Title:

The neurocognitive architecture of fluid ability in children and adolescents

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Abstract

Fluid ability is the capacity to solve novel problems in the absence of task-specific knowledge, and is highly predictive of outcomes like educational attainment and psychopathology. Here, we modelled the neurocognitive architecture of fluid ability in two cohorts: CALM (N=551, aged 5-17) and NKI-RS (N=335, aged 6-17). We used multivariate Structural Equation Modelling to test a preregistered 'watershed model' of fluid ability. We found that the watershed model fit the data well for both samples: White matter contributed to working memory and processing speed, which, in turn, contributed to fluid ability (R^2(CALM)=51.2%, R^2(NKI-RS)=78.3%). The relationship between cognitive abilities and white matter differed with age and showed a dip in strength around ages 7-12 years. Speculatively, this age-effect may reflect a reorganization of the neurocognitive architecture around pre- and early puberty. Overall, these findings highlight that fluid ability is part of a complex hierarchical system of partially independent effects.

Introduction

Fluid reasoning or fluid intelligence (g_f) is a core part of human cognition and refers to the

capacity to solve novel problems in the absence of task-specific knowledge. It is highly

predictive of a number of important life span outcomes, including educational attainment

(Primi, Ferrão, & Almeida, 2010; Roth et al., 2015), psychopathology (Gale, Batty, Tynelius,

Deary, & Rasmussen, 2010) and mortality (Batty, Deary, & Gottfredson, 2007; Calvin et al.,

2011). However, our understanding of the neurocognitive architecture of g_f remains limited.

Longstanding debates have considered, for instance, how g_f relates to more fundamental

cognitive functions such as working memory and processing speed, and how all of these

cognitive functions relate to brain structure and function (Chuderski, 2013; Ferrer et al.,

2013; Fry & Hale, 2000; Kyllonen & Christal, 1990).

Working memory refers to the ability to hold and manipulate information in the mind short-

term. It has been suggested that working memory is a key determinant of g_f by limiting

mental information processing capacity (Chuderski, 2013; Fukuda, Vogel, Mayr, & Awh,

2010). Proponents of this working memory account of g_f cite high correlations between the

two domains ranging from 0.5 to 0.9 in meta-analyses (Ackerman, Beier, & Boyle, 2005;

Oberauer, Schulze, Wilhelm, & Süß, 2005). Such high correlations have led some to suggest

that g_f and working memory are, in fact, isomorphic (Kyllonen & Christal, 1990). However,

more recent work has highlighted that this isomorphism only arises under conditions of high

time constraints for g_f tasks (Chuderski, 2013). This suggests that g_f and working memory are,

in fact, separable constructs and underlines the importance of processing speed for g_f.

Processing speed, the speed of mental computations, is thought to be rate-limiting to g_f and

is therefore often proposed to be a particularly good predictor of g_f (Ferrer et al., 2013; Kail,

Lervåg, & Hulme, 2015; Kail & Salthouse, 1994; Salthouse, 1996; Schubert, Hagemann, &

Frischkorn, 2017). Proponents of the processing speed account cite moderate but robust

correlations between g_f and processing speed of 0.2 in meta-analyses (Sheppard & Vernon,

2008) as well as longitudinal evidence, particularly from ageing populations (Coyle, Pillow,

Snyder, & Kochunov, 2011; Finkel, Reynolds, McArdle, & Pedersen, 2005; Kail et al., 2015).

Salthouse (1996) argued in the context of cognitive aging, that processing speed determines

high-level cognitive performance because slow processing means that relevant sub-

operations cannot be completed in a set amount of time or are not available for successful

integration. A complementary explanation of individual differences in g_f proposes that

processing speed may be a direct reflection of fundamental neuroarchitectonic properties of

the brain, such as myelination or white matter microstructure (Chevalier et al., 2015; Lu et

al., 2011).

White matter shows protracted development throughout childhood and adolescence, which

extends into the third decade of life (Mills et al., 2016). White matter tracts can be

characterised *in vivo* using diffusion-tensor imaging (DTI), which is sensitive, but not

necessarily specific, to white matter microstructural properties such as myelination or axonal

density (Jones, Knösche, & Turner, 2013; Wandell, 2016). Fractional anisotropy (FA) is the

most commonly investigated DTI measure and quantifies the directionality of water diffusion

in different white matter tracts (Pfefferbaum et al., 2000; Wandell, 2016). Working memory,

processing speed and g_f have each been linked to individual differences in FA in several white

matter tracts (Bathelt, Johnson, Zhang, & Astle, 2018; Kievit, Davis, Griffiths, Correia,

CamCAN, et al., 2016; Vestergaard et al., 2011). Studies of FA have shown that g_f is

associated with FA in the cingulum (Bathelt et al., 2018), that working memory is linked to

the superior longitudinal fasciculus (Vestergaard et al., 2011) and cingulum (Bathelt et al.,

2018), while processing speed has been associated mainly with the forceps minor (Kievit,

Davis, Griffiths, Correia, CamCAN, et al., 2016), the anterior thalamic radiation, and the

uncinate fasciculus (MacPherson et al., 2017).

