Dissecting heterogeneous cell-populations across signaling and disease conditions with PopAlign

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Abstract

Single-cell measurement techniques can now probe gene expression in heterogeneous cell populations from the human body across a range of environmental and physiological conditions. However, new mathematical and computational methods are required to represent and analyze gene expression changes that occur in complex mixtures of single cells as they respond to signals, drugs, or disease states. Here, we introduce a mathematical modeling platform, PopAlign, that automatically identifies subpopulations of cells within a heterogeneous mixture, and tracks gene expression and cell abundance changes across subpopulations by constructing and comparing probabilistic models. We apply PopAlign to discover specific categories of signaling responses within primary human immune cells as well as patient-specific disease signatures in multiple myeloma that are obscured by techniques like tSNE. We anticipate that PopAlign will enable large scale studies of natural and engineered cell populations as they respond to drugs, signals or physiological change.

Introduction

All physiological processes in the body are driven by heterogeneous populations of single cells [1, 2, 3]. Single-cell measurement technologies can now profile gene expression in thousands of cells from heterogeneous cell populations across different tissues, physiological conditions, and disease states. However, converting single cell data into models that provide a'population-level understanding of processes like an immune response to infection or cancer progression remains a fundamental challenge. All human tissues contain many different subpopulations of cells, and each subpopulation can undergo distinct changes in gene expression and cellular abundance in response to signals, drugs, or environmental conditions. New conceptual and mathematical frameworks are required to model and track the changes that occur within distinct subpopulations of cells within a heterogeneous tissue as they respond to perturbations or succumb to disease.

In this paper, we introduce a computational framework, PopAlign, that identifies, aligns, and tracks subpopulations of single cells within a heterogeneous cell population profiled by single cell mRNA-seq [4, 5, 2, 6]. Mathematically, PopAlign constructs a probabilistic model of each cell population across a series of samples. PopAlign (a) automatically identifies and models subpopulations of cells (b) aligns cellular subpopulations across experimental conditions (signaling, disease) and (c) quantifies changes in cell abundance and gene expression for all aligned subpopulations of cells.

The key conceptual advance underlying PopAlign is representational: we model the distribution of gene expression states within a heterogeneous cell population using a probabilistic mixture model that we infer from single cell data. PopAlign identifies and represents subpopulations of cells as independent Gaussian densities within a reduced gene expression space. PopAlign, then, makes quantitative statistical alignments between subpopulations across samples, and thus enables targeted and quantitative comparisons in gene expression state and cellular abundance. Probabilistic modeling is enabled by a novel low dimensional representation of cell-state in terms of a set of gene expression features learned from data [7, 8, 9].

PopAlign identifies shifts in cell-state that are obscured or even lost by existing techniques like tSNE. Geometric methods like tSNE can reveal subpopulations of cells within a mixture as isolated clusters in gene expression space [10, 11, 12, 13]. However, geometric methods do not provide a natural language for mathematically representing a subpopulation of cells or statistical metrics for quantifying shifts in population structure across experimental samples. PopAlign solves this problem by modeling subpopulations of cells as parameterized probability densities within a common low dimensional space, and then, uses statistical metrics to track quantitative changes within subpopulation specific densities across experimental samples.

We assess the accuracy and generality of PopAlign using twelve datasets from a mouse tissue survey (Tabula Muris) as well as new experiments on human peripheral blood cells in health and a diseased state, multiple myeloma [14]. We show that PopAlign can identify and track cell-states across a diverse range of tissues, signaling experiments, and human disease states. The probabilis-

tic models have high representational accuracy and identify biologically meaningful cell-states from data. In the immune system, we show that PopAlign identifies cell-type specific responses to cytokines in primary human immune cells. Finally, PopAlign extracts general and treatment specific signatures of disease progression from multiple myeloma patient samples. Moving forward, PopAlign sets the stage for the analysis of large-scale experimental screens of drugs and genetic perturbations on heterogeneous cell populations extracted from primary human tissue samples.

Key Contribution

- Probabilistic modeling of cell populations enables alignment of cellular subpopulations across experimental samples, and concurrent tracking of gene expression and cell abundance changes.
- Application of method to data sets from mouse tissues and primary human cells demonstrates accuracy of models and ability to track cell-state specific gene expression changes in response to signals and disease states.

Results

PopAlign represents heterogeneous cell populations with probabilistic mixture models

We develop a mathematical and computational framework (PopAlign) that (i) identifies and aligns cell-states across paired populations of single cells (a reference population and a test population), and then (ii) quantifies shifts in cell-state abundance and gene expression between aligned populations (Figure 1). The method has three steps: probabilistic mixture model construction, model alignment, and parameter analysis. PopAlign can be applied to analyze gene expression and population structure changes in heterogeneous populations of cells as they respond to signals, drugs, and disease conditions.

We consider two populations of cells, a test and a reference population, ($\boldsymbol{D}^{\text{test}}$ and $\boldsymbol{D}^{\text{ref}}$), that are profiled with single cell mRNA-seq (Figure 1a). Profiling of each population generates a set of gene expression vectors, e.g. $D^{\text{Test}} = \{\boldsymbol{g_i}\}_{i=1}^{k}$ where $\boldsymbol{g} = (g_1, g_2, \cdots, g_n)$, is an n dimensional gene expression vector that quantifies the abundance of each mRNA species in single cell $\boldsymbol{g}$ and k is the number of profiled single cells.

To compare the reference and test cell populations, we first, construct a probabilistic model of the gene expression distribution for each set of cells (Figure 1b). The high dimensional nature of gene expression ($n \sim 20,000$) space makes the inference and interpretation of probabilistic models challenging. Therefore, we represent each cell, not as a vector of genes, but as a vector

of gene expression programs or gene expression features that are extracted from the data using non-negative matrix factorization (NMF) (See Methods Section 1), so that each single cell is represented as a vector $\boldsymbol{c} = (c_1, c_2 \cdots c_m)$ of m feature coefficients, c_i, which weight the magnitude of gene expression programs (see Methods) in a given cell. NMF has been shown to produce useful features for gene expression analysis because they have positive entries, and so we can naturally think about the gene expression state of a cell, $\boldsymbol{g}$, as being assembled from a linear sum of positive gene expression programs.

Following dimensionality reduction, for a given cell population, we think of cell states as being sampled from an underlying joint probability distribution over this gene feature space, $P(\boldsymbol{c})$, that specifies the probability of observing a specific combination of gene expression features/programs, $\boldsymbol{c}$, in the cell population. We estimate a probabilistic model, $P^{\text{test}}(\boldsymbol{c})$ and $P^{\text{ref}}(\boldsymbol{c})$, for the reference and test cell populations that intrinsically factors each population into a set of distinct subpopulations each represented by a Gaussian probability density (density depicted as individual 'clouds' in Figure 1b):

$$P^{\text{test}}(\boldsymbol{c}) = \sum_{i=1}^{l} w_i \, \phi_i^{\text{test}}(\boldsymbol{c}) \tag{1}$$
$$\text{where } \phi_i^{\text{test}}(\boldsymbol{c}) = \mathcal{N}(\boldsymbol{c}; \boldsymbol{\mu_i}, \boldsymbol{\Sigma_i})$$

where $\mathcal{N}(\boldsymbol{c}; \mu_i, \Sigma_i)$ are multivariate normal distributions with weight w_i; centroids $\boldsymbol{\mu_i}$ and covariance matrices $\boldsymbol{\Sigma_i}$. The distributions $\phi_i^{\text{test}}(\boldsymbol{c}) = \mathcal{N}(\boldsymbol{c}; \boldsymbol{\mu_i}, \boldsymbol{\Sigma_i})$, mixture components, represent individual subpopulations of cells; l is the number of Gaussian densities in the model. We estimate the parameters of the mixture model $(\{\boldsymbol{\mu_i}, \boldsymbol{\Sigma_i}, w_i\})$ from single cell data using the expectation-maximization algorithm [15, 6] with an additional step to merge redundant mixture components to compensate for fitting instabilities (See Methods - Merging of redundant mixture components).

Gaussian densities provide a natural representation of a transcriptional state in gene expression space which is consistent with measured gene expression distributions as well as empirical models of transcription [16, 17, 18]. The parameters associated with each Gaussian density ,$(\boldsymbol{\mu_i}, \boldsymbol{\Sigma_i}, w_i)$, have a natural correspondence to the biological structure and semantics of a cellular subpopulation. The relative abundance of each subpopulation corresponds to the weight $w_i \in [0, 1]$; the average cell gene expression state of each subpopulation corresponds to the (m dimensional) Gaussian centroid vector $\boldsymbol{\mu_i}$, and the shape or spread of the subpopulation is captured by the covariance matrix $\boldsymbol{\Sigma_i}$. Intuitively, the local Gaussian densities provide a natural 'language' for comparisons between samples. Each Gaussian is a region of high density in gene feature space, and we compare cell populations by asking how the density of cells shifts across experimental conditions.

Statistical alignment of cellular subpopulations between samples

To compare the test and reference models, we 'align' each mixture component in the test population model, $\phi_i^{\text{test}}(\boldsymbol{c}) \in \{\phi_i^{\text{test}}(\boldsymbol{c})\}$, to a mixture component, $\{\phi_i^{\text{ref}}(\boldsymbol{c})\}$, in the reference population model

(Figure 1c). Alignment is performed by finding the 'closest' reference mixture component in gene feature space. Mathematically, to define closeness, we use, Jeffrey's divergence, a statistical metric of similarity on probability distributions. Specifically, for each $\phi_i^{\text{test}} \in \{\phi_i^{\text{test}}(\boldsymbol{c})\}$, we find an $\phi_j^{\text{ref}} \in \{\phi_j(c)\}^{\text{ref}}$, the closest mixture in the reference set:

$$\underset{\phi_j^{\text{ref}}(\boldsymbol{c}) \in \{\phi_j^{\text{ref}}(\boldsymbol{c})\}}{\arg\min} \; D_{\text{JD}} \left(\phi_i^{\text{test}}(\boldsymbol{c}) \parallel \phi_j^{\text{ref}}(\boldsymbol{c})\right), \tag{2}$$

where the minimization is performed over each $\{\phi_i^{\text{ref}}(\boldsymbol{c})\}$ in the set of reference mixtures, and D_{JD} is the Jeffrey's divergence (14). Intuitively, for each test mixture, we find the reference mixture ϕ_j that is closest in terms of position and shape in feature space. For each alignment, we then calculate an explicit p-value from an empirical null distribution $P(D_{JD})$ that estimates the probability of observing a given value of D_{JD} in an empirical data set of all subpopulation pairs within a single cell tissue database (See Methods 3).

