Genomic basis of transcriptome dynamics in rice under field conditions

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How genetic variations affect gene expression dynamics of field-grown plants remains unclear. Using statistical analysis of large-scale time-series RNA-sequencing of field-grown rice from chromosome segment substitution lines (CSSLs), we identified 1675 expression dynamics quantitative trait loci (edQTLs) leading to polymorphisms in expression dynamics under field conditions. Based on the edQTL and environmental information, we successfully predicted gene expression under environments different from training environments, and in rice cultivars with more complex genotypes than the CSSLs. Overall, edQTL' identification helped understanding the genetic architecture of expression dynamics under field conditions, which is difficult to assess with laboratory experiments[1].The prediction of expression based on edQTL and environmental information will contribute to crop breeding by increasing the accuracy of trait prediction under diverse conditions.

Organisms respond to fluctuations in field environments variably depending on their genetic backgrounds, developmental stages, and physiological status. This variability can cause missing heritability in crop breeding and low accuracy in medicine[2–4]. Environmental stimuli can induce transcriptional responses directly and/or indirectly. Measuring transcriptome dynamics is a comprehensive method for assessing environmental responses and their alteration due to genetic, developmental, and

physiological factors. The expression quantitative trait loci (eQTL) approach is frequently

used to assess the association between genetic variation and gene expression

polymorphism[5–9]. Because this approach requires transcriptome data for the complete set

of a QTL mapping population under the given conditions[5–9], only a limited range of

environmental conditions are covered. Contrarily, statistical models based on

meteorological information, circadian clock, and developmental age have succeeded in

describing transcriptome dynamics in the field[10–12], although these models can only

predict the transcriptome of one or few genotypes used as training data[10–12]. Thus, no

study has clarified the relationship between genetic variation and transcriptome dynamics

under fluctuating field conditions.

Here, we first defined the loci that determine polymorphisms in gene expression

dynamics as 'edQTL'. edQTL can explain polymorphisms in gene expression dynamics

under a broad range of environmental conditions, while eQTL can only explain them

under the investigated environment (Fig. 1a). Our approach consists of four steps: time-

series RNA-sequencing (RNA-Seq), prediction model development, edQTL detection,

and evaluation of edQTL (See Methods, Fig. 1b, and Supplementary Fig. 1). We used the

rice (*Oryza sativa* L.) cultivars 'Koshihikari' (a leading *japonica* cultivar) and 'Takanari'

(a high-yield *indica* cultivar)[13,14] that would present substantial polymorphisms in

expression dynamics. The 78 reciprocal CSSLs and two backcross inbred lines (BILs)

developed from the parental lines were also used[14–16] (Fig. 1b and Supplementary Fig.

2a). To refelct the effect of plant age (days after seeding) in our model, we prepared four

different sets of rice plants transplanted at two-weeks intervals. Sixteen sets (days) of

bihourly sampling for 24 h were conducted from May to September 2015 (cropping

season in Japan). At each sampling time, the youngest fully expanded leaves were

collected from two different genotypes at each transplant set while withered plants were

excluded from the sampling, which could not be sampled (Fig. 1c, Supplementary Fig.

2b, and Table S2). The 926 individual leaves were applied for RNA-Seq, to obtain

transcriptome data. After filtering samples and genes (Supplementary Fig. 3a, b) and

confirming genotypes (Supplementary Fig. 3c and 4), 23,924 expressed genes from 854

samples were used. Correlation analysis of the transcriptome data showed an obvious

transcriptome-wide diurnal variation in the expressed genes (Fig. 1d).

We then developed two prediction models describing transcriptome dynamics under

fluctuating field environments in 'Koshihikari' and 'Takanari'. In this step, we used the

statistical modeling tool 'FIT'[10] for predicting transcriptome dynamics under field

conditions. As most of the genome in the CSSLs was not substituted (Supplementary Fig

2a), the expression dynamics of most genes was expected to be identical to that of the

background parent. Thus, the RNA-Seq data of 'Koshihikari'- and 'Takanari'-background

samples were used to develop the prediction model for 'Koshihikari' and 'Takanari' lines,

respectively. Circadian clock and meteorological data (air temperature and global solar

radiation) (Supplementary Fig. 5a) were considered in model development and "scaled

age" was used to adjust differences in heading date among rice genotypes (See Methods

and Supplementary Fig. 5). We found polymorphisms in the predicted expression

dynamics of 3696 genes (15.4% of the expressed genes; Supplementary Fig. 3e and Table

S3). For instance, the 'Koshihikari' model for *Os09g0343200* showed an obvious diurnal

oscillation in expression, whereas the 'Takanari'-model showed a constant low-level

expression (Fig. 2a). Notably, in some 'Takanari'-background CSSLs, SL1329 and

SL1330, the expression of *Os09g0343200* resembled that of the 'Koshihikari' model

rather than the 'Takanari'-model (Fig. 2a), suggesting that genetic substitution in the

chromosome 9 affected *Os09g0343200* expression.

The CSSLs had been previously genotyped using 141 simple sequence repeat (SSR)

markers (Supplementary Fig. 2a)[14]. As so, at the edQTL detection step, we searched for

the SSR markers explaining the expression dynamics polymorphisms. We evaluated the

decrease in the residual error of each gene assuming edQTL around each SSR marker

(See Methods and Supplementary Fig. 6a). For *Os09g0343200*, the sum of residual errors

significantly decreased only by assuming edQTL on chromosome 9 (Fig. 2b and

Supplementary Fig. 6b), indicating that this is a *cis*-edQTL. The prediction models chosen

for each sample based on edQTL genotypes explained the expression dynamics of

Os09g0343200 better than their background genotypes (Fig. 2c). Overall, edQTL were

identified for 1675 genes (45.3% of genes with expression dynamics polymorphism

between 'Koshihikari' and 'Takanari', false discovery rate = 0.05; Supplementary Fig.

6f), including 222 genes affected by *trans*-edQTL (Fig. 2d). Forty-three genes were

affected by multiple edQTL. For 33 of these genes, the sum of residual errors based on

all edQTL was smaller than that of the most significant edQTL alone (Supplementary Fig.

7). Detailed results for individual genes can be found in our database

(https://ps.agr.ryukoku.ac.jp/osa_edqtl). A *cis*-edQTL could explain the expression

dynamics polymorphism of *Os01g0537250* that be specifically observed in young plants

(Fig. 2e,f). The expression depended on time of day in both parental models, but scaled

age was only important in the 'Koshihikari' model (Fig. 2e,f and Supplementary Fig. 8).

The edQTL-affected 393 genes depending on the time of day and 275 genes depending

on scaled age (Supplementary Fig. 9).

