The effect of stride length on lower extremity joint kinetics at various gait speeds

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Abstract

Robot-assisted training is a promising tool under development for rehabilitation of walking function following neurological injury. The challenges in developing the controllers for gait rehabilitation devices that promote desired changes in gait is complicated by the limited understanding of the human response to robotic input. A possible method of controller formulation can be based on the principle of bio-inspiration, where a robot is controlled to apply the change in joint moment applied by human subjects when they achieve a gait feature of interest. However, it is currently unclear how lower extremity joint moments are modulated by even basic gait spatio-temporal parameters.

In this study, we investigated how sagittal plane joint moments are affected by a factorial modulation of two important gait parameters: gait speed and stride length.

We present the findings obtained from 20 healthy control subjects walking at various 13 treadmill-imposed speeds and instructed to modulate stride length utilizing real-time 14 visual feedback. Implementing a continuum analysis of inverse-dynamics derived joint 15 moment profiles, we extracted the global effects of gait speed and stride length on joint 16 moment throughout the gait cycle. Moreover, we utilized a torque pulse approximation 17 analysis to determine the timing and amplitude of torque pulses that approximate the 18 difference in joint moment profiles between stride length conditions, at all gait speed 19 conditions. 20

Our results show that gait speed has a significant effect on the moment profiles in all 21 joints considered, while stride length has more localized effects, with the main effect 22 observed on the knee moment during stance, and smaller effects observed for the hip 23 joint moment during swing and ankle moment during the loading response. Moreover, 24 our study demonstrated that trailing limb angle, a parameter of interest in programs 25 targeting propulsion at push-off, was significantly correlated with stride length. As such, 26 our study has generated candidate assistance strategies based on pulses of torque 27 suitable for implementation via a wearable exoskeleton with the objective of modulating 28 stride length, and other correlated variables such as trailing limb angle.

Introduction

Stroke is the leading cause of disability in the industrialized world [1]. Following a stroke, survivors often experience hemiparesis; characterized as diminished strength and inability to normally contract muscles on the side of the body contralateral to the lesioned hemisphere [2]. This leads approximately 30% of chronic stroke survivors to experience impaired gait and require walking assistance [3]. As such, restoration of ambulation capacity is a top priority for individuals recovering from a stroke [4].

Standard of care ambulation therapy approaches include treadmill and overground training with and without body weight support. These are highly repetitive movement exercises requiring observation and cues from a physical therapist [5,6]. In recent years, more technologically advanced gait retraining tools have been devised. These include electromechanically-assisting gait retrainers which offer the benefit of reduced physical demand on the physical therapists [5]. Also, functional electrical stimulation (FES)

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protocols have been utilized to supplement survivors' diminished capacity to contract muscles during walking [7]. The primary goal of these treatments and indicator of gait performance improvement is gait speed (GS), which is associated with a better quality of life [8,9]. A secondary goal and indicator of function is ambulation endurance, as often measured by the 6-minute walking test, which is significantly associated with scores of community integration [10].

Currently, it is not well understood how the modulation of assistance provided by a 49 robot or FES unit during gait retraining will lead to improvements in clinical outcome measures. Anterior-posterior ground reaction force, the propulsive force of the foot 51 against the ground during gait, and propulsive impulse, the propulsive force integrated 52 over time, are measures of interest in stroke [11–13]. Previous work has shown that 53 propulsive impulse is positively correlated with GS and negatively correlated with the severity of hemiparesis [11]. Further investigation revealed that posture of the trailing limb at push-off is positively correlated with propulsive impulse in paretic, non-paretic, 56 and healthy control limbs [14]. The posture of the trailing limb at push-off is quantified by one kinematic parameter, known as trailing limb angle (TLA), defined as the angle of the line connecting the hip joint center and foot center of pressure at the instant of peak propulsive force, relative to the global vertical axis [13]. One study examined the 60 relationship between long distance walking function, TLA, and propulsive force in a 61 cohort of chronic stroke patients following twelve weeks of gait retraining. Subjects in 62 this study who experienced improvements in TLA and in propulsive force experienced 63 gains in long distance walking function as measured by the 6-minute walking test [12]. 64 In healthy control subjects, it was observed that when increasing GS, the increase in 65 TLA contributes twice as much as the increase in ankle moment to the resulting increase in propulsive force [15]. Recently, it was observed that in GS modulation by 67 stroke patients, TLA and ankle moment contribute in a 4:1 ratio to propulsive force in the paretic leg and in a 3:1 ratio in the non-paretic leg [13]. Therefore, TLA has been 69 advanced as a variable of interest for improving walking function in stroke patients. 70 However, it is unclear if a gait retraining therapy that targets modulations of TLA will 71 lead to increases in clinical outcome measures. 72

Our research group is exploring the use of robotic-assisted gait retraining to directly ⁷³ target and modulate TLA. A possible controller could be composed of torque pulses ⁷⁴ applied at specific instants during the gait cycle, with the advantage of not constraining 75 gait to follow prescribed trajectories [16]. This approach has been shown in previous studies to be a successful method of robot-assisted gait training [17, 18]. However, in absence of models of the human response to a robotic input, it would be difficult to define parameters for such a controller acting on multiple degrees of freedom. A possible 79 method of controller formulation can derive from the principle of bio-inspiration, where a robot is controlled to apply the difference in joint moment applied by human subjects 81 when they achieve a desired gait feature (in this case modulation of TLA), relative to 82 their normal walking condition. Once the effects of the variable of interest have been 83 identified, a rehabilitation robot could be controlled in either assistive, resistive, or perturbation mode to deliver different forms of robot-assisted training [19]. To support the development of such a controller, we first required knowledge about the joint moments applied by healthy control subjects to modulate TLA at a range of GSs. Since TLA has not been a primary measure of interest in the biomechanics literature, we extended our search to a more common variable, likely correlated to TLA, such as stride length (SL). ٥n

