Neural Correlates of Causal Confounding

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Abstract

As scientists, we are keenly aware that if putative causes perfectly co-vary, the independent influence of neither can be discerned – a "no confounding" constraint on inference, fundamental to philosophical and statistical perspectives on causation. Intriguingly, a substantial behavioral literature suggests that naïve human reasoners, adults and children, are tacitly sensitive to causal confounding. Here, a combination of fMRI and cognitive computational modeling was used to investigate neural substrates mediating such sensitivity. While being scanned, participants observed and judged the influences of various putative causes with confounded or non-confounded, deterministic or stochastic, influences. During judgments requiring generalization of causal knowledge from a feedback-based learning context to a transfer probe, activity in the dorsomedial prefrontal cortex (DMPFC) was better accounted for by a Bayesian causal model, sensitive to both confounding and stochasticity, than a purely error-driven algorithm, sensitive only to stochasticity. Implications for the detection and estimation of distinct forms of uncertainty, and for a neural mediation of domain general constraints on causal induction, are discussed.

Introduction

Consider two scenarios involving a target cause, C, an alternative cause, A, and the presence (+) or absence (-) of some effect. In the first scenario, the effect occurs when C and A are presented in combination, but not when A is presented alone [A-, $AC+$]. In the second scenario, the A- trials are removed, so that both C and A occur only in combination with each other [$AC+$]. How would your judgment about the influence of C differ across these two scenarios? Philosophers, scientists and statisticians alike recognize that perfect co-variation of C and A in the second scenario occludes the independent influences of each putative cause, rendering any judgment about their respective effects fraught with uncertainty. Importantly, naïve human reasoners also appear to, tacitly, apply this domain-general constraint on causal induction, as evidenced by behavioral data showing sensitivity to confounding in a wide range of causal and predictive judgments, by children as well as adults (Spellman, 1996; Kushnir and Gopnik, 2005; Meder et al., 2006; Schulz and Bonawitz, 2007; Schulz et al., 2007; Liljeholm, 2015). Very little is known, however, about the neural substrates mediating sensitivity to, and uncertainty associated with, causal confounding. The current study aims to identify such neural computations.

A second source of uncertainty in causal and predictive inference, that has been thoroughly investigated both behaviorally and neutrally, is the stochasticity, or variance, of the outcome variable, which is greatest when the probability distribution over possible outcome states is uniform. As with causal confounding, behavioral sensitivity to outcome stochasticity is well established (e.g., Holt & Laury, 2002). Moreover, several neuroimaging studies have implicated the dorsomedial prefrontal

cortex (DMPFC), anterior insula, thalamus and dorsolateral prefrontal cortex (DLPFC)

in the neural representation of outcome stochasticity (Volz et al., 2003; Huettel et al.,

2005; Grinband et al., 2006; Abler et al., 2009).   An important open question is

whether, and which of, these neural regions also implement signals associated with

causal confounding. In the current study, participants were scanned with functional

MRI while observing and judging the influences of various putative causes with

confounded or non-confounded, deterministic or stochastic, influences.

Note that, if a particular set of confounded causes always occur in the same

configuration, and the only goal is to predict the state of the outcome based on that

particular configuration, then confounding does not pose a problem. Indeed, most

individual, non-confounded, causes can presumably themselves be broken into a set

of always co-occurring, and thus confounded, elements.   The critical question, thus, is

whether it is necessary to tease apart the influences of individual elements, such as

when a confounded cause suddenly occurs on its own, or in a novel combination with

some other cause.   In the current study, participants were required both to make

predictions about the outcome given a particular, recurring, configuration of

confounded causes and, in other instances, to make explicit judgments about the

individual influences of those same causes.   Only in the latter case would uncertainty

due to confounding be warranted.

Formally, two dominant approaches to predictive uncertainty can be discerned in the

psychological literature.   First, in associative learning theory, an error-driven

representation of uncertainty about a cue's predictive strength is captured by the

"Pearce-Hall" (PH) algorithm, which relates the "associability" of a cue to a weighted

average of the absolute prediction error on previous trials involving that cue. Since

the frequency and size of prediction errors increases with unpredictability, this quantity

is proportional to the stochasticity of the outcome. Conversely, in Bayesian Causal

Models, uncertainty about the predictive strength of a cause is reflected in the entropy

of the posterior distribution over its possible strengths, which depends on the variance

of the effect variable, but also on the independent occurrence of alternative causes.

While both the associative and Bayesian causal model predicts sensitivity to

stochasticity, only the causal model accounts for uncertainty due to causal

confounding. Here, a combination of neuroimaging and cognitive modeling was used

to dissociate neural signals scaling with error-driven and causal uncertainty. In

particular, judgments requiring generalization of causal knowledge from a feedback-

based learning context to a transfer probe were expected to elicit strong neural

responses to both stochasticity and confounding.

Materials & Methods

Participants. 20 healthy normal volunteers (mean age = 20.9 ± 2.4, range: 18-27, 12

females) participated in the study. A power analysis performed on data from a pilot

study on error-driven uncertainty, described in detail below, indicated that a sample

size of 16 would yield a power of 0.9 at a Gaussian random field theory corrected

threshold of 0.05 in regions of interest. One participant was excluded prior to any

analyses due to excessive head movement (> 6mm), leaving a sample size of 19. The

volunteers were pre-assessed to exclude those with a history of neurological or

psychiatric illness. All subjects gave informed consent and the study was approved by

the Institutional Review Board of the University of California, Irvine.

Task & procedure. Participants were scanned with functional MRI while performing a

causal induction task in which they assumed the role of a research scientist assessing

the influence of various allergy medicines on headache – a potential side effect. At

the beginning of the study, participants were instructed that each medicine could

either produce headache or have no influence on headache (i.e., there were no

preventive causes) and, further, that the influence of a given medicine might be

stochastic, so that even if that medicine was indeed a cause of headache, it may still

not produce headache every time it was administered. Three target medicines, C, D

and S, mnemonically labeled here to indicate "confounded", "deterministic" and

"stochastic" influences respectively, occurred only in combination with some

alternative medicine during feedback-based learning. Specifically, with +, - and ±

respectively indicating a 1.0, 0 and 0.5 probability of the effect, and with lower case

letters indicating non-target causes, there were 5 different types of medicine

treatments: e-, s'±, De+, Se± and Cc'+. None of the allergy patients had headache in

the absence of any medicine, which was explicitly stated in the initial instructions as

well as apparent on several "no medicine" (n-) trials.