In the edQTL evaluation step, we first compared the predicted performances based

on the edQTL model and background genotypes model (BG model) under different

environments from where training data were obtained for model development.

'Koshihikari', 'Takanari', and the CSSLs of two transplant sets were cultivated in a field

different from that used in 2015 and sampled in August 2016; 139 RNA-Seq datasets were

obtained (Supplementary Fig. 2c and Table S2). Environmental factors differed between

the 2015 and 2016 fields (Supplementary Fig. 10a). For 91.8% of the 1675 genes affected

by the edQTL, the prediction was improved by the edQTL model compared with the BG

model (Fig. 3a,b and Supplementary Fig. 10b). Furthermore, the prediction of the

expressed genes for the validation dataset showed comparable accuracy to that of the

training dataset (Supplementary Fig. 11). Thus, the identified edQTL could explain the

expression dynamics polymorphisms in different years and locations. For further

verification , two BILs between 'Koshihikari' and 'Takanari' (HP-a and HP-b)[15,16] were

cultivated at Takatsuki in 2015, and five and six RNA-Seq datasets were obtained for HP-

a and HP-b, respectively (Supplementary Fig 2b and Table S2). These lines carried

'Koshihikari' alleles at 16.3% and 19.9% of their markers with the genetic background of

'Takanari' (Supplementary Fig 2a). To evaluate the performance of the edQTL-based

prediction, we calculated the sum of prediction errors of all edQTL-influenced genes

using edQTL-, 'Koshihikari'-, and 'Takanari'-models. Overall, the edQTL model

provided the best prediction (Fig. 3c,d). Permutation analysis of markers in HP-a and HP-

b genomes showed the significant advantage of the edQTL model (p < 0.001) (Fig. 3c,d).

For instance, as the genotype of the *trans*-edQTL for *Os03g0388300* comprised the

'Koshihikari' allele in the genome of HP-a and the 'Takanari' allele in the genome of HP-

b (Fig. 3e), 'Koshihikari'- and 'Takanari'-models were used to predict *Os03g0388300*

expression in the edQTL models for HP-a and HP-b, respectively. The expression of

Os03g0388300 fluctuated with time and it was higher in 'Koshihikari' than in 'Takanari'

(Supplementary Fig. 12a). The prediction of *Os03g0388300* expression based on the

edQTL models was better than the prediction based on the BG models (Fig. 3f).

Regarding *OsKS3* (*Os04g0611700*)[17] in HP-b, the prediction based on the edQTL model

was worse than that based on the BG model (Supplementary Fig. 12b,c). True edQTL

might therefore exist around SSR markers 52 or 53, which were not substituted with the

'Koshihikari' allele in HP-b (Supplementary Fig. 12b). Such difficulty may occur in some

edQTL-based predictions for the BILs, because several edQTLs identified by the CSSLs

can be unlinked to the substituted genome regions in the BILs. Finally, we concluded that

our approach successfully identified loci linked to expression dynamics polymorphisms

in field conditions.

The edQTL approach scans a broader range of conditions (Fig. 1a) but it is less

sensitive when focusing on a specific condition. This is because eQTL are expected to

show the same effects on gene expressions in all samples used in a study. Contrarily, some

edQTL were expected to affect gene expressions in some specific samples, depending on

their environments. The lower sensitivity of the edQTL approach relative to that of the

eQTL approach might explain the smaller fraction of *trans*-edQTL compared with that of

trans-eQTL in previous studies (13.3% and 62–71%[6,8,9], respectively) because *trans*-

edQTL/eQTL have generally smaller effect sizes than *cis*-edQTL/eQTL[6,8].

Phenotypic plasticity plays key roles in plant's environmental adaptation[18,19].

Nevertheless, as this is difficult to study in natural environments, the genetic architecture

of phenotypic plasticity under field conditions remains largely unveiled. The edQTL

identified by our novel method improve the understanding of the genetic architecture

underlying the expression dynamics polymorphisms between 'Koshihikari' and 'Takanari'

in the field. In addition, our method showed that rice gene expression dynamics can be

predicted based on genotypic and meteorological information. In crop breeding,

polymorphisms of environmental responses result in unexpected performance of a bred

crop under environments differing from experimental fields. Because the transcriptome

can be beneficial for trait prediction[20–23] and our approach allows transcriptome

predictions under various conditions, it contributes for crop breeding and for

understanding plant systems.

METHODS

Overview of edQTL identification and verification. Our approach consisted of four

steps (Fig. 1b). First, the parent rice lines and their descendants were cultivated in a paddy

field and sparsely sampled at several time points for RNA-Seq (Time-series RNA-Seq,

Fig. 1b and Supplementary Fig. 1). Second, parental prediction models were developed

to describe environmental responses in 'Koshihikari' and 'Takanari' in terms of gene

expression (Prediction model development, Fig. 1b). Most gene expression dynamics

would be identical among the CSSLs and the parent with the same background genotype.

Thus, we used 'Koshihikari'- and 'Takanari'-background CSSLs as well as their parents

to develop the parental models (Fig. 1b and Supplementary Fig. 1). Third, the dependency

of expression dynamics polymorphisms on genetic variation was statistically evaluated

based on comparisons between predictive gene expression and observed gene expression

in CSSLs assuming that genetic variation of SSR markers leads to expression dynamics

polymorphisms (Supplementary Fig. 6) (edQTL detection, Fig. 1b). Fourth, by integrating

the parental models and edQTL information, gene expression dynamics was predicted

based on environmental and genotypic information (Evaluation of edQTL).

Plant materials. We cultivated the following rice (*O. sativa*) lines: the *japonica* variety

'Koshihikari', the *indica* variety 'Takanari', 78 reciprocal CSSLs (40 'Koshihikari'-

background lines, except for SL1213, and 38 'Takanari'-background lines, except for

SL1306)[15], and two 'Takanari'-background BILs (HP-a and HP-b)[15,16]. Each variety and

line were sown in nursery trays. Approximately one month after sowing, seedlings were

transplanted to a paddy field at Takatsuki, Japan (34°51′19″N, 135°37′51″E) in 2015 and

at Kizugawa, Japan (34°4′05″N, 135°50′33″E) in 2016. To consider the effect of plant age

in the prediction of gene expression dynamics, four and two transplant sets were prepared

in 2015 and 2016, respectively (Fig. 1c and Supplementary Fig. 2b,c). Seed sowing and

transplanting were conducted according to the following schedules. In 2015, transplant

set 1: Seed sowing date "2015-04-03", Transplanting date "2015-05-01"; transplant set 2:

Seed sowing date "2015-04-17", Transplanting date "2015-05-08"; transplant set 3: Seed

sowing date "2015-05-01", Transplanting date "2015-05-22"; transplant set 4: Seed

sowing date "2015-05-15", Transplanting date "2015-06-05". In 2016, transplant set 1:

Seed sowing date "2016-04-21", Transplanting date "2016-05-16"; transplant set 2: Seed

sowing date "2016-05-19", Transplanting date "2016-06-08".