The joint kinetics associated with GS modulation in healthy control subjects have been thoroughly elucidated in the literature, where an increase in GS is generally 92 associated with increase in magnitude of peak joint moments. A very early investigation 93 examining knee kinetics found increasing GS to be strongly correlated with an increase in peak knee extension moment [20]. Further work found peak hip flexion and extension moments, knee flexion and extensions moments, and ankle plantar and dorsiflexion moments to all increase with GS. However, these changes in joint kinetics to increasing 97 speed were primarily observed at the hip, particularly in extension, and secondarily at the ankle for purposes of support [21]. Most recently, an increase in GS was observed to 99 be associated with an increase in peak hip extension moment during loading response, 100 knee flexion moment in late stance and peak ankle plantarflexion moment [22]. 101

However, fewer investigations have focused on the joint kinetics associated with the modulation of spatiotemporal parameters such as SL or TLA. Summed joint work has been observed to be strongly correlated with SL in both young and old adults, where young adults primarily utilized swing phase hip work to modulate SL and old adults utilized ankle and knee joint work [23]. An early investigation found stance phase peak knee extension moment to be strongly correlated with increasing SL [20]. A more recent and in-depth investigation found that as SL increased, peak ankle plantarflexion moment, plantarflexion moment at 40% of stance, and peak knee extension moment all increased, while peak knee flexion moment and peak hip flexion moment decreased [24].

Thus far, the factorial modulation of both GS and SL and resulting hip, knee, and 111 ankle kinetics has not been investigated; as such it is unclear how lower extremity joint 112 moments are modulated by both gait parameters. Addressing this gap of knowledge, we 113 designed an experimental study to establish the effects of GS and SL on the resulting 114 hip, knee, and ankle joint moments. The findings are intended to inform the design of a 115 robotic assistance controller that delivers pulses of torque to the lower extremity joints 116 with optimal timing and amplitude to induce desirable modulations of gait 117 spatiotemporal parameters. 118

Materials and methods

Subjects

20 healthy adults (10 males, 10 females) were recruited to participate in this study (protocol approved by the University of Delaware Institutional Review Board, protocol no. 619724). Subjects — age (mean \pm std) 21.55 \pm 2.50 yrs, height 1.73 \pm 0.08 m, body mass 69.20 \pm 8.73 kg — were naive to the purpose of the study, and free of orthopedic neurological disorders affecting walking function. Subjects were required to wear their own comfortable athletic shoes and lightweight clothing for the walking experiment.

Setup

Subjects walked on an instrumented dual-belt treadmill (Bertec Corp., Columbus OH, USA), as shown in Fig. 1, while wearing thirty-six reflective spherical markers (4 on the pelvis, 4 per thigh, 4 per shank, 2 per knee, and 6 on ankle/foot). An eight camera for subjects 1-14, and ten camera Vicon T40-S passive motion capture system (Oxford Metrics, Oxford, UK), for subjects 15-20, were used to measure marker position in space. Marker data were acquired at 100 Hz, while the treadmill analog force/torque 129

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data were acquired at 2 kHz. A 24-in screen was placed at approximately 1500 mm 135 anteriorly from the center of the treadmill, and was used in biofeedback conditions. The 136 screen provided visual feedback about the SL measured at the previous gait cycle 137 (starting and ending at right heel strike) which was updated within 20 ms after each 138 right heel strike. In this experiment, SL for cycle k was defined based on the right heel 139 strike time t and anteroposterior coordinate in the laboratory frame x and constant 140 velocity of the treadmill v: 141

$$SL_{(k)} = x_{(k+1)} - x_{(k)} + v \left(t_{(k+1)} - t_{(k)} \right)$$
(1)

Visual feedback of $SL_{(k)}$ was provided in terms of the height of a bar, while the desired SL was displayed as a horizontal line with dashed lines indicating the $\pm 10\%$ range. The bar indicating $SL_{(k)}$ was color coded to indicate whether the measured value was within $\pm 10\%$ of the desired value. During biofeedback sessions, subjects were instructed to modify the length of their strides to achieve the target range, while walking at treadmill-imposed speeds.

Procedures

Subjects were exposed to a total of fifteen experimental conditions, determined as the 149 combinations of two factors: i) GS, with five levels, and ii) SL, with three levels. Factor 150 levels were defined in terms of percent change relative to subject self-selected (ss) values 151 to accommodate inter-subject variability in gait parameters. Moreover, to account for 152 the correlation between GS and SL [25], we first measured self-selected stride length 153 (ss-SL) at all speeds, and defined biofeedback-modulated SL conditions as percent 154 changes of SL relative to the ss-SL at any given speed. This experimental setup allowed 155 us to investigate joint kinetics underscoring an increase or decrease of SL relative to the 156 subject's self-selected stride length, at all speeds. 157

Self-selected gait speed

A preliminary set of trials were conducted to calculate the subject's self-selected gait ¹⁵⁹ speed (ss-GS). Subjects were asked to walk on the treadmill moving at an initial speed ¹⁶⁰ of 0.5 m/s, with the treadmill speed gradually increased by intervals of 0.03 m/s, and to ¹⁶¹

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indicate when ss-GS was reached. The same procedure was repeated by starting with the treadmill at 1.8 m/s, and decreasing treadmill speed in increments of 0.03 m/s, until the subject indicated that ss-GS had been reached. This procedure was repeated three times and the ss-GS was calculated as the average between the six measured treadmill speed values.

Fig 1. The experimental setup - subject walking on instrumented treadmill while wearing retroflective markers captured by infrared camera system. Real-time visual feedback provided on screen cues imposed stride length condition.