Note that medicines C and c' both occurred *only* in combination with one another, so

that target medicine C was perfectly confounded with its compound counterpart.

Medicine D also occurred only in combination with another medicine, e, however,

"elemental" medicine e also occurred by itself, allowing for an estimation of the

independent causal influence of medicine D across the constant presence of e (i.e.,

across e- and De+ trials). Medicine S was identical to medicine D, except that the

probability of headache given the Se compound was 0.5 rather than 1.0; a

comparison of medicines S and D, therefore, identifies a difference between

deterministic and stochastic causation. Finally, medicine s' also produced headache

with a probability of 0.5 but unlike medicine S, never occurred in combination with

any other medicine, so that contrasting S with s' identifies differences due to

compound presentation. From a modeling perspective, an algorithm that relies

solely on outcome variance to track uncertainty would only respond to medicines S

and s'. In contrast, a computation that treats *both* confounding and stochasticity as

sources of uncertainty should generate an increased signal to medicine C as well as

S and s': These divergent predictions are respectively instantiated by the error-

driven and Bayesian causal model specified in the subsequent section.

Figure 1. Task illustration. A. Prediction and outcome screens from a feedback-based learning trial. Screens were separated by a 2 sec. pause during which the word "Wait" (not shown in figure) was displayed.   B. Causal query regarding the individual influence of a particular medicine on headache.

On each trial of feedback-based learning (see Figure 1A), participants were

presented with an individual allergy patient and were told either that the patient had

not received any medicine or that a particular medicine, or combination of

medicines, had been administered. Medicines were color coded and labeled

accordingly (i.e., Medicine "B" was blue) with color assignments counterbalanced

across participants. On the first screen, the allergy patient's state was obscured by

a question mark and participants were asked to indicate whether or not the patient

had a headache, pressing the ''Y'' key for ''yes'' and the ''N'' key for ''no''. After a

prediction was made, there was a brief (2000ms) pause during which the word "wait"

was displayed on the screen, followed by a screen revealing the allergy patient's

state (i.e., with or without headache).

Each of 6 distinct trial types was presented 8 times in each of 4 sessions, for a total

of 192 trials, with sessions separated by 2-minute breaks, during which the scanner

was turned off. Each session was further divided into 4, non-delineated, blocks, with

each type of trial occurring once in random order in each block. All trials were

separated by a jittered 4-second ITI. In every other block, each trial (except for n-

trials) was followed by a query regarding the individual causal influence of the target

medicine present on that trial (i.e., C, D or S), or of medicine e or s' on non-target

trials. Specifically, participants were asked to choose between the following three

options regarding the relevant medicine: "is a cause of headache", "is not a cause

headache" and an ambivalent "may or may not be a cause of headache" (Figure

1B). Categorical response options, rather than the rating scales commonly

employed in the causal learning literature (Lu et al., 2008; Liljeholm & Cheng, 2009;

Liljeholm, 2015) were used for two main reasons: First, because the limited number

of response buttons in the scanner would have necessitated a scale slider, the initial

positioning of which might potentially bias ratings, and second, to reduce movement

due to back-and-forth scale scrolling.

Computational models. Two formal accounts of predictive strength and uncertainty

were implemented. First, an error-driven, associative learning algorithm updates

predictive strength, on each trial, using the difference between the observed state of

the outcome and the expected state based on all present cues:

$$V_{c,n+1} \leftarrow V_{c,n} + \alpha_{c,n}\left(\lambda - \sum_i V_i\right)_n \qquad (1)$$

where V is the predictive strength of a particular cue, c, ΣV_i is the summed strength of

all cues present on trial n, and λ is the observed state of the outcome on that trial. The

associability of cue c on trial n, $\alpha_{c,n}$, is defined as:

$$\alpha_{c,n} = \gamma \left| \lambda - \sum_i V_i \right|_{n-1} + (1-\gamma) a_{c,n-1} \qquad (2)$$

where *n-1* references to the previous trial involving cue c, and γ is a free parameter

accounting for the weighting of that previous trial relative to preceding ones. In

addition to scaling the influence of the prediction error on predictive strength, α is

taken to reflect current levels of uncertainty about the cue's influence (e.g., Esberg &

Haselgrove, 2011). The strengths and uncertainties associated with each cue were

initialized to 0.5. On feedback-based trials with target causes, which always occurred

in compound with an alternative cause, the α of the target cause (always greater than

or equal to that of its compound counterpart) was used to indicate uncertainty, while

the summed strength, ΣV, was used to predict the outcome, and to compute the

prediction error.

A second formal account is provided by a Bayesian Causal Model, in which reasoners

make inferences over causal structures potentially responsible for the observed data

(Griffiths & Tenenbaum, 2005; Lu et al., 2008; Liljeholm, 2015). Here, as illustrated in

Figure 2, three possible graphs (*Graphs 0-2*) were defined for a candidate cause, C, an

alternative cause, A, and a constantly present background cause, B, where a causal

link to the effect exists for neither C nor A (G_0), C only (G_1), or both C and A (G_2).

Figure 2. Possible causal structures involving a background cause, B, a target cause, C,
an alternative cause, A, and an effect, E. A link to the effect always exists for B, and
may exist for neither C nor A (*Graph$_0$*), C only (*Graph$_1$*), or both C and A (*Graph$_2$*).

Each graph has a set of parameters θ, which are strengths w_i associated with causal

links in each graph (i.e., w_B, w_C and w_A for links associated with B, C and A respectively

in G_2). Sequential estimates of these parameters were modeled, for each candidate

cause, using the smallest possible *focal set* (Cheng & Novick, 1992) – a set of events

across which alternative causes can be assumed to occur with the same probability in

the presence and absence of the relevant cause. Thus, estimates of the strength of

the "no medicine" background cause was modeled using only n- trials in G_0. Estimates

of elemental (non-target) candidate causes s' and e were respectively modeled based

on trials with those causes and the "no medicine" trials (i.e., n- and s'± for candidate

s', and n- and e- for candidate e), using G_1. Finally, for each target candidate cause,

which always occurred in compound with some alternative cause, estimation was

modeled given G_2, using the trials relevant for the particular target cause (i.e., n- and

Cc' for target C; n-, e- and De+ for target D, and n-, e- and Se± for target S).

