig 3E.

Fig 5. Numerical sequencing simulations applied to 6 hypothetical communities with different lognormal distributions that were defined by the parameter, *a*, from equation 10 (A). The number of sequences necessary to sequence a target fraction of $|K_{MG}|$ (dashed contours) as a function of the Pielou evenness index, *J*, for a given lognormal community structure (B).

Fig 6. Numerical sequencing simulations show the number of bases (color bar) required to sequence a target fraction of a genome as a function of that genome's relative abundance in the community metagenome. Genomes evaluated were 0.5×10^6 (A), 2×10^6 (B), 5×10^6 (C), 10×10^6 (D), and 20×10^6 (E) base pairs long.

Fig 7. Complete-genome equivalents as a function of a target genome's relative abundance when sequenced to a target fraction (0.5) for human gut, maize soil, estuarian sediment, and surface ocean sequence datasets. Sequencing effort was converted to genome relative abundance sequenced to target fraction using the general additive model presented in Fig 6. The translucent shaded areas correspond to uncertainty from the target genome size (1 Mbp, or the lower-bound, to 20 Mbp, or the upper-bound) while the solid lines correspond to genome sizes 5 Mbp.

Fig 8. A cartoon illustrating an example microbial community (G), the metagenomes for

genomes ($g_{MG,i}$) (as defined in equation 5) within G, and the overall metagenome for the given

microbial community (G_{MG}). In this example, there are 6 genomes ($s=6$) and a total of 13 individual

microbes. (A) Black circles represent individual microbes whose genomes are averaged together, g.

The average genome, g, are indicated by different color inner-circles. (B) Individual average genomes

can be sequenced at K unique positions depending on the characteristic read length, k, of a sequencer.

(C) All unique positions that can be sequenced for a given genome, g, defines the metagenome, g_{MG},

for the ith genome, g_i. (D) Replacing all individual genomes in (A) with metagenomes, g_{MG}, gives the metagenome of the microbial community, G_{MG}.

Sample Microbial Community

Sequencing Effort

Bioinformatic Pipeline

Metabolic Characterization

<u>Knowns</u>

1) Target Environment
2) Sequencing Depth
3) Software/Assembly Parameters
4) Reference Databases

Community Structure

<u>Unknowns</u>

1) Depth and Coverage
2) Pipeline Efficiency
3) Pipeline Accuracy

A

μ

λ

D

High Evenness: 1.06x+1.84
Medium Evenness: 1.06x+3.37
Low Evenness: 1.06x+4.03
Lognormal: 1.07x+4.72

Equation 2

Equation 3

Equation 4

Equation 5

Equation 6