

34
35
36
37
38
39
40
41
42
43
44
45
46
47
48
49
50
51
52
53
54
55
56
57
58
59
60
61
62
63
64
65
66

A crucial question is how evolution of resistance depends on antibiotic concentration. A major step towards answering this question was the formulation of the mutant selection window (MSW) hypothesis, which proposes that resistance arises within a predictable range of antibiotic concentrations [2] [3]. Specifically, the lower boundary of this window is the minimum inhibitory concentration (MIC) of the sensitive strain, i.e. the concentration that is sufficient to prevent its growth. The upper boundary, the so-called “mutant prevention concentration”, was originally defined as the minimal concentration at which no colonies are formed when plating a culture of at least 10^{10} cells [2] [3], and gradually became equated with the MIC of the most resistant strain available [4] [5]. The MSW hypothesis has become an important paradigm in the field and a framework for suggesting dosing strategies to avoid emergence of resistance [4] [5] [6].

In recent years, it has been recognized both on theoretical grounds and with accumulating empirical evidence that higher resistance can also be selectively favoured below the MIC of a sensitive or less-resistant strain [7] [8] [9] [10]. This selection occurs because sub-MIC concentrations of antibiotic can already be sufficiently detrimental to render the sensitive strain less fit than the resistant strain. Indeed, the “minimal selective concentration” (MSC) favouring resistance can be as low as a few percent of the sensitive strain’s MIC [9]. Thus, the window of antibiotic concentrations over which resistance can emerge is potentially very wide.

Selection is expected to operate efficiently when both sensitive and resistant cell populations are reasonably large, resulting in an increase in frequency of the fitter strain (i.e. of the resistant strain at concentrations beyond the MSC). However, a crucial distinction must be drawn between situations in which resistance is already present at appreciable frequency, and cases in which it emerges *de novo* [11]. If a resistant strain is not already present, it must first arise (e.g. by mutation or horizontal gene transfer), and then increase in frequency from a single cell to a sizeable population. Although the spread of a common resistant strain is a more or less deterministic process driven by selective differences, the establishment of an initially rare resistant strain is highly stochastic [11]. That is, due to chance events on the individual cell level (division and death), a resistance mutation present in only a few cells may be lost even when it is selectively favoured. Indeed, predicting the probability that beneficial mutations escape stochastic loss when rare is a classical problem in population genetics [12].

67 While a large body of work has addressed the question of antibiotic concentrations at which
68 resistance is selectively favoured, it is largely unknown how antibiotic concentration impacts the
69 stochastic establishment phase of *de novo* resistant mutants. Theory predicts that the probability of
70 establishment gradually decreases as cell division rate decreases or death rate increases, even while
71 the expected population net growth rate remains positive, albeit reduced [12]. Thus, if antibiotics
72 have detrimental effects on resistant cells below the resistant strain's MIC, it is possible that small
73 initial populations of cells are stochastically lost and thus establishment fails, despite being within
74 the MSW.

75

76 We set out to test this idea in the bacterium *Pseudomonas aeruginosa* challenged with the
77 aminoglycoside antibiotic streptomycin. We used a strain carrying a plasmid that confers
78 streptomycin resistance through production of an enzyme (aadA5) that deactivates the antibiotic
79 intracellularly. To study stochastic establishment, we tested growth in hundreds of cultures
80 inoculated with resistant cells at extremely low density (around a single cell per culture). Consistent
81 with an important role for stochasticity, under identical culture conditions some replicate cultures
82 grew and others did not. By fitting a mathematical model to these data, we were able to quantify
83 the probability that a single resistant cell establishes a large surviving population across a range of
84 antibiotic concentrations. This establishment probability began to decrease at concentrations below
85 the resistant strain's MIC (MIC_R), suggesting that antibiotics are exerting detrimental effects on
86 resistant cells. This decline occurred at surprisingly low antibiotic concentrations: significant effects
87 were seen as low as $1/16 \times MIC_R$, and we estimate that establishment from single cells succeeds less
88 than 5% of the time at $1/8 \times MIC_R$. We show that these results are consistent with a simple model of
89 how inoculum size affects population growth, whereby each individual cell has a small, independent
90 chance of establishment. Collectively, our results imply that *de novo* emergence of resistance may
91 be unlikely throughout a substantial portion of the mutant selection window.

92

93

94 **Results**

95

96 ***Resistance is selectively favoured over a wide range of streptomycin concentrations***

97

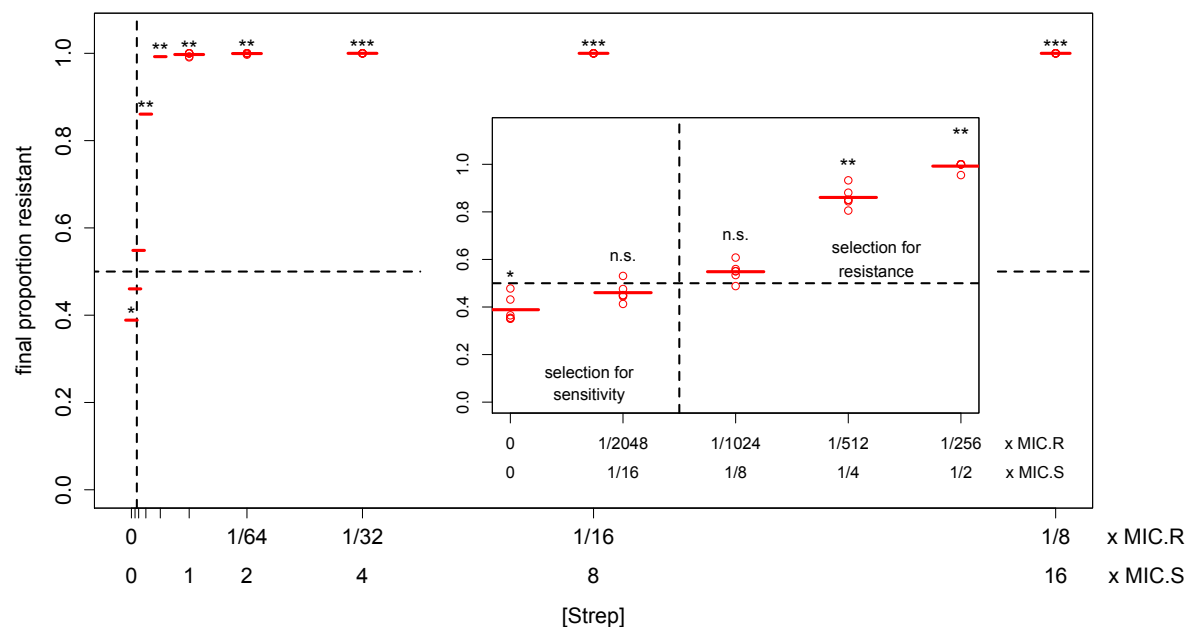
98 We measured the impact of antibiotic exposure on selection for resistance by directly competing
99 fluorescently labelled sensitive and resistant strains across a gradient of streptomycin
100 concentrations, ranging from $1/16$ to 16 times the MIC of the sensitive strain ($MIC_S = 16 \mu\text{g/ml}$; [13]),

101 corresponding to 1/2048 to 1/8 times the MIC of the resistant strain ($MIC_R = 2048 \mu\text{g/ml}$; [13]).
 102 Strains were inoculated at approximately equal starting frequencies, with a total density around $5 \times$
 103 10^5 CFU/ml, and their final frequencies were evaluated by flow cytometry after 24h (**Methods** and
 104 **Suppl. Fig. 1**). An increase in resistant frequency to significantly above 50% thus indicates selection
 105 for resistance. We were unable to quantitatively estimate the selection coefficient of the resistant
 106 strain, because virtually no sensitive cells were recovered after culturing with streptomycin above
 107 MIC_S , leading to 100% final resistance frequency (**Fig. 1**, **Suppl. Fig. 2** and **Suppl. Table 1**). Strikingly,
 108 resistance was still strongly selected as low as $1/4 \times MIC_S$ (two-sided t-test: $p = 5e-6$), while
 109 sensitivity was only significantly favoured in the complete absence of streptomycin ($p = 4e-3$). Based
 110 on the point at which final resistance frequency crosses the 50% mark, we estimate that the MSC of
 111 streptomycin lies between $1/16$ and $1/8 \times MIC_S$, in agreement with previous results for these strains
 112 [13].

113

114 These results demonstrate that the resistant strain, once present at an appreciable population size,
 115 is expected to spread by outcompeting the sensitive strain over a wide range of streptomycin
 116 concentrations, ranging from well below ($1/4x$ or less) to well above (at least $16x$) the MIC of the
 117 sensitive strain (MIC_S).