The likelihoods $P(d|\theta, G_i)$ were computed using a noisy-OR parameterization (Cheng,

1997). Specifically, summarizing data d by contingencies $N(e,c,a)$, the frequencies of

each combination of the presence versus absence of the effect, target cause and

alternative cause, the likelihood term for G_2 is:

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$$P\left(d|w_B, w_C, w_A, G_2\right) = \left(\begin{array}{c} N\left(c^-,a^-\right) \\ N\left(e^+,c^-,a^-\right) \end{array}\right) \times \left(\begin{array}{c} N\left(c^+,a^-\right) \\ N\left(e^+,c^+,a^-\right) \end{array}\right) \times \left(\begin{array}{c} N\left(c^-,a^+\right) \\ N\left(e^+,c^-,a^+\right) \end{array}\right) \times \left(\begin{array}{c} N\left(c^+,a^+\right) \\ N\left(e^+,c^+,a^+\right) \end{array}\right) w_B^{N\left(e^+,c^-,a^-\right)}$$

$$(1-w_B)^{N\left(e^-,c^-,a^-\right)}\left[w_B+w_C-w_B w_C\right]^{N\left(e^+,c^+,a^-\right)}\left[1-w_B-w_C+w_B w_C\right]^{N\left(e^-,c^+,a^-\right)}\left[w_B+w_A-w_B w_A\right]^{N\left(e^+,c^-,a^+\right)}\left[1-w_B-w_A+w_B w_A\right]^{N\left(e^-,c^-,a^+\right)} \quad (3)$$

$$\left[w_B+w_C+w_A-w_B w_C-w_B w_A-w_C w_A+w_B w_C w_A\right]^{N\left(e^+,c^+,a^+\right)}\left[1-w_B-w_C-w_A+w_B w_C+w_B w_A+w_C w_A-w_B w_C w_A\right]^{N\left(e^-,c^+,a^+\right)}.$$

where $\left(\begin{smallmatrix} n \\ k \end{smallmatrix}\right)$ denotes the number of ways of picking k unordered outcomes from n

possibilities. $N(c^+)$ indicates the frequency of events in which the target cause is

present, with analogous definitions for the other $N(.)$ terms. The likelihood terms for G_0

and G_1 are similarly specified, where frequencies $N(.)$ are summed across the

presence and absence of relevant events, and $w_i=0$ for any cause that does not have

link to the effect in the relevant graph.

The marginal posterior distribution over strengths for a particular cause is obtained by

applying Bayes' rule, and integrating out the parameters for other causes in the graph,

such that, for G_2,

$$P(w_C \mid d, G_2) = \int_0^1 \int_0^1 \frac{P(d \mid w_B, w_C, w_A, G_2) P(w_B, w_C, w_A \mid G_2)}{P(d)} dw_B \, dw_A,$$ (4)

where $P(d \mid w_B, w_C, w_A, G_2)$ is the likelihood term, $P(w_B, w_C, w_A \mid G_2)$ refers to the prior

probabilities of causal strength parameters, and $P(d)$ is the normalizing term, denoting

the probability of the observed data.

Uncertainty about the strength of a particular cause, the focus of the study, was

modeled as the Shannon entropy $H(w_c)$ of its marginal posterior distribution $P(w_C \mid d)$:

$$H(w_C) = -\int_0^1 P(w_C \mid d) \ln P(w_C \mid d) dw_C.$$ (5)

Point estimates of causal strength for each elemental and target cause were modeled

as the mean of the relevant marginal posterior distribution (Lu et al., 2008; Liljeholm,

2015), given by:

$$\bar{w}_C = \int_0^1 w_C P(w_C \mid d) dw_C.$$ (6)

The model was sequentially implemented on a trial-by-trial basis, such that, for each

cause, likelihoods were computed on each trial of feedback-based learning that

yielded information relevant to that cause, using only the data point provided on that

trial, with the posterior on the previous trial being used as the prior to obtain the

posterior on the current trial. On the first trial in which data relevant to a particular

candidate cause was presented, the priors were assigned independent uniform

distributions. Finally, as an analog to the model-free prediction error, the Kullback-

Leibler (KL) divergence between the prior (Q) and posterior (P) on each trial,

commonly referred to as "Bayesian surprise" (Itti & Baldi, 2006) was computed as:

$$KL(P\|Q) = \int_0^1 p(w_C \mid d) \log \frac{p(w_C \mid d)}{q(w_C \mid d)} dw_C.$$

(7)

As noted, target causes only occurred in compound with alternative causes during

feedback-based learning.   On such trials, the entropy and KL divergence of the

marginal distribution for the target cause was used to indicate uncertainty and surprise

respectively, while predictions regarding the occurrence of the effect were generated

using a noisy-OR integration of causal strengths.

Both models assumed that, during feedback-based learning, participants selected

"Yes"/"No" responses to questions about whether headache would occur using

probabilities generated by a softmax distribution, in which a free noise parameter

controls the influence of predictive strength (i.e., the strength with which the

medicine(s) present on a given trail predicted headache; noisy-logical $P(e|w_i)$ and ΣV

in causal and error-driven models respectively) on choice behavior. An additional

noise parameter controlled the influence of uncertainty (i.e., α and $H(w_c)$ in error-

driven and causal models respectively) on the proportion of ambivalent "May or may

not be a cause" judgments during queries about the individual influence of each

medicine. The first two instances of each trial type, and the first causal judgment for

each medicine, were excluded from model fitting to eliminate transient noise due to

task adjustment, and to minimize the influence of assumptions about priors. Free

parameters were fit to behavioral data by minimizing the negative log likelihood of

observed responses for each individual and the Bayesian Information Criterion (BIC)

was used for model selection.

