MICOM: metagenome-scale modeling to infer metabolic interactions in the gut microbiota.

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Abstract

Compositional changes in the gut microbiota have been associated with a variety of medical conditions such as obesity, Crohn's disease and diabetes. However, connecting microbial community composition to ecosystem function remains a challenge. Here, we introduce MICOM - a customizable metabolic model of the human gut microbiome. By using a heuristic optimization approach based on L2 regularization we were able to obtain a unique set of realistic growth rates that corresponded well with observed replication rates. We integrated adjustable dietary and taxon abundance constraints to generate personalized metabolic models for individual metagenomic samples. We applied MICOM to a balanced cohort of metagenomes from 186 people, including a metabolically healthy population and individuals with type 1 and type 2 diabetes. Model results showed that individual bacterial genera maintained conserved niche structures across humans, while the community-level production of short chain fatty acids (SCFAs) was heterogeneous and highly individual-specific. Model output revealed complex cross-feeding interactions that would be difficult to measure *in vivo*. Metabolic interaction networks differed somewhat consistently between healthy and diabetic subjects. In particular MICOM predicted reduced butyrate and propionate production in a diabetic cohort, with restoration of SCFA production profiles found in healthy subjects following metformin treatment. Overall, we found that changes in diet or taxon abundances have highly personalized effects. We believe MICOM can serve as a useful tool for generating mechanistic hypotheses for how diet

32   and microbiome composition influence community function. All methods are implemented in the

33   open source Python package, which is available at https://github.com/micom-dev/micom.

Introduction

The composition of the gut microbiome can influence host metabolism (1) and has been

associated with a variety of health conditions such as obesity, Crohn's Disease, diabetes and

colorectal cancer (2–6). However, the causal roles played by the gut microbiota in host physiology

and disease remain unclear. Several studies have mapped individual gut microbial genes to

functions (7–9). However, these mappings are largely qualitative, as the presence of a particular

gene does not guarantee expression of a functional enzyme. An alternative strategy to quantify

the metabolic capacity of a microbial community is to use computational models for inferring fluxes

in biochemical networks (10, 11). While direct experimental measurement of fluxes by carbon or

nitrogen labeling is costly, one can readily estimate the metabolic fluxes of a model organism

using genome-scale metabolic models. For individual bacteria, metabolic modeling using flux

balance analysis (FBA) has been a valuable tool for exploring metabolic capacities under varying

conditions and has been used extensively in basic research, biochemical strain engineering, and

in vitro models of bacterial interactions (12–15). In FBA, fluxes are usually approximated from a

genome-scale model containing all known biochemical reactions by maximizing the production of

biomass under constraints mirroring enzymatic, thermodynamic and environmental conditions

(13). For instance, one can restrict metabolic import fluxes to those whose substrates are present

in the growth medium (12, 14, 16) in order to simulate a particular nutrient environment. Extending

FBA to microbial communities is challenging due to the necessity of modeling metabolic

exchanges between many taxa and selecting an appropriate objective function to account for

potential tradeoffs between species and community growth rates.

Maximizing the community growth rate is at odds with maximizing individual species

growth rates. Multi-objective methods, like OptCom, attempt to find the joint maximum of

individual and community growth rates (17). However, these multi-objective methods are limited

to smaller-sized communities. The human gut microbiome, on the other hand, may contain up to

several hundred distinct species (18). An additional challenge is the integration of relative

abundances obtained from 16S amplicon or metagenomic shotgun sequencing into a community

FBA model. This is particularly important for accurately inferring the metabolic exchanges taking

place between different species within the community. A very abundant species should import

and export much greater absolute quantities of metabolites than a very rare species, which in turn

impacts the resulting community-level biochemical fluxes. Despite the challenges, genome-scale
metabolic modeling of microbial communities holds great promise as a tool for estimating the
metabolic potential of an individual's gut microbiome. In particular, this approach could yield
valuable insights into possible metabolic mechanisms underlying host disease states.

Here, we present a computational approach that efficiently extends metabolic modeling to
entire microbial communities. Using a two-step optimization procedure, we were able to simulate
growth and metabolic exchange fluxes for metagenome-scale metabolic models of ecologically
diverse bacterial systems. Additionally, we explicitly included microbial abundances from
metagenomic shotgun sequencing and realistic dietary inputs in order to make quantitative,
personalized, metabolic predictions. This entire strategy is implemented in an open-access
Python software package called "MICOM" (**MI**crobial **COM**munity).

We tested our approach by applying MICOM to a balanced data set of 186 Danish and
Swedish individuals, including healthy controls, patients with type 1 diabetes, and patients with
type 2 diabetes with and without metformin treatment. We show that individual bacterial growth
rates vary greatly across samples and are correlated with independently measured replication
rates. We quantified exchanges between the gut microbiota and gut lumen and studied the effect
of the microbiota composition on the production of short chain fatty acids (SCFAs) across samples
from healthy and diabetic individuals. Overall, we found that MICOM predicted a bimodal usage
pattern of dietary metabolites, ecological interactions between microbes tended to be community-
specific and largely competitive, key gut genera associated with health participated in the largest
number of ecological interactions, inferred SCFA production was lower in diabetic patients, and
targeted dietary or probiotic interventions had unique functional consequences for each individual.
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Results

A regularization strategy for microbial community models.

Metabolic modeling is commonly applied to model a single strain of bacteria in log-phase, where
the growth rate is approximately constant and the log of the bacterial abundance increases
linearly with time. Modeling bacterial growth in natural environments is often more complex than
this, but some information on environmental context can be extracted from the relative
abundances of bacterial taxa. Within a single individual, and in the absence of persistent dietary

changes, gut microbial relative abundances tend to fluctuate around a fixed median value over
month-to-year timescales (19–21). This is consistent with a steady state model where bacterial
growth is in equilibrium with a dilution process that continuously removes biomass from the
system (22). Under this approximation bacterial growth rates are constants μ_i (in 1/h), which is
compatible with the assumption of FBA. All bacteria in the microbial community then contribute to
the production of total biomass, with an overall growth rate constant μ_c. The community growth
rate μ_c is obtained from the individual growth rates μ_i by a weighted mean, with the relative
contribution of species i (a_i) to the total biomass serving as the weight (17, 22).