118



119

120 **Figure 1: Evaluating selection for resistance by competition assays across streptomycin concentrations.** Concentration on
 121 the x-axis is scaled by the previously measured MIC values of these strains (sensitive: $MIC_S = 16 \mu\text{g/ml}$; resistant: $MIC_R =$
 122 $2048 \mu\text{g/ml}$; [13]). Cultures were inoculated with a 1:1 volumetric mixture of overnight cultures of the sensitive (DsRed)
 123 and resistant (YFP) strains, at total density similar to a standard MIC test. After 24h, cultures were sampled by flow
 124 cytometry to estimate the proportion of each strain (**Methods**). The horizontal dashed line at 0.5 indicates the starting

125 proportion, which would be maintained if resistance were selectively neutral. Final proportions falling below this line
126 indicate selection favouring the sensitive strain, while those above this line indicate selection favouring the resistant strain.
127 The vertical dashed line indicates the approximate position of the minimum selective concentration (MSC). The inset shows
128 a zoomed-in version at low streptomycin concentration, on a log scale. Points (shown only in the inset for lower
129 concentrations) represent six biologically independent replicates at each concentration, with line segments indicating their
130 mean (see also **Suppl. Table 1**). Asterisks indicate that the mean final proportion of the resistant strain significantly differs
131 from 0.5 in a two-sided t-test at each streptomycin concentration, with a Bonferroni correction for multiple testing (n.s.: p
132 $> 0.05/7$; *: $p = 4e-3$; **: $p \leq 5e-6$; ***: lack of variation among replicates precludes a t-test).

133

134

135 ***Outgrowth of resistance from single cells is significantly inhibited within the MSW***

136

137 Although a wide range of streptomycin concentrations select for already-established resistant
138 strains, it does not follow that streptomycin resistance will emerge *de novo* at these concentrations.
139 To test how antibiotic exposure affects establishment of the resistant strain, which we define as the
140 growth of a large population starting from a single cell, we conducted a large-scale “seeding
141 experiment” (see also [14]). In this experiment (**Fig. 2**), we inoculated a large number of cultures at
142 very low densities, averaging <1 to around 3 cells per culture. Importantly, the number of cells
143 inoculated into each culture is random, and can be described by a Poisson distribution (**Suppl. Fig.**
144 **3**). One implication of this protocol is that many cultures are not inoculated with any cells; our
145 approach is to account for this variation statistically. We conducted seeding experiments at three
146 different inoculation densities and across a range of streptomycin concentrations (up to $1/8 \times \text{MIC}_R$).
147 We then scored the number of cultures showing growth based on reaching a threshold optical
148 density up to 3 days post-inoculation.

149

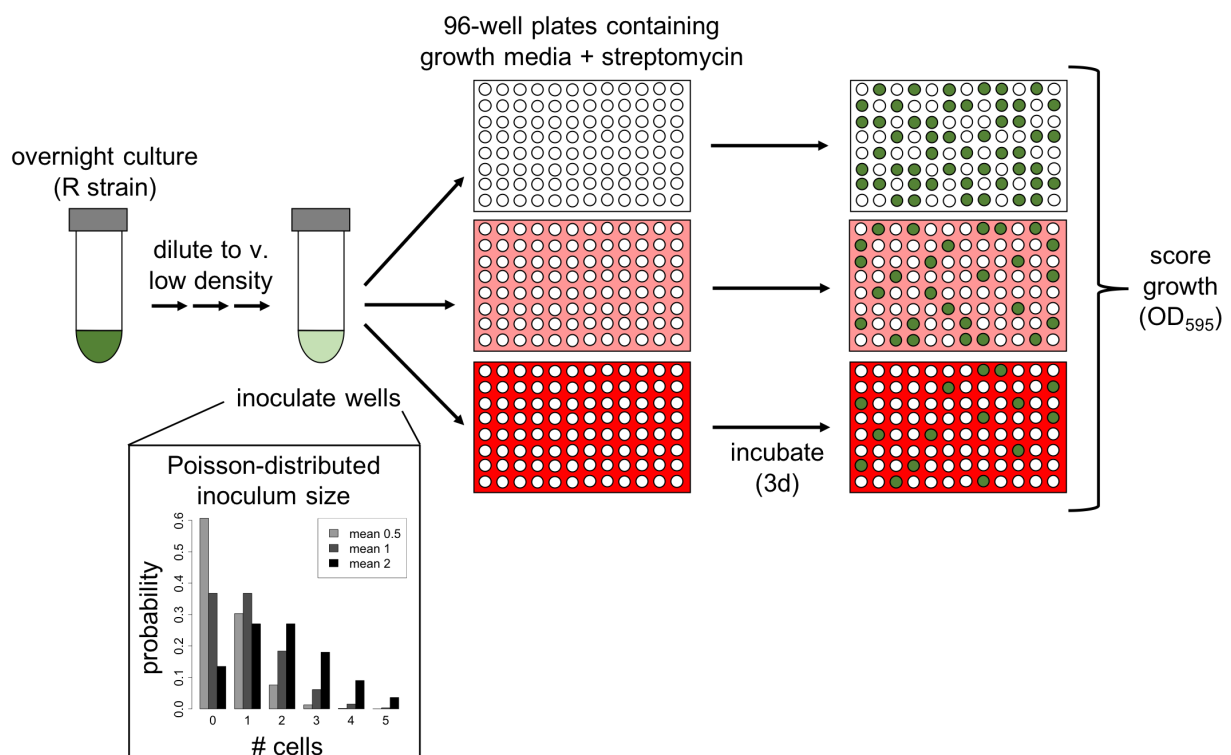
150 Cell division and death in individual bacterial cells are probabilistic events, with the consequence
151 that a single cell will not necessarily give rise to a successful lineage of cells. This chance of stochastic
152 loss applies even under benign culture conditions, but is expected to be exacerbated by stressors
153 such as antibiotic exposure that reduce the probability of successful cell division compared to cell
154 death. In the context of our experiment, a culture could fail to grow either because the inoculum did
155 not contain any cells, or because cells in the inoculum failed to give rise to a surviving lineage. To
156 discriminate between these possibilities, we fit a stochastic model to infer the per-cell probability of
157 establishment (i.e. probability that a single cell produces a large number of surviving descendants
158 leading to detectable growth) from the number of replicate cultures showing growth, across
159 streptomycin concentrations (**Methods**). We cannot estimate an absolute establishment probability
160 because this quantity cannot be separated from the effect of inoculum size: inoculating on average

161 two cells, each with 50% chance of establishment, will give equivalent results to one cell with 100%
162 chance of establishment. Therefore, all results are expressed in terms of relative per-cell
163 establishment probability, denoted \tilde{p}_c , which is normalized by the result in streptomycin-free
164 culture medium. Thus, $\tilde{p}_c=1$ by definition in streptomycin-free conditions, while we expect
165 streptomycin treatment to yield $\tilde{p}_c \leq 1$; however, values larger than one can arise due to sampling
166 error. Since the distribution of inoculum size is the same in streptomycin-treated and streptomycin-
167 free cultures, any significant differences in estimated \tilde{p}_c can be attributed to effects of streptomycin
168 on the dynamics of resistant cells.

169

170 Exposure to very low concentrations of streptomycin (up to $1/32 \times \text{MIC}_R$) had no detectable impact
171 on the probability of establishment of a single resistant cell, but this probability declined sharply as
172 concentration increased (**Fig. 3** and **Suppl. Table 2**). Strikingly, \tilde{p}_c is already significantly reduced by
173 $1/16 \times \text{MIC}_R$ streptomycin and is only marginally above zero (3-5%) at $1/8 \times \text{MIC}_R$. Thus, the chance of
174 *de novo* establishment of the resistant strain is nearly abolished under antibiotic concentrations well
175 below its MIC, where resistance is strongly favoured in competition with the sensitive strain
176 (compare **Fig. 3** to **Fig. 1**).