Non-biofeedback conditions

After determination of ss-GS, five walking trials were conducted consisting of ninety seconds of acquired data in the absence of biofeedback. In each trial, treadmill speed was imposed at one of five percentages of the subject's ss-GS [80%, 90%, 100%, 110%, 120%] in a randomized order. For each GS, ss-SL was calculated as the mean SL measured at that treadmill speed and utilized for the definition of subsequent desired SL values at each GS.

Biofeedback conditions

After determination of ss-SL for all five GS conditions, ten additional walking trials 175 were conducted consisting of ninety seconds of data acquisition, two for each treadmill 176 speed value, using biofeedback to cue a desired SL. For each GS, the desired SL was set 177 to be either 17% greater or 17% smaller than the ss-SL at that GS, in a random order. 178 The range of change in SL values was specified based on previous studies showing 179 feasibility of achieving distinguishable gait kinetics when SL was modulated by 17% of 180 the ss value [25]. The investigator initiated data acquisition for each condition when the 181 subject sufficiently achieved the cued SL condition specified via biofeedback. 182

Data analysis

Pre-processing

Raw marker trajectories were labeled offline. Marker position and force/torque data 185 were fed into a standard Visual3D pre-processing pipeline, which included i) noise gating 186

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of measured force with a 25 N threshold *ii*) low-pass filtering of marker and force/torque data (Butterworth filter at 6 Hz and 30 Hz cut-off frequency, respectively), *iii*) 188 interpolation of missing marker data with a third order polynomial fit for a maximum 189 gap size of five samples, *iv*) application of the subject-specific model for calculation of 190 joint angles and moments based on inverse kinematics and inverse dynamics algorithms. 191

In a custom MATLAB (MathWorks, Natick, MA, USA) script, hip, knee, and ankle 192 joint angles and moments for the right leg in the sagittal plane were extracted and 193 filtered with a 2^{nd} order low-pass zero-shift Butterworth filter with a cut-off frequency 194 of 15 Hz. Gait cycles were segmented between subsequent heel strike events, defined as 195 the instants at which the vertical ground reaction force changed in value from zero to 196 positive, and remained positive for a minimum of 400ms. Due to events such as marker 197 occlusion or subjects' foot stepping on contralateral force plates, acquired data were 198 manually screened and some gait cycles were excluded from the analysis. A minimum of 199 25 segmented gait cycles were linearly resampled in the [0.100] % gait cycle domain \tilde{t} 200 and averaged at each point in gait cycle to yield an average hip, knee, and ankle joint 201 moment profile $\tau(t)$ for each of the 15 experimental conditions for each of the 20 202 subjects. 203

Prior to obtaining average group moment profiles, joint moment profiles were non-dimensionalized. In agreement with [26], the non-dimensional joint moment $\tilde{\tau}$ was calculated for each joint as $\tilde{\tau}(\tilde{t}) = \frac{\tau(\tilde{t})}{WL_l}$, where W is body weight in N, and L_l is leg length in m, measured as the distance between the hip joint center and the floor during straight-leg standing.

Protocol validation

A non-dimensional GS was defined as the Froude number $Fr = GS/\sqrt{L_lg}$, where g is 210 the acceleration of gravity. Although several other factors such as body mass and 211 athletic fitness condition account for the variability in ss-GS across individuals [27], the 212 Froude number has been extensively used to describe the conditions underlying the 213 transition from walking to running in several species [28]. As such, we used the Froude 214 number as an index of across-subject dynamic similarity in ss-GS: a smaller variance of 215 Froude numbers within a group of individuals should reflect consistent gait kinetics. We 216 calculated the coefficient of variation $CV_{Fr} = \frac{\sigma_{Fr}}{\mu_{Fr}}$ as the ratio between the standard 217

deviation and the mean of Froude numbers corresponding to the ss-GS condition, and compared it to alternative indices, $CV_{ss-GS} = \frac{\sigma_{ss-GS}}{\mu_{ss-GS}}$ that uses ss-GS, and $CV_{ss-GS_0} = \frac{\sigma_{ss-GS_0}}{\mu_{ss-GS_0}}$ that uses ss-GS normalized by leg length. 220

Two gait parameters were also calculated; SL was measured using eq. (1), while TLA was calculated as the angle relative to the vertical axis of the line connecting the hip joint center and the position of the center of pressure at the instant of maximum anterior ground reaction force [29]. We conducted a linear correlation between normalized SL (SL₀ = SL/ L_l) and TLA to validate our protocol as suitable to inform the design of TLA-oriented robot-aided training protocols

Continuum analysis

We sought to determine if the two factors GS and SL had any significant effect on the 228 sagittal plane moment profiles for the hip, knee, and ankle joint $\bar{\tau}(\tilde{t})$, and, if so, at which 229 phase of a gait cycle was a significant effect of either factor measured. We conducted an 230 analysis for the main effects of the two factors, GS and SL, by implementing a 231 repeated-measure 2-way ANOVA on the mean joint moment profiles measured from 232 each subject and experimental conditions, spanning exhaustively the 15 combinations of 233 factors. ANOVA was conducted to test the null hypothesis that neither factor (GS and 234 SL), nor their interaction, induce a significant effect on joint moment at any time point. 235