The model-derived probabilities of uncertain "May or may not be a cause" judgments

are plotted in Figure 3 for the causal and error-driven account of uncertainty

respectively. Note that, as was foreshadowed in the description of the study design,

the Bayesian causal model predicts high levels of uncertainty for both confounded, C,

and stochastic, S and s', causes, while the error-driven algorithm predicts that only

stochasticity will generate high levels of uncertainty.

Figure 3. Mean (normalized) probabilities of uncertain "May or may not be a cause" judgments derived from the Bayesian causal model and error-driven algorithm respectively. Confounded, deterministic and stochastic target causes (C, D & S) occurred only in compound with alternative causes during feedback-based learning, while non-target causes occurred only individually (s'), or individually as well as in compound (e). Error bars=SEM.

Imaging procedure and analysis. A 3 Tesla scanner (Phillips Achieva) was used to

acquire structural T1-weighted images and T2*-weighted echoplanar images

(repetition time = 2.65 s; echo time = 30 ms; flip angle = 90°; 45 transverse slices;

matrix = 64 × 64; field of view = 192 mm; thickness = 3 mm; slice gap = 0 mm) with

BOLD contrast. The first four volumes of images were discarded to avoid T1

equilibrium effects. All remaining volumes were corrected for differences in the time of

slice acquisition, realigned to the first volume, spatially normalized to the Montreal

Neurological Institute (MNI) echoplanar imaging template, and spatially smoothed with

a Gaussian kernel (8 mm, full width at half-maximum). High-pass filtering with a cutoff

of 128 s was used. All effects were reported at a whole brain FWE cluster corrected

threshold of $p < 0.05$, calculated using the Statistical nonParametric Mapping toolbox

(SnPM13; http://warwick.ac.uk/snpm; Nichols & Holmes, 2002), with 5000

permutations, no variance smoothing and an uncorrected height threshold of $p <$

0.001.

To address collinearity due to the shared prediction by causal and error-driven

accounts that uncertainty increases with increased stochasticity, the two algorithms

were analyzed using separate 1st level General Linear Models (GLM), and were

directly contrasted at the group level using Bayesian Model Selection. Specifically, for

each participant, two GLMs were specified, one for the causal and one for the error-

driven computational model.   In each such GLM, three regressors respectively

modeled 1) the onset of the prediction screen on each trial during feedback-based

learning, 2) the onset of the outcome screen on each trial during feedback-based

learning, and 3) the onsets of the screens soliciting judgments about the influences of

individual medicines. In one GLM, prediction-screen onsets and causal judgment onsets were modulated by trial-by-trial error-driven estimates of uncertainty and strength, while outcome-screen onsets were modulated by trial-by-trial absolute values of the prediction errors. In the second GLM, prediction-screen and causal judgment onsets were modulated by trial-by-trial estimates of uncertainty and strength by the causal model while outcome-screen onsets were modulated by trial-by-trial estimates of Bayesian surprise. In both models, the first two instances of each trial type, and corollary causal judgments, were modeled by separate regressors, to eliminate transient noise due to task adjustment and reduce the influence of, non-modeled, prior beliefs. Moreover, in each model, two onset regressors respectively modeled the responses on feedback-based and causal query trials, three regressors indicated separate sessions, and six additional regressors accounted for the residual effects of head motion. Except for those indicating sessions and head motion, all regressors were convolved with a canonical hemodynamic response function. No orthogonalization was applied.

To directly contrast error-driven and causal accounts, BMS analyses was performed using a set of GLMs with the same onsets as those specified above, but with a single parametric modulator (e.g., error-driven uncertainty at prediction screen onsets) entered for each GLM. Thus, model comparisons isolated the relative contribution of a particular computational variable to neural activity. First-level Bayesian estimation procedure was used to compute a log-model evidence map for every subject and each GLM and inferences at the group level were modeled by applying a random effects approach (Rosa et al., 2010) at every voxel of the log evidence data. Group-

level exceedance probability maps, reflecting the probability that one model is more likely than the other, were thresholded to identify voxels in which the exceedance probability was greater than 0.95. Classical inferences were then assessed using BMS masks, within which a particular computational variable was more likely than its competing counterpart, at an exceedance probability greater than 0.95: For example, significant effects of error-driven uncertainty during the prediction period of feedback-based learning trials were reported only for those regions in which error-driven uncertainty during this event period was more likely than causal uncertainty during the same period, according to the BMS.

A pilot study on error-driven uncertainty. Previous research has implicated the anterior insula and dorsal medial frontal cortex in outcome stochasticity, using a range of decision-making tasks (Mohr et al., 2010; Volz et al., 2003; Huettel et al., 2005; Grinband et al., 2006). To ensure that such effects, used here as a benchmark against which to assess neural representations of uncertainty due to confounding, also emerge in our causal learning task, a pilot study (n=10) was conducted that was highly similar to that reported here, with the following exceptions: First, only stochasticity, not confounding, was manipulated across putative causes, second, both generative and preventive causal influences were included and, finally, judgments of individual causal influences were solicited at the end of, rather than throughout, each scanning session, and measured on a scale raging from -100 (strongly removes headache) to +100 (strongly produces headache). The error-driven model of uncertainty was implemented as described above, fit to behavior during feedback-based learning, and regressed against the BOLD data. A single region of interest

(ROI) was constructed from the anterior insula and dorsal medial frontal cortex, using

the "Willard" functional parcellation atlas (Richiardi et al., 2015). A power analysis

performed with NeuroPower (http://neuropowertools.org; Durnez et al., 2015), using a

screening threshold of z=2.3, revealed that a sample size of 16 would yield a power of

0.9 to detect effects of error-driven predictive uncertainty in this ROI at a Gaussian

random field corrected alpha level of 0.05.

Results

Behavioral results

All statistical tests of behavioral data were planned comparisons, employing two-tailed

t-tests, and were calculated using *n*=19. Confidence intervals (*CI*) and effect sizes

(Cohen's *d*) are reported for all comparisons. The distributions of causal judgments

across response options are shown, with associated response times, for each putative

cause in Table 1. Note that the proportion of ambivalent "May or may not be a cause"

judgments (3rd row in Table 1) were greatest for the confounded cause (C),

intermediate, although still substantial, for the two stochastic causes (S and s'), and

virtually absent for deterministic causes (D and e), exactly as predicted by the causal,

but not the error-driven, model (cfr Figure 3; see Liljeholm, 2015 for similar results).

