

## Original Article

### **Risky behaviors and Parkinson's disease: A Mendelian randomization study in up to 1 million study participants**

**Running head:** Risky behaviors and Parkinson's disease

Sandeep Grover\*

Institut für Medizinische Biometrie und Statistik

Universität zu Lübeck

Universitätsklinikum Schleswig-Holstein, Campus Lübeck

Lübeck

Germany

[sandeep.grover@imbs.uni-luebeck.de](mailto:sandeep.grover@imbs.uni-luebeck.de)

Fabiola Del Greco M\*

Institute for Biomedicine

Eurac research

Via Galvani 31, Bolzano

Italy

[fabiola.delgreco@eurac.edu](mailto:fabiola.delgreco@eurac.edu)

Meike Kasten

Institute of Neurogenetics

Department of Psychiatry and Psychotherapy

Universität zu Lübeck

Lübeck

Germany

[meike.kasten@neuro.uni-luebeck.de](mailto:meike.kasten@neuro.uni-luebeck.de)

Christine Klein

Institute of Neurogenetics

Universität zu Lübeck

Lübeck

Germany

[christine.klein@neuro.uni-luebeck.de](mailto:christine.klein@neuro.uni-luebeck.de)

Christina M. Lill

Genetic and Molecular Epidemiology Group

Lübeck Interdisciplinary Platform for Genome Analytics (LIGA)

Institutes of Neurogenetics & Cardiogenetics

Universität zu Lübeck

Lübeck

Germany

[christina.lill@uni-luebeck.de](mailto:christina.lill@uni-luebeck.de)

Inke R. König\*\*

Institut für Medizinische Biometrie und Statistik

Universität zu Lübeck

Universitätsklinikum Schleswig-Holstein, Campus Lübeck

Lübeck

Germany

[inke.koenig@imbs.uni-luebeck.de](mailto:inke.koenig@imbs.uni-luebeck.de)

\*\*Corresponding author

Address for correspondence:

Inke R. König

Institut für Medizinische Biometrie und Statistik

Universität zu Lübeck

Universitätsklinikum Schleswig-Holstein, Campus Lübeck

Lübeck

Germany.

Email: [inke.koenig@imbs.uni-luebeck.de](mailto:inke.koenig@imbs.uni-luebeck.de)

\* Contributed equally

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## Abstract

**Objective:** Dopaminergic neurotransmission is known to be a potential modulator of risky behaviors including substance abuse, promiscuity, and gambling. Furthermore, observational studies have shown associations between risky behaviors and Parkinson's disease; however, the causal nature of these associations remains unclear. Thus, in this study, we examine causal associations between risky behavior phenotypes on Parkinson's disease using a Mendelian randomization approach.

**Methods:** We used two-sample Mendelian randomization to generate unconfounded estimates using summary statistics from two independent, large meta-analyses of genome-wide association studies on risk taking behaviors (n=370,771-939,908) and Parkinson's disease (cases: n=9581, controls: n = 33,245). We used inverse variance weighted as the main method for judging causality.

**Results:** Our results support a strong protective association between the tendency to smoke and Parkinson's disease (OR=0.714 per log odds of ever smoking; 95% CI=0.568-0.897; p-value=0.0041; Cochran Q test; p-value=0.238; I<sup>2</sup> index=6.3%). Furthermore, we observed risk association trends between automobile speed propensity as well as the number of sexual partners and Parkinson's disease after removal of overlapping loci with other risky traits (OR=1.986 for each standard deviation increase in normalized automobile speed propensity; 95% CI=1.215-3.243; p-value=0.0066, OR=1.635 for each standard deviation increase in number of sexual partners; 95% CI=1.165-2.293; p-value=0.0049).

**Interpretation:** These findings provide support for a causal relationship between general risk tolerance and Parkinson's disease and may provide new insights in the pathogenic mechanisms leading to the development of Parkinson's disease.

## Introduction

Parkinson's disease (PD) is the second most prevalent neurodegenerative disorder characterized pathologically by progressive loss of dopaminergic neurons in the substantia nigra<sup>1</sup>. The currently available treatment options for PD are symptomatic only. The lack of disease-modifying or protective treatments is at least in part due to the fact that the exact disease mechanisms are currently only partly understood.

The vast majority of PD cases are caused by the combined action and likely interaction of genetic variants as well as environmental and lifestyle exposures<sup>2-6</sup>. Several common habitual agents like smoking, coffee and alcohol drinking have shown protective associations with PD in large scale meta-analyses of observational studies<sup>7</sup>. Several recent studies have further shown beneficial effects of cannabidiol, a non-psychotomimetic compound derived from cannabis on non-motor symptoms in PD patients<sup>8</sup>. It is noteworthy that several impulse control disorders (ICDs) such as gambling, hypersexuality and compulsive eating are observed more frequently in PD patients with some studies reporting up to 40% prevalence of ICDs in PD patients<sup>9</sup>. However, it is believed that these symptoms may be the result of dopamine agonist therapy prescribed to PD patients<sup>10</sup>. The imminent challenge in this context is to decipher whether these PD-associated environmental/lifestyle/behavioral variables contribute to or are an effect of the disease.

The recent development of the Mendelian randomization (MR) approach allows to judge causality based on genetic data generated in observational studies. Specifically, this relies on the utilization of genetic variants as proxy markers of risk factors<sup>11</sup> and takes care of confounding by exploiting random allocation of genetic variants at birth. We have also seen a surge of MR studies in the field of PD specifically exploring the causal role of several circulating biomarkers<sup>12-18</sup>. For example, a recent MR study reported a significant causal association with a lifelong PD risk reduction of 3% per 10 µg/dl increase in serum iron levels<sup>14</sup>. Most recently, another study further reported a risk reduction of 18% with a lifetime exposure of 5kg/m<sup>2</sup> higher BMI<sup>15</sup>.

To date, the majority of studies in the field of PD have focused on modifiable environmental factors only, and MR studies exploring the role of behavioral phenotypes are lacking. Henceforth, our primary aim was to investigate the willingness to take risk as a causal factor in the development of PD. For this, we applied a two-sample MR to investigate whether people with risk taking tendency have an altered risk for PD<sup>19</sup>. A recent GWAS identified 611 independent loci with several measures of risky behaviors including general risk tolerance, adventurousness, and risky behaviors in the driving, drinking, smoking, and sexual domains<sup>20</sup>. We used all reported loci to mimic the random allocation of loci among PD cases and controls in available data from a large, recent GWAS on PD<sup>4,21</sup>. As secondary analyses, we further considered the wider literature to support inferences drawn from our primary analysis using previously reported GWAS on similar habitual behaviors such as smoking phenotypes, coffee consumption, alcoholism, cannabis dependence and gambling<sup>22-27</sup>.

## **Methods**

### **Study design and identification of datasets**

We conducted a two-sample MR using summary level estimates to explore the causal role of several risky behaviors on PD<sup>28</sup>. We identified genetic instruments that influence risky behaviors using a recently published meta-analysis of GWAS datasets on risky behaviors<sup>20</sup>. The study reported statistically significant associations of 611 independent loci ( $p\text{-value} < 5 \times 10^{-8}$ ) in a discovery cohort in up to 939,908 individuals of European ancestry with six highly correlated risky behavior phenotypes including general risk tolerance, adventurousness, automobile speeding propensity, drinks per week, ever versus never smoking and number of sexual partners. The study further defined general risk tolerance as the willingness to take risks, “adventurousness” as the self reported tendency to be adventurous vs. cautious, “automobile speeding propensity” as the tendency to drive faster than the speed limit, “drinks per week” as the average number of alcoholic drinks consumed per week, “ever smoker (tendency to smoke)” as whether one has ever been a smoker, and lastly “number of sexual partners” as the lifetime number of sexual partners.