Sampling and RNA extraction. Sixteen sets (2015) and three sets (2016) of bihourly

sampling for 22 h, from 16:00 on one day to 14:00 on the next, were conducted on the

following dates (Supplementary Fig. 2b,c and Table S2): 5–6 May, 16–17 June, 22–23

June, 29–30 June, 6–7 July, 13–14 July, 20–21 July, 27–28 July, 3–4 August, 19–20

August, 24–25 August, 31 August to 1 September, 7–8 September, 14–15 September, 21–

22 September, and 28–29 September in 2015; and 4–5, 11–12, and 25–26 August in 2016.

We applied a stratified randomization strategy to the sampling schedule to avoid biased

sampling of each line over seasons. We separated individual plants into four and two

groups in each transplant set in 2015 and 2016, respectively, each containing the same

number of individuals per line. Then, the order of sampling was randomized in each group.

According to the sampling schedule, two plants from the 82 genotypes of each transplant

set were sampled. Because aged rice withered, several samples were missed at the latest

cultivation (Supplementary Fig. 2b). Transplant set and sampling time for each sample

are listed in Table S2. The youngest fully expanded leaf from each plant was collected,

immediately frozen in liquid nitrogen, and stored at -80 °C until RNA isolation for RNA-

Seq. Individual plants were only sampled once to avoid wounding response. Thus, all

RNA-Seq data were obtained from independent plants. Frozen samples were

homogenized with TissueLyser II (Qiagen, Hilden, Germany), and total RNA was then

extracted using the Maxwell 16 LEV Plant RNA Kit (Promega, Madison, WI, USA) and

Maxwell 16 Automated Purification System (Promega). The concentration of RNA was

measured using the Quant-iT RNA Assay Kit, broad range (Thermo Fisher Scientific,

Waltham, MA, USA).

RNA-Seq library preparation and sequencing. An automated liquid handling system

Freedom EVO 150 (TECAN, Zurich, Switzerland) and a thermal cycler ODTC 384

(INHECO, Martinsried, Germany) were utilized to prepare the RNA-Seq library of 384

samples simultaneously. The protocol of RNA-Seq library preparation is described below.

We enzymatically degraded abundant RNAs such as rRNAs in the leaves, which cause

wasteful consumption of sequence reads (Supplementary Fig. 13). Aliquots of 252 and

192 types of 100-µM 60-mer antisense DNAs (IDT, San Jose, CA, USA) covering rRNAs

and the top eight most frequently-detected transcripts (*Osp1g00110.1*, *Osp1g00180.1*,

Osp1g00420.1, *Osp1g00170.1*, *Osp1g00340.1*, *Osp1g00330.1*, *Osp1g00600.1*, and

Os11t0707000-02) in the RNA-Seq dataset of *O. sativa* leaf were mixed in equal amounts

(SDRNA oligo pool)[24,25]. One microgram of total RNA, 1.5 µL of the SDRNA oligo pool,

3 µL of 5× hybridization buffer (0.5 M Tris-HCl (pH 7.4) and 1 M NaCl), and RNase-

free water were mixed to a total volume of 10 µL. The mixture was incubated to anneal

the antisense DNAs with the target transcripts using the thermal cycler ODTC 384 with

the following program: 95 °C for 2 min, slow cooling to 45 °C (0.1 °C/s), and 45 °C for

5 min. Subsequently, 1 µL of Tli Thermostable RNase H (TaKaRa, Kusatsu, Japan), 2 µL

of 10× RNase H digestion buffer [500 mM Tris-HCl (pH 7.4), 1 M NaCl, and 200 mM

MgCl], and 7 µL of RNase-free water were added and mixed well. This mixture was

incubated at 45 °C for 30 min to selectively digest the RNA of the DNA/RNA hybrid

double strand. Then, 24 μL of AMPure XP (Beckman Coulter, Brea, CA, USA) beads

was added and the mixture was purified in a 384-well magnetic plate following the

manufacturer's instructions. The RNA was then eluted with 14 μL of RNase-free water.

Ten microliters of the RNA solution, 1 μL of DNaseI (5 unit/μL), 1 μL of 10× DNaseI

buffer (both Promega), 1 μL of 100 mM DTT (Invitrogen, Carlsbad, CA, USA), and 7 μL

of RNase-free water were mixed well. This mixture was incubated at 37 °C for 30 min to

degrade the SDRNA oligo pool. Subsequently, 24 μL of AMPure XP beads was added

and purified according to the manufacturer's instructions. The RNA was eluted with 14

μL of RNase-free water. Five microliters of the purified mRNA obtained was mixed with

4 μL of 5× SS buffer (Invitrogen), and 1 μL of 100 mM DTT. Fragmentation of mRNA

was carried out at 94 °C for 4.5 min and immediately cooled on ice. Subsequently, 0.6

μL of 100 μM random primer $(N)_6$ (Promega) and 0.9 μL of RNase-free water were added

to the mixture, which was incubated at 50 °C for 5 min and immediately chilled on ice to

relax the secondary structure of the RNA. The fragmented RNA with random hexamers

and the reverse transcription master mix [1 μL of 100 mM DTT, 0.4 μL of dNTPs (25

mM each; Promega), 0.1 μL of SuperScript IV (Invitrogen), 0.2 μL of actinomycin D

(1000 ng/μL) (Nacalai Tesque, Kyoto, Japan), and 5.9 μL of RNase-free water] were

mixed. For the reverse transcription step, the mixture was incubated at 25 °C for 10 min,

followed by 50 min at 50 °C. SuperScript IV was inactivated by heating the mixture at

75 °C for 15 min. Subsequently, 24 μL of AMPure XP and 12 μL of 99.5% ethanol were

added, and the mixture was purified according to the manufacturer's protocol. The reverse

transcription product was eluted with 14 μL of RNase-free water. The purified DNA/RNA

hybrid solution without beads and the second-strand synthesis master mix [2 μL of 10×

Blue Buffer (Enzymatics, Beverly, MA, USA), 1 μL of dUTP/NTP mix (Fermentas,

Burlington, Canada), 0.5 μL of 100 mM DTT, 0.5 μL of RNase H (Enzymatics), 1 μL of

DNA polymerase I (Enzymatics), and 5 μL of RNase-free water] were mixed. This

mixture was incubated at 16 °C for 4 h, and then purified with 24 μL of AMPure XP

according the manufacturer's protocol. The purified dsDNA was eluted with 14 μL of

RNase-free water. Five microliters of the dsDNA solution was used in the following step.