The literature thus shows strong bivariate links between g_f, working memory, processing

speed and white matter. There are, however, several methodological and practical issues,

which we seek to address here: First, there is limited systematic evidence on the concurrent

relationships between g_f, working memory, processing speed and white matter. This leaves

the relative contributions of processing speed and working memory to g_f unclear, which, in

turn, poses challenges for the design of effective interventions for populations with learning

problems and disabilities. Second, studies usually use a single task as an indicator (i.e.

observed variable) as a proxy for complex and abstract constructs (i.e. latent variables) such

as processing speed, working memory, and g_f. This raises questions about the generalizability

of findings (Noack, Lövdén, & Schmiedek, 2014). Third, our understanding of how the

relationships between relevant cognitive domains and between brain and cognition change

with age remains limited (Achterberg, Peper, van Duijvenvoorde, Mandl, & Crone, 2016;

Tamnes, Roalf, Goddings, & Lebel, 2017). Influential work by Garrett (1946) suggested that

relationships between *cognitive abilities* may change with age such that cognitive abilities

differentiate (i.e. decrease in covariance) over development (Garrett, 1946; Li et al., 2004),

or, conversely, de-differentiate (i.e. increase in covariance) with age (Blum & Holling, 2017),

while others still found no changes in covariance with age at all (de Mooij, Henson, Waldorp,

& Kievit, 2018; Molenaar, Kő, Rózsa, & Mészáros, 2017; Tucker-Drob, 2009). Age-related

changes in the relationship between *brain and cognition* is even less well understood.

Johnson (Johnson, 2000, 2001, 2011) proposed the interactive specialization theory, which

holds that brain maturation involves activity-dependent reorganization and specialization of

networks. Critically, interactive specialization predicts that the same cognitive function may

be represented by different structures at different ages. Brain - cognition relationships may

therefore change fundamentally and dynamically with age (Johnson, 2011). At present,

however, there are very few empirical investigations of potential non-monotonic, dynamic

changes in brain-cognition relationship in general and for the neurocognitive architecture of

fluid ability in particular (de Mooij et al., 2018).

To address these issues, we here used structural equation modelling (SEM) to investigate the

multivariate associations between latent variables of g_f, working memory, processing speed,

and observed variables of white matter microstructure (FA) and age in two large samples: the

Centre for Attention, Leaning and Memory sample (CALM, N = 551, aged 5 − 17 years

(Holmes, Bryant, & Gathercole, 2018)), which consists of children referred to a clinic for

having problems which attention, learning and memory, and the Enhanced Nathan Kline

Institute − Rockland Sample (NKI-RS, N = 335, aged 6 − 17 years (Nooner et al., 2012)), a

community-ascertained sample.

To investigate the neurocognitive architecture of g_f in a principled way, we used a watershed

model of individual differences. This model was originally developed for understanding

complex psychopathological phenotypes (Cannon & Keller, 2006) and then repurposed for

the investigation of individual differences in g_f in an ageing sample (Kievit, Davis, Griffiths,

Correia, CamCAN, et al., 2016). Based on the metaphor of a watershed, the model predicts a

hierarchical many-to-one mapping of partially independent effects such that upstream

tributaries (e.g. brain structure) contribute to intermediate cognitive phenomena (cognitive

endophenotypes), which then contribute to downstream, complex cognitive phenomena

(Cannon & Keller, 2006; Kievit, Davis, Griffiths, Correia, CamCAN, et al., 2016). Kievit et al.

(2016), for instance, showed that individual differences in white matter anatomy predicted individual differences in processing speed, which in turn predicted fluid intelligence scores and a cohort of participants aged 18 - 87 years. Here, we extend the watershed model to include working memory as well as processing speed as cognitive endophenotypes and tested this model in two heterogeneous developmental cohorts. See Figure 1 for a representation of the model. Using structural equation modeling trees (Brandmaier, von Oertzen, McArdle, & Lindenberger, 2013), a novel, decision-tree-based exploratory technique, we also investigated whether the associations in the watershed model change with age.

Based on the watershed model we made the following preregistered predictions (http://aspredicted.org/blind.php?x=u5pf6z):

1.   Working memory, g_f and processing speed are separable constructs.

2.   Individual differences in g_f are predicted by working memory and processing speed.

3.   White matter microstructure is a multi-dimensional construct.

4.   There is a hierarchical relationship between white matter microstructure, cognitive endophenotypes (working memory and processing speed) and g_f, such that white matter contributes to working memory and processing speed, which, in turn contribute to g_f.

5.   The contribution of working memory and processing speed to g_f changes with age.

Figure 1. The Watershed Model. Schematic representation of the watershed model developed by Cannon & Keller (2006) and adapted for the present study. Fluid ability is hypothesized to be the downstream product of working memory and processing speed, which are, in turn, the product of white matter contributions. Figure adapted from Kievit et al. (2016).

Results

Measurement Models of Cognition and White Matter

Cognition. To examine the neurocognitive architecture of g_f, we built up the watershed model in steps, starting with the measurement model of g_f, working memory and processing speed. We first examined whether these three cognitive skills are best described as three separate, two separate or a single cognitive factor (Figure 2).