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$$\mu_C = \sum_i a_i \mu_i \tag{1}$$

Even though FBA can be used to obtain the maximum community growth rate, one can see from
Equation 1 that there is an infinite combination of different individual growth rates μ_i for any given
community growth rate μ_c (see Fig. 1A for an example). Various strategies have been employed
in order to deal with this limitation, where the simplest strategy is to report any one of the possible
growth rates distributions for μ_i. Other approaches attempt to find the set of growth rates that
maximize community growth and individual growth at the same time (17), but this is
computationally intensive and may not scale well to the species-diverse gut microbiome (18, 23).
Thus, we formulated a strategy that allows us to identify a realistic set of individual growth rates
μ_i and scales to large communities. The simplest case of a microbial community is a community
composed of two identical clonal strains, each present in the same abundance. Assuming that
the maximum community and individual growth rates are equal to 1.0 there are now many
alternative solutions giving maximal community growth (Fig. 1A). However, the two populations
are identical and present in the same abundance, so one would expect that both grow at the same
rate. In order to enforce a particular distribution of individual growth rates one can try to optimize
an additional function over the individual growth rates μ_i. This is known as regularization and a
feasible regularization function should enrich for biologically relevant growth rate distributions. As
a heuristic, our minimal requirement for a feasible regularization function was consistency with
the observed metagenomic abundances. This means that a taxon that is observed in the data
should be able to grow. Thus the growth rate of a taxon should be non-zero if its abundance is
non-zero. We show in Supplemental Text S1 that no linear regularization function can comply
with that requirement whereas a simple quadratic regularization, also known as L2 regularization,
does fulfill that requirement (24, 25). L2 regularization is known to distribute magnitude over all

variables, which is also consistent with a maximization of individual growth rates and thus forms
a heuristic for the simultaneous maximization of individual and community growth rates.

L2 regularization can be readily integrated into FBA as a quadratic optimization problem, which
is not necessarily true for any generic function. In the previous example of two identical strains
only the L2 norm correctly identifies the solution where both strains grow at the same rate as
optimal. Additionally, the L2 norm has a unique minimum. Thus, there is only one configuration of
individual growth rates μ_i that minimizes the L2 norm for a given community growth rate μ_c. In
practice, maximal community growth might only be achievable if many taxa are excluded from
growth, for instance by giving all resources to a fast-growing subpopulation. Again, this is
inconsistent with reality if one has prior knowledge that the other taxa are present in the gut and
should be able to grow. Instead of enforcing the maximal community growth rate one can limit
community growth to only a fraction of its maximum rate, thus creating a tradeoff between optimal
community growth and individual growth rate maximization. Community growth maximization
requires full cooperativity, whereas the L2 norm minimization represents selfish individual growth
maximization. Thus we call our two-step strategy of first fixing community growth rate to a fraction
of its optimum and then minimizing the L2 norm of individual growth rates a "cooperative trade-
off". Even though it is difficult to formulate a closed form solution for this two-step optimization,
one can obtain a solution for the second optimization (minimization of regularization term) when
dropping additional constraints for growth rates (see Supplemental Text S1 for derivation). In that
case, growth rates are given by:

150                                $$\mu_i = \frac{\alpha \mu_c}{a^T a} a_i .$$ (2)

Thus, optimal growth rates will be approximately correlated with abundance where the slope
depends on the abundance distribution and the maximum community growth rate.

We found that computation time generally scaled well with the community size (with most
individual optimizations taking less than 5 minutes) when using interior point methods, which are
known to provide better performance for larger models (26). However, we found that it was difficult
to maintain numerical stability with large community models. None of the tested solvers were able
to converge to optimality when solving the quadratic programming problem posed by the L2 norm
minimization (see Methods). Thus, we used a crossover strategy to identify an optimal solution to
the L2 minimization (see Methods).

Figure 1: Regularization of growth rates. (A) Regularization values for a toy model of two identical *E. coli*

populations. Two alternative solutions are shown with different individual growth rates and respective

values of regularization optima. Here L1 denotes minimizing the sum of growth rates whereas L2 denotes

minimization of the sum of squared growth rates. Only L2 regularization favors one over the other and

identifies the expected solution where both populations grow with the same rate. (B) Effect of different

trade-off values (fraction of maximum community growth rate) on the distribution of individual genus

growth rates. Zero growth rates were assigned a value of 10^{-16} which was smaller than the observed non-

zero minimum. Growth rates smaller than 10^{-6} were considered to not represent growth (shaded area).

(C) Pearson correlation between replication rates and inferred growth rates under varying trade-off

values. "None" indicates a model without regularization returning arbitrary alternative solutions (see

Methods). The dashed line indicates a correlation coefficient of zero.

Regularization by cooperative trade-off yields realistic growth rate

estimates.

In order to test whether cooperative trade-off yields realistic growth rates, we implemented and

applied it to a set of 186 metagenome samples from Swedish and Danish individuals (27),

consisting of healthy individuals, individuals with type 1 diabetes and individuals with type 2

diabetes stratified by metformin treatment (a known modulator of the gut microbiome) (28).