Note also that the distribution of "Is a cause" and "Is not a cause" judgments differed

markedly across stochastic and confounded causes, such that participants almost

never select the latter option for the confounded cause, consistent with the normative

increase in the likelihood that target cause C is causal given Cc+ trials, but distributed

their responses fairly evenly across these two options for stochastic causes, reflecting,

perhaps, the outcome on trials immediately preceding each judgment.

Planned comparisons revealed that the mean proportion of ambivalent causal

judgments were significantly lower for the deterministic (D) target cause than for both

confounded (C), t(18)=6.78, *p*<0.001, *CI*=[0.35 0.67], *d*=1.79, and stochastic (S),

t(18)=4.09, *p*<0.001, *CI*=[0.14 0.44], *d*=1.11, target causes. The difference between

the confounded and stochastic cause was also significant, with the mean proportion of

ambivalent judgments being greater for C than for S t(18)=2.65, *p*<0.05, *CI*=[0.0.05

0.40], *d*=0.68. Recall that all target causes, including S, occurred only in compound

with alternative causes during feedback-based learning. The proportion of ambivalent

judgments for the two stochastic causes, one always occurring in isolation (s') and the

other always in compound (S), was marginally significant, *p=0.08*.

Judgment response times, averaged across response options for each cause, were

significantly faster for judgments about the deterministic target cause D than

stochastic target cause S, t(18)=2.20, *p*<0.05, *CI*=[0.03 1.02], *d*=0.35, but did not differ

significantly between D and C, nor between C and S, *p>0.1*. Moreover, response

times did differ significantly across the two stochastic causes, being significantly

slower for target cause S, which never occurred in isolation during feedback-based

learning, than for cause s', t(18)=3.42, *p*<0.005, *CI*=[0.29 1.21], *d*=0.52.

With respect to model fitting, the mean best fitting values of the free parameters for

the error driven model were 0.35±0.27 for the weighting parameter, 3.25±0.95 for the

prediction noise parameter, and 1.08±0.88 for the judgment noise parameter.   The

mean best fitting values of the free parameters for the causal model were 4.13±1.44

for the prediction noise parameter and 1.03±1.05 for the judgment noise parameter.

The mean BIC was significantly lower, indicating superior performance, for the causal

model (220.00±44.15) than for the error-driven model (237.09±35.58), $t(18)=2.51$,

p<0.05, *CI*=[2.81 31.37], *d*=0.43.

Neuroimaging results

All results reported below survived inclusive masking with voxels identified by

Bayesian model selection analyses. A complete list of the non-masked significant

effects of each model is provided in Tables 2 and 3. Bar graphs show effect sizes

extracted from 8mm spheres centered on peak MNI coordinates, using rfxplot

(Glascher, 2009).

Uncertainty signals during causal judgments. During causal queries, as participants

made judgments about the causal influences of individual medicines, selective and

significant effects of the causal model of uncertainty emerged in the medial and lateral

superior frontal gyrus and the right dorsolateral prefrontal cortex (Figure 4). Note that

the effect size of BOLD responses (bar graphs in Figure 4) across putative causes

were greatest for the confounded cause (C), intermediate for stochastic causes (S and

s'), and smallest for deterministic causes (D and e), exactly as predicted by the

causal, but not the error-driven, model (cfr Figure 3). To further probe these effects,

response times (RTs) were included as a parametric modulator of judgment

uncertainty in 1st level, individual-subject, models, and as a between-subjects

covariate in group-level analyses. Neither of these additional analyses revealed any

modulatory influence by RTs on the neural effects of the causal model of judgment

uncertainty. No effects of the error-driven measure of uncertainty during causal

judgments survived whole-brain correction.

Figure 4. Selective effects of the causal model of uncertainty during judgments about the causal influences of individual medicines, showing effects in the dorsomedial and dorsolateral prefrontal cortex (DMPFC and DLPFC), and the lateral Superior Frontal Gyrus (LSFG). Bar plots show effect sizes extracted from spheres centered on peak coordinates for each queried medicine. Error bars=SEM.

Uncertainty signals during prediction. During the prediction period of feedback-based

learning trials, the error-driven, but not causal, measure of uncertainty was

significantly positively correlated with activity throughout the inferior occipital and

lingual gyri, as well as with activity in the supplementary motor area and right anterior

insula (Figure 5). Moreover, activity negatively correlated with error-driven predictive

uncertainty was observed in the posterior middle temporal gyrus.

Figure 5. Selective effects of the error-driven model of uncertainty during the prediction period of feedback-based learning trials, showing effects in the right anterior insula (aINS), posterior middle temporal gyrus (pMTG) and lingual gyrus (LG). Bar plots show effect sizes extracted from spheres centered on peak coordinates for each type of medicine trial. Error bars=SEM.

These effects of the error-driven account of uncertainty appear relatively robust, as

they closely mirror those from the pilot study, in which, at an uncorrected threshold of

0.005, extensive (cluster size > 100) effects of error-driven predictive uncertainty

emerged in the dorsal medial frontal cortex, bilateral anterior insula, and lingual gyrus

during feedback-based learning. In contrast to the widespread effects of error-driven

predictive uncertainty during feedback-based learning, selective and significant effects

of the causal measure of uncertainty emerged only in the right dorsal thalamus.

Figure 6. Strength predictions derived from the causal and error-driven model, together with effect sizes extracted from spheres centered on peak coordinates in the precentral gyrus, for each type of medicine trial during feedback-based learning. Error bars=SEM.

Strength & surprise signals. While uncertainty in causal inferences, due to

confounding or stochasticity, is of primary interest here, estimates of the strength with

which a cause, or combination of causes, predicts the effect, and of the surprise at the

outcome on each trial of feedback-based learning, were also modeled for

completeness. During the prediction period of feedback-based learning trials, the

causal measure of predictive strength was selectively and significantly positively

correlated with activity in the calcarine and superior occipital gyrus, as well as

negatively correlated with activity in the left lateral superior frontal and precentral gyri

(Figure 6). In contrast, a selective and significant negative correlation with the error-

driven measure of predictive strength emerged in the posterior MTG. No significant

effects of either the causal or error-driven account of predictive strength were found

during causal judgments.