Figure 2. Different Measurement Models of Cognition. Abbreviations: WM: working memory, PS: processing speed

The Single-Factor Model (Figure 2) showed acceptable absolute fit (Table 1) but poor comparative fit: The Akaike weight (AIC_{weight}) of the Single-Factor Model was approximately 0% for both samples, indicating an extremely low probability that the Single-Factor Model was the data-generating model compared to more complex alternatives. Two-Factor Model A (Figure 2, treating processing speed and g_f as a unitary factor) showed good absolute fit (Table 1) but also approximately 0% probability of being data-generating model. Two-Factor Model B (Figure 2, treating working memory and g_f as a unitary factor) fit the data well (Table 1). Two-Factor Model B showed low probability of being the data-generating model for NKI-RS (AIC_{weight} = 3.40%) but a high probability of being the data-generating model for CALM (AIC_{weight} = 72.85%). The Three-Factor Model (Figure 2) showed 96.60% probability of being data-generating model for NKI-RS and 27.15% probability for CALM.

    Table 1. Model Fit of Competing Measurement Models

	Single-Factor Model	Two-Factor Model A	Two-Factor Model B	Three-Factor Model
		CALM		
	$\chi^2(20) = 70.28$, $p < .001$	$\chi^2(19) = 67.99$, $p < .001$	$\chi^2(19) = 41.66$, $p = .002$	$\chi^2(18) = 41.74$, $p = .001$
	RMSEA = .068	RMSEA = .068	RMSEA = .047	RMSEA = .049
	[.051-.085]	[.052 - .086]	[.027 - .066]	[.030 - .068]
	CFI = .963	CFI = .964	CFI = .983	CFI = .983
	SRMR = .047	SRMR = .043	SRMR = .032	SRMR = .032
	AIC = 9697.18	AIC = 9696.44	AIC = 9668.58	AIC = 9670.55
	BIC = 9800.66	BIC = 9804.24	BIC = 9776.37	BIC = 9782.66
		NKI-R		
	$\chi^2(35) = 109.96$, $p < .001$	$\chi^2(34) = 108.15$, $p < .001$	$\chi^2(34) = 64.85$, $p = .001$	$\chi^2(32) = 54.15$, $p = .009$
	RMSEA = .080	RMSEA = .081	RMSEA = .052	RMSEA = .045
	[.064 - .097]	[.064 - .098]	[.033 - .071]	[.024 - .065]
	CFI = .936	CFI = .936	CFI = .974	CFI = .981
	SRMR = .045	SRMR = .044	SRMR = .035	SRMR = .030
	AIC = 7155.64	AIC = 7155.74	AIC = 7109.43	AIC = 7102.74
	BIC = 7270.07	BIC = 7273.98	BIC = 7227.67	BIC = 7228.60

Note. See Figure 2 for the configuration of different models. Abbreviations: root mean square error of
approximation (RMSEA), comparative fit index (CFI), standardized root mean square residual (SRMR), Akaike
Information Criterion (AIC), Bayesian Information Criterion (BIC)

Even though working memory, processing speed and g_f were highly correlated in both

samples (Table 2), this pattern of results indicates that processing speed formed a clearly

separable factor from working memory and g_f in both samples. Working memory and g_f were

also clearly separable in NKI-RS, but not in CALM. To facilitate comparison across samples and

in accordance with our preregistered analysis plan we used the Three-Factor Model (Table 2,

Supplementary Table 1) in all subsequent analyses.

Table 2. Covariance between Cognitive Measures in the Three-Factor Model

Sample	Path	Standardized Estimate
CALM	g_f <-> memory	0.71, $z = 28.42$, $p < .001$
	g_f <-> speed	0.55, $z = 12.20$, $p < .001$
	memory <-> speed	0.79, $z = 19.35$, $p < .001$
NKI-RS	g_f <-> memory	0.91, $z = 19.51$, $p < .001$
	g_f <-> speed	0.81, $z = 24.73$, $p < .001$
	memory <-> speed	0.87, $z = 17.43$, $p < .001$

Note. See Supplementary Table 1 for factor loadings.

White matter. We examined whether white matter microstructure could be adequately

captured by a single factor by examining model fit. The single-factor model of white matter

microstructure did not fit the data well (CALM: $\chi^2(35)$ = 124.63, p < .001; RMSEA = .125 [.103

- .147]; CFI = .933; SRMR = .039; NKI-RS: $\chi^2(35)$ = 132.33, p < .001; RMSEA = .204 [.167 - .242];

CFI = .885; SRMR = .023). This indicates that white matter microstructure could not be

reduced to a single 'global FA' dimension in our samples, in line with Lövdén et al. (2013) and

Kievit et al. (2016). We therefore modelled each of the ten white matter tracts separately in

all subsequent models.

The Relationship between Working Memory, Processing Speed and g_f

Using Multiple Indicators, Multiple Causes (MIMIC) models (Jöreskog & Goldberger, 1975;

Kievit et al., 2012), we examined the relationship between proposed cognitive

endophenotypes (processing speed and working memory) and g_f. The MIMIC model fit well

for both samples (CALM: $\chi^2(18)$ = 41.74, p = .001; RMSEA = .049 [.030 - .068]; CFI = .983;

SRMR = .032, NKI-RS: $\chi^2(32)$ = 54.15, p = .009; RMSEA = .045 [.024 - .065]; CFI = .981; SRMR =

.030). To understand whether working memory and processing speed made significant joint

and independent contributions to g_f, we compared a freely-estimated MIMIC model to one in

which regression paths from processing speed and working memory to g_f were constrained

to zero. The freely-estimated MIMIC model fit better than a model in which these regression

paths were constrained to zero (CALM: $\Delta\chi^2(2)$ = 250.20, p < .001; NKI-RS: $\Delta\chi^2(2)$ = 199.67, p <

.001). This indicates that working memory and processing speed jointly contributed to g_f. In

CALM ($\Delta\chi^2(1)$ = 15.53, p < .001), but not NKI-RS ($\Delta\chi^2(1)$ = 3.25, p = .072), the freely-estimated

MIMIC model also fit better than a model in which paths from processing speed and working

memory to g_f were constrained to be equal. Working memory showed a greater effect (as

indicated by standardized path estimates) than processing speed in both samples, but the

difference was more pronounced in CALM (Table 3). Contrary to our prediction, therefore,

processing speed seemed to make a negligible contribution to g_f above and beyond working

memory.