Since the dependent variable $\bar{\tau}(\tilde{t})$ is one-dimensional (1D) smoothed time series 236 including highly temporally correlated data, and not a zero-dimensional scalar quantity 237 (e.g. peak torque, range of motion, etc.), definition of confidence intervals and control of 238 false positive rates (FPR) requires proper correction for multiple comparisons that 239 accounts for the temporal correlations in the input time series [30]. We used the 240 software SPM1D, a parametric statistical testing method developed for nD time 241 series [31], to control for FPR in the analysis of normalized joint moment profiles $\tilde{\tau}(t)$, 242 and quantify the effect of both factors (GS and SL) and of their interaction on the 243 dependent variable in different phases of the gait cycle. 2-way repeated-measure 244 ANOVA was conducted using the SPM1D software package, using SPM1D's function 245 anova2rm [31]. Inference was conducted setting a corrected type-I error rate $\alpha = 0.05$ 246 based on SPM1D's correction based on random field theory (RFT) to estimate the 247 smoothness in the input data. 248

After the main effect analysis, we conducted pairwise comparisons to establish the 249 specific effect of SL on the measured joint moment profiles, testing for the null 250 hypothesis that the mean profiles measured at the same speed for nominal and 251 bio-feedback modulated SL conditions were not different from one another at any time 252 point. This resulted in two comparisons (ss vs. increased SL, and ss vs. decreased SL) 253 per speed, per joint, for a total of $n_{comp} = 30$ pairwise comparisons. Pairwise 254 comparisons were conducted using two-tailed paired t-tests using SPM1D function 255 ttest_paired, and defining thresholds t-scores for significance at $\alpha = 0.05$ using a 256 Bonferroni correction $(n = n_{comp})$ on the paired difference thresholds calculated by 257 SPM1D. 258

Torque pulse approximation

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We then conducted a second analysis with the specific purpose of deriving candidate 260 primitives for a controller based on the application of pulses of torque as select phases of 261 the gait cycle. As such, we sought to approximate the effect of SL modulation on joint 262 moment profiles with a series of rectangular pulses of torque. First, the normalized 263 difference $\Delta \bar{\tau}(\tilde{t})$ of subject specific average joint moment profiles $\tau(\tilde{t})$ between 264 conditions of positive or negative SL, and ss SL, at each gait speed j were extracted as: 265

$$\Delta \bar{\tau}_{\pm}^{(j)}(\tilde{t}) = \frac{\tau_{\pm \text{SL}}^{(j)}(\tilde{t}) - \tau_{\text{ss-SL}}^{(j)}(\tilde{t})}{\operatorname{Max}\left(\tau_{\text{ss-SL}}^{(j)}(\tilde{t})\right) - \operatorname{Min}\left(\tau_{\text{ss-SL}}^{(j)}(\tilde{t})\right)}$$
(2)

The rectangular torque pulses $P_{\pm}(\tilde{t})$ were defined using a constant duration of 10% 266 of gait cycle, variable time of application α_l , and amplitude A_l as in the equation: 267

$$P_{\pm}(\tilde{t}) = \sum_{l=1}^{N} A_{l,\pm} \operatorname{rect}\left(\frac{\tilde{t} - \alpha_{l,\pm}}{10}\right)$$
(3)

For this specific work, we examined the one and two pulse (N = 1, 2) approximation 268 of the function $\Delta \bar{\tau}_{\pm}^{(j)}(\tilde{t})$, and used nonlinear constrained optimization—MATLAB 269 function's fmincon—to find the values of parameters A and α that minimize the norm 270 of residuals $P_{\pm}(\tilde{t}) - \Delta \bar{\tau}_{\pm}^{(j)}(\tilde{t})$. In this optimization, the domain for α_l was defined as the 271 set of integers between 5% and 95%, providing a quantization in the time of application 272 of torque pulses equal to 1% of a gait cycle duration. For each joint, we divided the 273 estimated pulses into two different groups depending on whether their amplitude was in the positive, or negative direction (i.e., positive extension for the hip and knee joint, positive plantarflexion for the ankle joint). 276

With the purpose of identifying location and amplitude of application of pulses of 277 torque that would approximate the SL-specific difference between joint moment profiles 278 measured at all speeds, we performed statistical analyses to determine if any of the 279 outcome measures, pulse magnitude and location, were significantly modulated by any 280 of the three independent variables (i.e., joint, direction of SL modulation, and sign of 281 applied pulse). For the purposes of our analysis, pulse magnitude is the absolute value 282 of pulse amplitude. Four separate linear mixed effects models (SAS V9.4, SAS Institute, 283 Cary, NC) were performed on the one and two pulse approximations for both the pulse 284 magnitude and location data sets to test the null hypothesis that no independent 285 variable had an effect on the outcome measures. The models included fixed effects for 286 each of the independent variables as well as all two-way and one three-way interaction 287 between them. Heterogeneity due to trials completed under different gait speed 288 conditions, and multiple pulses in the case of the two pulse approximation were 289 accounted for by the inclusion of random effects. Correlation between multiple 290 measurements taken on the same subject were accounted for by the inclusion of a 291 repeated measure effect. Upon comparing nested model Akaike information criterion 292 (AIC) values; the lowest AIC value came from the unstructured covariance structure 293 and was therefore selected for the final models. In case of effects statistically significant 294 at the $\alpha < 0.05$ level, effects and interactions were further investigated through post hoc 295 Tukey-Kramer tests ($\alpha = 0.05$). 296

Results

Protocol validation

The mean Froude numbers with 95% confidence intervals, across gait speed conditions, are shown in the left of Fig. 2. The use of the Froude did not reduce the across-subject variability in ss-GS, with $CV_{Fr} = 0.106$, slightly greater than $CV_{ss-GS} = 0.104$, and both smaller than $CV_{ss-GS_0} = 0.121$. All differences account for an effect size that can

be considered very small. To determine if the imposed biofeedback effectively 303 modulated SL in healthy subjects, we assessed the distribution of SL values across all 304 three feedback conditions at the five different gait speeds, shown in the center of Fig. 2. 305 The change in mean SL from ss-SL across all ten feedback conditions for all subjects 306 was equal to $\pm 14.94\%$, close to the target $\pm 17\%$ value. The maximum standard 307 deviation of SL₀ values for all five non-feedback conditions, averaged across all subjects, 308 was relatively small ($\sigma_{\rm max} = 3.86\%$). Based on these measures, we conclude that the 309 protocol significantly modulated values of SL and GS, such that statistical analysis may 310 be performed. 311