Relative abundances and cleaned coverage profiles for a total of 239 bacterial genera and 637

species were obtained with SLIMM (29) from previously published metagenomic reads (27, 29)

as described in the Methods section. We used ratios in coverage between replication initiator and

terminus as a measure for replication rates, which have been reported to be good proxies for

bacterial growth rates *in vivo* (30). This provided a set of 1571 strain level replication rate

measurements across the 186 samples that were used for validation of the inferred growth rates

(1062 and 1113 on genus and species levels respectively, see Methods). Abundance profiles for

all identified genera across all samples were connected with the AGORA models, a set of
manually curated metabolic models which currently comprises 818 bacterial species (31). 93 gut-
associated genera within the AGORA reconstructions represented more than 96% of
metagenomic reads across the 186 samples (91% vs 94%, see Table 1 "genus" row). Even
though the cooperative tradeoff strategy is applicable to species or even strain-level data, the
AGORA reconstructions accounted for only 52% of all bacterial species in the data set. Thus, we
decided to perform community model construction separately on species level as well as the
genus level, which covered a larger fraction of the observed microbiome. To accomplish this,
individual strain models from AGORA were pooled into the higher phylogenetic ranks (see
Methods). After removing low abundance taxa (<0.1% for genera and <0.01% for species), the
resulting communities contained between 12 and 30 taxa at the genus level and between 23 and
81 taxa at the species level. Each taxon was represented by a full genome-scale metabolic model
and connected by exchange reactions with the gut lumen, thus yielding two sets of 186 complete
metagenome-scale metabolic models (one set for the species level and one for the genus level).
We used the relative read abundances as a proxy for the relative biomass of each taxa in each
sample (see Methods). Even though relative abundances from shotgun metagenomes are not an
exact representation of bacterial mass (in grams dry weight), we argue that the discrepancy
between the two is probably much smaller than the variation in taxon abundances, which spans
several orders of magnitude (18).

taxa	unique taxa	assigned reads	with model
kingdom	1	100% ± 0%	100% ± 0%
phylum	22	100% ± 0%	99% ± 0%
class	32	100% ± 0%	99% ± 0%
family	102	100% ± 0%	91% ± 0%
genus	239	94% ± 5%	91% ± 5%
species	637	79% ± 9%	52% ± 9%

Table 1: Distribution of taxa assignments across ranks. Only reads classified as bacteria were considered.
Shown are the number of unique taxa for each rank together with the percentage of mapped reads that
could be uniquely assigned to taxa within the rank, as well as the percentage of reads whose taxa had at

least one representative in the AGORA genome-scale metabolic models. Percentages are shown as mean
± standard deviation across the 186 samples.

The data on 186 individuals used in this analysis did not include diet, metabolomics, or data on
total microbial load. Thus, we were limited to study metabolic effects that are driven by microbiota
composition alone and not by additional factors such as diet or total bacterial biomass. To use a
moderately realistic set of import constraints for the community models, we modeled all individuals
as consuming an average Western diet (32). Import fluxes for external metabolites were based
on a reported set of fluxes for an average Western diet (31, 33). To account for uptake in the
small intestine, we reduced all import fluxes for metabolites commonly absorbed in the small
intestine by a factor of 10.

To evaluate the performance of the cooperative tradeoff we compared the inferred growth rates
with the replication rates obtained directly from sequencing data. First, to establish a baseline we
ran an optimization that only maximized the community growth rate and used the distribution of
growth rates returned by the solver when applying no regularization. This was followed by
applying the cooperative trade-off strategy with varying levels of suboptimality ranging from 10%
to 100% of the maximum community growth rate. As stated above, we observed that simply
optimizing the community growth rate with no regularization of the individual growth rates led to
solutions where only a few taxa grew with unreasonably high growth rates (doubling times shorter
than 5 minutes), whereas the rest of the microbial community had growth rates near zero
(compare Fig. 1B with strategy marked by "none"). Consequently, the resulting model growth

rates were uncorrelated with replication rates (mean Pearson rho=-0.02). Adding the L2 norm

minimization while maintaining maximum community growth allowed more genera to grow (see

Fig. S1) but yielded growth rates that were anticorrelated with replication rates (mean r = -0.11).

Lowering the community growth rate to suboptimal levels strongly increased the growing fraction

of the population (Fig. S1) and led to a much better agreement with replication rates for tradeoff

values smaller than 70% (mean Pearson rho ≅ 0.4). Calculating correlations across all samples

rather than within samples showed a similar tendency, with no regularization showing no

correlation with replication rates (r = -0.05, p = 0.07) and increased agreement up to a tradeoff of

50% (r = 0.21, p = 2e-12). The lower magnitude correlations in the across samples setting is likely

due to differences in diet or bacterial load across people that were not taken into account. Overall,

the best agreement with the observed replication rates across and within samples was observed

at 50% sub-optimal community growth. We observed similar performance with the species level

models (Fig. S2). However, the best agreement with in vivo replication rates was observed for a

tradeoff parameter of 0.7 (Fig. S2C). Because genus level model performed equally well as

species level models but represented a higher percentage of observed reads (compare Table 1)

we decided to continue all further analysis with the genus models and a tradeoff parameter of 0.5.

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252   Figure 2: Non-zero growth rates (> 10^{-6}) across genera obtained by cooperative trade-off (50% maximum
253   community growth rate). Each small filled point denotes a growth rate in one of the 186 samples and
254   larger points with white fill denote the mean growth rate for the genus (see Methods). Genera are sorted
255   by mean growth rate from left (lowest) to right (largest).

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Growth rates are heterogeneous and depend on the community composition

A tradeoff of 50% maximal community growth led to good agreement with replication rates. Bacterial communities showed an average doubling time of about 6 hours, where individual genera had an average doubling time of 11 hours. Community growth did not vary substantially across samples (0.246 +- 0.002 1/h) indicating that each individuals' microbiota were almost equally efficient at converting dietary metabolites into biomass at the community level. However, we found that individual genus-level growth rates often varied over five orders of magnitude (Fig. 2). *Bacteroides* was predicted to be the fastest growing genus overall and was closely followed by *Eubacterium*, which is consistent with the ubiquitous presence of these abundant taxa in microbiome samples (34, 35).