177



178

179 **Figure 2: Design of seeding experiment to estimate establishment probability.** An overnight culture of the resistant strain
180 is highly diluted and used to inoculate 96-well plates containing growth media (LB) with streptomycin at various
181 concentrations (higher concentration indicated by darker red). The number of cells inoculated per well follows a Poisson

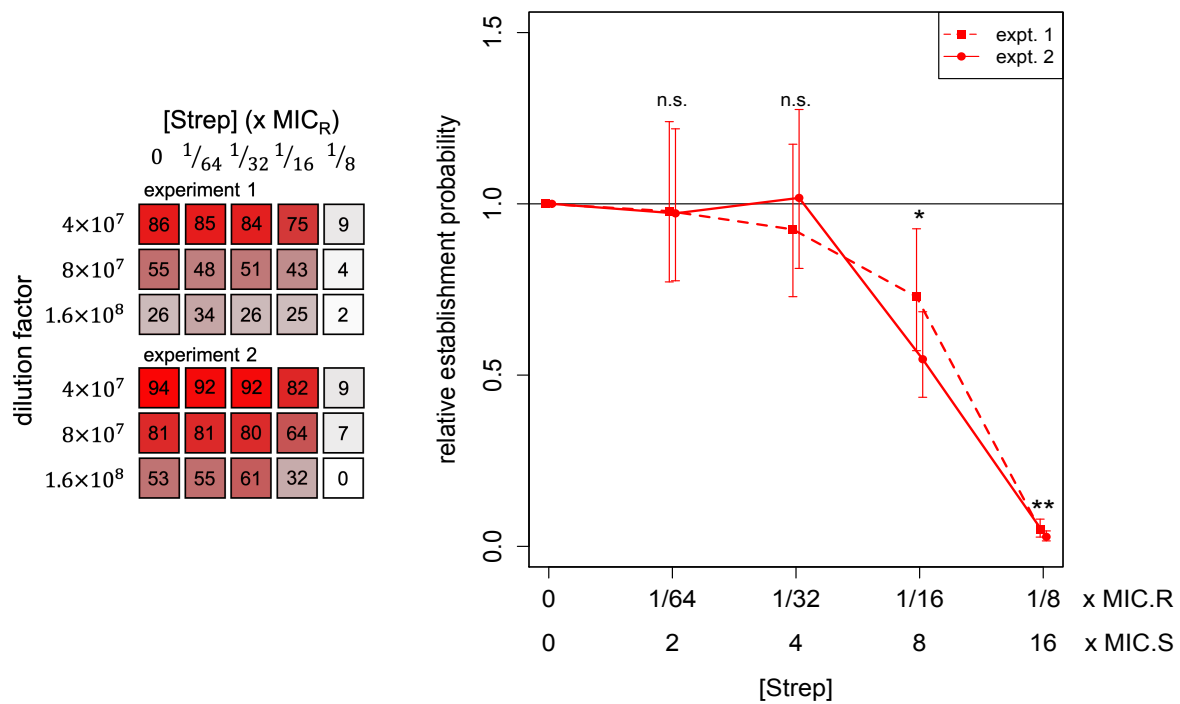
182 distribution (examples illustrated in the inset plot for mean inoculum size of 0.5, 1, or 2 cells per well). Plates are incubated
 183 up to 3d and optical density is measured to score growth in wells ($OD_{595} > 0.1$; indicated by dark green colouring). The final
 184 number of replicate cultures showing growth at each inoculum size and streptomycin concentration is used to estimate the
 185 per-cell establishment probability by fitting a mathematical model (**Methods**).

186

187

188 (a)

(b)



189

190 **Figure 3: Establishment probability of single resistant cells, estimated from seeding experiments. (a)** A visual
 191 representation of the growth data, indicating the number of replicate cultures (out of 96) that grew in each test condition.

192 **(b)** The estimated relative per-cell establishment probability \tilde{p}_c of the resistant strain, scaled by the probability in
 193 streptomycin-free medium, plotted as a function of streptomycin concentration, scaled by the MICs of the resistant and
 194 sensitive strains (MIC_R and MIC_S , respectively). Results are shown for two separate experiments. Plotted points indicate the
 195 maximum likelihood estimate of \tilde{p}_c and error bars indicate the 95% confidence interval, using the fitted model selected by
 196 the likelihood ratio test (Model B' in experiment 1 and Model C' in experiment 2, both of which pool data across three
 197 inoculation densities; see **Suppl. Text**). Significance of the streptomycin effect is determined by fitting a generalized linear
 198 model to the population growth data (n.s.: not significant, $p > 0.05$; * $p = 0.01$ in expt. 1 and $2e-7$ in expt. 2; ** $p < 2e-16$ in
 199 both experiments).

200

201

202 **MIC depends on inoculum size**

203

204 We hypothesized that the resistant strain's low probability of establishment even far below its MIC
 205 could be explained by an inoculum size effect. Specifically, MIC is standardly measured with an

206 inoculation density of 5×10^5 CFU/mL [15], which corresponds to 10^5 CFU inoculated per well in our
207 experimental set-up. In contrast, in the seeding experiments above, we inoculated on the order of 1
208 CFU per well. MIC for many antibiotics is known to increase with inoculation densities higher than
209 standard [16] [17] [18]. Although less well-explored, it has also occasionally been noted that MIC can
210 decrease when lower inoculum sizes are used [19] [20].

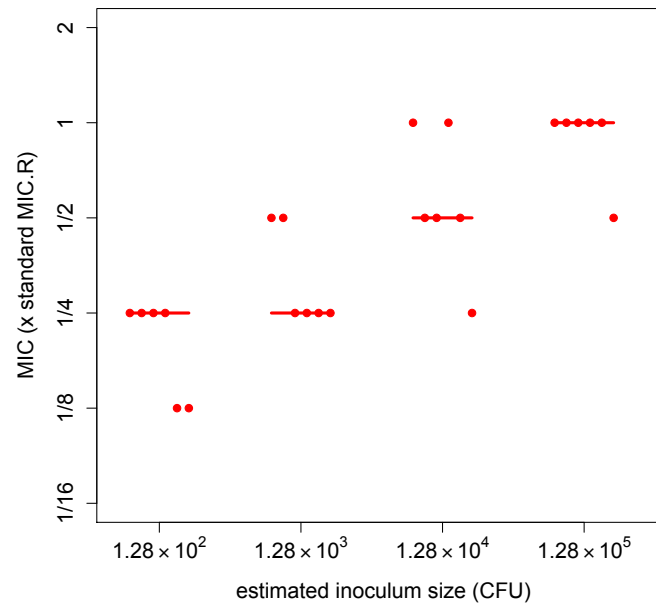
211

212 To test the hypothesis that MIC is inoculum size-dependent, we conducted MIC assays with
213 inoculation densities ranging over three orders of magnitude, from approximately 5×10^2 CFU/ml up
214 to the standard 5×10^5 CFU/ml (i.e. 10^2 - 10^5 CFU/well). We found that MIC indeed increased with
215 inoculum size (**Fig. 4**). This pattern is apparent regardless of whether growth is scored at 20h, as per
216 the standard MIC assay protocol [15], or up to 3d post-inoculation, as in our seeding experiments,
217 although the number of cultures showing detectable growth and thus the observed MIC tends to
218 increase over time (**Suppl. Fig. 4**).

219

220 The observed increase in MIC with inoculum size could have been due either to the increase in
221 absolute number of individuals in the inoculum (i.e. CFU), or in their density (i.e. CFU per unit
222 volume). These two possibilities are not typically distinguished in the literature; however, they lead
223 to distinct interpretations. If demographic stochasticity is the dominant force, we expect absolute
224 numbers to matter, whereas if ecological effects (competition or cooperation) affect establishment,
225 cell density per unit volume could be more important. To disentangle these two factors, we
226 repeated the MIC assay co-varying inoculation density and culture volume. This experiment
227 confirmed that absolute inoculum size has a strong effect on MIC, but revealed that when
228 controlling for absolute size, inoculation density per unit volume does not have a significant effect
229 on the MIC, within the range that we tested (**Suppl. Fig. 5**).

230



231

232

233

234

235

236

237

238

239

240

241

242

243

244

245

246

247

248

249

250

251

252

253

254

Figure 4: MIC as a function of inoculum size. Test cultures at two-fold concentration steps of streptomycin were inoculated at four different densities. MIC was evaluated as the minimal tested concentration that prevented detectable growth up to 3d post-inoculation; a qualitatively similar pattern arose if growth was evaluated at 20h (**Suppl. Fig. 4**). The y-axis is scaled by the previously measured MIC of the resistant strain at standard inoculation density (MIC_R), as before. The points represent six biologically independent replicates at each inoculum size, with the line segments indicating their median.