We further extracted summary estimates of the identified genetic variants from the discovery cohort of a recent meta-analysis of GWAS on 9581 PD cases and 33,245 controls of European ancestry<sup>4</sup>. For this, we used data available on the PDGene database (<http://www.pdgene.org>)<sup>21</sup>. Genetic instruments were identified for smoking (cigarettes per day), smoking initiation, smoking cessation, cannabis dependence, pathological gambling, alcohol and coffee consumption from independent GWAS as a part of our secondary analyses<sup>22-27</sup>.

### **Prioritization of genetic variants and power analysis**

We systematically screened all the identified loci for a possible direct involvement in PD. For this, we used data available via PDGene to extract the list of loci shown to be significantly associated with PD ( $p\text{-value} < 5 \times 10^{-8}$ )<sup>21</sup>. We further checked overlapping loci for a relevant role in the pathogenesis of PD using a literature search. If substantial evidence was found, the respective loci were excluded from the list of genetic variants (SNPs, genetic instrument) of the respective behavioral traits.

The SNPs constituting each genetic instrument were checked for strong linkage disequilibrium (LD). We used the rAggr database to look for correlated variants in individuals with European descent from the 1000 Genomes Phase 3 data (<http://raggr.usc.edu>; date last accessed June 22, 2018) and excluded one of the variants for pairs with  $R^2$  greater than 0.25. Finally, if SNPs were not available in the PD GWAS dataset, we identified proxy SNPs using an  $R^2$  cut-off of 0.9 based on the rAggr database as above.

The strength of the prioritized genetic instrument was judged using F-statistics as explained earlier<sup>14</sup>. We computed the variance in exposures explained by prioritized genetic instruments ( $R^2$ ) of the genetic instruments using effect estimates and the standard error of individual SNPs as described elsewhere<sup>29</sup>. Lastly, power calculations were done using the method described by Brion *et al.*, which is available online (<http://cnsgenomics.com/shiny/mRnd/>)<sup>30</sup>.

## Estimation of causal effects

In cases where the genetic instruments comprised a single SNP, we used the Wald ratio estimate along with the Delta method to obtain the related estimate of the variance. In cases where the genetic instruments consisted of multiple SNPs, we used the inverse variance weighted (IVW) fixed effect method as the main method to estimate the effect of genetically predicted behavioral phenotypes on PD by combining the genetic loci-specific Wald ratio estimates. We specifically employed the IVW method using second order weights because casual estimates generated through this method are expected to provide a more accurate reflection of the variance of the Wald ratio estimate<sup>31</sup>.

However, in the absence of reliable information on functional pathways, proportion and direction of pleiotropic genetic variants, additional MR methods including MR-Egger, Weighted median and Weighted mode-based method were also employed to check the consistency of direction of effect estimates<sup>19, 32-34</sup>. Unlike IVW, which assumes no intercept term in the model, the MR-Egger method provides less biased causal estimates in the presence of directional pleiotropy and considerable heterogeneity assuming absence of measurement error (NOME assumption)<sup>19</sup>. However, the MR-Egger method is more sensitive to unobserved associations of genetic variants with confounders of the exposure-outcome association and requires a greater sample size for the same underlying variance in exposure<sup>33</sup>. Both IVW and MR-Egger methods further assume that the pleiotropic effects of genetic variants are independent of their associations with the exposure known as the InSIDE assumption. In the case of violation, the Weighted median method may provide consistent causal estimates even if up to 50% of genetic variants do not conform to the InSIDE assumption. Also the Weighted mode-based method may provide consistent causal estimates, in particular, even when the NOME assumption was not met, but assuming that the most frequent value of the bias of the Wald ratio estimates is zero.<sup>34</sup>

Within every MR method, we computed casual estimates as odds ratio (OR) for PD per unit log of odds of the categorical behavioral phenotypes or OR per unit standard deviation (SD) of the continuous behavioral phenotypes. And lastly, to address the issue of multiple testing, results were considered statistically significant at the 5% level after a conservative Bonferroni correction of the significance level, therefore if  $p\text{-value} < 8.3 \times 10^{-3}$  ( $0.05/6$  independent primary MR hypotheses).

### **Assessment of pleiotropy**

We used the Cochran Q-statistic and  $I^2$  for the IVW method using second order weights as main methods to identify pleiotropic variants<sup>35</sup>. Furthermore, results from the less powerful MR-Egger's test were also used to explore heterogeneity including the test for deviation of the intercept from the null for MR-Egger's model using the  $\chi^2$ -test for independence<sup>33</sup>. We further used Ruckers Q' statistic to describe heterogeneity around MR-Egger fit<sup>36</sup>. The appropriate use of the main MR method for interpretation of causal estimate in the present study was judged by calculating the ratio between Rucker's Q' and Cochran's Q statistics<sup>37</sup>. As a rule of thumb, the IVW method is recommended as the main method for judging the causal effect if the ratio approaches one.

To evaluate heterogeneity graphically, funnel plots were constructed that plot the spread of the inverse of the standard error of the respective Wald ratio estimates of each individual SNP around the MR estimates. Also, scatter plots of effect estimates of individual SNPs with outcome vs. effect estimates of individual SNPs with exposure are provided as a comparative visual assessment of the causal estimates generated from different MR methods. We further constructed radial MR plots which have been recently suggested as a more suitable approach for visual detection of outliers compared to traditional scatter plots, specifically when the difference between IVW and MR-Egger estimates is large<sup>38</sup>.

### **Sensitivity analyses**

A leave-one-out sensitivity analysis was conducted to check for a disproportionate influence of individual SNPs on overall causal effect estimates using the IVW method. We used forest plots to visually assess the results of the analysis and further identify the outliers.

Since all the behavioral traits are highly correlated and are expected to exhibit shared genetic influence, we conducted a sensitivity analysis by including only genetic loci specific to each individual behavioral trait. We used an  $R^2 > 0.8$  to consider loci to be overlapping with other loci in an independent genetic instrument. Such an approach may help us to judge the reliability of independent

associations of observed phenotypic traits. We further adopted a conservative approach by using loci unique to each phenotypic trait ( $R^2 < 0.01$ ) at the cost of reduced power.

We used PhenoScanner database to identify potential pleiotropic variants by checking significant associations of loci prioritized in the present study with phenotypes from previously published GWAS (<http://phenoscanner.medschl.cam.ac.uk>)<sup>39</sup>. We further checked GWAS listed in the GWAS Catalog (<https://www.ebi.ac.uk/gwas/>) to search for any missed hits. The identified variants were then grouped into categories depending on their association with potential confounder phenotypes and were checked for an influence on the causal estimate using leave-out approach.

We further evaluated the biological relevance of different brain regions in contribution to the overall causal estimate through analysis of gene expression data for the available loci from our different genetic instruments. Gene expression data was extracted from the Genotype-Tissue Expression (GTEx) Project comprising data on a total of 12 different brain regions ([www.gtexportal.org](http://www.gtexportal.org); date last accessed June 22, 2018)<sup>40</sup>. The identified loci were then grouped into categories as per their expression in specific brain regions and checked for an influence on the causal estimate after their exclusion using a leave-out approach.

## Results

### Prioritization of genetic instruments and power analysis

The descriptive statistics of the genetic instruments selected for the MR analyses are presented in **Table 1**. The data used for the analyses are given in the **Supplementary Table**.

Two SNPs from the genes *MAPT* (rs62062288; p-value with PD= $3.1 \times 10^{-21}$ ) and *HLA-DQB* (rs3021058; p-value with PD= $2.7 \times 10^{-3}$ ) were excluded from the genetic instrument for automobile speeding propensity phenotype based on a potential direct involvement in PD. Two additional *MAPT* SNPs were also present in the respective genetic instrument for the phenotypes drinks per week (rs62055546; p-value with PD:  $5.2 \times 10^{-21}$ ) and number of sexual partners (rs62063281; p-value with PD= $1.73 \times 10^{-21}$ ) and were not carried forward to further analyses. One SNP was observed to be in

complete LD with another SNP for the adventurousness phenotype and was excluded. The final number of available SNP data varied from 35 (out of 42) for automobile speeding propensity to 213 (out of 223) for smoking tendency with F-statistics of the pooled genetic instrument ranging from 1406.4 (for automobile speeding propensity) to 9639.6 (for ever vs. never smoking).