Tracking cell-state shifts through mixture model parameters

Following mixture alignment, we analyze quantitative differences in mixture parameters between the reference and test models to track shifts in gene expression state, gene expression covariance, and cellular abundances across the identified subpopulations of cells(Figure 1d). Mathematically, for each aligned mixture pair, $(\phi_i^{\text{test}}, \phi_j^{\text{ref}})$ with parameters $\{\boldsymbol{\mu_i}^{\text{ref}}, \boldsymbol{\Sigma_i}^{\text{ref}}, w_i^{\text{ref}}\}$ and $\{\boldsymbol{\mu_j}^{\text{test}}, \boldsymbol{\Sigma_j}^{\text{test}}, w_j^{\text{test}}\}$, we calculate:

$$\Delta\mu_i = ||\boldsymbol{\mu_i}^{\text{ref}} - \boldsymbol{\mu_j}^{\text{test}}||_2 \tag{3}$$

$$\Delta\Sigma_i = D_{\text{C}}(\boldsymbol{\Sigma_i}^{\text{ref}}, \boldsymbol{\Sigma_j}^{\text{test}}) \tag{4}$$

$$\Delta w_i = |w_i^{\text{ref}} - w_j^{\text{test}}| \tag{5}$$

where $\Delta\mu_i$ measures shifts in mean gene expression; Δw_i quantifies shifts in cell-state abundance; $\Delta\Sigma_i$ quantifies shifts in the shape of each mixture including rotations and changes in gene expression variance (see Methods) [19]. We calculate these shifts in parameters for all mixture pairs to assess the impact of signaling conditions or environmental changes on the underlying cell population.

PopAlign identifies and aligns cell-states across disparate mouse tissues

To test the accuracy and generality of PopAlign, we first constructed and aligned probabilistic models across a wide range of mouse tissues from a recent public study (Tabula Muris) [14].

The Tabula Muris study contains single cell data collected from 12 different tissue samples with $\sim 40,000$ cells total.

For all tissues analyzed, the probabilistic mixture models produce an accurate and interpretable decomposition of the underlying cell states (SI Figure 3). Accuracy of the models can be assessed by comparing the synthetic (model generated) data to raw experimental data held out from model training. PopAlign models generate synthetic data that replicates the geometric structures and statistical variations found in the tissue data in tSNE or PCA plots with quantitative error of $\sim 12\%$ error (See methods; Figure 2; 3a,b; SI Tissues; see methods).

In addition to providing an accurate representation, the mixture models decompose the cell populations into a biologically interpretable set of cellular subpopulations represented by individual $\phi_i(\boldsymbol{c})$, the mixture components (Figure 3c,d). The PopAlign mixture components, $\{\phi_i(\boldsymbol{c})\}$ commonly contain cells of a single cell 'type' as defined by the Tabula Muris project labels and, in example tissues, PopAlign extracts known tissue resident cell-types including (Figure 3c,d) basal cells, luminal cells, macrophages, and T-cells (in mammary gland) and skeletal muscle cells, mesenchymal stem cells, endothelial cells, and macrophages (in limb muscle). Broadly, across all tissue models, 70% of the mixture components classified for a single cell-type provided by Tabula Muris (SI Figure 3).

Through alignment of model components across tissues, PopAlign enables high-level comparisons of tissue composition. By aligning Mammary Gland to Limb Muscle (Fig. 3e), we identified 'common' cell-types between the two tissues including B-cells (p=0.0006), T-cells (p=0.001), endothelial cells (p=0.0013), and macrophages (p=0.004, 0.0076) (SI Fig 4), and also revealed tissue scale differences in relative abundance. T-cells are highly prevalent ($w = .3$ in the mammary gland but rare in the limb muscle $w = .05$) (Fig. 3g); endothelial cells are highly abundant in the limb muscle ($w = .32$), but rare in the mammary gland ($w = .06$) (Fig. 3g). Between shared cell types, such as macrophages, we reveal common programs such as FC-receptor Signaling and Lysosome, as well as tissue-specific gene expression programs such as TGF-Beta, Phagocytosis, and Leukocyte Chemotaxis(Fig. 3h). PopAlign can, thus, give insight into the underlying composition of a tissue, shedding light onto principles of tissue organization with respect to tissue function.

PopAlign identifies subpopulation specific responses to cytokines in primary immune cell populations

A key application of PopAlign is to study heterogeneous cell populations in the human body as they respond to environmental change, drug treatments, and disease. The human immune system is an important application domain for PopAlign as an extremely heterogeneous physiological system that is central for disease and cell engineering applications [2, 20, 21, 22, 23]. As a test case, we asked whether PopAlign could identify subtype-specific signaling responses in a complex mixture of primary human immune cells. Thus, we treated primary human peripheral blood mononuclear

cells (PBMCs) (Fig. 4a) with two lymphokines - granulocyte-macrophage colony-stimulating factor (GM-CSF)(Fig. 4b) and interferon-gamma (IFNG) (Fig. 4c) that are known to have a specific impact on monocytes [24, 25] (Methods).

Broadly, PopAlign found that both lymphokine signals impact gene expression in monocytes (canonical and non-canonical) while driving minimal gene expression changes in T-cells. In the reference PBMCs, we identified five independent cell densities (Fig 4a), and classified them using canonical markers from the literature (Fig. 4d). We aligned the signal treated cell-populations (Fig. 4c-d) to the PBMC data (Fig. 4a), finding that, for both signals, the monocyte populations (mixtures PBMC-1, PBMC-3) had the largest shifts in their average gene expression state ($—\Delta\mu—> 0.04$) (Fig 4f). While T-cell abundances changed (numbers), their mean cell states ($\boldsymbol{\mu}$) had the lowest magnitude of gene expression shift.

Additionally, the PopAlign framework exposed that GM-CSF and IFNG exerted qualitatively different classes of transcriptional effects. In the GM-CSF condition, both canonical and non-canonical monocytes (mixture 1 and 3) respond to the signal 'additively' by activating a new gene expression program called Chemokine Secretion/Leukocyte Migration pathway (Figure 4g bottom, blue arrows) that includes genes such as CCL2 (Fig 4g,h, iii), and cystatin B (CSTB) while losing the expression of 'Lysosome' program, which includes genes such as LYZ and granulysin (GRN) (Figures 4h: v, vi). We call the GM-CSF induced change in gene expression 'additive' because both classes of monocytes respond with the same transcriptional changes, but these changes are induced in addition to their baseline gene expression state (Fig 4g - black arrows).

Alternately, IFNG signaling generates a 'convergent' signaling response inducing both canonical and non-canonical monocytes to a single new transcriptional state. IFNG leads to a single new monocyte subpopulation that aligns most closely with 'non-classical' monocyte population(Fig 4b), but also has a high secondary alignment score for the 'classical' monocyte population (Fig. 4c, 4e - red x). We infer based on the disappearance of the 'classical monocytes' and the high score of the secondary alignment, that IFNG drives both subtypes towards a new transcriptional state that has up-regulation of pathways involved in antigen presentation and phagocytosis - 'ER Phagosome', 'IFN-Gamma signaling', 'Antigen Processing and Presentation' [24]. Genes upregulated include the known targets of IFNG STAT1(Fig. 4h -i) and also PSME2 (Fig. 4h - ii), a proteasome gene involved in the non-lysosomal protein degradation pathway. The induced programs are consistent with literature showing that IFNG converts monocytes to macrophages, cells that recycle cellular debris within the body [26, 27].

In this way PopAlign reveals cell-type specific signaling responses within a complex cell mixture subject to a globally applied signal. PopAlign, further, enabled us to define distinct categories of signaling responses, 'additive' and 'convergent', within a highly heterogeneous population of primary cells.

tSNE obscures association between cell-states following signal addition

PopAlign extracts and quantifies population level changes in the signaling data set that are obscured by existing geometric techniques like tSNE. tSNE provides a low dimensional but non-linear embedding of data points that seeks to preserve geometric features of the high dimensional data while separating distinct cell-states into clusters [10]. Thus, in the tSNE embedding of the signaling data, the relationships between cell-states (subpopulations) is often obscured because gene expression changes can result in the formation of new, separate, and 'distant' clusters (Fig 5). For example, in response to IFNG, the monocyte populations become separated in the two dimensional embedding (Figure 5b) making association of these clusters pre- and post-signal complicated. Further, tSNE does not provide natural metrics for quantifying distinct parameters of a point cluster, so that abundance, mean, and shape all must be defined and quantified through heuristic approaches. Unlike tSNE, PopAlign provides a precise framework for defining populations of cells and also for calculating statistical similarity scores for individual subpopulations across samples by using the Jeffrey's divergence to measure changes in gene expression state, abundance, and covariance.

PopAlign discovers general and treatment-specific signatures of multiple myeloma

Given the success of the PopAlign framework in extracting cell-type specific responses in the immune signaling data, we applied the method to study underlying changes in cell state due to a disease process. As a model system, we applied PopAlign to compare human PBMC samples from healthy donors to patients being treated for multiple myeloma (MM). Multiple myeloma is an incurable malignancy of blood plasma cells in the bone marrow. Both the disease and associated treatments result in broad disruptions in cell function across the immune system [28, 29, 30, 31] further contributing to disease progression and treatment relapse. In MM patients, immune cells with disrupted phenotypes can be detected in the peripheral blood[32, 30, 33]. An ability to monitor disease progression and treatment in the peripheral blood could therefore provide a powerful new substrate for making clinical decisions.

We obtained samples of frozen PBMCs from two healthy and four multiple myeloma patients undergoing various stages of treatment (SI Table 1). We profiled $> 5,000$ cells from each patient, and constructed and aligned probabilistic models to one reference healthy population (Figure 6a-f).

PopAlign identified several common global signatures in the MM samples at the level of cell-type abundance and gene expression. Across all samples, we find previously known signatures of multiple myeloma including a deficiency in B cells [30, 34, 35], and an expansion of monocyte/myeloid derived cells [32], and critically, new impairments in T-cell functions.

Plotting Δw across all patients, we find high-level changes in subpopulation abundances, which are known to be prognostic of disease progression [33]. We find that all MM patients experience a contraction in B cell numbers (Fig. 4b), and 2 out of 4 see a dramatic expansion ($\Delta w >> .10$) of

monocytes (Fig. 4c). Changes in T cell levels, however, can be highly variable, with outlier patient MM4 experiencing a large increase in effector T cell ($\Delta w = .2$), and a complete elimination of resting T cells ($\Delta w = .2$). For this patient, who was receiving a thalidomide-derived drug therapy, these deviations are consistent with thalidomide's known stimulatory effects on T-cells [36].