Population growth can be explained by an independent chance of each cell to establish

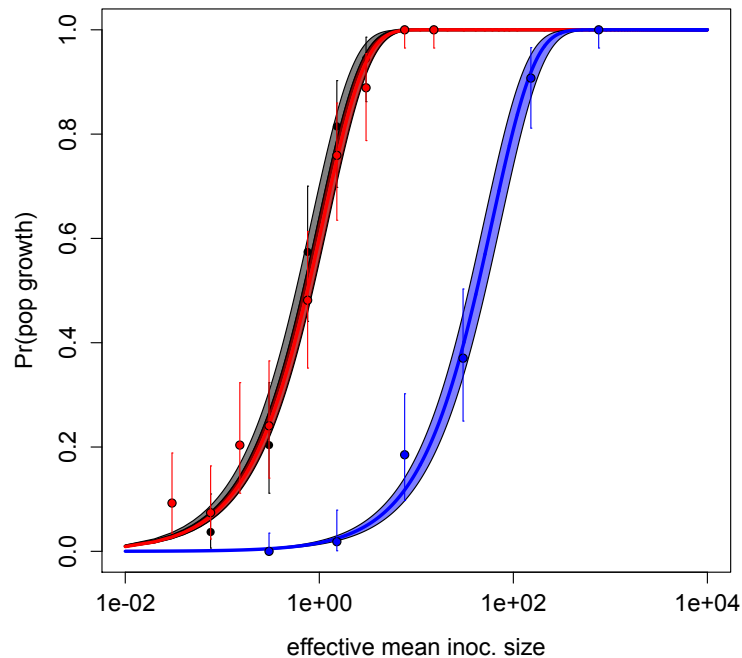
The above MIC tests revealed that the absolute number of cells in the inoculum has a strong effect on whether the population eventually shows detectable growth at streptomycin concentrations within the MSW. The simplest explanation for this effect is that at a given streptomycin concentration, each inoculated cell has a certain probability of establishing a surviving lineage (leading to detectable population growth), independently of all other cells in the inoculum. This independence assumption leads to a simple stochastic model capturing the effect of inoculum size on the probability of population growth (**Methods, Eqn. 2**). Note that this independent model of cellular replication need not be true; for example, if cells secrete an enzyme that breaks down an antibiotic extracellularly, then the chance that an individual cell establishes a successful lineage could increase with initial cell density due to cooperative resistance. On the other hand, if cells strongly compete for limiting resources or secrete toxins, the per-cell chance of establishment could decrease with initial cell density.

255 To formally test the null hypothesis that each inoculated cell has an independent probability of
256 establishment, we again conducted high-replicate population growth experiments. These were
257 similar to the seeding experiments, but used multiple inoculum sizes spanning approximately three
258 orders of magnitude, in order to capture the full range of response (i.e. 0-100% of cultures showing
259 growth) at a given streptomycin concentration. We tested two concentrations ($1/16$ and $1/8 \times \text{MIC}_R$)
260 for which growth often failed in the seeding experiment, but succeeded at standard inoculation
261 density in the MIC tests. In parallel, we tested growth in streptomycin-free media in order to
262 calibrate the “effective” mean inoculum size (**Methods** and **Suppl. Fig. 6**).

263

264 We found good agreement between the null model and our experimental data at all tested
265 streptomycin concentrations, as we would expect if establishment were independent (main
266 experiment, **Fig. 5**, and repeat experiments, **Suppl. Fig. 7**). More precisely, in no case did the null
267 model show significant deviance from the observed proportion of populations that grew (according
268 to the likelihood ratio test; see **Suppl. Text** for full results), and thus we accept it as the most
269 parsimonious explanation for the data. At $1/16$ and $1/8 \times \text{MIC}_R$ we obtain estimates of relative
270 establishment probability, \tilde{p}_c , comparable to those from the seeding experiments (**Suppl. Table 2**).

271



272

273 **Figure 5: Null model of the inoculum size effect fit to population growth data.** Probability of population growth of the
274 resistant strain is plotted as a function of effective mean inoculum size (\bar{N}_{eff} , calibrated by the results in streptomycin-free
275 media; see **Suppl. Fig. 6**). Black: streptomycin-free; red: streptomycin at $1/16 \times \text{MIC}_R$; blue: $1/8 \times \text{MIC}_R$. The solid line shows
276 the best fit of the null model (using the maximum likelihood estimate [MLE] of \tilde{p}_c) and the shaded area corresponds to the
277 95% confidence interval (CI). Points and error bars indicate the MLEs and 95% CIs in the full model, i.e. treating each

278 inoculum size separately (here MLE simply equals the proportion of experimental replicate cultures showing growth).
279 According to the likelihood ratio test, the null model deviance from the full model is not significant at any streptomycin
280 concentration (streptomycin-free: $p=0.55$; $1/16 \times \text{MIC}_R$: $p=0.28$; $1/8 \times \text{MIC}_R$: $p=0.71$).

281

282

283 ***Sub-MIC streptomycin concentrations extend lag phase and induce cell death***

284

285 To explain the reduced establishment probability of the resistant strain when exposed to sub- MIC_R
286 concentrations of streptomycin, we hypothesized that even these low doses must compromise cell
287 division rate and/or viability. As a simple test of this idea, we measured the relative abundance of
288 dead cells in cultures of the resistant strain grown under various concentrations of streptomycin. We
289 found that the fraction of dead cells after 7h of treatment, as determined by propidium iodide
290 staining, increases from an average of 3-4% in streptomycin-free conditions to >20% at $1/8 \times \text{MIC}_R$
291 streptomycin (**Fig. 6a**, **Suppl. Fig. 8** and **Suppl. Table 3**). Note that this is a conservative measure of
292 cell death, because this assay only detects cells that have compromised membrane permeability,
293 and does not detect, for example, lytic cell death.

294

295 To gain further insight into how the toxic effects of streptomycin impact the population dynamics of
296 the resistant strain, we quantified viable cell density (colony-forming units in plated samples) over
297 the first few hours of sub- MIC_R streptomycin treatment. In this experiment, cultures were inoculated
298 with approximately 100 CFU/well, such that populations are small enough for demographic
299 stochasticity to be relevant, yet large enough to be detectable by plating samples.

300

301 We found that streptomycin treatment has a significant effect on population growth (ANOVA, main
302 effect: $p < 2e-16$), and this effect varies over time (ANOVA, interaction term: $p < 2e-16$). Following
303 inoculation into fresh medium, the resistant strain experienced a lag phase of approximately 2
304 hours. Control cultures that were exposed to streptomycin-free medium then exited lag phase and
305 began to grow exponentially. The lowest tested dose of streptomycin ($1/32 \times \text{MIC}_R$) had no
306 significant effect on these dynamics (Dunnett's test: $p = 0.87$); however, $1/16 \times \text{MIC}_R$ was already
307 sufficient to slow population growth ($p = 4e-4$). Nonetheless, all replicate populations (48/48 at each
308 concentration, pooled across all sampled treatment plates) eventually grew, as detected by OD, at
309 streptomycin concentrations up to $1/16 \times \text{MIC}_R$. Meanwhile, higher doses of streptomycin had
310 dramatic effects ($1/8 \times$ or $1/4 \times \text{MIC}_R$ vs. control: $p < 1e-4$), with cultures exhibiting an extended lag
311 phase of at least 7-8 hours, in which viable cell density initially declined. After further incubation (up
312 to 3 days), 25% of cultures (15/60) exposed to $1/8 \times \text{MIC}_R$ eventually showed growth to saturation,

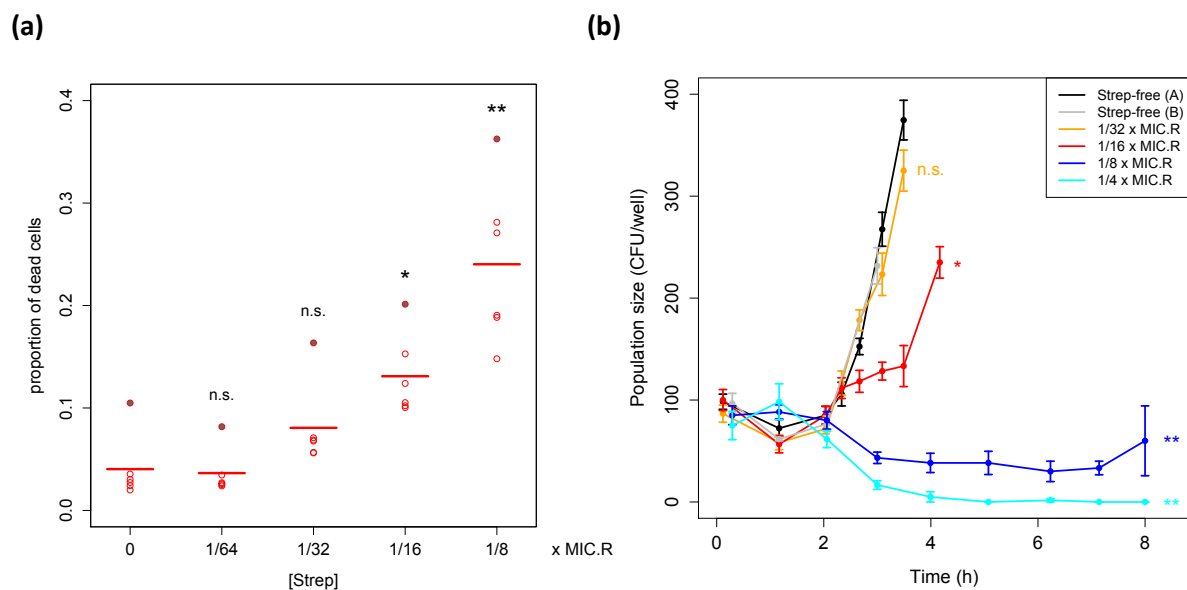
313 while the remaining 75% of cultures (45/60) failed to reach detectable OD. At $1/4 \times \text{MIC}_R$, no viable
314 cells were sampled in most replicates from 4h on, and only 1/60 replicates reached detectable OD
315 within 3 days.