Table 3. Regression Paths in the MIMIC Model.

Sample	Path	Standardized Estimate
CALM	speed -> g_f	-0.01, z = -0.16, p = .876
	memory -> g_f	0.72, z = 7.65, p < .001
NKI-RS	speed -> g_f	0.06, z = 0.21, p = .208
	memory -> g_f	0.86, z = 1.81, p = .070

The Watershed Model: Relationships between Cognition and White Matter

Next, we built the full watershed model including white matter, working memory, processing

speed and g_f. We examined path estimates and alternative configurations of the model to

assess whether the main prediction from the watershed model holds, namely, that white

matter contributes to working memory capacity and processing speed, which, in turn,

contribute to g_f.

We found largely converging results across samples. The watershed model showed good fit in

CALM (χ^2(78) = 107.78, p = .014; RMSEA = .026 [.012 - .038]; CFI = .981; SRMR = .043) and

acceptable fit in NKI-RS (χ^2(112) = 219.22, p < .001; RMSEA = .053 [.043 - .064]; CFI = .928;

SRMR = .088). White matter explained large amounts of variance in working memory (R^2_{CALM}

= 32.3%; R^2_{NKI-RS} = 46.1%) and processing speed (R^2_{CALM} = 38.2%; R^2_{NKI-RS} = 54.4%), which, in

turn, explained even more variance in g_f (R^2_{CALM} = 51.2%; R^2_{NKI-RS} = 78.3%).

Comparing the freely estimated watershed model to alternative, constrained, models

showed that white matter contributed significantly to memory and processing speed.

Specifically, a model in which paths from white matter to processing speed were constrained

to zero fit worse than the freely-estimated model (CALM: $\Delta\chi^2(10) = 50.26$, $p < .001$; NKI-RS:

$\Delta\chi^2(10) = 27.19$, $p = .002$), as did a model in which paths from white matter to working

memory were constrained to zero (CALM: $\Delta\chi^2(10) = 52.26$, p < .001; NKI-RS: $\Delta\chi^2(10) = 25.85$,

$p = .004$).

White matter tracts differed in their relationship to cognition. A model in which paths from

white matter to cognition were constrained to be equal, fit worse than the freely-estimated

model for CALM ($\Delta\chi^2(18) = 47.76$, $p < .001$) and NKI-RS ($\Delta\chi^2(18) = 30.42$, $p = .034$), indicating

that the role of white matter microstructure in supporting working memory and processing

speed differed across tracts. For CALM, only the anterior thalamic radiation contributed

significantly to processing speed while the superior longitudinal fasciculus, forceps major and

cingulum were significantly and positively related to working memory (Figure 3). For NKI-RS,

the superior longitudinal fasciculus was significantly and positively related to processing

speed and working memory (Figure 4). Two tracts showed an unexpected, strongly negative

(< -1), relationship: the forceps minor for CALM and the inferior fronto-occipital fasciculus for

NKI-RS. We found that these negative estimates occurred only when all other brain to

cognition pathways were also estimated: When estimated on their own, path estimates

became positive (forceps minor to working memory: standardized estimate = 0.36, $z = 4.05$, p

< .001; inferior fronto-occipital fasciculus to working memory: standardized estimate = 0.14, z

= 0.859, $p = .390$; inferior fronto-occipital fasciculus to processing speed: standardized

estimate = 0.26, $z = 1.41$, $p = .158$). This pattern of results suggests that the negative

pathways were potentially due to modelling several, highly-correlated paths at the same time

(Jöreskog, 1999).

Figure 3. The Watershed Model in CALM. See Supplementary Table 3 for regression estimates. Residual covariances between white matter tracts were allowed but are not shown for simplicity. Abbreviations: uncinate fasciculus (UNC), superior longitudinal fasciculus (SLF), inferior fronto-occipital fasciculus (IFOF), anterior thalamic radiations (ATR), cerebrospinal tract (CST), forceps major (FMaj), forceps minor (FMin), dorsal cingulate gyrus (CG), ventral cingulate gyrus (CH), inferior longitudinal fasciculus (ILF).