Especially in patients with apparently normal abundances (i.e. Δw are small), uncovering subpopulation-specific changes in transcription can point to specific modes of immune dysfunction. We use PopAlign to find that monocyte subpopulations in patients acquire immunosuppressive phenotypes, evidenced by upregulated expression of CD11b and CD33. Both genes are specific markers of myeloid derived suppressor cells[37] which are negative regulators of immune function associated with cancer. By plotting the monocyte-specific mean gene expression values for both CD11b and CD33, we see that all patients except patient MM3 score highly for both MSDC markers. (Fig 6k). Patients with high MDSC populations typically have a poor prognosis, underscoring the need to monitor MDSC populations in patients.

Importantly, we also find that naive and effector T cells across all multiple myeloma patients have transcriptional defects in pathways essential for T cell function. By plotting $\Delta\mu$, we show that both populations of T cells experience large mean transcriptional shifts, compared to T cells from our second healthy donor, healthy2 (Fig. 6m). By examining the $\mu's$ in terms of gene expression vectors (Fig. 6n), we find that in multiple myeloma, T cells reduce their expression of two key features - Leukocyte Motility, and Cytotoxic Lymphocyte Killing. Surprisingly, the impact on motility is apparent even on the expression of beta-actin (ACTB) (Fig. 6o), a core subunit of the actin cytoskeleton, and which was the top hit in the Leukocyte Motility feature. We find similar declines in the distribution of Perforin 1 (PFN1), a pore-forming cytolytic protein that was found as a top hit in the Cytotoxic Lymphocyte program (Fig. 6p).

Our analysis establishes that we can extract consistent and also patient-specific transcriptional signatures of human disease and treatment response from PBMCs. Interpreting these signatures in the context of disease progression or drug response can provide insight into treatment efficacy and can form the basis of a personalized medicine approach. Our framework enables new applications by providing a highly scalable way of extracting, aligning, and comparing these disease signatures, across many patients at one time.

Discussion

In this paper, we introduce PopAlign, a computational and mathematical framework for tracking changes in gene expression state and cell abundance in a heterogeneous cell populations across experimental conditions. The central advance in the method is a probabilistic modeling framework that represents a cell population as a mixture of Gaussian probability densities within a low dimensional space of gene expression features. Models are aligned and compared across experimental samples, and by analyzing shifts in model parameters, we can pin-point gene expression and cell

abundance changes in individual cell populations.

PopAlign constitutes a conceptual advance over existing single cell analytical methods. PopAlign is explicitly designed to track changes within complex cell populations. Since human diseases like cancer and neurodegeneration arise due to interactions between a wide variety of cell-types within a tissue, population level models will be essential for building a single cell picture of human disease and for understanding how disease interventions like drug treatments impact the wide range of cell-types within a tissue.

Mathematically, existing single cell analysis methods rely on heuristic cluster based analysis to extract subpopulations of cells. Fundamentally, such approaches lack well defined statistical metrics for making comparisons across samples. By conceptualizing a single-cell population as a probability distribution in gene expression space, we define a discrete mathematical objects whose parameters can be interpreted, and which can be used to explicitly calculate quantitative statistical metrics for subpopulation alignment. Our probabilistic representation allows us to quickly and scalably learn drug responses even on a complex mixture of cells, in 'one shot'. Computationally, PopAlign scales to analyze many samples that each contain thousands of single cells. We anticipate applying the method to analyze data from drug screens where the goal is to identify cell-population specific transcriptional responses.

In the future, we hope that PopAlign can be used as a part of a work-bench for single cell analysis and treatment of human disease. By applying PopAlign to data sets from the human immune system, we highlight the potential power of PopAlign for identifying drug/signal targets and for deconstructing single cell disease states. PopAlign identified cell-type specific signatures of disease treatment in multiple myeloma patients exposing a potential defect in T-cell activation and motility in three patient samples. This result points to a potential use of PopAlign for guiding treatment interventions by exposing the spectrum of transcriptional states within a diseased tissue and revealing the impact of drug treatments on diseased cell-states as well as the cellular microenvironment and immune cell-types. Such insights could lead to single cell targeting of drug combinations to treat human disease as an essentially population level phenomena.

Methods

Mathematical framework

We consider two populations of cells, a reference population, ($\boldsymbol{D}^{\text{test}}$ and $\boldsymbol{D}^{\text{ref}}$), and a test population. Following profiling by single cell mRNA-seq, each population of cells is a set of gene expression vectors, $\boldsymbol{D} = \{\boldsymbol{g_i}\}_{i=1}^{k}$ where k is the number of cells in the population, and $\boldsymbol{g} = (g_1, g_2, \cdots, g_n)$, is an n dimensional vector that quantifies the abundance of each mRNA species. While raw mRNA-seq measurements generate integer valued gene count data, due to measurement

noise and data normalization, we consider $\boldsymbol{g}$ to be embedded in an n dimensional Euclidean vector space, gene expression space, $\boldsymbol{g} \in \mathbb{R}^n$. The high dimensional nature of gene expression space poses the key challenge for construction and interpretation of statistical models.

We think of the gene expression vectors, $\{\boldsymbol{g_k}\}$ as being distributed according to an underlying probability density function, $P(\boldsymbol{g})$, that quantifies the probability of observing a particular joint gene expression state, $\boldsymbol{g} = (g_1, g_2, \cdots, g_n)$, in a given cell population. Our broad goal is to estimate a statistical model of $P(\boldsymbol{g})$,based upon single cell measurements. The model provides a parametric representation of the gene expression density in each condition. Then, we seek to use this representation to track changes in the structure of the cell population across conditions.

In general, the mathematical challenge we face is model estimation in the high dimensional nature of gene expression space. For human cells $n > 20,000$, and single cell profiling experiments can routinely probe 10,000 cells per sample. The number of parameters in our probabilistic models scales quadratically with n, and mixture model learning has data requirements that are exponential in n [8]. Therefore, we first reduce the dimensionality of the problem by building models in a common low dimensional space defined by gene expression programs discovered from pooled data across all samples.

Data normalization

Single cell gene expression data must be normalized to 1) to account for the variation in the number of transcripts captured per cell and 2) to balance the wide disparity in the scale of values across different genes due to measurement noise and gene drop-out.

The total number of transcripts captured for each single cell can vary from 1000 to 100,000 unique transcripts per cell. Technical variability in reagents and library prep steps can have a large impact on the number of transcripts retrieved per cell. To scale out these differences, we divide each gene expression value g_i by the total number of transcripts and then multiply by a scaling factor β.

Additionally, across genes, mean transcript values can span 5 orders of magnitude. Transforming the data using the logarithm brings values across all genes close in scale, while also reducing the skew in the data distributions. The equation for transforming a raw gene expression value g_i (for a single gene) into a normalized gene expression value, g_i' is:

$$g_i' = \log(\beta \frac{g_i}{\sum_i^n g_i} + 1) \tag{6}$$

where n is the total number of genes, and β is a scaling factor, and we add a 1 pseudo-count to each gene expression value. We found that by setting $\beta = 1000$ to be roughly the minimum number of total transcript counts in a cell (1000 transcripts), we achieve a smooth transition in the distribution of transformed g_i' when raw g_i values step from 0 to 1. Gene expression values are thus denoted in units of log(TPT+1) where TPT is transcripts per thousand.

Extraction of gene feature vectors with matrix factorization

We circumvent the curse of dimensionality ([7]) by building models in a common low-dimensional space defined by gene expression features or programs. Mathematically, we represent the transcriptional state of each single cell, $\boldsymbol{g}$, as a linear combination of gene expression feature vectors, $\{\boldsymbol{f_i}\}$:

$$\boldsymbol{g} = \sum_{i=1}^{m} c_i \, \boldsymbol{f_i} \tag{7}$$

where $\mathbf{f_i} \in \mathbb{R}^{\mathbf{n}}$ specifies a gene expression feature, and c_i is a coefficient that encodes the weighting of vector f_i in $\boldsymbol{g}$, the gene expression state of a single cell. The key result in [7] is that a cell's gene expression state, $\boldsymbol{g}$ can be represented as a linear combination of m gene expression module vectors, $\boldsymbol{f_i}$ where $m << n$. This insight allows us to construct a low dimensional representation of a cell population and, then, to estimate statistical models within the low dimensional space. [7, 38, 39].

The gene features, $\{\boldsymbol{f_i}\}$ can be extracted using a wide range of matrix factorization and machine learning technique including Singular Value Decomposition, and its matrix factorization relatives like sparse PCA as well as methods like layered neural networks [7]. We use a technique called orthogonal non-negative matrix factorization [40] (oNMF) to define a space of orthogonal gene expression features vectors. Like other linear dimensionality reduction techniques, such as PCA, oNMF factors the original data matrix, $\boldsymbol{D}^{\text{train}}$ (SI Fig. 1a) into two matrices $\boldsymbol{D} \approx \boldsymbol{F}\ \boldsymbol{C}$ (SI Fig. 1b,c). Factorization occurs through minimization of an objective function with positivity and orthogonality constraints:

$$\begin{aligned} &\arg\min_{\boldsymbol{F},\boldsymbol{C}} ||\boldsymbol{D}^{\text{train}} - \boldsymbol{F}\boldsymbol{C}||_2 \\ &\text{subject to } \ \boldsymbol{F}^T\boldsymbol{F} = I, C_{ij} \geq 0, F_{ij} \geq 0, \end{aligned} \tag{8}$$

The optimization minimizes the (Frobenius) norm of the difference between the training data, $\boldsymbol{D}^{\text{train}}$, and its factored representation $\boldsymbol{F}\boldsymbol{C}$. The columns of $\boldsymbol{F}$ contain gene features, $\boldsymbol{f_i}$. The matrix $\boldsymbol{C}$ is m by k, where k is the number of single cells in $\boldsymbol{D}^{\text{train}}$. Each column of $\boldsymbol{C}$ encodes the weighting of the m gene features across a given single cell. The entries of $\boldsymbol{F}$ and $\boldsymbol{C}$; are constrained to be positive, and the columns of $\boldsymbol{F}$ (the gene features) are constrained to be orthogonal. $\boldsymbol{F}$ is an n by m matrix (genes by features). Each column contains n weights where each weight corresponds to the weight of a given gene in that feature, $\boldsymbol{f_i}$.

Standard non-negative matrix factorization has been shown to provide a useful set of features for gene expression analysis because feature vectors have positive entries, and so we can naturally think about the gene expression state of a cell, $\boldsymbol{g}$, as being assembled as a linear sum of positive gene expression programs.

In PopAlign, we incorporate orthogonality in $\boldsymbol{F}$ as a secondary constraint to aid interpretation. Empirically, we found that orthogonality aids in interpretation of the features as well as in model construction because the orthogonal gene expression features are interpretable as non-overlapping

components (SI Fig. 1b) that can individually be analyzed by gene set enrichment analysis. Second, orthogonality tended to force individual cell states to be represented by more than one feature which aided stability during model parameter estimation.