316

317 In summary, these population dynamics experiments reveal that sub-MIC streptomycin treatment
318 has the effect of extending the lag phase, before populations eventually either grow to saturation or
319 die out. The failure of many populations to establish at sub-MIC_R concentrations can be explained by
320 a significant fraction of cells dying rather than successfully dividing, which can lead to stochastic loss
321 of initially small populations.

322

323



324

325 **Figure 6: Sub-MIC_R streptomycin treatment has profound effects on resistant cell dynamics. (a) Proportion of dead cells**

326 **after 7h treatment in sub-MIC_R streptomycin.** The proportion of dead cells was estimated using live-dead staining and

327 flow cytometry (**Methods**). Note that this is a minimal estimate of the proportion of dead cells, since propidium iodide

328 staining only detects cells with compromised membranes. The points represent six treatment replicates at each

329 concentration and the line segments indicate their mean. Differences from the streptomycin-free control cultures were

330 assessed using a one-way ANOVA followed by a post-hoc Dunnett's test (n.s.: not significant, $p > 0.05$; *, $p = 9e-3$, **, $p <$

331 $1e-4$). The significant effects do not change if we exclude an outlier replicate (shaded-in points) showing consistently

332 elevated dead fractions (**Suppl. Table 3**). (b) **Viable cell population dynamics in sub-MIC_R streptomycin.** The plotted points

333 with connecting lines indicate mean population size of six independent replicate cultures per streptomycin concentration

334 and per sampling time point (or twelve replicates for each streptomycin-free control; see **Methods**); the error bars indicate

335 standard error. See **Suppl. Fig. 9** for a version showing all individual replicates. Population size was estimated by plating

336 and truncated when plated colonies became too dense to count. Significance of each streptomycin concentration

337 compared to the streptomycin-free control was assessed by a post-hoc Dunnett's test (n.s.: not significant, $p=0.87$; * $p =$

338 $4e-4$; ** $p < 1e-4$).

339

340

341 Discussion

342

343 In order for a novel resistant strain to emerge, it must successfully expand from a single cell to form
344 a large population of cells. At the level of individual cells, division and death occur as probabilistic
345 events, and demographic stochasticity should therefore be critical to the fate of newly arisen
346 resistant strains. Our key finding is that these stochastic effects impose a significant barrier to the
347 emergence of antibiotic resistance. This result was highlighted by our seeding experiment, in which
348 cultures were inoculated with approximately one resistant cell at streptomycin concentrations well
349 below the MIC of the resistant strain. Strikingly, the chance of population growth was significantly
350 reduced by streptomycin concentrations as low as $1/16 \times \text{MIC}_R$, and nearly abolished (<5% chance)
351 by $1/8 \times \text{MIC}_R$ (**Fig. 3**). This result can be explained by detrimental effects of sub-MIC antibiotic
352 exposure, including a significant mortality rate on the individual cell level (**Fig. 6**), leading to frequent
353 stochastic extinction of small populations even at concentrations where the expected net population
354 growth rate is positive (albeit slowed). One way to quantify the importance of stochasticity is to
355 measure the difference between the range of antibiotic concentrations at which resistance is
356 selectively favoured and the range at which resistance is likely to establish when rare. In this case,
357 we estimate that resistance is selectively favoured, and will on average grow, between $1/512 \times$ and 1
358 $\times \text{MIC}_R$, or $1/4 \times$ and $128 \times \text{MIC}_S$ (**Fig. 1** and **Fig. 4**); in contrast, resistance is very unlikely to emerge
359 from a single cell at concentrations $\geq 1/8 \times \text{MIC}_R$ (**Fig. 3**). Therefore, stochastic effects impose a
360 major constraint on the emergence of resistance across nearly 90% of the mutant selection window.

361

362 It is possible that survival of resistant lineages, while appearing stochastic at the macro-scale, is
363 governed by intrinsic differences among cells in their phenotypic level of resistance [21] [22]. Our
364 model is still valid in case of such phenotypic heterogeneity, but the per-cell establishment
365 probability must then be interpreted as an average among cells (**Suppl. Text**). Further investigation
366 of the cell-level factors that influence the fate of lineages represents an interesting avenue for future
367 work.

368

369 Standardized MIC assays are a widely used and crucial tool for repeatable evaluation of the
370 susceptibility of bacterial strains to antibiotics [15]. However, the discrepancy between the antibiotic
371 concentration required to reliably prevent growth of a larger population (as in a standard MIC assay)
372 and the concentration that is sufficient to drastically reduce the probability of growth of a smaller
373 population, highlights that MIC depends on inoculum size (**Fig. 4**). To clarify this effect, we rigorously
374 quantified the probability of population growth from inoculum sizes spanning three orders of

375 magnitude, in the presence and absence of streptomycin. We found that a simple null model, in
376 which each individual cell independently has some antibiotic concentration-dependent probability of
377 establishing a surviving lineage, is sufficient to explain our data (**Fig. 5**). In other words, we need not
378 invoke any interactions among cells (e.g. competition or cooperation) to explain the low probability
379 of emergence of resistance when starting from small inoculum sizes. Moreover, this simple
380 stochastic “numbers game” provides a unifying explanation for all of the preceding results. When
381 only a single cell is inoculated, as in the seeding experiments, outgrowth is very unlikely at as little as
382 $1/8^{\text{th}}$ of this strain’s standard MIC. However, the cumulative effect of many individuals, each with a
383 small chance of establishment, virtually guarantees population growth from a sufficiently large
384 inoculum size, as in standard MIC assays.

385

386 The “inoculum effect” has previously been recognized primarily in reference to the reduced efficacy
387 of antibiotic treatment (i.e. increase in MIC) as bacterial population density increases beyond that
388 used in standard MIC assays [16] [17]. This effect can arise via various mechanisms, including
389 titration of antibiotics out of the media [17] and enzymatic inactivation of antibiotics by resistant
390 cells [23] [24] [18] [25] [13]. In contrast to this density-dependent, cooperative regime at high
391 bacterial density, the inoculum effect that we found at very low bacterial population size is a
392 stochastic phenomenon driven by cells acting effectively independently. There are hints of this low
393 inoculum size effect in earlier literature [19] [20], but to our knowledge we are the first to provide a
394 rigorous explanation in terms of stochastic population dynamics. These two types of inoculum effect
395 are distinct: the cooperative growth regime arises for certain antibiotics at high bacterial
396 concentrations (cells/ml), whereas we expect the stochastic growth regime to arise very generally at
397 low absolute numbers of cells. That is, we expect the functional relationship between probability of
398 population growth and absolute inoculum size (**Eqn. 2**) to hold widely (with the scaling parameter \tilde{p}_c
399 depending on the bacterial strain and antibiotic); however, the relationship between MIC and
400 inoculum size will also depend on the relationship between \tilde{p}_c and antibiotic concentration.
401 Moreover, measured MIC will become increasingly variable at low inoculum size, calling for a more
402 precise measure of inhibition, such as the IC_{99} [26]. In these two facets of the inoculum effect, cells
403 at the limit of low density (cells/ml) where cooperative interactions are abolished [18] are
404 distinguished from low absolute numbers of cells, where stochastic effects become relevant. Both
405 types of inoculum effect raise important considerations for antibiotic dosing. Cooperative growth
406 has implications for treatment of high-density bacterial infections [24] [27] [28], whereas the
407 stochastic effects we studied here could have implications for dosing strategies to avoid emergence
408 of *de novo* resistant mutants.

409

410 Although the role of demographic stochasticity in the fate of *de novo* mutations has long been
411 recognized in theoretical population genetics, until very recently it had never been addressed
412 empirically [12]. Our study joins a small handful of others that have now experimentally quantified
413 establishment probability from single cells [29] [26] [30] [31]. Two of these studies [26] [31] likewise
414 addressed establishment of bacterial cells across antibiotic concentrations, but using different
415 methods to ours (see **Suppl. Text** for a more detailed comparison). Bacterial evolution of resistance
416 to antibiotic treatment is also a prime example of the more general phenomenon of evolutionary
417 rescue, whereby adaptation prevents extinction of populations facing severe environmental change
418 [32]. Indeed, our equation for the probability of population growth (**Eqn. 2**) is identical in form to
419 general approximations for the probability of rescue [33] (see **Suppl. Text** for further discussion).
420 More broadly, the concepts and methods developed here could be applied to a variety of situations
421 where growth depends on establishment of rare cells and is thus highly stochastic, for instance the
422 onset of invasive bacterial infections [34] or the outgrowth of bacteria in food products from small
423 initial contaminants [35].