Our power analyses suggest that our study has approximately 80% power to detect a true OR of 1.349 or 0.698 for PD per SD of the continuous phenotype assuming that the proportion of the continuous phenotype explained by the genetic instrument is  $\geq 1\%$  at a type-1 error rate of 0.05.

### **Estimation of causal effects and assessment of pleiotropy**

The causal effect estimates using different MR methods are provided in **Table 2**. Using the IVW method, a genetically increased risk of tendency to smoke was associated with a reduced risk of PD per unit increase in log odds of ever smoking (OR: 0.714 per log odds of ever smoking; 95% CI=0.568-0.897; p-value=0.0041). Results from the Weighted median MR analysis showed similar results (OR: 0.707 per log odds of ever smoking ; 95% CI=0.601-0.832). There was minimal evidence of heterogeneity of causal effects between individual variants ( $I^2 = 6.30\%$ ; Cochran's Q test p-value=0.2367), which was confirmed using MR-Egger's Intercept test (p-value=0.6619). Corresponding plots used for the assessment of pleiotropy are shown in **Figure 1**. We did not detect any potential outlier or pleiotropic variant in the association analysis of the tendency to smoke phenotype and PD.

For other risk taking behaviors including general risk tolerance, automobile speeding propensity and the number of sexual partners, we observed a trend towards positive associations (general risk tolerance: OR: 1.620 per log odds of general risk tolerance; 95% CI=1.046-2.511; p-value=0.0311; automobile speeding propensity: OR=2.043 for each SD increase in normalized automobile speed propensity; 95%CI=1.076-3.876; p-value=0.0299; number of sexual partners: OR: 1.473 for each SD increase in the number of sexual partners; 95%CI=1.079-2.010; p-value=0.0152). A triangulation flowchart summarizing the findings of the study in the context of the MR workflow is given in **Figure 2**.

## Sensitivity analyses

### *Leave-one-out analysis*

Based on the forest plots, we observed outlier SNPs for the phenotypes general risk tolerance (rs993137), adventurousness (rs10433500) and drinks per week (rs1229984) (data not shown).

Leaving out outlier SNPs for each of the respective phenotypes did not alter the results substantially (general risk tolerance OR=1.763 per log odds of general risk tolerance; 95% CI=1.133-2.744; p-value=0.0125, adventurousness OR=1.164 per log odds of adventurousness; 95% CI=0.858-1.580; p-value=0.3270, drinks per week OR=1.368 for each SD increase in the number of drinks per week; 95% CI=0.978-1.915; p-value=0.0669).

### *Genetic overlap between risky behaviors*

We identified a reduction in the number of unique SNPs in the genetic instruments for each of the phenotypes using two different LD cut-offs ( $R^2 \leq 0.8$  and  $R^2 \leq 0.01$ ) (**Table 3**). However, irrespective of cut-off criteria, there was no change in the protective association of the tendency to smoke phenotype with PD (number of SNPs in the genetic instrument with  $R^2 \leq 0.8$ : 195, OR=0.713 per log odds of ever smoking; 95% CI=0.557-0.913; p-value=0.0037; number of SNPs in the genetic instrument with  $R^2 \leq 0.01$ : OR=0.719 per log odds of ever smoking; 95% CI=0.547-0.945; p-value=0.0185). Furthermore, consistent with the risk-increasing trend observed for general risk tolerance and the number of sexual partners, we observed a stronger causal association with PD for both phenotypes after reducing the number of SNPs from 117 to 94 ( $R^2 \leq 0.8$ ) for general risk tolerance and 109 to 94 ( $R^2 \leq 0.8$ ) for the number of sexual partners (OR=1.986 per log odds of general risk tolerance; 95% CI=1.215-3.243; p-value= 0.0066, OR=1.635 for each SD increase in number of sexual partners; 95% CI=1.165-2.293 ; p-value=0.0049). The associations persisted using a stringent lower  $R^2$  cut-off of 0.01 (p-value=0.0440 and p-value=0.0484).

### *Genetic variants associated with potential confounders*

We did a comprehensive screening of the the PhenoScanner database for potential associations of genetic loci used in the current study and reported to be associated with other phenotypes. The identified associated phenotypes were then investigated for association with PD based on a thorough literature search. Using this strategy, we identified rheumatoid arthritis, years of educational attainment, adiposity related traits, age at menarche and type I diabetes as potential confounders<sup>41-45</sup> (**Figure 2**). We identified eight genetic variants or loci from our genetic instrument for the tendency to smoke phenotype associated with different confounding traits. SNP rs12042017 has been previously reported to be associated with years of educational attainment (p-value=4.48x10<sup>-10</sup>). SNPs rs13396935 and rs6265 were observed to be associated with several adiposity related measures. The proxy variant rs1514174 of rs4650277 (R<sup>2</sup>=0.99) was further associated with BMI (p-value=2.99x10<sup>-27</sup>). Rheumatoid arthritis, age at menarche and type I diabetes were further identified as potential confounders associated with rs2734971, rs4650277 and rs1701704 (proxy for rs772921 with complete LD). Our sensitivity analysis excluding each of the SNPs or combinations of SNPs based on their common associated trait showed no overall influence on causal effect estimate for the tendency to smoke phenotype (data not shown).

### ***Genetic variants involved in brain expression***

Using brain-specific expression quantitative trait loci (eQTL) retrieved from GTEx, we identified 27 different SNPs from the genetic instrument for the tendency to smoke phenotype with varied influence in different brain regions (data not shown). Surprisingly, the corresponding candidate genes were least represented in the substantia nigra, while as many as 10 genetic variants were observed to significantly influence gene expression in cerebellar hemisphere as well as cerebellum. Our sensitivity analysis showed that excluding genetic variants mapping to genes over-expressed in the cerebellum had maximum influence on the overall causal effect estimation (OR=0.761; 95% CI=0.606-0.957; p-value=0.0197). A similar influence was observed after excluding all the genetic variants mapping to genes expressed in brain (OR=0.735; 95% CI=0.581-0.930; p-value=0.0106). Our sensitivity analysis thereby suggested an important role of the cerebellum in the smoking tendency phenotype. Our literature search for the excluded genetic variants in the tendency to smoke genetic

instrument influencing expression in cerebellum for potential influence on other biological pathways rules out pleiotropic effect of these variants.

### Secondary MR analysis

The descriptive statistics of the genetic instruments selected for the secondary MR analyses are presented in **Table 4**. The causal effect estimates are shown in **Table 5**. We first employed genetic instruments for different traits representative of the smoking phenotype (ever smoker vs. never been a regular smoker, former vs. current smoker and cigarettes per day)<sup>22-23</sup>. A previous meta-analysis of GWAS on the ever smoker phenotype in 143,023 individuals of European ancestry identified genetic variants from the *BDNF* gene to be associated with the ever smoker phenotype<sup>22</sup>. Since all the variants were in high LD with each other, we used rs6265, a non-synonymous variant with a functional effect on gene expression, as a proxy for all other variants for the MR analysis. We failed to observe any association (OR=0.545; 95% CI=0.230-1.291; p-value=0.1681). Interestingly, a genetic instrument (rs3025343) for former smokers vs. current smokers showed a trend towards risk for predisposition to PD (OR=1.874; 95% CI=1.003-3.499; p-value=0.0487). We further extracted a genetic instrument comprised of four uncorrelated genetic variants for a related phenotype of cigarettes per day from a meta-analysis of GWAS on 86,956 individuals, again failing to observe any trend (OR=0.989; 95% CI = 0.870-1.124; p-value=0.7995)<sup>23</sup>.