To perform oNMF, we select m, the number of features to be extracted through an optimization that balances accuracy and dimensionality explicitly. In oNMF (as opposed to PCA and SVD), m is a parameter given to the optimization. Choosing m involves balancing the tension between the 'expressiveness' in the feature set and its dimensionality. Higher m reduces the error in the representation while also breaking up blocks of genes into smaller modules that represent independent gene expression pathways with finer granularity. However, as m increases, the typical computational and sampling challenges associated with high dimensionality emerge.

Practically, we balance this tension in PopAlign by constructing a loss function with a penalty that increases with m:

$$\arg\min_{m} f(m) = \ ||\boldsymbol{D}^{\text{train}} - \boldsymbol{F_m C_m}||_2 + m^{\alpha}. \tag{9}$$

For each value of m, we perform oNMF on $\boldsymbol{D}^{\text{train}}$ yielding $\boldsymbol{F_m}$ and $\boldsymbol{C_m}$, and thus an error $||\boldsymbol{D}^{\text{train}} - \boldsymbol{F_m C_m}||^2$. This error is, then, incremented by the term m^{α} which penalizes higher values of m and hence the dimensionality of the feature set. We set $\alpha = .7$ based upon numerical experimentation on model data sets (SI Figure 2). For any choice of m, we can estimate the accuracy of the representation by plotting reconstructed data FC (SI Figure 1d) against normalized data D (SI Figure 1a). The SI shows such plots and the PopAlign software package outputs these plots by default.

Practically, given data sampled a set of cell populations,(eg $D^{\text{test}_1}, D^{\text{test}_2}, D^{\text{ref}}$) we pool data from all cell populations into a training data set, $\boldsymbol{D}^{\text{train}}$, and perform oNMF. If we are analyzing a large number of data sets or sets with many single cells, we generate $\boldsymbol{D}^{\text{train}}$ by sampling 500-1,000 cells uniformly at random from the reference and test cell populations and selecting a m via ((9)).

Because the feature vectors are not always purely orthogonal, we recast the complete dataset into the feature space using a non-negative least squares. Specifically, for each gene expression profile, $\boldsymbol{g}$, we find, $\boldsymbol{c}$ via:

$$\begin{aligned} &\arg\min_{\boldsymbol{c}} ||\boldsymbol{g} - \boldsymbol{F}\ \boldsymbol{c}||_2 \\ &\boldsymbol{c_i} \geq 0 \end{aligned} \tag{10}$$

where $\boldsymbol{F}$ is a fixed feature set learned from oNMF on $\boldsymbol{D}^{\text{train}}$. The gene expression vector g is thus compressed into a k-dimensional vector c_i that provides a high-level programmatic representation of cell state in terms of gene expression 'features'. Finally, we interpret the biological meaning

of the feature vectors in terms of annotated gene expression programs using gene set enrichment analysis.

Using matrix factorization, we map a cell population, $D = \{\boldsymbol{g}\}$, from an $n \sim 20,000$ dimensional gene expression space into a gene feature space that is often of order $10 - 20$ dimensions [7].

$$D = \{\boldsymbol{g_i}\} \rightarrow \{\boldsymbol{c_i}\},$$

where $\{\boldsymbol{c_i}\}$ are $m \times 1$ dimensional vectors that now represent the cell population in the reduced gene feature space.

oNMF error analysis

The error associated with each feature set F_m was assessed by comparing data entries between a cross validation dataset $\boldsymbol{D_x}$ and its reconstructed matrix $\boldsymbol{F_m C_x}$. We binned the data in $\boldsymbol{D_x}$ into bins of equal width ~ 0.1 (in units of log(TPT+1)). We retrieved data values from each bin, $(\boldsymbol{D_x})_{\text{bin}}$, and then plotted their means against the means of corresponding data values in the reconstructed matrix, $(\boldsymbol{F_m C_x})_{bin}$ (SI Figure 1d). The standard error in each bin is calculated as the mean squared deviation of the reconstructed data from the original data:

$$\sigma_{bin} = \sqrt{\frac{\sum_i^n ((\boldsymbol{F_m C_x})_i - (\boldsymbol{D_x})_i)^2}{n}}, \ \ \forall\, i \in bin$$

To quantify the amount of dispersion relative to the mean, we also calculate the coefficient of variation for each bin:

$$CV_{\text{bin}} = \frac{\sigma_{bin}}{\overline{(\boldsymbol{D_x})_i}}, \ \ \forall\, i \in bin$$

We find empirically that the average CV is $\sim 30 - 35\%$ across all bins for most feature sets.

Representing a cell population as a Gaussian Mixture Model in gene feature space

Following the feature based representation, we construct a statistical model of each given cell population within the reduced gene feature space. Mathematically, we have exchanged a probability distribution in gene expression space for a probability distribution in gene feature space:

$$P(\boldsymbol{g}) \rightarrow P(\boldsymbol{c}), \tag{11}$$

where $\boldsymbol{g}$ is $n \times 1$ and $\boldsymbol{c}$ is $m \times 1$, and $m << n$. We can now estimate a statistical model of $P(\boldsymbol{c})$.

To account for the heterogeneity of cell-states within a tissue, we model cell-populations using Gaussian mixture models. Gaussian densities provide a natural representation of a transcriptional

state in gene expression space which is consistent with measured gene expression distributions as well as empirical models of transcription [16, 17, 18]. Theoretical models of stochastic transcription commonly yield univarite gene expression distributions where mRNA counts are Poisson or Gamma distributed. Normal distributions provide a reasonable approximation to these distributions with a computationally tractable inference procedure.

We represent the cell population as a mixture of Gaussian densities, so that for a given cell population D, we construct:

$$\widehat{P(\boldsymbol{c})} = \sum_{i=1}^{l} w_i \, \phi_i(\boldsymbol{c}) \tag{12}$$
$$\phi(\boldsymbol{c}) = \mathcal{N}(\boldsymbol{c}; \boldsymbol{\mu_i}, \boldsymbol{\Sigma_i})$$

where $\widehat{P(\boldsymbol{c})}$ is a mixture of Gaussian densities, $\mathcal{N}(\boldsymbol{c}; \boldsymbol{\mu_i}, \boldsymbol{\Sigma_i})$, with centroid, μ_i; covariance matrix, ,Σ_i; and scalar weighting w_i. $\boldsymbol{\mu_i}$ is a vector in the m dimensional feature space, and $\boldsymbol{\Sigma_i}$ is a symmetric $m \times m$ matrix. l is the number of Gaussian mixtures or components in the statistical model. The Gaussian mixture model (GMM) represents a cell populations as a mixture of individual Gaussian densities. Biologically, we think of each density as parameterizing a subpopulation of cells.

We can estimate the parameters $\boldsymbol{\mu_i}$, $\boldsymbol{\Sigma_i}$, and w_i based upon training data, using maximum likelihood estimation with likelihood function:

$$\mathcal{L}(\{\mu_i, \Sigma_i, \pi_i\} | \boldsymbol{D'}) = \sum_{j=1}^{k} \log P(\boldsymbol{c_j}; \{\boldsymbol{\mu_i}, \boldsymbol{\Sigma_i}, w_i\}), \tag{13}$$

where $\boldsymbol{c_j}$ are single cell profiles drawn from a cell population $\boldsymbol{D'}$ and cast into feature space; k is the number of single cells in the cell population $\boldsymbol{D'}$. For a given experimental data set, $\mathcal{L}$ defines a function over the space of model parameters. To select model parameters given data, we can attempt to maximize the value of $\mathcal{L}$. In general for Gaussian Mixture models, likelihood maximum is complicated by the geometry of $\mathcal{L}$ which is not concave and can have multiple local and global maxima [41]. $\mathcal{L}$ can be maximized approximately using expectation maximization.

Expectation-maximization is a heuristic algorithm that finds (local) maximum likelihood parameters. Although it is known to have fundamental problems - including weak performance guarantees and a propensity to overfit data, new methods [8] place constraints that are invalid for our application (such as shared covariance matrices). We find empirically that the EM algorithm performs well, learning low-error representations of $\boldsymbol{c}$ (SI Fig. S2), and that we can overcome fitting instabilities by algorithmically merging components. Practically, we perform expectation maximization

using sci-kit learn. We regularize the variance of individual mixtures to constrain variance to be non-zero to avoid fitting instabilities. We determine mixtures number through the Bayesian Information Criterion (BIC) which optimizes a trade-off between model complexity and accuracy on training data.

Merging of redundant mixture components

One drawback of the EM algorithm is its propensity to fit 'redundant' mixture components to the same local density with significant overlap. For our application, this redundancy complicates interpretation and comparisons across samples. We overcome this problem by taking advantage of mixture model properties to algorithmically merge redundant mixtures using the Jeffrey's divergence.

For multivariate Gaussian distributions, the Jeffrey's divergence has a closed analytic form.

$$D_{\text{JD}}(\phi_i \| \phi_j) = \frac{1}{2}(D_{\text{KL}}(\phi_i \| \phi_j) + D_{\text{KL}}(\phi_j \| \phi_i)) \tag{14}$$

where ϕ_0 and ϕ_1 are two independent components from the same mixture model (12), and μ and Σ are their associated parameters.

D_{KL} is the Kullback Leibler divergence and has a convenient parametric form for Gaussian distributions:

$$D_{\text{KL}}(\mathcal{N}_i(\boldsymbol{c}; \boldsymbol{\mu_i}, \boldsymbol{\Sigma_i}) \| \mathcal{N}_j(\boldsymbol{c}; \boldsymbol{\mu_j}, \boldsymbol{\Sigma_j})) = \frac{1}{2}\left(\text{tr}\left(\Sigma_i^{-1}\Sigma_j\right) + (\mu_i - \mu_j)^{\text{T}}\,\Sigma_i^{-1}(\mu_i - \mu_j) - k + \ln\left(\frac{\det \Sigma_i}{\det \Sigma_j}\right)\right)$$

For each mixture model, we iteratively attempt to merge component pairs with the lowest Jeffrey's divergence and accept mergers that increase the BIC of the model given the data. With each merge step, model parameters for candidate pair are recalculated, and the updated model is accepted or rejected based on the new BIC. Mergers are performed until the first rejection. This procedure removes redundant mixture components from the model.