424

425 In summary, our study highlights the stochastic nature of *de novo* emergence of antibiotic
426 resistance, and sheds new light on the question of what concentrations of antibiotics drive
427 resistance evolution. Importantly, we found that concentrations within the traditional mutant
428 selection window (MSW), above the MIC of the sensitive strain (MIC_S) but still well below the MIC of
429 the resistant strain (MIC_R), can be sufficient to severely limit *de novo* establishment of resistance
430 from single cells. Thus, our findings suggest that moderate doses (between MIC_S and MIC_R) may be
431 more effective than previously thought at preventing emergence of resistance, in cases where the
432 chance of transmitted resistance is negligible. Taken together with our own and previous evidence
433 of sub- MIC_S selection for resistance [7] [8] [9] [10] [13], our work contributes to an emerging picture
434 that resistance tends to arise at a range of concentrations shifted lower than the traditional MSW.
435 Importantly, however, we studied establishment of the resistant strain in isolation; it remains to be
436 determined how interactions with an initially dominant sensitive population shape establishment of
437 resistance during antibiotic treatment. A key direction for future work will thus be to integrate the
438 dual effects of competition and direct inhibition by the antibiotic, which operate with varying
439 strengths across antibiotic concentrations [6]. By better understanding these limiting forces on the
440 evolution of resistance, we may ultimately be able to design effective dosing strategies that avoid
441 emergence of resistance.

442

443 **Methods**

444

445 ***Bacterial strains, media and culture conditions***

446

447 Bacterial strains: We used a set of *Pseudomonas aeruginosa* PAO1 strains described previously [13].

448 These include both streptomycin-sensitive and -resistant strains, which are isogenic except that the

449 resistant strains carry the clinically derived, non-conjugative plasmid Rms149 [36]. Streptomycin

450 resistance is conferred by the *aadA5* gene, which is located within a class 1 integron, and codes for

451 an enzyme that adenylates streptomycin [37]. For both plasmid carriers (resistant) and non-carriers

452 (sensitive), we have strains with a chromosomal YFP label, a DsRed label, or unlabelled. The

453 competition experiment was conducted with the DsRed-labelled sensitive and YFP-labelled resistant

454 strains. The live-dead staining experiment was conducted with the unlabelled resistant strain. All

455 other experiments were conducted with the YFP-labelled resistant strain. The minimum inhibitory

456 concentrations of streptomycin in LB broth were previously determined to be $MIC_S = 16\mu\text{g/ml}$ for

457 the sensitive strain and $MIC_R = 2048\mu\text{g/ml}$ for the resistant strain [13].

458

459 Media and antibiotics: For culture media we used LB broth containing 5g/L NaCl (Sigma-Aldrich,

460 product no. L3022). For plating to assess colony-forming units we used LB Agar, Vegitone, containing

461 5g/L NaCl and 15g/L agar (Sigma-Aldrich product no. 19344). Streptomycin was prepared as stock

462 from streptomycin sulfate salt (Sigma-Aldrich product no. S6501), stored according to supplier

463 directions, and added to media on the day of experiments. Bacterial cultures were diluted in

464 phosphate buffered saline (PBS) prepared from tablets (Sigma-Aldrich product no. P4417). Note that

465 treatment cultures were generally set up with 90% media plus 10% inoculating culture by volume;

466 thus, the final concentrations of LB and streptomycin in the treatments are 90% of the prepared

467 media values denoted on plots.

468

469 Culture conditions: All cultures were incubated at 37°C, shaking at 225rpm. Overnight cultures were

470 inoculated directly from freezer stocks into 2ml of LB and then incubated for approximately 16h.

471 Overnight cultures were then diluted in PBS and used to inoculate treatment plates. All experimental

472 treatments were conducted in 200 μl cultures in flat-bottom 96-well microtitre plates unless

473 otherwise noted.

474

475 ***Competition experiment***

476

477 Competitions were carried out between the YFP-labelled resistant strain and the DsRed-labelled
478 sensitive strain, in streptomycin-free media and streptomycin at 1/16 to 16 x MIC_s in 2-fold
479 concentration steps. Previous work demonstrated that these two markers have equivalent effects on
480 bacterial fitness on LB-agar [13]. Treatment cultures were inoculated with a 1:1 volumetric mixture
481 of overnight cultures of the two strains, with expected total inoculum size around 10⁵ CFU. We
482 tested six replicate mixed cultures along with two of each pure culture as controls. After 24h,
483 treatment cultures were diluted 500-fold and 66µl of each diluted culture was sampled by flow
484 cytometry (BD Accuri C6 Flow Cytometer, fast fluidics, discarding events with forward scatter FSC-H
485 < 10,000 or side scatter SSC-H < 8000). By gating based on fluorescence and correcting for the
486 overlap of each strain into the opposing gate using the pure culture controls (see **Suppl. Text** and
487 **Suppl. Fig. 1**), we estimated the final frequency of each strain.

488

489 ***Seeding experiment***

490

491 Experimental protocol: An overnight culture was serially diluted up to 10⁶-fold, then in three further
492 independent steps, diluted to 4 x 10⁷-, 8 x 10⁷, and 1.6 x 10⁸-fold. Each of the latter diluted cultures
493 was used to inoculate an entire 96-well test plate for each tested streptomycin concentration, at
494 20µl per well. Test plates were incubated and scored for growth after approximately 1, 2, and 3d.

495

496 Scoring culture growth: We evaluated culture growth by measuring optical density (OD₅₉₅) using a
497 BioTek Synergy 2 plate reader. Lids on 96-well plates were briefly removed for the reading;
498 comparison to controls mock-inoculated with PBS indicated that contamination was rare. We set a
499 threshold OD of 0.1 to score as growth, as background OD in controls was typically <0.05. OD
500 showed a clearly bimodal distribution after a sufficiently long growth period; thus, the precise choice
501 of threshold is unimportant.

502

503 Estimating establishment probability: From the number of replicate cultures showing growth by 3d
504 at each condition, we obtained a likelihood-based point estimate and confidence interval for the
505 single-cell establishment probability, relative to streptomycin-free controls, by fitting a
506 mathematical model to the data (see section “Mathematical model of establishment” below).

507

508 Generalized linear model: To evaluate the effect of streptomycin on establishment, culture growth
509 data was additionally fit with a generalized linear model using the built-in R function ‘glm’. Number
510 of replicate cultures showing growth was treated as binomial data, with inoculating dilution factor

511 and streptomycin concentration taken as explanatory variables, using the complementary log-log
512 link function (see **Suppl. Text** for details).

513

514 ***MIC assay with varying inoculation density***

515

516 Test cultures were inoculated with 20 μ l of overnight culture diluted 10³-, 10⁴-, 10⁵-, or 10⁶-fold.
517 These dilution factors were chosen such that the highest inoculation density was expected to be
518 similar to a standard MIC test (5 x 10⁵ CFU/ml; [15]); actual density was estimated by plating out
519 diluted overnight culture on LB-agar and counting colony-forming units. We tested six replicates per
520 inoculation density and per streptomycin concentration, ranging from 1/16 x MIC_R to 2 x MIC_R in
521 two-fold steps, plus streptomycin-free controls. Culture growth was evaluated by OD at
522 approximately 20h, 2d, and 3d, according to the same procedure as for the seeding experiments. In
523 rare cases, MIC within a replicate was ambiguous due to growth failing in a culture at lower
524 streptomycin concentration but succeeding at the next highest step before being definitively
525 abolished. In these cases, we took MIC to be the higher concentration, at and beyond which no
526 further growth was observed.

527

528 ***Testing the null model of the inoculum size effect***

529

530 To test the relationship between inoculum size and probability of population growth, we evaluated
531 growth at selected streptomycin concentrations using many different dilution factors for
532 inoculation: five in streptomycin-free conditions and 6-10 in each streptomycin concentration. These
533 dilution factors were chosen for each streptomycin concentration based on the results of previous
534 experiments, with the aim of capturing the range over which the proportion of replicate cultures
535 showing growth increases from near 0 to near 1. In each case, 54 replicate cultures were evaluated
536 per inoculum size. Culture growth was evaluated daily by OD, according to the same procedure as
537 for the seeding experiments, up to 5d. The final number of replicate cultures showing growth was
538 used for model fitting (see section “Mathematical model of establishment” below).