End-repair, A-tailing, and adapter ligation were carried out using the KAPA Hyper prep

kit (KAPA BIOSYSTEMS, Wilmington, MA, USA) with 1/10× volume of the solutions

according the manufacturer's protocol. One microliter of 0.1 μM Y-shape adapter[24] was

used in the adapter ligation step for 15 min. Then, size selection of the ligation product

was performed with 5.5 μL of AMPure XP. The purified dsDNA was eluted with 10 μL

of RNase-free water and a second round of size selection was performed with 10 μL of

AMPure XP. The size-selected ligation product was eluted with 15 μL of 10 mM Tris-

HCl (pH 8.0). One microliter of uracil DNA glycosylase (UDG; Enzymatics) was added

to the size-selected ligation product. The mixture was incubated at 37 °C for 30 min to

exclude the second-strand DNA. For library amplification, 2 μL of the UDG-digested

DNA, 1 μL of 2.5 μM index primer[24], 1 μL of 10 μM universal primer[24], 0.5 μL of RNase-

free water, and 5 μL of KAPA HiFi HotStart ReadyMix (2×) (KAPA BIOSYSTEMS)

were mixed. The DNA fragments with the adapters and an index sequence were amplified

using a thermal cycler with the following program: denaturation at 94 °C for 2 min, 16

cycles at 98 °C for 10 s, 65 °C for 30 s, and 72 °C for 30 s for amplification, and 72 °C

for 5 min for the final extension. Two rounds of size selection were then performed to

remove adapter dimers with equal volume of AMPure XP to the library solution. The

purified library was eluted with 10 μL of RNase-free water. The concentration of each

library was measured by qPCR using the KAPA SYBR FAST qPCR Master Mix (2×) Kit

(KAPA BIOSYSTEMS) and the LightCycler 480 II (Roche Diagnostics, Basel,

Switzerland) to mix equal amounts of libraries. One microliter of the mixed library was

used for electrophoresis with the Bioanalyzer 2100 and the Agilent High Sensitivity DNA

kit (Agilent Technologies, Santa Clara, CA, USA) to check for quality. Sequencing of

50-bp single-ends using HiSeq 2500 (Illumina, San Diego, CA, USA) was carried out by

Macrogen (Seoul, South Korea).

Calculation and normalization of RNA-Seq count data. All obtained reads were

processed with Trimmomatic (version 0.33)[26] using the following parameters:

TOPHRED33 ILLUMINACLIP:TruSeq3-SE.fa:2:30:10 LEADING:19 TRAILING:19

SLIDINGWINDOW:30:20 AVGQUAL:20 MINLEN:40. This procedure removed

adapter sequences (ILLUMINACLIP:TruSeq3-PE.fa:2:30:10), as well as leading and

trailing low quality (Q < 19) or N bases (LEADING:19 TRAILING:19). It also trimmed

the reads when the average quality per base dropped below 20 with a 30-base wide sliding

window (SLIDINGWINDOW:30:20). Thus, trimmed reads with length > 39 nucleotides

and average Q score > 19 were obtained. These trimmed reads were then mapped to the

reference sequences of IRGSP-1.0_transcript[27,28], rRNAs, and transcripts coded in the

mitochondria and chloroplast genomes (NC_001320.1, NC_011033.1) using RSEM

(version 1.3.0)[29] and Bowtie (version 1.1.2)[30] with default parameters. Expected counts

of each gene in the RSEM outputs were used in R (version 3.4.2)[31] for the following

analysis. According to a previous study[32], we removed the effects of misassigned reads in

multiplex sequencing experiments with HiSeq 2500 as:

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$$c_{i,m,n} = \begin{cases} \tilde{c}_{i,m,n} - 0.0005 \times C_{i,n} & (\tilde{c}_{i,m,n} > 0.0005 \times C_{i,n}), \\ 0 & (\tilde{c}_{i,m,n} \leq 0.0005 \times C_{i,n}), \end{cases} \tag{1}$$

where $c_{i,m,n}$ and $\tilde{c}_{i,m,n}$ denote corrected and raw read counts of gene i in sample m in

RNA-Seq library pool n, respectively. $C_{i,n}$ denotes the sum of read counts of gene i in

RNA-Seq library pool n.

Then, read counts of gene i in sample m in different libraries were merged,

followed by read count normalization based on total read counts of transcripts except for

the antisense-oligo targets (reads per million: rpm):

$$r_{i,m} = \frac{c_{i,m}}{C_m} \times 10^6, \tag{2}$$

$$y_{i,m} = \log_2\left(r_{i,m} + 0.1\right), \tag{3}$$

where, $c_{i,m}$, $r_{i,m}$, and $y_{i,m}$ denote merged corrected read counts, normalized read counts,

and the log2-transformed value of read counts of gene i in sample m, respectively. C_m

denotes total read counts of transcripts except for the antisense-oligo targets in sample m.

Of the 926 samples from 2015 and 143 samples from 2016, we used 887 and 139

RNA-Seq transcriptome datasets, respectively, with more than 10^5 total read counts for

all genes, except for the targets of selective depletion (Supplementary Fig. 3a). We then

filtered out rarely detected genes (number of samples with read count > 0; $\leq 20\%$ of all

samples in the 2015 dataset) from the following analyses to focus on the 23,924 expressed

genes only (Supplementary Fig. 3b).

Identification of specific 'Koshihikari' and 'Takanari' single nucleotide

polymorphisms. The paired-end reads of 'Koshihikari' (SRR1630927) and 'Takanari'

(DRR018360) genomes were trimmed with Trimmomatic (version 0.33)[26] using the

parameters stated in the previous section, except for TruSeq3-PE.fa for adapter trimming.

The reads were then mapped to the reference genome (IRGSP-1.0, plastid: NC_001320.1,

mitochondrion: NC_011033.1)[27] using Bowtie2 (version 2.2.9) with option -N1, which

allows a mismatch in seed alignment [33]. By using SAMtools (version 1.3.1)[34], the reads

mapped at multiple loci and the duplicated reads, except for each representative read,

were discarded. Picard (version 2.1.0) (https://broadinstitute.github.io/picard/) was used

to add read group information to each SAM file. All SAM files of each sample were

merged and a VCF file was generated using SAMtools (version 1.3.1)[34] and BCFtools

(version 1.1)[34]. Single nucleotide polymorphisms (SNPs) were filtered using VCFtools

(version 0.1.14)[35] with option --minDP5, which indicates that all alleles supported by

more than four reads are reported. The VCF file generated was analyzed using R (version

3.4.2)[31] and the R package "vcfR" (version 1.5.0)[36]. We selected 15,117 homozygous

SNPs distinguishing 'Koshihikari' and 'Takanari' genotypes for the following analysis of

genotype confirmation.