Sampling data from Gaussian mixture models

Given parameters, $\{(\boldsymbol{\mu_i}, \boldsymbol{\Sigma_i}, w_i)\}_{i=1}^{l}$ for a given mixture model. We can 'generate' synthetic data from the model through a simple sampling procedure. A given model has, l mixtures, and we first select a mixture from the set of l mixtures with probabilities weighted by w_i. Following selection of a mixture, j, we use standard methods to draw a 'point' in the m dimensional feature space from $\mathcal{N}_|(\boldsymbol{c}, \boldsymbol{\mu_j}, \boldsymbol{\Sigma_j})$.

Analysis of model error

Model error was assessed by comparing the distribution of model generated data to the empirical data. To avoid under-sampling, we performed error analysis within two dimensional projections of the data and averaged error over all 105 projections. Briefly, each projection was binned into 25 bins. In each bin, the deviation between the model generated data and empirical data was calculated as percent error, $\frac{N_d - N_m}{N_d}$, where N_d and N_m are, respectively, the number of data of empirical and model generated data points in a given bin. Total error within each projection is calculated as a weighted average of this percent error over all bins in a projection:

$$\text{error} = \sum_i \frac{N_i}{N_\text{T}} \frac{|N_i - \widehat{N_i}|}{N_i},$$

where N_i is the number of experimental data points in bin i, and $\widehat{N_i}$ is the number of model generated data points in bin i. N_T is the total number of experimental data points. This metric weights the per-bin fractional error by the probability density of each binned region in the projection.

Alignment of models across reference and test populations

To compare the test and reference models, we 'align' each mixture component in the test population model, $\phi_i^{\text{test}}(\boldsymbol{c}) \in \{\phi_i^{\text{test}}(\boldsymbol{c})\}$, to a mixture component, $\{\phi_i^{\text{ref}}(\boldsymbol{c})\}$, in the reference population model (Figure 1c). Alignment is performed by finding the 'closest' reference mixture in gene feature space. Mathematically, to define closeness, we use, Jeffrey's divergence, a statistical metric of similarity on probability distributions. Specifically, for each $\phi_i^{\text{test}} \in \{\phi_i^{\text{test}}(\boldsymbol{c})\}$, we find an $\phi_j^{\text{ref}} \in \{\phi_j(c)\}^{\text{ref}}$, the closest mixture in the reference set:

$$\underset{\phi_j^{\text{ref}}(\boldsymbol{c}) \in \{\phi_j^{\text{ref}}(\boldsymbol{c})\}}{\arg\min} \; D_{\text{JD}} \left(\phi_i^{\text{test}}(\boldsymbol{c}) \parallel \phi_j^{\text{ref}}(\boldsymbol{c})\right), \tag{15}$$

where the minimization is performed over each $\{\phi_i^{\text{ref}}(\boldsymbol{c})\}$ in the set of reference mixtures, and D_{JD} is the Jeffrey's divergence (14). Intuitively, for each test mixture, we find the reference mixture ϕ_j that is closest in terms of position and shape in feature space.

For each alignment, we then calculate an explicit p-value from an empirical null distribution $P(D_{JD})$ that estimates the probability of observing a given value of D_{JD} in an empirical data set of all subpopulation pairs within a single cell tissue database. Aligning subpopulations using Jeffrey's divergence incorporates information about the shape and position of the probability distributions (i.e. covariance), while disregarding the relative abundance of each subpopulation.

Scoring alignments

To assign a p-value to alignments, we calculate the probability of observing two cell-states with a

given Jeffrey's divergence by chance using an empirical null distribution generated from the tissue data set. Specifically, we constructed a mixture model for all tissue pairs in Tabula Muris. Then, we calculate all pair-wise Jeffrey's divergence scores for all the underlying mixtures. This calculation gives us a global distribution over D_{JD} for cell-states in the mouse. This distribution provides a null distribution for typical statistical closeness between cell-states in feature space.

Model interpretation through parameter analysis

Following mixture alignment, we analyze quantitative differences in mixture parameters between the reference and test sample to track shifts in gene expression state, gene expression covariance, and cellular abundances across the identified cell-states in the cell population.

Specifically, for each aligned mixture pair, $(\phi_i^{\text{test}}, \phi_j^{\text{ref}})$ with parameters $\{\mu_i^{\text{ref}}, \Sigma_i^{\text{ref}}, w_i^{\text{ref}}\}$ and $\{\mu_j^{\text{test}}, \Sigma_j^{\text{test}}, w_j^{\text{test}}\}$, we calculate:

$$\Delta\mu_i = ||\boldsymbol{\mu_i}^{\text{ref}} - \boldsymbol{\mu_j}^{\text{test}}||_2$$
$$\Delta\Sigma_i = D_{\text{C}}(\boldsymbol{\Sigma_i}^{\text{ref}}, \boldsymbol{\Sigma_j}^{\text{test}})$$
$$\Delta w_i = |w_i^{\text{ref}} - w_j^{\text{test}}|$$

where $\Delta\mu_i$ measures shifts in mean gene expression; Δw_i quantifies shifts in cell-state abundance; $\Delta\Sigma_i$ quantifies shifts in the shape of each mixture including rotations and changes in gene expression variance . We calculate these shift parameters for all mixture pairs, and then analyze the shifts to assess the impact of signaling conditions or environmental changes on the underlying cell population.

$$D_{\text{C}}(\Sigma_i, \Sigma_j) = \sqrt{\sum_{z=1}^{m} \ln^2 \lambda_z(\Sigma_i, \Sigma_j)}$$

where λ_z is a generalized eigenvalue of $\boldsymbol{\Sigma_i}$ and $\boldsymbol{\Sigma_j}$ or a solution to $\boldsymbol{\Sigma_i}\, v = \lambda_z \boldsymbol{\Sigma_j} v$, and m is again the number of gene features.

Classification of mixtures using marker genes

For the Tabula muris data set, cells and mixtures were classified using cell-type annotations provided by the study. Mixtures were classified according to the cell-type with the maximum abundance within a given mixture (Figure 2).

For immune cell experiments, we classified the independent mixture components as effector T cells (CD3+/CD57+), naive T-cells(CD3+/CD28+), erythrocytes (HBB+), canonical monocytes

(CD14+/CD16low), and nonclassical monocytes(CD14low/CD16++)[42] (Fig 4e). We aligned populations in our test samples - GM-CSF (Fig 4b) and IFNG (Fig. 3c) - to the reference populations (control) by finding pairs of components that minimize the Jeffrey's divergence (Fig 4d).

Software implementation

PopAlign has been implemented within our LOWMIX software infrastructure package. It is written in Python3 and requires common scientific computing libraries (numpy, matplotlib, pandas, seaborn, tables, MulticoreTSNE, adjustText) that can be easily installed with pip and our requirements file. A guide on how to get set up and install dependencies is provided on the packages Github page. The software runs on local machines, as well as on Amazon Web Services headless EC2 instances, providing a powerful setting for large-scale analyses. The architecture of this package is built around classes that store experimental samples as objects and provide a set of specific methods to perform tasks such as normalization, dimensionality reduction, model construction, model alignment, and parameter comparison.

Experimental methods

Single-cell RNA-sequencing. Cryopreserved PBMCs (healthy and disease samples) from Hemacare were thawed in a 37C waterbath for 2 minutes after which the cells were transferred to a 15mL conical tube. Prewarmed RPMI1640 was then added to the 15mL conical to a final volume of 10mL and centrifuged for 2 minutes at 300RCF to pellet the cells. Supernatant was removed and cells were resuspended to 1 million cells/mL in RPMI1640 supplemented with 10% FBS and 17,400 cells were loaded into each TENX lane.

PBMC signaling experiment. Cryopreserved PBMCs sourced from Hemacare (~ 10 million cells, source ID: D23127, Lot: 17043194) were thawed in a 37C waterbath for 2 minutes after which the cells were transferred to a 15mL conical tube. Prewarmed RPMI1640 was then added to the 15mL conical to a final volume of 10mL and centrifuged for 2 minutes at 300RCF to pellet the cells. Supernatant was removed and cells were resuspended in 10mL of RPMI1640 supplemented with 10% FBS and 1% pen/strep. The cell suspension was then plated onto a 100mm low attachment plate and placed in a CO2 incubator at 37C for 5 hours for the cells to recover.

For each condition, 700,000 cells were aliquoted into a well of a 96well plate. These cells were cultured for 13hours under two different conditions: 1) IFN- @1ng/mL, 2) GM-CSF 1ng/mL, each with a single additional replicate. After incubation, the plate was sealed with an aluminum adhesive and centrifuged for 3minutes at 300RCF. Supernatant was removed and replaced with 40L of TrypLE and incubated for 2 minutes before transferring cells to two 1.5mL Eppendorf tubes, one for each culture condition.

Both tubes were centrifuged and the cells were resuspended to 5 million cells/mL in PBS+ 1% FBS. 180L from each cell suspension processed using BDgenomics multiplexing kit and 8,000 cells from each condition were combined and run through a single TENX lane.

Figure Captions

Figure 1. Summary of PopAlign framework PopAlign provides a scalable method for deconstructing quantitative changes in population structure including cell-state abundance and gene expression across many single cell experimental samples. (a) Users input PopAlign single-cell gene expression data from a 'Reference' sample, and at least one 'Test' sample, which are each a collection of n-dimensional gene expression vectors $\boldsymbol{g}$, shown as single dots. (b) For each sample, PopAlign estimates a low-dimensional probabilistic model that represents the distribution of gene expression states as a mixture of local Gaussian densities ϕ_i with parameters encoding subpopulation abundance (w_i), mean gene expression state (μ_i), and population spread (Σ_i). PopAlign reduces the dimensionality of the input data by representing each gene expression vector as set of m gene expression features ($m = 10-20$), thus representing each cell as an m-dimensional vector of coefficients $\boldsymbol{c}$. (c) Each ϕ_i^{test} in the test population is aligned to the closest ϕ_i^{ref} in the Reference sample by minimizing Jeffrey's divergence. (d) Following alignment, the parameters of aligned subpopulation pairs are compared to identify subpopulation-specific shifts in cellular abundance Δw, shifts in mean gene expression state $\Delta\mu$ and shifts in subpopulation shape $\Delta\Sigma$.

Figure 2. PopAlign models represent experimental data with high qualitative and quantitative accuracy (a) Experimental data for ~ 3600 bone marrow cells projected into an $m = 15$ dimensional gene feature space. 2D plots show single cells projected along gene feature pairs (c_i, c_j), and a single selected 3D projection (inset) is shown. Blue axis denote shared axis between 2D and 3D plot. (b) Model generated data for the same 2D and 3D feature space projections shown in (a). In the 3D projection (inset), each maroon circle denotes the centroid (μ) of a Gaussian mixture component where the circle radius is proportional to w_i, the mixture weight. In all cases, the model generated data replicates the qualitative geometric structures in the experimental data. (c) Model error across 2D projections from (a) and (b) quantified by analyzing percent deviation in point density for experimental vs model generated data. Quantification of error is performed using a numerical error metric based on binning the 2D projections (See Methods - Analysis of model error). The quantitative error in model-generated is on average 11.3% (red line) compared with 8.5% error when comparing random sub-samples of experimental data (blue line).