Testing alternative hierarchies in the model showed that the watershed model had a high probability of being the data-generating model for CALM (AIC_{weight} = 49.64%) but not NKI-RS (AIC_{weight} = 0.10%). For NKI-RS, a model with an additional level, such that g_f was regressed on working memory, which was regressed on processing speed, which was then regressed on white matter showed the highest probability (AIC_{weight} = 96.95%). This model showed a low probability for CALM, however (AIC_{weight} = 0.19%). For CALM, reversing the hierarchy of levels of the watershed model such that white matter was regressed on processing speed and working memory, which are then regressed on g_f, showed as similarly high probability as the original watershed model (AIC_{weight} = 49.64%). This was not the case for NKI-RS (AIC_{weight} = 0.10%). A model, in which the hierarchy between cognitive endophenotypes and g_f was

254 inverted, showed comparatively poor fit for CALM (AIC$_{weight}$ = 0.00%) and NKI-RS (AIC$_{weight}$ =

255 2.85%). Similarly, a model in which g_f was directly regressed on white matter, working

256 memory and processing speed, showed no clear advantage over the watershed model for

257 CALM (AIC$_{weight}$ = 0.53%) or NKI-RS (AIC$_{weight}$ = 0.00%).

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Figure 4. The Watershed Model in NKI-RS. See Supplementary Table 4 for regression
estimates. Residual covariances between white matter tracts were allowed but are not
shown for simplicity.

We carried out a series of supplementary and exploratory analyses to examine whether

possible confounders influenced our models. These analyses showed that our findings were

robust to the inclusion of covariates such as scanner motion or socio-economic status. They

were also robust across genders and participants taking or not taking medication. There were

no differences between participants with and without diagnosed disorders for CALM, while

potential differences cannot be ruled out for NKI-RS, likely due to the low number of

diagnosed participants (*N* = 106) (see Supplementary Analyses).

In summary, we found that the watershed model performed well for both cohorts and

explained large amounts of variance in g_f. As predicted by the watershed model, white matter

contributed to working memory and processing speed, which, in turn, contributed to g_f.

Age-Related Differences in the Neurocognitive Architecture of g_f

We investigated whether there were age-related differences in the neurocognitive

architecture of g_f. An examination of cross-sectional age differences in g_f, working memory

and processing speed showed that factor scores increased with age for all three cognitive

phenotypes (Figure 5).

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Figure 5. Cognitive Factor Scores by Age.

Using SEM trees, we investigated whether the associations among cognitive and neural

dimensions varied with age. As per our preregistration, we investigated potential age-

differences in the relationships between g_f, working memory and processing speed. In

additional, exploratory, analyses we investigated potential age differences in significant

watershed paths between white matter and cognitive endophenotypes (Figure 3 and 4).

SEM trees combine SEMs with decision tree methods, whereby separating a dataset into

subgroups (in this case age groups) if SEM parameter estimates of interest differ sufficiently

(Brandmaier et al., 2013). SEM trees allowed us to investigate age as a potential moderator

without imposing a-priori categorical age splits. It is an inherently data-driven technique and

the results obtained from SEM trees depend on parameters passed to the algorithm such as

the minimum number of individuals allowed per group (here $N = 50$) or the number of levels

per branch of the tree allowed, which determines the maximum number of possible

subgroups. To counterbalance the inherently exploratory nature of SEM Trees and maximise

the robustness of our results, we implemented cross-validated model selection and a

conservative, Bonferroni-corrected alpha level of .001. We initially allowed for a maximum

depth of one level per tree branch, which could result in no more than two age groups. This

yielded inconsistent results for CALM and NKI-RS (Supplementary Table 4). To test whether

these inconsistencies were an artefact of allowing for only two groups, we repeated our

analysis and allowed for a maximum depth of two levels per branch, which could result in up

to four age groups. This analysis yielded consistent results between CALM and NKI-RS (Table

4). This pattern of results indicates that the initial parameters of our analysis caused us to

miss relevant age differences.

Table 4. SEM Tree Results for the Watershed Model. The table shows differences in

parameter estimates for paths of interest (as shown in Figure 3 and 4) depending on

participants' age in years.

Path	Est. Before	Age Split 1	Est. Betw.	Age Split 2	Est. Betw.	Age Split 3	Est. After
CALM							
memory <–> speed	0.85	**8.46**	0.97	**9.46**	0.74	-	-
memory –> g_f	0.83	**9.38**	0.42	**10.04**	1.14	**10.88**	0.94
speed –> g_f	0.04	**6.88**	-0.19	**11.21**	0.17	-	-
SLF –> memory	0.67	**7.21**	0.18	**11.21**	0.76	-	-
FMaj –> memory	0.59	**7.71**	0.14	**9.29**	0.33	**11.13**	0.74
CG –> memory [1]	0.64	**6.96**	0.09	**11.04**	0.70	-	-
ATR –> speed	0.96	**7.13**	0.68	**7.96**	0.17	**11.96**	0.65
NKI-RS							
memory <–> speed	0.90	**9.82**	0.48	**14.72**	1.11	-	-
memory –> g_f	1.10	**8.59**	0.59	**12.67**	1.03	-	-
speed –> g_f	0.53	**8.59**	-0.12	**12.96**	0.52	-	-
SLF –> memory	2.15	**8.30**	1.47	**12.15**	1.93	-	-
SLF –> speed	3.12	**8.63**	1.83	**15.09**	2.31	-	-

As shown in Table 4, we observed age-related shifts in brain-behaviour associations

throughout childhood and adolescence. For both samples and all but one path, there was an

initially strong relationship between variables, then a dip around ages 7 - 9 for CALM and age

8 for NKI-RS, followed by an increase in path strength around ages 11 to 12. Speculatively,

this pattern of results is consistent with an interpretation of a reorganization of

neurocognitive faculties in late childhood, followed by a consolidation of neurocognitive

pathways around the onset of adolescence.