Figure 3: Co-dependencies of growth rates. (A) Relationship between abundance and growth rate across samples. Shown in the larger scatter plot are growth rates and abundances for the first 10 samples. Each dot denotes one genus in one sample and is colored by sample provenance. Dashed lines denote the linear relationship between growth rates and abundances predicted by Equation 2 for each sample. The black box demonstrates how varying slopes (i.e. as community evenness declines, so does the within-sample slope) can result in negative correlation between abundance and growth rate across samples. Smaller inset scatter plot shows data from all samples (Pearson rho=0.69, n=39,815). (B) Growth rate interactions between genera as estimated by genus knockouts. Shown are only interaction that induce a mean growth rate change of 0.1 across all samples (i.e. ubiquitous interactions). Color of edges indicates change of growth rate and type of interaction. Red edges denote competition where removal of one genus increases

the growth rate of the other and blue edges denote cooperation or syntropy where the removal of one genus

lowers the growth rate of the other. Nodes are colored by the degree (number of total connections) from

lime (few) to dark blue (many).

In the absence of additional constraints, L2 regularization will result in growth rates that are

linearly dependent on the taxa abundances (see Equation 2 and Supplemental Text S1).

However, this requires some simplifying assumptions that may not be met in the particular

constraints of the full metabolic community models. We compared growth rates estimates

obtained from numerical optimization with the approximation from Equation 2. We found that

growth rates obtained with the cooperative tradeoff usually followed the derived linear

relationship, albeit with a large variation (mean R^2 = 0.94, sd = 0.34, Fig. 3A). Deviations from

that relationship were mostly observed for small growth rates (see Fig. 3A) which could not reach

the suggested growth rate due to additional constraints on growth. Thus, the linear relationship

between growth rates and abundance holds for most growth rates but is likely inaccurate for very

small growth rates. It is important to note that even though abundances are positively correlated

with growth rates within a single individual this is not true across samples where one can observe

a negative correlation for abundant taxa (see Fig. 3A). This is a consequence of the coefficient in

Equation 2, which depends on the actual abundance distribution as well. In particular, the slope

of the linear relationship between abundance and growth rate will be the greatest if all taxa have

equal abundances and take its lowest value when one taxon dominates.

We observed a wide variation in individual taxon growth rates across samples. Because all of the

community models were constrained under the same diet this phenomenon was due to microbiota

composition only. To explain this variation in individual growth rates, we hypothesized that

different genera might influence each other's growth rate, either by competition or by cooperation.

In order to quantify growth rate interdependencies we performed *in silico* knockouts for each

genus in each sample and tracked the change in growth rates for all remaining genera in the

sample (see Methods). Here we found that the growth rate of each genus was influenced by

another genus in at least one of the 186 samples. As would be expected for bacterial species

competing for the same resources, most interactions were competitive (red edges in Fig. 3B).

However, we observed a distinct subset of bacteria that were interconnected by a network of

cooperative interactions, including *Akkermansia* and *Faecalibacterium* (blue edges in Fig. 3B,

also Fig. S3). Strikingly, genera participating in many interactions across all samples, such as

Bacteroides, *Eubacterium*, *Akkermansia*, *Alistipes* and *Faecalibacterium*, are known to be
ubiquitous members of the gut microbiome and are often associated with health (6, 36–40). We
found that the prevalence and strength of interactions were highly dependent on the composition
of the microbiome. The vast majority of strong growth interactions were present in only 1-5
samples, whereas all other samples showed very few strong interactions (Fig. S3). This result is
perhaps unsurprising, as many strong species-species interactions are thought to be destabilizing
to ecological communities (41, 42).
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Analysis of exchange fluxes reveals differential use of diet components and niche partitioning in the microbiota

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One of the major modes of interaction between the gut microbiota and the host is by means of
consumption or production of the metabolite pool in the gut. In our simulations all individuals were
under the same average Western diet (see Methods), which imposes an upper bound for the flux
of metabolites into the gut lumen. However this does not determine *a priori* which components of
the diet are consumed at what rate in each sample, because individual microbiota may consume
less than what is imposed by that maximum diet flux. We quantified this effect by obtaining all
import and export fluxes for each individual genus across all samples (1,613 exchange reactions
in each of 62 genera) as well as metabolite exchanges between the microbiota and the gut lumen
(152 metabolites). This was done in the absence of a metabolic model for enterocytes,
colonocytes or goblet cells due to the lack of a curated metabolic reconstruction and validated
objective function for those cells. A unique set of exchange fluxes was obtained by considering
the set of exchange fluxes with smallest total import flux for the growth rates obtained by the
cooperative trade-off (see Methods). This assumes that the microbiota competes for resources
with the host gut and will thus favor an efficient import that yields the maximum growth rate. This
also corresponds to the particular distribution of import fluxes an individual microbiota is most
adapted to.
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Even though the minimization of total import fluxes favors simpler media compositions, most
samples showed a diverse consumption of metabolites from the gut, particularly in the wide array
of carbon and nitrogen sources (Fig. 4A). There was a large set of metabolites that were
consumed across all samples but we also observed many metabolites with differential import

fluxes across individuals. In particular, we observed a bimodal distribution where microbiota either

consumed fibers and starches or branched chain amino acids (see indicated metabolites in Fig.

4A). This bimodal pattern did not correlate with health or disease states. As expected, all

communities showed net anaerobic growth.

Given the observed heterogeneity in taxon growth rates and the large number of interactions we

were interested in looking at the uptake rates of metabolites from the gut lumen by each genus.

There was overlap of metabolite usage across genera. On average 32% of the metabolites were

shared between any two genera in any sample (standard deviation of 13%, see Methods).

Figure 4: Microbiota import fluxes across samples. Exchange fluxes were calculated as the smallest set of

import fluxes that could maintain the genera growth rates obtained by the cooperative trade-off. (A) Import

fluxes across samples. Rows were normalized to their absolute maximum and colors denote the import

rate ranging from 0% to 100% maximum import. Metabolite groups of interest are marked by blue and red

lines (BCAA = branched chain amino acids). Column headers are colored by host metabolic health state

(CTRL = metabolically healthy, T1D = type 1 diabetes, T2D = type 2 diabetes, metformin +/- =

with/without metformin treatment). (B) Growth niche map for gut genera. Import fluxes for each genus in

each sample were reduced to two dimensions using t-SNE. Each point denotes a genus in one sample

and is colored and named by its genus.