539

540 ***Fraction of dead cells by live-dead staining***

541

542 For this experiment, we used the resistant strain with no fluorescent label, to avoid interfering with
543 the signal from the live-dead stain. We inoculated treatment cultures (six replicates per
544 streptomycin concentration) with 10³-fold diluted overnight culture, as in the MIC assay at standard

545 inoculation density. After 7h of treatment, we diluted test cultures a total of 100-fold and stained
546 with thiazole orange and propidium iodide (BD Cell Viability Kit, product no. 349483). In parallel, we
547 diluted and stained media and heat-killed cultures as controls. We sampled 50 μ l per diluted culture
548 using flow cytometry (with the same settings as in the competition experiment). The staining and
549 flow cytometry steps were carried out in groups containing one replicate per concentration plus
550 controls, to avoid potentially toxic effects of stain exposure over prolonged times (**Suppl. Text**). To
551 better discriminate cells from background in the flow cytometry data, we first gated on events
552 according to forward and side scatter before defining clusters of dead (membrane-compromised)
553 and intact cells based on fluorescence; see **Suppl. Text** and **Suppl. Fig. 8** for details.

554

555 ***Viable cell density dynamics***

556

557 We tracked the number of viable cells over time in streptomycin-free media and at 1/32, 1/16, 1/8,
558 and 1/4 x MIC_R streptomycin. An independent test plate was used for sampling at each time point.
559 Six replicate cultures per concentration, per test plate (twelve replicates for streptomycin-free) were
560 inoculated with 20 μ l of 5 x 10⁵-fold diluted overnight culture. Streptomycin treatments were split
561 across two sets of plates, each of which included a streptomycin-free control; set A additionally
562 included the two lower streptomycin concentrations and set B included the two higher. These sets
563 were sampled at generally different time points. At each sampling time, undiluted cultures were
564 plated in arrays of 4 μ l spots, repeated across 5 square LB-agar plates for a total sampled volume of
565 20 μ l per culture. LB-agar plates were incubated at 37°C for the rest of the day of the experiment,
566 moved to the bench overnight to prevent overgrowth, then incubated again the next day at 37°C
567 until colonies were visible, but still separated, for counting. Population size in each replicate culture
568 was then estimated by scaling up total CFU count by a factor 10 (since 1/10th of the culture volume
569 was sampled). Comparing streptomycin-free controls from sets A and B indicated that the plate set
570 effect was non-significant (ANOVA: $p = 0.103$); thus, controls were pooled for further analysis of the
571 streptomycin effect.

572

573 ***Mathematical model of establishment***

574

575 Model: We used a simple stochastic model to describe population (i.e. bacterial culture) growth as a
576 function of inoculum size and streptomycin concentration. We denote by p_w the probability that a
577 population inoculated with a small number of cells grows to a large size (detected by OD as
578 described above). Among a set of n independent replicate populations (inoculated at the same

579 density and cultured in the same conditions), the number of populations showing growth is thus
580 described by a Binomial(n, p_w) distribution. Under the assumptions that (i) the number of cells in the
581 inoculum that establish surviving lineages is Poisson-distributed with some mean α (an assumption
582 supported by the observed distribution of colony-forming units counted in highly diluted cultures;
583 **Suppl. Fig. 3**), and (ii) population growth is observed provided at least one cell establishes, we have
584 the relationship:

585

$$586 \quad p_w = 1 - e^{-\alpha} \quad (\text{Eqn. 1})$$

587

588 Relative per-cell establishment probability $\tilde{p}_c(x)$ is generally calculated as the ratio of the mean
589 number of established cells at any given streptomycin concentration x to that in the absence of
590 streptomycin: $\tilde{p}_c(x) = \alpha(x)/\alpha(0)$.

591

592 In the full model (statistically speaking), p_w , or equivalently α , is taken to be distinct for each
593 inoculum size and each streptomycin concentration. Simpler nested models make additional
594 assumptions (see **Suppl. Text** for details). In particular, the null model of the inoculum size effect
595 supposes that cells behave independently of one another, such that the true per-cell probability of
596 establishment, p_c , depends only on streptomycin concentration. Then the mean number of
597 established cells, α , is directly proportional to the mean inoculum size, \bar{N} (where the actual
598 inoculum size is Poisson-distributed about this mean). In this null model, the probability of
599 population growth can be expressed as:

600

$$601 \quad p_w = 1 - e^{-\bar{N}p_c}$$
$$602 \quad = 1 - e^{-\bar{N}_{eff}\tilde{p}_c} \quad (\text{Eqn. 2})$$

603

604 where $\bar{N}_{eff} = \alpha(0) = \bar{N}p_c(0)$ denotes the “effective mean inoculum size”, defined as the mean
605 number of established cells in the absence of streptomycin; and the relative establishment
606 probability at streptomycin concentration x can simply be expressed as $\tilde{p}_c = p_c(x)/p_c(0)$. Note
607 that by scaling up \bar{N}_{eff} by the dilution factor(s) applied to the inoculating culture, we have an
608 estimate of bacterial density in this culture, determined equivalently to the historical “most
609 probable number” method [38] [39]. This definition of effective inoculum size based on cells that
610 grow in benign conditions is no different in principle to the usual quantification of “viable” cells
611 according to successful formation of a colony; we simply assess growth in liquid rather than on solid

612 medium. We also note that if cells (or possibly clumps of cells acting as individual units) are
613 heterogeneous, then p_c should be interpreted as the mean establishment probability (**Suppl. Text**).

614

615 Likelihood-based model fitting and comparisons: These stochastic models were fit to experimental
616 population growth data using likelihood-based methods. Specifically, we obtained a maximum
617 likelihood estimate and a 95% confidence interval (determined by the range of parameter values
618 that would not be rejected by a likelihood ratio test at 5% significance level) on the parameter p_w ,
619 which can be transformed to an estimate for α . In the case of relative establishment probability,
620 $\tilde{p}_c(x) = \alpha(x)/\alpha(0)$, the (profile likelihood) confidence interval takes into account the uncertainty
621 in both numerator (i.e. results at streptomycin concentration x) and denominator (i.e. results in
622 streptomycin-free conditions). The fit of nested models is compared using the likelihood ratio test
623 (LRT) at 5% significance level, i.e. a χ^2 test on model deviance with degrees of freedom equal to the
624 difference in number of fitted parameters between the two models.

625

626 To test the null model of the inoculum size effect, we assume that mean inoculum size \bar{N} is precisely
627 inversely proportional to the dilution factor applied to the inoculating culture, i.e. we neglect any
628 experimental error in the dilution steps relative to one another. Effective mean inoculum size, \bar{N}_{eff} ,
629 is estimated by fitting Eqn. 2 to population growth data in streptomycin-free media. Per-cell relative
630 establishment probability \tilde{p}_c then remains as the single free parameter to fit at each tested
631 streptomycin concentration. The goodness of fit of the null model (**Eqn. 2**) is assessed for each test
632 concentration separately. Here, the relevant comparison in the LRT is to the full model (**Eqn. 1**),
633 where a distinct parameter p_w (or equivalently α) is estimated for each inoculum size, without any
634 assumed relationship among them.

635

636 All model fitting was implemented in R, version 3.3.1 (The R Foundation for Statistical Computing,
637 2016).

638

639

640 **Acknowledgements:** H.K.A. was supported by an Early Postdoc.Mobility Fellowship
641 (P2EZP3_165188) and an Advanced Postdoc.Mobility fellowship (P300PA_177789) from the Swiss
642 National Science Foundation. R.C.M. was supported by Wellcome Trust Grant 106918/Z/15/Z.

643

644 **Author contributions:** H.K.A. and R.C.M. conceived of the study and designed experiments. H.K.A.
645 carried out experiments and data analysis with advice from R.C.M. H.K.A. developed the

646 mathematical model, wrote the code and carried out model fitting. H.K.A. and R.C.M. wrote the
647 paper.

648

649 **Data availability:** The data generated in this study will be deposited in a public repository upon
650 manuscript acceptance.

651

652 **Code availability:** Custom R scripts for likelihood-based model fitting and comparisons will be made
653 freely available upon manuscript acceptance.