Figure 3. Probabilistic models identify, align, and dissect cellular subpopulations across disparate tissues Experimental single-cell data (black) for two tissues, mammary gland (a) and limb muscle (b), are plotted together with PopAlign model-generated data (teal) using a 2D t-SNE transformation. Both experimental datasets contain ~ 3600 cells. For each tissue, mixture model centroids (μ) are indicated as numbered disks. (c-d) For models from both tissues, mixture components (x-axis) are scored using cell type annotations supplied by Tabula Muris (y-axis). Each

cell of the heatmap represents the percentage of cells associated with each mixture component that have a specific cell type label. Columns (but not rows) sum to 1. (e) Alignments between mixture component centroids (μ) from the reference population (Mammary Gland) and the test population (Limb Muscle) are shown as connecting lines. All m-dimensional mu vectors are transformed using principal components analysis (PCA) and plotted using the first 3 PCS. Width of each line is inversely proportional to the p-value associated with the alignment (see Legend). (f) Null distribution of Jeffrey's divergence used to calculate p-values. Jeffrey's divergence was calculated for all possible pairs of mixture components from models of all tissues from Tabula Muris. (g) We rank aligned subpopulations in terms of maximum Δw and show top two pairs. These subpopulations are identified as T cells and endothelial cells, and are highlighted using a blue dotted line in (e). We find that T cells are highly abundant in Mammary gland while endothelial cells are highly abundant in muscle. (h) Comparing subpopulation centroids (μ) for macrophages in terms of gene expression features. Macrophages in Mammary gland and Limb Muscle share common features (black font), but also have tissue-specific features (red font). Corresponding alignments are highlighted in (e) with a gray dotted line.

Figure 4. PopAlign identifies cell-type specific signaling responses in primary human immune cells (a) Experimental data from reference PBMC sample (~ 1300 cells) are projected into a 17-dimensional feature space, but are plotted in a 3D subset of features that are found to be impacted by signals. Mixture model centroids (μ) are indicated as numbered disks. (b) Classification of mixture model components using mean gene expression values for canonical markers of immune cells. Classical monocytes are CD14+, non-classical monocytes are CD14low/CD16+, effector T cells are CD3D+/CD57+, naive T cells are CD3D+/CD28+, B cells are CD19+, and Erythrocytes are HBB+. All gene means are calculated across cells scored by that mixture component, and independently scaled by its max value across all mixture component means (including subpopulations for signaling conditions). (c) Experimental data from cells exposed to GM-CSF, with mixture component centroids denoted in orange. (d) Experimental data from cells exposed to IFNG, with mixture component centroids denoted in light blue. Both test samples (GM-CSF and IFNG) are aligned to the reference population, with specific alignments for reference mixture components that classify as monocytes (overlaid and shown with connected lines). Shifts in the position of the mixture component $\mu's$ reveal how GM-CSF and IFNG exert differential effects on monocyte populations. (e,f) For each mixture component from test samples (x-axis), Jeffrey's divergence against all reference components is plotted on a log10 scale, with lowest divergence values giving the best alignment. Colored symbols correspond to alignments displayed in (c) and (d). Black cross in (d) and (f) denotes a high secondary alignment for IFNG mixture 1, suggesting that both PBMC-1 and PBMC-3 converge to the same final population in response to IFNG. (g) Scatterplot of parameter changes $\Delta\mu$ and Δw for all subpopulation alignments across both test samples. Monocytes have above average ($\Delta\mu \geq 0.04$) for both test samples. (h) Heatmap of aligned monocyte mixture component centroids (μ) in terms of gene expression features. IFNG (upper) induces expression of an ER Phagosome feature (blue arrow). GM-CSF alters expression of two features (Chemokine Activity and Lysosome) but does not affect each subpopulation's baseline levels for 3 monocyte specific features (black arrows). (i) Abundances (w) are shown for both sets of aligned subpopulations. (j) Gene expression distributions for all monocyte subpopulations

for specific genes from (h) that are affected by signals (blue arrows): (i) STAT1, a signal transducer (ii) PSME2, a proteasome activator, (iii) CCL2, an inflammatory chemokine, (iv) CSTB, a protease inhibitor, (v) LYZ, an antimicrobial enzyme, and (vi) GRN, a lysosomal protein.

Figure 5. t-SNE analysis obscures the relationships between aligned subpopulations (a-c) tSNE plots of single cell data for three signaling conditions from Figure 4: (a) PBMC, (b) GM-CSF, and (c) IFNG. Mixture component $\mu's$ (disks) for PBMC mixture components 1 and 3 overlaid on plots (b) and (c) to aid interpretation. PopAlign generated alignments for monocyte mixtures from figure (3) are indicated as lines. tSNE was performed on data pooled from all experiments, but experimental conditions are shown individually for clarity. Plot shows that gene expression changes due to signaling can result in relative cluster shifts within the tSNE plots that obscure subpopulation alignment. For example, in (b) the monocyte populations aligned by PopAlign (Figure 4) become separated by a cluster of T-cells (mixture components 3 and 4).

Figure 6. Discovering signatures of disease and treatment in PBMCs from multiple myeloma patients (a-f) Experimental single cell mRNA-seq data from two healthy donors and four multiple myeloma patients (MM1-4) are projected into 16-dimensional gene feature space. 3D plots show single cells in a subset of three gene features that highlight separation between different immune cell types. Mixture model centroids (μ) are indicated as numbered disks. Subpopulations in test samples are aligned to the reference (a) and changes in abundance (Δw) are plotted for (g) B cells, (h) monocytes, i) naive T cells, and j) effector T cells, showing general and patient specific changes. (k) Mean gene expression levels for two markers of myeloid derived suppressor cells (MDSC), CD33 and CD11b, are plotted for all monocyte subpopulations. Error bars denote confidence interval of the mean. (m) $|\Delta\mu|$ for naive T cell and effector T cell populations relative to healthy1. (n) Heatmap of mixture component μ vectors in terms of feature coefficients c_i for aligned naive and effector T cells across samples. MM subpopulations exhibit reduced expression of two features (red font): Leukocyte motility and Cytotoxic Lymphocyte Killing. (o) Distribution of beta-actin (ACTB) expression for all effector T cell subpopulations across samples. Violin shows distribution, and mean is denoted by white circle. (p) Distribution of perforin 1 (PFN1) expression for all effector T cell subpopulations across samples. For single gene plots (k), (o), (p), units are in terms of normalized and transformed gene expression $(\log(\text{TPT}+1))$ where TPT denotes transcripts per thousands.

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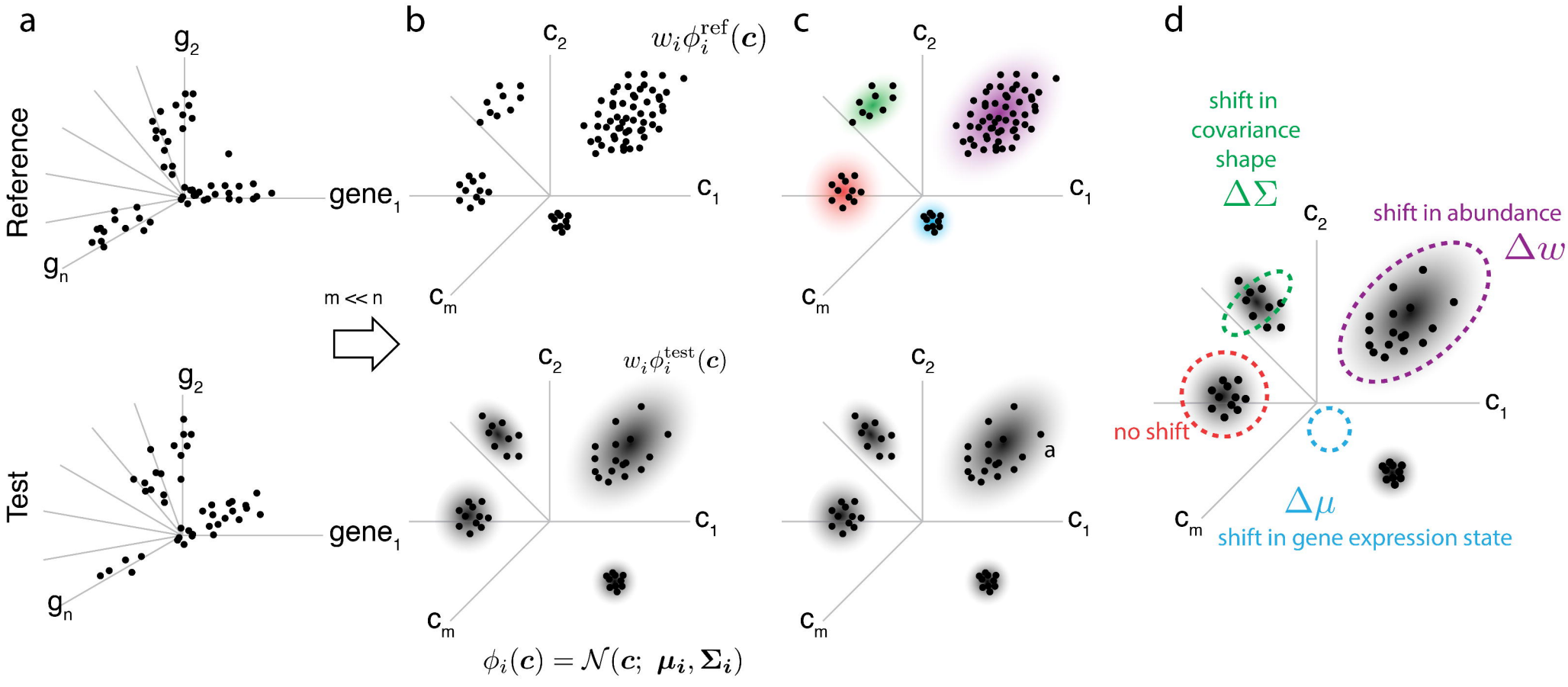
Figure 1