Consequently, more than two thirds of metabolites were used differentially between pairs of

genera. To visualize the structure of metabolite consumption by individual bacterial genera in the

gut we used t-SNE dimensionality reductions on the individual genus-specific import fluxes (43).

This revealed clear genus-specific niche structure across samples, where individual genera could

be uniquely identified by their particular set of import fluxes (Fig. 4B). Here, taxa closer to one
another overlap more in consumed metabolites (Fig. 4B). For instance, *Bacteroides* and
Prevotella were relatively close to each other in the center of the map, which may help explain
the observed tradeoff between *Bacteroides* and *Prevotella* abundances across humans (44).
Blautia, *Desulfovibrio*, *Bifidobacterium* and *Eubacterium* had the most unique growth niches
overall. Genus identity alone explained 61% of the variance in import fluxes (Euclidean
PERMANOVA p=0.001). Thus, there was extensive growth niche partitioning between bacterial
genera.

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Figure 5: SCFA fluxes. (A) Production capacities of the major SCFAs stratified by population. Fluxes
denote total amount of SCFA produced by 1g of bacterial biomass in the gut. Stars denote significance
under Welch's t-test (* p<0.05, ** p<0.01). (B) Genus-specific fluxes for the three major SCFAs. Shown
are only genera with a relative abundance >1%. Fluxes denote total production/consumption for each
genus (see Methods) and are directed towards exports. Thus, positive fluxes denote production of the
metabolite and negative fluxes consumption. Genera are ordered by average relative abundance (relative
abundances shown in the first column) from top to bottom.

SCFA production is driven by extensive cross-feeding within the microbiota
and can be modulated by personalized interventions

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Given the association between SCFAs and disease phenotypes we investigated the degree of
SCFA production by the model microbiota (2, 45, 46). Intestinal cells have access to the full pool
of SCFAs in the gut lumen and would probably take up a significant fraction of those extracellular
SCFAs. Thus, the total export flux of any SCFA into the gut lumen by all taxa in a specific model
is a measure of host-available SCFA production by the microbiota (see Methods for details on

computation). Overall SCFA production for the major SCFAs showed large variations even in
healthy individuals, which indicates a large impact of gut microbiota composition on SCFA
availability. In particular we observed that Swedish individuals showed higher SCFA production
rates than the Danish individuals in the study. Butyrate production was diminished by about 2-fold
in Danish individuals with type 1 diabetes (Welch's t-test, p=0.004) but not in Swedish individuals.
Danish and Swedish individuals with type 1 diabetes had microbiota that produced more acetate
than healthy individuals (Welch's t-test, p=0.02 and 0.003, respectively; Fig. 5A). Metformin
treatment had a moderate effect in increasing butyrate productions in both cohorts, however this
effect was strongest when comparing Danish individuals with T1D and metformin-treated danish
individuals with T2D (Welch's t-test, p=0.003). Higher production of SCFAs was usually
accompanied by an increased consumption of SCFAs within the gut microbiota (Fig. S4). This is
consistent with prior findings in Danish and Chinese populations (4, 27, 47).

Decreases in butyrate production were usually accompanied by increases in acetate production.
This appeared to indicate SCFA cross-feeding within the microbiota, which we confirmed by
comparing the total production and consumption fluxes for each bacterial genus across all
samples (Fig. 5B). We observed that butyrate was almost exclusively formed in an acetate-
dependent manner from acetyl-CoA, which is the most prevalent butyrate-production pathway in
bacteria (48). In particular, production of butyrate depended on acetate production in the
community. This was enabled by an extensive cross-feeding between the genera. All SCFAs were
produced by a heterogeneous set of taxa, with acetate and propionate production being spread
out across most taxa in the system and butyrate produciton being somewhat more restricted to a
smaller set of taxa (Fig. 5B). Across samples, the most efficient butyrate producers were
Faecalibacterium, *Coprococcus*, *Roseburia*, *Anaerostipes* and *Lachnoclostridium*, all of which are
known butyrate producers (48, 49). However, the models also predicted consumption of acetate
by *Bacteroides* and consumption of butyrate by *Eubacterium*, which is not commonly observed *in*
vivo. Production of SCFAs was complemented by several other genera, generating a network of
SCFA cycling within the microbiota. SCFA production by any genus showed high variation across
samples and in some cases would even switch between consumption and production of a
particular SCFA, which shows how specific SCFA production is to a particular microbial
community (compare Fig 5B). Net production of SCFAs was low compared to overall production
(Fig. 5A and S4B) which indicates that most SCFAs in our models were cycled within the bacterial
community.

Finally, as a proof of concept for the utility of MICOM, we aimed to quantify the impact of targeted

interventions on the net consumption or production of SCFAs by the microbiota. For this we chose

three Swedish samples (normal, T2D metformin-, T2D metformin+). The impacts of a particular

univariate interventions were then quantified by using elasticity coefficients (50, 51), which are

dimensionless measures of how strongly a parameter affects a given flux (see Methods).

Univariate interventions included increasing the availability of a single metabolite in the diet or

increasing the abundance of a single bacterial genus. We observed that the effect of these single

interventions were very heterogeneous across all 3 samples (see Fig. 6). The strongest and most

common observed effects were to diminish overall SCFA production. However, we observed a

few interventions that were able to weakly increase SCFA production. There was a distinct set of

metabolites that would increase butyrate production in the T2D individuals but not in the healthy

individual. This was also dependent on metformin status. For instance, Arabinan increased

butyrate production in the metformin positive individual and D-Xylose increased butyrate

production in the metformin negative individual, but not *vice versa*. Thus, MICOM is able to

explore the potential functional consequences of targeted dietary or probiotic interventions, which

can differ greatly depending on the context of the microbiota in which the interventions are made.

Figure 6: Effect of interventions on SCFA production in 3 samples. Each row denotes a SCFA in a specific

individual and each column either denotes a diet component or bacterial genus. Colors denote the elasticity

(i.e. the percent change in SCFA production given a percent increase in the specific effector). Red denotes

interventions that would increase SCFA production and blue interventions that would decrease production.

Shown are only interventions with non-zero elasticities in at least one sample.