Discussion

Here we used multivariate statistical techniques to investigate the neurocognitive

architecture of g_f in two large developmental cohorts and, for the first time, investigated how

the neurocognitive architecture of g_f changes dynamically with age. We tested a

preregistered watershed model of g_f, which predicts a hierarchy of partially independent

effects. We found that the watershed model performed well for both the atypical (CALM)

and typical cohort (NKI-RS): White matter contributed to working memory and processing

speed, which, in turn, contributed to g_f and explained 51% of variance therein for the CALM

sample and 78% of variance for the NKI-RS. The watershed model performed well against

most alternative models tested. Models were robust across genders, participants taking or

not taking medication and when controlling for socio-economic status and scanner motion.

Investigations of age effects showed that the relationship between cognitive abilities and

white matter dipped in strength around ages 7 - 12. Speculatively, this age-effect may reflect

a reorganization of the neurocognitive architecture during pre-puberty and early puberty

(Byrne et al., 2017). These findings have implications for understanding and targeting

cognitive impairments in populations with learning difficulties.

Our watershed model consisted of three levels: g_f was conceptualized as the most down-

stream point, with working memory and processing speed as intermediate tributaries, and

white matter microstructural tracts as upstream sources. Previous studies suggested that

matter microstructure is best characterised by a single, 'global FA' factor (Penke et al., 2010)

while others have contended that association patterns among different white matter tracts

are more complex (Kievit, Davis, Griffiths, Correia, CamCAN, et al., 2016; Lövdén et al., 2013).

Here we found strong evidence for a multifactorial view of white matter tracts – for both

samples, a unidimensional model of white matter fit poorly, and for CALM multiple tracts also

showed partially independent contributions to distal cognitive outcomes. In line with

previous research (Bathelt et al., 2018; Kievit, Davis, Griffiths, Correia, Cam-CAN, et al., 2016;

MacPherson et al., 2017), we found that the anterior thalamic radiation was related to

processing speed, as were the forceps major, forceps minor and the cingulum to working

memory. Of note, the superior longitudinal fasciculus was consistently associated with

working memory in both samples. For NKI-RS, the superior longitudinal fasciculus was also

associated with processing speed. The superior longitudinal fasciculus is a large, bilateral

association fibre connecting temporal, occipital, parietal and frontal regions (Kamali,

Flanders, Brody, Hunter, & Hasan, 2014). It may therefore be particularly suited to

contributing to cognitive processes that rely on general multiple-demand systems such as g_f

(Parlatini et al., 2017).

The cognitive levels of the watershed model highlighted a close relationship between

working memory and g_f. Previous studies had variably suggested that g_f and working memory

(Fukuda et al., 2010; Kyllonen & Christal, 1990), or g_f and processing speed (Ferrer et al.,

2013; Kail et al., 2015; Kail & Salthouse, 1994; Salthouse, 1996; Schubert et al., 2017) may be

most closely related (but see Chuderski, 2013). We found that all three cognitive factors were

highly correlated for both samples. Nonetheless, processing speed formed a cognitive factor

clearly separable from working memory and g_f. Working memory and g_f, in turn, were

separable in NKI-RS but not in the CALM sample. This close relationship between g_f and

working memory was also evident in other models where processing speed and working

memory were used as predictors of g_f. Working memory showed a stronger relationship to g_f

than processing speed. Contrary to the predictions of the watershed model, processing

speed became non-significant after controlling for working memory for CALM. This indicates

that working memory was more strongly related to g_f than processing speed and supports

the notion that mental information processing capacity is a key determinant of g_f (Fukuda et

al., 2010; Kyllonen & Christal, 1990).

The associations in the watershed model differed between ages in a complex, non-monotonic

fashion. Previous research had suggested either a decrease in covariance among cognitive

domains with age (age differentiation) (Garrett, 1946; Li et al., 2004), an increase in

covariance with age (age de-differentiation) (Blum & Holling, 2017), or no changes with age

(de Mooij et al., 2018; Molenaar et al., 2017; Tucker-Drob, 2009). These investigations have

traditionally focussed on relations between cognitive domains, however, not on relationships

between brain and cognition (although see de Mooij et al., 2018 for a recent example in an

aging cohort). Possible linear and non-linear changes in brain-behaviour mapping with age

have remained mostly unexplored (Tamnes et al., 2017). Using structural equation modelling

trees, a novel decision-tree-based exploratory technique, we here found evidence of complex

developmental differences that were consistent across samples and relationships in the

watershed model: Initially strong path estimates showed a pronounced decrease in strength

around ages 7 - 9, followed by a renewed increase in the strength, even surpassing initial

levels, around ages 10 - 15. In line with the interactive specialization theory (Johnson, 2000,

2001, 2011), we speculate that this dip may reflect a remapping of the relationships between

brain substrates and cognitive abilities during late childhood and early adolescence, manifesting

in weaker covariance between cognitive domains and between brain and cognition.

On a physiological level, this effect could be driven by neuroendocrine changes during pre-

and early puberty. Puberty is driven by a complex and only partially understood set of

hormonal events including gonadarche and andrenarche (Blakemore, Burnett, S., & Dahl, R.