Figure 7. Outcome surprise signals derived from the causal and error-
driven model (top), together with effect sizes (bottom) extracted from
spheres centered on peak coordinates in the anterior caudate nucleus
(aCN) and anterior insula (aINS), for each type of medicine trial during
feedback-based learning. Error bars=SEM.

For measures of surprise during the outcome period of feedback-based learning trials,

the absolute prediction error generated by the error-driven account was selectively

and significantly positively correlated with activity in the bilateral anterior insula, the

medial and lateral dorsal prefrontal cortex and the inferior parietal gyrus. In addition,

significant negative correlations with this measure were identified in the middle

occipital gyrus, fusiform gyrus, the superior parietal lobule, precentral gyrus and

calcarine sulcus. In contrast, a selective and significant negative correlation with the

causal measure of (Bayesian) surprise was observed in the anterior caudate.

Discussion

Sensitivity to the "no confounding" constraint – a requirement that putative causes do

not co-vary, fundamental to philosophical and statistical perspectives on causation –

has been reliably demonstrated in a range of behavioral judgments by children as well

as adults (Spellman, 1996; Kushnir and Gopnik, 2005; Meder et al., 2006; Schulz and

Bonawitz, 2007; Schulz et al., 2007; Liljeholm, 2015). The current study used fMRI

and cognitive computational modeling to investigate neural substrates mediating

uncertainty associated with causal confounding. While being scanned, participants

studied and judged the influences of various putative causes with confounded or non-

confounded, deterministic or stochastic, influences. During judgments requiring

generalization of causal knowledge from a feedback-based learning context to a

transfer probe, activity in the dorsomedial prefrontal cortex (DMPFC) was better

accounted for by a Bayesian causal model, sensitive to both confounding and

stochasticity, than a purely error-driven algorithm, sensitive only to stochasticity.

The DMPFC has been implicated in encoding uncertainty across a variety of tasks

(Volz et al., 2003; Grinband et al., 2006; Xue et al., 2008; Michael et al., 2015). For

example, scanning participants with fMRI as they classified stimuli according to a

variable category boundary, Grinband et al. (2006) found that activity in the DMPFC

scaled with the proximity of a stimulus to a category boundary. An important aspect of

this and other demonstrations of the involvement of the DMPFC in uncertainty is that

the level of uncertainty is directly related to the degree of experienced errors: that is,

the greater the uncertainty on a given trial, the more likely it is that the prediction on

that trial will be incorrect, as indicated by explicit feedback. In contrast, during

feedback-based learning in the current task, the outcome on confounded trials was

deterministic. Consequently, the increased activity in the DMPFC, as participants

generated judgments about confounded cause, cannot be attributed to a history of

errors. Instead, these results suggest that the DMPFC encodes a more abstract

representation of uncertainty, divorced from the immediate consequences of choice,

and irrespective of whether the specific source of uncertainty is confounding or

stochasticity.

Several neuroeconomic studies (Hsu et al., 2005; Krain et al., 2006; Huettel et al.,

2006; Bach et al., 2009; Levy et al., 2010; Pushkarskaya et al., 2015) have contrasted

gambles involving stochasticity with "ambiguous" gambles, for which outcome

probabilities are completely unknown, with some finding greater activity in the DMPFC

in response to ambiguity than stochasticity (e.g., Hsu et al., 2005). One might argue

that, in the current study, the influence of the confounded medicine is unknown, just

as the outcome probabilities of ambiguous gambles in previous studies were

unknown. But what does it mean to "know" something? Critically, all target medicines

occurred only in compound with alternative medicines during feedback-based

learning: consequently, when presented individually during causal queries, each was

equally novel and, presumably, equally unknown. In other words, differences in

ambiguity between confounded and other target causes cannot be attributed to a

generalization decrement based on changes in stimulus features. Moreover, ambiguity

due to small samples and uncertainty due to confounding call for very different plans

of exploratory action: while in former case, repeated observation of the same stimulus

can resolve the uncertainty, the latter case requires an intervention that unconfounds

the stimulus configuration. Further work is needed to assess the overlap between

neural representations of confounding and ambiguity.

In addition to stochasticity and ambiguity, the DMPFC has been heavily implicated in

cognitive control (Ridderinkhof et al., 2004; Fellows & Farah, 2005; Egner, 2009;

Taren et al., 2011; Shenhav et al., 2013), raising the question of whether the currently

observed effects in this region might reflect similar processes.   In particular,

judgments about individual causal influences may be associated with efforts to

retrieve instances of the queried cause occurring in isolation from episodic memory –

a search that is futile, and thus presumably more effortful, for confounded causes.

However, such retrieval-induced control demands are unlikely to account for the

current effects in the DMPFC for several reasons: First, effort associated with episodic

memory retrieval does not address responses to stochasticity – in contrast, the causal

uncertainty account explains why DMPFC responses are greater for *both* confounded

and stochastic causes, relative to non-confounded and deterministic causes. It may

be argued, of course, that stochasticity elicits additional control processes mediated

by the DMPFC, such as, for example, response conflict (Botnivick et al., 2001;

Ridderinkhof et al., 2004; Kerns et al., 2004). Recall, however that *none* of the target

causes (C, S and D) were presented in isolation during feedback-based learning and

that, consequently, *all* target causes should involve some level of searching for non-

presented information during causal judgments. If DMPFC activity reflected this

search process as well as response conflict due to stochasticity, one would expect a

difference between neural responses to causes S and s', both of which are stochastic,

but only one of which, s', was presented in isolation during feedback-based learning,

eliminating the need to search for an un-experienced memory. Notably, while RTs

(Table 1), commonly used as a measure of cognitive load, are indeed consistent with

the notion of an increased control demand for cause S over s', the DMPFC responses

(Figure 4) are not. Finally, and perhaps most conclusively, the inclusion of RTs as a

parametric modulator of judgment uncertainty in 1st level models, and as a covariate in

group analyses, did not have any impact on the neural effects of model-based

judgment uncertainty. Taken together, these aspects of the results make causal

uncertainty the more parsimonious explanation of the DMPFC effects observed here.