We further used a recently published GWAS on alcohol consumption using data from 112,117 individuals from the UK biobank<sup>24</sup>. With seven genetic variants as a genetic instrument, we were able to replicate our finding of absence of causal association of alcohol consumption with PD (OR=1.389; 95%CI=0.110-17.563; p-value=0.7621). We further failed to observe any causal association of number of regular coffee cups per day with PD (OR=1.032 ; 95% CI=0.653-1.632; p-value=0.8405)<sup>27</sup>.

We additionally investigated a protective causal role of cannabis dependence in PD by exploiting a recent GWAS study on cannabis dependence in 2080 cannabis-dependent cases and 6435 cannabis-exposed controls of European descent<sup>25</sup>. The study reported a potential role of a cluster of highly linked 26 SNPs spanning a region on chromosome 10. The study further identified a putative

functional SNP rs1409568 among this cluster responsible for the observed phenotypic association. Our investigation of a causal role of rs1409568 did not show evidence for a role of cannabis dependence in PD predisposition (OR=0.973; 95%CI=0.811-1.167; p-value=0.7681). A recent GWAS in 1531 Caucasians further reported absence of any significant SNPs with pathological gambling<sup>26</sup>. We generated a genetic instrument based on the top hits from the study with a P-value cut off of  $<10^{-4}$  employing 45 uncorrelated genetic variants. We observed no causal association of PD with pathological gambling (OR=1.004; 95% CI=0.991-1.018; p-value=0.5120).

## Discussion

To the best of our knowledge, this is the first comprehensive study exploring the role of risky behaviors as causal factors for PD using a MR approach. The study suggests that the tendency-to-smoke trait is causally related to PD with individuals who started smoking being protected against PD. Our sensitivity analysis further demonstrated robustness of the reported association in the absence of any detectable pleiotropic effect. Furthermore, our secondary MR analysis did not show any causal association of other habitual traits including other smoking phenotypes such as number of cigarettes per day, cannabis dependence, pathological gambling, alcohol and coffee consumption with PD.

Numerous observational studies have previously shown an inverse association of smoking with PD. A meta-analysis merging smoking status trait from 33 different populations demonstrated a risk reduction by 36% for ever- vs never-smokers with consistent results in both case-control and cohort studies<sup>2</sup>. Other epidemiological studies suggested significant gene-by-smoking interaction effects in PD<sup>46,47</sup>. In our study, we observed a PD risk reduction of 31% for ever smokers vs never smokers. Although risk reduction effects demonstrated in an observational and in an MR study may not be comparable, a consistency in the direction of protective associations by both the approaches is an important finding.

To validate our results, we performed secondary MR analyses using other habit-related behaviors from other GWAS. The lack of association with a previously reported genetic instrument for ever smoker instrument as well as former smoker vs. current smoker may be explained by lower

power of the GWAS with only one significant variant contributing to the instrument for both MR analyses. We also did not observe an association of PD risk and the number of cigarettes per day. One explanation would be that this continuous phenotype mainly just reflects the tobacco and nicotine exposure, whereas the ever vs never smoking might rather be a sign of risk taking behaviour. Our results thereby clearly imply the need for careful dissection of different smoking phenotypes. This will help understanding the causal role of the tendency to smoke on PD and reveal further insight into the development of the disease.

As outlined in the results section, our MR results on coffee consumption (cups per week) and alcohol consumption (drinks per week) also did not show significant causal associations. However, we cannot exclude that the analysis of coffee and alcohol consumption as quantitative traits may have the same limitations as the analysis of cigarettes smoked per day. Lastly, lack of a causal role of cannabis dependence observed in the present study needs to be further evaluated with stronger genetic instruments.

Absence of association with gambling in our analysis, however, could be attributed to the winner's curse as SNP-exposure estimates used for calculation of causal estimates may be overestimated due to limited power of the study on gambling phenotype. Another important finding of our comprehensive MR analyses was absence of any causal role of drinks per week with PD. Using data from the UK biobank, we were able to replicate our finding of absence of a causal association of alcohol consumption with risk of PD. However, our study suggests a potential causal association of the number of sexual partners and PD risk. To the best of our knowledge, no epidemiological population-based study has yet examined the role of promiscuity on PD risk. Therefore, our MR results need to be interpreted cautiously and independent lines of validation of this association are required to confirm these results.

An important limitation of our current study is that we could not directly assess associations of individual genetic variants with potential confounders of association between risk behavior and PD due to the lack of knowledge of potential confounders and unavailability of individual-level data. Nevertheless, our sensitivity analysis demonstrated that exclusion of loci being associated with PD-

associated phenotypes from the MR analyses had no effect on the overall association. We could not further provide data on the degree of sample overlap among GWAS datasets on exposure and outcome in our two-sample MR design. A considerable overlap could bias the results towards the estimates generated through observational studies. However, this potential limitation could not have any impact on our results as the IVW method using second order weights employed in the current study is known to address this bias. And lastly, before drawing conclusions on the role of risky behavior on PD, we must recognize a critical limitation of our study that we could not do a stratified MR analysis based on dopaminergic treatment in cases as dopaminergic agonists are known to modulate risky behavior in PD patients.

Despite these limitations, to our knowledge, our study represents the most comprehensive MR study to date on risky behavior phenotypes and PD. An extensive sensitivity analysis including use of genetic instruments specific to individual phenotypic traits, use of previous studies, literature search for potential pleiotropic variants and brain expression analysis collectively demonstrate a strong causal protective role of smoking tendency on PD. Furthermore, the role of automobile speeding propensity as a causal risk factor emphasizes the need for a stratified MR based on dopamine-agonist treatment. The present study also demonstrates that careful interpretation of pleiotropic signals and sensitivity analysis based on biological function could lead to fine filtering of GWAS signals. Such an approach may assist in differentiating between mediators and exposures, thereby helping us to construct the causal pathways leading to PD<sup>48</sup>.

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### **Author's contributions**

SG and CML conceived the idea of the study. SG designed the study, performed the data extraction from the literature and the statistical analyses, wrote the first draft, and revised the final draft of the manuscript. IK supervised study design and statistical analyses. FDG contributed to statistical analysis plan and conducted independent blinded statistical analyses. IK, FDG and CML participated in improving the design of the study, helped to draft the manuscript, and have been involved in revising the manuscript. MK and CK have been involved in revising the manuscript. All authors read and approved of the final manuscript.

#### **Availability of data and materials**

Not applicable

#### **Ethics Approval and consent to participate**

Not applicable

#### **Consent for publication**

Not applicable

#### **Competing Interests**

The authors declare that they have no competing interests.

#### **List of figures and Tables**

**Figure 1:** Causal association analysis and assessment of pleiotropy for the ever smoking phenotype with Parkinson's disease

- A. Scatterplot showing causal effect estimates computed using various MR methods.
- B. Funnel plot showing the extent of heterogeneity among the individual Wald ratio estimates.
- C. Plot of Cochran's Q estimates for individual SNPs constituting the genetic instrument for ever smoker phenotype using IVW method employing second order weights.
- D. Radial MR plot showing the distribution of weights contributed by individual SNPs in the causal effect estimation by IVW method employing second order weights.