E., 2010; Sisk & Zehr, 2005). Gonadarche begins with the secretion of gonadotropin-releasing

hormone from the hypothalamus around ages 10 - 11 and closely tracks the overt bodily

changes of puberty (Dorn, 2006). Andrenarche, beginning with the maturation of the

andrenal gland, starts as early as six years of age, and is increasingly recognized as a

complimentary driver of puberty and brain development (Byrne et al., 2017). It is possible

that the hormonal changes of andrenarche and early gonadarche may lead to a level of

neural reorganization, which may initially weaken relationships in the watershed model. The

sweeping bodily, social and cognitive changes happening in early adolescence may then drive

a consolidation of the neurocognitive architecture of g_f.

On a more general level, this result highlights the existence of potential non-linear changes in

brain-behaviour mapping during childhood and adolescence and underlines the value of

modern statistical approaches, such as SEM Trees, for the study of age-related differences.

Nonetheless, it is worth noting that this exploratory result will need to be replicated in future

confirmatory studies with fine-grained data on puberty and larger sample sizes. The latter will

also allow for detailed investigations of potential gender differences.

Testing our model in two different samples allowed us to address several critical questions:

First, participants from both samples completed a small set of common and a larger set of

different cognitive tasks. Therefore, the results obtained here are likely not task-specific, but

rather can be expected to generalize to the domains of working memory, processing speed

and g_f (Noack et al., 2014). Second, by comparing an atypical (CALM) and a typical (NKI-RS)

sample, we demonstrated that the watershed model performed well for different

populations. We are not able to make general claims about potential differences between

typical and atypical populations, however. CALM and NKI-RS were collected in countries with

different socio-economic conditions (United Kingdom and United States of America) and

were of different sample sizes. It will therefore be necessary to replicate these findings in

large typical and atypical cohorts collected in the same setting. Another limitation of our

study is that our samples were cross-sectional, and not longitudinal. This means that were

able to investigate individual differences in g_f, but we could not assess intra-individual

changes during childhood and adolescence. Although the relatively narrow age range makes

large cohort effects unlikely, it may still be that there were differences in recruitment and

selection that varied across the age range. As such, the cross-sectional nature of our samples

limits our ability to make inferences about developmental dynamics and causality. The

watershed model predicts that white matter changes would precede changes in cognitive

endophenotypes, which then precede changes in g_f, but this remains to be tested by future

longitudinal studies such as ABCD (Volkow et al., 2018).

The findings from our study have implications understanding and targeting cognitive

impairments in populations with learning difficulties. First, the close relationship between

working memory and g_f found here, and in other studies (Fukuda et al., 2010; Kyllonen &

Christal, 1990), indicates that children and adolescents struggling with reasoning tasks in

school are likely to also struggle in terms of working memory. Either reducing working

memory load or training working memory and fluid ability capacity in such populations may

therefore be promising lines of inquiry for intervention studies. It is worth highlighting,

however, that cognitive training studies have so far shown little generalizability: Training

abstract reasoning, a common measure of g_f, has not resulted in robust increases in working

memory (Knoll et al., 2016) and working memory training has not been shown to transfer to

reasoning skills or school performance (Dunning, Holmes, & Gathercole, 2013; Schwaighofer,

Fischer, & Bühner, 2015). Similarly, transfer from processing speed to reasoning seems to be

limited (Mackey, Hill, Stone, & Bunge, 2011). The results obtained here suggest that

interventions may increase their chance of success by implementing programs of sufficient

complexity to affect the entire neurocognitive architecture of effects (see also Kievit et al.,

2016). The level of intensity required to produce sustained benefits may also need to be

similarly demanding as education itself, which shows robust effects in increasing general

cognitive abilities over time (Ritchie & Tucker-Drob, 2018). This work and work by others

(Noack et al., 2014) also highlights the value of assessing, modeling and potentially

intervening on multiple tasks, rather than relying on a single task to capture complex

cognitive domains such as g_f. Finally, the age-related differences in the relationships of the

watershed model observed here also highlight that some interventions may work best at

particular developmental phases.

Materials and Methods

Samples

We analysed data from the CALM and NKI-RS sample, as described in detail by (Holmes et al.,

2018) and (Nooner et al., 2012) respectively. We had also preregistered to analyse data from

the ABCD cohort (Volkow et al., 2018). The latter cohort contains only data for 9 - and 10 -

year olds at present, however, which limits comparability to CALM and NKI-RS, and makes it

unsuitable for investigations of developmental differences. We therefore opted to not

analyse ABCD data here and instead recommend a replication of the analyses presented here

in ABCD once longitudinal data is available. The CALM sample consists of children and

adolescents referred by health and educational professionals as having difficulties in

attention, learning and/or memory. The NKI-RS is a community-ascertained, lifespan sample,

and representative of the general population of Rockland, New York, and the United States as

a whole, in terms of ethnicity, socioeconomic status etc. For NKI-RS, we included data for

participants under the age of 18 only to match the age range of CALM and excluded data that

were completed more than half a year after enrolment. The latter was implemented to

ensure that age at assessment did not differ substantively between cognitive measures. The

final samples included 551 participants from CALM (30.85% female, aged 5.17 − 17.92,

$N_{Neuroimaging}$ = 165) and 335 participants from NKI-RS (43.48% female, aged 6.06 − 17.92,

$N_{Neuroimaging}$ = 67). See Table 5 for prevalence of relevant disorders and problems in the

samples.