Fig 2. (Left) the distribution of mean Froude numbers with 95% confidence intervals at the various treadmill imposed GSs, (center) normalized stride lengths measured at various speeds and biofeedback conditions. The box plot shows the median as a horizontal line, and the box at 25% and 75% percentiles, with whiskers extending to $\pm 3\sigma$, and (right) mean trailing limb angle and mean normalized stride length for each of the 15 conditions for each of the 20 subjects. Linear regression indicates that there is a strong correlation (r = 0.87) between the two measures.

We calculated the mean TLA and SL_0 value for each of the 15 conditions for all subjects to assess the correlation between these measures; see right side of Fig. 2. Linear regression demonstrated a strong correlation (r = 0.87) between SL_0 and TLA, which indicates that subjects also modulated TLA while achieving biofeedback cued SL modulation. The group analysis of joint moments in the three SL conditions across five GS conditions is depicted in Fig. 3

Fig 3. Effect of gait speed (GS) and stride length (SL) modulation on the normalized joint moments $\bar{\tau}$. Joints are organized by row, GS are organized by columns, relative to the subject-specific ss-GS. Conditions corresponding to cued SL values are superimposed on each plot. Lines indicate the group mean, with the shaded region indicating the standard error.

Continuum analysis

The continuum analysis showed an effect of GS on the normalized joint moment profiles, where a significant effect of GS was detected for a total 82.2%, 78.8% and 64.1% of the gait cycle for the hip, knee, and ankle joint respectively, as shown in Fig. 4. The effect of SL on hip joint moment was significant for four short clusters in early to midstance, at push-off, and two clusters spanning the majority of swing for a total duration of a

significant effect of SL on hip joint moment of 61.4% of the gait cycle. A stronger effect, 324 both in magnitude and duration, was measured at the knee joint with the first two 325 clusters spanning early stance. The effect of SL on knee joint moment was highly 326 significant from midstance until the end of gait cycle; for a total of 87.7% of the gait 327 cycle with a significant effect. A strong effect of SL was detected at the ankle joint for 6 328 clusters; at weight acceptance, push-off and four clusters covering approximately half of 329 swing for a total of 39.3% of the gait cycle. The interaction between the two factors was 330 significant for the hip for four clusters; mainly during the transition from stance to 331 swing and during late swing for a total of 27.1% of the gait cycle. For the knee, the 332 interaction was significant for 5 clusters; 2 short intervals during early stance, late 333 stance, early swing, and midswing for a total of 33.2% of gait cycle. For the ankle, the 334 interaction was only significant for two clusters, late stance and mid swing for a total of 335 18.1% of gait cycle. Pairwise comparisons of joint moment profiles measured at nominal 336 and biofeedback-modulated SL values are shown in Fig. 5-7 for all GSs. 337

Fig 4. Main effects of gait speed (GS) and stride length (SL), and of their interaction, on the normalized joint moment profiles during normal walking, as described by the 1D time series of F-scores extracted by the 2-way repeated measure ANOVA. The threshold F score for each experimental condition is reported by the red dashed line, and values above (shaded in green) correspond to a significant group effect of the factor, at the corresponding gait cycle instant, for a corrected type I error rate $\alpha < 0.05$.

Pairwise comparisons for the hip joint show that during increased SL conditions (Fig. 338 5, right), hip flexion moment during early swing and hip extension moment during late 339 swing decreased. The first comparison reached significance in three out of five GS 340 conditions, while the second effect was significant at all GS values. No effect of SL on 341 hip joint moment during stance were observed in more than one GS condition. A 342 similar pattern is observed when SL is decreased via biofeedback (Fig. 5, left). 343

Fig 5. T-scores resulting from pairwise comparisons of normalized hip torque moment at normal and modulated SL (columns), measured at the same GS, for each GS (row). Red dashed lines show the threshold t value that provides a corrected type I errfor rate $\alpha = 0.05$, extracted using a Bonferroni correction that accounts for all pairwise comparisons $n_{comp} = 30$.

For the knee, during increased SL conditions (Fig. 6, right), knee extension moment ³⁴⁴ increased in early stance, while knee flexion moment increased in late stance. During ³⁴⁵ the swing phase, knee extension moment decreased in early swing, and knee flexion ³⁴⁶ moment decreased in late swing. The effects reported were significant at the group level at all GSs. A similar pattern was observed for a decrease of SL, with smaller effects for the increased knee extension at early stance (a significant effect was measured only in four out of five GS conditions, Fig. 6, right).

Fig 6. T-scores resulting from pairwise comparisons of normalized knee torque moment at normal and modulated SL (columns), measured at the same GS, for each GS (row). Red dashed lines show the threshold t value that provides a corrected type I error rate $\alpha = 0.05$, extracted using a Bonferroni correction that accounts for all pairwise comparisons $n_{comp} = 30$.

For the ankle, during increased SL conditions (Fig. 7, right), ankle dorsiflexion moment increased at early stance, while no effect on plantarflexion moment was measured at push-off. A similar pattern was observed for a decrease of SL (Fig. 7, right), with a greater effect measured in terms of increased plantarflexion moment at early swing (significant at all GSs). In two out of five GS conditions, the decreased SL condition exhibited a reduced ankle plantarflexion moment during push-off significant at the group level.