654

655

656 **References**

657

- [1] World Health Organization, "Antimicrobial Resistance: Global Report on Surveillance," WHO Press, Geneva, 2014.
- [2] X. Zhao and K. Drlica, "Restricting the selection of antibiotic-resistant mutants: a general strategy derived from fluoroquinolone studies," *Clin. Infect. Dis.*, vol. 33 (Suppl 3), pp. S147-S156, 2001.
- [3] X. Zhao and K. Drlica, "Restricting the selection of antibiotic-resistant mutant bacteria: measurement and potential use of the mutant selection window," *J. Infect. Dis.*, vol. 185, pp. 561-565, 2002.
- [4] K. Drlica and X. Zhao, "Mutant selection window hypothesis updated," *Clin. Infect. Dis.*, vol. 44, pp. 681-688, 2007.
- [5] R. Canton and M.-I. Morosini, "Emergence and spread of antibiotic resistance following exposure to antibiotics," *FEMS Microbiol. Rev.*, vol. 35, pp. 977-991, 2011.
- [6] T. Day, S. Huijben and A. F. Read, "Is selection relevant in the evolutionary emergence of drug resistance?," *Trends Microbiol.*, vol. 23, pp. 126-133, 2015.
- [7] M.-C. Negri, M. Lipsitch, J. Blázquez, B. R. Levin and F. Baquero, "Concentration-dependent selection of small phenotypic differences in TEM beta-lactamase-mediated antibiotic resistance," *Antimicrob. Agents Chemother.*, vol. 44, pp. 2485-2491, 2000.
- [8] A. Liu, A. Fong, E. Becket, J. Yuan, C. Tamae, L. Medrano, M. Maiz, C. Wahba, C. Lee, K. Lee, K. P. Tran, H. Yang, R. M. Hoffman, A. Salih and J. H. Miller, "Selective advantage of resistant strains at trace levels of antibiotics: a simple and ultrasensitive color test for detection of antibiotics and genotoxic agents," *Antimicrob. Agents Chemother.*, vol. 55, no. 3, pp. 1204-1210, 2011.
- [9] E. Gullberg, S. Cao, O. G. Berg, C. Ilbäck, L. Sandegren, D. Hughes and D. I. Andersson, "Selection of resistant bacteria at very low antibiotic concentrations," *PLoS Pathogens*, vol. 7, p. e1002158, 2011.
- [10] D. I. Andersson and D. Hughes, "Evolution of antibiotic resistance at non-lethal drug concentrations," *Drug Resistance Updates*, vol. 15, pp. 162-172, 2012.
- [11] P. Abel zur Wiesch, R. Kouyos, J. Engelstädter, R. R. Regoes and S. Bonhoeffer, "Population biological principles of drug-resistance evolution in infectious diseases," *Lancet Infect. Dis.*, vol. 11, pp. 236-247, 2011.

- [12] Z. Patwa and L. M. Wahl, "The fixation probability of beneficial mutations," *J. R. Soc. Interface*, vol. 5, pp. 1279-1289, 2008.
- [13] I. Frost, W. P. J. Smith, S. Mitri, A. San Millan, Y. Davit, J. M. Osborne, J. M. Pitt-Francis, R. C. MacLean and K. R. Foster, "Cooperation, competition and antibiotic resistance in bacterial colonies," *ISME J.*, vol. 12, pp. 1582-1593, 2018.
- [14] I. Levin-Reisman, I. Ronin, O. Gefen, I. Braniss, N. Shoshani and N. Q. Balaban, "Antibiotic tolerance facilitates the evolution of resistance," *Science*, vol. 355, pp. 826-830, 2017.
- [15] European Committee for Antimicrobial Susceptibility Testing (EUCAST), "Determination of minimum inhibitory concentrations (MICs) of antibacterial agents by broth dilution," *Clin. Microbiol. Infect.*, vol. 9, pp. 1-7, 2003.
- [16] I. Brook, "Inoculum effect," *Rev. Infect. Dis.*, vol. 11, pp. 361-368, 1989.
- [17] K. I. Udekwe, N. Parrish, P. Ankomah, F. Baquero and B. R. Levin, "Functional relationship between bacterial cell density and the efficacy of antibiotics," *J. Antimicrob. Chemother.*, vol. 63, pp. 745-757, 2009.
- [18] T. Artemova, Y. Gerardin, C. Dudley, N. M. Vega and J. Gore, "Isolated cell behavior drives the evolution of antibiotic resistance," *Molec. Syst. Biol.*, vol. 11, p. 822, 2015.
- [19] K. C. Haltalin and J. D. Nelson, "In vitro susceptibility of Shigellae to sodium sulfadiazine and to eight antibiotics," *JAMA*, vol. 193, pp. 705-710, 1965.
- [20] S. E. McLinn, J. D. Nelson and K. C. Haltalin, "Antimicrobial susceptibility of Hemophilus influenzae," *Pediatrics*, vol. 45, pp. 827-838, 1970.
- [21] I. El Meouche, Y. Siu and M. J. Dunlop, "Stochastic expression of a multiple antibiotic resistance activator confers transient resistance in single cells," *Sci. Rep.*, vol. 6, p. 19538, 2016.
- [22] T. Bergmiller, A. M. C. Andersson, K. Tomasek, E. Balleza, D. J. Kiviet, R. Hauschild, G. Tkačik and C. C. Guet, "Biased partitioning of the multidrug efflux pump AcrAB-TolC underlies long-lived phenotypic heterogeneity," *Science*, vol. 356, pp. 311-315, 2017.
- [23] E. A. Yurtsev, H. X. Chao, M. S. Datta, T. Artemova and J. Gore, "Bacterial cheating drives the population dynamics of cooperative antibiotic resistance plasmids," *Molec. Syst. Biol.*, vol. 9, p. 683, 2013.
- [24] N. M. Vega and J. Gore, "Collective antibiotic resistance: mechanisms and implications," *Curr. Opin. Microbiol.*, vol. 21, pp. 28-34, 2014.
- [25] R. A. Sorg, L. Lin, G. S. van Doorn, M. Sorg, J. Olson, V. Nizet and J.-W. Veening, "Collective resistance in microbial communities by intracellular antibiotic deactivation," *PLoS Biol.*, vol. 14, p. e2000631, 2016.
- [26] M. F. Schenk, I. G. Szendro, J. Krug and J. A. G. M. de Visser, "Quantifying the adaptive potential of an antibiotic resistance enzyme," *PLoS Genet.*, vol. 8, p. e1002783, 2012.
- [27] H. R. Meredith, J. K. Srimani, A. J. Lee, A. J. Lopatkin and L. You, "Collective antibiotic tolerance: mechanisms, dynamics and intervention," *Nat. Chem. Biol.*, vol. 11, pp. 182-188, 2015.
- [28] J. Karslake, J. Maltas, P. Brumm and K. B. Wood, "Population density modulates drug inhibition and gives rise to potential bistability of treatment outcomes for bacterial infections," *PLoS Comput. Biol.*, vol. 12, p. e1005098, 2016.
- [29] D. R. Gifford, J. A. G. M. de Visser and L. M. Wahl, "Model and test in a fungus of the probability that beneficial mutations survive drift," *Biol. Lett.*, vol. 9, p. 20120310, 2012.
- [30] D. R. Gifford and R. C. MacLean, "Evolutionary reversals of antibiotic resistance in experimental populations of *Pseudomonas aeruginosa*," *Evolution*, vol. 67, pp. 2973-2981, 2013.

- [31] J. Coates, B. R. Park, D. Le, E. Simsek, W. Chaudhry and M. Kim, "Antibiotic-induced population fluctuations and stochastic clearance of bacteria," *eLife*, vol. 7, p. e32976, 2018.
- [32] H. K. Alexander, G. Martin, O. Y. Martin and S. Bonhoeffer, "Evolutionary rescue: linking theory for conservation and medicine," *Evol. Appl.*, vol. 7, pp. 1161-1179, 2014.
- [33] G. Martin, R. Aguilée, J. Ramsayer, O. Kaltz and O. Ronce, "The probability of evolutionary rescue: towards a quantitative comparison between theory and evolution experiments," *Phil. Trans. R. Soc. B*, vol. 368, p. 20120088, 2013.
- [34] R. E. Moxon and P. A. Murphy, "Haemophilus influenzae bacteremia and meningitis resulting from survival of a single organism," *PNAS*, vol. 75, no. 3, pp. 1534-1536, 1978.
- [35] K. Koutsoumanis, "A study on the variability in the growth limits of individual cells and its effect on the behavior of microbial populations," *Int. J. Food Microbiol.*, vol. 128, pp. 116-121, 2008.
- [36] A. S. Haines, K. Jones, M. Cheung and C. M. Thomas, "The IncP-6 plasmid Rms149 consists of a small mobilizable backbone with multiple large insertions," *J. Bacteriol.*, vol. 187, pp. 4728-4738, 2005.
- [37] D. Sandvang, "Novel streptomycin and spectinomycin resistance gene as a gene cassette within a class 1 integron isolated from Escherichia coli," *Antimicrob. Agents Chemother.*, vol. 43, pp. 3036-3038, 1999.
- [38] W. G. Cochran, "Estimation of bacterial densities by means of the "most probable number"," *Biometrics*, vol. 6, pp. 105-116, 1950.
- [39] M. A. Hurley and M. E. Roscoe, "Automated statistical analysis of microbial enumeration by dilution series," *J. Appl. Bacteriol.*, vol. 55, pp. 159-164, 1983.

658

659