**Figure 2:** Triangular representation of results from the present MR study

Abbreviations in the figure:

SNP: Single Nucleotide Polymorphism

N<sub>snp</sub>: Number of SNPs in the genetic instrument for each respective risky behavior

$\beta_{\text{snp}}$ : Regression coefficient corresponding to each specific SNP for respective arm of the triangle corresponding to each relationship

se<sub>snp</sub>: Standard error estimates corresponding to each specific SNP for the respective arm of the triangle corresponding to each relationship

GRT: General risk tolerance

ADV: Adventurousness

ASP: Automobile speeding propensity

DPW: Drinks per week

ES: Ever vs never smoking

NSP: Number of sexual partners

OR<sub>GRT</sub>, OR<sub>ADV</sub>, OR<sub>ES</sub>: Expressed as odds ratio per log odds of respective risky behavior

OR<sub>ASP</sub>, OR<sub>DPW</sub>, OR<sub>NSP</sub>: Expressed as odds ratio per SD of respective risky behavior

### **Table 1.** Summary of genetic instruments used in the present Mendelian randomization analysis.

PD = Parkinson's disease

Maximum sample size in PD cohort = 42,286 individuals.

\*p-value < 5x10<sup>-8</sup> for association with risky behavior in GWAS (20)

\*\*Excluded from further analysis (identified using pd.org database)

Automobile speed propensity: MAPT (rs62062288); HLA-DQB (rs3021058)

Drinks per week: MAPT (rs62055546)

Number of sexual partners: MAPT (rs62063281)

\*\*\*Excluded from further analysis

Adventurousness: rs1492436 and rs35377646 (r<sup>2</sup>=1.0 with similar effect estimates on exposure).

rs1492436 was selected for the present study based on better variance explained by the SNP

### **Table 2.** Causal effect estimates using different Mendelian randomization methods and heterogeneity analysis of causal effect estimates for risk taking behaviors.

P-values for effects on PD marked in bold show statistical significance after Bonferroni corrections with a cut-off p-value of 0.05/6 = 0.0083

P-values for test on heterogeneity marked in bold are below 0.05.

General risk tolerance: OR per log odds of general risk tolerance

Adventurousness: OR per log odds of adventurousness

Automobile speeding propensity: OR per SD of normalized Automobile speeding propensity

Drink per weeks: OR per SD of number of drinks per week

Ever smokers: OR per log odds of ever smoking

Number of sexual partners: OR per SD of number of sexual partners

PD = Parkinson's disease

\* Computed using second order weights

### **Table 3.** Causal effect estimates using unique loci among different phenotypic traits .

PD = Parkinson's disease

\* Computed using second order weights

### **Table 4.** Summary of genetic instruments used in the Mendelian randomization analysis based on risky- and habit-related behaviors from previous GWAS in European populations.

PD = Parkinson's disease

\* For 1 SNP (rs4105144), no proxy was available in PD dataset

\*\*For 2 SNPs, no proxy was available

\*\*\* Two different sets of SNPs were used. rs2472297 and rs2470893 showed moderate LD (r<sup>2</sup>=0.658) and hence 2 separate MR analyses were conducting including one SNP at a time.

### **Table 5.** Causal effect estimates for habit-related behaviors from previously published GWAS.

PD = Parkinson's disease

\* Computed using second order weights

\*\*Computed using Wald estimate with delta method  
\*\*\*Excluding high LD SNP rs2470893  
\*\*\*Excluding high LD SNP rs2472297

**Supplementary Table.** List of summary estimates used for the calculation of causal estimates for the primary MR analysis.

Exp: Exposure or phenotype  
GRT : General risk tolerance  
ADV: Adventurousness  
ASP : Automobile speeding propensity  
DPW: Drinks per week  
ES: Smoking tendency (Ever vs never smoking)  
NSP: Number of sexual partners  
Chr: Chromosome  
Pos: Position (GRCh37.p13)  
EA (exp): Effect allele of the SNP in the exposure dataset  
OA (exp): Other allele of the SNP in the exposure dataset  
EAF: Effect allele frequency  
Gene: Gene or nearby gene  
 $\beta$  (exp): Effect estimate of SNP from regression analysis of the SNP with respective exposure  
se (exp): Standard error of SNP from regression analysis of the SNP with respective exposure  
p-value (exp): p-value of SNP from regression analysis of the SNP with respective exposure  
EA (out): Effect allele of the SNP in the PD dataset  
OA (out): Other allele of the SNP in the PD dataset  
 $\beta$  (out): Effect estimate of SNP from regression analysis of the SNP with PD  
se (out): Standard error of SNP from regression analysis of the SNP with PD  
p-value (out): p-value of SNP from regression analysis of the SNP with PD  
PD = Parkinson's disease

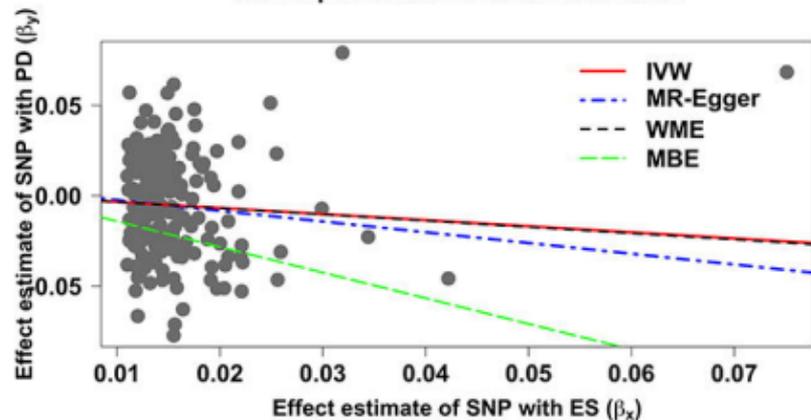
## References

1. de Lau LM, Breteler MM. Epidemiology of Parkinson's disease. *Lancet Neurol.* 2006 Jun;5(6):525-35.
2. Noyce AJ, Bestwick JP, Silveira-Moriyama L, et al. Meta-analysis of early nonmotor features and risk factors for Parkinson disease. *Ann Neurol.* 2012 Dec;72(6):893-901.
3. Lill CM. Genetics of Parkinson's disease. *Mol Cell Probes.* 2016 Dec;30(6):386-96.
4. Nalls MA, Pankratz N, Lill CM, et al. Large-scale meta-analysis of genome-wide association data identifies six new risk loci for Parkinson's disease. *Nat Genet.* 2014 Sep;46(9):989-93.
5. Chang D, Nalls MA, Hallgrimsdottir IB, et al. A meta-analysis of genome-wide association studies identifies 17 new Parkinson's disease risk loci. *Nat Genet.* 2017 Oct;49(10):1511-6.
6. Nalls MA, Blauwendraat C, Vallerga CL, et al. Parkinson's disease genetics: identifying novel risk loci, providing causal insights and improving estimates of heritable risk. *bioRxiv.* 2018.
7. Noyce AJ, Nalls MA. Mendelian Randomization - the Key to Understanding Aspects of Parkinson's Disease Causation? *Mov Disord.* 2016 Apr;31(4):478-83.
8. Peres FF, Lima AC, Hallak JEC, Crippa JA, Silva RH, Abilio VC. Cannabidiol as a Promising Strategy to Treat and Prevent Movement Disorders? *Front Pharmacol.* 2018;9:482.
9. Cossu G, Rinaldi R, Colosimo C. The rise and fall of impulse control behavior disorders. *Parkinsonism Relat Disord.* 2018 Jan;46 Suppl 1:S24-S9.
10. Moore TJ, Glenmullen J, Mattison DR. Reports of pathological gambling, hypersexuality, and compulsive shopping associated with dopamine receptor agonist drugs. *JAMA Intern Med.* 2014 Dec;174(12):1930-3.
11. Ziegler A, Mwambi H, König IR. Mendelian Randomization versus Path Models: Making Causal Inferences in Genetic Epidemiology. *Hum Hered.* 2015;79(3-4):194-204.