Fig 7. T-scores resulting from pairwise comparisons of normalized ankle torque moment at normal and modulated SL (columns), measured at the same GS, for each GS (row). Red dashed lines show the threshold t value that provides a corrected type I error rate $\alpha = 0.05$, extracted using a Bonferroni correction that accounts for all pairwise comparisons $n_{comp} = 30$.

Torque pulse approximation

Figures 8 and 9 show the distribution of torque pulse magnitudes grouped by joint for	35
both positive and negative SL modulations for the one and two pulse approximations,	36
respectively.	36

Fig 8. Histogram of one pulse approximation normalized amplitudes, sorted by SL modulation and joint.

Fig 9. Histogram of two pulse approximation normalized amplitudes, sorted by SL modulation and joint.

The results of the linear mixed effects model analyses for torque pulse magnitude are shown in Tables 1 and 2. A significant effect of factors joint and pulse sign, and a ³⁶³ significant interaction between factors joint and SL modulation were observed for the 364 one and two pulse approximations. The interaction between factors joint, SL 365 modulation, and pulse sign was significant for the two pulse approximation at the 366 selected significance level. The Tukey-Kramer post hoc test for the joint and SL 367 modulation interaction, shown in Tables 3 and 4, was used to establish the presence of 368 significant differences in pulse magnitude means between joints for SL modulation 369 conditions, separately. The mean and standard deviation of the joint magnitudes are 370 shown in Figures 10 and 11. In the one and two pulse approximations, for both SL 371 modulation conditions, the normalized torque pulse magnitudes of the knee joint were 372 greater than both the hip and ankle joints. The only significant difference between hip 373 and ankle joint magnitudes existed for the positive SL modulation condition for the two 374 pulse approximation. 375

Table 1. Magnitude linear mixed effects model results for the one torque pulse approximation

Factor	DF	F Value	Prob >F
Joint	2	130.63	< 0.001
SL Mod	1	0.50	0.482
Pulse Sign	1	5.44	0.020
Joint x SL Mod	2	12.41	< 0.001
Joint x Pulse Sign	2	1.15	0.318
SL Mod x Pulse Sign	1	0.03	0.862
Joint x SL Mod x Pulse Sign	2	2.14	0.120

 Table 2. Magnitude linear mixed effects model results for the two torque pulse approximation

Factor	DF	F Value	Prob >F
Joint	2	154.81	< 0.001
SL Mod	1	0.21	0.645
Pulse Sign	1	24.68	< 0.001
Joint x SL Mod	2	10.75	< 0.001
Joint x Pulse Sign	2	2.91	0.055
SL Mod x Pulse Sign	1	2.05	0.152
Joint x SL Mod x Pulse Sign	2	3.99	0.019

Table 3. Magnitude Tukey-Kramer post hoc test results for the one torque pulse approximation

	Mean (Std Err Dif)		
SL Mod	Hip - Knee	Knee - Ankle	
+SL	-0.172(0.014)	$0.194 \ (0.014)$	
-SL	-0.094 (0.011)	$0.010\ (0.010)$	$0.104\ (0.012)$

	Mean (Std Err Dif)			
SL Mod	Hip - Knee	Hip - Knee Hip - Ankle		
+SL	-0.115 (0.009)	$0.019 \ (0.005)$	$0.135\ (0.009)$	
-SL	-0.069 (0.007)	$0.015 \ (0.006)$	$0.085\ (0.008)$	

 Table 4. Magnitude Tukey-Kramer post hoc test results for the two torque pulse approximation

Fig 10. Pulse magnitude by joint and SL modulation for the one pulse approximation (mean \pm standard deviation). Asterisks indicate pairwise comparisons significant at the p < 0.05 corrected level.

Fig 11. Pulse magnitude by joint and SL modulation for the two pulse approximation (mean \pm standard deviation). Asterisks indicate pairwise comparisons significant at the p < 0.05 corrected level.

Figures 12 and 13 show the distributions of torque pulse location in gait cycle. All pulses were combined across the twenty subjects and five gait speed conditions and grouped by joint and SL modulation condition for the one and two pulse approximations, seperately. For representation purposes, within each histogram, the pulses are divided into positive and negative groups according to the sign of pulse amplitude $\Delta \bar{\tau}^{(j)}_{\pm}(\tilde{t})$, and then further divided into two more categories (i.e. small and large) based on whether their magnitude was smaller or larger than the group median.

Fig 12. Histogram of one pulse approximation locations in gait cycle, sorted by pulse amplitude sign and magnitude.

Fig 13. Histogram of two pulse approximation locations in gait cycle, sorted by pulse amplitude sign and magnitude.

The linear mixed effects model results for torque pulse location in gait cycle are shown in Tables 5 and 6. The analyses yielded a highly significant effect of joint and interactions between joint and pulse sign, SL modulation and pulse sign, and joint, SL modulation, and pulse sign for both pulse approximations. The factor of pulse sign was only significant for the one pulse approximation and the factor of SL modulation and interaction of joint and SL modulation were only significant for the two pulse approximation.