Discussion

There is a large amount of sequencing data on microbial communities available today. This is

mostly due to the falling cost of 16S rRNA sequencing or shallow shotgun sequencing (52). There

is wide interest in extracting information from sequencing data that goes beyond bacterial

proportions (53). Metabolic modeling incorporates a rich knowledge base from genomics and

biochemistry and is a valuable resource for adding value to existing datasets. Specifically, MICOM

allows for the integration of genome-scale metabolic models, dietary information in the form of

import flux bounds, and abundance data from metagenomic shotgun or marker gene sequencing.

This framework enables *in silico* mechanistic predictions concerning ecological interactions within

microbial communities, inferred exchanges between microbial communities and their

environment, and mechanistic hypotheses for how metabolic interactions can be modulated by

changes in the environment. The design of a reasonable metagenome-scale metabolic models is

challenging due to apparent tradeoffs between individual and community growth rates and issues

with computational tractability. Here, we provided a viable strategy that allows for complex

analysis of the metabolic consequences of variation in microbial community composition. Our

regularization strategy allows for fast identification of unique sets of individual growth rates, which

operate in biologically realistic ranges. Our assumption that there is a tradeoff between community

growth rate and individual taxon growth rates is supported by the observation that most microbial

communities are composed of a large number of species with non-negligible abundances.

Individual growth rates for bacterial genera varied greatly across samples (Fig. 2) and were tightly

coupled to genus abundances within a sample (Fig. 3A). However, there may be other

regularization strategies that provide better agreement with the underlying biology. Our validation

strategy using replication rates obtained directly from metagenomic data provides a simple

framework to test new regularization functions in the future. It seems that the large variation of

growth rates can be explained by a dependency of the growth rate on the presence of other

bacteria in the sample (Fig. 3B). Thus, bacterial growth in the gut microbiota is not only dictated

by abundance but also by taxon-taxon interactions.

Our predictions are somewhat limited by a variety of factors. For instance, the lack of metabolic

models for the major cell types of the gut epithelium (especially goblet cells, enterocytes and

colonocytes) and sample-specific metabolite availability in the gut lumen limits the accuracy of

MICOM's predictions. Additionally, the use of representative models for bacterial genera may also

have caveats. Available metabolic reconstructions are often based on laboratory strains that may

not represent the exact metabolic capacity of strains in the human gut. Thus, reconstructions may

lack certain metabolic pathways present in the sample and yield inaccurate results, especially

when not applying appropriate bounds for the underlying diet (33, 54). Model-predictions may

become more quantitative as better personalized data becomes available. The incorporation of

personalized data on diet and a better grasp on what fraction of metabolites are absorbed in the
small intestine should help to improve model-based predictions. Personalized reconstruction of
microbial community metabolic models directly from metagenomic sequencing data may provide
more accurate predictions as well, but this approach is currently limited by insufficient sequencing
coverage for low abundance taxa.

MICOM provided valuable ecological insights into the gut microbiota. For instance, the
cooperative tradeoff in Equation 2 indicates that a more diverse microbiome (i.e. higher evenness)
results in higher individual growth rates (on average) due to the magnitude scaling of the
abundance vector (denominator in Equation 2). We also found strong niche partitioning in the
model, where taxa showed minimal overlap with each other in resource utilization space. This
minimal overlap implies that there is likely an upper bound on alpha diversity in the gut, due to
the fact that growth niches eventually saturate and limit the number of taxa that can engraft. Even
though only about a third of metabolites were consumed by any pair of taxa in the models, this
small amount of niche overlap still resulted in resource competition between taxa. This was
particularly true for dominant taxa (e.g. *Bacteroides*), which tended to show competitive
interactions with many other genera, likely due to the comparatively higher resource requirements
of these abundant taxa for maintaining growth. This community-wide resource competition fits
well with the observed growth dependency on amino acid import fluxes across all taxa (Fig. 3A),
which is consistent with prior work that suggest that nitrogen may be the global limiting factor for
microbial growth in the gut (55). Finally, the methods here extend to any ecosystem containing
many microbial taxa and is applicable to abundance data summarized at various phylogenetic
ranks. As such MICOM can be employed to perform functional analyses on a wide range of
microbial ecosystems.

It has been difficult and time-consuming to obtain empirical evidence for the mechanistic basis of
gut-microbiota interactions. MICOM provides a high-throughput platform for generating
mechanistic hypotheses and running *in silico* experiments that would be impossible to perform *in*
vivo. Thus, we feel that the major application for MICOM is to provide detailed functional
hypotheses that can serve as targets for experimental validation. For example, MICOM reveals
widespread SCFA cross-feeding in the gut microbiota. The mere presence of butyrate producers
was not enough for stable SCFA production - acetate production was also required. Furthermore,
MICOM generated personalized predictions for how dietary and probiotic interventions influenced
SCFA production capacity. This basic approach could be extended to any number of clinically-

relevant metabolites. Thus, we hope that the method presented here will aid researchers in
leveraging existing gut microbiome data to design and test personalized intervention strategies.
530

Methods

Data availability and reproducibility

All data to reproduce the manuscript, intermediate results as well as Python scripts to reproduce
the figures in this manuscript are available in a data repository at
https://github.com/resendislab/micom_study. Metagenomic reads for the 186 individuals were
obtained from Forslund et. al. (27) and can be downloaded from the Sequence Read Archive
(https://www.ncbi.nlm.nih.gov/sra) with the SRA toolkit
(https://www.ncbi.nlm.nih.gov/sra/docs/toolkitsoft/). A full list of run accession IDs for the
individual samples is provided in the data repository ("recent.csv"). All algorithms and methods
used here were implemented in a Python package and can be easily applied to different data
sets. The Python package "MICOM" (**MI**crobial **COM**munities) along with documentation and
installation instructions are available at https://github.com/resendislab/micom. MICOM is based
on the popular COBRApy Python package for constraint-based modeling of biological networks
and is compatible with its API (56). The cooperative trade-off strategy as described here was
introduced to MICOM in version 0.9.0. The AGORA reference reconstructions with an already
applied average Western diet can be downloaded from https://vmh.uni.lu/#downloadview. Several
methods used in MICOM require an interior point solver with capabilities for quadratic
programming problems (QPs) for which there is currently only commercial software available.
MICOM supports CPLEX (https://cplex.org) and Gurobi (https://gurobi.org) both of which have
free licenses for academic use. Intermediate results that required those solvers are also provided
in the data repository to permit reproduction of our major conclusions.