Genotype confirmation. In previous studies, the genotypes of the CSSLs and BILs were

examined based on the genotypes provided by 141 SSR markers[14]. To confirm the

genotype of each sample, we estimated each SSR marker genotype based on the SNPs

obtained by RNA-Seq. First, to type 'Koshihikari'- or 'Takanari'-type SNPs in each

RNA-Seq sample, the trimmed RNA-Seq reads of each sample were mapped to the

reference genome as described in the previous section using Bowtie2 (version 2.2.9).

Then, the SNPs were filtered using VCFtools (version 0.1.14)[35] with option --minDP3,

which indicates that all alleles supported by more than two reads are reported. To estimate

each SSR marker genotype in our samples, 'Koshihikari'- and 'Takanari'-type SNPs were

counted from the SSR marker position until one of the counts reached "7" (Supplementary

Fig. 4a,b). In the cases where the SNPs around an SSR marker position were too sparse,

all the SNPs in ±1.5 Mbp from each SSR marker were counted (Supplementary Fig. 4a,c).

Then, each estimated SSR marker was categorized into 'Koshihikari'-type, 'Takanari'-

type, or undetermined based on most of its SNP counts (Supplementary Fig. 4a-c). Any

samples of the estimated genotypes not perfectly matching the SL1201, SL1229, or

SL1321 genotypes might be due to the low density of SNPs around the substituted SSR

markers or to doubtful SNP calls. For the samples with estimated genotypes similar to

one of these lines, imputation of estimated SSR marker genotypes was conducted

(Supplementary Fig. 4d). Finally, each sample was labeled with the genotype of the most

similar SSR marker composition. Thus, 865 samples were confirmed as 'Koshihikari',

'Takanari', or CSSLs and 11 samples as BILs in 2015, and 139 samples as 'Koshihikari',

'Takanari', or CSSLs in 2016 (Supplementary Fig. 3d).

Correlation plot of transcriptomes. Pearson's correlation coefficients of transcriptome

data for all pairwise comparisons of the 854 samples from 2015 were calculated as:

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$$\rho_{m,n} = \frac{\sum_i (y_{i,m} - \bar{y}_m)(y_{i,n} - \bar{y}_n)}{\sqrt{\sum_i (y_{i,m} - \bar{y}_m)^2}\sqrt{\sum_i (y_{i,n} - \bar{y}_n)^2}}, \tag{4}$$

where $\rho_{m,n}$ denotes the Pearson's correlation coefficient between samples m and n. Mean

log2-transformed rpms are denoted as $\bar{y}_m$ and $\bar{y}_n$, respectively.

The heatmap representing the correlations in the time-series order was drawn

with the "image.plot" function in the R package "fields" (version 9.0)[37].

Meteorological data. Data of average air temperature per 10 min at the Hirakata Weather

Station (34°48′5″N, 135°33′6″E) in 2015 for Takatsuki and at Kyotanabe Weather Station

(34°49′8″N, 135°45′6″E) in 2016 for Kizugawa were obtained from the Japan

Meteorological Agency. Data of air temperature per minute for the FIT (version 0.0.4)[10]

library were prepared by linear interpolation of the data per 10 min with the "approxfun"

function in R (version 3.4.2)[31]. Data of global solar radiation per minute at the Osaka

Weather Station (34°40′9″N, 135°37′51″E) in 2015 for Takatsuki and at the Nara Weather

Station (34°40′4″N, 135°50′2″E) in 2016 for Kizugawa were also obtained from the Japan

Meteorological Agency.

Conversion of age to scaled age. To consider the difference in plant's age at heading

among genotypes, we defined scaled age. Under the assumption that rice ages linearly

before heading at various speeds among genotypes, flowers at scaled age 85, and ages at

the same speed after flowering, the scaled age of sample m was defined as:

382
$$a_m = \begin{cases} 85 \dfrac{d_m - d_m^S}{d_m^F - d_m^S} & (d_m \le d_m^F), \\ d_m - d_m^F + 85 & (d_m > d_m^F), \end{cases} \tag{5}$$

where d_m is the sampling date, d_m^S is the seeding date, and d_m^F is the flowering date

of each line with genotype and seeding date identical to that of sample m.

Calculation of precision weights for each observed normalized read count (log2rpm).

Precision weights were utilized in the training step of FIT (version 0.0.4)[10] to deal with

over-dispersion of RNA-Seq data[10]. Precision weight for each observed log2rpm value

was calculated based on the residuals from the smoothed time-series as in the voom

method[10,38] (Supplementary Fig. 14). First, for each gene, spline interpolation against

time and observed log2rpm was conducted with the "splinefun" function in R, to infer the

function value $s_{iG}(t)$ for interpolation of observed log2rpm of gene i at time t in $G \in$

$\{K, T\}$, where K and T denote 'Koshihikari' and 'Takanari'-background genotypes,

respectively (Supplementary Fig. 14a). Then, residual standard deviation of gene i in G-

background genotype $\sigma_{i,G}$ was calculated as:

395
$$\sigma_{i,G} = \sqrt{\frac{1}{|M_G|} \sum_{m \in M_G} \left(s_{i,G}(t_m) - y_{i,m} \right)^2}, \tag{6}$$

396   where t_m denotes sampling time of sample m. M_G and $|M_G|$ denote indices of samples

397   whose background genotypes are G and its size, respectively.

398   The mean log2-transformed read count of gene i in G-background genotype was then

399   calculated as:

400
$$\bar{c}_{i,G} = \frac{1}{|M_G|} \sum_{m \in M_G} y_{i,m} + \frac{1}{|M|} \sum_{m \in M} \log_2 C_m + 1 - \log_2 10^6 \,, \tag{7}$$

where M denotes sample indices in both background genotypes.

A LOWESS curve[39] was then fitted to the square-root of residual standard deviations as

a function of mean log2-transformed read counts, resulting in $l()$, which represents the

trend between log2-transformed read counts and square-root of residual standard

deviations (Supplementary Fig. 14b).

Similar to $\bar{c}_{i,G}$, smoothed log2rpm of gene i in sample m was converted to

smoothed log2-transformed read counts of gene i in sample m as:

408
$$\hat{c}_{i,m} = s_{i,G_m}(t_m) + \log_2 C_m + 1 - \log_2 10^6 \,, \tag{8}$$

409   where G_m denotes the background genotype of sample m.