Table 5. Prevalence of Relevant Disorders and Problems

Variable	Percentage CALM	Percentage NKI-RS
ADHD	31.94	17.01
Dyslexia	5.81	5.67
Autism	6.72	0.60
Mood disorder	0.54	0.90
Anxiety disorder	2.36	18.21
Medicated[1]	10.53	17.01
Speech/language problems	38.11	19.40

Note. [1] unspecified medication for NKI-RS, ADHD-medication for CALM

Cognitive Tasks

We included cognitive tasks measuring the domains of g_f, working memory or processing

speed for CALM and NKI-RS. See Table 6 for the complete list of tasks used and the

Supplementary Material for detailed task descriptions. See Supplementary Figure 1 and 2 for

raw scores on all tasks. The tasks modelled here were preregistered for CALM but not NKI-RS.

White Matter Microstructure

We modelled mean FA for ten tracts of the Johns Hopkins University (JHU) white matter

tractography atlas (Hua et al., 2008) averaged over the hemispheres (Figure 5). See

Supplementary Methods for details of MRI data acquisition, quality control and DTI data

processing.

Table 6. Cognitive Tasks Modelled

	CALM	NKI-RS
Working memory	AWMA Digit Recall (forward digit span)	WISC-R Forward Digit Span
	AWMA Backward Digit Span	WISC-R Backward Digit Span
	AWMA Dot Matrix	-
	AWMA Mr X	-
	-	CNB N-back task
g_f	WASI-II Matrix Reasoning	WASI-II Matrix Reasoning
	-	WASI-II Block Design
	-	WASI-II Similarities
	-	CNB Verbal Reasoning
Processing speed	DKEFS Trail-Making	DKEFS Trail-Making
	PhAB Rapid Naming	-
	TEA-Ch RBBS	-
	-	CNB Motor Speed
	-	CNB Sensory Motor Speed

Note. Abbreviations: AWMA - Automated Working Memory Assessment (Alloway, 2007), CNB - Computerized
Neurocognitive Battery (Gur et al., 2001), DKEF - Delis-Kaplan Executive Functioning System (Delis, Kramer,
Kaplan, & Holdnack, 2004), PhAB – Phonological Assessment Battery (Gallagher & Frederickson, 1995), TEA-Ch
RBBS - Test of Everyday Attention for Children, Red & Blues, Bags & Shoes subscale (Manly et al., 2001), WASI -
Wechsler Abbreviated Scale of Intelligence - Second Edition(Wechsler, 2011), WISC-R - Wechsler Intelligence
Scale for Children - Revised (Kaufman, 1975).

479 Data Processing and Structural Equation Modelling

Covariance matrices and scripts replicating key analyses can be obtained from:

https://github.com/df1234/gf_development. We modelled raw scores for g_f and working

memory tasks, as preregistered. Raw scores on processing speed tasks were transformed,

which was not preregistered. First, we inverted response time scores to obtain more intuitive

measures of speed for all but the CNB Motor Speed task, for which raw scores were already a

measure of speed. To satisfy SEM assumptions of normality, we additionally applied a log-

transformation to reaction time tasks. For the CNB Motor Speed task only, we removed

values ± 2 *SD* of the mean (*N* = 6) because of the presence of outliers.

We modelled the associations between cognition and white matter microstructure using

Confirmatory Factor Analysis and SEM in R (R core team, 2015) using the package lavaan

(Rosseel, 2012). All models were fit using maximum likelihood estimation with robust Huber-

White standard errors and a scaled test statistic. Missing data was addressed using full

information maximum likelihood estimation. Model fit was inspected using the chi-square

test, RMSEA, CFI and SRMR. Good fit was defined as approximately RMSEA < 0.05, CFI > 0.97

and SRMR < 0.05; acceptable fit as approximately RMSEA = 0.08 - 0.05, CFI = 0. 95 - 0.97,

SRMR = 0.05 - 0.10 (Schermelleh-Engel, Moosbrugger, & Müller, 2003). Comparative model

fit for nested models was assessed using the chi-square test. Non-nested models were

compared using the AIC_{weight}, which indicates the probability of a model being the data-

generating model compared to all other models tested (Wagenmakers & Farrell, 2004). Effect

sizes were evaluated by inspecting R^2 and standardized parameter estimates manifest

variables. Absolute standardized parameter estimates above 0.10 were defined as small

effects, 0.20 as typical and 0.30 as large (Gignac & Szodorai, 2016).

We used SEM Tress to investigate whether the associations among cognitive and neural

measures differed with age. SEM Trees use decision tree methods to hierarchically split a

dataset into subgroups if parameter estimates differ significantly based on a covariate of

interest - in this case age (Brandmaier et al., 2013). We first ran the watershed model in

OpenMx (Boker et al., 2011) and then passed this model object to semtree to compute the

SEM Trees. We ran one SEM Tree for each parameter of interest (e.g. the covariance

between working memory and processing speed). All other parameters in each semtree

object were set to be invariant across groups to ensure that splits were specific to the

parameter of interest. We used a 10-fold cross-validation estimation method as recommend

by Brandmaier et al. (2013). For the path from the cingulate to working memory only we

used 5-fold cross-validation because the model did not converge using 10-fold cross-

validation. Minimum sample size in age group was set to $N = 50$ to ensure reliable estimation

of standard errors. Note that this choice effectively limited search space for potential splits to

ages 6.58 - 12.42 years for CALM and 8.08 - 15.49 years for NKI-RS

Figure 5. White Matter Tracts Modelled in the Analyses.

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Conflicting Interests

531   The authors report no conflict of interest.

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