Metagenomic shotgun data analysis

All metagenomic analyses were performed in R using an in-house pipeline which is available as
an open source package along with documentation at https://github.com/resendislab/mbtools.
Sample FASTQ files were downloaded using the SRA toolkit and trimmed and filtered using the
DADA2 "filter_and_trim" function (57) with a left trimming of 10 bp, no right trimming, a quality

cutoff of 10 and a maximum number of 2 expected errors under the Illumina model. Abundances
across different taxa levels were then obtained using SLIMM (29) which was chosen because it
supported one of the largest references (almost 5,000 reference bacterial genomes). In brief, all
sample FASTQ files were first aligned to the SLIMM reference using Bowtie2 saving the 100 best
matches for each read. Taxa abundance profiles were then obtained using SLIMM with a window
size of 100bps and assembled into a single abundance file. SLIMM coverage profiles resolved to
single strains where then used to infer replication rates using the iRep method (58). In brief,
coverage profiles were first smoothed with a rolling mean over 5kbp windows and only genomes
with at least a mean coverage of 2 and with at least 60% of total length covered were considered.
Coverage values were log-transformed, sorted, and the lowest and highest 10% of the data points
were removed to obtain the linear part of the curve. Replication rates were then inferred from the
slope of a regression on that linear part. An estimate for the minimum coverage was then obtained
from the intercept of the regression and only replication rates for strains with a minimum coverage
>2 were kept. No correction for GC content was performed. Before model construction, genus-
level and species-level quantifications for each sample were matched separately to the AGORA
models by name. The final quantification and mapping is provided in the data repository
("genera.csv" and "species.csv" at https://github.com/micom-dev/paper).

Strategies used in MICOM

576 Flux balance analysis obtains approximate fluxes for a given organism by assuming a steady
577 state for all fluxes in the biological system and optimizing an organism-specific biomass reaction.
578 Using the stoichiometric matrix S which contains reaction in its columns and metabolites in its
579 rows this can be formulated as a constrained linear programming problem for the fluxes v_i (in
580 mmol/[gDW h]):

$$maximize \; v_{bm}$$
$$s.t. \, Sv \; = \; 0$$
$$lb_i \; \geq v_i \geq ub_i$$

The biomass reaction v_{bm} is usually normalized such that it will produce 1g of biomass which
results in a unit 1/h corresponding to the growth rate μ of the organism. The upper and lower
bounds (lb_i and ub_i, respectively) impose additional thermodynamic constraints on the fluxes or
restrict exchanges with the environment (in the case of exchange fluxes). In order to describe a
community model containing several organisms each with a particular abundance a_i (in gDW) one

usually embeds each organism in an external compartment which represents the community
environment (for instance the gut lumen for models of the gut microbiota). Adding exchanges for
the environment compartment and exchanges between a particular organism and the
environment one obtains a community model with the following constraints:

595
$$maximize\ \mu_c = \sum_i a_i \cdot \mu_i$$

596
$$s.t.\ \forall i: Sv = 0$$

597
$$\mu_i = v_i^{\ bm} \geq \mu_i^{\ min}$$

598
$$lb_i \geq v_i \geq ub_i$$

599
$$lb_i^{\ ex} \geq a_i \cdot v_i^{\ ex} \geq ub_i^{\ ex}$$

600
$$lb_i^{\ m} \geq v_i^{\ m} \geq ub_i^{\ m}$$

Here, a_i denotes the relative abundance of genus i, μ_i its growth rate, v_i^{bm} its biomass flux, μ_i^{min} a
user specified minimum growth rate, v_i^{ex} the exchange fluxes with the external environment, and
lb and ub the respective lower and upper bounds. Additionally, μ_c denotes the community growth
rate and v_i^{m} the exchanges between the entire community and the gut lumen. The described
constraints are identical to the ones employed in SteadyCom (22, 29). We assigned an upper
bound of 100 mmol/[gDW h] for the internal exchange fluxes v_i^{ex}. Assuming a total microbiota
biomass of 200 g and a representative bacterial cell dry weight of 2 pg (59), this corresponds to
a maximum import or export of more than 100,000 molecules/[cell s]. Diet derived lower bounds
with values smaller 10^{-6} mmol/[gDW h] were set to zero as they would have been lower than the
numerical tolerance of the solver. Taxa with relative abundances a_i smaller 10^{-3} for the genus
models or 10^{-4} for the species models were discarded since they would not be able to affect the
external metabolite levels in a significant way but do increase computation time. Internal fluxes v_i
received respective bounds of 1000.0 (or 0 if irreversible) making them essentially unbounded.
The described constraints are applied to all optimization problems in MICOM and will be further
called the "community constraints".
616
617 The cooperative trade-off method consists of two sequential problems. First, maximize the
618 community growth rate μ_c to obtain μ_c^{max}. Using a user specified trade-off α now solve the following
619 quadratic minimization problem:

620
$$minimize\ \sum_i \mu_i^{\ 2}$$

621
$$s.t.\ \mu_c \geq \alpha \cdot \mu_c^{\ max}$$

622   *and community constraints*

624   The knockout for a genus i was performed by setting all fluxes belonging to this genus along with

625   its exchanges with the external environment to zero (lb=0 and ub=0). This is followed by running

626   cooperative trade-off on the knockout model and comparing the growth rates after the knockout

627   with the ones without the knockout.