410   Finally, precision weight for observed log2rpm of gene i in sample m was calculated as

411   (Supplementary Fig. 14c):

412
$$w_{i,m} = \frac{1}{l(\hat{c}_{i,m})^4} \,. \tag{9}$$

Prediction model development for 'Koshihikari' and 'Takanari'. A prediction model

for gene expression dynamics in 'Koshihikari' and 'Takanari' was developed with the R

package "FIT" (version 0.0.4)[10]. In this section, we describe how "FIT" developed a

prediction model.

There are several strategies for the integrative analysis of the transcriptome and

meteorological data to investigate environmental responses in field conditions[40–44]. In the

FIT package, we employed the statistical modeling approach for a comprehensive

relationship among these data and for understanding the effects of age and circadian clock

of plants. We used the transcriptome data in combination with the meteorological data,

namely air temperature and global solar radiation, which were measured at a

meteorological station near the field from where the samples were acquired. Because

there was no remarkable disease or pest damage during cultivation and no difference in

fertilization among samples, the model did not include these effects.

The log2rpm of the normalized expression for gene i of sample m was described by a

simple linear model:

428
$$y_{i,m} = \alpha_i + \boldsymbol{x}_{i,m}^T \boldsymbol{\beta}_i + \epsilon_{i,m}, \tag{10}$$

where α_i, $\boldsymbol{x}_{i,m}$, and $\boldsymbol{\beta}_i$ are a constant term, a vector of eight explanatory variables, and

regression coefficients, respectively. The third term $\varepsilon_{i,m}$ is the independently and

identically distributed noise drawn from a Gaussian distribution. The explanatory

variables were plant's scaled age, circadian clock, optional variable, response to

environmental stimuli, and the interactions between scaled age and circadian clock and

between scaled age and response to environmental stimuli. The optional variable was not

used in this study. The plant's scaled age corresponded to the scaled number of days after

transplanting. See section "Conversion of age to scaled age" for the definition of scaled

age. The values of the scaled age were normalized to mean 0 and variance 1. The circadian

clock, with the peak at an arbitrary time, was represented by two variables corresponding

to its cosine and sine components:

440
$$\cos\left(\frac{2\pi t_m}{24}\right),$$
$$\sin\left(\frac{2\pi t_m}{24}\right),$$
(11)

where t_m is the time at which sample m was obtained. Similarly, the interaction between

scaled age and circadian clock of a plant was represented by two variables.

Although we employed a linear model for simplicity, the actual gene expression can

exhibit more complex responses. The complexity of fluctuations in gene expression is

mainly explained by the nonlinearity of the response to the environmental stimuli, which

is the cumulative sum of the nonlinearly transformed environmental stimuli for a given

period. Such cumulative sum of temperature, for example, is designated as cumulative

temperature and it is widely used for the prediction of reproduction timing and vegetation

in ecology and of harvest timing in agriculture[45,46]. By applying this to gene expression,

the environmental response of gene i in sample m can be described as:

$$\hat{e}_{i,m} = \sum_{\tau=t_m-p_i}^{t_m} g_i(\tau) \max\left(0, \tanh\left(\pm \exp\left(\gamma_i^{(f)}\right)\left(e_\tau - \theta_i^{(f)}\right)\right)\right) \times \sqrt{\exp\left(-2\gamma_i^{(f)}\right) + 1}, (12)$$

where p_i is the period during which gene expression was affected by an environmental

stimulus, e_τ is the value of a meteorological parameter at time τ normalized to mean 0 and

variance 1, $\theta_i^{(f)}$ is the response threshold to the stimulus, and $\gamma_i^{(f)}$ is the parameter that

controls the shape of the response (Supplementary Fig. 15). Here, the meteorological data

consisted of air temperature and global solar radiation. Of these parameters, the one

yielding the highest likelihood was selected for each gene in the optimization procedure.

In the limit $\gamma_i^{(f)} \to \infty$, a gene responds to an environmental stimulus in a dose-

independent manner, that is, the response is constant if a meteorological parameter

exceeds the threshold (Supplementary Fig. 15). Conversely, as this parameter approaches

minus infinity, the response also approaches the dose-dependent manner (Supplementary

Fig. 15). The sign of the value in the hyperbolic function determines whether the gene

responds to an environmental stimulus larger than (positive) or smaller than (negative)

the threshold. This sign was decided in the optimization procedure based on the parameter

likelihood. The last term $\sqrt{\exp\left(-2\gamma_i^{(f)}\right) + 1}$ ensures that the scale of the response is

largely independent of this parameter. The function $g_i(\tau)$ is a gate function that explains

the time-of-day-specific environmental responses. A gate function is defined as:

468
$$g_i(\tau) = n_i \left(\tanh \left(\exp \left(\gamma_i^{(g)} \right) \left(\cos \left(2\pi(\tau - \psi_i) \right)/24 - \theta_i^{(g)} \right) \right) \right), \quad (13)$$

where $\gamma_i^{(g)}$ controls the shape of the gate as $\gamma_i^{(f)}$ in the environmental response, $\theta_i^{(g)}$

determines the opening length of the gate, and $n_i()$ is the function which normalizes an

input value $[0,1]$. This gate function approaches a sinusoidal and rectangular wave with

the small and large values of $\gamma_i^{(g)}$, respectively. If the value of $\theta_i^{(g)}$ is smaller than -1,

the gate is a constant value in the limit $\gamma_i^{(g)} \to \infty$, that is, the gate is always open

(Supplementary Fig. 15).

It can be considered that some explanatory variables might contribute to neither the

explanation nor the prediction of the expression of many genes. Thus, we simultaneously

preformed the optimization and variable selection of the regression coefficients by using

group lasso[47]. Let I be the index set for scaled age, environmental response, and the

interaction between scaled age and environmental response. The cost function to be

minimized is thus defined as:

481
$$L(\boldsymbol{\beta}) = \sum_{m \in M} \left(\hat{y}_{i,m} - y_{i,m} \right)^2$$

482
$$+ \lambda \left(\sum_{k \in I} \zeta_k |\beta_k| + \zeta_c \sqrt{\beta_{cos}^2 + \beta_{sin}^2} + \zeta_{dc} \sqrt{\beta_{dcos}^2 + \beta_{dsin}^2} \right), (14)$$

where λ and $\boldsymbol{\zeta}$ are the regularization parameter and the adaptive weights, respectively.

The values of the adaptive weights are decided as in the adaptive lasso[48] or the adaptive

group lasso[49]. Further, the adaptive weights were used to absorb differences in the degrees

of freedom by multiplying the covariates of the environmental response and interaction

between the environmental response and scaled age by seven, as seven free parameters

were considered. We selected the value of λ as the largest value for which the cross

validation (CV) error was smaller than the sum of the minimum CV error and its standard

error.