Single-cell Gene Expression Data

Low-dimensional Probabilistic Model

Subpopulation Alignment

Compare parameters between alignments

Figure 2

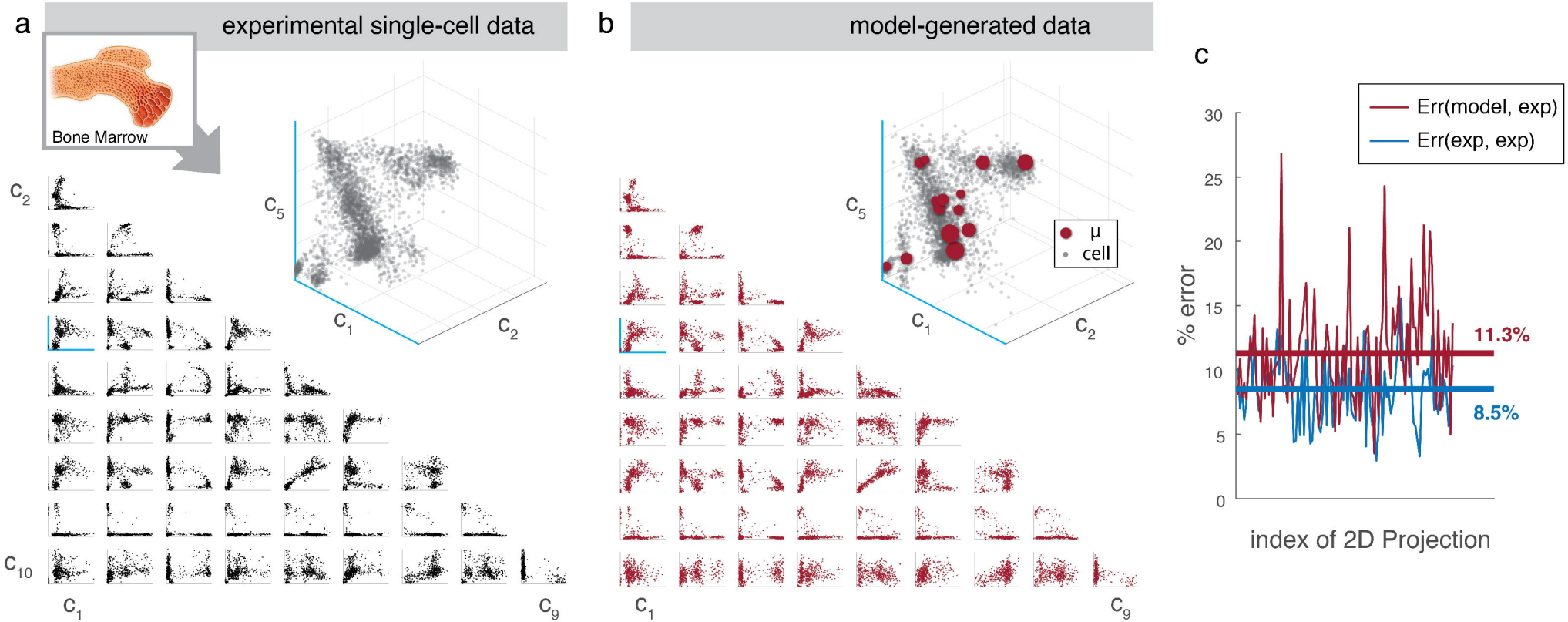

a
experimental single-cell data
Bone Marrow
C_2
C_5
C_1
C_2
C_{10}
C_1
C_9
b
model-generated data
C_5
μ
cell
C_1
C_2
C_1
C_9
c
Err(model, exp)
Err(exp, exp)
% error
30
25
20
15
10
5
0
11.3%
8.5%
index of 2D Projection

Figure 3

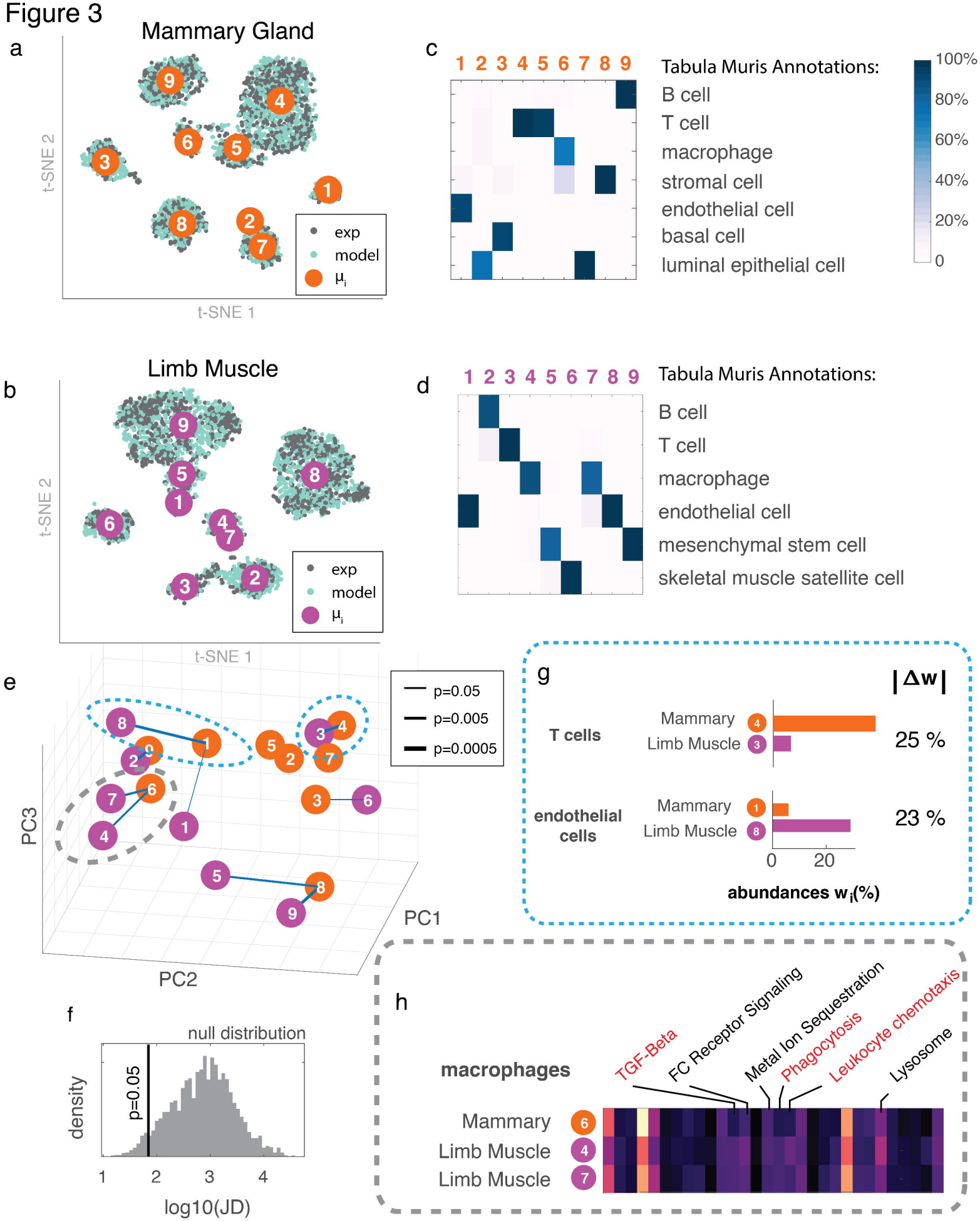

a
Mammary Gland
t-SNE 2
t-SNE 1
exp
model
μ_i
c
1 2 3 4 5 6 7 8 9
Tabula Muris Annotations:
B cell
T cell
macrophage
stromal cell
endothelial cell
basal cell
luminal epithelial cell
100%
80%
60%
40%
20%
0
b
Limb Muscle
t-SNE 2
t-SNE 1
exp
model
μ_i
d
1 2 3 4 5 6 7 8 9
Tabula Muris Annotations:
B cell
T cell
macrophage
endothelial cell
mesenchymal stem cell
skeletal muscle satellite cell
e
PC3
PC2
PC1
p=0.05
p=0.005
p=0.0005
g
|Δw|
T cells
Mammary
Limb Muscle
25 %
endothelial cells
Mammary
Limb Muscle
23 %
0
20
abundances w_i(%)
f
null distribution
density
p=0.05
1
2
3
4
log10(JD)
h
macrophages
TGF-Beta
FC Receptor Signaling
Metal Ion Sequestration
Phagocytosis
Leukocyte chemotaxis
Lysosome
Mammary
Limb Muscle
Limb Muscle