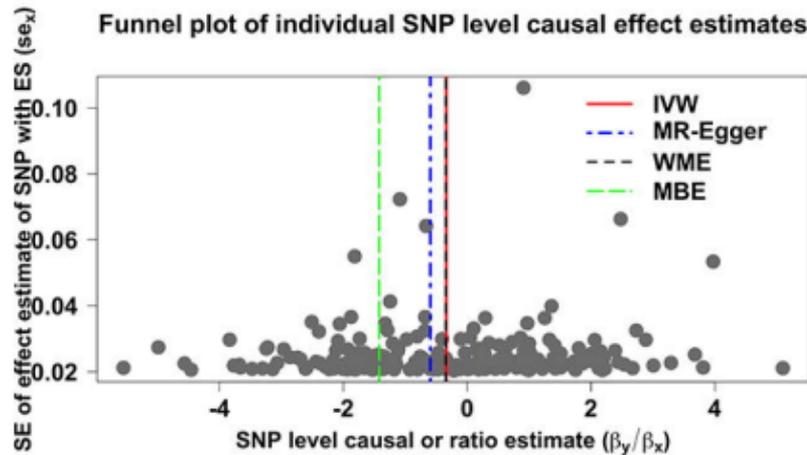
12. Prins BP, Abbasi A, Wong A, et al. Investigating the Causal Relationship of C-Reactive Protein with 32 Complex Somatic and Psychiatric Outcomes: A Large-Scale Cross-Consortium Mendelian Randomization Study. *PLoS Med.* 2016 Jun;13(6):e1001976.
13. Benn M, Nordestgaard BG, Frikke-Schmidt R, Tybjaerg-Hansen A. Low LDL cholesterol, PCSK9 and HMGCR genetic variation, and risk of Alzheimer's disease and Parkinson's disease: Mendelian randomisation study. *BMJ.* 2017 Apr 24;357:j1648.
14. Pichler I, Del Greco MF, Gogele M, et al. Serum iron levels and the risk of Parkinson disease: a Mendelian randomization study. *PLoS Med.* 2013;10(6):e1001462.
15. Noyce AJ, Kia DA, Hemani G, et al. Estimating the causal influence of body mass index on risk of Parkinson disease: A Mendelian randomisation study. *PLoS Med.* 2017 Jun;14(6):e1002314.
16. Larsson SC, Singleton AB, Nalls MA, Richards JB, International Parkinson's Disease Genomics C. No clear support for a role for vitamin D in Parkinson's disease: A Mendelian randomization study. *Mov Disord.* 2017 Aug;32(8):1249-52.
17. Kobylecki CJ, Nordestgaard BG, Afzal S. Plasma urate and risk of Parkinson's disease: A mendelian randomization study. *Ann Neurol.* 2018 Jul 16.
18. Kia DA, Noyce AJ, White J, et al. Mendelian randomization study shows no causal relationship between circulating urate levels and Parkinson's disease. *Ann Neurol.* 2018 Jul 16.
19. Burgess S, Bowden J. Integrating summarized data from multiple genetic variants in Mendelian randomization: bias and coverage properties of inverse-variance weighted methods. *arXiv:151204486* 2015.
20. Linnér RK, Biroli P, Kong E, et al. Genome-wide study identifies 611 loci associated with risk tolerance and risky behaviors. *bioRxiv.* 2018.
21. Lill CM, Roehr JT, McQueen MB, et al. Comprehensive research synopsis and systematic meta-analyses in Parkinson's disease genetics: The PDGene database. *PLoS Genet.* 2012;8(3):e1002548.
22. Tobacco and Genetics Consortium. Genome-wide meta-analyses identify multiple loci associated with smoking behavior. *Nat Genet.* 2010 May;42(5):441-7.
23. Thorgeirsson TE, Gudbjartsson DF, Surakka I, et al. Sequence variants at CHRN3-CHRNA6 and CYP2A6 affect smoking behavior. *Nat Genet.* 2010 May;42(5):448-53.
24. Clarke T-K, Adams MJ, Davies G, et al. Genome-wide association study of alcohol consumption and genetic overlap with other health-related traits in UK Biobank (N=112,117). *bioRxiv.* 2017.
25. Agrawal A, Chou YL, Carey CE, et al. Genome-wide association study identifies a novel locus for cannabis dependence. *Mol Psychiatry.* 2018 May;23(5):1293-302.
26. Lang M, Lemenager T, Streit F, et al. Genome-wide association study of pathological gambling. *Eur Psychiatry.* 2016 Aug;36:38-46.
27. Coffee and Caffeine Genetics Consortium, Cornelis MC, Byrne EM, et al. Genome-wide meta-analysis identifies six novel loci associated with habitual coffee consumption. *Mol Psychiatry.* 2015 May;20(5):647-56.
28. Grover S, Del Greco MF, Stein CM, Ziegler A. Mendelian Randomization. *Methods Mol Biol.* 2017;1666:581-628.
29. Li B, Martin E. An approximation to the F-distribution using the chi-square distribution. *Computational Statistica and Data Analysis.* 2002;40(1):6.
30. Brion MJ, Shakhbazov K, Visscher PM. Calculating statistical power in Mendelian randomization studies. *Int J Epidemiol.* 2013 Oct;42(5):1497-501.
31. Bowden J, Del Greco MF, Minelli C, et al. Improving the accuracy of two-sample summary data Mendelian randomization: moving beyond the NOME assumption. *bioRxiv.* 2017.
32. Bowden J, Davey Smith G, Haycock PC, Burgess S. Consistent Estimation in Mendelian Randomization with Some Invalid Instruments Using a Weighted Median Estimator. *Genetic epidemiology.* 2016 May;40(4):304-14.

33. Bowden J, Davey Smith G, Burgess S. Mendelian randomization with invalid instruments: effect estimation and bias detection through Egger regression. *Int J Epidemiol.* 2015 Apr;44(2):512-25.
34. Hartwig FP, Davey Smith G, Bowden J. Robust inference in summary data Mendelian randomization via the zero modal pleiotropy assumption. *Int J Epidemiol.* 2017 Dec 1;46(6):1985-98.
35. Del Greco MF, Minelli C, Sheehan NA, Thompson JR. Detecting pleiotropy in Mendelian randomisation studies with summary data and a continuous outcome. *Stat Med.* 2015 Sep 20;34(21):2926-40.
36. Rucker G, Schwarzer G, Carpenter JR, Binder H, Schumacher M. Treatment-effect estimates adjusted for small-study effects via a limit meta-analysis. *Biostatistics.* 2011 Jan;12(1):122-42.
37. Bowden J, Del Greco MF, Minelli C, Davey Smith G, Sheehan N, Thompson J. A framework for the investigation of pleiotropy in two-sample summary data Mendelian randomization. *Stat Med.* 2017 May 20;36(11):1783-802.
38. Bowden J, Spiller W, Del Greco MF, et al. Improving the visualization, interpretation and analysis of two-sample summary data Mendelian randomization via the Radial plot and Radial regression. *Int J Epidemiol.* 2018 Jun 28.
39. Staley JR, Blackshaw J, Kamat MA, et al. PhenoScanner: a database of human genotype-phenotype associations. *Bioinformatics.* 2016 Oct 15;32(20):3207-9.
40. GTEx Consortium. Human genomics. The Genotype-Tissue Expression (GTEx) pilot analysis: multitissue gene regulation in humans. *Science.* 2015 May 8;348(6235):648-60.
41. Rughjerg K, Friis S, Ritz B, Schernhammer ES, Korbo L, Olsen JH. Autoimmune disease and risk for Parkinson disease: a population-based case-control study. *Neurology.* 2009 Nov 3;73(18):1462-8.
42. Kotagal V, Bohnen NI, Muller ML, et al. Educational attainment and motor burden in Parkinson's disease. *Mov Disord.* 2015 Jul;30(8):1143-7.
43. Park K, Oeda T, Kohsaka M, Tomita S, Umemura A, Sawada H. Low body mass index and life prognosis in Parkinson's disease. *Parkinsonism Relat Disord.* 2018 May 15.
44. Cereda E, Barichella M, Cassani E, Caccialanza R, Pezzoli G. Reproductive factors and clinical features of Parkinson's disease. *Parkinsonism Relat Disord.* 2013 Dec;19(12):1094-9.
45. De Pablo-Fernandez E GR, Pakpoor J, Noyce AJ, Warner TT. Association between diabetes and subsequent Parkinson disease: A record-linkage cohort study. *Neurology.* 2018.
46. Chuang YH, Lee PC, Vlaar T, et al. Pooled analysis of the HLA-DRB1 by smoking interaction in Parkinson disease. *Ann Neurol.* 2017 Nov;82(5):655-64.
47. Lee PC, Ahmed I, Lorient MA, et al. Smoking and Parkinson disease: Evidence for gene-by-smoking interactions. *Neurology.* 2018 Feb 13;90(7):e583-e92.
48. Grover S, Del Greco MF, König I. Evaluating the current state of Mendelian randomization studies: A protocol for a systematic review on methodological and clinical aspects using neurodegenerative disorder as a model outcome. *Syst Rev.* 2018.