The values of the parameters related to environmental responses, which are

$p_i, \theta_i^{(f)}, \gamma_i^{(f)}$, and the sign of the value in the hyperbolic function in Eq. (12), and

$\theta_i^{(g)}, \gamma_i^{(g)}$, and ψ_i in Eq. (13), are selected by the Nelder-Mead algorithm[50]. Because the

likelihood function is complex and has multiple local maxima, a grid search was

performed before optimization. In each step of the optimization by the Nelder-Mead

algorithm, we performed an ordinary least square regression instead of group lasso (Eq.

(14)) to reduce the computational cost.

'Koshihikari' and 'Takanari'-prediction models were developed based on time-

series RNA-Seq data of the 'Koshihikari'-background CSSLs and 'Koshihikari' and of

the 'Takanari'-background CSSLs and 'Takanari' in 2015, respectively (Supplementary

Fig. 2b). Values of the expressed genes (a matrix of log2rpm for each sample), sample

attributes (a matrix of sampling year, month, day, hour, and min and scaled age for each

sample), weather data (a matrix of air temperature and global solar radiation at every

minute in 2015), and precision weights for the matrix of log2rpm of each sample

(Supplementary Fig. 16) were prepared. Details on the parameters for 'FIT' are listed in

https://github.com/naganolab/Rice_edQTL-analysis_and_edQTL-based-

prediction/analysis.R. As a result, the functions $f_{i,G}()$ were yielded for gene i. A log2rpm

of gene i of sample m was predicted as (Supplementary Fig. 16):

$$\hat{y}_m = f_{i,G_m}(t_m, a_m, \boldsymbol{e}),\qquad(15)$$

where $\boldsymbol{e}$ denotes meteorological data (air temperature and global solar radiation) at every

minute.

Detection of genes with expression dynamics polymorphism. To increase detection

power in multiple tests, we focused on genes with expression dynamics polymorphism.

Euclidean distances of gene i between 'Koshihikari'- and 'Takanari'-prediction models

(d_i) were calculated as:

$$d_i = \sqrt{\sum_{m \in M} \left(f_{i,K}(t_m, a_m, \boldsymbol{e}) - f_{i,T}(t_m, a_m, \boldsymbol{e}) \right)^2}.\qquad(16)$$

As a result, 3696 genes with $d_i > 40$ were defined as genes with expression dynamics

polymorphism (Supplementary Fig. 3e).

Detection of edQTL. According to background genotypes, the sum of residual errors

between the observed log2rpm and predicted log2rpm of gene i in the 'Koshihikari'- or

'Takanari'-models (E_i) were calculated as (Supplementary Fig. 3a):

$$E_{i,m} = \left| \hat{y}_{i,m} - y_{i,m} \right|, \tag{17}$$

$$E_i = \sum_{m \in M} E_{i,m}, \tag{18}$$

where E_{im} denotes the residual error between the observed log2rpm and predicted

log2rpm of gene i in sample m on the assumption that the 'Koshihikari'- or 'Takanari'-

type gene expression dynamics shown by sample m is determined based on background

genotype G_m. Similarly, on the assumption that edQTL for gene i exist around SSR marker

g, the sum of residual errors between the observed log2rpm and predicted log2rpm of

gene i in the 'Koshihikari'- or 'Takanari'-models $\tilde{E}_{i,g}$ were calculated as

(Supplementary Fig. 3a):

$$\tilde{E}_{i,m,g} = \left| f_{i,Q_{g,m}}(t_m, a_m, \boldsymbol{e}) - y_{i,m} \right|, \tag{19}$$

$$\tilde{E}_{i,g} = \sum_{m \in M} \tilde{E}_{i,m,g}, \tag{20}$$

where $\tilde{E}_{i,m,g}$ denotes the residual error between the observed log2rpm and predicted

log2rpm of gene i in sample m and $Q_{g,m}$ denotes genotype of SSR marker g in sample

m. Thus, $f_{i,Q_{gm}}(t_m, a_m, \boldsymbol{e})$ denotes predicted log2rpm in sample m assuming the

existence of edQTL for gene i around SSR marker g. The improvement in residual errors

assuming that edQTL for gene i exists around g was therefore calculated as

(Supplementary Fig. 3a,b):

$$I_{i,g} = E_{i,m} - \tilde{E}_{i,m,g}. \tag{21}$$

To evaluate the statistical significance of the improvement in the sum of residual

errors, 1000 permutations of samples with 'Koshihikari'- and 'Takanari' genetic

background were performed (Supplementary Fig. 6e). For each permutation, the

improvement in residual errors on the assumption that edQTL for gene i exist around SSR

marker g was calculated (Supplementary Fig. 6b). Then, we generated the null

distribution of improvements in residual errors by fitting log-normal distribution to

obtained improvements using the fitdist function in the R package "fitdistrplus"[51], with

the fitting method of quantile matching of 70% and 1% (Supplementary Fig. 6c). For

almost all genes with expression dynamics polymorphism, the 99.9 percentiles of the

fitted log-normal distributions were higher than (2863/3696 genes) or almost the same as

those of the distribution obtained by the permutations (Supplementary Fig. 6d), indicating

that the p-value obtained from log-normal distribution was conservative. Thus, the p-

value for the improvement in the sum of residual errors for gene i on the assumption that

an edQTL affecting gene expression dynamics of gene i exists around SSR marker g was

obtained based on the null distribution.

The adjustment for multiple comparisons against p-values was performed with

the Benjamini-Hochberg method[52] using the p.adjust function in R. Finally, we defined

the markers with the smallest adjusted p-value (< 0.05) in each peak for each gene as the

edQTL affecting each gene expression dynamics.

edQTL-based prediction of gene expression dynamics (edQTL model). To predict

gene expression with edQTL information, the 'Koshihikari'- or 'Takanari'-model was

selected based on the genotype of the edQTL for edQTL-affected genes and of genetic

backgrounds for the other expressed genes. Let $\mathcal{G}_i$ denote the set of edQTLs affecting

gene i. We selected the model for prediction of log2rpm of gene i in sample m as:

$$g_{i,m} = \begin{cases} \tilde{G}_m, & (\exists g \in \mathcal{G}_i, Q_{g,m} \neq G_m), \\ G_m, & \text{otherwise}, \end{cases} \tag{22}$$

565   where $\widetilde{G}_m$ denotes the opposite genotype to that of the genetic background in sample m,