The three-way interaction between joint, SL modulation, and pulse sign was broken down using the Tukey-Kramer post hoc test to find the significant differences between torque pulse location, for a given joint, under different combinations of factors SL

DF	F Value	Prob >F
2	49.32	< 0.001
1	2.14	0.146
1	31.48	< 0.001
2	0.53	0.589
2	13.61	< 0.001
1	166.41	< 0.001
2	78.42	< 0.001
	$ \begin{array}{c} 2 \\ 1 \\ 1 \\ 2 \\ 2 \\ 1 \\ 1 \end{array} $	$\begin{array}{c ccccccccccccccccccccccccccccccccccc$

Table 5. Location linear mixed effects model results for the one torque pulseapproximation

 Table 6. Location linear mixed effects model results for the two torque pulse approximation

Factor	DF	F Value	Prob >F
Joint	2	55.62	< 0.001
SL Mod	1	8.62	0.004
Pulse Sign	1	3.01	0.083
Joint x SL Mod	2	3.30	0.039
Joint x Pulse Sign	2	22.11	< 0.001
SL Mod x Pulse Sign	1	176.74	< 0.001
Joint x SL Mod x Pulse Sign	2	147.21	< 0.001

modulation and torque pulse sign. We were especially interested in testing whether 393 there were specific instants of time where the application of a positive torque pulse would modulate SL in a certain direction, and, simultaneously, where the application of 395 a negative torque pulse would modulate SL in the opposite direction. As such, for each 396 joint, we conducted two pairwise comparisons: one to compare the location variables 397 measured for positive pulse torque sign and positive SL modulation with the variables 398 measured for negative pulse torque sign and negative SL modulation, and a second one 399 to compare the location variables measured for positive pulse torque sign and negative 400 SL modulation with the variables measured for negative pulse torque sign and positive 401 SL modulation. All of these pairwise comparisons are reported in Tables 7 and 8. In the 402 one pulse approximation, both pairwise comparisons for the hip and ankle joint yielded 403 a relatively large difference in mean location, with three out of the four comparisons 404 statistically significant. On the contrary, the pairwise comparisons for the knee yielded 405 small (1 - 2% gait cycle duration) and statistically insignificant mean differences. This 406 indicates that for the knee joint, clustering of torque pulses by location was symmetrical 407 in reversed stride length conditions, with negative pulses in negative SL conditions 408 clustering around a similar value as positive pulses in positive SL conditions, and 409 positive pulses in negative SL conditions clustering around a similar value as negative 410 pulses in positive SL conditions, while the same effect was not measured for the hip and ankle joints. However, this pattern was not observed in the two pulse approximation; in which one out of the four hip and ankle joint comparisons and one of the two knee joint comparisons were statistically significant. For the one and two pulse approximations, all knee joint mean comparisons were below 10% gait cycle, the width of the torque pulses used for the approximation. 410

 Table 7. Location Tukey-Kramer post hoc testing for the one torque pulse approximation

	Mean[1] - Mean[2] (Std Err Dif)	
Ioint	[1] + SL + Sign	[1] - SL + Sign
Joint	[2] -SL - Sign	[2] + SL - Sign
Hip	-9 (7)	-19 (5)
Knee	-1 (2)	-2 (2)
Ankle	-11 (3)	-13 (3)

 Table 8. Location Tukey-Kramer post hoc testing for the two torque pulse approximation

	Mean[1] - Mean	[2] (Std Err Dif)
Loint	[1] + SL + Sign	[1] - SL + Sign
Joint	[2] -SL - Sign	[2] + SL - Sign
Hip	-3 (4)	-20 (4)
Knee	8 (2)	6(2)
Ankle	-1 (2)	-3 (2)

Discussion

We exposed subjects to a factorial modulation of gait speed (GS) and stride length (SL) 418 and utilized inverse dynamics to estimate the lower extremity joint moments in the 419 sagittal plane. With our protocol, we modulated SL of individuals significantly between 420 conditions, with a mean change in SL equal to $\pm 15\%$ of the self-selected value, close to 421 the target $\pm 17\%$ with a relatively low standard deviation of 3.86%. Furthermore, 422 inter-individual variability in self-selected gait speed (ss-GS) was reasonably small, with 423 a coefficient of variation $CV_{\rm ss-GS} = 0.103$. Based on these measures, it is apparent that 424 our protocol significantly modulated both SL and GS, such that statistical analysis may 425 be performed to assess changes in joint kinetics arising from exposure to these conditions. 426 Our analysis showed a strong correlation (r = 0.87) between SL and TLA, indicating 427 that TLA was indirectly modulated through the explicit cueing of SL modulation. Our 428

data analyses focused primarily on the effects on joint kinetics introduced by modulation of SL at various GSs and secondarily on the effects introduced by GS.

The most prominent effects of SL modulation on joint moment were observed for the 431 knee joint. Simple visual inspection of the normalized joint moment profiles, Fig. 3, 432 clearly indicates the effect of SL on joint moment, where stance phase peak extension 433 moment and peak flexion moment increase with increasing SL as well as increasing GS. 434 Increasing peak knee extension moment with SL was also observed previously [20,24] 435 while increasing peak flexion moment is in contrast with the previously observed 436 decreasing peak knee flexion moment [24]. This contrasting result could be attributed 437 to important differences with the experimental paradigm pursued in [24], where SL and 438 cadence, and not SL and gait speed, were cued. These observations are validated by the 439 continuum analysis, where a main effect of SL was observed for 87.7% of the gait cycle, 440 the highest percentage of all three joints. The significant effect of SL on peak flexion 441 and extension moments during stance is supported by the pairwise comparisons between 442 joint torque measured at different SL conditions (Fig. 6). Here, significant effects at 443 early stance support an increase in knee extension moment with increasing SL, and 444 significant effects in late stance support an increase in knee flexion moment with 445 increasing SL. These effects introduced by increases in SL are also captured by the 446 torque pulse approximations through visual inspection of the pulse approximation 447 histograms 12 and 13 and the findings of their associated linear mixed effects model 448 pulse location analyses. Our findings indicate that an increase in SL is associated with 449 positive pulses of torque – an increase in extension moment – in early stance, and 450 negative pulses – an increase in flexion moment – in late stance. The reverse pattern is 451 observed for SL decrease in which there are negative pulses of torque – decreasing 452 extension moment – in early stance, and positive pulses of torque – decreasing flexion 453 moment – in late stance. This pattern is supported by the Tukey-Kramer post hoc one 454 pulse approximation results for the joint, SL modulation, and pulse sign interaction 455 effect of the pulse location linear mixed effects model analyses. Furthermore, for the one 456 pulse approximation, there is no significant difference in location between pulses of 457 positive SL and negative sign and pulses of negative SL and positive sign. The lack of 458 significant difference in the location of these specific pulse groups supports the 459 observation of a systematic pulse pattern at the knee which reverses in sign with 460