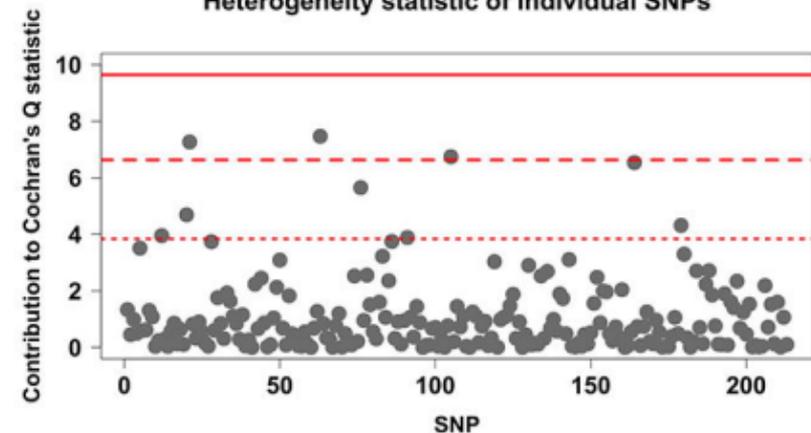
Scatterplot of causal effect estimates



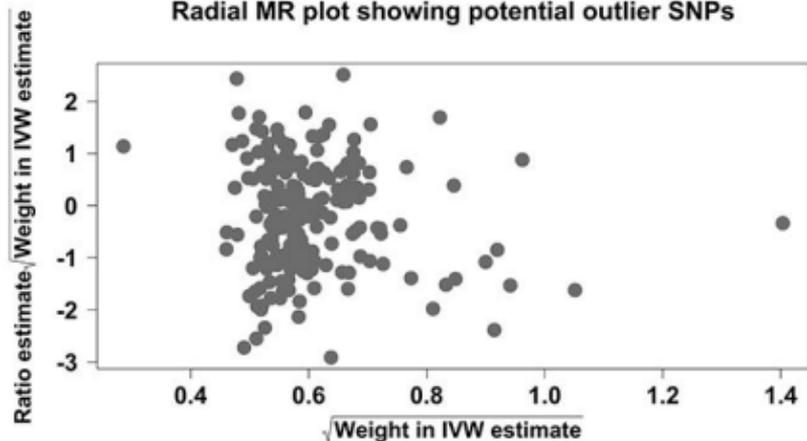
Funnel plot of individual SNP level causal effect estimates



Heterogeneity statistic of individual SNPs

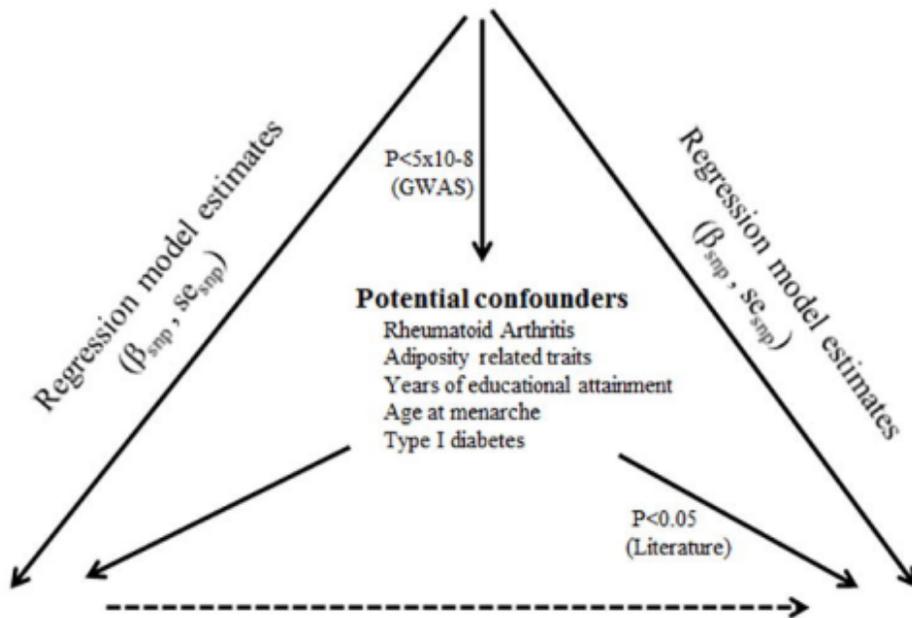


Radial MR plot showing potential outlier SNPs



**SNPs**  
(Genetic Instruments)

GRT (N<sub>snp</sub> =117); ADV (N<sub>snp</sub>=151); ASP (N<sub>snp</sub>=35)  
DPW (N<sub>snp</sub>=70); ES (N<sub>snp</sub>=213); NSP (N<sub>snp</sub>=109)



**Risky Behaviors**

GRT (N=939908); ADV (N=557923); ASP (N=404291)  
DPW (N=414343); ES (N=518633); NSP (N=370711)

**Causal estimates using IVW method**

OR<sub>GRT</sub> = 1.620; 95% CI = 1.046-2.511; P = 0.0311  
OR<sub>ADV</sub> = 1.091; 95% CI = 0.810-1.470; P = 0.5620  
OR<sub>ASP</sub> = 2.043; 95% CI = 1.076-3.876; P = 0.0299  
OR<sub>DPW</sub> = 1.150; 95% CI = 0.868-1.525; P = 0.3247  
OR<sub>ES</sub> = 0.714; 95% CI = 0.568-0.897; P = 0.0041\*  
OR<sub>NSP</sub> = 1.473; 95% CI = 1.079-2.010; P = 0.0152

**Parkinson's Disease**

PD (Cases N = 9581, Controls N = 33245)

**Table 1.**

<b>Risky behavior</b>	<b>Maximum sample size (N)</b>	<b># SNPs associated with risky behavior*</b>	<b># SNPs with direct influence on PD **</b>	<b># SNPs in high LD (<math>R^2 &gt; 0.25</math>)***</b>	<b># proxy SNPs (<math>R^2 &gt; 0.9</math>)</b>	<b># SNPs in genetic instrument</b>	<b>F-statistic</b>
General risk tolerance	939,908	124	0	0	10	117	5064.4
Adventurousness	557,923	167	0	1	16	150	6874.5
Automobile speeding propensity	404,291	42	2	0	2	35	1406.4
Drinks per week	414,343	85	1	0	9	70	4342.4
Tendency to smoke	518,633	223	0	0	11	213	9639.6
Number of sexual partners	370,711	118	1	0	12	109	4515.6