Solvers and Numerical stabilization

Most genome-scale metabolic models usually do not treat more than 10,000 variables in the

corresponding linear or quadratic programming problems. However, in microbial community

models we usually treat 10s to 100s of distinct genome-scale models, which makes the

corresponding problem much larger. Unfortunately, many open and commercial solvers have

difficulties solving problems of that scale, so we also implemented strategies to increase the

success rate of those optimizations. All linear and quadratic programming problems were solved

using interior point methods as those were much faster than simplex methods for problems with

more than 100,000 variables. Here, we used CPLEX (https://www.ibm.com/analytics/cplex-

optimizer) but also tested all methods with Gurobi (https://www.gurobi.com). Since growth rates

tend to be small we also multiplied the objectives used in the cooperative tradeoff (maximization

of community growth rate and minimization of regularization term) with a scaling factor in order to

avoid near-zero objective coefficients. A scaling factor in the order of the largest constraint

(1000.0) seemed to work well. Nevertheless, the default interior point methods for quadratic

problems in CPLEX or Gurobi were usually not capable of solving the minimization of the

regularization term to optimality and usually failed due to numerical instability. The solutions

reported by the aborted optimization run were usually close to the optimum, but tended to violate

some numerically ill-conditioned constraints. To alleviate this problem, we implemented a

crossover strategy where we took the solution of the numerically ill-conditioned quadratic interior

point method as a candidate solution set μ_i^{ca}. Based on that we then optimized the following linear

programming problem in order to restore feasibility:

650
$$maximize\ \mu_c = \sum_i a_i \cdot \mu_i$$

651
$$s.t.\ \mu_i \leq \mu_i{}^{ca}$$

652   *and community constraints*

653 Linear interior point methods are usually numerically stable so this linear programming problem
654 can usually be solved to optimality. The maximization together with the new constraints will push
655 the individual growth rates towards the candidate solution as long as it is numerically feasible.

Minimal media and exchange fluxes

657 By convention MICOM formulates all exchange fluxes in the export direction so that all import
658 fluxes are positive and export fluxes are negative. Based on this, the minimal medium for a
659 community was obtained by minimizing the total import flux:

$$minimize\ v_{tot} = \sum_i \{|v_i{}^m|, v_i{}^m < 0\}$$

$$s.t.\ \forall i:\ \mu_i \geq \mu_i{}^{ct}$$

$$\mu_c \geq \alpha \cdot \mu_c{}^{max}$$

$$and\ community\ constraints$$

664 Here $\mu_i{}^{ct}$ denotes the optimal genera growth rates obtained by cooperative trade-off. The
665 community exchanges were then obtained by extracting all $v_i{}^m$, whereas genus-specific
666 exchanges were given by all $v_i{}^{ex}$ as defined earlier.
667
668 Overall production fluxes were calculated as

$$v_{tot}{}^m = \sum_{i,v_i{}^m > 0} a_i \cdot v_i{}^m,$$

670 where $v_i{}^m$ denotes an exchange flux for the metabolite m in taxon i. Overall consumption rates
671 were calculated in a similar manner but restricting fluxes to ones with $v_i{}^m < 0$ (imports).
672

Single target intervention studies

674 We used elasticity coefficients (50, 51) to evaluate the sensitivity of exchange fluxes to changes
675 in exchange flux bounds (*ergo* diet changes) or changes in genus abundances. The logarithmic
676 formulation of elasticity coefficients is given by
677

$$\varepsilon_p{}^v = \frac{\partial\ ln\ |v|}{\partial\ ln\ |p|}$$

679
where v denotes the exchange flux of interest and p the changed parameter. Since the absolute
value removes information about the directionality of the flux this was logged separately to

maintain this information. We used a value of 0.1 as differentiation step size in log space, which
corresponds to a bound or abundance increase of about 10.5% in the native scale. To enable
efficient computation, elasticity coefficients were grouped by the p parameter, then the
cooperative trade-off was run once without modification, the p parameter was increased, the
cooperative trade-off was run again, and differentiation was performed for all exchange fluxes at
once.

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Author contributions

696 CD developed/implemented the methods and performed the analysis. SMG helped design the
697 metagenomic analyses and growth rate validations. ORA developed the methods and designed
698 the meta-analysis. All authors wrote the manuscript.

Supplementary Material

701 **Supplementary Text S1.** Closed form solutions for regularization.

703 Additional materials and data are available at https://github.com/micom-dev/micom_study.

 Supplemental Figures

705
706 Figure S1. (A) Pearson correlations between inferred growth rates and replication rates across all
707 samples (n=1,062). Stars denote significance of a Pearson test (ns - p>0.05, * - p<0.05, ** - p<0.01, *** -
708 p<0.001). The dashed line denotes a zero correlation coefficient. (B) Fraction of observed genera growing
709 in each sample. Each dot denotes a single sample. A trade-off of "none" means optimization without L2
710 regularization and only maximizing the community growth rate.

712
Figure S2. Metrics on the species level. (A) Effect of different trade-off values (fraction of maximum
community growth rate) on the distribution of individual species growth rates. Zero growth rates were
assigned a value of 10^{-16} which was smaller than the observed non-zero minimum. Growth rates smaller
than 10^{-6} were considered to not represent growth. (B) Pearson correlation between replication rates and
inferred species growth rates under varying trade-off values. "None" indicates a model without
regularization returning arbitrary alternative solutions (see Methods). The dashed line indicates a
correlation coefficient of zero. (C) Pearson correlations between inferred species growth rates and

replication rates across all samples (n=1,113). Stars denote significance of a Pearson test (ns - $p > 0.05$, *

- $p < 0.05$, ** - $p < 0.01$, *** - $p < 0.001$). The dashed lines denotes a zero correlation coefficient.

Figure S3. Strong interactions are heterogeneously distributed across samples. (A) Histogram of the

number of samples a particular interaction occurs in. Lower numbers on the x-axis denote sample-specific

interactions and higher numbers denote ubiquitous interactions. (B) Interactions observed across all

samples stratified by genus and type.

Figure S4. Total consumption and net flux of SCFAs by the microbiota. (A) Total consumption flux of

SCFAs. (B) Net production of SCFAs (i.e. difference of total production and total consumption).

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