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reversal of SL modulation direction. Another indication of the effect of SL modulation 461 on knee joint moment derives from the fact that the one and two pulse approximations 462 yielded a significant effect of the joint and SL modulation interaction on magnitude 463 (p < 0.001) with Tukey-Kramer post hoc tests indicating the knee having the greatest 464 normalized pulse magnitude (Tables 3 and 4 and Figures 10 and 11). Overall, our 465 findings indicate that the effects of SL on the knee joint moment can be described as 466 follows: knee extension moment at early stance and knee flexion moment at late stance 467 are increased for an increase in SL, while knee extension moment at early stance and 468 knee flexion moment at late stance are decreased for a decrease in SL. 469

Significant effects of SL modulation were also observed at the hip joint. As indicated 470 by the main effect of SL quantified by the continuum analysis, there are two major 471 intervals of significance: early swing and late swing. These effects in swing are visible 472 upon inspection of the group moment profiles where an increase in SL is associated with 473 a decrease in flexion moment during early swing and a decrease in extension moment 474 during late swing. The statistical significance of these observations is clearly indicated 475 by the pairwise comparisons of SL modulation at different speeds shown in Fig. 5. Here, 476 in all pairwise comparisons, an increase in SL is associated with a significant interval in 477 early swing and in late swing. Another indication on how SL modulates hip moment 478 during the swing phase is provided by the torque pulse approximation histograms, 479 which consistently depict a grouping of negative pulses in late swing for positive change 480 in SL and positive pulses in late swing for negative change in SL. However, the pattern 481 is less clear than the one seen at the knee joint because of the small magnitude of those 482 pulses occurring in the swing phase, which even though they are representative of a 483 statistically significant effect, they account for a small amplitude (see distribution of 484 larger pulses in Fig. 12 and 13). As such, the effect in SL obtained for a change in 485 magnitude of the applied torque is not symmetrical (Tables 7 and 8), as demonstrated 486 by the Tukey-Kramer post hoc tests for both pulse approximations. These groupings of 487 pulses in late swing are likely associated with the decrease in hip extension moment 488 during late swing associated with increasing SL. 100

Relatively small effects of SL modulation were observed at the ankle joint. Most 490 prominently, during loading response, dorsiflexion moment increases with an increase in 491 SL as also can be observed in the group moment profiles. This is supported by the main 492 effect of SL measured via the continuum analysis, where the first 10% of gait cycle 493 shows a significant effect of GS on joint moment. This is further supported by the observation of a significant increase in dorsiflexion moment (negative change) for 495 increases in SL, in nine of the ten pairwise comparisons at all speeds, shown in Fig. 7. Another observed effect through the pairwise comparisons conducted via the continuum 497 analysis is the increase in plantarflexion moment at early swing with increasing SL. This effect can be confirmed through visual inspection of the group moment profiles. In 499 contrast with previous work [24], we did not observe a consistent increase in peak 500 plantarflexion moment with increasing SL, with only the three highest speed conditions 501 showing a significant effect for the transition from -17% ss to ss-SL, and no significant 502 effects measured for an increase in stride length over the self-selected value. 503

Conclusion

Our study has measured the effects of stride length (SL) on the lower extremity joint moment profiles at different speeds, demonstrating several consistent effects in our 506 population. The main effects of increasing SL at the knee include an increase in knee 507 extension moment at early stance and an increase in flexion moment at late stance. At 508 the hip, the main effects of increasing SL are a decrease in flexion moment during early 509 swing and a decrease in extension moment during late swing. For an increase in SL, the 510 ankle primarily exhibits an increase in dorsiflexion moment during loading response. 511 Given the observed linear relationship between SL_0 and TLA, pulse torque 512 approximation patterns associated with SL modulation are also associated with TLA 513 modulation. These findings suggest that a possible joint moment assistance strategy 514 based on pulses of torque applied primarily at the hip and the knee joint could induce 515 modulations in both SL and TLA. According to our analysis, the application of positive 516 pulses of torque in early stance and a negative pulse in late stance to the knee appear to 517 be suitable candidate assistance strategies to support an increase of SL and TLA during 518 walking. If pulse torque assistance is to be applied at an additional joint, pulse torque 519 assistance could be applied at the hip with a negative pulse applied during late swing. 520

This study has some limitations. The methods pursued in this paper are based on a 521 group analysis of joint moment profiles measured via inverse-dynamics. As such, it is 522

possible that the most successful assistance strategies may significantly change between different individuals. Therefore, the group analysis based assistance strategy candidate could be best utilized as an initial estimate and assistance strategies could be iteratively optimized for each subject using human-in-the-loop optimization, like it has been done for single-joint assistance schemes [17].

Moreover, the proposed assistance strategy candidate is based on the assumption 528 that human contribution will not change when an assistive torque is applied via a 529 wearable exoskeleton, such that the combination of torques applied by the two agents 530 would result in a simple summation. However, it is well known that the human 531 neuromuscular system is non-linear [32] and it involves complex feedback loops [33]. As 532 such, the response to a torque perturbation at a specific instant in gait cycle will be 533 difficult to predict. Given the difficulty of formulating a model of the human response 534 to these assistance strategies, once again the results of this analysis work could be used 535 as an initial estimate to be iteratively optimized for each subject using 536 human-in-the-loop optimization. 537

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