**Table 2.**

Risky behavior	MR methodology	Causal effect estimates (PD as outcome)			Tests of heterogeneity	
		OR*	95%CI	P-value	Test	P-value
General risk tolerance	Inverse variance weighted (2nd order weights)	1.620	1.046-2.511	0.0311	MR-Egger intercept (p-value)	0.2114
	MR-Egger	0.406	0.042-3.934	0.4335	I <sup>2</sup> (%)	8.34%
	Weighted median method	1.122	0.820-1.535	0.7148	Cochrane Q-test (p-value)	0.2367
	Weighted mode method (NOME assumptions)	0.708	0.141-3.550	0.6751	Rucker's Q <sup>2</sup> -test (p-value)	0.2507
					Rucker's Q <sup>2</sup> statistic /Cochrane Q statistic	1.0141
Adventurousness	Inverse variance weighted (2nd order weights)	1.091	0.810-1.470	0.5620	MR-Egger intercept (p-value)	0.0796
	MR-Egger	0.382	0.111-1.318	0.1192	I <sup>2</sup> (%)	15.85%
	Weighted median method	0.879	0.712-1.085	0.5427	Cochrane Q-test (p-value)	0.0579
	Weighted mode method (NOME assumptions)	0.607	0.252-1.460	0.2667	Rucker's Q <sup>2</sup> -test (p-value)	0.0730
					Rucker's Q <sup>2</sup> statistic /Cochrane Q statistic	1.0194
Automobile speeding propensity	Inverse variance weighted (2nd order weights)	2.043	1.076-3.876	0.0299	MR-Egger intercept (p-value)	0.0617
	MR-Egger	0.125	0.006-2.654	0.1757	I <sup>2</sup> (%)	28.75%
	Weighted median method	2.738	1.856-4.040	0.0140	Cochrane Q-test (p-value)	0.0594
	Weighted mode method (NOME assumptions)	5.202	0.673-40.193	0.1232	Rucker's Q <sup>2</sup> -test (p-value)	0.1213
					Rucker's Q <sup>2</sup> statistic /Cochrane Q statistic	1.1189
Drinks per week	Inverse variance weighted (2nd order weights)	1.150	0.868-1.525	0.3247	MR-Egger intercept (p-value)	<b>0.0408</b>
	MR-Egger	0.791	0.498-1.257	0.3159	I <sup>2</sup> (%)	0.00%
	Weighted median method	0.845	0.676-1.055	0.4513	Cochrane Q-test (p-value)	0.6032
	Weighted mode method (NOME assumptions)	0.855	0.543-1.344	0.4992	Rucker's Q <sup>2</sup> -test (p-value)	0.7071
					Rucker's Q <sup>2</sup> statistic /Cochrane Q statistic	1.0672
Tendency to smoke	Inverse variance weighted (2nd order weights)	0.714	0.568-0.897	<b>0.0041</b>	MR-Egger intercept (p-value)	0.6619
	MR-Egger	0.552	0.185-1.646	0.2849	I <sup>2</sup> (%)	6.30%
	Weighted median method	0.707	0.601-0.832	0.0339	Cochrane Q-test (p-value)	0.2389
	Weighted mode method (NOME assumptions)	0.242	0.083-0.703	<b>0.0098</b>	Rucker's Q <sup>2</sup> -test (p-value)	0.2276
					Rucker's Q <sup>2</sup> statistic /Cochrane Q statistic	1.0010
Number of sexual partners	Inverse variance weighted (2nd order weights)	1.473	1.079-2.010	0.0152	MR-Egger intercept (p-value)	0.3621
	MR-Egger	0.674	0.113-4.006	0.6615	I <sup>2</sup> (%)	21.22%
	Weighted median method	1.365	1.103-1.688	0.1465	Cochrane Q-test (p-value)	<b>0.0308</b>
	Weighted mode method (NOME assumptions)	1.270	0.446-3.618	0.6547	Rucker's Q <sup>2</sup> -test (p-value)	0.0298
					Rucker's Q <sup>2</sup> statistic /Cochrane Q statistic	1.0064

**Table 3.**

Risky behavior	MR methodology	Number of SNPs ( $R^2 \leq 0.80$ across phenotypes)	Causal effect estimates (PD as outcome)			Number of SNPs ( $R^2 \leq 0.01$ across phenotypes)	Causal effect estimates (PD as outcome)		
			OR (95% )*	95%CI	p-value		OR (95% )*	95%CI	p-value
General risk tolerance	Inverse variance weighted*	94	1.986	1.215-3.243	<b>0.0066</b>	66	1.821	1.017-3.261	<b>0.0440</b>
Adventurousness	Inverse variance weighted*	125	1.169	0.837-1.633	0.3564	80	1.002	0.672-1.493	0.9915
Automobile speeding propensity	Inverse variance weighted*	31	2.279	1.130-4.597	<b>0.0229</b>	19	2.498	0.893-6.992	0.0780
Drinks per week	Inverse variance weighted*	67	1.094	0.819-1.459	0.5370	52	1.091	0.800-1.488	0.5755
Ever smoker	Inverse variance weighted*	195	0.713	0.557-0.913	<b>0.0075</b>	154	0.719	0.547-0.945	<b>0.0185</b>
Number of sex partners	Inverse variance weighted*	94	1.635	1.165-2.293	<b>0.0049</b>	50	1.585	1.003-2.502	<b>0.0484</b>

**Table 4.**

Habitual phenotype	Phenotypic definition	Reference	Maximum sample size	GWAS study cohort used for extraction of effect estimates	GWAS cut-off used for prioritization of SNPs (p-value)	# significant SNP/s	# proxy SNPs (R <sup>2</sup> >0.9)	# SNPs in high LD (R <sup>2</sup> >0.25) excluded	# SNPs in genetic instrument
Smoking	Cigarettes per day	Thorgeirsson <i>et al.</i> 2010 [23]	86,956	Pooled	5x10 <sup>-8</sup>	6	0	1	4*
Smoking initiation	Ever smoker vs. never been a regular smoker	Tobacco Genetic Consortium 2010 [22]	143,023	Pooled	5x10 <sup>-8</sup>	8	0	7	1
Smoking cessation	Former vs. current smoker	Tobacco Genetic Consortium 2010 [22]	64,924	Pooled	5x10 <sup>-8</sup>	1	0	NA	1
Cannabis dependence	Cannabis dependence (DSMIV) vs. Individuals not meeting cannabis dependence criteria with a history of at least once in lifetime use of cannabis	Agrawal <i>et al.</i> 2018 [25]	8515	Discovery	1x10 <sup>-6</sup>	26	0	25	1
Pathological Gambling	Individuals diagnosed with pathological gambling (DSMIII/IV) vs. population controls	Lang <i>et al.</i> 2017 [26]	1431	Discovery	1x10 <sup>-4</sup>	57	1	12	45
Alcohol consumption	Units of alcohol consumed in the previous week	Clarke <i>et al.</i> 2017 [24]	112,117	Discovery	5x10 <sup>-8</sup>	10	1	1	7**
Coffee consumption	Regular coffee cups consumed per day	Coffee and Caffeine Genetic Consortium 2015 [27]	91,642	Discovery	5x10 <sup>-8</sup>	6	0	2	4***

**Table 5.**

<b>Risky behavior</b>	<b>Causal effect estimate (PD as outcome) (Inverse variance weighted*)</b>		<b>Test of heterogeneity</b>
	<b>OR (95% CI)</b>	<b>p-value</b>	<b>MR Egger Intercept test (p-value)</b>
Cigarettes per day	0.989 (0.870-1.124)	0.7995	0.5668
Ever smoker vs. never been a regular smoker	0.545 (0.230-1.291)	0.1681**	NA
Former vs. current smoker	1.874 (1.003-3.499)	0.0487**	NA
Cannabis dependence	0.973 (0.811-1.167)	0.7681**	NA
Pathological gambling	1.004 (0.991-1.018)	0.5120	0.6212
Alcohol consumption	1.389 (0.110-17.563)	0.7621	0.4981
Coffee consumption***	1.032 (0.653-1.632)	0.8405	0.1654
Coffee consumption****	1.078 (0.568-1.767)	0.6599	0.1332