As with the DMPFC and DLPFC, the anterior insula has been implicated in outcome

stochasticity, or "risk", in several previous neuroimaging studies (e.g., Volz et al.,

2003; Huettel et al., 2005; d'Acremont et al., 2013; Preuschoff et al., 2008; Critchley et

al., 2001). Here, the dorsal anterior right insula was selectively recruited by the error-

driven model of uncertainty during the prediction period of feedback-based learning

trials. Plots of effect sizes for each type of trial (in Figure 5) suggest that these

responses were indeed driven by stochasticity–based uncertainty about the state of

the outcome on any given trial. Moreover, the fact that, unlike the DMPFC, activity in

the anterior insula did not scale with either an error-driven or causal measure of

uncertainty during judgments about individual causal influences indicates that

representations in this region pertain primarily to the immediate consequences of

choice, rather than generalizable knowledge. Consistent with this interpretation, the

bilateral anterior insula also responded selectively to stochastic conditions during the

outcome period of feedback-based learning trials, suggesting that this region is

recruited by both the anticipation and detection of errors. While, as noted, it is difficult

to discern based on the current results whether the unsigned error signals observed in

the anterior insula are Bayesian or model-free, they are consistent with a substantial

literature implicating this region in expectancy violations across a wide range of tasks

(Klein et al., 2007; Metereau et al., 2012; Allen et al., 2016; Bastin et al., 2016).

Selective effects of the error-driven account of predictive uncertainty were also

identified throughout large portions of the posterior lingual gyrus. In a recent

neuroimaging study, Causse et al., (2013) assessed decision making under

uncertainty using an aviation task in which participants had to either land a plane or

abort landing, given certain (100% or 0%) vs. uncertain (50%) information regarding

landing conditions. Consistent with the current results, they found greater activity in

the posterior lingual gyrus during high than during low uncertainty conditions.   Our

results are also consistent with those of Payzan-LeNestour et al. (2016), who found

that the lingual gyrus scaled specifically with uncertainty due to outcome stochasticity,

as opposed to uncertainty due to changes in the statistical structure of the

environment or due to ambiguity. Payzan-LeNestour et al. argued that the lingual

gyrus might not typically emerge in studies assessing outcome stochasticity because

those studies did not dissociate stochasticity from other uncertainty components. A

similar case can be made here, that the inclusion of an alternative, generative, model

that captures additional aspects of uncertainty may have increased sensitivity to

detect stochasticity-specific signals in the lingual gyrus.

An important aspect of the current study is the direct comparison of a Bayesian

inference model with an error-driven, model-free, algorithm. While not uncommon

(Courville et al, 2006; Hampton et al., 2006; Payzan-LeNestour & Bossaerts, 2011;

Prevost et al., 2013), the contrasting of these quite different computational frameworks

somewhat complicates interpretation. In particular, it raises the question of whether

selective effects of the causal account of uncertainty indeed reflect sensitivity to

confounding, or instead some extraneous, perhaps incidental, feature of the different

modeling frameworks. Here, a significant clue is provided by neural and behavioral

responses to each queried putative cause, shown in Figure 4 & Table 1 respectively,

which clearly correspond to the increased uncertainty associated with confounded and

stochastic causes predicted by the causal model (cfr Figure 3). It is worth noting that,

far from incidental, the failure of the error-driven account to predict sensitivity to

confounding is intrinsic to the model-free approach, due to its lack of a representation

of the independence of causal influences. Critically, a configural solution (e.g.,

Pearce, 1987; 2002), such that individually queried causes are treated as non-

overlapping with the compounds in which they occur during feedback-based learning,

would fail to account for both behavioral and neural differences between confounded

and deterministic target causes, both of which occurred only in compound with

alternative causes during feedback-based learning.

While the effects of causal uncertainty in the DMPFC likely reflect sensitivity to causal

confounding, other neuroimaging results are clearly unrelated and, in some cases,

largely incidental to the specific model implementation. In particular, as can be seen

in Figures 6 and 7, the ordinal mean values derived from causal and error-driven

accounts across trial types during feedback-based learning are identical for both

strength and surprise measures, suggesting that model-selection may reflect more

granular differences. For example, because of the Σ term in Equations 1 and 2 of the

error-driven account, the variance in the occurrence of the outcome across e-, De+

and Se± trials results in a constantly fluctuating strength, and persistent prediction

errors, even on deterministic e- and De+, trials. In contrast, in the implementation of

the causal model, such non-conditional variance is eliminated by the use of a focal set

(e.g., Cheng & Novick, 1992; Spellman, 1996). Behaviorally, the focal set assumption

is strongly supported by the lack of variability in judgments about non-target cause e.

This cause, which is paired with the outcome on half of trials in which it occurs, but

never when occurring in isolation, is almost exclusively judged to be non-causal,

suggesting that participants conditioned their inferences about its influence on trials

across which alternative causes could be assumed to be held constant. Nonetheless,

at the neural level, several regions may instead have been tracking the unconditional

outcome variance computed by the error-driven model.

As the mantra "covariation does not equal causation" implies, requirements for causal

induction extend far beyond mere statistical regularity. Countless studies have

demonstrated neural and behavioral effects of error-driven uncertainty, whether due to

stochasticity, insufficient samples or changes in the statistical structure of the

environment.　The current study reveals that, given violation of a basic boundary

condition on causal inference, and in spite of the same constraints on perceptual

generalization, differences in neural representations of uncertainty can emerge across

equally deterministic, static and well-sampled contingencies. These results contribute

to a growing literature on the role of generative models in neural representations of

uncertainty, and shed light on a possible neural implementation of domain general, a

priori, constraints on causal induction.

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760   Figure Captions

Figure 1.   Task illustration.   A.   Prediction and outcome screens from a feedback-

based learning trial.   Screens were separated by a 2 sec. pause during which the

word "Wait" (not shown in figure) was displayed.   B. Causal query regarding the

individual influence of a particular medicine on headache.

Figure 2. Possible causal structures involving a background cause, B, a target cause,

C, an alternative cause, A, and an effect, E.   A link to the effect always exists for B,

and may exist for neither C nor A (Graph$_0$), C only (Graph$_1$), or both C and A (Graph$_2$).