566   that is, $\tilde{G}_m = K$, if sample m is 'Takanari'-background, and vice versa. Then, predicted

567   log2rpm of gene i in sample m was calculated as:

$$\hat{y}_{i,m}^{\text{ED}} = f_{i,g_{i,m}}(t_m, a_m, \boldsymbol{e}). \tag{23}$$

569   **Statistical test for the accuracy of gene expression dynamics prediction in BILs**. The

570   sums of prediction errors of all genes affected by each edQTL for the BILs based on the

571   'Koshihikari' model (E_K), 'Takanari' model (E_T), and edQTL model (E^{ED}) were

572   calculated as:

$$E_G = \sum_{i \in I^{\text{ED}}, m \in M^{\text{B}}} \left| f_{i,G}(t_m, a_m, \boldsymbol{e}) - y_{i,m} \right|, \tag{24}$$

$$E^{\text{ED}} = \sum_{i \in I^{\text{ED}}, m \in M^{\text{B}}} \left| \hat{y}_{i,m}^{\text{ED}} - y_{i,m} \right|, \tag{25}$$

where M^{B} denotes the indices of BIL samples and I^{ED} denotes the genes affected by the edQTL.

To evaluate the statistical significance of improvement in the sums of prediction errors, 10,000 permutations were performed for the markers in each BIL genome. In each permutation, the sums of prediction errors of all genes affected by the edQTL for the BILs were calculated with the edQTL model. Then, the value corresponding to percentile 0.1 of the sums of prediction errors was calculated.

Code availability. The scripts used in this study are available from https://github.com/naganolab/Rice_edQTL-analysis_and_edQTL-based-prediction.

Data availability. All datasets generated and/or used in this study are available from PRJDB7234.

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Figures

Figure 1 | Concept and workflow of expression dynamics quantitative trait loci (edQTL) detection. a, Conceptual differences in the scopes of expression quantitative trait loci (eQTL) and edQTL. **b**, Workflow of edQTL detection and its evaluation using chromosome segment substitution lines (CSSLs). **c**, Summary of the sampling design for edQTL detection. The left panels show plots of meteorological data [air temperature (Temp., °C) and global solar radiation (Rad., kJ m^{-2} min^{-1})] in Takatsuki from May to September in 2015. Vertical red lines represent the sampling time points. **d**, Pearson's correlations among the 854 samples used for developing the prediction model based on expression data of 23,294 expressed genes. White lines indicate the border of each

739   bihourly sampling set.

Figure 2 | Expression dynamics quantitative trait loci (edQTL) detection revealed

cis- **and** *trans*-**edQTL that were involved in environmental responses. a, c, e,**

Observed and predicted expressions (log$_2$rpm) of *Os09g0343200* (**a,c**) and

Os01g0537250 (**e**) in Takatsuki in 2015. The predicted expressions were calculated using

'FIT' based on scaled age and environmental information (time, air temperature, and

global solar radiation).Blue and red/pink lines indicate predicted expression levels in

'Koshihikari' and 'Takanari' in transplant sets 1, 2, 3, and 4, respectively. Blue and

red/pink points indicate the expression level obtained by RNA-sequencing for samples in

transplant sets 1, 2, 3, and 4 of individuals with 'Koshihikari'-background genotype (BG)

and 'Takanari'-BG in (**a**) or 'Koshihikari'- and 'Takanari'-type edQTL in (**c,e**),

respectively. In (**c,e**), strong colored points emphasized by the arrowheads indicate

samples harboring edQTL different from their BGs, and light-colored points indicate the

samples harboring edQTL identical to their BGs. The upper gray bar indicates dark

periods (global solar radiation < 0.3 kJ m^{-2} min^{-1}). **b,f**, edQTL regulating *Os09g0343200*

(**b**) and *Os01g0537250* (**f**). **d**, Position of the 1675 edQTL are shown as red bars (false

discovery rate $= 0.05$). X-axis and Y-axis represent the positions of markers with edQTL

and the positions of genes influenced by edQTL, respectively.

Figure 3 | Expression dynamics quantitative trait loci (edQTL)-based prediction of

gene expression dynamics under different environments and in cultivars with more

complex genotypes than the chromosome segment substitution lines (CSSLs). a,

Examples of prediction of expression dynamics in Kizugawa in 2016 based on

environmental information and edQTL in transplant set 1. Points in intense colors

emphasized by arrowheads indicate samples harboring edQTL different from their

background genotypes and light colors indicate samples harboring edQTL identical to

their background genotypes. The upper gray bars indicate dark periods (global solar

radiation < 0.3 kJ m^{-2} min^{-1}). **b,** Effects of edQTL over the prediction of gene expression

dynamics in Kizugawa in 2016. Genes influenced by edQTL are in order of position on

the chromosomes along the horizontal axis. **c,d,** Prediction accuracy of edQTL model for

HP-a (**c**) and HP-b (**d**). The blue, red, and orange vertical lines indicate the sums of

prediction errors based on 'Koshihikari', 'Takanari', and edQTL models. The histogram

shows the distribution of the sums of prediction errors based on the edQTL model in

10,000 permutations of markers in HP-a or HP-b genomes. The dashed vertical line

indicates the 0.1% percentile of the distribution. **e,** edQTL for *Os03g0388300* and

genotypes of HP-a and HP-b. Dark blue points indicate significant 'Koshihikari'-type

markers. **f,** Prediction of *Os03g0388300* expression in HP-a and HP-b. Intense color lines

are applied models for HP-a and HP-b.

Supplemental Tables

Table S1 SSR marker composition in the rice lines used in the present study

"Position" indicates the position of each SSR marker on the chromosomes in the

IRGSP1.0 reference genome. A and B indicate 'Koshihikari'- and 'Takanari'-type alleles,

respectively.

Table S2 Sample attributes used in this study

"TransPlantSet" indicates the transplant sets in each year. Details of seeding and

transplant dates in each set are described in the methods. "Year", "month", "day", and

"hour" columns indicate the sampling date. The "LineName" column indicates the

predicted genotypes based on SNPs detected in each RNA-Seq. The "field" column

indicates the location of the fields.

Table S3 *Oryza sativa* genes and their expressions

The "Expression" column indicates whether the genes are "expressed genes" or not. "Yes"

indicates that "expressed genes" were detected in more than 20% of all samples. The

"Polymorphism" column indicates whether the genes show expression dynamics

polymorphism between 'Koshihikari' and 'Takanari' or not. "Yes" indicates that the

Euclidian distances of predicted expression dynamics between 'Koshihikari' and

'Takanari' were more than 40. The "edQTL", "cis.edQTL", and "trans.edQTL" columns

indicate whether the genes have the designated edQTL or not.