Figure 4

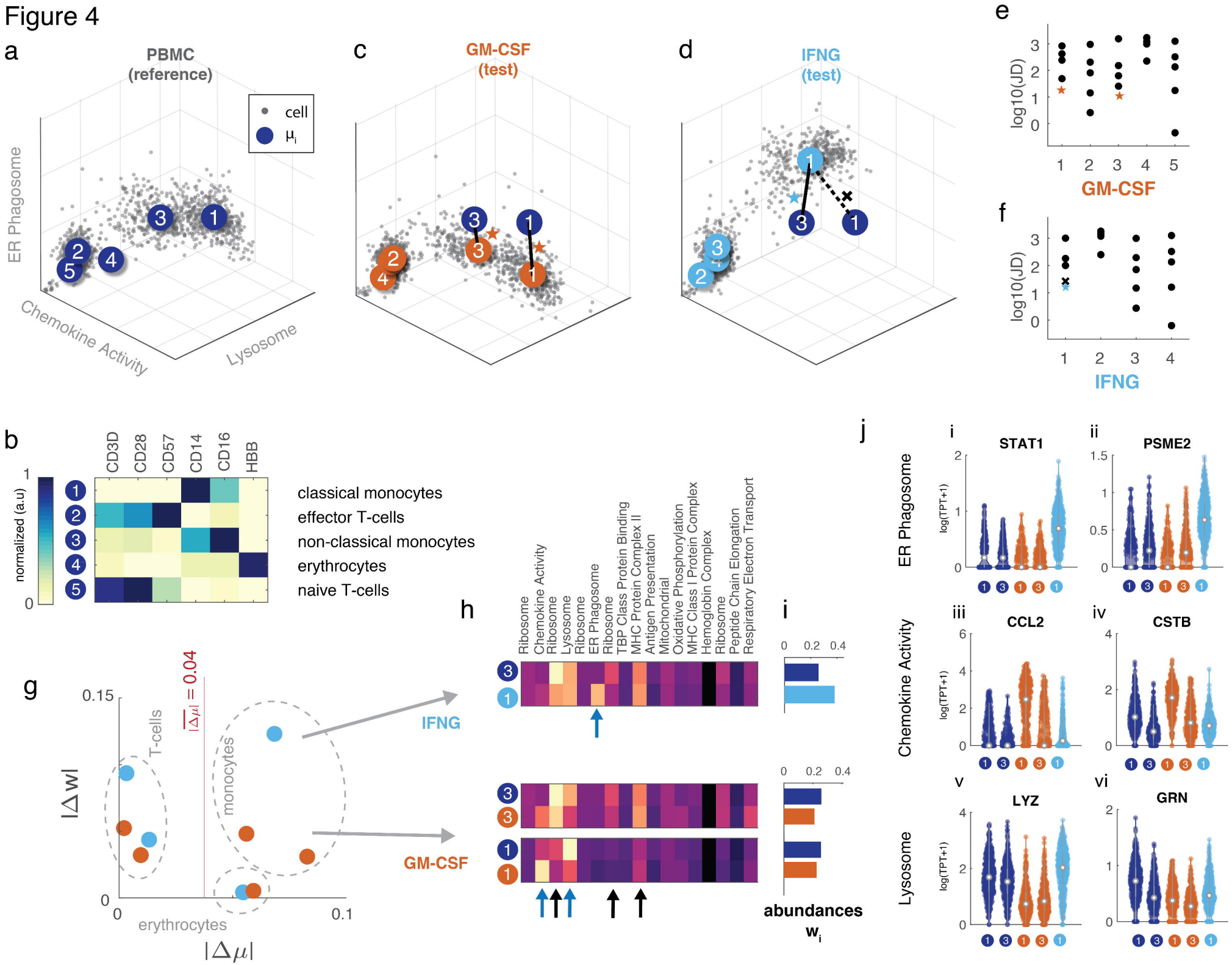

a
PBMC
(reference)
cell
μi
ER Phagosome
Chemokine Activity
Lysosome
c
GM-CSF
(test)
d
IFNG
(test)
e
log10(JD)
GM-CSF
f
log10(JD)
IFNG
b
CD3D
CD28
CD57
CD14
CD16
HBB
normalized (a.u)
classical monocytes
effector T-cells
non-classical monocytes
erythrocytes
naive T-cells
g
|Δw|
|Δμ|
T-cells
monocytes
erythrocytes
|Δμ| = 0.04
IFNG
GM-CSF
h
Ribosome
Chemokine Activity
Ribosome
Lysosome
Ribosome
ER Phagosome
Ribosome
TBP Class Protein Binding
MHC Protein Complex II
Antigen Presentation
Mitochondrial
Oxidative Phosphorylation
MHC Class I Protein Complex
Hemoglobin Complex
Ribosome
Peptide Chain Elongation
Respiratory Electron Transport
i
abundances
wi
j
ER Phagosome
Chemokine Activity
Lysosome
STAT1
PSME2
CCL2
CSTB
LYZ
GRN
log(TPT+1)

Figure 5

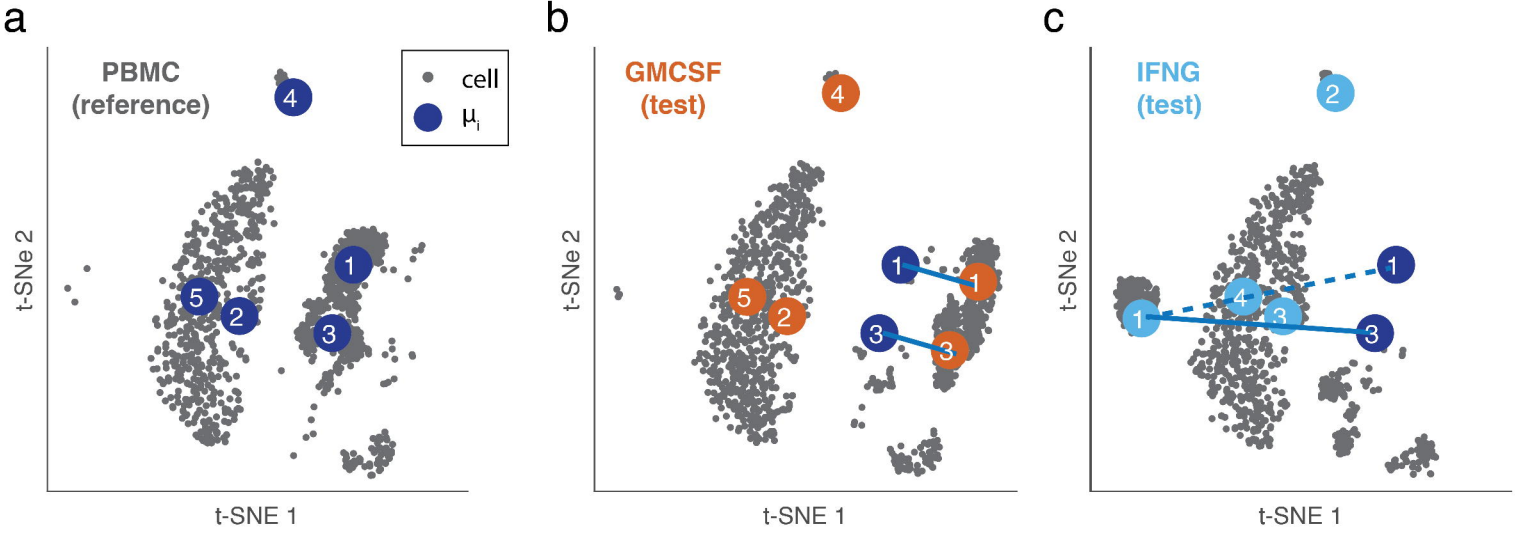
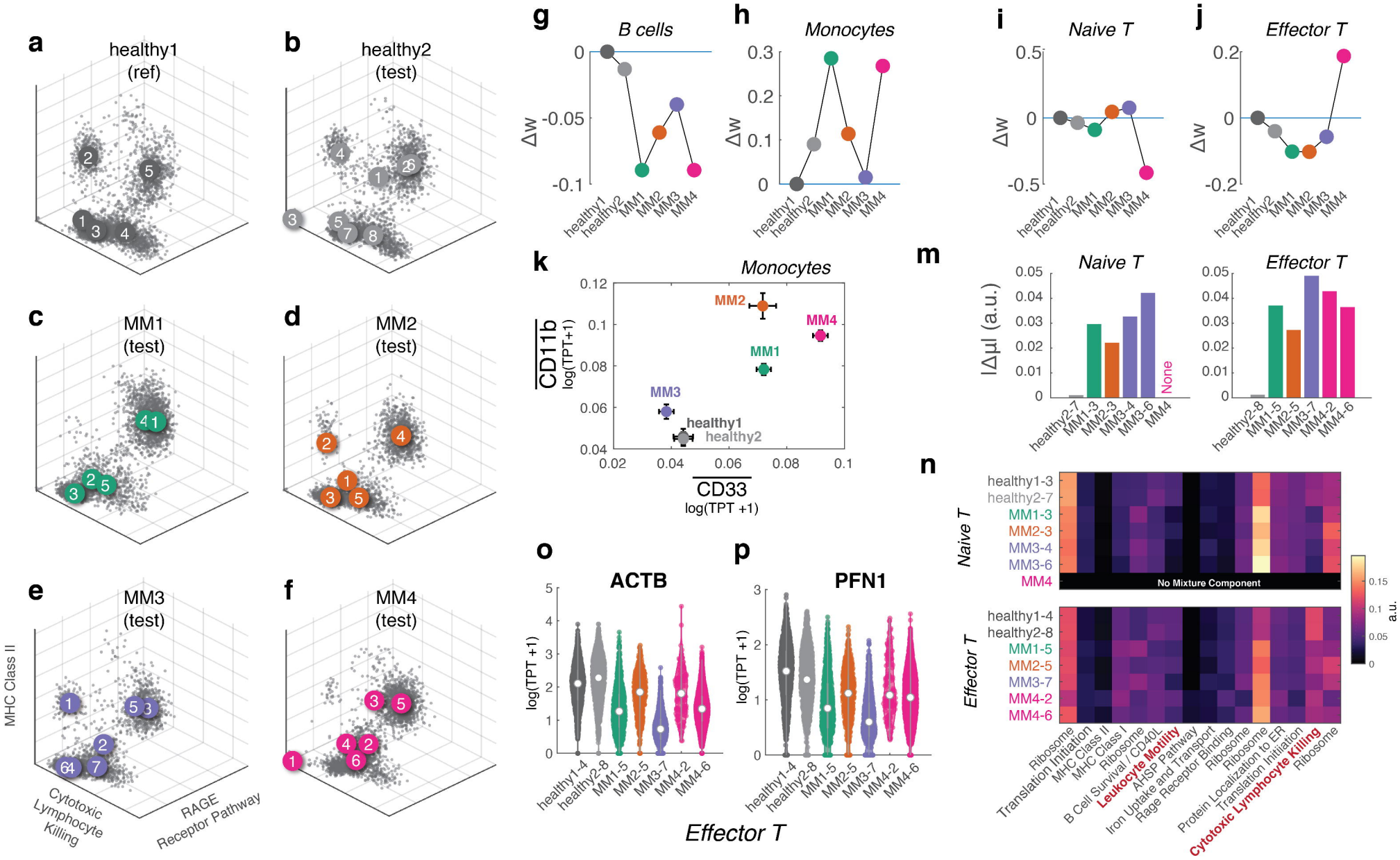

a
PBMC
(reference)
cell
μ_i
t-SNe 2
t-SNE 1
b
GMCSF
(test)
t-SNe 2
t-SNE 1
c
IFNG
(test)
t-SNe 2
t-SNE 1

Figure 6


a
healthy1 (ref)
b
healthy2 (test)
c
MM1 (test)
d
MM2 (test)
e
MM3 (test)
f
MM4 (test)
MHC Class II
Cytotoxic Lymphocyte Killing
RAGE Receptor Pathway
g
B cells
h
Monocytes
i
Naive T
j
Effector T
Δw
healthy1
healthy2
MM1
MM2
MM3
MM4
k
Monocytes
CD11b
log(TPT+1)
CD33
log(TPT +1)
m
Naive T
Effector T
|Δμ| (a.u.)
None
healthy2-7
MM1-3
MM2-3
MM3-4
MM3-6
MM4
healthy2-8
MM1-5
MM2-5
MM3-7
MM4-2
MM4-6
n
Naive T
Effector T
healthy1-3
healthy2-7
MM1-3
MM2-3
MM3-4
MM3-6
MM4
No Mixture Component
healthy1-4
healthy2-8
MM1-5
MM2-5
MM3-7
MM4-2
MM4-6
a.u.
Ribosome
Translation Initiation
MHC Class II
MHC Class I
Ribosome
B Cell Survival / CD40L
Leukocyte Motility
AHSP Pathway
Iron Uptake and Transport
Rage Receptor Binding
Ribosome
Ribosome
Protein Localization to ER
Translation Initiation
Cytotoxic Lymphocyte Killing
Ribosome
o
ACTB
p
PFN1
log(TPT +1)
Effector T