Figure 3. Mean (normalized) probabilities of uncertain "May or may not be a cause"

judgments derived from the Bayesian causal model and error-driven algorithm

respectively. Confounded, deterministic and stochastic target causes (C, D & S)

occurred only in compound with alternative causes during feedback-based learning,

while non-target causes occurred only individually (s'), or individually as well as in

compound (e). Error bars=SEM.

Figure 4. Selective effects of the causal model of uncertainty during judgments about

the causal influences of individual medicines, showing effects in the dorsomedial and

dorsolateral prefrontal cortex (DMPFC and DLPFC), and the lateral Superior Frontal

Gyrus (LSFG). Bar plots show effect sizes extracted from spheres centered on peak

coordinates for each queried medicine. Error bars=SEM.

Figure 5. Selective effects of the error-driven model of uncertainty during the

prediction period of feedback-based learning trials, showing effects in the right anterior

insula (aINS), posterior middle temporal gyrus (pMTG) and lingual gyrus (LG). Bar

plots show effect sizes extracted from spheres centered on peak coordinates for each

type of medicine trial. Error bars=SEM.

Figure 6. Strength predictions derived from the causal and error-driven model (top),

together with effect sizes (bottom) extracted from spheres centered on peak

coordinates in the precentral gyrus, for each type of medicine trial during feedback-

based learning. Error bars=SEM.

Figure 7. Outcome surprise signals derived from the causal and error-driven model

(top), together with effect sizes (bottom) extracted from spheres centered on peak

coordinates in the anterior caudate nucleus (aCN) and anterior insula (aINS), for each

type of medicine trial during feedback-based learning. Error bars=SEM.

793
794
795 Table 1. Mean proportions of "Is a cause", "Is not a cause" and "May or may
796 not be a cause" judgments, and associated response times, with standard
797 deviations, for target causes, C, D and S, and non-target causes, s' and e'.

	C	D	S	s'	e
Proportions					
Is a cause	0.36±0.33	0.84±0.25	0.38±0.25	0.40±0.27	0.02±0.03
Is not a cause	0.03±0.04	0.06±0.09	0.23±0.11	0.30±0.17	0.96±0.05
May be a cause	0.61±0.35	0.10±0.21	0.39±0.31	0.30±0.28	0.02±0.04
Response Times					
Is a cause	2.92±1.88	2.52±1.23	4.03±3.62	2.82±2.59	2.86±2.11
Is not a cause	4.26±2.69	3.94±3.36	4.19±3.64	2.72±1.23	2.44±1.15
May be a cause	3.79±2.68	3.76±1.46	2.99±1.94	2.59±1.32	4.08±4.78

800 Table 2: Non-masked significant, whole-brain corrected, neural effects of the
801 causal model of uncertainty, strength and surprise, during causal judgments and
802 feedback-based learning, with peak MNI coordinates and cluster sizes at p<0.001.
803 R/L = Right/Left, (-) = negative correlation.

Contrast	Region	Peak MNI	Cluster Size
Judgment Uncertainty	DMPFC	-8, 28, 46	1900
	R. IPG	50, -52, 40	292
	Dorsolateral PFC	42, 34, 30	201
Prediction Uncertainty	Fusiform	28, -72, -12	7195
	DMPFC	-10, 16, 52	1387
	R. Insula	-30, 18, 4	726
	L. Insula	34, 30, -4	470
	Thalamus	6, -22, 2	867
Prediction Strength	Calcarine/Occipital	18, -96, 4	236
Prediction Strength (-)	Lingual Gyrus	14, -62, -2	2232
	Superior Frontal Gyrus	-26, -8, 58	458
	Superior Parietal Lobule	-28, -50, 62	304
	Inferior Parietal Gyrus	-46, -26, 36	336
	Supp. Motor Area	8, 2, 52	297
Surprise	Right Insula	36, 18, -2	652
	Left Insula	-28, 22, -6	482
	Dorsomedial PFC	4, 46, 34	1146
	Left Dorsolateral PFC	-46, 4, 42	192
	Right Dorsolateral PFC	44, 24, 30	325
Surprise (-)	Supramarginal Gyrus	-50, -26, 38	1858
	Caudate	18, 32, 0	505
	Mid Cingulum	-10, -30, 42	472
	Postcentral Gyrus	50, -20, 48	2399
	Parahippocampal	34, -32, -12	227
	Precuneus	16, -52, 14	221

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307 Table 3: Non-masked significant, whole-brain corrected, neural effects of the
308 error-driven model of uncertainty, strength and surprise during feedback-based
309 learning, with peak MNI coordinates and cluster sizes at p<0.001. R/L =
310 Right/Left, (-) = negative correlation.

Contrast	Region	Peak MNI	Cluster Size
Prediction Uncertainty	Calcarine/Lingual Gyrus	14, -88, 0	3116
	Supp. Motor Area	10, 20, 48	153
	Left Insula	-24, 22, 2	65
	Right Insula	34, 28, -2	109
Prediction Uncertainty (-)	Middle Temporal Gyrus	-46, -62, 20	341
	Middle Temporal Gyrus	-56, -22, -10	84
	Ventromedial PFC	-2, 30, -14	157
	Middle Temporal Gyrus	46, -58, 22	64
Prediction Strength	Lingual Gyrus	-20, -78, -10	2801
Prediction Strength (-)	Superior Parietal Lobule	-18, -52, 64	407
	Supramarginal Gyrus	-60, -28, 30	390
	Precentral Gyrus	26, -12, 50	240
	Supramarginal Gyrus	58, -38, 28	728
	Middle Temporal Gyrus	-54, -60, 2	603
	Precentral Gyrus	-22, -14, 64	383
	Thalamus	-12, -10, -4	214
Prediction Error	Right Insula	38, 18, -4	912
	Left Insula	-40, 20, 4	945
	Precentral Gyrus	-42, -2, 40	713
	Anterior Cingulate	4, 38, 30	2021
	Inferior Parietal Gyrus	-28, -52, 42	246
	Dorsolateral PFC	44, 28, 26	885
Prediction Error (-)	Middle Occipital Gyrus	-42, -74, 26	735
	Fusiform Gyrus	-34, -42, -12	253
	Postcentral Gyrus	54, -26, 52	906